1 TITLE

Large scale genome-wide association study in a Japanese population identified 45 novel susceptibility loci for 22
 diseases

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135 INTRODUCTORY PARAGRAPH

136	The overwhelming majority of participants in current genetic studies are of European ancestry ^{1–3} , limiting our
137	genetic understanding of complex disease in non-European populations. To address this, we aimed to elucidate
138	polygenic disease biology in the East Asian population by conducting a genome-wide association study (GWAS)
139	with 212,453 Japanese individuals across 42 diseases. We detected 383 independent signals in 331 loci for 30
140	diseases, among which 45 loci were novel ($P < 5 \times 10^{-8}$). Compared with known variants, novel variants have lower
141	frequency in European populations but comparable frequency in East Asian populations, suggesting the advantage
142	of this study in discovering these novel variants. Three novel signals were in linkage disequilibrium ($r^2 > 0.6$) with
145	(p R220W of ATG16L2 a autophagy-related gene) associated with coronary artery disease. We further
145	investigated enrichment of heritability within 2,868 annotations of genome-wide transcription factor occupancy, and
146	identified 378 significant enrichments across nine diseases (FDR < 0.05) (e.g. NF-κB for immune-related diseases).
147	This large-scale GWAS in a Japanese population provides insights into the etiology of common complex diseases
148	and highlights the importance of performing GWAS in non-European populations.

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MAIN TEXT

151	We conducted a genome-wide association study (GWAS) of 42 diseases in a Japanese population, comprising
152	179,660 patients who participated in the BioBank Japan Project (BBJ) ^{4,5} and 32,793 population-based controls
153	(Supplementary Table 1). The 42 diseases encompassed a wide-range of disease categories; 13 neoplastic
154	diseases, five cardiovascular diseases, four allergic diseases, three infectious diseases, two autoimmune diseases,
155	one metabolic disease, and 14 uncategorized diseases. By including patients with unrelated diagnoses into control
156	samples, we maximized the power of our GWAS (Supplementary Table 1 and Supplementary Figure 1). We
157	employed a generalized linear mixed model in our association analysis using SAIGE ⁶ . Following imputation with
158	1000 Genomes Project Phase 3 reference data (1KG Phase3) ⁷ , we tested 8,712,794 autosomal variants and
159	207,198 X chromosome variants for association with 42 diseases. We estimated their heritability using linkage
160	disequilibrium score regression (LDSC) analysis ⁸ (Supplementary Table 2). Consistent with a recent finding in the
161	European population ⁹ , incorporating the baselineLD model ¹⁰ into the LDSC analysis improved heritability estimation
162	in our GWAS (Methods; Supplementary Figure 2 and Supplementary Table 2). Although we observed high genomic
163	inflation factors (λ_{GC}) for some diseases (e.g. λ_{GC} = 1.3 for type 2 diabetes (T2D); Supplementary Table 2), LDSC
164	analysis indicated that the majority of the inflated chi-squared statistics originated from polygenic effects rather than
165	confounding biases (e.g. intercept = 1.01 for T2D; Supplementary Table 2). Overall, we detected significant
166	associations for 30 diseases at 309 autosomal loci (outside of the HLA region) and nine loci on the X chromosome
167	($P < 5 \ge 10^{-8}$) (Supplementary Table 3 and 4). Associations at the HLA region have been investigated in detail in a
168	separate article ¹¹ .

169	We further performed conditional analyses in these 318 loci to explore disease-associated variants
170	independent of the lead variants. We detected 52 additional independent signals for 11 diseases ($P < 5 \ge 10^{-8}$)
171	(Supplementary Table 5). The largest number of independent signals in a single locus was seven, found in the
172	FAM84B/POU5F1B locus associated with prostate cancer and in the KCNQ1 locus associated with T2D.
173	For 35 diseases for which we have both male and female patients, we conducted male- and female-
174	specific GWAS. We detected 13 additional loci for 10 diseases which were not identified in a sex-combined
175	analysis ($P < 5 \ge 10^{-8}$) (Supplementary Table 6). We tested heterogeneity between effect size estimates for males
176	and females using Cochran's Q test. This analysis found seven loci showing significant differences in effect size
177	estimates between sexes (P values of heterogeneity (P_{het}) < 0.05/13); two asthma loci, a cataract locus, a cerebra
178	aneurysm locus, and a lung cancer locus were specifically associated with females; and a coronary artery disease
179	(CAD) locus and a T2D locus were specifically associated with males.
180	In total, we detected 383 independent signals in 331 loci outside of the HLA region for 30 diseases, of
181	which 45 loci were novel (Figure 1a, Table 1, Supplementary Table 3, 4, and 6). Five novel disease-associated
182	variants were rare variants (MAF < 0.01), and four of them had large effect sizes (odds ratio > 2, Figure 1b). To
183	understand the characteristics of novel and known disease-associated variants, we examined their allele
184	frequencies in East Asian and European populations of 1KG Phase3. Allele frequencies of novel and known
185	variants were of comparable level in East Asian populations ($P = 0.35$, Figure 1c). However, novel variants have
186	lower allele frequencies than known variants in European populations ($P = 0.0030$, Figure 1d). Although both of
187	novel and known variants have lower allele frequencies in European populations than in East Asian populations,

