Genetic basis of growth, phenology and susceptibility to biotic stressors in maritime

pine

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Genetic basis of several traits in maritime pine

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Abstract

Forest ecosystems are increasingly challenged by extreme events, e.g. drought, storms, pest and pathogenic fungi outbreaks, causing severe ecological and economical losses. Understanding the genetic basis of adaptive traits in tree species is of key importance to preserve forest ecosystems, as genetic variation in a trait (i.e. heritability) determines its potential for human-mediated or evolutionary change. Maritime pine (*Pinus pinaster* Aiton), a conifer widely distributed in southwestern Europe and northwestern Africa, grows under contrasted environmental conditions promoting local adaptation. Genetic variation at adaptive phenotypes, including height, growth phenology and susceptibility to two fungal pathogens (*Diplodia sapinea* and *Armillaria ostoyae*) and an insect pest (*Thaumetopoea pityocampa*), were assessed in a range-wide clonal common garden of maritime pine. Broad-sense heritability was significant for height (0.219), growth phenology (0.165-0.310) and pathogen susceptibility (necrosis length caused by *D. sapinea*, 0.152; and by *A. ostoyae*, 0.021) measured after inoculation under controlled conditions, but not for pine processionary moth incidence in the common garden. The correlations of trait variation among populations revealed contrasting trends for pathogen susceptibility to *D. sapinea* and *A. ostoyae* with respect to height. Taller trees showed longer necrosis length caused by *D. sapinea* while shorter trees were more affected by *A. ostoyae*. Moreover, maritime pine populations from areas with high summer temperatures and frequent droughts were less susceptible to *D. sapinea* but more susceptible to *A. ostoyae*. Finally, an association study using 4,227 genome-wide SNPs revealed several loci significantly associated to each trait (range of 3-26), including a possibly disease-induced translation initiation factor, eIF-5. This study provides important insights to develop genetic conservation and breeding strategies integrating species responses to biotic stressors.

Introduction

Forest ecosystems are challenged worldwide by changing environmental conditions (Turner, 2010). Warmer and drier climates are expected to increase fire risk, droughts and insect outbreaks while warmer and wetter climates will probably increase storm and pathogen incidence on forests (Seidl *et al.*, 2017), leading to episodes of high tree mortality (Allen *et al.*, 2010) and severe economic losses (Hanewinkel *et al.*, 2013). Changing environmental conditions can also cause range shifts in previously locally restricted pests and pathogens or shifts to increased pathogenicity (Desprez-Loustau *et al.*, 2006). Thus, understanding forest tree genetic variation in disease (and other stressor) responses, in relationship to growth and phenology, is crucial to develop informed forest restoration, conservation and management strategies. Moreover, genes underlying adaptive traits, such as growth or disease response, can serve tree breeding and increase forest productivity, e.g. targeting resistance to drought or against pathogens in forest plantations (Neale & Kremer, 2011).

Forest trees are long-lived, sessile organisms. They are characterized by outcrossing mating systems, high standing genetic variation, large effective population sizes, and the production of vast numbers of seeds and seedlings exposed to strong selection (Petit *et al.*, 2004; Petit & Hampe, 2006). High genetic and phenotypic differentiation has been observed in tree species along environmental gradients (e.g. Savolainen *et al.*, 2007, 2013) or between contrasting habitats, indicating local adaptation (e.g. Lind *et al.*, 2017). Common garden experiments (i.e. experiments evaluating trees from a wide range of populations under the same environmental conditions) provide valuable insights on the phenotypic and genotypic variation of forest trees (Morgenstern, 2011). They revealed genetic differentiation for adaptive traits (such as flushing, senescence or growth) along latitudinal and altitudinal gradients (Mimura & Aitken, 2007; Vitasse *et al.*, 2009). Geographical variation has also been found for disease resistance against certain pests (Menéndez-Gutiérrez *et al.*, 2017) and pathogens (e.g. Hamilton *et al.*, 2013; Freeman *et al.*, 2019). Phenological traits, such as time to flowering, or spring bud burst and autumn leaf senescence times, are often genetically

correlated with disease susceptibility, providing hints on the different kind of resistance or

avoidance mechanisms found in forest trees (Elzinga *et al.*, 2007).

Disease resistance is generally thought to be the result of selective pressures exerted by the

pathogen, in areas where host and pathogen have co-existed during considerable periods of

time, under the co-evolution hypothesis (e.g. Burdon & Thrall, 2000; Ennos, 2015). In this

line, geographical variation in disease resistance has been interpreted in some cases as a

result of past heterogeneous pathogen pressures within the range of a given host species

(Ennos, 2015; Perry *et al.*, 2016). However, the past distribution of pathogen species is often

unknown (Desprez-Loustau *et al.*, 2016), therefore, other processes than co-evolution, such

as "ecological fitting" or "exaptation" should not be excluded (Agosta & Klemens, 2008).

These biological processes have been suggested when, for example, variability in disease

resistance was observed in tree species with no co-evolutionary history with a pathogen

(Leimu & Koricheva, 2006; Freeman *et al.*, 2019). Such resistance may have evolved in

response to other pathogens but had broad-range efficacy, even to a novel pathogen.

Generic mechanisms of resistance in conifers include the production of large amounts of

non-volatile compounds (resin acids) that can act as mechanical barriers to infections (Shain,

1967; Phillips & Croteau, 1999) and volatile compounds (such as monoterpenes or phenols)

that can be toxic to fungi (Cobb *et al.*, 1968; Chou & Zabkiewicz, 1976). The composition of

secondary metabolites can show marked geographic variation in tree species (Meijón *et al.*,

2016). The evolution of plant defences against biotic stressors can also be shaped by

differences in resource availability and environmental constraints throughout the host's

species distribution. Depending on resource availability, plants have evolved distinct

strategies by investing either more in growth, to increase competition ability, or more in

chemical and structural defences, to better respond to herbivores and pathogens (Herms &

Mattson, 2004). Typically, faster growing trees invest more in inducible defences while slow

growing trees invest more in constitutive defences (Moreira *et al.*, 2014).

Many quantitative traits in forest species, including height, growth phenology and disease resistance, show significant heritability and often stronger differentiation (Q_{ST}) among populations than neutral genetic markers (F_{ST}) (Hamilton *et al.*, 2013; see review in Lind *et al.*, 2018), a common indication of adaptive divergence. Major-effect genes for growth, e.g. *korrigan* in *Pinus pinaster* (Cabezas *et al.*, 2015), and resistance genes against forest pathogens, e.g. against the fusiform rust disease in *Pinus taeda* (Kuhlman *et al.*, 2002) and against white pine blister rust in several other North American pine species (Sniezko, 2008; Weiss *et al.,* 2020), have been identified in forest trees. However, most adaptive traits have a highly polygenic basis of quantitative inheritance, typically involving many loci with rather small effects (Goldfarb *et al.*, 2013; De la Torre *et al.*, 2019). Most association genetic studies in forest trees focused on wood property and growth traits to assist tree breeding (e.g. Pot *et al.*, 2005; Neale *et al.*, 2006; Beaulieu *et al.*, 2011). Also, genetic association approaches identified some loci associated to other ecologically important traits, such as cold hardiness (e.g. Eckert *et al.*, 2009; Holliday *et al.*, 2010), drought tolerance (reviewed in Moran *et al.*, 2017), or disease resistance (e.g. Liu *et al.*, 2014; Resende *et al.*, 2017). Compared to more intensively studied traits, association studies addressing biotic interaction traits, including responses to pests and pathogenic fungi, are still limited to few tree species.

Our study focused on maritime pine (*Pinus pinaster* Aiton, Pinaceae), a long-lived conifer from southwestern Europe and northern Africa with a wide ecological amplitude. We assessed height, growth phenology (bud burst and duration of bud burst) and susceptibility to pests/pathogens in a clonal common garden, which allowed us to explore variation in disease response and genetic correlations with other traits in range-wide populations of maritime pine. Considering disease and growth traits together is relevant from an evolutionary and ecological perspective, and can have important implications in terms of management, especially in breeding programs. We selected three important disease agents: two fungal pathogens, *Diplodia sapinea* (Botryosphaeriaceae) and *Armillaria ostoyae*

(Physalacriaceae), as well as the pine processionary moth, *Thaumetopoea pityocampa* (Thaumetopoeidae)*,* a main defoliator of pine forests.

