An association between the gut microbiota and immune cell dynamics in humans.

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ABSTRACT

The gut microbiota influences development and homeostasis of the mammalian immune system[1–3], can alter immune cell compositions in mice[4–7], and is associated with responses to immunotherapy that rely on the activity of peripheral immune cells[8–12]. Still, our understanding of how the microbiota modulates immune cells dynamics remains limited, particularly in humans where a lack of deliberate manipulations makes inference challenging. Here we study hundreds of hospitalized—and closely monitored—patients receiving hematopoietic cell transplantation as they recover from chemotherapy and stem cell engraftment. This aggressive treatment causes large shifts in both circulatory immune cell and microbiota populations, allowing the relationships between the two to be studied simultaneously. We analyzed daily changes in white blood cells from 2,235 patients, and 10,680 longitudinal microbiota samples to identify bacteria associated with those changes. Bayesian inference and validation across patient cohorts revealed consistent associations between gut bacteria and white blood cell dynamics in the context of immunomodulatory medications, clinical metadata and homeostatic feedbacks. We contrasted the potency of fermentatively active, obligate anaerobic bacteria with that of medications with known immunomodulatory mechanism to estimate the potential of the microbiota to influence peripheral immune cell dynamics. Our analysis establishes and quantifies the link between the gut microbiota and the human immune system, with implications for microbiota-driven modulation of immunity.

MAIN TEXT

Experiments in mice provide evidence that the mammalian intestinal microbiome influences the development[1–3] and homeostasis of its host's immune system[4–7,13–15]. In humans, inflammatory bowel diseases correlate with functional dysbiosis in the gut microbiota[16,17]. Children born preterm and at term have different gut microbiome compositions and differ in the development of immune cell populations in their blood[18]. The composition of the gut microbiota may also influence the success of immunotherapies[8–11]. Immune checkpoint inhibitor therapy relies on activation of circulating T-cells and its success has, independently, been associated with abundances of intestinal anaerobic genera such as *Akkermansia*[9] and *Faecalibacterium*[10]. It is therefore an intriguing prospect to augment treatments such as cancer immunotherapy[19], including the burgeoning field of chimeric antigen receptor (CAR) T-cell therapy[20], by leveraging microbiome-driven immune system modulation.

Our understanding of how the microbiota influences the dynamics of immune cells in humans, and how this compares to deliberate immunomodulatory interventions nevertheless remains limited. Experiments with animals may not always be sufficient to study mechanisms of microbiome-immune interactions and translate them to human biology as the microbial ecology in the gut of an animal model may be different from humans receiving treatment[21]. On the other hand, studies directly in patients may be criticized when they have small subject numbers, are cross-sectional, lack statistical power, or disregard key confounders such as medications[21].

To overcome these limitations, we conducted a large-scale longitudinal study of the gut microbiota and day-by-day changes in circulatory immune cell counts. We investigated immune reconstitution dynamics after allogeneic hematopoietic cell therapy (HCT) within all 2,926 patients who underwent HCT at Memorial Sloan Kettering for various hematological malignancies, including leukemia, between 2003 and 2019 (**Figure 1A, Table S1**). The conditioning regimen of radiation and chemotherapy administered to HCT patients is the most severe perturbation to the immune system deliberately performed in humans and thus offers a unique opportunity to investigate dynamic links between the gut microbiota and the immune system directly in humans. Conditioning depletes white blood cell counts (**Figure 1A**) and can lead to prolonged periods of neutropenia (<500 neutrophils per µl blood). Immune reconstitution begins after transplanted stem cells have matured sufficiently to

67   release granulocytes from the bone marrow (neutrophil engraftment is defined as 3 consecutive days

68   with >500 neutrophils per µl blood, **Figure 1A-C**). The blood of each patient is carefully monitored

Figure 1: Immune reconstitution and microbiome dynamics after allogeneic hematopoietic cell transplantation (HCT). A) Three major periods of HCT—immunoablation during chemotherapeutic conditioning before HCT, defined as day 0, post-HCT neutropenia, and reconstitution following neutrophil engraftment—lead to recovery trajectories with large variability between patients. Shown are the mean counts (shaded: ±1 standard deviation, σ) of neutrophils, lymphocytes and monocytes per day relative to HCT from patients transplanted between 2003 and 2019 (A), contrasted with two individual patients (B,C) representative of the recovery trajectories for different stem cell graft source; patient 1 who received a peripheral blood stem cell graft, PBSC (line with circles: patient data, solid line and shaded region: mean±1 standard deviation of all PBSC patients), and patient 2 who received a graft of umbilical cord blood (line with circles: patient data, solid line and shaded region: mean±1 standard deviation of all cord patients). Fecal samples collected and analyzed by 16S rRNA gene sequencing reveal the loss of microbial diversity reported previously[23,24] (D, line: mean per day, shaded: ±1 standard deviation; E,F: individual patient measurements) and commensal families (G, line: mean relative abundances of bacterial families, shaded: ±1 standard deviation; H,I: individual patient measurements), often replaced by Enterococcaceae domination. J,K). Administration of immunomodulatory medications for the two example patients.

throughout this recovery, and medications are administered to modulate the immune cell dynamics,

including granulocyte-colony stimulating factor (GCSF) to increase neutrophil counts, and

immunosuppressants such as mycophenolate mofetil or tacrolimus to prevent complications such as

graft-vs-host disease (**Figure 1 J,K**). To investigate if the composition of the gut microbiota is

associated with the dynamics of circulating white blood cells, we analyzed detailed blood and clinical

metadata of our patients between 3 days before HCT and until 100 days post neutrophil engraftment

(excluding pediatric patients, and other exclusion criteria: N=2,235, supplementary methods, **Figure**

S1). During this period patients are monitored carefully, and our analysis included over 140,000 host

phenotype measurements in the form of complete blood counts which quantify the most abundant

white blood cells—neutrophils, lymphocytes, monocytes, eosinophils—as well as platelet counts

(**Figure 1, S1**). We started collecting patients' fecal microbiota data in 2009[22], and by now obtained

10,680 high frequency, longitudinal microbiota compositions.

HCT patients lose gut microbiota biodiversity and commensal microbial families during their

treatment (**Figure 1D-I**); this figure generated from N=1,294 HCT patients clarifies the preliminary

trends observed in previous studies from smaller datasets[23,24]. We have shown recently that mice with

a depleted intestinal flora had worse recoveries of white blood cells after bone-marrow

transplantation[25]. In our patients, microbial diversity usually recovers slowly during white blood cell

reconstitution (**Figure 1D**); however, microbiota recovery as well as immune reconstitution can vary

strongly between patients and treatment types (**Figure 1B,C,J,K, S1**). This variation is illustrated by

the distinct trajectories of patient 1 who received a graft of peripheral blood stem cells (PBSC),

retained high microbiota diversity and engrafted earlier (**Figure 1B,E,H**), and patient 2 who received

a graft of umbilical cord blood (cord), lost microbiota diversity and engrafted later (**Figure 1C,F,I**).

Low microbiota diversity at the time of neutrophil engraftment has been associated with 5-fold

increased transplant-related mortality[26], suggesting that that the joint recovery of the microbiota and

white blood cells in circulation is critical for clinical outcomes.

