Animal personality shapes evolutionary trajectories via conserved genetic architecture.

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Abstract

Consistent individual differences in behavior, i.e. animal personality, are often

correlated within broader syndromes[1] that can cause evolution in one trait to drive

correlated changes in other behaviors via genetic correlations[2]. Despite abundant

demonstration that these behavioral syndromes are common across taxa[3,4], it has

not been clear whether syndromes arise due to genetic and cellular mechanisms or

because particular combinations of behavioral responses are favored. Here we show

that populations of field crickets (*Gryllus integer*) exhibit a genetically conserved

behavioral syndrome structure despite differences in average behaviors. We found

that the distribution of genetic variation and covariance among behavioral traits was

consistent with genes and cellular mechanisms underpinning behavioral syndromes

rather than historic correlated selection. Moreover, divergence among populations'

average behaviors was constrained by the genetically conserved behavioral

syndrome. By constraining the divergence of populations, animal personality may

therefore also affect the ability of species to behaviorally diverge and therefore

speciation. This suggests important ways in which labile traits—traits that are highly

plastic, like behavior, physiology, and life-history traits—might affect evolutionary

outcomes.

Behavior is frequently assumed to have been shaped by selection[5] and thus populations are

expected to differ in a range of behaviors based on local selective pressures. This

perspective implies that behaviors are able to evolve independently, an assumption

increasingly challenged by the ubiquity of behavioral syndromes—correlations among

behaviors[6]—which have been documented across taxonomic groups[3,4] and are comprised

of both genetic and environmental contributions[2,7].

Given the contribution of genetic correlations to behavioral syndromes[7,8],

syndromes have the potential to constrain the ability of populations to diverge and respond

to local selective pressures[9]. Specifically, based on quantitative genetic theory, if

syndromes reflect pleiotropic effects—where a single gene affects multiple behaviors—

populations will be constrained to diverge along shared evolutionary pathways. Further,

this divergence is predicted to occur in the direction in trait space that contains the most

variation[10] (Figure 1). Consequently, if syndromes have a constraining effect on evolution,

the pattern of correlations among traits will be conserved among populations (Figure 1).

Alternatively, if genetic correlations underpinning syndromes are the result of selection

historically favoring particular trait combinations (i.e. selection-induced linkage

disequilibrium[11]), the divergence of populations will be relatively unconstrained as these

genetic correlations are expected to rapidly break down when selection changes[12].

These two quantitative genetic explanations for behavioral syndromes have explicit

analogs in the behavioral literature: whether syndromes emerge from pleiotropy or other

mechanistic effects is termed the "constraints hypothesis," as opposed to selection-induced

linkage disequilibrium, which is termed the "adaptive hypothesis" (Figure 1, see also[13]).

Knowing whether the constraints or adaptive hypothesis drives the emergence of

behavioral syndromes is of critical importance. These hypotheses differentially affect

evolutionary outcomes ranging from responses to environmental changes to speciation

dynamics. For example, in *Anolis* lizards, constraints imposed by morphological genetic

correlations have shaped divergence and the phenotypes possible during adaptive

radiations[14]. Whether behavioral syndromes have similar effects is an open question.

Unfortunately, while some studies have compared phenotypic or among-individual

correlations across populations[15-19], population comparison of behavioral syndromes at the

additive genetic level has been restricted to a single pairwise comparison of two

populations[13]. Consequently, the generality of the constraints versus adaptive hypotheses

is unclear and the general effect of behavioral syndromes on population divergence is

unknown.

Here we tested whether behavioral syndromes have diverged at the genetic level

among populations of the field cricket, *Gryllus integer*. We critically evaluated predictions of

the adaptive and constraints hypotheses (Figure 1) via comparisons of behavioral genetic

(co)variance matrices, i.e. **G**, estimated for multiple populations of *G. integer* (Figure 2). We

collected crickets from four populations, one in California (USA), one in Arizona (USA) and

two in New Mexico (USA; Figure 2, Table S1) during the summer of 2017. Collected females

(F_P) were individually housed under standard conditions and allowed to oviposit into a

cotton substrate (see Methods and Table S1). Offspring (F_0) were then reared to maturity,

after which known matings of males and females within population were conducted[20],

producing a subsequent generation of crickets (F_1) of known parentage. This was repeated

for a second generation (F_2). We measured seven behaviors of F_0 , F_1 , and F_2 crickets (965

individuals in total, Table S1) in three ecologically relevant behavioral assays: latency to

emerge from shelter, activity/exploration of a novel area, and locomotive response to cues

of predator presence (see Methods). Based on the known relatedness among individuals

we estimated **G** for each population. We then compared the four populations to determine

whether the differences among **G**s were consistent with expectations of either the adaptive

or constraints hypotheses based on the distribution of genetic variation in multivariate

trait space[10,14,21-23] and based on whether correlations degraded across generations as

expected under the adaptive hypothesis[12].

