Rare variants imputation in admixed populations: 1 Comparison across reference panels and bioinformatics tools. 2 3 Sanjeev Sariya<sup>a,b</sup>, Joseph Lee<sup>a,b</sup>, Richard Mayeux<sup>a,b,c</sup>, Badri N. Vardarajan<sup>a,b</sup>, Dolly Reyes-4 Dumever<sup>a,b</sup>, Jennifer Manly, Adam Brickman, Rafael Lantigua<sup>d</sup>, Martin Medrano<sup>e</sup>, Ivonne Z. 5 Jimenez-Velazquez<sup>f</sup> and Giuseppe Tosto<sup>a,b,c</sup> 6 7 a. Taub Institute for Research on Alzheimer's Disease and the Aging Brain, College of 8 Physicians and Surgeons, Columbia University. 630 West 168th Street, New York, NY 9 10032. 10 b. The Gertrude H. Sergievsky Center, College of Physicians and Surgeons, Columbia 11 University. 630 West 168th Street, New York, NY 10032. 12 c. Department of Neurology, College of Physicians and Surgeons, Columbia University and 13 the New York Presbyterian Hospital. 710 West 168th Street, New York, NY 10032 14 d. Medicine College of Physicians and Surgeons, and The Department of 6Epidemiology, 15 School of Public Health, Columbia University, New York, NY, USA 16 17 e. School of Medicine, Pontificia Universidad Catolica Madre y Maestra, Santiago, **Dominican Republic** 18 f. Department of Medicine, Geriatrics Program, School of Medicine, University of Puerto 19 Rico, San Juan, Puerto Rico 20 21 22 23 24 Word counts: Abstract 344; Main Text 3,821; Tables: 4; Figures: 2; References: 35 25 26 27 28 29 Correspondence: 30 Giuseppe Tosto MD PhD 31 Taub Institute for Research on Alzheimer's Disease and the Aging Brain 32 630 West 168<sup>th</sup> Street 33 New York, NY 10032 34 35 Tel: +1 212 305 9274 36 Email: gt2260@cumc.columbia.edu 37

#### 38 Abstract.

**Background**: Imputation has become a standard approach in genome-wide association studies 39 (GWAS) to infer in silico untyped markers. Although feasibility for common variants imputation 40 is well established, we aimed to assess rare and ultra-rare variants' imputation in an admixed 41 42 Caribbean Hispanic population (CH). 43 **Methods**: We evaluated imputation accuracy in CH (N=1,000), focusing on rare  $(0.1\% \le \text{minor})$ 44 45 allele frequency (MAF)  $\leq 1\%$ ) and ultra-rare (MAF < 0.1%) variants. We used two reference panels, the Haplotype Reference Consortium (HRC; N=27,165) and 1000 Genome Project 46 (1000G phase 3; N=2,504) and multiple phasing (SHAPEIT, Eagle2) and imputation algorithms 47 48 (IMPUTE2, MACH-Admix). To assess imputation quality, we reported: a) high-quality variant counts according to imputation tools' internal indexes (e.g. IMPUTE2 "Info" ≥80%). b) 49 50 Wilcoxon Signed-Rank Test comparing imputation quality for genotyped variants that were 51 masked and imputed; c) Cohen's kappa coefficient to test agreement between imputed and whole-exome sequencing (WES) variants; d) imputation of G206A mutation in the PSEN1 52 53 (ultra-rare in the general population an more frequent in CH) followed by confirmation genotyping. We also tested ancestry proportion (European, African and Native American) 54

against WES-imputation mismatches in a Poisson regression fashion.

56

**Results**: SHAPEIT2 retrieved higher percentage of imputed high-quality variants than Eagle2
(rare: 51.02% vs. 48.60%; ultra-rare 0.66% vs 0.65%, Wilcoxon p-value < 0.001). SHAPEIT-</li>
IMPUTE2 employing HRC outperformed 1000G (64.50% vs. 59.17%; 1.69% vs 0.75% for high-

60 quality rare and ultra-rare variants, respectively; Wilcoxon p-value < 0.001). SHAPEIT-

61 IMPUTE2 outperformed MaCH-Admix. Compared to 1000G, HRC-imputation retrieved a

62	higher number of high-quality rare and ultra-rare variants, despite showing lower agreement
63	between imputed and WES variants (e.g. rare: 98.86% for HRC vs. 99.02% for 1000G). High
64	Kappa ( $K = 0.99$ ) was observed for both reference panels. Twelve G206A mutation carriers were
65	imputed and all validated by confirmation genotyping. African ancestry was associated with
66	higher imputation errors for uncommon and rare variants (p-value < 1e-05).
67	
68	Conclusion: Reference panels with larger numbers of haplotypes can improve imputation quality
68 69	<b>Conclusion</b> : Reference panels with larger numbers of haplotypes can improve imputation quality for rare and ultra-rare variants in admixed populations such as CH. Ethnic composition is an
69	for rare and ultra-rare variants in admixed populations such as CH. Ethnic composition is an
69 70	for rare and ultra-rare variants in admixed populations such as CH. Ethnic composition is an important predictor of imputation accuracy, with higher African ancestry associated with poorer

74 Keywords: Rare variants, Imputation, Admixed population, GWAS, 1000G

#### 75 Introduction

Genome-wide association studies (GWASs) are a major tool to identify common variants 76 associated with complex diseases. GWAS can include 550K to over 2M Single Nucleotide 77 Polymorphisms (SNPs) (Ha et al., 2014) to cover the human genome evenly. Although GWAS 78 has shown to be a robust method to identify disease loci of interest, they rarely point to a causal 79 80 coding variant. In fact, microarray SNP chips for GWAS are optimally designed to uncover common variants, often associated with small effect sizes mostly located in intronic and 81 intergenic regions. The focus of genetic investigations has since shifted toward rarer alleles with 82 83 larger effect sizes (Gibson, 2012). With the changing paradigm, imputation of rare variants has become an important topic to enhance the genome coverage in GWAS. Imputation is a process 84 of inferring untyped SNP markers in the discovery population by using densely typed SNPs in 85 external reference panel(s). These 'in silico' markers increase the coverage of association tests 86 while conducting genome-wide association analysis. In addition, large number of SNPs facilitate 87 88 meta-analysis when merging data from different study cohorts. The quality of imputation essentially depends on two parameters: available reference datasets 89 and algorithms that employ those reference datasets. Previous studies have shown that 90 91 imputation quality depends on how well reference panels reflect the study population. To 92 respond to the needs, the 1000 Genome project (1000G), now in its third phase release, has 93 proven to be one of the most frequently used reference panels (Genomes Project et al., 2015). 94 Using these composite reference panels, a number of studies (Pei et al., 2010; Howie et al., 2012; Verma et al., 2014; Liu et al., 2015) have compared imputation accuracy using different 95 96 imputation tools and algorithms, although the results are equivocal. Few studies (Browning and 97 Browning, 2009; Zheng et al., 2012; Zheng et al., 2015) assessed the impact of reference panel

