ANOMALOUS PERCEPTION OF BIOLOGICAL MOTION IN AUTISM:

A CONCEPTUAL REVIEW AND META-ANALYSIS

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Abstract

Despite its popularity, the construct of biological motion (BM) and its putative anomalies in autism spectrum disorder (ASD) are not completely clarified. Here, we propose a new model describing distinct levels of BM processing, and we present a meta-analysis investigating BM perception in ASD. We screened 114 articles testing BM perception in ASD and typical developing peers. A general meta-analysis including all the selected studies (N=27) showed BM processing moderate deficit in ASD, but high heterogeneity. This heterogeneity was explored in different additional meta-analyses where studies were grouped according to different levels of BM processing (*first-order/direct/instrumental*) and the manipulation of low-level perceptual features (*spatial/temporal*). Results suggest that the most severe deficit in ASD is evident when perception of BM is serving a secondary purpose (e.g., inferring intentionality/action/emotion) and, interestingly, that temporal dynamics could be an important factor in determining BM processing anomalies in ASD. In conclusion, this work questions the traditional understanding of BM anomalies in ASD and claims for a paradigm shift that deconstructs BM into distinct levels of processing and specific spatio-temporal subcomponents.

Keywords: social cognition, biological motion, temporal processing, visual perception.

Public Significance statement

Since the seminal study by Johansson (1973), the construct of "biological motion" (BM) has gained a considerable success in a wide range of disciplines. In particular, BM processing has been considered a putative marker for social difficulties in neurodevelopmental conditions such as autism spectrum disorder (ASD). Our work aims to quantitatively test the solidity of this view through a meta-analytic approach and also to better define anomalies in BM perception according to distinct levels of complexity and specific spatio-temporal features. Interestingly, we do it by challenging the traditional approach to the conception of BM. This novel conceptualization has intriguing clinical and theoretical insights.

1. Introduction

Our brain is constantly facing a plethora of sensory stimuli that needs to be properly sampled and organized to construct a meaningful perceptual experience. Since seminal studies on point-light displays investigating the perception of distinct types of motion patterns such as walking, running and dancing (Herrington et al., 2011), the robust tuning of the human visual system to biological motion (BM) has represented an intriguing challenge for scientists (Neri et al., 1998; Pavlova, 2012; Proffitt et al., 1984; Simion et al., 2008). Compared to full body motion, stimuli created with point-light displays (PLD) permit to profitably separate motion processing from other features such as color or shape, representing an essential tool to investigate the principles behind our perceptual tuning to BM. In light of its implications for developmental, clinical and cognitive psychology and neuroscience, the study of BM processing raised considerable attention in recent years (Rosa Salva et al., 2015; Thompson and Parasuraman, 2012). In particular, its putative implication in social cognition have been widely debated in the context of Autism Spectrum Disorder (ASD) raising the hypothesis that BM processing could be considered a marker or an intermediate phenotype of ASD (Pavlova, 2012). The present work has multiple aims. First, to test whether and eventually to what extent BM processing is generally anomalous in ASD. Second, to offer a refined theoretical and conceptual model for deconstructing BM into distinct levels of processing and according to distinct perceptual manipulations. Third, to present a quantitative meta-analysis of putative BM processing anomalies in ASD in line with these new conceptual classifications.

2. Autism spectrum disorder: symptomatology, pathogenesis and putative markers.

Being characterized by early onset, lifelong and debilitating behavioral symptomatology, ASD represents a serious concern for families, clinicians and generally for societies (Buescher et al., 2014; Lord et al., 2018). To date, clinical diagnosis of ASD is based on behavioral symptoms and according to DSM 5 they include restricted and repetitive patterns of behaviors, interests or activities, and difficulties in the social communication/interaction domain (Bullivant and Loewenstein, 1968). Hyper- or hypo-reactivity to sensory stimuli or unusual interest for specific sensory aspects are also considered important features of the disorder (Robertson and Baron-Cohen, 2017). One of the major concerns in dealing with ASD is the heterogeneity in severity and behavioral manifestations and difficulties in early detection when behavioral signs are less evident (Lai et al., 2014). These aspects significantly impact on the clinicians' ability to provide early and specific intervention (Constantino and Charman, 2016). Although the growing body of studies targeting neural structures implicated in ASD pathophysiology has improved our insights on the condition (Fatemi et al., 2012; Kucharsky Hiess et al., 2015; Paul et al., 2014; Wang et al., 2014), the neurobiology of ASD is far to be clarified (Di Martino et al., 2014; Ecker and Murphy, 2014; Haar et al., 2016; Lai et al., 2014; Stoner et al., 2014). Similarly, neither brain functional and connectivity approaches (Khan et al., 2015; Lewis et al., 2017) nor genetics (Geschwind, 2011; Richards et al., 2015) have definitely solved the puzzle. Efforts in understanding ASD are hampered by substantial disconnection between neurobiological findings aiming to characterize the neural/genetic components associated with the condition, and its core behavioral phenotype. Thus,

research investigating endophenotypes of ASD, i.e. measurable and quantifiable links between neurobiological underpinnings and behavioral symptoms (Gottesman and Gould, 2003; Kendler and Neale, 2010), represent a critical challenge for the future (Homberg et al., 2016; Klin et al., 2015). In recent years, a growing body of studies on endophenotypes of ASD has been performed exploring complex functions such as visual attention (Gammer et al., 2015; Keehn et al., 2013; Robertson et al., 2013; Ronconi et al., 2018c, 2018b, 2013), visual perception (Ewing et al., 2013; Foss-Feig et al., 2013; Kröger et al., 2014; Vlamings et al., 2010), motor cognition (Ansuini et al., 2018; Becchio et al., 2007; Casartelli and Parma, 2017; Cattaneo et al., 2007; Rochat et al., 2013) and sensory processing (Collignon et al., 2013; Endevelt-Shapira et al., 2018; Leslie D Kwakye et al., 2011; Parma et al., 2013; Robertson and Baron-Cohen, 2017; Rozenkrantz et al., 2015).

A significant interest in the literature has been directed to the idea that individuals with ASD have a different perceptual experience of the world and, even more critical for our aims, that sensory/perceptual anomalies in ASD are not just a secondary effect of reduced social interactions. Accordingly, sensory/perceptual anomalies would represent a key and primary component both in terms of symptoms development and pathophysiology (Robertson and Baron-Cohen, 2017). A specific focus has been oriented on BM processing in ASD. Following the first report of BM impairment in ASD (Moore et al., 1997), the number of studies investigating this question has constantly increased over the years (see Figure 1). However, factors such as heterogeneous experimental designs, weak understanding of potential confounds and inconsistency among the classes of information conveyed by BM may rise various misunderstandings. To date, there is a general and quite vague agreement on the supposed BM anomalies in ASD, which however lacks both a clear

epistemological awareness of the construct of BM and a quantitative evaluation of these alleged anomalies (Blake et al., 2003; Pavlova, 2012; Pelphrey and Carter, 2008; Redcay, 2008).

3. Biological motion processing across phylogenesis and ontogenesis

3.1. Biological motion processing across phylogenesis

The suitability of BM as a putative endophenotype of ASD would require a co-segregation with the genetic background of the condition. Preliminary evidence exploring this issue has been proposed in the literature testing autistic traits normally distributed within the general population or infants at high familial risk of ASD (Di Giorgio et al., 2016; Klin et al., 2015; Puglia and Morris, 2017; Thurman et al., 2016; van Boxtel et al., 2017; van Boxtel and Lu, 2013) (see also the notion of "broader autism phenotype", Billeci et al., 2016; Piven, Palmer, Jacobi, Childress, & Arndt, 1997). Recently, a genetic study including 117 same-gender twin pairs by Wang and colleagues (2018) has recently observed that the ability to process the inherent kinematic properties of BM is heritable and negatively correlated with autistic traits. Another strategy to mechanistically investigate how this building block of social cognition evolves is to evaluate the observed heritable traits of BM processing across phylogeny. A large and converging body of evidence indicates that BM is an evolutionarily well-conserved ability. Rapid and effortless detection, recognition and interpretation of BM represents an adaptive tool for vertebrates to support efficient interaction with other living beings (Thompson and Parasuraman, 2012). The ability to follow the movements of conspecifics, learning by imitation and disentangling avoidant vs. prosocial interaction are just some of the many examples showing the functional role

of BM processing in supporting social behaviors (Klin et al., 2015). Comparative studies demonstrate sensitivity to BM processing in a wide range of species including primates, cats, rats, several birds and fishes. Pigeons recognize PLD of walking and pecking birds (Juon et al., 2010) and can also discriminate the PLD of a pigeon vs. a toy dog (Omori and Watanabe, 1996). Medaka fish is induced to shoaling behavior by the observation of the motion of conspecifics, whereas deviation from normal and familiar motion does not promote shoaling behavior (Nakayasu and Watanabe, 2014). Cats can discriminate a walking cat from similar spatially scrambled displays (Blake, 1993). A certain degree of sensitivity to BM is reported also in rats, even though they do not seem to be able to generalize their trained ability to distinguish between left- and rightward walking humans (MacKinnon et al., 2010) to a novel PLD (e.g. a walking cat). Very limited ability in transferring the recognition to novel PLD is reported also in baboons (Parron et al., 2007). Specific neural populations in the macaque temporal cortex (superior temporal polysensory area) encode without training forward vs. backward human locomotion, indicating a certain degree of tuning to BM decodable at the neural level. Nonetheless, Rhesus monkeys need extensive training to discriminate forward vs. backward human walking (Vangeneugden et al., 2009). In recent years, also an increasing interest for chicks and their predisposed/learned preference for specific stimuli (BM, but also face-like configurations, self-propelled objects, etc.) has been reported (Di Giorgio et al., 2016). Newly hatched chick, lacking any previous visual experience being reared and hatched in darkness, exhibit a spontaneous preference in approaching BM rather than non-BM patterns. In addition, this predisposition has been demonstrated to be generalizable to other vertebrates, such as cats, who can be considered potential predators for hen (Vallortigara et al., 2005). Taken

together, these comparative pieces of evidence strongly support the idea that some degree of tuning to BM is phylogenetically robust, although not all species show the same efficiency in generalizing BM predisposition to other stimuli (e.g., to other species). Understandably, humans are no exception.

