Is it reasonable to account for population structure in genome-wide association studies?

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7 Abstract

8 Population structure is widely perceived as a noise factor that undermines the quality of association between an SNP variable and a phenotypic variable in genome-wide association studies (GWAS). 9 10 The linear model for GWAS generally accounts for population-structure variables to obtain the adjusted phenotype which has less noise. Its result is known to amplify the contrast between 11 12 significant SNPs and insignificant SNPs in a resultant Manhattan plot. In fact, however, conventional GWAS practice often implements the linear model in an unusual way in that the 13 14 population-structure variables are incorporated into the linear model in the form of continuous variables rather than factor variables. If the coefficients for population-structure variables change 15 across all SNPs, then each SNP variable will be regressed against a differently adjusted phenotypic 16 variable, making the GWAS process unreliable. Focusing on this concern, this study investigated 17 whether accounting for population-structure variables in the linear model for GWAS can assure 18 the adjusted phenotypes to be consistent across all SNPs. The result showed that the adjusted 19 phenotypes resulting across all SNPs were not consistent, which is alarming considering 20 21 conventional GWAS practice that accounts for population structure.

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25 Introduction

26 Genome-wide association studies (GWAS) aim to identify single nucleotide 27 polymorphisms (SNPs) whose allelic variation is significantly tied to phenotypic variation. In principle, the tie between the allelic variation and phenotypic variation can be measured based on 28 the variance among the phenotypic averages for all scores per each SNP (Kim, 2017; Kim, 2018a). 29 Greater variance indicates a stronger tie. Conventional GWAS practice has been largely conducted 30 using statistical methods such as the linear model and the linear mixed model (LMM). To date, the 31 use of the LMM has been widely encouraged because of the general perception that accounting 32 for a kinship matrix can reduce the noise between a phenotypic variable and an SNP variable, by 33 correcting the bias that genetic relationship among entities in a population introduces (Yu et al. 34 2006; Bradbury et al, 2007; Kang et al, 2008; Lipka et al, 2012; Hoffman, 2013; Kim et al, 2018b). 35 36 Recently, however, Kim (2019) demonstrated that the use of a kinship matrix actually makes the LMM unreliable. In this regard, this study excluded the LMM. 37

38 Conventional GWAS practice based on the linear model often regresses each SNP variable along with population-structure variables against a phenotypic variable, one by one across all SNPs. 39 Therein, the use of population-structure variables aims to obtain an adjusted phenotype calculated 40 by subtracting the estimated population-structure effect from the phenotype (Yu et al, 2006; 41 42 Bradbury et al, 2007; Kang et al, 2008; Lipka et al, 2012; Hoffman, 2013; Kim et al, 2018b). For reliable GWAS practice, it is crucial to assure the adjusted phenotypes resulting across all SNPs 43 are consistent. Otherwise, every SNP variable will be regressed against a differently adjusted 44 phenotypic variable, which consequently confounds GWAS results. This study investigated 45 whether accounting for population structure in the linear model for GWAS assures the adjusted 46 phenotypes resulting across all SNPs to be consistent 47

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52 Materials and Methods

53 Rice data set

This study used a rice data set comprising SNP data, principle component analysis (PCA) data and phenotypic data. The data set was originally used for GWAS by Zhao et al. (2011) and freely available to public at <u>http://ricediversity.org/data/index.cfm</u>. Therefore, more information about the data set can be found from the related paper. In the original data, 413 entities were genotyped with 36,901 SNPs. The number of SNPs was reduced to 12,983 by screening with a criterion of the minor allele frequency (MAF) of 0.1. The phenotype chosen for this study was seed length.

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61 Statistical model

62 The two linear models were established as follows:

$$63 y = \mu + \beta_1 x_{SNP} + \varepsilon (1)$$

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$$y = \mu + \beta_1 x_{SNP} + \beta_2 x_{PCA1} + \beta_3 x_{PCA2} + \beta_4 x_{PCA3} + \beta_5 x_{PCA4} + \varepsilon$$
(2)

where y = the phenotypic observation; $\mu =$ the phenotypic mean; $x_{SNP} =$ the SNP variable; x_{PCA1} = the PCA1 variable; $x_{PCA2} =$ the PCA2 variable; $x_{PCA3} =$ the PCA3 variable; $x_{PCA4} =$ the PCA4 variable; $\varepsilon =$ the error term; $\beta_1 =$ the coefficient for x_{SNP} ; $\beta_2 =$ the coefficient for x_{PCA1} ; $\beta_3 =$ the coefficient for x_{PCA2} ; $\beta_4 =$ the coefficient for x_{PCA3} ; $\beta_5 =$ the coefficient for x_{PCA4} .