98 size and input data's features - such as density of SNPs - to impute rare variants, suggesting larger size of reference panels work better. Surakka and colleagues (Surakka et al., 2016) 99 assessed accuracy of imputed SNPs by evaluating rate of false polymorphisms in a Finnish 100 101 population using global reference panels – Haplotype Reference Consortium (HRC) release 1, 1000G phase 1 and a local reference panel. They concluded that higher false positive rate was 102 103 observed in imputation from global reference panels compared to imputation performed using a 104 local panel. Other studies (Huang et al., 2015; Das et al., 2016) found imputation accuracy increases with higher number of haplotypes, specifically for variants with MAF  $\leq 0.5\%$ . For 105 106 Hispanic populations, Nelson and colleagues (Nelson et al., 2016) compared imputation performances with 1000G phase 1 (N=1,092) vs. 1000G phase 3 (N=2,504), concluding that 107 108 phase 3 improved accuracy for variants with MAF < 1% by . Further, Nagy and colleagues (Nagy 109 et al., 2017) showed that HRC reference panel provides new insight for novel variants particularly for rare variants in a family-based Scottish study cohort. Aforementioned studies 110 111 highlighted the need of a larger sized reference panel to improve imputation quality. Herzig and colleagues (Herzig et al., 2018) assessed tools for haplotype phasing and their impact on 112 imputation in a population isolate of Campora in southern Italy, and showed that SHAPEIT2, 113 114 SHAPEIT3 and EAGLE2 were highly accurate in phasing; MINIMAC3, IMPUTE4 and 115 IMPUTE2 were found to be reliable for imputation. Roshyara and colleagues (Roshyara et al., 116 2014) compared MaCH-Admix, IMPUTE2, MACH, MACH-Minimac in different ethnicities by 117 evaluating accuracy of correctly imputed SNPs; MaCH-Minimac outperformed SHAPEIT-IMPUTE2 in subsamples of different ethnic groups. These studies demonstrated how employed 118 119 imputation algorithm determines quality of inferred SNPs.

- 121 However, no study to our knowledge has evaluated reference panels in tandem with different
- imputation algorithms to assess imputation quality of inferred SNPs based on MAF in a three-
- 123 way admixed population. Based on these findings, we assessed imputation quality, focusing on
- 124 rare and ultra-rare variants, in a large dataset of Caribbean Hispanics (CH) leveraging available
- 125 GWAS and sequencing data available for our cohort.
- 126
- 127
- 128

## 129 Methods

We will refer SNPs with MAF between 1-5% as "uncommon," 0.1-1% as "rare," and ≤0.1% as
"ultra-rare". We considered SNPs with IMPUTE-Info metric ≥0.40 as "good-quality" and ≥0.80
as "high-quality".

134	<u>GWAS samples and genotyping</u> . We selected randomly 1,000 Caribbean Hispanics as part of an
135	original genotyped cohort of 3,138 individuals: genotyped data can be downloaded at dbGaP
136	Study Accession: phs000496.v1.p1. 719 individuals were derived from Estudio Familiar
137	Investigar Genetica de Alzheimer (EFIGA), a study of familial LOAD; and 281 individuals from
138	the multiethnic longitudinal cohort, Washington Heights, Inwood, Columbia Aging Project
139	(WHICAP). The information on study design, recruitment and GWAS methods for the EFIGA
140	and WHICAP study was previously described in Tosto, G., et al (Tosto et al., 2015).
141	
142	<u>GWAS quality control (QC)</u> . Genotyped data underwent quality control using PLINK (v1.90b4.9
143	64-bit) (Purcell et al., 2007). Briefly, we excluded SNPs with missing rate $\geq$ 5% followed by
144	exclusion of SNPs with MAF $\leq$ 1%. We then removed SNPs with P-value < 1e-6 for Hardy-
145	Weinberg Equilibrium. Samples with missing call rate $\geq 5\%$ were excluded from analysis.
146	
147	Global Ancestry estimation and selection of "true Hispanics". Prior to imputation, we estimated
148	global ancestry using the ADMIXTURE (v.1.3.0) software (Alexander et al., 2009; Zhou et al.,
149	2011). We conducted supervised admixture analyses using three reference populations: African
150	Yoruba (YRI) and non-Hispanic white of European Ancestry (CEU) from the HAPMAP project
151	as representative of African and European ancestral populations; and eight Surui, 21 Maya, 14
152	Karitiana, 14 Pima and seven Colombian individuals from the Human Genome Diversity Project

153	(HGDP) were used to represent native American ancestry (Li et al., 2008). We used ~80,000
154	autosomal SNPs that were: I) genotyped in all three datasets (Caribbean Hispanics, 1000G and
155	HGDP); II) common (i.e. MAF >5 %); and III) in linkage equilibrium. Supervised admixture
156	analyses with the three reference populations (YRI, CEU, and Native Americans) revealed that
157	European lineage accounted for most of the ancestral origins (59%), followed by African (33%)
158	and native American ancestry (8%). We then selected only individuals with at least 1% of all
159	three ancestral populations.
160	
161	<u>Reference panels.</u> HRC reference panel contained over 39M SNPs from 27,165 individuals who
162	participated in 17 different studies (Table 1). The data were downloaded from the Wellcome
163	Trust Sanger Institute (WTSI).
164	1000G phase 3 reference panel contained over 81M SNPs from 2,504 individuals
165	(https://mathgen.stats.ox.ac.uk/impute/1000GP_Phase3.tgz). It includes 26 ethnic groups, with
166	most variants rare, approximately 64 million had MAF <0.5%; approximately 12 million had a
167	MAF between 0.5% and 5%; and approximately 8 million have MAF $>5\%$ . In order to perform
168	imputation with MaCH-Admix, 1000G Phase 3 pre-formatted data were downloaded from
169	ftp://yunlianon:anon@rc-ns-
170	ftp.its.unc.edu/ALL.phase3_v5.shapeit2_mvncall_integrated.noSingleton.tgz that contained over
171	47M SNPs.
172	The subsequent analyses were restricted to autosomal chromosomes, only.
173	
174	Phasing and Imputation Procedures. We compared SHAPEIT2 (Delaneau et al., 2013) and Eagle2
175	(Loh et al., 2016) by phasing and then imputing (see next section) a single chromosome

176 (Chromosome 21), using both reference panels. We refer to SHAPEIT2 as SHAPEIT when used177 in tandem with IMPUTE2 for the remainder of paper.

