Environmental correlates of genetic variation in the invasive and largely panmictic

European starling in North America

Running title: Starling genotype-environment associations

Natalie R. Hofmeister[1,2*], Scott J. Werner[3], and Irby J. Lovette[1,2]

[1] Department of Ecology and Evolutionary Biology, Cornell University, 215 Tower Road,

Ithaca, NY 14853

[2] Fuller Evolutionary Biology Program, Cornell Lab of Ornithology, Cornell University, 159

Sapsucker Woods Road, Ithaca, NY 14850

[3] United States Department of Agriculture, Animal and Plant Health Inspection Service,

Wildlife Services, National Wildlife Research Center, 4101 LaPorte Avenue, Fort Collins, CO

80521

*corresponding author: nrh44@cornell.edu

ABSTRACT

Populations of invasive species that colonize and spread in novel environments may

differentiate both through demographic processes and local selection. European starlings

(*Sturnus vulgaris*) were introduced to New York in 1890 and subsequently spread

throughout North America, becoming one of the most widespread and numerous bird

species on the continent. Genome-wide comparisons across starling individuals and

populations can identify demographic and/or selective factors that facilitated this rapid

and successful expansion. We investigated patterns of genomic diversity and

differentiation using reduced-representation genome sequencing (ddRADseq) of 17

winter-season starling populations. Consistent with this species' high dispersal rate and

rapid expansion history, we found low geographic differentiation and few F_{ST} outliers even

at a continental scale. Despite starting from a founding population of approximately 180

individuals, North American starlings show only a moderate genetic bottleneck, and

models suggest a dramatic increase in effective population size since introduction. In

genotype-environment associations we found that ~200 single-nucleotide polymorphisms

are correlated with temperature and/or precipitation against a background of negligible

genome- and range-wide divergence. Local adaptation in North American starlings may

have evolved rapidly even in this wide-ranging and evolutionarily young population. This

survey of genomic signatures of expansion in North American starlings is the most

comprehensive to date and complements ongoing studies of world-wide local adaptation in

these highly dispersive and invasive birds.

Keywords: panmixia, gene flow, local adaptation, invasion genetics

INTRODUCTION

Studies of local adaptation have long bridged the interface between ecological and

evolutionary questions by exploring how populations adapt to differing environmental

conditions. Traditionally, high degrees of local adaptation were expected to be present only

in fairly isolated populations—those free from the homogenizing effects of high gene

flow—with a long history in those locations, providing the time thought to be necessary for

local adaptation to evolve (Lenormand, 2002). We now know that local adaptation occurs

frequently even in systems with high gene flow (Yeaman & Whitlock 2011; Tigano &

Friesen, 2016) and often rapidly after colonization of a novel environment (Prentis *et al.*,

2008). We continue to find evidence for rapid local adaptation in systems as divergent as

cane toads (Rollins *et al.*, 2015), sticklebacks (Lescak *et al.*, 2015), honeybees (Avalos *et al.*,

2017), steelhead trout (Willoughby *et al.*, 2018), deer mice (Pfeifer *et al.*, 2018), and many

more. These studies show that many taxa can adapt rapidly to local conditions in response

to the new selection regimes they encounter as they expand their range.

Invasive species that have recently expanded into new locations provide tractable

opportunities to investigate local adaptation as it originates (Colautti & Lau, 2015).

Invasions typically expand from the founding population(s) following a predictable spatial

and temporal pattern (reviewed in Excoffier *et al.*, 2009). After successful colonization of a

new habitat, many invasive species show a demographic boom that may be facilitated by

their ecological release in the new environment; ecological release is the concept that

introduced species are often released from top-down limitations to their population

growth, such as predators or pathogens in their native range (Riccardi *et al.*, 2013).

Theory predicts that this rapid population growth will plateau as the population

approaches carrying capacity in a region, but successful invasive species may continue to

expand their range and thus maintain a high rate of overall population growth.

When population density increases and demographic rates change, introduced

species may rapidly evolve traits that enable them to spread (Szűcs *et al.*, 2017): for

example, invasive cane toads in Australia evolved a suite of morphological, physiological,

and behavioral traits that facilitated their expansion in only 80 years (Rollins *et al.*, 2015).

Increased dispersal may evolve in concert with the demographic boom, such that certain

traits are selected for in successful invasive species. Flexible dispersal strategies can result

in gene flow that counteracts inbreeding depression and increases adaptive potential

(Garant *et al.,* 2007; Rius & Darling, 2014). If particular traits enable individuals to disperse

more easily to their preferred habitat, gene flow may be directional and even adaptive

(Edelaar & Bolnick, 2012; Jacob *et al.,* 2017). Invasions thus allow us to observe

interactions between demography and the early processes of selection (Dlugosch *et al.,*

2015) as populations experience new environments. Importantly, these eco-evolutionary

interactions depend in part on the genetic variation in the established population.

