

1 Polygenic risk score of alcohol consumption 2 predicts alcohol-related morbidity and all- 3 cause mortality

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23

24 Abstract

25

26 **Objective:** To develop a highly polygenic risk score (PRS) for alcohol consumption and study whether
27 it predicts alcohol-related morbidity and all-cause mortality.

28

29 **Design:** Biobank-based prospective cohort study

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31 **Setting:** FinnGen Study (Finland)

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33 **Participants:** 96,499 genotyped participants from the nationwide prospective FinnGen study and
34 36,499 participants from prospective cohorts (Health 2000, FINRISK, Twin Cohort) with detailed
35 baseline data and up to 25 years of follow-up time.

36

37 **Main outcome measures:** Incident alcohol-related morbidity and alcohol-related or all-cause
38 mortality, based on hospitalizations, outpatient specialist care, drug purchases, and death reports.

39

40 **Results:** In 96,499 FinnGen participants there were in total 4,785 first-observed incident alcohol-
41 related health events. The PRS of alcohol consumption was associated with alcohol-related
42 morbidity and the risk estimate (hazard ratio, HR) between the highest and lowest quintiles of the
43 PRS was 1.67 [95 % confidence interval: 1.52-1.84], $p=3.2*10^{-27}$). In 28,639 participants with
44 comprehensive baseline data from prospective Health 2000 and FINRISK cohorts, 911 incident first
45 alcohol-related events were observed. When adjusted for self-reported alcohol consumption,
46 education, marital status, and gamma-glutamyl transferase blood levels, the risk estimate between
47 the highest and lowest quintiles of the PRS was 1.58 (CI=[1.26-1.99], $p=8.2*10^{-5}$). The PRS was also
48 associated with all-cause mortality with a risk estimate of 1.33 between the highest and lowest
49 quintiles (CI=[1.2-1.47], $p=4.5e-08$) in the adjusted model. In all 39,695 participants with self-
50 reported alcohol consumption available, a 1 SD increase in the PRS was associated with 11.2 g (=0.93
51 drinks) higher weekly alcohol consumption ($\beta=11.2$ [9.85-12.58 g], $p = 2.3*10^{-58}$).

52

53 **Conclusions:** The PRS for alcohol consumption associates for both alcohol-related morbidity and all-
54 cause mortality. These findings underline the importance of heritable factors in alcohol-related
55 behavior and the related health burden. The results highlight how measured genetic risk for an
56 important behavioral risk factor can be used to predict related health outcomes.

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61 Introduction

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63 Alcohol drinking is a major dose-dependent contributor to morbidity and mortality. Globally, 3
64 million annual deaths (5 % of all deaths) result from alcohol consumption, and is also linked to more
65 than 200 disease and injury outcomes.(1) As ethanol is a psychoactive substance with addictive
66 properties,(2) alcohol consumption can lead to the development of alcohol use disorders (AUDs),
67 globally prevalent mental disorders of pathological addictive or abusive drinking patterns, which are
68 linked to worse health outcomes, negative socioeconomic effects, and increased mortality.(3) There
69 is a strong connection between the health burden and the level of alcohol consumed, (4) and in
70 total alcohol has been estimated to be the most damaging of all substances of abuse, in terms of
71 harm caused to self and others.(5)

72

73 Alcohol-related behaviors are also affected by genetic factors and the estimated heritability of
74 alcohol consumption in twin studies has ranged between 35% and 65% (weighted average 37%) (6)
75 and its single nucleotide polymorphism -based heritability has been estimated to be 10%. (7) Recent
76 large-scale genome-wide association studies (GWAS) have identified multiple loci associated with
77 alcohol consumption, underlining the importance of large study populations for unraveling the
78 genetic architecture underlying alcohol-related traits. (7, 8) Similarly, GWAS of alcohol dependence
79 and the Alcohol Use Disorders Identification Test (AUDIT) scores have shown the traits to be
80 genetically distinct but positively correlated.(9,10)

81

82 Polygenic risk scores (PRS) derived from GWAS summary statistics have showcased improved
83 performance in disease prediction. (11) PRSs for known risk factors have also been shown to
84 associate with the related disease (12) and recently associations between multiple risk factor PRSs
85 and related traits were confirmed and reported. (13, 14) However, the link between PRSs for
86 behavioral traits and associated health outcomes remains poorly understood.