178

179 Imputation was carried out using two bioinformatics tools: IMPUTE2 (Howie et al., 2009) and

180 MaCH-Admix (Liu et al., 2013). For both, imputation quality ranged from 0 to 1, with 0

indicating complete uncertainty in imputed genotypes, and 1 indicating no uncertainty in

182 imputed genotypes.

IMPUTE2 (version 2.3.2). IMPUTE2 uses an MCMC algorithm to integrate over the space of 183 184 possible phase reconstructions for genotypes data. We conducted imputation in non-overlapping 1MB chunk regions; chunk coordinates were specified using the "*-int*" option. Other options 185 were used with default parameters (Supplementary section 1). Briefly, we used a default 186 187 250KB buffer region to avoid quality deterioration on the ends of chunk region. "-Ne" value as 2000 suggested for robust imputation which scales linkage disequilibrium and recombination 188 189 error rate. MaCH-Admix. We used MaCH-Admix because it uses a method based on IBS 190 matching in a piecewise manner. The method breaks genomic region under investigation into small pieces and finds reference haplotypes that best represent every small piece, for each target 191 192 individual separately. MaCH-Admix imputes in three steps: phasing, estimation of model parameter that includes error rare and recombination rate and lastly, haplotype-based imputation. 193 MaCH-Admix (version Beta 2.0.185) was run on default parameters of 30 rounds, 100 states (--194 195 autoFlip flag). Details can be found in supplementary file (section 1). We initially compared performance between MaCH-Admix and IMPUTE2 using the 1000G reference panel for 196 197 Chromosome 21 only. We then proceeded to impute all remaining chromosomes with the tool 198 that performed better.

199

200	Imputation Performance Metrics. IMPUTE2 uses "Info" parameter to report imputation quality
201	that measures relative statistical information about SNP allele frequency from imputed data. It
202	reflects the information in imputed genotypes relative to the information if only the allele
203	frequency were known. "Info" metric is used to filter poorly imputed SNPs from IMPUTE2 and
204	is reported for all imputed SNPs. In addition, IMPUTE2 uses an internal metric known as R <sup>2</sup> ,
205	reported for genotyped SNPs only: it measures squared correlation between genotyped SNPs and
206	the same SNPs that have been first masked internally and then imputed. MaCH-Admix uses Rsq
207	to report imputation quality. The R <sup>2</sup> metric is also known as variance ratio, calculated as
208	proportion of empirically observed variance (based on the imputation) to the expected binomial
209	variance p(1-p), where p is the minor allele frequency. A threshold of 0.30 is recommended to
210	filter out poorly imputed SNPs.
211	Despite quality measures from IMPUTE2 and MaCH-Admix being highly correlated (Marchini
212	and Howie, 2010), we calculated a <i>r2hat</i> score to generate a single common metric to assess
213	imputation quality across the software (Hancock et al., 2012) (v109,
214	http://www.unc.edu/~yunmli/tgz/r2_hat.v109.tgz).
215	We compared performance of MaCH-Admix and SHAPEIT-IMPUTE2 by: a) Reporting raw
216	SNP counts based on quality (MaCH-Admix "Rsq" and IMPUTE2 "Info"); b) Comparing r2hat
217	for overlapping imputed SNPs from both tools; c) Conducting a Wilcoxon Signed-Rank Test (R
218	v3.4.2) on <i>r2hat</i> value of overlapping SNPs.
219	
220	We compared performance of Eagle2 and SHAPEIT2 phasing tools in tandem with IMPUTE2 as
221	imputation tools across reference panels by: a) Comparing their respective IMPUTE2 R <sup>2</sup> : b)

10 | Page

222	Conducting a Wilcoxon Signed-Rank Test on R <sup>2</sup> value; c) Reporting raw counts of imputed
223	SNPs based on IMPUTE2 "Info" metric and stratified by MAF bins (e.g. common, rare, ultra-
224	rare).
225	In all comparisons, the MAFs are estimated from imputed data according to the reference panel
226	employed. We retained monomorphic SNPs in our analyses for several reasons. A monomorphic
227	SNP in one study might not be monomorphic in other cohorts. This has profound affects, for
228	example, when performing meta-analysis across different studies. In addition, monomorphic
229	SNPs provide information about MAF across studies. Without the information it is difficult to
230	tell, for instance, if a SNP is monomorphic or failed quality control in that study.
231	
232	Agreement between Imputed and Sequence data. To further test the quality of imputation -
233	without relying on software's internal metrics (i.e. "Info" and $R^2$ ) - we calculated genotyped
234	concordance between imputed and WES data using the VCF-compare tool (v0.1.14-12-
235	gcdb80b8) (Danecek et al., 2011). First, we converted posterior probabilities obtained from
236	imputation into genotype data using the PLINK software (v1.90b4.9) by applying a threshold of
237	0.9 (supplementary section 1), such that SNPs that failed on this criterion were left uncalled.
238	For example, an imputed SNP with $P(G=0,1,2)=(0.01,0.9,0.09)$ would be called as a '1'
239	(heterozygous), whereas an imputed SNP with $P(G=0,1,2) = (0.2, 0.6, 0.2)$ would be left
240	uncalled. We restricted the comparison to overlapping SNPs between HRC, 1000G reference
241	panels and whole-exome sequencing (WES) data for Chromosome 14 only, on SNPs with 0%
242	missingness (plinkmissing flag) in WES data. We also assessed variants' agreement according
243	to different MAF bins for "high-quality" ("Info" $\geq 0.8$ ) SNPs. The output resulted in number of
244	variant "mismatches", i.e. the count of allele not matching between imputed and sequenced
	<b>11  </b> P a g e

