> Ancestry-agnostic estimation of DNA sample contamination from sequence reads.

Fan Zhang^{1,2} (fanzhang@umich.edu) Matthew Flickinger^{1,3} (mflick@umich.edu) InPSYght Psychiatric Genetics Consortium Gonçalo R. Abecasis^{1,3} (goncalo@umich.edu) Michael Boehnke^{1,3} (boehnke@umich.edu) Hyun Min Kang^{1,3*} (hmkang@umich.edu)

- 1. Center for Statistical Genetics, University of Michigan, Ann Arbor, MI
- 2. Department of Computational Medicine and Bioinformatics, University of Michigan Medical School, Ann Arbor, MI
- 3. Department of Biostatistics, School of Public Health, University of Michigan, Ann Arbor, MI

*Correspondence to:

hmkang@umich.edu

1 Abstract

2	Detecting and estimating DNA sample contamination are important steps to ensure high quality
3	genotype calls and reliable downstream analysis. Existing methods rely on population allele
4	frequency information for accurate estimation of contamination rates. Correctly specifying
5	population allele frequencies for each individual in early stage of sequence analysis is impractical or
6	even impossible for large-scale sequencing centers that simultaneously process samples from
7	multiple studies across diverse populations. On the other hand, incorrectly specified allele
8	frequencies may result in substantial bias in estimated contamination rates. For example, we
9	observed that existing methods often fail to identify 10% contaminated samples at a typical 3%
10	contamination exclusion threshold when genetic ancestry is misspecified. Such an incomplete
11	screening of contaminated samples substantially inflates the estimated rate of genotyping errors
12	even in deeply sequenced genomes and exomes.
13	We propose a robust statistical method that accurately estimates DNA contamination and is
14	agnostic to genetic ancestry of the intended or contaminating sample. Our method integrates the
15	estimation of genetic ancestry and DNA contamination in a unified likelihood framework by
16	leveraging individual-specific allele-frequencies projected from reference genotypes onto principal
17	component coordinates. We demonstrate this method robustly and accurately estimates
18	contamination rates across different populations and contamination rates. We further demonstrate
19	that in the presence contamination, quantitative estimates of genetic ancestry (e.g. principal
20	component coordinates) can be substantially biased if contamination is ignored, and that our
21	proposed method corrects for this bias. Our method is publicly available at
22	http://github.com/Griffan/verifyBamID

23 Introduction

24	Sample contamination is a common problem in DNA sequencing studies. Contamination may
25	occur during sample shipment (due to spillage across wells, pipetting errors, insufficient dry ice),
26	library preparation (due to gel cut-through in fragment size selection or unexpected switch
27	between barcoded adaptors in-vitro), in-silico demultiplexing from a sequenced lane into barcoded
28	samples, or on many other unexpected occasions. Even modest levels of contamination (e.g. 2-5%)
29	within a species substantially increase genotyping error, even for deeply sequenced genomes ¹ .
30	Accurate estimation of DNA contamination rates allow us to identify and exclude contaminated
31	samples from downstream analysis, and genotypes of moderately contaminated samples (e.g.
32	<10%) can be improved by accounting for contamination in genotype calling ¹ .
33	Previously we developed methods and a software tool, <i>verifyBamID</i> ² , to estimate DNA
34	contamination from sequence reads given known population allele frequencies of common
35	variants. Many investigators and most major sequencing centers use <i>verifyBamID</i> as a part of their
36	standard sequence processing pipeline. However, we have shown that verifyBamID can
37	underestimate DNA contamination rates if the assumed population allele frequencies are
38	inaccurate ² . Such an underestimation can be avoided if correct population allele frequencies are
39	provided in an ideal circumstances. However, in early stage of sequence analysis, performing a
40	tailored customization of quality control (QC) steps for each sequenced genome based on their
41	ancestry is not always feasible or or sometimes impossible. Such a tailored customization requires
42	planned coordination between sequencing centers and study investigators prior to sequencing to
43	share the self-reported ancestry (which is not always accurate) or estimated ancestry from external

genotypes (which is not always available). Modifying the QC pipeline to accommodate studyspecific or sample-specific parameters may not be a possible option for large sequencing centers.
Even if such a tailored customization of QC pipeline is possible, preparing per-sample ancestry prior
to QC may delay time-sensitive issues in the sequencing procedure. If contamination rates can be
accurately estimated without having to know the ancestry or allele frequencies a priori this will
simplify the sequence analysis pipeline and expedite the QC.

50 Here we describe a novel method to robustly detect and estimate DNA contamination by modelling the probability of observed sequence reads as a function of "individual-specific allele 51 52 frequencies" that account for genetic ancestry of a sample. Instead of assuming that the population allele frequencies are known, we represent individual-specific allele frequencies as a function of 53 genetic ancestry using principal component coordinates and the reference genotypes from a 54 diverse population, e.g. Human Genome Diversity Project (HGDP)³ or 1000 Genomes⁴. We then 55 jointly estimate genetic ancestry and contamination rates of a sequenced individual based on a 56 57 mixture model, without requiring the assumption that population allele frequencies are known. 58 Our method enables accurate ancestry-agnostic estimation of contamination through a unified 59 likelihood framework that incorporates genetic ancestry and contamination together. We show that our method provides (1) comparable or more accurate estimates of genetic ancestry than existing 60 methods such as TRACE/LASER^{5,6} even in the absence of contamination and (2) reduced bias in 61 62 contamination rate estimates compared to our previous method requiring known population allele frequencies using in silico contaminated datasets and sequenced genomes from the InPSYght 63 psychiatric genetics sequencing study. 64

65 Material and Methods

66	We aim to jointly estimate sample contamination rates and genetic ancestry from sequence
67	reads without specifying population allele frequencies. First, we describe our previous mixture
68	model to estimate contamination rates assuming population allele frequencies are known. Second,
69	we introduce a model for sequence reads using population allele frequencies as a function of
70	genetic ancestry represented in principal component coordinates. Third, we extend the model to
71	enable joint estimation of contamination rates and genetic ancestry. Fourth, we evaluate our
72	methods using in silico contaminated samples and whole genome sequence data from the InPSYght
73	study.

74 Likelihood-based mixture model for DNA sequence contamination

In our previous contamination detection methods², we assumed that the DNA sequence reads 75 from an intended sample are contaminated by sequence reads from at most one contaminating 76 77 sample from the same population, and that the population allele frequencies of all analyzed genetic 78 variants are known. For each bi-allelic variant i $(1 \le i \le m)$, let $b_{ij} \in \{R, A, O\}$ $(1 \le j \le D_i)$ be the 79 observed base call representing the reference allele (R), alternate allele (A), or other allele (O) for the *j*-th read that overlaps the variant; D_i is the observed sequence depth at variant *i*. Let $e_{ii} \in$ 80 81 {0,1} be a random variable indicating whether a sequencing error did (1) or did not (0) occur for observed base b_{ij} ; we assume e_{ij} follows a Bernoulli distribution with success probability $10^{-\frac{4ij}{10}}$ 82 where Q_{ii} is a phred-scale base quality score of b_{ii} . In the absence of contamination, if the true 83 genotype $g_i \in \{0,1,2\}$ represents the count of alternate alleles of the sequenced sample, then 84

Pr $(b_{ij}|g_i^s, e_{ij})$ can be easily represented as in Table 1, making the simplifying assumption of equally likely errors across four possible nucleotides.

We assume that the observed sequence reads are a $(1 - \alpha)$: α mixture of intended and contaminating reads given a contamination rate $0 \le \alpha \le 1$. Let g_i^1 and g_i^2 represent the true genotypes of the intended and contaminating samples at variant i, respectively. Then the mixture model likelihood of each observed base becomes $\Pr(b_{ij}|g_i^1, g_i^2, e_{ij}; \alpha) = (1 - \alpha)\Pr(b_{ij}|g_i^1, e_{ij}) + \alpha\Pr(b_{ij}|g_i^2, e_{ij})$ (1) Assuming a homogenous population with known population allele frequency f_i and Hardy-Weinberg Equilibrium (HWE), $\Pr(g_i^2; f_i)$ follows a Binomial(2, f_i) distribution. Under the

94 simplifying assumption of independent variants, the likelihood of the contamination rate becomes

95
$$L(\alpha) = \prod_{i=1}^{m} \sum_{g_i^1} \sum_{g_i^2} \left\{ \prod_{j=1}^{D_i} \sum_{e_{ij}} \Pr(b_{ij} | g_i^1, g_i^2, e_{ij}; \alpha) \Pr(e_{ij}) \right\} \Pr(g_i^2; f_i) \Pr(g_i^1; f_i)$$
(2)

96 The maximum likelihood estimate (MLE) of contamination rate $\hat{\alpha}$ can be obtained using Brent's 97 algorithm⁷.

As we previously reported², this model assumes correctly specified population allele frequencies 99 f_i .

100 Likelihood-based estimation of genetic ancestry (in the absence of contamination)

- 101 We extend this model to incorporate genetic ancestry. The key idea of this extension is to use
- 102 the individual-specific allele frequency (ISAF)^{8,9} to model the likelihood of the sequence reads.
- 103 Several methods, including Spatial Ancestry Analysis (SPA)¹⁰ and logistic factor analysis (LFA)⁹,

previously proposed modelling allele frequency as a function of genetic ancestry via principal
 component (PC) coordinates.

