1	Global Effects of Feature-based Attention		
2	Depend on Surprise		
3	(Abbreviated title: Feature-based Attention Depends on Surprise)		
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Abstract

25 Recent studies have shown that prediction and attention can interact under various 26 circumstances, suggesting that the two processes are based on interdependent neural 27 mechanisms. In the visual modality, attention can be deployed to the location of a task-relevant 28 stimulus ('spatial attention') or to a specific feature of the stimulus, such as colour or shape, 29 irrespective of its location ('feature-based attention'). Here we asked whether predictive 30 processes are influenced by feature-based attention outside the current spatial focus of 31 attention. Across two experiments, we recorded neural activity with electroencephalography 32 (EEG) as human observers performed a feature-based attention task at fixation and ignored a 33 stream of peripheral stimuli with predictable or surprising features. Central targets were 34 defined by a single feature (colour or orientation) and differed in salience across the two 35 experiments. Task-irrelevant peripheral patterns usually comprised one particular conjunction 36 of features (standards), but occasionally deviated in one or both features (deviants). Consistent 37 with previous studies, we found reliable effects of feature-based attention and prediction on 38 neural responses to task-irrelevant patterns in both experiments. Crucially, we observed an 39 interaction between prediction and feature-based attention in both experiments: the neural 40 effect of feature-based attention was larger for surprising patterns than it was for predicted 41 patterns. These findings suggest that global effects of feature-based attention depend on 42 surprise, and are consistent with the idea that attention optimises the precision of predictions 43 by modulating the gain of prediction errors.

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Significance Statement

45 Two principal mechanisms facilitate the efficient processing of sensory information: *prediction*46 uses prior information to guide the interpretation of sensory events, whereas *attention* biases

47 the processing of these events according to their behavioural relevance. A recent theory 48 proposes to reconcile attention and prediction under a unifying framework, casting attention as 49 a 'precision optimisation' mechanism that enhances the gain of prediction errors. Crucially, 50 this theory suggests that attention and prediction interact to modulate neural responses, but this 51 hypothesis remains to be tested with respect to feature-based attention mechanisms outside the 52 spatial focus of attention. Here we show that global effects of feature-based attention are 53 enhanced when stimuli possess surprising features, suggesting that feature-based attention and 54 prediction are interdependent neural mechanisms.

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Introduction

56 Selective attention mechanisms enhance the processing of sensory stimuli that are relevant for 57 guiding behaviour (Desimone & Duncan, 1995; Posner, 1994). Visual processing can be biased toward stimuli at a relevant location, commonly known as 'spatial attention', or toward stimuli 58 59 that possess task-relevant features, known as 'feature-based attention' (Carrasco, 2011). Monkey neurophysiology studies (Martinez-Trujillo & Treue, 2004) and human neuroimaging 60 61 studies (Gledhill, Grimsen, Fahle, & Wegener, 2015; Saenz, Buracas, & Boynton, 2002; 62 Serences & Boynton, 2007) have demonstrated that neural responses to stimuli at taskirrelevant locations are enhanced when they possess task-relevant features, demonstrating that 63 64 the effects of feature-based attention are global and dissociable from those of spatial attention. 65 In humans, neural responses to visual stimuli at task-irrelevant locations can be enhanced when they possess surprising features (e.g., colour, orientation, motion), demonstrating that top-66 67 down 'prediction' mechanisms also exert a global effect on incoming sensory signals (Friston, 68 2005; Stefanics, Kremlácek, & Czigler, 2014). At present, it is unknown whether the global effects of visual feature-based attention can interact with those of prediction. Here we used 69

electroencephalography (EEG) to measure neural responses to peripheral visual stimuli that
were predictable or surprising along two feature dimensions (orientation and colour), and tested
whether attending to a particular feature at fixation modulated the effect of prediction on neural
responses to peripheral stimuli at task-irrelevant locations.

Predictive coding theories propose that top-down prediction signals effectively 'silence' 74 75 bottom-up sensory signals that match the predicted content, leaving only the remaining 76 prediction error to propagate forward and update a model of the sensory environment (Friston, 2005; Rao & Ballard, 1999). In addition to predicting the *content* of sensory signals, an optimal 77 78 inference system should also estimate the level of uncertainty about its predictions (i.e., inverse 79 precision; Hohwy, 2012). Recently, it has been proposed that selective attention mechanisms 80 fulfil this role, optimising the expected precision of predictions by enhancing the activity of 81 units encoding prediction errors for attended stimuli (Feldman & Friston, 2010; Friston, 2009, 82 2010). Recent studies have supported this theory by demonstrating that selective attention and 83 prediction can interact under various circumstances (Auksztulewicz & Friston, 2015; Jiang, 84 Summerfield, & Egner, 2013; Kok, Rahnev, Jehee, Lau, & De Lange, 2012; Marzecová, 85 Widmann, SanMiguel, Kotz, & Schröger, 2017; Smout, Tang, Garrido, & Mattingley, 2019). 86 However, selective attention mechanisms encompass distinct information-processing 87 subcomponents (e.g., spatial attention, temporal attention) across sensory modalities (e.g., 88 auditory, visual) and it is important to establish which of these subcomponents interacts with 89 prediction and in what manner. In the visual domain, previous studies that reported an 90 interaction between attention and prediction typically presented stimuli at task-relevant 91 locations (Jiang et al., 2013; Kok, Rahnev, et al., 2012; Marzecová et al., 2017; Smout et al., 92 2019). One previous study found an effect of feature-based attention on mismatch responses to 93 stimuli at task-irrelevant locations, but this study presented clearly visible targets that likely

did not necessitate a tight focus of spatial attention on the central stimulus stream (Czigler &
Sulykos, 2010). Thus, it remains unclear whether prediction can interact with global featurebased attention mechanisms that modulate neural responses to stimuli outside the spatial focus
of attention.

98 Here we tested whether feature-based attention modulates the effect of prediction at task-99 irrelevant locations by comparing event-related potentials evoked by peripheral stimuli that 100 either matched ('congruent') or mismatched ('incongruent') a cued feature of the target at fixation. Participants searched for targets at fixation while predictable or surprising task-101 102 irrelevant stimuli were presented in the periphery. We conducted two experiments that differed 103 in the salience of central targets and distractors to investigate whether the strength of the top-104 down feature-set modulates the neural interaction between feature-based attention and 105 prediction.

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Methods

107 Experiment 1: Effects of feature-based attention at fixation on neural responses to

108 predicted and surprising peripheral stimuli

109 Participants

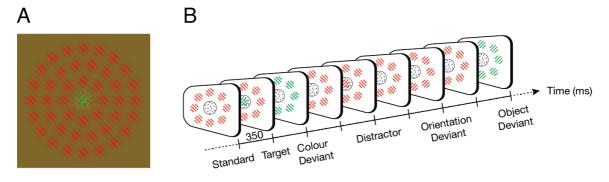
110 Twenty-four healthy adults (13 female, 11 male, age = 22.08 ± 2.38 years) with normal or 111 corrected-to-normal vision were recruited for participation via an online portal at The 112 University of Queensland. The study was approved by The University of Queensland Human 113 Research Ethics Committee, and all participants provided written, informed consent before 114 commencing the experiment.

