How many cubs can a mum nurse? Maternal age and size influence litter size in polar bears

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Abstract. Life history theory predicts that females' age and size affect the level of maternal

investment in current reproduction, balanced against future reproductive effort, maintenance

and survival. Using long-term (30 years) individual data on 193 female polar bears (*Ursus*

maritimus), we assessed age- and size-specific variation on litter size. Litter size varied with

maternal age, younger females had higher chances of losing a cub during their first months of

life. Results suggest an improvement of reproductive abilities early in life due to experience

with subsequent reproductive senescence. Litter size increased with maternal size, indicating

that size may reflect individual quality. We also found an optimum in the probability of

having twins, suggesting stabilizing selection on female body size. Heterogeneity was

observed among the largest females, suggesting that large size comes at a cost.

1. Introduction

Life history theory predicts that an optimal level of parental investment should maximize current reproductive success (RS) balanced against maintenance, survival, and future reproduction [1,2]. Among mammals, capital breeders are characterized by high maternal investment [10]. Lactation imposes high energetics demands on mothers [10] whereas energy is stored before breeding when foraging is constrained during reproduction [2]. Mothers' traits, namely age and body condition, should influence their ability to provide for their young, therefore influencing RS.

RS should increase with age due to an increase in maternal allocation to reproduction as residual reproductive value decreases ("terminal investment" hypothesis [1]). However, recent studies suggest a decline in RS with old age in wild vertebrates because of fewer resources to allocate to reproduction ("reproductive senescence" hypothesis [5]). An increase in RS has also been observed early in life due to increasing breeding abilities with experience ("constraint hypothesis" [6]). Moreover, irrespective of age, RS can vary with female body size and mass [3,4]. Larger size might benefit reproduction by improving or reflecting foraging abilities [7] and lactation [8]. RS might therefore increase until an optimal maternal size or age and then potentially decrease or level off because maintenance costs exceed the benefits associated with higher size or experience of older females [3,9]. However, to date, the influence of maternal traits on reproductive outputs in mammalian capital breeders has received little attention [3].

Using long-term individual data, we assessed age- and size-specific variation in female polar bears (*Ursus maritimus*) relative to litter sizes, which vary from one to three young, in the Svalbard population. We assumed high maternal investment because polar bears i) live in an extreme environment, ii) rely only on stored fat reserves during pregnancy and for the first four months of lactation, and iii) continue to care for and feed young during two and a half

years [11]. We expected litter size to increase until an optimal age and size, due to experience

and individual quality, and then decline for the largest, and for senescent, individuals.

2. Materials and methods

(a) Data collection

We live-captured polar bears from 1992 to 2017 at Svalbard, Norway, from late March

to beginning of May – just after females have emerged from maternity dens with their cubs

[11] – using methods described in Stirling et al. 1989 [12]. The age of first reproduction for

Svalbard females is usually six years of age [11]. Age was estimated using a premolar tooth

extracted from sub-adult and adult bears (cubs were of known age based on size) [13]. Body

straight length (cm), hereafter size, was measured as the dorsal straight-line made from the tip

of the nose to the caudal end of the tail bone with bears laying sternally recumbent. Litter size

(one to three) was recorded upon capture as the number of cubs-of-the-year a mother had

reared to that time.

(b) Statistical analyses

We analysed litter size as a function of mother's traits using multinomial regression

[14] that extends standard logistic regression to more than two outcomes. Litters of one were

chosen as the reference category. We did not consider whole litter loss here, and we included

only females that were observed with at least one young in the analyses. Separate odd-ratios

(OR) were determined for the relative risk of a litter size of "two" *versus* "one", and the

relative risk of a litter size of "three" *versus* "one", as a function of the covariates. Parameters

α and γ respectively represent the OR of twins and triplets, and will give an estimated

probability of having one, two or three cubs as a function of the covariates. We tested for

linear and quadratic effects of maternal age and size. Because fieldwork was spread over

more than a month and that mortality rates for cubs within their first year can be high [11], we

considered capture date (as an ordinal date with 1st January being 1 in normal years and 0 in

leap years) to account for a possible loss of young between den emergence and observation. A

yearly random effect was included on α and γ to account for environmental variation. Twenty-

seven models (Table 1) were fitted with a Bayesian approach using Markov Chain Monte

Carlo (MCMC) techniques in JAGS [15]. We used non-informative normal prior distributions

for the regression coefficients and a uniform prior distribution for the standard deviation of

the random effect. We ran two MCMC in parallel with different initial values, 200,000

iterations each and an initial burn-in of 40,000 iterations. One out of ten values were kept. We

assessed convergence by visual inspection and by using the Gelman and Rubin R-hat

diagnostic (R-hat < 1.1 [16]). For model comparison, we used the Deviance Information

Criteria (DIC [17]) and considered the model with lowest DIC as being best supported by the

data. Model selection consisted of 5 steps. In step 1, we compared different model structures

for the intercept (same or different intercept on α and γ). In step 2 (respectively 3 and 4) we

tested for the effect of maternal age (respectively maternal size and capture date). Each tested

variable could influence differently both OR (1: one common coefficient, 2: two distinct ones,

3: one for α or 4: one for γ). In step 5, we compared models with additive effects and

interactions between the previously selected variables. For ease of interpretation, we

considered young females to be aged between 6 and 9 years old because this should be age at

their first reproduction; passed 15 years old, we considered females as being old because

previous studies suggested reproductive and body senescence around that age [18]; last, we

considered prime-aged females as being aged from 10 to 15 years old. Using these cut-offs,

we fitted an additional model including age as a factor in the model best supported by the data

to assess differences in each age class (Table 2).