245	variants per individual. To measure interrater reliability we computed Cohen's kappa coefficient
246	(McHugh, 2012) for both the reference panels against WES data. Kappa coefficient $\leq 0$
247	indicates no agreement, 0.01–0.20 as none to slight, 0.21–0.40 as fair, 0.41–0.60 as moderate,
248	0.61–0.80 as substantial, and 0.81–1.00 as almost perfect agreement.
249	
250	Effects of Ancestry on Imputation Quality. To assess how ancestry affected imputation quality,
251	we conducted a Poisson regression using R. We used percentage of global ancestry (European
252	(CEU), Native (NAT) and African (YRI) as predictors, and total number of mismatches as the
253	outcome; analyses were restricted to "high-quality" SNPs, only.
254	
255	Imputation of G206A Mutation in PSEN1. To evaluate imputation performance of a specific rare
256	variant, we examined a founder mutation, p.Gly206Ala (G206A - rs63750082) in the PSEN1
257	gene (PSEN1-G206A) (Athan et al., 2001; Lee et al., 2015). The PSEN1-G206A mutation is a
258	rare variant observed primarily in Puerto Ricans with familial early onset Alzheimer's disease
259	(EOAD), but it is rare in Puerto Ricans and other populations with late-onset Alzheimer's
260	disease (LOAD) (Arnold et al., 2013). The mutation was present in the 1000G phase 3 reference
261	panel with an allele frequency of 0.001, but was absent in the HRC reference panel. To verify
262	whether individuals who were found to carry the PSEN1-G206A mutation based on 1000G-
263	imputation, they were genotyped using the KASP genotyping technology by LGC genomics
264	(https://www.lgcgroup.com), which uses allele-specific PCR for SNP calling. Agreement
265	between imputed and genotype data for the PSEN1-G206A mutation was then assessed. We also
266	tested the effect on imputation quality based on different IMPUTE2-parameters settings, more
267	specifically by modifying the chunk size (i.e. 1MB vs. 5 MB).

## 268 **Results**

269

270	Comparison of Phasing Tools: Eagle2 vs. SHAPEIT2. To select the optimal tool for phasing, we
271	compared SHAPEIT2 with Eagle2 using Chromosome 21 with 13,066 genotyped SNPs by
272	performing subsequent imputation with IMPUTE2 on phased outputs, and using both reference
273	panels. We found SHAPEIT2 better than Eagle2 when evaluated based on mean $R^2$ and "Info"
274	metric using either the reference panels. For instance, using the 1000G, we observed higher
275	mean R <sup>2</sup> for data phased with SHAPEIT2 as compared to Eagle2 (0.92 vs. 0.91; Wilcoxon p-
276	value < 0.001). Similarly, when HRC panel was employed, mean $R^2$ of 0.89 was observed for
277	SHAPEIT2 against 0.85 for Eagle2 (Wilcoxon signed-rank test p-value < 0.001).
278	SNP count comparison details can be found in Supplementary Table 1-2. Regardless of the
279	reference panel employed, we observed higher percentage of "high-quality" rare and ultra-rare
280	SNPs for SHAPEIT-IMPUTE2 than Eagle2-IMPUTE2. For instance, 1000G-imputation
281	retrieved 51.02% of "high-quality" rare SNPs using SHAPEIT-IMPUTE2 vs. 48.38% with
282	Eagle2-IMPUTE2. Detailed comparisons for different MAF bins and quality threshold can be
283	found in Supplementary Section 2. Nevertheless, we found Eagle2 faster than SHAPEIT2 when
284	computation times were compared; for instance, with HRC Eagle2 was ~6 times faster than
285	SHAPEIT2 (Supplementary Table 3). We therefore imputed the remaining chromosomes on
286	phased output from SHAPEIT2. Comparison of phasing tools by assessing switch error rate was
287	beyond the scope of this paper due to limited resources, for e.g., availability of phased reference
288	panel for an admixed population.
200	

290	MaCH-Admix vs. IMP	PUTE2. W	<i>We</i> found that	SHAPEIT-IMI	PUTE2 pe	erformed bett	er than MaCH-
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- Admix. For Chromosome 21, we imputed 1,104,648 and 646,594 SNPs for SHAPEIT-
- 292 IMPUTE2 and MaCH-Admix, respectively; 549,091 SNPs were overlapping. For SHAPEIT-
- IMPUTE2 we observed 446,591 bi-allelic SNPs with "Info"  $\ge 0.40$ , in contrast with 598,943
- SNPs with  $Rsq \ge 0.30$  from MaCH-Admix (Supplementary Table 4). SNP counts for different
- 295 MAF bins based on platform-specific quality index can be found in **Supplementary Table**
- **5**.When the two outputs were compared in terms of *r2hat*, SHAPEIT-IMPUTE2 showed a
- higheraverage r2hat of 0.62 against 0.36 from MaCH-Admix (Wilcoxon signed-rank test p-value
- 298 < 0.001). Also, MaCH-Admix was 109 times slower than IMPUTE2. (Supplementary Table 6),
- thus, comparison between different panels using MaCH-Admix were excluded due to limited
- 300 resources. For the remaining of this manuscript, we focused on imputation employing SHAPEIT-
- 301 IMPUTE2, only.
- 302
- 303 Comparison between HRC and 1000G using SHAPEIT-IMPUTE2. Using SHAPEIT-IMPUTE2,

we imputed 81,240,392 and 38,532,090 SNPs across all autosomal chromosomes with 1000G

- and HRC reference panels, respectively (**Table 2**).
- Overall, we observed slightly higher mean  $R^2$  with 1000G than with HRC panel (0.94 vs. 0.92;
- 307 Wilcoxon p-value< 0.001). Nevertheless, when the analyses were restricted to only "good-" and
- <sup>308</sup> "high-quality" SNPs, HRC consistently performed better: 60.82% of HRC-imputed SNPs were
- 309 "good-quality" and 48.87% were "high-quality" (Wilcoxon signed-rank test p-value < 0.001). On
- the contrary, 40.32% of 1000G imputed SNPs were "good-quality" and 30.11% were "high-
- 311 quality".