188	novel variants have larger inter-population differences than known variants ($P = 0.0047$, Supplementary Figure 3).
189	To estimate population specificity in our GWAS results, we compared our results with those reported in previous
190	European GWAS. We utilized publicly available GWAS summary statistics of European populations for 10 diseases
191	(Methods), and tested for consistency in direction of effect between populations at 11 novel and 146 known
192	disease-associated variants from our GWAS; 10 out of 11 novel and 141 out of 146 known variants were replicated
193	in the same allelic direction in European GWAS (binomial test <i>P</i> values were 0.011 and 1.1 x 10^{-35} , respectively;
194	Supplementary Figure 4). In addition, 595 out of 665 disease-associated variants detected in European GWAS
195	were replicated in the same allelic direction in our GWAS (binomial test P values = 1.1×10^{-104} ; Supplementary
196	Figure 4). These findings suggested that genetic etiologies around the disease-associated variants are generally
197	shared across populations, and the higher allele frequencies at novel associated variants in our East Asian cohort
198	increased the efficiency of the variant discovery. This highlights the importance of performing GWAS in non-
199	European populations.
200	We next investigated the potential impact of the disease-associated variants on protein functions
201	(Supplementary Table 7). Nine novel variants were in linkage disequilibrium (LD) with missense variants ($r^2 > 0.6$ in
202	the 1KG Phase3 East Asian populations) (Table 2). Among them, three missense variants are monomorphic in
203	European populations (1KG Phase3); p.R220W of ATG16L2 associated with CAD; p.V326A of POT1 associated
204	with lung cancer; and p.E62G of PHLDA3 associated with keloid (Figure 2 and Supplementary Figure 5). First,
205	ATG16L2 is an autophagy-related gene. Although p.R220W of ATG16L2 is not the lead variant at this locus,
206	conditioning on this variant cancelled the signal of the lead variant (Figure 2a). Previous GWAS for CAD in

207	European populations did not detect significant associations at this locus ¹² (Figure 2a). These findings suggested
208	that p.R220W of ATG16L2 which is absent in Europeans may be the causal variant. p.R220W of ATG16L2 is also
209	associated with Crohn's disease in a Chinese population ¹³ , and ATG16L2 is highly expressed in immune cells.
210	Therefore, dysregulated autophagy in immune cells might have an important role in CAD. Second, POT1 is a
211	member of the telombin family and this protein binds to telomeres, regulating telomere length. Missense variants of
212	POT1 have been described to be responsible for several familial cancers ^{14–16} . Together with a known association at
213	the TERT locus (Supplementary Table 3), we provide additional evidence that telomere dysregulation is pathogenic
214	for lung cancer. Intriguingly, this association was discovered in the female-specific GWAS, and a significant
215	heterogeneity in effect sizes between males and females was observed ($P_{het} = 7.7 \times 10^{-4}$) (Figure 2b; Supplementary
216	Table 6). This finding might help to understand the inter-sex differences in the etiology of lung cancer. Third,
217	p.E62G of <i>PHLDA3</i> is predicted to have a deleterious effect to its protein function (SIFT score ¹⁷ =0; CADD
218	score ¹⁸ =33), and we detected a large effect size for keloid (odds ratio = 9.56; 95% CI 5.91-15.45). PHLDA3 is
219	known to be a suppressor of AKT ¹⁹ , and upregulated AKT signaling pathway is related to increased collagen
220	production from dermal fibroblasts ²⁰ . Therefore, damaged PHLDA3 may activate the AKT pathway, promoting the
221	development of keloid. Together, our study successfully identified novel potential causal genes which would be
222	hard to be discovered by GWAS in European populations due to restrictive European allele frequencies.
223	We also investigated the potential impacts of the disease-associated variants on the mRNA levels using
224	two databases of expression quantitative trait locus (eQTL) analysis ^{21,22} . Out of the remaining 36 novel variants
225	whose functions were not explained by missense variants,11 variants were in LD with at least one eQTL variant (r ² 9

226	> 0.6), regulating the expression of 17 genes in total (Supplementary Table 8); <i>P2RY13, SIAH2</i> , and <i>SIAH2-AS1</i> for
227	breast cancer; ATP2B1, BET1L, POC1B, and ZNF767 for cerebral aneurysm; CCAT1 and CD40 for Graves's
228	disease; GABPB2 for osteoporosis; MOV10 and WNT2B for pancreatic cancer; CYP2A6 and CYP2B7P1 for
229	peripheral artery disease; ZMIZ1-AS1 for prostate cancer; and STIM1 and TRIM21 for urolithiasis. Intriguingly, the
230	eQTL signals for ATP2B1 which are in LD with a novel variant of cerebral aneurysm (rs11105352) is highly specific
231	to arterial tissues (Figure 3). Since the loss of ATP2B1 in vascular smooth muscle cells induced blood pressure
232	elevation in mice ²³ , decreased expression of <i>ATP2B1</i> in arteries might induce hypertension, which leads to
233	increased risk of cerebral aneurysm.
234	To understand differences in the genetic risks between males and females, we assessed genetic
235	correlations using LDSC ²⁴ between the results of sex-specific GWAS for the 20 diseases (see Methods for
236	selection of diseases). Although most correlations are close to one, correlation of asthma was significantly smaller
237	than one ($P = 2.2 \times 10^{-3} < 0.05/20$; Supplementary Figure 6). This finding suggested that genetic risks of asthma is
238	slightly different between males and females. To explore the biological mechanism underlying this finding, we
239	estimated the enrichment of the heritability of male or female asthma in the 220 cell-type specific regulatory regions
240	using stratified LD-score regression (S-LDSC) ²⁵ . We found significant enrichments for either of male or female
241	asthma in three annotations; Th0, Th1, and colonic mucosa ($P < 0.05/220$; Supplementary Figure 6). Among them,
242	the colonic mucosa annotation showed significant heterogeneity in the enrichment of heritability (P_{het} = 0.006 <
243	0.05/3). Recent studies suggested that host-microbiome interactions at intestinal mucosa (gut-lung axis) have

- important roles in the development of asthma^{26,27}, and our study suggested that the importance of the gut-lung axis
- in asthma might be different between males and females.