Maritime pine has a highly fragmented natural range in the western Mediterranean Basin, the Atlas Mountains in Morocco, the Atlantic coast of southern France and the west coast of the Iberian Peninsula, and grows from sea level to over 2,000 m altitude. Genetic diversity and population structure is high in natural populations of maritime pine, especially in the Iberian Peninsula, possibly due to its long-term persistence in this region (Petit *et al.*, 1995; Salvador *et al.*, 2000; Bucci *et al.*, 2007; Jaramillo-Correa *et al.*, 2015). In addition, traits, such as stem form, height (González-Martínez *et al.*, 2002), metabolite content (Meijón *et al.*, 2016), drought response (Aranda *et al.*, 2010; Gaspar *et al.*, 2013) and pest and disease resistance (Schvester, 1982; Desprez-Loustau & Baradat, 1991; Burban *et al.*, 1999; Elvira-Recuenco *et al.*, 2014), are highly variable in maritime pine and often strongly differentiated across geographical regions. Maritime pine has also been widely planted and is currently exploited for timber and paper, for example, covering ~1.03 million ha in in southwestern France (including the Landes region), one of the largest plantation forests in Europe (Memento FCBA 2020, https://www.fcba.fr/ressources/memento-2020/). Despite the ecological and economical importance of maritime pine natural forests and plantations, only a few genetic association studies have been developed in this species. Lepoittevin *et al.* (2012) identified two loci associated to growth and wood cellulose content, respectively. Cabezas *et al.* (2015) revealed four SNPs in *korrigan* (a pine gene orthologous to an Arabidopsis degrading enzyme cellulase) also significantly associated to growth traits (total height and polycyclism). Finally, Bartholomé *et al.* (2016) reported four loci for stem straightness and three loci for height growth. In addition, Budde *et al.* (2014) were able to predict 29% of the phenotypic variation in a fire adaptive trait (proportion of serotinous cones) in eastern Spain based on 17 SNP loci. However, none of these studies targeted phenology or biotic interaction traits, such as disease resistance.

Diplodia sapinea is the causal agent of several diseases, such as tip-blight, canker or root

collar necrosis in needles, shoots, stems and roots of conifers, eventually leading to mortality

in case of severe attacks (Piou *et al.*, 1991; Luchi *et al.*, 2014). The pathogenicity of *D.*

sapinea is associated to environmental conditions. It can remain in an endophytic form, i.e.

without causing any symptoms, until stressful environmental conditions, such as drought

(Stanosz *et al.*, 2002; Desprez-Loustau *et al.*, 2006), hail storms (Zwolinski *et al.*, 1990), or

changes in the nitrogen concentration of the soil (Piou *et al.*, 1991; Stanosz *et al.*, 2004),

weaken the host and trigger *D. sapinea* pathogenicity. Trees from all ages are affected

(Chou, 1978; Georgieva & Hlebarska, 2017), though seedlings and old trees show increased

susceptibility (Swart & Wingfield, 1991). The fungus can be found in many conifers,

especially in the genus *Pinus.* Iturritxa *et al.* (2013) classified maritime pine as moderately

susceptible. The species was first described in Europe in 1823 under the name *Sphaeria*

sapinea, and then received many synonyms (Piou *et al.*, 1991). Recent surveys showed that

D. sapinea is currently broadly distributed in all pine forests throughout the world, though its

origin is debated (Burgess *et al.*, 2004; Brodde *et al.*, 2019). Serious damage associated with

D. sapinea in Europe has only been reported in the last decades but it may become a serious

threat to pine forests, as climate change will certainly favor pathogen activity by increasing

temperature, and the frequency and intensity of drought events (Woolhouse *et al.*, 2005;

Desprez-Loustau *et al.*, 2006; Boutte, 2018). In this line, recent outbreaks associated with *D.*

sapinea in northern Europe suggest an ongoing northward expansion (Brodde *et al.*, 2019).

Armillaria ostoyae is a root pathogen that causes white rot and butt rot disease in conifers,

leading to growth deprivation, high mortality and major losses in timber wood; hence, its

economic importance (Cruickshank, 2011; Heinzelmann *et al.*, 2019). The species can be

traced back to six millions years ago, both in Eurasia and North America (Tsykun *et al.*, 2013;

Koch *et al.*, 2017). *Armillaria ostoyae* has been reported in all the conifer forests of the

Northern Hemisphere but it appears to be replaced by *A. mellea* (Marxmüller & Guillaumin,

2005) in the Mediterranean due to higher temperatures and drought. It is likely that *A. ostoyae*

would have co-existed for a long time with maritime pine in Europe (Tsykun *et al.*, 2013) and consequently it could have been affected by the same extinction-recolonization events associated to past climatic changes as its host (Labbé *et al.*, 2017a). *Armillaria ostoyae* is one of the most common fungal species in maritime pine forests, being particularly dangerous as it can act as both a parasite and a saprophyte (Cruickshank *et al.*, 1997; Labbé *et al.*, 2017b), i.e. the death of its host does not prevent its spread. In maritime pine, the severity of *A. ostoyae* symptoms is related to host age, with higher mortality in young trees (Lung-Escarmant & Guyon, 2004; Labbé *et al.*, 2015). Climate change is predicted to foster an increased impact of *A. ostoyae* on conifer forests in the coming years (Kubiak *et al.*, 2017).

Thaumetopoea pityocampa is considered the most severe defoliator insect in pine forests in southern Europe and northern Africa (Jactel *et al.*, 2015), leading to severe growth loss (Jacquet *et al.*, 2013). The species typically reproduces in summer followed by larval development during autumn and winter. Caterpillars and moths of *T. pityocampa* are sensitive to climatic and environmental conditions, and this pest is expected to expand its range following events of climate warming (Battisti *et al.*, 2006; Toïgo *et al.*, 2017).

The specific objectives of our study are to: 1) estimate phenotypic variability and heritability within and among range-wide populations of maritime pine for height, growth phenology and susceptibility to pests/pathogens; 2) test for adaptive divergence across the maritime pine range for these traits (i.e. Q_{ST} vs. F_{ST} approach); 3) analyze the pattern of trait correlation within and among populations, in particular to identify putative correlations that could be useful/detrimental for conservation and breeding programs; and 4) identify loci associated to disease-related, growth and phenology traits using genotype-phenotype association. Altogether, our approach, combining the evaluation of a clonal common garden and a genotyping array, produced relevant insights on the evolution, genetic basis, and architecture of adaptive traits in maritime pine, an ecologically and economically important forest tree species.

Material and Methods

Plant material and common garden measurements

A clonal common garden (CLONAPIN) was planted in 2011 in Cestas, southwestern France (for details see Rodríguez-Quilón, 2017). It includes trees from 35 populations of maritime pine covering the whole natural species distribution (see Table S1.1, Supporting Information for number of individuals and genotypes, and population coordinates of 32 populations included in this study), representing all known differentiated gene pools in the species (Central Spain, Southeastern Spain, Iberian Atlantic, French Atlantic, Corsica and Morocco; see Jaramillo-Correa *et al.*, 2015). The common garden design consisted of eight randomized complete blocks, with one clonal copy (ramet) of each genotype replicated in each block. For the pathogen inoculation experiments, we chose genotypes from populations representing each of the six gene pools. Because of the higher logistical effort involved in inoculation protocols with respect to other measurements (see below), it was not possible to include all common garden genotypes in these experiments (see Table S1.1, Supporting Information).

Height, growth phenology (bud burst and duration of bud burst), and incidence of processionary moth (*T. pityocampa*), were measured in all individuals from 5-8 blocks, depending on the trait (sample size of 1,392-3,204 trees, see Table S1.1, Supporting Information). Pathogen susceptibility was assessed in a subset of genotypes using excised branches collected from the clonal trial (sample size of 180-453 branches, see Table S1.1, Supporting Information and below). Tree height was measured in 2015, four years after the establishment of the trial. Bud burst stage was evaluated using a phenological scale ranging from 0 to 5 (0: bud without elongation during winter, 1: elongation of the bud, 2: emergence of brachyblast, 3: brachyblast begins to space, 4: elongation of the needles, 5: total elongation of the needles; see Figure S2.1, Supporting Information). The Julian day of entry in each stage (S1 to S5) was scored for each tree. Julian days were converted into accumulated degree-days (0°C basis) from the first day of the year, to take into account the

between year variability in temperature. The number of degree-days between stages 1 and

4 defines the duration of bud burst. Both growth phenology phenotypes, bud burst and

duration of bud burst, were assessed in 2015 and 2017. The presence or absence of pine

processionary moth nests (*T. pityocampa*) in the tree crowns was assessed in March 2018.

Experimental evaluation of susceptibility to Diplodia sapinea

Inoculations were carried out on excised shoots taken from pines in the common garden (for

a detailed laboratory protocol see Supporting Information S3.1). We used the pathogen strain

Pier4, isolated from *P. nigra* cones in Pierroton, France (May 2017), and maintained on malt-

agar medium. The identity of this strain as *D. sapinea*, was confirmed by sequencing the ITS

region, amplified using the primers ITS1-F and ITS4 (Gardes & Bruns, 1993), and blasting it

against the NCBI nucleotide database (Benoît Laurent, personal communication). Only

current-year shoots at the phenological stages 3 to 5, i.e. with fully elongated buds but not

fully mature, were collected, otherwise sampled randomly (see Supporting Information S2.1).

