

# Biased competition in semantic representation during natural visual search

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## 1 **Abstract**

2 Humans divide their attention among multiple visual targets in daily life, and visual search gets more difficult as  
3 the number of targets increases. The biased competition hypothesis (BC) has been put forth as an explanation for  
4 this phenomenon. BC suggests that brain responses during divided attention are a weighted linear combination of  
5 the responses during search for each target individually. Furthermore, this combination is biased by the intrinsic  
6 selectivity of cortical regions. Yet, it is unknown whether attentional modulations of semantic representations of  
7 cluttered and dynamic natural scenes are consistent with this hypothesis. Here, we investigated whether BC accounts  
8 for semantic representation during natural category-based visual search. Human subjects viewed natural movies, and  
9 their whole-brain BOLD responses were recorded while they attended to “humans”, “vehicles” (i.e. single-target  
10 attention tasks), or “both humans and vehicles” (i.e. divided attention) in separate runs. We computed a voxelwise  
11 linearity index to assess whether semantic representation during divided attention can be modeled as a weighted  
12 combination of representations during the two single-target attention tasks. We then examined the bias in weights of  
13 this linear combination across cortical ROIs. We find that semantic representations during divided attention are linear  
14 to a substantial degree, and that they are biased toward the preferred target in category-selective areas across ventral  
15 temporal cortex. Taken together, these results suggest that the biased competition hypothesis is a compelling account  
16 for attentional modulations of semantic representation across cortex.

## 17 **Significance Statement**

18 Natural vision is a complex task that involves splitting attention between multiple search targets. According to the  
19 biased competition hypothesis (BC), limited representational capacity of the cortex inevitably leads to a competition  
20 among representation of these targets and the competition is biased by intrinsic selectivity of cortical areas. Here  
21 we examined BC for semantic representation of hundreds of object and action categories in natural movies. We  
22 observed that: 1) semantic representation during simultaneous attention to two object categories is a weighted linear  
23 combination of representations during attention to each of them alone, and 2) the linear combination is biased toward  
24 semantic representation of the preferred object category in strongly category-selective areas. These findings suggest

25 BC as a compelling account for attentional modulations of semantic representation across cortex in natural vision.

## 26 **Introduction**

27 In daily life, humans frequently search for a multitude of objects in their visual environment. Yet, attending to multiple  
28 objects becomes more difficult as the number of targets increases. Psychophysical studies showed that reaction time  
29 and error rate during visual search systematically increase with growing number of items to be attended (Eckstein  
30 et al., 2000; Wolfe, 2012; Reynolds and Chelazzi, 2004; Luck et al., 1997). The biased competition hypothesis  
31 (BC) has been proposed to account for this performance decline (Duncan, 1984). BC reasons that the brain has  
32 limited representational capacity. Thus, simultaneous search for multiple visual objects results in a competition among  
33 their representations. Moreover, task demands due to spatial (Keitel et al., 2013; Kastner et al., 1998), feature-based  
34 (McMains and Kastner, 2011; Bichot et al., 2005; Boynton, 2005), and object-based (Gentile and Jansma, 2010;  
35 Reddy et al., 2009) attention can bias this competition in favor of the target (Desimone, 1998).

36 Several neuroimaging studies provided evidence for competition among cortical representations of multiple objects  
37 across visual cortex in the absence of specific task demands (Kastner et al., 1998; MacEvoy and Epstein, 2009;  
38 Gentile and Jansma, 2010; Nagy et al., 2011; Baeck et al., 2013; Jeong and Xu, 2017). Gentile and Jansma (2010)  
39 measured average BOLD responses in fusiform face area (FFA) while subjects viewed a single face image or a pair  
40 of face images. Response to a pair of faces was lower than the summation of the responses when each of the faces  
41 were presented individually. Similarly, Nagy et al. (2011) presented four equispaced isolated images of faces or noise  
42 images. The number of face images was increased systematically from zero to four. They reported that responses in  
43 FFA and lateral occipital complex (LOC) to multiple faces were lower than the summation of the responses to individ-  
44 ual faces. MacEvoy and Epstein (2009) suggested a linear model of representational competition among objects. The  
45 authors presented either a single or a pair of isolated images of objects from four categories (shoes, chairs, cars, or  
46 brushes). Using multivoxel pattern analysis, they showed that the response pattern in object-selective areas in ventral  
47 temporal cortex when subjects viewed pairs of objects can be approximated by the mean of response patterns when  
48 they viewed each of the objects in isolation. These results were interpreted to imply the existence of competition

49 among representations.

50 Recent studies also provided evidence for top-down influences in BC during divided attention to multiple objects  
51 (Reddy et al., 2009; Gentile and Jansma, 2010). Reddy et al. (2009) studied BOLD responses while subjects attended  
52 to a single or a pair of object categories among four alternatives (faces, houses, shoes, or cars). A multivoxel pattern  
53 analysis in category-selective areas in ventral temporal cortex revealed that the response pattern during divided at-  
54 tention to two object categories was a weighted linear combination of response patterns while attending to individual  
55 targets. Furthermore, the authors reported that in PPA (preferentially responsive to houses) and in FFA (preferentially  
56 responsive to faces), combination weights were biased toward the preferred object category. Similarly, Gentile and  
57 Jansma (2010) reported that during attention to one of the faces in a pair of face images, FFA responses were biased  
58 toward the responses recorded when the target was presented in isolation.

59 Previous studies on BC have provided evidence for competition in representation of isolated static objects. Yet,  
60 natural scenes are intrinsically dynamic and cluttered with many objects and actions. It has recently been suggested  
61 that thousands of visual categories are embedded in a continuous semantic space across cortex (Huth et al., 2012), and  
62 that category-based visual search causes broad modulations in these semantic representations (Çukur et al., 2013). It  
63 is currently unknown whether BC can account for modulations in semantic representation during natural visual search  
64 for object categories.

65 Here we questioned whether BC can account for semantic representation of object-action categories during divided  
66 attention in natural visual search. To address this question, we conducted a functional magnetic resonance imaging  
67 (fMRI) experiment (Fig.1). Five human subjects viewed 72 min of natural movies while performing three separate  
68 tasks in different runs: attend to “humans”, attend to “vehicles”, and attend to “humans and vehicles” (i.e., divided  
69 attention). Whole-brain BOLD responses were recorded and category responses for 831 objects and actions were  
70 measured separately for each task and each individual subject (Nishimoto et al., 2011). Semantic tuning was estimated  
71 via principle component analysis on the response profiles. To test whether the semantic tuning during divided attention  
72 can be approximated by a weighted linear combination of the tuning during single-target tasks, ordinary least squares

73 was used among semantic tuning profiles for the three tasks. To reveal the interactions between the attentional bias in  
74 semantic tuning and the intrinsic selectivity of brain areas, the semantic tuning distribution during divided attention  
75 was regressed onto tuning distributions during the two single-target tasks.

## 76 **Materials and Methods**

77 **Subjects.** Five healthy adult volunteers (four males, one female) with normal or corrected-to-normal vision partic-  
78 ipated in this study: S1 (age 32), S2 (age 28), S3 (age 28), S4 (age 28), S5 (age 28). Data were collected at the  
79 University of California, Berkeley. The main experiment contained three sessions. Functional localizers were col-  
80 lected in two sessions. The protocols for these experiments were approved by the Committee for the Protection of  
81 Human Subjects at the University of California, Berkeley. Written informed consent was obtained from all subjects  
82 before scanning.

83 **Stimuli.** Continuous natural movies were used as the stimulus in the main experiment. Three 8 min movies were  
84 compiled from short clips. Movie clips were selected from a wide variety of sources as detailed in Nishimoto et  
85 al. (2011). High-definition movie frames were cropped into a square frame and downsampled to  $512 \times 512$  pixels,  
86 covering a  $24^\circ \times 24^\circ$  field of view. Subjects were directed to fixate on a color square of  $0.16^\circ \times 0.16^\circ$  at the center.  
87 The color of the fixation dot was changing at 1 Hz to ensure visibility. The stimulus was presented at a rate of 15 Hz  
88 using an MRI-compatible projector (Avotec) and a custom-built mirror arrangement.

