A macro-ecological approach to predators' functional response

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Abstract

Predation often deviates from the law of mass action: many micro- and meso-scale experiments have shown that consumption saturates with resource abundance, and decreases due to interference between consumers. But does this observation hold at macro-ecological scales, spanning many species and orders of magnitude in biomass? If so, what are its consequences for large-scale ecological patterns and dynamics?

We perform a meta-analysis of predator-prey pairs of mammals, birds and reptiles, and show that predation losses appear to increase, not as the product of predator and prey densities following the Lotka-Volterra (mass action) model, but rather as the square root of that product. This suggests a phenomenological power-law expression of the effective cross-ecosystem functional response. We discuss whether the same power-law may hold dynamically within an ecosystem, and assuming that it does, we explore its consequences in a simple food chain model. The empirical exponents fall close to the boundary between regimes of donor and consumer limitation. Exponents on this boundary are singular in multiple ways. First, they maximize predator abundance and some stability metrics. Second, they create proportionality relations between biomass and productivity, both within and between trophic levels. These intuitive relations do not hold in general in mass action models, yet they are widely observed empirically.

These results provide evidence of mechanisms limiting predation across multiple ecological scales. Some of this evidence was previously associated with donor control, but we show that it supports a wider range of possibilities, including forms of consumer control. Limiting consumption counter-intuitively allows larger populations. It is worthwhile to reconsider whether the observed functional response arises from microscopic mechanisms and constraints, or could hint at selective pressure at the population level.

1 Introduction

Dynamical food web models must summarize, at the population level, the complexity of predator and prey behavior, physiology and ecology, often aggregating over both space and time. Since Lindeman [1], all models of trophic dynamics have shared the same fundamental structure, i.e. a balance of energy gains and losses from predation and from non-trophic processes. Ecologists have typically circumvented the hidden complexity and specificity of a given ecosystem by summarizing it in a simple function, the predator functional response, which describes how fluxes between trophic levels (predation) depend on stocks at each level (predator and prey abundances). The most basic assumption is *mass action* – encounter probability scaling as the product of predator and prey density – but natural settings often deviate from this baseline [2]. The expression of the functional response has been the topic of extensive empirical and theoretical inquiries, and even major debates [3;4]. But there is no reason to expect that a single form of functional response, whether mechanistic or phenomenological, will be appropriate to model all ecological processes [5;6]. We may have to derive different expressions for different questions and scales, from short-term laboratory experiments [7],

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through multi-generation predator-prey cycles[8], to macro-scale food web properties such as biomass pyramids and stability[9].

Here we propose a new, empirically-motivated and phenomenological formulation of the functional response at large spatio-temporal scales. By contrast with mechanistic theories[10–13], our aim is not to ascertain the most realistic expression of predation rates, but to find a simple expression that can mimick observed rates across multiple scales, and then to understand its consequences for other food chain properties.

We first perform a meta-analysis of predation rates observed for higher vertebrates in the field, covering a range of species, study durations and areas. In the existing literature, most measurements of the functional response (predation rate as a function of predator and prey density) come from feeding experiments, restricted to time scales far shorter than the predator's generation time[11;14]. While these experimental measurements are valuable for mechanistic models of specific populations, they may be misleading when extrapolating to macro-scale dynamics[5]. For the latter, we need a phenomenological expression that exhibits some form of scale invariance, so that it may hold across different levels of aggregation. We propose a power-law expression similar to the Cobb-Douglas production function in economics[15]. We then find best-fit exponents suggesting a square root dependence of predation on both predator and prey density. These exponents are starkly lower than those expected from the simple mass action (Lotka-Volterra) model, suggesting that hidden mechanisms strongly limit predation.

This empirical trend can only be interpreted as a meaningful functional response if it can be used to predict large-scale dynamics. This raises multiple questions. Knowing that local populations undergo many complex processes on short time scales, is it possible to write an effective model that approximates the dynamics of abundances aggregated over years and over whole landscapes? What kind of equations and nonlinearities should be used in this effective model? These questions have been raised in prior studies, mainly within the theoretical literature[16–19], which have shown that effective models can indeed be constructed in some cases, and may be quite different from the dynamics known at small scales. Being phenomenological, such models often cannot easily be interpreted in terms of concrete mechanisms or instantaneous demographic events. They generally do not exhibit the intuitive properties of classical functional responses[20]. But when successful, they can provide useful predictions. The present study should be viewed as exploratory, investigating one hypothetical way of describing large-scale, long-term dynamics by a simple set of effective differential equations.

An important question remains: how can we parameterize our effective model from observational data, collected within or across ecosystems? Indeed, the empirical relations between predation rates and densities of predators and prey do not necessarily translate directly to an effective functional response, i.e. a true dynamical dependence between these three quantities. Other latent factors could be varying between measurements, and creating spurious relations. Nevertheless, we find hints that such confounding variation is weaker than the dependence that we wish to capture. We therefore assume that the observed scaling law can be inserted directly into a simple dynamical model, in order to explore its potential ecological consequences.

The empirical scaling exponents occupy a special position in our model, as they correspond to a maximum in predator abundance and some metrics of stability. While our sublinear model suggests that predation frequency is much lower than in a classic Lotka-Volterra model, it also leads to larger standing predator populations. This apparent paradox is closely related to the "hydra effect", a term that originally applies to situations where increasing individual predator mortality leads to larger predator populations[21;22], but can be extended to any situation, such as ours, where consumers that seem individually less fit (e.g. consume less prey than their competitors) may become more abundant in the long term[23].

A second unexpected consequence is the emergence of consistent scaling relationships between predator and prey densities, and between production and biomass within each trophic level. Such scaling laws are widely found in data and postulated in ecosystem models[19;24]. They are notably used in contexts of ecosystem management, such as fisheries[25;26]. But these relationships do not

emerge spontaneously from arbitrary dynamical models, and only prevail under particular ecological conditions[9]. Our fitted exponents recover simple proportionality rules that are usually associated with donor control.

Our work raises the possibility that consumptive saturation and interference may happen at all scales, not only the classic scale of individual predator behavior[2], and that this prevalence is tied to the empirical robustness of various allometric relationships, and the maximization of predator abundance and persistence.

2 Methods

2.1 Dataset

We compiled data from 32 observational studies (details in Appendix S4) that reported kill rates k_{12} (number of prey killed per predator per year) in the field, as well as prey density N_1 and predator density N_2. For each species, we also collected typical body mass measurements w (in kg) and mass-specific metabolic rate measurements m (in W/kg) from the AnAge database[27] and two studies by White and Seymour[28] and Makarieva et al.[29], as these traits were rarely reported for the populations studied in the field[30]. Whenever multiple values where found, we used the average values for each species.

The final dataset comprises 46 predator-prey pairs including 26 predator species and 32 prey species. Predator species belong to classes Mammalia, Aves and Reptilia, while prey species belong to classes Mammalia, Actinopterygii, Aves and Malacostraca.

For each predator-prey pair, we computed the rate of biomass loss in the prey population due to predation (in kg/km^2/year) as

$$C = k_{12} w_1 N_2 \tag{1}$$

and population biomass densities $B_1 = w_1 N_1$ and $B_2 = w_2 N_2$ in kg/km^2.

2.2 Food chain model

We investigate a simple food chain model of arbitrary length, where we follow the biomass density B_i of trophic level i as it changes through time,

$$\frac{dB_i(t)}{dt} = -\underbrace{L_i(t)}_{\text{internal losses}} + \underbrace{P_i(t)}_{\text{production}} - \underbrace{C_i(t)}_{\text{predation losses}} \tag{2}$$

where L_i and P_i represent internal losses and biomass production at level i, while C_i represents losses from predation by level $i+1$. Production at level 1 arises from autotrophic growth

$$P_1(t) = g_1 B_1(t) \tag{3}$$

while at higher levels, we assume

$$P_{i+1}(t) = \varepsilon C_i(t) \tag{4}$$

where ε (taken here to be constant for simplicity[1]) is the conversion rate between the biomass lost through predation at level i and produced at level $i+1$. Internal losses can arise from individual metabolic costs and mortality, μ_i, and from self-regulation or density-dependent processes such as competition and pathogens, D_i,

$$L_i(t) = \mu_i B_i(t) + D_i B_i^2(t). \tag{5}$$

[1]The conversion rate ε is known to vary, but its range is limited compared to other quantities which can span orders of magnitude, see discussion in Barbier and Loreau[9]. We choose to ignore ecological contexts in which production may be density-dependent at fixed consumption, e.g. reduced offspring number in denser groups, which could be represented by a function $\varepsilon(B_{i+1})$.

