

1 **Title:** The fitness cost of mismatch repair mutators in *Saccharomyces cere-*
2 *visiae*: partitioning the mutational load

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7 **Running title:** The fitness cost of mismatch repair mutators

8 **Abstract:** Mutational load is the depression in a population's mean fitness
9 that results from the unceasing influx of deleterious mutations. Here, we
10 directly estimate the mutational load in a population of mismatch repair-
11 deficient *Saccharomyces cerevisiae*. We partition the load into two compo-
12 nents. To estimate the load due to nonlethal mutations, we measure the
13 competitive fitness of hundreds of randomly selected clones from both mis-
14 match repair-deficient and -proficient populations. Computation of the mean
15 clone fitness for the mismatch repair-deficient strain permits an estimation
16 of the nonlethal load, and the histogram of fitness provides an interesting
17 visualization of a loaded population. In a separate experiment, in order to
18 estimate the load due to lethal mutations (i.e. the lethal mutation rate), we
19 manipulate thousands of individual pairs of mother and daughter cells and
20 track their fates. These two approaches yield point estimates for the two
21 contributors to load, and the addition of these estimates (0.016 and 0.009
22 respectively) is approximately equal to the separately measured short-term
23 competitive fitness deficit for the mismatch repair-deficient strain. This
24 correspondence suggests that there is no need to invoke direct fitness ef-
25 fects to explain the fitness difference between mismatch repair-deficient and
26 -proficient strains. These results enhance our understanding of mutational
27 load, a classic population genetics concept, and we discuss their implications
28 for the evolution of mutation rates.

29 **Keywords:** mutational load, experimental evolution, evolution of mutation
30 rates, yeast

31 Introduction

32 An evolving population experiences a continual influx of mutations, the
33 vast majority of which, excluding neutral mutations, are deleterious (Fisher
34 1930). A deleterious allele in a haploid population will attain an equilib-
35 rium frequency that is the quotient of the mutation rate to that allele and
36 the selection coefficient against it (Danforth 1923). The influx of deleteri-
37 ous mutations causes a depression in the population’s mean fitness that is
38 termed the mutational load (Muller 1950), and the load at equilibrium is
39 equal to the deleterious mutation rate (Haldane 1937). Because all popula-
40 tions experience mutation, all populations experience load, and a substantial
41 proportion of the genetic variance for fitness in natural populations is due
42 to mutational load (Charlesworth 2015).

43 Mutational load is closely connected to the evolution of mutation rates.
44 Consider an asexual population in which there is genetic variation for the
45 mutation rate: within such a population, distinct lineages with differing mu-
46 tation rates experience differing loads and therefore possess differing mean
47 fitnesses. In this way a downwards selective pressure on the mutation rate is
48 realized. This pressure is indirect in the sense that modifiers of the mutation
49 rate are subject to selection without affecting any physiological property re-
50 lated to fitness. The existence of ancient and highly conserved systems for
51 replication fidelity (including proofreading, mismatch repair, and nucleotide
52 excision repair) attests to the persistence of this selective pressure.

53 In evolving populations, lineages with higher mutation rates (“mutators”)

54 are continually produced by mutation to any of numerous mutation-rate-
55 affecting loci. In the absence of beneficial mutations, the expected frequency
56 of mutators within a population depends on the increase in the deleterious
57 mutation rate caused by the mutator allele, the rate of mutation from wild
58 type to mutator, and the mean selective effect of newly arising deleterious
59 mutations (Johnson 1999; Desai and Fisher 2011). Investigations of natu-
60 ral and clinical isolates of *Escherichia coli* and other bacteria have shown
61 that mutators one to two orders of magnitude in strength, often defective in
62 mismatch repair, are present at low but notable frequencies in many pop-
63 ulations (Jyssum 1960; Gross and Siegel 1981; LeClerc et al. 1996; Matic
64 et al. 1997; Oliver et al. 2000; Denamur et al. 2002; Richardson et al.
65 2002; Trong et al. 2005; Denamur and Matic 2006; Gould et al. 2007;
66 Raynes and Sniegowski 2014). Evolution experiments conducted with *E.*
67 *coli* have demonstrated that mutators can displace wild types by virtue of
68 their increased access to beneficial mutations (Cox and Gibson 1974; Chao
69 and Cox 1983; Sniegowski et al. 1997; Giraud et al. 2001; Shaver et al.
70 2002; de Visser and Rozen 2006). Similar findings have been reported for
71 *Saccharomyces cerevisiae* (Thompson et al. 2006; Raynes et al. 2011, 2018).
72 However, in contrast to findings in prokaryotes, mismatch repair-deficient
73 (henceforth *mmr*) or other types of strong mutators have not been found
74 in natural *S. cerevisiae* populations (but see Bui et al. 2017; Raghavan et
75 al. 2018), though weaker variation for the mutation rate has been detected
76 (Gou et al. 2019). One explanation for this difference could be that *mmr*
77 mutators experience higher load, compared to the wild-type, in *S. cerevisiae*

78 than they do in *E. coli*. Indeed, it has been observed by several investigators
79 that haploid *mmr* *S. cerevisiae* strains decline in frequency in the short term
80 when co-cultured with wild-type strains (Thompson et al. 2006; Raynes et
81 al. 2011, 2018; Bui et al. 2017), even if they eventually out-adapt the wild
82 type. While this short-term deficit of the fitness of *mmr* mutators relative
83 to the wild type has been attributed to increased mutational load in the
84 *mmr* strain, the evidence that this is the case has been mostly circumstan-
85 tial (but see Wloch et al. 2001) because it is generally difficult to rule out
86 an additional direct fitness effect of any allele thought to cause an indirect
87 fitness effect (Raynes and Sniegowski 2014).

