1	Self-reported Health is Related to Body Height and Waist Circumference in
2	Rural Indigenous and Urbanized Latin-American Populations
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20 Abstract

21 Body height growth is a life history component. It involves important costs for its expression and 22 maintenance, which may originate trade-offs on other costly components such as reproduction or 23 immunity. Although previous evidence has supported the idea that human height could be a sexually 24 selected trait, the explanatory mechanisms that underlie this selection is poorly understood. Moreover, 25 despite the association between height and attractiveness being extensively tested, whether immunity 26 might be linking this relation is scarcely studied, particularly in non-Western samples. Here, we tested 27 whether human height is related to health measured by both, self-perception, and relevant nutritional and 28 health anthropometric indicators in three Latin-American populations that widely differ in 29 socioeconomic and ecological conditions: two urbanized samples from Bogota (Colombia) and Mexico 30 City (Mexico), and one isolated indigenous population (Me'Phaa, Mexico). Using Linear Mixed 31 Models, our results show that, for both men and women, self-rated health is best predicted by an 32 interaction between height and waist, and that the costs associated to a large waist circumference are 33 differential for people depending on height, affecting taller people more than shorter individuals in all 34 population evaluated. The present study contributes with information that could be important in the 35 framework of human sexual selection. If health and genetic quality cues play an important role in human 36 mate choice, and height and waist interact to signal health, its evolutionary consequences, including its 37 cognitive and behavioral effects, should be addressed in future research.

38

39 Introduction

40	In modern Western societies, it has been seen that while women usually show a marked
41	preference for men significantly taller, over significantly shorter, than average [1,2], men are more
42	tolerant in choosing women who are taller or shorter than average [3]. This is consistent with the idea
43	that male height can be adaptive [4] and that sexual selection favors taller men, possibly because it
44	provides hereditary advantages, such as genetic quality for the offspring [5,6], or direct benefits,
45	provisioning resources and protection for women and their children [7]. This because height has been
46	proposed as an indicator of resource holding potential (RHP), in terms of social dominance and
47	deference [8,9], and socioeconomic status [5,10].
48	Supporting this idea, it has been found a direct linear relationship between male height and
49	reproductive success, which would not apply to women, and suggest unrestricted directional selection,
50	that would work to favor even very tall men, but not to very tall women [11]. In fact, it has been
51	reported that taller men (but not extremely tall men) are more likely to find a long-term partner and have
52	several different long-term partners [12], while the maximum reproductive success of women is below
53	female average height [13]. Furthermore, heterosexual men and women tend to adjust the preferred
54	height of hypothetical partners depending on their own stature [14]. In general, heterosexual men and
55	women prefer couples in which the man is taller than the woman, and women show a preference for
56	facial cues that denote a taller man [15].
57	Although previous evidence has supported the idea that human height could be a sexually
58	selected trait, the explanatory mechanisms that underlie this selection is poorly understood.
59	One possibility can be addressed in the framework of the Life-History theory [16], and the

60 immunocompetence handicap hypothesis (ICHH [17–19]). Body height growth is a life history

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component [1,20], that involves important costs for its expression and maintenance, which may originate
trade-offs on other costly components such as reproduction [21] or immunity [22].

63 The costs in height can be measure in terms of survival and physiological expenditure [22]. For 64 example, it has been shown that shorter people are more likely to be more longevous and less likely to 65 suffer from age-related chronic diseases [22,23]. With some exceptions, we have a limited number of 66 cell replications during our lifetimes. A minimal increment in body height necessary involves more 67 cells, maybe trillions, and more replications during the life. This higher number of cell replications 68 demands greater number of proteins to maintain taller, larger bodies [22], which together with an 69 increase on free radicals generated by the corresponding energy consumption, may lead to greater 70 likelihood of DNA damage [24], thus increasing the incidence of cancer and reducing longevity [22].

71 Trade-offs between these life-history components could be mediated by sexual hormones. Trade-72 off with reproduction occurs because at the beginning of sexual maturity sexual hormones are 73 responsible to reallocate energetic and physiological resources to this function, instead of somatic 74 growth. For instance, an increment in estrogen production leads to the onset of menstrual bleeding in 75 women, but also slows the process of growth, and eventually causes it to cease [25]; estrogen stimulates 76 mineral deposition in the growth plates at the ends of the long bones, thus terminating cell proliferation, 77 and resulting in the fusion of the growth plates to the shaft of the bone [26, see also 27]. In turn, trade-78 off with immunity occurs because the same increment in sexual steroids, usually has suppressive effects 79 on several immune components [17]. For example, testosterone may increase the severity of malaria, 80 leishmaniasis, amebiasis [28], and perhaps tuberculosis [see 29,30].