Alternatively, we later consider the possibility of a power-law expression $L_i(t) = l_i B_i^\delta(t)$ which can interpolate between these two types of losses.

In the following, we only consider stationary properties of the dynamics (2), and thus drop the time-dependence of all quantities, which are assumed to take their equilibrium values.

2.3 Functional response

Throughout this study, we consider phenomenological functional responses, emerging at the landscape level from unknown, potentially complex, underlying spatio-temporal dynamics[18]. Consequently, these functions may not follow traditional expectations and intuitions for mechanistic functional responses[20] – for instance, they might not scale additively with the number of prey or predator individuals or species.

We particulary investigate a power-law dependence of predation losses (see discussion in Box 1)

$$C_i \equiv C(B_i, B_{i+1}) = A B_i^\beta B_{i+1}^\gamma \tag{6}$$

with $\beta, \gamma \in [0, 1]$, and A a constant which contains the attack rate, i.e. the basic frequency of encounters for a pair of predator and prey individuals.

In this expression, $\beta = \gamma = 1$ recovers the classic mass action (Lotka-Volterra) model. An exponent $\beta < 1$ is our counterpart to the well-studied phenomenon of *prey saturation*: predation increases sublinearly with prey density, meaning that the ability of predators to catch or use each prey decreases with their availability[2]. On the other hand, $\gamma < 1$ indicates *predator interference*: larger predator densities lead to less consumption per capita[14]. In the following, we will consider various arguments suggesting that empirical exponents may instead approximately satisfy the relationship $\beta + \gamma \approx 1$, which would imply a simpler ratio-dependent form

$$C(B_i, B_{i+1}) \sim B_i^\beta B_{i+1}^{1-\beta} = B_{i+1} \left(\frac{B_i}{B_{i+1}} \right)^\beta \tag{7}$$

This relationship requires a strong effect of either or both interference and saturation.

These two predation-limiting phenomena are commonly reported and have been studied with various mathematical models, but classic functional responses are not appropriate for fitting data such as ours, compiled from different systems and scales. For instance, we can consider the standard saturating (Michaelis-Menten or Holling Type 2) response[2] and its extension, the DeAngelis-Beddington model[31;32] with both saturation H and interference I

$$C(B_i, B_{i+1}) = \frac{A B_i B_{i+1}}{1 + H B_i + I B_{i+1}} \tag{8}$$

where H/A is traditionally defined as the handling time. This expression is applicable to short-term feeding experiments, where consumption is observed for a particular predator-prey species pair whose densities are imposed. The parameters H and I are contextual: they are expected to depend on species traits[11], but also on the spatial structure and area of each study site. We discuss in Results and Appendix the difficulties associated with using this expression in a cross-ecosystem context[2].

2.4 Links between statistical and dynamical laws

We seek to identify the dynamical relationship $C(B_1, B_2)$ between predation losses and the biomasses of predator and prey in empirical data. In our theoretical analysis, we also discuss model predictions for another important observation, the scaling relations between production P and biomass B within and between trophic levels[19].

[2]While it may produce similar results as (6), it only does so for carefully tuned H and I, and in a fully-saturated regime ($HB_i + IB_{i+1} \gg 1$) where this model does not seem meaningful as a mechanistic account of predation.

Box 1: Dimensional analysis and functional response

The difficulty of building a phenomenological cross-scale functional response is that it should account for both:

- between-systems parameter differences (e.g. different study areas, spatial structures, species traits),
- the within-system range of dynamical variables B_i that can lead to saturation and interference if they cross system-dependent thresholds.

We first appeal to dimensional analysis[33]. From (2) we see that predation losses C_i have the dimensions of biomass density over time, e.g. kg/km^2/year. A classic nondimensionalization choice[34] is to express these losses using predator density and mass-specific metabolic rate (see below for other options)

$$f_i = \frac{C_i}{m_{i+1} B_{i+1}} \tag{9}$$

where f_i is the normalized functional response, incorporating all biological mechanisms. Since f_i has no dimension, it can only depend on dimensionless quantities. If we assume that all variation in predation is determined solely by species densities and metabolic rates, we can only construct two dimensionless ratios π_1 and π_2 and f_i must take the form

$$f_i = f\left(\pi_1 = \frac{m_i}{m_{i+1}}, \pi_2 = \frac{B_i}{B_{i+1}}\right) \tag{10}$$

with some arbitrary function f. The choice $f(\pi_1, \pi_2) = 1$ leads to consumer dependence ($C = m_{i+1} B_{i+1}$), while $f(\pi_1, \pi_2) = \pi_1 \pi_2$ leads to donor dependence ($C_i = m_i B_i$). More generally, the expression (10) is strongly reminiscent of ratio-dependent functional response[35], but it does not posit a specific functional form.

To deviate from this ratio-dependent expression, we must use additional parameters to construct other dimensionless quantities. Setting aside metabolic scaling for now, the functional responses (6) and (8) in the main text involve B_i and B_{i+1} separately, suggesting the form

$$f_i = f\left(\frac{B_i}{B_{i,\min}}, \frac{B_{i+1}}{B_{i+1,\min}}\right). \tag{11}$$

Indeed, it is plausible for nonlinear density-dependence in f to involve system-specific thresholds $B_{i,\min}$ (see discussion in Appendix S1). These thresholds can be derived from other parameters that characterize each system, such as movement range or body mass. Another possibility, not investigated here, is that other biological rates may intervene if metabolic rate cannot be used as a universal "clock" for trophic processes[36].

Dimensionless ratios can capture important differences between systems, and a number of heuristics and mathematical theories exist to guide their choice[37]: for instance, it is often useful to find ratios that are significantly larger than 1 in some systems and smaller than 1 in others, as this can indicate qualitatively different regimes.

Yet, the problem of relating cross-ecosystem statistical laws to underlying dynamical models is subtle, due to the potential covariance between variables and parameters. As a simple example, consider the scaling between production and biomass in one trophic level, assuming that the underlying dynamical law is $P(B) = rB$ with r a fixed growth rate. When we consider many ecosystems at equilibrium, each with a different growth rate, their values of r and B will likely be correlated[3]. If they covary perfectly, the apparent trend will be $P \sim B^2$ and will not be representative of the true dynamical law.

The same issues occur for all observational scalings, including the scaling of predation losses with prey and predator densities: if we try to fit the relationship $C_{12} = AB_1^\beta B_2^\gamma$ in (6), we may encounter the problem of correlations between attack rates and population densities. We now discuss how to address this problem in the context of an empirical estimation of β and γ.

2.5 Empirical analysis

The empirical identification of a relationship between predation losses and species densities such as

$$C_{12} \sim B_1^\beta B_2^\gamma$$

can be affected by two sources of error: colinearity between the known variables B_1 and B_2, and colinearity between either of these variables and other unknown factors that may appear in the full expression of C_{12}.

The first problem is that the colinearity between B_1 and B_2 (Fig. 1d) may obscure their respective contributions to $C_{12} = C(B_1, B_2)$. We can overcome this difficulty by performing a commonality analysis, which is a statistical test based on variance partitioning. We use the function *regr* of the R package *yhat* to check the existence of suppression i.e. the distorsion of regression coefficients due to the colinearity between predictors[38]. This test shows the absence of suppression, allowing us to directly interpret the respective contributions of B_1 and B_2 to $C(B_1, B_2)$ below.