For the inoculation, we removed a needle fascicle in the middle of each shoot with a scalpel.

A 5 mm diameter plug of malt-agar taken at the active margin of a *D. sapinea* culture was

put on the wound, mycelium-side down, and then wrapped in cellophane. Control shoots

were treated in the same manner but with plugs of sterile rather than colonized malt-agar.

The shoots were put in water and kept in a climatic chamber set at 20°C with a daily cycle of

12h of light and 12h of dark (Blodgett & Bonello, 2003; Iturritxa *et al.*, 2013). Six days after

the inoculation, we removed the cellophane and measured the lesion length around the

inoculation point with a caliper. The shoots were not lignified and the lesions were visible.

However, the surface was superficially stripped to see the limit of the lesion when it was not

visible otherwise. Needle discoloration was also observed, and evaluated using a scale from

0: no discoloration to 3: all needles along the necrosis showed discoloration (see Figure S3.1,

Supporting Information). To confirm that discoloration was caused by the pathogen, one

discolored needle from one branch per population was placed on a malt-agar Petri dish to

grow. After 3 days, *D. sapinea* could be visually identified in each Petri dish.

For this experiment, we sampled a total of 453 branches, from 151 genotypes (i.e. one branch from each of three replicate trees per genotype) in ten populations, representing all differentiated gene pools known in maritime pine (see Jaramillo-Correa *et al.*, 2015). Every day between June 12th and July 31st 2018, one lateral branch per tree was cut from the previous year whorl in 30 of the selected trees (see above) and taken to the laboratory for inoculation. Inoculations were performed on the leader shoot of the current whorl of the excised branch.

Experimental evaluation of susceptibility to Armillaria ostoyae

For the inoculation with *A. ostoyae*, we used the pathogen strain A4, collected from a dying maritime pine tree in La Teste (Gironde, France) in 2010 (Labbé *et al.*, 2017b). For the experiment, two plugs of 5 mm diameter of malt agar with the *A. ostoyae* mycelium were put on the top of a mixture of industrial vegetable soup (Knorr 9 légumes©, Heilbronn, Germany), malted water and hazelnut wood chips in a 180 mL plastic jar (Heinzelmann & Rigling, 2016) (for a detailed laboratory protocol see Supporting Information S3.2). The lid was closed loosely enough to allow some oxygen flow. The jars were placed during three months in a heat chamber set at 23°C and 80% humidity before inoculation. The basal part of the sampled shoots (ca. 8 cm) was placed in the center of the mycelial culture in the heat chamber, maintaining the same temperature and humidity settings as for the mycelium growth, but adding an additional 12h cycle of light/dark. Only the jars showing a minimum jar occupation by *A. ostoyae* of 60% were used. After 3 weeks, wood colonization success for each sample was evaluated visually by confirming the presence of mycelium under the bark. At this point, the needles and the phloem of the branches were still green and thus the sample was considered adequate for susceptibility measurements. The length of the colonizing mycelium and length of the lesion in the sapwood (i.e. wood browning, hereafter referred to as necrosis) were measured. In the jar, we also visually evaluated the level of humidity of the medium (dry, medium, and very humid) and *A. ostoyae* growth. Controls were prepared in

the exact same manner, but with plugs of sterile malt-agar as opposed to those colonized by

A. ostoyae.

For this experiment, we randomly sampled ten maritime pine genotypes from six populations,

one for each of the six differentiated gene pools represented in the CLONAPIN common

garden. As root inoculations are destructive and impossible to implement directly in the

common garden, we selected fully elongated current year shoots (bud stage 4 and 5), with a

maximum diameter of 15 mm and a minimum length of 10 cm, as proxies (Matusick *et al.*,

2016). A total of 180 branches from 60 genotypes (i.e. one branch from each of three

replicate trees per genotype) were measured, cut and taken to the laboratory to be

inoculated, on October 3rd-4th 2018.

Climatic Data

Summary climate data for the years 1950–2000 were retrieved for 32 variables from

Worldclim (Hijmans *et al.*, 2005) and a regional climatic model (Gonzalo, 2007) for the 11

non-Spanish and the 22 Spanish populations, respectively. Climate variables included

monthly mean, highest and lowest temperatures, and mean monthly precipitation. Gonzalo's

(2007) model was favored for climate data in Spain because it considers a much denser

network of meteorological stations than Worldclim, which is known to underperform,

particularly for precipitation estimates, in this region (see Jaramillo-Correa *et al.*, 2015).

DNA extraction and SNP genotyping

Needles were collected from one replicate per genotype (*N*=416, including all genotypes

used for pathogen susceptibility assays) and desiccated using silica gel. Genomic DNA was

extracted using the Invisorb® DNA Plant HTS 96 Kit/C kit (Invitek GmbH, Berlin, Germany).

An Illumina Infinium SNP array developed by Plomion *et al.* (2016) was used for genotyping.

Apart from potentially neutral genetic polymorphisms, this array comprises SNPs from

candidate genes that showed signatures of natural selection (Eveno *et al.*, 2008; Grivet *et*

al., 2011), significant environmental associations with climate at the range-wide spatial scale

(Jaramillo-Correa *et al.*, 2015), or differential expression under biotic and abiotic stress in

maritime pine (Plomion *et al.*, 2016). After standard filtering followed by removal of SNPs

with uncertain clustering patterns (visual inspection using *GenomeStudio v. 2.0*), we kept

5,176 polymorphic SNPs, including 4,227 SNPs with a minor allele frequency (MAF) above

0.1.

Quantitative genetic analyses

To estimate the genetic variance components of the analyzed traits, we ran two different sets

of mixed-effect models. The first one included three hierarchical levels, to account for the

strong population genetic structure in maritime pine: gene pool, population nested within

gene pool, and genotype (clone) nested within population and gene pool, as follows:

$$y_{igjk} = \mu + block_i + gp_g + gp(pop)_{gj} + gp(pop(genotype))_{gjk} + \varepsilon_{igjk} \qquad (1)$$

$$y_{igjk} = \mu + block_i + cov + gp_g + gp(pop)_{gj} + gp(pop(genotype))_{gjk} + \varepsilon_{igjk} \qquad (2)$$

where for any trait y_{igjk}, μ denotes the overall phenotypic mean, *block$_i$* represents the fixed

effect of experimental block *i*, *gp$_g$* is the random effect of gene pool *g*, *gp(pop)$_{gj}$* is the random

effect of pop *j* nested within gene pool *g*, *gp(pop(genotype))$_{gjk}$* is the random effect of genotype

k nested within population *j* and gene pool *g*, and ε is the overall residual effect. In model 2,

cov represents the covariates implemented when modeling the presence of pine

processionary moth nests (i.e. tree height in 2015) and necrosis caused by *A. ostoyae* (i.e.

a categorical evaluation of jar humidity). Using these models, the effects at the gene pool

level had extremely wide confidence intervals, probably because some of them were

represented by only a few highly contrasted populations. Thus, we also ran a set of models

without the gene pool effect, as follows:

$$y_{ijk} = \mu + block_i + pop_j + pop(genotype)_{jk} + \varepsilon_{ijk} \qquad (3)$$

$$y_{ijk} = \mu + block_i + cov + pop_j + pop(genotype)_{jk} + \varepsilon_{ijk} \qquad (4)$$

This set of models produced more accurate BLUP estimates, avoiding also confounded

effects between gene pool and population, and was thus preferred for some of the

subsequent analyses (see below).

All mixed-effect models were fitted in a Bayesian framework using Markov chain Monte Carlo

(MCMC) methods implemented in the R package *MCMCglmm* (Hadfield, 2010) using R

v.3.4.1 (R Development Core Team, 2017). All analyzed traits presented a Gaussian

distribution with the exception of presence of pine processionary moth nests and needle

discoloration caused by *D. sapinea* infection that followed a binomial distribution and were

modeled with *logit* and *probit* link functions, respectively. Multivariate-normal prior distribution

with mean centered around zero and large variance matrix (10^8) were used for fixed effects

with the exception of the model for needle discoloration caused by *D. sapinea* where a

gelman prior for V was set, as suggested by Gelman *et al.* (2008) for ordinal traits. Inverse

Wishart non-informative priors were used for the variances and covariances with a matrix

parameter V set to 1 and a parameter n set to 0.002 (Hadfield, 2010). Parameter expanded

priors were used to improve the convergence and mixing properties of the chain as

suggested by Gelman (2006) for models on the presence of pine processionary moth nests,

needle discoloration caused by *D. sapinea*, and necrosis caused by *A. ostoyae*. Parameter

estimates were not sensitive to change in the priors. The models were run for at least 750,000

iterations, including a burn-in of 50,000 iterations and a thinning interval of 500 iterations.

