Contribution of Retrotransposition to Developmental Disorders

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Abstract

Mobile genetic Elements (MEs) are segments of DNA which, through an RNA intermediate, can generate new copies of themselves and other transcribed sequences through the process of retrotransposition (RT). In humans several disorders have been attributed to RT, but the role of RT in severe developmental disorders (DD) has not yet been explored. As such, we have identified RT-derived events in 9,738 exome sequenced trios with DD-affected probands as part of the Deciphering Developmental Disorders (DDD) study. We have ascertained 9 *de novo* MEs, 4 of which are likely causative of the patient's symptoms (0.04% of probands), as well as 2 *de novo* gene retroduplications. Beyond identifying likely diagnostic RT events, we have estimated genome-wide germline ME mutagenesis and constraint and demonstrated that coding RT events have signatures of purifying selection equivalent to those of truncating mutations. Overall, our analysis represents a comprehensive interrogation of the impact of retrotransposition on protein coding genes and a framework for future evolutionary and disease studies.

Main

In humans, three classes of Mobile genetic Elements (MEs) – *Alu*, long interspersed nuclear element 1 (L1), and SINE-VNTR-*Alu* (SVA) – are still active and can generate new copies, known as Mobile Element Insertions (MEIs), throughout their host genome[1]. The L1 replicative machinery can also facilitate the duplication of non-ME transcripts, typically protein-coding genes, through the mechanism of retroduplication to generate processed pseudogenes (PPGs)[2]. Combined, these two processes constitute retrotransposition (RT) in the human genome, with new (*de novo*) MEI variants previously estimated to occur in every 1 out of 18.4 to 26.0 births[3]. On a population level, each individual human genome harbors ~1,200 polymorphic variants, with the smallest ME, *Alu*, generally contributing 75% of total RT polymorphisms[4-6].

To date roughly 130 pathogenic variants caused by RT activity have been documented[7]; however, the majority of these deleterious events have been discovered in isolated cases. Neither MEIs nor PPGs are analyzed as part of routine clinical sequencing and thus represent a largely unassessed category of genetic variation in many disorders. Furthermore, of the clinically relevant RT-attributable cases thus identified, few (~14/123; 11.4%) are caused by new mutational events and are instead typically attributable to rare inherited polymorphisms[7]. Additionally, of the large disease-focused whole genome sequencing (WGS) projects which have ascertained MEIs, all have focused on autism[8,9] and have failed to identify likely causative RT-derived variants. In fact, in the largest and most recent WGS study investigating the role of large structural variants in the genetic architecture of autism, the authors failed to identify a single *de novo* MEI in a coding exon, deleterious or otherwise, in 829 families[9]. This finding is likely a result of several factors, predominant among them the low frequency of cases attributable to gene disruption by MEIs in autism[10], due in part to a low ME mutation rate[3] and lack of a sufficiently large sample size[8,9,11]. As such, it is not precisely known at what rate *de novo* ME variants are generated in the human genome, the functional consequences of such variants, the role that they play

in the etiology of rare disease, and if routine clinical sequencing should assess patient

genomes for deleterious RT events.

We analyzed the whole exome sequencing (WES) data produced by the Deciphering

Developmental Disorders (DDD) study to systematically assess the role of RT in severe

developmental disorders (DDs). The DDD data have already been investigated for

pathogenic single nucleotide variants (SNVs), small insertions and deletions (InDels), large

copy number variants (CNVs), and other classes of structural variation[12-18]. Approximately

24% of DDD cases harbour a pathogenic *de novo* mutation in a gene known to be

associated with developmental disorders[12]. The DDD cohort should thus be relatively

enriched for highly penetrant *de novo* RT events in comparison to recent studies on

autism[8,9]. With a cohort of 9,738 trios (n = 28,132 individuals) whole exome sequenced, the

DDD study presents a powerful opportunity to identify, and ascertain the role in DD of,

pathogenic *de novo* RT events that impact coding sequences.

Results

Generation of a genome-wide dataset of RT variants

To assess 9,738 DDD study trios for RT events we utilized two separate computational approaches to identify both MEIs and PPGs. First, we used the Mobile Element Locator Tool (MELT)[5] to identify *Alu*, L1, and SVA variants located within the WES bait regions (Methods). The second is a new bespoke tool developed to identify PPGs from WES data (Methods, Supplemental Fig. 1). Due to cross-hybridisation between a PPG and the exome baits targeting the donor gene, we anticipated that we should be able to detect PPGs genome-wide, not just the subset that insert within the WES bait regions. Our PPG detection tool ascertained putative PPGs by identifying multiple discordant read pairs mapping to different exons of the same transcript, before then typing all individuals for the presence/absence of the PPG using discordant read-pairs and split reads. The tool was optimized by comparing against previously described PPG polymorphisms in the 1000 genomes project (1KGP; see below).

	Total Sites Cohort-Wide	Mean Sites Per Unaffected Parent	Total de novo Sites
Alu	917	23.6±4.2	7
LINE-1	167	2.8±1.5	2
SVA	45	0.2±0.5	0
Total - MEI	*1,129*	*26.6±4.7*	*9*
Processed Pseudogenes (PPGs)	576	6.9±2.6	2
Total – MEI + PPG	*1,705*	*33.5±7.3*	*11*

Table 1: RT variant discovery in the DDD

Quantification of the four different classes of retrotransposons discovered as part of this study.

Grey-highlighted rows indicate totals across the classes listed above.

As our study is the first to discover MEIs directly from WES on a large scale with

MELT, we first utilized matched sample WGS data to determine if MELT could ascertain MEI

variants reliably from WES data. We compared MEI variants identified by MELT within the

DDD WES data to both WGS data generated for 90 overlapping DDD individuals and

population MEI data previously generated from the 1000 Genomes Project Phase 3

(1KPG)[4-6] WGS data. The latter comparison was to ensure that the number of exonic MEIs

identified within DDD WES data was concordant with expectations at the individual and

population level. When comparing our WES genotypes to WGS in the same individuals, we

had a genotype concordance rate of 94.46% (93.93% *Alu*, 97.29% L1, 98.25% SVA) among

calls with at least 10X coverage in our WES data. In total, we were able to re-identify 1,355

(1,289 *Alu*, 160 L1, 1 SVA) MEI genotypes, or 84.5% of all heterozygous or homozygous

genotypes identifiable with WGS in WES bait regions (Supplemental Table 1). Based on

these findings we were confident that MELT was appropriately calibrated to ascertain MEIs

in WES data.

