

1 Reproduction has different costs for 2 immunity and parasitism in a wild 3 mammal

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22 tradeoff, wild mammal

23

24 Abstract

- 25 1. Life history theory predicts that reproductive investment draws resources away from
26 immunity, resulting in increased parasitism. However, studies of reproductive tradeoffs
27 rarely examine multiple measures of reproduction, immunity, and parasitism. It is
28 therefore unclear whether the immune costs of reproductive traits correlate with their
29 resource costs, and whether increased parasitism emerges from weaker immunity.
- 30 2. We examined these relationships in wild female red deer (*Cervus elaphus*) with
31 variable reproductive investment and longitudinal data on mucosal antibody levels and
32 helminth parasitism. We noninvasively collected faecal samples, counting propagules
33 of strongyle nematodes (order: Strongylida), the common liver fluke *Fasciola hepatica*
34 and the red deer tissue nematode *Elaphostrongylus cervi*. We also quantified both total
35 and anti-strongyle mucosal IgA to measure general and specific immune investment.
- 36 3. Contrary to our predictions, we found that gestation was associated with decreased
37 total IgA but with no increase in parasitism. Meanwhile, the considerable resource
38 demand of lactation had no further immune cost but was associated with higher counts
39 of strongyle nematodes and *Elaphostrongylus cervi*. These contrasting costs arose
40 despite a negative correlation between antibodies and strongyle count, which implied
41 that IgA was indicative of protective immunity.
- 42 4. Our findings suggest that processes other than classical resource allocation tradeoffs
43 are involved in mediating observed relationships between reproduction, immunity, and
44 parasitism in wild mammals. In particular, reproduction-immunity tradeoffs may result
45 from hormonal regulation or maternal antibody transfer, with parasitism increasing as
46 a result of increased exposure arising from resource acquisition constraints. We
47 advocate careful consideration of resource-independent mechanistic links and
48 measurement of both immunity and parasitism when investigating reproductive costs.

49 Introduction

50 Costly traits are central to the fields of life history theory and ecoimmunology. Tradeoffs arising
51 between reproductive investment and other aspects of life history are a fundamental prediction
52 of the former (Harshman & Zera, 2007; Stearns, 1989; Williams, 1966), while the latter
53 examines the ecology of costly immune responses (Graham et al., 2011; Sheldon & Verhulst,
54 1996). Because reproduction and immunity compete for host resources, in resource-limited
55 environments, animals that invest in reproduction should have fewer resources to allocate to
56 immune defences (Deerenberg, Arpanius, Daan, & Bos, 1997; French, Denardo, & Moore,
57 2007; Sheldon & Verhulst, 1996). If immunity is protective, these individuals will experience
58 higher parasitism as a result. Intuitively, traits with higher resource demands should result in
59 the diversion of more resources away from immunity, leading to higher parasite burdens.
60 However, recent advances have demonstrated that the interrelationships between host
61 resources, immunity, and parasitism can be unexpectedly complex (Cressler, Nelson, Day, &
62 Mccauley, 2014). Few studies in wild mammals have examined tradeoffs with multiple
63 reproductive traits, so it is unclear whether different components of reproduction have different
64 costs for immune defence, and whether their costs are proportional to their resource demand.
65 Furthermore, studies of reproductive tradeoffs rarely quantify both immunity and parasitism to
66 examine the importance of susceptibility versus exposure in driving higher parasite intensities
67 in reproductive females (Bradley & Jackson, 2008; Knowles, Nakagawa, & Sheldon, 2009).
68 Here, we examine the partitioning of reproductive costs for multiple measures of immunity and
69 parasitism to investigate the possible mechanisms governing a reproduction-immunity-
70 parasitism tradeoff in a wild mammal.

71 Mammalian reproduction is a complex, multi-stage process, featuring extensive maternal
72 investment which varies in intensity through the reproductive period (Langer, 2008;
73 Maestriperi & Mateo, 2009). As such, different components of reproduction vary substantially
74 in their resource and fitness costs. In particular, lactation is a highly energetically demanding
75 process which carries costs for immunity, parasitism or fitness in a range of mammals

76 (Beasley, Kahn, & Windon, 2010; Christe, Arlettaz, & Vogel, 2000; Clutton-Brock, Albon, &
77 Guinness, 1989; Froy, Walling, Pemberton, Clutton-brock, & Kruuk, 2016; Jones, Sakkas,
78 Houdijk, Knox, & Kyriazakis, 2012; Rödel, Zapka, Stefanski, & von Holst, 2016; Woodroffe &
79 Macdonald, 1995). Meanwhile, only one of these studies uncovered an immunological cost of
80 gestation (Christe et al., 2000), which generally requires fewer resources than does lactation
81 (Clutton-Brock et al., 1989). However, although experimentally modifying resource availability
82 can affect the severity of reproduction-immunity tradeoffs (French et al., 2007; Jones et al.,
83 2012), this is not always the case (Stahlschmidt, Rollinson, Acker, & Adamo, 2013). Similarly,
84 studies in birds have questioned whether the energetic costs of immunity are sufficient to drive
85 tradeoffs (Eraud, Duriez, Chastel, & Faivre, 2005; Svensson, Råberg, Koch, & Hasselquist,
86 1998). Such findings imply that reproduction-immunity tradeoffs can be linked mechanistically
87 as well as through resource reallocation. Potential such links include production of reactive
88 oxygen species, reduction in immunologically active fat stores, or resource-independent
89 hormonal regulation (Speakman, 2008; Svensson et al., 1998).

90 Different components of mammalian reproduction can have qualitatively different effects on
91 host immunity as well as varying quantitatively in terms of their resource demand. For
92 example, pregnancy necessitates modulation of the immune system to avoid mounting an
93 immune response to the developing foetus, which will directly affect anti-parasite defence
94 (Weinberg, 1984, 1987). Similarly, lactation draws immune molecules away from the mother
95 for transfer to offspring, reducing their availability for the mother's own defence (Grindstaff,
96 Brodie, & Ketterson, 2003; Hasselquist & Nilsson, 2009). Reproduction also induces a suite
97 of physiological and behavioural changes which will affect susceptibility and exposure to
98 parasites indirectly: for example, it has been suggested that bats compensate for the energetic
99 demand of lactation by reducing costly grooming behaviour, with ectoparasite burden
100 increasing as a result (Speakman, 2008). It is unclear how such mechanistic links between
101 components of reproduction and immunity interact with resource allocation to influence
102 immune defence and parasite intensity in wild mammals.

103 The wild red deer (*Cervus elaphus*) in the North block of the Isle of Rum exhibit a well-studied
104 life history tradeoff, in which reproduction substantially decreases the mother's probability of
105 overwinter survival and reproduction the following year (Clutton-Brock et al., 1989; Froy et al.,
106 2016). However, not all components of reproduction are equally costly: gestation has a
107 negligible detectable fitness cost compared to that of lactation (Clutton-Brock et al., 1989).
108 Moreover, while giving birth late and caring for a male calf compared to a female calf are
109 associated with decreased maternal fitness, their effects are small compared to the cost of
110 lactation itself (Froy et al., 2016). The study population has a high prevalence of
111 gastrointestinal helminths, and parasite burdens can be quantified noninvasively through
112 faecal propagule counts (Albery et al., 2018). Mucosal antibodies, and especially the IgA
113 isotype, are important effectors of ruminant adaptive immunity to gut helminths (Butler, 1969;
114 McRae, Stear, Good, & Keane, 2015). Mucosal IgA can be quantified in wild ruminant faeces,
115 correlating positively with the same isotype measured in plasma or serum and negatively with
116 helminth faecal egg counts (Watt, Nussey, Maclellan, Pilkington, & McNeilly, 2016).

117 In this study, we measured both total and helminth-specific mucosal IgA and propagule counts
118 of multiple helminth species, using faecal samples collected from the Isle of Rum study
119 population. We quantified the associations between several reproductive traits of known
120 fitness cost and subsequent measures of immunity and parasitism. We also examined
121 covariance between IgA and parasites to discern whether increased IgA was associated with
122 decreased parasite intensity independently of shared reproductive and seasonal effects,
123 implicating IgA as an indicator of protective immunity. We predicted that reproductive
124 investment would be associated with reduced antibody levels and increased parasite counts,
125 and that aspects of reproduction previously found to be more costly for fitness, especially
126 lactation, should likewise be more costly in terms of both immunity and parasitism.
127 Furthermore, providing parasitism is mediated by immune susceptibility, aspects of
128 reproduction that are costly for immunity should have similar costs in terms of parasitism.

