

1 **A spatial analysis of childhood cancer and industrial air pollution in a**
2 **metropolitan area of Colombia**

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15 **Authors' contributions:** AMVG participated in the design of the research, performed the
16 non-spatial and spatial statistical analyses and drafted the manuscript. LARV participated in
17 the design of the research, review of the analyzes, coordination of the study. FEMC and
18 RGOM contributed with support the geographical data processing, visualization and
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21 have approved the final manuscript.

22 **Abstract**

23 Background: Air pollutants are considered carcinogenic to humans. In some European
24 countries, an association with childhood cancer in children has been established. This
25 phenomenon has not been addressed in Latin America, despite the spatial variability of air
26 pollutants that may limit the extrapolation of the results to other geographical areas.

27 Objective: To conduct a spatial analysis of the relationship between childhood cancer and
28 air pollution from industrial sources in a metropolitan area of Colombia.

29 Methods: Incident cases of childhood cancers were obtained from the Population-based
30 Cancer Registry of the Bucaramanga Metropolitan Area (2000-2015). Local and focused
31 cluster tests were used for the detection of spatial clusters and the Poisson multivariable
32 model was used to evaluate the combined effects of spatial variables.

33 Results: The Kulldorff's focused test found a significant spatial cluster ($p=0.001$) around one
34 industrial agglomerate and the multivariable model results suggests that the distance effect
35 is modified by the directional effect of the wind.

36 Conclusion: A spatial cluster of incident cases of childhood cancer occurred in the
37 municipality of Bucaramanga. Our finding supports the hypothesis that childhood cancer
38 might be related with industrial air pollution exposure in a Latin American city.

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40 **Keywords:** Childhood cancer, spatial analysis, air pollution

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44 Introduction

45 Childhood cancer (CC), defined here as cancer in children aged 0-14 years, is
46 considered a group of rare diseases. Nevertheless, about one new case of CC is diagnosed
47 in the world every two minutes [1] and it is the leading cause of non-accidental death for
48 this age group [2]. Worldwide, approximately 215,000 cancers are diagnosed annually in
49 children younger than 15 years [2], and in Latin America and the Caribbean 17,500 children
50 are newly diagnosed every year. It is estimated that 84% of all cases of CC in the world occur
51 in low-income and middle-income countries, where 90% of the child population live [3].

52 Knowledge regarding the etiology of CC is limited. Some authors argue that
53 environmental factors can be the cause of 85-96% of all cancers in childhood [4] [5] [6] [7].
54 In recent years, studies conducted in England, Spain and the United States found an
55 association between CC, especially leukemia and CNS tumors, and air pollution exposure [8]
56 [9] [10]. Air pollution is considered the world's largest single environmental health risk as
57 the World Health Organization (WHO, 2016) affirmed that 92% of the world population lives
58 in places where air quality levels are poor [11].

59 During 2013, the International Agency for Research on Cancer (IARC) found that
60 there is sufficient evidence about the carcinogenicity of outdoor air pollution in general and
61 fine particulate matter (PM_{2,5}) [12]. Research published in Europe and North America, have
62 found association between CC and proximity to industrial sources that emit pollutants into
63 the air [5][9][10][13][14]; also, since the beginning of the 20th century, geographic clusters
64 of cancer have been identified, mainly for leukemia [15]. The mixture of pollutants varies
65 widely in space and time, reflecting the heterogeneity of their sources, the influence of

66 climate, atmospheric transformations, and other factors, and therefore health effects might
67 also differ widely in space and time and cannot be generalized.

68 In Colombia, a developing country located at the north extreme of South America,
69 there are no exact estimates of the incidence of CC since 90% of the regions of the country
70 do not have cancer registries, which limits making truthful national statistics [16].
71 Population-based registries with validated high quality data by the IARC are located only in
72 the metropolitan areas of Cali, Manizales, Pasto and Bucaramanga [17]. In Colombia,
73 outdoor air pollution is considered one of the three most hazardous environmental
74 exposures [18] and annual mean concentrations of particulate matter in capital cities
75 exceeds the ones reported for most European and US cities [19]. Previous studies in
76 Colombia have identified adverse effects of ambient air pollutants on respiratory and
77 cardiovascular mortality and morbidity [20,21].

78 Industrial sources are important contributors of outdoor air pollution and location
79 of industrial facilities is often close to residential areas [22]. To our knowledge there are no
80 studies in Colombia or other Latin American country that evaluate the geographic location
81 of incident cases of CC with proximity to industrial sources of air pollution, despite the
82 contribution of industrial emissions to outdoor air pollution in developing countries.
83 Therefore, the aim of this study was to assess the spatial relationship between CC and air
84 pollution from industrial sources in the metropolitan area of Bucaramanga, Colombia,
85 during 2000-2015.

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88 **Methods**

89 **Study area and population**

90 Colombia is the fourth largest country in South America located at the top corner of
91 the continent. The study area was the Bucaramanga Metropolitan Area (BMA) that it is
92 located in the Northeast Andean region of Colombia with an extension of approximately
93 1,479 km² covering four municipalities: Bucaramanga, Floridablanca, Giron, and
94 Piedecuesta. The study population were the children under 15 years old living in BMA;
95 according to the National Administrative Department of Statistics (DANE for its acronym in
96 Spanish) for the year 2015, the total population in BMA was 1,122,961 people and the
97 population under 15 years was 258,097 children.