Cross cultural studies represent an intriguing approach in elucidating the neurobiological predisposition and the environmental influences to the processing of BM. Exploring how the Mundurukù, an isolated Amazonian indigene group, process BM stimuli suggested that even this population, without prior exposure to representational depictions of human or animal activity (e.g. PLD), perceived the coherent, global shape depicted in BM stimuli and experienced the classic inversion effects when such stimuli are turned upside down (Pica et al., 2011).

In light of the robust phylogenetic roots of BM, it is not surprising that atypical BM processing in distinct clinical conditions such as ASD has been linked to cascade effects in the development of more complex abilities. However, before analyzing in more detail anomalies of BM processing in ASD we first consider the ontogenesis of this ability in humans.

3.2. Biological motion processing: an early emerging ability

We may expect that a phylogenetically highly preserved mechanism, such as BM processing, also shows an early developmental progression during the ontogenesis. However, it still remains to be clarified whether, and eventually to what extent, the tuning to BM is dependent on experience. A seminal study in infancy indicates that at 3 months of age infants are able to discriminate BM stimuli vs. their scrambled and upright

counterparts (Zhang et al., 2013). This trend is observed also at 4-5 months of age (Fox and McDaniel, 1982; Zhang et al., 2013). More recently, a convincing argument in favor of the hypothesis that BM processing is largely independent from experience was provided by a study testing the ability of newborns to discriminate BM stimuli (Simion et al., 2008). The sensitivity to BM was tested in 2-day-old babies using hen-walking animations as compared to nonbiological stimuli (i.e., the same pattern of elements moving in a random manner) and inverted (upside -down) hen-walking stimuli. The use of hen stimuli allowed to exclude any possible (although remote) learning mechanisms taking place in the first 2 days of life. The results indicate that newborns manifest a spontaneous preference for BM stimuli, supporting the idea that this processing is largely (or even completely) independent from experience (Simion et al., 2008). Although the authors could not rule out the possible role of fetal experience in shaping BM predispositions (see for example Pavlova, 2012), these findings supported a certain degree of innate or nearly innate tuning of human perceptual system to BM.

The developmental trajectory of the sensitivity to BM stimuli has also been extensively explored in comparison to global motion (Hadad et al., 2011). Testing different groups of children (aged 6-8, 9-11 and 12-14 years, respectively) and adults, Hadad and colleagues (2011) claimed that developmental trajectories for coherent and BM are similar. In both cases, the authors found that thresholds achieved adult-like profiles around the age of 14, following a sort of monotonic developmental trajectory of improvement from childhood. In contrast, testing before surgery the BM processing and global motion in patients that suffered from early visual deprivation because of dense bilateral congenital cataracts, showed that marked deficits in global motion are not accompanied with similar difficulties

in BM processing (Hadad et al., 2012). Importantly, a recent EEG study showed that the BM perception is not only preserved at the behavioral level, but also that the same neural processing is involved. Indeed, in patients with congenital cataracts, the N1 component of event-related potentials was modulated post-surgery by the processing of BM as in the control group (Bottari et al., 2015). Spared abilities in the processing of BM after early visual deprivation are surprising in light of the data reporting impairments in similar cohorts of patients both in global motion (Lewis et al., 2002) and in holistic face processing (Le Grand et al., 2004; Robbins et al., 2010).

Taken together, these data indicate a composite picture in which clear similarities (Hadad et al., 2011) and marked differences (Hadad et al., 2012) between basic and more complex mechanisms (e.g., global motion and BM processing) coexist (see also for example Mather, Battaglini, & Campana, 2016). Such coexistence should stimulate further studies aiming at characterizing shared developmental patterns vs. neuroplastic mechanisms. For shared patterns, a classical approach refers to the well-established neurocostructivist view supporting the idea that very early, not-detailed, domain-general functions impact on the development of more complex, domain-specific, and even "social" functions (Karmiloff-Smith, 1998; Ronconi et al., 2016). For compensatory mechanisms, examples from heterogeneous clinical pictures may provide intriguing insights (Hadad et al., 2012; for example, related to other functions, see also Casartelli & Parma, 2017; Rezlescu, Barton, Pitcher, & Duchaine, 2014). This overview of both BM phylogenesis and ontogenesis suggests that oversimplifying this construct, as if it was to convey a monolithic process, can be misleading and a paradigm shift in approaching BM is needed.

4. **Motor, spatio-temporal and social aspects in biological motion processing**

It has been hypothesized that the ability to detect, recognize and interpret BM is related, in different ways, to the expertise in performing similar movements or actions and, to some extent, also to one's social abilities (Pavlova, 2012). Beyond the attractiveness of this view, whether and how the ability to process BM is linked *tout court* to the expertise in producing similar movements or actions requires a cautious interpretation. Recent experimental and theoretical investigations on mirror mechanisms clearly indicate that humans benefit from their own motor representations to understand actions of others directly ("motorically", "from the inside"; Rizzolatti & Sinigaglia, 2016). Noteworthy, mirroring phenomena may occur at distinct level of abstractness, as if we are able to represent distinct components of an action in a motor way (Rizzolatti et al., 2001; Rizzolatti and Sinigaglia, 2016), and they have been hypothesized to be anomalous in ASD (Ronconi et al., 2016; Trevarthen and Delafield-Butt, 2013).

Two studies reporting seemingly diverging results may help in furnishing further insights on this point. First, adolescents born preterm and suffering from periventricular leukomalacia (PVL), a clinical condition characterized by white matter lesions near the lateral ventricles, were tested in their ability to process BM (Krägeloh-Mann et al., 1999, 1993; Pavlova, 2012; Pavlova et al., 2003; Pelphrey et al., 2003). These participants varied in their motor ability, ranging from complete walking disability to typical or near-typical motor patterns (Pavlova et al., 2003). Impairment in the execution of actions with the lower and upper limbs correlated with the lateral extent of periventricular lesions, but did not correlate with the performance at the BM task. Interestingly, sensitivity to the processing of BM negatively correlated with the volume of the lesions over the parieto-occipital

complex, whereas no correlation over the frontal and temporal regions was found (Pavlova et al., 2003). Second, the link between action perception and action execution was explored by Casile and Giese (2006), who tested whether the acquisition of novel motor behaviors could improve the processing of BM stimuli. The authors demonstrated that nonvisual motor training selectively improves the visual recognition of novel BM patterns, thus suggesting that motor learning, independently of visual feedback given that participants were blindfolded, impacts on BM recognition (Casile and Giese, 2006). Taken together, these two studies outline a more articulated picture. The hypothesis that BM recognition is linked to some extent to the abilities in producing specific motor patterns is fascinating, but its simplistic interpretation should be overcome (Casile and Giese, 2006). At the same time, BM processing seems to benefit from more efficient and unexpected compensatory mechanisms, in turn suggesting that functional networks supporting BM processing are more complex (Pavlova et al., 2003 see also Sokolov, Erb, Grodd, & Pavlova, 2014).

Another source of debate in the literature is how spatio-temporal features of BM stimuli impact on BM processing itself. BM processing entails both spatial and temporal dynamics, but the exact contribution of these components and individual tolerance to their perturbation have not been clarified in ASD (Hadad et al., 2011; Troje and Westhoff, 2006). This may offer insights toward the understanding of apparently controversial findings reporting in the ASD literature, given that being tolerant to spatial and/or temporal perturbations are supported by potentially distinct neural mechanisms. Recent evidence refined traditional views about timing, attributing also to the cerebellum a key role in it (Bareš et al., 2018). Nowadays, cerebellar contribution to non-motor functions such as sensory and perceptual ones are well-established (Baumann et al., 2015; Buckner, 2013).

Interestingly, also the contribution of the cerebellum in the pathophysiology of ASD (Fatemi et al., 2012) and – more specifically – in ASD sensory/perceptual anomalies have been recently explored (Casartelli et al., 2018; Wang et al., 2014).

The complexity of networks supporting BM processing and neuroplasticity mechanisms that compensate, at least partially, the recognition of BM may be considered cues of the functional relevance of BM. A classical view suggests that the perception of BM is tightly linked with individual abilities in the social domain, leading back this hypothesis to the evolutionary well-established benefit in distinguishing efficiently biological from nonbiological motion. This "social" interpretation of BM posits that performance in BM tasks may serve as benchmark for unveiling difficulties in social functioning, for example in neurodevelopmental condition such as ASD (Pavlova, 2012; Rosa Salva et al., 2015). Beyond the considerable success of this view (Kaiser and Pelphrey, 2012), refined theoretical models and robust quantitative approaches (i.e., meta-analysis) are lacking and whether and how BM processing is linked to social functioning is not clear. One possibility is to interpret BM processing as a building block of more complex social abilities in the context of the so-called "social brain". Poor performance in BM discrimination in ASD would then be a marker of their social functioning difficulties (Kaiser and Pelphrey, 2012; Pavlova, 2012). The exact nature of the "social brain" construct (Elsabbagh and Johnson, 2016; Gammer et al., 2015; Whiten and van de Waal, 2017) and of other related constructs such as "theory of mind" or "mindreading" (Leslie et al., 2004; Meunier, 2017; Premack and Woodruff, 1978), which gained considerable attention in the last decades in a wide range of disciplines, remains however often vague and inconsistent among studies. The

methodologies employed for exploring these constructs are often heterogeneous and related neurobiological underpinnings are unclear (Schaafsma et al., 2015).

