

1 **Title:** Species abundance distributions should underpin ordinal cover-abundance  
2 transformations

3 **Running Title:** Ecological analysis using global plant data

4 **Abstract**

5 The cover and abundance of individual plant species have been recorded on ordinal  
6 scales for millions of plots world-wide. Many ecological questions can be addressed  
7 using these data. However ordinal cover data may need to be transformed to a  
8 quantitative form (0 to 100%), especially when scrutinising summed cover of multiple  
9 species. Traditional approaches to transforming ordinal data often assume that data  
10 are symmetrically distributed. However, skewed abundance patterns are ubiquitous  
11 in plant community ecology. A failure to account for this skew will bias plant cover  
12 estimates, especially when cover of multiple species are summed. The questions  
13 this paper addresses are (i) how can we estimate transformation values for ordinal  
14 data that accounts for the underlying right-skewed distribution of plant cover; (ii) do  
15 different plant groups require different transformations and (iii) how do our  
16 transformations compare to other commonly used transformations within the context  
17 of exploring the aggregate properties of vegetation? Using a continuous cover  
18 dataset, each occurrence record was mapped to its commensurate ordinal value, in  
19 this case, the ubiquitous Braun-Blanquet cover-abundance (BBCA) scale. We fitted a  
20 Bayesian hierarchical beta regression to estimate the predicted mean (PM) cover of  
21 each of six plant growth forms within different ordinal classes. We illustrate our  
22 method using a case study of 2 809 plots containing 95 812 occurrence records with  
23 visual estimates of cover for 3 967 species. We compare the model derived  
24 estimates to other commonly used transformations. Our model found that PM  
25 estimates differed by growth form and that previous methods overestimated cover,

26 especially of smaller growth forms such as forbs and grasses. Our approach reduced  
27 the cumulative compounding of errors when transformed cover data were used to  
28 explore the aggregate properties of vegetation and was robust when validated  
29 against an independent dataset. By accounting for the right-skewed distribution of  
30 cover data, our alternate approach for estimating transformation values can be  
31 extended to other ordinal scales. A more robust approach to transforming floristic  
32 data and aggregating cover estimates can strengthen ecological analyses to support  
33 biodiversity conservation and management.

34

35 **Keywords:** aggregated, beta regression; Braun-Blanquet; growth form; midpoint;  
36 ordinal transformation; species abundance distribution; sPlot; summed foliage cover;  
37 VegBank; vegetation cover.

38

39 **Abbreviations:** Braun-Blanquet cover-abundance - BBCA

40

## 41 **Introduction**

42 Field-based assessment of the cover and abundance of individual plant species is  
43 complex. Observers making on-ground visual estimates of plant cover need to  
44 account for, and assess, foliage cover of different densities, dimensions, shapes and  
45 structures across multiple species, growth forms and strata. So too, counting cryptic,  
46 clonal, or copious numbers of plants can be complicated. Owing to this complexity,  
47 vast numbers of floristic plots across many continents have been surveyed using  
48 ordinal scales (Schaminée et al. 2009; Dengler et al. 2011; Chytrý et al. 2016).  
49 Whilst, in Braun-Blanquet (1932) originally described an abundance-dominance  
50 scale, the practical, on-ground application of this scale is to assess plant cover, and  
51 where cover is less than 5%, abundance is also assessed. The Braun-Blanquet  
52 cover-abundance (BBCA) scale is perhaps the most common ordinal scale used in  
53 plant ecology. For example, within the vegetation plot database sPlot v2.1  
54 ([www.idiv.de/splot](http://www.idiv.de/splot)), more than 745 000 plots (66%) have recorded plant occurrence  
55 using Braun-Blanquet cover-abundance (sPlot extract supplied by Borja Jiménez-  
56 Alfaro, 19<sup>th</sup> September 2017). This volume of data is testament that ground-based  
57 visual assessments of cover-abundance using ordinal scales provide a cost-  
58 effective, rapid and non-destructive approach to gathering the data needed to  
59 summarise the composition and structure of plant communities. These data  
60 represent a wealth of investment in field effort and have supported major advances  
61 in vegetation classification, mapping and distribution modelling.  
62 The ever-growing access to global vegetation plot databases (Dengler et al. 2011;  
63 Schaminée et al. 2011) has opened pan-continental opportunities to explore many  
64 uses of floristic data. Some ecological questions may best be addressed using  
65 aggregate properties of vegetation, such as the summed total foliage cover within a

66 plot or across strata, the total summed cover or abundance of exotic or invasive  
67 species, or the relative cover or abundance of plants within different functional,  
68 taxonomic or growth form groups. Summing cover to derive aggregate properties of  
69 floristic data have a multitude of uses in ecology including assessing presence and  
70 diversity of faunal habitat, as covariates in species' distribution models (SDMs), for  
71 assessing the spatial and temporal status of ecosystem baselines, predicting the  
72 effects of shifts in climate, land use and land cover, or measuring site-scaled  
73 responses to disturbance (e.g. Scholes & Biggs 2005; McElhinny et al. 2006; Pereira  
74 et al. 2010). Aggregate properties of vegetation data are particularly relevant to  
75 exploring ecological questions concerning the patterns, processes and prognoses at  
76 a range of spatial scales in contemporary and predicted future landscapes.

77 There are many applications where ordinal data have been used successfully, such  
78 as ordination, classification, modelling or mapping of vegetation communities (e.g.  
79 Podani 2005; Podani 2006; Lyons et al. 2016) and for modelling the cover of single  
80 species (e.g. Damgaard 2014; Irvine et al. 2016). However, ordinal scaled cover  
81 observations of individual species cannot be summed (Guisan & Harrell 2000;  
82 Podani 2006; Chen et al. 2008b) and need to be transformed into a continuous scale  
83 prior to aggregating.