81 Therefore, as consequence of these life-history trade-offs, height could be considered as a 82 reliable indicator of individuals' condition in terms of (1) the amount and quality of nutritional resources 83 that were acquired until sexual maturity, (2) the RHP to obtain resources for the somatic maintenance in

adult stage, and (3) the current immunocompetence to afford the immune cost imposed by sexual
steroids. Thus, according with ICHH height can be used for potential partners to receive information
about the quality of potential mate; only high-quality individuals could afford to allocate resources to
better immunity and attractive secondary sexual traits simultaneously [18], which would result in
increased sexual preference towards taller individuals.

89 Despite the association between height and attractiveness being widely tested, whether immunity 90 might be linking this relation is poorly studied. Moreover, most studies have been done using high-91 income developed populations (often samples characterized as Western, Educated, and from 92 Industrialized, Rich, and Democratic [WEIRD] societies [31]), which has led to a lack of information of 93 what is occurring in other populations with important socio ecological differences. Considering these 94 ecological pressures is important because although genetic allelic expression could be the main factor 95 that determines individual height differences [25], height is also the most sensible human anatomical 96 feature that respond to environmental and socioeconomic conditions [21,32]. For instance, variation in 97 height across social classes is known to be greater in poorer countries [33], but much reduced where 98 standards of living are higher [34]. Economic inequality not only affects population nutritional patterns, 99 which are especially important during childhood to stablish adult height, but also the presence of 100 infectious diseases [35]. Childhood disease is known to adversely affect growth: mounting an immune 101 response to fight infection increases metabolic requirements and can thus affect net nutrition, and hence 102 reduce productivity. Disease also prevents food intake, impairs nutrient absorption, and causes nutrient 103 loss [36,37]. Therefore, comparing with high-income, developed populations, habitants from sites with 104 stronger ecological pressures imposed by pathogens, or greater nutritional deficiencies, would face 105 greater costs to robustly express this trait, and in consequence could show a stronger sexual selective 106 pressure over height, since it would more accurately signal growth rates, life-history trajectories, and

107	health status. This phenotypic variation is described as developmental plasticity, which is a part of the
108	phenotypic plasticity related to growth and development, in response to social, nutritional, and
109	demographic conditions, among others [38]. In fact, during the last century, and given a general
110	improvement in nutrition, height has increased around the world [39], but maintaining the level of
111	dimorphism in favor of men.
112	Colombia and Mexico are two of the most socioeconomically heterogeneous countries in the
113	world; although both countries have a high Human Development Index [40], and have relatively good
114	health compared to global standards, attaining respective scores of 68 and 66 in the Healthcare Access
115	and Quality (HAQ) Index [41], Colombia and Mexico have GINI coefficients of 50.8 and 43.4,
116	respectively, making them the 12 th and 43 th most unequal countries in the world (GINI index – World
117	Bank estimate; https://data.worldbank.org/indicator/SI.POV.GINI). These national-level statistics,
118	however, hide important within-country differences. In particular, in Latin-America people in rural areas
119	tend to be poorer and have less access to basic services such as health and education than people in
120	urban areas.
121	According to data from the World Bank and the Colombian National Administrative Department
122	of Statistics, in 2017 Colombia was the second most unequal country in Latin-America after Brazil; in
123	rural areas 36% of people were living in poverty, and 15.4% in extreme poverty, while in urban areas
124	these values were only 15.7% and 2.7%, respectively [for a summary, see 42].
125	In addition to rural communities, in Latin-America, indigenous people tend to have high rates of
126	poverty and extreme poverty [43], and have poorer health [44] less susceptible to improve by national
127	income growth [45]. In Mexico, there are at least 56 independent indigenous peoples, whose lifestyle
128	practices differ in varying degrees from the typical "urbanized" lifestyle. Among these groups, the
129	Me'Phaa people, from an isolated region known as the "Montaña Alta" of the state of Guerrero, is one
	6

