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3	Caterpillars on a phytochemical landscape
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28 Summary

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30	• Modern metabolomic approaches that generate more comprehensive phytochemical profiles
31	than were previously available are providing new opportunities for understanding plant-animal
32	interactions. Specifically, we can characterize the phytochemical landscape by asking how many
33	individual compounds affect herbivores and how compounds covary among plants.
34	• Here we use the recent colonization of alfalfa (Medicago sativa) by the Melissa blue butterfly
35	(Lycaeides melissa) to quantify primary and secondary plant metabolites and the performance of
36	caterpillars as affected by both individual compounds and suites of covarying phytochemicals.
37	• We find that survival, development time and adult weight are all associated with variation in a
38	large number of compounds, including biomolecules associated with plant cell function as well
39	as putative anti-herbivore action. The dimensionality of the plant-insect interface is high, with
40	clusters of covarying compounds in many cases encompassing divergent effects on different
41	aspects of caterpillar performance.
42	• Individual compounds with the strongest associations tend to be secondary metabolites,
43	including alkaloids, phenolic glycosides and saponins. The saponins are represented in our data
44	by more than 25 individual compounds with beneficial and detrimental effects on caterpillars,
45	which highlights the value of metabolomic data as opposed to approaches that rely on total
46	concentrations within defensive classes.
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48	Keywords: Lycaeides melissa, Medicago sativa, metabolomics, plant defense, specialization
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51 Introduction

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One of the conceptual pillars of trophic ecology is the idea that herbivores must overcome the 53 barrier of plant secondary chemistry before extracting the nutrients necessary for growth and 54 reproduction (Feeny et al., 1992). The success of this idea is reflected in areas of research that 55 include coevolution (Agrawal et al., 2012), ecological specialization (Dyer, 1995), and nutrient 56 flow in ecosystems (Olson, 1963). In most cases, progress has been made by chemical ecologists 57 focusing on small subsets of the secondary metabolites produced by plants and consumed by 58 herbivores. The focus on a few charismatic molecules or classes of compounds, such as 59 furanocoumarins (Berenbaum, 1983) or cardiac glycosides (Zalucki et al., 2001), was at least in 60 part necessitated by early methods in natural products chemistry that were targeted and not easily 61 optimized for the discovery of large suites of co-occurring primary and secondary metabolites 62 (Maag et al., 2015; Dyer et al., 2018). As technological limitations have dissipated, the 63 opportunity now exists for a deeper understanding of the challenges faced by herbivores, with 64 the possibility of discovering, among other things, novel compounds and synergistic interactions 65 among compounds (Prince & Pohnert, 2010; Richards et al., 2010; Sardans et al., 2011). More 66 generally, an important task is to quantify the phytochemical dimensionality of the antagonistic 67 interaction between plants and herbivores, with an eye towards understanding constraints on the 68 evolution of both players (Fordyce & Nice, 2008; Macel et al., 2010) and predicting the 69 formation of new plant-herbivore interactions (Erbilgin, 2018). Here we use the example of a 70 specialized herbivore and a novel host plant to investigate the phytochemical landscape from the 71 perspective of developing caterpillars. 72

The Melissa blue butterfly, Lycaeides melissa, specializes on larval host plants in the pea 73 family (Fabaceae), primarily in the genera Astragalus and Lupinus. Within the last 200 years, L. 74 melissa has colonized introduced alfalfa, Medicago sativa (Fabaceae), at least twice and 75 probably multiple times (Chaturvedi et al., 2018), forming a heterogeneous patchwork of 76 association throughout the range of the butterfly in western North America. M. sativa is a 77 suboptimal host, relative to native hosts that have been examined, and populations of the 78 butterfly that persist on *M. sativa* show evidence of loss of preference for native hosts (Forister *et* 79 al., 2012), reduced caterpillar performance on native hosts, and a slight increase in ability to 80 develop on the suboptimal novel host (Gompert et al., 2015). The genetic architecture of host 81

use in this system is known to be polygenic and characterized by loci with conditionally neutral 82 (host-specific) effects (Gompert et al., 2015). What is needed next is an understanding of how 83 many independent or covarying phytochemical compounds have consequential effects on 84 caterpillars eating the novel host. For example, will the trajectory of further local adaptation by 85 L. melissa to M. sativa be a matter of evolving the ability to detoxify one or a large number of 86 compounds? We would also like to know how key compounds covary among individual plants, 87 which should help us understand the puzzle faced by ovipositing females encountering an array 88 of co-occurring *M. sativa* chemotypes. Here we use a common garden approach (to minimize 89 non-genetic phenotypic variation among plants) and caterpillars individually reared in a 90 controlled environment to address these questions while describing the effects of metabolomic 91 variation in M. sativa on L. melissa. 92