246	To acquire more insights to disease biology, we estimated the heritability enrichments in the binding sites
247	of a variety of transcription factors (TFs) using S-LDSC. We included TF binding sites defined by 2,868 publicly
248	available chromatin immunoprecipitation sequencing (ChIP-seq) datasets for 410 unique TFs (Supplementary
249	Table 9). To make mutually comparable data, we began our analysis from the raw sequencing data, and defined TF
250	binding sites using a uniform protocol (Methods). Using LD-scores of all TF binding sites, we grouped them into 15
251	clusters (cluster name was defined by the most dominant TF), and performed uniform manifold approximation and
252	projection (UMAP) ²⁸ to project all TF binding sites into a two-dimensional space (Methods; Figure 4a and
253	Supplementary Figure 7). To scale the performance of this analysis, we first analyzed previously reported GWAS
254	for red blood cell-related traits ²⁹ where the critical role of <i>GATA1</i> was supported by multiple pieces of evidence ^{30–34} ,
255	and we successfully recapitulated this biology (Figure 4b). We then applied this analysis to our 24 GWAS results
256	(see Methods for selection of diseases), and detected 378 significant enrichments for nine disease (FDR < 0.05)
257	(Figure 4c, Supplementary Figure 8, and Supplementary Table 10). Biologically plausible TFs were highlighted by
258	this analysis; <i>RELA</i> , a subunit of NF-κB, for atopic dermatitis, RA, and Graves' disease; sex hormone receptors
259	(AR and ESR1) for prostate cancer; and FOXA2, which regulates insulin secretion in pancreatic beta-cells ³⁵ , for
260	T2D (Figure 4c). This analysis also suggested that NKX3-1, a prostate-specific homeobox gene, has an important
261	role in the biology of prostate cancer (Figure 4c). In addition to this polygenic analysis, the importance of NKX3-1
262	was also suggested by the regional analysis integrating eQTL databases; the risk allele of prostate cancer at

263	<i>NKX3-1</i> locus (rs4872174-C) was suggested to decrease the expression of <i>NKX3-1</i> (Supplementary Table 8).
264	Consistently, loss of NKX3.1 expression in human prostate cancers was reported to be correlated with tumor
265	progression ³⁶ . Together, our results confirmed and expanded our current understanding of complex traits in the
266	context of TF activity.
267	In summary, we conducted a large-scale GWAS of 42 diseases in a non-European population and
268	provided rich public resources for genetic studies. Our study provided multiple insights into the etiology of complex
269	traits by integrating annotations of missense variants, eQTL variants, and transcription factor binding site tracks.
270	Currently, genetic studies are overwhelmed by European-descent samples, and the clinical translation of genetic
271	findings would be far more beneficial to European individuals than other populations ¹ . Our study contributed to
272	broaden the population diversity in genetic studies and should potentially mitigate the problems originating from this
273	imbalance.