129 Methods

130 Study system, sampling and parasitology

131 The study population is located in the North block of the Isle of Rum National Nature Reserve
132 in the Inner Hebrides, Scotland (57°N 6°20'W). The resident population comprises ~350
133 animals at any one time, and is regularly censused to keep track of each individual and its life
134 history. See Clutton-Brock *et al.* (1982) for a full summary of the project and the deer
135 reproductive cycle. Briefly, the deer mate in September and October and give birth in May-
136 June, after an approximately 235 day gestation. Females do not reproduce every year, and
137 produce a maximum of one calf per year. During the calving season, daily monitoring of
138 pregnant females enables the recording of precise birth dates. Neonates are caught, sexed,
139 weighed and individually marked, enabling life-long individual identification. Calves are
140 dependent on their mothers for much of their first year. Regular population censusing
141 throughout the year and winter mortality searches allow dates of death to be reliably assigned
142 to the nearest month for the vast majority of individuals. Most calf deaths occur either within
143 the first few weeks of life, or in the following winter ~6-9 months later. Females that
144 successfully raise a calf to the age of one, or that lose the calf in its first winter, have lower
145 rates of overwinter survival and reproduction the following year compared to those that do not
146 reproduce that year or that lose their calf in the summer (Clutton-Brock *et al.*, 1989; Froy *et*
147 *al.*, 2016). Many calves die over the winter, but the mothers of these calves have paid the cost
148 of lactation associated with feeding them until the winter, whether or not the calf survive.
149 Therefore these females are treated as a single category here (Clutton-Brock *et al.*, 1989;
150 Froy *et al.*, 2016).

151 We collected faecal samples from female deer across the annual reproductive cycle. As a
152 “deer year” runs from May to April, this study examines the effects of reproduction over a year,
153 beginning in May, on egg counts and antibody levels until the following April. A description of
154 the sample collection procedure can be found in Albery *et al.* (2018). Sampling occurred over
155 seven two-week sampling trips spanning April 2016-April 2018, in August (“summer”),

156 November (“autumn”) and April (“spring”). Note that our dataset included an April sampling
157 trip from the deer reproductive cycle starting May 2015, without an accompanying August and
158 November trip from this reproductive cycle. Figure 1 illustrates how sampling relates to
159 different aspects of reproductive investment by female deer across the annual cycle. We
160 classify a female’s reproductive status for a given year as “No Calf”, “Calf Died” and “Calf
161 Survived” (see Figure 1). “No Calf” samples were collected from females that did not
162 reproduce in the calving season preceding the sampling trip; “Calf Died” samples were
163 collected from females that gave birth to a calf in the preceding calving season which died
164 before October 1st of that year; and “Calf Survived” samples were collected from females that
165 gave birth to a calf in the preceding calving season which survived past October 1st of that
166 year. We excluded females that were reproducing for the first time from our analyses, as their
167 reproductive success is heavily confounded with their young age (mean age 4.21 years). In
168 addition, females may or may not become pregnant during the autumn rut. Samples were
169 therefore assigned a pregnancy status, beginning in November, based on whether or not the
170 female gave birth to a calf in the following spring (Figure 1).

171 In total 837 faecal samples were collected noninvasively from 140 mature females. In the
172 evening after collection, samples were homogenised by hand and subsampled, with 1-15g
173 frozen at -20°C for antibody quantification and the remainder refrigerated at 4°C for
174 parasitological analysis. Subsamples were transferred to Edinburgh at these temperatures.
175 Parasite propagule counts were carried out as previously described, without correcting for dry
176 weight, and included counts of strongyle nematodes, the common liver fluke *Fasciola hepatica*
177 and the tissue nematode *Elaphostrongylus cervi* (Albery et al., 2018). Final sample sizes for
178 each variable are displayed in Table S11.

179 Antibody extraction and quantification

180 Faecal antibodies were quantified using a protocol modified from Watt *et al.* (2016). Faecal
181 matter was stored at -20°C until extraction. Extractions occurred either in January-March 2017
182 (session “A”, samples collected April-November 2016; N=132), January 2018 (“B”, samples

183 collected April-November 2016; N=212) or within the sampling trip (“C”, samples collected
184 April 2017-April 2018, N=460). 0.6g (+/- 0.005g) of the homogenate was weighed out into an
185 Eppendorf tube and mixed thoroughly with 0.9ml of protease inhibitor solution (cOmplete™
186 Mini Protease Inhibitor Cocktail tablets, Roche, Basel, Switzerland; 1 tablet mixed with 7ml
187 Phosphate Buffered Saline). The mixture was left to stand for a minimum of 5 minutes to allow
188 the protease to act and then centrifuged at 10,000g for 5 minutes. The supernatant was
189 removed using a micropipette and stored in a separate Eppendorf tube at -20°C until ELISA.
190 We measured two antibodies by ELISA: total IgA and anti-*Teladorsagia circumcincta* third
191 larval stage IgA (anti-Tc IgA). *T. circumcincta* is an abundant and important sheep strongyle,
192 and anti-Tc IgA correlates negatively with strongyle (order: Strongylida) faecal egg count in
193 wild Soay sheep (Watt et al., 2016). *T. circumcincta* is also present in the Rum deer
194 (unpublished data). ELISA plates were coated the night before using sheep-derived capture
195 antibodies (Bethyl Laboratories, Montgomery, TX) for total IgA and with third larval stage
196 antigen for anti-Tc IgA (Moredun Research Institute, Penicuik, Scotland). For total IgA the
197 samples were diluted in the ratio 1:64; due to lower antibody concentrations undiluted
198 supernatant was used for the anti-Tc IgA assay. After this stage, the ELISA protocol was
199 carried out as described in Watt *et al.* (2016). The total IgA dilution was selected by carrying
200 out serial dilutions on a set of samples and selecting the dilution at which different
201 concentrations of antibodies were deemed to have the widest spread of optical densities.
202 ELISA readings diluted linearly as expected. Samples were corrected using controls according
203 to the calculation: Final OD=(sample OD-mean plate negative OD)/(mean plate positive OD-
204 mean plate negative OD). All samples were run on duplicate plates, which were highly
205 correlated (R=0.98 across all duplicates). The mean value for the two duplicates was taken
206 for each sample and used for analysis.

207 **Statistical analysis**

208 We used four sets of Generalised Linear Mixed Models (GLMMs) to test how reproductive
209 traits were associated with antibody levels and parasite intensity. Analyses were carried out

210 in R version 3.5.0 (R Core Team 2018) with the package MCMCglmm (Hadfield, 2010). All
211 models were run for 2.6×10^6 iterations with a 2000 iteration thinning interval and a 6×10^5
212 iteration burn in period. P_{MCMC} values for differences between factor categories were
213 calculated using the proportional overlap of estimates' posterior distributions, divided by half
214 the number of iterations.

215 Full models

216 We first constructed five univariate GLMMs using the full dataset (837 samples from 140
217 individuals). Three models used an overdispersed Poisson distribution, with strongyle, *F.*
218 *hepatica* and *E. cervi* intensity as response variables. Models initially included the following
219 fixed effects, without interactions: Year (factor with three levels representative of the deer
220 reproductive cycle beginning in 2015, 2016 and 2017); Season (factor with three levels:
221 Summer, Autumn and Spring); Age in years (continuous); and Reproductive Status (factor
222 with three levels: No Calf, Calf Died and Calf Survived). Individual identity was fitted as a
223 random effect. All continuous variables except parasite counts were scaled to have a mean of
224 0 and a standard deviation of 1 before analysis.

225 The two remaining models examined antibodies as response variables. As mucosal antibodies
226 are vulnerable to degradation by temperature-dependent faecal proteases, we had to account
227 for the extraction session and time to freezing and extraction (Figure SI5-6). There was an
228 uneven distribution of year, season, and status categories across different extraction sessions,
229 so that these variables could not all be fitted in the same model. Therefore, to control for
230 collection factors and quantify reproductive status effects conservatively we first transformed
231 antibody levels to approximate normality (square-root transform for total IgA and cube-root
232 transform for anti-Tc IgA), and fitted a linear model with fixed effects including extraction
233 sessions performed at different times (factor with three levels); day of collection within a
234 sampling trip (continuous integers, range 0-11); time elapsed from sample collection until
235 freezing (continuous, in hours). The scaled residual values from these models (mean=0,

236 SD=1) were used as the response variables in two Gaussian GLMMs with the same fixed and
237 random effects as the parasite GLMMs.

238 Previous work on the Rum deer revealed extensive seasonal fluctuations in parasite count
239 (Albery et al., 2018). We therefore followed up the above five models by fitting a season by
240 reproductive status interaction in order to investigate whether the effects of reproductive status
241 varied by season. Each model was compared with and without this interaction to investigate
242 whether it affected Deviation Information Criterion (DIC) values as a measure of model fit
243 (threshold values for distinguishing between models $\Delta\text{DIC}=2$) or changed model estimates.

244 **Pregnancy models**

245 Pregnancy may directly affect immunity, and effects attributed to reproductive status could be
246 due to correlated variation in pregnancy status over the sampling year. To check this we ran
247 a second set of models investigating the role of pregnancy status. This used a subset of
248 samples collected in November and April (518 samples from 122 individuals), as mating
249 occurs in the early autumn and females could not be pregnant in August. These five models
250 featured the same explanatory variables as the full status models, with only two levels in the
251 season category (Autumn and Spring), and with Pregnancy included as a binary variable. We
252 compared these models with and without the pregnancy term as a fixed effect to investigate
253 whether its inclusion changed reproductive status effect sizes or affected model fit.