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94 Materials and Methods

95 Plants and caterpillars

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Plants used in this project were grown at the University of Nevada, Reno, Main Station 97 experimental farm. The common garden was planted in 2016 with seeds collected the previous 98 year from 45 plants (previously studied by Harrison et al. (Harrison *et al.*, 2018)) growing in a 99 fallow field in north-western Nevada on the western edge of the Great Basin Desert. The focal 100 butterfly, L. melissa, was present in the source field but has not yet colonized the university farm 101 where experimental plants were grown. The 45 maternal plants each contributed 15 offspring to 102 a randomized grid design in the common garden, irrigated with broadcast sprayers in 2016 and 103 drip in 2017, without supplemental fertilization. Out of each maternal family, a single plant was 104 randomly selected for use in the rearing experiment reported here as a way to capture as much 105 genetic and phenotypic variation as possible. 106

On 17 and 18 July 2017, a total of 45 *L. melissa* females were collected from an alfalfaassociated population near Verdi, NV, and confined to oviposition arenas (500 mL plastic cups) in groups of three with host plant leaves and supplied with Gatorade on mesh lids. After three days, eggs were removed from leaves, pooled, and kept at room temperature until hatching, at which time caterpillars were placed individually in Petri dishes (100 x 25 mm) with leaves of a particular *M. sativa* individual (which became the only plant from which they were fed

throughout the experiment). Ten caterpillars were assigned to each of the 45 experimental *M*.

sativa plants (for a total of 450 independently-reared caterpillars) and kept in a growth chamber

set to 25 C and a 12 hour light / 12 hour dark cycle. From each caterpillar we recorded survival

to adult, date of eclosion (if successful) and adult weight to the nearest 0.01 mg on a Mettler

117 Toledo XP26 microbalance.

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119 Phytochemistry and plant traits

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Metabolomic variation among individual plants was characterized with liquid chromatography-121 mass spectrometry (LC-MS, (Jorge et al., 2016)) using leaves collected on a single day at the 122 start of the rearing experiment. Dried, ground leaves (10 mg) were extracted in 2 mL of 70% 123 aqueous ethanol and injected into an Agilent 1200 analytical high performance liquid 124 chromatograph paired with an Agilent 6230 Time-of-Flight mass spectrometer via an 125 electrospray ionization source. Resulting chromatograms were analyzed using MassHunter 126 Quantitative Analysis (v.B.06.00, Agilent, Santa Clara, CA); see electronic supporting 127 information for additional phytochemical protocols. Major classes of compounds were identified 128 using characteristic relative mass defects (Ekanayaka et al., 2015). Leaf protein content was 129 quantified with three replicates (~2 mg each) per plant using the Bicinchoninic acid assay (Pierce 130 Biotechnology, Waltham, MA). Before grinding, five dried leaflets from each sample were 131 weighed to the nearest 0.1 mg, scanned, and area was measured using ImageJ (v.1.52a); specific 132 leaf area (SLA) was calculated as leaf area divided by dry mass. Finally, leaf toughness was 133 measured on fresh material in the common garden, at the start of the experiment (mid-July, when 134 leaves were also sampled for chemistry and protein) and at the end of the experiment (mid-135 August), from three leaves per plant at each date, with a penetrometer (Chatillon 516 Series) 136 through the center of the middle leaflet, as in (Harrison et al., 2018); the three leaves were 137 selected haphazardly, avoiding the oldest and youngest leaves. The six leaf toughness 138 measurements per plant were averaged for a single toughness measure used in analyses. 139 140

141 Analyses of plant traits and caterpillar performance

142 Overview

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Our analytical strategy to understand the association between phytochemical variation and 144 caterpillar performance followed two complementary paths, one focusing on dimension 145 reduction and feature selection to produce relatively simple models, and the other on the 146 estimation of effects of all individual compounds. For the first path, involving dimension 147 reduction and feature selection, we utilized an approach developed for gene transcription studies 148 that identifies groups or modules of correlated variables with hierarchical clustering (Langfelder 149 & Horvath, 2008); after clustering, we reduced the number of independent variables by selecting 150 among modules and other plant traits (specific leaf area, protein and leaf toughness) using lasso 151 regression (Ogutu et al., 2012). Selected modules (and other plant traits) were analyzed in 152 Bayesian linear models that are useful in this context because they allowed us to quantify our 153 confidence in the sign of effects (positive or negative) as continuous probabilities (as opposed to 154 relying on arbitrary significance cutoffs). For the second analytical path, we utilized ridge 155 regression (Ogutu *et al.*, 2012) to estimate effects for all compounds simultaneously, which 156 allowed us to investigate the distribution of effects among compounds and classes of compounds. 157 Both analytical paths incorporated cross-validation during the lasso and ridge regressions, and as 158 a means of evaluating the predictive success of the Bayesian models. We also used 159 randomization tests to compare the performance of modules and individual compounds with 160 randomly-chosen suites of compounds. 161

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163 Dimension reduction and feature selection