89 **Experiment design.** The main experiment was performed in a single session consisting of nine 8 min runs. Four  
90 mutually exclusive classes of stimuli (only humans, only vehicles, both humans and vehicles, and neither humans  
91 nor vehicles) were randomly interleaved and evenly distributed within and across the runs. Subjects were instructed  
92 to covertly search for the target categories in the movies. A cue word was displayed before each run to indicate the  
93 attention task: “humans”, “vehicles”, or “humans and vehicles”. In the attend to humans task, subjects searched for  
94 human categories (e.g. woman, man, boy). For the attend to vehicles task, subjects searched for vehicle categories  
95 (e.g. car, truck, bus). Subjects searched for targets from either of the human or vehicle categories in the divided

96 attention task. To minimize subject expectation bias, the order of search tasks was interleaved across runs. To maintain  
97 vigilance, subjects were asked to press a button when they detected a target on the screen. BOLD responses were  
98 recorded from the whole brain. To minimize the effect of transient confounds, data from the first 20 seconds and the  
99 last 30 seconds of each run were discarded. These procedures resulted in 690 data samples for each attention task.

100 **Definition of functional areas.** Functional regions of interest (ROIs) were identified in individual subjects using  
101 functional localizers (Huth et al., 2012). Localizer experiments for category-selective areas (fusiform face area, FFA;  
102 extrastriate body area, EBA; parahippocampal place area, PPA; retrosplenial cortex, RSC; lateral occipital complex,  
103 LOC) were performed in six 4.5 min runs of 16 blocks (Huth et al., 2012). Subjects passively viewed 20 random static  
104 images from one of the objects, scenes, body parts, faces, or spatially scrambled objects groups in each block. Each  
105 image was shown for 300 ms following a 500 ms blank period. Scene-selective ROIs (PPA, RSC) were identified as  
106 voxels with positive scene versus objects contrast ( $t$ -test,  $p < 10^{-4}$ , uncorrected). FFA, EBA, and LOC were defined  
107 using face-versus-object, body-part-versus-object and object-versus-scrambled-object contrasts respectively ( $t$ -test,  
108  $p < 10^{-4}$ , uncorrected). Localizer experiment for attentional-control areas (intraparietal sulcus, IPS; frontal eye field,  
109 FEF; frontal operculum, FO) contained one 10 min run. In each 20 sec block, either a self-generated saccade task  
110 or a resting task was prescribed (Corbetta et al., 1998). The run contained 30 blocks. Attentional-control areas were  
111 localized using saccade-versus-rest contrast ( $t$ -test,  $p < 10^{-4}$ , uncorrected). ROIs were refined to voxels with contrast  
112 level more than half of the maximum near a 2 mm neighborhood of the cortical surface.

113 **MRI protocols.** Data were collected using a 3T Siemens Tim Trio MRI scanner (Siemens Medical Solutions) using  
114 a 32-channel receiver coil. Functional data were collected using a T2\*-weighted gradient-echo echo-planar-imaging  
115 pulse sequence with the following parameters: TR = 2 sec, TE = 33 msec, water-excitation pulse with flip angle =  
116 70°, voxel size = 2.24 mm × 2.24 mm × 4.13 mm, field of view = 224 mm × 224 mm, 32 axial slices. To construct cor-  
117 tical surfaces, anatomical data were collected using a three-dimensional T1-weighted magnetization-prepared rapid-  
118 acquisition gradient-echo sequence with the following parameters: TR = 2.3 sec, TE = 3.45 msec, flip angle = 10°,  
119 voxel size = 1 mm × 1 mm × 1 mm, field of view = 256 mm × 212 mm × 256 mm.

120 **Data preprocessing.** Functional images collected in the main experiment were motion corrected. Using the SPM12  
121 software package (Friston et al., 1995), the functional images were aligned to the first image from the first session of  
122 the main experiment. Three subjects that participated in this study were common with a previous study (Çukur et al.,  
123 2013). Functional images of these subjects were aligned to the reference images from that previous study. Non-brain  
124 tissues were removed using the brain extraction tool (BET) from the FSL software package (Smith, 2002). Within  
125 each run, low-frequency drifts were removed from BOLD responses in each voxel using a second order Savitzky-  
126 Golay filter over a 240 sec temporal window. The resulting voxelwise time series were z-scored to attain zero mean  
127 and unity variance.

128 **Voxelwise category model.** For each voxel, response profiles were determined by fitting category models that rep-  
129 resented hundreds of objects and actions in natural movies. As a first step, each 1sec clip of the movie stimulus was  
130 manually labeled for presence of 831 distinct object and action categories (Huth et al., 2012). Presence of superordi-  
131 nate categories was inferred from the terms in the WordNet lexicon, a lexical database that groups words based on their  
132 semantic relationships (Miller, 1995). This procedure yielded time courses for 831 model features (i.e. categories).  
133 Each time course was then downsampled to 0.5 Hz to match the acquisition rate of fMRI. Separate finite impulse  
134 response (FIR) filters were used for each model feature to capture the hemodynamic response. Filter delays were set  
135 to 4, 6, and 8 secs. This is equivalent to concatenating feature vectors that are delayed by two, three, and four samples.  
136 To prevent head-motion and physiological noise confounds, estimates of these nuisance factors were regressed out  
137 of the BOLD responses. Six affine motion time courses estimated during the motion-correction stage were taken  
138 as the head-motion regressors. Two regressors to capture respiration and nine regressors to capture cardiac activity  
139 were estimated using the data collected via a pulse oximeter and a pneumatic belt during the main experimental runs  
140 (Verstynen and Deshpande, 2011).

141 To reduce spurious correlations between model features and global motion-energy of the movie stimulus, a nuisance  
142 regressor was included that reflected the total motion-energy. The motion-energy time course was formed by tak-  
143 ing the mean motion-energy in each one second movie clip. Movie frames were transformed into the International

144 Commission on Illumination LAB color space, and the luminance channel was extracted. The luminance was then  
145 passed through the motion-energy filter bank. The motion-energy filter bank contained 2139 Gabor filters. Filters  
146 were computed at eight directions (0 to 315°, in 45° steps), three temporal frequencies (0, 2, and 4 Hz) and six spatial  
147 frequencies (0, 1.5, 3, 6, 12, and 24 cycles/image). Filters were placed on a square grid spanning the 24° × 24° field of  
148 view. Finally, the motion-energy time course was assessed by squaring and summing outputs of quadrature filter pairs,  
149 and the results were passed through a logarithm compressive nonlinearity and temporally downsampled to match the  
150 fMRI acquisition rate (Nishimoto et al., 2011).

151 To account for potential correlations between target detection and BOLD responses, a target-presence regressor was  
152 included in the model. The target-presence regressor contained category regressor for “person” during attend to  
153 humans task and the category regressor for “conveyance” during attend to vehicles task. The target-presence regressor  
154 during divided attention task contained the binary union of the “person” and “conveyance” category regressors. The  
155 described regressors were aggregated and used as the stimulus matrix.

156 **Model fitting and testing.** Voxelwise models were fit using regularized linear regression with an  $\ell_2$  penalty to avoid  
157 overfitting. To prevent bias, model fitting for the three attention tasks was performed concurrently. To do this, the  
158 stimulus and BOLD response matrices were aggregated across tasks (Fig.2). Note that this procedure ensures that the  
159 same regularization parameter will be used in each voxel across the three tasks. Furthermore, using the aggregated  
160 stimulus matrix enables employing the target regressor.