312 Further, we evaluated performance for uncommon, rare and ultra-rare SNPs. For "good-" and "high-quality" SNPs, HRC outperformed 1000G. For example, HRC panel produced 62.85% of 313 "high-quality" rare SNPs, whereas 1000G had 53.83% (Table 3). When average imputation 314 315 "Info" quality was evaluated, HRC-imputation again performed better than with 1000G 316 (Wilcoxon p-value< 0.001) (Figure 1). 317 Next, we restricted our analyses to *overlapping* SNPs across the two reference panels only, based 318 on their chromosome and position mapping, reference and non-reference alleles. For "good-"and 319 320 "high-quality" SNPs, imputation in both panels performed similarly (**Table 2**). When restricted to uncommon, rare and ultra-rare SNPs, we observed higher percentage of "good-" and "high-321 322 quality" SNPs for HRC panel as compared to 1000G reference panel (**Table 3**). For example, 323 7.44% of HRC-imputed ultra-rare SNPs were "good-quality" vs. 4.95% with the 1000G. 1.69% of HRC-imputed ultra-rare SNPs were "high-quality" vs. 0.75% with the 1000G. Further, 324 Wilcoxon test on "Info" value of "high-quality" ultra-rare SNPs (2,972) again showed better 325 performances when HRC was employed vs. 1000G (P-value < 0.001). Complete list of counts 326 and percentages across reference panels, MAF bins and quality score can be found in **Table 3**. 327 328 The case of G206A and the effect of chromosomal chunk size on imputation quality. SNP 329 rs63750082 is absent from HRC panel therefore no imputation was achieved. Using 1000G 330 331 reference panel, 12 individuals were imputed as G206A carriers. SNP rs63750082 was imputed with an IMPUTE2 "Info" score of 0.48 using 1MB as chromosomal region parameter. When we 332 increased the chunk size to 5MB, IMPUTE-Info score drastically improved to 0.94 (Figure 2). 333 334 Those patients labeled as mutation-carriers according to imputation were then genotyped: all 12

were confirmed to be G206A carriers, therefore achieving a perfect imputation prediction (100%agreement) for that specific SNP.

337

Genotype Concordance and Kappa Coefficient. Out of the 1,000 individuals included in our 338 study, 262 had whole exome sequencing (WES) data available (Raghavan et al., 2018). We had 339 340 14,157 overlapping SNPs in WES, HRC and 1000G reference panels with 0% missingness in WES data on Chromosome 14; SNPs imputed with each reference panel were compared against 341 342 WES data separately. When concordance was evaluated, HRC panel performed slightly poorer, despite showing higher number of "high-quality" variants as compared to 1000G (Table 4). 343 Using 1000G, we observed 3,542 rare and 35 ultra-rare "high-quality" SNPs; across 262 344 345 samples, we counted 1,245 ((1,245/(3,542\*262))\*100=0.13%) and 10 (0.10%) mismatches for rare and ultra-rare respectively. Using HRC, we retrieved 3,759 rare and 93 ultra-rare "high-346 quality" variants; we observed 2,439 (0.24%) and 32 (0.13%) mismatches for rare and ultra-rare 347 variants, respectively. Details about pipeline can be found in **supplementary section 3**. 348 349 Next, we computed Cohen's kappa coefficient (K) for 14,157 imputed SNPs common in WES and the two reference panels. For both HRC and 1000G-imputation, we observed Kappa (K) of 350 ~0.99 for both rare and ultra-rare "high-quality" variants(Table 4). Details about pipeline can be 351 352 found in supplementary section 4.

353

Effects of Ancestry on Imputation Quality. We evaluated the effect of individual ancestral
 component separately on SNP mismatches for Chromosome 14 on 262 individuals. For both
 reference panels we found that higher African ancestry (YRI) was associated with higher number
 of mismatches (Supplementary table 7). For instance, with 1000G reference panel, for rare
 variants ("Info" ≥0.80), we observed an estimate of 1.46 (P-value<0.001) for YRI component</li>

- 359 (indicating that for each unit increase in YRI ancestry, it results in 1.46 additional mismatches).
- 360 Details on confidence intervals and robust standard errors can be found in **supplementary file**
- 361 (Table 7 and Section 5). We did not observe significant effect of ancestry on "high-quality"
- 362 ultra-rare variants in both panels.

## 364 **Discussion**

This study examined imputation performances in a cohort Caribbean Hispanics, focusing on 365 uncommon, rare and ultra-rare variant, by comparing different phasing and imputation tools, as 366 367 well as evaluating the effects of different reference panels. Overall, uncommon and rare variants can be well imputed in this population, characterized by a unique genetic background. Caribbean 368 369 Hispanics are admixed with 59% of their genetic component from European, 32% African, and 8% Native American ancestry (Tosto et al., 2015). Due to their genetic makeup and unique 370 linkage disequilibrium patterns, admixed populations offer unique opportunity in studying 371 372 complex diseases. First, disease prevalence varies across ethnic groups (Igartua et al., 2015) and certain admixed populations show higher incidence rates and prevalence (e.g. Alzheimer's 373 374 disease, diabetes etc.) or lower ones (e.g. multiple sclerosis). Second, variants that are ethnic-375 specific may explain a higher prevalence of the disease of interest in admixed groups. 376 In the present study, we examined multiple parameters of imputation using the Caribbean 377 Hispanics population. First, we found that imputation using SHAPEIT-IMPUTE2 phasing 378 generated better results than Eagle2-IMPUTE2, and SHAPEIT-IMPUTE2 is superior to MaCH-379 380 Admix in terms of imputation performances and process time. Using SHAPEIT-IMPUTE2, 1000G SNPs outnumbered HRC panel because of the higher 381 number of SNPs included in the reference panel itself. However, when we restricted our analyses 382 383 to overlapping "good-" and "high-quality" SNPs (i.e. those variants that most likely would be included in association analyses), HRC-imputation outperformed 1000G with higher. The 384 385 superior performance of HRC over 1000G was confirmed also when we focused on uncommon,

rare and ultra-rare SNPs only. Our findings confirm data in literature, i.e. reference panels with

387 higher number haplotypes perform better in different scenarios.

388 Additional investigations are needed in order to apply our findings to other admixed and non-

389 admixed populations.

African ancestry.