As predicted under the constraints hypothesis, the populations exhibited conserved

behavioral syndrome structure. This conclusion is supported by three primary lines of

evidence: First, and most importantly, the four populations shared three subspaces in

multivariate space that explained 99% of the observed genetic variation (Table S2; all

Bayesian probabilities < 0.6 for differences). These subspaces were primarily characterized

by variation in latencies to emerge from shelter, distance traveled during exploration, and

in response to cues of predator presence (Table S2). These subspaces describe the

structure of the species' behavioral syndrome and their being shared demonstrates that the

orientation of genetic variation was conserved among the populations. Second, we found

that the g_{max}'s (the dimension in which most behavioral variation was expressed[10]) of the

Aguila and Dunnigan and Socorro and Aguila populations were strongly correlated (Figure

2). This, again, demonstrates that behavioral syndrome structure is shared among the

populations. Third, under the adaptive hypothesis, genetic correlations are expected to

decrease by about 50% each generation with a corresponding decrease in phenotypic

correlations (Appendix S1); these correlations were, instead, constant (Figure 3).

Despite this conservation of behavioral syndrome structure, populations did exhibit

some degree of divergence. Importantly, however, this divergence was constrained by the

structure of behavioral syndromes (Table S3). Specifically, the populations have diverged

in their multivariate averages (Figure 4, Table S2) and in the magnitude of genetic variation

present (Figure S3), but this divergence aligned with the shared behavioral syndrome (i.e.

shared subspaces) of the populations (Table S3). The divergence in magnitude of genetic

variation was driven by the three easternmost populations having less genetic variation

than the Dunnigan, CA population (Figure S1). Whether this represents a loss of variation

for the three eastern populations or the accumulation of variation for the western

population is not currently clear.

We thus have three key results demonstrating conservation of behavioral

syndromes at the genetic level despite differences among populations in average behavior.

This conservation of behavioral syndrome structure has had the effect of constraining

population divergence. Jointly these results provide strong support for the constraints

hypothesis for the expression of behavioral syndromes. However, the first two results—

shared subspaces and correlated $\mathbf{g}_{max}$'s—could also result under the adaptive hypothesis if

the selective pressures each of the populations experienced were the same. We consider

this unlikely because of the degree of geographic separation among populations (Figure 2)

and because $\mathbf{g}_{max}$ was most similar among populations that were geographically most

separated (Figure 2). Moreover, our third main result directly contradicts the adaptive

hypothesis: if trait correlations, like those of behavioral syndromes, arise due to the

adaptive hypothesis and therefore selection-induced linkage disequilibrium, they are

expected to rapidly degrade under random mating[11,12]. In direct contradiction of this

expectation we observed that correlations did not decrease across generations (Figure 3).

Put another way, our first two results—which showed that the multivariate composition of

behavioral syndromes was shared among populations—are consistent with the predictions

of the constraints hypothesis. Next, our third result—the maintenance of behavioral

correlations despite random mating—demonstrates the failure of predictions made by the

adaptive hypothesis.

In contrast to our findings, a previous study with stickleback[13] found that two

populations differed in the magnitude of heritabilities and genetic correlations between

behaviors, albeit with overlapping confidence intervals. This earlier result is suggestive of

behavioral syndromes emerging from correlated selection but predates the development of

methods for the comparison of the full multivariate structure among populations as done

here. Unfortunately, examination and comparison of pairwise genetic correlations can be

misleading in regard to evolutionary constraints. For example, while two traits may be

uncorrelated, each may be strongly correlated to others and differ in the magnitude of

genetic variation available, thereby generating constraints not apparent from correlations.

Instead, it is the overall multivariate structure of variation that shapes evolutionary

responses and leads to quantitative constraints on evolution[24,25]—constraints that we have

evaluated for behavioral syndromes for the first time here.