To detect a directional and causal link between the microbiota and white blood cells, we first

used data from a recent prospective randomized trial of autologous fecal microbiota transplantation

(auto-FMT), which is a microbiota manipulation experiment done directly in our patients[23]

(supplementary methods). Twenty-four patients (**Figure 2A, Table S2)** underwent randomization,

resulting in 10 untreated control and 14 treated patients, including patient 3 in **Figure S2**. To

investigate if auto-FMT affected white blood cell reconstitution, we compared the 24 patients'

neutrophil, lymphocyte, monocyte (**Figure 2**) and total white blood cell counts (**Figure S3**) post-

engraftment (i.e. when the transplanted hematopoietic cells begin producing new white blood cells,

Figure 2A,S3). FMT procedures were conducted at variable time points relative to neutrophil

engraftment, but overall, we observed higher counts of each white blood cell type in patients who

received an auto-FMT during the first 100 days post neutrophil engraftment (p<0.001, **Figure 2B,S3-**

S6).

Figure 2: Neutrophil, lymphocyte and monocyte counts increased in FMT-treated patients
during the weeks following the treatment compared to control patients. A) Absolute counts of
neutrophils (blue), lymphocytes (green) and monocytes (orange) in 24 patients enrolled in a randomized
controlled trial to receive an autologous fecal microbiota transplant post-neutrophil engraftment (10 control:
black vertical line, 14 FMT treated: red vertical line,). B) Weekly mean cell counts aligned to the date of
randomization into FMT treatment arm (red) or control (black). Line: mean per week, shaded region: 95%-CI.
C) Results from a linear mixed effects model with random effects per patient and per day relative to neutrophil
engraftment confirms that neutrophil, lymphocyte and monocyte counts are higher in patients receiving auto-
FMT after their treatment as compared to control patients after the randomization date; bars and confidence
intervals for the averages of observed white blood cell counts without FMT and post-FMT (***: p<0.001).
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The increased white blood cells in patients receiving auto-FMT could be due to the

reconstitution of a complex microbiota that we saw in these patients[23] and the associated metabolic

capabilities[6,7,25], or it could be a systemic response to a severe therapy which introduced billions of

intestinal organisms at once via an enema (no enema was administered to control patients[23]).

Moreover, while the mixed-effects model accounted for patient-specific HCT treatments in this

randomized patient cohort, chance differences in extrinsic factors such as different

immunomodulatory drug exposures may have affected this result due to the small cohort size.

Nonetheless, observing that auto-FMT recipients had increased white blood cell counts supports the

notion that the microbiota can modulate the peripheral immune system. High counts of lymphocytes

during immune reconstitution has been associated with improved clinical outcomes[27]. Additionally, in

our HCT patients, a higher average level of white blood cells measured across a period of 100 days

after neutrophil engraftment (supplementary methods) confirms a positive association with 3-year

survival (hazard ratio: 0.91, p:0.04). Determining which taxa modulate immune dynamics could open

new ways to improve robust immune reconstitution, which is critical for clinical outcomes[27–29].

To address this question, we next investigated the link between the gut microbiota and the

dynamics of white blood cell recovery in our large observational cohort of HCT patients. Homeostasis

of circulatory white blood cell counts is a complex, dynamic process: neutrophils, lymphocytes and

monocytes are formed and released into the blood *de novo* by differentiation of hematopoietic

progenitor cells from the bone marrow, and they can be mobilized from thymus and lymph nodes

(lymphocytes), spleen, liver and lungs (neutrophils); they can also migrate from the blood to other

tissues when needed[30]. These processes are dynamic sources and sinks of circulatory white blood

cells, and they can be modulated by drugs administered to patients receiving HCT. To identify factors

associated with these dynamic source- and sink-processes—including the microbiota—we developed

a two-stage approach analyzing the changes of white blood cell counts between two days (i.e. the

rates of cell count increases and decreases). Stage 1 served as a feature selection stage where we used

data of 1,096 patients (after filtering for qualifying samples and applying exclusion criteria, see

supplementary methods) *without* available microbiome information to identify associations between

clinical metadata, including immunomodulatory medications (methods), and changes of white blood

cell counts from one day to the next (**Figure 3A**). Stage 2 was performed on data from an independent

cohort of 841 different patients at MSK from whom concurrent microbiome samples were available.

Stage 2, our main analysis, sought to reveal associations between the abundance of microbial taxa and

the daily changes in blood cell counts in context of immunomodulatory medications, additional

clinical metadata and the current state of the blood itself.

In stage 1 we calculated the changes in neutrophil, lymphocyte and monocyte counts during

patients' recovery from >20,000 pairs of post-engraftment blood samples separated by a single day

(**Figure 3B, S7-9,** supplementary methods). Using a cross-validated feature selection approach, we

detected medications and HCT treatment parameters that were associated with different rates of

change in neutrophil, lymphocyte and monocyte counts, including, as expected, GCSF and the graft

stem cell sources (**Figure S7-S9, Table S1**).

During stage 2 we sought to identify associations between bacterial taxa of the gut microbiota

and the dynamics of immune cells in circulation. For this, we performed Bayesian inferences using

data from different sets of patients *with* available microbiome samples. Stage 1 had identified—as

expected—that stem cell graft sources are associated with immune reconstitution kinetics (e.g. cord

on average slower compared with PBSC[31]), and we therefore stratified our patients by graft source in

stage 2. The model of stage 2 now included microbial genera as predictors of observed changes in

white blood cell counts, in addition to the medications selected in stage 1, clinical features

(conditioning intensity, age, sex), and the current state of the blood in the form of counts of

neutrophils, lymphocytes, monocytes, eosinophils, and platelets. The data comprised 841 patients, but

approximately 60% of the stool samples paired with a daily change in white blood cell counts were

taken before neutrophil engraftment (**Figure 3B, Table S1,** supplementary methods), i.e. when blood

cells counts were zero. In total, we analyzed 2,615 post-engraftment observations of changes in

neutrophil counts during immune reconstitution (lymphocytes: 2,006, monocytes: 2,534) with paired

stool samples which provided a large sample of observed white blood cell dynamics (**Figure 3B,**

Table S3,S4). We first focused on the data from the largest (**Figure 3C**) cohort—patients who

received a PBSC graft—and withheld the other cohorts (bone marrow, BM; T-cell depleted graft

(*ex-vivo*) by CD34+selection, TCD; and cord) to use as independent validation cohorts. For this

validation, we analyzed TCD, BM, and cord patients' data in the same way as PBSC patients' data

and compared the resulting posterior coefficient distributions (methods). We assigned coefficients

obtained from the PBSC cohort a validation score (v-score) between 0 and 3, representing the number

of times that the focal coefficient was validated in the other cohorts; but, conservatively, the score was

always set to zero if we observed counter-evidence among any of the other data sets, i.e. evidence that

coefficients had the opposite sign, ensuring only the most consistent associations were considered as

validated. Finally, we analyzed data from another patient cohort consisting of 493 bone marrow

transplantation recipients treated at Duke University including 9,603 blood samples and a total of 629

microbiota samples from 218 patients, albeit with lower sampling density, and we used the results

from this analysis for further validation.