79   Recent work in invasion genetics aims to tease apart how selection and demography

might resolve a paradox of invasion (Estoup *et al.,* 2016, Schrieber & Lachmuth, 2017):

many invasive species experience genetic bottlenecks as a result of an initial founder effect,

but often thrive and spread despite this loss of standing genetic diversity. Theory predicts

that introductions will typically result in an initial contraction in population size and/or

genetic diversity (Dlugosch & Parker, 2008). However, bottlenecks of genetic variation

clearly do not limit the success of many invasive species (Schmid-Hempel *et al.,* 2007;

Dlugosch & Parker, 2008; Facon *et al.,* 2011). Invaders might adapt through soft sweeps

that reduce genetic diversity while selecting for adaptive variants from standing or novel

genetic variation, which is especially likely in the case of ecological adaptation (Messer &

Petrov, 2013). Furthermore, some invasions may increase rather than reduce genetic

diversity, as when multiple invasions from different source populations introduce

previously isolated alleles and thereby facilitate admixture (Dlugosch & Parker, 2008). The

new conditions can also select among standing variation, where the presence of certain

genetic variants in the native range accelerates adaptation upon introduction (Tsutsui *et*

al., 2000; Schlaepfer *et al.,* 2009; Hufbauer *et al.,* 2011). In sum, although genetic diversity

in introduced populations is often viewed as a pre-requisite to adaptation, changes in

genetic diversity alone do not explain invasiveness (Uller & Leimu, 2011). Regardless,

chance may be just as important as selection in an invasive species' establishment and

spread (Gralka & Hallatschek, 2019).

99   The European starling (*Sturnus vulgaris*) stands out as an exceptionally successful

avian colonist and invasive species. In North America, an estimated 200 million starlings

currently range from northern Mexico to southern Alaska (Linz *et al.,* 2007). Introduced to

New York City in 1890, starlings nearly covered the continent within a few generations by

expanding up to 91 km each year (Bitton & Graham, 2014). The 1890 introduction is widely

accepted as the first successful establishment of starlings in North America, but several

populations were introduced in Cincinnati, OH (1872), Quebec, Canada (1875), Allegheny,

PA (1897), and Springfield, MA (1897), with the second-most successful having been

introduced to Portland, OR in 1889 (Forbush, 1915; Kalmbach & Gabrielson, 1921).

Records indicate that none of the earlier starling introductions survived more than a few

years after colonization, but it is possible that some populations in the western U.S.

persisted without record (Kessel, 1953).

During the starling expansion, ongoing migration and dispersal might have also

influenced patterns of genetic variation. In North America, some—but not all—starling

populations migrate (Dolbeer, 1982). Previous studies indicate that there is considerable

variation in migratory distances within flyways (Burtt & Giltz, 1977). Models of molt origin

indicate that starlings in the western U.S. may disperse or migrate shorter distances

(Werner, Fischer, & Hobson, 2020). The models used feathers from the same individuals

sampled for this genetic survey, and they initially hypothesized that starlings in the south

may be less likely to migrate long distances. Instead, they found that longitude better

explains the differences in molt origin among starling populations: starling feathers

collected west of -90° longitude were more likely to have originated nearby. In other

words, starlings that migrate in eastern North America likely experience greater gene flow

among sampling locations (states, in our study), but overall starlings tend to move only

regionally and not continent-wide.

Early genetic work based on a small set of allozyme markers indicated near-random

mating at a continental scale in North American starlings, with large demes

(subpopulations) and high dispersal rates (Cabe, 1998; Cabe, 1999). Here we use robust

genomic markers to explore the genomic and demographic patterns of range expansion in

North American starlings with three specific aims: (1) to characterize genome-wide levels

of diversity and differentiation among starlings; (2) to examine how genetic variation

changes across the contiguous United States; (3) to test for a genetic bottleneck; and (4) to

test for signatures of selection associated with environmental gradients. We also interpret

our results in the context of range-wide data on starling migration and dispersal (Werner,

Fischer, & Hobson, 2020), as this movement certainly influences population structure.

Recent developments in genotype-environment association methods can identify polygenic

traits which are locally adapted to overlapping environmental gradients (Forester *et al.,*

2018; Capblancq *et al.,* 2018), whereas traditional outlier-based methods may not recover

evidence of local adaptation in a species that likely has low overall levels of genetic

diversity. Especially with these more sensitive methods, the same low genetic diversity

resulting from a founder effect upon colonization could improve our ability to discriminate

signatures of selection from low background divergence across the genome (Dlugosch *et*

al., 2015). This work employs modern genomic and analytical tools to examine the

evolutionary history of this remarkably successful avian invasion.

METHODS

Sample collection and processing

All starlings sampled were collected during the non-breeding season, when large flocks

aggregate at high-quality food sources. These samples therefore do not represent discrete

breeding populations but rather a mixture of birds from the surrounding region, potentially

including migrants from more northerly areas. Starlings (N = 166) were collected in

January-March 2016 and 2017 from 26 dairies and feedlots by the U.S. Department of

Agriculture's Wildlife Services personnel in Arizona, California, Colorado, Idaho, Illinois,

Iowa, Kansas, Missouri, Nebraska, Nevada, New Hampshire, New Mexico, New York, North

Carolina, Texas, Washington, and Wisconsin. USDA Wildlife Services personnel collected

birds from 2-5 sites in each state, where each collection site was >5 km apart. USDA

personnel then euthanized whole adult males by lethal gunshot, avicide, or live traps,

stored at these birds at 0°C until tissue sampling, and recorded the latitude and longitude

of each collection site using a handheld GPS. The collection and use of starlings for this and

related studies were approved by the U.S. Department of Agriculture, National Wildlife

Research Center's Institutional Animal Care and Use Committee (QA-2572, S.J. Werner -

Study Director; Werner, Fischer, & Hobson, 2020; Table S1).