98 **Data Sources**

99 The study used secondary retrospective data available in different regional and
100 national institutions: The Population-based Cancer Registry of the Bucaramanga
101 Metropolitan Area (RPC-AMB for its acronym in Spanish) was the source for cancer cases
102 data, The DANE was the source for estimations of population and geographical data at small
103 geographical level (census sector), the Corporation for the Defense of the Bucaramanga
104 Plateau (CDMB for its acronym in Spanish) was the source for the identification and location
105 of industrial facilities emitting air pollutants to the air, and planning office of each
106 municipality was the source for the socioeconomic information at small area level.

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109 **Childhood cancer data**

110 The CC data were obtained from the RPC-AMB, which is one of the four Population-
111 based Cancer Registry of regions of Colombia considered by the IARC as high quality.
112 Incident confirmed cases of CC diagnosed from January 1, 2000 to December 31, 2015 were
113 included in the study. The information obtained for CC cases were: the type of cancer, sex,
114 date of birth, date of diagnosis, and place of residence at the time of diagnosis. In the RPC-
115 AMB and for the present research, it is considered a resident who is verified as having been
116 residing 6 months or more in the BMA before the first time they were diagnosed with CC.

117 **Population data and geographical data**

118 The population data for the children's population in the BMA were taken from the
119 official DANE population estimates based on the 2005 national census. The projected infant
120 population under 15 years was used as the population at risk to calculate CC rates during
121 the study period. The DANE population estimates information is available with free access
122 with a minimum level of geographic disaggregation up to block level that is added in an
123 ascending way at the level of sections, sectors, communes, urban/rural areas and
124 municipalities. For the analysis, the census sectors were used as geographical grouping unit
125 (average population of each census sector is 1857 children under 15 years old); these
126 geographical units have a unique identification number assigned by DANE and that was kept
127 as the identification number for each unit of analysis.

128 The geographic coordinates (latitude and longitude) of each centroid of the
129 geographic units (census sector) were estimated using the estimation method with
130 weighting by population location and therefore the centroids are not geographical centers

131 of polygons necessarily. For the analysis of proximity to sources of industrial pollution, the
132 distance and direction between each "putative source" and the centroids of the census
133 sectors were calculated using distance (in meters) and angle (in geodetic degrees)
134 calculation tool of ArcGIS[®] [23]. The map of the BMA with resolution at the block and
135 sector census level was obtained from the Geoportal cartographic data tool of DANE [24]
136 and the spatial data were created in ArcGIS[®] using the projection of Colombia in mode
137 Custom Azimuth Equidistant and Datum WGS 1984.

138 **Data sources of industrial sources and socioeconomic status**

139 The industrial facilities emitting air pollutants to the air were identified with the fixed
140 sources inventory existing in the local environmental authority, the CDMB. There were 32
141 industrial facilities that were located spatially and displayed four identifiable industrial
142 areas around and within BMA. The central point (centroid) of each identified industrial area
143 was used as the location of the "putative source of industrial pollution" and was calculated
144 as the geographic center of the area, georeferenced in latitude and longitude coordinates.

145 The socioeconomic strata (SES) is a DANE classification of the socioeconomic
146 resources of a residential place [25] and was used as proxy of socioeconomic status. The
147 predominant SES of each census sector was assigned according to the SES of the
148 neighborhoods that make up each census sector. This information was obtained from the
149 planning office of each municipality.

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152 **Data analysis**

153 **Exploratory analysis**

154 Crude rates (CR) of annual and cumulative CC incidence were calculated by census
155 sector. The DANE population estimates for each geographical unit were used as population
156 denominators of the annual and accumulated rates. Standardized rates were calculated by
157 age and sex and their 95% confidence intervals, using the direct method and as a standard
158 population, the population of Colombia was used by five-year groups according to the 2005
159 DANE Census. To reduce the heterogeneity in estimating the risk of CC, the calculation of a
160 standardized morbidity ratio (SMR) of the incident CC was made using a Bayesian smoothing
161 technique. Crude rates, standardized rates and Bayesian SMR were calculated in Stata 15[®]
162 and then visualized on choropleth maps using ArcGIS 12[®].

163 **Hypothesis tests**

164 A cancer spatial cluster is defined as a region with a statistically significant excess in
165 the number of cancer cases that occurs [26]. We used the Kulldorff circular space-time
166 scanning test to identify spatial and spatio-temporal local clusters of CC. Three different
167 hypothesis focused tests were used to identify potential spatial clusters of CC around the
168 four identified industrial pollution areas. Based on the spatio-temporal parameters,
169 retrospective space-time scanning analysis was applied to identify the geographic areas and
170 time periods of potential clusters. The Kulldorff 's circular spatial scanning test in its focused
171 version having as reference grid or "putative source" the coordinates of the centroid of each
172 industrial area identified in the BMA; this test was implemented using the SaTSCan 9.1[®]
173 software, with a Poisson probability distribution model, scanning for high rates and with a

174 maximum cluster size of 25% the population at risk [27]. The Stone's test was used to
175 evaluate wheather the risk of the disease is altered with the distance to the putative source;
176 this test was implemented using the "DCluster" programming package encoded in the R
177 software [28]. Finally, the Lawson's directional score test was used to evaluate the variation
178 in direction of the clusters using a score statistic for a single parameter that has a Chi2
179 distribution with one degree of freedom [29]; the score calculation was implemented in a
180 Microsoft Excel[®] spreadsheet.