Let G be an $m \times n$ genotype matrix (where g_{ij} = 0, 1, or 2 is the number of non-reference 106 107 alleles at variant *i* in individual *i*) of a genetically diverse reference panel of size n, such as 1000 Genomes or HGDP. We define ISAF f_i ($0 \le f_i \le 1$) for variant *i* as a weighted average of genotypes 108 from the reference panel ($f_i = \sum_{r=1}^n w_r G_{ir}$), where $0 \le w_r \le 1$ and $G_{ir} \in \{0,1,2\}$ for individual *r*. 109 For a homogenous population, $w_r = \frac{1}{2n}$ results in a *pooled allele frequency* across all individuals in 110 the reference panel. If each individual can be categorically represented as a one of k mutually 111 exclusive subpopulations, the *population-specific allele frequency* for the subpopulation $s \in$ 112 $\{1,2, \dots, k\}$ can be represented as $w_r = \frac{I(s_r=s)}{2n_s}$, where and $s_r \in \{1,2, \dots, k\}$ represents the 113 114 subpopulation that individual <u>r</u> belongs to, and n_s represents the size of subpopulation s . More 115 generally, if individual's genetic ancestry is represented as continuous variables (such as PCs, SPAs, or LFAs), the individual-specific allele frequency (ISAF) can be represented as a function of the 116 continuously represented genetic ancestry^{9,5}. 117

The estimated ISAF can be viewed as (one half times the) genotype dosages approximated from a fixed number(=K) of factors, such as PCs, SPAs, or LFAs. In our method, we used a linear model to estimate ISAF from PCs, similar to previous studies^{8,9}. Given the reference panel genotype matrix G, let $\frac{1}{2}\hat{G}$ be the *ISAF matrix* as a function of top K factors. ISAF matrix $\frac{1}{2}\hat{G}$ should well approximate $\frac{1}{2}G$. For example, under a linear model, typical principal component analysis takes the singular value decomposition (SVD) of the mean-centered genotype matrix $\overline{G} = G - 2\mu \mathbf{1}_n^T = UDV^T$, where $\mu = \frac{1}{2n}G\mathbf{1}_n$ is the pooled allele frequencies and $\mathbf{1}_n$ is the column-vector of ones. Using the top K eigenvalues and corresponding eigenvectors $U^{(K)}$, $D^{(K)}$, $V^{(K)}$ from the SVD, it is known that $\hat{G} = \frac{1}{2}U^{(K)}D^{(K)}[V^{(K)}]^T + \mu \mathbf{1}_n^T$ minimizes $\|G - \hat{G}\|_2 = \sum_{i,j}(G_{ij} - \hat{G}_{ij})^2$ among all possible rank *K* matrices¹¹, making it a good proxy for the ISAF matrix.

For a new individual *s* with genetic ancestry represented as $x_s \in \mathbb{R}^k$ in the PC (eigenvector) space of the reference panel, the ISAF for *i*-th variant can be modelled as $f_i(x_s) = \frac{1}{2}u_i^{(K)}D^{(K)}x_s^T + \mu_i$, where $u_i^{(K)}$ is *i*-th row of $U^{(K)}$ and μ_i is the *i*-th element of μ . To avoid boundary condition, we constrain $\frac{\varepsilon}{2n} \le f_i(x_s) \le 1 - \frac{\varepsilon}{2n}$ for a fixed ε (we used $\varepsilon = 0.5$ in our experiments). Then the overall likelihood of an individual's genetic ancestry x is

133
$$L(\boldsymbol{x}_{s}) = \prod_{i=1}^{m} \sum_{g_{i}} \left\{ \prod_{j=1}^{D_{i}} \sum_{e_{ij}} \Pr(b_{ij} | g_{i}, e_{ij}) \Pr(e_{ij}) \right\} \Pr(g_{i}; f_{i}(\boldsymbol{x}_{s}))$$
(3)

where g_i represents the unobserved genotype of the sequenced sample at variant *i*. The maximumlikelihood genetic ancestry coordinates can be estimated as $\hat{x}_s = \operatorname{argmax}_{x_s \in \mathbb{R}^k} L(x_s)$ using the Nelder-Mead¹² algorithm, starting with PC coordinates of a randomly selected individual from the reference panel. In our experiments, we always obtained consistent estimates of \hat{x}_s regardless of start values.

139 Joint estimation of genetic ancestry and DNA contamination

Because our goal is to obtain unbiased estimates of the DNA contamination rate α agonistic to genetic ancestry, we propose to jointly estimate α and ancestry by combining the models described in the previous sections. Let $x_1, x_2 \in R^K$ be the genetic ancestries of the intended and contaminating samples. Then the likelihood under the combined model is

144
$$L(\alpha, \mathbf{x_1}, \mathbf{x_2}) = \prod_{i=1}^{m} \sum_{g_i^1} \sum_{g_i^2} \left\{ \prod_{j=1}^{D_i} \sum_{e_{ij}} \Pr(b_{ij} | g_i^1, g_i^2, e_{ij}; \alpha) \Pr(e_{ij}) \right\} \Pr(g_i^1; f_i(\mathbf{x_1})) \Pr(g_i^2; f_i(\mathbf{x_2}))$$

145 When the contamination rate $\alpha \approx 0$, the parameters corresponding to x_2 do not contribute 146 (much) to the likelihood and the estimates of x_2 may be unstable. To address this problem, we 147 initially assume that the intended and contaminating samples are from the same population $x_1 =$ 148 x_2 ('equal-ancestry' model) and then repeat the analysis allowing for $x_1 \neq x_2$ ('unequal-ancestry' 149 model). The dimension of parameter space for the unequal-ancestry model is 2k + 1. We choose 150 final parameter estimates between the two models based on Akaike Information Criterion (AIC)¹³.

151 Evaluation on *in-silico* contaminated data based on 1000 Genomes project samples

We constructed in-silico contaminated DNA sequence reads using aligned low-coverage whole 152 genome sequence reads from the 1000 Genomes phase 3 project⁴. We filtered out unmapped and 153 154 mark-duplicated reads and then randomly sampled aligned sequence reads proportional to the 155 intended contamination rates $\alpha \in \{0.01, 0.02, 0.05, 0.1, 0.2\}$. To match the mixing proportion of 156 sequence reads originated from intended and contaminating to be $(1 - \alpha)$: α , each read was sampled with probability $(1 - \alpha)$ and $\frac{B_1}{B_2}\alpha$ from each sample, where B_1 and B_2 are number of 157 158 aligned bases from unique reads from intended and contaminating samples. We selected four populations, CHS (Han Chinese South), GBR (British in England and Scotland), MXL (Mexican 159 160 Ancestry from Los Angeles USA), YRI (Yoruba in Ibadan, Nigeria), and arbitrarily selected 10 pairs of 161 individuals with similar sequencing depths within the same population and across populations. To 162 estimate genetic ancestry and/or contamination rate for these in-silico contaminated sequence 163 reads, we used a reference panel of 938 HGDP³ individuals across 1,000, 10,000 and 100,000

164 randomly chosen SNPs (pooled MAF > 0.5%), avoiding variants masked by the 1000 Genomes

165 Project⁴ (See Web Resource). When we compared estimated genetic ancestry with *LASER*, we used

166 the same set of selected SNPs and sequence reads as input. For TRACE, we used genotypes from the

- 167 phase 3 release (for 1000 Genomes) or an interim callset from the *GotCloud* software tool¹⁴ (for
- 168 InPSYght, see next section for details) on the same SNP set.

169 Experiment with real sequence data from the InPSYght study

170 Next, we applied our method to 500 deeply sequenced (mean depth 32x) genomes from the first two batches of the InPSYght study. For each sample, we evaluated the results from the six 171 models: (1) the original *verifyBamID* using pooled allele frequencies; the original *verifyBamID* using 172 173 (2) African, (3) East Asian, and (4) European allele frequencies; (5) the new verifyBamID2 under the equal-ancestry model; and (6) verifyBamID2 under the unequal-ancestry model. To calculate 174 175 pooled, population-specific, and individual-specific allele frequencies, we used the 1000 Genomes 176 phase 3 reference panel (n=2,504), randomly selecting 100,000 SNPs among the sites also 177 polymorphic in Illumina Human Omni 2.5 array, with the same filtering criteria (MAF > 5% and 1000 178 Genomes mask) as above.

180 **Results**

181	We assessed our new methods in the following steps. First, in the absence of contamination, we
182	demonstrate that our estimation of genetic ancestry provides comparably accurate estimates of
183	genetic ancestry as other state-of-art methods. Second, in the presence of contamination, we
184	demonstrate that joint estimation of genetic ancestry and contamination substantially improves the
185	estimation accuracy of both parameters. Third, using in-silico contaminated samples, we
186	demonstrate that our methods robustly provide more accurate estimates than previous methods
187	across various combinations of genetic ancestries and contamination rates. Fourth, from the
188	analysis of deeply sequenced genomes in the InPSYght study, we demonstrate that our new
189	methods deliver more accurate contamination estimates than the previous methods.
190	

191 New model-based methods accurately estimate genetic ancestry.

In the absence of contamination, widely used methods such as LASER and TRACE are known to 192 193 estimate genetic ancestry accurately. Because we propose using a new model-based approach to 194 estimate the genetic ancestry (jointly with contamination rates), we first compared the accuracy of 195 our new method, in the absence of contamination, with LASER and TRACE. We randomly chose 500 196 ethnically diverse samples from the 1000 Genomes Project low-coverage (4X) genomes, and 500 197 African American samples from the deeply sequenced (32x) genomes from the InPSYght project. 198 We estimated their genetic ancestries using 100,000 SNPs from the HGDP reference panel (see 199 Methods for details) and compared their genetic ancestry estimates obtained by LASER (using the 200 same sequence data), and TRACE (using the hard-call genotypes). As illustrated in Figure 1A, 1C, 1E,

201 the estimated PC coordinates of the 1000 Genomes individuals are located close to their

202 corresponding HGDP populations across all three methods. Compared to TRACE and LASER, we

203 observed that the estimated genetic coordinates from *verifyBamID2* were the closest to the

204 centroid of corresponding HGDP population (Table 2) in 4 of the 5 populations (all except TSI).

205 These results suggest that our method provides estimates at least as precise compared to those for

206 other state-of-the-art methods.

207 Genetic ancestry estimates may be confounded by DNA contamination.

208 Next, we constructed *in-silico* contaminated sequenced data from the 1000 Genomes Project

and estimated contamination parameters and genetic ancestries jointly. We observed that when

210 sequences are contaminated between different continental populations, the genetic ancestry esti-

211 mates in PC coordinates drift towards the contaminating population when contamination is ignored

212 (Figure 2A) or when assuming that intended and contaminating samples originated from the same

213 population (Figure 2B). As the contamination rate increases, drift increases.

However, when we accounted for possible differences in genetic ancestries between the two

215 intended and contaminating samples using our new methods, PC coordinates remained similar to

- those for uncontaminated samples (Figure 2E), and contaminated samples constructed from indi-
- viduals that belong to the same population (Figure 2B, 2D, 2F).

