# **1** Stochastic bacterial population dynamics prevent the emergence of

# 2 antibiotic resistance from single cells

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### 12 Abstract:

- 13 A better understanding of how antibiotic exposure impacts the evolution of resistance is crucial for 14 designing more sustainable treatment strategies. The conventional approach to relating antibiotic 15 dose to resistance evolution within a bacterial population is to measure the range of concentrations 16 over which resistant strain(s) are selectively favoured over a sensitive strain – the "mutant selection 17 window". Here, we instead investigate how antibiotic concentration impacts the initial 18 establishment of resistance from single cells, mimicking the clonal expansion of a resistant lineage 19 following mutation or horizontal gene transfer. Using two Pseudomonas aeruginosa strains carrying 20 distinct resistance plasmids, we show that single resistant cells have <5% probability of outgrowth at 21 antibiotic concentrations as low as 1/8<sup>th</sup> of the resistant strain's minimum inhibitory concentration. 22 This low probability of establishment is due to detrimental effects of antibiotics on resistant cells, 23 coupled with the inherently stochastic nature of cell division and death on the single-cell level, which leads to loss of many nascent resistant lineages. Our findings suggest that moderate doses of 24 25 antibiotics, within the traditional mutant selection window, may be more effective at preventing de 26 *novo* emergence of resistance than predicted by deterministic approaches. 27 28 Keywords: antimicrobial resistance, Pseudomonas aeruginosa, minimum inhibitory concentration 29 (MIC), inoculum effect, mathematical model, extinction probability, demographic stochasticity, 30 evolutionary rescue 31 32 Significance statement: 33 The emergence of antibiotic resistance poses a critical threat to the efficacy of antibiotic treatments. 34 A resistant bacterial population must originally arise from a single cell that mutates or acquires a
- resistance gene. This single cell may, by chance, fail to successfully reproduce before it dies, leading
- 36 to loss of the nascent resistant lineage. Here we show that antibiotic concentrations that selectively
- 37 favour resistance are nonetheless sufficient to reduce the chance of outgrowth from a single cell to a
- 38 very low probability. Our findings suggest that lower antibiotic concentrations than previously
- 39 thought may be sufficient to prevent, with high probability, emergence of resistance from single
- 40 cells.

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# 41 Introduction

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43 Antibiotics have had a huge impact on human health by reducing the burden associated with 44 bacterial infections, and the use of antibiotics now underpins many areas of medicine. 45 Unfortunately, antibiotic treatment is also associated with the evolution of resistance [1], resulting 46 in poorer patient outcomes [2]. A better understanding of how antibiotic dosing affects resistance 47 evolution could aid the design of more effective treatment strategies that suppress pathogenic 48 bacteria without driving the emergence of resistance. 49 50 To date, in vitro work addressing how antibiotics affect evolution of resistance has focused on 51 identifying the range of antibiotic concentrations at which resistant mutants are selectively favoured

52 by antibiotic treatment, known as the "mutant selection window" (MSW) [3, 4, 5, 6]. Here we will 53 refer to any strain with reduced susceptibility relative to a wild-type ("sensitive") strain simply as 54 "resistant", as is common in evolutionary microbiology literature (e.g. [7, 8, 9]), as opposed to 55 defining resistance with respect to clinical breakpoints. Originally, the lower boundary of the MSW 56 was approximated by the minimum inhibitory concentration of the sensitive strain (MIC<sub>s</sub>) [3, 4, 5, 6], 57 i.e. the lowest antibiotic concentration that abolishes its growth in a standardized assay (such as 58 [10]). However, more recent work has emphasized that a resistant strain can be selectively favoured 59 down to a minimal selective concentration (MSC) that is often well below MIC<sub>s</sub> [11, 12, 13, 7, 8]. The 60 upper boundary of the MSW is conventionally defined by the lowest antibiotic concentration that 61 prevents growth of all mutant subpopulations [3, 4]. This upper bound is often equated with the 62 minimum inhibitory concentration of the most resistant single-point mutant [5] or of a specific 63 resistant strain under study [13, 8], which we denote MIC<sub>R</sub>. The MSW therefore ranges between 64 antibiotic concentrations that are so low that they are unlikely to have any clinical benefit (below 65  $MIC_s$ ) and very high concentrations (up to  $MIC_R$ ) that may be difficult to achieve in practice because 66 of physiological constraints on the accumulation of antibiotics in tissues (pharmacokinetics) and

- 67 toxic side-effects of antibiotics [14, 15].
- 68

Selection operates efficiently when both sensitive and resistant populations are large, resulting in an increase in frequency of the fitter strain in an antibiotic dose-dependent manner. Correspondingly, the mutant selection window is typically measured by direct competition between large numbers of cells (typically >10<sup>4</sup> colony-forming units, CFU) of both resistant and sensitive strains across a gradient of antibiotic concentrations (e.g. [13]). This provides a powerful approach to measure the impact of antibiotic dose on selection for established resistant strains, i.e. those that are already

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75 reasonably prevalent. However, de novo emergence of resistant strains, when not initially present, 76 should be subject to stochastic processes [16] that are not reflected by the MSW, nor captured by 77 this experimental design. First, resistance must stochastically arise in a sensitive cell by mutation, or 78 by acquisition of a resistance gene through horizontal gene transfer. Next, the single resistant cell 79 thus generated must survive and successfully divide to produce daughter cells that likewise survive, 80 and so on to generate a large number of resistant descendant cells. The latter process, which we will 81 refer to throughout as "establishment" of resistance [16], will be our focus here. Importantly, due to 82 the stochastic nature of cell divisions and deaths on the individual level, establishment is not 83 guaranteed, even under conditions in which the resistant strain is selectively favoured [17]. 84 Moreover, if antibiotics even partially inhibit the resistant strain below its  $MIC_{R}$ , the chance that a 85 resistant cell dies or fails to divide, and thus the risk that a resistant lineage is stochastically lost, 86 should increase in the presence of antibiotics at concentrations within the MSW. Despite the 87 substantial body of work addressing the selection of resistance, very little work has addressed the 88 stochastic establishment phase (see however [18, 19, 20]). 89 90 We set out to quantify stochastic establishment using in vitro experiments with Pseudomonas

91 aeruginosa, an important opportunistic pathogen that evolves resistance at an exceptionally high 92 rate during infections [1, 21]. To isolate the establishment phase, we inoculated hundreds of 93 cultures, each with a very small number of resistant cells (on average, approx. 1-3), and assessed 94 culture growth. We tested two strains carrying non-conjugative plasmids (Rms149 and PAMBL2) that 95 confer resistance to streptomycin and meropenem, respectively, across a gradient of the 96 corresponding antibiotic concentrations within the MSW. By fitting mathematical models to these 97 data, we estimated the probability of establishment, i.e. of detectable culture growth due to clonal 98 expansion from a single resistant cell, as a function of antibiotic concentration. Our key finding is 99 that the establishment probability of resistant cells drastically declines at concentrations well below 100 the MIC of the resistant strain, reaching  $\leq 5\%$  at  $1/8^{th}$  of MIC<sub>R</sub> in both systems. These concentrations 101 lie well above the corresponding MIC values of the sensitive strain and within the conventional 102 mutant selection window. Our results highlight that antibiotic selection pressure is not a sufficient 103 condition for *de novo* emergence of resistance starting from single cells. Accounting for the 104 demographic stochasticity inherent to the outgrowth of mutant lineages substantially narrows the 105 window of concentrations at which resistant mutants are likely to establish, suggesting that 106 moderate antibiotic dosing may be an effective strategy to prevent the emergence of resistance. 107

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### 109 **Results**

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### 111 Establishment of resistance is inhibited by sub-MIC<sub>R</sub> antibiotic concentrations