161 A nested cross-validation (CV) procedure was used to estimate response profiles for each voxel. Data were segmented  
162 into 58 25sec blocks. In each of the 20 outer folds, 6 blocks were randomly held-out as validation data and the  
163 remaining blocks were used for parameter optimization and fitting models on the inner folds. In each of the 20  
164 inner folds, blocks were randomly shuffled and split to 40 blocks as training data and 12 blocks as test data. Models  
165 were fit on the training data for regularization parameters in the range  $[2^{-3}, 2^{20}]$ . Using the weights found for each  
166 regularization parameter, responses were predicted for the test data. Prediction scores were separately computed for  
167 each voxel, taken as the Pearson’s correlation between actual and predicted responses. Prediction scores were then



168 averaged across the inner CV folds. Regularization parameters maximizing the average prediction score were selected  
169 in each voxel. Nuisance regressors were discarded from further analyses. Afterwards, optimized parameters were used  
170 to fit models on the union of training and test data in each outer fold, yielding category response profiles. To assess  
171 model performance, responses were predicted for the validation data using the fit models and prediction scores of each  
172 voxel were averaged across the attention tasks. Finally, response profiles and prediction scores for each voxel were  
173 averaged across the outer folds. Model fitting was performed using custom-written software in Matlab (MathWorks  
174 MA).

175 **Significance of the fit models.** Significance of the estimated category response profiles was assessed using a boot-  
176 strapping procedure on the validation data held-out in the model fitting stage. A 20-fold CV procedure was imple-  
177 mented to assess significance of the fit models. In each fold, the corresponding fit models and validation data from the  
178 model fitting stage were used. Responses were predicted for the validation data using the fit models. Prediction score  
179 was calculated in each fold and averaged across CV folds to get the mean score. In each fold, predicted responses  
180 were resampled 500 times with replacement. We assumed the null hypothesis that attention does not alter response  
181 profiles of voxels. Thus, under the null hypothesis, response profiles would be the same across the three attention tasks  
182 (Çukur et al., 2013). To get the prediction score distribution under the null hypothesis, resampled predicted responses  
183 were shuffled across the attention tasks, and prediction scores were computed. The  $p$ -value was taken as the fraction  
184 of samples for which the average prediction score across CV folds was lower than that under the null hypothesis.

185 **Semantic representation of objects and actions.** To obtain a quantitative description of the individual subjects'  
186 semantic spaces, principal component analysis (PCA) was performed on estimated response profiles in individual  
187 subjects (Huth et al., 2012). To prevent bias, PCA was performed on response profiles for the two single-target  
188 attention tasks pooled together. The collection of principal components (PCs) that described at least 90% of the  
189 variance in response profiles was selected. This resulted in  $L \in [36, 47]$  PCs for the five subjects. The semantic  
190 representation of individual object or action categories can then be assessed in individual subjects by projecting the  
191 response profiles onto the PCs.

192 **Linearity of semantic tuning during divided attention.** To test whether semantic tuning during divided attention  
193 could be predicted using a weighted linear combination of semantic tuning during the two single-target tasks, we  
194 compared voxelwise semantic tuning across search tasks. A mask vector,  $\bar{m} \in \mathbb{R}^{1 \times 831}$ , was used to select the categories  
195 of interest among 831 categories. Elements of  $\bar{m}$  were one for categories of interest and zero elsewhere. Masked  
196 response profiles,  $\bar{w}'_{ti}$ , were obtained by element-wise multiplication

$$\bar{w}'_{ti} = \bar{w}_{ti} \odot \bar{m} \quad (1)$$

197 where  $\bar{w}_{ti}$  is the response profile for voxel  $i$  and task  $t \in \{H, V, B\}$  denoting attend to “humans”, attend to “vehicles”,  
198 and attend to “both humans and vehicles” (Fig.2), and  $\odot$  represents element-wise multiplication. Masked response  
199 profiles were then projected onto the PCs to assess semantic tuning profiles,  $\bar{S}_{ti}$

$$\bar{S}_{ti} = \bar{w}'_{ti} T \quad (2)$$

200 where  $T \in \mathbb{R}^{831 \times L}$  is the matrix of  $L$  PCs. Semantic tuning profile during divided attention was predicted as a weighted  
201 linear combination of semantic tuning during the two single-target tasks using ordinary least-squares. A voxelwise lin-  
202 earity index (LI) was then quantified as the Pearson’s correlation coefficient between measured and predicted semantic  
203 tuning during divided attention ( $\hat{S}_{Bi}$ ; Fig.4a)

$$LI_i = \text{corr}(\bar{S}_{Bi}, \hat{S}_{Bi}) \quad (3)$$

204 An LI of 1 indicates that semantic tuning during divided attention can be completely described as a weighted linear  
205 combination of the tuning profiles during the two single-target tasks. Whereas, an LI of 0 means that semantic tuning  
206 during divided attention can not be described linearly in terms of tuning profiles during the two single-target tasks. To  
207 study the linearity of the semantic representation during divided attention in an ROI, LIs were averaged across voxels  
208 with significant prediction scores within the ROI.

209 **Bias in semantic representation during divided attention.** We questioned whether semantic representation during  
210 the divided attention task was biased toward any of the single-target attention tasks. To address this issue, we stud-  
211 ied the distribution of semantic representation within an ROI for each individual task. Semantic tuning profiles of

212 significantly predicted voxels within each ROI were pooled to obtain the distribution of tuning profiles

$$S_t = [\bar{S}_{t1} | \bar{S}_{t2} | \dots | \bar{S}_{tn}], \quad t \in \{H, V, B\} \quad (4)$$

213 where  $S_H$ ,  $S_V$ ,  $S_B$  represent distribution of tuning profiles for attend to “humans”, attend to “vehicles”, and attend to  
214 “both humans and vehicles” tasks, and  $n$  is the number of significantly predicted voxels within the ROI. Note that  $S_t$   
215 can also be expressed as

$$S_t = [\bar{P}_{t1}^T | \bar{P}_{t2}^T | \dots | \bar{P}_{tL}^T]^T, \quad t \in \{H, V, B\} \quad (5)$$

216 where  $\bar{P}_{tj} \in \mathbb{R}^{1 \times n}$  is a row vector that represents the projections of the response profiles for task  $t \in \{H, V, B\}$  on the  
217  $j^{\text{th}}$  PC across ROI voxels. To emphasize semantic axes that explain higher variance, projections ( $\bar{P}_{tj}$ ) were weighted  
218 by the explained variance of the corresponding PCs. This yielded the semantic tuning distribution,  $S'_t$ . The tuning  
219 distribution during divided attention was then regressed onto the distributions during the two single-target tasks

$$S'_B = b_H S'_H + b_V S'_V \quad (6)$$

220 The bias index (BI) was quantified (Fig.5a) as

$$BI = \frac{b_H - b_V}{|b_H| + |b_V|} \quad (7)$$

221 Bias toward the attend to humans task would yield  $BI \in (0, 1]$ . A BI of 0 means that the tuning distribution during  
222 divided attention is not biased toward any of the single-target attention tasks. Whereas, a BI of 1 means that tuning  
223 distribution during divided attention is completely biased toward attend to humans task. Similarly, bias toward the  
224 attend to vehicles task would yield  $BI \in [-1, 0)$  where a BI of  $-1$  means that tuning during distribution divided  
225 attention is completely biased toward attend to vehicles task. Note that since the response profiles for the three tasks  
226 were projected onto the same PCs the calculated BI is immune to changes in the direction of PCs.

227 **Visualization on cortical surfaces.** Cortical flatmaps were generated by projecting voxelwise results on individual  
228 subject flattened cortical surfaces. Cortical surfaces were constructed for each subject using T1-weighted anatomical  
229 brain scans. Freesurfer software was used to construct surfaces (Reuter et al., 2012). Surfaces were then flattened  
230 using Pycortex (Gao et al., 2015).

## 231 **Results**

### 232 *Representation of categories during visual search*

233 A recent study showed that thousands of object and action categories in natural scenes are embedded in a semantic  
234 space the axes of which are mapped continuously across the cortex (Huth et al., 2012). Moreover, in a previous study  
235 from our lab we reported that category-based attention warps this space to expand the representation of targets (Çukur  
236 et al., 2013). Yet, attentional modulation of semantic representations during attention to multiple targets remains un-  
237 derstudied. To investigate this issue, we estimated voxelwise tuning for hundreds of object and action categories across  
238 neocortex. Five human subjects viewed 72 minutes of natural movies while they attended to “humans”, “vehicles” (i.e.  
239 single-target attention tasks), or “both humans and vehicles” (i.e. divided attention task) in separate runs. Separate  
240 models were fit for each voxel and attention task. This enabled us to measure responses to 831 distinct object and  
241 action categories in the three attention tasks (Fig.2, see *Materials and Methods*). We find that the category model  
242 accurately predicts responses in many voxels across ventral-temporal, parietal and prefrontal cortices (Fig.3).