Breast muscle tissue was sampled using biopsy punches (Integra Miltex, York, PA)

and frozen in 95% ethanol. Samples were shipped on dry ice, and DNA was extracted using

a Qiagen DNeasy kit following the manufacturer's protocol (Qiagen, New York, NY). DNA

concentration of each sample was quantified using a Qubit 2.0 fluorometer (Thermo Fisher

Scientific, New York, NY). Following the protocol of Peterson *et al.* (2012), we generated a

reduced-representation genomic dataset of double-digested, restriction-site associated

DNA markers as described in Thrasher *et al.* (2017) using the restriction enzymes SbfI and

MspI and adaptors P1 and P2. We sequenced 100bp, single-end reads on an Illumina HiSeq

2500. We trimmed and filtered for quality using the FASTX-Toolkit

(http://hannonlab.cshl.edu/fastx_toolkit). We then used the process_radtags commands in

STACKS v 1.19 (Catchen *et al.,* 2013) to demultiplex the remaining sequences. In subsequent

filtering steps, we retained reads only if the following conditions were met: reads passed

the Illumina chastity filter, contained an intact *Sbf*I RAD site, contained one of the unique

barcodes, and did not contain Illumina indexing adaptors.

We assembled sequences to an *S. vulgaris* reference genome (Hofmeister, Rollins *et*

al., in prep) using the ref-map option in STACKS. Individual reads were mapped to the

reference genome using BOWTIE2 version 2.2.8 (Langmead & Salzberg, 2012) using the

"very sensitive local" set of alignment pre-sets. The bioinformatics pipeline used for the

reference-based assembly has the advantage of using fewer similarity thresholds to build

loci. We required that a SNP be present in a minimum of 80% of the individuals (-r 0.8)

with a minimum stack depth of 10 reads at a locus within an individual (-m 10) for it to be

called. We removed two individuals, one with >50% missing data and one with >50%

relatedness (measured using the unadjusted AJK statistic and calculated within *vcftools*),

leaving 158 individuals remaining in the study. We also used *vcftools –hwe* option to

remove any SNPs out of Hardy-Weinberg equilibrium (e.g., an exact test that compared

expected and observed heterozygosity gave a P-value less than 0.001). About 6% of

sequenced variants were out of HWE, and this filter left 14,134 variants remaining. For all

genotype-environment analyses we used all SNPs in a given stack, but for STRUCTURE and

other analyses sensitive to linkage disequilibrium we used only the first SNP in each stack;

for unphased genomic data like the RAD loci analyzed here, this strategy of exporting one

SNP per stack is often used as an indirect method of controlling for linkage.

(1) Patterns of genetic diversity and differentiation

194   We estimated per-locus measures of genetic diversity and genome-wide

differentiation using the *populations* option in Stacks (Catchen *et al.*, 2013). We used

vcftools to calculate F_{ST} among population pairs and heterozygosity and nucleotide

diversity (π) within populations (Danecek *et al.,* 2011). We determined expected

heterozygosity at the population level using the Hs() function in the R package *adegenet*

(Jombart ,2008), and tested for differences in observed and expected heterozygosity of

each locus using a paired t-test in base R (R Core Team, 2019). We investigated genetic

structure within North American starlings using an analysis of molecular variance

(AMOVA) in the R package *poppr* (Kamvar *et al.,* 2014). We tested whether most genetic

variation was observed among individuals or among sampling sites ("populations"). To

determine significance, we compared observed variation at each hierarchical level to the

randomly permuted distance matrices for that particular level using 1000 simulations in

the function *randtest()* in the R package *adegenet* (Jombart, 2008), hypothesizing that the

observed variance is greater than expected within individuals and less between individuals

and between populations.

We tested for isolation by distance (IBD) using a simple Mantel test in *adegenet*

(Jombart, 2008): for these data, the assumption of stationarity likely holds, given that

North American starlings appear to be in mutation-migration-drift equilibrium (Guillot &

Rousset, 2013). Geographic distances among sampling locations follow a bimodal

distribution where locations are either very near or far from each other, and thus we

caution that the Mantel test may not be an appropriate test of isolation by distance

(Supplementary Information). Nevertheless, we report the results of a Mantel test with a

Spearman correlation and 10000 permutations. Finally, we test for isolation by

environment (IBE) using the R package *vegan* to compare environmental, geographic, and

genetic distances (Oksanen *et al.,* 2019). Controlling for geographic distance, we ran partial

Mantel correlograms using Euclidean environmental distance matrices built from all

bioclimatic variables retained in the *Selection analyses* described in Section 4 below.

(2) Population structure

We first used non-parametric approaches to determine whether starlings clustered by

sampling location, using principal components analysis in *SNPRelate* (Zheng *et al.*, 2012)

and discriminant analysis of principal components (DAPC) in *adegenet* (Jombart, 2008). We

then tested for population structure using STRUCTURE (Pritchard *et al.,* 2000) by

simulating 10 runs of each K. Although we sampled starlings from 17 locations, we

hypothesized that North American starlings would cluster into at most eight populations

(K=1-8) based on the non-parametric tests described above. To select the best-supported

K, we used the Evanno method implemented in STRUCTURE HARVESTER v0.6.94 (Earl &

vonHoldt, 2011). We averaged results across the 10 runs using the greedy algorithm in the

program CLUMPP v1.1.2 (Jakobsson & Rosenberg, 2007), and visualized results using

DISTRUCT v1.1 (Rosenberg, 2003). Given that evidence of population structure depends on

the filtering thresholds selected, we ran this model-based approach using a very strict

minimum minor allele frequencies (MAF=0.3) and a more relaxed minimum frequency

(MAF=0.1) (Linck & Battey, 2017). STRUCTURE results did not differ substantially among

MAF thresholds, and we used a filtered dataset with a minimum minor allele frequency of

0.1 in the demographic analyses below.