We identified 1,129 MEI variants and 576 polymorphic PPGs, with each individual's

exome containing on average 33.5±7.3 variants (Table 1). All MEIs were genotyped across

all individuals to form a comprehensive catalogue of RT-derived variation within and

adjacent to (±50 bp) sequences targeted in the WES assay (Methods), including coding

exons and targeted non-coding elements (Table 1; Fig. 1). The average time to assess a

single family for RT-derived events was approximately 15 minutes and the rate of false

findings was low (1 incorrect candidate *de novo* variant per every 295 patients; either a false

positive variant [1 per 649] or false negative genotype [1 per 541] in at least one parent). We

investigated whether false negative parental genotypes could be mosaic. Using a long-PCR

coupled to long-read sequencing assay (see methods), we were able to phase four of our

false-negative *Alu* variants to nearby SNVs. With our approach, we could demonstrate two

insertions were mosaic (chr3:g.101378585_101378586ins*Alu*(1_280),

chr15:g.91145646_91145647ins*Alu*(1_281)), one was part of a complex event of unknown

mosaicism (*FCGBP*:c.4743_4744ins*Alu*(1_110)), and another constitutive. Therefore these

false apparent *de novo* calls can be generated by a variety of causes, including parental

mosaicism (Supplemental Fig. 2).

As expected, the total number of variants per individual for each RT class (Fig. 1a-d)

as well as combined number of RT events (Fig. 1e) approximated a Poisson distribution (see

methods; Supplemental Fig. 3). The vast majority of variants are rare (allele frequency <

1×10^{-4}; Fig. 1f), with >65% of *Alu* and L1 variants identified in fewer than 4 unrelated

individuals. SVA and PPGs appear to be moderately under ascertained compared to *Alu* and

L1 at lower AFs, with >50% of variants identified in the lowest AF bin and length estimates

for the three MEI classes largely fit the findings of previous studies (Fig. 1g)[4,5].

We next sought to ensure that our total number of ascertained RT variants, both on a

population and individual basis, accorded with previously published WGS data[4,5]. On a

population level, WES did not appreciably limit our overall sensitivity compared to WGS

sampled data. When we compared a downsampled version of our call set to the 1KGP, our

total number of *Alu* and SVA variants fell within the expected distribution, while L1 was close

to expectation (Supplemental Fig. 4).

Figure 1: The DDD RT call set

(**a-e**) Histograms of total number of variants per individual for the four classes of RT events

identified in the DDD cohort (*Alu* – blue; L1 – green; SVA – orange; PPGs – red; combined RT

events – grey) in size one bins. (**f**) Allele frequency distributions for the RT classes depicted in **a-**

e in $\log_{10}$ allele frequency bins. (**g**) Insert size estimates provided by MELT for the MEI classes

ascertained in this study in $\log_{10}$ insert size bins. All plots only include variants from unaffected

parents.

To assess the quality of the PPG call set, we compared PPG allele counts (i.e. total

number of individuals with a retro-duplication of a given gene) to a recent assessment of

PPGs in samples sequenced as part of the 1KGP[6]. Generally, PPGs identified in both data

sets shared similar relative allele counts (r^2 = 0.64) and variants identified in this study but

missing from Zhang et. al.[6] are typically rare (Supplemental Fig. 5). To further validate our

approach and ensure that the identified PPG donor genes fit with previously identified

patterns of germline PPG formation[2,19], we assessed each donor gene for both functional

annotation and expression across 30 tissue types analyzed by the GTEx consortium[20]. The

major functional cluster (DAVID[21] enrichment score 8.82) belonged to genes involved in the

ribosomal and translational machinery, consistent with previous findings involving fixed

PPGs in the human genome[2]. Our expression analysis likewise confirmed previous

findings[19], and shows that donor genes that give rise to PPGs are more highly expressed in

a large number of tissues compared to non-retroposed genes (Wilcoxon rank sum $p < 1\times10^{-3}$

for all tissues; Supplemental Fig. 6). Additionally, while it could be assumed that increased

germ-line expression of a gene may play a role in increased probability of PPG generation,

when we compared PPG donor gene expression in the testis and ovary to that in other

tissues, the majority of tissues (20/28, identical tissues for ovary and testis) showed

statistically identical patterns of donor gene expression (Wilcoxon rank sum $p > 1\times10^{-3}$;

Supplemental Fig. 6).

Coding RT burden and constraint

As expected for WES, the vast majority (84.9%) of detected MEIs impacted the

coding or intronic sequence of a protein-coding gene or a regulatory element targeted in the

augmented WES assay described in Short et. al.[15] (Fig. 2a). Due to the large numbers of

individuals with WES data in this study, we have ascertained over five-fold more exonic

variants than the largest previously published study (Supplemental Fig. 7)[4,5]. Nonetheless,

the number of MEIs identified in this study, based on the proportion of the genome assayed,

likely represent only 2.2% of MEI variants genome-wide in these individuals.

Our large collection of coding variants allowed us to examine the evolutionary forces

acting on coding MEI variation (Fig. 2b). To examine selective constraint, we utilized two

common measures: the proportion of variants observed in only one individual (e.g.

singletons)[22] and the proportion of variants found in genes likely to be intolerant of loss of

function (LoF) as determined by the pLI score[23]. To avoid issues of relatedness and the

potential for clinical ascertainment bias for pathogenic MEIs in individual DD patients, only

the 17,032 unaffected parents sequenced as part of DDD were included in our analysis.

MEIs which directly impact exons are under strong selective constraint, indistinguishable

from that of both nonsense and essential splice site SNVs (Fig. 2b). Interestingly, we did not

find any sign of selection acting on intronic MEIs as they appear to be constrained similarly

to synonymous SNVs. In contrast to previous studies[24,25], we did not find a statistically

significant (χ^2 p < 0.05) bias towards intronic MEIs inserted in the antisense orientation of

the gene in which they are found (Supplemental Fig. 8). This is likely not a repudiation of

such work, but attributable to the relatively small number of intronic events we identified as

part of our analysis compared to WGS[4,24] or reference genome-based[25] studies. To put our

findings on exonic MEI constraint into perspective with other forms of variation, every human

genome will harbor approximately one (0.76±0.62 per individual) MEI which directly impacts

protein-coding sequence. Since MEIs are similar to nonsense SNVs in terms of

deleteriousness (Fig. 2), MEIs thus make up 0.6%[22,26] of all coding protein truncating

variants (PTVs; among SNVs, InDels, and large CNVs) in each individual human genome.

Figure 2: Coding constraint on MEIs

(**a**) Cumulative consequence annotations for *Alu*, L1, and SVA MEIs. The majority of variants

identified in this study fell within the non-coding space (either an enhancer or intron) (**b**) Comparison

of constraint between MEIs and SNVs in unaffected parents. To compare the impact of exonic and

intronic *Alu* (blue) and all MEIs (grey) to varying classes of SNVs (black), we used two metrics: the

proportion of variants in genes that have been identified as LoF intolerant as gauged by pLI-score[22]

(x-axis) and the proportion of variants identified in only one individual (i.e. singletons; y-axis). Error

bars indicate 95% confidence intervals based on population proportion; confidence intervals were

calculated for SNVs, but are too small to appear at the resolution displayed in this figure.