130 of the groups whose lifestyle most dramatically differs from the westernized lifestyle typical of more 131 urbanized areas [46]. Me'Phaa communities are small groups, composed of fifty to eighty families, each 132 with five to ten family members. Most communities are based largely on subsistence farming of legumes 133 such as beans and lentils, and the only grain cultivated is corn. Animal protein is acquired by hunting 134 and raising some fowl, but meat is consumed almost entirely during special occasions and is not part of 135 the daily diet. There is almost no access to allopathic medications, and there is no health service, 136 plumbing, or water purification system. Water for washing and drinking is obtained from small wells. 137 Most Me'Phaa speak only their native language [47]. In consequence, these communities have some of 138 the lowest income and economic development in the country, and the highest child morbidity and 139 mortality due to chronic infectious diseases [46]. 140 These three Latin-American populations can provide an interesting indication about how 141 regional socioeconomic conditions, and the intensity of ecological pressures by pathogens, may 142 modulate the function of height as an informative sexually selected trait of health and individual 143 condition. Therefore, the aim of the present study was to evaluate whether human height is related to 144 health measured by both, self-perception, and relevant nutritional and health anthropometric indicators 145 in three Latin-American populations that widely differ in socioeconomic and ecological conditions: two 146 urbanized samples from Bogota (Colombia) and Mexico City (Mexico), and one isolated indigenous 147 population (Me'Phaa, Mexico).

148 Materials and Methods

149 **Ethics Statement**

All procedures for testing and recruitment were approved by El Bosque University Institutional Committee on Research Ethics (PCI.2017-9444) and National Autonomous University of Mexico Committee on Research Ethics (FPSI/CE/01/2016). All participants read and signed a written informed consent.

154 **Participants**

A total of 251 (120 women and 131 men) adults took part in the study. They were from three

156 different samples: (1) Mexican indigenous population, (2) Mexican urban population, and (3)

157 Colombian urban population.

158 The first sample consisted of 75 subjects (mean age \pm SD = 33.60 \pm 9.51 years old) from the 159 small Me'Phaa community - "Plan de Gatica" from a region known as the "Montaña Alta" of the state 160 of Guerrero in Southwest Mexico. In this group, 24 participants were women (33.46 ± 8.61) and 39 were 161 men (33.74 ± 10.41) , who were participating in a larger study about immunocompetence. Both sexes 162 were aged above 18 years old. In Mexico, people from this age is considered as Adult. We collected all 163 measurements in the own community. Me'Phaa communities are about 20 kilometers apart, and it takes 164 about three hours traveling on rural dirt roads to reach the nearest large town, about 80 km away. 165 Mexico City is about 850 kilometers away and the trip takes about twelve hours by road. This 166 community has the lowest income in Mexico, the highest index of child morbidity and mortality by 167 gastrointestinal and respiratory diseases (children's age from 0 to 8 years old, which is the highest 168 vulnerability and death risk age; [46]), and the lowest access to health services. These conditions were 169 determined by last 10 years of statistical information obtained from the last record of the national system 170 of access to health information in 2016 [46].

171	The second sample consisted of 66 subjects (20.67 \pm 2.32) over 18 years old of general
172	community from Mexico City, of whom 36 were women (20.2 ± 2.27) and 30 were men (21.13 ± 2.36).
173	Finally, the third sample consisted of 122 undergraduate students with ages ranging from 18 to 30 years
174	old (30.23 \pm 4.27), 60 were women (20.2 \pm 2.27) and 62 were men (21.13 \pm 2.36) from Bogota,
175	Colombia. All urban participants were recruited through public advertisements.
176	Participants from both urban population samples were taking part in two different, larger studies
177	in each country. In Colombia, all data were collected in the morning, between 7 and 11 am, because
178	saliva samples (for hormonal analysis), as well as voice recordings, odor samples, and facial
179	photographs, were also collected as part of a separate project. Additionally, women in the Colombian
180	sample were not hormonal contraception users, and all data were collected within the first three days of
181	their menses.

Participants who were under allopathic treatment, and hormonal contraception female users from
both countries were excluded from data collection. All participants completed a sociodemographic data
questionnaire, which included medical and psychiatric history.

185 **Procedure**

All participants signed the informed consent and completed the health and background questionnaires. For participants from the indigenous population, the whole procedure was carried out within their own community, and participants from the urban population attended a university laboratory from each country on individual appointments.

190 First, participants were asked to complete the health and sociodemographic data questionnaires.191 Subsequently the anthropometric measurements were taken.