218 Robust, accurate, ancestry-agnostic estimation of DNA contamination.

- 219 Next, we evaluated the effect of genetic ancestry misspecification in estimating DNA
- 220 contamination rates. We constructed contaminated samples between various combinations of

221	populations, and compared the accuracy of estimated contamination rates using both the original
222	methods which assume known allele frequencies and the new methods which estimate
223	contamination rate and genetic ancestry jointly.
224	When contamination happens within the same population, running original methods with
225	correct continental population allele frequencies specified provided accurate contamination
226	estimates (Figure 3A, 3E, 3I). However, using pooled allele frequencies, which would be a default
227	option when it is infeasible to specify population information <i>a priori</i> before sequencing,
228	consistently underestimated contamination rates. Bias was particularly large when intended individ-
229	uals were of African ancestry.
230	Specifying incorrect population allele frequencies results in even larger contamination
231	estimation bias. For example, using African allele frequencies on East Asian samples resulted in an
232	average estimate of 2.9% contamination for samples with contamination 10% (Table S1), implying
233	that a large fraction of 10% contaminated samples within East Asian ancestry would not have been
234	flagged for contamination-based exclusion at the contamination-exclusion threshold of 1-3% used
235	by many studies e.g. the Trans-Omics Precision Medicine (TOPMed) study ¹⁵ .
236	Our results consistently demonstrated that the ancestry-agnostic method provides as accurate
237	estimates as the original methods specified with correct population labels (Figure 3A, 3E, 3I, Table
238	S1), and the estimates are substantially better than those from pooled allele frequencies or
239	incorrectly specified allele frequencies.
240	When the intended and contaminating populations are different, we observed that

241 contamination is sometimes overestimated due to increased fraction of heterozygous genotypes

242	than expected by a given contamination rate under single population model. Our method based on
243	unequal-ancestry model outperforms all the other methods in terms of overall bias and Mean
244	Squared Error(MSE) (Figure 3, Table S4), correcting for both upward and downward biases in
245	various ancestry combinations. For example, the relative deviation of estimated to intended
246	contamination rate (i.e. $ \hat{lpha}/lpha-1)$ is reduced by 80% (73-88%) compared to the original
247	verifyBamID with various population allele frequencies, suggesting reduced bias. MSE is also
248	reduced by 92% (86-97%). This robustness reflects the ability to incorporate differences in
249	population allele frequencies between intended and contaminating individuals (Figure 3B, 3C, 3D,
250	3F, 3G, 3H, Table S1).
251	We also examined the accuracy of our methods for admixed populations by performing a similar
252	experiment using the Mexican population (MXL) and obtained consistent results (Supplementary
253	Table S2).
254	Results with deep whole genome sequence data from the InPSYght study.
255	Next, we applied our methods to 500 African American samples from the InPSYght study (see

256 Methods). Consistent with the results from our *in silico* contamination studies, we observed that

- the average contamination rate was 1.1-fold higher with newer method (0.36% for unequal-
- ancestry, 0.37% for equal-ancestry) compared to the original method with pooled allele frequency
- 259 (0.33%) (Figure 4). The number of samples with estimated contamination rate >1% increased from
- 260 16 (original method with pooled allele frequency) to 21 (unequal-ancestry method) or 23 (unequal-
- ancestry method), suggesting our new method more rigorously screens for contaminated samples.

262 All 500 deeply sequenced genomes in InPSYght study are reported to be African Americans, and 263 indeed the estimated PC coordinates for all 500 individuals under all three methods lie between European and African samples. Compared to other methods to estimate genetic ancestry, our 264 265 estimates resulted in tighter clustering along the European-African segment than LASER, and 266 similarly tight clustering to TRACE (Figure 1B, 1D, 1F). For example, the correlation coefficient 267 between the PC1 and PC2 coordinates were 0.927 for LASER, 0.981 for TRACE, and 0.985 for 268 verifyBamID2, corroborating that verifyBamID2 results in more precise estimate of African ancestry 269 along the European-African segment in PC coordinates.

270 Impact of number of markers on accuracy, computational cost, and memory requirements.

As we have shown previously², there are trade-offs between computation cost and accuracy of contamination estimates. Using as many as 100,000 variants results in accurately estimated intended contamination rate. For example, MSE of relative deviation (i.e. $|\hat{\alpha}/\alpha - 1|$) was 0.02, 0.01, 0.01 when the intended contamination was 1%, 2%, and 5%, respectively. When we use 10,000 variants, the MSEs modestly increased to 0.11, 0.04, and 0.01, respectively. When we use only 1,000 variants, MSEs further increased to 0.69, 0.25, 0.11, suggesting that the estimates may not be precise for low contamination rate when using only 1,000 variants. (Supplementary Table S3).

We also evaluated the computational cost and memory consumption of *verifyBamID2* on whole genome sequence data with various coverages. For the BAM files from the 1000 Genomes whole genome sequence data (4.3-5.1x coverage), the average wall-clock running time was 5.5 minutes with a single thread and peak memory consumption was 505 MB when using 10,000 markers in a server with Xeon 2.27GHz processor. When using 100,000 markers, the average wall-clock running time was

- 283 20.5 minutes with a single thread and 8.0 minutes with four threads, and peak memory consumption
 284 was 528 MB.
- 285 For deep genome data from the InPSYght study (31x coverage) stored in CRAM format, the
- average wall-clock time was 17.3 minutes and peak memory consumption was 514 MB when using
- 287 10,000 markers. For 100,000 markers the average wall-clock time was 155.6 minutes (single thread)
- 288 or 96 minutes (four threads) and peak memory consumption was 548 MB.

289 **Discussion**

290 Contamination detection is an essential step in the sequence analysis process that has important 291 effects on following downstream analyses. Early and accurate estimation of DNA contamination can 292 prevent wasted effort, time, and money by identifying the problems early on before too many 293 samples are sequenced using contamination-prone protocols. Our previous method enabled such a 294 timely contamination detection from sequence data and population allele frequencies at known 295 variant sites, without requiring independent SNP genotype data. Our new method maintains these 296 advantages, and in addition provide three more. First, because our joint analysis method is agnostic 297 to genetic ancestry, it eliminates sample-to-sample variation in the parameter settings for the 298 contamination checking procedure, simplifying the sequence analysis pipeline. Second, it provides 299 more robust contamination estimates against potentially misspecified population allele frequency of the intended (or contaminating) samples when relying on the reported ancestry information. Third, 300 301 it provides accurate estimates of genetic ancestries for both intended and contaminating samples. 302 This enables additional sanity checking of the sequence data, such as determining whether a 303 sequenced sample matches its expected (participant-reported) ancestry. It also facilitates 304 incorporating ancestry information in the variant calling and downstream analysis, and allows us to 305 track the source of contamination more precisely when contamination occurs.

Our method can be used not only to detect and estimate contamination, but also to estimate genetic ancestry from sequence data. Relatively few methods, such as *LASER*^{5,6} and *bammds*¹⁶, exist for estimating genetic ancestry from sequence data while several methods have been developed for array-based genotypes, such as *EIGENSOFT*¹⁷, *FRAPPE*¹⁸, *ADMIXTURE*¹⁹, and *TRACE*⁶. We have

310 demonstrated that our method provides ancestry estimates as or more accurate than *LASER*, 311 particularly when the sequenced samples are contaminated between different ancestries.

By jointly estimating genetic ancestry and contamination, we are able to accurately estimate contamination without requiring ancestry information *a priori*. Since obtaining population allele frequency information may be infeasible or even impossible at the time of sequencing, it is important to highlight that our ancestry-agnostic approach provides more timely and accurate feedback to the sequencing facilities. Our ancestry-agnostic approach also simplifies the sequence analysis pipeline, because the same input arguments can be applied across all samples regardless of their genetic ancestry

319 The key idea of using individual-specific allele frequencies (ISAF) to account for population 320 structure in genetic analysis has been suggested previously in the context of characterizing 321 population structure or identifying highly differentiated variants across populations^{8,9}. To the best 322 our knowledge, our method describes the first likelihood-based model utilizing ISAF to represent high 323 throughput sequence reads under population structure and/or contamination. While previous studies proposed logistic models as alternative to linear model^{8,9}, we used linear models (bounded 324 325 by minimum and maximum value) between allele frequencies and population structure represented 326 by Singular Value Decomposition (SVD) on the genotype matrix. We made this choice because the 327 logistic model is computationally more intensive, and the linear model is accurate for the common 328 variants we use, as demonstrated by the previous studies⁹.

Because we use Nelder-Mead optimization for maximum likelihood estimation, it is possible that the estimates do not converge to the global maximum, especially when many principal components are used. We observed that estimating the full unequal-ancestry model parameters sometimes does 332 fail to converge especially when there is little or no contamination, due to the limited identifiability of the genetic ancestry of contaminating samples in this situation. Starting by estimating 333 contamination rate and shared genetic ancestry parameters using the equal-ancestry model, and 334 335 using those estimates as start values for the unequal-ancestry model to allow different ancestries 336 between the intended and contaminating samples dramatically improved convergence; in fact, the 337 method converged to consistent estimates across multiple starting points within 1,000 iterations in 338 all our benchmark cases, in both real and in-silico contaminated data. When the contamination rate 339 is extremely small (e.g. <0.1%), estimation of genetic ancestry of contaminating samples can still be 340 challenging. We allow unequal ancestries between intended and contaminating samples only when 341 the likelihood substantially improves beyond AIC threshold between equal ancestry and unequal 342 ancestry models. This procedure effectively removed all outlier estimates of genetic ancestries of 343 contaminating samples in our experiments.

There are other possible useful extensions to our joint contamination and estimation method. We are extending these methods to detect and estimate contamination for RNA-seq and other epigenomic sequence data. The same model has potential applications in other areas, such as cancer single cell transcriptomics²⁰.

We expect that our new *verifyBamID2* software will facilitate more accurate, convenient, and timely quality control of sequence genomes. Our software tool is publicly available at http://github.com/Griffan/verifyBamID. Our GitHub repository provides reference files that can be used as test input for our methods. These files contain key input files required for *verifyBamID2*, including variant loadings, supporting various genome builds (GRCh37 and GRCh38), and various numbers of variants.

354 Web Resources

- 355 1000 genomes project genome mask file:
- 356 (ftp://ftp.1000genomes.ebi.ac.uk/vol1/ftp/release/20130502/supporting/accessible_genome_mask
- 357 s/StrictMask/)

358

359 Acknowledgements

- 360 This work was supported by NIH grants HG009976 (to M.F. and M.B.), HL137182 (to H.M.K. and F.Z),
- 361 HG007022 (to G.R.A), MH105653 (to M.B, InPSYght Consortium, and H.M.K.)

362

363 **References**

- 1. Flickinger, M., Jun, G., Abecasis, G.R., Boehnke, M., and Kang, H.M. (2015). Correcting for sample
- 365 contamination in genotype calling of DNA sequence data. Am. J. Hum. Genet. *97*, 284–290.
- 366 2. Jun, G., Flickinger, M., Hetrick, K.N., Romm, J.M., Doheny, K.F., Abecasis, R., Boehnke, M., and
- 367 Kang, H.M. (2012). Detecting and estimating contamination of human DNA samples in sequencing
- and array-based genotype data. 839–848.
- 369 3. Cavalli-Sforza, L.L. (2005). The Human Genome Diversity Project: past, present and future. Nat
- 370 Rev Genet *6*, 333–340.
- 4. The 1000 Genomes Project Consortium (2015). A global reference for human genetic variation.