A more important difficulty stems from correlations with other parameters, as suggested in the previous section. If we suppose that, in each ecosystem k, predation losses follow

$$C^{(k)}(B_1, B_2) = A^{(k)} B_1^\beta B_2^\gamma \tag{13}$$

where we are given no information on the ecosystem-dependent parameter $A^{(k)}$, we can only estimate β and γ from a simple empirical fit if $A^{(k)}$ is weakly correlated with the observed values of B_1 and B_2. This could be the case if $A^{(k)}$ varies less between ecosystems, or has less impact, than other factors that control B_1 and B_2, such as primary production or mortality, or external factors such as immigration, out-of-equilibrium dynamics, or other interactions.

We employ two distinct methods to identify the exponents β and γ in our proposed power-law scaling relationship.

The first approach is a direct fit, with the aim of minimizing residual variance: given that C, B_1 and B_2 span multiple orders of magnitude, any combination of these variables that exhibits much less spread suggests an interesting regularity, even though the precise values of the exponents may be incorrectly identified due to colinearities and should be interpreted with care.

[3] Consider for instance the simple case of logistic growth:

$$\frac{dB}{dt} = rB\left(1 - \frac{B}{K}\right) = \underbrace{rB}_{P\ \text{(production)}} - \underbrace{DB^2}_{L\ \text{(internal losses)}} \tag{12}$$

We can imagine two scenarios that lead to different scalings between P and B across ecosystems at equilibrium (each with $P - L = 0$):

- $K = r/D$ changes across systems while r is constant: $P(B) = rB$ so $P \sim B$,
- r changes across systems while D is constant: $P = L(B) = DB^2$ so $P \sim B^2$ (at equilibrium, $B \sim r$)

This example highlights the importance of knowing which parameter (here r or D) varies most between ecosystems to predict the empirical scaling law, and from it deduce the true dynamical relationship $P(B) = rB$.

The second approach is an attempt to account for at least some unknown factors that may undermine a naive parameter estimation. We focus on the dynamical food chain model (2) and the power-law response (6). We then simulate this model for many random combinations of parameter values (including exponents, but also attack rate, growth rate, mortality, etc.). We only retain model realizations that give triplets (B_1, B_2, C) sufficiently close to those found in our empirical data, giving us a distribution of exponents compatible with this evidence, as well as an estimation of how other parameters may affect observed species densities and predation rates.

2.5.1 Direct fit

We note that the factor $A^{(k)}$ is not unitless. This may be problematic when fitting data collected at different scales. Following the discussion in Box 1, we propose to expand $A^{(k)}$ in a way that accounts for metabolic scaling and nonlinearity thresholds, i.e.

$$C^{(k)}(B_1, B_2) = a^{(k)} m_2 B_2 \left(\frac{m_1}{m_2}\right)^{\nu} \left(\frac{B_1}{B_{1,\min}}\right)^{\beta} \left(\frac{B_2}{B_{2,\min}}\right)^{\gamma-1} \tag{14}$$

where $m_2 B_2$ sets the dimensions of C and all other factors are dimensionless. In particular, $a^{(k)}$ is an ecosystem-dependent parameter that can be interpreted as a random effect. We will see in the Results section that many terms in this complex equation appear to have a negligible contribution, leading back to our simple theoretical expression (6).

We have no direct measurement of the reference densities $B_{i,\min}$, which are necessary to construct dimensionless ratios, and which we interpret as thresholds for nonlinearity (saturation or interference). Therefore, we must make some assumptions on these values, guided by biological intuitions. The simplest possible assumption is that these thresholds vary independently of other parameters, and can be included into the random effect $a^{(k)}$. As a second option, if these density thresholds were proportional to individual body masses, $B_{i,\min} \sim w_i$, the relevant variables for our fit would be population densities N_i rather than biomass densities. A third option, $B_{i,\min} \sim 1/m_i$, would instead suggest focusing on $E_i = m_i B_i$ which has commonly been interpreted as an "energy" density[39]. For each of these choices, the exponents ν, β and γ may then be identified by a least-squares linear regression in log space. To identify the correct definition of $B_{i,\min}$, we test all three possibilities, and ask which one leads to the best predictions, i.e. the lowest amount of residual variation in $a^{(k)}$.

2.5.2 Dynamical model estimation

Using our dynamical model (2) with (5) and (6), we generate predator-prey pairs with different parameter values drawn from a broad prior distribution, and compute their equilibrium biomass densities B_1 and B_2 and predation losses C_{12}. We then retain only model realizations for which the three values (B_1, B_2, C_{12}) are close to those observed in our data. We repeat this process until we have retained 2500 model realizations, and use the parameter values of these realizations to compute a joint posterior distribution for all our parameters, in particular the exponents β and γ that we wish to identify.

For two species, our model has 8 parameters, which we draw independently in each model realization. The exponents β and γ are drawn uniformly in the range $[0, 1.5]$ in order to confirm that they are expected to be less than 1 (the Lotka-Volterra or mass action limit). Other parameters are drawn uniformly on a log scale, over a span of 10 orders of magnitude: the attack rate $A \in [10^{-5}, 10^{5}]$, prey productivity $g \in [10^{-4}, 10^{6}]$, predator mortality $\mu_2 \in [10^{-8}, 10^{2}]$, and prey and predator self-regulation $D_1, D_2 \in [10^{-5}, 10^{5}]$. Finally, the biomass conversion efficiency is drawn in the realistic interval $\varepsilon \in [10^{-2}, 1]$, see discussion in Barbier and Loreau[9].

We retain only model realizations in which both species survive, and filter the remaining according to their similarity to empirical data. To do so, we train a Kernel Density Estimator (KDE, from the Python library *scikit-learn* using classes *GridSearchCV* and *KernelDensity*) on triplets $(\log_{10} B_1, \log_{10} B_2, \log_{10} C_{12})$ in the empirical data. We then compute the same triplet in model

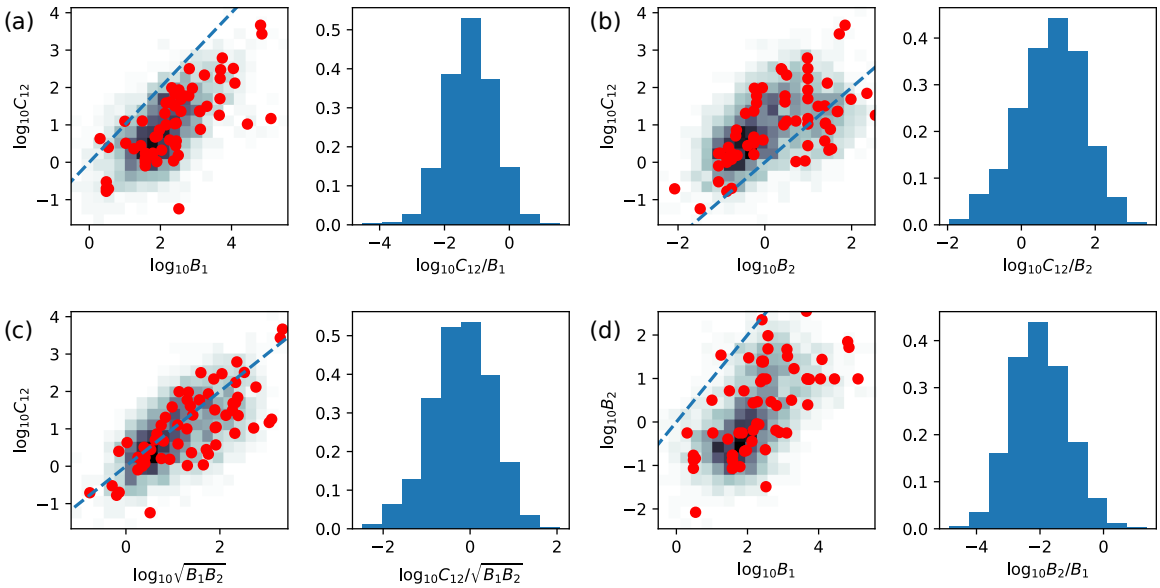

Figure 1: Empirical and simulated scaling relationships betwen biomass densities B_1 and B_2 and predation losses C_{12} in logarithmic scale. We show two panels for each pair among (a) C_{12} and B_1, (b) C_{12} and B_2, (c) C_{12} and $\sqrt{B_1 B_2}$, and (d) B_2 and B_1. **Left panels:** Each point corresponds to one study in our meta-analysis, and the dashed line represents the 1:1 relationship. The background color indicates the density of simulated realizations of our dynamical model, retaining only realizations that approximate the empirical distribution (see Sec. 2.5.2). **Right panels:** Histogram of ratio between the two quantities in our simulated ecosystems.