87

88 The assessment of potential health risks related to alcohol has so far relied on traditional risk factors,
89 including family history, without explicit measurement of genetic risk. Here we developed a highly
90 polygenic risk score for alcohol consumption and studied whether alcohol-related polygenic burden
91 predicts alcohol-use disorders and other alcohol-related morbidity and mortality in Finnish biobank

92 cohorts (n=96,499) linked to electronic health records. Furthermore, we studied whether the PRS
93 for alcohol consumption predicts alcohol-related outcomes beyond self-reported alcohol
94 consumption and other related risk factors, thus providing more objective information independent
95 of individual reporting bias or temporal fluctuations.

96

97 **Materials and methods**

98

99 **Study sample and definition of alcohol-related morbidity**

100 The data is comprised of 96,499 Finnish individuals from FinnGen Data Freeze 2, which includes
101 prospective epidemiological and disease-based cohorts as well as hospital biobank samples
102 (Supplementary Table 1). The data were linked by the unique national personal identification
103 numbers to national hospital discharge, death, and medication reimbursement registries. Additional
104 details and information on the genotyping and imputation are provided in the online-only
105 Supplementary Information.

106

107 Alcohol-related baseline measures were available for a subset of the FinnGen dataset consisting of
108 national population survey cohorts: FINRISK, collected in 1992, 1997, 2002, 2007 and 2012 and
109 Health 2000, collected in 2000. The baseline data included self-reported information assessed by
110 questionnaires, anthropometric measures, and blood samples. More detailed descriptions of the
111 FINRISK and Health 2000 studies have been published previously.(15,16)

112

113 Additionally, three Finnish twin cohorts, FinnTwin12, NAG-FIN, and Old Twin were pooled and
114 analyzed as one dataset. For these datasets, cohort baseline data was available, but the cohorts
115 were not linked to electronic health records. For details regarding the twin datasets, see the online
116 descriptions (<https://wiki.helsinki.fi/display/twineng/Twinstudy>). (17,18)

117

118 Using nationwide registries for deaths (1969-2016), hospital discharges (1969-2016), outpatient
119 specialist appointments (1998-2016) and drug purchases (1995-2016), we combined 21 somatic and
120 psychiatric alcohol-related diagnoses into a composite disease endpoint, harmonizing the
121 International Classification of Diseases (ICD) revisions 8, 9 and 10, and ATC-codes (**Supplementary**
122 **Table 1**). These registries spanning decades were electronically linked to the cohort baseline data

123 using the unique national personal identification numbers assigned to all Finnish citizens and
124 residents.

125

126 **Genotyping and imputation**

127

128 FinnGen, FINRISK, Health2000, and Finnish Twin Cohort samples were genotyped with Illumina and
129 Affymetrix genomewide SNP arrays. Individuals with non-European ancestry or uncertain sex were
130 excluded. The details about the genotype calling, quality controls and imputation are provided in
131 the Supplementary Information.

132

133 **Polygenic risk scores**

134

135 Summary statistics from the largest existing GWAS meta-analysis on alcohol consumption (8) were
136 used for constructing the PRS. To avoid overfitting, a separate ad hoc meta-analysis was performed
137 by GSCAN, excluding all Finnish and 23andMe samples (n=527,282 after exclusions). The LDpred-inf
138 method (19) was used to account for linkage disequilibrium (LD) among loci, with whole-genome
139 sequencing data on 2,690 Finns serving as the external LD reference panel. The final scores were
140 generated with PLINK2 (20) by calculating the weighted sum of risk allele dosages for each variant.
141 The number of variants in the final scores was 1.1×10^6 .

142

143 **Statistical analysis**

144

145 The Cox proportional hazard model was used to estimate survival curves, hazard ratios (HRs) and
146 95 % confidence interval (95% CI) in the survival analyses where age was used as the time scale. R's
147 cox.zph function was used to test whether the proportional assumption criteria applied in our
148 models. Linear regression in FINRISK and Health 2000 and linear mixed model in the Twin Cohort
149 was used for estimating the relationship between the PRS and alcohol consumption. Logistic
150 regression in the FINRISK and Health 2000 cohorts and linear mixed model in the Twin Cohort was
151 used to estimate the relationship between alcohol abstinence and the PRS.