115 Stimuli and apparatus

116 Participants were positioned at a viewing distance of 57 cm and seated in a comfortable 117 armchair in an electrically shielded laboratory. Stimuli were presented on a 61 cm LED monitor 118 (Asus, VG248QE) with a 1920 x 1080 pixel resolution and refresh rate of 120 Hz, using 119 PsychToolbox presentation software (Kleiner, M., Brainard, D., Pelli, D., Ingling, A., Murray, 120 R., & Broussard, 2007) for Matlab (v.15b) running under Windows 7 with an NVidia Quadro 121 K4000 graphics card. The intensity of the green phosphor was adjusted per participant to 122 produce subjective equiluminance with that of the red phosphor at full intensity. The 123 equiluminance point was determined prior to the experiment using minimum motion photometry (Anstis & Cavanagh, 1983), with intensity values determined by two interleaved 124 125 adaptive staircases (1 up-1 down, stopping after 15 reversals).

126 Central and peripheral stimuli were sinusoidal Gabors (diameter: 4.72°, spatial frequency: 0.94 c/°, 100% contrast) with one of two orientations (tilted 45° clockwise or counterclockwise of 127 128 vertical) and one of two colours (red or green). Central stimuli were superimposed over a red-129 green noise patch (diameter: 4.72) and onset every 700 - 1400 ms for 66.67 ms. Twenty percent 130 (20%) of central stimuli were targets (approximately 28 targets and 114 distractors per block). 131 Multi-element peripheral stimuli ('patterns') were arranged in three concentric circles (radii: 4.72°, 8.49°, 12.26°; containing 8, 14, and 20 Gabors, respectively; Figure 1A). Peripheral 132 patterns were presented every 350 ms for 66.67 ms (428 events per block) on top of a 133 134 background that alternated between uniform red and green pixels at the screen refresh rate (120 135 Hz), producing a subjective percept of a uniform brown background. During each block, peripheral patterns were more likely to contain one of the four possible feature conjunctions 136 137 (e.g. clockwise-tilted red Gabors; 76% of presentations, *standards*), with the other three feature

- 138 conjunctions being rare and of equal likelihood (8% each, *deviants*). Standards were pseudo-
- 139 randomized across blocks, and the order of deviants was pseudo-randomized within blocks.



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141 Figure 1. Stimulus display and task. (A) Stimulus display, showing a green clockwise-tilted target 142 within the central noise patch and a peripheral pattern consisting of three concentric rings of red 143 counterclockwise-tilted Gabors. (B) Simplified task diagram. In this example trial, participants 144 monitored for green Gabors (targets) and ignored red Gabors (distractors) within the central noise 145 patch. Peripheral patterns typically contained red counterclockwise-tilted Gabors (standards), but 146 occasionally contained red clockwise-tilted Gabors (orientation deviant), green counterclockwise-tilted 147 Gabors (colour deviant), or green clockwise-tilted Gabors (object deviant, i.e. deviating in both colour 148 and orientation). In this example trial, the colour and object deviants shared features with the target 149 (i.e., green) and would thus be labelled 'congruent'.

150 *Procedure*

Participants were asked to fixate on a central dot and click a mouse button as soon as they detected a target in the stream of central stimuli, continuously throughout blocks (duration: 150 s), while ignoring central distractors and peripheral patterns (*Figure 1B*). In each block, Gabor targets were designated as either (1) red, (2) green, (3) clockwise-tilted, or (4) counterclockwise-tilted. Note that each condition dictated two of the four possible feature conjunctions as targets and two as distractors (e.g., if searching for clockwise-tilted targets, both red and green clockwise-tilted Gabors were valid targets).

- 158 Participants completed two practice blocks with auditory feedback after each response, before
- being fitted with the EEG cap and electrodes (see *EEG Data Acquisition*). Participants then
- 160 completed 16 test blocks with target type and standard pattern features pseudorandomized

across blocks (6848 peripheral patterns per session). Feedback on mean reaction time and the
number of hits and false alarms was provided between blocks.

163 Behavioural Data Analysis

164 We investigated whether the feature-congruence and predictability of peripheral patterns 165 affected participants' detection of central targets. Targets were sorted into prediction 166 conditions according to whether the preceding pattern (i.e., the peripheral stimulus presented 167 up to 700 ms prior to peak target contrast) was a standard ('predicted') or a deviant ('surprising'), and *feature-congruence conditions* according to whether the preceding 168 169 peripheral pattern matched the features of the central target ('congruent') or distractor 170 ('incongruent'). Participant responses were scored as hits if they occurred within 1 s of the 171 onset of a target. Successive responses within this window were ignored, as were any responses 172 that occurred within 250 ms of a preceding response. Because we observed differences in hit rates and reaction times between target feature conditions (i.e., the feature that participants 173 174 searched for at fixation, e.g. 'red'), we first normalised hit rates and reaction times within each 175 target feature condition, separately for feature-congruence and prediction conditions, and then 176 collapsed across the target feature conditions. The resulting normalised hit rates and reaction 177 times were then subjected to two-way repeated measures ANOVAs to assess the effects of 178 peripheral pattern prediction (two levels: predicted, surprising) and feature-congruence (two 179 levels: congruent, incongruent) on target detection.

180 *EEG Data Acquisition*

Participants were fitted with a 64 Ag-AgCl electrode EEG system (BioSemi Active Two:
Amsterdam, Netherlands). Continuous data were recorded using BioSemi ActiView software
(http://www.biosemi.com), and were digitized at a sample rate of 1024 Hz with 24-bit A/D

184 conversion and a .01 - 208 Hz amplifier band pass. All scalp electrode offsets were adjusted 185 to below $20\mu V$ prior to beginning the recording. Pairs of flat Ag-AgCl electro-oculographic 186 electrodes were placed on the outside of both eyes, and above and below the left eye, to record 187 horizontal and vertical eye movements, respectively.

188 EEG Preprocessing

189 EEG recordings were processed offline using the EEGlab toolbox in Matlab (Delorme & 190 Makeig, 2004). Data were resampled to 256 Hz and high-pass filtered with a passband edge at 191 0.5 Hz (1691-point Hamming window, cut-off frequency: 0.25 Hz, -6 db). Raw data were 192 inspected for the presence of faulty scalp electrodes (none were found). To clean the data, we 193 applied an iterative process of artifactual epoch and component rejection using independent 194 component analyses (ICA). The data were segmented into 350 ms epochs surrounding Gabor 195 onsets (50 ms pre- and 300 ms post-stimulus) and baseline activity prior to stimulus onset was 196 removed from each epoch. Epochs were subjected to ICA, and the SASICA plugin for EEGlab 197 (Chaumon, Bishop, & Busch, 2015) was used to identify blink, saccade, and focal trial 198 components. Epochs were rejected if they met any of the following criteria: (1) blink 199 component activity greater than $\pm 10 \ \mu V$ between -50 and 150ms; (2) saccade component 200 activity greater $\pm 5 \,\mu V$ between 0 and 350 ms; (3) focal component activity exceeding a joint 201 probability threshold of ± 7 SD (5.5% of epochs were removed due to blink, saccade, or focal 202 activity). The remaining epochs were then subjected to ICA for a second time, and SASICA 203 was used again to identify artifactual components.

For further analysis, the resampled raw data were band-pass filtered between 0.5 and 40 Hz (1691-point Hamming window, cut-off frequencies: 0.25 and 40.25 Hz, -6 db) and segmented into 550 ms epochs surrounding Gabor onsets (100 ms pre- and 450 ms post-stimulus). Epochs

207 containing artefacts (identified previously using the first ICA) were removed. Independent 208 component weights from the second ICA were applied to this new dataset and artefactual 209 components (identified previously using the second ICA) were removed. Baseline activity in 210 the 100 ms prior to stimulus onset was removed from each epoch.