181 **Multivariable models**

182 For the estimation of multivariable regression models with spatial data, we used the
183 approach proposed by Lawson [30]. In the basic or null Poisson model, the log of the
184 expected cases based on the number of populations by geographic area is used as an offset.
185 Subsequently, the spatial functions (distance, direction and their interactions) are included
186 in the multivariable model; to model direction, the functions of sine (longitude) and cosine
187 (latitude) were used. These multivariable generalized linear models (GLM) for CC were
188 constructed separately for each putative source of industrial pollution. Finally, the effects
189 of the spatial functions were adjusted by the SES of the census sectors. The Akaike
190 information criterion (AIC) and the deviance of the potential models were used as the
191 parameters for model selection. The residuals of the models were estimated, and the spatial
192 autocorrelation of the models was evaluated using the global Moran's index. These
193 analyzes will be performed using Stata 15[®].

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196 **Missclassification bias and sensitivity analyses**

197 A sub-sample of the CC were contacted by the RPC-AMB in order to assess the
198 quality of the residential address at the time of diagnosis. We inquired about the address
199 at the time of diagnosis and mobility in the two years prior to the time of diagnosis. With
200 this information, the proportion of misclassification in the subsample was estimated to
201 estimate the potential misclassification of the exposure in the study.

202 Sensitivity analysis was carried out for the local and focused cluster Kulldorf's test.
203 In both cases, the methodology was to execute the analysis with six circular search windows
204 with different parameters, with the following specifications: radius of 0.5 km, 1 km, and 3
205 km and 5%, 10% and 25% of population, Poisson distribution, non-overlap parameter and
206 level of significance of ≤ 0.05 . Conducting multiple analyses using different search windows
207 may strengthen findings and provide confidence about the detected CC clusters.

208

209 **Ethical considerations**

210 Ethical approval for this study was obtained from the ethics committee of the
211 Industrial University of Santander. The information received from the RPC-AMB did not
212 contain identifying information about participants. The data was analyzed aggregated and
213 not at the individual level, the georeferencing corresponds to the count of cases of CC for
214 each census sector and it is not the exact georeferencing of the place of residence of the
215 cases. For the analysis of misclassification where a subsample was recontacted, the
216 recontact was managed with the caregiver of the minor through the attending physician.
217 Informed consent was obtained from all participants of this subsample.

218 **Results**

219 **Descriptive analysis of childhood cancer**

220 From 2000 to 2015, 679 children under 15 years of age with a median age of 6 years
221 (Interquartile range [IQR] 3 - 11 years) were diagnosed with some subtype of CC in BMA and
222 43.89% were girls. The median annual CC was 44 cases (IQR 38,5 – 46 cases). The two
223 municipalities with the highest proportion of cases were Bucaramanga (54.05%) and
224 Floridablanca (25.77%). The most commonly diagnosed types of cancer were: 1) leukemia,
225 myeloproliferative diseases and myelodysplastic diseases (38.29%); 2) neoplasms of the
226 central nervous system (CNS) and intracranial and intraspinal neoplasms (16.05%); and 3)
227 lymphomas and reticuloendothelial neoplasms (13.4%). The overall CR of incident CC for
228 BMA over the study period was 158.12 cases per million in children aged under 15 years;
229 the overall annual CR increased from a minimum of 103.53 in 2000 to a maximum of 195.62
230 in 2013. The overall direct sex and age standardized rate (SASR) for BMA was 158.92 cases
231 per million (IC 95%: 147.18 – 171.34); according to the 95% confidence interval (95% CI),
232 the overall SASR of the municipality of Bucaramanga was significantly higher than that of
233 Giron and Piedecuesta. Table 1 shows the overall crude and age-standardized incidence
234 rates by sex for AMB and each municipality.

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239 **Table 1. Overall incidence rates for BMA and its municipalities by sex, 2000-2015**

| Region | Crude rate | | Age-standardized rate | | | |
|----------------------|---|-------------|---|-------------------|-------------|-------------------|
| | per million children under 15 years old | | per million children under 15 years old | | | |
| | <i>Female</i> | <i>Male</i> | <i>Female</i> | <i>95% CI</i> | <i>Male</i> | <i>95% CI</i> |
| AMB | 141.72 | 173.85 | 142.35 | (126.60 – 159.43) | 174.76 | (157.64 – 193.23) |
| Bucaramanga | 156.44 | 203.30 | 157.01 | (133.24 – 183.79) | 203.97 | (177.31 – 233.51) |
| Floridablanca | 140.89 | 190.69 | 142.32 | (111.33 – 179.24) | 193.29 | (157.41 – 234.87) |
| Giron | 110.23 | 103.31 | 110.33 | (77.25 – 152.76) | 103.29 | (71.94 – 143.66) |
| Piedecuesta | 120.27 | 110.30 | 120.67 | (83.06 – 169.46) | 109.80 | (75.09 – 155.04) |