5. Deconstructing biological motion processing: a multi-level conceptual model

As outlined in previous sections, BM perception is phylo- and onto- genetically robust and it has been hypothesized to be a putative marker for ASD (Di Giorgio et al., 2016; Klin et al., 2015; Pavlova, 2012). In addition, humans seem to have preserved a certain degree of confidence also in processing perturbed (or scrambled) versions of BM stimuli, both in temporal and spatial version (Troje and Westhoff, 2006). However, the use of the construct of BM remains often quite vague and ambiguous. We propose a conceptual distinction that aims at deconstructing the notion of BM that in turn should help in classifying distinct (sub)components that play a role in BM processing. A similar risk could be observed also in other complex constructs such as "theory of mind" (Baron-Cohen et al., 1995, 1985; Fishman et al., 2014), "social brain" (Adolphs, 2009; Elsabbagh and Johnson, 2016) or "imitation" (Casartelli and Molteni, 2014; Casartelli and Parma, 2017; Caspers et al., 2010; Lesourd et al., 2018). When these constructs are treated as monolithic ones, we may run into misleading consequences. From a conceptual and epistemological perspective, they may convey different meanings and could translate into highly heterogeneous experimental protocols. The ambiguity in converting in operational terms these constructs is a key node because it can threat reproducibility of findings and reliability of results (Schaafsma et al., 2015).

Our model distinguishes between different levels of BM processing, which mediate distinct levels in BM processing. Specifically, we propose a three-levels model: I) first-order processing of BM, II) direct processing of BM, and III) instrumental processing of BM.

I) First-order processing of BM refers to the implicit detection of BM that implicates a behavioral response which does not entail any explicit recognition or categorization of BM. Experimenters infer that participants have detected a BM stimulus by simply evaluating their spontaneous behavioral response (e.g., looking time in preferential looking paradigms).

II) Direct processing of BM involves a form of categorization of distinct features of stimuli. This level of processing needs at least a basic degree of recognition of BM, despite an explicit recognition and categorization is not always necessary (e.g. BM Vs. non-BM; BM "A" Vs. BM "B"; BM "A" Vs. scrambled BM "A"; right walking BM Vs. left walking BM). The efficiency in direct processing of BM is usually measured in terms of accuracy or reaction times.

III) Instrumental processing of BM implicates that one takes advantage of the recognition and categorization of a BM stimulus for a secondary purpose, such as inferring other's intentionality or emotional state, and interfering with or predicting someone else's action. Here, recognition and categorization of BM stimulus, which are generally explicit, are tools to disclose distinct and possibly more complex information. Accuracy and reliability of such instrumental use of BM is usually based on the identification in a multiple alternative forced choice task or identification with free description.

Our three-level model aims to disentangle distinct levels of information conveyed by BM. In addition, the present work aims to preliminarily assess whether different spatial, temporal or spatio-temporal features of BM stimuli that can selectively affect BM processing in ASD. Altogether, the present meta-analysis will develop according to the following four steps:

a) A general analysis exploring all studies using PLD to investigate BM in ASD;

b) An analysis testing BM processing in ASD according to our three-levels model (i.e., first-order, direct and instrumental BM processing);

c) An analysis investigating the contribution of spatial and temporal confounds in how individuals with ASD process BM;

d) An analysis comparing ASD vs TD performance in non-BM stimuli.

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6. Methods

6.1. Literature Search

Following the PRISMA-P guidelines (Moher et al., 2015), we have established a protocol detailing the a priori rationale, the methodological and analytical approaches to which we adhered during the preparation of the present meta-analysis. This protocol does not represent an update of any other reports and it is not part of a systematic review on this topic.

An automatic literature search began on September 2017 and the last search was performed in March 2018. We searched for all studies published in English in both PubMed and PsycInfo, by performing a search on titles, abstracts and keywords. For direct replication, we include the three Boolean queries used to the search the PubMed database: 1) (("biology"[MeSH Terms] OR "biology"[All Fields] OR "biological"[All Fields]) AND ("motion"[MeSH Terms] OR "motion"[All Fields])) AND ("autistic disorder"[MeSH Terms] OR ("autistic"[All Fields] AND "disorder"[All Fields]) OR "autistic disorder"[All Fields] OR "autism"[All Fields]); 2) "Human motion"[All Fields] AND ("autistic disorder"[MeSH Terms] OR ("autistic"[All Fields] AND "disorder"[All Fields]) OR "autistic disorder"[All Fields] OR "autism"[All Fields]); 3) point-light[All Fields] AND ("autistic disorder"[MeSH Terms] OR ("autistic"[All Fields] AND "disorder"[All Fields]) OR "autistic disorder"[All Fields] OR "autism"[All Fields]). In addition, we also performed a manual literature search which encompassed a search of the reference lists of the review articles and the original research papers retrieved via the Boolean queries. One paper was retrieved by using this method. Altogether the automatic and the manual literature searches

resulted in 114 unduplicated published articles. No reviews, book chapters, master and doctoral theses, conference presentations or unpublished data were hereby included.

6.2. Screening Process

The studies retrieved from the automatic and manual literature searches were independently assessed and screened by two authors (A.F. and L.Ra.). At the end of the first round of screening, they evaluated together the differences emerged from the two separate searches and discussed with the other authors the final set of papers to be included, based on the inclusion criteria as detailed below. The *n* reported represents the number of studies excluded at each step. Please, also refer to Figure 2.

1) Screening: from the analysis of title and abstract, the empirical studies that did not include a task in which the processing of BM was assessed via PLD in both an ASD and a TD group were excluded from the initial paper count (*n*=55).

2) Eligibility: from reading the whole publication, studies were further excluded if they did not meet the same criteria of the previous screening step (*n*=16).

3) Qualitative assessment: from reading the whole publication, the articles were further excluded if they did not meet the following criteria:

   a. If the study sample was the subset of a bigger sample presented in a different paper, the study with the smaller sample was discarded (*n*=2; i.e. Klin & Jones, 2008; Koldewyn, Whitney, & Rivera, 2011).

   b. If the study did not investigate behavioral measures such as accuracy, sensitivity (d'), threshold estimation, reaction times (RTs), or percentage of preferential looking (or first orienting) in a BM task (*n*=8).

c. If the dependent variables and/or task employed were not comparable to the majority of the studies included. Specifically, these three studies were removed for the following reasons: i) Krüger and colleagues (2017) used a confidence rating as the dependent variable, no accuracy or RTs were reported; ii) Van Boxtel and colleagues (2016)used as the dependent variable the difference in the point of subjective equality (PSE) between the walking and running conditions, and thus the data related to BM performance could not be retrieved; iii) Swettenham and colleagues (2013) used an attentional cuing paradigm in which cues were BM stimuli and reported RTs for valid and invalid conditions, which do not primarily map onto BM processes.

d. If the study did not have sufficient statistical information (either present in the paper, supplementary materials or provided after contacting the authors) (n=3).

-------------------------------------Please insert here Figure 2-------------------------------------

6.3. Coding and Reliability

The final sample of studies were coded based on the following rules:

1. Level of BM processing, which led to the creation of three main categories; 1) First-order; 2) Direct recognition; 3) Instrumental recognition of BM;

2. Type of scrambled stimulus used in tasks that assess the perception of BM vs. its scrambled version: 1) no scrambled; 2) spatial; 3) temporal; 4) spatial and temporal scrambled.

3. Type of motion stimuli presented: 1) BM, when the data refer to the performance following the presentation of solely BM stimuli; 2) BM vs. non-BM, when the dependent variable refers to a task in which the perception of BM stimuli is contrasted with the perception of non-BM stimuli; 3) non-BM, when it is possible to separately extract data concerning only the performance of non-BM stimuli (i.e., the scrambled version of the BM stimulus or the motion of an object, such as a truck or a bicycle). Disagreement and inconsistencies between the two raters (A.F. and L.Ra., N=7/114, 6% of all assessments) were questioned with all authors to gather consensus.

In addition, for all the studies we collected:

1. The relevant sample sizes of both ASD and the TD groups;

2. The percentage of male participants in each of the two groups;

3. The mean age of each of the two groups;

4. The test employed to assess general cognitive level of participants, and the relative mean score for each of the two groups;

5. The type of behavioral measure collected: accuracy rates (including error rates) or percentage of looking time. Since RTs were present only in 3 of the studies included in this meta-analysis, we preferred to exclude them to increase the homogeneity and the comparability of measures;

6. Whether behavioral tasks are performed during an EEG or fMRI acquisition;

7. A brief qualitative description of the task and the displayed stimuli.

All this relevant information about the studies selected to be included in the final analyses is reported in Table 1. Moreover, to calculate the Cohen's *ds*, descriptive statistics of the ASD and TD groups' performances or the associated test statistics were extracted. If further groups were tested in the study, then the statistics relative to the other groups were not considered.

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6.4. Meta-Analytic Procedures

6.4.1. **Effect size calculation.**

For each observation, using the descriptive or test statistics present in the papers selected, we calculated Cohen's *d* as a measure of the size of the group effect (ASD vs. TD) on the variable of interest. Only two studies explicitly reported Cohen's *d*. For all other studies, we extracted Cohen's *d* from mean and standard deviations (or standard errors), or from the *t*-test value (or the associated *p*-value) of the reported statistical analyses. The mean and the standard deviation were available in most of the studies, therefore we used the following equation (1) to compute the Cohen's d[1] (see (Borenstein et al., 2009; Lakens, 2013):

$$d = \frac{M_1 - M_2}{\sqrt{\dfrac{(n_1 - 1)SD_1^2 + (n_2 - 1)SD_2^2}{n_1 + n_2 - 2}}}$$

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[1] For very small sample sizes (i.e., $n_1 < 10$ and $n_2 < 10$), Cohen's *d* overestimates both the effect size and its variance, therefore the use of Hedges' *g* is sometimes preferred over the use of Cohen's *d* as a measure of the effect size (see Lakens, 2013). However, as reported by Borenstein et al. 2009, p. 28), for sample sizes larger than 10 the difference between Cohen's *d* and Hedges' g is trivial, which led us to favor the use of the easily interpretable and very well-known Cohen's *d*.