211 Event-Related Potential and Bayes Factor Analyses

212 Peripheral patterns were sorted into prediction conditions based on whether they were 213 standards (repeated at least 4 times; 'predicted') or deviants ('surprising'), and attention 214 conditions based on whether they shared features with central targets ('congruent') or 215 distractors ('incongruent') in the central task. Trials in each attention and prediction condition 216 were averaged within participants to produce event-related potentials (ERPs) for each 217 individual. Statistical analyses of condition ERPs were conducted using two-tailed cluster-218 based permutation tests across participants (Monte-Carlo distribution with 5000 permutations, 219 $p_{cluster} < 0.05$; sample statistic: dependent samples *t*-statistic, aggregated using the maximum 220 sum of significant adjacent samples, p_{sample} <.05) in the Fieldtrip toolbox for Matlab 221 (Oostenveld, Fries, Maris, & Schoffelen, 2011). Statistical analyses of univariate condition 222 averages were conducted using paired-samples *t*-tests and Bayesian analyses. The Bayes factor 223 analyses allowed for quantification of evidence in favour of either the null or alternative 224 hypothesis, with $BF_{10} > 3$ indicating substantial support for the alternative hypothesis and BF_{10} 225 < 0.33 indicating substantial support for the null hypothesis. Bayes factors were computed 226 using the Dienes (2014) calculator in Matlab.

227 Experiment 2: Replication with individually thresholded manipulation of feature-based 228 attention at fixation

In Experiment 1, the high contrast targets were detected at near-ceiling levels. To investigate whether the neural interaction between feature-based attention and prediction is sensitive to the strength of the top-down feature set, we conducted a second experiment in which central targets and distractors were individually thresholded to be less salient. Except for the minor methodological differences noted below, Experiment 2 was the same as Experiment 1 and thus afforded an opportunity to replicate the original results in a separate group of participants.

235 *Methods*

236 A new cohort of 24 healthy adults with normal (or corrected-to-normal) vision was recruited 237 to participate in Experiment 2 (12 female, 12 male, age = 22.17 ± 2.88 years, mean \pm SEM). 238 The stimuli and apparatus were identical to those used in Experiment 1 (Figure 1), except that 239 in Experiment 2 the central stimuli (targets and distractors) were presented at lower contrast 240 and with a sinusoidal onset and offset profile (total duration: 700 ms). The peak contrast of the 241 central stimuli was determined during the two practice blocks, using a transformed and 242 weighted up/down adaptive staircase configured to approximate 83% detection of targets 243 (up/down step ratio: 1/3, up/down size ratio: .1/.07; García-Pérez, 1998). Blocks lasted for 150 244 s (as per Experiment 1) for all except two participants, for whom blocks lasted for 120 s (due 245 to time constraints for these two individuals). Participant responses were scored as hits if they 246 occurred within 1 s of the peak target contrast (i.e., within 1.35 s of target onset, accounting 247 for the 350 ms on-ramp). During EEG preprocessing, we interpolated 11 faulty electrodes 248 (across 5 participants) using the average activation across neighbouring electrodes (defined by 249 the EEGlab Biosemi 64 template) and removed 4.1% of epochs due to blink, saccade, or focal 250 component contamination.

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Results

252 Experiment 1: Effects of feature-based attention at fixation on neural responses to

253 predicted and surprising peripheral stimuli

254 Feature-Congruent Peripheral Patterns Interfere with Target Detection at Fixation

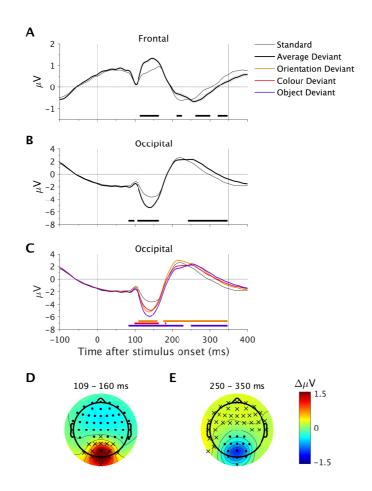
255 We first asked whether the congruence between peripheral pattern and central target features 256 affected participants' detection of central targets shortly after pattern onset. There was no 257 significant main effect of feature-congruence on normalized hit rates (congruent = $94.41 \pm$ 258 1.28%, 0.06 ± 0.19 z-normalised, incongruent = $93.63 \pm 1.39\%$, -0.08 ± 0.20 z-normalised, 259 F(1,23) = 3.33, p = .081, $\eta_p^2 = .013$). There was a significant main effect of feature-congruence on normalised reaction times, however, with participants responding more slowly to central 260 261 targets preceded by congruent peripheral patterns (438.80 ± 9.43 ms, mean \pm SEM; 0.04 ± 0.19 262 z-normalised) than to those preceded by incongruent peripheral patterns (436.09 ± 9.53 ms; - 0.04 ± 0.20 z-normalised, F(1,23) = 5.70, p = .026, $\eta_p^2 = 0.20$). This finding suggests that 263 participants were more distracted by peripheral patterns with task-relevant (congruent) 264 features, relative to those with task-irrelevant (incongruent) features, and is consistent with the 265 theory that involuntary orienting to task-irrelevant stimuli is contingent on attentional control 266 267 settings (Folk, Remington, & Johnston, 1992).

268 Peripheral Pattern Prediction Does Not Affect Target Detection

In a second analysis we asked whether the predictability of peripheral patterns affected behavioural responses to subsequent central targets. There was no significant effect of peripheral pattern prediction on normalised hit rates (predicted = 94.39 ± 1.22%, 0.04 ± 0.18, surprising = 93.65 ± 1.47%, -0.06 ± 0.21, F(1,23) = 1.63, p = .215, $\eta_p^2 = .07$) or normalised reaction times (predicted = 437.12 ± 9.66 ms, -0.01 ± 0.20, surprising = 437.76 ± 9.32 ms, 0.01 274 $\pm 0.19, F(1,23) = 0.27, p = .605, \eta_p^2 = .01)$, and no interaction between prediction and feature-275 congruence on either normalised hit rates ($F(1,23) = 0.31, p = .582, \eta_p^2 = .01$) or normalised 276 reaction times ($F(1,23) = 0.15, p = .701, \eta_p^2 = .01$). These findings suggest that the 277 predictability of peripheral patterns did not modulate the extent to which participants were 278 distracted from their task at fixation.