152

153 Age, sex, genotyping array, and the first ten principal components of ancestry were used as core
154 covariates. Additionally, body mass was used as a covariate in the model estimating the PRS-alcohol

155 consumption relationship. Self-reported weekly average alcohol consumption from the past year
156 (when unavailable, the past week's consumption) was used as the estimate for alcohol
157 consumption. In the fully adjusted survival model analyses, $\log(x+1)$ -transformed alcohol
158 consumption-estimate, current smoking status, binary higher education status, binary
159 marital/cohabitation status and GGT (Gamma Glutamyl Transferase) blood levels at baseline served
160 as covariates.

161

162 In the survival analyses, all prevalent cases and individuals with covariate missingness were
163 excluded. The PRS was normalized and included as a continuous variable in the models. In the
164 survival analysis the highest and lowest genetic risk for alcohol consumption were compared using
165 PRS quintiles.

166

167 In analyses using baseline consumption data, the analyses were performed separately in the Health
168 2000, FINRISK Study, and Twin Cohorts and then meta-analyzed using fixed effects model.

169

170 In risk prediction, FINRISK cohorts with at least 10 years of follow-up (from 1992 to 2002) were used
171 to train the model, and the predictive performance was tested in the Health 2000 cohort. The
172 maximal follow-up window was restricted to 10 years. The change in the predictive performance
173 was assessed by comparing models with and without the PRS using the correlated C-index approach
174 (21) along with calculating the continuous reclassification improvement (NRI) (22) and integrated
175 discrimination improvement (IDI).(23) The Hosmer-Lemeshow goodness-of-fit test was used to test
176 model calibration.

177

178 **Patient and public involvement**

179

180 No patients were involved in setting the research question or the outcome measures, nor were they
181 involved in developing plans for recruitment, design, or implementation of the study. No patients
182 were asked to advise on interpretation or writing up of results. There are no plans to disseminate
183 the results of the research to study participants or the relevant patient community.

184

185

186 **Results**

187

188 **Cohorts**

189 Our primary dataset (FinnGen) is comprised of 96,499 unrelated individuals (54,262 women) with a
 190 total of 55,484,114 person-years of registry-based follow-up and 4,785 first-observed alcohol-
 191 related major health events. Alcohol consumption estimates were available for a total of 39,695
 192 individuals from the prospective cohorts (FINRISK, Health 2000 and Twin Cohort). Two cohorts,
 193 FINRISK and Health 2000, have full registry data and information on self-reported alcohol
 194 consumption and related baseline data and consist of 28,639 individuals (94.5% of the participants
 195 after excluding 964 prevalent alcohol-related morbidity cases), with 424,053 person-years of
 196 registry-based follow-up and 988 first ever alcohol-related events. The interview-based DSM-IV
 197 AUD-status was available in a subset of the Twin cohort for 713 cases and 1460 controls.

198

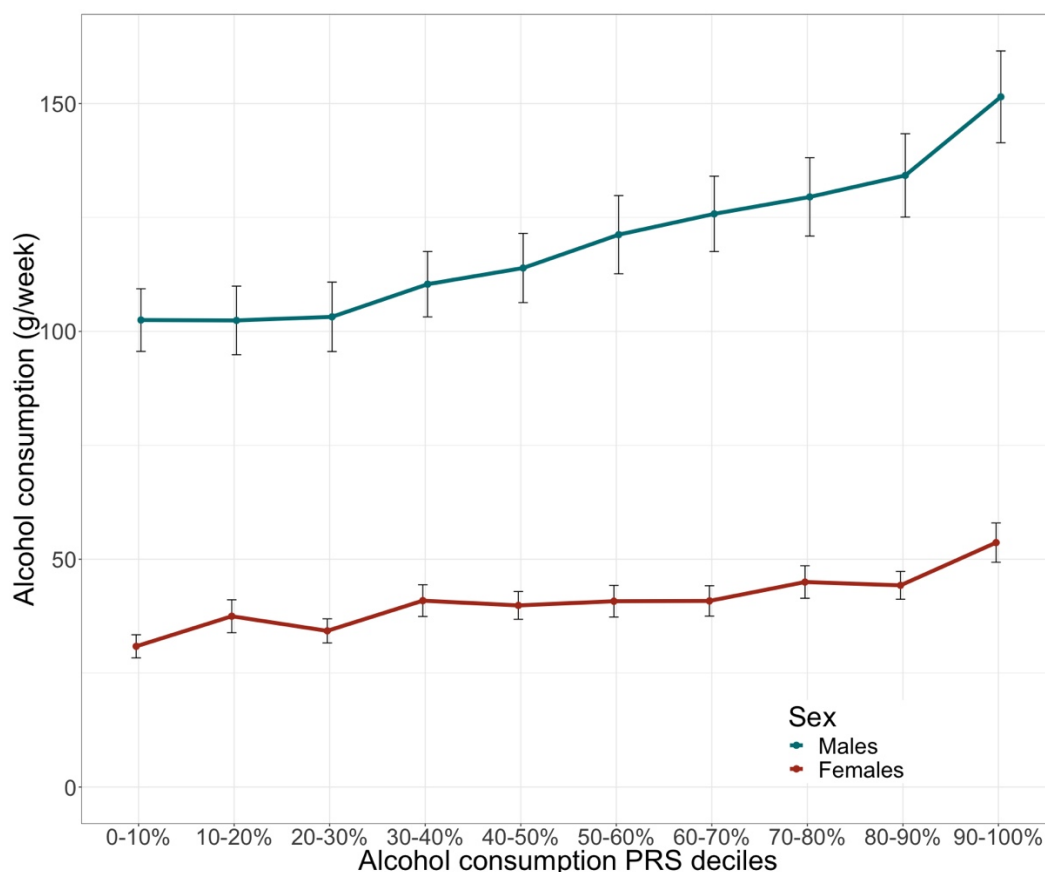
199 **Table 1.** Population characteristics of FinnGen, FINRISK, Health 2000, and Twin Cohort -datasets.
 200