240

241 BMA is divided into 151 census sectors, in 121 (80.13%) there was at least one case of CC
 242 during 2000-2015. The overall CR in the census sector varied from a maximum of 866.35 to
 243 a minimum of 36.72 cases per million, with a median of 177.04 (IQR: 103.1 - 259.7). The
 244 SASR were between 33.78 (95% CI: 4.09 - 125.89) and 795.67 cases per million (95% CI:
 245 20.14 - 4650.31), with a median of 179.18 (IQR: 103.2 - 263.6). A census sector located in
 246 the northwest of the municipality of Bucaramanga, had the overall SASR significantly higher
 247 than the overall SASR of BMA. The Bayesian SMR in the census sectors ranged from 0.63 to

248 1.49 (median: 1.02, IQR: 0.87 - 1.14). Figure 1 show the choropleth maps of the SASR of CC
249 for BMA over the study period.

250 **Figures 1. Age and sex standardized rates of childhood cancer in the urban area of the**
251 **municipalities of the BMA during 2000 to 2015.**

252 **Descriptive analysis of industrial sources**

253 The BMA inventory included a total of 38 fixed industrial sources, 78.95% were
254 constituted before the year 2000, 7.89% in 2004 and 13.16% between 2007 and 2009. Four
255 main areas of industrial facilities agglomerations were identified, to which the centroid and
256 the coordinates were calculated; these four industrial areas and their centroids were used
257 as a source of exposure for subsequent analyzes. The first agglomerate was referenced as
258 industrial area A with latitude 7.138530546 and longitude -73.13197221 and located in the
259 census sector 680011000000000306. The second referenced as industrial area B with
260 latitude of 7.095961661 and longitude -73.16810226 and located in the census sector
261 683071000000000001. The third area is the denominated C with latitude 7.081772454 and
262 longitude -73.15087541 and located in the census sector 680011000000001196. The fourth
263 industrial area, the D has latitude 7.06152597 and longitude -73.09061454, located in the
264 census sector 682761000000000009 (Fig 2).

265 **Hypothesis tests**

266 Of the total of cases, 598 (88.07%) had data for georeferencing. The Kulldorff circular
267 scan local test identified a spatio-temporal cluster of statistically significant of CC cases
268 ($p < 0.001$), with a radius of 1.64 km and estimated relative risk of 1.79 in the municipality of
269 Bucaramanga; furthermore, it identified a space-time cluster of CC cases covering 18 census

270 sectors in the municipality of Bucaramanga, the high-risk period was the year 2009 and
 271 estimated relative risk of 6.21 (p:0.02). Around the industrial agglomerate A, the focused
 272 cluster test of Kulldorff found a significant cluster (p:0.001) with a radius of 2.65 km and
 273 estimated relative risk of 1.41 in the municipality of Bucaramanga, being coincident with
 274 the cluster identified by the local test. No significant clusters were found around the other
 275 three points of industrial area with this test. Stone's test did not detect significant clusters
 276 for the points of industrial agglomerates. Lawson's directional score was significant from
 277 industrial areas A, B and C. Table 2 summarizes the results of the three focused tests located
 278 for the four industrial areas.

279

280 **Table 2. Results of focused hypothesis tests for clusters of childhood cancer in proximity**
 281 **to industrial areas in BMA, 2000-2015**

| Industrial areas | Kulldorf's circular spatial | | Stone's test for | | Lawson's directional | |
|------------------|-----------------------------|---------|------------------|---------|----------------------|---------|
| | scan test | | distance decline | | score test | |
| | Statistic | p-value | Statistic | p-value | Statistic | p-value |
| A | 12.27 | <0.001 | 2.44 | 0.09 | 6.82 | 0.01 |
| B | No clusters found | | 1.02 | 0.89 | 13.17 | <0.001 |
| C | 0.002 | 0.86 | 1.45 | 0.25 | 6.93 | 0.01 |
| D | 2.11 | 0.08 | 1.19 | 0.29 | 0.94 | 0.33 |

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283

284 **Multivariable Analysis**

285 The Moran coefficient was 0.059 ($p < 0.001$) for the total number of CC events, which
286 means that there was a low spatial autocorrelation among the census sectors. The
287 distribution of the SMR as outcome variable of the multivariable model (mean: 1.09,
288 variance: 1.00) showed to be appropriated for the proposed Poisson model.

289 The results of the multivariable Poisson models support a consistent significant
290 inverse effect of the distance to the industrial area A on the SMR (Table 3). The model that
291 evaluated only the directional effect had a significant result for the positive effect of the
292 longitude, suggesting a dominant effect toward the East from the industrial area A. The
293 model that evaluates the combined effect of distance and direction suggested a significant
294 effect by decreasing the distance and directional towards the East from A. When the
295 interaction terms between the distance and the address parameters were added to the
296 model and adjusted by the SES, the model improved the deviance suggesting the best model
297 quality. In this model the inverse effect of the distance to industrial area A remained
298 significant, and the statistical significance of the interaction terms suggest that the distance
299 effect is modified by the directional effect (that is, the smaller the distance the stronger is
300 the directional effect), the coefficients and signs of the interaction terms suggest that the
301 predominant directionality is towards the SouthEast direction from A (Table 3).