88 In this work, we establish, by short-term competitive fitness assays and
89 in agreement with prior studies, that an *mmr* haploid *S. cerevisiae* strain
90 is substantially less fit than an otherwise isogenic *MMR+* (i.e. wild-type)
91 strain. This fitness difference could be caused solely by load, or solely due
92 to some direct fitness effect of the *mmr* phenotype; or it could be some
93 combination of the two. We develop separate assays to measure the compo-
94 nents of load due to nonlethal and lethal deleterious mutations. To estimate
95 the load caused by nonlethal deleterious mutations, we randomly sampled
96 hundreds of clones from *mmr* and wild-type populations and measured the
97 competitive fitness of each. The resulting histogram of the distribution of
98 fitness of the *mmr* population provides an illustration of the effect of a high
99 mutation rate on population mean fitness. We find that the means of these
100 distributions differ, indicating substantial load for the *mmr* strain, but not

101 fully accounting for the total observed fitness difference between *mmr* and
102 wild type strains. To estimate the lethal mutation rate under the two dif-
103 ferent mutational regimes, we manipulate single cells to track the fate of
104 mother/daughter duos. We show that these two separately measured com-
105 ponents of load—due to nonlethal and lethal mutations—approximately sum
106 to the measured fitness difference between the strains; hence we find no rea-
107 son to suppose a direct fitness effect for *mmr*. Investigations with diploid
108 versions of our strains provide support for this conclusion. We discuss some
109 implications of these findings for continued experimental and theoretical in-
110 vestigations into the evolution of mutation rates.

111 **Materials and Methods**

112 **Data analysis and figure production**

113 Data processing and analysis was performed in R (R Core Team 2019) and
114 RStudio (RStudio Team 2015). Graphical output was produced using the
115 package `ggplot2` (Wickham 2016).

116 **Strains**

117 yJHK112, a haploid, prototrophic, heterothallic, MATa, BUD4-corrected,
118 ymCherry-labeled W303 strain, was used as the haploid wild type in all
119 work described in this paper. yJHK111, labelled with ymCitrine (a variant
120 of YFP) and otherwise isogenic to yJHK112, was used as the “reference

121 strain” in all haploid fitness competitions. These strains have been pre-
122 viously described (Koschwanez et al. 2013) and were generously provided
123 by the laboratory of Andrew Murray, Harvard University, Cambridge, MA.
124 An *msh2* Δ derivative of yJHK112, in which the wild-type *MSH2* allele was
125 replaced with a kanMX geneticin (G418) resistance cassette (Wach et al.
126 1994), was used as the haploid *mmr* mutator strain in all work described
127 in this paper. This strain was generously provided by Yevgeniy Raynes of
128 the laboratory of Dan Weinreich, Brown University, Providence, RI and has
129 been previously described (Raynes et al. 2018). The kanMX cassette has
130 been shown to not have a negative effect on growth (Baganz et al. 1997;
131 Goldstein and McCusker 1999).

132 We constructed diploid versions of each of the three above strains by trans-
133 forming (Gietz and Schiestl 2007) each with plasmid pRY003, temporarily
134 providing a functional HO locus allowing mating type switching and sub-
135 sequent mating. pRY003 was a gift from John McCusker (Addgene plas-
136 mid #81043; <http://n2t.net/addgene:81043>; RRID:Addgene_81043). The
137 diploid state of resulting isolates was confirmed by (1) ability to produce
138 tetrads after plating to sporulation media; (2) by flow cytometry for total
139 genomic content (following Gerstein et al. 2006); and (3) by the presence
140 of a PCR product for both the *MATa* and *MAT α* loci. The *mmr* diploids
141 would not sporulate, but were confirmed to be diploids by the other two
142 methods.

143 **Growth conditions**

144 The liquid media for all fitness competitions was synthetic dextrose (SD)
145 minimal media containing yeast nitrogen base at a concentration of 6.7 g/L
146 and glucose at a concentration of 1.5 g/L (0.15%), supplemented with tetra-
147 cycline (15 mg/L) and ampicillin (100 mg/L). Fitness competitions were
148 conducted in volumes of 200 ul media in deep polypropylene 96-well plates
149 (Nunc 260251) sealed with flexible caps (Nunc 276002) and shaken at 1000
150 rpm with an orbit of 3mm (Corning LSE 6780-4) at a temperature of 30 C.
151 Initial growth in liquid for the lethal event assays was performed in SD as
152 described above but without antibiotics, in flasks shaken at 200 rpm at 30C.
153 Growth on agar SD (2% glucose, no antibiotics) plates for the lethal assays
154 took place at room temperature, approximately 24 C.

155 **Competitive fitness assays and isolation of clones**

156 Short-term competitive fitness assays between wild-type and *mmr* genotypes
157 were conducted as follows, with the interval between each consecutive day
158 spanning 24 hours. Day 1: wild-type, *mmr*, and the *YFP+* reference strain
159 were inoculated from frozen stock into single wells. Day 2: each strain was
160 transferred to a new well with fresh media, diluting 1/100. Day 3: com-
161 peting strains were mixed 1:1 by volume and transferred to new wells with
162 fresh media, diluting 1/100, to create between 6 to 8 replicate competitions.
163 Day 4: competitions were transferred to new wells with fresh media, diluting
164 1/100, and the frequencies of the competitor and reference strain were as-

165 said by flow cytometry (Guava EasyCyte). Discrimination between strains
166 was performed on the SSC/GRN scatter plot. Day 5: the frequencies of the
167 competitor and reference strain were again assayed by flow cytometry. The
168 population density at the end of a 24-hour cycle was $\sim 2 \times 10^7$ cells/mL; the
169 census population size was thus $\sim 4 \times 10^6$ at transfer and $\sim 4 \times 10^4$ just after
170 transfer.