where M_1 is the mean of the dependent variable for the first group (e.g., ASD), M_2 is the mean of the dependent variable for the second group (e.g., TD), SD_1 and SD_2 are respectively the standard deviations of the mean for the first and the second group, n_1 and n_2 refer to the number of participants in the first and the second group, respectively. We arbitrarily set the direction as positive indicating a better performance for the TD group as compared to the performance of the ASD group (e.g., lower error rate). According to the guidelines of Cohen (1988), an absolute effect size of 0.2-0.3 is considered a small effect, ~0.5 a medium effect and >0.8 a large effect. The variance associated to each Cohen's d was computed following equation (2) (Borenstein et al., 2009):

$$Var(d) = \frac{n_1 + n_2}{n_1 n_2} + \frac{d^2}{2(n_1 + n_2)}$$

In some cases, two or more Cohen's ds referring to the same study were included in the same meta-analysis (see below). When more than one task was included in the same (sub)group of the meta-analysis, we pooled those Cohen's ds in order to compensate for the possible underestimation of the variance of the mean effect size resulting from the inclusion of multiple effect sizes from a single study in the same meta-analysis (see Borenstein et al., 2009; Tanner-Smith, Tipton, & Polanin, 2016). The Cohen's ds referring to the same study are pooled using the formulas reported by (Borenstein et al., 2009, p. 230) (equations 3 and 4):

$$\bar{d} = \frac{1}{m}\left(\sum_{j}^{m} d_j\right)$$

$$Var(\bar{d}) = \left(\frac{1}{m}\right)^2 \left(\sum_{j}^{m} Var(d_j) + \sum_{j \neq k}^{m} \left(r_{jk}\sqrt{Var(d_j)}\sqrt{Var(d_k)}\right)\right)$$

where $\bar{d}$ and $Var(\bar{d})$ are the average Cohen's d and its variance, m is the number of Cohen's ds that were pooled, d_j and d_k are the Cohen's ds referring to measures j and k, and r_{jk} is the correlation between measures j and k, which was arbitrarily set to 0.5[2]. The maximum number of Cohen's ds that were pooled is 10. This number varies depending on the meta-analysis performed.

In each of the four meta-analyses (see below), multiple Cohen's ds from the same study were pooled when they referred to tasks that were coded in the same manner. For instance, in (Annaz et al., 2010) there are two tasks, one measuring the sensitivity (d') for the perception of BM in normal PLD vs. scrambled stimuli, and the other establishing the thresholds for the detection of a PLD walker in noise. Since both tasks were coded as *direct processing*, the two Cohen's d were always pooled together. This was also the case of the "action", "subjective state", and "emotion" conditions in Hubert et al. (2007), that were all coded as tasks of *instrumental processing*. In some cases, multiple Cohen's ds from the same study could be kept separate or pooled together depending on whether the categorical distinction between the tasks was relevant or irrelevant for the meta-analysis at hand. For instance, in Nackaerts et al. (2012) the Cohen's ds referring to the task categorized as *direct processing* and the task categorized as *instrumental processing* were pooled in the first general meta-analysis, but they were kept separate in the meta-analysis testing the effects of the level of BM processing.

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[2]Note that, when a large number of studies include a large number of multiple measures, the robust variance estimation method (Tanner-Smith et al., 2016) is sometimes used to obtain unbiased estimates of the average Cohen's d and of its variance. However, because in the current context the number of studies including multiple effect size measures was relatively small (i.e., 10/27 in meta-analysis 1), and because there were no more than three effect size measures for each of these studies (more often there were two), we deemed it preferable using a simpler random-effects model with appropriate methods for aggregating effect sizes from the same study (Borenstein et al., 2009).

6.4.2. **Mixed effects models.**

Statistical analyses were conducted in R, using the package *metafor* (Zarate and Manji, 2008). The meta-analysis performed was analyzed using a mixed effects model (Zarate and Manji, 2008), estimating the regression coefficients for each moderator included as well as the error between experiments. No within error estimate was included in the model, because we only included one Cohen's *d* per study. The rationale for this choice was that the majority of data from the studies presented a single effect size and a traditional random-effect model would suffice (Borenstein et al., 2009; see also Note 2). A simple model with one moderator variable is given in equation (5):

$$y_i = b_0 + b_1 x_i + e_i$$

for $i \in \{1,...,n\}$

y_i is the Cohen's *d* for the ith study

b_0 is the fixed intercept for the regression model

b_1 is the fixed slope for the regression model

x_i is the predictor for the ith study

$e_i \sim N(0, \sigma)$ is a Gaussian error term.

The variance components of the model including the predictor variables refer to the unexplained variance. To answer specific questions, we run four separate meta-analyses:

1. ***Biological motion in ASD vs. TD***. We included all the studies that were selected, and, for each of them, a single Cohen's *d* was inserted. Since the cognitive functioning was assessed with different tests across studies, as reported in Table 1, we computed a Cohen's *d* per study to have a comparable measure of the cognitive

level to be used as a possible moderator. A positive Cohen's *d* indicates that the cognitive level of the TD group is greater than the cognitive level of the ASD group. We included this moderator in the meta-analytic models as a way to assess if the ASD diagnosis has a specific effect on the processing of BM, which cannot simply be attributed to a reduction in cognitive level.

2. ***Level of processing of biological motion in ASD vs. TD***. We included one Cohen's *d* per study in each of the three levels of BM processing identified. 'Level of processing' was included in the meta-regression analysis as a categorical predictor, and no further moderating effects were considered here, due to the limited number of studies included in each group.

3. ***Impact of low-level features of BM scrambled stimuli in ASD vs. TD***. We included one Cohen's *d* for each study that used the scrambled version of the BM stimulus. We considered two types of scrambled stimuli according to the manipulation, spatial or temporal, employed to generate them. To this aim, this meta-analysis includes a categorical moderator, 'type of scrambled', which consists of two levels: 1) *spatial scrambled*, the difference between the scrambled and the BM stimulus is limited to a difference in their spatial features. In other words, in the scrambled version each dot of the PLD was randomly displaced in a different spatial position; 2) *temporal scrambled*, the difference between the scrambled and the BM stimulus is limited to a difference in their temporal features. In other words, each dot in the PLD scrambled stimulus was presented in the same spatial position but the motion of each single dot was temporally out of phase as compared to the BM stimulus. Studies with both spatial and temporal manipulations were excluded since they

were not numerous enough to create a stand-alone group (n=2). Considering the limited number of studies per level, we refrained from including other moderators.

4. ***Non-biological motion in ASD vs. TD***. This meta-analysis is limited to the studies that test, together with the perception of a BM stimulus, also the processing of another stimulus that can be defined as a non-BM stimulus (i.e., an object as a bike or a truck presented in PLD; (Hubert et al., 2007; Wright et al., 2014) or of a scrambled version of the BM stimulus (Herrington et al., 2007; Kröger et al., 2014; Murphy et al., 2009). It is important to note that we do not consider this analysis to be ultimately conclusive, because an extensive analysis of this type should include also other types of non-BM tasks such as coherent motion, form-from-motion, etc. However, we still believe that this aspect is worth of investigation to provide an initial characterization of possible dissociation between BM and non-BM processing in ASD vs. TD.

All tests were conducted with a significance level of 5%. The results will be presented including the following measures: weighted mean effect sizes (Cohen's *d*), 95% confidence intervals, I^2 heterogeneity values, and *p* values. All error bars in forest plots are 95% confidence intervals and were generated by applying customized R scripts.

6.4.3. **Moderator correlations.**

Being aware of the need to control for confounded moderators (Field & Gillet, 2010), we opted to use ANOVAs (performed in R with the *aov* function) in order to evaluate possible systematic relationships between the categorical predictor (i.e. 'level of processing', 'type of scrambled') and the quantitative predictor (i.e. 'cognitive level').

This procedure allowed to define whether the cognitive level differed systematically across the *levels of processing* and *type of scrambled*.

6.4.4. **Publication bias.**

To determine the presence of publication bias for the main effect of each meta-analysis, we plot each study's effect size (Cohen's *d* ASD vs. TD) against its own standard error in a funnel plot. Publication bias is computed as a negative association between the sample size and the effect size across studies, since to be significant a result coming from a small sample must have a large effect size. More specifically, publication bias occurs when studies with small samples are more likely to be published when they report statistically significant results (i.e., a large effect size), as compared to when they do not (i.e., a small effect size). The publication bias implies an overestimation of the true effect size, for which appropriate correction methods are available (see below).

7. **Results**

7.1. **Biological motion in ASD vs. TD.**

The analysis including all papers contrasting the processing of biological motion in ASD vs. TD revealed a medium effect of the difference in performance between the two groups, Cohen's $d = 0.57$, $SE = 0.14$, CI 95% = [0.30, 0.85], $z = 4.08$, $p < .0001$. In other words, there is a medium effect size indicating that individuals with ASD have more difficulties in BM processing than TD individuals (Figure 3). However, the I^2 coefficient reveals that 79% of the variability of the effect sizes can be attributed to the heterogeneity among the true effects (Test for Heterogeneity: $Q_{(25)} = 108.89$, $p < .0001$). Adding cognitive level to

the model as a moderator revealed no statistically significant influence on the effect sizes, $b = 0.23$, $SE = 0.24$, CI 95% = [-0.23, 0.70], $z = 0.98$, $p = .33$.

--------------------------------------Please insert here Figure 3--------------------------------------

A visual inspection of the funnel plot resulting from this meta-analysis suggests a quite symmetrical distribution of the studies. To confirm whether or not a publication bias is present in this case, we computed a trim and fill test (Taylor and Tweedie, 1998). Such test estimates the absence of three studies on the left side ($SE = 3.39$) and none on the right side. When computing a new average effect size including the effect sizes of the three studies estimated to be absent, the average effect size was still significantly different from zero, Cohen's $d = 0.45$, $SE = 0.15$, CI 95% = [0.16, 0.73], $z = 3.06$, $p = .002$ (Figure 4). Thus, since following the inclusion of the estimated missed studies, the effect size was still medium and significant, the difference in the processing of BM between ASD and TD groups cannot be considered dependent on publication bias.