84 Approaches to transforming Braun-Blanquet cover-abundance (BBCA) ordinal data  
85 have been proposed by Tüxen and Ellenberg (1937) and Braun-Blanquet (1964)  
86 (see Table 3 in van der Maarel 1979). In addition, van der Maarel (1979) proposed  
87 the ordinal transform value (OTV) with different scale adjustments, as a solution for  
88 converting ordinal data to percentage cover values. All these methods tend to  
89 transform data to the approximate midpoint of the ordinal class range for  
90 observations of cover greater than 5%. For classes with cover less than 5%, the

91 transformation values appear arbitrary and differ considerably (Table 1 columns 4–  
92 6).

93 Transforming data to the approximate midpoint of the class ranges assumes that  
94 data are symmetrically distributed within each class. Yet, patterns in plant  
95 abundance including density, biomass (Chiarucci et al. 1999; Morlon et al. 2009),  
96 frequency (Chiarucci et al. 1999), percentage cover (Damgaard 2009), size, energy  
97 use and productivity (Whittaker 1965) have all been shown to have a right-skewed  
98 distribution; skewed species abundance distributions occur in every known multi-  
99 species community (McGill et al. 2007). Midpoint transformations are inflexible to the  
100 underlying distribution of cover data and assume that the distribution does not vary  
101 across species, groups of plant entities (such as growth forms, life forms, functional  
102 or taxonomic groups), vegetation types or biomes. Due to the prevalence of right-  
103 skewed distribution, we predict that midpoint transformations overestimate cover and  
104 the compounding of these errors will result in gross overestimation of summed cover  
105 for aggregated properties.

106 Here we develop a flexible approach to estimate cover transformations for ordinal  
107 scaled data that can then be used to provide accurate estimates of summed  
108 vegetation cover. The method we describe is applicable to data in any ordinal scale,  
109 can be extended to allow for differences in vegetation type or among biomes and  
110 can accommodate alternative aggregate properties of plant data such as growth  
111 forms, life forms, functional or taxonomic groups. To demonstrate the potential  
112 applicability of our approach we build and then validate the model using two  
113 separate and independent datasets.

114 Given that diverse architectures and spatial arrangements of foliage lead to varied  
115 patterns of plant cover (Damgaard 2013), we also predict that different plant growth

116 forms will require different transformation values. Growth forms are practical and  
117 observable entities that can inform site-based assessment and monitoring, are  
118 recognizable from remotely-sensed imagery and are used to report on broad-scale  
119 biodiversity assessment or baselines (Pereira et al. 2013) with which we can  
120 measure change in cover (Pettorelli et al. 2014; Abelleira Martínez et al. 2016).

121

## 122 **Materials and Methods**

123 We outline the key steps required to estimate transformation values within ordinal  
124 classes for different plant groups. A pre-requisite for our method is cover data that  
125 have been collected on a continuous cover scale, ideally sourced from the same  
126 study region and vegetation types as the ordinal cover data. To prepare the input  
127 data for the model, ordinal values need to be mapped, *a posteriori*, to this continuous  
128 cover data as an intermediary variable (Figure 1, Step 1). Models, with a beta  
129 distribution, are then used to predict the mean cover of each plant group within each  
130 ordinal cover class. This predicted mean cover is the transformation value (Figure 1,  
131 Step 2). Using a case study, we explore summed cover estimates for different plant  
132 groups and evaluate the performance of the ordinal cover transformations. We  
133 compare our transformation to existing approaches in the context of summed cover  
134 for plant groups (Figure 1, Step 3). We evaluate the robustness of our predicted  
135 mean transformations on an independent dataset (Figure 1, Step 4).

136 *Estimate mean cover using parameters of a beta distribution*

137 We used a generalised linear mixed model (GLMM) with a beta distribution to derive  
138 estimates of the mean vegetation cover, within an ordinal class, given a plant's  
139 growth form and random variation owing to plot identity. Individual species cover are  
140 continuous proportional estimates, and once suitably transformed, fall within the  
141 known range ( $0 < y < 1$ ). Linear regression with a normal distribution is inappropriate  
142 for the analysis of proportions, such as percent plant cover, because data often  
143 violate assumptions such as normality and homogeneity of errors and furthermore  
144 fitted values can fall outside of the range  $[0, 1]$  (Ferrari & Cribari-Neto 2004). A  
145 common approach to address these problems is to apply arcsine or logit  
146 transformations to the response variable, prior to regression (Warton & Hui 2011),  
147 although the results can be difficult to interpret (Ferrari & Cribari-Neto 2004).  
148 Numerous authors have instead demonstrated that percent plant cover are more  
149 appropriately analysed by assuming that cover approximates a two-parameter beta  
150 distribution (Ferrari & Cribari-Neto 2004; Chen et al. 2008a; Cribari-Neto & Zeileis  
151 2010; Herpigny & Gosselin 2015). Beta distributions are attractive because fitted  
152 values are constrained between the interval  $0 < y < 1$  and they can accommodate  
153 asymmetrical distributions with left- or right-skew. This flexibility makes beta  
154 distributions highly suitable for modelling diverse and often asymmetrical plant cover  
155 data (Cribari-Neto & Zeileis 2010).  
156 We present a Bayesian GLMM with a logit link to estimate the parameters of the beta  
157 distribution and allowed these parameters to vary among ordinal classes and plant  
158 growth forms. Estimates of these parameters were used to derive the predicted  
159 mean (PM) for each plant growth form in each ordinal class.

160

161 The proportional vegetation cover is given by the two-parameter beta distribution;

162  $Proportion_{ij} \sim Beta(a_{ij}, b_{ij})$

163 Where  $a_{ij}$  and  $b_{ij}$  are shape parameters for species  $j$  in plot  $i$ , and  $i = 1, \dots, n$  plots. The

164 shape parameters are further defined as

165  $a_{ij} = \theta \times \pi_{ij}$

166  $b_{ij} = \theta \times (1 - \pi_{ij})$

167 where  $\theta$  allows for potential overdispersion to be incorporated in the model (Zuur et  
168 al. 2013).