While we were unable to perform similar population genetic analyses for PPG

events, due to the difficulty of resolving the putative insertion site with WES data and thus

distinguishing between different PPGs for the same donor gene, we were able to assess the

propensity for specific genes to give rise to PPGs based on their selective constraint. We

observed that PPG donor genes were significantly enriched for genes that are highly

intolerant of loss of function variation (pLI > 0.9). High pLI genes make up 25.3% of donor

genes, compared to 17.6% of all protein-coding genes (χ^2 p = 2.4x10^{-6})[22]. This observation is

likely driven by loss of function intolerant genes being more likely to be highly expressed in

multiple tissues[22], similar to genes known to have been retroduplicated (Supplemental Fig.

6)[19]. This observation implies that PPG events rarely strongly perturb the function of their

donor gene – despite several previously documented instances of PPGs impacting

expression or functionality of their donor gene[27].

Discovery and clinical annotation of *de novo* RT variants in DD

Using the computational approaches outlined above we identified a total of 11 germ-

line *de novo* RT variants (Table 2). Our findings include coding, noncoding, pathogenic and

benign variants, as well as, to our knowledge, the first *de novo* MEI identified in a pair of

monozygotic twins (Supplemental Fig. 9). All *de novo* RT variants were confirmed via a PCR

assay specific to the RT class (Fig. 3; Supplemental Fig. 9; Supplemental Table 2) and,

where possible, inspected for poly(A) tail and target site duplication – hallmarks of *bona fide*

RT activity[28]. We identified no *de novo* RT variants which localized to the non-coding

elements included on the WES capture, which falls in line with expectations based on

mutation rate estimates (Fig. 4b). We also attempted to determine the parental origin of each

RT event using SNVs located on sequencing reads which support the RT insertion (Table 2).

Of the 11 *de novo* RT events, we were able to phase three variants, all to the father. While

this finding is not statistically significant (χ^2 p = 0.083), it fits with previous findings that the

majority of *de novo* structural variants[9], and indeed most variant classes[29], are attributable to

paternal origin.

Nine of our validated *de novo* mutations were MEIs (7 *Alu*, 2 L1), or a rate of

approximately one *de novo* event per every 1,000 patient exomes sequenced (9/9,738). As

expected, based on both the total number of polymorphisms[3-5] and mutation rate (Table 1;

Supplemental Table 4), we identified more *Alu de novo* variants than the other RT classes.

We also identified 2 PPG germ-line *de novo* variants, or approximately one new PPG per

every 5,000 patient whole genomes sequenced (2/9,738). As a further quality control for

PPGs, we capillary sequenced all resulting PCR products to confirm the gene of origin

(Supplemental Data 1) and performed WGS to identify the PPG insertion site. We were able

to localize the *SERINC5* PPG to an ~50Kbp intron of the gene *CLIC4* and the *SLC35F2*

event to an intergenic region between the genes *MAK* and *GCM2* (Fig. 3e). Neither of the

events directly impacted coding sequence and *CLIC4* is neither under strong selective

constraint nor known to have any link with DD.

Insertion Coord.	RT Type	Genomic Compartment	ENSEMBL Gene ID	HGNC Gene ID	pLI	DDG2P Annotation	Decipher ID[30]	Diagnostic?	Parental Origin	Notes
chr3:9495459	*Alu*	Exonic	ENSG00000168137	*SETD5*	1.000	confirmed,monoallelic	280818	True	Father	
chr5:176638159	*Alu*	Exonic	ENSG00000165671	*NSD1*	1.000	confirmed,monoallelic	259118	True	Unknown	Included in Wright et. al.[31]
chr6:159190834	*Alu*	Exonic	ENSG00000092820	*EZR*	0.988	None	300984	False	Unknown	
chr7:77552086	*Alu*	Exonic	ENSG00000006576	*PHTF2*	0.024	None	271388	False	Father	
chr3:135913800	*Alu*	Intronic	ENSG00000174579	*MSL2*	0.890	None	292325	False	Unknown	
chr3:148614204	*Alu*	Intronic	ENSG00000163751	*CPA3*	<0.001	None	270426; 270428	False	Unknown	Monozygotic twins
chr3:172480619	*Alu*	Intronic	ENSG00000114346	*ECT2*	<0.001	None	307591	False	Unknown	
chr12:46246325	L1	Exonic	ENSG00000189079	*ARID2*	1.000	probable,monoallelic	264759	True	Unknown	
chr5:88100580	L1	Exonic	ENSG00000081189	*MEF2C*	0.004	confirmed,monoallelic	285645	True	Unknown	
chr6:10847968	Retrogene-SLC35F2	Intergenic	N/A	N/A	N/A	N/A	291670	False	Unknown	
chr1:25074202	Retrogene-SERINC5	Intronic	ENSG00000169504	*CLIC4*	0.009	None	301168	False	Father	

Table 2: Confirmed germ-line de novo variants in the DDD study

Relevant clinical and annotation information for MEI and PPG d*e novo* variants identified as part of this study. Location of the insertion event is given in hg19

reference coordinates (Insertion Coord.). A "True" value in the "Diagnostic" column indicates, at the time of publication, that this variant intersected a known

DD gene and was deemed likely to be involved in the patient's phenotype by the referring clinician. "False" does not indicate whether or not, with additional

future evidence, the gene may become associated with DD and the variant thus deemed diagnostically relevant. If applicable, ENSEMBL[32] gene IDs indicate

the gene impacted, not the gene from which the event is derived (i.e. for PPGs).

Each *de novo* mutation was then compared to known DD-associated genes (using

the Developmental Disorders Genotype-to-Phenotype database – DDG2P) to identify

potentially pathogenic variants (Table 2). Of the mutations identified, four directly inserted

into coding exons of DD-associated and fetal brain-expressed (see methods; Supplemental

Table 5)[33] genes (Fig. 3, Table 2) with all four found in genes statistically enriched for PTVs[12]

and therefore likely to operate by a LoF mechanism. We did not identify any intronic *de novo*

mutations likely to be pathogenic (Fig. 3a-d; Supplemental Fig. 9). An additional mutation

inserted into the coding sequence of a strongly LoF-intolerant gene, *EZR* (pLI = 0.99;

Supplemental Fig. 9), but we could not directly attribute it to the patient's phenotype due to

lack of significant enrichment for PTVs, although there is prior evidence for a role in a

familial DD syndrome[34]. The four mutations in DD-associated genes were reported to the

referring clinician for clinical interpretation based on both initially reported and updated

phenotypes (Supplemental Table 3). Three out of four reported mutations (*NSD1*, *MEF2C*,

ARID2) were subsequently deemed to be likely causative of the patient's phenotype

(Supplemental Table 3) by the referring clinician. To determine if our patient with a likely

pathogenic *NSD1* variant phenocopies the already described aberrant blood DNA

methylation signature typical of Sotos syndrome patients[35], we compared our patient's

genome-wide methylation status to several already published Sotos syndrome cases (see

methods)[35]. Our patient strongly clusters with other confirmed Sotos syndrome patients and

confirms our clinical assessment and diagnosis of this particular patient (Supplemental Fig.