372 Nature *526*, 68–74.

- 373 5. Wang, C., Zhan, X., Bragg-Gresham, J., Kang, H.M., Stambolian, D., Chew, E.Y., Branham, K.E.,
- Heckenlively, J., Fulton, R., Wilson, R.K., et al. (2014). Ancestry estimation and control of
- population stratification for sequence-based association studies. Nat. Genet. *46*, 409–415.
- 376 6. Wang, C., Zhan, X., Liang, L., Abecasis, G.R., and Lin, X. (2015). Improved ancestry estimation for
- both genotyping and sequencing data using projection procrustes analysis and genotype
- 378 Imputation. Am. J. Hum. Genet. *96*, 926–937.
- 379 7. Brent, R.P. (1974). Algorithms for minimization without derivatives. IEEE Trans. Automat. Contr.
- 380 8. Conomos, M.P., Reiner, A.P., Weir, B.S., and Thornton, T.A. (2016). Model-free estimation of
- recent genetic relatedness. Am. J. Hum. Genet. *98*, 127–148.
- 382 9. Hao, W., Song, M., and Storey, J.D. (2015). Probabilistic models of genetic variation in structured
- populations applied to global human studies. Bioinformatics *32*, 713–721.
- 10. Yang, W.W.-Y., Novembre, J., Eskin, E., and Halperin, E. (2012). A model-based approach for
- analysis of spatial structure in genetic data. Nat. Genet. 44, 725–731.
- 11. Pearson, K. (1901). LIII. On lines and planes of closest fit to systems of points in space. Philos.
- 387 Mag. Ser. 6 2, 559–572.
- 388 12. Nelder, J.A., and Mead, R. (1965). A simplex method for function minimization. Comput. J. 7,
 389 308–313.
- 390 13. Akaike, H. (1974). A new look at the statistical model identification. IEEE Trans. Automat. Contr.
- 391 14. Jun, G., Wing, M.K., Abecasis, G.R., and Kang, H.M. (2015). An efficient and scalable analysis
- 392 framework for variant extraction and refinement from population-scale DNA sequence data.

- 393 Genome Res. 25, 918–925.
- 394 15. Natarajan, P., Peloso, G.M., Zekavat, S.M., Montasser, M., Ganna, A., Chaffin, M., Khera, A. V.,
- Zhou, W., Bloom, J.M., Engreitz, J.M., et al. (2018). Deep-coverage whole genome sequences and
- blood lipids among 16,324 individuals. Nat. Commun.
- 397 16. Malaspinas, A.S., Tange, O., Moreno-Mayar, J.V., Rasmussen, M., DeGiorgio, M., Wang, Y.,
- 398 Valdiosera, C.E., Politis, G., Willerslev, E., and Nielsen, R. (2014). bammds: a tool for assessing the
- ancestry of low-depth whole-genome data using multidimensional scaling (MDS). Bioinformatics
- *30,* 2962–2964.
- 401 17. Price, A.L., Patterson, N.J., Plenge, R.M., Weinblatt, M.E., Shadick, N. a, and Reich, D. (2006).
- 402 Principal components analysis corrects for stratification in genome-wide association studies. Nat.
- 403 Genet. *38*, 904–909.
- 404 18. Tang, H., Peng, J., Wang, P., and Risch, N.J. (2005). Estimation of individual admixture: Analytical
 405 and study design considerations. Genet. Epidemiol. *28*, 289–301.
- 406 19. Alexander, D.H., Novembre, J., and Lange, K. (2009). Fast model-based estimation of ancestry in
 407 unrelated individuals. Genome Res. *19*, 1655–1664.
- 408 20. Kang, H.M., Subramaniam, M., Targ, S., Nguyen, M., Maliskova, L., McCarthy, E., Wan, E., Wong,
- 409 S., Byrnes, L., Lanata, C.M., et al. (2018). Multiplexed droplet single-cell RNA-sequencing using
- 410 natural genetic variation. Nat. Biotechnol.
- 411

True Genotype <i>g</i> _i	Base Calling Error Event <i>e</i> ij	Pr(<i>b_{ij}</i> = R)	Pr(<i>b_{ij}</i> = A)	Pr(b _{ij} = O) ^b
g _i = RR ^a	<i>e_{ij}</i> = 0	1	0	0
	<i>e</i> _{<i>ij</i>} = 1	0	1/3	2/3
a – DAa	$e_{ij} = 0$	1/2	1/2	0
$g_i = RA^a$	<i>e</i> _{<i>ij</i>} = 1	1/6	1/6	2/3
α. – ΔΔa	<i>e</i> _{<i>ij</i>} = 0	0	1	0
$g_i = AA^a$	<i>e</i> _{<i>ij</i>} = 1	1/3	0	2/3

Table 1. Conditional probability $P(b_{ij} | g_i, e_{ij})$ of read b_{ij} given true genotype g_i and the variable representing the event of base calling error e_{ij} , as described in (Jun et al 2012²)

^a RR, RA, AA: homozygous reference, heterozygous, and homozygous non-reference genotypes

^b O: alleles other than R or A; assumes four possible alleles (bases)

Popula	Population Label		LASER	verifyBamID2	
1000G	HGDP	TRACE	LAJER	venjybumiDz	
СНВ	Han-NChina	1.68	2.61	0.40	
CHS	Han	1.70	1.24	1.18	
TSI	Tuscan	1.52	2.16	1.81	
YRI	Yoruba	2.32	1.73	0.42	
JPT	Japanese	1.54	1.03	1.22	

Table 2. Distance between estimated PCA coordinates of HGDP and 1000G populations^{*}

*Distances were measured between the mean PCA coordinates across the population in HGDP (estimated from the array data of Wang et al.⁶) and the mean PCA coordinates estimated from 1000 Genomes low coverage sequence data of the corresponding population, projected onto the same PCA coordinates using *TRACE*, *LASER*, or *verifyBamID2* (assuming no contamination). Bold face represents the smallest distance among the three methods for each population.

Sample Population		Original N	lodel (Fixed	Allele Freq	Equal-Ancestry	Unequal-Ancestry		
Intended	Contaminating	European	East Asian	African	Pooled	Model	Model	
GBR	GBR	4.73%	3.19%	2.67%	3.76%	4.63%	4.63%	
CHS	CHS	1.90%	4.73%	1.25%	2.38%	4.73%	4.76%	
YRI	YRI	1.78%	1.58%	4.44%	2.45%	4.40%	4.40%	
CHS	YRI	3.33%	6.91%	2.27%	4.10%	6.71%	4.81%	
YRI	CHS	2.79%	2.55%	6.29%	3.76%	5.99%	4.67%	
GBR	YRI	6.13%	4.16%	3.60%	5.04%	5.90%	4.83%	
YRI	GBR	2.81%	2.57%	6.38%	3.80%	6.01%	4.63%	
CHS	GBR	2.87%	6.33%	1.98%	3.55%	6.13%	4.83%	
GBR	CHS	5.32%	3.78%	3.05%	4.32%	5.16%	4.67%	

Table 3. Average contamination estimates for 5% contaminated samples (size r	ı=10).
--	--------

416 Average contamination estimates of *in-silico* contaminated samples when the true contamination

417 rate is 5%. Each mixing configuration (e.g. GBR+CHS) contains 10 samples that are constructed with

418 95% reads coming from the intended sample and 5% reads from the contaminating sample. The

estimated contamination rates are obtained using the original version *verifyBamID* by specifying
 prior allele frequencies as European, East Asian, African, and Pooled , respectively. Bold represents

421 the closest estimate to the true value of 5%.

422 Figure Legends

423

424 Figure 1.

Evaluation of estimated genetic ancestry coordinates, in the absence of contamination, between *TRACE, LASER*, and *verifyBamID2* on samples from the 1000 Genomes low coverage genome (n=500, diverse ancestry) sequence data (A,C,E) and from the InPSYght deep genome (n=500, African Americans) sequence data (B,D,F). Panels A and B show results from *TRACE*, C and D from *LASER*, and E and F from *verifyBamID2* (assuming no contamination). Each point represents a sample and each color represents a population ancestry with the exception that grey point represents PCA coordinates of reference (HGDP) samples.

432

433 Figure 2.

434 Impact of DNA sample contamination on the estimation of genetic ancestry. Each point represents a 435 sample. Each grey point represents reference (HGDP) sample and its PCA coordinates, similar to 436 Figure 1. Each colored point represents in-silico contaminated samples across various 437 contamination rates and populations. In panels A, C, and E, European (GBR) and East Asian (CHS) 438 samples are contaminated with African (YRI) samples at different contamination rates (i.e. 439 between-ancestry contamination). In panel B, D, and F, European (GBR) and East Asian (CHS) 440 samples are contamination with another sample in the same population (i.e. within-ancestry 441 contamination). Different colors represent different contamination rates ranging from 1% to 20%. 442 Upper panels (A, B) show verifyBamID2 estimates without modelling contamination, middle panels 443 (C, D) verifyBamID2 estimates under the assumption that intended and contaminating populations are identical (i.e. equal-ancestry model), lower panels (E, F) verifyBamID2 estimates under the 444

- assumption that intended and contaminating populations can be different (i.e. unequal-ancestrymodel).
- 447
- 448 Figure 3.

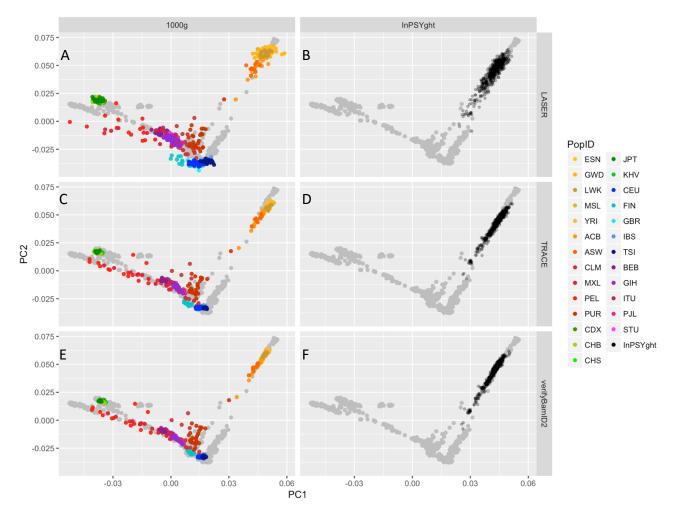
449 Comparison of different models to estimate contamination rates. Horizontal (x) axis shows intended contamination rate, vertical (v) axis shows the ratio of estimated to intended contamination rates. 450 451 Each color represents different models to estimate contamination rates. EUR AF, EAS AF, and 452 AFR AF represent original verifyBamID using European, East Asian, and African allele frequencies 453 across the continental population using the 1000 Genomes data. Pooled AF represents the original 454 verifyBamID using aggregated allele frequencies across all 2,504 individuals in the 1000 Genomes 455 Project. Equal Ancestry represents the verifyBamID2 assuming that intended and contaminating 456 samples belong to the same population. Unequal Ancestry represents verifyBamID2 allowing 457 different genetic ancestry between intended and contaminating sample (recommended setting).