--------------------------------------Please insert here Figure 4--------------------------------------

7.2. Level of processing of biological motion in ASD vs. TD.

To investigate the role of the level of processing needed to accurately perceive BM in ASD vs. TD, we computed a mixed effect model with 'level of processing' as a moderator. This moderator has three levels: i) first-order processing; ii) direct processing; iii) instrumental

processing, indicating the three subgroups in which tasks were divided. The I^2 coefficient indicates that the heterogeneity among the true effects is ~73% (Test for Residual Heterogeneity: $QE_{(28)}$ = 103.03, p <.0001). The results reveal a significant effect of the moderator, $QM_{(3)}$ = 29.61, p <.0001. Thus, to further explore each level of the moderator we calculated separate meta-analyses for each level of processing. When analyzing the *first-order processing* level, we obtained Cohen's d = 0.55, SE = 0.27, CI 95% = [0.02, 1.09], z = 2.04, p =.04, and the heterogeneity was comparable to that of the full model (I^2 = 79%, $Q_{(6)}$ = 27.04, p <.0001, Figure 5). When analyzing the *direct processing* level, we obtained Cohen's d = 0.40, SE = 0.18, CI 95% = [0.05, 0.74], z = 2.24, p =.03, and the heterogeneity was comparable to that of the full model (I^2 = 76%, $Q_{(15)}$ = 61.29, p <.0001, Figure 6). Instead, the *instrumental processing* level is the one showing the greatest difference in the processing of BM between ASD and TD, Cohen's d = 1.04, SE = 0.19, CI 95% = [0.68, 1.40], z = 5.62, p <.0001, and the heterogeneity was reduced as compared to that of the full model (I^2 = 52%, $Q_{(7)}$ = 14.70, p =.04, Figure 7).

Moreover, in order to test if there is a significant difference between the magnitude of these estimated effect sizes, we performed three z-tests between the Cohen's ds of the different levels of processing using the Bonferroni correction method. Only the difference between *direct* and *instrumental* level was statistically significant (direct vs. instrumental: z = -2.53, p =.03; first-order vs direct: z= -0.46, p ~ 1; first-order vs. instrumental: z = -1.48, p =.42). To evaluate possible systematic differences in cognitive level across the levels of BM processing, we performed a one-way ANOVA on the Cohen's ds referring to the cognitive level with 'level of processing' as a factor. No statistically significant differences were

retrieved, $F_{(1, 24)} = 2.11$, $p =.16$, $\eta_p^2 = 0.89$, showing that the differences in level of processing retrieved are not confounded by cognitive level.

------------------------------------Please insert here Figure 5------------------------------------

------------------------------------Please insert here Figure 6------------------------------------

------------------------------------Please insert here Figure 7------------------------------------

7.3. Impact of low-level features of BM scrambled stimuli in ASD vs. TD

With this third meta-analysis, we aim at further characterizing the low-level perceptual aspects potentially affecting the processing of BM in ASD vs. TD. Here, we focus only on studies that include both BM stimuli and non-BM scrambled stimuli. In other words, with this meta-analysis we inquire whether there are specific spatial or temporal features of the stimuli in motion that affect ASD and TD individuals differently.

The I^2 coefficient indicates that the heterogeneity among the true effects is ~58% (Test for Residual Heterogeneity: $QE_{(11)} = 25.95$, $p =.007$). Interestingly, results indicate that the moderator is significant, $QM_{(2)} = 12.02$, $p =.003$. To further investigate each level of the moderator we calculated separate meta-analyses for each moderator level. When considering only the studies with a BM task that includes a spatially scrambled stimulus, results indicated no significant difference in the performance of ASD vs. TD (Cohen's $d = 0.38$, $SE = 0.23$, CI 95% $= [-0.07, 0.83]$, $z = 1.67$, $p =.10$, Figure 8). The confidence interval includes 0, suggesting that we cannot exclude that the effect size for this subgroup of

studies is null, and the heterogeneity is still moderate and significantly different from zero ($I^2 = 60\%$, $Q_{(6)} = 14.84$, $p =.02$).

On the contrary, the model including the temporally scrambled stimulus showed a significant medium-large effect size (Cohen's $d = 0.67$, $SE = 0.21$, CI 95% = [0.26, 1.08], $z = 3.20$, $p =.001$, Figure 9), and the included studies can be considered a homogenous group ($I^2 = 55\%$, $Q_{(5)} = 11.11$, p =.05). Differences in the processing of scrambled stimuli found are not confounded by the cognitive level, as it was shown by a one-way ANOVA on Cohen's ds referring to cognitive level with 'type of scramble' as a factor, $F_{(1, 9)} = 0.009$, $p =.93$, $\eta_p^2 = 0.002$.

----------------------------------Please insert here Figure 8-------------------------------------

----------------------------------Please insert here Figure 9-------------------------------------

7.4. Non-biological motion in ASD vs. TD.

From the included papers, we are here selecting only the studies including tasks that contrast BM stimuli with non-BM ones, to test whether the deficit is selective to BM stimuli or generalizes also to non-BM stimuli. Results indicate that no significant difference between the performance of individuals with ASD and TD is evident, (Cohen's $d = 0.26$, $SE = 0.29$, CI 95% = [-0.31, 0.83], $z = 0.89$, $p =.38$, Figure 10). The total heterogeneity estimated with I^2 is ~75% ($Q_{(4)} = 19.80$, $p =.0005$).

-------------------------------------Please insert here Figure 10-----------------------------------

-------------------------------------Please insert here Table 2-----------------------------------

8. Discussion

In this article, we conducted the first meta-analytic investigation examining the abilities of individuals with ASD in the processing of BM stimuli. The results of the general meta-analysis that included all the studies that matched our selection criteria showed a moderate ($d = 0.57$) deficit for individuals with ASD as compared to TD individuals (Table 2). This deficit was not influenced by the cognitive level of participants and was not due to a publication bias. Nonetheless, we observed a high heterogeneity among studies ($I^2 = 79\%$), and this is an interesting aspect because it suggests that possible sources of variability are present in the group of studies included. We believe this represents a confirmation of the idea that the perception of BM is a complex ability, which cannot always be reduced to a monolithic process. Indeed, historically the BM literature has been characterized by remarkably heterogeneous experimental protocols. This can create ambiguity in transferring in operational terms the construct under investigation, which in turn can undermine reproducibility of studies, as previously proposed for example for the investigation of the "theory of mind" construct (Schaafsma et al., 2015). Thus, proven that the general meta-analysis points toward a significant impairment in BM perception in participants with ASD, we further developed our meta-analytic investigation in the direction of testing the possible factors that contribute to the high heterogeneity observed. Among these factors, we considered the distinct levels of processing defined in our three-level model (i.e., *first-order*, *direct* and *instrumental* BM processing), the type of manipulation used to create the comparison scrambled stimulus (spatial vs. temporal), and the selectivity of the impairment for biological as opposed to non-biological motion.

8.1.Is biological motion perception in ASD anomalous across all levels of processing?

A more precise analysis testing BM processing in ASD according to a three-level model that distinguished among *first-order*, *direct* and *instrumental* processing of BM confirmed that such a deficit was present independently of the specific level of information processing required in the BM task. Nonetheless, this moderator revealed a significant effect on the observed effect size of the impairment in BM processing. Individuals with ASD showed the largest deficit as compared to TD participants when the experimental task required the *instrumental* recognition of the BM ($d = 1.04$), while this deficit was moderate ($d = 0.55$) for tasks requiring the *first-order* processing of BM and only small-to-moderate ($d = 0.40$) when testing task requiring the *direct* recognition of BM. It is worth noting that the category of first-order BM usually includes samples of younger children which performed preferential looking tasks. Moreover, for this category the deficit is moderate, but there is only a single study (Wright et al., 2016) in which we observed a negative effect size. This study however tested a preferential looking paradigm in children at the age of 6, and thus deviates from the typical preferential looking experiment that usually includes younger children or infants. Nonetheless, post-hoc tests revealed that *direct* and *instrumental* processing were significantly different in terms of effect size. The differences in the observed deficit as a function of the levels of BM processing were not ascribable to differences in cognitive level across the three levels above. The study of BM anomalies in ASD has so far been accumulating a conspicuous body of evidence, but often without putting the necessary effort in trying to understand the different sub-components required to efficiently process the distinct types of BM tasks. In this meta-analytic investigation, we

made a first attempt in this direction by distinguishing three levels in the BM processing. By applying this model, we were able to selectively identify which are potential reasons driving such differences in processing, which in turn can be used to inform diagnosis and treatment. The most severe deficit in participants with ASD emerges when the recognition and categorization of a BM stimulus is serving a secondary purpose, such as when inferring someone else's intentions, actions or emotional states; in other words, when correctly perceiving BM is a pre-requisite to disclose distinct and possibly more complex information. On the contrary, the least severe deficit is observable in participants with ASD when a more basic recognition of BM is tested, the level that we referred to as *direct* recognition of BM (e.g., BM vs. non-BM/scrambled stimulus; BM stimulus "A" vs. BM stimulus "B", leftward vs. rightward moving stimulus, etc.). This less severe deficit could be due, at least in part, to compensatory mechanisms that can help participants with ASD to correctly process BM when the task is not excessively complex. This is in line also with the age range of participants tested, who are school age children, adolescents and adults. Overall, these results stress the need for a careful evaluation of what is at the core of BM processing and what are instead additional, secondary and higher-level information that can be extracted from a PLD depicting a stimulus moving in a biologically plausible fashion. Persevering with such an ambiguity would undermine the study of this construct as a putative endophenotype for ASD with negative clinical repercussions.