realizations, and its score s given by the KDE. We retain each model realization with probability $e^{\theta(s-s_{\max})}$, given $s_{\max}$ the highest score given among empirical triplets. We choose $\theta = 2$ as we find empirically that it gives the best agreement between simulations and data for the variance and range of B_1, B_2, C_{12} and their ratios.

Finally, we compute the posterior distribution of parameters, i.e. histograms of the parameter values found in accepted model realizations, and study their correlations with each other and with dynamical variables B_1, B_2 and C_{12}.

3 Results

3.1 Meta-analysis of kill rates

3.1.1 Results from direct fit

The full expression to be fitted (14) breaks down predation losses C_{12} into three different contributions: the observed species densities B_i raised to the exponents β and γ; the metabolic rates m_i, which are often assumed to define the time scale for predation processes; and the unknown typical density scales $B_{i,\min}$, which we interpret as thresholds for nonlinearity (e.g. the scale for interference or saturation).

Our first result is that there does not seem to be a systematic effect due to typical density scales $B_{i,\min}$. As explained in Sec. 2.5.1, our main three choices for these parameters correspond to using biomass densities B_i, number densities N_i or "energy" densities $E_i = m_i B_i$ as our fitting variables. To showcase these different options, we detail in Table 1 the residual standard deviation for various examples of power-laws involving these variables. These examples are only a few representative

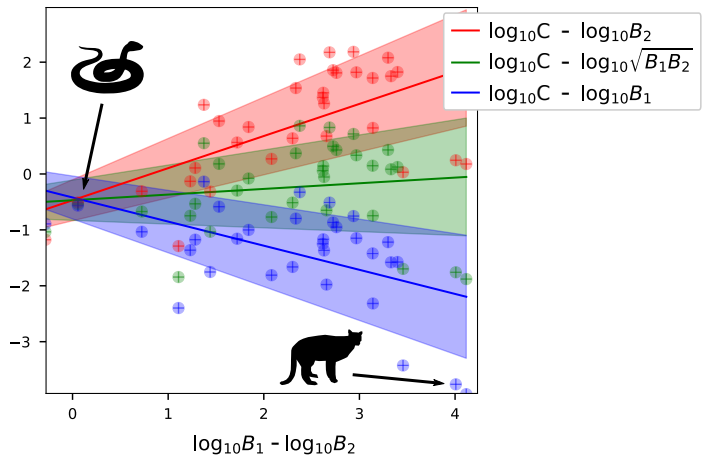

Figure 2: Residual variation with $\log_{10} B_1/B_2$ when assuming strict donor dependence ($C \sim B_1$), a symmetric square-root law ($C \sim \sqrt{B_1 B_2}$), or strict consumer dependence ($C \sim B_2$). Residues around the square-root law show no trend ($p = 0.4$, $R = 0.08$), and have minimal variance (see Table 1).

Expression y	$\frac{C}{B_1}$	$\frac{C}{\sqrt{B_1 B_2}}$	$\frac{C}{B_2}$	$\frac{C}{N_1}$	$\frac{C}{\sqrt{N_1 N_2}}$	$\frac{C}{N_2}$	$\frac{C}{E_1}$	$\frac{C}{\sqrt{E_1 E_2}}$	$\frac{C}{E_2}$	$\frac{C}{B_1^\beta B_2^\gamma}$
$std(\log_{10} y)$	0.90	**0.77**	0.97	1.6	1.5	1.5	0.93	**0.77**	0.9	**0.74**

Table 1: Residual standard deviation for ratios of predation losses and biomass density B_i, population density N_i, or "energy" density $E_i = m_i B_i$. See Fig. 2 for a visualization of the residual variation with B_1/B_2 for the first three expressions. For the rightmost expression, we use the best-fit exponents $\beta = 0.47$, $\gamma = 0.30$, and see that this contributes only a small reduction of residual variation compared with the square root expression (second column).

points in our systematic exploration of (14), but they illustrate our finding that number densities give the poorest results, while biomass or energy densities give comparable results. For simplicity, we then focus on biomass densities B_i, corresponding to the assumption that the reference density $B_{i,\min}$ is independent of body size or metabolism[4].

Our second result is the limited impact of metabolic rates on predation losses. From dimensional analysis (Box 1), we expected a baseline scaling $C_{12} \sim m_1$ or m_2. But we find that including m_i as a prefactor has no significant impact on any of our parameter estimates. Furthermore, irrespective of our choice of variable, our estimates for exponent ν in (14) are all of order $\nu \approx 0.1$, suggesting a weak and perhaps spurious dependence in metabolic rates. We therefore eliminate m_i from our expression (14). We performed a similar analysis (not shown here) using body masses w_i, with comparable results.

Given these two simplifications, we are left with the simpler relationship (6) i.e.

$$C_{12} \sim A B_1^\beta B_2^\gamma$$

to be fitted. We find the best-fit exponents $\beta = 0.47 \pm 0.17$ and $\gamma = 0.30 \pm 0.19$. The limited extent of our dataset implies large uncertainty on the exact exponents, but they are starkly sublinear.

[4]This is consistent with ssuming that the thresholds $B_{1,\min}$ and $B_{2,\min}$ are equal or proportional to the minimal sustainable population densities for prey and predators. Cross-species data from Hatton et al. [19] and Stephens et al. [40] suggest that the minimal numeric density is roughly inversely proportional to body mass, $N_{i,\min} \sim 1/w_i$. Therefore, $B_{i,\min} = w_i N_{i,\min}$ is expected to be size-independent, and exhibits no known link to metabolism.

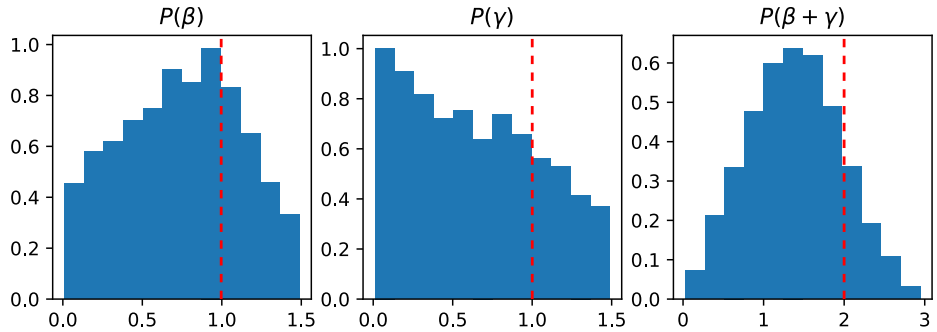

Figure 3: Posterior distributions of exponents β and γ in power-law functional response (6) for simulations selected to reproduce the empirical distribution of densities B_1 and B_2 and predation losses C_{12} (Fig. 1). The prior distribution was uniform over the interval $[0, 1.5]$. The posterior distributions remain broad, showing that exponents are only weakly constrained by our empirical evidence (with almost no correlation between exponents, $R = 0.11$). The most likely values are $\beta \approx 0.9, \gamma \approx 0$ which approximate the classic model of linear donor-dependence (Box 2), while the median values are more symmetrical, $\beta \approx 0.8, \gamma \approx 0.5$ and hence $\beta + \gamma \approx 1.3$. Both show significant departure from the mass action (Lotka-Volterra) hypothesis $\beta = \gamma = 1$ (dashed lines).