Behaviors like those measured here—exploratory behaviors and responses to

predation threat—are frequently assumed to have been under selection and their

responses to selection have been assumed to be unconstrained. Our results suggest an

unrecognized and important role for behaviors like these and for behavioral syndromes in

the evolution of populations. That behavioral syndromes are important for trait divergence

among populations implies that they may also affect speciation patterns. While correlations

among morphological traits have been implicated in constraining speciation[14,26,27] and

sexually selected behaviors have been argued to play a role in pre-zygotic reproductive

isolation in allopatry[28], the potential for behaviors like those examined here to influence

speciation has not been critically evaluated. Moreover, genetic correlations are common

across trait types[2,29,30] and, given that genetic variance is rarely limited[31], these

correlations should be expected to quantitatively constrain divergence and responses to

selection even if they do not act as absolute constraints. Because behaviors exhibit

heritabilities similar to those observed for life-history and physiological traits[29,32,33], our

results also suggest that correlations among labile traits—traits that are highly plastic—

may actually play a role in shaping differences among populations and species. Our findings

suggest that a more diverse range of traits and behaviors may have shaped population

divergence and speciation patterns than previously considered.

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Author Contributions

N.A.D. and A.H. conceived the project and supervised the gathering of data, R.R analyzed the
data, and N.A.D., A.H., and R.R. wrote the manuscript.

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Figures.

Figure 1. Two contrasting hypotheses can explain the presence of genetic correlations among behavioral traits (i.e. behavioral syndromes): Genetic constraints arising from pleiotropy and shared molecular mechanisms should lead isolated populations to express the same behavioral syndrome (top left panel). As a result, the vector correlations between major axes of genetic variation ($\mathbf{g_{max}}$) are predicted to be approaching 1 (bottom left panel). Alternatively, selection-induced linkage disequilibrium should lead to differing orientation and strength of behavioral syndromes when selective pressures differ among populations (top right panel). The vector correlation between $\mathbf{g_{max}}$'s should therefore be below 1 (bottom right panel).

Figure 2. The genetic architecture of behavioral syndromes is shared in four isolated populations of *Gryllus integer*. Values along the diagonal (peach shading) represent describe the multivariate structure of behavioral variation: first the size (total genetic variance), second the shape (percent of variance explained by the major axis of genetic variation, $\mathbf{g_{max}}$), and, third, orientation (vector correlation between $\mathbf{g_{max}}$ and conserved genetic subspaces $\mathbf{h_{1\text{-}3}}$). Off-diagonal elements (green shading) represent the correlation between the $\mathbf{g_{max}}$ of each population (top row) and the probability that alignment differed from 1: ** P < 0.01, * P < 0.05.

Figure 3. A) Additive genetic (r_A) and B) phenotypic correlations (r_P) remained stable over the course of three successive generations compared to theoretical expectations based on selection-induced linkage disequilibrium and random mating (r_A: P_{mcmc} for difference from expectations under selection-induced linkage disequilibrium > 0.85; r_P: P_{mcmc} F_2 > 0.80). Error bars correspond to 95% credibility intervals around the posterior mean.

Figure 4. Evolutionary divergence in the structure of behavioral syndromes occurs along shared axes of genetic variation. A-C) Correlations between pairs of traits that exhibit the greatest variation in divergence (Table S2). Points represent breeding values for each individual within a population centered around the population mean for that trait. >50% of divergence was in latency to emerge from shelter by antipredator response activity D) Population specific divergence in average behaviors. Population-specific **G** matrices were visualized by transforming estimated breeding values for each trait based on the

319   divergence among populations. Ellipses represent the 95% confidence ellipses for each

320   population centered at the multivariate species mean (DUN: Dunnigan CA, AG: Aguila AZ,

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Methods

Cricket collection

We collected adult female crickets from four populations throughout the southwest and western US: Socorro, NM; Las Cruces, NM; Aguilla, AZ; and Dunnigan, CA (Figure 1) during the summer of 2017. Crickets from these locations are formally recognized as all members of *Gryllus integer* but additional splitting out of subspecies or different species based on population genetic structure is currently being considered (Weissman, personal communications). Around 50 females on average were collected from each population (Table S1) and taken to animal housing facilities at North Dakota State University. Females were housed individually in 0.71 L containers and provided with ad libitum food (Purina Chick Starter) and water (water was provided in glass vials capped with cotton). Each cricket was also provided with a small piece of cardboard egg carton for shelter. The cricket housing room was maintained on a 12:12 dark:light cycle reversed such that the room was dark during daytime hours. The housing room was kept at ~27C.