8.2. What are the lower-level perceptual features linked to biological motion anomalies in ASD?

After assessing the role of different levels of BM processing in influencing the strength of the deficit that characterizes individuals with ASD, we assessed another fundamental aspect of BM that is the role of lower-level perceptual features. Indeed, in many studies, BM stimuli are compared to scrambled stimuli obtained by varying the spatial or temporal properties of the PLD. Experiments using a *spatial* scrambled make use of comparison stimuli that are characterized by the displacement of all or part of the dots constituting the target PLD in a different spatial position but where each dot maintains the original temporal phase. Experiments using a *temporal* scrambled, instead, make use of comparison stimuli where each dot composing the target PLD remains in the same relative spatial position within the original configuration, but the motion of each of these dots is shown temporally out of phase. By testing the role of the type of scrambled stimuli used as a comparison, our aim was to inquire whether the deficits observable in individuals with ASD are dependent on the specific spatial or temporal properties of the target stimulus that are experimentally manipulated. We found that such a moderator had a significant influence in the effect size of the impaired BM processing in ASD. Studies employing a spatial scrambled as a comparison stimulus revealed no significant difference in the performance between individuals with ASD and TD controls. On the contrary, a significant impaired performance was evident for studies employing a temporal scrambled as a comparison stimulus, with a moderate-to-large effect size ($d = 0.67$). Although the sample size for each of this meta-analysis was limited (N=7 spatial; N=6 temporal), they furnish an initial interesting insight on the deficit underlying the perception of BM in ASD, which, we believe, deserves further investigation in future studies.

It is increasingly evident that simple spatio-temporal visual dynamics are essential building blocks for social behavior. A recent study in juvenile zebrafish showed that collective behavior, like shoaling, can be driven by simple specific motion cues that are constituted by black dots that mimic the precise zebrafish motion, while other naturalistic cues such as fish-like shape, pigmentation pattern, or non-visual sensory modalities are not required (Larsch and Baier, 2018). Anomalies in the processing stages connected to the tracking of low-level motion dynamics might be disruptive for detection, recognition and categorization of BM stimuli. Our results suggest that a candidate core mechanism playing a role in BM processing anomalies in ASD is linked to the temporal properties of moving dots embedded in a PLD. This is consistent with several experimental pieces of evidence and theoretical proposals suggesting that one of the core anomaly of individuals with ASD is related to temporal binding of uni- and -multi- sensory information (Brock et al., 2002; Feldman et al., 2018; Wallace and Stevenson, 2014; Zhou et al., 2018), with accumulating evidence showing that this would be a common endophenotype also for other neurodevelopmental disabilities (Wallace and Stevenson, 2014; Zhou et al., 2018). Specifically, individuals with ASD have repeatedly shown an extended temporal binding window (TBW; an epoch of time within which stimuli from different sensory modalities are highly likely to be bound) which would bring to an imprecise temporal processing of sensory stimuli (Zhou et al., 2018). The majority of this evidence comes from audio-visual multisensory paradigms (for reviews see, Feldman et al., 2018; Wallace & Stevenson, 2014; Zhou et al., 2018), where researchers assess the temporal precision needed to perceive auditory and visual stimuli as synchronous. Nonetheless, initial evidence for temporal processing anomalies in unisensory perception has been documented in

individuals with ASD in both uni-sensory auditory (Foss-Feig et al., 2017; Leslie D. Kwakye et al., 2011) and visual (Parks, 1965; Thompson et al., 2015) perception, suggesting a common underlying deficit in the fundamental coordination of brain networks that are responsible for the timing of sensory information processing (Burr and Morrone, 2006). It is also worth noting that the extension of TBWs has a high degree of malleability and they can shrink or expand as a function of the complexity and the low-level statistics of the incoming sensory information. For example, TBWs change as a function of spatial and temporal relationship among stimuli, such that for example information coming from different sensory modalities are integrated across a larger temporal window when they originate from the same spatial position (Zampini et al., 2005) or when processing predictable/rhythmic stimuli as compared to irregular ones (Denison et al., 2012). TBWs can shrink or expand also as a result of synchronization of brain rhythms (i.e. entrainment) (Ronconi et al., 2018a; Ronconi and Melcher, 2017), or after few days of perceptual learning trainings (Powers et al., 2012, 2009; Stevenson et al., 2013). Intriguingly, the hypothesis of a temporal processing anomaly as one of the key aspects responsible for BM deficits observed in ASD would be consistent with the involvement of the cerebellum and its bidirectional connectivity with the right posterior superior temporal sulcus (STS), which represent an important part of the network involved in the processing of BM (Jack et al., 2017; Sokolov et al., 2018, 2014, 2012, 2010) as we will discuss in the next paragraph.

8.3. Is the impairment observed in ASD specific to biological motion stimuli?

We conducted a meta-analysis of the subgroup of studies that tested performance in BM vs. non-BM tasks. The aim of this analysis was to assess whether the deficit of participants

with ASD was selective for BM stimuli or generalized also to non-BM stimuli. We observed no evidence of impairments in the processing of non-BM stimuli in individuals with ASD; this suggests that the deficit is selective for BM stimuli and to their specific spatio-temporal pattern, whereas other types of complex stimuli in motion do not elicit a similar impairment. Among the non-BM tasks tested, we can find recognition/categorization of an object as a bike or a truck presented in PLD (Hubert et al., 2007; Wright et al., 2014) but also recognition/categorization of a scrambled version of the BM stimulus (Herrington et al., 2007; Kröger et al., 2014; Murphy et al., 2009). It is clear however that testing the scrambled version of BM stimuli is not ideal to claim that the deficit observed in individuals with ASD is specific to stimuli moving in a biological fashion. Consistently, there is evidence showing that when neurotypical adults cannot extract from a non-BM stimuli a form, non-BM stimuli themselves are more difficult to discriminate relative to BM stimuli. Specifically, these studies show that the presence of motion stimuli conveying specific forms nullifies the advantage in processing BM vs. non-BM stimuli. In other words, the presence of a form that can be extracted from non-BM stimuli seems to be a crucial aspect when comparing the recognition/categorization of BM and non-BM stimuli (Hiris, 2007). This would mean that BM is not generally easier to detect than equally-structured non-BM. To claim that individuals with ASD show a deficit which is specific to BM, the processing of non-BM stimuli should involve a similar extraction and interpretation of a form-from-motion. In the present meta-analysis there were only two studies that tested non-BM processing using tasks where the form was preserved (Hubert et al., 2007; Wright et al., 2014). These two studies, however, found no

clear evidence of impairments in participants with ASD, independently of the nature of the moving stimuli (BM vs. non-BM) employed.

We conclude that from the evidence available at the time of the present meta-analysis, whether anomalies observed in individuals with ASD are specific to BM is still an open question. Thus, although interesting, the comparison between the processing of BM and non-BM stimuli needs a more careful investigation in future studies.

8.4. Neural and oscillatory correlates of BM processing in individuals with ASD

Looking at studies that investigate the neural correlates of BM processing in primates, neurotypical human, and in the ASD population, it seems to be consistent with the idea that BM is hardly conceivable as a monolithic process, given that the network of brain areas that are responsible for BM processing, although not yet completely understood, is highly complex and probably not unitary.

Neuroimaging studies in the typical population show that the activity of the posterior STS (pSTS) is specifically associated to BM stimuli presented as PLD. The pSTS involvement is often found to be lateralized in the right hemisphere (Brancucci et al., 2009; Grosbras et al., 2012; Grossman et al., 2000; Pelphrey et al., 2003). While this region is activated for the observation of a wide range of human movements (Grosbras & Paus, 2006; Thompson, Hardee, Panayiotou, Crewther, & Puce, 2007), it has not been revealed to be involved in studies examining the perception of static bodies or body parts, suggesting that pSTS is involved specifically in extracting human body motion features rather than forms (Grosbras et al., 2012). Congruently, studies in non-human primates have shown the presence of neurons that selectively respond to various biological actions (e.g. walking, turning of the

head, bending of the torso, moving of the arms, etc.; Jellema & Perrett, 2003). However, these neurons receive convergent information from the dorsal visual stream areas MT and MST, and also from the ventral stream area in the inferior temporal (IT) regions (Grosbras et al., 2012). The homologue of pSTS in the primate brain is also connected with the inferior parietal cortex (Seltzer and Pandya, 1994). A recent meta-analytic investigation of human brain regions activated in different types of BM processing (i.e. whole body, hand and face movements) confirmed common convergent activations independent of the specific motion category in the right pSTS region. However, the pattern of activations included also bilateral regions at the junction between middle temporal and lateral occipital gyri (Grosbras et al., 2012).

Other evidence in favor of a highly complex network of brain regions supporting the processing of BM that goes well beyond pSTS comes from functional connectivity studies. For example, a study by Sokolov et al. (2012) that used fMRI to test healthy participants during a BM task, in combination with functional connectivity analysis and dynamic causal modeling, showed that the left lateral cerebellum (lobules Crus I and VIIB) exhibited increased activation and connectivity with right pSTS during the task. More recently, Sokolov et al. (2018) used novel dynamic causal modeling informed by probabilistic tractography to highlight that the network for BM processing is organized in a parallel rather than hierarchical manner. Specifically, the authors showed that although the right pSTS serves as an integrator within temporal areas, this area does not appear to be the main gate in the functional integration of the activity of the occipito-temporal and frontal regions. Indeed, both fusiform gyrus and middle temporal areas were also connected to the right inferior frontal gyrus and insula during processing of BM stimuli, indicating multiple

parallel pathways. BM-specific loops of effective connectivity were found also between the left lateral cerebellar lobule Crus I and right pSTS, as well as between the left Crus I and right insula. According to their conclusions, this parallel processing that is dependent on the activity of different cortico-cortical and cortico-cerebellar pathways may help to explain why BM processing is rather resilient to focal brain damage but is severely impaired in neuropsychiatric conditions with distributed network alterations, such as in ASD.

In addition, it has been suggested that the cerebellum uses a common computational algorithm not only upon motor but also non-motor functions, such as perceptual ones (D'Angelo and Casali, 2012; Knolle et al., 2013). The idea that the cerebellum may play a critical role as internal "timing device" for motor and non-motor functions has been widely explored in the literature (Bareš et al., 2018). Notably, this idea may fit with the evidence indicating that the cerebellum is implicated in the pathophysiology of ASD (Fatemi et al., 2012; Wang et al., 2014) and may also fit with our present results showing a specific role for temporal scrambled stimuli in explaining the anomalies in BM processing observed in individuals with ASD.