	FinnGen	FINRISK	Health 2000	Twin Cohort
N	96,499 (54,262 women)	23,824 (12,513)	5,945 (3,260)	9,926 (5,036)
N (incident events)	4,785	817	171	NA
Age (years)	57.5(end of follow-up)	48.6 (baseline)	54.0 (baseline)	49.3 (baseline)
Alcohol drinking (g/week)	NA	76.5	73.9	84.8
Non-drinkers	NA	2,874(12%)	1,298 (22%)	30(0.3%)
Current smokers	NA	5,929 (25%)	1,538 (26%)	3,478 (35%)
Higher education	NA	8,612 (36%)	1,721 (28%)	1,192 (12%)
Marriage or co-habitation	NA	17,468 (73%)	4,151 (68%)	6,661 (67%)
GGT (U/I)	NA	33.8	36.6	NA

201 **Alcohol consumption**

202

203 In a meta-analysis of the three cohorts with alcohol consumption estimates available (n=39,888),
204 the PRS for alcohol consumption was strongly associated with self-reported alcohol consumption.
205 A one SD increase in the PRS was associated with an 11.2 g (= 0.93 drinks á 12g) increase in weekly
206 pure alcohol intake (beta=11.2 [9.85-12.6 g], $p = 2.3 \cdot 10^{-58}$) (**Fig 1, cohort-specific figures in the**
207 **supplementary material**). Adding the PRS to the model improved r^2 by ~0.6 percentage points (from
208 9.17 % to 9.80 %). In addition, the PRS was negatively associated with alcohol abstinence (reported
209 alcohol consumption 0). In FINRISK and Health2000, a 1 SD increase in the PRS for alcohol
210 consumption was associated with a 13.7% reduced likelihood of being a nondrinker (OR=0.863
211 [0.833-0.895], $p=6.1 \cdot 10^{-16}$) while this was not case in the Twin Cohort where there were only 30
212 nondrinkers (OR=0.999[0.998-1.00] $p=0.31$). Cohort-specific figures for sex-specific alcohol
213 consumption are in the supplementary material.