Four chains per model were run to test for parameter estimates convergence. Gelman-Rubin

criterion Potential Scale Reduction Factor (psrf) was consistently below 1.01 (Gelman &

Rubin, 1992) (see Table S4.1, Supporting Information, for further details on model

specifications).

Variance components from the first set of models (including gene pool effect) were then

used to compute broad-sense heritability (H^2) as:

$$H^2 = \frac{\sigma^2_{genotype}}{(\sigma^2_{genotype} + \sigma^2_e)} \tag{5}$$

where $\sigma^2_{genotype}$ is the variance among genotypes within populations and gene pools and

σ^2_e the residual variance. In addition, we computed the variance ratios associated to the

population (H^2_p) and gene pool (H^2_{gp}) effects as:

$$H^2_p = \frac{\sigma^2_{pop}}{(\sigma^2_{pop}+\sigma^2_{genotype}+\sigma^2_e)}$$ (6)

$$H^2_{gp} = \frac{\sigma^2_{gp}}{(\sigma^2_{gp}+\sigma^2_{pop}+\sigma^2_{genotype}+\sigma^2_e)}$$ (7)

where σ^2_{pop} is the variance among populations within gene pools, σ^2_{gp} is the variance among

gene pools. When appropriate, we included an extra term in the denominator to account for

implicit *logit* and *probit* link function variance ($\pi^2/3$ and $+1$, respectively; Nakagawa &

Schielzeth, 2010). Genetic differentiation among populations for the analyzed traits (Q_{ST}) was

calculated as in Spitze (1993), using the second set of models (without gene pool effects,

equations 3 and 4), as we were interested in the differentiation at the population level:

$$Q_{ST} = \frac{\sigma^2_{pop}}{\sigma^2_{pop}+2\sigma^2_{genotype}}$$ (8)

where σ^2_{pop} is the overall variance among populations and $\sigma^2_{genotype}$ is the variance among

genotypes within populations. Finally, we estimated the global F_{ST} using all available SNP

genotypes in SPAGeDi 1.5 (Hardy & Vekemans, 2002). The difference between global F_{ST}

and Q_{ST} values for each adaptive trait was considered significant when the 95% credible

intervals (CI) did not overlap.

Genetic correlations were calculated with the Pearson's coefficient of correlation using the

Best Linear Unbiased Predictors (BLUPs) of the genotype (clone) effect (Henderson, 1973;

Robinson, 1991) for each trait. The same estimate was used to characterize correlations of

trait variation at population level considering the combined genotype and population effects.

BLUPs were obtained from the model including genotype and population effects but not the

gene pool effects (second set, equations 3 and 4), as they produced more precise

estimates. In addition, we employed linear mixed models to test the effect of environmental

variables on the population BLUPs while taking into account the assignment of each

population to one of the six gene pools as random factor. First, we tested each model

against a null-model without fixed effects to identify the significant associations. Then, for

significant environmental associations, we calculated R^2 as the proportion of variance

explained by the fixed predictors using Nakagawa & Schielzeth's approach (Johnson, 2014;

Nakagawa & Schielzeth, 2013) implemented in the function 'r2beta' of the *r2glmm* package

(Jaeger, 2017).

Genotype-phenotype association

First, we used a mixed linear regression approach (MLM, Yu *et al.*, 2006) implemented in

Tassel v. 5.0 (Bradbury *et al.*, 2007) to identify single SNPs associated to each of the

phenotypes. Phenotypic values were estimated using the BLUPs accounting for both

population and genotype effects from the models without gene pool effects (second set,

equations 3 and 4). Ancestry proportions of each sample to six genetic clusters (see

Jaramillo-Correa *et al.*, 2015) were computed using STRUCTURE (Pritchard *et al.*, 2000).

These ancestry proportions were included as covariates in the MLM. A covariance matrix

accounting for relatedness between all sample pairs was estimated using Loiselle's kinship

coefficient (Loiselle *et al.*, 1995) in SPAGeDi 1.5 and was included as random effect.

Negative kinship values were set to zero following Yu *et al.* (2006). Only loci with a *P*-value

below 0.005 in the Tassel analyses and with a minimum allele frequency above 0.1 were

used for further analyses. Second, for the identified associations, we used a Bayesian mixed-

effect association approach (Bayesian Association with Missing Data, BAMD; Quesada *et*

al., 2010; Li *et al.*, 2012) in R to estimate single-locus genotype effects under three genetic

models accounting for additive, over-dominance and dominance effects, respectively (as in

Budde *et al.*, 2014). As in the previous models, the STRUCTURE ancestry proportions were

used as covariates and the relatedness matrix as random factor. Mean allelic effects (γ) and

95% confidence intervals were obtained from the distribution of the last 20,000 iterations

(50,000 in total). Only those SNPs with credible intervals not overlapping zero were

considered to have a significant (non-zero) effect on the trait.

Functional annotations, SNP motives and BLAST results for significant genotype-phenotype

associations were retrieved from Plomion *et al.* (2016). Finally, in order to inspect visually

geographical patterns, the minimum allele frequency of significantly associated SNPs was

estimated in each population using SPAGeDi 1.5 and plotted in a map.

Results

Phenotypic variability, broad-sense heritability and genetic differentiation

Broad-sense heritability was highest for bud burst in 2015 (H^2: 0.310, CI [0.272-0.358]),

intermediate for height (H^2: 0.219 [0.186-0.261]) and lowest for necrosis length caused by *A.*

ostoyae (H^2: 0.021 [0.004-0.121], Table 1). Broad-sense heritability of susceptibility to *D.*

sapinea, assessed as the necrosis length, was also significant (H^2: 0.152 [0.038-0.292]), with

trees from northern Africa and southern Spain showing shorter necrosis length than trees

from Atlantic populations (Figure 1a). Broad-sense heritability of needle discoloration caused

by *D. sapinea* had similar magnitude, but was not significant (H^2: 0.123 [0.000-0.233]).

Necrosis length caused by *A. ostoyae* was most pronounced in southern populations,

especially in Morocco and southern Spain, while less pathogen growth was observed in

northern populations, e.g. in St. Jean de Monts from the French Atlantic gene pool (Figure

1b). Incidence of pine processionary moth nests in the common garden was not heritable

(H^2: 0.001 [0.000-0.207]). The tallest trees were found in populations from the Atlantic French

and Atlantic Iberian gene pools, and for one of the Corsican populations (Pinia), whereas the

shortest ones originated from southeastern Spain and Morocco (Figure S5.1, Supporting

Information).

Most traits showed strong differences among gene pools, although the variance ratio

associated to the gene pool effect was characterized by wide confidence intervals (Table 1).

Based on overall population genetic differentiation, we observed significant Q_{ST} values

(ranging from 0.191 for bud burst in 2017 and duration of bud burst in 2015, to 0.636 for

necrosis length caused by *D. sapinea*) indicating strong population differentiation (Table 1).

Global F_{ST} calculated using the available SNPs was 0.109 ([0.013; 0.325], p-value < 0.001)

which is significantly lower than the Q_{ST} estimates obtained for height and necrosis length

caused by *D. sapinea* (Table 1).

Correlations between traits and with environmental variables

All genetic correlations (i.e. those considering only the genotype effect) involving height and

growth phenology were significant while disease susceptibility traits did not show significant

genetic correlations except for necrosis length and needle discoloration caused by *D.*

sapinea (Table 2). The strongest genetic correlation was observed between bud burst and

duration of bud burst in 2015 (0.798, p-value < 0.001). Once the population effect was

considered, a significant negative trait covariation was found between necrosis length

caused by each of the two fungal pathogens (-0.692, p-value < 0.001; Table 2, Figure 2). We

also observed significant population correlations with height, negative for necrosis length

caused by *A. ostoyae* (-0.653, p-value < 0.001) and positive for necrosis length caused by

D. sapinea (0.679, p-value < 0.001).

Moreover, using mixed linear models that account for gene pool as random factor, we

observed a negative relationship for maximum temperatures in the summer months with

height, duration of bud burst in 2017, and necrosis length and needle discoloration caused

by *D. sapinea* (Table 3). The strongest relationship was found for height and maximum

temperatures in July (R^2 = 0.709 [0.556-0.831], p-value <0.001) while the variance of e.g.

needle discoloration caused by *D. sapinea* explained by maximum temperatures in July was

lower (R^2 = 0.524 [0.128-0.841], p-value = 0.004; Table 3, Figure 3). Also, geographical

variables, such as altitude, latitude and longitude showed significant correlation with all

phenotypes. Interestingly, necrosis length caused by *A. ostoyae* showed a latitudinal cline

470 (R^2 = 0.750, p-value = 0.016) but only a marginally significant positive correlation with

471 maximum temperatures in July (R^2=0.514, p-value= 0.066).

472 *Genotype-phenotype association*

473 Between three and 28 SNPs were significantly associated with each of the phenotypic traits

474 evaluated under different genotype effect models (see Table S6.1, Supporting Information).