171 The change in frequencies between Days 4 and 5 was used to calculate a se-
172 lection coefficient s . Under a continuous model of growth (Crow and Kimura
173 1970 p. 193)

$$s = \frac{1}{t} \ln \frac{p_t(1 - p_0)}{p_0(1 - p_t)}$$

174 from which a relative fitness $w = 1 - s$ follows. The number of generations,
175 t , was assumed to be the logarithm base 2 of 100, or approximately 6.64. p_0
176 and p_t are the starting and ending frequencies of the genotype being tested
177 (i.e. the frequencies at Days 4 and 5 in the above procedure).

178 We conducted fitness competitions using this procedure in 8 separate blocks,
179 each with multiple replicates as described above. Each block was begun on
180 a different date. For each block, we computed the fitness of the *mmr* strain
181 relative to the fitness of the wild-type strain by subtracting their mutual
182 relative fitnesses to the reference strain. Each block included competitions
183 in both haploid and diploid genotypes. Our final point estimate of the
184 fitness difference between *mmr* and wild-type strains is the mean difference

185 across all blocks, and the 95% confidence intervals (as shown in Fig. 2) were
186 computed from the set of point estimates according to the t-distribution.

187 In two of the six fitness competition blocks, we randomly sampled individual
188 clones. To do so, we additionally propagated the haploid wild-type and *mmr*
189 strains on Day 3 in addition to mixing them 1:1 with the reference strain.
190 Then, on Day 4, we plated these cultures, diluting appropriately, to YPD
191 agar plates. After sufficient incubation, the resulting colonies were picked by
192 pipet tip into wells containing 200 ul YPD, grown for 24 hours, and frozen
193 down by mixing 1:1 with 30% glycerol and storing at -80 C until needed for
194 fitness assays. The random selection of colonies was ensured by either (1)
195 picking all colonies on a given plate or (2) picking colonies concentrically
196 from a randomly placed point.

197 Fitness assays for sets of isolated clones were conducted on a single 96-well
198 plate, which allowed us to assay the fitnesses of 88 clones (some wells being
199 reserved for various purposes) or fewer per run. We followed essentially
200 the same procedure as the 5-day competition described above, except that
201 frequencies were estimated at Day 3 and Day 4 instead of Day 4 and Day 5.
202 This modification was made because some clones had such reduced fitnesses
203 that an extra day of growth after mixing 1:1 with the reference strain caused
204 the starting frequency of the clone to depart too greatly from 50%.

205 **Calculation of nonlethal load**

206 Mutational load is classically defined (Bürger 1998) as

$$L = \frac{w_{max} - \bar{w}}{w_{max}}$$

207 where \bar{w} is the mean fitness of the population and w_{max} is the fitness of the
208 fittest genotype.

209 We measured all fitnesses relative to a common fluorescent strain, as de-
210 scribed above. We define the unloaded fitness of each genotype as equal to
211 1 and we expect no beneficial mutations to rise to appreciable frequency in
212 the short course of these experiments. Hence, $w_{max} = 1$ and thus

$$L = 1 - \bar{w}$$

213 We expect measured selection coefficients to be distributed approximately
214 normally around the true value, because of various sources of error includ-
215 ing binomial sampling error, drift, instrument noise, environmental pertur-
216 bations to individual wells (within-batch effects) and among 96-well plates
217 (across-batch effects). To eliminate across-batch effects, for each run we ad-
218 justed all measured fitnesses by a constant c such that the mode fitness is
219 1 (equivalently, such that the mode selection coefficient is zero). Once this
220 adjustment has been made,

$$L = 1 - \text{mean}(W)$$

221 where W is the vector of all sampled clone fitnesses, or equivalently

$$L = \text{mean}(S)$$

222 where S is the vector of all sampled clone selection coefficients. A 95%
223 confidence interval for the nonlethal load was computed by bootstrapping
224 from the measured fitnesses of all sampled clones.

225 **Lethal event assay**

226 Strains were inoculated from frozen stock into 6 mL SD in a flask, and
227 transferred to fresh media after 24 hours, diluting 1/100. After an overnight
228 of growth, a streak from the culture was made onto an SD agar plate and
229 five single cells with nascent buds were physically isolated by means of a
230 Zeiss (West Germany) micromanipulating microscope fitted with a Singer
231 Instruments (Somerset, UK) dissecting needle. These cells were periodically
232 checked over the next few hours and the daughters physically separated once
233 developed. These daughters became the founders of microcolonies that were
234 allowed to grow at room temperature for ~20 hours, reaching an average
235 size of 23.3 cells (22.8 for the wild type, 23.8 for *mmr*, difference not signif-
236 icant at $p > 0.6$). These microcolonies were then dissected into a gridlike
237 arrangement of single cells (step 1 in Fig. 1). These cells were then checked
238 at intervals of one to two hours and daughters separated as soon as possi-
239 ble (step 2 in Fig. 1). The colonies resulting from these mother/daughter
240 duos were checked at ~24, ~48, and in some cases ~72 and ~96 hours after
241 separation. A lethal event was recorded when the growth of a mother or

242 daughter lineage ceased. In such cases cessation of growth was sometimes
243 immediate and sometimes occurred after a few generations. In the latter
244 cases the growth was generally markedly slowed by the first observation. In
245 a few cases, slow but unceasing growth was noted: these are presumed to be
246 cases in which a strongly deleterious mutation occurred, though we stress
247 that this assay was not designed to detect nonlethal deleterious mutations.
248 As described in the Discussion, the difference in the rate of lethal events
249 between the wild type and *mmr* strains was used as the estimate for the
250 lethal mutation rate for the *mmr* strain. A 95% confidence interval for this
251 difference in rates was computed by the `prop.test` function in R (Newcombe
252 1998).