Breeding design

Females collected from the field (generation P) were assumed to have mated prior to capture (with the possibility of multiple mating, as is common in the genus[34]) and were allowed to oviposit in water vials while in their containers. Offspring of these females (termed generation F_0 as sires were unknown) hatched in their dam's container and were moved to individual housing prior to maturation. We assayed the behavior of 387 F_0 individuals (see below) upon maturation (Table S1). After behavioral trials, F_0 individuals were assigned to breeding pairs such that individual males were mated to multiple

randomly assigned females from the same population but different dams according to a

standard full-sib, half-sib breeding design[20]. Matings were conducted as follows: females

were moved from their normal housing containers to a larger container (34.6 × 21 × 12.4

cm) along with their food dish, water vial, and egg carton shelter. After the female had been

transferred, the assigned male was likewise moved to the large container, also with their

food dish, water vial, and egg carton. The male and female remained in these containers for

24 hours to allow sufficient time for courtship and multiple mating. After 24 hours the male

and female crickets were returned to their original containers. If males were to be mated

with additional females, they were allowed a minimum of 24 hours before repeating the

above procedure. These F_0 females were subsequently allowed to oviposit into water vials

within their containers. Resulting F_1 offspring were moved to individual housing prior to

maturation and had their behaviors assayed upon maturation. After behavioral assays, F_1

individuals were likewise paired with F_1 individuals of the same population but different

sires in the same manner and resulting F_2 offspring moved to individual housing and had

their behavior measured upon maturation. This resulted in the behavioral testing of 395 F_1

individuals and 163 F_2 individuals (Table S1). Across the three generations this

represented behavioral testing of 946 individual crickets.

Behavioral testing

All behavioral tests followed standard procedures previously validated in the literature for

Gryllid crickets[35-41]. Below, we briefly describe these tests and their ecological relevance.

Latency to emerge from shelter

Gryllid crickets, including *G. integer*, use small burrows and natural cracks for refugee from

predators and to which they retreat when under threat. As a result, latency to emerge from

shelter after disturbance can be considered a proxy for risk-taking behavior or

"boldness"[37]. Here, we conducted latency tests wherein individuals were transferred from

their home containers to small artificial burrows (40 cm^3) placed within a 34.6 × 21 cm

arena. These artificial burrows were capped so that individuals could not immediately

emerge. Crickets were forced to remain in the artificial burrow for two minutes after which

the cap was removed. Crickets were then allowed six minutes and thirty seconds to emerge

from the artificial burrow. During this test we recorded how long it took for an individual

emerge (in seconds). Individuals that did not emerge were given a maximum latency of 390

seconds.

Open field exploratory behavior

Open field tests are a classic behavioral assay across taxa[42] which measure the exploratory

propensity of individuals[43], including for crickets[39-41]. Individuals that move through more

of the arena are considered more thorough explorers[43]. Here we used open field tests to

measure activity and exploratory propensity in a 30 × 30 cm plexiglass arena. Individuals

were introduced into the arena under a small container and allowed to rest for 30 seconds

after introduction. At the end of this 30 seconds, the container was removed and the cricket

was allowed to explore the arena for 3 minutes and 40 seconds. The arena was cleaned

with isopropyl alcohol between trials to remove any chemosensory cues from the arena.

From this test we used Ethovision XT to record the total distance the individual moved

during the trial (cm), the number of unique zones of the arena an individual visited during

the trial, and the variance in velocity of individuals $(cm/s)^2$. This latter measure indicates

whether an individual's speed of exploration was constant (low velocity variance) or

whether individuals had frequent activity bursts punctuated by long bouts of inactivity

(high velocity variance).

Response to cues of predator presence

How individuals respond to cues of predator presence often varies within and among

populations and is likely to covary with fitness[16,44]. Crickets respond to chemical cues of

predator presence by either freezing or increasing activity depending on whether

confronted by predator cues of sit-and-wait or active predators[45,46]. Here we used a

behavioral assay to measure response to cues of predator presence previously validated

with another Gryllid species[39,40]. Specifically, individuals were introduced into a 15 cm

diameter circular arena (7.5 cm height) the floor of which is covered with dry filter paper

that had been soaked with diluted excreta from leopard geckos (*Eublepharis macularius*).

All leopard geckos were fed a diet of *G. integer* with occasional diet supplementation of

mealworms (i.e. larval *Tenebrio molitor*) and the related decorated cricket (*Gryllodes*

sigillattus). Crickets were introduced to a portion of the arena without predator cue under

a small shelter. After a 30s rest period, the shelter was removed and the individual allowed

to freely move throughout the arena for 3 minutes and 40 seconds. We then used

Ethovision XT to record the total distance an individual moved during the trial (cm). Total

distance moved during the predator cue trial, the latency to first movement (in seconds),

and the variance in velocity were used in subsequent analyses.