169  $\pi_{ij}$  is modelled with a logit link

170  $logit(\pi_{ij}) = \eta_{ij}$

171 The model consists of regression parameters ( $\beta$ ) for each ordinal class, plant growth

172 form and their interactions, plot level random intercepts and variance ( $z_i$ ):

173  $\eta_{ij} = X_{ij} \times \beta + z_i$

174  $z_i \sim N(0, \delta^2_{plot})$

175 Where  $z_i$  is a random intercept for plot,  $X_{ij}$  are the matrix of all covariates (ordinal

176 classes and their interaction with plant growth form) and  $\beta$  are the regression

177 parameters for each covariate. That is, for each ordinal class 1...6, separate  $\beta$  were

178 estimated for each plant growth form. For a simplified example with two growth forms

179 and two ordinal classes this can also be expressed as:

180  $\eta_{ij} = \beta_0 + \beta_1 \times fOrdinalClass_{ij} + \beta_2 \times fGrowthForm_{ij} + \beta_3 fOrdinalClass_{ij} \times fGrowthForm_{ij} +$

181  $z_i$

182 Where  $\beta_0$  = predicted value of logit transformed cover if species  $j$  belongs to the

183 “reference” growth form and its’ value in plot  $i$  has the “reference” level ordinal cover-

184 abundance class.



185  $\beta_1$  = departure of the predicted value for species  $j$  from  $\beta_0$  if the observation is of  
186 another ordinal cover-abundance class.

187  $\beta_2$  = departure of predicted value from  $\beta_0$  if species  $j$  belongs to another growth form.

188  $\beta_3$  = departure of predicted value from  $\beta_0 + \beta_1 + \beta_2$  when neither growth form nor  
189 ordinal cover-abundance class are of the reference level.

190 In this example,  $fOrdinalClass_{ij}$  and  $fGrowthForm_{ij}$  are binary dummy variables  
191 coding growth form and cover-abundance scale categories, thus  $X_{ij}$  is a vector  
192 containing values for these dummy variables (including their products) for species  $j$   
193 in plot  $i$ .

194 We included plot as a random intercept because although we assumed each plot  
195 should follow the characteristic skewed species abundance curve, we expected  
196 variation among plots and hence differences in the average cover of any given  
197 ordinal class and plant growth form.

198 This basic model structure can be easily expanded to accommodate other possible  
199 sources of variation, such as among vegetation types or owing to the richness of  
200 plant species within a plot. In this case study, we decided not to include additional  
201 covariates to minimise computational demands and simplify model interpretation and  
202 operational complexity.

203 The model was fit via Markov chain Monte Carlo optimization in JAGS ([http://mcmc-](http://mcmc-jags.sourceforge.net)  
204 [jags.sourceforge.net](http://mcmc-jags.sourceforge.net)) via the R2jags package (Su & Yajima 2015) within R 3.5.0 (R  
205 Core Team 2018). Posterior parameter estimates and back-transformed predicted  
206 means were derived from 3 chains, with a burn-in of 3000 iterations, 15 000  
207 subsequent iterations per chain and with a thinning rate of 15. Autocorrelation and  
208 mixing were visually inspected. The interaction models were compared to additive

209 models using Deviance Information Criteria. Appendix S1 contains R code for our  
210 models.

### 211 *Case study – New South Wales, Australia*

212 We illustrate our model with a case study where we have used 1-6 BBCA as our  
213 ordinal scale and grouped plants into six growth form categories. Following is a brief  
214 description of how we prepared the case study dataset to build our model. We note  
215 that randomly generated data from an appropriate beta distribution (for similar  
216 example see Damgaard 2014) could also be used to demonstrate our approach.  
217 However, we chose to use a large archival dataset from a range of bioclimatic  
218 regions and vegetation types to demonstrate that, despite the underlying variation,  
219 our approach still led to robust estimates of summed cover.

#### 220 *1. Preparation of observed percentage cover dataset*

221 To demonstrate our modelled approach, we sourced case study data from archival  
222 quantitative floristic data that met three considerations: (i) each species record  
223 included a visual estimate of foliage cover on a continuous scale from 0.1% to 100%  
224 and a count of abundance where cover was less than 5%; (ii) in each plot, full  
225 species inventories were recorded from a fixed-area (400 m<sup>2</sup>) and (iii) sites covered  
226 a wide geographic distribution (Appendix S2—Figure 1) and included a wide range of  
227 vegetation types with different structural complexity including rainforests, forests,  
228 woodlands, shrublands, grasslands and wetlands (Keith 2004). A total of 2809 geo-  
229 referenced plots containing 95 812 occurrence records with visual estimates of cover  
230 for 3967 species met these criteria and were exported from the NSW BioNet Atlas  
231 database ([www.bionet.nsw.gov.au](http://www.bionet.nsw.gov.au)).

#### 232 *Analysis of the empirical cover distribution*

233 To confirm our assumption of right-skewed distribution of cover data we plotted our  
234 data and used the 'skewness' function in the e1071 package (Meyer et al. 2017)  
235 within R 3.3.3 (R Core Team 2018) to calculate the adjusted Fisher-Pearson  
236 skewness coefficient ( $G_1$ ) (Joanes & Gill 1998) for the whole distribution, and for  
237 distributions within each BBCA class. Skewness is a diagnostic tool usually used to  
238 test the symmetry of the data distribution. Here, we interpret skewness coefficients  
239 as being strongly and positively skewed when the  $G_1$  coefficient is greater than 0.5  
240 (Bulmer 1979; Doane & Seward 2011).

## 241 2. *Preparation of plant group entities*

242 All taxa were allocated to one of six growth form categories: tree, shrub, grass and  
243 grass-like (hereafter referred to as grass), forb, fern and other (remaining growth  
244 forms) (Oliver et al. submitted). For each growth form in each plot, total cover was  
245 estimated by summing the observed quantitative estimates of cover and the  
246 estimates of cover derived from the transformations of the ordinal data.