Notably, as a verification of our approach, we detected associations between the

administration of immunomodulators and increased or decreased rates of immune cell count changes

consistent with the known biological mechanism of these medications (**Figure 3C,S10-S13**). The

strongest across all predictors is the well-known neutrophil-increasing effect of GCSF[32]; GCSF

administration—used to accelerate recovery from chemotherapy-induced neutropenia[32]—was

associated with a +140% increase in the rate of neutrophil changes from one day to the next ([+114%,

+170%], 95 percent probability density interval [HPDI95]). This finding was observed in all MSK

validation data sets (v-score=3, **Figure 3D**), as well as among Duke University patients (**Figure

S14,S15**). We furthermore found a GCSF-associated increase of +43% ([+30%, +58%]HPDI95, v-

score=3) in monocyte rates, and, although smaller, in lymphocyte rates (+16%, [+5%, +27%]HPDI95,

v-score=3). Both neutrophil and lymphocyte rates decreased following the exposure to antihistamine

or immunosuppressive medications (cetirizine -18%, [-35%, +5%]HPDI95, mycophenolate

mofetil -8% [-15%,+1%]HPDI95, respectively). Finally, less intensive chemotherapeutic conditioning

regimens (non-ablative conditioning and reduced intensity) were associated with larger lymphocyte

and monocyte count growth rates during immune reconstitution (**Figure S10C**)

Beyond associating medications in agreement with their known biological mechanism, our

analysis detected associations between the current count of white blood cells and their rate of change:

a negative association among lymphocytes, negative associations between counts of neutrophils and

lymphocytes with the rates of monocytes, and a negative association between the counts of platelets

and lymphocytes and the rates of neutrophils (**Figure 3E**). Conversely, we found positive associations

between monocytes and the rates of each of the investigated white blood cell subsets. These

associations, derived from daily counts of white blood cells, could reflect a complex network

underlying the regulation of blood immune cell composition[30]. More importantly, the associations

quantified for medications and potential homeostatic feedbacks provided a benchmark against which

we could compare associations from gut microbial taxa.

Figure 3: Bayesian inference conducted in context of the host state (white blood cell and platelet counts)
and clinical variables including immunomodulatory medications reveals the microbiota potential to affect

daily changes in circulatory white blood cell counts. A) Cartoon of the model: observed changes in white blood cell counts between two consecutive days are associated with the current state of the host in the form of blood cell counts in circulation, the administration of immunomodulatory medications, patient and clinical metadata, and the state of the microbiome. B) Visualization of the dynamic white blood cell data; scatter plot of the principal components (PC) of observed daily changes of neutrophils, lymphocytes and monocytes without (grey, see **Figures S7-9**) and with (orange) concurrent longitudinal microbiome data (bold: post engraftment sample counts). C) PBSC patients (N=312) provided the most blood samples with simultaneous microbiome data (n=995) relative to TCD, BM and cord patients, who were used as validation data sets. D-F) PBSC patient data inference results; bars show the posterior parameter estimate distributions (thin: 95% highest posterior density intervals, HPDI95, thick: HPDI50) for the jointly inferred associations between treatments (D), white blood cells counts (E), and fecal microbiota genus log-relative abundances (F) with the observed daily changes in neutrophils, lymphocytes and monocytes. v-score: number of validation cohort confirming associations, *always* set to zero if invalidated in any of the TCD, BM, or cord cohorts (additional coefficients in **Figure S10**). G) 100 microbiota samples with highest (left) or lowest (right) relative abundances of *Faecalibacterium*, *Ruminococcus 2* and *Akkermansia*. H) Simulation of the neutrophil population over time in presence of GCSF with microbiota compositions sampled either from those high (blue) or low (red) in *Faecalibacterium*, *Ruminococcus 2* and *Akkermansia* relative abundance as shown in G); line: median of 1,000 simulations, shaded regions: quartile range of simulated neutrophil trajectories. I) In absence of GCSF, equivalent simulations to H) predict that the time to reach neutrophil counts $>2K*\mu l^{-1}$ for the first time after HCT when the microbiota is high (red) in *Faecalibacterium*, *Ruminococcus 2* and *Akkermansia* compared with when these genera are low (blue) will decrease from 6.8 (95%-confidence interval, CI: [6.5, 7]) days to 4.4, (CI: [4.3, 4.5]).

We identified microbial genera that consistently associated with increases or decreases in white blood cell counts by first using data from the PBSC patients and then validating the associations in the other cohorts (**Figure 3F**). Higher abundances of *Faecalibacterium* (+8%, [+1%, +14%]HPDI95 per $\log_{10}$), *Ruminococcus 2* (+5%, [0%, +10%]HPDI95) and *Akkermansia* (+4%, [+1%, +7%]HPDI95) were associated with greater neutrophil increases, whereas increased *Rothia* (-3%, [-7%, 0%]HPDI95), and *Clostridium sensu stricto 1* (-3%, [-6%, 0%]HPDI95) relative abundances associated with reduced neutrophil rates. These results were validated in univariate analyses conducted in the Duke University cohort (**Figure S14, S15**). We also conducted the inference using total genus abundances as predictors instead of relative abundances; this analysis confirmed *Faecalibacterium* as most strongly associated with neutrophil dynamics (**Figure S16**, supplementary methods). *Staphylococcus* was positively associated with lymphocyte rates (+4%, [+1%, +6%]HPDI95) and, again, *Ruminococcus 2* was also associated with faster lymphocyte increases (+5%, [+1%, +9%]HPDI95). Both *Faecalibacterium* as well as *Ruminococcus 2* also associated with increases in monocytes, and while this association was validated in other cohorts (v-score 3 and 1, respectively), there was higher uncertainty of the association estimate (HPDI50>0). Again, *Clostridium sensu strictu 1* (-3% [-5%, -1%]HPDI95) associated consistently with decreased rates of monocytes. The associations we identified—and validated in other cohorts—between

microbial taxa in the gut and daily changes in white blood cell counts support the idea that

hematopoiesis and mobilization respond to the composition of the gut microbiome, influencing

systemic immunity[33].

Most of the taxa that strongly associated with white blood cell dynamics were obligate

anaerobes. *Rothia*, was a notable exception: this aerobic genus is typically found in the oral cavity[34]

but can become an opportunistic pathogen in immunosuppressed patients and is not known to provide

metabolic functions to the host[35]. Some obligate anaerobes, on the other hand, produce short-chain

fatty acids[36,37] and bacterial cell-wall molecules[1,38,39] that modulate immune responses and

granulopoiesis[6]. Nutritional support from the intestinal microbiota improved hematopoietic

reconstitution in a mouse model[25]. To identify a similar association in our patients, we estimated a

microbiota potency by multiplying the $\log_{10}$-relative abundances of microbial genera in a sample with

their corresponding posterior coefficients. We analyzed shotgun metagenomics sequences from 124 of

the samples and observed that samples with positive microbiota potency were associated with

enrichment in cholate degradation and vitamin-B1 synthesis related pathways, as well as butanoate

formation (**Figure S17**). Our findings are in line with evolutionary theory[40] that essential but broadly

available microbial traits such as the production of B vitamins, secondary bile acid metabolism, and

fermentation to short-chain fatty acids[41] could be co-opted by the host's immune system as part of the

homeostatic interplay between immune system and a complex microbiota[42,43]. For example,

Ruminococcus 2 is a genus that contains *R. bromii*, a keystone species necessary for the efficient

release of energy from complex starch in the normal diet[44]. Reassuringly, the genera

Faecalibacterium, *Ruminococcus 2* and *Akkermansia* that we associated with faster rates of white

blood cells (**Figure 3F**) were among those best reconstituted by auto-FMT[23], potentially explaining

why we found higher counts of neutrophils, monocytes and lymphocytes in patients who received the

auto-FMT (**Figure 2B,C**).