Statistical analyses

G *matrix estimation*

We used multi-response mixed effect animal models[47] implemented using the MCMCglmm

package in R[48] to estimate genetic variances and covariances. We included the effects of

temperature of the behavioral arena room, sex of the individual, and mass of the individual

as fixed-effects and the individual relatedness matrix (based on the known pedigree) as a

random effect. Traits for which variances and covariances were estimated were: (i) the

latency that an individual emerged from the shelter during the trial (censored gaussian),

(ii) the distance moved during the open field trial (gaussian), (iii) the number of unique

zones an individual visited during the open field trial (Poisson), (iv) the log-transformed

variance in velocity during the open field trial (gaussian), (v) the square-root transformed

distance an individual moved during the predator cue response trial (gaussian), (vi) the

latency to initiate movement in the antipredator response trial (Poisson) and (vii) the log-

transformed variance in velocity during the antipredator response trial (gaussian). Multi-

response models were fit individually by population with each population's variances and

covariances estimated from the posterior of an MCMC chain of 4.8×10^6 iterations, with an

800,000 burn-in period and a thinning interval of 4,000. A prior that was minimally

informative for both variances and covariances was used. All variances and covariances

were estimated at the additive genetic level and on the latent scale (Table S4).

Population comparisons of **G**

To compare behavioral syndrome structure at the additive genetic level we:

429   (i)   compared alignment of dominant eigenvectors among populations (i.e. $\mathbf{g_{max}}$,

430   Table S5[10]);

431   (ii)   tested whether populations exhibited shared subspaces of **G** [21];

432   (iii)   tested how populations differed in their variances and covariances (i.e.

433   genetic covariance tensor analysis[21]).

We followed the recommendations of Aguirre et al. (2014) in that all tests were based on

the full MCMC posterior distributions, null distributions for population comparisons were

based on randomizations of breeding values, and angles estimated against null

expectations of 0 (see also[49]). Specifically, to compare whether eigenvectors were

significantly aligned, we generated a random distribution of vector correlations following

[14]. The critical values of this distribution were 0.93 (P < 0.001), 0.85 (P < 0.01), 0.71 (P <

0.05) and 0.62 (P < 0.1). To assess the significance of eigenvalues of **H** and **E** against

random expectations, we calculated the largest posterior quantiles for which these

distributions did not overlap (Figures S2 and S3 respectively). This threshold serves as a

Bayesian probability in favor of the observed distribution being generated by patterns

other than chance (hereafter, P_{mcmc}). We interpret these probabilities on the following

scale: P_{mcmc} < 0.7: poor evidence of difference compared to random expectations; P_{mcmc} >

0.8: moderate evidence of difference compared to random expectations; P_{mcmc} > 0.9 strong

evidence of difference compared to random expectations; P_{mcmc} > 0.95: very strong

evidence of difference compared to random expectations.

Supplementary Materials

Appendix S1. Comparing observed patterns of phenotypic correlations with expectations
based on SILD and random mating.

Genetic correlations are expected to decline every new generation when they arise from

selection-induced linkage disequilibrium (SILD). Conner [12] showed that in absence of both

selection and linkage, the magnitude of genetic correlations under a SILD model and

random mating should halve with every new generation. In contrast, genetic correlations

caused by pleiotropic effects are expected to remain stable even after multiple generations

of random mating. From this we expect the additive genetic correlation in the F_2 generation

to be half as large as that of the F_1 generation:

$$(1)\; r_{\text{A, F}_2} = \frac{1}{2} r_{\text{A, F}_1}$$

To test this hypothesis, we estimated the observed genetic covariance based on

multivariate animal models including the data of the F_0 and F_1 and F_1 and F_2 generations

respectively. These models were specified similarly to the multivariate models described in

the main text, but population origin was added as a fixed effect to control for population

differences in average trait expression. We transformed the genetic covariance matrix to a

correlation matrix and calculated the observed average correlation coefficients, $r_{\text{A } observed}$,

for every generation. We then compared, $r_{\text{A } observed}$ to the expected average correlation

shown in equation (1). Since our multivariate models were specified within a Bayesian

framework, we performed these operations on every slices of the posterior estimates to

obtain 95 % credible intervals of observed and expected estimates and base our inference

on the overlap between the two 95 % credible intervals.