Although it is not within the scope of the present article to provide an extensive review of the neuroimaging literature that investigates the neural correlates of impaired BM processing in ASD, it is interesting to notice that consistent report in the literature show that during the processing of BM stimuli individuals with ASD show hypoactivation in the pSTS (Ahmed and Vander Wyk, 2013; Freitag et al., 2008; Herrington et al., 2007; Kaiser et al., 2010) and in regions of the prefrontal cortex (Kaiser et al., 2010; Koldewyn et al., 2011). However, only few studies so far have tested the effective connectivity among the

key regions of the cortical and subcortical (i.e. cerebellum) network responsible of the processing of BM (Alaerts et al., 2017; Jack et al., 2017; McKay et al., 2012). These studies suggest that a defective temporo-parietal (McKay et al., 2012) and cerebellar-pSTS connectivity (Jack et al., 2017) could contribute to the anomalies in the processing of BM reported in ASD and reviewed in the present meta-analysis. Although these studies provide initial important evidence, the results of the present meta-analysis suggest that distinct levels of BM processing can be differentially affected in participants with ASD, and that the manipulation of low-level perceptual features may result in markedly distinctive behavioral outcomes. These aspects should be taken into account in future neuroimaging studies that will address the neural correlates of the anomalies in the processing of BM in ASD. More generally, even from a neural perspective, BM should not be considered as a monolithic process, but as constituted by different operations, that are likely supported by different parallel neural routes.

9. Conclusions

Taken together, this work represents to our knowledge the first meta-analytic investigation that assesses in a detailed manner the putative impairment in the processing of BM in individuals with ASD. Despite some of our analyses are probably influenced by the limited sample of studies available and by the high heterogeneity of participants' characteristics, we made an effort toward interpreting the multifaceted aspects of the vast literature that use BM as a proxy for social perception deficits in ASD. While on the one hand a significant impairment in BM processing emerged in individuals with ASD, on the other hand the idea that this deficit is selective for BM and does not extend to non-BM stimuli is, at present, not properly supported by the literature. Moreover, we observed a high heterogeneity due to the different experimental protocols employed. A major source of heterogeneity can be individuated in the distinct levels (*first-order, direct, instrumental*) of BM processing, given that the most severe deficit in participants with ASD emerges when recognition and categorization of a BM stimulus is serving a secondary purpose (e.g. inferring other's intentionality, action or emotional state). How more severe difficulties reported in the *instrumental* recognition of BM may represent just a byproduct of more basic anomalies should be clarified in future studies. Finally, an initial result points toward the importance of lower-level temporal perceptual features in determining the deficit in the perception of BM in ASD.

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Authors' contribution

A.F., V.P., L.C. and L.Ro. have conceived the idea of the paper. A.F. and L.Ra. executed the literature search, coded the papers and calculated the effect sizes, independently. A.F. and M.V. defined the analytical protocol and performed the analyses. A.F., V.P., L.C. and L.Ro. drafted the article, and all authors critically revised it. All authors approved the final version of this article.

Figure 1. Number of articles published per year as indexed in PubMed with the general search using the keywords "biological motion autism".

Figure 2. Flow diagram of the selection of studies included in the current meta-analysis. The number of studies included or excluded at each step of the evaluation process is indicated in each of the diagram boxes.

Figure 3. Forest plot of the differences in the BM performance between the ASD and TD groups. Positive values indicate a better performance for the TD group as compared to the performance of the ASD group.

Figure 4*. Funnel plot indicating the publication bias in studies of BM processing in ASD vs. TD. White circles represent the missing studies identified by the trim and fill test.*

Figure 5. Forest plot of studies investigating the difference in first-order level of BM processing between ASD and TD group (i.e. positive values indicate a better performance for the TD group as compared to the performance of the ASD group).

Figure 6. *Forest plot of studies investigating the difference in direct level of BM processing between ASD and TD group (i.e. positive values indicate a better performance for the TD group as compared to the performance of the ASD group).*

Instrumental processing of BM

Figure 7. Forest plot of studies investigating the difference in instrumental level of BM processing between ASD and TD group (i.e. positive values indicate a better performance for the TD group as compared to the performance of the ASD group).

Spatial scrambled

Figure 8. Forest plot of studies investigating the difference in BM performance between ASD and TD group that employed spatial scrambled stimuli (i.e. positive values indicate a better performance for the TD group as compared to the performance of the ASD group).

Figure 9. *Forest plot of studies investigating the difference in BM performance between ASD and TD group that employed temporal scrambled stimuli (i.e. positive values indicate a better performance for the TD group as compared to the performance of the ASD group).*

Figure 10. Forest plot of studies investigating the difference in the processing of non-BM between ASD and TD group (i.e. positive values indicate a better performance for the TD group as compared to the performance of the ASD group).

Table 1. The table contains the final set of studies included in the meta-analysis. For each study descriptive statistics are reported for both ASD and TD groups (since different tests were used in order to assess the participants' cognitive level, in the subcolumn 'Test' the name of the test is specified). When the data were not availaible in the paper N.A.is reported. The column 'Meta-analysis and related categoization' contains four subcolumns where we reported in which meta-analysis each study was included (1=All BM studies, 2=Level of processing, 3=Low-level features, 4=Non-BM), and the relative categorization: Level of processing(First-order, Direct, Instrumental), Type of scrambled (Spatial, Temporal, Spatial and Temporal) and Type of stimuli(BM=only biological motion stimuli, non-BM=only non-Biological motion stimuli, BM vs. non-BM when a biological stimulus has to be discriminated from a non-biological one). In the 'Task characteristics' column we reported which behavioral measure was assessed in the task (Accuracy, % Looking), if the task was performed during an EEG or a fMRI acquisition, and a brief description of the implemented task.

Table 2. *Each row reports the main result for each of the four meta-analyses performed. Which meta-analysis, the subgroup, and the computed Cohen's d are reported in the first, second and third column, respectively. The number of * represents the level of significance: *** p<.001; ** p<.01; * p<.05.*

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STUDY	PARTICIPANTS		% MALE		AGE Mean±SD		COGNITIVE LEVEL Mean±SD			META-ANALYSIS AND RELATED CATEGORIZATION				TASK CHARACTERISTICS		
	ASD	TD	ASD	TD	ASD	TD	Test	ASD	TD	Meta-analysis performed (1=All BM studies, 2=Level of processing, 3=Low-level features, 4=Non-BM)	Level of Processing	Type of scrambled	Type of stimuli	Behavioral measure	Neuro acquisition	Qualitative description of task
Blake et al., 2003	12	9	100%	100%	range 8-10	8.42±1.89	Kaufman Assessment Battery for Children, or K-ABC	7.9±1.8	N.A.	1, 2, 3	Direct	Temporal	BM vs. non-BM	Accuracy	/	Twenty-five human actions (e.g., running, kicking, climbing, throwing, jumping) and their scrambled version were presented. Each participant verbally reported whether the sequence was "a person" or "not a person."
Herrington et al., 2007	10	10	100%	100%	27.60±7.10	25.6±4.9	Wechsler Abbreviated Scale of Intelligence	109±N.A.	119±N.A.	1, 2, 3	Direct	Spatial	BM	Accuracy	fMRI	Participants indicated with pressing the right or left botton whether a PLD human image or its scrambled version is walking either to the left or to the right side of the screen.
										4	/	Spatial	non-BM	Accuracy	fMRI	
Hubert et al., 2007	19	19	89%	89%	21.5±6.08	24.3±8.5	Wechsler Scale (WAIS or WISC)	83.3±15.9	N.A	1, 2	Instrumental	/	BM	Accuracy	/	Participants were asked to describe orally what is happening in a series of movies that shows 10 common actions , 5 subjective states, 5 emotional states, and 5 objects.
										4	/	/	non-BM	Accuracy	/	
Freitag et al., 2008	13	15	87%	87%	17.6±3.6	18.6±1.2	Wechsler scale (riportati anche valori di IQ verbal e performance separatamente)	101.2±21.1	112.1±N.A.	1, 2	Direct	Spatial and Temporal	BM vs. non-BM	Accuracy	/	Participants were asked to press different buttons if they see either a walking person (40 famele and 40 male) or a scrambled version.
Atkinson, 2009	13	16	92%	88%	30.9±13.8	26.7±12.8	The ASD Wechsler Scale – Third Edition.IQ for the TD individuals were derived using the revised version of the National Adult Reading Test, which has good convergent validity withWAIS IQ scores	106.2±12.2	106.6±8.5	1, 2	Instrumental	/	BM	Accuracy	/	Participants were asked to indicate which emotion label, among 5 emotions (anger, disgust, fear, happiness and sadness), or which action label, among 8 human actions (digging, kicking, knocking, pushing, bending to touch toes, hopping,walking on the spot, and star-jumping or jumping jacks), best described the displayed movie clip.
Klin et al., 2009	21	39	100%	######	2.21±0.54	1.99±0.66	Visual Reception subtest of the Mullen Scales of Early Learning	2±0.95	2.1±0.91	1, 2	First-order	/	BM vs. non-BM	% Looking	/	In a preferential looking paradigm an animations consisted of children's games (e.g., playing "peek-a-boo" or "pat-a-cake") was presented in half of the screen. On the other half of the screen, the same animation was presented inverted (shown upside-down) and played in reverse order. Only the one (forward) audio soundtrack was presented.
Murphy et al., 2009	16	16	81%	81%	25.56±7.67	26.40±2.8	Raven	43.73±9.64	56.21±N.A.	1, 2	Direct	Spatial and Temporal	BM	Accuracy	/	A human walker or its scrambled version was displayed in a noisy dots background. On each trial participants were asked to respond as quickly as possible with two differnet buttons whether the stimulus was moving to the left or the right.
										4	/	Spatial and Temporal	non-BM	Accuracy	/	
Annaz et al., 2010	23	34	100%	100%	8.83±1.83	8.25±2.25	British Picture Vocabulary Scale Pattern II	7.17±1.67	8.42±2.25	1, 2, 3	Direct	Temporal	BM vs. non-BM	Accuracy	/	In the first task a human action (i.e., walking, running, throwing, kicking and star-jumping) or its scrambled version was displayed; the participant has to identify if the stimulus was moving like a person or not. In the second task a walker human or its scrambled version were presented side-by-side in a noisy dots background; the participant has to select the panel (left or right) where the dots look like a person.
							Construction subtest from the british ability scale II	8.92±3.08	8.25±2.08							
Koldewyn et al., 2010	25	32	92%	94%	15.12±2.64	15.78±2.41	Wechsler Scale (WASI)	107.8±16	121.3±2.41	1, 2	Direct	/	BM	Accuracy	/	Participants were required to report in which direction the figure was walking (left or right). The walker was embaded in a mask that could have different levels of coherence.
Saygin et al., 2010	16	20	81%	65%	33.75±12.7	37.75±11.35	IQ	112.19±16.25	113.16±12.35	1, 2	Direct	/	BM	Accuracy	/	Participants were required to report in which direction the figure was walking (left or right). The walker was embaded in noise dots, each of them with the same trajectory of one of the dots in the walker.
Annaz et al., 2012	17	17	100%	100%	5.5±1.08	5.58±1.08	British Picture Vocabulary Scale Pattern II	5.17±1.08	6.58±1.17	1, 2	First-order	Temporal	BM vs. non-BM	% Looking	/	In a preferential looking paradigm a human walking and its scrambled version were displayed side by side for 6 sec.