The associations we reveal are interpretable as potential effectors on sources and sinks of

white blood cell counts in circulation. Intestinal bacteria may affect white blood cell counts in

circulation by influencing either their sources in the bone marrow or their cytokine profiles[45] and

proliferation rates in the blood, their sinks in different organs, or both. The human immune system in

turn can interact with the microbiota and modulate its composition, for example via

immunoglobulin A responses targeting specific bacteria as studied in mice[43,46,47]. To investigate a

reverse effect of the peripheral immune cell system onto bacterial populations, we employed an

analogous approach to the stage 1 analysis of white blood cell dynamics. Dynamics of white blood

cells can be estimated from changes in absolute cell counts, and to obtain the necessary measurements

in absolute bacterial abundances, we measured total bacterial 16S rRNA gene copies per gram of stool

for a subset of our samples (3,995 samples from 481 patients). Using absolute abundances of bacteria

as predictors in addition to medications, we jointly inferred the association network of dynamics

between the gut bacterial ecosystem and the peripheral immune system. All of our patients receive

antibiotics on some days during their treatment[24] and their strong effects on microbiota dynamics

were the dominant effects that survived cross-validated regularized elastic net regression (**Figure

S18**). Relaxing the strength of the regularization (methods), however, revealed several bi-directional

relationships between immune cells in circulation and bacterial dynamics in the gut (**Figure S19**). Of

note, we detected a negative association of absolute *[Ruminococcus] gnavus* group abundance with

lymphocytes dynamics, confirming our main result based on relative bacterial abundances (**Figure 3**).

In the reverse direction, we saw a positive association of *[Ruminococcus] gnavus* group dynamics

with lymphocyte counts. This result agrees with findings that *Ruminococcus gnavus* thrives in and

promotes inflammatory conditions such as Crohn's disease and other inflammatory bowel diseases

(IBD)[48]; our analysis suggests it may drive high neutrophil to lymphocyte ratios that are broadly

characteristic for poor disease outcomes in IBD[49,50] and beyond[51,52].

Overall, our analysis identified that the microbiome is associated with immune cell dynamics

in addition to medications. The effects should be interpreted as net effects since they do not

distinguish, for example, how the microbiota impacts *de novo* hematopoiesis in isolation from its

impact on other sources and sinks. Unlike the plausible role of obligate anaerobe fermenters in

augmenting hematopoiesis via nutritional support[25], the positive association detected between

Staphylococcus and lymphocyte dynamics could instead result from reduced extravasation of T cells

from circulation into the gut epithelium[53], especially since high abundances of *Staphylococcus* are

associated with low gut microbiota diversity (p<0.001, **Figure S20**), which indicates a depleted

microbiota.

Nevertheless, our approach allows us to leverage the chronology of events and assess

"mathematical causality"[54]. Of course, due to the observational nature of these data there are risks of

confounding that could explain some of the associations found, but the close temporal

correspondence[54] between microbiota and blood cell dynamics, and the validation across cohorts

reduces the number of plausible confounders. Our results, therefore, quite naturally suggest candidate

microbial taxa to manipulate if we seek to steer complex hematopoietic dynamics and utilize the

microbiota as an immunomodulatory component of the human body. Intriguingly, members of

Faecalibacterium and *Ruminococcus* in one study[10], and *Akkermansia* in another[9], were identified as

enriched in patients with better responses to anti–PD-1 immunotherapy, which suggested a

disagreement between the two studies[55]. Our results, however, revealed *Faecalibacterium*,

Ruminococcus 2, and *Akkermansia* as the most strongly associated taxa with increases in white blood

cell counts from one day to the next. Therefore, our results agree with the findings of both anti PD-1

therapy studies that these taxa are associated with immune modulation in humans. Our results also

allow us to compare the potency of manipulating these intestinal commensals to that of

immunomodulatory drugs. While these genera are common in the gut microbiota of healthy people[17],

the relative abundance of each genus can drop below detection in our patients during the intestinal

damage related to HCT[24]. Therefore, realistic ranges of 3-5 orders of magnitude in bacterial log-

relative abundances (**Figure 3G, S21**) can yield effect sizes similar to that of homeostatic feedbacks

between white blood cells and several immunomodulatory medications (e.g. a change in

Ruminococcus 2 from below detection to 1% relative abundance associated with a +67% and +63%

increases in neutrophil and lymphocyte rates, respectively). Therefore, while the effect sizes of

intestinal bacteria at first may appear smaller than those of immunomodulatory drugs, the herein

estimated homeostatic effects of gut bacteria may not be that small since their coefficients refer to

changes in exponential rates of white blood cells and accumulate each day. To better demonstrate how

this accumulation of effects would work, we conducted simulations of the inferred dynamic system of

white blood cells using our posterior coefficient distributions (methods). We simulated 1,000 time

series for microbiota compositions either chosen from the 100 samples highest or lowest in

Faecalibacterium, *Ruminococcus 2* and *Akkermansia* (**Figure 3G**), in presence (**Figure 3H**) or

absence (**Figure 3I**) of GCSF administration. Simulations predict that a microbiota enriched in these

genera could accelerate immune reconstitution, and reduce the time until neutrophils reach $>2K*\mu l^{-1}$

in absence of GCSF by 2.4 days, from predicted 6.8 (CI: [6.5, 7]) to 4.4 days (CI: [4.3,4.5]) days. Gut

bacteria, in concert and over time, could therefore have significant impact on systemic immunity even

in individuals with less severely injured microbiomes.

In sum, our work links the gut microbiota to the dynamics of the human immune system via

peripheral white blood cell populations. Our analysis uses white blood cells counted directly from

patients, which are coarse-grained clinical analyses conducted at large scale but lack details such as

lymphocyte and other immune cell subsets. Nonetheless, because it is in humans, this study fills an

important gap at a critical time for microbiome research when the clinical relevance of animal models

of microbiome-immune interaction has been questioned[21]. By studying a large number of patients

over time, we were able to infer and quantify for the first time the association of microbiota

components on systemic immune cell dynamics, and our results help to consolidate previous

findings[10,9] that seemed in conflict with each other[55]. Our study demonstrates that the composition of

the microbiota does indeed modulate systemic immune cell dynamics, a link that could be used in the

future to improve immunotherapy and help identify microbiota treatments for inflammatory

diseases[9,10,56–60].

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356   **Bibliography**

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Supplementary Information

Appended and available online.

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Author Contributions

J.S. and J.B.X. wrote the manuscript. J.S. and J.B.X. designed the analyses with expert help from R.N.. J.U.P. and Y.T. contributed to the clinical data preparation, B.P.T. provided the 16S data processing pipelines, A.D. provided the shotgun processing pipelines. All authors contributed to the writing and interpretation of the results.

SUPPLEMENTARY INFORMATION

Methods

Complete blood count collection and characterization

Absolute white blood cells count data were obtained from routine complete blood counts ordered by

clinicians during normal clinical practice, used to obtain informative diagnostic and monitoring

information. Blood samples received in the clinical hematology laboratory were analyzed using

Sysmex XN automated hematology analyzers (Sysmex, Lincolnshire, IL) and, when needed based on

specific flags and parameters as per MSKCC standard operating procedures, were validated manually

using the Sysmex DI-60 Slide Processing System or CellaVision DM9600 Automated Digital

Morphology System (Sysmex, Lincolnshire, IL).