253 **Fluctuation assays**

254 To measure the mutation rate to 5-fluoroorotic acid (5-FOA) resistance, we
255 employed the following procedure. Strains of interest were inoculated into
256 10 mL YPD, grown in flasks with shaking at 30 C for ~24 hours, and then
257 transferred to 30 mL fresh YPD diluting such that ~200 cells were passaged,
258 in replicates of 5. After ~48 hours of growth, each replicate was plated
259 without dilution to SD + 5-FOA (1 g/L) agar plates to estimate density
260 and absolute number of resistants, and plated with a 10^{-5} dilution to YPD
261 agar plates to estimate total population density and absolute number. Plates
262 were counted after ~48 hours of growth and mutation rates were estimated
263 using the maximum likelihood method of Gerrish (2008). For each round

264 of fluctuation tests, we estimated mutation rates for both wild-type and
265 (putative) *mmr* strains simultaneously in order to minimize the influence of
266 any uncontrolled sources of variation.

267 **Results**

268 **Mutation rate elevation in *mmr* strain**

269 To confirm that the *mmr* (*msh2* Δ) strain is a mutator, we conducted fluctu-
270 ation tests using resistance to 5-FOA as the selectable phenotype. Averaged
271 across replicate fluctuation tests, we found a 20.8-fold increase (95% CI:
272 13.4- to 28.3-fold) in the mutation rate for the *mmr* strain relative to the
273 wild-type (Fig. S1). This is likely an underestimate of the effective genome-
274 wide increase in the mutation rate because *mmr* mutators have a greatly
275 elevated indel rate for homopolymeric runs (Lang et al. 2013), of which
276 *URA3*, the main locus involved in this fluctuation test, is relatively devoid
277 (author's own calculations).

278 **Fitness disadvantage of *mmr* compared to wild-type**

279 We competed *mmr* and wild-type strains against a common YFP+ reference
280 strain. We found the mutator to be less fit than the wild-type, with an
281 average fitness deficit, expressed as a selection coefficient per generation, of
282 2.3% (Fig. 2).

283 **Estimation of nonlethal load**

284 We randomly sampled individual clones from both *mmr* and wild-type pop-
285 ulations and measured the competitive fitness of each clone. The sampled
286 fitness distributions are shown in Fig 3. The *mmr* strain's fitness distribu-
287 tion has a prominent left tail of less fit individuals. We calculated the load
288 as the difference between the mode and the mean fitness; this is ~zero for
289 the wild-type strain and ~1.7% for the *mmr* strain.

290 **Estimation of lethal mutation rate**

291 To assay lethal mutation rates, we manipulated single *S. cerevisiae* cells,
292 separating mother/daughter duos and tracking events in which one member
293 of the duo failed to found a colony. The procedure is shown in Fig. 1 and
294 explicated more fully in the methods. Assaying over 2200 duos for each
295 strain, we found a rate of lethal events of 0.6% in the wild-type strain and
296 1.5% in the *mmr* strain (Table 1). Photographs of representative lethal
297 events are shown in Fig. S3.

298 In our assay we followed separated duos that contained a suspected lethal
299 until growth ceased. In some cases growth never ceased, but doubling times
300 were very slow compared to the usual growth rate; such cases were not
301 counted as lethal events but are tallied separately in Table 1. We also de-
302 tected cases in which both members of a duo were lethal, or both showed
303 strongly reduced growth, and also cases in which the mother cell never di-
304 vided. Because such events probably stemmed from a mutation that oc-

305 curred prior to the division that created the duo under observation, we
306 excluded these events from our analysis.

307 **Results in diploids**

308 From the haploid strains, we constructed diploid *mmr* and wild-type strains.
309 We calculate the nonlethal load in the *mmr* diploid strain as 0.3% (Fig S2).
310 We also measured the difference in population fitness between wild-type and
311 *mmr* diploid strains. We found that the *mmr* diploid is less fit than the wild-
312 type diploid by a selection coefficient of 1.7% (Fig. 4). This difference is
313 less than the fitness difference between wild-type and *mmr* haploid strains
314 by 26% ($p < 0.004$).

315 Discussion

316 Prior work has shown that, over the short term, haploid *mmr S. cerevisiae*
317 strains decline in frequency when competed with a strain that is wild-type
318 for the mutation rate (Thompson et al. 2006; Raynes et al. 2011, 2018; Bui
319 et al. 2017). Consistent with these findings, we find a fitness disadvantage,
320 expressed as a selection coefficient, of 2.3% for *mmr* haploids in short-term
321 fitness competitions (Fig. 2). The magnitude of this selective disadvantage
322 is similar to that in other reports, including Raynes et al. (2011) (2.4%
323 cost), Raynes et al. (2018) (3.3% cost), and Wloch et al. (2001) (4.6% cost,
324 though this is a noncompetitive measure of absolute growth rate).