Study	N1	N2	%1	%2	Age1	Age2	Test	IQ1	IQ2	Exp	Task	Domain	Condition	Measure	Method	Description
							Construction subtest from the british ability scale II	6.33±1.25	6.41±1.08	1	First-order	/	BM vs. non-BM	% Looking	/	In a preferential looking paradigm a human walking and a moving object (spinning top) were displayed side by side for 6 sec.
McKay et al., 2012	10	10	100%	100%	28.6±6.92	27.9±7.37	Wechsler Scale (WASI)	125±7.01	124.8±6.7	1, 2, 3	Direct	Spatial	M vs. non-B	Accuracy	/	A walker or its scrambled version was presented embaded in a number of noise dots. Participants were required to report in which direction the stimulus is moving.
Nackaerts et al., 2012	12	12	58%	58%	34.9±8.5	31.5±6.3	Wechsler Scale (Ward 7- subtest short-form)	111.5±18.2	115.5±15.1	1, 2, 3	Direct	Spatial	BM vs. non-BM	Accuracy	/	A human action (walking, jumping on the spot or kicking) with an emotional states (neutral, happy, sad or angry) or a scrambled version was displayed and the participant had to indicate as fast and accurate as possible whether the stimulus was ''a person'' or ''not a person''.
										1, 2	Instrumental	/	BM	Accuracy	/	A prime and then a target stimulus were shown. Participant had to judge if the target (a human walking, jumping on the spot or kicking who could be happy, sad, angry or neutral) showed the same emotion of the prime stimulus.
Price et al., 2012	14	16	100%	87.5%	14.14±4.8	14.08±4.61	Wechsler Scale (WAS)	60.14±9.44	57.81±9.29	1, 2	Direct	/	BM	Accuracy	/	The stimuli were 28 typical gaits half coherent - gait of one person, half incoherent leg movements from two different walkers and 27 depicting gaits with limps.Participants made two judgments of each PLD walker: judging whether the stimulus depicted a coherent or incoherent gait, and whether each walker depicted a normal gait or a limp.
Rutherford & Troje 2012	13	14	100%	100%	29±6	31±9	Wechsler Scale (WAIS)	98.1±13.61	94.6±10.03	1, 2, 3	Direct	Spatial	BM vs. non-BM	Accuracy	/	The stimuli represent human, cat or pigeon walking (upright or inverted) or their scrambled version, the stimulus was embedded into the scrambled-walker mask. Two tasks were performed: in the first, the subject had to select if the coherent walker was displayed in the first or in the second stimulus; in the second, the subjcet should judge the directions of the stimulus.
Falck-Ytter et al., 2013	10	14	80%	79%	3.42±0.44 (Visual 1.89±0.64; Language 1.44±0.63)	3.55±0.43 (Visual 3.69±0.93; Language 3.55±0.6)	Visual / Language			1, 2	First-order	/	M vs. non-B	% Looking	/	In a preferential looking paradigm a video of a human action (e.g., clapping hands) and its up-side down version were shown side by side on the screen.
Alaerts et al., 2014 - 2017	15	15	100%	100%	21.7±4	23.3±2.9	Wechsler Scale	107.9±13.9	114.8±12.8	1, 2	Instrumental	/	BM	Accuracy	fMRI	Participants had to indicates if a human (walking, jumping or kicking) is happier, sadder, angrier or not different from the previous one; the first stimulus showed always a neural emotional state.
	15	14 (one removed from the analysis)								1, 2, 3	Direct	Spatial	BM vs. non-BM	Accuracy	fMRI	Participants were presented an intact BM (walking, jumping or kicking in four different 'emotional states' neutral, happy, sad, angry) or the scrambled version. They have to indicate as fast and accurate as possible whether the presented PLD represented 'a person' or 'not a person'.
Kröger et al., 2014	17	21	100%	100%	11.9±2.2	11.63±2.4	Raven	77.6±30.2	59.1±38.2	1, 2	Direct	Spatial	BM	Accuracy	EEG	30 different walkers and 30 scrambled motion stimuli were shown. Participants had to indicate if they have seen a walker or just a motion pattern.
										4	/	Spatial	non-BM	Accuracy	EEG	
Wright et al., 2014	19	21	84%	62%	6.96±1.5	6.49±1.77	Non verbal mental age	103.14±14.59	7.3±2.84	1, 2	Direct	/	BM	Accuracy	/	After the presentation of the moving stimulus (i.e., a human, a cat, a truck or a bicycle), the children have to verbally answer at the question "What is that?"
										4	/	/	non-BM	Accuracy	/	
Cusack et al., 2015	15	15	100%	100%	16.09±2.24	15.54±2.15	Wechsler Abbreviated Scale	103.14±11.59	104.79±9.14	1, 2	Instrumental	/	BM vs. non-BM	Accuracy	/	Participant have to discriminate if two agents were fighting or dancing in a normal or upside-down condition
										1, 2, 3	Direct	Temporal	BM vs. non-BM	Accuracy	/	1.Participants were asked to discriminate between a BM and its scrambled version. 2. Participants were asked to select the BM, in which limbs were intact vs. its version with scrambled limbs. Both tasks have an upright and inverted conditions.
Wang et al., 2015	21	21	81%	76%	67.05±15.21	59.38±16.99	verbal score Language test	28.05±30.27	60.11±24.83	1, 2, 3	First-order	Spatial	M vs. non-B	% Looking	/	In a preferential looking paradigm each trial present a human or a pigion walking and its scrambled version side by side for 10 sec.
										1, 2	Instrumental	/	BM	Accuracy	/	In this task the subject saw a short movie where one of 12 different human actions (e.g., standing up, sitting down, walking, running) was presented. Participants had to verbal respond as soon as they have indentifed the action displayed.
Fujioka et al., 2016	21	35	100%	100%	27.6±7.7	25.2±4.5	Wechsler Scale (WAIS-WISC)	99.8±13.5	N.A.	1, 2	First-order	/	BM vs. non-BM	% Looking	/	In a preferential looking paradigm a movie presents upright and inverted BM simultaneously side by side for 20 sec.

Study																Description
von der Lühe et al., 2016	16	16	75%	62.5%	41.56±9.15	36.19±12.11	WST (Wortschatztest, German multiple-choice vocabulary test)	16.88±15.5	115.31±8.43	1, 2, 3	Direct	Temporal	M vs. non-B	Accuracy	/	The stimlus was composed by two intervals both containing an agent A, and only one of them containing the agent B; the other interval contained a scrambled version of the agent B. Agent B or its scrambled version was masked by dots. Participants had to decided which interval contains the agent B.
										1, 2	Instrumental	/	BM	Accuracy	/	Participants performed an explicit intention recognition task on 21 videos depicting two PLD agents (14 comunicative and 7 indipendent actions).
Wright et al., 2016	18	18	83%	83%	6.61±1.29	6.41±1.75	verbal IQ	102±19.3	105.22±11.44	1, 2	First-order	/	BM vs. non-BM	% Looking	/	In a preferential looking paradigm participants watched a video with one animate (i.e., human or cat) and one inanimate (i.e., trucks or bicycle) PLD moving side by side on the screen.
							non verbal IQ	102.82±16.7	106.61±9.88							
Burnside et al., 2017	15	16	100%	50%	5.22±1.36	3.98±1.49	DAS - Non Verbal mental	4.36±1.87	4.28±1.56	1, 2, 3	First-order	Temporal	BM vs. non-BM	% Looking	/	In a preferential looking paradigm a walking human and its scrambled version were presented side by side on the screen for 6 sec.
							PPTV verbal metal age	4.06±1.85	4.72±1.79							
Turi et al., 2017	19	18	84.21%	77.78%	11.49±2.24	11.94±2.73	Wechsler Scale (WAS)	105.9±16.68	107.6±9.21	1, 2	Instrumental	/	BM	Accuracy	/	Subjects observed a movies of a hand grasping an invisible object, and had to guess whether the goal of the movement was towards a small cylinder or large cube. The movement could be presented with an egocentrinc or an allocentric view.

META-ANALYSIS	SUBGROUP	EFFECT SIZE [CI]
Biological motion (all studies)	/	0.57 [0.30, 0.85] ***
Level of BM processing	*First-order*	0.55 [0.02, 1.09] *
	Direct processing	0.40 [0.05, 0.74] *
	Instrumental processing	1.04 [0.68, 1.40] ***
Impact of low-level features of BM scrambled stimuli	*Spatial Scrambled*	0.38 [-0.07, 0.83]
	Temporal Scrambled	0.67 [0.26, 1.08] **
Non-biological motion	/	0.26 [-0.31, 0.83]