458 Each panel represents different combinations of intended (row) and contaminating (column) 459 populations, in the order of GBR, CHS, and YRI.

460

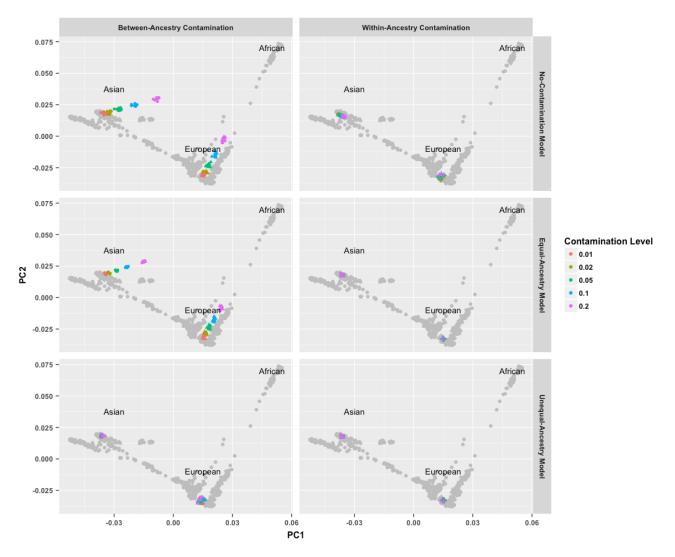
461 Figure 4

462 Comparison of contamination estimation between using *verifyBamID* and *verifyBamID2* on 500 463 InPSYght samples. All subjects are African Americans. Each dot represents the pair of contamination 464 rate estimates using different methods. The left panel shows the estimated contamination rates of 465 the original *verifyBamID* with pooled allele frequencies, which is the default setting of *verifyBamID* in 466 x-axis. Y-axis shows *verifyBamID2* with unequal-ancestry model (y-axis). Each point represents a 467 sequenced subject. The right panel compares the estimated contamination rates between two 468 models (unequal-ancestry vs. equal-ancestry) of *verifyBamID2* on the same dataset.



470 Figure 1.

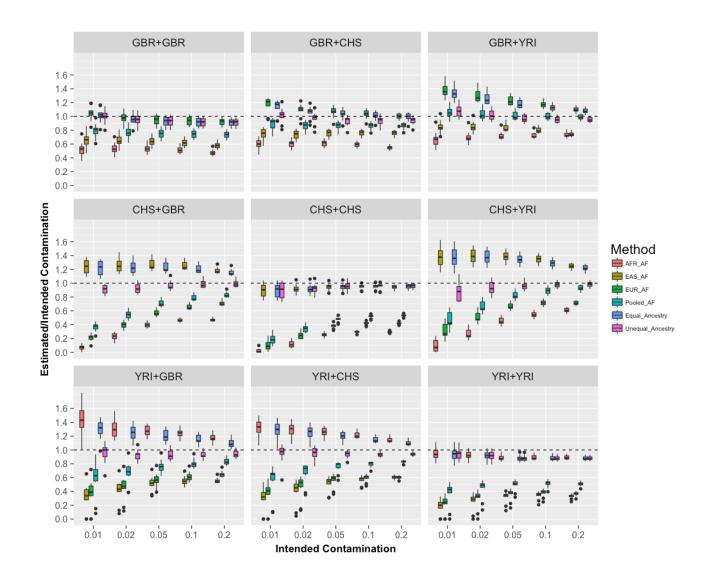
Evaluation of estimated genetic ancestry coordinates, in the absence of contamination, between *TRACE, LASER*. and *verifyBamID2* on samples from the 1000 Genomes low coverage genome (n=500, diverse ancestry) sequence data (A,C,E) and from the InPSYght deep genome (n=500, African Americans) sequence data (B,D,F). Panel A and B show results from *TRACE*, C and D from *LASER*, and E and F from *verifyBamID2* (assuming no contamination). Each point represents a sample, each color represents a population ancestry with the exception that grey point represents PCA coordinates of reference (HGDP) samples.



479

480 Figure 2.

481 Impact of DNA sample contamination on the estimation of genetic ancestry. Each point represents a sample. Grey point represents reference (HGDP) sample and its PCA coordinates, similar to Figure 1. 482 483 Each colored point represents in-silico contaminated samples across various contamination rates and 484 populations. In panel A, C, E, European (GBR) and East Asian (CHS) samples are contaminated with 485 African (YRI) samples at different contamination rates (i.e. between-ancestry contamination). In 486 panel B, D, F, European (GBR) and East Asian (CHS) samples are contamination with another sample 487 in the same population (i.e. within-ancestry contamination). Different colors represent different 488 contamination rate ranging from 1% to 20%. Upper panels (A, B) show verifyBamID2 estimates 489 without modelling contamination. Middle panels (C, D) show verifyBamID2 estimates under the 490 assumption that intended and contaminating populations are identical (i.e. equal-ancestry model). 491 Lower panels (E, F) show verifyBamID2 estimates under the assumption that intended and 492 contaminating populations can be different (i.e. unequal-ancestry model).

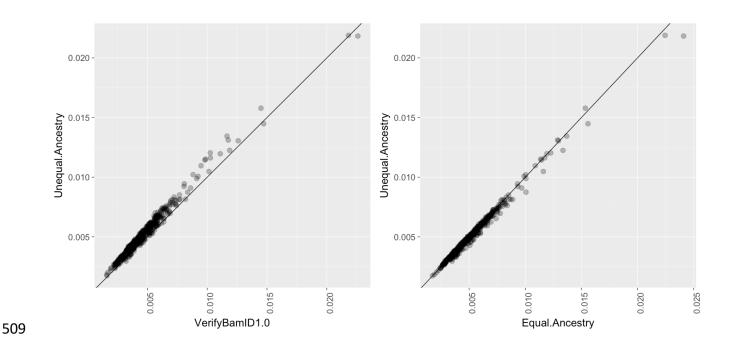


494

495

496 Figure 3.

497 Comparison of different models to estimate contamination rates. Horizontal (x) axis shows intended 498 Contamination rate, vertical (v) axis shows the ratio of estimated to intended contamination rates. 499 Each color represents different models to estimate contamination rates. EUR AF, EAS AF, AFR AF represents old verifyBamID using European, East Asian, and African allele frequencies across the 500 501 continental population using the 1000 Genomes data. Pooled AF represents the old verifyBamID 502 using aggregated allele frequencies across all 2,504 individuals in the 1000 Genomes Project. "Equal Ancestry" represents the verifyBamID2 assuming that intended and contaminating samples 503 504 belong to the same population. "Unequal Ancestry" represents verifyBamID2 allowing different 505 genetic ancestries between intended and contaminating samples (recommended setting). Each panel 506 represents different combinations of intended (row) and contaminating (column) populations, in the 507 order of GBR, CHS, and YRI.



- 510 Figure 4.
- 511 Comparison of contamination estimation between using *verifyBamID* and *verifyBamID2* on 500
- 512 InPSYght samples. All subjects are African Americans. Each dot represents the pair of contamination
- rate estimates using different methods. The left panel shows the estimated contamination rates of
- 514 the original *verifyBamID* with pooled allele frequencies, which is the default setting of *verifyBamID*
- 515 in x-axis. Y-axis shows *verifyBamID2* with unequal-ancestry model (y-axis). Each point represents a
- 516 sequenced subject. The right panel compares the estimated contamination rates between two
- 517 models (unequal-ancestry vs. equal-ancestry) of *verifyBamID2* on the same dataset.

518 Supplementary Materials

519 **Supplementary Table S1**: Mean estimated contamination rates of *in-silico* contaminated population

520 across different intended contamination rate, populations of intended and contaminating samples,

521 and the estimation methods.

Popul		Intended % Contam.	Equal- Ancestry	Unequal- Ancestry	Pooled AF	EUR AF	EAS AF	AFR AF
Intended	Contam.		(VB2)	(VB2)	(VB1)	(VB1)	(VB1)	(VB1)
		1%	1.0%	1.0%	0.8%	1.0%	0.6%	0.5%
		2%	1.9%	1.9%	1.5%	2.0%	1.3%	1.1%
GBR	GBR	5%	4.6%	4.6%	3.8%	4.7%	3.2%	2.7%
		10%	9.2%	9.2%	7.4%	9.4%	6.2%	5.2%
		20%	18.3%	18.3%	14.7%	18.5%	11.6%	9.5%
		1%	1.1%	1.0%	0.9%	1.2%	0.7%	0.6%
		2%	2.1%	1.9%	1.7%	2.2%	1.5%	1.2%
GBR	CHS	5%	5.2%	4.7%	4.3%	5.3%	3.8%	3.1%
		10%	10.1%	9.4%	8.6%	10.4%	7.6%	5.9%
		20%	19.8%	18.7%	17.3%	19.9%	15.1%	10.9%
		1%	1.3%	1.1%	1.1%	1.4%	0.8%	0.7%
	YRI	2%	2.5%	2.0%	2.1%	2.6%	1.7%	1.4%
GBR		5%	5.9%	4.8%	5.0%	6.1%	4.2%	3.6%
		10%	11.3%	9.5%	10.0%	11.7%	8.0%	7.3%
		20%	21.6%	19.1%	19.7%	22.0%	14.8%	14.6%
	GBR	1%	1.2%	0.9%	0.4%	0.2%	1.2%	0.1%
		2%	2.5%	1.8%	1.1%	0.8%	2.5%	0.5%
CHS		5%	6.1%	4.8%	3.6%	2.9%	6.3%	2.0%
		10%	12.0%	9.9%	7.9%	6.6%	12.5%	4.6%
		20%	23.0%	19.8%	16.6%	14.2%	23.6%	9.4%
		1%	0.9%	0.9%	0.2%	0.1%	0.9%	0.0%
		2%	1.8%	1.8%	0.7%	0.5%	1.8%	0.2%
CHS	CHS	5%	4.7%	4.8%	2.4%	1.9%	4.7%	1.3%
		10%	9.5%	9.5%	5.2%	4.2%	9.5%	2.9%
		20%	19.1%	19.1%	10.6%	8.4%	18.9%	5.9%
		1%	1.4%	0.9%	0.5%	0.3%	1.4%	0.1%
		2%	2.8%	1.9%	1.3%	1.0%	2.8%	0.5%
CHS	YRI	5%	6.7%	4.8%	4.1%	3.3%	6.9%	2.3%
		10%	12.9%	9.8%	8.9%	7.2%	13.5%	5.4%
		20%	24.3%	19.6%	18.6%	14.2%	24.9%	12.2%
YRI	GBR	1%	1.3%	1.0%	0.6%	0.4%	0.4%	1.4%