214

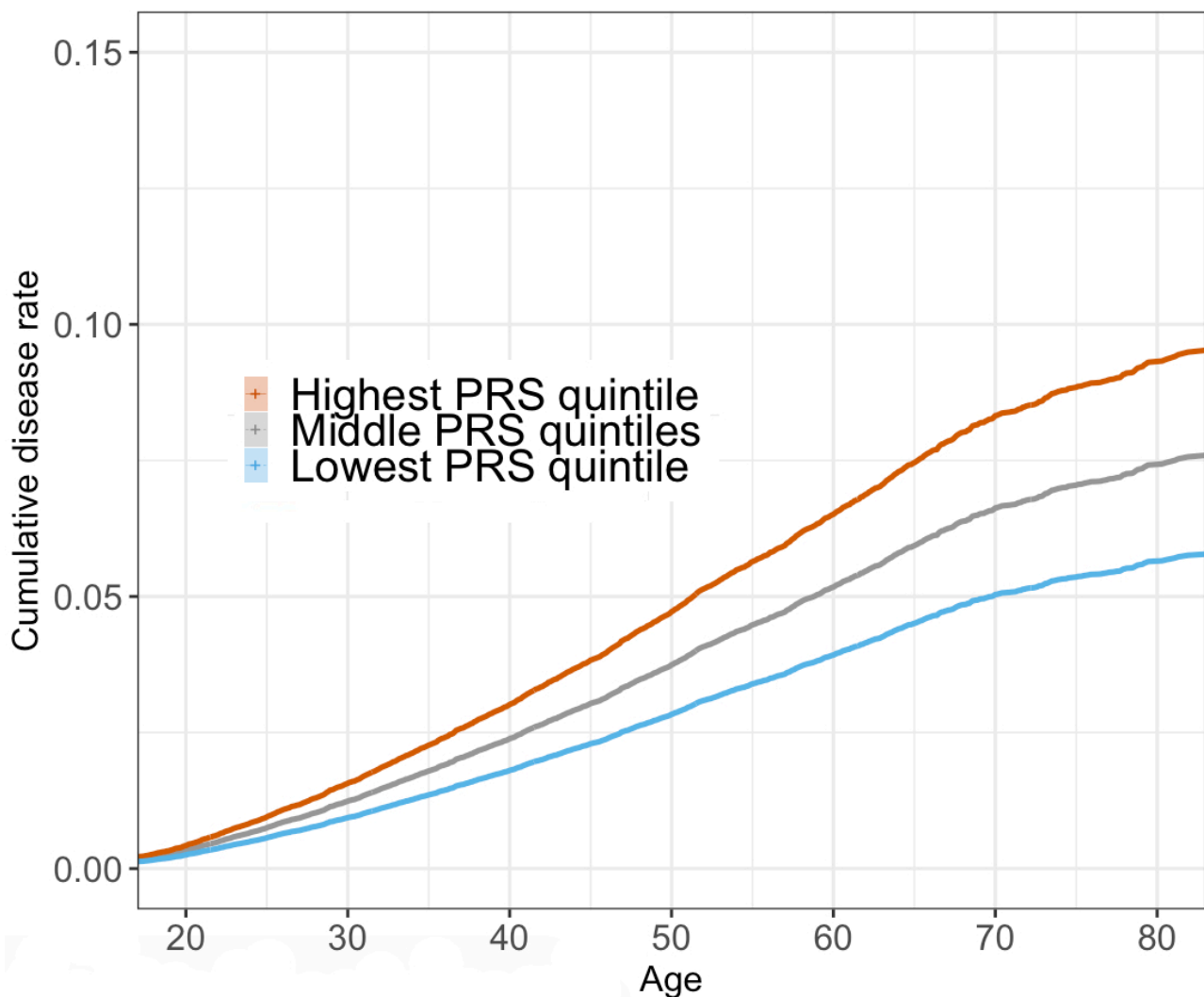
215

216 **Figure 1.** Alcohol drinking (g/week) for the deciles of the alcohol consumption polygenic risk score
217 shown for males (n= 18,887) and females (n= 20,808) with 95% confidence interval error bars
218 (n=39,695)

219

220 Alcohol-related morbidity

221 The PRS for alcohol consumption was strongly associated with increased risk for lifelong major
222 alcohol-related events derived from electronic health-records in the FinnGen dataset (n=96,499,
223 cases = 4,785) (**Figure 2**). The difference in the risk for alcohol-related morbidity events between
224 the lowest and highest risk quintiles in the PRS was 67 % (HR=1.67 [1.52-1.84], $p=3.2 \times 10^{-27}$) and a 1
225 SD increase in the PRS was associated with a 21 % increase in risk (HR=1.21 [1.18-1.25], $p=1.4 \times 10^{-40}$).
226 The association was similar in both males (HR=1.21 [1.17-1.26], $p=1.8 \times 10^{-28}$) and females
227 (HR=1.22 [1.16-1.29], $p=4.3 \times 10^{-14}$).



228

229 **Figure 2.** The FinnGen dataset was divided into three groups consisting of the lowest quintile, three
230 middle quintiles and the highest quintile of the alcohol consumption PRS. The cumulative disease
231 rate of alcohol-related morbidity is displayed as a function of age (n=96,499).