325 The deleterious mutations that cause load include both lethal and non-lethal
326 mutations. There is no fundamental theoretical distinction between these
327 two classes of mutation insofar as their contribution to load is concerned:
328 in many population genetic models, all members of an asexual population
329 who are not of the least-loaded class are considered to be doomed (Rice
330 2002). However, their different manifestations require different experimental
331 techniques. We therefore developed separate approaches to measure these
332 two components of load.

333 Load due to nonlethal mutations

334 We measured the short-term competitive fitnesses of hundreds of randomly
335 selected clones. The histogram and QQ plot for the haploid wild-type popu-
336 lations (Fig. 3A) suggest that, apart from one less-fit clone, the distribution

337 of fitness for the wild-type strain is essentially normal. The normality of the
338 distribution is consistent with nearly all wild-type clones having the same
339 genotype and thus the same expected fitness, along with many small sources
340 of error in estimation of fitness. One such source of error is drift over the
341 course of the short-term fitness competition. The formula derived by Gallet
342 et al. (2012) suggests that the expected variance in fitness measurement due
343 to drift given our experimental parameters is $\sim 2 \times 10^{-6}$. A larger source of
344 variance is due to sampling error: in the fitness competitions, we estimate
345 the relative frequencies of the competitors at two time points, sampling
346 ~ 8000 cells per time point. We carried out simulations that suggest that the
347 expected variance due to sampling error is $\sim 2.3 \times 10^{-5}$. These two sources of
348 variance, summed, make up about 20% of the observed variance in selection
349 coefficient. The remainder of the variance probably stems from small-scale
350 environmental variation and other unknown sources of error.

351 In contrast to the results in wild types, the fitness histogram and QQ plot for
352 the haploid *mmr* strain (Fig. 3B) are not reflective of a normal distribution.
353 Instead, a prominent left tail of less fit clones demonstrates the effect of
354 mutational load. The mean selection coefficient is -1.7% (95% CI: -1.1% to -
355 2.2%), which is the quantification of the reduction in population mean fitness
356 due to nonlethal load. This reduction accounts for a substantial portion
357 ($\sim 75\%$) of the measured competitive fitness difference (Fig. 2) between the
358 two strains.

359 We note that our estimate of the nonlethal load (1.7%) is likely an under-

360 estimate of the load that would be realized at a longer time scale, because
361 load is maximized when equilibrium (mutation-selection balance) is reached
362 and the populations may not have been at equilibrium when clones were
363 randomly sampled. This same consideration also applies to our short-term
364 fitness assays which showed (Fig. 2) a fitness deficit of 2.3% for the *mmr*
365 strain relative to the wild type.

366 Selection coefficients of about the magnitude we observe here cause changes
367 in relative frequency that are extremely rapid in evolutionary terms. For ex-
368 ample, a selective deficit of 2% would cause a decline from 50% frequency to
369 20% frequency in 70 generations. Observing a rapid initial decline of haploid
370 *mmr S. cerevisiae* strains in competition with wild-types, some investigators
371 (e.g. Grimberg and Zeyl 2005) have attributed the observed fitness differ-
372 ence to an unknown direct cost (i.e. a pleiotropic effect) while others (e.g.
373 Raynes et al. 2018) have assumed that mutational load fully explains the
374 dynamics. The question has remained open, in part because it is nearly
375 impossible to definitively rule out a direct fitness effect of being *mmr*—any
376 attempt to measure such an effect will be confounded by the indirect fitness
377 effects. By quantifying the indirect fitness effects (i.e. load) we seek to deter-
378 mine if a direct effect need be invoked to explain the observed experimental
379 dynamics.

380 It is not surprising that the nonlethal load accounts for only a portion of the
381 observed fitness difference. The nonlethal load assay relies on the growth of
382 deleterious mutants in order to measure their fitness and thus cannot detect

383 mutants who do not grow, i.e. lethal mutations. In order to measure this
384 portion of the load, we designed an assay in which lethal events are directly
385 observed.

386 **Load due to lethal mutations**

387 The lethal mutation rate has long been a matter of interest (e.g. Dobzhansky
388 and Wright 1941). By observing thousands of mother-daughter pairs (duos),
389 we found a rate of lethal events of 0.015 and 0.006 for the *mmr* and wild-type
390 strain, respectively.

391 The wild-type rate, 0.006, is on the order of estimates for the genomic muta-
392 tion rate itself (Drake 1991; Lynch et al. 2008; Zhu et al. 2014; Sharp et al.
393 2018) and therefore cannot possibly reflect the rate of lethal mutations. Our
394 interpretation is that, for the wild-type, all or most observed lethal events
395 were not caused by genomic mutations and are instead best considered to
396 be non-mutational deaths, perhaps caused by fine-scale environmental fluc-
397 tuations, experimental manipulation, or other stochastic sources of insult
398 and stress. Observations of relatively high rates of cell death, too high to be
399 due to lethal mutation, are not uncommon. Replicative aging studies of *S.*
400 *cerevisiae* often observe low but substantial rates of cell death even in very
401 young mother cells (e.g. Chiocchetti et al. 2007; Shcheprova et al. 2008).
402 Rates of cell death on the order of our observed rate for the wild-type strain
403 have also been observed in young bacterial cells (Wang et al. 2010), sug-
404 gesting that relatively high rates of non-mutational, non-age-related deaths

405 are common among microbes. Our assay design ensured that colonies were
406 young (the oldest cell in a microcolony was on average ~ 4.3 generations old)
407 and we did not observe a bias in lethal events towards mothers (Table S1),
408 so we do not attribute the observed lethal events to senescence. In fact, we
409 observed, across both strains, a bias towards the lethal event occurring in
410 the daughter cell. This difference was not statistically significant ($p = 0.14$),
411 although within the *mmr* strain only we observed 10 lethal events in mothers
412 and 23 in daughters ($p = 0.04$). The observed bias towards daughters dy-
413 ing, if not a sampling effect, could be attributable to smaller daughter cells
414 being relatively more vulnerable to stress. Indeed, increased vulnerability of
415 daughters to environmental sources of stress has been previously reported
416 (Knorre et al. 2010).