Equation 1 regresses the SNP variable against the phenotypic variable. Equation 2 regresses the SNP variable along with the four PCA variables (x_{PCA1} , x_{PCA2} , x_{PCA3} , x_{PCA4}) against the phenotypic variable. This means that Equation 2 regresses the SNP variable against the adjusted phenotypic variable obtained by accounting for the four PCA variables. Equation 3 highlights the adjusted phenotypic variable:

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$$y - \beta_2 x_{PCA1} - \beta_3 x_{PCA2} - \beta_4 x_{PCA3} - \beta_5 x_{PCA4} = \mu + \beta_1 x_{SNP} + \varepsilon$$
 (3)

Equation 3 is compatible with Equation 2 and represents the adjusted phenotypic variable as $y - \beta_2 x_{PCA1} - \beta_3 x_{PCA2} - \beta_4 x_{PCA3} - \beta_5 x_{PCA4}$.

77 Manhattan plot

The F test was implemented as a significance test, from which P values were obtained. The P 78 values transformed by $-log_{10}$ were drawn in a Manhattan plot. It is important to note that the P 79 values resulting from the linear model for GWAS are prone to genomic inflation. Prior to 80 81 confirming the resultant Manhattan plot, therefore, it is necessary to calculate the genomic inflation factor (λ_{GC}). The situation of $\lambda_{GC} > 1$ indicates the genomic inflation, which means that the 82 resultant P values are overly estimated compared with the χ^2 -distribution (van Iterson et al, 2017). 83 This study adjusted the genomic inflation using the genomic control. More information about the 84 genomic control can be found in previous studies (Devlin and Roeder, 1999; Yang et al, 2011; van 85 Iterson et al, 2017). 86

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88 Integrity validation of accounting for population structure in GWAS

89 Equation 3 (compatible with Equation 2) regresses each SNP variable against an adjusted phenotypic variable. As GWAS handle numerous SNPs one by one at a time, it is important to 90 assure that the adjusted phenotypes resulting across all SNPs are consistent. Otherwise, each SNP 91 92 variable will be regressed against a differently adjusted phenotypic variable. The consistency 93 among the adjusted phenotypes resulting across all SNPs can be achieved, only if every coefficient per each PCA variable is consistent across all SNPs. To check the consistency among the adjusted 94 phenotypes resulting across all SNPs, this study calculated Pearson coefficients between the 95 phenotype and every adjusted phenotype. 96

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98 Data set and R code

All computations were conducted using R (R Core Team, 2016). The data set and R scripts used

100 in this study are freely available at <u>https://github.com/bongsongkim/Population.Structure.GWAS</u>.

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103 **Results**

104 Validation of consistency across all adjusted phenotypes

Table 1 summarizes the coefficients per each PCA variable, resulting from applying all SNPs to Equation 3. Figure 1 represents the estimated coefficients per each PCA variable, showing large variation. Figure 2 represents the estimated Pearson correlation coefficients between the phenotype and every adjusted phenotype, illustrating the adjusted phenotypes resulting across all SNPs are not consistent. This means that each SNP variable is regressed against a differently adjusted phenotypic variable.

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Table 1. Summary of coefficients per each PCA variable in relation to Equation 3.

	Min.	1 st Qu.	Median	Mean	3 rd Qu.	Max
β_2	-7.833	-2.246	-2.162	-2.109	-2.024	3.629
β_3	-4.438	-1.097	-1.020	-1.007	-0.944	3.271
eta_4	-13.610	-9.229	-9.194	-9.180	-9.160	0.659
eta_5	-9.308	3.016	3.087	3.025	3.121	8.704

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Figure 1. (A) Estimated coefficients for the PCA1 variable, (B) estimated coefficients for the
PCA2 variable, (C) estimated coefficients for the PCA3 variable, (D) estimated coefficients for
the PCA4 variable.