275 **TABLE**

276

277 Table 1. 45 novel loci detected in this GWAS.

Variant	Chr.	Position	REF	ALT	Gene	OR	L95	U95	Р
ex-combined	l analy	sis							
rs3835894	2	204576923	С	CA	CD28	1.09	1.06	1.13	3.22E-08
rs10797119	9	92202495	т	С	GADD45G SEMA4D	1.11	1.07	1.15	1.45E-08
rs146383533	3	150480100	Т	TTTTC	SIAH2	0.86	0.83	0.90	1.04E-11
rs332827	1	61743160	G	Α	NFIA	1.06	1.04	1.08	3.22E-08
rs9720071	7	155086439	A	С	HTR5A INSIG1	0.94	0.91	0.96	2.03E-08
rs11235571	11	72411843	G	Α	ARAP1	0.90	0.87	0.93	2.64E-09
rs75812946	14	86192480	G	А	FLRT2	1.35	1.22	1.50	3.41E-09
rs12226402	11	215904	G	А	SIRT3	1.34	1.23	1.45	1.57E-12
rs78535549	12	20156904	С	т	AEBP2 PDE3A	0.85	0.81	0.90	7.97E-09
rs11105352	12	90026462	G	A	ATP2B1	0.85	0.81	0.90	1.22E-08
rs139337062	6	123096973	AAAAC	A	FABP7 PKIB	0.70	0.62	0.80	2.98E-08
rs2129981	4	111704199	G	т	C4orf32 PITX2	1.09	1.06	1.12	9.60E-09
rs140989504	11	100439234	С	т	CNTN5 TMEM133	1.37	1.22	1.52	2.17E-08
rs11066008	12	112140669	Α	G	ACAD10	1.29	1.21	1.37	4.34E-17
rs10673095	8	128203024	т	TTC	FAM84B POU5F1B	0.81	0.76	0.87	2.11E-09
rs11065783	12	111396249	A	G	CUX2 MYL2	1.34	1.24	1.44	7.23E-14
rs1569723	rs1569723 20 44742064 C		С	Α	CD40 NCOA5	1.20	1.13	1.28	4.06E-09
rs8107030	19	39736719	A	G	IFNL2 IFNL3	1.44	1.28	1.62	7.96E-10
	Variant ex-combined rs3835894 rs10797119 rs1463835333 rs332827 rs9720071 rs9720071 rs11235571 rs12226402 rs12226402 rs12226402 rs139337062 rs139337062 rs139337062 rs13993504 rs11065783 rs11065783 rs1569723 rs8107030	Variant Chr. ex-combined analy rs3835894 2 rs10797119 9 rs146383533 3 rs332827 1 rs9720071 7 rs11235571 11 rs75812946 14 rs758535549 12 rs1105352 12 rs139337062 6 rs140989504 11 rs140989504 12 rs11066008 12 rs1106673095 8 rs11065783 12 rs1569723 20 rs8107030 19	VariantChr.Positionex-combinedanalysisrs38358942204576923rs10797119992202495rs1463835333150480100rs332827161743160rs97200717155086439rs112355711172411843rs758129461486192480rs758129461486192480rs758355491220156904rs111053521290026462rs1393370626123096973rs14098950411100439234rs106730958128203024rs1106578312111396249rs15697232044742064rs81070301939736719	Variant Chr. 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Interstitial lung disease	rs6477542	9	109507432	С	т	TMEM38B	1.34	1.21	1.48	6.90E-09	
		-			-	ZNF462					
Keloid	rs192314256	1	201437730	Т	С	PHLDA3	9.56	5.91	15.45	3.28E-20	
Osteoporosis	rs578031265	2	168857545	С	Т	STK39	10.16	4.74	21.74	2.38E-09	
Pancreatic cancer	rs60579835	1	113113194	С	т	ST7L	1.48	1.29	1.69	1.00E-08	
Drestete sensor		10	00005007	0	۸	RPS24	1 1 5	1 1 0	1.00	1.675 10	
Prostate cancer	1811002805	10	00020097	G	A	ZMIZ1	1.15	1.10	1.20	1.07E-10	
Drestete sensor	ro1007E77	01	16071100	٨	т	NRIP1	0.00	0.94	0.02	1 825 00	
Prostate cancer	181997577	21	16371102	A	I	USP25	0.00	0.64	0.92	1.03E-09	
Prostate cancer	rs138426	22	38871891	G	Α	KDELR3	1.19	1.13	1.25	2.04E-12	
		0	040700540	–	~	CXCR2	1.00	1.0.1	4.00	0.705.00	
Type 2 diabetes	1820403309	2	210702040	I	C	TNS1	1.06	1.04	1.00	3.70E-00	
	7704000	-	07020270	–	~	MEF2C	4.05	1.0.1	4.07	1 115 00	
Type 2 diabetes	rs7721099	Э	87936379	I	U	TMEM161B	1.05	1.04	1.07	1.41E-09	
Turne O diabataa	TO 200525972	E	122652913	GT	~	CEP120	0.04	0.88	0.94	4.005.00	
Type 2 diabetes	18200525673	5			G	PRDM6	0.91			4.90E-09	
T O P I I I	00040	-	00740040	Т	~	STEAP1	4.00	1.04	4	4 005 00	
l ype 2 diabetes	1539210	/	09740010		C	ZNF804B	1.06	1.04	1.08	1.28E-09	
Type 2 diabetes	rs2277339	12	57146069	Т	G	PRIM1	1.06	1.04	1.08	4.53E-08	
Turne O diabataa	ro17105012	11	77375691	<u> </u>	٨	IRF2BPL	1.05	1 03	1.07		
Type 2 diabetes	1517105012	14		C	A	LRRC74A	1.05	1.03		1.54E-06	
Urolithiasis	rs4148155	4	89054667	А	G	ABCG2	1.12	1.08	1.16	1.70E-08	
Urolithiasis	rs12290747	11	3939650	т	С	STIM1	0.89	0.85	0.92	3.24E-09	
Arrhythmia	rs73205368	Х	23399501	Т	С	PTCHD1	1.08	1.06	1.10	4.25E-15	
Contrin consort	TO 100 FE 28	v	100540014	т	<u> </u>	GUCY2F	0.02	0.90	0.04		
Gastric cancer	191200020	^	106542014	1	C	IRS4	0.92	0.89	0.94	2.00E-10	
Loci detected in se	ex-specific ar	nalysi	S								
A a thurs a		0	00045500	^	~	LRRC3B	0.00	0.00	0.04		
Astrima	189836823	3	20840023	A	G	NEK10	0.86	0.82	0.91	5.19E-09	
Asthma	rs13227841	7	73279482	т	С	WBSCR28	0.86	0.81	0.90	2.04E-09	
Cataract	rs557090273	22	28942263	G	А	TTC28	2.71	1.89	3.87	4.52E-08	
- · ·		_			~	ZNF467	4.00				
Cerebral aneurysm	rs85591 <i>1</i>	7	149197364	A	G	ZNF746	1.22	1.14	1.31	2.72E-08	
Lung cancer	rs75932146	7	124487025	A	G	POT1	2.29	1.71	3.05	2.21E-08	
Osteoporosis	rs2864700	1	151028929	С	Т	CDC42SE1	1.18	1.11	1.25	2.03E-08	

Peripheral artery	ro70490749	10	41414401	C	^	CYP2A7	1 0 1	1.13	1.30	3.64E-08
disease	1572400740	19	4 14 1440 1	G	A	CYP2B6	1.21			
Type 2 diabetes	rs58202132	8	121764521	С	А	SNTB1	1.08	1.05	1.11	1.28E-08
Type 2 diabetes	rs2526678	11	61623793	G	Α	FADS2	0.93	0.91	0.96	2.44E-08
Type 2 diabetes	rs202209118	18	42361423	Т	TCC	SETBP1	1.16	1.10	1.22	7.78E-09

Summary data of the novel disease-associated variants are provided. For variants detected in sex-specific GWAS,
 statistics of sex with significant associations are provided. REF, reference allele; ALT, alternative allele; OR, odds
 ratio relative to the alternative allele; L95, lower 95% confidence interval; U95, upper 95% confidence interval;

281 COPD, chronic obstructive pulmonary disease. *, SIAH2 locus was also detected in an accompanying GWAS

282 project of breast cancer (Lee et al. in submission).

Table2. Missense variants in LD with nine novel disease-associated variants.