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We chose an approach that reduces the number of independent variables while allowing us to 165 learn something about the correlational structure of the data, specifically unsupervised 166 hierarchical clustering as implemented in the blockwiseModules function of the WGCNA 167 package (Langfelder & Horvath, 2008) in R (Team & R Development Core Team, 2016). 168 Among the options in the pipeline, we used positive correlations among variables ("signed" 169 network type), merge cut height at 0.25, and correlations raised to the power of five (which is 170 where the scale free topology index reached a plateau). Through experimentation, we found that 171 our results with LC-MS data were robust to variation in these choices, including the choice of 172 signed or unsigned networks. After an initial round of clustering, we took a remaining 19 173 unassigned compounds and put them through a second round of clustering (although the majority 174

of consequential compounds were identified in the first round). One output of the WGCNA 175 procedure is the first eigenvector from each cluster of compounds, which reduced our number of 176 predictor variables by a factor of ten. The resulting eigenvectors plus protein, SLA (specific leaf 177 area) and leaf toughness were then put through the feature reduction step of lasso regression 178 (Ogutu et al., 2012), a penalized regression that allows beta coefficients to be constrained to zero 179 (thus excluding variables). We used the cv.glmnet function of the glmnet package (Friedman et 180 al., 2016) with cross-validation during error reduction set to leave out one plant (and associated 181 caterpillars) at each iteration. The variables selected by the lasso were then put into a Bayesian 182 linear model to estimate coefficients and associated credible intervals using JAGS (version 3.2.0) 183 run in R with the rjags package (Plummer & others, 2003). Two Markov chains were run for 184 10,000 steps for each analysis (no burn in was required) and chain performance was assessed by 185 plotting chain histories, and calculating the Gelman and Rubin convergence diagnostic and 186 effective sample sizes (Gelman et al., 1992; Brooks & Gelman, 1998). For all models, 187 uninformative priors for the regression coefficients were modeled as a normal distribution with a 188 mean of zero and variance of 100 (variance = 1/precision). We quantified our confidence in the 189 sign of coefficients (positive or negative) as the fraction of the posterior samples that were less 190 than zero (for coefficients with a median negative value) or greater than zero (for coefficients 191 with a median positive value). 192

All analyses were done using the R statistical language (Team & R Development Core 193 Team, 2016) on scaled (z-transformed) predictor variables, and both lasso and Bayesian models 194 used binomial (for survival), Poisson (for development time) and Gaussian (for adult weight) 195 errors. The latter two analyses (development time and adult weight) included sex as a factor. 196 The analysis of development time also included adult weight as a covariate; while (reciprocally) 197 the analysis of adult weight included development time as a predictor. These variables are 198 negatively correlated (at -0.52), and they function as useful covariates of each other, allowing us 199 to investigate the possibility of unique plant effects on weight gain and development time, which 200 could not be discovered if, for example, these variables were combined into a single performance 201 index. 202

The success of models developed with the dimension reduction and feature selection pipeline was judged in two ways. We used a cross-validation procedure in which we left out one plant (and associated caterpillars) in each iteration of the Bayesian model and then used the estimated coefficients (for phytochemical variables and other plant traits) to predict the
performance of the unobserved caterpillars. After 45 iterations (one for each plant), we
calculated a simple correlation coefficient between the observed and predicted performance of
caterpillars across plants. In addition, we repeatedly resampled the original LC-MS data to
match the structure of the reduced set of predictor variables to ask to what extent randomly
assembled modules could outperform the empirically-derived modules.

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213 Individual compound effects

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The second path of our two-part analytical strategy involved simultaneous estimation of the 215 effects of all individual chemical compounds on caterpillar survival, development time and adult 216 weight. For this approach, we again used penalized regression (in the glmnet package (Friedman 217 et al., 2016)), but this time with ridge regression (instead of lasso) which constrains beta 218 coefficients to avoid variance inflation but does not eliminate variables. As with the analyses 219 above, ridge regression was done using error structures appropriate to the specific response 220 variables, and included additional covariates where possible (in models of development time and 221 adult weight). The resulting coefficients associated with all individual compounds were 222 examined as a second perspective on the modules examined in the first set of analyses, and were 223 used to ask to what extent individual compound effects could be predicted by the degree to 224 which they vary among individual plants as quantified with the simple coefficient of variation. 225 To assess confidence in the results of ridge regressions, we used a bootstrap approach, repeatedly 226 resampling the data and estimating coefficients 1000 times, noting the compounds whose 227 bootstrap confidence intervals did or did not overlap zero (Delaney & Chatterjee, 1986). We 228 also allowed for the discovery of interactions among compounds using penalized regression on 229 all individual compounds and all pairwise interactions between compounds. For ease of 230 interpretation, this final analysis of potential interactions used lasso (not ridge) regression, letting 231 the coefficients for many of the individual compounds and pairwise interactions go to zero. 232