Because phenotypic correlations result from the combined influence of genetic and

environmental influences, it is also possible to compare the expected decline in phenotypic

correlations under SILD and random mating with the phenotypic correlations observed in

every generation. The expected relationship between the phenotypic covariance in

generation F_0, F_1 and F_2 can therefore be expressed by the following sets of equations:

$$(2)\ \mathbf{P_{F_0}} = 2\ \mathbf{G_{F_1}} + \mathbf{E}$$

$$(3)\ \mathbf{P_{F_1}} = \mathbf{G_{F_1}} + \mathbf{E}$$

$$(4)\ \mathbf{P_{F_2}} = \frac{1}{2}\mathbf{G_{F_1}} + \mathbf{E}$$

Where $\mathbf{P_{F_x}}$, and $\mathbf{G_{F_x}}$ represent the phenotypic and genetic covariance matrices for

generation F_x respectively and $\mathbf{E}$ represents the overall environmental covariance matrix

estimated across all generations. Note that these formulas only apply to the covariance

among traits (i.e. the off-diagonal elements) while the genetic variances (i.e. the diagonal

elements) are kept constant across generations. $\mathbf{G_{F_1}}$ was used as our reference here

because this generation had the largest sample size and thus had the most influence on the

estimation of $\mathbf{G}$. We estimated $\mathbf{G_{F_1}}$ from the animal model used to estimate equation (1) and

we estimated $\mathbf{E}$ using a model containing the data from all populations for every

generations.

We then compared these expected covariances to the observed phenotypic

covariance estimated for each generation separately by fitting multivariate models

excluding pedigree information so that only the phenotypic covariance was estimated. As

above, we converted each covariance matrix to a correlation matrix and compared the

observed average correlation coefficient $r_{P\ observed}$, with the average correlation coefficient r_P

$_{expected}$ for every generation.

Table S1. Number of sires, dams and phenotyped offspring for each population and
generation. The F_0 population represents the first population established in the laboratory
after collection of gravid females in the field. N indicates the number field collected females
for the F_P generation.

Population and generation	No. of sires	No. of dams	No. offspring in pedigree	No. phenotyped offspring
Dunnigan (CA) (N = 38)				
F_0	0	8	30	30
F_1	6	8	36	36
F_2	6	8	35	35
Total	12	24	101	101
Aguila (AZ) (N = 71)				
F_0	0	32	144	143
F_1	16	21	92	92
F_2	2	2	16	16
Total	18	55	252	251
Socorro (NM) (N = 65)				
F_0	0	19	112	112
F_1	10	20	126	125
F_2	4	4	8	8
Total	14	43	246	245
Las Cruces (NM) (N = 38)				
F_0	0	13	105	102
F_1	20	47	157	142
F_2	13	16	104	104
Total	33	76	366	349

Table S2. Eigenvectors of phenotypic divergence (**d**), conserved genetic variation (**h**) and divergence in **G** (**e**). Traits legend: Latency = latency to exit from the shelter, OF.Dist = distance travelled in the open-field test, UZ = number of unique zones explore in the open-field arena, OF.Var.Velo = variance in velocity in the open-field test, AP.Dist = distance travelled in the antipredator response test, AP.Lat.Mov = latency to initiate movement in the antipredator response test, AP.Var.Velo = variance in velocity in the antipredator response test.

Traits	d_1	d_2	h_1	h_2	h_3	E1 (53 %) e_{11}	E2 (31 %) e_{21}	e_{22}
Latency	**-0.32**	**0.90**	**-0.31**	**0.86**	**0.47**	**0.63**	0.08	**0.77**
OF.Dist	**-0.80**	**-0.34**	0.18	**-0.37**	**0.88**	-0.13	**-0.79**	**0.42**
UZ	-0.02	-0.06	0.02	-0.02	0.06	-0.01	-0.08	0.06
OF.Var.Velo	-0.09	-0.06	0.02	-0.02	0.07	-0.02	-0.10	0.06
AP.Dist	**-0.48**	-0.04	**0.93**	**0.35**	0.02	**-0.74**	**-0.59**	**-0.45**
AP.Lat.Mov	-0.05	0.24	-0.10	-0.01	0.02	0.16	0.07	0.15
AP.Var.Velo	-0.07	-0.03	0.05	0.02	0.01	-0.05	-0.05	-0.02
% Variance explained	58.2	31.4	33.1	33.0	32.9	97.4	69.5	30.4