We notice that these exponents lay close to the line $\beta + \gamma = 1$, i.e. a simpler ratio-dependent model $C_{12} \sim B_2(B_1/B_2)^\beta$ which we could have assumed for dimensional reasons (Box 1). If we impose this relationship by setting $\gamma = 1 - \beta$, we find a similar best-fit value $\beta = 0.51 \pm 0.12$. Due to the colinearity between B_1 and B_2 ($R = 0.56$, Fig. 1d), we might expect any model of this form to provide a comparable fit, ranging from $\beta = 1$ (donor dependent, see Box 2) to $\beta = 0$ (consumer dependent). Nevertheless, our colinearity analysis strongly indicates $\beta \approx \frac{1}{2}$ as the most likely value. This is visualized by the fact that C/B_1 and C/B_2 both exhibit a residual variation in B_1/B_2 ($p < 10^{-6}$, $R = 0.61$ and $R = -0.52$ respectively, Fig. 2) whereas our best fit exhibits no residual variation ($p = 0.4$, $R = -0.08$), suggesting that neither strict donor nor consumer dependence prevails in our data.

3.1.2 Results from dynamical model estimation

We show in Fig. 3 the results of our second method of parameter estimation, described in Sec. 2.5.2, which consists in simulating our dynamical model with random parameter combinations, and retaining those that produce densities and predation losses comparable to those of the data.

This second approach also supports the idea that $\beta + \gamma$ (whose distribution is centered close to 1) is significantly reduced compared to the mass action expectation $\beta + \gamma = 2$. However, we notice a stark asymmetry in the posterior distributions of β and γ: the distribution of β has a mode below but close to 1, while the distribution of γ is strictly decreasing from 0. This might lend some credence to purely donor-dependent trophic fluxes, a hypothesis which lacked support in our direct fit approach above, see Fig. 2.

We must note that this model exploration is still missing many other factors that may affect B_1 and B_2 and their relationship to C_{12}, such as non-equilibrium dynamics or interactions with other prey and predator species, among others. These factors may have a greater impact on predators, whose populations are typically smaller, slower to attain equilibrium and more easily perturbed. We could thus expect that measured predator abundances are less strictly related to C_{12}. This decorrelation may be sufficient to explain the apparent bias toward small γ. We discuss below the hypothesis that our results are mainly driven by an artificial donor-dependence imposed by the scale of observation in the data.

Box 2: Definitions

We introduce two sets of terms to distinguish:

- what sort of *dependence* is assumed for predation losses C

  donor-dependent predation A model where $C(B_1, B_2) = C(B_1)$

  consumer-dependent predation A model where $C(B_1, B_2) = C(B_2)$

  ratio-dependent predation A model where $C(B_1, B_2) = C(B_1/B_2)$

- which of predation or non-trophic losses is the limiting factor for abundances:

  depleted regime A solution of our model where prey stocks are strongly depleted, so consumers compete over the flux of prey production. If we increase individual consumption (attack rate and exponents), fewer predators coexist at equilibrium.

  regulated regime A solution of our model where prey levels are mainly set by their own self-regulation, so that equilibrium predator abundance increases with attack rate and exponents.

These definitions should be distinguished from terms of *control* (e.g. donor control, bottom-up and top-down control), which are often used in more complex statements about the dynamics, e.g. how the system will respond to a perturbation, see discussion in Barbier and Loreau [9]. While these notions have strong overlaps (e.g. the regulated regime only allows bottom-up control), we employ a distinct terminology throughout this study to avoid possible confusions surrounding these terms.

3.1.3 Comparison with a classic functional response

The DeAngelis-Beddington (DAB) model (8), or Holling Type 2 with predator interference, can be fit to our data, but while it may be useful as a mechanistic model for a single system [6], it is unsatisfactory in our macro-ecological approach. We can either take handling time H and predator interference I as constant parameters across systems, or try to construct them as functions of other known parameters for each system (e.g. body sizes or metabolic rates). We show in Appendix S4 that we were unable to identify any expression that allows a meaningful fit: we always find that almost all predator-prey pairs are in the saturated part of the function, i.e. $HB_1 + IB_2 \gg 1$, meaning that we recover a simple ratio-dependent expression rather than a true saturating function.

Furthermore, we show that the colinearity $B_1 \sim B_2$ can make it difficult to differentiate between an additive law, $HB_1 + IB_2$, and a multiplicative law with exponents adding up to one, $B_1^{1-\beta} B_2^{\beta}$, in the denominator of the functional response. This has been put forward as an important criticism of the Cobb-Douglas production function [41] which resembles our power-law model (6). But we find here that prey and predator are equally limiting. This requires a fine-tuning of the ratio of handling time to interference, H/I, so that HB_1 and IB_2 are comparable in magnitude. This is found in our best fit of the DAB model (8) but is difficult to justify mechanisticially, whereas the same reality may be more clearly and simply encapsulated by $\beta \approx 0.5$.

3.2 Theoretical properties and consequences of empirical exponents

3.2.1 Population sizes

We consider two trophic levels, and investigate equilibrium predator density B_2^* in Fig. 4a as a function of both exponents β and γ. We observe two distinct regimes: there is only one stable equilibrium throughout the parameter space, but it is dominated either by species self-regulation, or by trophic regulation.

These regimes can be understood by first assuming that self-regulation can be neglected at both levels. In that case, our dynamical model (2) with the power-law functional response (6) has the

equilibrium solution (see Appendix S1)

$$B_1^* \approx \left(\frac{g_1^{1-\gamma}\left(\mu_2/\varepsilon\right)^{\gamma}}{A_{12}}\right)^{\frac{1}{\beta+\gamma-1}} \qquad B_2^* \approx \left(\frac{g_1^{\beta}\left(\mu_2/\varepsilon\right)^{1-\beta}}{A_{12}}\right)^{\frac{1}{\beta+\gamma-1}} \qquad \text{if } D_i \to 0, \tag{15}$$

where we see that $\beta+\gamma=1$ leads to a singularity, meaning that self-regulation cannot be neglected for these exponents.

The first regime is found when $\beta+\gamma>1$, and we call it the *depleted* regime (see Box 2 and Fig. 4b). In this case, the solution (15) is a good approximation at low self-regulation (small D_i). We notice that this solution has two counter-intuitive properties: predator biomass B_2 decreases with attack rate A_{12} and exponents β and γ, and increases with predator mortality μ_2 (if there is saturation, $\beta<1$). This is because prey stocks are brought to low levels, so predators simply divide between themselves the flux of prey production: the more each predator consumes per capita, the fewer predators can coexist. As we discuss below, this inverse relationship between individual fitness (high attack rate, low mortality) and population abundance has been demonstrated in various models and called the "hydra effect" [21;22].

If $\beta+\gamma<1$, the solution without self-regulation (15) becomes unstable. In the stable equilibrium, the biomass of prey will be mainly determined by their self-regulation i.e. $B_1^* \approx g_1/D_1$, leading to what we call the *regulated* regime (Box 2 and Fig. 4c). In that case, predator biomass

$$B_2^* \approx \left(\frac{\varepsilon A_{12}}{\mu_2}\right)^{\frac{1}{1-\gamma}} \left(\frac{g_1}{D_1}\right)^{\frac{\beta}{1-\gamma}} \tag{16}$$

increases with attack rate and with exponents β and γ. The fact that the previous equilibrium (15) is unstable reveals a type of Allee effect: if initial prey levels are below a threhsold, predators will consume all the prey (consumption will decrease too slowly to adjust to the falling resource level) then collapse.