(3) Demographic modeling

We used a traditional model-based approach (fastsimcoal2) to explicitly test for a

bottleneck in North American starlings (Excoffier *et al.,* 2013). We first estimated the

folded SFS using a Python script from Simon Martin (available at

https://github.com/simonhmartin/genomics_general). We modified the simple

1PopBot20Mb.tpl and .est files provided with the software as follows: the size of the

bottleneck (NBOT: 10-1000), start of bottleneck (TBOT: 10-300), ancestral size (ANC:

1000-100000), current size of the population (NCUR: 100-100000), and end of bottleneck

(TENDBOT: TBOT + 500). We performed 100 runs of this model and assessed model fit

using delta-likelihood for each run of the model; since all models use the same numbers of

parameters, the Akaike information criteria do not change between models. We also ran a

demographic model that takes linkage into account, the Stairway plot method, but given

that this method does not explicitly test for a genetic bottleneck, we describe the method

and results in the Supplementary Information.

(4) Selection analyses

We first identified environmental variation at each sampling location using the R package

raster (Hijmans & van Etten, 2012), extracting all 19 bioclimatic variables for each set of

sampling coordinates from WorldClim 2.0 at a resolution of 5 min on June 16, 2018 (Fick &

Hijmans, 2017). We recognize that these environmental conditions represent those

experienced during the non-breeding season, and therefore they do not represent the full

range of conditions experienced by the starlings sampled; see the Supplementary

Information for a discussion of our assumptions about this environmental sampling. We

used redundancy analysis (RDA) to examine how loci may covary across multiple

environmental gradients (Forester *et al.,* 2018). RDA is especially powerful when testing

for weak selection, and detects true positives in large data sets more reliably than other

multivariate methods like Random Forest (Forester *et al.,* 2018). Because RDA requires no

missing data, we first imputed genotypes where missing sites were assigned the genotype

of highest probability across all individuals—a conservative but quick imputation method.

Before imputation and after filtering, about 31% of genotypes were missing from the

dataset of 15,038 SNPs (not filtered for MAF), and we required that all imputed SNPs were

present in at least 80% of individuals. We then used the R package *vegan* to run the RDA

model (Oksanen *et al.*, 2019); for a full description of this method, see Forester *et al.,* 2018.

Briefly, RDA uses constrained ordination to model a set of predictor variables, and

unconstrained ordination axes to model the response (genetic variation). RDA infers

selection on a particular locus when it loads heavily onto one or more unconstrained

predictor axes, and we retained five variables with relatively low variance inflation factors:

BIO1 or Annual Mean Temperature (VIF=3.54), BIO7 or Temperature Annual Range (4.55),

BIO12 or Annual Precipitation (8.69), BIO16 or Precipitation of Wettest Quarter (7.91), and

elevation (2.19). We tested for significance using the *anova.cca* function within the *vegan*

package, and also permuted predictor values across individuals to further check

significance of the model. We identify candidate loci as those that are 3 or more standard

deviations outside the mean loading. The R scripts for all RDA analyses and figures were

written by Brenna Forester (available at https://popgen.nescent.org/2018-03-

27_RDA_GEA.html).

RESULTS

(1) Patterns of genetic diversity and differentiation

We identified 15,038 SNPs at a mean of 27X sequencing depth across 17 sampling

locations. Using a dataset filtered by a minor allele frequency of 0.1, we find that genome-

wide F_{ST} is extremely low (0.0085), and measures of genetic diversity do not vary

substantially among sampling locations (Table 1). Across all of these wintering starling

populations, the highest pairwise F_{ST} was 0.0106 between birds from the adjacent states of

Arizona and New Mexico. Using a haplotype-based statistic of differentiation, ϕ_{ST} among

populations shows an absence of genetic structure (ϕ_{ST} = 0.0002). Hierarchical AMOVAs

reveal that 94% of the observed genetic variance is explained by variation within

individuals, and the remaining variance reflects differences among individuals in the same

population, with negligible variation explained at the between-population level (Figure 1C-

D). Across the genome, F_{ST} and nucleotide diversity are exceptionally low (Table 1,2).

Genome-wide heterozygosity is moderate at 0.339, and across loci observed heterozygosity

differs significantly from expected (t = 66.6, df = 3569, P<0.001, Table 1).

	df	Sum of squares	variance components	% variance explained
Between populations	1	1354	0.264	0.021
Within populations	156	205297	73.75	5.936
Within samples	158	184622	1168	94.04
Total	*315*	*391273*	*1242*	

302   **Figure 1.** Population structure. A) Principal components analysis on 6,287 SNPs, explaining
303   1.03-1.08% of genetic variation observed. B) STRUCTURE analyses with K=2 and K=3 (best
304   supported). C) Significance testing of hierarchical AMOVAs: the histogram shows expected
305   variance components based on 999 simulations, and the black diamond is the observed
306   variance component. D) AMOVA results.
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310   **Table 1.** Population genetic summary statistics. For each sampling location, the table
311   shows the number of private alleles, observed and expected heterozygosity, nucleotide
312   diversity, and the inbreeding coefficient F_{IS}.
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Pop.	No.	Private	H_{obs}	H_{exp}	π	F_{IS}
NY	10	0	0.34	0.25	0.35	0.037
NH	11	1	0.33	0.26	0.35	0.065
NC	11	1	0.33	0.25	0.35	0.060
IL	7	1	0.33	0.26	0.35	0.056
IA	9	0	0.33	0.24	0.36	0.078
WI	8	0	0.33	0.22	0.35	0.050
MO	9	0	0.33	0.23	0.35	0.053
NE	11	0	0.34	0.22	0.35	0.042
CO	8	1	0.33	0.24	0.35	0.066
KS	9	2	0.35	0.23	0.35	0.012
TX	9	1	0.33	0.24	0.35	0.052
NM	7	2	0.33	0.29	0.36	0.068
AZ	10	1	0.32	0.24	0.35	0.069
NV	10	0	0.33	0.31	0.36	0.068
CA	11	0	0.33	0.25	0.35	0.061
ID	8	0	0.33	0.22	0.35	0.047
WA	10	2	0.32	0.30	0.35	0.082