254 **Calving trait models**

255 We next used a restricted dataset consisting of individuals that had given birth in the year of
256 sampling (571 samples from 116 individuals) to investigate whether specific traits associated
257 with a calving event influenced antibody levels and parasitism. We fitted the same fixed and
258 random effects as the full model set, but with only two factor levels in the reproductive status
259 category (Calf Died and Calf Survived), and including several variables relating to each birth:
260 Parturition Date (continuous, centred around median birth date that year); Birth Weight

261 (continuous, in kilograms, calculated from a projection using capture weight and age in hours,
262 slope 0.01696 kg/h); Calf Sex (Female or Male).

263 Multivariate model

264 Multivariate mixed-effects models can be used to investigate covariance between measures
265 of immunity and parasitism, while accounting for fixed effects. To test whether antibodies and
266 parasites were correlated we fitted a model with strongyles, *E. cervi*, total IgA and anti-Tc IgA
267 as response variables, with the same fixed effects as the full univariate models. Unstructured
268 variance/covariance matrices were fitted for random and error terms, allowing estimation of
269 the phenotypic correlations between the response variables. Phenotypic covariance between
270 two response variables A and B ($\text{Cov}_{\text{phenotypicA,B}}$) is calculated using the random (G) and
271 residual (R) variance structure of the model, with the formula
272 $\text{Cov}_{\text{phenotypicA,B}} = \text{Cov}_{\text{IndividualA,B}} + \text{Cov}_{\text{residualA,B}}$. Phenotypic correlation between two response
273 variables ($r_{\text{phenotypicA,B}}$) was calculated by dividing the phenotypic covariance by the square root
274 of the sum of the variance in both response variables:
275 $r_{\text{phenotypicA,B}} = \text{Cov}_{\text{phenotypicA,B}} / (\text{V}_{\text{phenotypeA}} + \text{V}_{\text{phenotypeB}})^{0.5}$. P_{MCMC} values for correlations were
276 calculated using the posterior distributions, dividing the number of iterations overlapping with
277 zero by half the total number of iterations.

278 Results

279 Reproductive investment was strongly associated with both lower antibody levels and
280 increased parasite counts, but patterns differed considerably between different response
281 variables (Figure 2, S11). Compared to “No Calf” individuals, “Calf Survived” status was
282 associated with higher intensity strongyle ($P_{\text{MCMC}} < 0.001$) and *E. cervi* infection ($P_{\text{MCMC}} = 0.01$),
283 and with lower total IgA ($P_{\text{MCMC}} = 0.016$) and anti-Tc IgA levels ($P_{\text{MCMC}} < 0.001$). “Calf Survived”
284 females also had higher parasite counts than “Calf Died” individuals ($P_{\text{MCMC}} < 0.001$ for
285 strongyles and *E. cervi*), but these reproductive status categories did not differ in total IgA
286 ($P_{\text{MCMC}} = 0.502$) or anti-Tc IgA ($P_{\text{MCMC}} = 0.336$; Figure 2-3). “Calf Died” individuals did not differ

287 from “No Calf” females in strongyle, *E. cervi* or anti-Tc IgA levels (Figure 2) but had lower total
288 IgA levels ($P_{\text{MCMC}}=0.018$). That is, “Calf Died” individuals had lower total IgA than “No Calf”
289 females, but with similar parasite intensities, while “Calf Survived” individuals had the same
290 antibody levels as “Calf Died” individuals, but with increased parasite intensities. *F. hepatica*
291 was not associated with reproductive status, but decreased with age ($P_{\text{MCMC}}=0.004$) as did *E.*
292 *cervi* ($P_{\text{MCMC}}<0.001$; Figure SI1, SI7).

293 Strongyles and both antibodies all exhibited the same seasonality, peaking in the spring and
294 being lowest in the autumn, with the summer intermediate (Figure 3, all differences
295 $P_{\text{MCMC}}<0.001$). *F. hepatica* was higher in the spring than in the summer or autumn
296 ($P_{\text{MCMC}}<0.034$), and *E. cervi* was lowest in the summer, with the autumn intermediate
297 ($P_{\text{MCMC}}<0.001$). There was also some between-year variation: strongyle levels increased
298 between 2015 and 2016, and again in 2017 (all $P_{\text{MCMC}}<0.001$), while total IgA levels decreased
299 in 2017 compared to 2015 and 2016 ($P_{\text{MCMC}}<0.024$). Anti-Tc IgA was also lower in 2017 than
300 2016 ($P_{\text{MCMC}}<0.001$). Inclusion of season-by-status interactions improved strongyle model fit
301 ($\Delta\text{DIC}=-3.79$), but did not improve the fit of any other models ($\Delta\text{DIC}<2$). Fixed status effects
302 remained largely unchanged in magnitude or significance, suggesting that the observed
303 associations with reproductive status were consistent across seasons (Figure 3). All
304 interaction terms implied an attenuation of reproductive status effects from summer through
305 winter to spring, rather than any major qualitative change in this association (Figure 3). Both
306 “Calf Died” and “Calf Survived” females had increased antibody levels and decreased parasite
307 intensities relative to “No Calf” females over this period. See Figure SI2 for a comparison of
308 the full model estimates and DIC changes when a season-by-status interaction was included.

309 Pregnancy models examining April and November samples revealed marginally lower total
310 IgA in pregnant females ($P_{\text{MCMC}}=0.034$, Figure 2, SI1, SI3). Including pregnancy status in our
311 models did not alter the direction or significance of reproductive status effects; in fact, in the
312 case of total IgA and anti-Tc IgA it increased the significance of the “Calf Survived” category’s
313 effect (Figure SI3). It also slightly improved the fit of the total IgA model ($\Delta\text{DIC}=-3.00$). No

314 other models were impacted notably by the inclusion of the pregnancy term, although it slightly
315 reduced the effect size of the “Calf Survived” category in influencing strongyle count (Figure
316 SI3). Although the “Calf Died” category was not significant in the total IgA pregnancy model
317 as it was in the full model, the fact that adding and removing pregnancy as a variable had very
318 little effect on the model estimate (Figure SI3) implies that this did not arise from confounding
319 effects of pregnancy.

320 None of the calving traits modelled (parturition date, calf birth weight or calf sex) were
321 associated with maternal parasite or antibody levels (Figure 2, SI1).

322 The fixed effects of the multivariate model were very similar to those of the full models (Figure
323 SI4). The raw correlations between the response variables of the model are displayed in
324 Figures 4 and SI8. Phenotypic correlations (R_p) derived from the variance structure of the
325 multivariate model are as follows. There were strong positive correlations between strongyles
326 and *E. cervi* ($R_p=0.26$, $P_{MCMC}<0.001$) and between total and anti-Tc IgA ($R_p=0.424$,
327 $P_{MCMC}<0.001$). Strongyle count was also weakly negatively correlated with total IgA ($R_p=-$
328 0.074 , $P_{MCMC}=0.016$, Figure SI8) and more strongly with anti-Tc IgA ($R_p=-0.142$, $P_{MCMC}<0.001$,
329 Figure 4).

330 Discussion

331 Lactation is associated with weaker immunity or increased parasitism in a range of mammals
332 (Festa-Bianchet, 1989; Jones et al., 2012; Rödel et al., 2016; Woodroffe & Macdonald, 1995).
333 In accordance with these studies, we found that lactating females had both decreased
334 antibody levels and increased parasite counts relative to non-reproductive females. In
335 contrast, gestation is rarely found to be costly for immunity or parasitism in mammals (Irvine,
336 Corbishley, Pilkington, & Albon, 2006; Rödel et al., 2016; Woodroffe & Macdonald, 1995), and
337 carries no detectable fitness cost in the Rum red deer (Clutton-Brock et al., 1989). Here, deer
338 that gave birth to a calf that died as a neonate, thereby incurring a limited lactation cost, had
339 lower total IgA levels than non-reproducing females. Gestation therefore carried an immune

340 cost in this study. We predicted that resource depletion incurred through investment in a given
341 reproductive trait would lead to reduced immune investment, and that this would lead to
342 increased parasite intensity (Knowles et al., 2009; Sheldon & Verhulst, 1996). Our results
343 deviated from our expectations in two ways: first, gestation's long-lasting immune cost was
344 not accompanied by increased parasite intensities. Second, the considerable additional
345 resource investment of prolonged lactation was not associated with additional immune costs
346 relative to gestation, but was instead associated with an increase in parasite intensity. These
347 results have two implications: reproduction-immunity tradeoffs were unlikely to be mediated
348 by simple resource reallocation, and reproduction-parasitism tradeoffs were unlikely to be
349 mediated solely by immunity – despite our observation that higher immune investment was
350 associated with lower parasite counts between individuals (Figure 4, SI8).