233

234 **Results**

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Of the 450 caterpillars that started the experiment, 261 were reared to eclosion as adults (a 236 mortality rate similar to previous work with this system (Gompert *et al.*, 2015)) on leaves from 237 45 individual alfalfa plants that were characterized for protein, leaf toughness, specific leaf area 238 and 163 individual metabolomic features (see Fig. 1 for variation among plants in caterpillar 239 performance and a subset of plant traits, and Supporting Information Table S1 for a list of 240 compounds). Hierarchical clustering identified 14 subsets (or modules) of compounds with 241 generally low correlations among modules and high correlations within modules (see Supporting 242 Information Figs. S1 and S2 for correlations within and among modules, and Fig. S3 for module 243 variation among plants). The correlational structure of the phytochemical data is illustrated as an 244 adjacency network in Fig. 2, where it can be seen that some modules (e.g., modules 1 and 2) 245 contain a great diversity of compound types, while other modules are made up of more narrow 246 classes (e.g. modules 7 and 8 are mostly saponins, a class of defensive secondary metabolites 247 (Levin, 1976)). From the 14 eigenvectors summarizing variation in the modules, as well as the 248 other plant traits, lasso regression produced a reduced set of potential predictors which were then 249 used in Bayesian multiple regression models that included between seven and nine independent 250 variables (Table 1). The models had reasonably high performance in leave-one-out cross-251 validation (correlations between observed and predicted values were between 0.50 and 0.59, 252 Table 1), and also in resampling analyses (Supporting Information Fig. S4), where a small 253 fraction (never more than 4%) of randomly-generated models exceeded the variance explained of 254 the models reported in Table 1. 255

Variation among plants in the suites of covarying compounds had large effects on 256 caterpillar performance: for example, the beta coefficient of -2.33 (on the log-odds scale) 257 associated with module 3 corresponds to a reduction in mean survival from 0.58 to 0.12 258 associated with a one unit change in variation associated with that phytochemical module (Table 259 1; note that in Table 1 and elsewhere negative coefficients for development time are associated 260 with fewer days, and thus can be thought of as potentially beneficial effects, in contrast to 261 negative coefficients for survival and weight that are detrimental to caterpillars). The 262 phytochemical predictor variables are eigenvectors from clustering analysis, and thus are not 263 entirely straightforward to interpret, especially when the clustering analysis was itself based on 264 z-transformed data. It is useful to note that our LC-MS data consists of peak areas divided by the 265 peak of an internal standard, and again divided by the dry weight of the sample (thus, in total, 266

referred to as "relative abundance per dry weight"). Variation in these numbers reflects variation
in concentrations within compounds (among plants), but care should be used in comparing
among compounds because of different ionization responses relative to the standard (thus the use
of z-transformation for among-compound analyses). Nevertheless, the effects reported in Table
1 reflect real variation in suites of compounds, as can be seen in correlations between
eigenvectors and individual compounds in Supporting Information Fig. S2, and in variation
among plants in average z-scores in Fig. S3.

Modules included in multiple regression models frequently had common effects across response variables (e.g., the positive association of module 10 with both survival and adult weight), with the exceptions of module 11 that had a solitary effect on survival and module 6 with an effect only on development time (although the probability of the latter having a negative effect was only 0.75). Specific leaf area had a negative effect on survival and adult weight, and the coefficients for specific leaf area (-0.31 for survival and -0.42 for weight) were of smaller magnitude than most phytochemical effects.

Module-based analyses (as in Table 1) focused on feature reduction with lasso regression; 281 as a complementary analytical approach, we used ridge regression on all of the compounds 282 (which estimates effects of individual compounds without excluding variables as in lasso 283 regression). Analyses of individual compounds by ridge regression (Fig. 3) were broadly 284 consistent with the strongest module-specific effects, as can be seen, for example, with module 285 10 having positive effects on survival and adult weight in module analyses (Table 1) and in 286 compound-specific analyses (Fig. 3). Similarly, the individual compounds in module 3 had 287 negative compound-specific effects on survival (Fig. 3), and that module had the strongest 288 negative effect on survival in the eigenvector-based analyses in Table 1. Not surprisingly, the 289 larger modules (with a greater number of covarying compounds, including many primary 290 metabolites) tended to have a more complex mix of positive and negative effects (for example, 291 modules 1 and 2, Fig. 3). For ease of interpretation, the coefficients from compound-specific 292 regressions of survival and development time (in Figs. 3 and 4) have been back-transformed to 293 be on the scales of probability and days (respectively), and displayed as changes relative to 294 intercepts. For example, a compound with a relatively large effect on survival in Fig. 3 could be 295 associated with a one half percent (0.005) reduction in the probability of survival relative to 296 average survival and while holding other compounds constant. 297