475 Here we only focus on SNPs that were significant under the additive genetic model, this

476 model being built on three genotypic classes and therefore considered the most robust.

477 Based on this model, five SNPs were associated to height, 27 SNPs were associated with

478 growth phenology (considering altogether the different phenology traits and measurement

479 years), and seven with pathogen susceptibility (Table 4). In total, four significantly associated

480 SNPs showed non-synonymous changes. Two non-synonymous SNPs were associated to

481 bud burst in 2017 (Figure S6.1, Supporting Information). In addition, one non-synonymous

482 SNP was associated to needle discoloration caused by *D. sapinea* and another one to

483 duration of bud burst in 2015 (Table 5 and Figure 4). All the remaining SNPs involved in

484 significant associations under the additive model were either non-coding or the effect of the

485 substitution was unknown (Table S6.1, Supporting Information). The geographical

486 distribution of minor allele frequency of the associated SNPs was quite variable and did not

487 reflect the population genetic structure of the species (Figure 5).

489 **Discussion**

490 In the current context of climate change, understanding the genetic basis of adaptive traits

491 in tree species is key for an informed forest management and the development of new

492 breeding applications. In this study, we assessed variation in maritime pine for height and

493 growth phenology traits, as well as for incidence of pine processionary moth nests and

494 response to two pathogenic fungi, *D. sapinea* and *A. ostoyae*, at the range-wide scale. Novel

495 inoculation protocols were developed for *D. sapinea* and *A. ostoyae* based on excised

branches and applied to trees grown in a clonal common garden. For the first time, we estimated broad-sense heritability of pine susceptibility for these two major fungal pathogens (evaluated as necrosis length), as well as the variance ratios associated to the population and gene pool effects. Most of these adaptive traits showed significant broad-sense heritability. Variation of pathogen susceptibility among geographical provenances, as well as duration of bud burst in 2017, followed a latitudinal gradient, possibly corresponding to a climatic gradient, but in opposite directions for the two pathogens. A genetic association approach revealed several loci significantly associated to height, growth phenology and pathogen susceptibility in maritime pine. This information will be useful for current efforts to implement breeding strategies based on genome prediction (i.e. genomic selection) in this species. The presence of pine processionary moth nests evaluated in the common garden was not heritable but future studies should consider the level of infestation and/or assess damage more precisely.

Genetic variation and differentiation, and correlations with climate

Broad-sense heritability estimates were in the range of previously published values in maritime pine or other forest tree species. The variance ratios associated to population and gene pool effects reflect genetic differences that are not necessarily due to selection but might also reflect other evolutionary processes at the population level, e.g. drift or different levels of gene flow. Nevertheless, they provide important insights on trait variation at the range-wide geographical scale. The variance ratio associated to the gene pool effect had very wide confidence intervals indicating that highly differentiated units were mixed at this level. Already Rodríguez-Quilón *et al.* (2016) described strong quantitative genetic differentiation for height and survival within gene pools of maritime pine. While gene pools may reflect different evolutionary histories, our results suggest that selection and hence adaptation at quantitative traits takes place at a finer spatial scale. Therefore, genetic differentiation for adaptive traits at the population level (e.g. as estimated by Q_{ST}) is expected to be more informative. Neutral genetic differentiation, i.e. F_{ST}, was moderate ($F_{ST} = 0.109$

[0.013; 0.325], p-value < 0.001) and significantly lower than Q_{ST} estimates obtained for height

and necrosis length caused by *D. sapinea*, which suggests that divergent selection is

promoting local adaptation in these traits (Whitlock & Guillaume, 2009; Lamy *et al.*, 2011).

Height is a crucial, frequently studied trait in forest trees, as increased height growth is a

main target of breeding programmes (e.g. Kremer & Lascoux, 1988; Cornelius, 1994). In our

study, we found moderate broad-sense heritability of 0.219 for height, well in line with

estimates in other conifer species e.g. ranging from 0.21 in *Pinus taeda* to 0.78 in *Picea abies*

(reviewed in Lind *et al.*, 2018) and from 0.148 to 0.282 in maritime pine saplings depending

on the common garden site and the provenance (Rodríguez-Quilón *et al.*, 2016). Height is

known to be a highly integrative trait closely responding to abiotic factors, such as climate

(Alía *et al.*, 2014; Jaramillo-Correa *et al.*, 2015), and has thus been used in combination with

genetic markers to identify relevant conservation units in maritime pine (Rodríguez-Quilón *et*

al., 2016). In our study, we showed that height is not only correlated with climate but also

with biotic factors, such as pathogen susceptibility (positively for *D. sapinea* and negatively

for *A. ostoyae*). This is crucial information, for example, for the Landes maritime pine

breeding programme, which is based on artificial selection for height growth and straightness,

and that may have selected at the same time for higher susceptibility to *D. sapinea*, a

pathogen that is expected to increase its presence in maritime pine productive forests in the

near future (see below).

Growth phenology traits (date and duration of bud burst) showed low to moderate broad-

sense heritability, depending on the year (2015 or 2017). Differentiation (Q_{ST}) for bud burst

varied from 0.191 to 0.275, which is comparable to a mean of 0.249 for bud flush averaged

over several forest tree species (reviewed in Alberto *et al.*, 2013). In our study, trees

originating from northern populations flushed later than trees from southern populations.

Similar clines have been observed for other conifers (reviewed in Alberto *et al.*, 2013), which

is not surprising, as spring phenology, such as flushing time, is known to be correlated with

climatic factors (e.g. Zohner & Renner, 2014). Such clines in growth phenology along climate

gradients suggested local adaptation due to divergent selection, which might suffer mismatch

with changing climate conditions (Badeck *et al.*, 2004; Lindner *et al.*, 2010, Alberto *et al.*,

2013). Spring phenology can also play a role in resistance to or avoidance of forest tree

pathogens (e.g. Swedjemark *et al.*, 1998; Ghelardini & Santini, 2009; Nielsen *et al.*, 2017).

In line with this, we found a positive genetic correlation between needle discoloration and

necrosis length caused by *D. sapinea* with growth phenology indicating that earlier flushing

trees with faster spring growth showed less severe disease symptoms. Krokene *et al.* (2012)

showed that the concentrations of starch and total sugars (glucose, fructose and sucrose) in

twigs of *Picea abies* change during shoot development, which affects pathogen-related

symptoms. In our study, inoculations were carried out on twigs with already elongated

needles, however, the chemical composition of twigs might differ with time since bud burst

and we cannot fully discard these effects in our assessment of disease incidence.

Broad-sense heritability for susceptibility to *D. sapinea* based on necrosis length (H^2=0.152)

was much lower than e.g. narrow-sense heritability of lesion length caused by *Fusarium*

circinatum (h^2=0.618, Elvira-Recuenco *et al.*, 2014). However, this low but significant

heritability still indicated the potential for an evolutionary response of maritime pine to *D.*

sapinea. Moreover, Q_{ST} was highest for necrosis length caused by this pathogen (Q_{ST}=0.636)

in our study, which was also significantly higher than neutral-marker F_{ST}. This result pointed

to adaptive divergence and substantial variability in susceptibility to *D. sapinea* across the

geographical range of maritime pine. This population variability could be harnessed in

maritime pine breeding programmes to increase resistance to the disease (Bouffier *et al.*,

2008, Ingvarsson & Dahlberg, 2019). Remarkably, we observed a negative correlation of trait

variation across populations for necrosis length caused by *D. sapinea* and *A. ostoyae*. This

direct comparison should be taken with caution as the inoculation experiments were carried

out at different time points (*D. sapinea* in summer and *A. ostoyae* in autumn), which is likely

to affect the host response. However, correlation with climate across populations for

susceptibility to both *D. sapinea* and *A. ostoyae* (as well as for other adaptive traits such as

height and growth phenology) indicated similar environmental clines driving differentiation at

these traits. Especially, maximum temperatures during summer months showed significant

correlations with genetic variability of susceptibility to pathogens across maritime pine

populations. Trees from populations with high maximum summer temperatures were less

susceptible to *D. sapinea*. This result can be interpreted in different ways: 1) If we assume

that *D. sapinea* is native in Europe, the pathogen pressure can be expected to be stronger

in southern regions, with a climate more favourable to *D. sapinea* pathogenic outbreaks,

triggered by stress in the host plant, especially by droughts (Luchi *et al.*, 2014). Maritime pine

populations growing in these regions (e.g. Morocco and southern Spain) would then be more

likely to have evolved resistance to the disease. In contrast, trees from populations where

severe drought periods have most likely not been common so far (e.g. Atlantic populations

from the Iberian Peninsula and France) would be more susceptible. 2) If maritime pine and