10). The fourth patient, with an *Alu* insertion in *SETD5* (Fig. 3a), has clinical features

(polydactyly and truncal obesity; Supplemental Table 3) more suggestive of a ciliopathy. As

such, the identified MEI is unlikely to be the sole cause for the patient's DD but may

contribute to a composite phenotype. Accordingly, we did investigate this patient's exome for

other potentially diagnostic variants, but did not identify additional *de novo* or inherited

variation that could plausibly be associated with this patient's phenotype.

Figure 3: RT-derived de novos in the DDD

We identified a total of nine *de novo* MEIs, four of which disrupted the protein-coding sequence of a

known DD gene: (**a**) *SETD5*, (**b**) *MEF2C*, (**c**) *ARID2*, and (**d**) *NSD1*. Shown in each panel is a

diagram of the affected gene (blue model) with the relevant insertion indicated with a colored bubble.

To the right are PCR validations confirming the *de novo* status of each mutation; a positive result is

indicated by a raised secondary band present only in the proband sample (red arrow). (**e**) Circos

diagram and PCR results for two identified germ-line *de novo* PPGs. For each *de novo* PPG shown is

a diagram of the donor gene (gene model), location of duplication as PPG (directional arrow), and

new insertion site. Exons from the donor gene included in the PPG are indicated by brackets

underneath the donor gene model. To confirm PPG presence, PCR was performed (Methods) on

proband, paternal, and maternal gDNA (sample in each lane is shown by pedigree). The band which

represents the PPG is marked with a red arrow and was confirmed via capillary sequencing

(Supplemental Data 1). Dashed lines indicate intergenic regions, all genes models are shown in

sense orientation, and PPG gene diagrams are not to scale.

We also examined our dataset for inherited rare pathogenic RT variants. We

evaluated variants inherited from an affected parent, bi-allelic inheritance (either a

homozygous MEI or a heterozygous MEI paired with another variant class), and X-linked

variants maternally inherited by affected males. We did not identify any rare MEI variants

inherited from an affected parent nor any compound heterozygous individuals with a rare

MEI and a non-MEI PTV (e.g. SNV/InDel) impacting the same gene. We did identify a single

proband-specific homozygous MEI inserted into an exon of *PAN2* which was unique to a

single family. This gene was recently identified as nominally significant (genome-wide p =

4.2×10^{-4}) in a study investigating the role of recessive variants in DD[13], although more data

are required to be confident of its association to DD. We also identified a total of 22 (14 *Alu*,

7 L1, 1 SVA) polymorphic MEIs on the X chromosome, of which 4 (3 *Alu*, 1 L1) directly

impacted protein-coding sequence. Of these variants, none were at a low enough allele

frequency to be reasonably DD-associated, were located within a gene associated with DD,

nor fit an inheritance pattern consistent with X-linked disease.

MEI mutation rate and enrichment of deleterious RT events in DDD

Based on our findings, in the coding and peri-coding portion of the genome, one out

of every 2,434 DD cases (0.04%±0.04; 95% CI) is directly attributable to RT-derived

mutagenesis. To determine both if our observed number of *de novo* variants meets

expectation and if our patient cohort is enriched for causal *de novo* RT events, we estimated

the population mutation parameter, Θ[36], from the unaffected parents in the DDD study and

from the 1KGP[4,5] (Supplemental Table 4). The resulting calculation gives very similar

estimates of MEI mutation rate (combined across *Alu*, L1, SVA) of between 1.4×10^{-11} (1KGP)

and 1.2×10^{-11} (DDD) variants per bp per generation (μ), or ~1 new MEI genome-wide per

every 12 to 14 births – largely concordant with prior estimates from smaller WGS

datasets[3,37].

Figure 4: Estimating enrichment of deleterious MEIs

Depicted are total number of expected (black) and observed (red) *de novo* mutations observed in

exons (**a**) and enhancers (**b**) for all, high pLI (pLI > 0.9), and known monoallelic DD (MA DDG2P)

genes. Expectation is based on the Poisson distribution of 100 simulations utilizing the neutral

mutation rate (1.2×10^{-11} μ). P-values are based on the Poisson distribution, and used to determine

statistical deviation of observed to expected *de novo* counts for exons and enhancers (right).

Using this genome-wide mutation rate, we estimated the number of expected

mutations in various genomic compartments, including within genes intolerant to PTVs and

within DD-associated genes (Fig. 4; Methods). We identified a significant enrichment of *de*

novo MEIs in dominant DD-associated genes ($p = 4.82 \times 10^{-5}$), but not in the much larger set

of LoF intolerant genes ($p = 0.05$). To ensure that this finding was not due to inaccurate

estimation of the genome-wide mutation rate, we also assessed the probability that four out

of six exonic *de novo* MEIs would fall within exons of dominant DD-associated genes by

chance, based on the proportion of the exome represented by these genes (and assuming

known DD-associated genes have the same MEI mutation rate as other genes) and likewise

found a significant enrichment ($p = 4.3 \times 10^{-5}$).

Discussion

Here we have described the development, validation and exemplification at scale of

an analytical pipeline for the rapid assessment of patient genomes for RT variants.

Alongside Torene et. al.[38], we have used these approaches to present the largest study

examining the coding genome for RT-derived variation to date (Table 1; Fig. 1). With this

dataset, we first demonstrated that exonic MEIs (regardless of insertion length) are under

selective constraint on par with protein-truncating SNVs (Fig. 2). We identified four likely

pathogenic RT mutations, two *Alu* and two L1 insertions (Fig. 3), all of which arose *de novo*

in known haploinsufficient DD-associated genes (Fig. 3a-d), implying that dominant loss-of-

function is the major mode of pathogenic exonic RT variation. Finally, we estimated the

genome-wide MEI mutation rate and used it to determine that DDD probands are enriched

for damaging RT variation within exons of dominant DD-associated genes (Fig. 4a).