279 Prediction Decreases Neural Activity

We next assessed the main effect of prediction on neural activity by comparing ERPs to 280 peripheral deviant patterns ('surprising' patterns, collapsed across orientation, colour, and 281 282 object deviants) and standard patterns that had been repeated at least 4 times ('predicted' 283 patterns). Relative to baseline, standards evoked smaller neural responses than deviants (Figure 284 2). Over posterior electrodes, the neural response to standards was significantly reduced 285 relative to deviants during both the early negative deflection (i.e. standards > deviants; 82 - 164 286 ms, p = .020) and the late positive deflection (i.e. deviants > standards < deviants; 242 - 348ms, p = .010; Figure 2B). Over frontal electrodes, the neural response to standards was 287 significantly reduced relative to deviants during both the early positive deflection (i.e. 288 289 standards < deviants; 90 - 230 ms, p < .001) and the late negative deflection (i.e. standards > deviants; 254 - 348 ms, p = .008; Figure 2A). These effects are consistent with the theory that 290 291 surprising stimuli (deviants) produce greater prediction errors than predicted stimuli (standards; Friston, 2005, 2009; Rao & Ballard, 1999). 292



293

294 Figure 2. Main effect of prediction in Experiment 1. (A-B) ERPs evoked by standards and deviants 295 (collapsed across deviant types) at frontal electrodes (Fz, F1, F3, AFz, AF3, AF4; A) and occipital 296 electrodes (Oz, O1, O2, POz, PO3, PO4; B). Shading indicates the within-subject standard error of 297 the mean, calculated relative to standards. Black bars along the x-axis denote significant timepoints at 298 the displayed electrodes (cluster-corrected). (C) ERPs evoked by standards and each of the three 299 deviant conditions. Shading indicates the within-subject standard error of the mean, calculated 300 separately for each deviant condition relative to standards. Yellow, red and purple bars along the x-301 axis denote significant differences between standards and each corresponding deviant condition 302 (cluster-corrected). (D-E) Headmaps show the effect of prediction (standard minus average deviant) 303 during the indicated time windows. Asterisks and dots denote electrodes with larger or smaller 304 responses, respectively, across at least 33% of the averaged time points (cluster-corrected).

305 We followed up this result with direct comparisons between standards and each type of deviant,

306 which revealed similar effects to those reported above for the average deviant condition

307 (*Figure 2C*). Early posterior negativities were smaller in response to standards than orientation

deviants (109 - 160 ms, p = .033), colour deviants (98 - 164 ms, p = .040), and object deviants

 $309 \quad (82 - 348 \text{ ms}, p < .001)$, and late posterior positivities were significantly smaller in response to

standards than orientation deviants (176 - 348 ms, p < .001) and object deviants (250 - 348 ms)

ms, p = .019). Early frontal positivities were smaller in response to standards than orientation deviants (98 – 164 ms, p = .002), colour deviants (86 – 238 ms, p = .001), and object deviants (102 – 238 ms, p < .001), and late frontal negativities were smaller in response to standards than orientation deviants (242 – 348 ms, p < .001) and object deviants (84 – 348 ms, p < .001).

315 Visual Mismatch Negativities Are Additive Across Feature Deviations

316 Because previous investigations have suggested that the visual mismatch negativity (vMMN) 317 is non-additive across feature deviations (Czigler & Sulykos, 2010), we also tested for differences between vMMNs evoked by each type of deviant (orientation, colour, or object). 318 319 We used a data-driven approach to identify spatiotemporal samples (electrodes x timepoints) 320 that were significantly different from standards in all three deviant conditions (electrodes: Pz, 321 P1, P2, P3, P4, POz, PO3, PO4, Oz, O1, O2, Iz; timepoints: 109 – 160 ms) and then averaged 322 across these samples to produce one amplitude value per deviant condition and participant. We 323 then compared each pair of deviant conditions with paired-samples *t*-tests and Bayesian analyses, using a uniform prior with upper and lower bounds set to the average vMMN 324 325 amplitude. As can be seen in *Figure 2C*, there was no difference between the orientation (-1.02) $\pm 0.13 \ \mu\text{V}$) and colour vMMN (-0.98 \pm Z0.14 μV , t(23) = -0.35, p = .733, $BF_{10} = 0.14$). In 326 327 contrast, the object vMMN (-1.64 \pm 0.16 μ V) was significantly larger than both the orientation vMMN (t(23) = -5.39, p < .001, $BF_{10} = 2.4 \times 10^5$) and the colour vMMN (t(23) = -7.06, p < 0.001328 .001, $BF_{10} = 6.3 \times 10^9$), suggesting that the vMMN is sensitive to features of the deviant 329 330 stimulus.

331 Feature-based Attention Decreases Neural Activity

We assessed the main effect of feature-based attention by comparing ERPs to peripheral patterns that shared features with targets ('congruent') or distractors ('incongruent') in the

- 334 central detection task. Congruent peripheral patterns evoked a smaller positivity over posterior
- electrodes than incongruent patterns late in the epoch (188 305 ms, p = .004; *Figure 3B*,*D*).
- 336 This effect was matched by a polarity-reversed activity profile over frontal electrodes (191 –
- 337 309 ms, p = .003; Figure 3A,D).

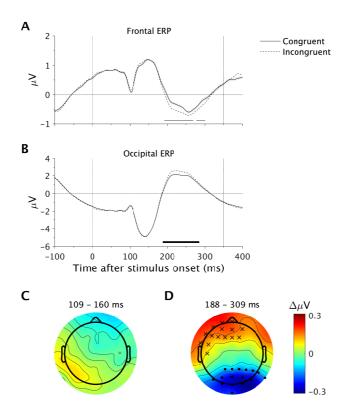




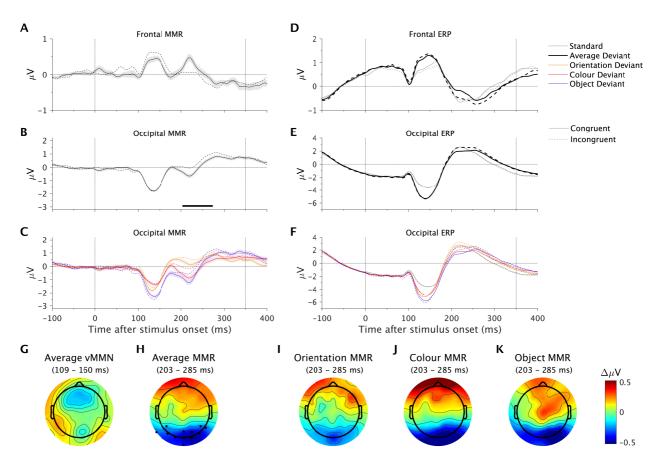
Figure 3. Main effect of feature-based attention in Experiment 1. (A-B) Congruent and incongruent ERPs are collapsed across prediction conditions separately for frontal electrodes (Fz, F1, F3, AFz, AF3, AF4; A) and occipital electrodes (Oz, O1, O2, POz, PO3, PO4; B). Shading indicates the withinsubject standard error of the mean. Black bars along the x-axis denote significant differences at the displayed electrodes (cluster-corrected). (C-D) Headmaps show the effects of feature-based attention (congruent minus incongruent) during the indicated time windows. Asterisks and dots denote

electrodes with larger, or smaller responses, respectively, in at least 33% of the averaged time points(cluster-corrected).

347 The Effect of Feature-based Attention Depends on Surprise

348 Next, we investigated the interaction between feature-based attention and prediction by 349 subtracting the standard ERP from the deviant ERP (i.e., the mismatch response, collapsed 350 across deviant conditions) and comparing difference waves between congruent and incongruent conditions (*Figure 4*). Over posterior electrodes, the mismatch response was more negative for congruent stimuli than for incongruent stimuli late in the epoch (203 - 285 ms, p)= .041; *Figure 4B*). Inspection of individual condition ERPs (*Figure 4E*) revealed that the significant interaction was driven by a larger (negative) effect of feature-based attention on the neural response to deviants, relative to standards.