243 We hypothesized that if the category responses get modulated by attention task, the fit models during the three attention  
244 tasks would predict better than a null model fit by pooling data across tasks. To investigate this issue, we compared  
245 the average prediction score in the three attention tasks with that of a null model. To obtain the null model, predicted  
246 responses were shuffled across tasks and prediction scores were computed afterwards. We found that attention signif-  
247 icantly modulates category responses in  $77.03 \pm 5.47\%$  of cortical voxels (mean $\pm$ s.d. across five subjects; bootstrap  
248 test,  $p < 0.05$ ). This result implies that attention modulates category responses across visual and non-visual areas.

### 249 *Linearity of semantic tuning during divided attention*

250 We hypothesized that if the attentional modulations during search for multiple targets are mediated by BC, semantic  
251 tuning profile during divided attention should be a weighted linear combination of tuning profiles during isolated at-  
252 tention to individual targets. To test this prediction, we projected category response profiles onto individual subjects’  
253 semantic spaces. The semantic space was constructed via principal component analysis (PCA) on voxelwise response  
254 profiles pooled across the single-target attention tasks. A collection of 36-47 PCs that accounted for over 90% of

255 the variance were selected. To obtain semantic representations of specific categories, the categories of interest were  
256 masked in the response profiles. Ordinary least-squares was then used to predict voxelwise semantic tuning profile dur-  
257 ing divided attention as a weighted linear combination of semantic tuning profiles during the two single-target attention  
258 tasks. Linearity index (LI) was taken as Pearson's correlation coefficient between the predicted and measured seman-  
259 tic tuning during divided attention (Fig.4a, see *Materials and Methods*). We find that LI for target categories (union  
260 of human and vehicle categories) is  $0.81 \pm 0.02$  in category-selective areas (FFA, PPA, EBA, and RSC; mean $\pm$ s.d.),  
261  $0.86 \pm 0.02$  in general object-selective area LOC (mean $\pm$ s.e.m), and  $0.88 \pm 0.02$  in attentional-control areas (IPS,  
262 FEF, and FO; mean $\pm$ s.d.; Fig.4b). LI is significantly greater than zero in all of the studied functional areas (bootstrap  
263 test,  $p < 10^{-4}$ ). These results imply that a substantial portion of semantic tuning during divided attention is linearly  
264 described as a weighted linear combination of tuning during attention to individual targets. In LOC, LI is significantly  
265 higher than that of category-selective areas (bootstrap test,  $p = 0.023$ ). Moreover, LI in attentional-control areas is sig-  
266 nificantly higher than that of category-selective areas (bootstrap test,  $p = 0.004$ ). These results suggest that semantic  
267 tuning for target categories better conform to the weighted linear combination model in general object-selective area  
268 and in later stages of visual processing compared to visual areas that have strong category preference.

269 A recent study from our laboratory showed that during category-based attention voxelwise tuning for nontarget cat-  
270 egories that are semantically similar to targets shifts toward target categories (Çukur et al., 2013). We thus asked  
271 whether the prediction performance of the weighted linear combination for nontarget categories that are semantically  
272 similar to targets is similar to that for targets. To answer this question, we calculated LI separately for nontarget  
273 categories that are semantically similar to targets (i.e. animals, social places, devices, and buildings), and for non-  
274 target categories that are dissimilar to targets (all categories except the union of animals, devices, buildings, social  
275 places, and target categories). LI for similar categories is  $0.64 \pm 0.02$  in category-selective areas,  $0.67 \pm 0.05$  in LOC,  
276 and  $0.71 \pm 0.02$  in attentional-control areas. LI for dissimilar categories is  $0.50 \pm 0.01$  in category-selective areas,  
277  $0.54 \pm 0.06$  in LOC, and  $0.56 \pm 0.03$  in attentional-control areas. LI for similar categories is higher than that for  
278 dissimilar categories in all functional ROIs (bootstrap test,  $p < 10^{-4}$ ). This finding suggests that during divided at-  
279 tention to multiple targets, the competition in representation of nontarget visual objects is carried over to objects that

280 are similar to targets (McMains and Kastner, 2010; Beck and Kastner, 2007). LI is significantly lower in category-  
281 selective areas than in LOC (bootstrap test,  $p = 0.048$  for similar categories,  $p = 0.001$  for dissimilar categories), and  
282 in attentional-control areas (bootstrap test,  $p = 0.007$  for similar categories,  $p = 0.019$  for dissimilar categories). This  
283 result implies that similar to target categories, semantic tuning for nontarget categories is more linear in LOC and in  
284 later stages of visual processing compared to visual areas that have strong category preference.

### 285 *Bias in semantic representation during divided attention*

286 BC observes that inherent selectivity of cortical areas during passive viewing can bias the competition in favor of the  
287 preferred target during visual search (Desimone and Duncan, 1995). To study the interactions between the attentional  
288 bias in semantic representation and the intrinsic category-selectivity of cortical areas, we expressed the semantic rep-  
289 resentation during divided attention task as a weighted linear combination of the semantic representations during the  
290 two single-target tasks in individual cortical areas. We then investigated whether weights in the weighted linear com-  
291 bination were biased toward any of the single-target attention tasks. Masked response profiles across voxels within the  
292 ROI were projected onto individual subjects' semantic spaces to assess the semantic tuning distribution. We regressed  
293 the semantic tuning distribution during divided attention onto distributions during the two single-target tasks. We then  
294 quantified a bias index (BI) using the regression weights. According to this index, bias in the semantic representation  
295 during divided attention task toward the attend to humans task versus attend to vehicles task was represented as posi-  
296 tive versus negative BI in the range  $[-1, 1]$  (Fig.5a, see *Materials and Methods*). We find that BI for target categories  
297 is  $0.32 \pm 0.09$  in human-selective areas (FFA and EBA), and  $-0.29 \pm 0.12$  in scene-selective areas (PPA and RSC;  
298  $\text{mean} \pm \text{s.d.}$ ; bootstrap test,  $p < 10^{-4}$ ; Fig.5b). BI is non-significant in attentional-control areas (IPS, FEF, FO) and the  
299 general object-selective area LOC (bootstrap test,  $p > 0.05$ ). These results suggest that the competition in represen-  
300 tation of target categories during divided attention is biased in favor of the preferred target in cortical areas that are  
301 strongly selective for targets. On the contrary, semantic representation is not biased in the areas without any specific  
302 category preference.

303 In a previous study we showed that attention shifts semantic tuning for both target and nontarget categories (Çukur

304 et al., 2013). Thus, we asked if there is any bias in representation of nontarget categories during divided attention.  
305 To answer this question we separately calculated BI for nontarget categories that are similar to targets and nontarget  
306 categories that are dissimilar to targets. BI for similar categories is  $0.38 \pm 0.25$  in human-selective areas, and  $-0.11 \pm$   
307  $0.38$  in scene-selective areas (mean $\pm$ std.; bootstrap test,  $p < 10^{-4}$ ; non-significant in RSC ( $p = 0.288$ )). BI is non-  
308 significant in attentional-control areas and in LOC (bootstrap test,  $p > 0.05$ ). Meanwhile, BI for nontarget dissimilar  
309 categories is  $0.32 \pm 0.37$  in human-selective areas, and  $-0.11 \pm 0.27$  in scene-selective areas (mean $\pm$ std.; bootstrap  
310 test,  $p < 0.05$ ; non-significant in EBA ( $p = 0.610$ ) and in RSC ( $p = 0.754$ )). BI is non-significant in attentional-control  
311 areas and in LOC (bootstrap test,  $p > 0.05$ ). These results indicate that representation of nontarget categories that are  
312 similar to targets is biased in favor of the preferred target in category-selective areas. Yet, representation of nontarget  
313 categories that are dissimilar to targets is only biased in areas that are strongly selective for the targets and not in areas  
314 that are selective for categories that are semantically similar to targets.