(2) Population structure

F_{ST} among populations is relatively low overall (Table 2): starlings sampled in

Arizona show the highest differentiation from other sampling locations (F_{ST} = 0.008-0.011),

and only a few other pairwise comparisons (NM-IL, NM-CO, and CO-WI) show an F_{ST} of

0.009 or higher. However, we do not recover clear population structure. The first two

principal components each explain about 1% of variation among individuals (Figure 1A),

and although STRUCTURE identified three populations at the best-supported value of K,

these predicted populations do not show obvious differences in ancestry proportions

(Figure 1B, Figure S1). Controlling for shared ancestry does not resolve population

structure, and instead provides support for uniform gene flow among individuals (Figure

S2). K-means clustering within DAPC also does not identify biologically relevant clusters.

	AZ	CA	CO	IA	ID	IL	KS	MO	NC	NE	NH	NM	NV	NY	TX	WA	WI
Arizona	28	0.0085	0.0090	0.0093	0.0090	0.0082	0.00912	0.012	0.0093	0.0090	0.0125	0.01110	0.010	0.010	0.012	0.010	0.015
California	5	40	0.0080	0.0065	0.0050	0.0050	0.00510	0.0050	0.0075	0.0060	0.0070	0.00510	0.0060	0.0060	0.0070	0.0060	0.00610
Colorado	0	7	43	0.0060	0.0087	0.0082	0.0070	0.0070	0.0060	0.00611	0.0077	0.0104	0.0072	0.0060	0.0070	0.0072	0.0090
Iowa	4	16	0	8	0.0060	0.00514	0.0084	0.00714	0.0052	0.0090	0.0076	0.00714	0.0060	0.0062	0.0070	0.0070	0.0086
Idaho	0	0	4	0	23	0.0070	0.0070	0.0060	0.0070	0.0054	0.0070	0.0080	0.0048	0.0060	0.0070	0.00619	0.0070
Illinois	5	2	5	2	0	12	0.0075	0.00617	0.0070	0.0060	0.0067	0.0117	0.0050	0.0060	0.0062	0.0070	0.00812
Kansas	2	12	2	0	0	5	7	0.00512	0.00629	0.0052	0.0065	0.0062	0.0050	0.0062	0.0072	0.0060	0.00610
Missouri	5	5	0	0	0	5	5	11	0.00516	0.0060	0.0065	0.00726	0.0060	0.0060	0.0060	0.0070	0.00711
North Carolina	0	0	0	0	0	0	16	0	32	0.0060	0.0070	0.0070	0.0060	0.0060	0.00744	0.0080	0.0064
Nebraska	0	8	24	0	16	0	0	0	0	12	0.00812	0.0064	0.0068	0.0070	0.0070	0.0070	0.0060
New Hampshire	28	6	0	3	0	0	3	3	0	6	9	0.0076	0.0073	0.0060	0.0070	0.0070	0.00713
New Mexico	11	0	0	0	0	0	0	0	11	11	0	67	0.0060	0.0080	0.0070	0.0070	0.0070
Nevada	0	3	4	1	18	0	1	0	0	11	3	1	14	0.0050	0.0070	0.00614	0.0061
New York	3	10	0	1	0	15	1	10	0	1	6	1	1	1	0.0070	0.0060	0.00521
Texas	5	0	0	0	0	0	0	3	26	0	0	2	0	0	65	0.0070	0.0080
Washington	0	0	0	0	0	0	0	0	0	7	0	0	60	0	0	7	0.0060
Wisconsin	10	0	2	2	2	8	6	4	0	2	2	0	0	10	0	0	31

*percent assigned to each state does not sum to 100 because Minnesota, Oregon, and Vermont are not included in the genetic dataset.

Table 2. Differentiation and movement among sampling locations. This table shows the percentage of birds assigned to each sampling location from every origin. Note that these values are directional: although 12 starlings that were collected in AZ originated in KS, only 2 starlings collected in KS originated in AZ. The black diagonal indicates the percentage of birds assigned to that collection state according to a discriminant function analysis of molt-origin presented in Werner et al. For example, 28% of birds collected in Arizona originated in that location. F_{ST} among locations is presented above the diagonal, and darker gray colors indicate higher F_{ST}.

Starling populations follow a clear pattern of isolation-by-distance (Mantel r = 0.139, P < 0.0001). Spatially explicit models of isolation-by-distance suggest a fairly uniform rate of migration range-wide, where local increases in migration rate are likely a model artifact (Supplementary Information). However, models of isolation by environment (IBE) show that the relationship between environmental and genetic distances is stronger than the simple geographic-genetic distance model. After controlling for geographic distance, we find that all bioclimatic variables tested show non-zero correlations between

environmental and genetic distances (Supplementary Information). There is a strong

positive relationship between genetic distance and both precipitation in the wettest

quarter (BIO16, Mantel r = 0.282, P = 0.001) and elevation (Mantel r = 0.146, P = 0.001).