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113 To elucidate the direct impact of antibiotics on resistant cells, we first investigated establishment of 114 a resistant strain in the absence of a sensitive strain. We first focused on the streptomycin-resistant 115 PA01:Rms149 strain. To estimate its probability of establishment, defined as outgrowth of a 116 detectable (i.e. large) population from a single cell, we conducted large-scale "seeding" experiments 117 (see also [20]). In this assay (Fig. 1), a highly diluted overnight culture of the resistant strain is 118 inoculated into fresh media in a large number of replicate cultures. The high dilution factors yield 119 average inoculum sizes of <1 to ~3 cells per culture. Importantly, however, the actual number of 120 cells inoculated into each replicate culture is random, and can be described by a Poisson distribution 121 (SI Appendix, Fig. S1). One implication of this protocol is that many cultures are not inoculated with 122 any cells, while others receive more than one cell; our modelling approach will account for this 123 variation statistically. We inoculated parallel replicate cultures in streptomycin-free media and at a 124 range of streptomycin concentrations below the MIC of the resistant strain, denoted MIC<sub>R</sub>, as 125 measured using standard protocols [10] (SI Appendix, Table S1). We then scored the number of 126 replicate cultures showing growth based on reaching a threshold optical density (OD) of 0.1 within 3 127 days post-inoculation.

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129 A culture could fail to grow either because the inoculum did not contain any cells, or because every 130 cell in the inoculum failed to give rise to a surviving lineage. To infer the probability that a single cell 131 yields detectable population growth (i.e. the per-cell establishment probability), we fit a 132 mathematical model, accounting for both the random inoculum size and demographic stochasticity, 133 to the observed number of replicate cultures showing growth (Materials and Methods). All 134 probabilities are normalized by the result in streptomycin-free media, which corresponds to scaling 135 inoculum size by the mean number of cells that establish in benign conditions (which we call the 136 "effective" inoculum size). Thus, relative establishment probability  $\tilde{p}_c$  equals one by definition in 137 streptomycin-free conditions, while we expect  $\tilde{p}_c \leq 1$  with streptomycin treatment; however, 138 values larger than one can arise due to sampling error.

139

140 Our seeding experiments revealed that the probability of establishment of a single resistant cell

declines with increasing streptomycin concentration (Fig. 2 and SI Appendix, Table S2). While

142 exposure to the lowest tested concentrations of streptomycin (up to 1/32 x MIC<sub>R</sub>) had no detectable

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143	impact on establishment, 1/16 x MIC <sub>R</sub> was already sufficient for significant declines, to $\tilde{p}_c$ of 55-73%
144	(maximum likelihood estimates in two independent experiments). At 1/8 x MIC <sub>R</sub> , $\tilde{p}_c$ dropped to just
145	3-5%. These results suggest that a <i>de novo</i> resistant mutant would only rarely establish at antibiotic
146	concentrations that are well below its $MIC_{R}$ , i.e. within the MSW.
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149 MIC depends on inoculum size

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151 The frequent failure of the resistant strain to grow in our seeding experiments at concentrations well 152 below its MIC is, at face value, surprising. We hypothesized that these results could be explained by 153 the difference in inoculum size between these assays. Specifically, standard MIC values are assessed 154 from an inoculation density of 5x10<sup>5</sup> CFU/mL [10], which corresponds to an inoculum size of 10<sup>5</sup> CFU 155 per 200µl culture on our microtitre plates. In contrast, our seeding experiments used an inoculum 156 size on the order of 1 CFU per culture. MIC for many antibiotics has been observed to increase with 157 higher-than-standard inoculation densities (CFU/ml) [22, 23, 24] which corresponds to higher 158 absolute inoculum size (CFU) for a fixed culture volume. Although less well-explored, it has also 159 occasionally been noted that MIC can decrease when lower absolute inoculum sizes are used [25, 160 26].

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162 To test the hypothesis that inoculum size influences MIC in the present system, we conducted a 163 modified MIC assay using the PA01:Rms149 strain with inoculum sizes ranging over three orders of 164 magnitude, from approximately  $10^2$  to  $10^5$  CFU per culture (corresponding to inoculation densities of 165  $5x10^2$  up to the standard  $5x10^5$  CFU/ml). We found that MIC indeed increases with inoculum size 166 (Fig. 3a). This pattern arises regardless of whether growth is scored at 20h, as per the standard MIC 167 assay protocol [10], or up to 3d post-inoculation, as in our seeding experiments, although the 168 number of cultures showing detectable growth, and thus the measured MIC, tends to increase over 169 time (SI Appendix, Fig. S2).

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Since all cultures contained the same volume in the above experiment, this pattern could be due to changes either in absolute inoculum size (i.e. CFU) or in inoculation density (i.e. CFU per unit volume). These two possibilities are not typically distinguished in the literature; however, they lead to distinct interpretations. If demographic stochasticity is the dominant force, we expect absolute numbers to matter, whereas if interactions among cells (e.g. competition or cooperation) affect establishment, cell density per unit volume could be more important. To disentangle these two

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177 factors, we repeated the MIC assay co-varying inocu	ulation density and culture volume. This
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- 178 experiment confirmed that absolute inoculum size has a strong effect on MIC. In contrast,
- inoculation density per unit volume does not have a significant effect within the range that we
- 180 tested, after controlling for absolute cell numbers (SI Appendix, Fig. S3).
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- 182

### 183 Population growth can be explained by an independent chance of each cell to establish

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185 Taken together, our seeding experiments and MIC assays reveal that the absolute number of cells in 186 the inoculum has a strong effect on whether the culture eventually shows detectable growth. The 187 simplest explanation for this result is that population growth can be attributed to the stochastic 188 outgrowth of one or more lineages, each initiated by a single cell in the inoculum, acting 189 independently. This independence assumption yields a "null model" that mathematically describes 190 the effect of inoculum size on the probability of outgrowth of a detectable population (Materials 191 and Methods, Eqn. 1). Here the probability of establishment of each cell in the inoculum ( $\tilde{p}_c$ ) is a 192 scaling parameter, which does not depend on inoculum size. Note that this null model would not 193 hold if interactions among cells substantially influenced their chances of successful replication. For 194 example, if cells secrete an enzyme that breaks down an antibiotic extracellularly, then the 195 establishment probability of each cell could increase with inoculum size. On the other hand, if cells 196 compete for limiting resources or secrete toxins, the per-cell establishment probability could 197 decrease with inoculum size. 198

To formally test the null model, we again conducted seeding experiments with the PA01:Rms149 strain, but now using many different inoculum sizes, spanning approximately three orders of magnitude. We tested two streptomycin concentrations (1/16 and 1/8 x MIC<sub>R</sub>) for which growth often failed from a single cell, but succeeded from standard inoculum size in MIC assays. In parallel, we tested growth in streptomycin-free media in order to estimate the effective mean inoculum size (SI Appendix, Fig. S4). This left one free parameter, the per-cell relative establishment probability  $(\tilde{p}_c)$ , to fit at each streptomycin concentration.