## 247 3. *Allocating an intermediary variable*

248 We created an intermediary variable by matching each quantitative estimate of cover  
249 for every floristic record ( $n = 95\ 812$ ) to its commensurate ordinal value. Any ordinal  
250 scale can be used to partition data, but here we demonstrate our approach by  
251 allocating data to 1–6 BBCA (Table 1). BBCA1 and BBCA2 were assigned based on  
252 their observed foliage cover (<5%) and abundance; where BBCA1  $\leq 10$  and BBCA2  
253  $> 10$  individuals. The pragmatic choice of ten individuals provides an explicit  
254 quantitative abundance threshold between classes BBCA1 and BBCA2. BBCA3–  
255 BBCA6 were assigned based on observed foliage cover ( $\geq 5\%$ ) (Mueller-Dombois &  
256 Ellenberg 1974). The ordinal dataset created by this process approximates the form  
257 of many data held within vegetation databases.

258 In our case study dataset, observations of 5% cover were more prevalent than  
259 expected from a typical theoretical beta distribution (Figure 2). This bias was  
260 detected in preliminary model convergence diagnostics and model fit suggested that,  
261 for our case study, it would be preferable to split the data and separately model (i)  
262 BBCA1 and BBCA2 bounded between 0 and less than 5% cover and (ii) BBCA3 to  
263 BBCA6 bounded between 5% and 100% cover inclusive. To ensure the response  
264 variable was bounded by 0 and 1, percent cover was transformed using  $(y-a)/(b-a)$   
265 where in (i)  $a = 0$  and  $b = 5$  and in (ii)  $a = 5$  and  $b = 100$  (Cribari-Neto & Zeileis  
266 2010). In the second model, the response variable was further transformed using  $(y * (n-1))/n$   
267 where  $n =$  sample size (Cribari-Neto & Zeileis 2010). This split-model  
268 approach may not be necessary for all datasets, especially where data are derived  
269 from less subjective cover methods (e.g. point intercept or pin frame) but is included  
270 here to support the handling of datasets with similar patterns in distribution (see  
271 Appendix S3—Figures S1–S4 for other datasets that appear to show similar pattern).

### 272 *Evaluation of past and proposed approaches to transforming ordinal data*

273 We transformed each of the 1–6 BBCA records using three different approaches  
274 outlined in Table 1. We then evaluated these past approaches proposed by Tüxen  
275 and Ellenberg (1937), Braun-Blanquet (1964) and van der Maarel (2007) to the PM  
276 estimated from a beta distribution.

277 For each plot, growth form cover and total cover were calculated by summing the  
278 observed continuous cover estimates (%) and the estimates of cover derived from  
279 the various transformations. Linear regression models with zero-intercept were fitted  
280 to the sum of observed continuous cover data ( $y$ ) and sum of transformed cover data  
281 ( $x$ ) cover data in R 3.5.0 (R Core Team 2018). We can justify using a regression  
282 through the origin because we are most interested in comparing the slope of the

283 regression line to the 1:1 line of best fit to determine if our PM models were over or  
284 underpredicting summed cover. We compared the root mean squared deviation  
285 (RMSD) (see Eq. 1) as an estimate of the deviation of the transformed cover values  
286 from the 1:1 line.

287 RMSD =

$$\sqrt{\frac{1}{n-1} \sum_{i=1}^n (\hat{y}_i - y_i)^2}$$

288 (Eq. 1)

289 Where  $\hat{y}_i$  are the predicted cover values;  $y_i$  are the observed cover values and n is  
290 the number of observations.

291 The RMSD estimate represents the mean deviation of transformed cover values with  
292 respect to the true observed cover values. We also compared estimates of the slope  
293 with lower and upper 95% confidence intervals expecting that robust transformations  
294 would result in a slope = 1 and transformations that overestimate summed cover will  
295 have a slope <1. We include the adjusted coefficient of variation ( $R^2$ ) to evaluate how  
296 much of the linear variation of observed cover values is explained by the variation of  
297 transformed cover values.

298 We note that RMSD is useful for evaluating models as it represents an absolute  
299 measure of fit to the 1:1 line and reports the prediction error in the same units as the  
300 data (i.e. summed cover). Whereas adjusted  $R^2$  gives a relative measure of  
301 proportion of total variance that is explained by the model on a scale between 0 and  
302 1.

303 We validated the PM transformation values on an independent dataset (2 227 sites  
304 with 51 497 observations) from West Virginia Natural Heritage Program (Vanderhorst  
305 et al. 2012) accessed from VegBank (Peet et al. 2013 accessed 28th Aug 2018).  
306 Whilst VegBank has a primary role for enabling the vegetation classification, large  
307 volumes of individual floristic observations are available for ecoinformatic synthesis  
308 and analysis. Owing to the ease of access and completeness of datasets stored in  
309 VegBank we were able to validate our model estimates on a geographically distinct  
310 dataset containing cover estimates of plants from entirely different vegetation  
311 communities. Details outlining the data preparation are included in Appendix S6.

## 312 **Results**

### 313 *The empirical cover distribution*

314 The source continuous cover data were right-skewed and dominated by low cover—  
315 85% of observations were between 0.1 and 4%, and 60% of these observations  
316 were of cover less than 1% (Figure 2). Data were heavily right-skewed for the whole  
317 distribution ( $G_1 = 5.62$ ) and right-skewed within five of the six BBCA classes (BBCA1  
318  $G_1 = 2.64$ , BBCA2  $G_1 = 1.57$ , BBCA3  $G_1 = 1.04$ , BBCA4  $G_1 = 0.61$  and BBCA6  $G_1 =$   
319  $0.95$ ). Only BBCA5 had a skewness coefficient less than 0.5 ( $G_1 = 0.36$ ). We also  
320 note potential observer bias for 5% cover. These patterns are similar to other visually  
321 estimated floristic cover data from other archived datasets (see Appendix S3—  
322 Figures 1-4).

323

324 *Estimate mean cover using parameters of a beta distribution*

325 Table 1 (columns 7 and 8) shows the predicted mean transformations and their lower  
326 2.5% and upper 97.5% credible interval for each ordinal class, independent of  
327 growth forms. The most marked differences are noted in BBCA2 and BBCA3, where  
328 the predicted means are well below the previous approaches. The predicted mean  
329 for class BBCA6 is lower than the midpoint but was derived from relatively few  
330 observations (n = 138).