351 If gestation's lack of detectable fitness cost in our study population (Clutton-Brock et al., 1989)
352 demonstrates a small resource cost, why was gestation associated with reduced total IgA
353 levels, and why did the additional resource cost of lactation not decrease antibody levels
354 further? First, it is possible that reproductive hormones suppress the immune system without
355 being sensitive to resource availability (Foo, Nakagawa, Rhodes, & Simmons, 2017;
356 Svensson et al., 1998). Similarly, gestation may lead to alterations in immune investment and
357 antibody production, so that lower IgA resulted from selective investment in alternative
358 immune cells or functions rather than from lower absolute resource investment in immunity.
359 Reproductive mammals are commonly found to exhibit different (rather than weaker)
360 immunity, but specific patterns of immune prioritisation are unpredictable. For example,
361 reproductive vampire bats (*Desmodus rotundus*) prioritise innate over adaptive immunity
362 (Becker et al., 2018), while reproductive rabbits (*Oryctolagus cuniculus*) exhibit reduced white
363 blood cell counts but stronger humoral immunity (Rödel et al., 2016). Assessing whether
364 reproductive deer invest preferentially in aspects of immunity other than mucosal antibodies
365 would therefore necessitate examining numerous additional immune measures – however, in

366 this study we were restricted to quantifying mucosal antibodies using noninvasive faecal
367 samples as the deer are rarely handled as adults (Clutton-Brock et al., 1982).

368 Alternatively, gestation and early lactation may necessitate export of IgA from the gut to the
369 blood for transfer to offspring (Jeffcoate et al., 1992; Sheldrake, Husband, Watson, & Cripps,
370 1984). In ungulates a substantial proportion of maternal antibody transfer occurs via the
371 colostrum in the first few days of life (Hurley & Theil, 2011). It is feasible that this diversion of
372 IgA from the gut occurs around parturition and is detectable for an extended period of time
373 without an underlying resource allocation tradeoff, creating lower IgA levels in all reproductive
374 females regardless of their calf's survival. The necessity of transferring immune effectors to
375 offspring may therefore be an important obligate mechanism contributing to reduced antibody
376 levels in reproductive wild mammals (Rödel et al., 2016). In a proposed mechanism for the
377 periparturient rise in helminth egg count in domestic sheep, exportation of IgA from the gut
378 around parturition releases helminths from immune control (Jeffcoate et al., 1992). However,
379 in this study, the lower total IgA and intermediate anti-Tc IgA levels in female deer that only
380 paid the cost of gestation were not accompanied by any change in parasitism. This is
381 surprising, given that the results of our multivariate model implied that both IgA measures are
382 representative of increased resistance to strongyles (Figure 4, SI8).

383 If antibody levels were indicative of investment in protective immunity, how were the deer that
384 paid the immune cost of gestation able to maintain low strongyle and *E. cervi* intensities? Or,
385 what produced the higher parasite counts in lactating individuals? Lactating females' anti-Tc
386 IgA levels were significantly lower than nonreproductive females', which could explain their
387 increased parasitism in the absence of a contrast with any other reproductive categories.
388 However, levels of total and anti-Tc IgA in lactating females were not detectably lower than
389 those exhibited by females that paid the cost of gestation (Figure 2). This disparity suggests
390 that additional processes such as exposure were important in driving the high parasite
391 intensities in lactating females (Knowles et al., 2009; Sheldon & Verhulst, 1996). The energetic
392 and resource demand of milk production necessitates substantially increased forage intake

393 and grazing time (Hamel & Côté, 2008, 2009), and may reduce feeding selectivity or the ability
394 to exhibit parasite avoidance behaviours (Hutchings, Judge, Gordon, Athanasiadou, &
395 Kyriazakis, 2006; Speakman, 2008). Thus, lactating females may suffer increased exposure
396 to infective larvae, resulting in higher parasite burdens. This mechanism offers an explanation
397 for our observation that lactation was associated with increased parasite counts, while
398 gestation was not, as individuals that lost their calf as a neonate were not then saddled with a
399 necessity for such high resource acquisition. Based on our results, we suggest that severe
400 effects of mammalian reproduction on parasite infection are partly mediated by exposure as a
401 result of constraints on resource acquisition, foraging selectivity, and antiparasite behaviours,
402 in addition to increased immune susceptibility.

403 Effects of foraging on exposure can profoundly affect epidemiological dynamics: for example,
404 in the water flea *Daphnia dentifera*, temperature-induced increases in food intake can increase
405 the magnitude of fungal pathogen epidemics (Shocket et al., 2018). Similar processes may
406 act in the deer, if spatiotemporal variation in climatic conditions, deer density, or food
407 abundance modify feeding behaviour or the threat of exposure. In particular, strongyle and *E.*
408 *cervi* parasitism will be further exacerbated in years and areas of the study system where deer
409 density is high and food availability is low (Wilson, Grenfell, Pilkington, Boyd, & Gulland, 2004).
410 It is possible that higher parasitism in reproductive individuals will reduce their fitness, thereby
411 producing lactation's fitness cost – and, by extension, gestation's lack of fitness cost – in this
412 system (Clutton-Brock et al., 1989; Froy et al., 2016; Harshman & Zera, 2007; Williams, 1966).
413 If exposure is determining parasitism and parasitism is reducing fitness, we would expect that
414 parasite-mediated life history tradeoffs would be exacerbated in years and areas of high deer
415 density, as more deer will translate to higher levels of pasture contamination (Wilson et al.,
416 2004). Future investigations in this system could address the hypothesized role of parasite
417 exposure and foraging behaviour in reproductive tradeoffs, using available census data
418 (Clutton-Brock et al., 1982; Froy et al., 2018) to examine how annual, seasonal and spatial

419 variation in habitat use and deer density correlate with environmental larval counts, parasite
420 intensity, and the severity of reproductive tradeoffs.

421 Reproductive tradeoffs are a potential driver of seasonal dynamics of immunity and parasitism,
422 in which periodic reproduction-associated relaxation of immunity leads to increased parasitism
423 (Martin, Weil, & Nelson, 2008). Our results do not support this mechanism for several reasons:
424 all status categories exhibited seasonality of antibodies, strongyles, and *E. cervi* rather than
425 only reproductive individuals; reproductive increases in parasitism were not linked to lower
426 immunity; and immunity did not correlate with resource availability, being highest in April, when
427 the deer are in poor condition, having just survived the winter. In fact, antibody levels and
428 strongyle counts correlated positively across seasons despite their negative correlation among
429 individuals. This suggests that seasonality of propagule output is adaptive for helminths,
430 facilitating highest transmission when environmental conditions are favourable and
431 immunologically naïve calves are present, and leading to seasonal upregulation of immunity
432 in warmer months to combat increased exposure (Møller, Erritzøe, & Saino, 2003; Wilson et
433 al., 2004).

434 This study describes unexpected and complex interrelationships between different
435 components of mammalian reproduction, immunity, and parasitism in the wild. We suggest
436 that classical resource allocation mechanisms which are often hypothesised to underlie
437 tradeoffs with immunity (e.g. Sheldon & Verhulst 1996; Deerenberg *et al.* 1997; French *et al.*
438 2007) are insufficient to explain many of the patterns seen in wild mammals, corroborating
439 findings in other taxa (Stahlschmidt *et al.*, 2013; Svensson *et al.*, 1998). As such, studies
440 examining such tradeoffs in mammals should consider mechanistic links between
441 reproduction and immunity, resource acquisition limitations, and exposure components of
442 parasitism, particularly by quantifying both immunity and parasitism simultaneously (Bradley
443 & Jackson, 2008; Graham *et al.*, 2011). The potential complexity of such interrelationships
444 may contribute to the relative rarity of conclusive evidence for reproduction-immunity-
445 parasitism tradeoffs in mammals.

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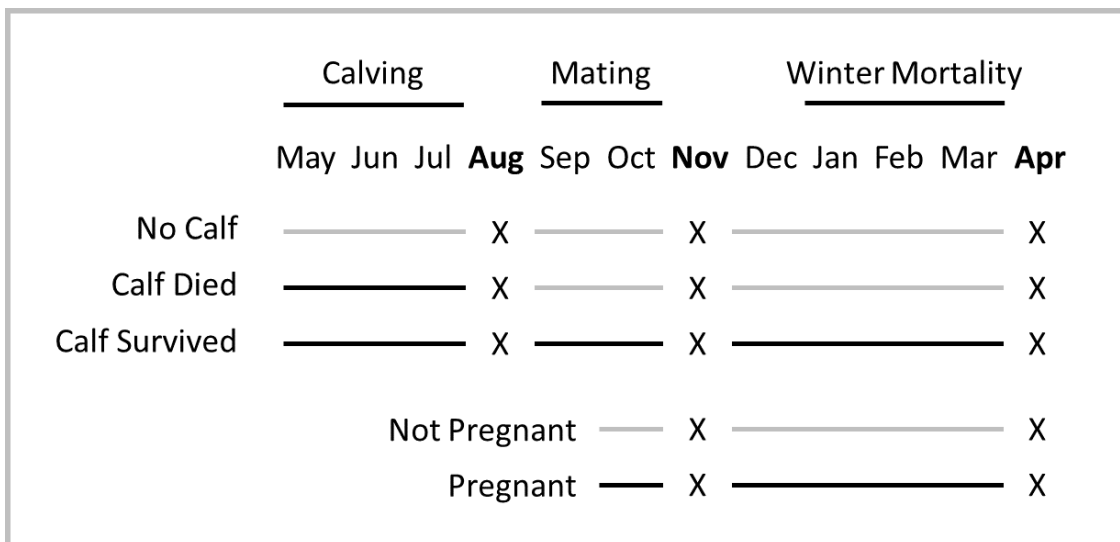
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Figures