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We found good agreement between the null model and our experimental data at all tested
streptomycin concentrations, consistent with the hypothesis that cells establish independently
(main experiment, Fig. 3b, and repeat experiments, SI Appendix, Fig. S5). More precisely, the null
model did not show significant deviance from the observed proportion of populations that grew

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211	(according to the likelihood ratio test), and thus we accept it as a parsimonious explanation for the
212	data. Furthermore, we obtain estimates of relative establishment probability, $\tilde{p}_c$ , at 1/16 and 1/8 x
213	MIC <sub>R</sub> similar to those from the previous seeding experiments (SI Appendix, Table S2).
214	
215	To summarize, the probability of culture growth at any given streptomycin concentration depends
216	on inoculum size, according to a simple quantitative relationship. Our experimental data are
217	consistent with a simple model of cells behaving independently, such that a fixed per-cell
218	establishment probability can explain our growth data across inoculum sizes. That is, cells are not
219	"more susceptible" to streptomycin at lower inoculum sizes, but rather, culture growth is less likely
220	to be observed simply because fewer cells are available to establish, and not all cells succeed. In
221	turn, the minimal concentration of streptomycin required to prevent growth in some proportion of
222	replicate cultures (i.e. the observed MIC) increases with inoculum size.
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225	Sub-MIC <sub>R</sub> streptomycin concentrations induce resistant cell death and extend lag phase
226	
227	We hypothesized that resistant cells sometimes failed to establish in our seeding experiments
228	because exposure to streptomycin compromised cell division rate and/or viability. As a simple test
229	of this idea, we measured the relative abundance of dead cells in cultures of the resistant strain
230	grown at sub-MIC $_{R}$ concentrations of streptomycin. We found that the fraction of dead cells after 7h
231	of treatment, as determined by propidium iodide staining, increased from an average of 3-4% in
232	streptomycin-free conditions to >20% at 1/8 x MIC <sub>R</sub> streptomycin ( <b>Fig. 4a</b> and <b>SI Appendix, Fig. S6</b>
233	and <b>Table S3</b> ). Note that this is a conservative measure of cell death, because this assay only detects
234	cells that have compromised membrane permeability, and not, for example, cells that have already
235	lysed. Furthermore, this assay provides only a snapshot in time.
236	
237	To gain further insight into how sub-MIC $_{ extsf{R}}$ streptomycin impacts the population dynamics of the
238	resistant strain, we quantified viable cell density over the first few hours after inoculation into
239	streptomycin-containing media. Cultures were inoculated with approximately 100 cells in this
240	experiment, to ensure that cell numbers were low enough for demographic stochasticity to be
241	relevant, yet large enough to be detectable using conventional plating methods.
242	
243	We found that streptomycin treatment has a significant effect on the growth of resistant cultures
244	(ANOVA, main effect: $p < 2e-16$ ), and this effect varies over time (ANOVA, interaction term: $p < 2e-16$

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245	16; Fig. 4b and SI Appendix, Fig. S7). Following inoculation, cultures exhibited a lag phase of
246	approximately 2 hours. Control cultures in streptomycin-free media then began to grow
247	exponentially. The lowest tested concentration of streptomycin (1/32 x $MIC_{R}$ ) had no significant
248	effect on these dynamics (Dunnett's test: $p=0.87$ ); however, $1/16 \times MIC_R$ was already sufficient to
249	slow growth (p=4e-4). Nonetheless, all replicate cultures (n=48 per concentration) eventually grew,
250	as detected by OD. Meanwhile, higher doses of streptomycin (1/8 x or 1/4 x $MIC_R$ ) had dramatic
251	effects on growth dynamics ( $p < 1e-4$ ), with cultures exhibiting an extended lag phase of at least 7-8
252	hours, in which viable cell density initially declined. After further incubation (up to 3 days), 25% of
253	cultures (15/60) exposed to $1/8 \times MIC_R$ eventually showed growth, while the remaining 75% (45/60)
254	failed to reach detectable OD. At 1/4 x $MIC_{R}$ , no viable cells were detected in most cultures from 4h
255	on, and only 1/60 cultures reached detectable OD within 3 days.
256	
257	In summary, sub-MIC $_{R}$ streptomycin treatment has the effect of extending the lag phase, before
258	cultures eventually either grow to saturation or die out. Failure to grow can be explained by
259	significantly elevated cell death rates beginning at $1/16 \times MIC_R$ , which can lead to stochastic loss of
260	initially small populations.
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278 resistant cells is not unique to our primary model system of PA01:Rms149 in streptomycin.

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281	The sensitive population modulates probability of establishment of resistant cells
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283	So far, we focused on the direct effects of antibiotics on resistant cells by conducting experiments
284	with monocultures of resistant strains. However, de novo resistance will actually arise within a
285	sensitive population, by mutation or transfer of a mobile genetic element into a sensitive cell.
286	Moreover, antibiotic treatment will only begin in clinical settings once the total pathogen population
287	is large enough to cause symptoms. We therefore asked whether the presence of a large sensitive
288	population affects the establishment of initially rare resistant cells during antibiotic treatment,
289	returning to the interaction between PA01 (sensitive) and PA01:Rms149 (streptomycin-resistant) as
290	a model system.
291	
292	We expect the sensitive population and the antibiotic to have interacting effects on establishment of
293	resistance. In particular, at sufficiently low antibiotic concentrations, a sensitive strain is generally
294	expected to outcompete a resistant strain due to the fitness cost associated with resistance [13, 7,
295	8]. We confirmed this expectation in our experimental system using a standard competition assay,
296	where both strains start from reasonably large inoculum sizes (SI Appendix, Fig. S8-S10 and Table
297	<b>S5</b> ). We found that the sensitive strain is favoured up to a minimum selective concentration (MSC)
298	between 1-2 $\mu$ g/ml streptomycin (equivalent to 1/32 – 1/16 x MIC <sub>s</sub> , or 1/2048 – 1/1024 x MIC <sub>R</sub> ), in
299	agreement with previous results for these strains [32]. We hypothesized that competition from the
300	sensitive strain would prevent establishment of resistance at streptomycin concentrations below the
301	MSC.
302	
303	As a simple test of this idea, we modified the seeding experiment to inoculate very few resistant
304	cells into a large sensitive population. Since bacterial densities in clinical infections can vary widely
305	[9, 33], we inoculated the sensitive strain at two different densities: approximately 5 x $10^5$ CFU/ml
306	(as in a standard MIC assay; labelled "low") and $5 \times 10^7$ CFU/ml (labelled "high"). The resistant strain
307	was seeded, with mean inoculum size on the order of one cell per culture, immediately thereafter.
308	
309	As hypothesized, we found that the presence of the sensitive population (at either density)
310	abolished establishment of resistant cells in the absence of streptomycin (Fig. 6 and SI Appendix,
311	Fig. S11 and Table S6). Meanwhile, at streptomycin concentrations above the MSC (1/256 to 1/8 x

312 MIC<sub>R</sub>, or 1/4 to 8 x MIC<sub>s</sub>), adding the sensitive population at low density had a negligible effect on

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313 the probability of establishment of resistant cells. At high density, the sensitive population also had 314 negligible effects on establishment of resistance at streptomycin concentrations up to  $1/16 \times MIC_{R}$  (4) 315 x MIC<sub>s</sub>). However, at  $1/8 \times MIC_{R}$  (8 x MIC<sub>s</sub>), the presence of a high-density sensitive population 316 increased the establishment probability from near zero to 65%. To confirm and further probe the 317 extent of this apparent protective effect, we repeated the experiment over a higher range of 318 streptomycin concentrations. The boost in establishment probability was repeatable and highly 319 significant at  $1/8 \times MIC_R$  (Wilcoxon rank-sum test, high- vs. zero or low-density sensitive: p < 5e-8 in 320 both experiments). However, at  $1/4 \times MIC_R$  (16 x MIC<sub>s</sub>), an apparent slight boost in establishment 321 probability was non-significant, and by  $1/2 \times MIC_R$  (32 x MIC<sub>s</sub>) the effect was abolished. Thus, a 322 sufficiently dense sensitive population can extend the range of streptomycin concentrations at 323 which the resistant strain is likely to emerge, but does not change the qualitative pattern of 324 stochastic establishment. 325

326

# 327 Discussion

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In order for resistance to emerge *de novo*, not only must a resistance gene arise in a bacterial population by mutation or horizontal gene transfer; this first resistant cell must also successfully expand to form a large population. Since any individual cell may fail to replicate, particularly in challenging environmental conditions, the expansion of newly arisen resistant strains is not guaranteed. Our key finding is that demographic stochasticity imposes a significant barrier to the emergence of resistance in the presence of antibiotics at concentrations within the mutant selection window.