(3) Demographic modeling

A traditional SFS-based model recovers a bottleneck in population size, as expected:

the estimated effective population size during the bottleneck is ~16,000 individuals (1% of

the ancestral population size). However, our demographic model suggests that the current

population size of North American starlings is nearly identical to its pre-bottleneck size

(Supplementary Information). In fact, every run of the model finds that starlings' current

effective population size (1.6 million individuals) is considerably higher than the estimated

N_e of the founding population. This model suggests that starlings experienced rapid

population growth after the initial founder effect, which may contribute to the overall lack

of evidence for inbreeding. We do detect very low levels of inbreeding within some

populations (Table 1, highest F_{IS} = 0.082 in Washington). Taken together, these models do

not suggest a classical founder effect, whereby effective population size remains very low

for many generations post-bottleneck.

(4) Selection analyses

Starlings encounter a range of precipitation, temperature, and elevation across their range

and redundancy analyses revealed the strongest signatures of local adaptation, showing

that 178 variants are correlated with environmental differences among populations (F =

1.022, P = 0.002, Figure 2). Populations living in warmer climates tend to cluster more

closely in the left quadrant and high elevation populations cluster in the middle right

quadrant. However, populations do not cluster based on geographic distance: for example,

starlings from TX and WA cluster closely due to shared genetic variants, even though the

two populations differ substantially in precipitation and temperature and are

geographically separated. After controlling for population structure, candidates for

selection are equally distributed among elevation, temperature- and precipitation-

associated predictors. Importantly, when environmental predictors are randomly shuffled,

the RDA model is not significant. Mean annual temperature (BIO1) opposes selective

pressure related to the range of temperatures experienced each year (BIO7), annual

precipitation (BIO12), precipitation in the wettest quarter (BIO16) and elevation (Figure

2). Genes under selection tend to have lower allele frequencies, and function in several

biological processes (Table S2).

Figure 2. Evidence for incipient local adaptation. Redundancy analyses indicate that 191 SNPs (small gray points) are associated with bioclimatic predictors (vectors). BIO1: mean annual temperature; BIO7: temperature annual range; BIO12: annual precipitation; BIO16: precipitation of wettest quarter.

DISCUSSION

Our whole-genome data reveal that invasive starling populations show moderate

genetic diversity across North America, rapid population growth after a genetic bottleneck,

negligible geographically structured genetic differentiation, significant but quite weak

isolation-by-distance, and intriguing associations of genetic variation with environmental

parameters that might result from adaptive processes. Genetic diversity is typically

predicted to be low in invasive populations (Dlugosch & Parker, 2008), and although

RADseq tends to estimate lower genetic diversity than the true diversity, this bias is

minimal for taxa known to be genetically depauperate (Cabe, 1998; Cariou *et al.,* 2016).

Within invasion genetics, one area of recent focus involves the spatial dynamics of invasion,

and we can explore how genetic variation varies across the post-invasion range of starlings

in North America. Although there is a significant (but low magnitude) signal of isolation-by-

distance, hierarchical AMOVAs find that variation within and among individuals explains

observed differentiation better than variation among populations. There is no evidence of

population structure, and while models indicate some subtle spatial patterns of genetic

variation, these model-based inferences likely reflect sampling artifacts (Supplementary

Information). Finally, after the initial founder effect, the effective population size has grown

by a ratio of 1:100. These patterns are consistent with the expectation that extensive gene

flow—as shown by extremely low F_{ST} among populations (Table 2)—maintains high

connectivity across North American starling populations. When interpreted within the

context of a complementary exploration of movements inferred from feather isotope

assays, this genomic survey confirms that dispersal and migration continue to influence the

genetic variation generated in just over a century of range expansion.

Long-distance dispersal may be common in introduced starlings on multiple

continents, where rates of dispersal are determined by demographic changes and

environmental quality. In South Africa, invasive starlings disperse when their natal

environment becomes crowded or unsuitable, as at the leading edge of their range

expansion (Hui *et al.,* 2012). In North America, the effective population size (N_e) of the

present-day invasive population has expanded dramatically, with models indicating that

current N_e is even larger than the N_e of the founding population. This explosive population

growth from hundreds to millions of individuals may have encouraged long-distance

dispersal away from the dense populations in eastern North America. In general, multiple

lines of evidence have shown that the dispersal rates and distances of juvenile starlings in

North America are remarkably high (Dolbeer, 1982; Cabe, 1999; Werner, Fischer, &

Hobson, 2020). Considered in concert with similar findings in the invasive South African

and Australian starling populations (Hui *et al.,* 2012; Phair *et al.,* 2017), it is clear that

dispersal similarly plays an ongoing and major role in the connectivity of North American

starling populations.

Although it is generally important, dispersal has not been uniform across time or

space in North American starlings. One open question in invasion genetics is whether

expanding populations maximize spatiotemporal fitness via spatial sorting and incipient

adaptation (Williams et al. 2019). Empirical studies of some invasions suggest that spatial

sorting is common, including in common mynas (Berthouly-Salazar *et al.*, 2012),

pumpkinseed fish (Ashenden *et al.,* 2017), and cane toads (Brown *et al.*, 2014). More

empirical work is needed to verify the conditions in which spatial sorting can lead to

lasting shifts in fitness (Lee, 2011; Phillips & Perkins, 2019; Williams *et al.*, 2019).