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305 **Table 3. Poisson models assessing distance and directional effects of childhood cancer**
 306 **rates around the industrial area A, 2000-2015**

| Model type | Variable | Industrial area A | | | | | |
|---------------------------|--------------------|-------------------|--------|---------|----------|------|---------|
| | | Coefficient | SE | p-value | Deviance | AIC | BIC |
| Distance decline | Distance | -0.00003 | 0.0000 | <0.001 | 224.16 | 4.11 | -517.41 |
| Directional effect | Latitude | -0.137 | 0.070 | 0.853 | 228.51 | 4.15 | -508.05 |
| | Longitude | 0.306 | 0.941 | 0.001 | | | |
| Distance Decline | Distance | -0.00003 | 0.0000 | <0.001 | 209.69 | 4.04 | -521.86 |
| plus Directional | Latitude | -0.124 | 0.077 | 0.109 | | | |
| | Longitude | 0.324 | 0.933 | 0.001 | | | |
| Effects | | | | | | | |
| Distance Decline | Distance | -0.0001 | 0.0000 | 0.01 | 198.01 | 3.99 | -523.51 |
| plus Directional | Latitude | -0.092 | 0.109 | 0.4 | | | |
| | Longitude | -0.109 | 0.157 | 0.488 | | | |
| Effects with | | | | | | | |
| | | | | | | | |
| interaction terms | Distance*Latitude | -0.00003 | 0.0000 | 0.422 | | | |
| | Distance*Longitude | 0.00008 | 0.0000 | 0.001 | | | |
| Distance Decline | Distance | -0.0001 | 0.0000 | 0.004 | 169.38 | 4.03 | -464.35 |
| plus Directional | Latitude | 0.104 | 0.132 | 0.430 | | | |
| | Longitude | -0.080 | 0.169 | 0.635 | | | |
| Effects with | | | | | | | |
| | | | | | | | |
| interaction terms | Distance*Latitude | -0.0001 | 0.0000 | 0.052 | | | |
| | Distance*Longitude | 0.00007 | 0.0000 | 0.01 | | | |
| adjusted by SES | | | | | | | |
| | SES | 0.048 | 0.042 | 0.250 | | | |

307 Sine=Longitude

308 Cosine=Latitude

309

310 In the model diagnosis, standardized Pearson, Deviance and Ascombe residuals fulfill the
311 assumption of normal and mean distribution close to zero. Finally, the Moran coefficient
312 for the residuals was 0.74 ($p < 0.001$), which means that there is a low spatial autocorrelation
313 of the residuals of the model.

314 For the industrial areas B, C and D the multivariable models adjusted by ESE that
315 incorporated the terms of interaction between distance and directionality were also those
316 that had the best criteria to be the best model and in no case the models had significant
317 results for the effect of distance or directional effect.

318

319 **Sensitivity analysis**

320 **Misclassification analysis**

321 A total of 40 CC cases were contacted. The median age was 6 years (IQR 2 - 9) and
322 86.84% were residents in neighborhoods of SES 1 to 3. Of the total, four (10%) had changed
323 their homes in the year prior to diagnosis. Of the four cases with mobility in the last year
324 prior to diagnosis, two had changed their housing within the same municipality, one
325 between the municipalities of the BMA and one to a municipality outside the BMA. A total
326 of 22 (55%) children resided in the same home since birth.

327 **Sensitivity analysis for Kulldorff tests**

328 In the local Kulldorff's test, a cluster of CC cases with statistical significance was identified in
329 the northwest zone of municipality of Bucaramanga, which was repeated in 5 of the 6
330 windows executed in the 1km circular window; 12 census sectors were identified that

331 extend to 26 in the 3km window and that were equivalent to the analysis with a population
332 threshold at risk of 5% and 25%, respectively. Figure 2 shows in red census sectors that were
333 detected repeatedly within 5 clusters, in yellow those detected within 3 clusters and in
334 green those detected within 2 clusters.

335

336 **Figure 2. Census sectors detected within the results of the six analyzes of focused**
337 **childhood cancer clusters through different search windows, Bucaramanga, Colombia,**
338 **2000-2015**

339

340 In the focused Kulldorf's tests, taking as reference the industrial area A, in the six
341 search windows a spatial cluster was identified, with two census sectors that were always
342 present in the six analyzes executed, 68001100000000306 (Sector that corresponds to the
343 area industrial A) and 68001100000000305. In the 0.5km circular window, two census
344 sectors were identified that extend to 41 in the 3km window and that were equivalent to
345 the analysis with a population threshold at risk of 5% and 25%, respectively.

346

347 **Discussion**

348 The incidence rates of CC in the BMA between 2000 and 2015 are higher than the
349 estimated world rates for the same period of time. The cumulative SASR 2000-2015 in the
350 BMA was 1.5 times higher than the world rate reported by the World Cancer Report in 2014
351 that was 100 CC cases per million [31]. However, the cumulative SASR in the BMA was
352 similar to the cumulative incidence rate of CC reported by the Population-based Cancer

353 Registry of Cali during 2002-2011 [32]. The cancer registry of Cali is the oldest in Latin
354 America and has international recognition for its high quality methods and data [33][34],
355 therefore, the similarities in the incidence of CC and the frequency of presentation of the
356 different types represent not only the quality and maturity of the RPC-AMB but also the
357 consistency of having higher CC rates in different cities in Colombia.