417 An *a priori* estimation of the wild-type lethal mutation rate can be made
418 as follows. Lang and Murray (2008) conducted careful estimations of the
419 rate of loss-of-function mutations to the *CAN1* locus in a similar (W303)
420 background as the strains used in this work. Multiplying this rate, 1.5×10^{-7} ,
421 by the number of genes thought to be essential for viability, ~ 1100 (Giaever
422 et al. 2002), gives an expected lethal rate in wild-type haploids of 1.7×10^{-5} .
423 This estimate is on the lower end but within the range of observed rates
424 of accumulation of recessive lethals in several experiments conducted with
425 diploids (Wloch et al. 2001; Hall and Joseph 2010; Nishant et al. 2010;
426 Zhu et al. 2014; Jasmin and Lenormand 2016). Such a rate would suggest
427 that we expected to observe about 0.35 lethal mutations in the wild-type

428 strain in our experiment; we actually observed 14. Therefore, we consider
429 the observed rate of lethal events (0.006) in the wild-type to be an estimate
430 of the rate of non-mutational deaths. The corresponding rate for the *mmr*
431 strain is 0.015 (difference significant at $p < 0.001$). Making the assumption
432 that non-mutational deaths equally affect both strains, we take the difference
433 between the wild-type and *mmr* lethal event rates, 0.009 (95% CI: 0.002 to
434 0.015), as the estimate of the lethal mutation rate in the *mmr* strain.

435 This empirical estimate falls within the range suggested by extrapolating
436 from fluctuation assays. Taking the Lang and Murray (2008) wild-type
437 *CAN1* loss-of-function rate and multiplying by the number of essential genes
438 and by the average fold increase for *mmr* strains in a collection of published
439 reports (44-fold; see Table S4) yields a rate of 0.007. In a slightly different
440 approach, taking the average *CAN1* loss-of-function rate of *mmr* strains
441 from published reports (1.5×10^{-5}) and again multiplying by 1100 essential
442 genes yields a rate of 0.017.

443 In many of the lethal events that we observed, growth did not immediately
444 cease but continued for a few generations (Table S2) before halting. Limited
445 growth for a few generations after an ultimately lethal mutation occurs has
446 previously been observed (Mortimer 1955). We also observed morphological
447 defects in several lethal events; one such instance is shown in the bottom
448 panel of Fig. S3.

449 **Diploid findings**

450 The distribution of fitness for *mmr* diploids (Fig. S2) suggests that they
451 are substantially less loaded than *mmr* haploids, as would be expected if
452 dominance attenuates the deleterious effects of new mutations. We calculate
453 the nonlethal load in *mmr* diploids as 0.3%, as opposed to 1.7% in *mmr*
454 haploids: that is, 82% of the load has gone away following diploidization.
455 One interpretation of this finding is that the average h value (weighted by
456 mutational effect size) is ~ 0.2 . Surprisingly, the wild-type diploids were
457 more loaded than the wild-type haploids (compare Fig. S2A to Fig. 3A).
458 We cannot fully explain this observation, but we note that a recent mutation
459 accumulation plus whole-genome sequencing (MA+WGS) study (Sharp et
460 al. 2018) found that wild-type diploids declined in fitness but haploids
461 did not, mostly due to a class of mutations (nondisjunctions leading to
462 aneuploidies) that occurred at a higher rate in diploids; we have possibly
463 observed a similar phenomenon here.

464 The relative difference in short-term competitive fitness between wild-type
465 and *mmr* strains is narrowed by 26% in diploids (Fig. 4). It is some-
466 what surprising that this figure is not larger, because most deleterious and
467 lethal mutations are expected to be closer to recessive than dominant, and
468 because the nonlethal loads are not very different between wild-type and
469 *mmr* diploids. One possibility is that the diploid mutator fixed a deleterious
470 mutation during the process of diploidization, which would account
471 for the discrepancy between the reduction in nonlethal load ($\sim 80\%$) and the

472 reduction in total fitness difference ($\sim 26\%$) in diploids compared to haploids.

473 **Considering the two loads together**

474 The total fitness difference between the haploid wild-type and *mmr* strains
475 could be a consequence of greater mutational load for the *mmr* strain, a
476 direct effect of the *msh2* Δ deletion, or a combination of the two. The ad-
477 dition of the lethal and nonlethal loads ($0.017 + 0.009 = 0.026$) is approxi-
478 mately 13% larger than the measured fitness difference (0.023), a difference
479 that may simply be due to sampling error. The broad equivalence of these
480 measurements, and the direction of the difference, are consistent with the
481 hypothesis that the total fitness difference is solely due to mutational load.
482 Hence, although we cannot strictly rule out the existence of a direct fitness
483 effect, these findings suggest that there is no need to invoke direct effects
484 in explaining the fitness difference between the *mmr* and wild-type haploid
485 strains.

486 The proportion of mutational load that is due to lethal mutations has been a
487 matter of longstanding interest (e.g. Greenberg and Crow 1960). Depending
488 whether we take the denominator to be the fitness difference measured in
489 the competitive fitness assays (2.3%), or the sum of the separately measured
490 nonlethal and lethal loads (2.6%), we found that the lethal load comprised
491 between $\sim 33\%$ and $\sim 40\%$ of the total load, which is in agreement with a
492 previous observation (Wloch et al. 2001). This high proportion may partly
493 be explained by the extremely elevated indel rate experienced by *mmr* *S*.