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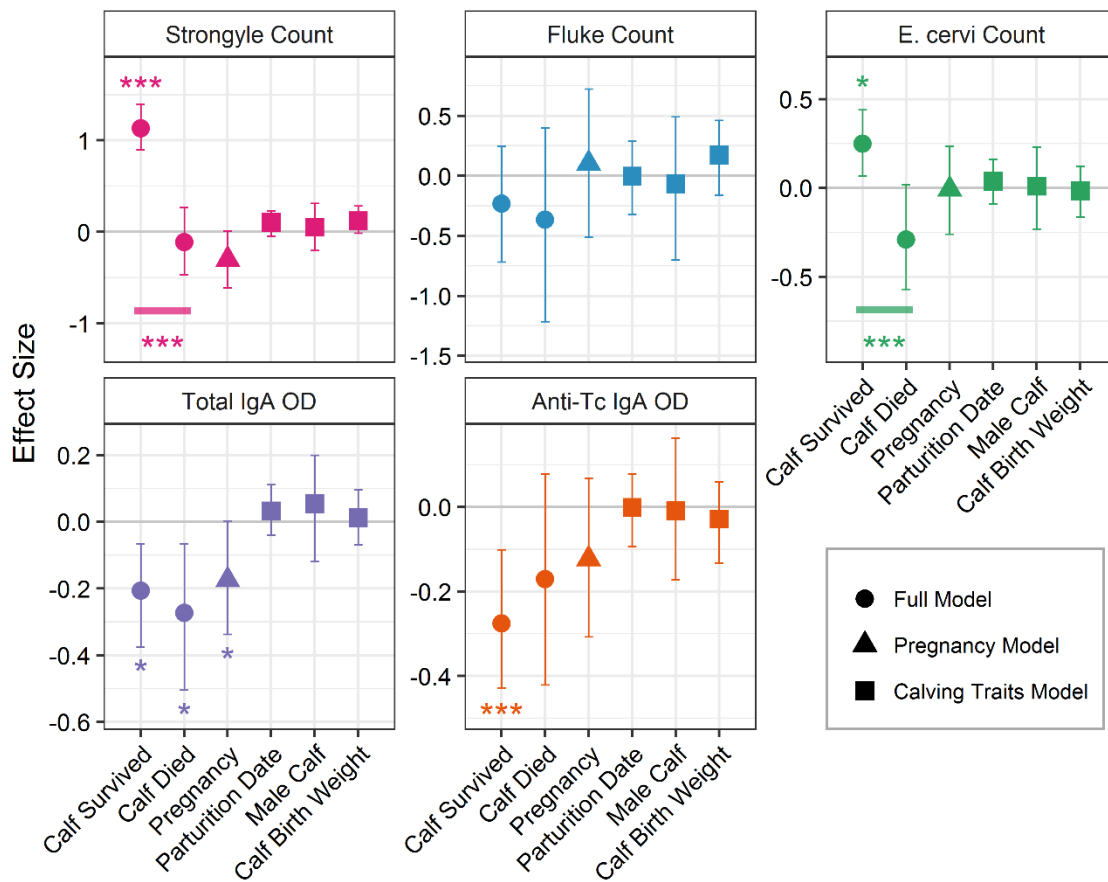


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613 Figure 1: The faecal sampling regime in the context of a deer reproductive cycle (“deer
614 year”). A cross (X) represents a two-week sampling trip. The deer year begins on May 1st
615 when calving begins; individuals are assigned to one of three reproductive status categories
616 (top three lines) according to the birth and survival of their calf over the course of the
617 following year. Individuals are also assigned a pregnancy status in November and April
618 based on reproduction in the following calving season (bottom two lines). Black lines
619 represent periods in which the calf is living or the female is pregnant; grey lines represent
620 periods in which the calf is dead or non-existent or the female is not pregnant.

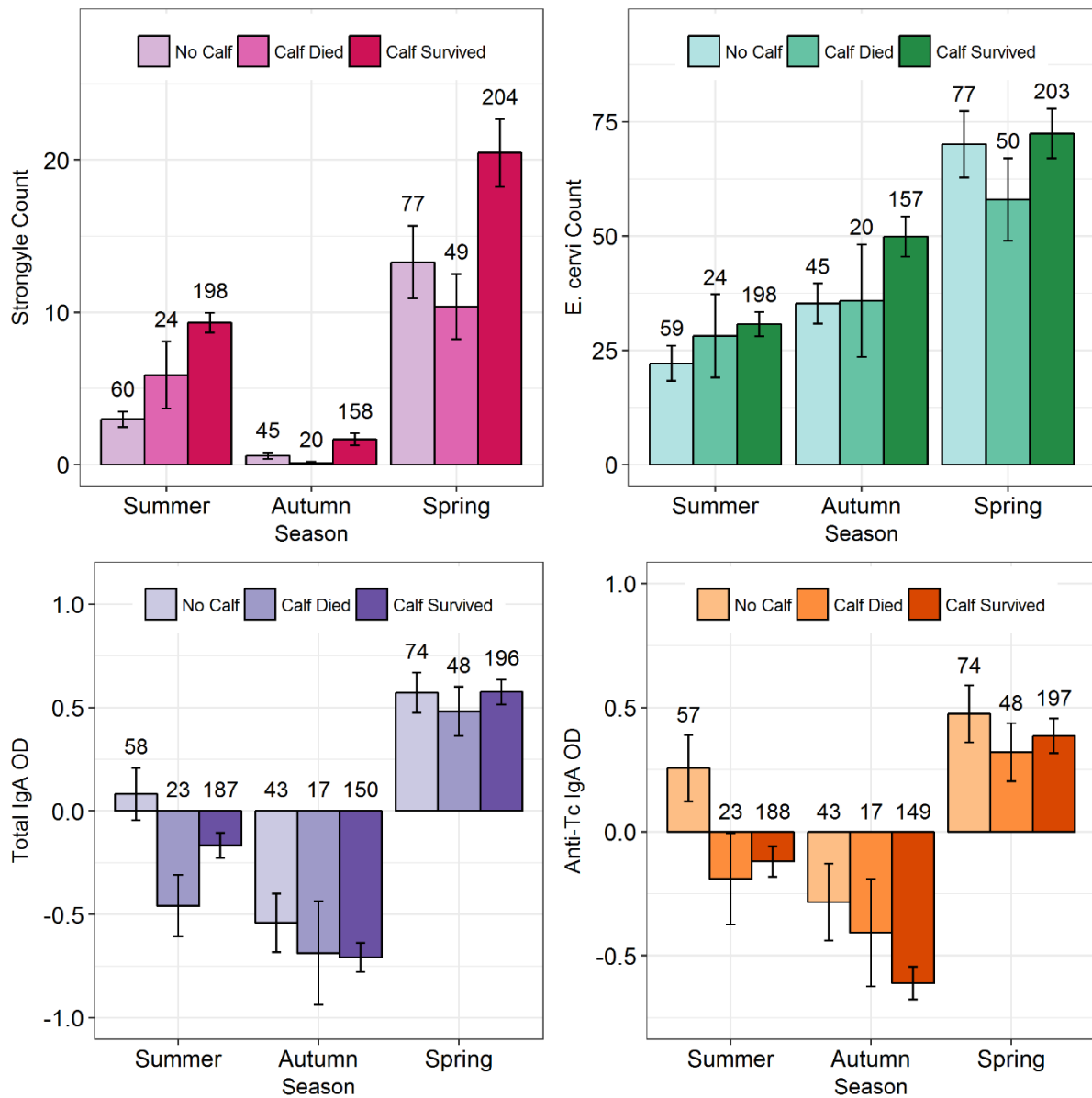
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626 Figure 2: Model outputs depicting the effects of reproductive traits, derived from all three
 627 univariate model sets. Points and error bars represent model estimates and 95% credibility
 628 estimates. Effect sizes for categorical variables (status, pregnancy and calf sex) denote
 629 differences from the first (absent) category of each, contained in the intercept (“No Calf”,
 630 “Not Pregnant” and “Female Calf” respectively). Effect sizes for continuous variables
 631 (parturition date and calf birth weight) represent the change in the response variable
 632 associated with a change of one standard deviation of the explanatory variable. Asterisks
 633 represent significant differences derived from MCMCglimm posterior distribution overlaps:
 634 ***, ** and * denote $P < 0.001$, $P < 0.01$ and $P < 0.05$ respectively. Bars denote differences
 635 between status categories.