358 The findings of this study suggest the presence of a CC cluster in proximity to the
359 industrial area denominated A. These results were consistent across the descriptive,
360 hypothesis testing and multivariable analysis findings, and remained statistically significant
361 in the same area when different scan windows were used in the sensitivity analysis. The
362 Lawson's directional score and in the Poisson model that evaluated distance and
363 directionality indicate that there is a spatial presentation of the cases that is mostly
364 concentrated at southeast of industrial area A, finding that might be explained by the
365 potential dispersion of air pollutants as the direction of the cluster is pointing towards the
366 predominant direction of wind in the area [29].

367 The Stone's test was the test that had no statistically significant findings for the
368 industrial area A; this negative result might be explained by the fact that the Stone's test is
369 executed under the assumption that there is a decreasing trend in risk symmetrically around
370 the radius as the distance to the source increases, and as shown in the multivariable analysis
371 and sensitivity analysis (Fig 2) the risk is not symmetric but with direction towards southeast
372 from A [35]. The statistically significant results in the industrial areas B and C for the
373 directional score of the Lawson test seem to be isolated findings as they were not consistent

374 with the multivariable analysis, where the interaction between distance and direction was
375 evaluated as suggested by Lawson [36].

376 The spatial association found in this research should not be taken as a causal
377 relationship between the industrial emissions of the industrial area A and the incident cases
378 of the CC cluster in this area of the municipality of Bucaramanga. In this case, the
379 relationship between exposure and outcome was modeled with respect to distance to the
380 source and directionality of the wind, but the exposure to the pollutants emitted were
381 measured. Likewise, the retrospective evaluation of the exhibition is difficult since it usually
382 involves different changes during the study period [37].

383 When a cluster of a disease is detected the Centers for Disease Control and
384 Prevention (CDC) proposes a four-stage approach, where the final stage includes proposing
385 a formal etiological investigation [38] and it is the proposal that is suggested as the next
386 step to this study. The increase of cases of IQ in this area can be due to multiple factors such
387 as the proximity to exit routes of the city with high flow of heavy duty vehicles of diesel
388 engines, which the exhaust of this type of engine is classified within the group 1 of
389 carcinogens by IARC, which has been found associated with CC [39]. As also for the
390 permanent transit of this type of automotive in this area, due to the loading and unloading
391 of merchandise in the industrial sources that are in this industrial agglomerate. Likewise, in
392 this area of industrial agglomerate the exhibition may not be related to air emissions but to
393 the occupational exposure of parents who live near the workplace; there are several
394 authors who have found a relationship between childhood cancer and maternal or paternal
395 occupation [40] [41] [42] [43] [44].

396 **Consistency with previous studies**

397 Different studies in Europe have found an association between exposure to air
398 pollution and CC. Steffen et al. carried out a multicenter case control study in France and
399 found an association between dwellings neighbouring a petrol station or a repair garage
400 during childhood and the risk of childhood leukemia (OR 4.0, 95% CI 1.5 to 10.3) [45]. During
401 2009, a study with the same aim was also conducted in France and the findings supported
402 the results and the associations established by Steffen [13].

403 In Spain several studies have been conducted assessing the association between
404 spatial location of industrial pollution sources and CC. During 2015, Ramis et al. published a
405 case-control study with spatial analysis, which included cases of leukemia, cancer of the
406 CNS and non-Hodgkin's lymphoma from five regions of Spain diagnosed from 1996 to 2011,
407 and reported absence of any statistically significant cluster [46]. Ortega-García and cols.,
408 analyzed the spatial distribution of all incident cases of CC diagnosed in Spain during 1998
409 to 2015, using focused cluster tests and found a possible association between proximity to
410 certain industries and CC [10]. Previous population case-control studies conducted in Spain
411 by same research team reported an increased risk of Leukemia in children (OR: 1.31, 95%
412 CI: 1.03-1.67) [5], and renal tumors (OR: 1.97, 95% CI: 1.13-3, 42) [47] associated with living
413 at a distance less than 2.5 km from industrial facilities; increased risk of neuroblastomas
414 associated with industrial facilities at 1 km (OR: 2.52; 95% CI: 1.20-5.30) and 2 km (OR: 1.99;
415 95% CI: 1.17-3.37) [48]; and increased risk of bone tumors in children under 15 years of age
416 associated with living less than 3 km from industrial sources (OR: 2.33; 95% CI: 1.17-4.63)

417 [49]. They found no excess risk for retinoblastoma, liver tumors, soft tissue sarcomas and
418 germ cell tumors and living close to industrial areas [50].