The total number of polymorphic, exonic RT variants identified in DDD is concordant

with previous studies characterizing MEI variation[3,5,39]. Pathogenic MEIs make up 0.04% of

diagnoses in the DDD study (4/9,738 probands), a small yet individually significant collection

of diagnostic variants. Reassuringly, our proportion of diagnostic variants in DDD is

statistically identical to the 7/11,011 (0.06%) diagnostic rate for neurodevelopmental disorder

patients as determined by Torene et. al.[38] (Fisher's exact test p = 0.56). Unlike Torene et.

al.[38], we did not identify a causative inherited MEI, although this difference is not statistically

significant (Fisher's exact test p = 0.51). We infer that despite making up a significant

proportion of reported MEI variants in the clinical literature[7], bi-allelic or X-linked MEI events

are a less frequent class of pathogenic variant in developmental disorders. This is in keeping

with recent estimates[13] that in a largely outbred clinical population, such as in the UK,

recessive disorders caused by coding variants account for a much smaller fraction of

patients than dominant disorders. We also attempted to investigate both the parental origin

of *de novo* variants (Table 2) and the mosaicism of MEIs with low variant allele frequency

(Supplemental Fig. 2). In both cases, additional work incorporating long-read technology is

needed to ascertain parental contribution and, in the case of parental mosaic MEIs, potential

recurrence risk.

Interestingly, it appears that the contribution of diagnostic RT variants may vary

among diseases. Wimmer et. al.[40] reported a total of 13 diagnostic, exonic MEI variants in

4,500 neurofibromatosis type I patients (0.3% of patients). This rate is seven times higher

than that observed in DDD or Torene et. al.[38] and was attributed to a potential RT mutation

"hotspot" associated with the canonical L1 endonuclease cleavage site of 3'-AA/TTTT-5'[41]

within the neurofibromatosis-associated gene, NF1. A recent study using a cell culture

retrotransposition assay to study MEI site preference was unable to identify a mutation

hotspot in NF1[42] and therefore suggests further work is needed to investigate the role of

sequence context in determining the overall genomic landscape of RT-mediated disease.

Analogously, inclusion of sequence context into the SNV mutation model noticeably

improved the ability to determine enrichment/depletion of deleterious SNVs within genes[22,23].

Our study is limited in that we only identified ~2% of the RT variants in each

individual human genome[4,5]. Despite a number of known disease-associated intronic MEIs in

the literature, we did not identify a pathogenic intronic MEI. As such, it remains an open

question as to what contribution RT mutations in the noncoding genome plays in the etiology

of DD. While it appears that the contribution of regulatory elements to DD is relatively small,

as defined by this (Fig. 4b) and other studies[15], previous work has identified a significant

signature of purifying selection against MEI events within 100bps of exons[25] – variants which

our study could potentially identify. As our data suggests that the majority of DD cases with

pathogenic coding MEIs are due to de novo insertions (Table 2; Fig. 3), we conjecture that

most additional DD-associated MEIs may be located in the introns of known DD-causing

genes and disrupt splicing – a known disease mechanism attributable to RT-derived

mutagenesis[7,40,43]. Simulations suggest that under a null genome-wide mutation model we

should expect to observe 12.5 (5.5-19.4, 95% CI) de novo intronic RT mutations in dominant

DD-associated genes in a population sample of 9,738 individuals. As such, a WGS study of

a clinical population of similar size to that analyzed here should be well powered to estimate

the pathogenic contribution of intronic MEIs.

De novo MEIs are typically readily interpretable with modest informatics expertise,

and represent a clinically relevant class of variation to assay in clinical bioinformatics

pipelines. While we ultimately find that the overall burden of RT-attributable disease is

relatively low in the human population, it is nonetheless an important consideration when

elucidating the genetic basis of disease in individual patients.

Online Methods

Patient recruitment and sequencing

A total of 13,462 patients were recruited from 24 clinical genetics centers from throughout the United Kingdom and the Republic of Ireland as previously described[44]. Informed consent was obtained for all families and the study was approved by the UK Research Ethics Committee (10/H0305/83, granted by the Cambridge South Research Ethics Committee and GEN/284/12, granted by the Republic of Ireland Research Ethics Committee). For the purposes of this study, individuals that were not recruited as part of a trio (e.g. individual patients or patients with just one parent), were included on the DDD sample blacklist, or failed to meet MELT QC requirements[5] were excluded from downstream analysis (leaving n = 9,738 probands; 28,132 individuals). Sequencing of families and alignment to version hg19 of the human reference genome with bwa[45] was performed as previously described[12].

Processed pseudogene pipeline development

PPGs, particularly young polymorphic events, share highly homologous sequence with the source gene from which they are derived. Consequently, the WES bait capture method will capture both DNA from the original "donor" gene and the new "daughter" copy. This allows, compared with our approach for MEI discovery, for ascertainment of PPGs genome-wide. While this approach does come with limitations, such as difficulty in identifying insertion variants, we can still determine events per individual.

Our discovery pipeline functions in two steps: first we collect read evidence on an individual level to determine which genes have been retroduplicated in that individual (Supplemental Fig. 1). Second, we determine presence/absence of each PPG in every individual in the DDD cohort based on the gene models built in the first step. In step one, we iterate over all genes in the ENSEMBL gene database which have a determined pLI score[22] and collect discordant read pairs (DRPs) which map between exons and have an insert size

>99.5% of all other reads in the sample. If more than four reads linking two exons are found,

the gene is considered to be retroduplicated elsewhere in the genome. In step two, for each

gene identified in step one, all evidence across all PPG positive individuals are pooled to

make a model of the PPG. This model is then used to check for DRP and split read pair

(SRP) evidence in all genomes. If an individual has at least 5 total read pairs of supporting

evidence with at least one SRP and one DRP, an individual is considered positive for the

given PPG. All genes and individuals were combined into a flat file listing presence or

absence of a given PPG in each individual. Source code and more information is available

online at github: https://github.com/eugenegardner/Retrogene.git.

MEI call set generation and consequence annotation

To identify MEIs in the DDD WES data we utilized the previously published Mobile

Element Locator Tool (MELT)[5]. MELT was run with default parameters (except the '-exome'

flag during IndivAnalysis) using 'Split' mode to generate a final unified VCF-format file[46] of all

28,132 unfiltered individuals independently for each MEI type (*Alu*, L1, SVA). Following initial

data set generation, we found that a subset of variants internal or adjacent to (±50bp) low

complexity repeats (defined here as a run of sequence ≥ 15bp composed of two or fewer

nucleotides) were likely false positive. As such, we added an additional filter to the final

MELT VCF, lc (low complexity), which removes such false positives from downstream

analysis. Variants that could not be genotyped in at least 25% of individuals, had ≤ 2 split

reads, had MELT ASSESS score < 3, or had any value in the VCF FILTER column other

than PASS or rSD were filtered.