356 We followed up this finding by averaging spatiotemporal samples spanned by the significant 357 effect, separately for each deviant condition. We then compared congruent and incongruent conditions with paired *t*-tests and Bayesian analyses (uniform prior with upper and lower 358 359 bounds set to the average amplitude across all conditions). Feature-based attention decreased 360 the mismatch response to all three deviant types (orientation: congruent = $0.13 \pm 0.10 \mu V$, 361 incongruent = $0.42 \pm 0.13 \mu V$, t(23) = -2.16, p = .041, $BF_{10} = 1.18$; colour: congruent = -0.45 $\pm 0.17 \,\mu\text{V}$, incongruent = $0.05 \pm 0.14 \,\mu\text{V}$, t(23) = -3.28, p = .003, $BF_{10} = 1.41$; object: congruent 362 363 $= -0.53 \pm 0.19 \,\mu\text{V}$, incongruent $= -0.06 \pm 0.14 \,\mu\text{V}$, t(23) = -2.92, p = .008, $BF_{10} = 1.28$; Figure 364 4C). These findings suggest that feature-based attention modulates the effect of prediction on 365 neural responses to stimuli at task-irrelevant locations, irrespective of the predicted feature (or 366 combination of features).



367

368 Figure 4. Interaction between feature-based attention and prediction in Experiment 1. (A-B) Average 369 mismatch response (MMR; average deviant minus standard) collapsed across frontal electrodes (Fz, 370 F1, F3, AFz, AF3, AF4; A) and occipital electrodes (Oz, O1, O2, POz, PO3, PO4; B). Solid lines 371 represent the congruent condition and dotted lines represent the incongruent condition. Shading 372 indicates the within-subject standard error of the mean. The black bar along the x-axis denotes 373 significant differences at the displayed electrodes (cluster-corrected). (C) Mismatch responses at 374 occipital electrodes for individual deviant conditions. (D-E) ERPs evoked by standards and deviants 375 (averaged across deviant types), shown separately for congruent (solid) and incongruent (dotted) 376 conditions. (F) ERPs for individual deviant conditions, shown separately for congruent (solid) and 377 incongruent (dotted) conditions. (G-H) Headmaps show the effect of feature-based attention 378 (congruent minus incongruent) on the average deviant mismatch response (average deviant minus 379 standard) during the early vMMN (G) and late interaction time windows (H). Dots denote electrodes 380 with significant differences in at least 33% of the averaged time points (cluster-corrected). (I-K) Effect 381 of feature-based attention (congruent minus incongruent) on the orientation mismatch response (I), 382 colour mismatch response (J) and object mismatch response (K) during the late interaction time 383 window. Note that cluster-based permutation tests were not conducted on these differences.

384 The Visual Mismatch Negativity is Not Modulated by Feature-based Attention

385 Because previous literature has provided evidence for an effect of feature-based attention on

- the vMMN (Czigler & Sulykos, 2010), we also used Bayes analyses to test for differences
- 387 between congruent and incongruent conditions during the (non-significant) vMMN time

388 period. Spatiotemporal samples spanning the common vMMN window (electrodes: Pz, P1, P2, 389 P3, P4, POz, PO3, PO4, Oz, O1, O2, Iz; timepoints: 109 – 160 ms) were averaged to produce 390 one amplitude value for each condition within participants. Congruent and incongruent 391 conditions were compared within deviant conditions using paired-samples *t*-tests and Bayes 392 analyses (uniform prior with upper and lower bounds set to the average amplitude across 393 conditions). We found no difference between congruent and incongruent vMMNs for any 394 deviant type (orientation: congruent = $-1.09 \pm 0.13 \mu$ V, incongruent = $-0.95 \pm 0.16 \mu$ V, t(23) =395 -1.23, p = .231, $BF_{10} = .26$; colour: congruent = $-1.00 \pm 0.16 \mu$ V, incongruent = -0.96 ± 0.14 396 μV , t(23) = -0.37, p = .713, $BF_{10} = 0.13$; object: congruent = $-1.55 \pm 0.19 \ \mu V$, incongruent = $-1.55 \pm 0.19 \ \mu V$, incongruent = $-1.55 \pm 0.19 \ \mu V$, incongruent = $-1.55 \pm 0.19 \ \mu V$, incongruent = $-1.55 \pm 0.19 \ \mu V$, incongruent = $-1.55 \pm 0.19 \ \mu V$, incongruent = $-1.55 \pm 0.19 \ \mu V$, incongruent = $-1.55 \pm 0.19 \ \mu V$, incongruent = $-1.55 \pm 0.19 \ \mu V$, incongruent = $-1.55 \pm 0.19 \ \mu V$, incongruent = $-1.55 \pm 0.19 \ \mu V$, incongruent = $-1.55 \pm 0.19 \ \mu V$, incongruent = $-1.55 \pm 0.19 \ \mu V$, incongruent = $-1.55 \pm 0.19 \ \mu V$, incongruent = $-1.55 \pm 0.19 \ \mu V$, incongruent = $-1.55 \pm 0.19 \ \mu V$, incongruent = $-1.55 \pm 0.19 \ \mu V$, incongruent = $-1.55 \pm 0.19 \ \mu V$, incongruent = $-1.55 \pm 0.19 \ \mu V$, incongruent = $-1.55 \pm 0.19 \ \mu V$, incongruent = $-1.55 \pm 0.19 \ \mu V$, incongruent = $-1.55 \pm 0.19 \ \mu V$, incongruent = $-1.55 \pm 0.19 \ \mu V$, incongruent = $-1.55 \pm 0.19 \ \mu V$, incongruent = $-1.55 \pm 0.19 \ \mu V$, incongruent = $-1.55 \pm 0.19 \ \mu V$, incongruent = $-1.55 \pm 0.19 \ \mu V$, incongruent = $-1.55 \pm 0.19 \ \mu V$, incongruent = $-1.55 \pm 0.19 \ \mu V$, incongruent = $-1.55 \pm 0.19 \ \mu V$, incongruent = $-1.55 \pm 0.19 \ \mu V$, incongruent = $-1.55 \pm 0.19 \ \mu V$, incongruent = $-1.55 \pm 0.19 \ \mu V$, incongruent = $-1.55 \pm 0.19 \ \mu V$, incongruent = $-1.55 \pm 0.19 \ \mu V$, incongruent = $-1.55 \pm 0.19 \ \mu V$, incongruent = $-1.55 \pm 0.19 \ \mu V$, incongruent = $-1.55 \pm 0.19 \ \mu V$, incongruent = $-1.55 \pm 0.19 \ \mu V$, incongruent = $-1.55 \pm 0.19 \ \mu V$, incongruent = $-1.55 \pm 0.19 \ \mu V$, incongruent = $-1.55 \pm 0.19 \ \mu V$, incongruent = $-1.55 \pm 0.19 \ \mu V$, incongruent = $-1.55 \pm 0.19 \ \mu V$, incongruent = $-1.55 \pm 0.19 \ \mu V$, incongruent = $-1.55 \pm 0.19 \ \mu V$, incongruent = $-1.55 \pm 0.19 \ \mu V$, incongruent = $-1.55 \pm 0.19 \ \mu V$, incongruent = $-1.55 \pm 0.19 \ \mu V$, incongruent = $-1.55 \pm 0.19 \ \mu V$, incongruent = $-1.55 \pm 0.19 \ \mu V$, incongruent = $-1.55 \pm 0.19 \ \mu V$, incongruent = $-1.55 \pm 0.19 \ \mu V$, incongruent = $-1.55 \pm 0.19 \ \mu V$, incongruent = $-1.55 \pm 0.19 \ \mu V$, incon 397 $1.73 \pm 0.15 \,\mu\text{V}, t(23) = 1.29, p = .208, BF_{10} = 0.33; Figure 4C).$

398 Taken together, the results from Experiment 1 suggest that feature-based attention modulates 399 the neural effect of prediction on neural responses to stimuli at task-irrelevant locations. This 400 interaction emerged after (but not during) the vMMN time period for all deviant types, from 401 approximately 200 ms after stimulus onset. We also found that the detection of high contrast 402 targets at fixation was slower following peripheral patterns with target features, relative to 403 those with distractor features, suggesting that feature-congruent peripheral patterns 'captured' attention to their location (Folk et al., 1992). Because our principle question of interest 404 405 pertained to the neural interaction between feature-based attention and prediction *outside* the 406 current spatial focus of attention, we conducted a second study in which target contrast was 407 individually titrated for each participant to increase the task difficulty and ensure that attention 408 remained fixed on the central target stream.