192Self-reported health

193	We used a Span	sh language val	lidated version	of the SF-36 c	uestionnaire [4	481.	The used	version

- 194 was validated in Colombia [49]. The SF-36 produces eight factors, calculated by averaging the recoded
- scores of individual items: 1) Physical functioning (items 3 to 12), 2) Role limitations due to physical
- health (items 13 to 16), 3) Role limitations due to emotional problems (items 17 to 19), 4)
- 197 Energy/fatigue (items 23, 27, 29 and 31), 5) Emotional well-being (items 24, 25, 26, 28 and 30), 6)
- 198 Social functioning (items 20 and 32), 7) Pain (items 21 and 22), and 8) General health (items 1, 33, 34,
- 199 35 and 36).

200 To calculate this factors, all items were recoded following the instructions on how to score SF-36

201 [48].We calculated final factor averaging the recoded items. To make this data compatible with the

202 Mexican database, and because item 35 cannot be answered by the Mexican Indigenous population, this

item was excluded and the health factor was calculated averaging only items 1, 33, 34, and 36.

204

Anthropometric measurements

All anthropometric measurements were measured three times, consecutively, and then averaged (for agreement statistics between the three measurements of each characteristic, see section 1.3 on S1 File). All participants were in light clothes and had their shoes removed. The same observer repeated the three measurements.

We measured the body height in centimeters, to the nearest millimeter, using a 220cm Zaude stadiometer, with the participant's head aligned according to the Frankfurt horizontal plane and with feet together against the wall.

Anthropomorphic measurements also included waist circumference (cm), weight (kg), fat
 percentage, visceral fat level, muscle percentage, and BMI. Circumference of waist was measured in

214 centimeters using a flexible tape, midway between the lowest rib and the iliac crest, and was recorded to 215 the nearest millimeter. These anthropomorphic measures have been used as an accurate index of 216 nutritional status and health, especially waist circumference. Metabolic syndrome is associated with 217 visceral adiposity, blood lipid disorders, inflammation, insulin resistance or full-blown diabetes, and 218 increased risk of developing cardiovascular disease [50,51, for a review see 52], including Latin-219 American populations [53]. Waist circumference has been proposed as a crude anthropometric correlate 220 of abdominal and visceral adiposity, and it is the simplest and accurate screening variable used to identify people with the presence of the features of metabolic syndrome [54,55]. Hence, In the presence 221 222 of the clinical criteria of metabolic syndrome, an increased waist circumference does provide relevant 223 pathophysiological information insofar as it defines the prevalent form of the syndrome resulting from 224 abdominal obesity [51].

Weight (kg), fat percentage, visceral fat level, muscle percentage and BMI were obtained using an Omron Healthcare HBF-510 body composition analyzer, calibrated before each participant's measurements were obtained.

228 Statistical analysis

To test the association between height and health, we fitted general a Linear Mixed Model (LMM). The dependent variable in this model were the self-reported health factor and the predictor variables included participant sex, age, population (indigenous, urban), height and waist as fixed, main effects, as well as anthropometric measurements (hip, weight, fat percentage, BMI and muscle percentage). Interactions between height and population, height and sex, and height and waist circumference were also included. Country was always included as a random factor, with random intercepts.

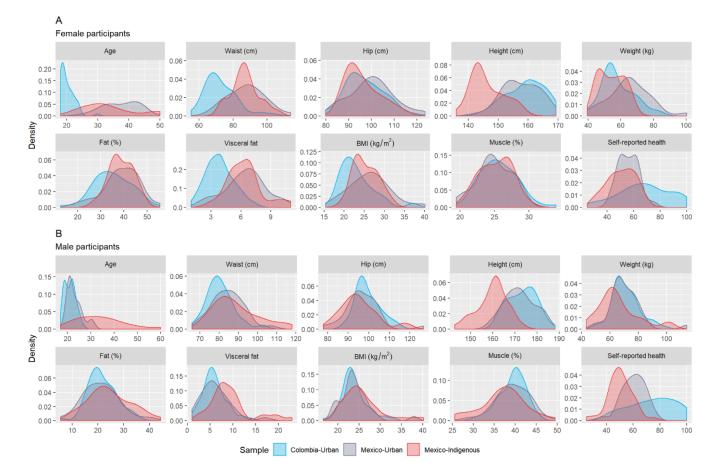
236	Although allowing slopes to vary randomly is recommended [56], we only included random
237	intercepts in the models because there is only one data-point per subject. Population (indigenous, urban)
238	was always included as a fixed effect because while there are important differences in health (and self-
239	reported health) between indigenous and urban populations in Latin-America, while no such differences
240	were expected by country. General LMM were fitted to test residual distribution. In all cases, residuals
241	were closer to a normal or gamma (inverse link) distribution, for each population/country. Models was
242	fitted using the <i>lmer</i> function from the <i>lmerTest</i> package [57;
243	https://www.rdocumentation.org/packages/lmerTest] in R, version 3.5.2 [58].
244	The most parameterized initial model was then reduced based on the Akaike Information
245	Criterion (AIC) and the best supported model (i.e. the model with the lowest AIC with a Δ AIC higher
246	than 2 units from the second most adequate model) is reported [see 59]. To accomplish this, we
247	implemented the ICtab function from the bbmle package [60;
248	http://www.rdocumentation.org/packages/bbmle]. Once a final model was selected, model diagnostics
249	were performed (collinearity, residual distribution, and linearity of residuals in each single term effect;
250	see section 3 in S1 File).