We show in Fig. 4 that predator abundance is maximized at the boundary between these two regimes. This boundary lies close to the line $\beta+\gamma=1$, where predation losses scale as

$$C(B_1,B_2) \sim B_1^{\beta} B_2^{1-\beta}$$

which is close to the exponents suggested by our meta-analysis. We however note that the regime boundary moves away from this line as we decrease attack rates A_{12} and increase self-regulation D_i at both levels. For sufficiently high self-regulation, even Lotka-Volterra dynamics can enter the regulated regime, which we interpreted as bottom-up control in previous work [9].

The overarching pattern is that the predator density first increases with attack rate A_{12} and exponents β and γ, when predation losses are still negligible for the prey, up to the regime boundary described above. Past this point, faster consumption leads to significant resource depletion and a drop in predator population [24], as competition increases more than growth.

3.2.2 Dynamical stability

We also observe how functional response exponents affect the stability of predator populations. Fig. 4(d,e) represents the invariability (inverse of the coefficient of variation over time) of predator abundance, computed in two scenarios: demographic stochasticity, where fluctuations (arising from birth and death processes) are proportional to $\sqrt{B_2}$, and environmental perturbations which affect the predators proportionally to their abundance B_2, see Arnoldi et al. [42]. Other scenarios and more stability metrics (including asymptotic resilience) are represented in Fig. S1 in Supporting Information, and show qualitatively similar patterns.

We make three main observations. First, the optimal parameter values for stability depend on the perturbation scenario (e.g. demographic or environmental noise) and the choice of stability metric, as having a larger abundance leads to higher stability in some metrics [42]. It is therefore important

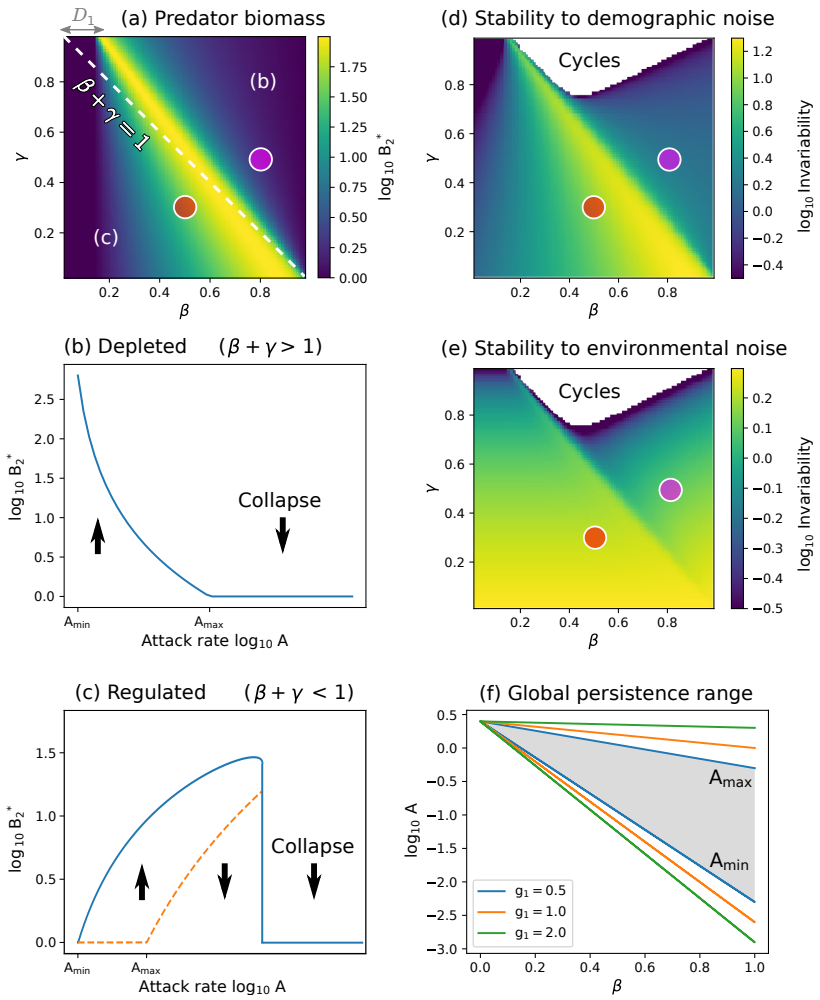

Figure 4: Equilibrium predator abundance and persistence in the dynamical predator-prey model ($D_i = 10^{-3}$ at both levels, $g_1 = 1$, $\mu_2 = 2$, $\varepsilon = 0.8$). On panels (a,d,e) the red dot indicates the exponents obtained by direct fit $\beta \approx 0.5$, $\gamma \approx 0.3$, while the purple dot indicates the median exponents from our dynamical model fit, $\beta \approx 0.8$, $\gamma \approx 0.5$. **(a)** Predator biomass density B_2^* (in log scale) varies with exponents β and γ in the power-law functional response (6). It is maximized on a boundary (lightest color on the graph) separating two different regimes (b and c). This boundary is close to the dashed line $\beta + \gamma = 1$, but deviates toward larger β due to prey self-regulation (the deviation increases with D_1). **(b)** For $\beta + \gamma > 1$, the resource is largely depleted by predation, and predator population decreases with attack rate, going extinct (for simplicity, set here to $B_2 < 1$) at $A_{\max}$. **(c)** For $\beta + \gamma < 1$, predator biomass (solid line) increases with attack rate, since the resource is mainly limited by its own self-regulation. The predator survives only if $A > A_{\min}$, but high attack rates $A > A_{\max}$ give rise to an unstable equilibrium (dashed line) under which populations can go extinct. The stable equilibrium disappears at higher A. **(f)** For the predator to persist in both regimes and for all initial conditions, we must have $A_{\min} < A < A_{\max}$ (shaded region). This interval expands with increasing β (less saturation) and basal growth g_1. **(d,e)** Predator invariability (1/coefficient of variation) around equilibrium (see e.g. Barbier and Loreau [9]; Arnoldi et al. [42]) in response to (d) demographic fluctuations and (e) environmental perturbations affecting the predator. For large γ and intermediate β, the equilibrum is unstable and replaced by limit cycles. This area is left blank as invariability cannot be computed analytically, but it remains well-defined in simulations.

to know which scenario and metric are relevant for empirical dynamics, and different ecosystems may possibly have different optima. Second, all the cases studied here display a ridge of increased predator stability at the maximum of its abundance, close to the transition line $\beta + \gamma = 1$, and a drop in stability right after this ridge. Third, while abundance depends almost symmetrically on β and γ, and is maximized close to the transition between regimes, stability favors systems closer to donor-dependence. Moving toward larger β along the line $\beta + \gamma = 1$ reduces the likelihood of cycles and widens the region of parameters with high stability both to demographic and environmental noise (Fig. 4d and e). More generally, stability to environmental perturbations can be improved by having lower γ than what would maximize abundance (i.e. going toward the bottom of Fig. 4e).

3.3 Scaling of biomass and production across levels

Previous literature has reported scaling laws between biomass and production within one trophic level, and between biomasses at different trophic levels[19]. We discussed in Sec. 2.4 the possible confounding factors that may prevent empirical scalings to be directly interpreted as dynamical relationships. In the following, we assume that there are no such confounding factors: this could be the case if, for instance, the main source of cross-ecosystem variation in B_i is taken to be the primary productivity g. We then show in Appendix S3 that our food chain model at equilibrium can give rise to a scaling of production with biomass

$$P_i \sim B_i^{\delta} \tag{17}$$

and the scaling of predator biomass with prey biomass

$$B_{i+1} \sim B_i^{\alpha}. \tag{18}$$

As discussed in Barbier and Loreau[9], neither of these laws can arise for more than two levels in a classic Lotka-Volterra model without self-regulation ($\beta = \gamma = 1$, $D_i = 0$).

There are only two ways in which these scaling laws can hold consistently across multiple trophic levels: either they are imposed by a narrow class of functional responses, or they originate from an equilibrium balance between trophic processes and other losses. We now show that only the second possibility is compatible with our data.