409 Experiment 2: Replication with individually thresholded manipulation of feature-based

410 **attention at fixation**

411 Peripheral Patterns Do Not Modulate Behaviour in a Demanding Feature-based Attention

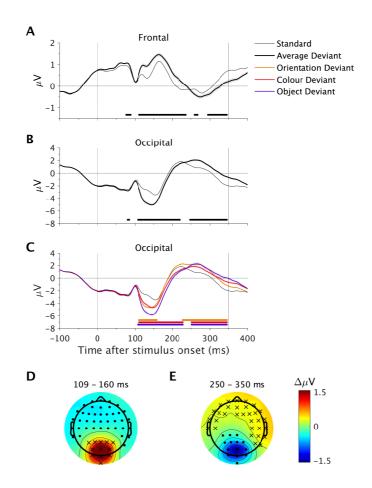
412 *Task*

In contrast to Experiment 1, there was no significant effect of feature-congruence on 413 414 normalised reaction times in Experiment 2 (congruent: 391.79 ± 11.12 ms, -0.01 ± 0.18 z-415 normalised; incongruent: 392.49 ± 11.56 ms, 0.02 ± 0.18 z-normalised; F(1,23) = 1.00, p =.329, $\eta_p^2 = .04$), suggesting that the top-down feature set modulates the effect of congruent 416 patterns on target detection and that the more difficult task in Experiment 2 contained spatial 417 attention to the central target stream. In line with Experiment 1, all other behavioural effects 418 419 were non-significant. Thus, there was no significant effect of feature-congruence on 420 normalised hit rates (congruent = $74.90 \pm 2.05\%$, -0.02 ± 0.16 z-normalised, incongruent = $75.23 \pm 2.14\%$, 0.01 ± 0.17 z-normalised, F(1,23) = 0.37, p = .547, $\eta_p^2 = .02$). In addition, there 421 was no significant effect of pattern prediction on normalised hit rates (predicted = $74.25 \pm$ 422 2.07%, -0.06 ± 0.16 z-normalised, surprising = $75.88 \pm 2.13\%$, 0.06 ± 0.17 z-normalised, 423 $F(1,23) = 3.64, p = .069, \eta_p^2 = 14$) or on normalised reaction times (predicted = 388.79 ± 11.11) 424 425 ms, -0.03 ± 0.18 z-normalised, surprising = 395.49 ± 11.58 ms, 0.04 ± 0.18 z-normalised, F(1,23) = 2.46, p = .130, $\eta_p^2 = .10$). Finally, there was no interaction between prediction and 426 feature-congruence on either normalised hit rates (F(1,23) = 2.11, p = .160, $\eta_p^2 = .08$) or 427 normalised reaction times ($F(1,23) = 1.50, p = .233, \eta_p^2 = .06$). 428

429 The Neural Interaction Between Feature-based Attention and Prediction Replicates With a

430 Demanding Feature-based Attention Set

431 The neural effects observed in Experiment 2 (see Figures 5-7) were highly similar to those in 432 Experiment 1 (see Figures 2-4). Prediction again modulated neural responses over posterior 433 electrodes early (standards > deviants; from 78 ms, p < .001) and late (standards < deviants; 434 246 - 348 ms, p = .014) in the epoch (*Figure 5B*), with opposite early (standards < deviants; 74 - 238 ms, p < .001) and late effects (standards > deviants; prior to 348 ms, p < .001) over 435 436 frontal electrodes (Figure 5A). Follow-up comparisons revealed similar effects of prediction 437 on each deviant type (Figure 5C). Over posterior electrodes, standards evoked smaller early negativities than all deviant types (orientation deviants: 109 - 160 ms, p = .037; colour 438 439 deviants: 90 - 348 ms, p < .001; object deviants; 78 - 348 ms, p < .001) and smaller late 440 positivities than all deviant types (orientation deviants: 227 - 348 ms, p = .001; colour deviants; 441 250 - 348 ms, p = .029; object deviants: 250 - 348 ms, p = .022). Over frontal electrodes, 442 standards evoked smaller early positivities than all deviant types (orientation: 47 - 156 ms, p 443 = .0012; colour: 98 – 242 ms, p = .002; object: 78 – 242 ms, p < .001) and smaller late 444 negativities than orientation deviants (234 - 348 ms, p < .001). As in Experiment 1, the vMMN 445 was sensitive to features of the deviant stimulus (*Figure 5C*), with object deviants evoking a 446 significantly larger vMMN (-1.86 \pm 0.27 μ V) than orientation deviants (-1.10 \pm 0.16 μ V, t(23) = -5.40, p < .001, $BF_{10} = 288,942.02$) and colour deviants (-1.05 ± 0.19 µV, t(23) = -6.17, p < -6.17447 448 .001, $BF_{10} = 22,207,026.78$). Again, there was no difference between orientation and colour 449 vMMNs (t(23) = -0.42, p = .679, $BF_{10} = 0.12$).

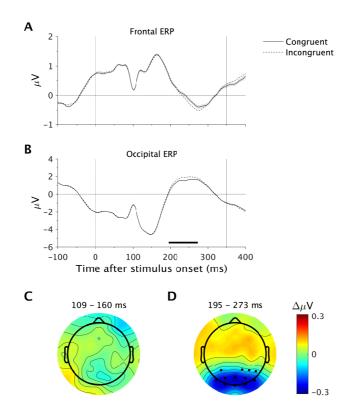


450

451 Figure 5. Main effect of prediction in Experiment 2. (A-B) ERPs evoked by standards and deviants 452 (collapsed across deviant types) at frontal electrodes (Fz, F1, F3, AFz, AF3, AF4; A) and occipital 453 electrodes (Oz, O1, O2, POz, PO3, PO4; B). Shading indicates the within-subject standard error of 454 the mean, calculated relative to standards. Black bars along the x-axis denote significant timepoints at 455 the displayed electrodes (cluster-corrected). (C) ERPs evoked by standards and each of the three 456 deviant conditions. Shading indicates the within-subject standard error of the mean, calculated 457 separately for each deviant condition relative to standards. Yellow, red and purple bars along the x-458 axis denote significant differences between standards and each corresponding deviant condition (cluster-corrected). (D-E) Headmaps show the effect of prediction (standard minus average deviant) 459 460 during the indicated time windows. Asterisks and dots denote electrodes with larger, or smaller 461 responses, respectively, across at least 33% of the averaged time points (cluster-corrected).