	_	2%	2.5%	1.9%	1.3%	0.9%	0.8%	2.6%
		5%	6.0%	4.6%	3.8%	2.8%	2.6%	6.4%
		10%	11.5%	9.3%	8.1%	6.2%	5.6%	12.5%
		20%	21.9%	18.8%	16.7%	13.0%	11.1%	23.5%
		1%	1.3%	0.9%	0.5%	0.4%	0.3%	1.3%
	CHS	2%	2.5%	1.9%	1.3%	0.9%	0.8%	2.6%
YRI		5%	6.0%	4.7%	3.8%	2.8%	2.5%	6.3%
		10%	11.5%	9.3%	7.9%	5.9%	5.6%	12.2%
		20%	22.0%	18.8%	16.6%	12.0%	12.1%	22.9%
		1%	0.9%	0.9%	0.4%	0.2%	0.2%	0.9%
		2%	1.8%	1.8%	0.9%	0.6%	0.5%	1.8%
YRI	YRI	5%	4.4%	4.4%	2.4%	1.8%	1.6%	4.4%
		10%	8.8%	8.8%	5.1%	3.8%	3.4%	8.9%
		20%	17.6%	17.6%	10.1%	7.3%	6.6%	17.9%

523

Equal-Ancestry Model:	Estimate from <i>verifyBamID2</i> assuming intended and contaminating samples have the same genetic ancestry (in PC coordinates)
Unequal-Ancestry Model:	Estimate from <i>verifyBamID2</i> allowing intended and contaminating samples to have different genetic ancestry
Pooled AF:	Estimate from original <i>verifyBamID</i> using allele frequency across all 1000 Genomes phase 3 samples
EUR AF:	Estimate from original <i>verifyBamID</i> using allele frequency across European subset of 1000 Genomes phase 3 samples
EAS AF:	Estimate from original verifyBamID using allele frequency across East Asian subset of 1000 Genomes phase 3 samples
AFR AF:	Estimate from original verifyBamID using allele frequency across Af- rican subset of 1000 Genomes phase 3 samples

Supplementary Table S2: Average of estimated contamination rates across 10 *in-silico* contaminated samples from Mexican population under different models. Results are similar as Europeans, except that unequal-ancestry model slightly reduces estimated contamination rate from equal-ancestry model, unlike GBR.

Intended % Contamination	Equal- Ancestry (VB2)	Unequal- Ancestry (VB2)	Pooled AF (VB1)	EUR AF (VB1)	EAS AF (VB1)	AFR AF (VB1)
1%	1.1%	1.0%	0.8%	1.0%	0.6%	0.3%
2%	2.1%	2.1%	1.6%	2.0%	1.4%	0.9%
5%	4.8%	4.8%	3.9%	4.6%	3.5%	2.5%
10%	9.3%	9.2%	7.8%	8.8%	6.8%	4.9%
20%	18.5%	18.3%	15.4%	17.0%	13.0%	9.4%

Sample Population		Marker	Intended Contamination Rate				
Intended	Contam.	Set	0.01	0.02	0.05	0.1	0.2
		1K	0.57(0.15)	0.88(0.38)	0.87(0.28)	0.92(0.18)	0.95(0.12)
GBR	GBR	10K	0.98(0.13)	0.95(0.11)	0.93(0.09)	0.91(0.08)	0.91(0.07)
		100K	1.00(0.10)	0.96(0.09)	0.93(0.08)	0.92(0.06)	0.91(0.05)
		1K	1.38(1.26)	1.09(0.63)	1.00(0.44)	0.95(0.41)	0.95(0.21)
CHS	CHS	10K	1.08(0.48)	1.03(0.26)	1.00(0.12)	1.01(0.08)	0.96(0.06)
		100K	0.89(0.12)	0.92(0.08)	0.95(0.07)	0.95(0.05)	0.96(0.04)
		1K	1.23(0.86)	0.92(0.46)	0.98(0.30)	0.95(0.16)	0.97(0.10)
YRI	YRI	10K	0.91(0.20)	0.87(0.17)	0.89(0.05)	0.88(0.04)	0.90(0.03)
		100K	0.94(0.08)	0.92(0.07)	0.88(0.04)	0.88(0.04)	0.88(0.03)
		1K	1.07(0.90)	1.03(0.61)	0.95(0.37)	0.97(0.22)	0.91(0.12)
CHS	YRI	10K	1.00(0.46)	0.99(0.22)	1.02(0.12)	1.02(0.08)	0.99(0.06)
		100K	0.88(0.14)	0.93(0.10)	0.96(0.06)	0.98(0.05)	0.98(0.04)
	CHS	1K	1.00(0.49)	1.00(0.35)	0.91(0.24)	1.00(0.17)	1.01(0.10)
YRI		10K	1.02(0.10)	1.00(0.07)	0.95(0.03)	0.94(0.03)	0.94(0.02)
		100K	0.94(0.15)	0.95(0.09)	0.93(0.05)	0.93(0.03)	0.94(0.03)
		1K	1.10(0.49)	1.10(0.28)	1.06(0.30)	0.98(0.18)	0.97(0.09)
GBR	YRI	10K	0.94(0.23)	0.98(0.10)	0.94(0.06)	0.93(0.04)	0.94(0.03)
		100K	1.07(0.09)	1.02(0.08)	0.97(0.06)	0.95(0.05)	0.95(0.04)
		1K	1.13(0.56)	0.78(0.36)	0.84(0.19)	0.93(0.11)	0.98(0.06)
YRI	GBR	10K	0.92(0.24)	0.89(0.15)	0.91(0.06)	0.93(0.05)	0.94(0.05)
		100K	0.95(0.15)	0.93(0.08)	0.93(0.08)	0.93(0.06)	0.94(0.06)
		1K	1.28(1.24)	1.12(0.70)	1.00(0.40)	0.95(0.21)	0.97(0.13)
CHS	GBR	10K	1.06(0.54)	1.01(0.33)	1.00(0.14)	1.00(0.07)	0.98(0.05)
		100K	0.91(0.06)	0.92(0.07)	0.97(0.07)	0.99(0.06)	0.99(0.05)
		1K	0.89(0.47)	0.83(0.42)	0.84(0.17)	0.91(0.14)	0.92(0.13)
GBR	CHS	10K	0.97(0.17)	0.93(0.11)	0.94(0.08)	0.94(0.06)	0.92(0.06)
		100K	1.01(0.12)	0.97(0.10)	0.93(0.08)	0.94(0.07)	0.94(0.06)

Supplementary Table S3: Comparison of mean contamination rate ratio (Estimated/Intended) using different size of marker set (under Unequal-Ancestry Model). The Numbers in parenthesis represent standard deviation.

527 Supplementary Table S4. A full table summarizing the contamination rate ratio (Estimated/Intended) across vari-

528 ous simulation parameters, populations, and estimation methods shown in Figure 3. 100K marker sets were used.

Sample P	opulation			Intended			
ln- tended	Contam.	Method	Allele Frequencies	% Contam	Mean	SD	MSE
GBR	GBR	VB1	AFR	1%	0.52	0.11	0.242
GBR	GBR	VB1	AFR	2%	0.53	0.09	0.223
GBR	GBR	VB1	AFR	5%	0.53	0.06	0.221
GBR	GBR	VB1	AFR	10%	0.52	0.05	0.237
GBR	GBR	VB1	AFR	20%	0.48	0.04	0.276
GBR	GBR	VB1	EUR	1%	1.04	0.10	0.012
GBR	GBR	VB1	EUR	2%	0.98	0.09	0.007
GBR	GBR	VB1	EUR	5%	0.95	0.07	0.008
GBR	GBR	VB1	EUR	10%	0.94	0.06	0.008
GBR	GBR	VB1	EUR	20%	0.92	0.05	0.008
GBR	GBR	VB1	EAS	1%	0.65	0.11	0.136
GBR	GBR	VB1	EAS	2%	0.65	0.09	0.132
GBR	GBR	VB1	EAS	5%	0.64	0.06	0.135
GBR	GBR	VB1	EAS	10%	0.62	0.05	0.148
GBR	GBR	VB1	EAS	20%	0.58	0.05	0.179
GBR	GBR	VB1	Pooled	1%	0.79	0.11	0.055
GBR	GBR	VB1	Pooled	2%	0.77	0.08	0.060
GBR	GBR	VB1	Pooled	5%	0.75	0.07	0.066
GBR	GBR	VB1	Pooled	10%	0.74	0.06	0.069
GBR	GBR	VB1	Pooled	20%	0.73	0.05	0.073
GBR	GBR	VB2	ISAF (Equal-Ancestry)	1%	1.02	0.11	0.010
GBR	GBR	VB2	ISAF (Equal-Ancestry)	2%	0.96	0.09	0.009
GBR	GBR	VB2	ISAF (Equal -Ancestry)	5%	0.93	0.07	0.010
GBR	GBR	VB2	ISAF (Equal -Ancestry)	10%	0.92	0.06	0.010
GBR	GBR	VB2	ISAF (Equal -Ancestry)	20%	0.91	0.05	0.010
GBR	GBR	VB2	ISAF (Unequal-Ancestry)	1%	1.00	0.10	0.009
GBR	GBR	VB2	ISAF (Unequal-Ancestry)	2%	0.96	0.09	0.009
GBR	GBR	VB2	ISAF (Unequal-Ancestry)	5%	0.93	0.08	0.011
GBR	GBR	VB2	ISAF (Unequal-Ancestry)	10%	0.92	0.06	0.010
GBR	GBR	VB2	ISAF (Unequal-Ancestry)	20%	0.91	0.05	0.010
GBR	CHS	VB1	AFR	1%	0.59	0.08	0.172
GBR	CHS	VB1	AFR	2%	0.60	0.06	0.162
GBR	CHS	VB1	AFR	5%	0.61	0.05	0.154
GBR	CHS	VB1	AFR	10%	0.59	0.04	0.169
GBR	CHS	VB1	AFR	20%	0.55	0.03	0.206