251 **Results**

All analysis, data manipulation, tables and figures, as well as the code to produce them, can be reproduced and explored in more detail using *R* scripts in *Markdown* format (S2 File) using the are available as Supplementary Files, as well as the output, S1 File (in HTML format), where all Supplementary tables and figures can also be found. All data are available at the Open Science Framework (https://doi.org/10.17605/OSF.IO/KGR5X).

257 Figure 1 shows the distribution of age, waist, height, visceral fat and self-reported health, which

strongly varies in both women (Fig 1A) and men (Fig 1B), sex, population (indigenous, urban) and





260

261 *Fig 1.* Distribution of all measured variables by sex, population and country. (A) Female participants. (B) Male

participants. For descriptives (mean, SD, median, minimum, and maximum values), see S2 Table (female participants) and
S3 Table (male participants), of the Supplemental Material.

264

265 To establish the relationship between height and self-reported health, we fitted three mixed266 models (Table 1).

267

	Model 1		Ν	Iodel 2		Model 3			
	Estimate	df	р	Estimate	df	р	Estimate	df	р
(Intercept)	-97.01	226.83	0.520	-166.17	233.81	0.198	-181.41	234.65	0.153
Age	0.07	224.16	0.660	0.11	231.11	0.488			
BMI (kg/m ²)	-0.03	226.02	0.990						
Fat (%)	-0.21	226.00	0.650						
Height (cm)	1.13	226.58	0.240	1.49	233.27	0.064	1.59	234.01	0.043
Height:PopulationUrban	0.30	226.00	0.300						
Height:SexMale	0.02	226.01	0.930						
Height:Waist	-0.02	226.37	0.180	-0.02	233.25	0.064	-0.02	234.00	0.041
Hip (cm)	-0.05	226.98	0.830						
Muscle (%)	-0.32	226.81	0.570						
PopulationUrban	-38.67	226.02	0.400	8.42	233.98	0.009	8.24	234.38	0.010
SexMale	3.18	226.09	0.940	6.01	233.07	0.034	5.82	234.00	0.039
Waist (cm)	2.60	226.19	0.220	2.66	233.18	0.094	2.91	234.01	0.061
Weight (kg)	0.03	226.06	0.970						

268 Table 1. Results of separate linear mixed models testing effects of independent variables on self-269 reported health.

270 271

Note. Indigenous population and females were used as reference for categorical predictors. Significant effects are in **bold**. For a full version of this table, including standard errors and t-values, see S7 Table, and for an ANOVA-like table of random 272 effects, see S8 Table in the Supplemental Material, available online.

273

274 In the first model we included, as predictors, all measured variables as main effects, as well as 275 the interactions between height and population, height and sex, and height and waist. In the second 276 model, we included age, height, population, sex, waist, and the interaction between height and waist. For 277 the final, third model, we removed age since this predictor did not have any influence on self-reported 278 health factor in the previous models.

279 These three models were compared using the Akaike Information Criterion (AIC) as well as

280 Akaike weights (w_i AIC), and Δ AIC (Table 2). The analyses revealed that Model 3 is not only the most

281 parsimonious model, but has a lower AIC and higher Akaike weight [see 59] than the previous two

282 models; in fact, Model 3 is 5.66 times more likely to be the best model compared to Model 2, and more

283 than 4000 compared to Model 1 (in comparison to Model 1, Model 2 is close to 750 times more likely to

284 be the best model).