To determine if variants per individual approximated a Poisson distribution, we

calculated the mean number of sites per unrelated individual (λ) independently for each RT

class and for all RT classes combined as displayed in Fig. 1a-e. Our calculated λ was then

supplied to the function *rpois* in R to generate a random deviate and plotted alongside

histograms identical to Figs 1a-e in Supplemental Fig. 3.

To generate consequences plotted in Fig. 2a, all MEIs were annotated using Variant

Effect Predictor v88 (VEP)[47] and intersected with bedtools intersect[48] to enhancers (one of

heart[49], VISTA[50], or highly evolutionarily conserved[51]) included on the DDD WES capture[15].

Only a single consequence was retained for each variant, with priority given to enhancer

annotation. Primary transcript as determined by VEP was used for all gene-based

consequences, pLI score[22] annotation, and DDG2P disease association (Table 2).

Assessment of Potentially Mosaic MEIs

To determine the mosaicmism of variants that were identified as "false-negative" *de*

novo during initial MELT analysis (i.e. genotyped as heterozygous in a single proband and

homozygous reference in both parents), we randomly selected 1,000 trios and determined

the total number of reads supporting both the insertion and reference allele (i.e. allelic

proportion) at all 917 *Alu* loci identified in this study. Alongside this data, we plotted the

parental allelic proportion at the 15 "false negative" *Alu* loci described above (Supplemental

Fig. 2a). To formalize this result and ensure that these specific loci do not have higher

genotyping or sequencing error than other loci we performed three analyses. First, we

compared the combined allelic proportion of all heterozygous variants to the combined allelic

proportion of both false negative parents and the associated heterozygous probands

(Supplemental Fig. 2b). Second, we examined the allelic proportion for all homozygous

reference genotypes at each false negative locus for evidence of higher error rates based on

higher average allelic proportion at those loci using a z-test (Supplemental Fig. 2c). Finally,

to assess if these loci had similar read start distributions in both parent and proband, we

retrieved all read start sites within 250 bps of each insert site and used a Kolmogorov-

Smirnov test to compare the two distributions. Only one locus,

chr11:g.95523918_95523919ins*Alu*(1_280), showed a significant difference at a threshold of

p = 0.05 (p = 2.70×10^{-4}).

Since these sites are not likely to be genotyping or sequencing artifact, we next

attempted to clarify the nature of a subset of variants with a custom long-range PCR and

Pacific Biosciences sequencing approach. The goal of this methodology was to attempt to

both validate and phase candidate mosaic MEIs to nearby (±5Kbps) SNVs. As above, we

limited our analysis to *Alu* MEIs as the large insert length of L1 (n = 1) and repeat heavy

SVA (n = 2) elements would likely prove refractory to long PCR with enough flanking

sequence to phase to nearby SNVs. To design locus specific primers, the genomic region

surrounding each target MEI was entered into Primer-BLAST

(www.ncbi.nlm.nih.gov/tools/primer-blast/) using default parameters. Each locus was

targeted with 3 primer pairs: 1 pair amplifying a 10kb region with the MEI lying centrally, and

two pairs of primers to amplify 5kb fragments extending up and downstream of the MEI.

10kb PCRs were performed using AccuTaq LA DNA Polymerase (D8045, Sigma-Aldrich)

with the following parameters: 98 degree denaturation for 1 minute, followed by 35 cycles of

94 degrees (30 seconds)/54 degrees (30 seconds)/68 degrees (11 minutes), with a final

extension of 18 minutes at 68 degrees. 5kb PCRs were performed as above, but with a

reduced 6 minute extension time at 68 degrees, and a final extension time of 10 minutes.

After quality checking the PCRs on control DNA, 100-200ng of parental DNA was amplified

in a 50 microlitre PCR reaction using the same parameters.

After confirming amplification by agarose gel, PCR products were isolated using

SPRI bead purification (Ampure-XP, Agilent), quantified, and mixed in equimolar amounts.

We then performed a standard PacBio library prep using SMRTbell Template Prep Kit 1.0

(part no. 100-991-900) without size selection. Sequencing was performed on the Pacific

Biosciences Sequel platform using Sequel Sequencing Kit 3.0 (101-597-800) and Sequel

SMRT cell 1M v3 (101-531-000). For error correction, raw sequencing reads were then

provided as input to pbccs v3.4.1 (https://github.com/PacificBiosciences/unanimity) with

default parameters other than minLength=100. Corrected reads were then aligned to the

human genome using blasr v5.3.3[52] and downsampled to between 2,000-3,000x coverage

per locus assessed.

MEIs were then phased to the nearest heterozygous SNV on a per-read basis using

a custom java pipeline. After dropping sites due to PCR errors or phasing ambiguity, we

were able to examine a total of 4 *Alu* MEIs (Supplemental Table 6). Total number of reads

supporting each of four possible phases (MEI+SNV+, MEI+SNV-, MEI-SNV+, MEI-SNV-)

were then quantified. As PCR can induce template switching errors resulting in variants

being sequenced on the same read despite in truth not being on the same haplotype (e.g.

MEI+SNV-)[53], we then determined if total number of reads supporting both possible true

phases (MEI+SNV+ and MEI-SNV+) were significantly different from error (MEI+SNV-) using

a Fisher's exact test. MEIs where both true phases were significantly different from error

were considered true mosaic. Primers and quantification of results are available in

Supplemental Table 6.

Quality control of RT data using WGS and 1KGP

To determine if our MEI WES call set was biased compared to WGS data, we

performed two independent comparisons: 1) to high coverage (>30x) WGS data generated

for a subset of DDD trios and 2) to a published collection of MEIs from 1KGP phase III[5].

For WGS quality-control, we used a subset of 30 DDD trios (n = 90 individuals) which

were previously whole genome sequenced. MEI discovery using MELT[5] on all 90 individuals

was performed and filtered identically to WES data. Genotypes identified in the WGS data

but not in WES were then separated based on coverage in the corresponding WES.

Genotypes in low coverage areas (<10x) were considered not possible (n.p), while variants

where coverage was greater than 10x are considered not detected (n.d). All remaining

genotypes were than compared for identity between WGS and WES results (Supplemental

Table 1).

To compare the DDD MEI call set to the 1KGP, we first filtered 1KGP calls to

variants with mean >10x coverage in 1,000 randomly sampled DDD WES individuals

(leaving 318 *Alu*, 81 L1, 26 SVA). We then randomly selected 2,453 DDD parents 1,000

times, retaining only loci present in downsampled individuals[4,5]. The resulting distribution

was then compared to the observed number of variants in the 1KGP-masked data to

generate z-scores independently for all three MEI types (Supplemental Fig. 4).