462 As in Experiment 1, congruent peripheral patterns evoked smaller positivities over posterior

- 463 electrodes than incongruent patterns late in the epoch (195 273 ms, p = .045; Figure 6B).
- 464 However, the polarity-reversed frontal effect observed in Experiment 1 was not significant in
- 465 Experiment 2 (230 281 ms, p = .127).



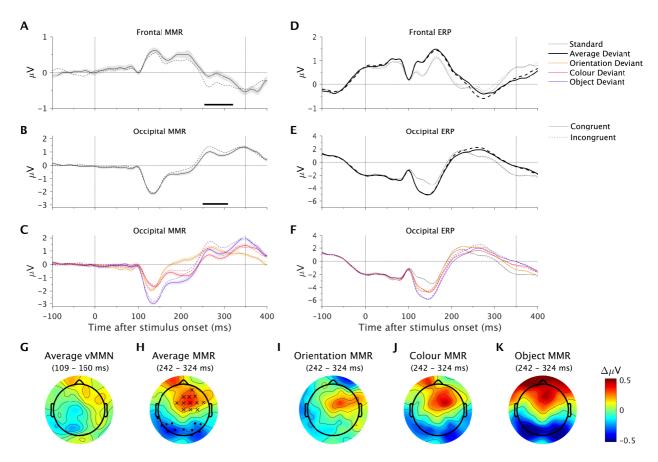
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Figure 6. Main effect of feature-based attention in Experiment 2. (A-B) Congruent and incongruent ERPs are collapsed across prediction conditions separately for frontal electrodes (Fz, F1, F3, AFz, AF3, AF4; A) and occipital electrodes (Oz, O1, O2, POz, PO3, PO4; B). Shading indicates the withinsubject standard error of the mean. The black bar along the x-axis denotes significant differences at the displayed electrodes (cluster-corrected). (C-D) Headmaps show the effects of feature-based attention (congruent minus incongruent) during the indicated time windows. Dots denote electrodes with smaller responses in at least 33% of the averaged time points (cluster-corrected).

474 Crucially, we replicated the significant interaction between feature-based attention and 475 prediction observed in Experiment 1 (Figure 7). Congruent mismatch responses (deviants 476 minus standards) were significantly smaller than incongruent mismatch responses over 477 posterior electrodes late in the epoch (242 - 320 ms, p = .048; Figure 7B, H). We also observed an additional polarity-reversed effect over frontal electrodes that was absent in Experiment 1 478 479 (254 - 324 ms, p = .026; Figure 7A, H). Follow-up analyses revealed that feature-based 480 attention significantly decreased the mismatch response to colour deviants (congruent = -0.01 $\pm 0.17 \,\mu\text{V}$, incongruent = $0.31 \pm 0.13 \,\mu\text{V}$, t(23) = -3.52, p = .002, $BF_{10} = 94.83$, Figure 7J) and 481 482 object deviants (congruent = $0.10 \pm 0.21 \mu$ V, incongruent = $0.61 \pm 0.14 \mu$ V, t(23) = -3.66, p =

483 .001, $BF_{10} = 51.42$, *Figure 7K*) but only trended in the same direction for orientation deviants 484 (congruent = $0.40 \pm 0.11 \mu$ V, incongruent = $0.60 \pm 0.10 \mu$ V, t(23) = -1.96, p = .062, $BF_{10} =$ 485 2.35, *Figure 7I*). Again, we found no effect of feature-based attention on the vMMN evoked 486 by any type of deviant (orientation: congruent = $-1.09 \pm 0.19 \mu$ V, incongruent = -1.11 ± 0.15 487 μ V, t(23) = .18, p = .857, $BF_{10} = .12$; colour: congruent = $-1.08 \pm 0.19 \mu$ V, incongruent = -1.02488 $\pm 0.22 \mu$ V, t(23) = -0.42, p = .676, $BF_{10} = .15$; object: congruent = $-1.96 \pm 0.28 \mu$ V, incongruent 489 $= -1.76 \pm 0.27 \mu$ V, t(23) = -1.67, p = .11, $BF_{10} = 0.43$; *Figure 7C*).

490 Overall, the findings from Experiment 2 replicate those from Experiment 1 to show that 491 feature-based attention and prediction interact in their modulation of neural responses to stimuli 492 at task-irrelevant locations. Similar to Experiment 1, this interaction emerged after (but not 493 during) the vMMN time period for all deviant types, from approximately 240 ms after stimulus 494 onset.



495

496 Figure 7. Interaction between feature-based attention and prediction in Experiment 2. (A-B) Average 497 mismatch response (MMR; average deviant minus standard) collapsed across frontal electrodes (Fz, 498 F1, F3, AFz, AF3, AF4; A) and occipital electrodes (Oz, O1, O2, PO2, PO3, PO4; B). Solid lines 499 represent the congruent condition and dotted lines represent the incongruent condition. Shading 500 indicates the within-subject standard error of the mean. The black bar along the x-axis denotes 501 significant differences at the displayed electrodes (cluster-corrected). (C) Mismatch responses at 502 occipital electrodes for individual deviant conditions. (D-E) ERPs evoked by standards and deviants 503 (averaged across deviant types), shown separately for congruent (solid) and incongruent (dotted) 504 conditions. (F) ERPs for individual deviant conditions, shown separately for congruent (solid) and 505 incongruent (dotted) conditions. (G-H) Headmaps show the effect of feature-based attention 506 (congruent minus incongruent) on the average deviant mismatch response (average deviant minus 507 standard) during the early vMMN (G) and late interaction time windows (H). Asterisks and dots denote 508 electrodes with larger, or smaller responses, respectively, in at least 33% of the averaged time points 509 (cluster-corrected). (I-K) Effect of feature-based attention (congruent minus incongruent) on the 510 orientation mismatch response (I), colour mismatch response (J) and object mismatch response (K) 511 during the late interaction time window. Note that cluster-based permutation tests were not conducted 512 on these differences.

513

Discussion

514 Here we investigated whether prediction interacts with feature-based attention outside the

515 spatial focus of attention. To achieve this, we measured neural responses to surprising and

516 predicted stimuli – deviants and standards, respectively – presented at task-irrelevant locations.

Task-irrelevant peripheral patterns shared features with either the targets (congruent) or distractors (incongruent) in a central search task. Across two experiments, we replicated the finding that feature-based attention decreased neural responses to surprising but not predicted task-irrelevant stimuli in the periphery of vision. This finding suggests that the global neural mechanisms of feature-based attention and prediction are interdependent, and supports the theory that attention increases the gain of prediction errors (Feldman & Friston, 2010).

523 Consistent with previous literature, prediction modulated early and late neural responses to stimuli in both experiments (*Figures 2 & 5*). Early responses (approximately 100 to 160 ms) 524 525 over posterior electrodes were more negative for surprising stimuli than predicted stimuli, 526 consistent with the commonly reported visual mismatch negativity (for a review, see Stefanics 527 et al., 2014). This finding is broadly consistent with the theory that top-down prediction signals 528 silence matching bottom-up sensory signals and leave only the remaining prediction error to 529 propagate forward (Friston, 2005, 2009; Rao & Ballard, 1999). Prediction also reduced the 530 later positive P3 component (from approximately 250 ms), consistent with the theory that this 531 component reflects involuntary orienting to novel stimuli (Friedman, Cycowicz, & Gaeta, 532 2001; Polich, 2007).