285

286 Table 2. Performance criteria of LME models.

Model	AIC	ΔΑΙC	df	$w_i(AIC)$
Model 3	1981.4		8	0.85
Model 2	1984.87	3.47	9	0.15
Model 1	1998.09	16.69	16	< 0.001

288

290

287 *Note.* Models are in descending order from the best, to the worst fitting. ΔAIC is the change in AIC between each model and the previous. Akaike weights w_i (AIC) are conditional probabilities for each model being the best model [59]. 289

Nevertheless, for Model 3 (the minimum adequate model), Variance Inflation Factors (VIF) 291 revealed extreme collinearity for height, waist, and the interaction between height and waist (VIF > 75292 in those cases; S9 Table). This problem, however, has solved after centering and rescaling both height

293 and waist measures (VIF < 3 in all cases; S10 Table). In addition, this centered and rescaled version of

294 Model 3 had no issues regarding its residual distribution (i.e. for all samples it resembled a normal

295 distribution) or linearity of residuals (see S2 Fig), and each single term predictor was linearly related to

296 self-rated health (see S3 Fig).

297 Furthermore, the final, centered and rescaled version of Model 3, had a lower AIC than model 3

298 (1962 vs 1981), and was over 1400 times more likely to be the best model, as revealed by Akaike

299 weights (see S11 Table).

300 The final model (Table 3; Fig 2) showed a significant, negative main effect waist circumference

301 $(t = -3.01, \beta = -3.27, p < 0.001)$, as well as a significant effect of population (urban samples rated their

302 health 8.24 points higher than indigenous participants; t = 2.60, p = 0.01), and sex (men rated their

303 health 5.82 points higher than women t = 2.07, p = 0.039). In addition, this model (Table 3) revealed that

304 Colombians reported better heath than Mexicans (Fig 2B).

305

Table 3. Results of the final linear mixed model testing effects of independent variables on self reported health

	Estimate	SE	df	t	р
(Intercept)	53.64	6.2	1.88	8.65	0.016
Height_cs	-0.17	1.57	234.16	-0.11	0.914
Waist_cs	-3.27	1.08	234.29	-3.01	0.003
SexMale	5.82	2.81	234	2.07	0.039
PopulationUrban	8.24	3.17	234.38	2.6	0.01
Height_cs:Waist_cs	-2.28	1.11	234	-2.06	0.041

308 *Note.* Indigenous population and females were used as reference for categorical predictors. Significant effects are in bold.

309 Both waist and height were centered and rescaled (identified by the suffix _cs).



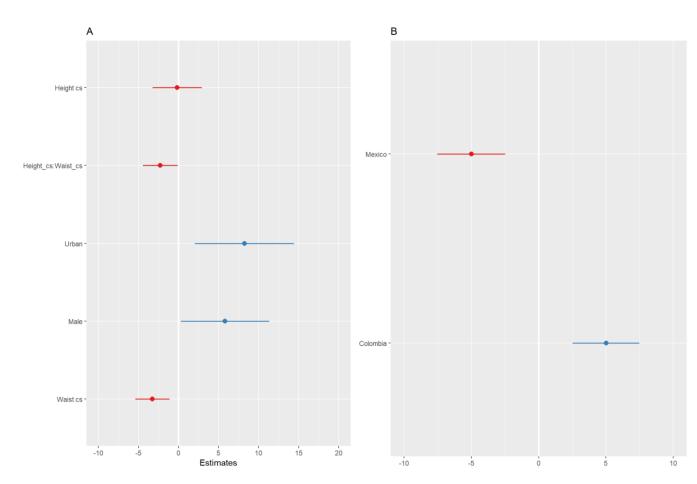
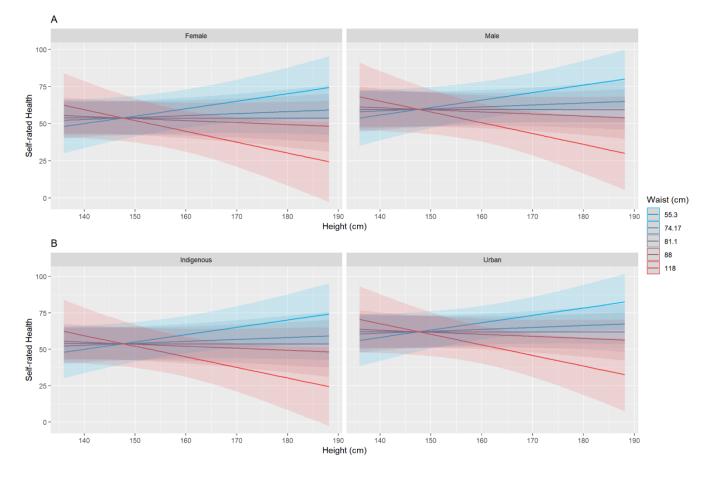




Fig 2. Final model estimates. Forest-plot of estimates for each fixed factor with 95% CI. (A) Fixed effects. (B) Random
effects. For categorical fixed predictors, indigenous population and female participants were used as reference. Both waist
and height were centered and rescaled (identified by the suffix _cs).