D. sapinea have not had sufficient time to co-evolve (e.g. if *D. sapinea* was not native to

Europe) or pathogen pressure was not strong enough, differences in susceptibility among

maritime pine populations might be due to exaptation or ecological fitting, i.e. due to

correlated traits selected for other functions (Agosta & Klemens, 2008). Populations of

maritime pine strongly vary geographically in many traits related to growth and response to

drought, along the gradient from North Africa to the Atlantic regions of the Iberian Peninsula

and France (Correia *et al.*, 2008; Aranda *et al.*, 2010; Corcuera *et al.*, 2012; Gaspar *et al.*,

2013; de la Mata *et al.*, 2014). Some of these traits may indirectly influence their susceptibility

to pathogens, as observed here for *D. sapinea*. For example, faster growing maritime pine

trees from northern populations are known to invest more in inducible defences while slow

growing trees from southern populations invest more in constitutive defences (López-Goldar

et al., 2018). The positive correlation between height and necrosis length caused by *D.*

sapinea might indicate that constitutive defences confer better resistance to this pathogen in

the southern populations. Also, Meijón *et al.* (2016) showed that the metabolomes in needles

of maritime pine trees from populations with distinct geographic origin (notably Atlantic versus

Mediterranean provenances) were quite differentiated, with flavonoids showing a significant

correlation with the water regime of the population of origin. However, the expression of metabolites is organ specific (de Miguel *et al.*, 2016) and knowledge about secondary metabolites involved in resistance to *D. sapinea* is still lacking.

A study on the invasive pathogen *Fusarium circinatum*, which did certainly not co-evolve with maritime pine, also revealed a geographic cline in susceptibility, with Atlantic maritime pine populations showing less susceptibility than Moroccan populations (Elvira-Recuenco *et al.*, 2014). A similar pattern was observed for *A. ostoyae* in our study. Our results indicated that maritime pine from southwestern France, where *A. ostoyae* outbreaks have been frequently reported (Labbé *et al.*, 2015), may have developed some resistance or might show exapted resistance to the disease. Considering the absence of reports of *A. ostoyae* from the south of the Iberian Peninsula (Marxmüller & Guillaumin, 2005), which is in line with the species' preference for humid forest sites (Cruickshank *et al.*, 1997; Heinzelmann *et al.*, 2019), trees in Morocco and southern Spain have most-likely never co-evolved with this pathogen. However, a study by Guillaumin *et al.* (2005) on the mortality of potted maritime pine plants revealed an opposite pattern to the one found in our study, with the Landes population in Atlantic France being the most susceptible and the Moroccan population the least susceptible to *A. ostoyae*. Moreover, Zas *et al.* (2007) found much higher (narrow-sense) heritability (h^2=0.35) for mortality due to *A. ostoyae* in an infested progeny trial of maritime pine seedlings than the values of broad-sense heritability of necrosis length found in our study. *Armillaria ostoyae* is a root pathogen and a critical point during natural infestation that could be key for resistance mechanisms is the penetration into the root (Prospero *et al.*, 2004; Solla *et al.*, 2011; Labbé *et al.*, 2017b), as the pathogen grows faster once it enters the organism and reaches the cambium (Solla *et al.*, 2002). This step was bypassed in our inoculation protocol on excised branches. In the future, it would therefore be interesting to carry out inoculations on potted seedlings or young trees from range-wide maritime pine populations to confirm the patterns observed in this study.

Suitable strategies to evaluate susceptibility to *D. sapinea* and *A. ostoyae* will become

increasingly important as climate change increases pathogen pressure. Droughts are

expected to become more frequent throughout Europe (IPCC, 2014), which will most likely

trigger *D. sapinea* outbreaks also in regions where the pathogen has not caused severe

disease symptoms so far. Recently, a northward expansion of *D. sapinea* outbreaks in

Europe, probably driven by higher spring temperatures, has been reported; these outbreaks

have caused severe damage on *P. sylvestris* in Sweden and eastern Baltic countries

(Adamson *et al.*, 2015; Brodde *et al.*, 2019). Because some of the largest and most

productive maritime pine forests are located in the Atlantic distribution of the species (e.g.

the Landes region in southwestern France), our results suggest that the expected increase

of drought events in these populations will most likely cause severe damages due to their

high susceptibility to *D. sapinea*. In the case of *A. ostoyae*, the main threat resides in the

condition of the host. As mentioned before, a weaker host will be more susceptible to the

fungus, and future extreme weather events are bound to weaken trees, also increasing

pathogenicity of *A. ostoyae* (Kubiak *et al.*, 2017). A mathematical model predicted a drastic

northward shift of *A. ostoyae* in the Northwestern United States for the years 2061-2080,

leading to increased mortality of stressed and maladapted trees (Hanna *et al.*, 2016). As

suggested by this study, trees maladapted to new temperatures are also expected to be more

susceptible to biotic stress.

Genotype-phenotype association and candidate genes

By using a single-locus, two-steps approach, we revealed significantly associated loci for all

heritable traits under study. However, genotype effects were small, pointing to a highly

polygenic nature of the studied traits, as often reported for adaptive traits in forest trees. In

addition, for susceptibility to *D. sapinea* and *A. ostoyae*, no resistance alleles with major

effects were detected. Nevertheless, information on genotype-phenotype associations can

make genome-based phenotypic predictions more efficient by providing a set of SNPs that

are more closely linked to causal SNPs than those selected just based on genome location

(e.g. Westbrook *et al.*, 2013). We retrieved annotations from Plomion *et al.* (2016) and found

four non-synonymous SNPs significantly associated to duration of bud burst in 2015 (one

locus), bud burst in 2017 (two loci) and needle discoloration caused by *D. sapinea* (one locus)

under the additive genetic model (see Table 5). The potential function of these genes has to

be interpreted with caution as this information usually derives from studies in distantly related

model species (typically *Arabidopsis thaliana*). Nevertheless, the locus (BX679001-1418),

which was significantly associated to needle discoloration caused by *D. sapinea*, possibly

codes for a translation initiation factor, eIF-5, that has previously been reported to be involved

in pathogen-induced cell death and development of disease symptoms in *A. thaliana*

(Hopkins *et al.*, 2008). This gene deserves further attention in future studies addressing the

genetic control of adaptive traits in conifers. Also, differential gene expression analyses could

shed light on the genes and pathways involved in pathogen resistance in maritime pine, when

applied on inoculated and control trees growing under controlled conditions in greenhouses

or common gardens.

Based on a well-replicated clonal common garden and state-of-the-art genotyping

technology, we were able to study key adaptive traits in maritime pine and found evidence

for non-synonymous mutations underlying genetic variation for some of these traits.

Association studies for highly polygenic traits are still challenging. Lind *et al.* (2017) reported

an average of 236 SNPs (out of 116,231 tested, i.e. 0.203 %) associated to each of four

fitness-related traits in *Pinus albicaulis* by detecting signals of significantly higher covariance

of allele frequencies. In the near future, multilocus association methods should be used to

reveal genome wide loci with non-zero effects for polygenic traits in forest trees (Goldfarb *et*

al., 2013; De la Torre *et al.*, 2019).

Conclusions

In our study, we evaluated phenotypic variability and broad-sense heritability for height,

growth phenology, pathogen susceptibility and incidence of processionary moth nests in a

range-wide clonal common garden of maritime pine. We revealed strong genetic divergence

at several adaptive traits, especially height and necrosis length caused by *D. sapinea*, across

maritime pine populations. We have also shown that several adaptive traits in maritime pine

were genetically correlated and that population variation was often established along climatic

clines, in particular with maximum temperatures. The evolution of suits of functional traits

along environmental clines is a common pattern in nature (e.g. Chapin *et al.*, 1993; Reich *et*

al., 1996). Currently, locally adapted populations are challenged by changing climate

conditions, and emergent pests and pathogens expanding their range (Seidl *et al.*, 2017).

Susceptibility to *D. sapinea* was highest in the Atlantic maritime pine populations where it is

expected to cause severe outbreaks due to increased incidence of drought events in the

future. Correlated selection for increased susceptibility needs to be considered in breeding

programmes aiming at increasing height growth in this species. In addition, opposing trends

in pathogen susceptibility across maritime pine populations, e.g. for *D. sapinea* and *A.*

ostoyae (this study), and for the invasive pathogen *F. circinatum* (Elvira-Recuenco *et al.*,

2014), challenge forest tree breeding and natural forest resilience. An improved

understanding of integrated phenotypes, including responses to known pests and pathogens,

and their underlying genetic architecture is fundamental to assist new-generation tree

breeding and the conservation of valuable genotypes. Coupled with early detection methods

(see e.g. Kenis *et al.*, 2018), knowledge on genetic responses to emerging pests and

pathogens will help to ensure the health of forests in the future. Finally, given the recent

development of efficient technologies to combine functional genetic variants (e.g. genome

editing), the identification of large numbers of polymorphisms associated with commercial

traits is expected to contribute to a new era in plant breeding (Breeding 4.0, see Wallace *et*

al., 2018).