494 *cerevisiae*. An MA+WGS experiment (Lang et al. 2013) in *mmr* haploids
495 found a ~14:1 indel:substitution ratio (~3:1 when restricted to coding se-
496 quences), while similar experiments with wild-type *S. cerevisiae* have found
497 indel rates that are 3-5% of the substitution rate (Zhu et al. 2014; Sharp
498 et al. 2018). Because frameshifts tend to be loss-of-function mutations, it
499 is not too surprising that the lethal mutation rate as a proportion of the
500 deleterious mutation rate is quite substantial.

501 The proportion of deleterious mutations that are lethal may also be inflated
502 by an underestimation of the total deleterious mutation rate. The load
503 is equal to the deleterious mutation rate only when the population is in
504 mutation-selection balance. This equilibrium is reached instantly for lethal
505 mutations, quickly for deleterious mutations of large effect, and very slowly
506 for deleterious mutations of slight effect (Johnson 1999). The populations
507 in our assays have likely not achieved mutation-selection balance for all
508 mutational classes in the ~35 generations of the fitness assay, in addition to
509 the generations that occurred before freezing (i.e. before the experiment).
510 Hence, our estimate of the total load (2.3%-2.6%) should be considered
511 an estimate of the lower limit for the deleterious mutation rate for *mmr*
512 haploids.

513 **Comparison to results in bacteria**

514 Insofar as *S. cerevisiae* and *E. coli* are two model organisms, from differ-
515 ent domains of life, with which many evolution experiments have been per-

516 formed, it is interesting to compare the loads of mismatch repair mutators in
517 both. It appears that in *E. coli* the relative fitness deficit for *mmr* strains is
518 smaller than it is in haploid *S. cerevisiae*. For instance, Shaver et al. (2002)
519 did not detect a fitness difference between *mmr* and wild-type strains, de
520 Visser and Rozen (2006) did not observe an initial decline in *mutS* frequency
521 when that genotype was competed with the wild type at different starting
522 ratios, and Boe et al. (2000) estimated at most a 1% selective disadvantage
523 for *mmr* mutators. In this context it is relevant to note that there are several
524 reports of *mmr* genotypes in natural *E. coli* isolates (LeClerc et al. 1996;
525 Matic et al. 1997; Denamur et al. 2002), as well as in other types of bacteria
526 (Oliver et al. 2000; Richardson et al. 2002; Trong et al. 2005; Gould et al.
527 2007). In *S. cerevisiae*, in contrast, no functionally *mmr* natural isolates
528 have yet been found (though see Raghavan et al. 2018). It is also interest-
529 ing to note that in a long term evolution experiment, spontaneously arising
530 *mmr* lineages fixed within the first ten thousand generations in 3 of 12 *E.*
531 *coli* populations (Sniegowski et al. 1997). Subsequently, after further tens
532 of thousands of generations, these *mmr* populations appear to have evolved
533 lowered mutation rates (Tenailon et al. 2016; Good et al. 2017). The evo-
534 lution of lowered mutation rates has also been observed in *mmr S. cerevisiae*
535 populations (McDonald et al. 2012) but on a shorter timescale, consistent
536 with the notion that the fitness cost of mismatch repair deficiency is higher
537 in *S. cerevisiae* than in *E. coli*.

538 Such observations suggest that *E. coli* are relatively more robust than *S.*

539 *cerevisiae* to the lack of a functional mismatch repair system, paying less
540 of a cost in load. One reason for this difference could be that the genomic
541 mutation rate in *E. coli* is lower than that of *S. cerevisiae*. An MA+WGS
542 experiment in wild-type *E. coli* found a genomic mutation rate of 0.001 (Lee
543 et al. 2012) while similar experiments in wild-type *S. cerevisiae* have found
544 rates of about 0.004 (Lynch et al. 2008; Zhu et al. 2014; Sharp et al. 2018).
545 Suppose that the deleterious mutation rate as a fraction of the genomic
546 mutation rate is approximately equal in the two organisms. Then, if the
547 relative fold increase in the deleterious mutation rate caused by the lack
548 of mismatch repair is also approximately equal in both, although the fold
549 increase in load between wild-type and *mmr* will be the same, the absolute
550 difference in load, which is what controls the evolutionary dynamics, will be
551 bigger in *S. cerevisiae* than in *E. coli*.

552 Another possible (and more intriguing) factor is differences in the spectrum
553 of mutations combined with differences in the prevalence of certain genomic
554 patterns. In both *E. coli* and *S. cerevisiae*, the indel rate is greatly increased
555 in *mmr* lineages, and the rate of indels is strongly elevated in homopolymeric
556 repeats (HPRs). Both the relative increase from wild-type to *mmr* and the
557 absolute indel rate in *mmr* are higher, and scale upwards faster with HPR
558 length, in *S. cerevisiae* than in *E. coli* (Schaaper and Dunn 1991; Tran et
559 al. 1997; Gragg et al. 2002; Lee et al. 2012; Lang et al. 2013). Examining
560 all coding sequences in the *E. coli* and *S. cerevisiae* genomes, we find that
561 there are significantly more homopolymeric repeats per coding genome, per