16S rRNA gene amplification and multiparallel sequencing

For each sample, duplicate 50-µl PCRs were performed, each containing 50 ng of purified DNA, 0.2

mM deoxynucleotide triphosphates, 1.5 mM MgCl2, 2.5 U Platinum Taq DNA polymerase, 2.5 µl of

10× PCR buffer, and 0.5 µM of each primer designed to amplify the V4-V5: 563F (5′-nnnnnnnn-

NNNNNNNNNNNN-AYTGGGYDTAAAGNG-3′) and 926R (5′-nnnnnnnn-NNNNNNNNNNNN-

CCGTCAATTYHTTTRAGT-3′). A unique 12-base Golay barcode (Ns) precedes the primers for

sample identification[61], and one to eight additional nucleotides were placed in front of the barcode to

offset the sequencing of the primers. Cycling conditions were 94°C for 3 min, followed by 27 cycles

of 94°C for 50 s, 51°C for 30 s, and 72°C for 1 min. For the final elongation step, 72°C for 5 min was

used. Replicate PCRs were pooled, and amplicons were purified using the QIAquick PCR Purification

Kit (Qiagen). PCR products were quantified and pooled at equimolar amounts before Illumina

barcodes and adaptors were ligated, using the Illumina TruSeq Sample Preparation protocol. The

completed library was sequenced on an Illumina MiSeq platform following the Illumina

recommended procedures with a paired-end 250 × 250 bp kit

Sequence analysis

The 16S (V4-V5) paired-end reads were merged and demultiplexed. Amplicon sequence variants

(ASVs) were identified using the Divisive Amplicon Denoising Algorithm (DADA2) pipeline

including filtering and trimming of the reads[62]. Reads were trimmed to the first 180 bp or the first

point with a quality score Q<2. Reads were removed if they contained ambiguous nucleotides (N) or

if two or more errors were expected based on the quality of the trimmed read. We assigned taxonomy

to ASVs using a 8-mer based classifier trained by IDTaxa[63] using the SILVA database[64]. We

determined the copy number of 16S rRNA genes per gram of stool for 4,158 of our samples as

reported previously[24], by quantitative PCR on total DNA extracted from fecal samples.

Quantification of total microbiota density per gram of stool and estimation of total genus

abundances.

Quantitative PCR (qPCR) was performed on DNA extracted from the 1g wet weight of a stool sample

using DyNAmo SYBR Green qPCR kit (Finnzymes) and 0.2 μM of the universal bacterial primer 8F

(5′-AGAGTTTGATCCTGGCTCAG) and the broad-range bacterial primer 338R

(5′-TGCTGCCTCCCGTAGGAGT-3′). Standard curves were prepared by serial dilution of the PCR

blunt vector (Invitrogen) containing 1 copy of the 16s rRNA gene. Cycling conditions were 95°C for

10 minutes followed by 40 cycles of 95°C for 30 seconds, 52°C for 30 seconds, and 72°C for 1

minute. We used the measurements of total 16S rRNA gene counts per gram of stool to multiply the

relative abundances of taxa obtained from 16S amplicon sequencing to obtain the estimate of their

total abundance per gram of stool (supplementary methods). Importantly, this does not account for

16S copy number variation between taxa, but the observed dynamic ranges in total abundances of taxa

in our data set span up to 9 orders of magnitude, exceeding the potential inaccuracies due to copy

number variation.

Diversity calculations

Microbiome alpha-diversity was measured by the inverse Simpson (IS) index of a sample. It was

calculated by $IS_i = \frac{1}{\sum_{j=1}^{N} p_{ij}^2}$, where p is the relative abundance of the jth ASV out of N total ASVs in

sample i.

Linear mixed-effects model of white blood cell counts

To study the effect of auto-FMT on white blood cells, we investigated the white blood cell counts of

24 enrolled patients of this trial from the day of neutrophil engraftment until 100 days after. FMT

occurred on different days relative to neutrophil engraftment. Thus, we performed an analogous

analysis to that conducted in the original publication that demonstrated how FMT re-established a

diverse microbiome in the post-FMT period[23]. To answer if white blood cell counts differed post-

FMT, we used a linear mixed effects model of white blood cell counts, y, modeled as a function of the

FMT treatment as well as patient and timepoint specific random effects. We included random

intercept terms for each day i and each patient j, and a fixed effects term for the post-FMT period with

associated coefficient "armpost", using the indicator variable "FMT", that is 1 when a patient was

from the FMT treated arm of the trial and day was greater than or equal to the day of the FMT

procedure. We conducted independent analyses for neutrophil, lymphocyte and monocyte counts.

This resulted in the following model of a cell count, y, for patient j on day i:

$$y_{ij} = \beta_0 + armpost * FMT_{ij} + day_i + patient_j + \varepsilon_{ij}, \qquad i = 0, \dots, D, j = 1, \dots, P$$

with prior distributions $day_i \sim \mathcal{N}(0, \sigma_{day}^2)$, and $patient_j \sim \mathcal{N}(0, \sigma_{patient}^2)$, independent error

$\varepsilon_{ij} \sim \mathcal{N}(0, \sigma^2)$ and fixed intercept β_0, for the D days post neutrophils engraftment and P patients,

(D=100, P=24). For convenience of those interested in reanalyzing our data, the part of our data

concerning the auto-FMT analysis is available in tidy format (supplementary data 11), and the

analysis code conducted in the R programming language is available as an exported notebook

(fmt_effect_on_wbc.pdf) on github: https://github.com/jsevo/wbcdynamics_microbiome/.[65] We

conducted an additional analysis with "day" as a continuous predictor which did not change our

conclusions (supplementary methods).

Dynamic systems analyses

We analyzed factors associated with the observed changes of absolute counts of neutrophils, lymphocytes and monocytes between two days. In the following we describe how chronology of events and biological samples were encoded, and the models used to infer a role of medications, clinical parameters and the microbiome on dynamics of white blood cells.

To reveal factors that associate with day-to-day changes in white blood cell counts, we started from a first-order differential equation of white blood cell (W) dynamics:

$$\frac{d(W)}{dt} = W \left(gr + \sum_{j=1}^{P} \beta_j X_j \right)$$

Where gr represents the intercept, i.e. the base line rate of change during immune reconstitution, and β_j are the to-be-estimated coefficients of the P predictors $X_j, j \in P$, of the white blood cell dynamics. This equation was then linearized to

$$\frac{d\, ln(W)}{dt} = gr + \sum_{j=1}^{P} \beta_j X_j$$

And we parameterized the corresponding discrete difference equation:

$$\frac{\Delta \ln(W)}{\Delta t} = gr + \sum_{j=1}^{P} \beta_j X_j$$

where Δln(W) is the log-difference between single days of neutrophils, lymphocytes or monocytes counts, and Δt=1 for all intervals. Predictors include the counts of neutrophils, lymphocytes, monocytes, eosinophils and platelets during an interval (homeostatic feedbacks), immunomodulatory medication and clinical observations such as a blood stream infection and the onset of graft versus host disease, HCT parameters such as graft types and conditioning regimens, and, additionally, the microbiota composition in "stage 2" of our analysis (supplementary methods for data exclusion and additional details on interval definitions). Importantly, by parameterizing a dynamic equation and analyzing *rates of change*, our coefficient estimates have an immediate causal interpretation within our modeling framework (i.e. a β_j>0 implies that higher levels of the corresponding X_j *increases* the

respective white blood cell type, W). To differentiate such results from other associations, analyses of

this type have been termed "mathematical causality"[54].