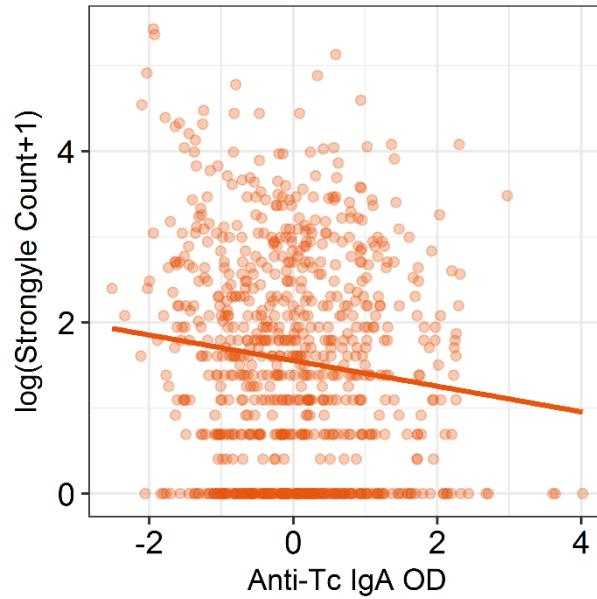


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638 Figure 3: Bar charts displaying raw mean (+/- SE) parasite counts and antibody levels of
 639 each reproductive status category in each season. Antibody measures were taken from the
 640 residuals of a model with square root-transformed (total IgA) or cube root-transformed (anti-
 641 Tc IgA) antibody OD as the response variable and including collection variables as fixed
 642 effects. Numbers above bars denote sample sizes.

643



644

645

646 Figure 4: Correlation between anti-*Teladorsagia circumcincta* IgA levels and strongyle count,
647 using the raw data. Individuals with higher anti-Tc IgA had lower strongyle counts
648 (multivariate model phenotypic correlation $R_p = -0.142$, $P_{\text{MCMC}} < 0.001$). The y axis is on the
649 $\log_e(\text{count}+1)$ scale to aid interpretation; the x axis data were taken from the residuals of a
650 model with cube root-transformed anti-Tc IgA as the response variable and including
651 collection variables as fixed effects. For this figure, these residuals were centred within
652 sampling trips to have a mean of 0 and a standard deviation of 1 to avoid a positive
653 correlation arising from shared seasonal and annual effects.

654

655 Supplementary information for Albery et
656 al., (2018): Reproduction has different
657 costs for immunity and parasitism in a wild
658 mammal

659 Section One: Table SI1

660

Factor	Prevalence (%)	Model Sample Sizes			Repeatability
		Full	Pregnancy	Calving Traits	
Strongyles	76	835	518	571	0.21 (0.16-0.28)
<i>F. hepatica</i>	33	824	517	571	0.11 (0.06-0.17)
<i>E. cervi</i>	95	833	518	571	0.39 (0.34-0.45)
Total IgA		796	499	550	0 (0-0.08)
Anti-Tc IgA		796	497	547	0.25 (0.17-0.32)

661

662 Table SI1: model sample sizes and repeatabilities (95% credibility intervals in brackets).

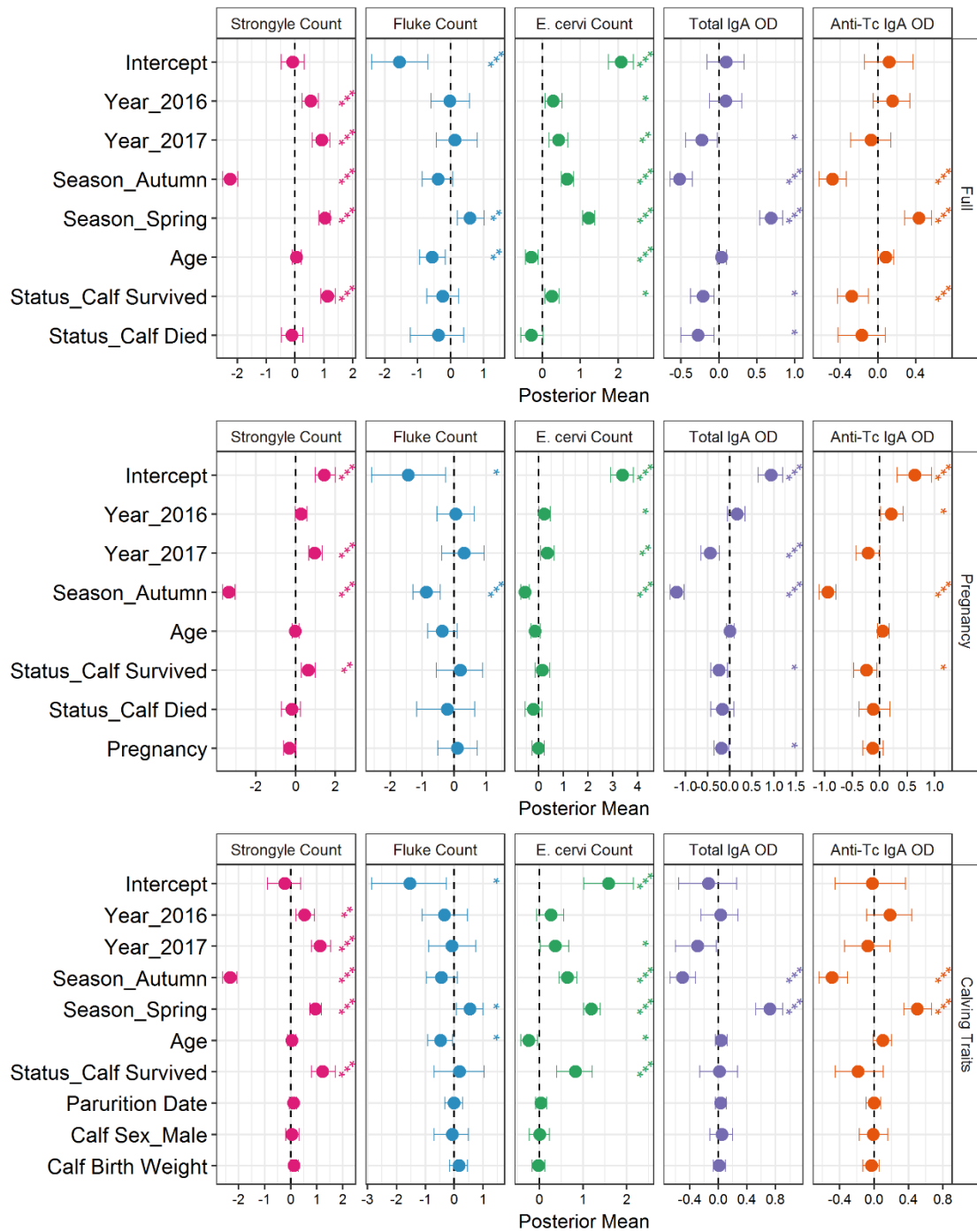
663 Section Two: Model output

664 This section includes model outputs for the fixed effects of all models we ran. We first include
665 the main models reported in the study (Figure SI1). We then compare these results with a set
666 of modifications that we investigated.

667 The next (Figure SI2) displays the full models including a season by status interaction to
668 display the way this affected the estimates, and to demonstrate the seasonal effects.
669 Generally, including a season by status interaction did not improve the model fit or change our
670 findings. The exception for this is the strongyle model (delta DIC = -3.79). There was, however,
671 a general trend for the differences between status categories to decrease over the autumn
672 and spring seasons as can be seen in Figure 3 in the main text.

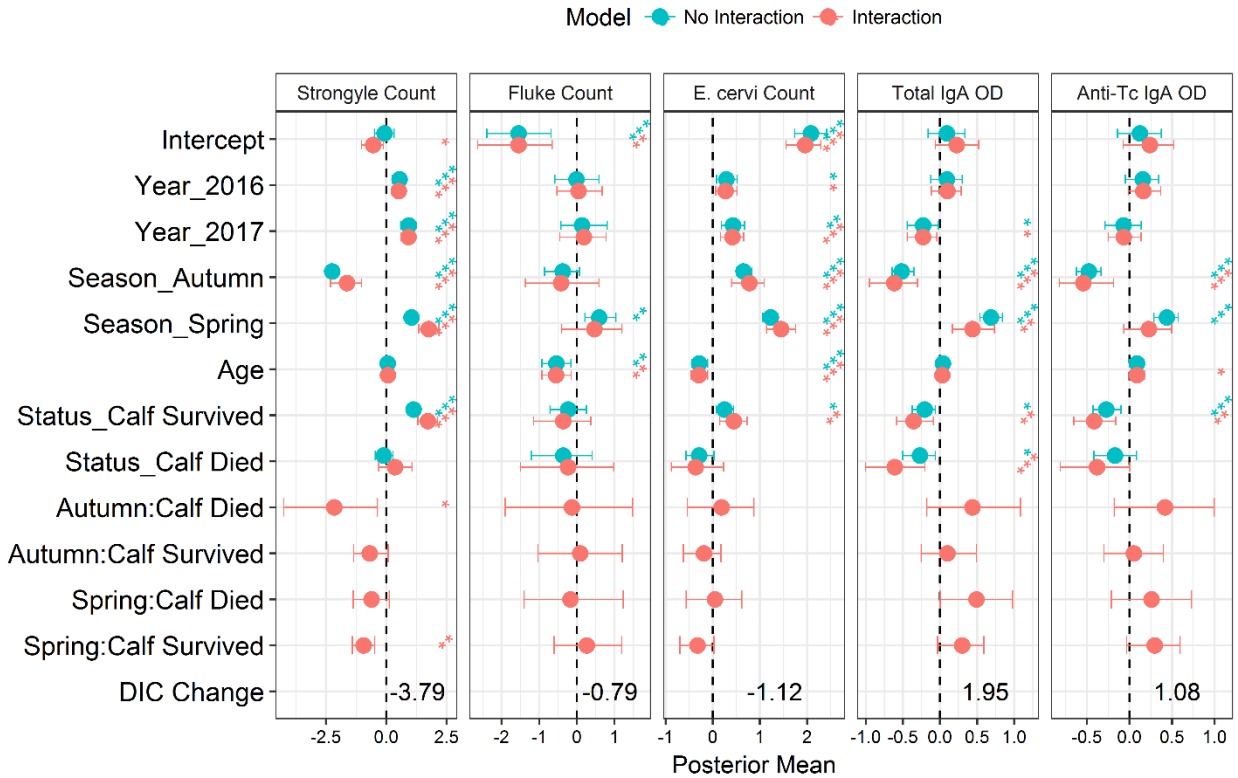
673 Figure SI3 displays the effects from pregnancy models when we either included or removed
674 pregnancy as a fixed effect, to investigate whether this affected the estimates of each status
675 category's effect. Inclusion of the pregnancy term slightly reduced the significance of the "Calf
676 Survived" effect in the strongyle model, and increased the effects of "Calf Survived" in both
677 the total IgA and anti-Tc IgA models. It also improved the fit of the total IgA model (delta DIC
678 = -3.71). Otherwise, pregnancy had little effect.