	GBR	CHS	VB1	EUR	1%	1.17	0.11	0.039
	GBR	CHS	VB1	EUR	2%	1.09	0.10	0.016
	GBR	CHS	VB1	EUR	5%	1.06	0.08	0.010
	GBR	CHS	VB1	EUR	10%	1.04	0.07	0.006
	GBR	CHS	VB1	EUR	20%	0.99	0.06	0.003
	GBR	CHS	VB1	EAS	1%	0.74	0.09	0.074
	GBR	CHS	VB1	EAS	2%	0.74	0.07	0.072
	GBR	CHS	VB1	EAS	5%	0.76	0.06	0.063
	GBR	CHS	VB1	EAS	10%	0.76	0.05	0.061
	GBR	CHS	VB1	EAS	20%	0.75	0.04	0.062
	GBR	CHS	VB1	Pooled	1%	0.89	0.09	0.019
	GBR	CHS	VB1	Pooled	2%	0.86	0.07	0.025
	GBR	CHS	VB1	Pooled	5%	0.86	0.06	0.022
	GBR	CHS	VB1	Pooled	10%	0.86	0.06	0.022
	GBR	CHS	VB1	Pooled	20%	0.86	0.05	0.021
	GBR	CHS	VB2	ISAF (Equal-Ancestry)	1%	1.13	0.11	0.028
	GBR	CHS	VB2	ISAF (Equal-Ancestry)	2%	1.06	0.09	0.011
	GBR	CHS	VB2	ISAF (Equal -Ancestry)	5%	1.03	0.08	0.007
	GBR	CHS	VB2	ISAF (Equal -Ancestry)	10%	1.01	0.07	0.004
	GBR	CHS	VB2	ISAF (Equal -Ancestry)	20%	0.99	0.06	0.004
	GBR	CHS	VB2	ISAF (Unequal-Ancestry)	1%	1.01	0.12	0.012
	GBR	CHS	VB2	ISAF (Unequal-Ancestry)	2%	0.97	0.10	0.010
	GBR	CHS	VB2	ISAF (Unequal-Ancestry)	5%	0.93	0.08	0.010
	GBR	CHS	VB2	ISAF (Unequal-Ancestry)	10%	0.94	0.07	0.008
_	GBR	CHS	VB2	ISAF (Unequal-Ancestry)	20%	0.94	0.06	0.007
	GBR	YRI	VB1	AFR	1%	0.67	0.11	0.119
	GBR	YRI	VB1	AFR	2%	0.70	0.09	0.096
	GBR	YRI	VB1	AFR	5%	0.72	0.06	0.082
	GBR	YRI	VB1	AFR	10%	0.73	0.05	0.077
	GBR	YRI	VB1	AFR	20%	0.73	0.04	0.074
	GBR	YRI	VB1	EUR	1%	1.38	0.10	0.150
	GBR	YRI	VB1	EUR	2%	1.30	0.09	0.097
	GBR	YRI	VB1	EUR	5%	1.23	0.07	0.055
	GBR	YRI	VB1	EUR	10%	1.17	0.05	0.032
	GBR	YRI	VB1	EUR	20%	1.10	0.04	0.011
	GBR	YRI	VB1	EAS	1%	0.85	0.10	0.032
	GBR	YRI	VB1	EAS	2%	0.85	0.08	0.028
	GBR	YRI	VB1	EAS	5%	0.83	0.06	0.031
	GBR	YRI	VB1	EAS	10%	0.80	0.04	0.042
	GBR	YRI	VB1	EAS	20%	0.74	0.03	0.069

GBR	YRI	VB1	Pooled	1%	1.05	0.09	0.010
GBR	YRI	VB1	Pooled	2%	1.03	0.08	0.007
GBR	YRI	VB1	Pooled	5%	1.01	0.06	0.003
GBR	YRI	VB1	Pooled	10%	1.00	0.05	0.002
GBR	YRI	VB1	Pooled	20%	0.99	0.04	0.002
GBR	YRI	VB2	ISAF (Equal-Ancestry)	1%	1.33	0.09	0.118
GBR	YRI	VB2	ISAF (Equal-Ancestry)	2%	1.26	0.09	0.074
GBR	YRI	VB2	ISAF (Equal -Ancestry)	5%	1.18	0.06	0.036
GBR	YRI	VB2	ISAF (Equal -Ancestry)	10%	1.13	0.05	0.018
GBR	YRI	VB2	ISAF (Equal -Ancestry)	20%	1.08	0.04	0.008
GBR	YRI	VB2	ISAF (Unequal-Ancestry)	1%	1.07	0.09	0.014
GBR	YRI	VB2	ISAF (Unequal-Ancestry)	2%	1.02	0.08	0.006
GBR	YRI	VB2	ISAF (Unequal-Ancestry)	5%	0.97	0.06	0.004
GBR	YRI	VB2	ISAF (Unequal-Ancestry)	10%	0.95	0.05	0.004
GBR	YRI	VB2	ISAF (Unequal-Ancestry)	20%	0.95	0.04	0.004
CHS	GBR	VB1	AFR	1%	0.07	0.04	0.868
CHS	GBR	VB1	AFR	2%	0.23	0.05	0.597
CHS	GBR	VB1	AFR	5%	0.40	0.04	0.366
CHS	GBR	VB1	AFR	10%	0.46	0.03	0.290
CHS	GBR	VB1	AFR	20%	0.47	0.02	0.282
CHS	GBR	VB1	EUR	1%	0.21	0.05	0.625
CHS	GBR	VB1	EUR	2%	0.40	0.05	0.364
CHS	GBR	VB1	EUR	5%	0.57	0.05	0.184
CHS	GBR	VB1	EUR	10%	0.66	0.04	0.119
CHS	GBR	VB1	EUR	20%	0.71	0.04	0.087
CHS	GBR	VB1	EAS	1%	1.24	0.11	0.069
CHS	GBR	VB1	EAS	2%	1.26	0.10	0.075
CHS	GBR	VB1	EAS	5%	1.27	0.08	0.077
CHS	GBR	VB1	EAS	10%	1.25	0.06	0.066
CHS	GBR	VB1	EAS	20%	1.18	0.05	0.035
CHS	GBR	VB1	Pooled	1%	0.36	0.06	0.409
CHS	GBR	VB1	Pooled	2%	0.55	0.06	0.207
CHS	GBR	VB1	Pooled	5%	0.71	0.05	0.086
CHS	GBR	VB1	Pooled	10%	0.79	0.05	0.047
CHS	GBR	VB1	Pooled	20%	0.83	0.04	0.031
CHS	GBR	VB2	ISAF (Equal-Ancestry)	1%	1.22	0.11	0.060
CHS	GBR	VB2	ISAF (Equal-Ancestry)	2%	1.23	0.10	0.060
CHS	GBR	VB2	ISAF (Equal -Ancestry)	5%	1.23	0.08	0.057
CHS	GBR	VB2	ISAF (Equal -Ancestry)	10%	1.20	0.06	0.044
CHS	GBR	VB2	ISAF (Equal -Ancestry)	20%	1.15	0.05	0.025

CLIC	CDD		ISAF (Unequal-Ancestry)	10/	0.01	0.00	0.011
CHS	GBR	VB2	ISAF (Unequal-Ancestry)	1% 2%	0.91	0.06	0.011
CHS	GBR	VB2	ISAF (Unequal-Ancestry)	2%	0.92	0.07	0.010
CHS	GBR	VB2	ISAF (Unequal-Ancestry)	5%	0.97	0.07	0.005
CHS	GBR	VB2	ISAF (Unequal-Ancestry)	10%	0.99	0.06	0.004
CHS	GBR	VB2	AFR	20%	0.99	0.05	0.003
CHS	CHS	VB1	AFR	1%	0.02	0.03	0.956
CHS	CHS	VB1	AFR	2%	0.12	0.05	0.782
CHS	CHS	VB1	AFR	5%	0.25	0.03	0.562
CHS	CHS	VB1	AFR	10%	0.29	0.02	0.500
CHS	CHS	VB1	EUR	20%	0.29	0.01	0.500
CHS	CHS	VB1		1%	0.10	0.07	0.815
CHS	CHS	VB1	EUR	2%	0.24	0.05	0.573
CHS	CHS	VB1	EUR	5%	0.38	0.03	0.385
CHS	CHS	VB1	EUR	10%	0.42	0.02	0.338
CHS	CHS	VB1	EUR	20%	0.42	0.02	0.337
CHS	CHS	VB1	EAS	1%	0.90	0.10	0.020
CHS	CHS	VB1	EAS	2%	0.92	0.07	0.011
CHS	CHS	VB1	EAS	5%	0.95	0.06	0.006
CHS	CHS	VB1	EAS	10%	0.95	0.04	0.004
CHS	CHS	VB1	EAS	20%	0.94	0.03	0.004
CHS	CHS	VB1	Pooled	1%	0.19	0.08	0.659
CHS	CHS	VB1	Pooled	2%	0.34	0.05	0.435
CHS	CHS	VB1	Pooled	5%	0.48	0.04	0.275
CHS	CHS	VB1	Pooled	10%	0.52	0.03	0.233
CHS	CHS	VB1	Pooled	20%	0.53	0.02	0.222
CHS	CHS	VB2	ISAF (Equal-Ancestry)	1%	0.90	0.11	0.021
CHS	CHS	VB2	ISAF (Equal-Ancestry)	2%	0.91	0.07	0.012
CHS	CHS	VB2	ISAF (Equal -Ancestry)	5%	0.95	0.06	0.006
CHS	CHS	VB2	ISAF (Equal -Ancestry)	10%	0.95	0.04	0.004
CHS	CHS	VB2	ISAF (Equal -Ancestry)	20%	0.95	0.03	0.003
CHS	CHS	VB2	ISAF (Unequal-Ancestry)	1%	0.89	0.12	0.024
CHS	CHS	VB2	ISAF (Unequal-Ancestry)	2%	0.92	0.08	0.012
CHS	CHS	VB2	ISAF (Unequal-Ancestry)	5%	0.95	0.07	0.006
CHS	CHS	VB2	ISAF (Unequal-Ancestry)	10%	0.95	0.05	0.004
CHS	CHS	VB2	ISAF (Unequal-Ancestry)	20%	0.96	0.04	0.003
CHS	YRI	VB1	AFR	1%	0.09	0.09	0.828
CHS	YRI	VB1	AFR	2%	0.27	0.08	0.543
CHS	YRI	VB1	AFR	5%	0.45	0.05	0.302
CHS	YRI	VB1	AFR	10%	0.54	0.04	0.210
CHS	YRI	VB1	AFR	20%	0.61	0.03	0.155