												All	ele freq	. in
												1K	G Phas	ie3
					Amino									
Disease	Variant	Chr.	Position	Gene	acid	REF	ALT	OR	L95	U95	Р	EAS	EUR	AFR
					change									
Loci detected	in sex-combi	ined an	alysis											
Coronary														
artery disease	rs11235604	11	72533536	ATG16L2	p.R220W	С	т	0.91	0.88	0.94	1.73E-08	0.100	0.000	0.000
Henatocellular														
	rs8103142	19	39735106	IFNL3	p.K70R	т	С	1.38	1.23	1.54	1.14E-08	0.083	0.312	0.695
carcinoma						_	~							
Keloid	rs192314256	1	201437730	PHLDA3	p.E62G	1	C	9.56	5.91	15.45	3.28E-20	0.010	0.000	0.000
Tune 2 diabetes	rs2303720	5	122682334	CEP120	p.R921H	C	т	0.91	0.89	0.94	1 77E-08	0 084	0.026	0.048
Type 2 diabetes	132000720	0	122002004	021 120	p.R947H	0		0.01	0.00	0.54	1.172-00	0.004	0.020	0.040
Type 2 diabetes	rs194520	7	89854446	STEAP2	p.F17C	т	G	1.06	1.04	1.08	1.23E-08	0.183	0.512	0.286
Type 2 diabetes	rs194524	7	89861832	STEAP2	p.R456Q	G	A	1.06	1.04	1.08	1.18E-08	0.183	0.512	0.269
Type 2 diabetes	rs2277339	12	57146069	PRIM1	p.D5A	т	G	1.06	1.04	1.08	4.53E-08	0.206	0.111	0.199
Urolithiasis	rs2231142	4	89052323	ABCG2	p.Q141K	G	т	1.12	1.08	1.16	1.74E-08	0.291	0.094	0.013
Loci detected	in sex-specif	ic analy	ysis											
Asthma	rs13246460	7	73249165	WBSCR27	p.R216W	т	A	0.89	0.85	0.93	9.23E-07	0.581	0.726	0.463
Asthma	rs13232463	7	73249299	WBSCR27	p.S171W	G	С	0.89	0.85	0.93	9.23E-07	0.581	0.726	0.463
Asthma	rs13241921	7	73254812	WBSCR27	p.Q107R	т	С	0.89	0.85	0.93	9.54E-07	0.581	0.726	0.464
Asthma	rs11770052	7	73275565	WBSCR28	p.I14N	т	Α	0.87	0.83	0.92	3.20E-08	0.630	0.726	0.285
Asthma	rs13227841	7	73279482	WBSCR28	p.W78R	т	С	0.86	0.81	0.90	2.04E-09	0.650	0.677	0.334
Lung cancer	rs75932146	7	124487025	POT1	p.V326A	Α	G	2.29	1.71	3.05	2.21E-08	0.003	0.000	0.000

Summary data of the missense variants in LD ($r^2 > 0.6$) with novel disease-associated variants are provided. For variants detected in sex-specific GWAS, statistics of sex with significant associations are provided. Rows

highlighted in bold indicate east Asian-specific variants (MAF = 0 in Europeans of 1KG Phase3). REF, reference

allele; ALT, alternative allele; Allele freq., allele frequency of alternative allele; OR, odds ratio relative to the

alternative allele; L95, lower 95% confidence interval; U95, upper 95% confidence interval; EAS, East Asian

290 populations; EUR, European populations; and AFR, African populations. *, the lead variant at this region

291 (rs8107030) is in LD (r² = 0.880) with ss469415590 (comprised of rs67272382 and rs74597329). ss469415590 is a

frameshift variant of *IFNL4*³⁷ (not listed as a functional gene in GRCh37 coordinates), and it might also explain this association.

294



298

299 Figure 1. Overview of 331 loci detected in this GWAS.

a, Phenogram³⁸ of 331 loci detected in this GWAS. Novel loci (\blacklozenge) were annotated by the closest gene names. Pleiotropic associations (see Methods for its definition) were plotted at the same position. **b**, Allele frequencies and the odds ratios of the lead variants at 331 loci detected in this GWAS. The odds ratio of the risk allele was used. *, loci detected in sex-specific GWAS. ¶, the lead variants are in LD with missense variants ($r^2 > 0.6$). **c**, **d**, Allele frequencies of the lead variants were compared between novel loci and known loci (**c**, East Asian populations; **d**,

305 European populations). The difference in MAF was tested by Mann–Whitney U test, and its *P* value was provided.

- 306 When MAF < 0.001, MAF was adjusted to 0.001 to fit in log scale. MAF_{EAS}, MAF in East Asian population (1KG
- 307 Phase3). MAF_{EUR}, MAF in European population (1KG Phase3).



308 309

Figure 2. Novel associations which can be explained by East Asian-specific missense variants.

311 Regional association plots are provided (**a**, coronary artery disease; and **b**, lung cancer). For coronary artery

312 disease (a), *P* values from conditional analysis and those in European GWAS¹² were plotted separately. For lung

313 cancer (**b**), *P* values from female- and male-specific GWAS were plotted separately.



Figure 3. A novel association of cerebral aneurysm can be explained by artery-specific eQTL signals for

317 **ATP2B1**.

315

a. Regional association plots of cerebral aneurysm GWAS at *ATP2B1* locus (top) and those of eQTL signals for

319 ATP2B1 in the tibial artery (bottom) are provided. The lead variant of GWAS (rs11105352; + dot) and the lead

320 variant of eQTL (rs2681492; ■ dot) are indicated by different shapes. Variants in LD with rs11105352 are

highlighted by red ($r^2 > 0.6$ both in East Asian and European populations of 1KG Phase3). **b**, Tissue-specificity of

eQTL signals for *ATP2B1* at rs2681492 (the lead variant of eQTL in the tibial artery (**a** dot in **a**)). *P* values in eQTL

323 analysis and M values (the posterior probability that an eQTL effect exist in each tissue tested in the cross-tissue

meta-analysis) in all tissues in GTEx project³⁹ are provided. Each dot indicates each tissue. All statistics of eQTL

- 325 analysis were derived from release v7 of GTEx project³⁹.
- 326









ό UMAP1 5

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329 Figure 4. Transcription factors whose binding sites were enriched for heritability of diseases.