679 Figure SI4 displays the results from the full models again, compared with the results from the
680 multivariate models. The models were extremely similar, with only small changes in effect
681 sizes and significance. This validates our use of the model to derive phenotypic correlations.



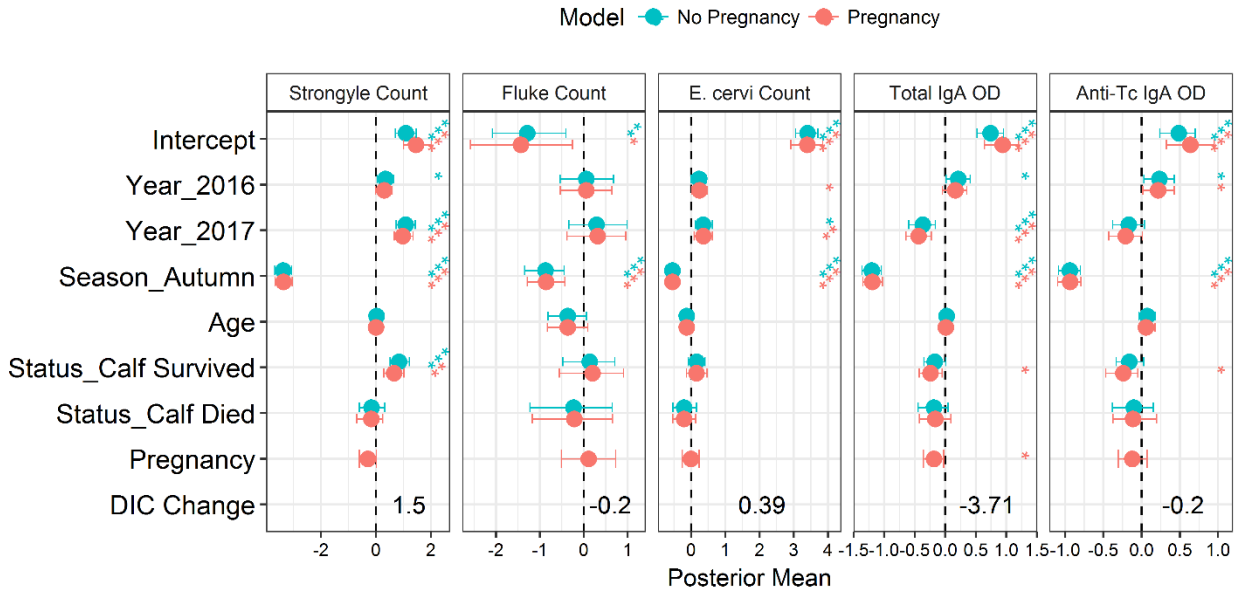
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683 Figure S11: Effect size estimates from the three model sets (full dataset, pregnancy models
 684 and calf traits models). Effect sizes for categorical variables denote differences from the first
 685 (absent) category of each, contained in the intercept. Effect sizes for continuous variables
 686 represent the change associated with a change of one standard deviation of the variable in
 687 question. Points and error bars represent model estimates and 95% credibility estimates for
 688 each of the five full models. Asterisks indicate the significance of variables: ***, ** and *
 689 indicate P < 0.001, P < 0.01 and P < 0.05 respectively.



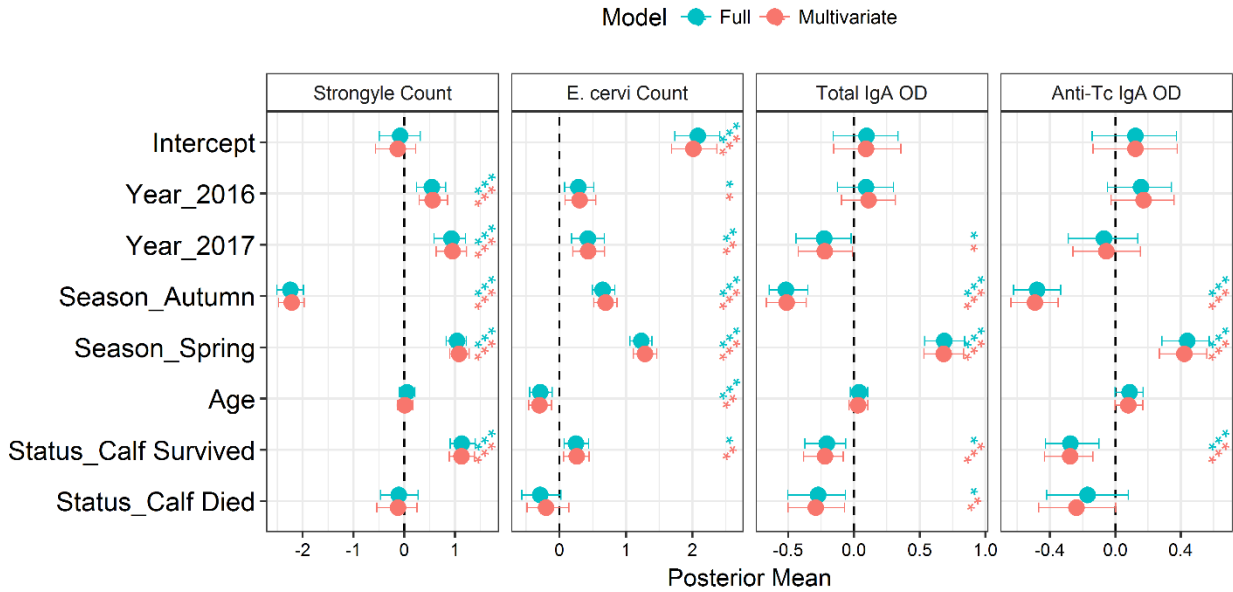
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691 Figure S12: Comparison between outputs from full models excluding and including a season
 692 by status interaction. Points and error bars represent model estimates and 95% credibility
 693 estimates for each of the five full models, with and without interactions. Effect sizes for
 694 categorical variables denote differences from the first (absent) category of each, contained in
 695 the intercept. Effect sizes for continuous variables represent the change associated with a
 696 change of one standard deviation of the variable in question. Asterisks indicate the
 697 significance of variables: ***, ** and * indicate $P < 0.001$, $P < 0.01$ and $P < 0.05$ respectively.
 698 DIC Change represents the change in DIC that occurred when an interaction was included.
 699 Including an interaction did not have a substantial effect on most of the original estimates,
 700 and only the fit of the strongyle model was significantly improved by its inclusion.



701

702 Figure S13: Comparison between outputs from pregnancy models, excluding and including
703 pregnancy as a covariate to investigate whether this changed the model estimates for the
704 reproductive status categories. Points and error bars represent model estimates and 95%
705 credibility estimates for each of the five full models, without and with pregnancy as a
706 covariate. Effect sizes for categorical variables denote differences from the first (absent)
707 category of each, contained in the intercept. Effect sizes for continuous variables represent
708 the change associated with a change of one standard deviation of the variable in question.
709 Asterisks indicate the significance of variables: ***, ** and * indicate $P < 0.001$, $P < 0.01$ and
710 $P < 0.05$ respectively. DIC Change represents the change in DIC that occurred when
711 pregnancy was included. Including pregnancy as a covariate slightly decreased the effect
712 size of “Calf Survived” for strongyles and increased it for total IgA and anti-Tc IgA, but
713 otherwise had little effect on the estimates. Pregnancy also significantly reduced total IgA
714 levels and improved the total IgA model fit.