562 gene, and per coding base in *S. cerevisiae* than in *E. coli* (Table S5). *S.*
563 *cerevisiae* that are *mmr* are therefore relatively more burdened by indels
564 than are *mmr* *E. coli* which could account for both the apparent larger fit-
565 ness difference between *MMR+* and *mmr* and the corresponding apparent
566 difference in occurrence in natural isolates. We caution that this particu-
567 lar analysis is speculative in nature at this time: one important caveat is
568 that, while this study and others have found large fitness differences between
569 wild-type and *mmr* haploids, *S. cerevisiae* spend most of their time in na-
570 ture as diploids, in which the fitness deficit of *mmr* lineages would be less
571 severe. However, while estimates of the rate of outcrossing in *S. cerevisiae*
572 are very low (Ruderfer et al. 2006), the rate of sporulation, which entails a
573 haploid stage, is not known, and evidence of extensive inbreeding and loss
574 of heterozygosity (Peter et al. 2018) suggest that it is relatively frequent.
575 Recessive deleterious mutations may thus be frequently exposed to selection
576 in natural *S. cerevisiae* populations by both the haploid life cycle stage and
577 loss of heterozygosity from inbreeding, suggesting that diploidy may not be
578 as much of a shield for *mmr* lineages as it otherwise would be. A second
579 caveat is that, even if there is no direct fitness effect of *mmr* in haploids,
580 there could be such an effect in diploids, perhaps due to misregulation of the
581 frequency of recombination events (reviewed in Surtees et al. 2004; George
582 and Alani 2012).

583 Conclusions and future directions

584 We have found that the indirect fitness effects of strong modifiers for muta-
585 tion rates are substantial in haploid *S. cerevisiae*, and that it is not necessary
586 to postulate direct fitness effects in order to explain the selective disadvan-
587 tage of the lack of a functional mismatch repair pathway. This finding is
588 probably most relevant to experimental inquiries of the dynamics of muta-
589 tion rate evolution in which *S. cerevisiae* are the model organism.

590 We have also reported findings relevant to fundamental questions about
591 mutational dynamics, including the lethal mutation rate and the relative
592 ratio of lethal and nonlethal deleterious mutations. By sampling the fitnesses
593 of many individuals we have clearly demonstrated mutational load in an *mmr*
594 population, and from the load we are able to estimate a lower limit for the
595 deleterious mutation rate. We sampled hundreds of clones and were able to
596 obtain a clear picture of the left tail of the fitness distribution for the *mmr*
597 strain, but not for the wild-type strain. If the fitnesses of tens of thousands
598 of clones could be measured, much could be learned about load and other
599 evolutionary dynamics at wild-type mutation rates; such experiments may
600 become possible as methods for high-throughput measurements continue to
601 advance.

602 A limitation of this study is that we captured a snapshot of mutational load
603 at a particular point in time in an evolving population. It would be inter-
604 esting to observe, at a fine scale, how the distribution of fitnesses changes
605 over time as a population approaches mutation-selection balance, adapts,

606 and experiences other population genetic processes.

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614 **Conflict of interest**

615 The authors declare that there are no competing financial interests in rela-
616 tion to the work described.

617 **Data archiving**

618 The authors plan to upload the raw data upon which this study is based to
619 Dryad.

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822 **Figure legends**

823 **Figure 1**

824 Schematic of lethal assay. The arraying in step 1 and separation in step
825 2 were performed by micromanipulation. Microcolonies in step 1 were on
826 average 23 cells large, and were founded by new daughter cells that had
827 themselves been isolated by micromanipulation.

828 **Figure 2**

829 The average competitive fitness deficit (in black, error bars are 95% confi-
830 dence intervals) of the mutator strain relative to the wild-type, expressed
831 as selection coefficient, is 2.3%. Fitness competitions were conducted in a
832 series of 8 blocks, shown in gray. The two strains were not competed directly
833 against each other; within each block, each was competed separately against
834 an otherwise isogenic *MMR+* *YFP+* strain. Each block contained between
835 5 and 8 replicate competitions.

836 **Figure 3**

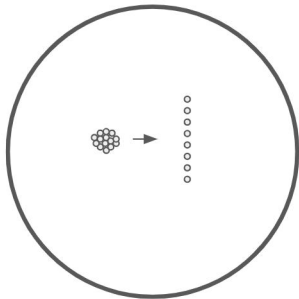
837 Distributions of fitness in haploid wild-types (**A**) and *mmr* mutators (**B**). We
838 measured the fitnesses of 327 wild-type and 313 *mmr* clones. Fitnesses were
839 measured in competition with an *MMR+* *YFP+* reference strain otherwise
840 isogenic to the wild-type, as described in Methods. Dashed vertical lines

841 indicate the mean. The load is ~ 0 for wild-types; for *mmr* mutators it is
842 $\sim 1.7\%$. QQ plots of the fitness distributions are shown as insets.

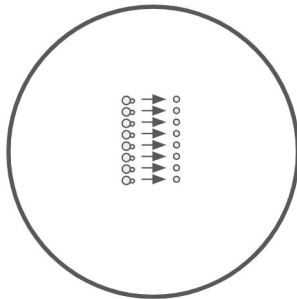
843 **Figure 4**

844 Change in *mmr* fitness disadvantage with ploidy. The fitness deficit, relative
845 to the wild type, decreases by 26% from haploid to diploid strains ($p <$
846 0.004).

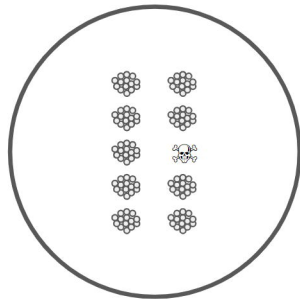
1. Array microcolonies



2. Separate new daughters



3. Score for lethal mutations



mmr fitness relative to wild-type