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Conflict of interests

The authors have no conflicts of interest to disclose.

Data availability: All data will be made available in the Dryad repository (https://datadryad.org) as soon as the manuscript has been accepted.

 Tables

 Table 1 Broad-sense heritability, variance ratios and genetic differentiation of adaptive traits in *Pinus pinaster*. Variability refers to the standard

 deviation of the raw phenotypic data. H^2, broad-sense heritability; H^2_p, variance ratio of the population effect (see main text for ratio definition);

 H^2_{gp}, variance ratio of the gene pool effect (see main text for ratio definition); Q_{ST}, population differentiation at quantitative traits; bb, bud burst;

 dbb, duration of bud burst; disc, needle discoloration; processionary, presence/absence of processionary moth nests; dd, degree-days. Heritability

 and variance ratios for incidence of the processionary moth were computed using height as a covariate. For *A. ostoyae* necrosis length, the

 population and gene pool levels were identical as only one population per gene pool has been analyzed. Values in bold are significant. Values in

 squared brackets indicate the 95% credible intervals.

	Phenotypic mean	Variability	H^2	H^2_p	H^2_{gp}	Q_{ST}
height (cm)	170.647	± 48.228	**0.219** [0.186-0.261]	**0.069** [0.028-0.132]	**0.254** [0.113-0.714]	**0.549** [0.392-0.662]
bb2015 (dd)	1311.95	± 82.411	**0.310** [0.272-0.358]	**0.058** [0.022-0.121]	**0.194** [0.06-0.581]	**0.275** [0.186-0.443]
dbb2015 (dd)	814.713	± 116.708	**0.165** [0.099-0.212]	0.040 [0.000-0.111]	0.000 [0.000-0.287]	**0.191** [0.100-0.332]
bb2017 (dd)	1286.245	± 79.853	**0.230** [0.181-0.261]	**0.031** [0.010-0.088]	0.066 [0.000-0.302]	**0.191** [0.106-0.404]
dbb2017 (dd)	901.149	± 78.922	**0.237** [0.195-0.297]	0.000 [0.000-0.063]	**0.231** [0.102-0.676]	**0.463** [0.293-0.579]
A. ostoyae necrosis (mm)	48.533	± 29.625	**0.021** [0.004, 0.121]	**0.066** [0.018, 0.203]	**0.066** [0.018, 0.203]	**0.217** [0.041-0.787]
D. sapinea necrosis (mm)	43.348	± 17.931	**0.152** [0.038-0.292]	0.000 [0.000-0.282]	0.002 [0.000-0.602]	**0.636** [0.349-1.000]

D. sapinea disc	*0 - no disc: 183* *1 - low: 123* *2 - medium: 141* *3 - high: 9*	0.123 [0.000-0.233]	0.001 [0.000-0.145]	0.002 [0.000-0.296]	0.093 [0.000-0.752]
Processionary	1 - presence: 48 0 - absence: 3282	0.001 [0.000-0.207]	0.001 [0.000-0.128]	-0.004 [0.000-0.380]	0.006 [0.000-0.985]

1233

Table 2 Pearson's correlation coefficients of the Best Linear Unbiased Predictors (BLUPs) of genotype (i.e. genetic correlations, above diagonal) and combined population and genotype effects (below diagonal). bb, bud burst; dbb, duration of bud burst; disc, needle discoloration. Significance levels after false discovery rate (FDR) correction: *<0.05; **<0.01; ***0.001.

	height	bb2015	dbb2015	bb2017	dbb2017	*D. sapinea* necrosis	*D. sapinea* disc	*A. ostoyae* necrosis
height		-0.172**	-0.132*	-0.178**	0.439***	0.111	0.082	-0.010
bb2015	0.194***		0.798***	0.392***	-0.292***	0.118	0.041	-0.017
dbb2015	0.117*	0.846***		0.380***	-0.233***	0.061	0.065	-0.110
bb2017	0.021	0.593***	0.547***		-0.154**	-0.023	0.069	-0.092
dbb2017	0.800***	0.147**	0.079	0.018		-0.032	-0.093	-0.030
D. sapinea. necrosis	0.679***	0.564***	0.511***	0.309***	0.639***		0.231*	-0.009
D. sapinea disc	0.414***	0.470***	0.492***	0.337***	0.392***	0.727***		-0.149
A. ostoyae necrosis	-0.653***	-0.321*	-0.284*	-0.056	-0.530***	-0.692***	-0.367**	

1240 **Table 3** Correlations between adaptive phenotypes and environmental variables, as identified using linear mixed-effects models that account for

1241 gene pool as a random factor. Only models that performed significantly better than a null model excluding the fixed effect are reported. R^2, the

1242 proportion of variance explained by the fixed predictor in the mixed-effects model. For necrosis length caused by *A. ostoyae*, a simple linear

1243 regression model was used as each gene pool was only represented by a single population. bb, bud burst; dbb, duration of bud burst; disc,

1244 needle discoloration; Tmax, maximum temperature; prec, precipitation.

1245

Trait	Environmental factor	R^2	P
height	altitude	0.650	<0.001
	Tmax July	0.709	<0.001
	Tmax August	0.201	0.023
bb2015	altitude	0.200	0.020
dbb2015	longitude	0.245	0.044
bb2017	longitude	0.262	0.028
dbb2017	latitude	0.290	0.042
	altitude	0.430	0.001
	Tmax July	0.212	0.044
	Tmax August	0.193	0.033
	prec in August	0.123	0.027
	prec in September	0.123	0.027
D. sapinea necrosis	latitude	0.341	0.031
	altitude	0.482	0.018
	Tmax July	0.465	0.034
	Tmax August	0.394	0.036
D. sapinea disc	latitude	0.284	0.013
	Tmax June	0.310	0.009
	Tmax July	0.524	0.004

	Tmax August	0.462	0.003
	Tmax September	0.315	0.004
A. ostoyae necrosis	latitude	0.750	0.016
	altitude	0.614	0.040

1246

Table 4 Single Nucleotide Polymorphisms (SNPs) significantly associated to height, growth phenology and pathogen susceptibility traits under

the additive genetic model, as identified by a two-step approach based on mixed-effects linear models (MLMs) implemented in Tassel and the

Bayesian framework in BAMD (BMLMs). Bayesian mean SNP-effects and 95% credible intervals (CIs) were obtained from the distribution of the

last 20,000 iterations in BAMD. Marker codes and linkage groups as reported in Plomion *et al.* (2016). bb, bud burst; dbb, duration of bud burst;

disc, needle discoloration; LG, linkage group; MAF, minimum allele frequency; unk, unknown; nc, noncoding; non-syn, non-synonymous; syn,

synonymous.

Trait	SNP name	SNP motif	Site annotation	LG	MAF	MLMs			BMLMs		
						F	P	R^2	Mean	lower 95% CI	upper 95% CI
height	BX249583-420	[A/G]	unk	1	0.266	7.383	0.001	0.027	3.956	1.162	6.762
	BX251999-509	[A/T]	unk	9	0.296	6.576	0.002	0.024	2.749	0.262	5.255
	CT2714-442	[T/A]	unk		0.381	5.683	0.004	0.021	-2.920	-5.335	-0.503
	F51TW9001AZG2W-933	[C/G]	unk	4	0.438	9.991	<0.001	0.036	3.480	0.931	6.083
	sp_v3.0_unigene17345-1191	[T/G]	nc	9	0.344	5.977	0.003	0.022	3.232	0.588	5.871
bb2015	BX249218-322	[A/C]	nc		0.315	6.499	0.002	0.03	7.230	2.107	12.440
	BX249671_307	[T/C]	unk	7	0.397	6.438	0.002	0.029	6.129	0.763	11.450
	BX253890-151	[T/G]	nc	12	0.157	6.350	0.002	0.029	10.95	4.298	17.552
	CL2033CT1302CN1398-513	[A/G]	nc	1	0.408	5.436	0.005	0.025	7.134	2.001	12.321
	CL544Contig1_03.Pipn-84	[T/G]	unk		0.135	9.107	<0.001	0.041	10.18	2.922	17.579
	FN692276-550	[A/G]	unk	12	0.402	7.903	<0.001	0.036	5.667	0.664	10.774
	i13066s710	[A/C]	nc		0.242	6.107	0.002	0.028	8.365	2.328	14.407
	i16267s380	[A/G]	unk	2	0.411	7.388	0.001	0.033	-10.55	-15.885	-5.168
	LP3-3-298	[C/G]	unk		0.143	5.538	0.004	0.025	8.510	1.696	15.302
bb2017	0_12730_01_contig1-159	[A/C]	unk	12	0.379	9.749	<0.001	0.044	4.003	1.509	6.499
	AL750545-695	[T/A]	non-syn	1	0.487	5.754	0.003	0.026	4.131	1.375	6.800
	AL750773_910	[T/A]	unk	3	0.499	5.580	0.004	0.026	-4.048	-6.810	-1.339
	BX249816-2143	[A/G]	non-syn	7	0.269	5.783	0.003	0.026	5.787	2.792	8.839