- a, All of the 2,868 sets of TF binding sites grouped into 15 clusters were plotted in the UMAP space. **b and c,** The
- 331 results of S-LDSC were plotted on the UMAP space. The significant results (FDR < 0.05) are highlighted by cluster-
- 332 specific colors. The names of the top five most significant TFs are also shown on the plot. **b**, The results of red
- blood cell-related traits. **c**, The results of diseases in this GWAS which had more than five significant TF binding
- 334 site tracks (the results of other diseases are provided in Supplementary Figure 8).
- 335
- 336

337 ONLINE METHODS

338 Subjects

339 All case samples in this GWAS were collected in the BioBank Japan Project (BBJ)^{4,5}, which is a biobank that 340 collaboratively collects DNA and serum samples from 12 medical institutions in Japan and recruited approximately 341 200,000 patients with the diagnosis of at least one of 47 diseases. Among them, cases with dyslipidemia were 342 excluded because it was already reported in our previous study²⁹. Amyotrophic lateral sclerosis and febrile seizure 343 were also excluded due to limited sample size. Cases with myocardial infarction, stable angina, and unstable 344 angina were re-classified into a single disease category (coronary artery disease). Thus, we analyzed 42 disease in 345 this study. For control samples, we used samples from the population-based prospective cohorts; the Tohoku 346 University Tohoku Medical Megabank Organization (ToMMo), Iwate Medical University Iwate Tohoku Medical 347 Megabank Organization (IMM)⁴⁰, the Japan Public Health Center-based Prospective Study and the Japan Multi-348 institutional Collaborative Cohort Study. In addition, we also included samples in BBJ without related diagnoses into 349 control group (Supplementary Figure 1). The sample sizes and the demographic data are provided in 350 Supplementary Table 1. All participating studies obtained informed consent from all participants by following the 351 protocols approved by their institutional ethical committees. We obtained approval from ethics committees of 352 RIKEN Center for Integrative Medical Sciences, and the Institute of Medical Sciences, The University of Tokyo.

353

354 Genotyping

355 We genotyped samples with the Illumina HumanOmniExpressExome BeadChip or a combination of the Illumina 356 HumanOmniExpress and HumanExome BeadChips. For quality control (QC) of samples, we excluded those with (i) 357 sample call rate < 0.98 and (ii) outliers from East Asian clusters identified by principal component analysis using the 358 genotyped samples and the three major reference populations (Africans, Europeans, and East Asians) in the 359 International HapMap Project⁴¹. For QC of genotypes, we excluded variants meeting any of the following criteria: (i) 360 call rate < 99%, (ii) P value for Hardy Weinberg equilibrium (HWE) < 1.0×10^{-6} , and (iii) number of heterozygotes 361 less than five. Using 939 samples whose genotypes were also analyzed by whole genome sequencing (WGS), we 362 added additional QC based on the concordance rate between genotyping array and WGS. Variants with a 363 concordance rate < 99.5% or a non-reference discordance rate $\ge 0.5\%$ were excluded. We note that the allele 364 frequency of rs671 (the East Asian-specific functional missense variant at ALDH2) substantially varies among the 365 domestic regions within Japan due to strong selection pressure⁴² and that genotypes of rs671 did not follow HWE. 366 We thus did not apply the HWE QC for rs671. We had confirmed the 100% concordance of rs671 genotypes 367 between the SNP microarray data used in this study and our internal WGS data (n = 2,798; see details in Matoba 368 N. et. al. manuscript in revision).

369

370 Imputation

We utilized all samples in the 1000 Genomes Project Phase 3 (version 5)⁷ as a reference for imputation. We first

- 372 prephased the genotypes with SHAPEIT2 (v2.778) and then imputed dosages with minimac3 (v2.0.1). After
- imputation, we excluded variants with imputation quality of Rsq < 0.7. For X chromosome, we performed

prephasing and imputation separately for males and females, and we excluded variants with imputation quality of
 Rsq < 0.7 in either of them.

376

377 Genome-wide association analysis

378 We conducted a GWAS by employing a generalized linear mixed model (GLMM) using SAIGE (v0.29.4.2)⁶. This 379 strategy enabled us to maintain related samples in our GWAS, and the sample sizes were increased by 6% on 380 average compared to removing related samples. Briefly, there are two steps in SAIGE. In step 1, we fit a null 381 logistic mixed model using genotype data, and we added covariates in this step (see below). In step 2, we 382 performed the single-variant association tests using imputed variant dosages. We applied the leave-one-383 chromosome-out (LOCO) approach. For the X chromosome, we conducted GWAS separately for males and 384 females, and merged their results by inverse-variance fixed-effect meta-analysis. We used only female control 385 samples for GWAS of female-specific diseases; breast cancer, cervical cancer, endometrial cancer, ovarian cancer, 386 endometriosis, and uterine fibroids. Similarly, we used only male control samples for GWAS of prostate cancer. We 387 incorporated age and top 5 principal component (PC) as covariates. We also used sex as covariate for GWAS of 388 diseases which include both of male and female samples. We created regional association plots by LocusZoom 389 (v1.2)⁴³.

We performed stepwise conditional analysis within ± 1 Mb from the lead variant; we repeated the association test by additionally incorporated the dosages of the identified variants as covariates in SAIGE step 1 until we do not detect any significant associations.

393 We also conducted male-specific and female-specific GWAS using the same pipeline as described above, 394 and estimated heterogeneity in the effect size estimates using Cochran's Q test.

We set a genome-wide significance threshold at $P = 5.0 \times 10^{-8}$. For each disease, we defined a significantly associated locus as a genomic region within ± 1 Mb from the lead variant. When a locus did not include any variants which were previously reported to be significantly associated with the same disease ($P < 5.0 \times 10^{-8}$), we defined it as a novel locus.