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337 We empirically demonstrated the importance of stochasticity with a simple "seeding experiment" 338 mimicking the growth of clonal resistant lineages founded by a single cell. First, to assess the direct 339 impact of antibiotics, we inoculated fresh antibiotic-containing media with approximately one 340 resistant cell per replicate culture and quantified the per-cell probability of establishing a detectable 341 population. Strikingly, this establishment probability dropped off at concentrations well below the MIC of the corresponding resistant strain (MIC<sub>R</sub>). For example, the establishment probability of 342 343 PA01:Rms149 was significantly reduced by streptomycin concentrations as low as  $1/16 \times MIC_{R}$ , and 344 dropped to <5% at 1/8 x MIC<sub>R</sub> (Fig. 2). Resistant cells failed to establish viable populations because of 345 the toxic effects of exposure to sub-MIC<sub>R</sub> concentrations of antibiotics (Fig. 4) coupled with the 346 inherently stochastic nature of individual cell death and division. Importantly, we were able to

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replicate our key finding of frequent stochastic loss using a different, meropenem-resistant strain
(PA01:PAMBL2; Fig. 5). This demonstrated that our results are not limited to a particular model
system, but are also relevant to bacterial pathogens of clinical concern: carbapenem-resistant *P. aeruginosa* is considered by the WHO to be of "critical priority" for antibiotic development [31].

352 In clinical settings, antibiotic treatment will typically begin only when the total bacterial population is 353 large enough to cause symptoms. Assuming resistance has not been transmitted, this population will 354 be predominantly antibiotic-sensitive, and de novo mutation or acquisition of a mobile element 355 conferring resistance will occur within a sensitive cell. We therefore next asked how the presence of 356 a large sensitive population would combine with the above effects of antibiotics to shape the 357 emergence of resistance from this first cell, again in the streptomycin model system. As predicted 358 from standard competition assays, emergence of resistance was abolished in the absence of 359 antibiotics (Fig. 6), presumably due to competitive suppression by the sensitive strain [8]. More 360 interestingly, a sufficiently dense sensitive population (inoculated at  $\sim 5 \times 10^7$  CFU/ml here) was able 361 to shift the range of concentrations at which resistance established upwards by approximately two-362 fold. We speculate that this apparent protection is due to sensitive cells absorbing antibiotics, thus 363 lowering their concentration in the media [23, 34], despite these concentrations being high enough 364 to cause decline of the sensitive population (i.e. > MIC<sub>s</sub>). A priori, one may not expect resistant cells 365 to "need" protection at sub-MIC<sub>R</sub> antibiotic concentrations. However, in the stochastic regime of 366 establishment, any increase in the probability of individual cells surviving and dividing can make a 367 qualitative difference to the fate of a rare resistant lineage. This protective effect depends on 368 bacterial density at the time of treatment, within a realistic range for some bacterial infections [9, 369 33]: when the sensitive population was inoculated at 100-fold lower density ( $^{5} \times 10^{5}$  CFU/ml), 370 protection was not apparent. We emphasize that although these experiments provide an initial 371 proof of concept, a complete investigation of the interacting effects of sensitive population density, 372 antibiotic dose and timing remains an important direction for future work. Importantly, however, 373 our main message continues to hold even in the more realistic context of resistance arising within a 374 sensitive population: stochastic loss of resistant cells is frequent at antibiotic concentrations within 375 the MSW.

376

The failure of resistant cells to establish successful lineages at concentrations well below the MIC<sub>R</sub>
shows a clear disconnect between antibiotic susceptibility of individual cells and populations. To
explain this effect rigorously, we quantified the probability of outgrowth of a detectable population
at a fixed streptomycin concentration, starting from inoculum sizes spanning three orders of

#### Stochastic bacterial population dynamics

13

381 magnitude. We fit these data to a mathematical model relating inoculum size to probability of 382 population growth, under the hypothesis that each cell in the inoculum behaves independently (Eqn. 383 **1**). This simple stochastic model, with a constant per-cell probability of establishment  $(\tilde{p}_c)$ , provides 384 a good explanation for inoculum size-dependent population growth in PA01:Rms149 (Fig. 3b). In this 385 model, individual cells are not "more susceptible" to antibiotic in smaller populations. Instead, the 386 cumulative effect of many cells, each with a small chance of establishment (e.g. <5% at  $1/8 \times MIC_{R}$  in 387 this system), virtually guarantees population growth from a sufficiently large inoculum size, 388 reconciling our results with the standard definition of the MIC. We thus emphasize that MIC is not 389 an innate property of a cell or strain, but rather an emergent property of a population of cells. We 390 also note that the inoculum size effect on MIC that we found here – a purely stochastic phenomenon 391 arising at low absolute numbers (CFU) – is distinct from the inoculum size effect already widely 392 recognized in the literature, which is seen at high cell density (CFU/ml) and attributed to various 393 density-dependent mechanisms, such as titration or enzymatic inactivation of antibiotics [35, 36, 24, 394 37, 32, 23, 38]. Although there are hints of the former absolute-number effect in earlier studies [25, 395 26], to our knowledge we are the first to provide a rigorous explanation in terms of stochastic

- 396 population dynamics.
- 397

398 The good fit of the null model, in which every cell has an equal probability of establishment, at first 399 seems to be at odds with the growing recognition that bacteria exhibit phenotypic heterogeneity, 400 which could affect individual cells' susceptibility to antibiotics [39, 40]. Indeed, it is entirely possible 401 that the resistant cells that successfully established in our experiments were those with a particular 402 metabolic state or gene expression level. However, this variability among cells would have no effect 403 on our experimental outcomes: under our mathematical model (Eqn. 1), the probability of observing 404 growth in a given number of replicate cultures is the same for any degree of cell-to-cell variation 405 around a fixed mean, assuming the susceptibilities of cells within an inoculum are independent from 406 one another (Suppl. Text, section 10.1). Thus, the establishment probability that we infer empirically 407 should more accurately be interpreted as a mean among cells.

408

Although the role of demographic stochasticity in the fate of *de novo* mutations has long been
recognized in theoretical population genetics, until very recently it had never been addressed
empirically [17]. Our study joins a small handful of others that have now experimentally quantified
establishment probability from single cells [41, 18, 42, 19], including two [18, 19] investigating
establishment of bacterial cells in the presence of antibiotics, using different methods to ours (see
Suppl. Text, section 10.2, for a more detailed comparison). Bacterial evolution of resistance to

#### Stochastic bacterial population dynamics

14

415 antibiotic treatment is also a prime example of the more general phenomenon of evolutionary 416 rescue, whereby adaptation prevents extinction of populations facing severe environmental change 417 [43]. Experiments quantifying the probability of rescue in initially large, but declining, populations 418 were pioneered around a decade ago in yeast exposed to high salt concentrations [44]. More recent 419 work clarified the quantitative relationship between initial population size and probability of rescue 420 [45], coinciding with our Eqn. 1 (see Suppl. Text, section 10.2, for further discussion). More broadly, 421 the concepts and statistical methods developed here are applicable to a variety of situations where 422 growth depends on success of rare cells and is thus highly stochastic, for instance the establishment 423 of productive infection in a host following pathogen transmission [46], the onset of invasive bacterial 424 infections [47], the outgrowth of bacteria in food products from small initial contaminants [48], or 425 the establishment of metastases from cancerous tumours [49]. 426

427 In summary, our study highlights the stochastic nature of *de novo* emergence of antibiotic

428 resistance. In a practical sense, this stochasticity implies that to accurately assess the risk of

429 resistance emerging, we must evaluate not only mutation rates, but also the probability that

430 resistant mutants escape extinction when rare [9], which will depend on the antibiotic dosing

regimen. Our results caution against the naïve use of the mutant selection window (MSW) for this

432 purpose. While a positive selection coefficient is a necessary condition for resistance to outcompete

433 an initially prevalent sensitive strain, it does not guarantee emergence when rare; indeed, we

434 showed that single resistant cells are frequently lost at antibiotic concentrations well within the

435 MSW. Thus, our findings suggest that moderate antibiotic doses may be more effective than

436 previously thought at preventing *de novo* emergence of resistance, especially in infections where

total pathogen density is relatively low. For antibiotics that have mutagenic effects, the chance of a

- resistant lineage arising in the first place might also be reduced at lower doses ([50]; see however
- 439 [51]). Furthermore, use of lower antibiotic doses could reduce both adverse effects on patients and440 release of antibiotics into the environment [52].
- 441
- 442

# 443 Materials and Methods

444

Further details of experimental protocols, data processing, mathematical models and statistical
methods are provided in the SI Appendix, Suppl. Text.