232 In the cohorts where alcohol consumption estimates and other related baseline data were available
233 at the cohort entry time, the PRS was associated with an increased risk of incident major alcohol-
234 related events and the association was maintained also in the fully adjusted model (n = 28,639,
235 cases = 911). In a meta-analysis of the two cohorts 1 PRS SD was associated with a 26% increased
236 risk of incident alcohol-related events when the consumption-estimate was not in the model
237 (HR=1.26 [1.18-1.34], $p=1.1*10^{-12}$) and with a 15 % increase when alcohol consumption was in the
238 model (HR=1.15 [1.08-1.22], $p=2.1*10^{-5}$). In a fully adjusted model, including marital status,
239 education, smoking status and GGT, the estimate were unchanged (HR=1.15 [1.08-1.22], $p=2.0*10^{-5}$) (Table 2).

241

242 **Mortality**

243

244 We observed a similar increase in the risk of alcohol-related and all-cause mortality. In FinnGen with
245 7,249 deaths one SD increase in the PRS for alcohol consumption was associated with 7 % increase
246 in the risk of death (HR= 1.07 [1.05-1.10], $p = 4.3*10^{-9}$). The risk estimate between the highest and
247 lowest 20 % in the PRS was 1.23 (HR = 1.23 [1.14-1.23], $p = 1.22*10^{-8}$). In our prospective cohorts,
248 with cause-of-death information available, 4,125 deaths were recorded (**Table 2**). For all-cause
249 mortality there was 11 % increase in the risk of death per 1 PRS SD in the basic model (HR=1.11[1.07-
250 1.14], $p = 3.2*10^{-10}$) and 9 % in the fully adjusted model (HR = 1.09 [1.06-1.12], $p=1.1*10^{-7}$). The risk
251 difference between the highest and lowest quintiles of the PRS was 33 % (HR=1.33 [1.2-1.47],
252 $p=4.5e-08$) in the fully adjusted model.

253

254 Of the 4,125 deaths 335 were known to be alcohol-related. Without alcohol consumption in the
255 model, the increase in alcohol-related mortality was 26 % per 1 PRS SD (HR=1.26 [1.13-1.4],
256 $p=3.7*10^{-5}$). When alcohol consumption was included in the model, the increase was 13 % (HR=1.13
257 [1.01-1.26], $p=0.027$) and in a model with all co-variates, 11 % (HR=1.11 [0.996-1.24], $p=0.058$)
258 (**Table 2**). Similarly, the PRS was associated with a higher risk of death from other than alcohol-
259 related causes (n=3,790) when fully adjusted for all covariates (HR=1.08 [1.05-1.12], $p=1.4*10^{-6}$).

260

261

262

263 **Table 2.** Cohort specific and meta-analyzed associations between the alcohol consumption polygenic
 264 risk score and alcohol-related **a)** morbidity and **b)** mortality. In the fully adjusted model age, sex,
 265 alcohol consumption, smoking, education, marital status and GGT (U/l) were used as non-genetic
 266 covariates.
 267

	FINRISK	Health 2000	Meta-analysis
a) Alcohol-related morbidity	Cases=817	Cases=171	Cases=988
Basic model with age and sex	HR=1.25 [1.16-1.34], p=5.9*10 ⁻¹⁰	HR=1.32 [1.13-1.53], p=0.00036	HR=1.26 [1.18-1.34], p=1.1*10 ⁻¹²
Model with alcohol consumption	HR=1.13 [1.06-1.21], p=0.00053	HR=1.23 [1.06-1.43], p=0.0081	HR=1.15 [1.08-1.22], p=2.1*10 ⁻⁵
Fully adjusted model	HR=1.14 [1.06-1.22], p=0.00027	HR=1.20 [1.03-1.4], p=0.022	HR=1.15 [1.08-1.22], p=2.0*10 ⁻⁵
b) Alcohol-related mortality	Deaths=264	Deaths=71	Deaths=335
Basic model with age and sex	HR=1.21 [1.07-1.37], p=0.0022	HR=1.41 [1.11-1.8], p=0.0045	HR=1.25 [1.12-1.4], p=5.9*10 ⁻⁵
Model with alcohol consumption	HR=1.08 [0.952-1.22], p=0.24	HR=1.34 [1.05-1.71], p=0.017	HR=1.13 [1.01-1.26], p=0.033
Fully adjusted model	HR=1.08 [0.957-1.23], p=0.21	HR=1.22 [0.965-1.55], p=0.096	HR=1.11 [0.996-1.24], p=0.058