However, we do know that adaptive dispersal strategies can facilitate range expansion in

Western bluebirds (Duckworth, 2008) and invasive beetles (Lombaert *et al.,* 2014; Ochocki

& Miller, 2017) among other species. For North American starlings, we can combine

historical records, isotopic and genetic evidence to make inferences about the possibility of

spatial sorting.

Feather isotope data from the same starlings sampled in this genetic survey

suggests that movements are not uniform continent-wide, and that both juvenile dispersal

and regional differences in migration could influence genetic variation (Werner, Fischer, &

Hobson, 2020). If starling movement was the primary driver of differentiation, then we

would expect F_{ST} among populations to be directly related to the number of birds collected

and assigned to that population. Starlings in the western U.S. appear to have differentiated

subtly from their eastern counterparts based on the higher F_{ST} between Arizona—and to

some degree, New Mexico and Colorado—and all other sampling locations (Table 2). Birds

collected in these southwestern states are also assigned to those states by discriminant

function analysis (Werner, Fischer, & Hobson, 2020); for example, 67% of starlings

collected in New Mexico were also assigned to that state (Table 2). We suggest that birds

assigned to the same state where they were collected may reside in that state year-round,

but we note that collecting feathers once in the lifespan of the bird does not allow us to

determine the bird's lifelong migration and dispersal. The highest rates of pairwise

differentiation occur between states where one of those "populations" shows a relatively

high collection-state assignment, which suggests that starling movement among sampling

locations does impact the genetic variation. This comparison is further supported by the

low but significant isolation-by-distance across the North American range; however,

environmental pressures explain genetic variation better than a simple model that

accounts for geographic distance. In addition, assignment to a state other than the

collection state seems to have only a minor effect on F_{ST}, even though we might expect that

movement among geographically distant sampling locations might lead to gene flow among

those locations. In Arizona—where F_{ST} is highest but still relatively weak—Werner et al.'s

model suggests that 25% of birds originated in New Hampshire, which is approximately

the percentage of birds assigned to Arizona (28%, Table 2). Unsurprisingly, geography,

environment, and demography all appear to influence genetic variation within North

American starling, and explicit tests for selection can help to clarify the relative weight of

each factor.

This species' general genome-wide panmixia across the continent allows for tests of

selection on loci that may be involved in local adaptation. Because genetic diversity and

differentiation are exceptionally low, it is relatively easy to identify sites that have

differentiated against this background of low genome-wide or population-specific

divergence (Dlugosch *et al.,* 2015). We find that almost 200 of the 15,038 RAD loci appear

to be under selection using a redundancy analysis. Only 13 of these SNPs overlap with the

SNPs identified by a latent-factor mixed model, and there is no overlap between the RDA

candidates and the differentiation-based scans (BayeScEnv and Bayescan). It is

unsurprising that each test identifies different candidates for selection, because the

assumptions underlying each are very different (for more details on our approach to

selection testing, see Supplementary Information). Rather than making inferences based on

the genes identified by these scans, we instead propose that genotype-environment

associations show that changes in precipitation and temperature can explain genetic

variation in North American starlings. However, we do suggest that rapid evolution in

genes underlying key physiological processes may have supported the starling's spread

across North America (Supplemental Information). Specifically, we hypothesize that aridity

and cold temperatures that are not experienced in the starling's native range exert enough

selective pressure on North American starlings to result in incipient local adaptation. While

this finding suggests that local adaptation may explain genetic variation within the North

American starling invasion, there are several relevant caveats to this approach.

Because there is evidence of a genetic bottleneck in the North American starling

invasion, it is possible that artifacts resulting from genetic drift could mimic these patterns

of local adaptation. We cannot rule out allele surfing during range expansion as an

explanation based on the methods currently available for testing genotype-environment

associations in a young and expanding system like this one. However, given how short the

timescale of divergence is in North American starlings, such drift effects are more likely if

genetic variation is (1) ancestral and (2) structured among populations. Gene flow among

starling populations is remarkably high, but we do not know if variants putatively under

selection were present in the ancestral population. These alleles under putative selection

do not approach fixation, as no putative outlier has an allele frequency greater than 0.28.

We note that concordance among environmental and genetic distances (e.g., partial Mantel

tests) indicate that spatial autocorrelation complicates our selection inferences. Under

these conditions, any evidence for selection is likely to be weak, and as always these

selection scans can generate false positives. However, RDA has the highest rate of true

positives and lowest of false positives, and although this method has not been tested in

such recent expansions, RDA is well-suited to systems where F_{ST} is very low (Meirmans,

498 2015, Forester *et al.*, 2018, Supplementary Information). Finally, we do not explicitly

499 control for linkage, and some allele frequency shifts could be explained by recombination

500 or by genetic hitchhiking.

501  Our study focuses on birds collected during the winter, which may limit our

502 inferences about population structure and selection to these wintering populations. As

503 discussed above, isotopic evidence suggests that starlings in the western U.S. tend to move

504 only regionally whereas birds sampled in the eastern U.S. undertake longer movements

505 (Table 2). This in turn suggests that starlings overwintering in the western U.S. are more

506 likely to breed nearby, and thus the environmental conditions may not change as

507 dramatically among wintering and breeding ranges. In addition, the environmental

508 conditions that we expect to drive selection—precipitation and temperature—vary most

509 substantially in the southwestern region: for example, the sampling location in Arizona is

510 consistently warmer (BIO1) and drier (BIO12 and BIO16) than other locations (Figure 2,

511 Supplementary Information). Western populations experience these environmental

512 conditions year-round, which could allow selection to drive advantageous alleles toward

513 fixation. Elsewhere in the U.S., starlings move more freely among states: individuals within

514 each sampling location may come from different breeding populations, and additional

515 sampling could reveal stronger population structure among true breeding populations.