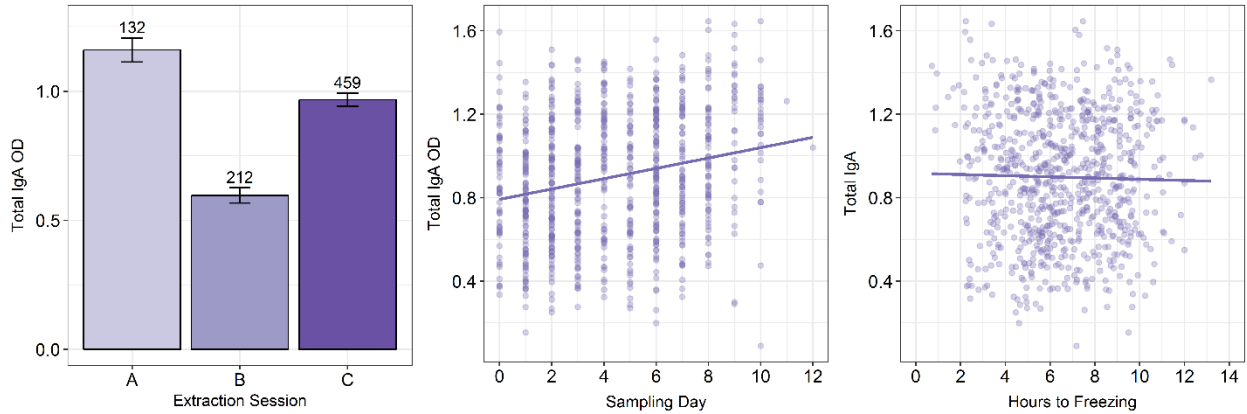
715

716 Figure SI4: Comparison of the fixed effect estimates from the full models and the multivariate
717 model. Points and error bars represent model estimates and 95% credibility estimates for
718 each of the five full models and the equivalent fixed effects in the multivariate model. Effect
719 sizes for categorical variables denote differences from the first (absent) category of each,
720 contained in the intercept. Effect sizes for continuous variables represent the change
721 associated with a change of one standard deviation of the variable in question. Asterisks
722 indicate the significance of variables: ***, ** and * indicate $P < 0.001$, $P < 0.01$ and $P < 0.05$
723 respectively.

724 Section Three: Additional Figures

725 This section contains some figures displaying patterns in the data which are of interest. This
726 includes the effects of faecal collection variables on antibody levels, age effects and
727 correlations between response variables.

728



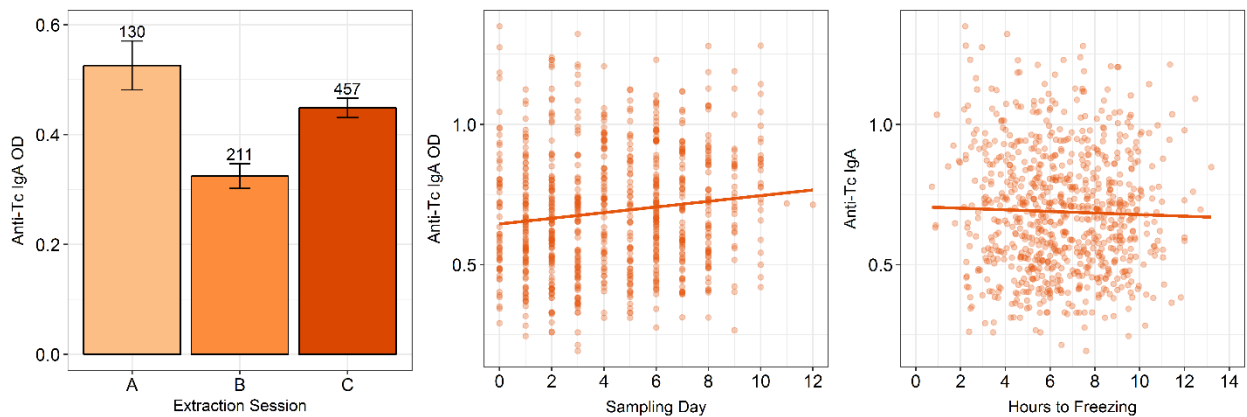
729

730

731 Figure SI5: Impact of faecal collection factors on total IgA level (extraction session, day of
732 collection and hours to freezing). Y axes for figures B and C have been square root
733 transformed.

734

735

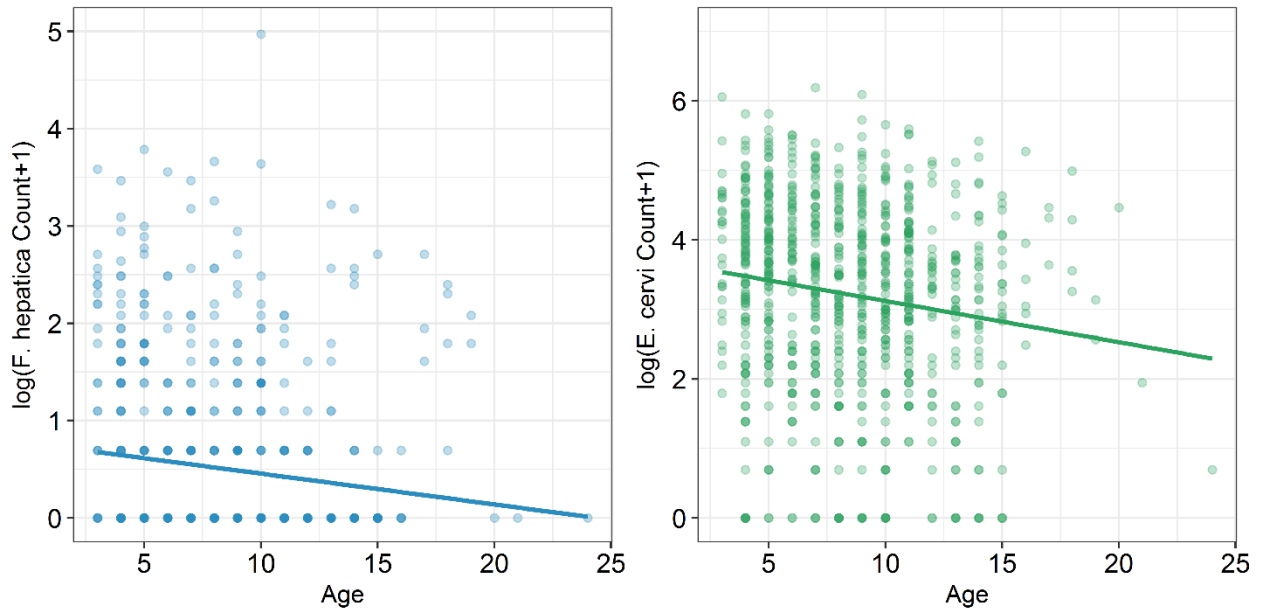


736

737

738 Figure SI6: Impact of faecal collection factors on anti-Tc IgA level (extraction session, day of
739 collection and hours to freezing). Y axes for figures B and C have been cube root
740 transformed.

741

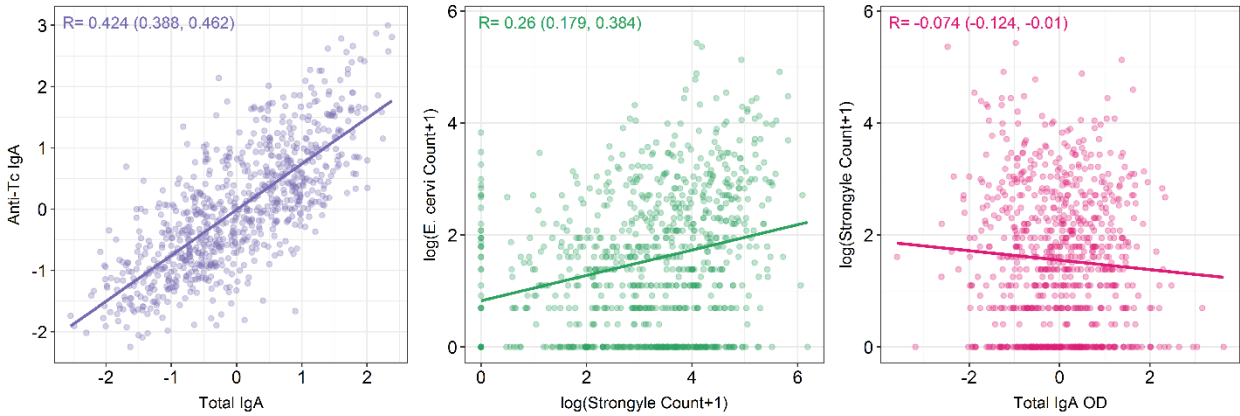


742

743 Figure S17: Age trends in parasite counts using the raw correlations (*F. hepatica* and *E.*
744 *cervi*). Y axes have been log(count+1)-transformed.

745

746



747

748 Figure S18: Correlations between response variables (Total IgA and anti-Tc IgA; Strongyles
749 and *E. cervi*; Total IgA and Strongyles). Model-derived phenotypic correlations (R_p) are
750 included, with 95% credibility intervals. Both antibody measures are based on the residuals
751 from a model including extraction session, day of collection and hours to freezing, with
752 transformed antibody OD as response variable (square root for the total IgA and cube root for
753 anti-Tc IgA). In the strongyle figure total IgA was scaled within each sampling trip to have a
754 mean of 0 and a standard deviation of 1 to avoid a positive correlation arising from shared
755 seasonal effects.