447

448 Bacterial strains, media, and culture conditions

#### Stochastic bacterial population dynamics

15

#### 449

450 Bacterial strains: The majority of our experiments, in streptomycin, were conducted with a set of 451 Pseudomonas aeruginosa PA01 strains studied previously [32]. The streptomycin-sensitive and -452 resistant strains are chromosomally isogenic, while resistant strains additionally carry the clinically 453 derived, non-conjugative plasmid Rms149 [53], which is stably maintained in PA01 at approximately 454 two copies per cell [54]. Streptomycin resistance is conferred by the aadA5 gene on Rms149, which 455 codes for an enzyme that adenylates streptomycin [55]. Both plasmid carriers (resistant) and non-456 carriers (sensitive) are available with either YFP or DsRed chromosomal fluorescent markers or with 457 no marker [32]. The live-dead staining experiment was conducted with the unlabelled resistant 458 strain. All other experiments reported in the main text were conducted with the YFP-labelled 459 resistant strain and, where applicable, the DsRed-labelled sensitive strain. We chose this pairing 460 because YFP provides a stronger signal, facilitating detection of the resistant strain in mixed cultures. 461 Previous work with these strains suggests that the two fluorescent labels have similar fitness effects 462 [32], and we confirmed that the label had no substantive effect on the MIC values of the sensitive 463 strain (**Suppl. Text**, section 2.1). For the seeding experiment in meropenem, we transformed the 464 plasmid PAMBL2 into the same PA01-YFP background (Suppl. Text, section 1). This plasmid, isolated 465 in 2007 from a patient in a Spanish hospital [27], confers meropenem resistance through three 466 copies of the *bla<sub>VIM-1</sub>* gene, which codes for a metallo-beta-lactamase [27, 28]. It is non-conjugative 467 [28] and stably maintained in PA01 at an average of 2-3 copies/cell [54]. MIC values of all relevant 468 strain-antibiotic pairs are reported in Table S1.

469

470 Media and antibiotics: We cultured bacteria in LB broth containing 5g/L NaCl (Sigma-Aldrich, product 471 no. L3022). To assess colony-forming units, we plated on LB Agar, Vegitone, containing 5g/L NaCl 472 and 15g/L agar (Sigma-Aldrich, product no. 19344). Streptomycin was prepared from streptomycin 473 sulfate salt (Sigma-Aldrich, product no. S6501) and meropenem was prepared from meropenem 474 trihydrate (Santa Cruz Biotechnology, Inc., product no. SC-485799). Stocks prepared in water were 475 stored according to supplier directions and added to media on the day of experiments. When high 476 antibiotic concentrations were required, stocks were instead prepared directly in LB on the day of 477 experiments to avoid excessive dilution of the media with water. Bacterial cultures were diluted in 478 phosphate buffered saline (PBS) prepared from tablets (Sigma-Aldrich product no. P4417). 479 Treatment cultures were set up with 90% media plus 10% inoculating culture by volume; thus, the 480 final concentrations of LB and antibiotics in the treatments are 90% of the prepared media values

denoted on plots.

#### Stochastic bacterial population dynamics

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483 <u>Culture conditions</u>: All cultures were incubated at 37°C, shaking at 225rpm. Overnight cultures were

- inoculated directly from freezer stocks into 2ml of LB in 14ml culture tubes and incubated for
- 485 approximately 16h. Overnight cultures were then diluted in PBS and used to inoculate treatment
- 486 plates. Unless otherwise noted, experimental treatments were conducted in 200µl cultures in flat-
- 487 bottom 96-well microtitre plates.
- 488

489 <u>Scoring culture growth</u>: In all experiments, we evaluated culture growth by measuring optical

- density (OD<sub>595</sub>) using a BioTek Synergy 2 plate reader, at room temperature. Lids on microtitre plates
- 491 were briefly removed in a non-sterile environment for the reading; comparison to controls mock-
- 492 inoculated with PBS indicated that contamination was rare (see **Suppl. Text** for detailed
- 493 quantification in each experiment). We set a threshold of OD<sub>595</sub> > 0.1 to score as growth, whereas
- 494 background OD in media-only controls was typically below 0.05. Final readings at 3d post-inoculation
- 495 were used for data analysis unless otherwise noted. By this time, growth had typically stabilized,
- 496 with OD much higher than the threshold.
- 497

#### 498 MIC assays

499

500 Standard MIC values for all applicable strain-antibiotic pairs (i.e. resistant Rms149-carrier against 501 streptomycin; resistant PAMBL2-carrier against meropenem; sensitive non-carrier against both 502 antibiotics) were determined under our culture conditions using the broth microdilution method. 503 Overnight cultures were diluted 10<sup>3</sup>-fold and inoculated into antibiotic-containing media at 504 20µl/well on 96-well test plates. This dilution factor consistently yielded an inoculation density close 505 to 5 x  $10^5$  CFU/ml, in accordance with standard protocol [10]; actual density was estimated by 506 plating. Test plates were incubated and scored for growth at approximately 20h (as per standard 507 protocol [10]), 2d, and 3d post-inoculation. For consistency with growth scoring in seeding 508 experiments, the standard MIC values (MICs and MICR) used to scale antibiotic concentrations on 509 plot axes are based on results at 3d. Consensus MIC values of all tested strain-antibiotic pairs, at 510 both 20h and 3d, are reported in Table S1, with results of individual replicates reported in the Suppl. 511 Text, section 2.1. For the YFP-labelled Rms149-carrying resistant strain, an additional MIC assay in 512 streptomycin was conducted varying inoculum size (Fig. 3a). Here, inoculations were conducted with 513 overnight culture diluted 10<sup>3</sup>-, 10<sup>4</sup>-, 10<sup>5</sup>-, and 10<sup>6</sup>-fold (see **Suppl. Text**, section 2.2 for details). 514

#### 515 Seeding experiments: resistant strains in isolation

#### Stochastic bacterial population dynamics

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517 Experimental protocol: A highly diluted overnight culture of the YFP-labelled resistant strain 518 (Rms149- or PAMBL2-carrier) was inoculated at 20µl/well into antibiotic-containing media on 96-519 well test plates. For experiments with PA01:Rms149 screening across many streptomycin 520 concentrations (Fig. 2), we used three dilution factors ( $4 \times 10^7$ ,  $8 \times 10^7$ , and  $1.6 \times 10^8$ -fold), each to 521 inoculate 96 replicate wells at each concentration. To test the null model of the inoculum size effect 522 (Fig. 4b), we screened fewer streptomycin concentrations across a larger number of dilution factors 523 (five in streptomycin-free conditions and six to ten in each streptomycin concentration), each with 524 54 replicates. These dilution factors were chosen differently for each streptomycin concentration to 525 capture the range over which the proportion of replicate cultures showing growth increased from 526 near 0 to near 1. For the experiment with PA01:PAMBL2 in meropenem (Fig. 5), we used two 527 dilution factors (5 x  $10^{7}$ - and 2 x  $10^{8}$ -fold), each with 96 replicates per concentration. In all cases, test 528 plates were incubated and scored for growth after approximately 1, 2, and 3d; for the null model 529 test, incubation and readings were continued up to 5d to confirm stabilization of growth. See Suppl. 530 **Text**, sections 4-5, for further details.