315

316	Moreover, a significant interaction between waist and height (Table 3; $t = -2.06$, $p = 0.041$) was
317	exposed, indicating that the associated health costs of a larger waist circumference were different for
318	people of different heights (Fig 3); the best predicted self-rated health was for tall participants with small
319	waists, and the worst was for (again) tall participants, but with large waist circumferences. The model
320	also revealed that for shorter people, there are no predicted significant associated costs of having a large
321	waist. In other words, the association between height and self-rated health is positive for people with
322	small waist circumferences, but negative for people with large waists.
323	In addition, age, waist circumference, height, visceral fat, BMI, and muscle percentage, were
324	significantly correlated with self-rated health ($r > 0.20$, in all cases), for men and women (for bivariate
325	Pearson correlations between all measured variables see S4 Table for all participants combined, S5
326	Table for women, and S6 Table for men).



327

Fig 3. Interaction between height and waist. Model predictions were split by (A) sex, and ((B) population. To simplify interpretation, raw (instead of centred and rescaled) values of height and waist were used. As waist reference, minimum, quartiles (lower, median and upper), and maximum waist circumference values were used, showed on a blue to red colour scale. For an interactive 3D plot of the interaction between height and waist, see S4 Fig, or the 3D animated version contained in S1 File.

333 **Discussion**

The present study provides new insights into the nature of the relationship between height and health, in both men and women, by studying three Latin American samples, which included urban and indigenous populations with marked differences in access to basic needs and services like food and health. 338 Contrary to our initial hypothesis, we did not find height by itself to be a significant predictor of 339 self-perceived health but by an interaction with waist circumference in all populations studied. Most 340 results in favor of a direct relationship between height itself and health were carried out more than 341 twenty years ago, in small samples, from modern societies, and in specific Western ethnic groups. New 342 studies with non-traditional population groups have failed to verify the positive relationship between 343 height and health, especially associated with cardiovascular and autoimmune diseases [61,62]. For 344 example, studies in groups of Native Americans, Japanese, Indians and Pakistanis showed that lower 345 people had a lower prevalence of cardiovascular disease than the highest people in each group [62]. 346 These findings were similar in a group of inhabitants of Sardinia, a European population with the lowest 347 physical stature recorded in Europe in recent years [61].

348 Interestingly, our results suggest that although there is a main effect of waist size on self-349 perceived health, the associated costs of a large abdominal circumference are differential depending on 350 stature; this is, waist circumference predicted self-reported health differently for people of different 351 heights: while being taller predicts better self-rated health for taller people with relatively small waists, 352 being taller was found to be associated with poorer perceptions of their of health in people with larger 353 waist circumferences. Furthermore, while there is a cost of abdominal and visceral adiposity for tall 354 people, there is no predicted cost for shorter persons. Therefore, these results argue the importance of 355 consider a phenotypic integration of different human features that could be involved in health or 356 physiological condition, when a possible sexually selected trait is being evaluated as a signal of 357 immunocompetence.

358 On the other hand, given that height is the most sensible human anatomical feature to 359 environmental and socioeconomic conditions [21,32], we expected stronger relation between health and 360 height for indigenous population, where the cost to produce and maintain this costly trait is greater than

361 for habitants from urbanized areas. Nevertheless, we did not find inter-population differences in the 362 magnitude of this relation, urban populations reported better health than the indigenous sample, and the 363 shortest participants tended to be from the indigenous Me'Phaa sample. These results could in fact 364 suggests different life history strategies. In harsh environments, compared to modern Western societies, 365 different life strategies could take place [63], like investing relatively less energy in growth and 366 reallocating it towards reproduction [21]. In addition, a relative increase in the intensity or number of 367 infectious diseases (including child disease, like in the case of the Me'Phaa) and a tendency to early 368 sexual maturity, could have negative effects on growth, resulting in lower average height values [64,65]. 369 These trends could be a compensation between life history components [25]. Finally, fast and prolonged 370 growth imply high costs for the organism [1]; rapid growth seems to influence mortality risk [66], and 371 growing for a longer time, delays the onset of reproduction, increasing the risk of dying and producing 372 fewer offspring [1]. This perspective of life strategies allows us to understand the relationship between 373 height, health, and reproduction. It suggests the importance of addressing factors such as ethnicity, 374 socio-economic status, level of urbanization, especially in populations where there is great heterogeneity 375 of access to food, health and pressure resources for pathogens, as in Latin American populations in 376 which this relationship has barely been directly explored.