Stage 1 analysis: Feature selection. Identifying medications and clinical observations associated with

white blood cell dynamics from patients without microbiome data

Stage 1 uses data of patients without any available microbiome samples and the following model of

white blood cell changes, y:

655
$$y = gr + \sum_{j=1}^{p} \beta_j X_j,$$

with intercept, gr. The predictors, X, include dummy variables for the HCT graft type, patients' age on

the date of HCT, sex, 13 most frequently observed positive blood cultures with remaining other blood

stream infections grouped into a separate category "other infections", an indicator for the onset of

graft versus host disease, administrations of 55 different, most common immunomodulatory

medications and platelet transfusion events, and HCT conditioning intensity regimens as well as the

log-transformed geometric mean counts of neutrophils, lymphocytes, monocytes, eosinophils and

platelets during the respective interval. We used elastic net regression[66] for feature selection using the

sklearn package for the Python programming language[67]. For elastic net regression with 50% L1-

penalty, predictors were scaled between zero and 1, and we used 10-fold cross validation (i.e. leaving

out 10% of patients at each cross-validation step) to choose the regularization strength, λ, solving for

666
$$argmin_{gr,\beta} \left\{ \frac{1}{2N} \sum_{i=1}^{N} (y_i - gr - \sum_{j=1}^{p} x_{ij}\beta_j)^2 + \frac{1}{2}\lambda \sum_{j=1}^{p} |\beta_j| + \frac{1}{2}\lambda \sum_{j=1}^{p} \beta_j^2 \right\}$$

667
Stage 1 yielded a sparse coefficient matrix of predictors used to design the model in stage 2.

Expanded analysis on patients with microbiome data – stage 2

To identify associations between microbiota and white blood cell dynamics, we conducted an

analogous, Bayesian regression using the package PyMC3 for the Python programming language[68].

Stage 1 identified important difference between transplant types, and we therefore stratified our data

into 4 cohorts according to their stem cell graft source. Using data independently from each cohort,

we applied "no U-turn" sampling[69] to produce 10,000 posterior samples from 5 independent MCMC

chains that parameterized the model:

677
$$y \sim \mathcal{N}(\mu, \sigma^2)$$

678
$$\mu = gr + \sum_{j=1}^{\hat{P}} x_j \beta_j$$

679   with uninformative prior distributions

680
$$gr \sim \mathcal{N}\ (mean = 0,\ standard\ deviation =\ 100)$$

681
$$\beta_j \sim \mathcal{N}\ (mean = 0,\ standard\ deviation =\ 100)$$

682
$$\sigma \sim HalfCauchy\ (\ beta = 2)$$

where y is the observed daily change of a focal white blood cell type as in stage 1 with normal

distributed mean, μ, and σ, the model uncertainty with a thick-tailed half Cauchy prior (importantly,

our posterior estimates do not depend on this choice as we obtain the same results with an inverse

Gamma prior, figure S19). μ was a function of the baseline growth rate, gr, and predictors, $\hat{P}$:

medications with non-zero coefficients in stage 1, the white blood cell counts, patient age and sex,

and HCT conditioning intensities; additionally, $\hat{P}$ now included the log-abundances of microbial

genera as measured by 16S sequencing from DNA in the stool collected on the second day of a daily

interval (see supplementary methods for details). We considered taxa that were among the 100 most

abundant, or had reached maximum relative abundances of at least 10%, and selected those who were

non-zero in more than 75% of our samples. White blood cell counts and microbiota data present

during a daily interval were log-transformed, and zeros were filled with half of the minimum observed

non zero counts (i.e. 0.5e3 and 2e-6, respectively). We focused on the largest cohort (PBSC) and

used the independent inference results from TCD, BM, and cord cohorts for validation.

Validation score

Coefficients learnt from the PBSC patient cohort were assigned a "validation score" based on the

results obtained from the other three MSK patient cohorts. Our requirements for validation were

conservative; we required evidence from our validation data sets as well as absence of counter

evidence. For regression results from each of the validation graft type cohorts, i.e. TCD, BM, and

cord, we checked if a coefficient had more than 75% probability (50%HPDI) to have the same sign as

the mean of the PBSC coefficient posterior for a given predictor. If so, this was considered evidence

of validation, and we summed the evidence over the three validation sets (i.e. maximum score of 3, 1

from each of TCD, BM, and cord cohorts). Conversely, if we found more than 75% probability

among any of the validation data sets that a given predictor had the *opposite* sign as the posterior

mean calculated from PBSC data, this was considered counter evidence and the validation score was

always set to zero.

Analysis of white blood cell dynamics with absolute bacterial abundances as predictors instead of

relative abundances

We conducted an ordinary least squares regression using the statsmodels package in the Python

programming language of the same model as in the main Bayesian analysis using total bacterial

abundances as predictors. This was only possible on a subset of 389 neutrophil, 331 lymphocyte and

376 monocyte rate observations from PBSC patients.

Forwards simulation of predicted immune system reconstitution kinetics

To assess the impact of the estimated microbiota coefficients on immune system dynamics, we

conducted 1,000 simulations of the system of 3 differential equations describing the dynamics of

neutrophils, lymphocytes and monocytes. We ran 1,000 simulations four times: in presence and

absence of GCSF, each with microbiota compositions enriched or depleted in *Faecalibacterium*,

Ruminococcus 2 and *Akkermansia*. To identify these compositions, we ranked the observed

microbiota compositions by these taxa, and chose randomly either from the top or bottom 100. The

coefficients for white blood cell interactions, interactions with the microbiota and the effect of GCSF

were sampled from our posterior coefficient distributions. Using these coefficients sampled at the start

of the simulation, and using 50 cells*μl^{-1} of neutrophils, lymphocytes and monocytes as initial values,

we simulated these differential equations forwards in time using the odeint function of the scipy

package for the Python programming language.

Validation on data from Duke University

We analyzed 9,603 blood samples with 25,581 associated administrations of immunomodulatory

medications, and 741 microbiota samples from Duke as an orthogonal data set to validate our

findings. The temporal resolution of this data was much lower, and after filtering for samples from the

relevant post neutrophil engraftment period, and by requiring daily intervals, 83 valid, complete data

points were available. Using these data, we correlated daily blood cell changes individually in

univariate, or jointly in a partial least squares regression, with those predictors that achieved more

than 95% probability density in the positive or negative domain in the PBSC data regression. For each

of these predictors, we present the sign of slopes and Bonferroni corrected *p*-values from individual

linear regressions.