531

532 Model fitting: The number of replicate cultures showing growth by 3d (or, additionally, by 5d for the 533 null model test), at each inoculating dilution factor and antibiotic concentration, was used for 534 subsequent model fitting. To estimate single-cell establishment probability and evaluate the null 535 model of the inoculum size effect, likelihood-based methods were used to fit a stochastic model of 536 population growth to these data (see *Mathematical model of establishment* below). In addition, to 537 evaluate the effect of antibiotic concentration on establishment, generalized linear models were fit 538 to data from the seeding experiments screening across streptomycin (Fig. 2) or meropenem (Fig. 5) concentrations. Using the built-in R function 'glm', growth data were treated as binomial, with 539 540 inoculating dilution factor and antibiotic concentration taken as explanatory variables, applying the 541 complementary log-log link function (Suppl. Text, section 12).

542

### 543 Seeding experiments: resistant strain in presence of sensitive population

544

545 Overnight culture of the DsRed-labelled PA01 sensitive strain was diluted 5-fold to obtain the "high 546 density" inoculating culture, and (in the first experiment only) further to 500-fold to obtain the "low 547 density" inoculating culture. Overnight culture of the YFP-labelled PA01:Rms149 resistant strain was 548 diluted up to  $5 \times 10^7$ -fold and  $2 \times 10^8$ -fold. These cultures were inoculated as follows into media at 549 various streptomycin concentrations on 96-well plates. Pure sensitive control cultures (24 replicates 550 per test condition) were inoculated with  $10\mu$ /well of the appropriate diluted culture plus  $10\mu$ /well

#### Stochastic bacterial population dynamics

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551 PBS. "Blank" wells to serve as background fluorescence controls were inoculated with 20µl/well PBS. 552 Seeding test plates were first inoculated with  $10\mu$ /well of either PBS (for pure resistant control 553 cultures), low-density or high-density sensitive culture. The resistant strain was inoculated at 554 10µl/well immediately thereafter (all sensitive and resistant culture inoculations were completed 555 within an hour). Seeding was conducted with 30-60 replicates per test condition and resistant 556 dilution factor (see **Suppl. Text**, section 6, for details). All test plates were then incubated as before, 557 with optical density (OD<sub>595</sub>) and fluorescence (excitation: 500+/-27 nm; emission: 540+/-25 nm) 558 measured at approximately 1, 2, and 3d post-inoculation. Among wells showing growth (OD>0.1), 559 we considered the YFP-labelled resistant strain to have established if fluorescence exceeded 5 x 10<sup>5</sup> 560 units, chosen by comparison to pure cultures. In each test condition, the number of replicates in 561 which resistance established was taken as data for model fitting, as in the previous seeding 562 experiments.

563

### 564 Fraction of dead cells by live-dead staining

Viable cell density dynamics

565

566 This experiment used the PA01:Rms149 resistant strain with no fluorescent label, to avoid 567 interfering with the signal from the stains. We inoculated streptomycin treatment cultures (six 568 replicates per concentration) with  $10^3$ -fold diluted overnight culture, as in the standard MIC assay. 569 After 7h of treatment, we diluted test cultures 100-fold and stained with thiazole orange and 570 propidium iodide (BD Cell Viability Kit, product no. 349483). In parallel, we diluted and stained 571 media and heat-killed cultures as controls. We sampled 50µl per diluted culture using flow 572 cytometry (BD Accuri C6 Flow Cytometer with fast fluidics, discarding events with forward scatter 573 FSC-H < 10000 or side scatter SSC-H < 8000). The staining and flow cytometry steps were carried out 574 in groups containing one replicate per concentration plus controls, to avoid potentially toxic effects 575 of stain exposure over prolonged times (Suppl. Text, section 7). To better discriminate cells from 576 background in the flow cytometry data, we first gated on events according to forward and side 577 scatter before defining clusters of dead (membrane-compromised) and intact cells based on 578 fluorescence; see Suppl. Text, section 7, and Suppl. Fig. 6 for details. 579

580 581

582 Using the YFP-labelled PA01:Rms149 strain, we tracked the number of viable cells over time in

- 583 streptomycin-free media (twelve replicates per time point) and at 1/32, 1/16, 1/8, and 1/4 x MIC<sub>R</sub>
- 584 streptomycin (six replicates per time point). An independent test plate was used for sampling at

#### Stochastic bacterial population dynamics

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each time point. Lower (Set A) and higher (Set B) streptomycin concentrations were split across separate plates and sampled at different times. Cultures were inoculated with  $20\mu$ l of 5 x  $10^5$ -fold diluted overnight culture. At each sampling time, we plated 5 x  $4\mu$ l spots of undiluted cultures (10%sampling by volume). The number of viable cells was estimated from total colony count following incubation. Comparison of streptomycin-free controls from both sets (A and B) indicated that the plate set effect was non-significant (ANOVA: p = 0.10); thus, controls were pooled for further analysis of the streptomycin effect (see **Suppl. Text**, section 8, for further details).

- 592
- 593 Mathematical model of establishment
- 594

595 <u>Model</u>: We denote by  $p_w$  the probability that a small number of inoculated cells grows into a large 596 population, i.e. that the culture reaches detectable OD as described above. Among a set of *n* 597 independent replicates, the number of cultures showing growth is thus described by a 598 Binomial( $n, p_w$ ) distribution.

599

In the "null" model, similar to previous work [46, 45], a simple expression for  $p_w$  is derived under the assumptions that: (i) the number of cells in the inoculum is Poisson-distributed with mean  $\overline{N}$ ; (ii) each cell, *independently*, establishes a surviving lineage with probability  $p_c$ , which depends only on antibiotic concentration x; and (iii) culture growth is observed provided at least one cell establishes a surviving lineage. Then the probability of observing culture growth, as a function of mean inoculum size and antibiotic concentration, can be written as follows (**Suppl. Text**, section 10):

- 606
- 607

608

 $p_w(\overline{N}, x) = 1 - e^{-\overline{N}p_c(x)}$  $= 1 - e^{-\overline{N}e_{ff}\tilde{p}_c(x)}$ (Eqn. 1)

609

610 In the second line, we have rewritten the expression in terms of the "effective mean inoculum size", 611  $\overline{N}_{eff} = \overline{N}p_c(0)$ , which is the mean number of established lineages in the absence of antibiotics; and 612 the "relative establishment probability",  $\tilde{p}_c(x) = p_c(x)/p_c(0)$ . Although we expect that  $p_c(0)$  is 613 close to 1,  $\overline{N}$  and  $p_c(0)$  play indistinguishable roles in this model, so that in practice we can only 614 estimate their product. This definition of effective inoculum size based on cells that grow in benign 615 conditions is similar to the usual quantification of "viable" cells according to successful formation of a colony; we simply assess growth in liquid rather than on solid medium. Scaling up  $\overline{N}_{eff}$  by the 616 617 dilution factor applied to the inoculating culture, we have an estimate of bacterial density in this 618 culture, equivalent to the historical "most probable number" method [56, 57]. If cells are

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phenotypically heterogeneous (i.e. vary in their propensity to establish), or if the individual units in
the inoculum are actually clumps of cells, then *p<sub>c</sub>* should be interpreted as the mean establishment
probability among individuals (**Suppl. Text**, section 10.1).