399

400 Estimation of heritability

We estimated heritability and confounding bias in our GWAS results with LDSC (v1.0.0)⁸ using the baselineLD model (v2.1)¹⁰ which include 86 annotations, including 10 MAF- and 6 LD-related annotations that correct for bias in heritability estimates⁹, and were calculated using 481 East Asian samples in 1KG Phase3. For the analysis using LDSC, we excluded variants in the HLA region (chr6:26 Mb-34 Mb). We also calculated heritability Z-score to assess the reliability of heritability estimation.

Absolute quantification of heritability estimation using GWAS results using GLMM can be biased because effective sample size could be different from the true sample size (relative quantification is not biased, and hence GWAS results using GLMM can be applied for genetic correlation analysis and S-LDSC safely). Therefore, to confirm the robustness of heritability estimation in our analysis, we also performed GWAS using generalized linear regression model (GLM). As simple GLM does not account for the bias caused by genetic relationships, we further 411 excluded related samples (Pi-hat by > 0.187), and we analyzed genotype data with PLINK (v1.90)⁴⁴ using the same

- 412 covariates as described above. Heritability estimates based on GWAS using two different methods (SAIGE vs
- 413 PLINK) were comparable level (Supplementary Table 2).
- 414

415 Comparison of GWAS results between populations

To compare the GWAS results of our study with those conducted in European populations, we prepared publicly available GWAS summary statistics for 10 diseases. Summary statistics for eight diseases were downloaded from GWAS Catalog (URL) and their names and their PMID were as follows; atrial fibrillation (30061737), breast cancer (29059683), coronary artery disease (29212778), glaucoma (29891935), ischemic stroke (29531354), prostate cancer (29892016), rheumatoid arthritis (24390342), and type 2 diabetes (30054458). Summary statistics of two diseases were downloaded from UK Biobank GWAS summary statistics at Neale Lab (URL) and their names and their phenotype code were as follows; asthma (22127), and congestive heart failure (I50).

423

424 Pleiotropy

We utilized the following variants detected in GWAS for each disease; (i) lead variants in the significantly associated loci, (ii) independent signals detected by conditional analysis, and (iii) lead variants detected in sex-

427 specific GWAS. We defined pleiotropic association when these variants were in LD ($r^2 > 0.6$). We calculated r^2 428 using East Asian samples in the 1KG Phase3⁷ by PLINK (v1.90)⁴⁴.

429

430 Functional annotation of associated variants

431 We utilized the same disease-associated variants as used in the previous section for this analysis. We calculated r^2 432 using East Asian samples (r²_{EAS}) and European samples (r²_{EUR}) in the 1KG Phase3⁷ by PLINK (v1.90)⁴⁴. We 433 annotated disease-associated variants with eQTLs detected in the Japanese population²¹ in the following 434 conditions; (i) the lead variants in eQTL study is in LD (r²_{EAS} > 0.6) with GWAS variants and (ii) Q value of the lead 435 variants in eQTL study is less than 0.05. We annotated GWAS variants with eQTL detected in the European 436 population (release v7 of GTEx project)³⁹ in the following conditions; (i) the lead variants of eQTL study is in LD 437 $(r^2_{EAS} > 0.6 \text{ and } r^2_{EUR} > 0.6)$ with GWAS variants and (ii) Q value of the lead variants in eQTL study is less than 438 0.05.

For the annotation of exonic nonsynonymous variants, we used ANNOVAR⁴⁵. We annotated GWAS variants with nonsynonymous variants when they are in LD ($r^2_{EAS} > 0.6$). GRCh37 coordinates were used in this study.

442

443 Genetic correlations between sex-specific GWAS

We estimated genetic correlations between our GWAS results by LDSC (v1.0.0)⁸ using East Asian LD scores which we presented in our previous study²⁹. We excluded variants in the HLA region (chr6:26 Mb-34 Mb). We analyzed 20 diseases based on two criteria; (i) heritability was reliably estimated (heritability Z-score > 2; Supplementary Table 2); and (ii) both of male and female patients were included.

448

449 Transcription factor binding sites

450 We obtained 3,158 raw human ChIP-seq data files in SRA format from the GEO database. We converted them to

- 451 FASTQ format using the fastq-dump function of SRA Toolkit. We performed QC of sequence reads using FastQC.
- 452 We mapped these reads to the genome assembly GRCh37 using Bowtie2 (v2.2.5) with default parameters. We
- 453 called peaks using MACS (v2.1) with default parameters (q < 0.01) and defined them as TF binding sites. We
- 454 excluded TF binding site tracks which do not have at least one binding region in every chromosome, and 2,868
- 455 genome-wide TF binding site tracks remained (Supplementary Table 9).
- 456

457 Stratified LD score regression

458 We conducted stratified LD score regression (S-LDSC)²⁵ to partition heritability. For S-LDSC analysis of sex-459 specific GWAS of asthma, we used 220 cell-type specific annotations used in previous articles^{25,29}. For other S-460 LDSC analysis, we used TF binding site tracks which were described in the previous paragraph. For all sites of TF 461 binding, we empirically extended sites by 500 bp at the both ends for this analysis. We computed annotation-462 specific LD scores using the 1000 Genomes Project Phase 3 (version 5) East Asian reference haplotypes⁷. We 463 estimated heritability enrichment of binding sites of each TF, while we controlled for the merged binding sites of all 464 TFs and the 53 categories of the full baseline model available at the authors' website (see URLs). We did not use 465 the baselineLD model (v2.1)¹⁰ in this analysis to avoid false negative findings. We excluded variants in the HLA 466 region (chr6:26 Mb-34 Mb). We analyzed 24 diseases whose heritability was reliably estimated (heritability Z-score 467 > 2; Supplementary Table 2). We calculated the P value of the regression coefficient. For each trait, we calculate 468 FDR using the Benjamini-Hochberg method. We set a significance threshold at FDR < 0.05 for this analysis.