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Overall, higher quality of imputation for rare and ultra-rare variants was also confirmed when we
tested results against sequencing data. Finally, higher YRI global ancestry was found to
significantly impair SNP imputation, suggesting that imputation quality decreases with increased

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Lastly, SHAPEIT-IMPUTE2 with 1000G reference panel was successful in identifying G206A 396 397 mutation carriers. We also noticed that imputation quality drastically improved when imputation was conducted using large (5MB) chunk size as compared to small (1MB) chunks. This seems to 398 contradict previous observation: Zhang et al. (Zhang et al., 2011) studied the effect of window 399 size on imputation in an African-Americans. They concluded that window size of 1MB could be 400 considered acceptable. Possible explanations for these different results might be the more 401 402 complex admixture of CH compare to AA (three-way vs. two-way admixed) and a more complex LD pattern for the G206A region. Ultimately, we recommend to consider a wider 403 404 window size to achieve high-quality imputation in specific variants that fail under default 405 settings.

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This work has limitations. First, we could carry out the comparison between the two referencepanels restricting the analyses to overlapping variants only, limiting our observation to a subset

- 409 of the variants included in the 1000G panel. This is a result of the HRC composition, which is
- 410 composed by several studies and ended up including only a consensus number of variants.
- 411 Second, we tested the agreement between imputed and sequenced variants in a smaller subset of
- 412 individuals that had both GWAS and WES data available.

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- 420 Conflict of Interest Statement
- 421 The authors declare no conflict of interests.

Alexander, D.H., Novembre, J., and Lange, K. (2009). Fast model-based estimation of ancestry in unrelated individuals. *Genome Res* 19(9), 1655-1664. doi: 10.1101/gr.094052.109.

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## 423 **References.**

- Arnold, S.E., Vega, I.E., Karlawish, J.H., Wolk, D.A., Nunez, J., Negron, M., et al. (2013). 426 427 Frequency and clinicopathological characteristics of presenilin 1 Gly206Ala mutation in Puerto Rican Hispanics with dementia. J Alzheimers Dis 33(4), 1089-1095. doi: 428 10.3233/JAD-2012-121570. 429 Athan, E.S., Williamson, J., Ciappa, A., Santana, V., Romas, S.N., Lee, J.H., et al. (2001). A 430 431 founder mutation in presenilin 1 causing early-onset Alzheimer disease in unrelated Caribbean Hispanic families. JAMA 286(18), 2257-2263. 432 433 Browning, B.L., and Browning, S.R. (2009). A unified approach to genotype imputation and 434 haplotype-phase inference for large data sets of trios and unrelated individuals. Am J435 Hum Genet 84(2), 210-223. doi: 10.1016/j.ajhg.2009.01.005. Danecek, P., Auton, A., Abecasis, G., Albers, C.A., Banks, E., DePristo, M.A., et al. (2011). The 436 437 variant call format and VCFtools. Bioinformatics 27(15), 2156-2158. doi: 10.1093/bioinformatics/btr330. 438 Das, S., Forer, L., Schonherr, S., Sidore, C., Locke, A.E., Kwong, A., et al. (2016). Next-439 440 generation genotype imputation service and methods. Nat Genet 48(10), 1284-1287. doi: 10.1038/ng.3656. 441 Delaneau, O., Zagury, J.F., and Marchini, J. (2013). Improved whole-chromosome phasing for 442 443 disease and population genetic studies. Nat Methods 10(1), 5-6. doi: 10.1038/nmeth.2307. 444 Genomes Project, C., Auton, A., Brooks, L.D., Durbin, R.M., Garrison, E.P., Kang, H.M., et al. (2015). A global reference for human genetic variation. Nature 526(7571), 68-74. doi: 445 446 10.1038/nature15393. 447 Gibson, G. (2012). Rare and common variants: twenty arguments. Nat Rev Genet 13(2), 135-448 145. doi: 10.1038/nrg3118. Ha, N.T., Frevtag, S., and Bickeboeller, H. (2014). Coverage and efficiency in current SNP 449 chips. Eur J Hum Genet 22(9), 1124-1130. doi: 10.1038/ejhg.2013.304. 450 Hancock, D.B., Levy, J.L., Gaddis, N.C., Bierut, L.J., Saccone, N.L., Page, G.P., et al. (2012). 451 Assessment of genotype imputation performance using 1000 Genomes in African 452 453 American studies. *PLoS One* 7(11), e50610. doi: 10.1371/journal.pone.0050610. 454 Herzig, A.F., Nutile, T., Babron, M.C., Ciullo, M., Bellenguez, C., and Leutenegger, A.L. (2018). Strategies for phasing and imputation in a population isolate. Genet Epidemiol 455 456 42(2), 201-213. doi: 10.1002/gepi.22109. 457 Howie, B., Fuchsberger, C., Stephens, M., Marchini, J., and Abecasis, G.R. (2012). Fast and accurate genotype imputation in genome-wide association studies through pre-phasing. 458 459 Nat Genet 44(8), 955-959. doi: 10.1038/ng.2354. Howie, B.N., Donnelly, P., and Marchini, J. (2009). A flexible and accurate genotype imputation 460 461 method for the next generation of genome-wide association studies. PLoS Genet 5(6),
- 462 e1000529. doi: 10.1371/journal.pgen.1000529.
  463 Huang, J., Howie, B., McCarthy, S., Memari, Y., Walter, K., Min, J.L., et al. (2015). Improved
  464 imputation of low-frequency and rare variants using the UK10K haplotype reference
  465 panel. *Nat Commun* 6, 8111. doi: 10.1038/ncomms9111.
  - 22 | Page