622

More generally, we need not assume that cells establish independently. If we suppose simply that
the number of established lineages is Poisson-distributed with some mean α (which is supported
empirically by the distribution of colony-forming units counted in highly diluted cultures; Suppl. Fig.
1), we have the relationship

- 627
- 628

$$p_w(\overline{N}, x) = 1 - e^{-\alpha(\overline{N}, x)}$$
(Eqn. 2)

629

630 where  $\alpha$ , and hence  $p_w$ , have an arbitrary dependence on mean inoculum size and antibiotic 631 concentration. In the statistical "full model", we estimate a distinct  $p_w$  (or equivalently  $\alpha$ , by the one-632 to-one mapping in **Eqn. 2**) in each test condition. Relative establishment probability is then generally 633 defined by  $\tilde{p}_c(\overline{N}, x) = \alpha(\overline{N}, x)/\alpha(\overline{N}, 0)$ . Nested models, including the null model above, make 634 additional assumptions about the form of  $\alpha$  (see **Suppl. Text**, section 10, for details).

635

636 Likelihood-based model fitting and comparisons: These stochastic models are fit to experimental 637 population growth data using likelihood-based methods (Suppl. Text, section 11). Specifically, under 638 each model we obtain a maximum likelihood estimate and a 95% confidence interval (determined by 639 the range of parameter values that would not be rejected by a likelihood ratio test at 5% significance 640 level) on the parameter  $p_w$ , which can be transformed to an estimate for  $\alpha$ . In the case of relative 641 establishment probability,  $\tilde{p}_c(x) = \alpha(x)/\alpha(0)$ , we use a profile likelihood confidence interval 642 accounting for the uncertainty in both numerator (i.e. results at antibiotic concentration x) and 643 denominator (i.e. results in antibiotic-free conditions). The fit of nested models is compared using 644 the likelihood ratio test (LRT) at 5% significance level, i.e. a  $\chi^2$  test on model deviance with degrees 645 of freedom equal to the difference in number of fitted parameters between the two models. 646 647 To test the null model of the inoculum size effect, we neglect any experimental error in preparing

648 overnight culture dilutions, and assume that mean inoculum size  $\overline{N}$  is inversely proportional to the 649 applicable dilution factor. Effective mean inoculum size,  $\overline{N}_{eff}$ , is estimated by fitting **Eqn. 1** to 650 population growth data in antibiotic-free media. Per-cell relative establishment probability  $\tilde{p}_c$  then 651 remains as the single free parameter to fit at each tested antibiotic concentration. The goodness of

652	fit of the null model (Eqn. 1) is assessed for each test concentration separately, using the LRT to
653	compare it to the fit of the full model ( <b>Eqn. 2</b> ).
654	
655	All model fitting was implemented in R, version 3.3.1 (The R Foundation for Statistical Computing,
656	2016).
657	
658	
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666	carried out experiments and data analysis with advice from R.C.M. H.K.A. developed the
667	mathematical model, wrote the code and carried out model fitting. H.K.A. and R.C.M. wrote the
668	manuscript.
669	
670	Data and code availability: Data generated in this study, as well as custom R scripts for likelihood-
671	based model fitting and comparisons, will be made available upon manuscript acceptance.
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673	
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Stochastic bacterial population dynamics

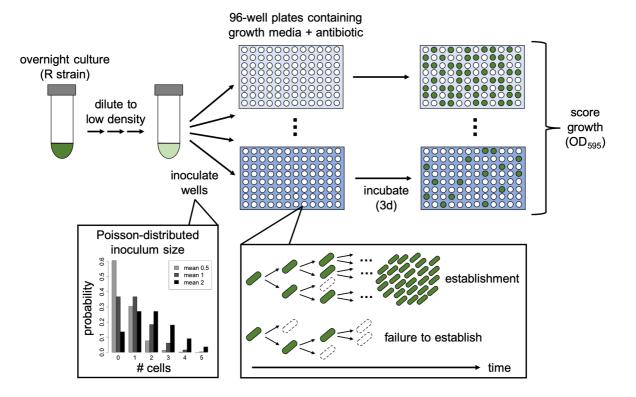
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#### Stochastic bacterial population dynamics

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# 678 Figures

## 679

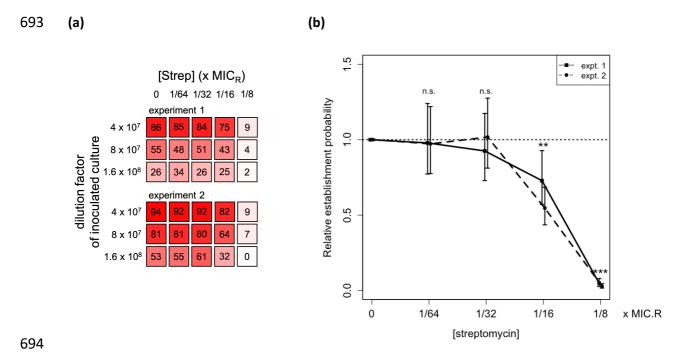


# 680 681

682 Figure 1: Design of seeding experiments to estimate establishment probability. An overnight 683 culture of the resistant strain is highly diluted and used to inoculate 96-well plates containing growth 684 media (LB broth) with antibiotic at various concentrations (shades of blue). The number of cells 685 inoculated per well follows a Poisson distribution (examples plotted for mean inoculum size of 0.5, 1, 686 or 2 cells per well). Within these culture wells, stochastic population dynamics imply that each 687 inoculated cell may either produce a large number of descendants (establishment) or produce 688 no/few descendants that ultimately die out (failure to establish). Plates are incubated for 3 days and 689 optical density is measured to score growth in wells (OD<sub>595</sub> > 0.1; dark green). The number of 690 replicate cultures showing growth is used to estimate the per-cell establishment probability at each 691 antibiotic concentration by fitting a mathematical model.

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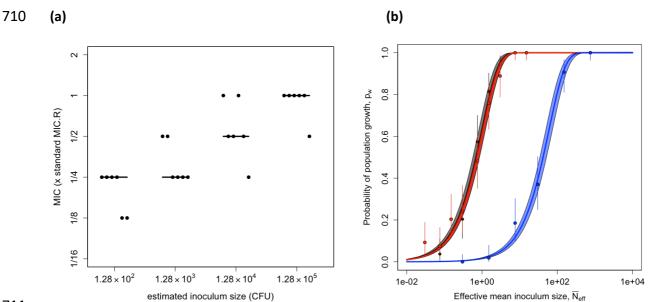
#### Stochastic bacterial population dynamics



695 Figure 2: Establishment probability of single PA01:Rms149 streptomycin-resistant cells, estimated 696 from seeding experiments. (a) Visual representation of the growth data, indicating the number of 697 replicate cultures (out of 96) that grew in each test condition up to 3d post-inoculation. (b) 698 Estimated relative per-cell establishment probability ( $\tilde{p}_c$ ), scaled by the probability in 699 streptomycin-free medium, as a function of streptomycin concentration, scaled by the standard MIC 700 value of the resistant strain ( $MIC_R = 2048 \mu g/ml$ ; **Table S1**). Results are shown for two separate 701 experiments. Plotted points indicate the maximum likelihood estimate of  $\tilde{p}_c$  and error bars indicate 702 the 95% confidence interval, using the fitted model selected by the likelihood ratio test (experiment 703 1: Model B', fixed environmental effect; experiment 2: Model C', the null model [Eqn. 1]. Both of 704 these models pool data across three inoculation densities; see Suppl. Text, section 10, for details). 705 Significance of the streptomycin effect is determined by fitting a generalized linear model to the 706 population growth data (n.s.: not significant, p > 0.05; \*\* p=0.01 in expt. 1, p=2e-7 in expt. 2, and p=2e-8 pooling both experiments; \*\*\* p < 2e-16 in both experiments; see **Suppl. Text**, section 14.1, 707 708 for full results).