419 **Strengths and limitations**

420 The main strength of this study is the quality of the CC data and the period of 16
421 years included for the analysis. The cancer registries are considered the best source of
422 information to measure cancer burden in a given time and region [15] [51], the RPC-AMB is
423 considered high quality by IARC due to the integrity, punctuality and accuracy of the data
424 provided, which translates into highly standardized and reliable data. Therefore, the
425 classification of the disease is considered adequate for the cases included in the analyzes
426 performed. In addition, it can be considered that the results of this research are not affected
427 by the mobility of the population, since the misclassification analysis conducted did not
428 found that the CC had a high proportion of migration and 55% of the cases in BAM lived in
429 the same residence where they received the diagnosis since birth.

430 Another important strength to note is the use of multiple analytical approaches and
431 tests used to evaluate the existence and characteristics of a potential CC cluster. The
432 combination of descriptive, hypothesis testing, and multivariable analysis improved the
433 quality of the analysis and gave consistency to the results.

434 The limitation of the pre-established shapes that the Kulldorff's tests used in this
435 study have been discussed in the literature [52][53][54][55]. These circular clusters in many
436 cases generate that areas of the detected spatial cluster correspond to a false positive, since
437 those areas with low risks are not a genuine part of the cluster [53][55]. To overcome this

438 limitation this study includes other hypothesis test and multivariable analysis that include
439 direction testing and their interaction with distance and results of the existence of a cluster
440 identified by scanning tests remained consistent with additional information related to the
441 direction of the cluster.

442 The correct classification of the exposure is a limitation and it is considered that to
443 a certain extent all the studies evaluating the effects of exposure to air pollution inevitably
444 have a degree of misclassification of exposure, since it is not feasible to conduct population-
445 based studies in which each person uses a personal measuring device for long time [56]. In
446 this study, the magnitude of the exposure was probably determined not only by the
447 distance to the industrial areas identified in the BMA, but also for quantity or air pollutant
448 emissions and the predominant wind direction in the industrial areas. Although the
449 predominant wind direction was taken into account for directional tests, the variability in
450 time of the wind direction was not included. Other important limitation of this study is that
451 exposure was based on proximity to industrial facilities only and did not include any
452 characterization or measurement of the pollutants emitted to the air by the industrial
453 sources. Therefore, proximity to industrial facilities was used in this study as proxy of
454 industrial air pollution exposure.

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458 **Conclusions**

459 The results of this study suggest the presence of a spatial cluster of incident cases of CC
460 during 2000-2015 in proximity to an industrial area in the municipality of Bucaramanga.
461 Given the ecological design of this research this finding cannot be assumed as a causal
462 relationship, but results of this study generate new research hypotheses to be evaluated in
463 further research in order to contribute to have a better understanding of the environmental
464 exposures related to the high incidence of CC in Bucaramanga and Colombia.