We also considered potential pairwise interactions among individual compounds, and 298 found few interactions that passed the filter of the penalized regression (Supporting Information 299 Table S2), at least relative to the large number of potential interactions. Saponins and alkaloids 300 tended to be overrepresented in the interactions that were detected, and phenolic glycosides were 301 involved in stronger negative interactions relative to other compounds (Fig. S5). We did not find 302 evidence that more or less variable compounds had differential effects on caterpillars, although 303 there was a trend towards both greater positive and greater negative effects being associated with 304 less variable compounds (Fig. S6). We saw some variation among classes of compounds in their 305 effects on caterpillars (Fig. 4). All classes included positive and negative effects, with saponins, 306 alkaloids and phenolic glycosides including some of the stronger negative effects of individual 307 compounds, while lipids and sterols tended towards positive associations with survival and 308 development (Fig. 4). 309

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311 Discussion

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The results reported here represent a dissection of the phytochemical landscape facing a 313 specialized insect herbivore attacking a novel host plant (Hunter, 2016). The phytochemical 314 landscape is both physical, referring to variation in compounds among individual plants in a 315 common garden, and hypothetical to the extent that effects of individual compounds on 316 caterpillars are estimated, although compounds are of course not encountered in isolation. Our 317 exploration of the phytochemical landscape facing L. melissa on M. sativa is necessarily a first 318 draft based on a snapshot in time, and the implications of this are discussed further (below) in the 319 context of within and among-population heterogeneity in the system. Nevertheless, models 320 including suites of covarying compounds and other plant traits had predictive success and 321 suggested different natural products affecting survival, development time and adult weight 322 (performance measured as adult weight is a proxy for fecundity (Forister *et al.*, 2009)). Previous 323 work with *M. sativa* and insect herbivores has focused on sapogenic glycosides (Levin, 1976), 324 and a simple outcome from our study could have been that one or a small number of saponins 325 have anti-herbivore properties that reduce fitness of our focal insect. Instead, we find large 326 numbers of compounds with potentially consequential effects on caterpillars (Fig. 3), and the 327

effects of those compounds were greater than the effects of leaf toughness and specific leaf area (Carmona *et al.*, 2011).

The precise identification of specific compounds is perhaps not as important as the result 330 that prominent classes of secondary metabolites, such as saponins and peptides, include 331 compounds with both positive and negative effects. Positive effects of these compounds are 332 potentially associated with feeding stimulation, as has been observed (along with other positive 333 effects) for other specialist herbivores and plant toxins (Seigler & Price, 1976; Smilanich et al., 334 2016). Negative effects of saponins on insects potentially include disruption of hormone 335 production (Chaieb, 2010), although exact modes of action on L. melissa will await further study. 336 Although most of the individual compounds with strong effects appear to be secondary 337 metabolites (including alkaloids and phenolic glycosides, as well as saponins and peptides), we 338 also find both positive and negative effects associated with variation in certain primary 339 metabolites (Fig. 4). For example, the larger detrimental effects of individual compounds on 340 caterpillars include phospholipids (Fig. 2). These could be direct effects if a compound is 341 suboptimal for development, or they could be associated with nutritional imbalance (Behmer, 342 2009), such that too much of one nutrient makes it difficult for caterpillars to consume a 343 balanced diet. It has been suggested that the presentation of unbalanced nutrition can be a kind 344 of anti-herbivore strategy (Berenbaum, 1995). Although this possibility has not been thoroughly 345 investigated in many systems with full metabolomic profiling, the idea that nutritional imbalance 346 could be as important as direct toxicity suggests that we might update theories of the evolution of 347 plant defense that were built on differential investment into simple categories of plant growth 348 versus defense (Stamp, 2003). 349

The finding that our specialist herbivore is affected by a wide range of metabolites, 350 primary and secondary, that vary greatly even within a single host population has implications 351 for our understanding of heterogeneity in the system and for the course of local adaptation of the 352 herbivore to the novel host. Lycaeides melissa typically colonizes weedy or feral patches of M. 353 sativa on roadsides or integrated into natural communities, and previous work has documented 354 dramatic variation among individual alfalfa locations (often in close proximity) in the extent to 355 which they can support caterpillar development (Harrison et al., 2016). Previous phytochemical 356 data with a lower resolution was less successful in explaining that variation (Harrison et al., 357 2016), and the results reported here suggest that among patch variation could be explained by 358

future studies using metabolomic data as used here. The within-population complexity described 359 in the current study combined with previous evidence for dramatic among-population variation 360 in *M. sativa* suitability for the focal herbivore also raises the possibility that the novel host 361 presents a multi-faceted and potentially ever-shifting target from the perspective of evolving 362 butterfly populations. In particular, it is possible that *M. sativa* defense against a specialist 363 herbivore might be realized through different combinations (within and among populations) of 364 individually-acting compounds, thus making it unlikely for butterflies in any one population to 365 possess an effective suite of alleles that improve fitness on the novel host. In this context, it is 366 interesting to note that a molecular genetic dissection of caterpillar performance in this system 367 found a large number (potentially hundreds) of individual loci associated with performance on 368 *M. sativa*, yet evolution in populations associated with the novel host is primarily associated with 369 a loss (through genetic drift) of the ability to eat a native host and only slight improvement in the 370 ability to eat the novel host (Gompert et al., 2015). 371