Joint analysis of the effect of antibiotics and white blood cell counts on the microbiota and the
microbiota and immunomodulatory medications on white blood cell counts

Analogous to *stage 1*, we performed cross-validated, regularized linear regressions (ElasticNet) using

the scikit-learn package for the Python programming language to jointly estimate the association

network between microbiota and circulatory white blood cells. For this, we constructed a block matrix

X of predictor matrices X_i that include the absolute bacterial abundances, drug data (antibiotics for

bacterial dynamics and immune modulators for white blood cell dynamics), as well as the counts of

white blood cells and a separate intercept term per block. Each block $X^l_{n_l, p_l}$, with n_l observations and p_l

predictors (l=0...k), on the diagonal of **X** corresponds to the indices of the observed daily log-changes

of one of the 41 bacterial genera considered in our main analysis or the log changes in neutrophil,

lymphocyte and monocyte counts from PBSC patients contained in **Y** (in total we calculated 15,833

rates from 256 patients). Our regression problem can thus be written as:

753
$$argmin_\beta (\mathbf{Y} - \mathbf{X}\boldsymbol{\beta}) \text{ where } \mathbf{X} = \begin{pmatrix} X^0_{n_0, p_0} & \cdots & 0_{n_0, p_k} \\ \vdots & \ddots & \vdots \\ 0_{n_k, p_0} & \cdots & X^k_{n_k, p_k} \end{pmatrix}$$

754 with k=44, i.e. 41 bacterial genera and 3 white blood cell types, the to-be estimated coefficient vector

755 **β** and **0** the zero matrix. This system is underdetermined and we therefore chose the same approach as

756 in stage 1, elastic net regression, for feature selection. Predictors were scaled between zero and 1, and

757 we used 3-fold cross validation, leaving out $1/3^{rd}$ of the patients at each iteration to identify a global

758 regularization strength, λ, solving for

759
$$argmin_\beta \left\{ \frac{1}{2\eta} \sum_{i=1}^{\eta} (y_i - \sum_{j=1}^{\rho} x_{ij}\beta_j)^2 + \frac{1}{2}\lambda \sum_{j=1}^{\rho} |\beta_j| + \frac{1}{2}\lambda \sum_{j=1}^{\rho} \beta_j^2 \right\}$$

where η is the total number of observed daily log changes in genera and white blood cells, and ρ the

total number of predictors. This yielded a strongly regularizing λ_s, and thus few predictors. To

characterize potential bidirectional relationships between white blood cell counts and the gut

microbiota, we iteratively reduced the regularization strength until the strongest interaction between

microbiota and white blood cell dynamics, i.e. *Faecalibacterium* with neutrophil dynamics, was

detected. We than re-ran the regression with this reduced regularization strength, λ_r.

Shotgun sequencing

Sequencing of 124 post-neutrophil engraftment was conducted on the Illumina HiSeq platform. For

details and the processing of the FASTQ files, see supplementary methods. We used the HUMAnN2

pipeline[70] with default settings for functional profiling of our samples, with the UniRef90 data base

and ChocoPhlAn for alignment, and we renormalized our samples by library depth to copies per

million. We used MetaCyc to obtain stratified and unstratified pathway abundances.

Statistical analysis of shotgun data

We calculated the predicted microbiota potency score for each sample and separately for neutrophils,

lymphocytes and monocytes, by multiplying the abundances of taxa in each of the 124 samples with

the corresponding posterior coefficients obtained from the PBSC inference. To distinguish the sets of

metabolic functions that separate samples with positive and negative predicted potencies, we

converted the pathway abundances into presence and absences profiles. We performed a linear

discriminant analysis between positive and negative potency samples with a least squares solver and

automatic shrinkage using the Ledoit-Wolf lemma using the sklearn package for the Python

programming language[67]. To assess differences in the presence or absence of pathways between

samples with positive and negative potency, we used Fisher's exact test for each pathway.

Supplementary Figures

Figure S1: White blood cell counts and platelet counts per graft source over the first 100 days post HCT per day relative to HCT; lines: mean, shaded: ± standard deviations).

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793 **Figure S2:** HCT patient who received an autologous fecal microbiota transplant (auto-FMT, dashed red line)
794 that restored commensal microbial families and ecological diversity in the gut microbiota, with concurrent cell
795 counts of peripheral neutrophils, lymphocytes and monocytes and immunomodulatory drug administrations.
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Figure S3: Total counts of white blood cells increased in FMT-treated patients in the weeks following the treatment compared to control patients. A) Total white blood cell counts in 24 patients enrolled in a randomized controlled trial to receive an autologous fecal microbiota transplant post-neutrophil engraftment (10 control: black vertical line, 14 FMT treated: red vertical line,). B) Weekly mean cell counts aligned to the date of randomization into FMT treatment arm (red) or control (black). Line: weekly mean, shaded region: 95%-CI C) Results from a linear mixed effects model with random effects per patient and per day relative to neutrophil engraftment confirms that the total white blood cell counts is higher in patients receiving auto-FMT after their treatment as compared to control patients after the randomization date, bars and confidence intervals for the averages of observed white blood cell counts without FMT and post-FMT (***: $p<0.001$).

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812 **Figure S4:** Neutrophil counts in 24. FMT trial patients. Thin lines: raw data (blue: post-FMT); thick black:
813 mean per day, thick blue: mean+post-FMT coefficient. Means and confidence intervals (shaded region) from
814 linear mixed effects model (methods).

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817 **Figure S5:** Lymphocyte counts in 24. FMT trial patients. Thin lines: raw data (blue: post-FMT); thick black:
818 mean per day, thick blue: mean+post-FMT coefficient. Means and confidence intervals (shaded region) from
819 linear mixed effects model (methods).
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823 **Figure S6:** Monocyte counts in 24. FMT trial patients. Thin lines: raw data (blue: post-FMT); thick black: mean
824 per day, thick blue: mean+post-FMT coefficient. Means and confidence intervals (shaded region) from linear
825 mixed effects model (methods).

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830 **Figure S7:** "Stage 1" regression on neutrophil dynamics on patients without microbiome data. Coefficients
831 from 10-fold cross-validated elastic net regression daily changes in neutrophils. gr: intercept; TCD: T-cell
832 depleted graft (ex-vivo) by CD34+selection; PBSC: peripheral blood stem cells; BM: bone marrow; cord:
833 umbilical cord blood; NONABL: Nonmyeloablative; REDUCE: reduced-intensity conditioning regimen; F:
834 female; N: patients, n: samples (daily changes in neutrophils).
835

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837 **Figure S8:** "Stage 1" regression on lymphocytes dynamics on patients without microbiome data. Coefficients
838 from 10-fold cross-validated elastic net regression daily changes in lymphocytes. gr: intercept; TCD: T-cell
839 depleted graft (ex-vivo) by CD34+selection; PBSC: peripheral blood stem cells; BM: bone marrow; cord:
840 umbilical cord blood REDUCE: reduced-intensity conditioning regimen; F: female. N: patients, n: samples
841 (daily changes in lymphocytes).
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844 **Figure S9:** "Stage 1" regression on lymphocytes dynamics on patients without microbiome data. Coefficients
845 from 10-fold cross-validated elastic net regression daily changes in lymphocytes. gr: intercept; TCD: T-cell
846 depleted graft (ex-vivo) by CD34+selection; PBSC: peripheral blood stem cells; BM: bone marrow; cord:
847 umbilical cord blood REDUCE: reduced-intensity conditioning regimen; F: female. N: patients, n: samples
848 (daily changes in lymphocytes).

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851 **Figure S10:** Additional coefficient estimates of medications (A), additional genera (B) and metadata (C) from
852 the Bayesian regression, see also Figure 3. REDUCE: reduced-intensity conditioning regimen; NONABL: non-
853 myeloablative conditioning regimen. F: female

Figure S11: Posterior sampling converged. Histograms of the ranked posterior draws from the model of neutrophil dynamics in PBSC patients (ranked over all chains), plotted separately for each chain (see supplementary methods), show no substantial differences between chains[71].

Figure S12: Posterior sampling converged. Histograms of the ranked posterior draws from the model of lymphocyte dynamics in PBSC patients (ranked over all chains), plotted separately for each chain (see supplementary methods), show no substantial differences between chains[71].

Figure S13: Posterior sampling converged. Histograms of the ranked posterior draws from the model of monocyte dynamics in PBSC patients (ranked over all chains), plotted separately for each chain (see supplementary methods), show no substantial differences between chains[71].