- Igartua, C., Myers, R.A., Mathias, R.A., Pino-Yanes, M., Eng, C., Graves, P.E., et al. (2015).
  Ethnic-specific associations of rare and low-frequency DNA sequence variants with
  asthma. *Nat Commun* 6, 5965. doi: 10.1038/ncomms6965.
- Lee, J.H., Cheng, R., Vardarajan, B., Lantigua, R., Reyes-Dumeyer, D., Ortmann, W., et al.
  (2015). Genetic Modifiers of Age at Onset in Carriers of the G206A Mutation in PSEN1
  With Familial Alzheimer Disease Among Caribbean Hispanics. *JAMA Neurol* 72(9),
  1043-1051. doi: 10.1001/jamaneurol.2015.1424.
- Li, J.Z., Absher, D.M., Tang, H., Southwick, A.M., Casto, A.M., Ramachandran, S., et al.
  (2008). Worldwide human relationships inferred from genome-wide patterns of variation. *Science* 319(5866), 1100-1104. doi: 10.1126/science.1153717.
- Liu, E.Y., Li, M., Wang, W., and Li, Y. (2013). MaCH-admix: genotype imputation for admixed
  populations. *Genet Epidemiol* 37(1), 25-37. doi: 10.1002/gepi.21690.
- Liu, Q., Cirulli, E.T., Han, Y., Yao, S., Liu, S., and Zhu, Q. (2015). Systematic assessment of
  imputation performance using the 1000 Genomes reference panels. *Brief Bioinform*16(4), 549-562. doi: 10.1093/bib/bbu035.
- Loh, P.R., Danecek, P., Palamara, P.F., Fuchsberger, C., Y, A.R., H, K.F., et al. (2016).
  Reference-based phasing using the Haplotype Reference Consortium panel. *Nat Genet* 48(11), 1443-1448. doi: 10.1038/ng.3679.
- Marchini, J., and Howie, B. (2010). Genotype imputation for genome-wide association studies.
   *Nat Rev Genet* 11(7), 499-511. doi: 10.1038/nrg2796.
- McHugh, M.L. (2012). Interrater reliability: the kappa statistic. *Biochem Med (Zagreb)* 22(3),
   276-282.
- Nagy, R., Boutin, T.S., Marten, J., Huffman, J.E., Kerr, S.M., Campbell, A., et al. (2017).
  Exploration of haplotype research consortium imputation for genome-wide association studies in 20,032 Generation Scotland participants. *Genome Med* 9(1), 23. doi: 10.1186/s13073-017-0414-4.
- 492 Nelson, S.C., Stilp, A.M., Papanicolaou, G.J., Taylor, K.D., Rotter, J.I., Thornton, T.A., et al.
  493 (2016). Improved imputation accuracy in Hispanic/Latino populations with larger and
  494 more diverse reference panels: applications in the Hispanic Community Health
  495 Study/Study of Latinos (HCHS/SOL). *Hum Mol Genet* 25(15), 3245-3254. doi:
  496 10.1093/hmg/ddw174.
- 497 Pei, Y.F., Zhang, L., Li, J., and Deng, H.W. (2010). Analyses and comparison of imputation498 based association methods. *PLoS One* 5(5), e10827. doi: 10.1371/journal.pone.0010827.
- Purcell, S., Neale, B., Todd-Brown, K., Thomas, L., Ferreira, M.A., Bender, D., et al. (2007).
  PLINK: a tool set for whole-genome association and population-based linkage analyses. *Am J Hum Genet* 81(3), 559-575. doi: 10.1086/519795.
- Raghavan, N.S., Brickman, A.M., Andrews, H., Manly, J.J., Schupf, N., Lantigua, R., et al.
  (2018). Whole-exome sequencing in 20,197 persons for rare variants in Alzheimer's disease. *Ann Clin Transl Neurol* 5(7), 832-842. doi: 10.1002/acn3.582.
- Roshyara, N.R., Kirsten, H., Horn, K., Ahnert, P., and Scholz, M. (2014). Impact of preimputation SNP-filtering on genotype imputation results. *BMC Genet* 15, 88. doi:
  10.1186/s12863-014-0088-5.
- Surakka, I., Sarin, A.-P., Ruotsalainen, S.E., Durbin, R., Salomaa, V., Daly, M., et al. (2016).
  The rate of false polymorphisms introduced when imputing genotypes from global imputation panels. *bioRxiv*. doi: 10.1101/080770.

- Tosto, G., Fu, H., Vardarajan, B.N., Lee, J.H., Cheng, R., Reyes-Dumeyer, D., et al. (2015). F box/LRR-repeat protein 7 is genetically associated with Alzheimer's disease. *Ann Clin Transl Neurol* 2(8), 810-820. doi: 10.1002/acn3.223.
- Verma, S.S., de Andrade, M., Tromp, G., Kuivaniemi, H., Pugh, E., Namjou-Khales, B., et al.
  (2014). Imputation and quality control steps for combining multiple genome-wide
  datasets. *Front Genet* 5, 370. doi: 10.3389/fgene.2014.00370.
- 517 Zhang, B., Zhi, D., Zhang, K., Gao, G., Limdi, N.N., and Liu, N. (2011). Practical Consideration
  518 of Genotype Imputation: Sample Size, Window Size, Reference Choice, and Untyped
  519 Rate. *Stat Interface* 4(3), 339-352.
- Zheng, H.F., Ladouceur, M., Greenwood, C.M., and Richards, J.B. (2012). Effect of genome wide genotyping and reference panels on rare variants imputation. *J Genet Genomics* 39(10), 545-550. doi: 10.1016/j.jgg.2012.07.002.
- Zheng, H.F., Rong, J.J., Liu, M., Han, F., Zhang, X.W., Richards, J.B., et al. (2015).
  Performance of genotype imputation for low frequency and rare variants from the 1000 genomes. *PLoS One* 10(1), e0116487. doi: 10.1371/journal.pone.0116487.
- Zhou, H., Alexander, D., and Lange, K. (2011). A quasi-Newton acceleration for highdimensional optimization algorithms. *Stat Comput* 21(2), 261-273. doi: 10.1007/s11222009-9166-3.

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## **Table 1**: SNP counts in HRC and 1000G reference panel.

Reference Panel	Individuals	Autosomal variants	<b>Bi-allelic SNPs</b>	Multi-allelic SNPs
1000G Phase 3	2,504	81,706,022	77,818,332	3,887,690
HRC	27,165*	39,131,600	39,131,600	NA

\*For Chromosome 1, the number of individuals were 22,691

Reference Panel	Mulfi-allelic SNPs				i-allelic SNI	Ps	Total SNPs				
	Total SNPsInfo* $\geq 0.40 \ (\%)$ Info $\geq 0.80$ $(\%)$		Total SNPs	Info ≥0.40 (%)	Info≥0.80 (%)	Total SNPs	Info ≥0.40 (%)	Info ≥0.80 (%)			
All SNPs											
1000G	3,319,815	2,586,342 (77.90)	2,061,295 (62.09)	77,920,577	31,423,926 (40.32)	23,468,086 (30.11)	81,240,392	31,423,926 (41.86)	25,529,381 (31.42)		
HRC	NA	NA	NA	38,532,090	32,090 23,436,980 18,833 (60.82) (48.8		38,532,090	23,436,980 (60.82)	18,833,790 (48.79)		
			SNPs o	overlapping	HRC & 10	00G					
1000G	NA	NA	NA	30,090,251	22,631,112 (75.21)	18,408,585 (61.17)	30,090,251	22,631,112 (75.21)	18,408,585 (61.17)		
HRC	NA	NA	NA	30,090,251	22,438,268 (74.56)	18,395,036 (61.13)	30,090,251	22,438,268 (74.56)	18,395,036 (61.13)		

# **Table 2**: Type of imputed SNPs across reference panels.

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**Table 3**: SNP Counts for all Bi-allelic uncommon, rare and ultra-rare SNPs.