The correlational structure of the phytochemical variation that we observed has 372 implications for the evolution of plant defense and the accumulation of insect herbivores on M. 373 sativa. Specifically, correlations among modules should make it possible to hypothesize 374 directions of least resistance for defense evolution. Compounds in module 9 had a negative 375 effect on survival (Table 1), and module 9 negatively covaried with module 10 (Supporting 376 Information Fig. S1), which itself had a positive association with caterpillar survival. Thus an 377 increase in module 9 and an associated decrease in 10 would be beneficial for the plant, at least 378 with respect to herbivory by our focal herbivore. Of course, most plants do not have the luxury 379 of optimizing defense against a single herbivore, and it is easy to imagine that improvements in 380 defense against one enemy could lead to increased attraction to another, especially given the 381 diversity of effects even within major classes studied here, including saponins and phenolic 382 glycosides. Compounds in the latter class (phenolics) were found to have strong positive and 383 negative effects on assemblages of arthropods associated with the maternal plants from which 384 seeds were collected to start the common garden used in the present study (Harrison et al., 2018). 385

The results reported here raise a number of avenues for future exploration, including the apparent overrepresentation of both saponins and alkaloids in interactions with other compounds (Fig. S5). Relative mass defect (RMD) is a useful tool for the categorization of compounds, but it has limitations in complex mixtures; we are developing methods that use other data from high

resolution mass spectrometry to further refine categorization of *Medicago* metabolites. Also, in 390 the present study, we have not attempted to separate constitutive and induced defenses (Jansen et 391 al., 2009) as the plants in the common garden were exposed to natural and continuous levels of 392 herbivory. We also acknowledge that feeding under laboratory conditions is of course not 393 natural, although we found in a previous study that genetic variants (in caterpillars) associated 394 with success in laboratory feeding trials were at least partially predictive of genetic variation 395 associated with alfalfa host use in the wild (Chaturvedi et al., 2018). Thus it is clear that 396 metabolomic data such as that analyzed here has the potential to both open up new avenues of 397 conceptual development in plant-insect interactions and to link micro-evolutionary trajectories 398 across hosts and herbivores. 399

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409 Author contributions

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411 MLF: designed experiment, conducted analyses, wrote first draft. SY: conducted experiment

and contributed to experimental design. CSP, CDD, BH: generated and interpreted

⁴¹³ phytochemistry and protein data. MLF, JGH, OS: developed and maintained common garden.

JAF, ZHM, CCN, LAR: contributed to analyses and experimental design. CAB, JAF, ZG, CCN:

415 contributed to experimental design. All authors: contributed to writing.

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516 Figure legends (also reproduced below individual figures)

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Fig. 1. Variation among plants in caterpillar survival (a), development time (b) adult weight (c), 518 three individual compounds (d-e) and two external plant traits, specific leaf area (g) and leaf 519 toughness (h). The three example compounds shown here (out of the 163 assayed) were among 520 the top five most influential compounds for survival, development time and adult weight: cpd. 9 521 is an alkaloid with a negative association with survival, cpd. 94 (a peptide) has a negative 522 association with development time, and cpd. 160 is a phospholipid with a negative association 523 with adult weight. Individual plants in all panels are organized from left to right by decreasing 524 caterpillar survival in the top panel (a). Standard errors are shown for panels b, c, g and h. The 525 units for d-e are compound relative abundance per dry weight of sample; the units for specific 526 leaf area are cm^2/mg , and grams/newton for leaf toughness. 527

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Fig. 2. Illustration of correlational structure among compounds: each node in the network is a 529 compound, and compounds are linked by a line if they are correlated among individual plants at 530 0.5 or above (links among compounds in modules 12-14 represent weaker correlations, greater 531 than 0.1; see main text for details). Two letter codes within nodes indicate compound classes, as 532 explained in the legend. Colors of nodes correspond to membership in modules as determined by 533 hierarchical cluster analysis; the color key to the 14 modules is shown in the lower left. Not shown are a 534 small number of compounds with weak connections to all other compounds, including two compounds 535 that were not included in any module (shown as module zero in Fig. 3). 536