331 Table 1: Class divisions for the 1–6 Braun-Blanquet ordinal cover-abundance (BBCA) scale (columns 1–3), previous proposals for  
 332 transforming them to percentage cover (columns 4–6), and proposed transforms (independent of growth form) based on estimating  
 333 the predicted mean (PM) from a beta distribution of observed quantitative cover data and the lower 2.5% and upper 97.5% credible  
 334 intervals. Number of observations (n) for BBCA1 (n = 54 811); BBCA2 (n = 26 968); BBCA3 (n = 11 946); BBCA4 (n = 1583);  
 335 BBCA5 (n = 366) and BBCA6 (n = 138).

Column 1	2	3	4	5	6	7	8
BBCA Class	Range of cover (%)	Qualitative abundance terms	Tüxen & Ellenberg (1937) <sup>1</sup>	Braun-Blanquet (1964) <sup>1</sup>	van der Maarel (2007) <sup>2</sup>	PM	Credible interval
1	<5	e.g. present, few, rare, erratic, occasional, uncommon	0.1	0.1	1	0.49	0.48–0.51
2	<5	e.g. common, abundant, many, several	2.5	5	2	0.74	0.72–0.76
3	5–25		15	17.5	17.5	8.95	8.84–9.07
4	26–50		37.5	37.5	35	38.77	37.97–39.57
5	51–75		62.5	62.5	70	62.43	60.69–64.13
6	76–100		87.5	87.5	140	81.24	79.10–83.26

336

337 <sup>1</sup> adapted from van der Maarel (1979).

338 <sup>2</sup> ordinal transform values (OTV) using 1.415 weighting factor (van der Maarel 2007).

339 Column 3 shows some of the qualitative descriptors used by field surveyors to divide observations between BBCA1 and BBCA2.



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340 Table 2: Proposed transformation values, tailored to different growth forms, based on estimates of the predicted mean (PM) from a  
 341 beta distribution of observed data. Lower 2.5% and upper 97.5% credible intervals (CI) are shown in square brackets; n = number  
 342 of individual observations for each Braun-Blanquet cover-abundance (BBCA) class.

		Growth Form					
		Tree	Shrub	Grass	Forb	Fern	Other
BBCA1	PM	0.78	0.52	0.45	0.42	0.43	0.45
	CI	(0.76 – 0.80)	(0.51 – 0.53)	[0.44 – 0.46]	[0.41 – 0.44]	[0.41 – 0.45]	[0.44 – 0.46]
	n	6465	13366	8731	16981	1863	7405
BBCA2	PM	1.78	0.96	0.82	0.58	0.75	0.74
	CI	[1.7 – 1.86]	[0.93 – 0.99]	[0.8 – 0.84]	[0.57 – 0.6]	[0.72 – 0.78]	[0.71 – 0.77]
	n	441	2821	8488	12065	1324	1829
BBCA3	PM	9.53	8.60	8.86	8.02	8.80	8.48
	CI	[9.38 – 9.7]	[8.42 – 8.76]	[8.71 – 9.01]	[7.81 – 8.24]	[8.46 – 9.19]	[8.23 – 8.74]
	n	4347	2070	3500	893	412	724
BBCA4	PM	38.06	39.18	39.32	37.86	39.30	38.72
	CI	[36.74 – 39.38]	[36.98 – 41.26]	[38.05 – 40.59]	[34.01 – 41.81]	[34.92 – 43.72]	[34.76 – 42.64]
	n	582	217	599	68	52	65
BBCA5	PM	61.71	63.15	62.67	62.80	62.01	62.11
	CI	[58.15 – 65.16]	[58.23 – 67.78]	[60.24 – 65.08]	[49.59 – 75.8]	[54.16 – 70.05]	[54.66 – 69.15]
	n	89	48	184	6	18	21
BBCA6	PM	80.80	83.9	80.81	78.87	77.59	80.55
	CI	[76.04 – 85.31]	[78.99 – 88.41]	[77.82 – 83.65]	[64.83 – 90.78]	[55.35 – 93.93]	[70.09 – 89.57]
	n	29	23	73	4	2	7

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344 *Evaluation of past and proposed transform values for summed growth form cover*

345 Estimates of the PM suggest that accounting for growth form within each ordinal class  
346 results in more robust summed cover estimates. Credible intervals suggest that in classes  
347 BBCA1 to BBCA3, trees typically have higher mean cover and warrant higher  
348 transformation values (Table 2). Credible intervals also suggest the need for separate  
349 transformation values for shrubs in BBCA1 and BBCA2 and a lower value for forbs in  
350 BBCA2 and BBCA3 (Table 2).

351 When these growth form specific transformations were evaluated using the summed cover  
352 estimates RMSD did not exceed 9.50 (trees) (Figure 3 and Appendix S4—Table 1). In  
353 contrast, estimates based on past transformations frequently resulted in RMSD exceeding  
354 10. Slope ranged from 0.91 (forbs) to 1.05 (others), whereas past transformations slopes  
355 were  $<0.85$ , suggesting considerable overestimation of summed cover (see Appendix S4—  
356 Table 1 and Appendix S5—Figures 1-4).

357 *Evaluation of past and proposed transform values for total summed cover*

358 Evaluation of summed total cover revealed that when transformations are tailored to growth  
359 forms, the PM performed better than existing approaches (Figure 3). The PM reduced the  
360 overestimation of total summed cover by up to 4 times. The evaluation of model fit for  
361 summed total cover using past approaches generally revealed a poorer model fit: RMSD  
362 ranged from 41.47–79.37 (PM = 18.21) (see Appendix S4—Table 1) and slope ranged from  
363 0.57 to 0.74 (PM = 1.01) and adjusted  $R^2$  ranged from 0.61-0.96 (PM = 0.97).