465

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467

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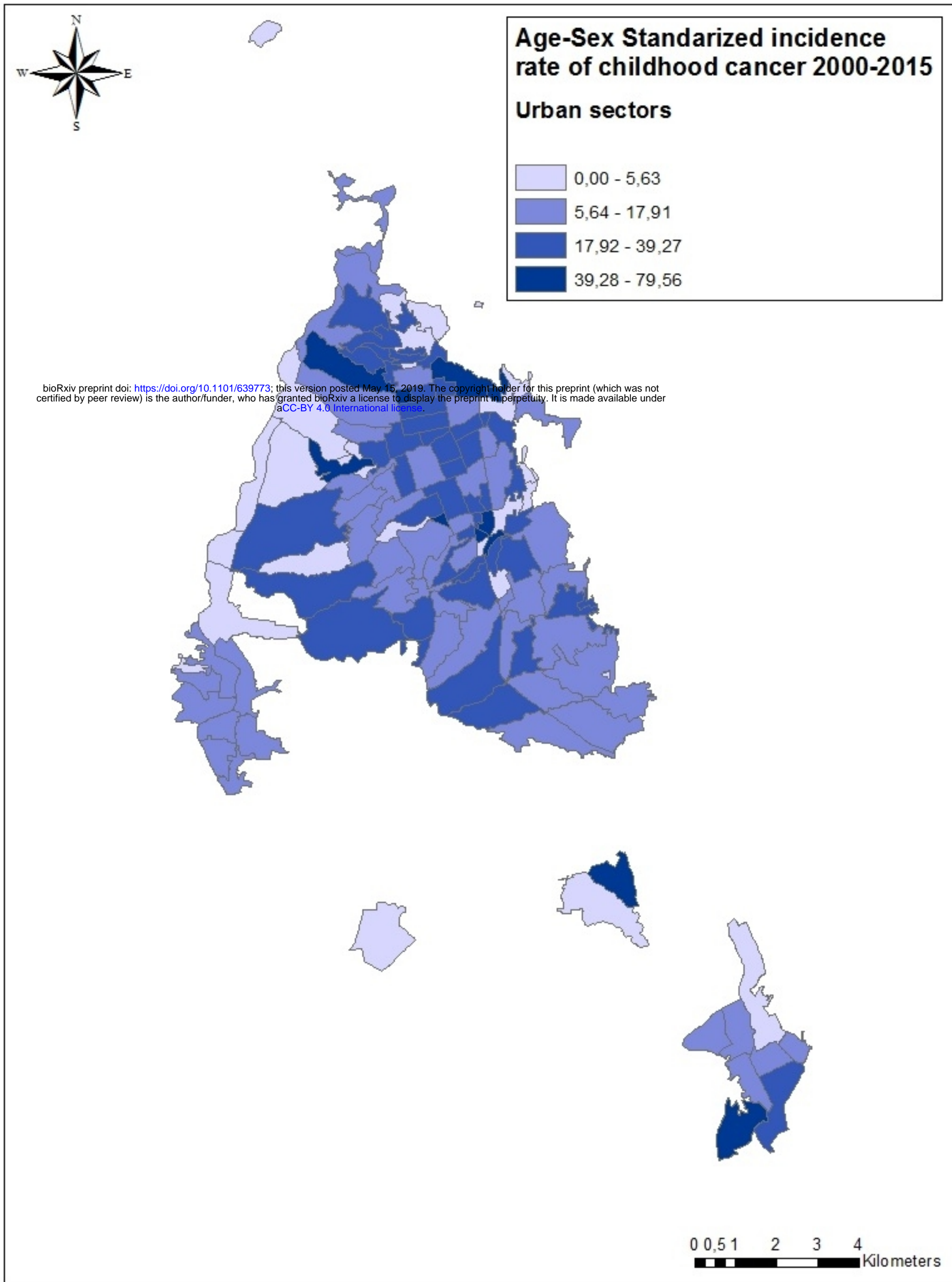
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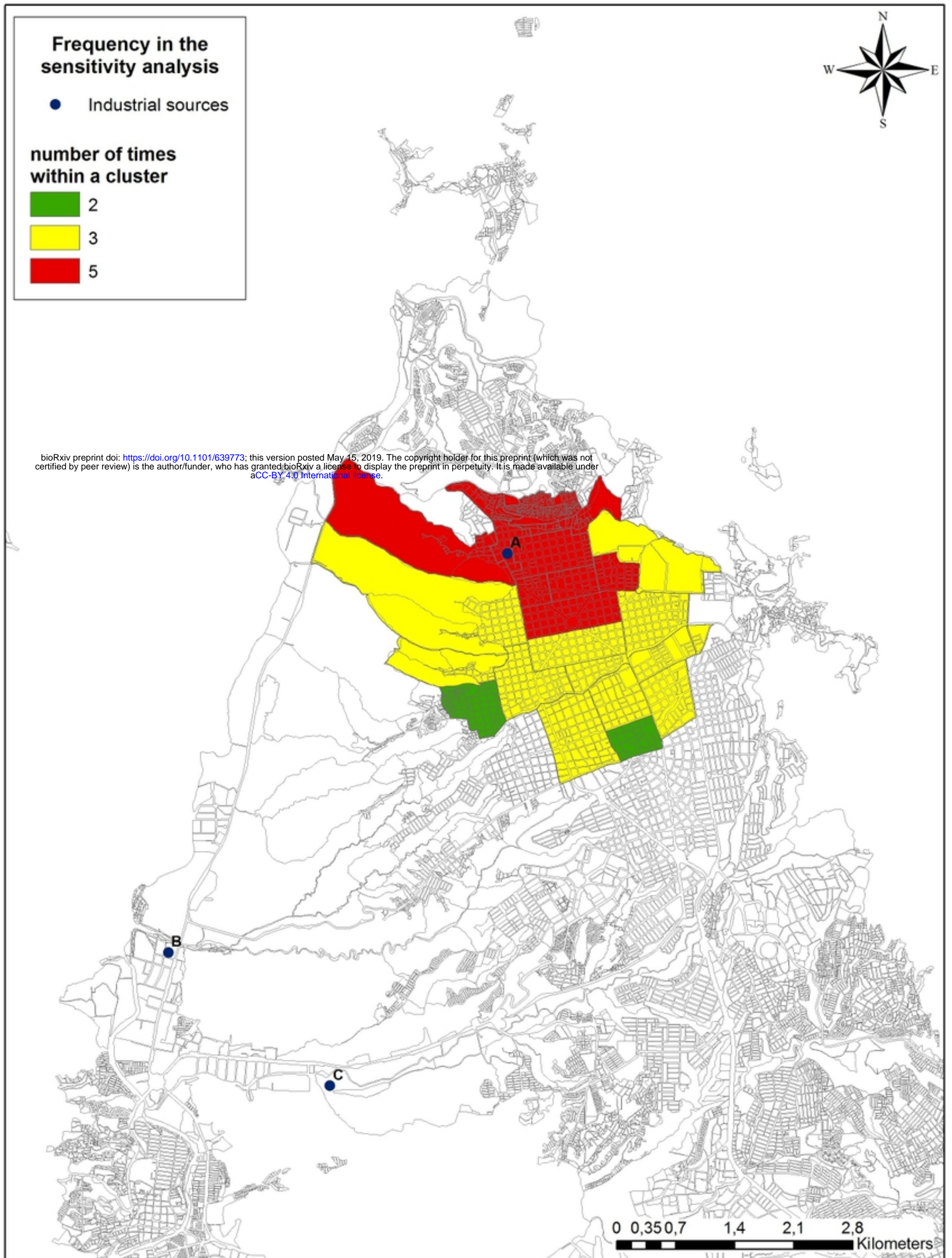
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Figure



Figure