CHS	YRI	VB1	EUR	1%	0.31	0.11	0.492	
CHS	YRI	VB1	EUR	2%	0.51	0.08	0.250	
CHS	YRI	VB1	EUR	5%	0.67	0.05	0.114	
CHS	YRI	VB1	EUR	10%	0.72	0.04	0.083	
CHS	YRI	VB1	EUR	20%	0.71	0.03	0.084	
CHS	YRI	VB1	EAS	1%	1.38	0.14	0.161	
CHS	YRI	VB1	EAS	2%	1.39	0.10	0.163	
CHS	YRI	VB1	EAS	5%	1.38	0.07	0.151	
CHS	YRI	VB1	EAS	10%	1.35	0.06	0.123	
CHS	YRI	VB1	EAS	20%	1.24	0.04	0.061	
CHS	YRI	VB1	Pooled	1%	0.47	0.13	0.298	
CHS	YRI	VB1	Pooled	2%	0.66	0.09	0.123	
CHS	YRI	VB1	Pooled	5%	0.82	0.06	0.036	
CHS	YRI	VB1	Pooled	10%	0.89	0.05	0.015	
CHS	YRI	VB1	Pooled	20%	0.93	0.04	0.007	
CHS	YRI	VB2	ISAF (Equal-Ancestry)	1%	1.37	0.14	0.157	
CHS	YRI	VB2	ISAF (Equal-Ancestry)	2%	1.38	0.11	0.154	
CHS	YRI	VB2	ISAF (Equal -Ancestry)	5%	1.34	0.07	0.122	
CHS	YRI	VB2	ISAF (Equal -Ancestry)	10%	1.29	0.06	0.085	
CHS	YRI	VB2	ISAF (Equal -Ancestry)	20%	1.22	0.05	0.048	
CHS	YRI	VB2	ISAF (Unequal-Ancestry)	1%	0.88	0.14	0.033	
CHS	YRI	VB2	ISAF (Unequal-Ancestry)	2%	0.93	0.10	0.014	
CHS	YRI	VB2	ISAF (Unequal-Ancestry)	5%	0.96	0.06	0.005	
CHS	YRI	VB2	ISAF (Unequal-Ancestry)	10%	0.98	0.05	0.002	
CHS	YRI	VB2	ISAF (Unequal-Ancestry)	20%	0.98	0.04	0.002	
YRI	GBR	VB1	AFR	1%	1.43	0.23	0.236	
YRI	GBR	VB1	AFR	2%	1.31	0.14	0.113	
YRI	GBR	VB1	AFR	5%	1.28	0.08	0.081	
YRI	GBR	VB1	AFR	10%	1.25	0.07	0.065	
YRI	GBR	VB1	AFR	20%	1.17	0.06	0.034	
YRI	GBR	VB1	EUR	1%	0.36	0.22	0.453	
YRI	GBR	VB1	EUR	2%	0.46	0.19	0.329	
YRI	GBR	VB1	EUR	5%	0.56	0.11	0.201	
YRI	GBR	VB1	EUR	10%	0.62	0.07	0.153	
YRI	GBR	VB1	EUR	20%	0.65	0.05	0.124	
YRI	GBR	VB1	EAS	1%	0.32	0.20	0.504	
YRI	GBR	VB1	EAS	2%	0.41	0.18	0.380	
YRI	GBR	VB1	EAS	5%	0.52	0.11	0.245	
YRI	GBR	VB1	EAS	10%	0.55	0.07	0.204	
YRI	GBR	VB1	EAS	20%	0.56	0.04	0.198	

YRI	GBR	VB1	Pooled	1%	0.57	0.26	0.248
YRI	GBR	VB1	Pooled	2%	0.67	0.17	0.138
YRI	GBR	VB1	Pooled	5%	0.76	0.10	0.066
YRI	GBR	VB1	Pooled	10%	0.80	0.07	0.043
YRI	GBR	VB1	Pooled	20%	0.83	0.05	0.030
YRI	GBR	VB2	ISAF (Equal-Ancestry)	1%	1.30	0.15	0.107
YRI	GBR	VB2	ISAF (Equal-Ancestry)	2%	1.25	0.11	0.072
YRI	GBR	VB2	ISAF (Equal -Ancestry)	5%	1.20	0.09	0.048
YRI	GBR	VB2	ISAF (Equal -Ancestry)	10%	1.15	0.07	0.028
YRI	GBR	VB2	ISAF (Equal -Ancestry)	20%	1.10	0.07	0.013
YRI	GBR	VB2	ISAF (Unequal-Ancestry)	1%	0.95	0.15	0.021
YRI	GBR	VB2	ISAF (Unequal-Ancestry)	2%	0.93	0.08	0.012
YRI	GBR	VB2	ISAF (Unequal-Ancestry)	5%	0.93	0.08	0.011
YRI	GBR	VB2	ISAF (Unequal-Ancestry)	10%	0.93	0.06	0.008
YRI	GBR	VB2	ISAF (Unequal-Ancestry)	20%	0.94	0.06	0.006
YRI	CHS	VB1	AFR	1%	1.32	0.13	0.120
YRI	CHS	VB1	AFR	2%	1.30	0.11	0.098
YRI	CHS	VB1	AFR	5%	1.26	0.07	0.071
YRI	CHS	VB1	AFR	10%	1.22	0.05	0.049
YRI	CHS	VB1	AFR	20%	1.14	0.04	0.022
YRI	CHS	VB1	EUR	1%	0.35	0.20	0.455
YRI	CHS	VB1	EUR	2%	0.46	0.18	0.319
YRI	CHS	VB1	EUR	5%	0.56	0.10	0.204
YRI	CHS	VB1	EUR	10%	0.59	0.06	0.168
YRI	CHS	VB1	EUR	20%	0.60	0.03	0.159
YRI	CHS	VB1	EAS	1%	0.29	0.17	0.528
YRI	CHS	VB1	EAS	2%	0.40	0.18	0.395
YRI	CHS	VB1	EAS	5%	0.51	0.11	0.252
YRI	CHS	VB1	EAS	10%	0.56	0.06	0.194
YRI	CHS	VB1	EAS	20%	0.61	0.03	0.157
YRI	CHS	VB1	Pooled	1%	0.55	0.24	0.259
YRI	CHS	VB1	Pooled	2%	0.66	0.16	0.136
YRI	CHS	VB1	Pooled	5%	0.75	0.09	0.068
YRI	CHS	VB1	Pooled	10%	0.79	0.04	0.044
YRI	CHS	VB1	Pooled	20%	0.83	0.02	0.030
YRI	CHS	VB2	ISAF (Equal-Ancestry)	1%	1.29	0.13	0.098
YRI	CHS	VB2	ISAF (Equal-Ancestry)	2%	1.25	0.11	0.074
YRI	CHS	VB2	ISAF (Equal -Ancestry)	5%	1.20	0.07	0.043
YRI	CHS	VB2	ISAF (Equal -Ancestry)	10%	1.15	0.04	0.023
YRI	CHS	VB2	ISAF (Equal -Ancestry)	20%	1.10	0.04	0.011

YR	I CHS	VB2	ISAF (Unequal-Ancestry)	1%	0.94	0.15	0.023
YR	I CHS	VB2	ISAF (Unequal-Ancestry)	2%	0.95	0.09	0.010
YR	I CHS	VB2	ISAF (Unequal-Ancestry)	5%	0.93	0.05	0.007
YR	I CHS	VB2	ISAF (Unequal-Ancestry)	10%	0.93	0.03	0.005
YR	I CHS	VB2	ISAF (Unequal-Ancestry)	20%	0.94	0.03	0.004
YR	I YRI	VB1	AFR	1%	0.95	0.09	0.010
YR	I YRI	VB1	AFR	2%	0.92	0.06	0.009
YR	I YRI	VB1	AFR	5%	0.89	0.05	0.014
YR	I YRI	VB1	AFR	10%	0.89	0.04	0.013
YR	I YRI	VB1	AFR	20%	0.89	0.03	0.012
YR	I YRI	VB1	EUR	1%	0.22	0.13	0.619
YR	I YRI	VB1	EUR	2%	0.29	0.14	0.518
YR	I YRI	VB1	EUR	5%	0.36	0.09	0.423
YR	I YRI	VB1	EUR	10%	0.38	0.06	0.391
YR	I YRI	VB1	EUR	20%	0.36	0.03	0.405
YR	I YRI	VB1	EAS	1%	0.18	0.11	0.680
YR	I YRI	VB1	EAS	2%	0.25	0.13	0.575
YR	I YRI	VB1	EAS	5%	0.32	0.09	0.474
YR	I YRI	VB1	EAS	10%	0.34	0.06	0.438
YR	I YRI	VB1	EAS	20%	0.33	0.03	0.452
YR	I YRI	VB1	Pooled	1%	0.36	0.18	0.433
YR	I YRI	VB1	Pooled	2%	0.44	0.14	0.333
YR	I YRI	VB1	Pooled	5%	0.49	0.08	0.267
YR	I YRI	VB1	Pooled	10%	0.51	0.05	0.242
YR	I YRI	VB1	Pooled	20%	0.51	0.03	0.245
YR	I YRI	VB2	ISAF (Equal-Ancestry)	1%	0.94	0.10	0.012
YR	I YRI	VB2	ISAF (Equal-Ancestry)	2%	0.92	0.06	0.011
YR	I YRI	VB2	ISAF (Equal -Ancestry)	5%	0.88	0.04	0.016
YR	I YRI	VB2	ISAF (Equal -Ancestry)	10%	0.88	0.04	0.015
YR	I YRI	VB2	ISAF (Equal -Ancestry)	20%	0.88	0.03	0.015
YR	I YRI	VB2	ISAF (Unequal-Ancestry)	1%	0.94	0.08	0.010
YR	I YRI	VB2	ISAF (Unequal-Ancestry)	2%	0.92	0.07	0.011
YR	I YRI	VB2	ISAF (Unequal-Ancestry)	5%	0.88	0.04	0.016
YR	I YRI	VB2	ISAF (Unequal-Ancestry)	10%	0.88	0.04	0.015
YR	I YRI	VB2	ISAF (Unequal-Ancestry)	20%	0.88	0.03	0.015