To compare our PPG dataset to Zhang et. al.[6], we downloaded provided

supplemental tables. We then summed the total number of unique events per person and

determined "allele counts" for each gene reported. Genes were then matched between our

call set and Zhang et. al.[6] using ENSEMBL gene identifiers and allele counts between each

data set were plotted to create Supplemental Fig. 5.

GTEx annotation of processed pseudogenes

To determine RNA expression levels of donor genes which gave rise to PPGs

identified in this study, we queried transcript per kilobase per megabase of sequencing

(TPM) scores for all genes in 30 tissues assessed by the current GTEx v7 release (available

at https://gtexportal.org/home/datasets). Only the 18,225 protein-coding genes which were

assessed for gene PPGs by our project were retained for subsequent analysis. TPM values

were then averaged across all GTEx individuals for a given tissue to generate a mean TPM

value as plotted in Supplemental Fig. 6. Nonparametric Wilcoxon rank-sum tests were

performed using the wilcox.test function in R with default parameters to generate p values

for both within tissue and between tissue comparisons.

SNV calling and quality control

To call SNVs from all DDD individuals we utilized GATK 3.5-0-ge91472d[54] in three

steps using default settings. First, we called variants in individual samples using

HaplotypeCaller. Next, individual VCF files were processed in 200 individual batches using

CombineGVCFs. Finally, all batched VCFs were passed to GenotypeGVCF to generate a

final joint-called VCF file. This file was then annotated using VEP v88[47] based on canonical

transcript. Unaffected parents (n = 17,032 individuals) were then extracted from this VCF

and only variants with an allele count greater than 1 in these individuals were retained.

For initial filtering, we removed SNVs with a VQSLOD < -2.0, read depth < 8, and

genotype quality < 20. We next performed more extensive QC using a 'missingness' score

identical to the method described in Martin et. al.[13]. In short, each genotype at a given

variant was assessed for genotype quality (GQ), depth (DP), and a binomial test for allelic

depth (i.e. number of alternate versus reference supporting reads; AD). If a given genotype

had GQ < 20, DP < 8, or AD p-value < 0.001 it was considered 'missing'. If more than 50%

of genotypes for a given variant were missing, the variant was subsequently filtered from

final analysis. Allele frequencies were recalculated based on included individuals while

accounting for missing genotypes.

Selective constraint of SNVs and MEIs

As sensitivity of variant discovery can bias our results, we generated an "accessibility

mask" of the DDD WES data where we expect our variant ascertainment sensitivity to be

>95% (Supplemental Fig. 11)[5]. Our mask thus includes only regions of the genome that

contain at least 10X average coverage in a mean cohort of 1,000 randomly selected

individuals for a total of 74.2Mbp, or ~2.3% of the genome (Supplemental Table 4). Using

this mask, we filtered our original 1,129 variants down to 828 (660 *Alu*, 109 L1, 31 SVA)

variants in unaffected parents (n = 17,032 individuals). Parents were determined to be

affected either by the referring clinician or, where ambiguous, through manual curation of

HPO terms for a matching parent-offspring phenotype.

Using this mask subset of variants, we determined genomic constraint as shown in

Figure 2b. Allele frequency values were recalculated for all variants, and a pLI score[22] for

each MEI was added as described above. MEIs which did not insert into a gene or inserted

into a gene without a calculated pLI score[22] were excluded from subsequent analysis. We

then calculated proportion of singleton variants and proportion of variants in high pLI genes

independently for *Alu* and, due to low overall numbers of the other MEI subtypes, for a

combined set of *Alu*, L1, and SVA. SNVs annotated as nonsense, missense, synonymous,

or splice acceptor/donor (splice in Fig. 2b) as determined by VEP v88[47] were extracted from

the SNV VCF files described above and used to calculate singleton and pLI proportion

identically to MEIs.

Fetal Brain Expression Data

To assess if genes disrupted by RT variants were expressed in the fetal brain, we

downloaded per-gene expression data generated by the BrainSpan consortium (file:

genes_matrix_csv.zip; http://www.brainspan.org/static/download.html)[33]. From this file, we

extracted columns representing data generated from fetuses 15 to 37 weeks post-

conception. This data was then provided as input to a custom python script which calculated

the median RPKM of all genes provided for the file. Genes with a median RPKM ≥ 1 were

considered expressed in the fetal brain (n = 15,474). Likely causal RT variants (n = 4) were

then intersected with this list to determine gene expression status. As RPKM ≥ 1 is a rather

liberal cutoff, we also ranked these genes compared to other genes in our expression list;

these values are reported in Supplemental Table 5.

MEI validation by PCR

To validate all 9 *de novo* MEI variants (Table 2) and the homozygous insertion in

PAN2 we used the following PCR protocol: primers were designed using Primer3 to make

products spanning the predicted insertion site (Supplemental Table 2). PCR was carried out

using Platinum™ Taq DNA Polymerase High Fidelity (Invitrogen); 20ng of genomic DNA

extracted from blood or saliva was amplified in the presence of 0.2 μM of each primer and 1

unit of Platinum™ Taq. Amplification was carried out using the following cycling conditions;

for Alu insertions: 2 min at 94°C, followed by 36 cycles of (30 sec at 94°C, 30 sec at 60°C

and 1 min at 68°C); for LINE1 insertions: 2 min at 94°C, followed by 36 cycles of (30 sec at

94°C, 30 sec at 60°C and 7 min at 68°C). PCR products were visualized using a 2% agarose

E-Gel® (Invitrogen).

Processed pseudogene validation by PCR and capillary sequencing

To validate the 2 *de novo* PPG variants (Table 2) we used the following PCR

protocol: primers were designed using Primer3 to make products within the exons of each

gene. Forward and reverse primers were then paired between exons to amplify across the

excised intronic regions (Supplemental Table 2). PCR was carried out using either

Platinum™ Taq DNA Polymerase High Fidelity (Invitrogen) or Thermo-Start Taq DNA

Polymerase (Thermo Scientific). Platinum™ Taq assay: 20ng of genomic DNA extracted

from blood or saliva was amplified in the presence of 0.2 μM of each primer and 1 unit of

Platinum™ Taq. Amplification was carried out using the following cycling conditions; 2 min at

94°C, followed by 36 cycles of (30 sec at 94°C, 30 sec at 60°C and 1 min at 68°C). Thermo-

Start Taq DNA Polymerase assay: 40 ng genomic DNA was amplified in the presence of 0.2

μM of each primer and 0.42 units of Thermo-Start Taq. Cycling conditions were as follows: 5

min at 95°C, 6 cycles of (30 sec at 95°C, 30 sec at 64°C and 1 min at 72°C), 6 cycles of (30

sec at 95°C, 30 sec at 62°C and 1 min at 72°C), 6 cycles of (30 sec at 95°C, 30 sec at 60°C

and 1 min at 72°C) followed by 36 cycles of (30 sec at 95°C, 30 sec at 58°C and 1 min at

72°C) with a final elongation of 10 min at 72°C. PCR products were visualized using a 2%

agarose E-Gel® (Invitrogen). PCR products were sequenced using either the forward or

reverse primer used in the amplification protocol by Eurofins GATC Biotech GmbH.