315 The target categories used here (i.e. humans and vehicles) show high semantic dissimilarity. This raises the possibility  
316 that the biases in semantic representation differ between human categories and vehicle categories. To examine this  
317 issue, we compared BI for human and vehicle categories separately. We find that BI for human categories is  $0.59 \pm 0.31$   
318 in human-selective areas, and  $-0.02 \pm 0.08$  in scene-selective areas (mean $\pm$ std.; bootstrap test,  $p < 10^{-4}$  in FFA, EBA;  
319 non-significant in PPA ( $p = 0.88$ ), and in RSC ( $p = 0.94$ )). Whereas, BI for vehicle categories is  $0.38 \pm 0.29$  in human-  
320 selective areas, and  $-0.50 \pm 0.01$  in scene-selective areas (mean $\pm$ std.; bootstrap test,  $p < 10^{-4}$ ; non-significant in EBA  
321 ( $p = 0.62$ )). BI in human-selective areas is significantly positive for both human categories and vehicle categories. Yet,  
322 it is stronger for human categories compared to vehicle categories. Whereas, in scene-selective areas, bias is significant  
323 only for vehicle categories. These results imply that scene-selective areas have a more dynamic representation of their  
324 nonpreferred object categories compared to human-selective areas during divided attention.

325 Among the attentional-control areas, BI is significant for both human and vehicle categories only in IPS (bootstrap test,  
326  $p < 10^{-4}$ ). In IPS, BI is  $-0.48 \pm 0.13$  for human categories, and  $0.60 \pm 0.14$  for vehicle categories (mean $\pm$ s.e.m).  
327 This result indicates that in IPS, representation of human categories is biased toward vehicles and representation  
328 of vehicle categories is biased toward humans. Previous studies suggest that IPS is involved in neural circuits that

329 enhance detection of distractors (Greenberg et al., 2010; Mevorach et al., 2010; Sakai et al., 2002; Bledowski et al.,  
330 2004). In line with these studies, this finding suggests that IPS is involved in category-based attention by enhancing  
331 the representation of distractors.

## 332 **Discussion**

333 In this study, we tested whether the biased competition hypothesis can account for modulation of semantic represen-  
334 tations during a divided attention task. We fit a category model to characterize category responses of single voxels  
335 during search for “humans”, “vehicles”, and “both humans and vehicles”. We found that the category model explains  
336 significant response variance in many voxels across ventral-temporal, parietal, and prefrontal cortices. We estimated  
337 the semantic space underlying category models, and then assessed semantic representations by projecting the category  
338 responses for the three search tasks onto the semantic space.

### 339 *Linearity of the semantic representation during divided attention*

340 We find that a large portion of the variance in semantic tuning during divided attention can be explained using a  
341 weighted linear combination of tuning during isolated attention to individual targets. We find that semantic tuning for  
342 target categories is more accurately predicted via the weighted linear combination relationship compared to semantic  
343 tuning for nontarget categories. In a recent study, we reported that attention shifts semantic tuning for target categories  
344 to a higher degree compared to that for nontarget categories (Çukur et al., 2013). Thus, our results can be attributed to  
345 the higher degree of attentional tuning shift for target categories compared to that for nontarget categories.

346 Several previous studies have investigated differences in the level of competition between strongly category-selective  
347 areas and areas without specific category preference. Reddy and Kanwisher (2007) and MacEvoy and Epstein (2009)  
348 showed that response patterns to a pair of objects can be better predicted by a linear combination of responses to con-  
349 stituent objects in LOC compared to in FFA or PPA. In line with these studies, here we find that semantic representation  
350 is more linear in LOC compared to that in category-selective areas. These results raise the possibility that semantic  
351 representation may also be more linear in attentional-control areas which are not selective for any specific categories



352 compared to strongly category-selective areas (Çukur et al., 2013). Here we find that semantic representations of  
353 target and nontarget categories during divided attention are better explained using the weighted linear combination of  
354 representations during search for individual targets in attentional-control areas compared to category-selective areas.  
355 Our results suggest that higher order areas that are not tuned for specific categories have a more flexible representation  
356 of natural scenes during divided attention.

### 357 *Bias in semantic representation during divided attention*

358 **Category-selective areas.** The bias in the competition among representation of multiple objects have been investi-  
359 gated in several previous studies. Reddy et al. (2009) reported that BOLD responses to multiple static images of faces,  
360 houses, shoes, or cars in category-selective areas, FFA and PPA, are biased toward the preferred target. In line with this  
361 finding, here we find that semantic representation of target categories during divided attention in category selective ar-  
362 eas FFA, EBA, RSC, and PPA is biased toward the preferred category. We have previously shown that category-based  
363 attention shifts semantic tuning not only for the target, but also for nontarget categories that are semantically similar  
364 to the target (Çukur et al., 2013). Consistent with this view, here we find that semantic representations of nontarget  
365 categories that are similar to targets in FFA and PPA are also biased toward the preferred target. Furthermore, repre-  
366 sentation of nontarget categories that are dissimilar to targets is biased in FFA and PPA, albeit to a lower degree. This  
367 finding suggests that in category-selective areas, semantic similarity to targets enhances the level of bias in semantic  
368 representation.

369 In human-selective areas, we find that the semantic representations of target categories (both humans and vehicles)  
370 during divided attention is biased toward the representation during the attend to humans task. Meanwhile, in vehicle-  
371 selective areas, the representation of vehicles but not humans is biased toward the representation during the attend  
372 to vehicles task. This suggests a differential role for human- and vehicle-selective areas in representing nonpreferred  
373 targets during divided attention. A potential explanation of this result is that vehicle-selective areas have a more  
374 dynamic representation of nonpreferred targets compared to human-selective areas (Grill-Spector et al., 2004).

375 **Attentional-control areas.** We did not observe bias in semantic representation of targets toward any of the target  
376 categories in the prefrontal areas that are considered to be part of the attentional-control network. This is expected  
377 considering the lack of tuning to specific categories in these areas (Huth et al., 2012). In IPS, we find that semantic  
378 representation of humans during divided attention is biased toward the representation during the attend to vehicles task  
379 and that the representation of vehicles during divided attention is biased toward the representation during the attend  
380 to humans task. Several previous studies suggest that areas in parietal cortex including IPS enhance visual search by  
381 maintaining the representation of distractors (Mevorach et al., 2010; Bledowski et al., 2004), in addition to spatial  
382 guidance of attention toward targets (Ptak, 2012; Preston et al., 2013). Consistent with this hypothesis, our results can  
383 be interpreted to imply that IPS facilitates natural visual search by maintaining representations of distractor categories.

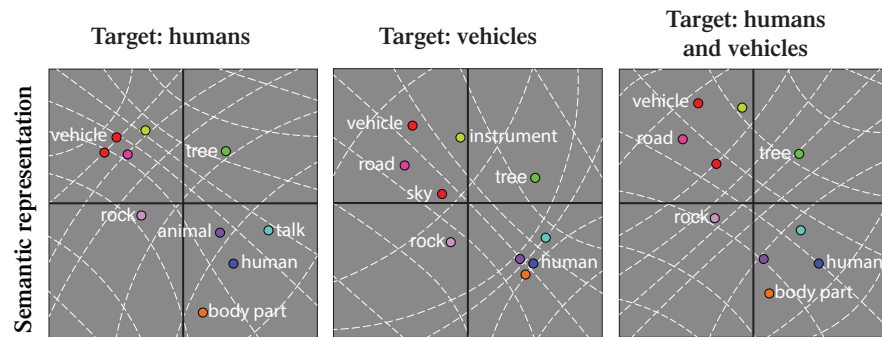
#### 384 *Limitations and future work*

385 Natural stimuli contain correlations among various levels of features (Hamilton and Huth, 2018). For instance, there  
386 might be correlations among low-level visual features of natural scenes and object categories within these scenes  
387 (Lescroart et al., 2015). Such correlations can then bias the category responses that we estimated here, which lead  
388 to a biased assessment of semantic representations. To minimize correlations between category features and global  
389 motion-energy of the movie clips, we used a motion-energy regressor in our modeling procedure. However, we do  
390 not rule out the possibility that there might be residual correlations between category features and intermediate fea-  
391 tures of the movie stimuli, such as object shape characteristics (Op de Beeck et al., 2008) and scene layout (Mullally  
392 and Maguire, 2011). Note that it is perhaps impossible to create natural stimuli completely free from these correla-  
393 tions. However, future studies may mitigate this problem by compiling more controlled natural stimuli that minimize  
394 stimulus correlations while maintaining high variance in individual categories.