3. Results

The best model (model 5.2, Table 1) included an interaction between a quadratic effect

of age and capture date within the field season on the probability of having a litter of two over

one cub.

Table 1: List of all models considered with comparison based on Deviance Information

Criteria (DIC) values, with ΔDIC for the difference between the model with two intercepts

only and the model under investigation. Parameters α and γ represent the odds ratio (OR) of

twins and triplets, respectively, while # par. is for the number of model parameters. For each

step, the model best supported by the data is in bold.

Maternal size influenced both the probability of having two over one cub, and three

over one cub. Concerning age, the probability of having twins increased and then decreased

after mid-season (day 105) (Figure 1, ESM Figure 1).

Figure 1: Estimated probability of having 1, 2 or 3 cubs as a function of capture date (in days

from March 21st) for (a) young (6 years old), (b) prime-aged (12 years old) and (c) old (18

years old) mothers and for a mean maternal size (194.8 cm). Predictions were obtained from

the best model (model 5.2 in Table 1). Solid lines are posterior means while dotted lines are

95% credible intervals.

Young females (6-9 years old (y.o.)) had a high probability of having twins shortly

after denning (day 90, $P(y=2) \approx 0.9$), but it declined within the field season, leading to a higher

probability of having just one cub alive towards the end of it (after day 125, $P(y=1) \approx 0.9$,

Figure 1a). Early in the season, prime-aged females (10-15 y.o.) also had a higher probability

of having two over one cub ($P(y=2)=0.7$ and $P(y=1)=0.3$, Figure 1b) while older females (>15

y.o.) had similar probabilities of having singletons or twins ($P(y=2) \approx P(y=1) \approx 0.5$, Figure 1c).

In contrast to the marked drop in spring litter size with time for young females, little variation

was observed for prime-aged ($0.7<P(y=2)<0.8$ and $0.2<P(y=1)<0.3$, Figure 1b) and old

females ($P(y=2) \approx P(y=1) \approx 0.5$, Figure 1c). Examining parameter estimates confirmed that the

interaction term between age and date was only significant for young mothers and not for

older ones (Table 2).

Table 2: Parameter estimates of the multinomial regression model derived from the model

best supported by the data (see Table 1) using age as a factor. Posterior mean, standard

deviation (SD) and 95% credible intervals are provided for the odds ratios (fixed effects) as

well as the variance of the random effect year. Rhat is the potential scale reduction factor (at

convergence, Rhat < 1.2).

Litter size globally increased with the size of a mother (Figure 2). A higher probability

of having singletons was found for smaller females (size<183cm, $0.5<P(y=1)<0.8$,

$0.2<P(y=2)<0.5$, $P(y=3)=0$), of having twins for medium-sized females (190cm<size<200cm,

$P(y=1)=0.3$, $P(y=2)=0.7$, $P(y=3)=0$), and of having singletons or triplets for larger females

(size>210 cm, $P(y=1) \approx 0.5$, $0<P(y=2)<0.2$, $0.2<P(y=3)<0.5$). However, sample size was very

low for triplets (n=8).

Figure 2: Estimated probability of having 1, 2, or 3 cubs as a function of maternal size for a

mean value of mother's age (11.4 years) and capture date (day 105 ~ 16th of April).

Predictions were obtained from the best model (model 5.2 in Table 1). Solid lines are

posterior means while dotted lines are 95% credible intervals.

4. Discussion

Our results showed an influence of maternal traits on litter size, an index of

reproductive success, suggesting that a mothers' ability to invest in reproduction and care for

their offspring varied with their age and size. The quadratic pattern of variation in litter size

observed with age supported the hypothesis of a benefit of gaining experience early in life [6]

until 12 y.o. and of reproductive senescence [5] starting from 15 y.o. These results contradict

the terminal investment hypothesis [1], and support results on Canadian polar bears showing a

decrease in litter size and maternal body senescence after 16 years [18].

Under the experience hypothesis, an improvement of female's hunting skills might

explain the improvement of reproductive investment early in reproductive life [19]. Other

studies on mammals [3,18] suggest that an age-related increase in reproductive success could

be linked to mass gain and hence resource availability. Further supporting the experience

hypothesis, we found that younger females (< 10 y.o.) had about the same probability as

prime-aged females to produce two cubs, but higher chances to lose one during the capture

spring season.

Under the senescence hypothesis, degradation of physiological functions with aging

[20] might impair females' fat stores accumulation, causing a simultaneous decrease in

females' mass and reproductive outcomes [18]. Older mothers might therefore acquire less

energy, and might have higher energy allocation needs toward self-maintenance, reducing

energy allocated toward reproduction [1]. Reproductive senescence has been documented in

many wild populations for several reproductive parameters such as litter size [4,18], offspring

mass [3,5,18], and survival [4].