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Fig. 3. Effects of individual compounds on survival, development time and adult weight, as estimated by 538 ridge regression (using binomial, Poisson and Gaussian models, respectively). The strength of effect for 539 each compound is indicated by the horizontal extent of each bar, and compounds are grouped by modules 540 541 (m1, m2, etc.); the order of compounds along the vertical axis is arbitrary within modules and fixed across columns. Orange colors indicate negative effects on survival, development and weight, while blue colors 542 are positive effects (note that negative effects for development time correspond to fewer days, or more 543 rapid development). The darker shades of orange and blue mark coefficients whose 95% confidence 544 intervals did not overlap zero in 1,000 bootstrap samples. Values for survival and development time have 545

been back-transformed from units on the log-odds and log scales to units of probability and days to 546 pupation, and are shown as changes from the mean or intercept values. For example, a negative (orange) 547 survival coefficient of 0.005 means a one-half percent reduction in average probability of survival 548 associated with variation in a particular compound. The fifteen compounds with the largest coefficients 549 (by absolute value) and bootstrap intervals not overlapping zero are labelled by compound classes (see 550 Fig. 2 for abbreviations) in each panel. Structural annotations are shown to the right for six compounds 551 based on matches from the METLIN metabolomics database, as follows by compound number: 154 552 (unidentified sterol); 9 (unidentified alkaloid); 60 (soyasaponin A3); 40 (unidentified saponin); 46 553 (medicagenic acid 3-O-triglucoside); 45 (medinoside E). Those same compounds are identified in 554 parentheses in the main panels next to the bars corresponding to their individual effects. 555 556 Fig. 4. Violin plots of compound-specific effects (coefficients from ridge regressions) summarized by 557 chemical classes. Plots show median (black dot), interquartile range (box) and 95% confidence intervals 558 (whiskers) surrounded by kernel density envelopes. Sample sizes for each category are shown above the 559

top panel ("Other" includes 1 sugar, 2 pigments and 2 halogenated compounds). Categories are arranged

⁵⁶¹ from left to right based on the gradient of median positive to negative effects on survival.

562 Coefficients for survival (a) and development time (b) have been back-transformed from the

⁵⁶³ units of log-odds and log to probability and days to pupation, respectively, and shown as

deviations from the mean or intercept value (as in Fig. 3). Note that negative effects for

development time correspond to fewer days (more rapid development).

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Table 1 Results from Bayesian regressions of module eigenvectors and covariates predicting caterpillar survival, development time and adult weight (as binomial, Poisson, and Gaussian regressions, respectively, with corresponding units in log-odds, log number of days, and milligrams).

	Survival	Development time	Weight
	coefficient (CI; prob.)	coefficient (CI, prob.)	coefficient (CI, prob.)
m2	0.37 (-1.87, 2.67; 0.63)	-0.07 (-0.26, 0.12; 0.78)	
m3	-2.33 (-3.94, -0.72; >0.99)		-2.33 (-4.87, 0.24; 0.96)
m6		-0.05 (-0.19, 0.09; 0.75)	
m9	-2.31 (-4.49, -0.15; 0.98)	0.064 (-0.11, 0.24, 0.78)	
m10	2.54 (0.81, 4.23; >0.99)		3.56 (1.22, 5.84; >0.99)
m11	2.01 (0.64, 3.42; >0.99)		
SLA	-0.31 (-0.55, -0.08; >0.99)		-0.42 (-0.78, -0.06; 0.99)
Protein	0.061 (-0.16, 0.28; 0.71)		0.08 (-0.26, 0.42; 0.68)
Tough.	0.036 (-0.17, 0.24; 0.36)	-0.002 (-0.02, 0.02; 0.57)	
Sex	/	0.06 (0.02, 0.10; >0.99)	1.10 (0.38, 1.83; >0.99)
mg	/	-0.03 (-0.05, -0.01; >0.99)) /
Days	/	/	1.10 (0.38, 1.83; >0.99)
Intercept	0.34 (0.14, 0.53; >0.99)	3.48 (3.45, 3.52; >0.99)	10.36 (9.81, 10.92; >0.99)
Validation	0.53	0.59	0.50

For each regression coefficient, numbers in parentheses are 95% credible intervals (the first two numbers) and the probability that the coefficient has the estimated sign (e.g., 0.63 for the m2 survival coefficient of 0.37 indicates a 63% probability that the m2 module has a positive effect on survival). Note that negative coefficients for development time indicate faster caterpillar development (fewer days) associated with variation in a particular compound. Modules (listed in the left column) are only shown if they were included in one of the three regressions following feature selection using lasso regression (see main text for additional details). Empty spaces in the Table appear if a particular module was selected through lasso regression for one or two analyses but not all three (m3, for example, was not selected by lasso regression for development time). Slash marks (/) indicate variables not considered for a particular analysis (e.g., sex, adult weight [mg] and development time [days] were not possible for the survival analysis because they are not observed on dead individuals). Values for "validation" shown in the last row are the correlation between observed and predicted values in cross-validation (Supporting Information Fig. S4).