364 Evaluation of the growth form specific PM transformation on an entirely independent  
365 validation dataset from West Virginia Natural Heritage Program (Vanderhorst et al. 2012)  
366 show that transformations were robust, although tended to underestimate summed cover of  
367 most growth forms (Appendix 6—Table 1). RMSD ranged between 1.59 (others) and 14.97  
368 (trees); slope ranged between 1 (others) and 1.12 (forbs) and adjusted  $R^2$  were high and

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369 ranged between 0.97 (trees and shrubs) and 0.93 (others). When compared to the  
370 transformation proposed by Tüxen and Ellenberg (1937), the PM transformation values  
371 were marginally better. Tuxen and Ellenberg (1937) transformation values tended to result  
372 in an overestimate summed cover of all growth forms (Appendix S7—Figures 2a-f); RMSD  
373 was consistently higher than PM transformations for all growth forms; slopes were further  
374 from 1 between 0.76 (forbs) and 0.95 (shrubs) and adjusted  $R^2$  ranged between 0.86  
375 (others) and 0.98 (trees) (Appendix 6—Table 1).

376 Evaluation of total cover, using the PM transformation values, showed RMSD was less than  
377 that estimated if the transformation was undertaken using estimates of Tüxen and Ellenberg  
378 (20.54 *cf.* 27.01) (Appendix 6—Table 1) and PM transformation values show a slight  
379 underestimation (slope = 1.1; adjusted  $R^2$  = 0.98) when tested on the independent dataset.

380 Scatter plots showing the relationships between visual estimates of summed cover for all  
381 six growth form groups using the PM model and for Tüxen and Ellenberg (1937)  
382 transformations are provided in Appendix S7—Figures 1a-f and Figures 2a-f.

### 383 **Discussion**

384 Transforming ordinal data to a quantitative form is common practice in plant ecology and  
385 extends across disciplines including restoration (Fill et al. 2017), classification (Cawsey et  
386 al. 2002; Faber-Langendoen et al. 2007; Wiser & De Cáceres 2013); and for assessing  
387 disturbance (Scott & Kirkpatrick 2008; Knapp & Ritchie 2016). Similarly, universal skewed  
388 patterns in the species abundance distribution are a long standing and well recognised  
389 pattern in ecology (e.g. MacArthur 1960). The data we present here are no exception. Yet  
390 the integration of these two concepts, underpinned by a robust modelling approach has  
391 received little attention, especially in the context of synthesizing information on aggregate  
392 properties of vegetation data. We demonstrate, using two large quantitative independent

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393 datasets that when the underlying right-skewed cover distribution is accounted for, a more  
394 robust set of transformations are generated. Where the aggregate properties of floristic data  
395 are of interest, our method, unlike previous approaches to transformation of ordinal data,  
396 does not overinflate cover.

397 Where possible, we advocate that others replicate this approach and source continuous  
398 cover data, so that the means within each ordinal class can be estimated accounting for the  
399 underlying distribution. Ideally, the continuous cover datasets will encompass the same  
400 temporal and spatial variation as that of the ordinal data. Notwithstanding these  
401 recommendations for best-practice, we have demonstrated our modelling approach can  
402 produce robust estimates of summed cover using floristic data from geographically distinct  
403 dataset containing observations of entirely unrelated vegetation communities. We expect  
404 the estimates of summed cover would further improve had we used representative data  
405 from that region and vegetation to model specific estimates of the parameters for the beta  
406 distribution. Undoubtedly there will be circumstances where appropriate continuous data  
407 will not be available and the parameters of the beta distribution cannot be estimated for a  
408 specific study or region. In these situations, adopting the PM transformations provided in  
409 Tables 1 and 2 would be preferable to application of ordinal class midpoints. When plant  
410 cover are right-skewed, midpoint transformations will bias and overestimate total cover.

411 Hierarchical models are useful for handling complex interactions in observational data.  
412 Despite the size of the initial dataset, some plant groups were poorly represented in the  
413 higher cover classes. By appropriately specifying the hierarchical model, estimates for  
414 these combinations could still be obtained, because they draw from the full model structure.  
415 We have identified that different growth forms have different cover distributions. Our  
416 empirical evidence strongly suggests that in plots where there are many small entities from  
417 the same growth form, such as for forbs and grasses, the cumulative cover of that growth

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418 form (when derived from transformations of ordinal data) may amplify and inaccurately  
419 describe the structural complexity of vegetation communities. Identifying and accounting for  
420 these distributions in other grouped entities has the potential to further improve summed  
421 cover estimates.

422 We also note potential observer bias for cover estimates of 5%. We acknowledge that  
423 visual estimates of cover and counts are subject to inter- and intra-operator error and bias  
424 (see Morrison 2016 for review) and this may account for the data digressions from a  
425 smooth shaped abundance curve. This is no doubt an artefact of observer preference for  
426 regular intervals when estimating cover, rather than a true representation of plant cover. In  
427 our case study analyses, the high frequency of estimates of 5% cover required a split-  
428 model approach where cover data were treated in two separate models. Given this right-  
429 skewed distribution and potential bias among disparate datasets (Appendix S3—Figures 1-  
430 4), we propose the split-model approach may serve wider applications. Simulated beta  
431 distribution data may not be entirely appropriate when using visually-estimated cover data,  
432 but may be useful where other less subjective methods for estimating cover are used (such  
433 as point-intercept methods). Given that visual estimates of cover-abundance are the  
434 assessment protocol for many floristic surveys, our approach offers a way these data can  
435 still be transformed and used with greater confidence, despite the underlying variability and  
436 bias. The approach we outline here can rapidly generate robust and defensible  
437 transformation estimates that are less prone to inflating summed cover estimates.