268

269

270 DSM-IV Alcohol-use disorder

271 The PRS was also associated with an interview-based DSM-IV alcohol use disorder diagnosis in the
 272 Nicotine Addiction Genetics Family cohort (440 cases, 1,140 controls) and a subset of FinnTwin16
 273 cohort (273 cases, 320 controls). A meta-analysis of the two cohorts (713 cases) resulted in a
 274 combined 20 % increase in the prevalence of AUD per 1 PRS SD (OR = 1.20 [1.11-1.31], p = 2.29*10⁻⁵)
 275 in the unadjusted model. Adjusting for marital status, education and smoking explained part of

276 the effect (OR=1.14 [1.02-1.28], p=0.023) and further adjusting with maximal amount of drinks
277 taken explained most of the effect (OR=1.06 [0.94-1.19], p=0.35).

278

279 **Prediction**

280 The predictive performance of the PRS was evaluated in the Health 2000 cohort (5,732 complete
281 cases, 110 events) with a follow-up-window of 10 years based on the Cox model trained in the
282 FINRISK cohort (18427 complete cases with ≥ 10 years of follow-up, 628 events). In a model not
283 including the alcohol consumption estimate, adding the PRS to the model increased the C-index by
284 0.020, from 0.69 to 0.71 ($p = 0.017$). Both IDI (0.00242 [0.00102-0.00383], $p = 7.3 \times 10^{-4}$) and NRI
285 (0.335 [0.146-0.523], $p = 5.1 \times 10^{-3}$) shifts were positive and statistically significant. When the log-
286 transformed alcohol consumption estimate was included, a modest improvement of prediction was
287 observed (C-index=0.0022 from 0.812 to 0.814, p-value=0.30; NRI=0.308 [0.119-0.497], $p = 0.0014$
288 and IDI=0.00173 [0.000726-0.00305], $p = 0.017$). Similarly, a modest gain was observed when adding
289 PRS to a model with all available covariates including also marital status, education status, smoking
290 status and GGT (C-index=0.00183 from 0.847 to 0.849, $p=0.44$; NRI=0.235 [0.0461-0.423], $p=0.015$;
291 IDI =0.00331[0.0000254-0.00659], $p = 0.048$).

292

293 **Discussion**

294 We developed a highly polygenic risk score for alcohol consumption by obtaining weights from a
295 recently published large-scale discovery sample and showed that the PRS was strongly associated
296 with alcohol consumption in independent biobank cohort samples. An increased polygenic burden
297 for alcohol consumption was associated with higher incidence of major alcohol-induced health
298 events. The associations remained significant when we accounted for self-reported alcohol
299 consumption and other relevant covariates; in a fully adjusted model the relative risk-estimate
300 between the highest and lowest quintiles of the polygenic risk score was 1.6. Furthermore, the PRS
301 was also associated with both alcohol-related, non-alcohol related and all-cause mortality.

302

303 **Comparison with other studies**

304

305 Our PRS shows the utility of genetic information for prediction of alcohol-related harm. The PRS,
306 developed from a genetic analysis of cross-sectional self-reported alcohol consumption, was

307 associated with future risk of major alcohol-related health events. While a large number of PRSs
308 have already been established for various traits and diseases (11), the development of PRSs for
309 behavioral traits, such as substance use, has until now been limited (24-27) and the studies have
310 not assessed their impact on future major health events.

311

312 **Implications**

313

314 Our results show that using a large sample size with long follow-up, we were able to build a PRS of
315 alcohol consumption that is associated not only with alcohol consumption in independent samples,
316 but also with future incident alcohol-related health events. In line with the knowledge that alcohol
317 consumption is a major contributor to the worldwide burden of death, especially among working-
318 age adults (1), we found the PRS to be associated also with all-cause mortality, further highlighting
319 the importance of alcohol drinking as a cause of premature death.