Sequence traces were aligned using SeqMan Pro 15 (Lasergene 15) and reads were

aligned to the human genome (hg19) using BLAT (UCSC)[55].

WGS of probands with *de novo* processed pseudogenes

To validate and determine the insertion site of the two identified *de novo* PPGs

(Table 2), we performed Illumina WGS on all individuals of each trio in which the *de novo*

event was identified (n = 6 individuals). Samples were first quantified with Biotium Accuclear

Ultra high sensitivity dsDNA Quantitative kit using Mosquito LV liquid platform, Bravo WS

and BMG FLUOstar Omega plate reader and cherrypicked to 500ng / 120ul using Tecan

liquid handling platform. Cherrypicked plates are then sheared to 450bp using a Covaris

LE220 instrument and subsequently purified using SPRI Select beads on Agilent Bravo WS.

Library construction (ER, A-tailing and ligation) was performed using 'NEB Ultra II custom kit'

on an Agilent Bravo WS automation system. Samples were then tagged using NextFLEX

Unique Dual Indexed adapter 1-96 barcodes at the ligation stage. Libraries were then

quantified by qPCR using Kapa Illumina ABI Sanger custom qPCR kits using a Mosquito LV

liquid handling platform, Bravo WS, and Roche Lightcycler. Libraries are then pooled in

equimolar amounts on a Beckman BioMek NX-8 liquid handling platform and normalised to

2.4nM for cluster generation on a c-BOT and then sequenced on the Illumina TenX

sequencing platform. Following sequencing, reads were aligned with BWA mem[45] (with

settings -t 16 -p -Y -K 100000000) to version hg19 of the human reference genome. Reads

were then manually inspected using the Integrative Genomics Viewer (IGV)[56] to confirm

presence, *de novo* status, and parent of origin of each PPG.

Methylation analysis

Methylation analyses were performed in accordance with the methods detailed in

Choufani, et al. [35]. In short, we downloaded publicly available methylation data from GEO

(accession GSE74432) generated by Choufani, et al. [35] and combined it with Illumina EPIC

methylation array data generated in-house for our *Alu NSD1* loss of function patient (Fig. 3b;

Table 2). Following data curation, we randomly sampled a subset of 15 cases (defined by

Choufani, et al. [35] as a carrier of an *NSD1* loss of function variant) and 15 controls to identify

a set of differentially methylated "signature" CpGs based on a 20% effect size cut-off and

0.01 FDR-corrected p-value which separated case and control individuals. Once a signature

was established, we used Pearson correlation to assess each sample's similarity to a mean

case and mean control based on differential methylation at signature CpG sites. Confirmed

pathogenic *NSD1* variants and control samples used to build the signature were not included

in the correlation analysis (Supplemental Fig. 10).

MEI mutation rate and burden

To determine the mutation rate independently for each MEI type (*Alu*, L1, SVA), we

utilized data generated by both DDD and the 1KGP[5]. For DDD data we filtered sites as

above based on our >10X coverage accessibility mask. For the 1KGP data[5], we created a

combined mask from three different data sources: 1) the pilot accessibility mask generated

by the 1KGP project phase III[57], which removes regions of the genome inaccessible to

variant calling, 2) reference ME sequences as identified by repeatmasker[58], as MELT is

unable to accurately ascertain MEIs in these regions, and 3) All sequence ±10Kbp from the

5' and 3' terminus of all protein-coding genes from RefSeq[59]. This mask was generated

separately for *Alu* and L1 and did not filter 1,113.0Mbp or 959.9Mbp of the genome,

respectively. The *Alu* mask was used for filtering SVA and both masks excluded both

allosomes. On masking the 1KGP data, we were left with a total of 10,930 autosomal MEIs

(8,554 *Alu*, 2,047 L1, 329 SVA). Following filtering of the DDD and 1KGP sets with their

corresponding masks, we used the Watterson estimator with an effective population size of

10,000 for all calculations to estimate the population mutation parameter, Θ[36], and mutation

rate, μ (Supplemental Table 4).

We next used our estimate of μ to determine the expected number of *de novo* MEI

mutations in exons, enhancers, and introns genome-wide. To calculate this value, we

extrapolated μ for 9,738 births (i.e. number of probands in our study), resulting in an

estimate of 686 *de novo* MEIs genome-wide. We then simulated this number of variants 100

times genome-wide using bedtools random[48] and annotated the resulting "variants"

identically to real variants reported in this study (e.g. as in Fig. 2a). The total number of

expected variants in the three categories depicted in Fig. 4 were then averaged across all

simulations to determine Poisson λ of *de novo* variants under neutral mutation rate. These

distributions were then statistically compared to the number of observed variants in DDD

using the *ppois* function in R (Fig. 4).

Data Availability

Sequencing, phenotype data, and variant calls for all data in this paper are accessible via

the European Genome-phenome Archive (EGA) under study number EGAS00001000775

(https://www.ebi.ac.uk/ega/studies/EGAS00001000775). Within this study, WES files of all

DDD families are provided as part of dataset EGAD00001004390 and RT variant calls and

WGS data of the two individuals with *de novo* PPG variants are provided as part of dataset

EGAD00001004586.

Author Contributions

E.J.G performed RT variant calling and annotation, PPG algorithm design, constraint and

burden testing, computational analysis of mosaic variants, and initial clinical annotation and

together with M.E.H. designed experiments, oversaw the study, and wrote the manuscript.

E. P. designed and performed PCR experiments. G.G. curated and prepared DDD

sequencing data. S.S.G and H.I designed and conducted molecular assays to assess

mosaic MEIs. P.J.S. assisted in estimating genetic burden of deleterious MEIs in the human

population. A.S. assisted with the design of the PPG discovery algorithm. P.D., K.E.S.,

E.J.G., and T.S. performed variant calling and quality control of SNVs. J.H. performed

methylation analysis of the *NSD1* variant patient. K.E.C, E.C., K.L.L., K.P., E.R., D.R.F, and

H.V.F prepared clinical assessments of patients and confirmation of molecular diagnoses as

they relate to patient phenotype.

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Competing Interests

M.E.H. is a co-founder of, consultant to, and holds shares in, Congenica Ltd, a genetics diagnostic company.

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