We showed that, on average, litter size increased with maternal size. The probability

of having triplets was only shown to depend on mother's size, although the sample size for

triplets was small. Large size might therefore be an index of individual quality, like in wolves

(*Canis lupus*) [4]. The increase in the probability of having twins, and decrease in probability

of having singletons, for females up to an optimum size, support this idea. Other traits

highlighting foraging capacities have been related to RS – e.g. body mass and condition in

bears and other species [3,18]. Among the largest females, the chances of having a singleton

or triplets were almost equal and increased, while that of having twins was low and

decreasing. Considering the decrease in the probability of having twins, and the increase in

the probability of singletons, stabilizing selection on adult female body size is likely to

happen: increased investment in growth likely comes at a cost in terms of somatic

maintenance.

Considering results on triplets, we suspect higher individual heterogeneity among the

largest females compared to smaller and medium sized ones. With a larger sample size, this

could be tested by assessing variance in litter size probability. Results on litter size probability

suggest that there is one group of 'good quality females' having triplets and another group of

'low quality females' having singletons. For the latter group, larger size could be associated

with a cost that may depend on other factors, such as body condition and environmental

quality. Large individuals occupying resource-poor habitats, or experiencing a year of

reduced resource availability, might not have enough resources to allocate to both their own

maintenance and care for triplets. Heterogeneity in individual quality may override

reproductive cost [22], and costs may be restricted to resource-limited contexts [23].

Overall, we found that litter size in polar bears increased with age of mothers early in

life until a plateau, followed by a decrease for old females. Because population growth mostly

depends on female's RS, itself influenced by maternal traits, our findings highlight the

importance of accounting for individual heterogeneity to understand the species response to

environmental perturbations. Future research will aim at understanding the determinants of

female polar bears' reproductive tactics by accounting for environmental conditions.

Influence of climate variability has been shown to affect reproductive parameters in several

populations, including Svalbard [24].

Table 1.

Steps	Models	α twins OR	γ triplets OR	# par.	DIC	ΔDIC
1: Intercept	**1.1**	**Intercept**	**Intercept '**	**2**	**365.8**	**0**
	1.2	Intercept	Intercept	1	515.7	150.5
2: Age	**2.1**	**Age^2**	**β_0'**	**4**	**353.5**	**-12.3**
	2.2	Age^2	Age^2'	6	356.5	-9.3
	2.3	Age^2	Age^2	4	359.4	-6.4
	2.4	β_0	Age'	3	364.7	-1.1
	2.5	Age	Age'	4	365.5	-0.3
	2.6	Age	β_0'	3	365.5	-0.3
	2.7	β_0	Age^2'	4	366.4	0.6
	2.8	Age	Age	3	367	1.2
3: Size	**3.1**	**Size^2**	**Size'**	**5**	**360.3**	**-5.5**
	3.2	β_0	Size'	3	361.4	-4.4
	3.3	Size^2	Size^2'	6	362	-3.8
	3.4	β_0	Size^2'	4	362.1	-3.7
	3.5	Size^2	β_0'	4	362.9	-2.9
	3.6	Size	Size'	4	363.2	-2.6
	3.7	Size	β_0'	3	366.8	1
	3.8	Size^2	Size^2	4	367.5	1.7
	3.9	Size	Size	3	367.9	2.1
4: Date	**4.1**	**Date**	**β_0'**	**3**	**363.7**	**-2.1**
	4.2	Date	Date	3	364.6	-1.2
	4.3	Date	Date'	4	365.8	0
	4.4	β_0	Date'	3	367.4	1.6
5: Additive effects and interactions	5.1	Age^2 + Date + Size^2	Size'	8	346.9	0
	5.2	**Age^2 x Date + Size^2**	**Size'**	**10**	**342.8**	**-4.1**
	5.3	Age^2 + Date * Size^2	Size'	10	348.6	1.7
	5.4	Age^2 x Size^2 + Date	Size'	12	354.1	7.2

212 **Table 2.**

	Parameter	Mean	SD	95% credible interval	Rhat
	Fixed effects				
α	Intercept	1.399	0.275	[0.881; 1.962]	1.001
	Age (young)	-0.653	0.350	[-1.352; 0.023]	1.001
	Age (old)	-1.183	0.421	[-2.015; -0.356]	1.001
	Date	-0.047	0.207	[-0.451; 0.357]	1.001
	Size	-0.005	0.158	[-0.312; 0.312]	1.001
	$Size^2$	-0.279	0.112	[-0.505; -0.062]	1.001
	Age (young) x Date	-0.760	0.390	[-1.550; -0.017]	1.001
	Age (old) x Date	-0.149	0.404	[-0.946; 0.638]	1.001
γ	Intercept'	-2.634	0.486	[-3.671; -1.772]	1.001
	Size'	0.732	0.358	[0.056; 1.466]	1.001
	Year random effect				
	Variance	0.398	0.246	[0.024; 0.936]	1.001

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