320

321 Our score provides a genetic basis for potentially identifying a subset of high-risk individuals even
322 early on in life, with potential for more targeted prevention of AUDs and other alcohol-related
323 morbidity. Prevention is a cost-effective and efficient strategy to reduce alcohol related harms (28)
324 and it is labeled one of the United Nations main health-related worldwide strategies of sustainable
325 development (<https://sustainabledevelopment.un.org/sdg3>). A higher genetic predisposition for
326 alcohol-related harms was detected both in the presence and absence of alcohol consumption data,
327 as our PRS predicted alcohol-related harms beyond self-reported alcohol consumption. Health
328 services are encouraged to support initiatives for screening and brief interventions for harmful
329 drinking (29) as an effective strategy for tackling alcohol-related harm.(30) Thus, genetic
330 information could potentially be used to improve the arsenal of possible strategies to detect high-
331 risk individuals for targets of brief interventions. The fact that individuals in the highest PRS quintile
332 showed an elevated risk for alcohol-related health events even in fully adjusted models could justify
333 the use of genetic information even in clinical settings where a detailed history of alcohol
334 consumption estimates, AUDIT-scores, or similar information are attainable. Communicating the
335 information of higher risk for alcohol-related harm to patients could serve as a motivator for
336 reducing drinking or committing to abstinence. The effect of being aware of one's negative
337 expectations could also be unwanted, as is thought to be the case in the stereotype threat
338 phenomenon. (31-34)

339

340 Self-reported alcohol consumption is known to be biased and problematic in terms of reliability and
341 validity for predicting alcohol-related risks (35,36). Some bias derives from true measurement error,
342 but another source is the lifelong temporal fluctuation of alcohol-drinking patterns not captured by
343 a measure at one single timepoint. Our PRS was associated with alcohol-related harms even when
344 adjusting for self-reported alcohol consumption estimate. One potential reason for this is that the
345 PRS contains information from the latent genetic predisposition for alcohol consumption, thus
346 overriding both the true measurement error and temporal fluctuations in alcohol drinking volume.

347

348 It has been hypothesized that alcohol consumption-based genetic discovery might inform more
349 about low-level drinking than about problematic drinking and AUDs.(37) However, we built a
350 polygenic risk score for alcohol consumption and successfully used it to predict alcohol-related
351 harms. Due to the robustness of a self-reported single timepoint alcohol consumption estimate and
352 the fact that different alcohol-related traits are to some degree genetically distinct,(9,10) it is
353 expected that a PRS developed directly for alcohol-related morbidity will outperform our PRS in
354 predicting alcohol-related health burden. Supporting this assumption, the general pattern is that
355 PRSs are more strongly associated with their respective diseases than with related phenotypes.
356 (13,14) Unfortunately, no high-quality summary statistics for alcohol-related harms including both
357 somatic and psychiatric outcomes yet exist; the performed GWAS have only covered alcohol
358 dependence (10) and been far smaller in size than our discovery sample of choice, thus making
359 future efforts for large-scale GWAS discovery based on alcohol-related harms more than necessary.

360

361 **Strengths and limitations**

362

363 Our polygenic risk score was derived using European ancestry discovery samples and tested in the
364 Finnish population. Its applicability in other populations therefore needs further evaluation as the
365 alcohol-related genetic mechanisms may vary between populations. However, it has to be noted
366 that the PRS derived from a non-Finnish sample performed well in the Finnish dataset, even though
367 Finns are somewhat genetically different from the rest of the Europeans (38).

368

369 Our design allowed us to study outcomes prospectively. Our registry-based follow-up captures
370 alcohol-related outpatient and inpatient visits, withdrawal treatment prescription for alcoholism,

371 and deaths, thus covering major alcohol-related health events over several decades. Nonetheless,
372 some of the milder cases of alcohol-related health problems could have gone undetected.

373

374 **Conclusions**

375

376 In conclusion, a polygenic risk score for alcohol consumption was associated with elevated risk for
377 incident alcohol-related health events and all-cause mortality. These findings underline the
378 importance of heritable factors driving alcohol-related behavior. A successful attempt to predict
379 alcohol-related health outcomes with a polygenic risk score shows promise in possible future
380 utilization of genetic information in risk estimation and prediction of alcohol-related harms.

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397 **Footnotes**

398

399 **Contributors**

400 All authors contributed to the study concept and design, analysis and interpretation of the data, as
401 well as to the critical revision of the manuscript for important intellectual content or additionally to
402 data acquisition. GSCAN provided the GWAS summary statistics. TK, TP, NJM and JKa performed the
403 statistical analyses with support from SRi, ASH, JKo and SRu. ASH connected the data to the
404 registries. TK drafted the manuscript. The corresponding author attests that all listed authors meet
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406

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425 The study was approved by the Coordinating Ethical Committee of the Helsinki and Uusimaa
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434 **Competing financial interests**

435 All authors have completed the ICMJE uniform disclosure form at
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