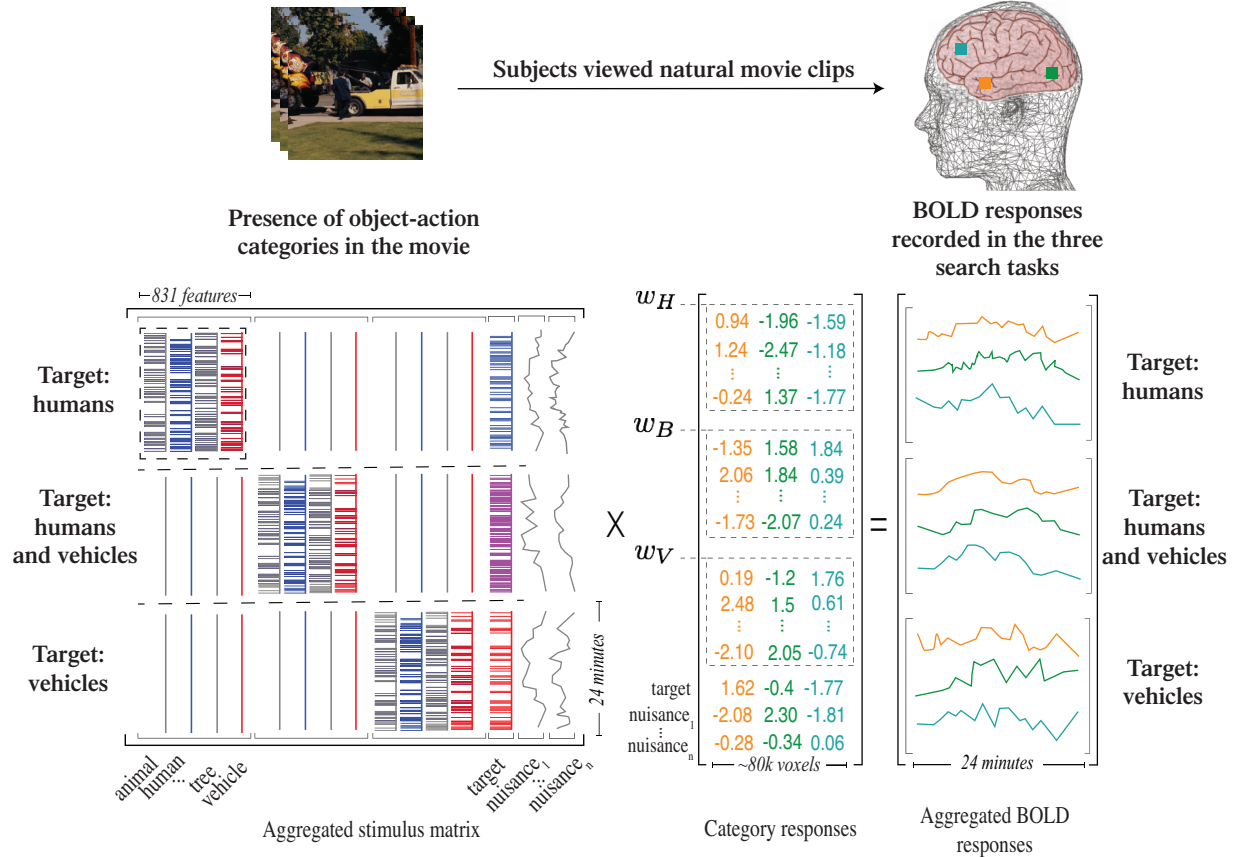
395 Here we examined weighted linear combination as a competition model for the attentional modulation of semantic  
396 representation during divided attention. However, nonlinear models of competition have also been proposed for single  
397 neuron responses in primates. Heuer and Britten (2002) found that in macaque area MT, responses to a pair of Gabor  
398 patches with high versus low contrast was similar to the responses to the Gabor patch with high contrast. Gawne and

399 Martin (2002) showed that in macaque area V4, responses to a pair of checkerboard patterns is similar to the maximum  
400 of the responses to individual patterns in the pair. Further work is needed to investigate possible nonlinear models of  
401 competition in semantic representations.

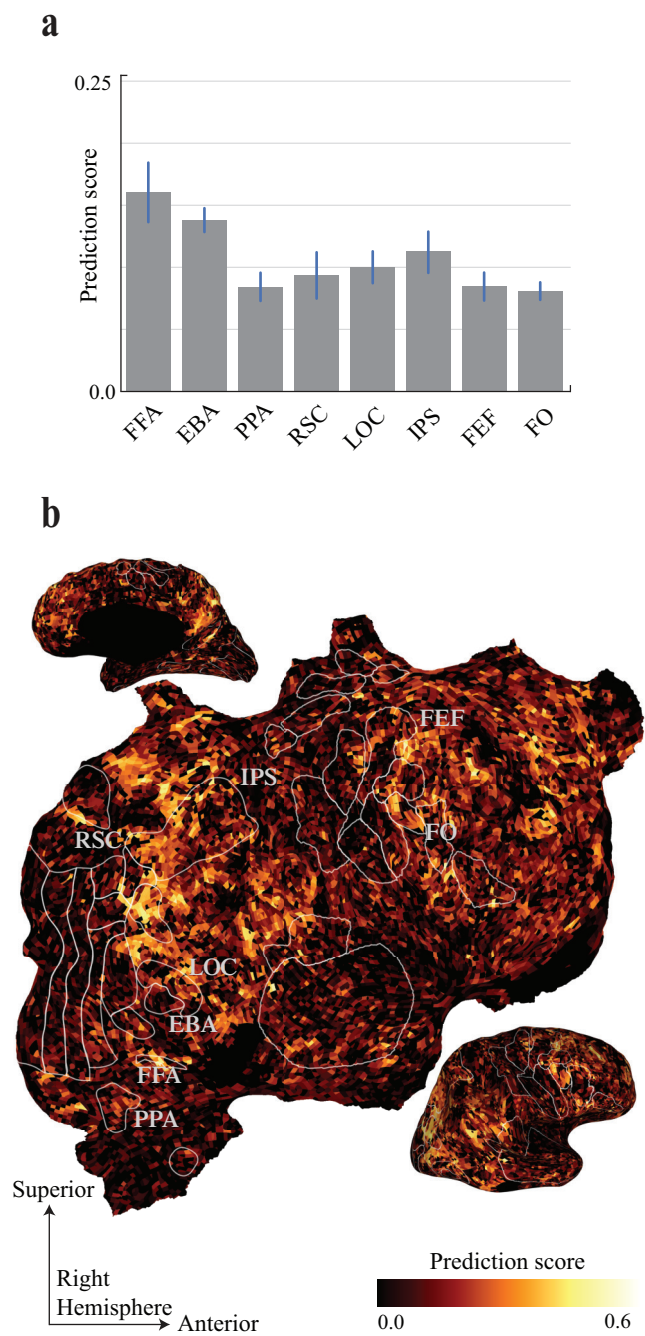
402 In conclusion, our results imply the presence of biased competition not only at the level of individual objects but  
403 also at the level of high-level semantic representations. This competition is evident not only among targets but also  
404 among nontarget categories in natural scenes. Yet, the linearity and bias in representation of nontargets depend on  
405 their semantic similarity to targets. Overall, these results help explain the human ability to perform concurrent search  
406 for multiple targets in complex visual environments.