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#### Stochastic bacterial population dynamics

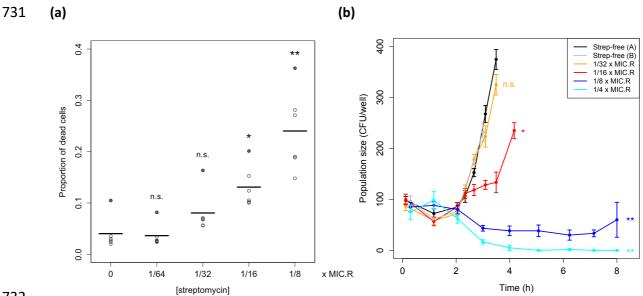




712 Figure 3: Inoculum size effects on MIC and probability of population growth of the resistant 713 PA01:Rms149 strain in streptomycin. (a) MIC as a function of inoculum size. Cultures were 714 inoculated with PA01:Rms149 at four different inoculum sizes. MIC was evaluated as the minimal 715 tested streptomycin concentration that prevented detectable growth up to 3d post-inoculation; a 716 qualitatively similar pattern arose if growth was evaluated at 20h (Fig. S2). The y-axis is scaled by the 717 MIC of this strain at standard inoculation density (MIC<sub>R</sub>). The points represent six replicates at each 718 inoculum size, with the line segments indicating their median. (b) Null model of the inoculum size 719 effect (Eqn. 1) fit to culture growth data. Probability of population growth  $(p_w)$  is plotted as a 720 function of effective mean inoculum size ( $N_{eff}$ , calibrated by the results in streptomycin-free media; 721 see Fig. S4). Black: streptomycin-free; red: streptomycin at  $1/16 \times MIC_R$ ; blue:  $1/8 \times MIC_R$ . These 722 results are based on growth in streptomycin up to 5d post-inoculation; see **Suppl. Text**, section 15, 723 for results at 3d post-inoculation. Points indicate the proportion of replicate cultures showing 724 growth, i.e. the maximum likelihood estimate (MLE) of  $p_w$  in the full model, with error bars indicating 725 the 95% confidence interval (CI). The solid line shows the best fit of the null model (i.e. Eqn. 1 726 parameterized with the MLE of  $\tilde{p}_c$ ) and the shaded area corresponds to the 95% CI. According to the 727 likelihood ratio test, the null model deviance from the full model is not significant at any 728 streptomycin concentration (streptomycin-free: p=0.55;  $1/16 \times MIC_R$ : p=0.28;  $1/8 \times MIC_R$ : p=0.71; see 729 Suppl. Text, section 15, for full results).

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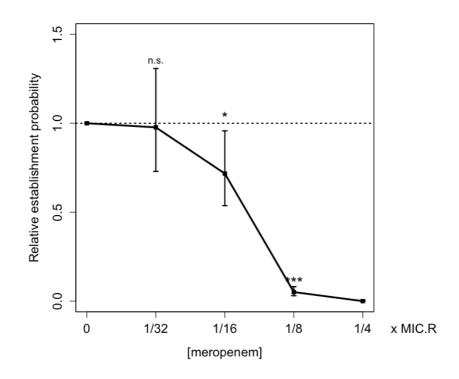
#### Stochastic bacterial population dynamics



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733 Figure 4: Effects of sub-MIC<sub>R</sub> streptomycin treatment on PA01:Rms149 resistant cell dynamics. (a) 734 Proportion of dead cells after 7h in sub-MIC<sub>R</sub> streptomycin. The proportion of dead cells in 735 streptomycin-treated cultures was estimated using live-dead staining and flow cytometry. Points 736 represent six independent treatment replicates at each concentration and line segments indicate 737 their mean. Differences from the streptomycin-free control cultures were assessed using a one-way 738 ANOVA followed by a post-hoc Dunnett's test (n.s.: not significant, p > 0.05; \*: p = 9e-3, \*\*: p < 1e-739 4). Effects identified as significant do not change if we exclude an outlier replicate (shaded-in points) 740 showing consistently elevated dead cell fractions (Table S3). (b) Viable cell population dynamics in 741 sub-MIC<sub>R</sub> streptomycin. Points with connecting lines indicate the mean number of viable cells across 742 six replicate cultures per streptomycin concentration, per sampling time point (or twelve replicates 743 for streptomycin-free controls); the error bars indicate standard error. Fig. S7 shows all individual 744 replicates. Viable cell numbers were estimated by plating undiluted culture samples; plots are 745 truncated when colonies became too dense to count. Significance of each streptomycin 746 concentration compared to the streptomycin-free control was assessed by a post-hoc Dunnett's test 747 (n.s.: not significant, *p*=0.87; \* *p*=4e-4; \*\* *p*<1e-4).

#### Stochastic bacterial population dynamics

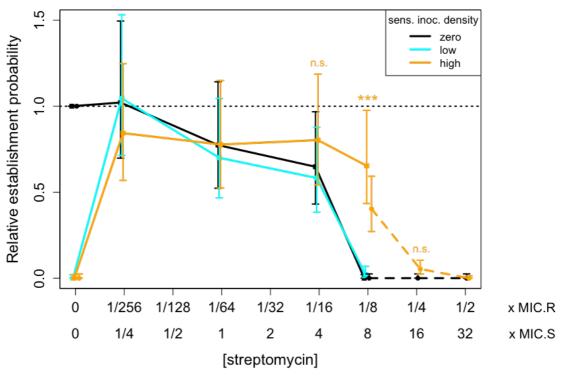


#### 749

750 Figure 5: Estimated relative per-cell establishment probability of the PA01:PAMBL2 meropenem-751 resistant strain as a function of meropenem concentration. Concentration is scaled by the standard 752 MIC of this strain in meropenem (MIC<sub>R</sub> = 512  $\mu$ g/ml; **Table S1**). Plotted points indicate the maximum 753 likelihood estimate of  $\tilde{p}_c$  and error bars indicate the 95% confidence interval, using the fitted model 754 selected by the likelihood ratio test (Model C', the null model [Eqn. 1], which pools data across two 755 tested inoculation densities). Significance of the meropenem effect is determined by fitting a generalized linear model (GLM) to population growth data (n.s.: not significant, p > 0.05; \* p = 0.02; 756 757 \*\*\* *p* < 2e-16; see **Suppl. Text**, section 14.2, for full results). 1/4 x MIC<sub>R</sub> meropenem was excluded 758 from the GLM because zero replicates established. 759

#### Stochastic bacterial population dynamics

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761 Figure 6: Impact of a large sensitive population on the establishment probability of a resistant cell. 762 The PA01:Rms149 resistant strain was seeded either alone (black) or into a low-density (cyan) or 763 high-density (orange) sensitive PA01 population, across a range of streptomycin concentrations. 764 Results are shown from two separate experiments, testing different subsets of conditions 765 (experiment 1 – data points in squares with solid line; experiment 2 – data points in circles with 766 dashed line). Within each experiment, the estimated relative establishment probability per resistant 767 cell ( $\tilde{p}_c$ ) in each condition is normalized by the result for the resistant strain alone in streptomycin-768 free media. Points indicate the maximum likelihood estimate of  $\tilde{p}_c$  and error bars indicate the 95% 769 confidence interval, using the fitted model selected by the likelihood ratio test (Model C', the null 770 model [Eqn. 1] for both experiments). At streptomycin concentrations of particular interest, the 771 number of replicates in which the resistant strain established in the presence of no or low-density 772 sensitive (pooled where applicable) vs. high-density sensitive was compared using a two-sided 773 Wilcoxon rank-sum test, with significance annotated on the plot  $(1/16x \text{ MIC}_{R}: \text{ experiment } 1, p=0.17;$ 774 1/8x MIC<sub>R</sub>: experiment 1, p < 2.2e-16 and experiment 2, p=4.8e-8; 1/4x MIC<sub>R</sub>: experiment 2, p=0.042, 775 not significant after Bonferroni correction); see **Suppl. Text**, section 16, for further details.