MAF		1000G		HRC						
MAT	Info ≥0	Info ≥0.40 (%)	Info ≥0.80 (%)	Info ≥0	Info ≥0.40 (%)	Info ≥0.80 (%)				
All SNPs										
[1% - 5%]	6,025,281	5,989,223 (98.90)	5,441,982 (90.31)	5,434,996	5,421,257 (99.84)	5,061,904 (93.13)				
[0.1% - 1%)	20,249,058	16,881,286 (83.36)	10,901,789 (53.83)	11,780,671	10,931,924 (92.79)	7,404,808 (62.85)				
[0-0.1%)	44,562,2051,490,434 (3.34)242,717 (0.544)15,055,433		828,256 (5.50)	174,673 (1.16)						
		SNPs ov	erlapping HRC	& 1000G						
[1% - 5%]	5,624,956	5,604,308 (99.63)	5,148,285 (91.52)	5,396,207	5,385,364 (99.79)	5,037,187 (93.34)				
[0.1% - 1%)	11,875,603	10,442,603 (87.93)	7,027,312 (59.17)	10,945,899	10,268,136 (93.80)	7,060,908 (64.50)				
[0-0.1%)	0-0.1%) 6,314,479 312,967 (4.95)		47,614 (0.75)	7,519,807	560,043 (7.44)	127,423 (1.69)				

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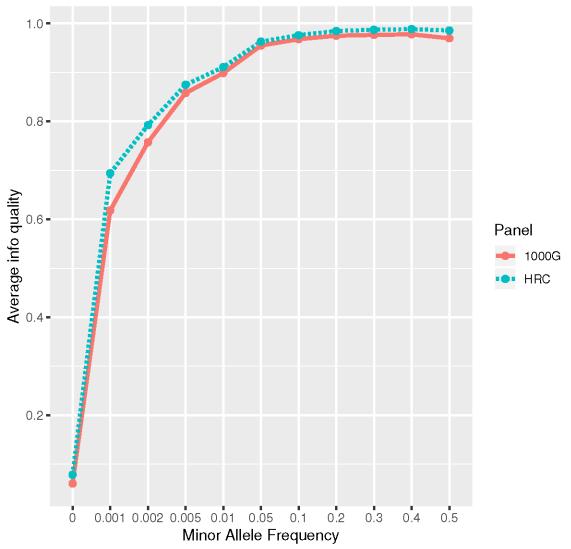
# 542 Table 4: Comparison for mismatch counts and Kappa (*K*) for HRC and 1000G using WES data543 on Chromosome 14.

MAF		10000	r J		HRC				
		Info ≥0.	80		Info ≥0.80				
	SNP	Total SNPs in	Mismatch	Kap	SNP	Total SNPs in	Mismatch	Kappa	
		all persons <sup>*</sup> pa		all persons <sup>*</sup>			(K)		
		_		(K)					
	2,354	610,550	7,397	0.99	2,264	587,961	8,963	0.99	
[1% - 5%]			(1.22%)				(1.52%)		
	3,542	926,109	1,245	0.99	3,759	982,734	2,439	0.99	
[0.1% - 1%)			(0.13%)				(0.24%)		
	35	9,163	10	0.99	93	24,348	32	0.99	
[0-0.1%)			(0.10%)				(0.13%)		

<sup>544</sup> \*Less value than 262\*SNP because imputed with poor posterior probability failed to be

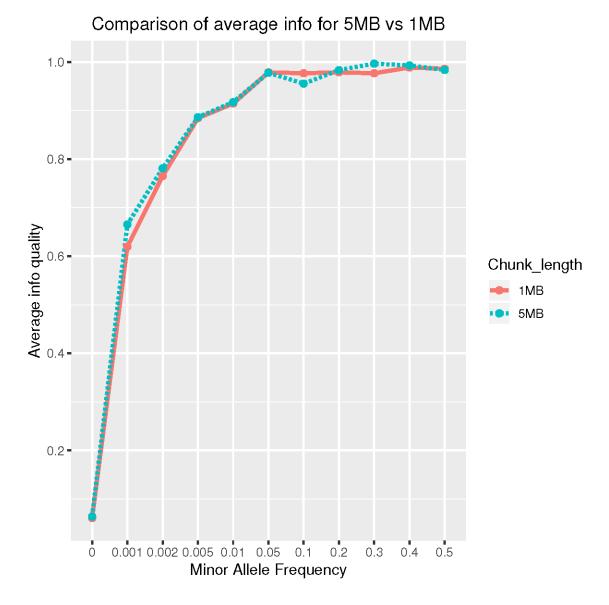
545 converted from .gen to plink format.

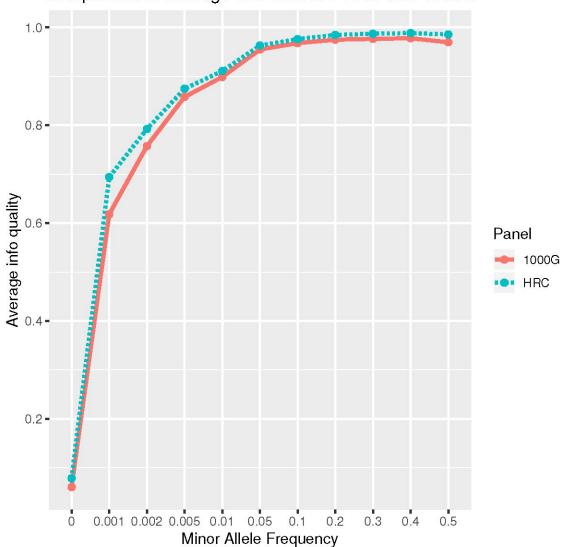
Figure 1: Comparison of average Info quality between HRC and 1000G reference panel for all autosomalchromosomes



Comparison of average info between HRC and 1000G







Comparison of average info between HRC and 1000G