The first possibility requires internal losses to be negligible, i.e. $L_i = 0$ in (2), and predation to be strictly donor-dependent or consumer-dependent ($\gamma = 0$ or $\beta = 0$). The functional response then completely determines the scaling between production and biomass,

$$P_i \sim B_i^{\beta} \qquad \text{or} \qquad P_i \sim B_{i-1}^{\gamma} \tag{19}$$

and since $L_i = 0$,

$$P_{i+1} = \varepsilon P_i \qquad \Rightarrow \qquad B_{i+1} \sim B_i \tag{20}$$

which entails a strict proportionality between predator and prey biomass (exponent $\alpha = 1$). This has been widely discussed by proponents of the ratio-dependent functional response[35], and the existence of such proportionality laws was put forward as evidence of donor control. But since we find nonzero β and γ, this possibility is excluded by our data.

The second possibility arises in our model with a power-law functional response (6) and internal losses given by $L_i \sim B_i^{\delta}$. If we assume that all three terms $C_i \sim L_i \sim P_i$ remain of comparable magnitude (i.e. that a finite fraction of biomass is always lost to predation and likewise to self-regulation), the internal losses will then determine the scaling between production and biomass:

$$L_i \sim B_i^{\delta} \quad \text{and} \quad L_i \sim P_i \qquad \Rightarrow \qquad P_i \sim B_i^{\delta}. \tag{21}$$

For instance, we have $\delta = 1$ for individual mortality, and $\delta = 2$ when a self-regulating process, such as intra-specific competition or pathogens, is the main contributor to internal losses. We show in

Appendix S3 that that a relationship emerges between the exponents defined here and the functional response exponents

$$\alpha(\delta - \gamma) = \beta. \tag{22}$$

Commonly assumed scaling exponents for biomass and production are $\alpha = \delta = 1$, i.e. simple proportionality between predator and prey biomass, and between production and biomass at each level. These relations are used in various mass-balanced models[43] or theory on biomass pyramids[25]. They entail $\beta = 1 - \gamma$, meaning that the relationship between predation losses and biomasses would be the ratio-dependent expression $C_{12} \sim B_2(B_1/B_2)^\beta$ which we have discussed in previous sections.

4 Discussion

The predator functional response, describing how predation rates depend on predator and prey abundances, is a crucial component of trophic ecology that has a considerable impact on predator-prey dynamics[2]. Long-standing debates have opposed advocates of various forms of functional response, such as prey-dependent or ratio-dependent expressions[3;4]. The current consensus seems to be that there is no correct universal expression; rather, different functional responses are appropriate in different settings, and they should be derived mechanistically for each studied ecosystem[4;6].

From a macro-ecological point of view, however, what matters is how the functional response scales up across several orders of magnitude in abundance, body size and area. Experiments typically measure predation rates for a given pair of species, and how they respond to varying the number of prey and predator individuals, on a short time scale and in an enclosed space. Here, we used kill rates recorded in the field to estimate how predation rates over long times (e.g. a year) vary between species (mainly terrestrial vertebrates) and across much larger spatial scales.

Our empirical analysis points to a phenomenological power-law, which differs from commonly-studied functional responses. If this law does hold dynamically within each ecosystem, it may have striking theoretical implications. The empirical exponents fall close to the boundary between two qualitatively distinct dynamical regimes (approximately the line $\beta + \gamma = 1$). This boundary has two properties in our food chain model: predator abundance is maximized, and simple allometric scaling laws appear between different ecosystem functions and between trophic levels. We may then ask whether these two important macroscopic properties arise accidentally from a functional response shaped by other constraints, or whether the functional response is in fact selected for its dynamical consequences.

4.1 Is the empirical functional response an artefact of donor-dependence?

Our theoretical model suggests that both saturation and interference can lead to similar qualitative consequences for many ecosystem properties. Yet, two main empirical results suggest that reality is closer to classic donor-dependence, i.e. assuming that trophic fluxes are only determined by the amount of resource, and in turn determine the abundance of consumers, and not the reverse. The first clue is the fact that C_{12}/B_1 is largely below 1, and the second is that exponent γ is consistently lower than β. As discussed previously, the precise values of these exponents are weakly constrained empirically and could be confounded by unknown latent factors, disguising an underlying process where $\beta = 1$ and $\gamma = 0$.

To properly assess the value of these clues, it is important to recognize that donor-dependence is not a necessity. Predation losses are in principle limited not by standing biomass, but by production,

$$C_i \lesssim P_i$$

which can be much larger than standing biomass. For instance, a Lotka-Volterra model with a classic biomass pyramid $B_{i+1} \sim B_i$ would predict the predation and production of all consumers to scale *quadratically* with biomass, as $B_{i+1}B_i \sim B_i^2$. In other words, more and more productive systems would exhibit more and more predation, but populations would only increase as the square

root of these fluxes, and prey density B_1 would not represent a bound on predation. A supralinear scaling of production and losses with biomass is not implausible. Analogous relationships have been proposed in economic systems, with empirical exponents ranging from 1.16 (economic growth as a function of city size[44]) to 1.6 (stock trading rate as a function of number of shares[45]). We could also invoke empirical facts that have been interpreted as evidence for the Allee effect[46], since this effect assumes that per-capita growth rate increases with density, and hence, that production P_i is supralinear with biomass B_i.

There is, however, an important reason for expecting donor-dependence to prevail, especially at large spatial and temporal scales where many ecological processes are aggregated: the possibility of a *bottleneck* in the dynamics. In various ecological contexts, only a fraction of possible prey (or predators) are available for predation events. This can include the existence of prey refuges in space[47], or the specialization of predators on particular age classes. The rate at which prey become available can be the limiting factor in the overall predation intensity, which therefore becomes purely donor-dependent. For example, if predators target prey of age one and above, these prey must have appeared in the previous year's census, imposing an insuperable bound $C_i < B_i$ (with a year's delay, ignored under equilibrium assumptions).

The evidence for or against this possibility is unclear here. In most of the data that we study, C_{12} is sufficiently small compared with B_1 that it could still plausibly increase with predator abundance, and there is no systematic preference for mature prey. On the other hand, our estimate of B_1 might miss other factors (such as refuges) limiting prey availability, as we discuss next. While a broader analysis across taxonomic groups, timescales and life histories would be needed to test for possible artefacts such as those suggested above, our provisional conclusion is that this empirical scaling of predation with densities is indeed meaningful, and goes beyond restating the classic donor-dependence hypothesis.

4.2 Is this functional response a consequence of biological constraints?

Nonlinear functional responses, with saturation and interference, are generally assumed to be imposed by a broad range of microscopic or mechanistic factors (physiology, behavior, spatial structure, individual trade-offs between competition and consumption, etc). In various ways, these factors limit the space or time available for encounters, leading to less predation than under mass action.

The most immediate candidate mechanism to explain a pattern that holds across ecosystems and scales is spatial structure. For instance, local prey refuges have been proposed as a classic explanation for a donor-dependent[47] or DeAngelis-Beddington response[48]. More generally, spatial structure has been shown to lead to power-law-like functional responses in agent-based or spatially explicit simulations[6;49]. But this seems to allow a wide range of possible exponents, and therefore, it does not provide an unequivocal explanation for our observations.

A functional response $C_{12} \sim \sqrt{B_1 B_2}$ ($\beta, \gamma \approx \frac{1}{2}$), similar to our measurements, could be imposed by spatial structure in two dimensions, if prey and predators occupy almost mutually exclusive regions and only meet at the boundaries (given that the periphery of a region scales like the square root of its area). This could happen due to local prey depletion by predators, a plausible explanation although it requires very strong negative correlations in space[50], and would predict a different scaling in three-dimensional settings[11] which remains to be tested.