438 We envisage that our method may be useful when combined with emerging technologies  
439 such as 3-dimensional LiDAR or radar sensors that can penetrate vegetation canopies and  
440 assess complex structural elements. Furthermore, where large-scale biodiversity  
441 assessments, that rely on terrestrial vegetation as indicators of change, seek to integrate  
442 site observations to validate or train imagery, vegetation cover data collected in an ordinal

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443 scale will be of little benefit. Summation of midpoint transformations have been widely used  
444 in remote-sensing applications, no doubt often resulting in overinflated cover estimates.  
445 While transformations derived from a beta distribution will reduce the problem of over  
446 inflation, summed cover estimates can still exceed 100%. Jennings *et al.* (2009) offer an  
447 approach to rescaling summed cover so that cover estimates do not exceed 100% and their  
448 approach may be useful where site-based data are integrated to inform remote sensing  
449 applications.  
450 Our approach to transforming ordinal estimates of cover using a beta distribution can  
451 extend the application of these data beyond the realm of vegetation classification and can  
452 salvage information from many millions of floristic records. We expect most large  
453 repositories of floristic data will contain cover estimates with multifarious and nuanced  
454 ordinal scales. Here we provide a method that can be applied to floristic data in different  
455 ordinal scales for transforming and integrating datasets with much greater confidence. We  
456 have demonstrated a pan-continental approach to transforming ordinal cover estimates  
457 needed to build robust and accurate aggregated cover estimates. We foresee this approach  
458 supporting the synthesis of multiple datasets containing legacy data collected in different  
459 ordinal scales, especially where the aggregate properties of vegetation cover for different  
460 plant groups are of interest. These transformations and the resultant aggregated properties  
461 of cover data can support a multitude of uses in ecology from site-scaled, to landscape-  
462 scaled and for global applications.

463

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473

#### 474 **Author Contribution Statement**

- 475 • conceived the ideas and designed methodology – ALL
- 476 • extracted the case study and validation data – MM
- 477 • designed predictive model – JD
- 478 • analysed and interpreted the data – ALL
- 479 • led the writing of the manuscript – MM
- 480 • contributed critically to drafting and revising the manuscript and gave final approval  
481 for publication – ALL

482

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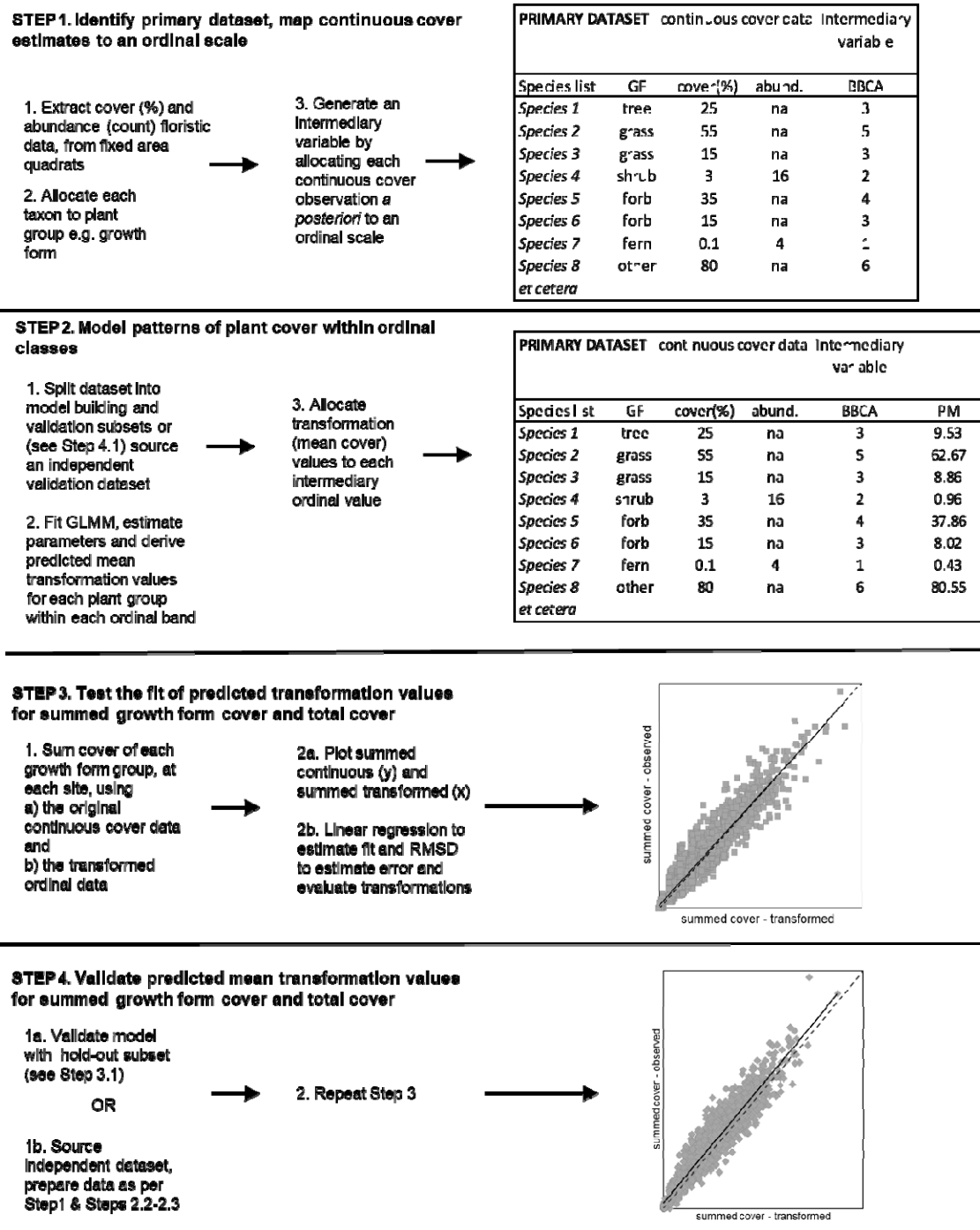
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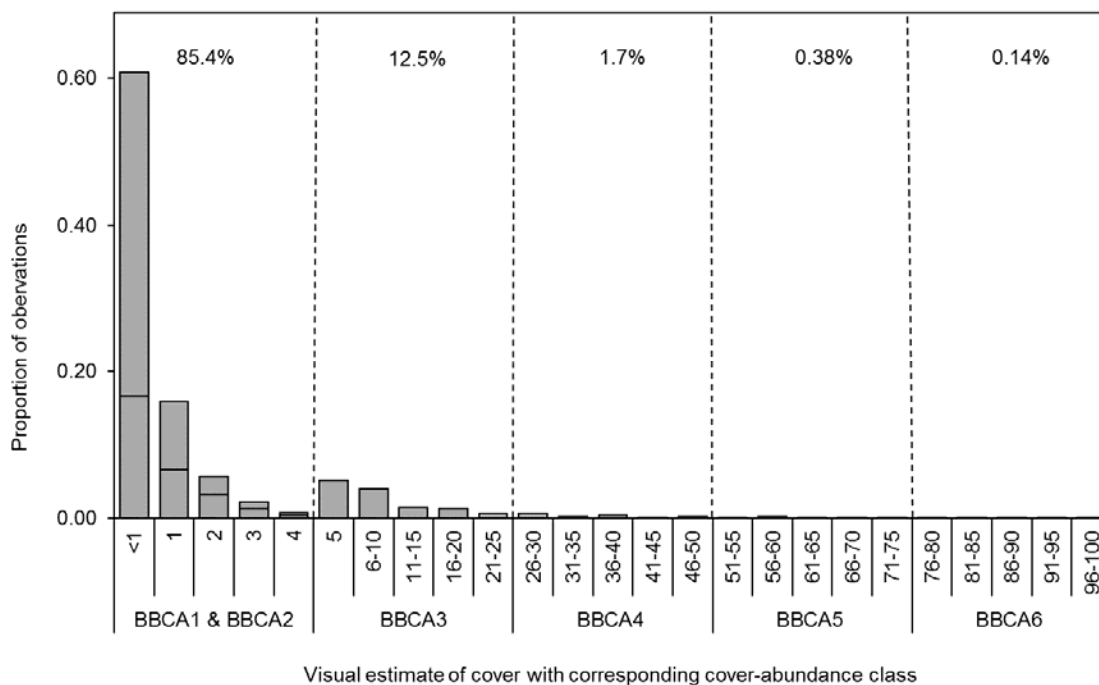
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- 621