Table S3. Vector correlation (r) between major axes of phenotypic divergence (d_1, d_2), subspaces of conserved genetic variation (h_1, h_2, h_3) and subspaces representing genetic divergence (e_{11}, e_{21}, e_{22}). Probability of significant alignment: *** $P < 0.001$, ** $P < 0.01$, * $P < 0.05$. Italics indicate significant alignment at $P < 0.1$.

	d_1	d_2	h_1	h_2	h_3
h_1	0.49	0.40			
h_2	0.15	0.89**			
h_3	0.85**	0.06			
e_{11}	0.25	*0.69*	0.92**	0.33	0.18
e_{21}	0.90**	0.40	0.72*	0.16	*0.67*
e_{22}	0.38	0.60	0.59	0.34	0.72*

	Latency	OF.Dist	UZ	OF.Var.Velo	AP.Dist	AP.Lat.Mov	AP.Var.Velo
Dunnigan (CA)							
Latency	33.66	0.58	0.66	0.61	*0.88*	**0.95**	0.77
OF.Dist	1.89	21.58	**0.95**	**0.99**	0.57	0.54	0.64
UZ	0.62	**1.96**	0.69	**0.99**	0.55	0.73	0.67
OF.Var.Velo	0.38	**2.18**	**0.34**	0.43	0.57	0.63	0.70
AP.Dist	*-14.73*	2.78	0.27	0.43	36.36	**0.94**	**0.99**
AP.Lat.Mov	**5.20**	0.02	0.27	0.10	***-5.44***	3.43	0.66
AP.Var.Velo	-0.90	0.41	0.08	0.08	***2.09***	-0.15	0.28
Aguila (AZ)							
Latency	20.67	0.60	0.57	0.54	0.76	0.68	0.61
OF.Dist	-2.01	18.48	**0.99**	**0.99**	0.69	0.79	0.60
UZ	-0.17	**1.48**	0.37	**0.99**	0.54	0.70	0.47
OF.Var.Velo	-0.13	**1.52**	**0.17**	0.30	0.61	*0.80*	0.65
AP.Dist	-4.28	2.80	0.10	0.21	17.89	*0.83*	**0.99**
AP.Lat.Mov	0.91	1.13	0.09	*0.15*	-1.27	1.61	0.69
AP.Var.Velo	-0.19	0.15	0.00	0.03	**0.91**	-0.06	0.16
Socorro (NM)							
Latency	20.57	0.54	0.65	0.54	0.51	0.55	0.46
OF.Dist	-1.91	26.26	**0.99**	**1.00**	**0.96**	***0.90***	***0.91***
UZ	-0.48	**2.16**	0.46	**1.00**	***0.91***	*0.89*	*0.82*
OF.Var.Velo	-0.23	**2.78**	0.29	0.41	***0.94***	***0.92***	***0.92***
AP.Dist	-0.74	**12.83**	***1.24***	***1.58***	24.39	**0.99**	**0.99**
AP.Lat.Mov	0.24	***-1.68***	-0.20	***-0.23***	**-2.54**	1.09	0.76
AP.Var.Velo	0.07	***0.97***	0.08	***0.13***	**1.68**	-0.09	0.22
Las Cruces (NM)							
Latency	22.00	0.55	0.64	0.68	0.74	0.69	0.77
OF.Dist	0.43	12.46	**0.99**	**1.00**	***0.92***	*0.83*	*0.84*
UZ	-0.17	**0.65**	0.21	**0.99**	*0.86*	*0.84*	0.77
OF.Var.Velo	0.26	**1.02**	**0.09**	0.18	*0.87*	0.76	***0.94***
AP.Dist	3.16	***4.90***	0.41	*0.54*	14.70	**0.97**	**1.00**
AP.Lat.Mov	-0.59	*-0.92*	-0.10	-0.09	**-1.67**	1.30	0.72
AP.Var.Velo	0.34	*0.38*	0.03	***0.07***	**0.90**	-0.06	0.13

Table S5. Eigen decomposition of **G** for each population. Bold values indicate loadings > 0.25 to help interpretation.