We cannot conclude, based on our data alone, that spatial heterogeneity is the main driver of our empirical functional response. Spatially explicit data may be necessary to test this hypothesis. If spatial structure does not impose an exponent, but allows a range of possibilities, it could provide the means through which predators and prey organize to achieve a functional response that is selected by other processes[51;52].

4.3 Can this functional response emerge from selection?

The functional response and exponents found here are close to those which theoretically maximizes predator abundance and some metrics of stability. Yet this comes at an apparent cost to the individ-

ual predator, which consumes less here than under a classic Lotka-Volterra (mass action) scenario. We recall the so-called hydra effect[21;22], a counter-intuitive but common phenomenon in resource-based dynamics: decreasing the fitness of individual consumers, e.g. increasing their mortality or reducing attack rates, can lead to larger populations in the long run, when this reduces competition more than it reduces growth. We showed that a similar effect applies here to consumptive saturation and interference, as represented by the power-law exponents β and γ in the functional response. Maximal predator abundance is reached for exponents close to the relationship $\beta + \gamma = 1$, which holds approximately in data[5].

This resonates with long-standing debates in ecology and evolution, in particular the group selection controversy[53] and its accounts of selection for population abundance[54] or persistence[55]. Should we conclude here that consumption rates are selected to optimize population-level properties? And would this optimization require population-level selection, or can it arise strictly at the individual level? We now show that these are valid possibilities in our setting.

A first possibility is that our proposed functional response is outcompeted at the individual level, but selected at the population level, due to some positive consequences of having larger abundances. Extending our model to include simple adaptive dynamics of attack rate and scaling exponents, we indeed find that this functional response is dominated by other strategies (Appendix S2). We observe maladaptive evolution: mutants with ever faster consumption will outcompete and replace residents, reaching ever lower equilibrium abundances, up until the point where the predators go extinct. A classic solution is to invoke competition between groups in a spatial setting[53;54], as a larger group can send out more propagules and will often disperse faster[56]. These intuitions pervade a large literature on the evolution of "prudent" predation strategies through group selection, also known as the milker-killer dilemma[51;52;57;58].

A less commonly considered possibility is that our functional response is favored even at the individual level, as a result of direct competitive interactions. The dominance of faster consumers relies on the assumption that competition takes place only through resource consumption. In the presence of other competitive interactions that are not tied to resource levels, a larger standing population can resist invasion both by mutants and by other species. We suggest in Appendix S2 that, for the resident population to be evolutionarily stable, this non-consumptive competition must induce higher mortality in individuals that *search* for more resources, regardless of their success. This feature seems plausible for territorial behavior and aggression as widely displayed by higher vertebrates [6], but could also be induced e.g. by allelopathy in other organisms. Under these conditions, slow consumers can avoid being overtaken by fast consumers, even within a single population.

As a further perspective, all our arguments so far have taken the viewpoint of the predator's strategy and abundance, despite the fact that our empirical results suggests a roughly equal role of predator and prey density in limiting predation. Our best-fit exponents situate real systems close to the limit between a regime where prey suffer from predation, and a regime where they are mainly self-regulated. It is thus possible that prey strategies, or a balance between prey and predator selective pressures, are at play.

We conclude that selection on the functional reponse exponents to maximize population size can easily arise in variations of our model, provided that we include non-consumptive processes (from dispersal to aggression). Some of these explanations involve competition between populations, while others take place within a single population, but all favor strategies that maximize predator abundance.

[5] More precisely, predator abundance is maximized at the boundary between two dynamical regimes, which is given by $\beta + \gamma = 1$ when species self-regulation is low, but moves toward higher exponents as we increase self-regulation and can be seen in Lotka-Volterra models[9].

[6] As an illustration, lions are both outliers in our dataset, since they display lower consumption rates than the general trend (see Appendix S5), and are a major cause of mortality among other carnivores[59]. These two features together may indicate a counter-intuitive predator strategy, which involves both consuming less prey and spending more time in non-consumptive competition, leading to and benefitting from population maximization.

4.4 Can this functional response explain other power-law relations between ecosystem functions?

Relations between production and biomass, and between predator and prey densities are widely assumed and observed to take linear or power-law forms[19;25]. Yet, we have shown here and in previous work[9] that such relations do not emerge universally from arbitrary dynamical models. They require particular ecological conditions, where internal losses L_i within trophic level i are comparable to, or larger than, predation losses C_i (both in their magnitude and in their density-dependence). For instance, in a Lotka-Volterra model, strong density-dependent self-regulation within a trophic level is required to observe linear (isometric) relations between different functions and between levels, creating classic biomass and energy pyramids[9]. With weak self-regulation, exponents $\beta + \gamma = 1$ in our proposed functional response recover the same linear relations and pyramids.

A recent meta-analysis by Hatton et al.[19] found sublinear scalings of production with biomass (with exponent α) and of predator density with prey density (with exponent δ), following $\alpha \approx \delta \approx \frac{3}{4}$, reminiscent of metabolic allometry. Assuming that this cross-ecosystem law also holds dynamically within one system, so that our model results apply, this requires strong interference and saturation, $\beta + \gamma < 1$. The empirical evidence is too weak to decide whether $\beta + \gamma \approx 1$ (leading to isometry) or $\beta + \gamma < 1$ (leading to allometry), but both possibilities fall within the range of our estimates.

In summary, the existence of well-defined scaling laws across trophic levels, between ecosystem functions such as biomass and production, is less self-evident than it may appear. We suggest here that some empirically-supported laws might be secondary consequences of the mechanisms that determine our phenomenological functional response.

4.5 Conclusions

We have shown that the intensity of predation across a range of spatial scales and species can be modeled as a power-law function of both consumer and resource densities, that deviates strongly from the mass action (Lotka-Volterra) assumption. Classic functional responses cannot be automatically assumed to be adequate for macroscopic data: descriptions that are valid at small scales may not always be used at larger scales, as dynamics are driven by very different mechanisms, from individual behavior all the way to landscape heterogeneity. Similar phenomena, such as saturation and predator interference, may nevertheless emerge from these different causes. Our empirical investigation should be extended to other taxonomic groups, and confronted to experimental evidence, to pave the way for a deeper understanding of the emergent functional response at macro-ecological scales.

This phenomenological power-law model can recover classic donor control as a special case, but our empirical estimates are also compatible with a more symmetrical expression where predator and prey are roughly equally limiting, leading to an unusual square root expression. This model is reminiscent of the Cobb-Douglas production function in economics[15], which similarly arises in a macroscopic context, and is also disputed as either a salient empirical fact or a simple consequence of colinearity and aggregation[41]. Nevertheless, the main observation in both fields is that two factors – labor and capital, or consumers and resources – are co-limiting, so that doubling the output requires doubling each of the two inputs.

Scaling laws measured across ecosystems do not always hold dynamically within a single ecosystem[60]: other latent variables could differ between ecosystems, altering the relationship between predation and abundances that we would observe locally. Our results are also limited by the fact that we consider only one interaction at a time, as we generally lack data for other prey and predators interacting with our focal species. Future work should employ data on temporal change to test whether our proposed functional response truly applies to the dynamics of an ecosystem. If it does, we have shown that our empirical exponents have important implications for various features of trophic systems: in particular, they are close to maximizing predator population size and some measures of stability. This could reflect an universal selection pressure acting upon consumption rates. We notably discussed the possibility that non-consumptive competition (e.g. territorial exclusion or allelopathy) would prevent the usual dominance of smaller populations of faster consumers,

and offer an explanation for the evolution of strategies that limit predation to maximize predator abundance.

We hope that future work can address two fundamental questions: (1) whether the functional response that we observe stems purely from aggregation or from other ecological phenomena, and (2) whether some universal principle, such as selection for larger populations, could explain this functional response and, through it, the emergence of other widespread macro-ecological scaling laws.

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Author contributions

All authors contributed to the design of the study and the writing of the manuscript. MB analyzed the theoretical model. LW designed the meta-analysis and assembled the data. LW and MB jointly analyzed the data.

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