469

470 Visualization of TF binding sites

471 There is a complex correlation structure among 2,868 TF binding site tracks used for S-LDSC analysis. In S-LDSC, 472 we regress GWAS chi-squared statistics on LD-scores of each TF binding site (TF LD-score), and hence we 473 focused on correlations between TF LD-scores, not correlations between TF binding sites. We first performed PCA 474 using all TF LD-scores. To classify them into mutually correlated TF groups, we performed k-means clustering 475 (k=15) using top 15 PCs. We named each cluster by the most dominant TF in each cluster (Figure 4). The list of 476 each TF binding site and its assigned cluster name was provided in Supplementary Table 9. We then performed 477 uniform manifold approximation and projection (UMAP)²⁸ using top 15 PCs to project all TF binding sites into a two-478 dimensional space. UMAP was conducted using R package umap (v.0.2.0.0). Our workflow was illustrated in 479 Supplementary Figure 7.

480

481 **Data availability**

482 GWAS summary statistics of the 40 diseases (all except breast cancer and coronary artery disease) are publicly

- 483 available at our website (JENGER; see URLs) and the National Bioscience Database Center (NBDC) Human
- 484 Database (Research ID: hum0014) without any access restrictions. For breast cancer and coronary artery disease,

- 485 we will deposit the results after acceptance to the journal due to accompanying projects. GWAS genotype data for
- 486 case samples were deposited at the NBDC Human Database (Research ID: hum0014).

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- 574
- 575

576	URLs.
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578	JENGER, <u>http://jenger.riken.jp/en/</u>
579	PLINK 1.9, https://www.cog-genomics.org/plink2
580	MACH, http://csg.sph.umich.edu//abecasis/MaCH/
581	Minimac, https://genome.sph.umich.edu/wiki/Minimac
582	SAIGE, https://github.com/weizhouUMICH/SAIGE
583	GWAS Catalog, https://www.ebi.ac.uk/gwas/
584	Neale Lab, http://www.nealelab.is/uk-biobank
585	PASCAL, https://www2.unil.ch/cbg/index.php?title=Pascal
586	ldsc, <u>https://github.com/bulik/ldsc/</u>
587	LD score, https://data.broadinstitute.org/alkesgroup/LDSCORE/
588	ANNOVAR, http://annovar.openbioinformatics.org/en/latest/
589	Locuszoom, http://locuszoom.sph.umich.edu/locuszoom/
590	R, https://www.r-project.org/
591	SRA Toolkit, https://www.ncbi.nlm.nih.gov/sra/docs/toolkitsoft/
592	FASTQC, https://www.bioinformatics.babraham.ac.uk/projects/fastqc/
593	Bowtie2, http://bowtie-bio.sourceforge.net/bowtie2/manual.shtml
594	MACS, https://github.com/taoliu/MACS
595	NBDC Human Database, https://humandbs.biosciencedbc.jp/en/.
596	1000 Genomes Project, <u>www.1000genomes.org/</u>
597	
598	ACKNOWLEDGMENTS:
599	We acknowledge the staff of BBJ for their outstanding assistance. We express our heartfelt gratitude to Tohoku

- We acknowledge the staff of BBJ for their outstanding assistance. We express our heartfelt gratitude to Tohoku
 University Tohoku Medical Megabank Organization (ToMMo), Iwate Medical University Iwate Tohoku Medical
 Megabank Organization (IMM), the Japan Public Health Center–based Prospective (JPHC) Study, and the Japan
 Multi-Institutional Collaborative Cohort (J-MICC) Study for their invaluable contributions to collecting control
 samples. We also express our gratitude to E.K. and H.S. for kindly sharing their results of ChIP-seq data analysis.
 - 604 We extend our appreciation to Y. Yukawa, Y. Yokoyama, and other members of the Laboratory for Statistical
 - 605 Analysis, RIKEN Center for Integrative Medical Sciences for their great support. This research was supported by
 - 606 the Tailor-Made Medical Treatment Program (the BioBank Japan Project) of the Ministry of Education, Culture,
 - 607 Sports, Science, and Technology (MEXT) and the Japan Agency for Medical Research and Development (AMED).
 - The JPHC Study has been supported by the National Cancer Center Research and Development Fund since 2011
 - and was supported by a Grant-in-Aid for Cancer Research from the Ministry of Health, Labour and Welfare of
 - 510 Japan from 1989 to 2010. The study of psychiatric disorders was supported by AMED under Grant Numbers
 - 611 JP18dm0107097, JP18km0405201 and JP18km0405208.
 - 612

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- 616 Y.Momozawa and M.Kubo performed genotyping. H.S and E.K. analyzed ChIP-seq data. M. Ikeda and N. I
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- 618 I.K, T.Tanaka, H.N, A.Suzuki, T.H, M.T, K.C, D.M, M.M, S.N, Y.D, Y.Miki, T.Katagiri, O.O, W.O, H.I, T.Yoshida, I.I,
- T.Takahashi, C.Tanikawa, T.S, N.Sinozaki, S.Minami, H.Yamaguchi, S.A, Y.T, K.Yamaji, K.T, T.F, R.T, H.Yanai,
- 620 A.M, Y.Koretsune, H.K, M.H, S.Murayama, K.Yamamoto, Y.Murakami, Y.N, J.I, T.Yamauchi, T.Kadowaki, M.Kubo,
- and Y.Kamatani contributed to the management of BBJ data. N.Minegishi, Kichiya Suzuki, K.Tanno, A.Shimizu,
- T.Yamaji, M.Iwasaki, N.Sawada, H.U, K.Tanaka, M.N, M.S, K.W, S.T, and M.Y contributed to the management of
- 623 cohort control data. S.R, J.I, T.Yamauchi, T.Kadowaki, M.Kubo, and Y.Kamatani jointly supervised this study.
- 624

625 **COMPETING FINANCIAL INTERESTS**:

- 626 The authors declare no competing financial interests.
- 627