	g_{max}	g_2	g_3	g_4	g_5	g_6	g_7	g_{max}	g_2	g_3	g_4	g_5	g_6	g_7
			Dunnigan (CA)							Aguila (AZ)				
λ	51.11	24.59	17.64	2.28	0.52	0.17	0.13	25.47	17.59	14.52	1.38	0.27	0.16	0.11
% Variance	53.00	25.50	18.29	2.37	0.54	0.17	0.14	42.81	29.56	24.40	2.32	0.46	0.26	0.18
Latency	**-0.66**	**-0.47**	**0.57**	-0.11	0.00	0.01	-0.01	**0.70**	**-0.53**	**0.48**	-0.04	0.00	0.00	0.00
OF.Dist	0.03	**-0.78**	**-0.61**	-0.02	-0.12	-0.05	-0.04	**-0.44**	**-0.84**	**-0.29**	-0.08	0.11	0.03	-0.01
UZ	0.00	-0.08	-0.04	0.12	**0.92**	**-0.33**	-0.15	-0.03	-0.07	-0.03	-0.02	**-0.88**	**0.42**	-0.20
OF.Var.Velo	0.00	-0.09	-0.05	0.05	**0.34**	**0.72**	**0.59**	-0.03	-0.07	-0.02	0.04	**-0.46**	**-0.79**	**0.40**
AP.Dist	**0.73**	**-0.40**	**0.54**	0.10	-0.03	-0.04	0.05	**-0.57**	0.00	**0.82**	0.08	-0.01	0.03	0.04
AP.Lat.Mov	-0.16	-0.01	0.00	**0.98**	-0.14	-0.05	0.05	0.04	-0.09	-0.07	**0.99**	0.01	0.04	-0.02
AP.Var.Velo	0.04	-0.03	0.02	0.08	0.07	**0.60**	**-0.79**	-0.03	0.00	0.04	0.00	0.00	**-0.45**	**-0.89**
			Socorro (NM)							Las Cruces (NM)				
λ	39.13	20.44	12.52	0.83	0.29	0.11	0.08	23.81	17.36	8.39	1.09	0.19	0.10	0.06
% Variance	53.31	27.85	17.06	1.13	0.39	0.14	0.11	46.69	34.05	16.46	2.13	0.37	0.19	0.11
Latency	0.10	**0.99**	-0.10	0.01	-0.02	0.01	0.00	**0.86**	**0.49**	-0.13	-0.01	0.01	0.02	0.00
OF.Dist	**-0.72**	**0.01**	**-0.68**	-0.03	0.10	0.05	0.05	0.23	**-0.61**	**-0.75**	-0.03	-0.07	0.04	0.03
UZ	-0.06	-0.01	-0.04	0.08	**-0.95**	0.19	0.20	0.01	-0.04	-0.03	0.04	**0.91**	**0.37**	0.18
OF.Var.Velo	-0.08	0.00	-0.06	0.04	**-0.26**	**-0.63**	**-0.72**	0.03	-0.05	-0.06	-0.01	**0.39**	**-0.67**	**-0.62**
AP.Dist	**-0.67**	0.14	**0.72**	-0.09	0.00	0.06	-0.05	**0.44**	**-0.62**	**0.64**	-0.10	-0.02	0.04	-0.04
AP.Lat.Mov	0.08	-0.01	-0.06	**-0.98**	-0.09	0.07	-0.09	-0.07	0.08	-0.04	**-0.99**	0.03	0.05	-0.02
AP.Var.Velo	-0.05	0.02	0.04	-0.12	-0.02	**-0.74**	**0.66**	0.03	-0.04	0.03	-0.05	0.11	**-0.64**	**0.76**

Figure S1. (a) Population coordinates along the leading eigentensors capturing most of the variation in divergence (**E**1 and **E**2). (b) Genetic variance in the direction of the leading eigenvectors of eigentensor **E**1 and **E**2. The largest population differences were found between the Dunnigan (DUN) and Las Cruces (LC) populations along **E**1 (Pmcmc > 0.85) and its leading eigenvector e_{11} (Pmcmc > 0.70), all other Pmcmc < 0.70.

Figure S2. Eigenvalues of **H** for a comparison of the first three eigenvectors of each **G** matrix compared against random expectations. None of the eigenvectors of **H** showed significant departure from random expectations (Pmcmc < 0.65), indicating that major axes of genetic variation are highly conserved among populations. (h1: Pmcmc = 0.6, h2: Pmcmc = 0.65, h3: Pmcmc = 0.65)

Figure S3. Eigenvalues of the first three positive eigentensors of **S**. Both $\mathbf{E}_1$, $\mathbf{E}_2$ and $\mathbf{E}_3$ were significantly different than expected from the random distribution (Pmcmc > 0.85), indicating that populations differed along $\mathbf{E}_1$, $\mathbf{E}_2$ and $\mathbf{E}_3$. (E1: Pmcmc = 0.85, E2: Pmcmc = 0. 85, E3: Pmcmc = 0. 85)