**Figure 1. Hypothesized changes in semantic representation during divided attention.** Previous studies have proposed that the human brain represents thousands of object and action categories by embedding them in a low-dimensional space based on their semantic similarity. It has been shown that attention warps semantic representation in favor of the target and semantically similar nontargets. Biased competition hypothesis predicts that semantic representation during divided attention is a weighted linear combination of representations during attention to individual targets.



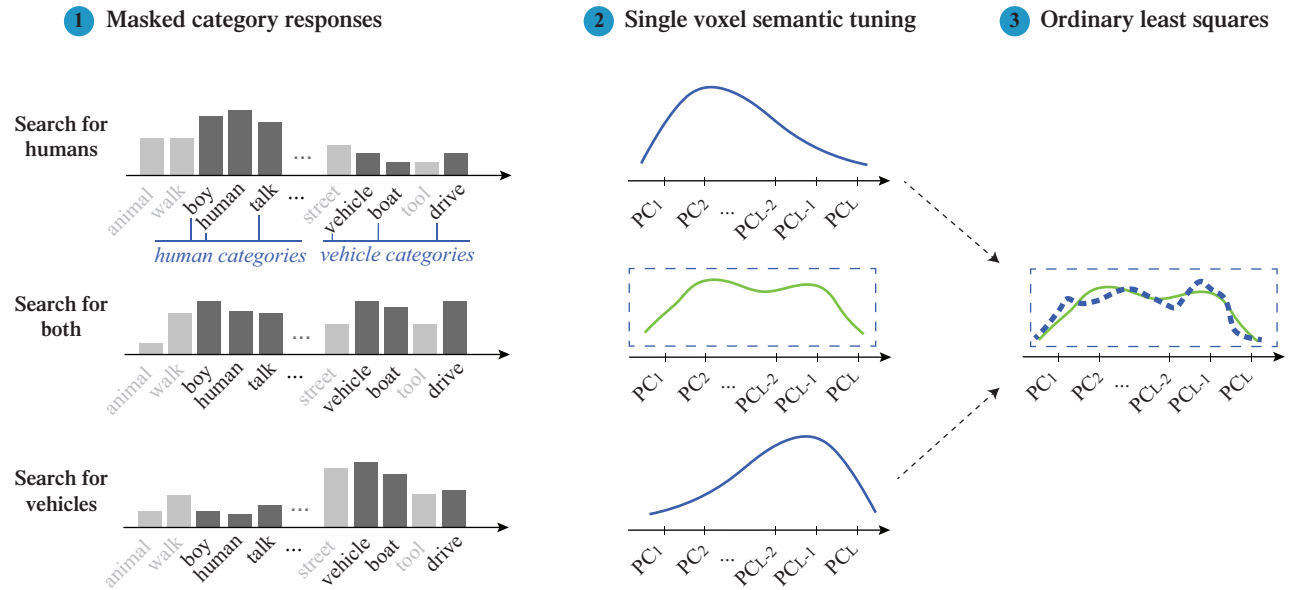
**Figure 2. Experimental and modeling procedures.** Subjects viewed 72 minutes of natural movies. BOLD responses were recorded using functional MRI (fMRI) while subjects performed covert search for “humans”, “vehicles” (i.e. single-target attention tasks), or “humans and vehicles” (i.e. divided attention task) in the movies. Movies were labeled for presence of object and action categories. Presence of superordinate categories was inferred using the WordNet lexicon that resulted in a total of 831 object and action categories. A target-presence regressor was used to account for BOLD response modulations resulting from detection of targets in the scenes. Target-presence regressor comprised of the human regressor (red series), vehicle regressor (blue series), and the binary union of the two (cyan series) to indicate the presence of “humans”, “vehicles”, and “humans and vehicles”, respectively. Nuisance regressors were used to account for physiological noise. Category models were fit independently for each voxel using regularized linear regression. Category responses represent the contribution of each of 831 object-action categories to BOLD responses. Models for the three attention tasks were fit simultaneously using the aggregated stimulus and BOLD response matrices ( $w_H$ ,  $w_V$ , and  $w_B$  for the attend to “humans”, attend to “vehicles”, and attend to “humans and vehicles” tasks, respectively).



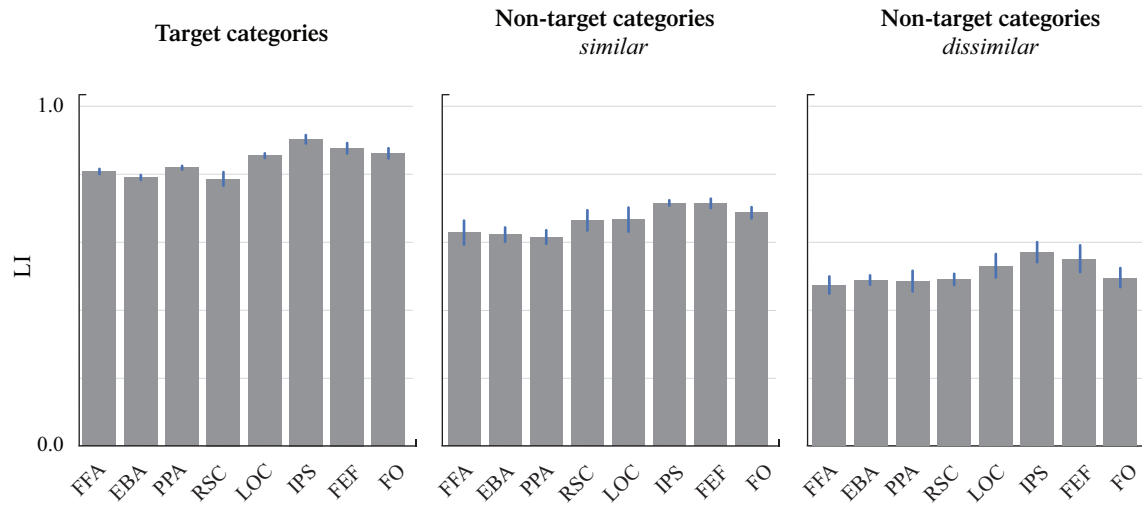
**Figure 3.** (Caption next page.)

**Figure 3. Prediction performance of the category model.** To assess the performance of the category model, BOLD responses were predicted using the estimated category responses in each voxel. Pearson's correlation coefficient between the predicted and measured BOLD responses was taken as the prediction score. **(a)** Prediction score in functional cortical areas (mean $\pm$ s.e.m. across five subjects). FFA, fusiform face area; EBA, extrastriate body area; PPA, parahippocampal place area; RSC, retrosplenial cortex; LOC, lateral occipital complex; IPS, intraparietal sulcus; FEF, frontal eye fields; FO, frontal operculum. **(b)** Cortical flat map of the prediction score for a representative subject. Prediction scores are shown in the right hemisphere. Voxels with high prediction scores appear in yellow color and voxels that have low prediction scores appear in dark gray color. Most voxels in the occipitotemporal cortex, parietal cortex, and prefrontal cortex are well modeled.

**a**



**b**



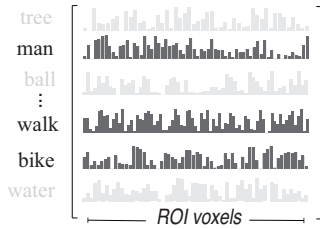
**Figure 4.** (Caption next page.)