	CT576106-142	[C/G]	unk	1	0.180	7.147	0.001	0.032	5.590	2.165	9.029
	CT577489-1569	[A/C]	unk		0.203	5.826	0.003	0.026	-4.606	-7.559	-1.602
	F7JJN6E01B7BCW-157	[A/G]	syn	5	0.117	5.591	0.004	0.025	7.708	3.801	11.533
	FM945796-840	[T/G]	unk		0.213	6.284	0.002	0.028	-4.613	-7.748	-1.473
	i10996s1211	[T/C]	unk		0.301	8.281	<0.001	0.037	4.295	1.467	7.149
dbb2015	CL2507CT3369CN3610-193	[T/A]	nc	1	0.387	5.406	0.005	0.025	8.981	2.203	15.730
	F51TW9001BWV4H-219	[T/C]	non-syn		0.462	5.941	0.003	0.028	7.101	0.458	13.72
	F51TW9002FPGRE-170	[A/G]	nc		0.346	5.510	0.004	0.026	9.467	2.507	16.580
dbb2017	AL749850_679	[A/G]	unk		0.402	5.449	0.005	0.021	-4.992	-9.318	-0.654
	BX681656-960	[A/C]	syn	5	0.481	8.050	<0.001	0.030	-6.416	-10.154	-2.669
	CT580064-331	[A/G]	unk		0.442	6.236	0.002	0.024	-4.363	-7.934	-0.822
	CT582680-451	[A/C]	unk		0.201	7.383	0.001	0.028	-10.48	-15.541	-5.470
	F51TW9001BAW7V-405	[A/G]	unk	12	0.163	9.088	<0.001	0.034	7.287	2.665	11.854
	i17647s350pg	[G/C]	unk		0.157	5.847	0.003	0.022	5.747	0.878	10.611
D. sapinea disc	BX679001-1418	[T/C]	non-syn	7	0.192	5.544	0.005	0.049	-0.054	-0.102	-0.006
	CT574726_30	[T/C]	unk		0.435	5.539	0.005	0.049	-0.051	-0.089	-0.014
	i11062s233	[G/C]	syn	7	0.293	5.596	0.005	0.050	-0.048	-0.090	-0.007
D. sapinea necrosis length	AL750492_595	[T/C]	unk	1	0.121	5.887	0.004	0.033	0.897	0.081	1.713
	F51TW9002FT2ZF-1060	[T/G]	unk	12	0.485	8.614	<0.001	0.048	-1.279	-1.992	-0.574
A. ostoyae necrosis length	F51TW9001AI9YZ_1847	[A/G]	unk	7	0.273	5.928	0.005	0.081	-0.735	-1.345	-0.129
	F51TW9001CXU1D_1264	[T/C]	unk	6	0.364	6.594	0.003	0.09	-0.997	-1.787	-0.225

1253

1254   **Table 5** Annotation for SNPs significantly associated with adaptive traits under the additive model and coding for a non-synonymous amino acid

1255   change, as retrieved from Plomion *et al.* (2016). bb, bud burst; dbb, duration of bud burst; disc, needle discoloration.

1256

Trait	SNP name	Motif	Protein change	Putative protein function
dbb2015	F51TW9001BWV4H-219	[T/C]	Asparagine - Serine	LANC-like domain containing protein
bb2017	BX249816-2143	[A/G]	Glutamate/Glutamine - Arginine	putative Cdc2-related protein kinase CRK2
bb2017	AL750545-695	[A/T]	Glutamate/Glutamine - Valine	Catalase
D. sapinea disc.	BX679001-1418	[T/C]	Isoleucine - Valine	Translation initiation factor eIF-5

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1258

 Figure captions

Figure 1 Stripchart of combined genotype and population BLUPs (Best Linear Unbiased

Predictors) for necrosis length caused *by D. sapinea* (a) and *A. ostoyae* (b). Populations

were assigned to one of the six gene pools identified by Jaramillo-Correa *et al.* (2015),

corresponding to the six colours in the figure, and ordered by latitude (north to south).

Symbols indicate different populations of the same gene pool. Black lines indicate the

average necrosis length in each population.

Figure 2 Correlation of trait variation across *Pinus pinaster* populations based on BLUPs

(Best Linear Unbiased Predictors) for necrosis length caused by *Diplodia sapinea* and

Armillaria ostoyae. A linear trend line is also shown (Pearson's correlation coefficient = -

0.692, p-value<0.001).

Figure 3 Stripchart of combined genotype and population BLUPs (Best Linear Unbiased

Predictors) for necrosis length caused *by D. sapinea* plotted against the maximum

temperature in July in the population of origin. Populations were assigned to one of the six

gene pools identified by Jaramillo-Correa *et al.* (2015), corresponding to the six colours in

the figure. Symbols indicate different populations of the same gene pool. Black lines indicate

the average necrosis length in each population.

Figure 4 Density plots of the effect sizes based on 20,000 BAMD simulations (left) and

genotypic effects (box plots, right) for two exemplary Single Nucleotide Polymorphisms

(SNPs) showing significant association with needle discoloration caused by *Diplodia sapinea*

(above) and duration of bud burst in 2015 (below) in *Pinus pinaster*.

Figure 5 Minor allele frequency distribution of SNP BX679001-1418 in natural populations

of *Pinus pinaster*. This locus was significantly associated to needle discoloration caused by

Diplodia sapinea, and possibly coded for a disease-induced translation initiation factor, eIF-

5.

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Supporting Information

Table S1.1 Number of genotypes from the CLONAPIN clonal common garden used to study adaptive traits in *Pinus pinaster.*

Figure S2.1 Phenological stages of bud burst.

Protocol S3.1 Laboratory protocol for *Diplodia sapinea* inoculation.

Figure S3.1 Pictures of the four scales of needle discoloration found along the necrosis caused by artificial *Diplodia sapinea* inoculation on excised branches of *Pinus pinaster*.

Protocol S3.2 Laboratory protocol for *Armillaria ostoyae* inoculation.

Table S4.1 MCMCglmm Bayesian model parametrization.

Figure S5.1 Stripcharts of trait Best Linear Unbiased Predictors (BLUPs) for each genotype by population.

Table S6.1 (see separate excel-file: TableS6.1.pdf) All significant genotype effects (including additive, dominance and overdominance effects) of Single Nucleotide Polymorphisms (SNPs, minor allele frequency (MAF) > 0.1) on height, growth phenology and pathogen susceptibility traits identified by a two-step approach based on mixed-effects linear models (MLMs) implemented in Tassel and the Bayesian framework in BAMD (BMLMs). Bayesian mean SNP-effects and 95% credible intervals (CIs) were obtained from the distribution of the last 20,000 iterations in BAMD. Marker names and linkage groups (LG) as reported in Plomion *et al.* (2016). Site annotations: nc, non-coding (untranslated regions or introns); non-syn, non-synonymous; syn, synonymous; unk, unknown. *N*, number of phenotypic observations included in the analyses.

Figure S6.1 Density plots of the effect sizes based on 20,000 BAMD simulations (left) and genotypic effects (box plots, right) for two non-synonymous Single Nucleotide Polymorphisms showing significant association with bud burst in 2017.

Figure 1 Stripchart of combined genotype and population BLUPs (Best Linear Unbiased Predictors) for necrosis length caused *by D. sapinea* (a) and *A. ostoyae* (b). Populations were assigned to one of the six gene pools identified by Jaramillo-Correa et al. (2015), corresponding to the six colours in the figure, and ordered by latitude (north to south). Symbols indicate different populations of the same gene pool. Black lines indicate the average necrosis length in each population.

Figure 2 Correlation of trait variation across *Pinus pinaster* populations based on BLUPs (Best Linear Unbiased Predictors) for necrosis length caused by *Diplodia sapinea* and *Armillaria ostoyae*. A linear trend line is also shown (Pearson's correlation coefficient = -0.692, p-value<0.001).

Figure 3 Stripchart of combined genotype and population BLUPs (Best Linear Unbiased Predictors) for necrosis length caused *by D. sapinea* plotted against the maximum temperature in July in the population of origin. Populations were assigned to one of the six gene pools identified by Jaramillo-Correa *et al.* (2015), corresponding to the six colours in the figure. Symbols indicate different populations of the same gene pool. Black lines indicate the average necrosis length in each population

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Figure 5 Minor allele frequency distribution of SNP BX679001-1418 in natural populations of *Pinus pinaster.* This locus was significantly associated to needle discoloration caused by *Diplodia sapinea*, and possibly coded for a disease-induced translation initiation factor eIF-5.