Although our study did not directly evaluate any immunological marker but a self-perception of health, the implementation of a physiological immune indicator of adaptive immune system appears to be consistent with our results. It has been found that men but not women show a curvilinear relationship between antibody response to a hepatitis-B vaccine and body height, with a positive relationship up to a height of 185 cm, but an inverse relationship in taller men [19]. In our three populations, the maximum height was lower than 185 cm, which could explain the linear but not curvilinear relation found. In addition, the fact that self-perception in our study and antibody response in previous studies are both

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positively associated with body height could contribute to the knowledge about the reliability of self perception of health as an indicator of immunological condition.

386 Finally, in relation with sex differences women reported lower health in average than men in all 387 communities, which is concordant with reports and normative SF-36 data in other populations, and 388 especially in younger people [e.g. 67,68]. These results could add support to the idea that height is a 389 reliable signal of health in men [25], while for women it could reflect reproductive success [69] in terms 390 of labor and birth, and to a lesser extend function as an indicator of health [70]. It has been seen that 391 taller women experience fewer problem during this process, because of a lower risk of a mismatch 392 between fetal head size and the size of the birth canal [70]. Nevertheless, this idea is only speculative 393 and more studies comparing health, reproductive success and female height need to be done.

The present study contributes with information that could be important in the framework of human sexual selection. If health and genetic quality cues play an important role in human mate choice [e.g. 71], and height and waist interact to signal health, its evolutionary consequences, including its cognitive and behavioral effects, should be addressed in future research. This could be done by studying the interaction between waist circumference and height, in relation to reproductive and/or mating success, as well as mate preferences and perceived attractiveness, in samples with both Westernized and non-Westernized lifestyles.

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577 Supporting Information

- 578 S1 File. HTML output for R Markdown. This file contains the script and output for all analyses, data
- 579 manipulation and compilation, tables and figures. This file was created using R scripts in Markdown
- 580 format (Rmd file) to promote transparency and ensure reproducibility.
- 581 **S2 File. R Markdown source file for HTML output.** R Markdown file used to generate S1 File.
- 582 S1 Table. Intraclass correlation of anthropometric characteristics measurements.
- 583 S2 Table. Descriptive statistics of measured variables of female participants.
- 584 **S3** Table. Descriptive statistics of measured variables of male participants.
- 585 **S4 Table.** Correlations between measured variables for all participants
- 586 S5 Table. Correlations between measured variables for female participants
- 587 S6 Table. Correlations between measured variables for male participants

- 588 S7 Table. Results of separate linear mixed models testing effects of independent variables on self-
- 589 **reported health.** Full table including standard errors and *t*-values.
- 590 S8 Table. ANOVA-like table with tests of random-effect terms.
- 591 S9 Table. Variance Inflation Factors of Model 3 predictors.
- 592 S10 Table. Variance Inflation Factors of the Final Model (Model 3 centered and rescaled)
- 593 predictors.
- 594 S11 Table. Information criteria for Model 3 and Model 3 (centered and rescaled).
- 595 S1 Fig. Sexual dimorphism of height, waist and health for all samples (A) Self-perceived health. (B)
- 596 Height. (C) Waist. Comparisons between female and male participants for each sample, were performed
- using *t*-tests, adjusted for multiple tests. **** p < 0.0001.
- 598 S2 Fig. Model diagnostics. (A) Residual distribution for each sample. (B) Linearity in each (single
- term) fixed factor. Centered and rescaled variables are identified by the suffix _cs.
- 600 S3 Fig. Single term predictor slopes. Slope of coefficients for each (single term) fixed predictor,
- against self-rated health (linear relationship between each model term and response). For Population, 1 =
- Indigenous, and 2 = Urban. For sex, 1 = female, and 2 = male. For simplicity, raw (instead of centered
- and rescaled) values of height and waist were used.
- 604 S4 Fig. Interaction between height and waist (interactive, animated 3D version). For simplicity, raw
- 605 (instead of centered and rescaled) values of height and waist were used. Click and drag the plot to
- 606 change its orientation. Scroll to zoom. In S1 File, where this figure is also included, you can also use the
- 607 buttons below the figure to control the animation.