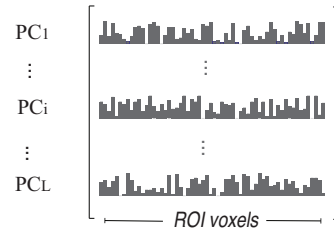
**Figure 4. Linearity of semantic tuning during divided attention.** (a) Masked category responses for the three attention tasks were projected onto subjects' semantic spaces to estimate the voxelwise semantic tuning profiles. The semantic spaces were estimated by performing principal component analysis (PCA) on response profiles pooled during the two single target attention tasks in individual subjects. The collection of principal components (PCs) explaining at least 90% of the variance in the data was selected. Semantic tuning profile during divided attention was predicted via a weighted linear combination of tuning profiles during the two single target tasks (dashed lines indicate predicted tuning; solid lines indicate measured tuning). Pearson's correlation coefficient between the predicted and measured semantic tuning during divided attention was taken as the linearity index (LI). (b) LI in functional cortical areas (mean $\pm$ s.e.m. across five subjects) for target categories (**left**), nontarget categories that are semantically similar to targets (**middle**), and nontarget categories that are dissimilar to targets (**right**). A substantial portion of semantic tuning during divided attention is described as a weighted linear combination even in the absence of target categories. For all cases, LI is significantly higher in attentional-control areas (IPS, FEF and FO;  $p = 0.004$  for target categories,  $p = 0.005$  for similar nontarget categories,  $p = 0.036$  for dissimilar nontarget categories) and in LOC ( $p = 0.023$  for target categories,  $p = 0.048$  for similar nontarget categories,  $p = 0.001$  for dissimilar nontarget categories) than in category-selective areas (FFA, EBA, PPA, and RSC).

**a**

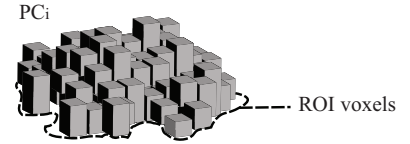
**1** Masked category responses



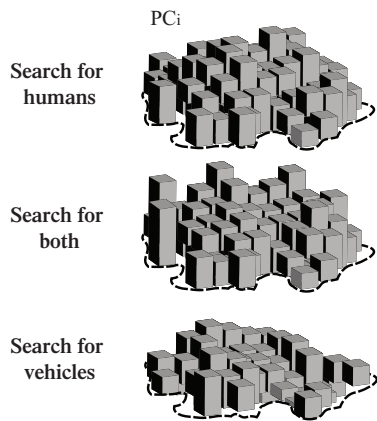
**2** Projection onto semantic space



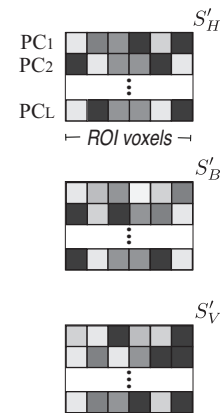
**3** Distribution of projections



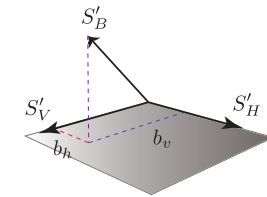
**4** Attentional modulation of projections



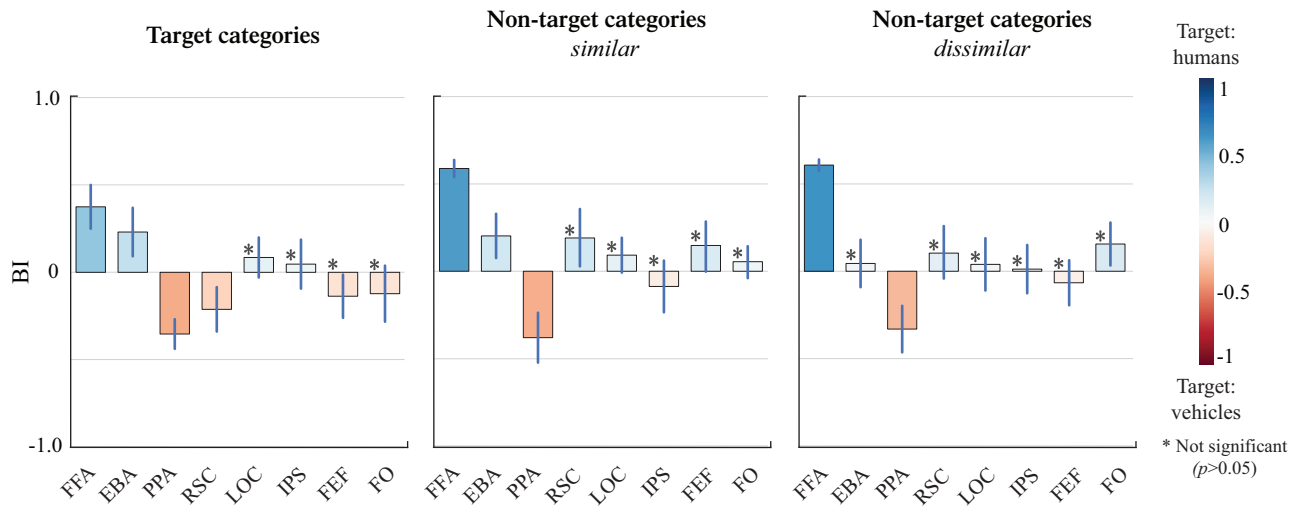
**5** Weighted projection profiles



**6** Linear regression

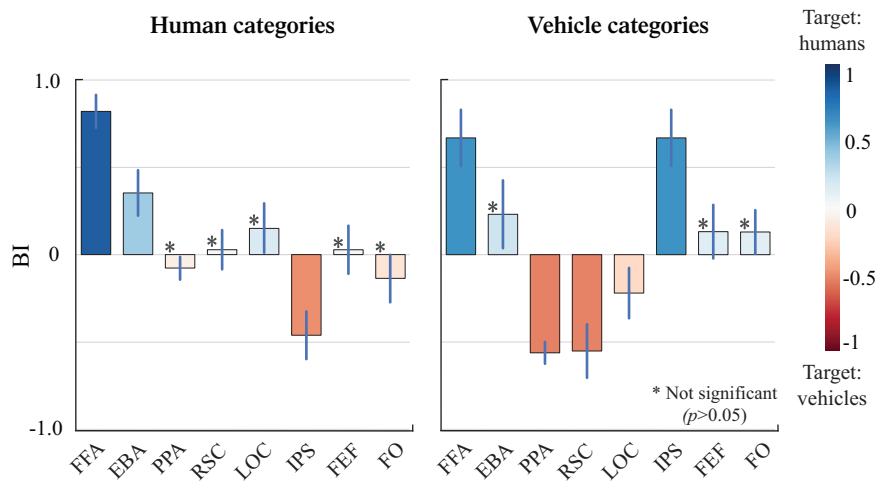


**b**



**Figure 5.** (Caption on next page.)

**Figure 5. Bias in semantic representation during divided attention.** (a) To assess the bias in semantic representation during divided attention, masked category responses for the three attention tasks were projected onto semantic spaces in individual subjects. Projections were weighted by the explained variance of PCs to yield the semantic tuning distributions. Semantic tuning distribution during divided attention was then regressed onto tuning distributions during the two single target tasks. Regression weights were used to calculate a bias index (BI). (b) BI in functional cortical areas (mean±s.e.m. across five subjects) for target categories (**left**), nontarget categories that are similar to targets (**middle**), and nontarget categories that are dissimilar to targets (**right**). Blue versus red bars indicate bias toward attend to humans versus attend to vehicles tasks. Semantic representation of target categories in category-selective areas (FFA, EBA, PPA, and RSC) is biased toward the preferred object-category. Bias is non-significant in attentional-control areas (IPS, FEF and FO) and general object-selective area LOC (bootstrap test,  $p > 0.05$ ). Representation of nontarget categories is also biased toward the preferred target in FFA and PPA.



**Figure 6. Bias in representation of human and vehicle categories.** BI for human categories (**left**), and vehicle categories (**right**) in functional cortical areas (mean±s.e.m. across five subjects). Representations of both human and vehicle categories in FFA are biased toward the attend to humans task. Meanwhile, representation of vehicles but not humans is biased toward the attend to vehicles task in PPA. Representations of both of the target categories are not significantly biased in attentional-control areas in prefrontal cortex. Representation in IPS is biased toward the distractor category.

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