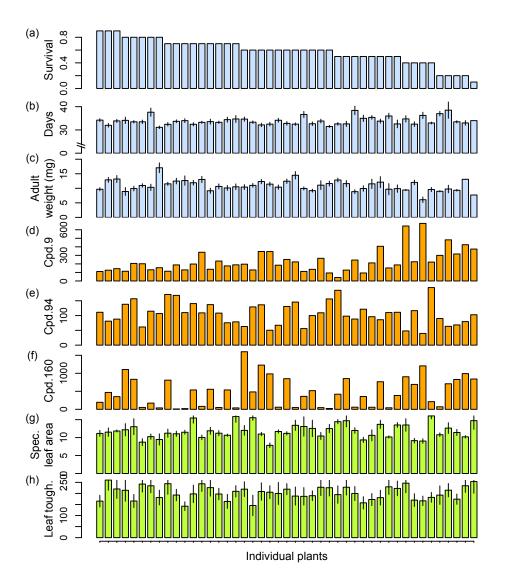


Fig. 1. Variation among plants in caterpillar survival (a), development time (b) adult weight (c), three individual compounds (d-e) and two external plant traits, specific leaf area (g) and leaf toughness (h). The three example compounds shown here (out of the 163 assayed) were among the top five most influential compounds for survival, development time and adult weight: cpd. 9 is a peptide with a negative association with survival, cpd. 94 (another peptide) has a negative association with development time, and cpd. 160 is a phospholipid with a negative association with adult weight. Individual plants in all panels are organized from left to right by decreasing caterpillar survival in the top panel (a). Standard errors are shown for panels b, c, g and h. The units for d-e are compound relative abundance per dry weight of sample; the units for specific leaf area are cm²/mg, and grams/newton for leaf toughness.

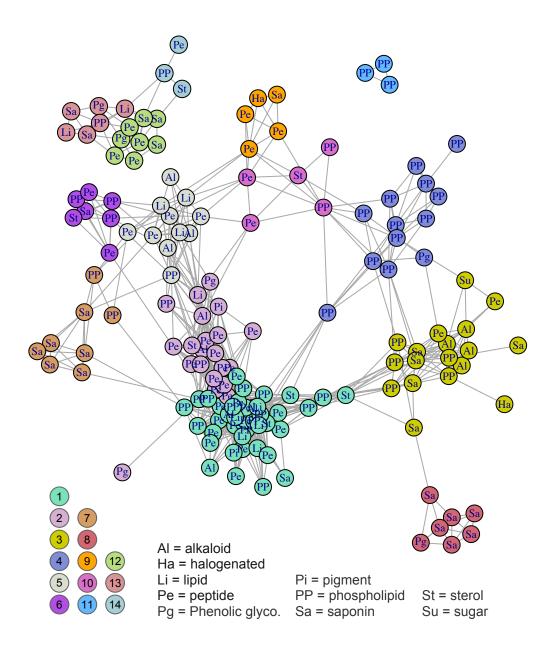


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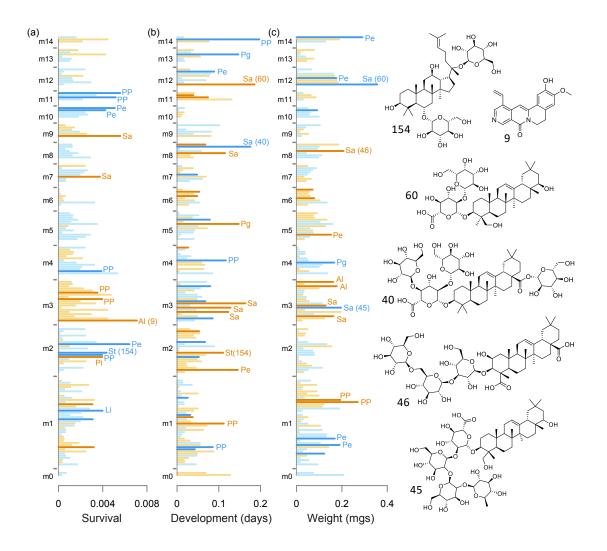


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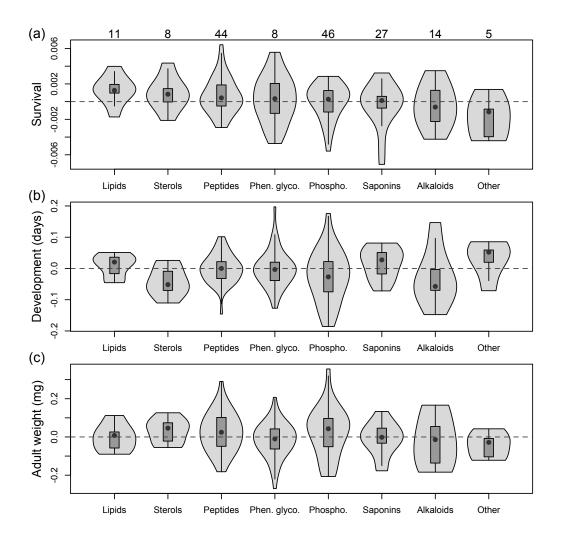


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