Figure S14: Validation analysis of predictors on white blood cell dynamics using data from patients treated at DukeHealth. Individual univariate regressions of microbiome and clinical predictors identified in stage 2 of our analysis on daily changes in neutrophils, lymphocytes and monocyte. Bonferroni corrected p-values: ***<0.001, **<0.01, *<0.05; p>0.05: n.s. Sign of coefficients from MSK PBSC patients for comparison.

Figure S15: Validation analysis of predictors on white blood cell dynamics using data from patients treated at DukeHealth. Partial least squares regression of microbiome and clinical predictors identified in stage 2 of our analysis on daily changes in neutrophils, lymphocytes and monocyte.

884 **Figure S16:** Validation analysis of the main model using absolute bacterial abundances as predictors instead of
885 relative abundances in Figure 3. Results show coefficients from a least squares regression for medications (A),
886 white blood cell feedbacks (B) metadata (C) and total genus abundances (D) of neutrophil, lymphocyte and
887 monocyte daily log-changes. This was only possible for only a subset of the data for which we obtained absolute
888 bacterial abundance estimates (methods), n: samples, N: patients.
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893 **Figure S17:** Functional analysis of microbiota samples. To distinguish samples predicted to increase rates of
894 white blood cells, a microbiota potency score was calculated from posterior coefficients (Figure 3, methods) and
895 the relative abundance of taxa in samples. Bars show linear discriminant analysis (LDA) scores of MetaCyc

pathway profiles from 124 shotgun sequenced samples that distinguished positive and negative potency samples the most (LDA-score magnitude in the 95th percentile). Highlighted pathways are discussed in the main text. For each pathway, we tested differences between positive and negative potency samples using Fisher's exact test; p-value <0.001: ***, <0.01:**, <0.05:*.

Figure S18: Jointly inferred association network between white blood cell and bacterial genus dynamics (methods). Strong regularization yields few non-zero coefficients and antibiotics dominate the dynamics.

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914 **Figure S19**: Jointly inferred association network between white blood cell and bacterial genus dynamics with
915 reduced regularization (methods) indicates potential bidirectional feedbacks, e.g. between lymphocytes and
916 *[Ruminococcus] gnavus group* (highlighter green boxes, and cartoon).
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922 **Figure S20:** The relative non-zero abundance of Staphylococcus is inversely related to microbiome alpha
923 diversity, shaded: 95% confidence intervals.
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929 **Figure S21:** Abundance profiles of the two genera, *Faecalibacterium* and *Ruminococcus 2*, most strongly
930 associated with white blood cell increase; number of times detected (left) and log_{10} abundance distribution when
931 above detection (right).
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936 **Figure S22:** Posterior association coefficients do not depend on the choice of prior for σ in the main Bayesian
937 model. Plotted are the posterior means from our main analysis against the equivalent inference with an inverse
938 Gamma prior (alpha=1, beta=1).
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942 Table S1: Data set summary and patient characteristics. HCT-graft types: TCD: T-cell depleted graft (*ex-vivo*)
943 by CD34+selection; PBSC: peripheral blood stem cells; BM: bone marrow; cord: umbilical cord blood;
944 Conditioning intensity: Bacigalupo classification, graded categories from most to least intense (ABLATIVE,
945 REDUCE, NONABL).
946

patients		2,926
HCT therapies*		3,060
blood samples	total	450,635
	between HCT-day -21 and HCT-day 183	193,396
Disease	Leukemia	1,635
	Non-Hodgkin's Lymphoma	415
	Multiple Myeloma	170
	Hodgkin's disease	88
	other	752
HCT graft type	TCD	1,106
	PBSC unmodified	959
	BM unmodified	617
	cord	378
Conditioning intensity	ABLATIVE	65%
	REDUCE	21%
	NONABL	13%
Gender	M	58%
	F	42%
Age of adults (years)	25%-tile	39
	mean	50
	75%-tile	62
Microbiome samples	total	12,633
	- from patients with blood data	10,680
	- of those, post engraftment	4,179
	- of those with daily change in WBC	2,615
	patients with microbiome sample	1,290

947 *) some patient received several HCTs
948
949
950 Table S2: Patient and HCT characteristics of 24 patients enrolled in the randomized controlled FMT trial.

	control	FMT treated
N patients	10	14
ABLATIVE	6	7
REDUCE	4	7
BM unmodified	1	3
PBSC unmodified	3	4
TCD	5	3
cord	1	4

959   Table S3: Patient and HCT characteristics of the subset of patients who donated microbiota samples.

patients		1,294
Disease distribution	Leukemia	51%
	Non-Hodgkin's Lymphoma	15%
	Multiple Myeloma	8%
	Hodgkin's disease	3%
	other	23%
HCT graft type	TCD	37%
	PBSC unmodified	38%
	BM unmodified	9%
	cord	16%
Conditioning intensity	ABLATIVE	55%
	REDUCE	34%
	NONABL	11%
Gender	M	59%
	F	41%
Age (years)	25%-tile	46
	mean	54
	75%-tile	65

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962   Table S4: Patient and HCT characteristics of the subset of patients who did not donate microbiota samples.

patients		1,010
Disease distribution	Leukemia	53%
	Non-Hodgkin's Lymphoma	17%
	Multiple Myeloma	5%
	Hodgkin's disease	5%
	other	20%
HCT graft type	TCD	40%%
	PBSC unmodified	31%
	BM unmodified	17%
	cord	12%
Conditioning intensity	ABLATIVE	65%
	REDUCE	14%
	NONABL	21%
Gender	M	58%
	F	42%
Age (years)	25%-tile	36
	mean	47
	75%-tile	59

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965   Table S5: Patient and HCT characteristics of the Duke University patient cohort.
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patients		493
Disease distribution	Lymphoma	11%
	Leukemia	50%
	Non-Hodgkin's Lymphoma	4%
	Multiple Myeloma	8%

		Hodgkin's disease	4%
		other	24%
HCT graft type		TCD	0%
		PBSC unmodified	72%
		BM unmodified	11%
		cord	16%
Conditioning intensity		ABLATIVE	92%
		NONABL	7%
Gender		M	65%
		F	35%
Age (years)		25%-tile	41
		mean	49
		75%-tile	57

Data availability

The data used in our study is organized in supplementary tables (data-tables.zip), with corresponding filenames (italic):

1. *cGENUS.csv*: relative taxon abundances in fecal microbiota samples from 12,633 stool samples
2. *cHCTMETA.csv*: HCT characteristics
3. *cINFECTIONS.csv*: positive blood culture results
4. *cMISAMPLES.csv*: NCBI SRA accession number, diversity (inverse Simpson index), total 16S (where available), stool consistency for each fecal microbiota sample
5. *cMED.csv*: medication data
6. *cPIDMETA.csv*: anonymized patient demographics
7. *cWBC.csv*: absolute counts of neutrophils, lymphocytes, monocytes, eosinophils, and platelets with indication if included in analyses
8. *cDUKE__GENUS.csv*: relative taxon abundances in fecal microbiota samples from 12,633 stool samples
9. *cDUKE__WBC.csv*: absolute counts of neutrophils, lymphocytes, monocytes, eosinophils, and platelets with indication if included in analyses
10. *cDUKE__MED.csv*: medication data
11. *cFMT_analysis*.csv: convenience table for Figure 2

Code availability

The relevant scripts for stage 1, stage 2, and the model assessing the effect of FMT on white blood cell counts are on github:https://github.com/jsevo/wbcdynamics_microbiome.