We also found that stimuli deviating in two feature dimensions (i.e., object deviants) evoked larger early negativities than stimuli deviating in only one feature dimension (i.e., orientation or colour deviants; *Figures 2 & 5*). This finding contradicts a previous study that found visual features elicit non-additive mismatch-related brain activity (Sulykos & Czigler, 2011), and suggests instead that the vMMN is sensitive to the extent of deviation across multiple feature dimensions. Importantly, object deviants in Sulykos & Czigler (2011) deviated in spatial frequency and orientation, whereas object deviants in the present study deviated in colour and

orientation. Thus, future studies should investigate the extent to which mismatch additivity inthe visual domain depends on the specific features involved.

542 We found that feature-based attention reduced neural responses to task-irrelevant peripheral 543 patterns from approximately 200 ms after stimulus onset (Figures 3 & 6), consistent with the commonly reported 'selection negativity' (Gledhill et al., 2015). This effect replicated with 544 545 low contrast stimuli that likely necessitated a tight focus of spatial attention (Experiment 2), 546 contradicting the finding that feature-specific modulation of the selection negativity is 547 contingent on spatial attention (Anllo-Vento & Hillyard, 1996; Hillyard & Münte, 1984) and 548 suggesting instead that late effects of feature-based attention are globally effective (Gledhill et 549 al., 2015). Interestingly, we found no difference between neural responses to congruent and 550 incongruent stimuli earlier in the epoch (Figures 3 & 6), in contrast to a previous study that 551 reported early effects of feature-based attention on neural responses to stimuli at task-irrelevant 552 locations (beginning within 100 ms of stimulus onset; Zhang & Luck, 2009). A critical difference between Zhang & Luck (2009) and the present study is that Zhang & Luck (2009) 553 554 had participants search for targets with specific feature conjunctions (luminance and colour), 555 whereas targets in our study were defined by only a single feature (colour or orientation). Thus, 556 it is possible that early effects of feature-based attention depend on the complexity of the 557 attentional set. This interpretation is consistent with a recent study in which we found that 558 neural responses to high-frequency flickering stimuli outside a search array (12.5 or 16.7 Hz, 559 corresponding to an 80 or 60 ms cycle) are enhanced by feature-based attention during 560 conjunction but not unique-feature search (Painter, Dux, Travis, & Mattingley, 2014).

561 Crucially, we found an interaction between feature-based attention and prediction in each of 562 the two experiments. Congruent stimuli evoked smaller posterior mismatch responses than

563 incongruent stimuli between approximately 200 and 300 ms after stimulus onset. Inspection of 564 the ERPs revealed that the effect of feature-based attention on neural responses was larger for deviants than it was for standards. This pattern of results is consistent with our recent finding 565 566 that attention enhances the processing of mismatch information from approximately 200 ms post-stimulus (Smout et al., 2019) and broadly supports the theory that attention enhances the 567 568 gain of prediction errors (Feldman & Friston, 2010). Neural responses to surprising stimuli 569 (deviants) are theorised to be modulated by attention because they contain prediction errors, 570 whereas neural responses to predicted stimuli (standards) are less affected because they contain 571 relatively few prediction errors. The present study extends this theory to suggest that featurespecific attentional modulation of prediction errors occurs even when the surprising stimuli are 572 573 task-irrelevant and presented outside the spatial focus of attention.

574 Interestingly, we found that feature-based attention had no effect on the earlier vMMN evoked 575 by deviants (109 - 160 ms). This pattern of findings contradicts a previous study that found the 576 vMMN evoked by peripheral stimuli was smaller (more positive) when participants searched 577 for a change in the deviating feature at fixation, relative to a different feature (Czigler & 578 Sulvkos, 2010). A subtle difference between the paradigms is that participants in Czigler and 579 Sulykos (2010) searched for a feature 'change' at fixation (e.g., a change in the target object 580 colour), whereas participants in our study searched for specific object onsets. Thus, it remains 581 possible that subtle differences in the configuration of the attentional set can influence the 582 timing and direction of the interaction between feature-based attention and prediction.

583 We manipulated target and distractor salience across the two experiments in order to 584 investigate whether the strength of the top-down feature set modulates the neural interaction 585 between prediction and feature-based attention. Although the pattern of neural effects did not

586 differ between the two experiments, we observed slightly different behavioural effects as a 587 function of task difficulty. Responses to highly salient targets (Experiment 1) that appeared 588 immediately after a congruent pattern were slower than those to targets that appeared after an 589 incongruent pattern. In contrast, there was no such effect of feature-congruence on responses 590 to less salient targets (Experiment 2). These findings are broadly consistent with contingent capture theory (Folk et al., 1992), which proposes that distracting stimuli within the spatial 591 592 focus of attention capture attention when they are congruent with the observers' current 593 attentional set. Since targets were easily detected in Experiment 1, it seems likely that some 594 amount of spatial attention 'leaked' to the peripheral stimuli, facilitating contingent capture. In 595 contrast, the higher task difficulty of Experiment 2 likely necessitated a tighter focus of 596 attention to the central stimuli, thus prohibiting a contingent capture effect.

597 We did not observe an effect of predictability of peripheral patterns on target detection, or an 598 interaction between pattern prediction and feature-congruence, in either experiment. This is 599 consistent with a previous study that failed to find any effect of pattern prediction on response 600 times to a central feature change target, nor an interaction with task set, at the level of single 601 trials (though note that this study did report sustained block-wise effects on behaviour; Czigler 602 & Sulykos, 2010). These findings suggest that the neural bias toward feature-congruent and 603 surprising stimuli at task-irrelevant locations, observed in the present study, does not interfere 604 with the concurrent processing of targets at task-relevant locations.

The present study contributes to a burgeoning literature on the relationship between prediction and attention. Whereas some studies have found an interaction between prediction and attention (Auksztulewicz & Friston, 2015; Jiang et al., 2013; Kok, Rahnev, et al., 2012; Marzecová et al., 2017; Smout et al., 2019), many others have reported only independent main

609 effects (e.g. Garrido, Rowe, Halász, & Mattingley, 2017; Hsu, Hämäläinen, & Waszak, 2014; 610 Kok, Jehee, & de Lange, 2012). We note that investigations to date have employed a wide variety of attention manipulations (e.g., feature-based, spatial, temporal) and prediction 611 612 manipulations (e.g., first-order, rule-based) across different sensory modalities (e.g., visual, 613 auditory). Thus, the equivocal pattern of findings to date may stem from distinct relationships 614 between different subprocesses of attention and prediction across the various modalities. In 615 particular, previous studies that found an interaction between visual attention and prediction 616 presented stimuli at attended locations (Jiang et al., 2013; Kok, Rahnev, et al., 2012; Marzecová 617 et al., 2017; Smout et al., 2019) or used paradigms that did not require focussed attention to 618 complete the task (Czigler & Sulykos, 2010), leaving open the possibility that spatial attention 619 is necessary in the interaction with prediction. The present study extends this literature by 620 demonstrating that visual predictions interact with feature-based attention to modulate neural 621 responses to stimuli outside the spatial focus of attention. The nature of this interaction is 622 consistent with the theory that attention optimises the expected precision of predictions by 623 modulating the gain of prediction errors (Feldman & Friston, 2010). Future research should 624 continue to parse 'attention' and 'prediction' into more precise taxonomies that reflect specific 625 mechanisms in the brain and investigate potential interactions between each of these subcomponents. This work could illuminate the extent to which predictive coding theory might 626 627 be considered a 'unified theory of the brain' (Friston, 2010).

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