1

2

3 Figure 1: Workflow showing the major elements required to estimate transformation  
 4 values for ordinal data using continuous cover estimates. Here we use Braun-  
 5 Blanquet cover-abundance (BBCA) 1-6 scale, although this approach could be  
 6 extended to any ordinal scale. Note, this flow diagram represents data from one plot,  
 7 but many plots are needed to obtain robust estimates of mean cover.



8

9 Figure 2: Distribution of visual estimates of cover for 95 812 observations, and their  
10 corresponding Braun-Blanquet cover-abundance (BBCA) class for our case study.

11 Dashed vertical lines show cut points between each BBCA class. Number of  
12 observations (n) for BBCA1 (n = 54 811); BBCA2 (n = 26 968); BBCA3 (n = 11 946);  
13 BBCA4 (n = 1 583); BBCA5 (n = 366) and BBCA6 (n = 138). Numbers between the  
14 dashed lines show the percentage of each class in the dataset. BBCA1 and BBCA2  
15 (both represent <math><5\%</math> cover) are shown as stacked histograms; BBCA1 ( $\le 10</math>  
16 individuals) sits above BBCA2 ( $> 10</math> individuals). See Appendix S3—Figures 1-4 for  
17 comparison with other archival datasets.$$



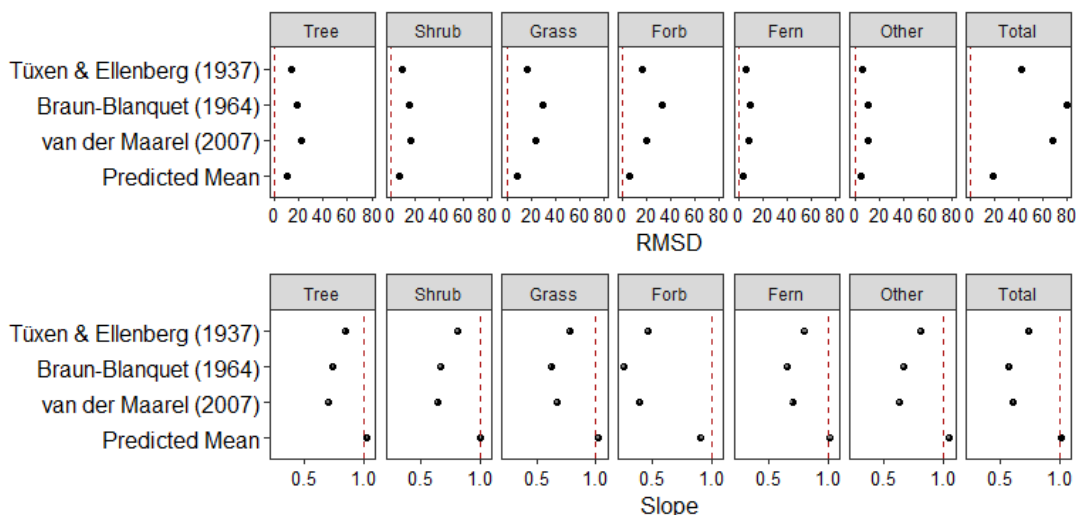


Figure 3: Results of linear regression with zero-intercept to compare root mean squared deviation (RMSD) and slope for each growth form and for total summed cover under previous transformations compared to the predicted mean transformation. Lower and upper 95% confidence intervals are shown for slope. Vertical dashed lines represent the perfect regression fit where  $RMSD = 0$  and slope = 1. (data table supplied in Appendix S4—Table 1). Number of observations ( $n$ ) for trees ( $n = 11\,953$ ); shrubs ( $n = 18\,545$ ); grasses ( $n = 21\,575$ ); forbs ( $n = 30\,017$ ); ferns ( $n = 3\,671$ ); other ( $n = 10\,051$ ) and total ( $n = 2\,809$ ). See Appendix S5 — Figures 1a-f to 4a-f for plots of all growth forms and three previous approaches to transformation proposed by Tüxen and Ellenberg (1937); Braun-Blanquet (1964) and van der Maarel (2007).