122 Figure 2. Pearson correlation coefficients between the phenotype and every adjusted phenotype.

125 Impact of accounting for population structure in GWAS

126 Figure 3 shows four Manhattan plots, for which the same SNP and phenotypic data were used. Figure 3A represents the Manhattan plot in relation to Equation 1, in which the resultant λ_{GC} was 127 3.688. Figure 3C is the same as Figure 3A in shape. However, Figure 3C meets $\lambda_{GC} = 1$ by 128 129 implementing the genomic control with Figure 3A. Figure 3E represents the Manhattan plot in relation to Equation 3, in which the resultant λ_{GC} was 1.433. Compared with Figure 3A, Figure 3E 130 has substantially lower λ_{GC} . This suggests that accounting for the four PCA variables was 131 impactful in diminishing the genomic inflation. Figure 3G was obtained by adjusting Figure 3E by 132 implementing the genomic control. This led to $\lambda_{GC} = 1$ in Figure 3G. It is apparent that Figure 3E 133 has clearer background than Figure 3A in relation to accounting for the four PCA variables. In this 134 regard, previous studies explained that accounting for population structure in the linear model for 135 136 GWAS eliminates the noise in SNP-phenotype associations, which results in clear background in a resultant Manhattan plot (Yu et al, 2006; Kang et al, 2008; Korte and Farlow, 2013; Sul et al, 137 2018; Barton et al, 2019). However, Figure 4 illustrates that significant SNP-phenotype 138 associations are not consistent between Figures 3C and 3G. This means that the clear background 139 140 was not from eliminating the noise in SNP-phenotype associations, but from defining new SNP-141 phenotype associations.

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Figure 3. (A) Manhattan plot obtained by not accounting for the four PCA variables ($\lambda_{GC} = 3.688$), (C) Manhattan plot obtained by adjusting Figure 3A with implementing the genomic control ($\lambda_{GC} = 1.000$), (E) Manhattan plot obtained by accounting for the four PCA variables ($\lambda_{GC} = 1.433$), (G) Manhattan plot obtained by adjusting Figure 3C with implementing the genomic control ($\lambda_{GC} = 1.000$).

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Figure 4. Correlation plot between the -log₁₀ (P) values obtained by not accounting for the four
PCA variables (Figure 3C) and the -log₁₀ (P) values obtained by accounting for the four PCA
variables (Figure 3G).

164 **Discussion**

It is generally perceived that accounting for population structure in GWAS improves the 165 quality of visual representation of a Manhattan plot by both suppressing genomic inflation and 166 reducing false-positive SNP-phenotype associations (Yu et al, 2006; Bradbury et al, 2007; Kang 167 et al, 2008; Lipka et al, 2012; Hoffman, 2013; Kim et al, 2018b). In fact, this study showed that 168 accounting for the four PCA variables was very effective in diminishing the genomic inflation. 169 Surprisingly, however, this study revealed that accounting for the four PCA variables breaks the 170 171 consistency among the adjusted phenotypes resulting across all SNPs. The loss of the consistency consequently causes each SNP variable to be regressed against a differently adjusted variable, 172 making the GWAS process unreliable. The use of population-structure variables in the linear 173 model for GWAS implies two errors. First, the linear model is misused. Considering that the linear 174 175 model is suited for analyzing data in experimental blocks, the use of continuous variables rather than factor variables necessarily causes an error. Second, the assumption for the relationship 176 between phenotype and population structure is unjustified. The linear model for GWAS generally 177 assumes that the population-structure variables additively contribute to the phenotypic variable. 178 However, how the population structure biologically influences the phenotype has yet been 179 180 unknown. Regardless of whether the additivity of the population-structure variables is true or false, the current way of accounting for population structure is inappropriate in that population-structure 181 effects vary across all SNPs. The abovementioned errors consequently lead to the loss of the 182 consistency among the adjusted phenotypes resulting across all SNPs and cause each SNP variable 183 184 to be regressed against a differently adjusted phenotypic variable.

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186 Conclusion

The linear model assures to preserve the consistency among the adjusted phenotypes resulting across all SNPs, only if factor variables such as years, locations, replications and treatments are used. This study concluded that the conventional way of accounting for population structure makes the GWAS process unreliable. This is because the population structure is represented as continuous variables. If population structure can be represented as factor variables, accounting for the population structure in the linear model for GWAS will be sound.

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