516 However, if our sampling overlooked some true populations, we would expect some signal

517 of population structure. Individual-based tests of population structure—e.g., those that do

518 not define possible populations *a priori*—do not recover any signals of population

519 structure. This sampling strategy uses the more vagile eastern populations as a comparison

520 to more-resident western populations that may also be under stronger selection. This

framing may suggest that population structure in the western U.S.—of which we find no

evidence—could explain allele frequency shifts that we infer to be selection, but when we

compare the relative importance of geographic and environmental distances in partial

Mantel tests, we find that environmental conditions better explain genetic variation. This

evidence supports our interpretation of selection as a major driver of genetic variation in

North American starlings.

A similar project on starlings in the Australian invasion—which colonized that

continent nearly concurrently with the North American invasion—found that geographic

but not environmental distance explains genetic patterns there (Cardilini *et al.,* 2020).

Starlings in the Australian range show substantial population structuring and significant

patterns of isolation-by-distance. Earlier work had shown that gene flow among Australian

starling populations is low (Rollins *et al.,* 2009), and phylogeographic patterns of

mitochondrial sequence variation confirm that starlings on the edge of the expansion front

in Western Australia have differentiated from those still living in the introduction site

(Rollins *et al.,* 2011). In fact, starlings at the expansion front may have rapidly adapted

during the Australian invasion (Rollins *et al.,* 2016): the proportion of adult starlings in

Western Australia carrying a novel mitochondrial haplotype has increased rapidly only at

this range edge. A genotyping-by-sequencing survey employing a much greater number of

SNP markers indicates three population subdivisions in Australia, where geographic

distance explains genetic differentiation in starlings better than does environmental

variation (Cardilini *et al.,* 2020). Global F_{ST} across all Australian populations is an order of

magnitude higher than the equivalent F_{ST} index across North America, despite similar areas

sampled. In Australia, the strong evidence for isolation-by-distance and founder effects

complicate attempts to disentangle selection from drift, yet despite their differences in

invasion dynamics, genotype-environment associations reveal signatures of selection in

both invasions. On both continents, starling genetic variation can be explained by extremes

in temperature and precipitation, and preliminary results of whole-genome resequencing

of native and introduced populations confirm that variability in temperature and

precipitation may shape observed genetic variation in starlings world-wide (Hofmeister et

al., *in prep*).

Our results contribute to the growing evidence of rapid adaptation in some

expanding populations, even in extremely young systems. Some studies of rapidly

expanding invasions find little evidence that adaptation may facilitate this expansion, as in

corals (Leydet et al. 2018). However, other work suggests a role for selection in supporting

rapid range expansion, such as in experimental studies of flour beetles (Szucs *et al.,* 2017)

and empirical work in guppies (Baltazar-Soares *et al.,* 2019). Invasion biologists have long

highlighted propagule pressure as a driver of invasion success, but the genetic composition

may be just as important as the size of the establishing population (Briski *et al.,* 2018). For

example, genetic bottlenecks in monk parakeets, another avian invader now distributed

world-wide, do not seem to inhibit invasion success (Edelaar *et al.,* 2015). Pre-adaptation

in the native range or selection during transport may facilitate the spread of invasive

species, and human commensalism may support establishment and spread, as shown in

house sparrows (Ravinet *et al.,* 2018) and common mynas (Cohen *et al.,* 2019), and

reviewed across alien bird species (Cardador & Blackburn, 2019). Empirical studies of

invaders like the ones described here also show how, in addition to genetic variation,

epigenetic shifts and/or plastic changes in gene expression may support the establishment

and expansion of invasive species (Marin *et al.,* 2019). In the well-studied house sparrow—

a system quite similar to starlings—epigenetic shifts may have supported invasions in

Africa (Liebl *et al.,* 2013) but not necessarily in Australia (Sheldon *et al.,* 2018). Taken

together, recent work suggests that we should consider a much wider range of

demographic and ecological processes that lead to adaptive evolution in invading

populations.

Invasive populations allow us to explore the genetic consequences of colonization

and establishment in novel environments. On a background of low genetic differentiation

and diversity, we find evidence of incipient genotype-environment associations in North

American starlings. Here we explore how genetic variation changes across the landscape,

but we cannot fully understand gene flow without studies of dispersal and migration of the

individuals that carry genes. Our results complement other recent studies that reveal

associations between climate variables and particular loci in North American vertebrates

(Schweizer *et al.,* 2015; Bay *et al.,* 2018). Finally, we suggest that our study adds to those

suggesting that rapid local adaptation can evolve even in dispersive and young populations.

582
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584
585   NRH, IJL designed research, SJW contributed samples, NRH analyzed data, NRH, IJL wrote
586   paper.
587

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605
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608 All scripts are archived on GitHub: https://github.com/nathofme/radseq-NAm
609 All data files that accompany the above scripts are archived on Dryad (doi:
610 10.5061/dryad.j424f07), and raw sequencing data archived in NCBI SRA (Bioproject
611 PRJNA545151).
612
613 **REFERENCES**
614

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