Title: The neural computation of human prosocial choices in complex motivational states

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Highlights

- Activating different social motives simultaneously can enhance prosocial choices
- Multi-motive combinations change initial prosocial biases
- Dorso-striatal activation increases with larger increase of prosocial bias
- Multi-motive combinations modulate relative response caution

Abstract

Motives motivate human behavior. Most behaviors are driven by more than one motive, yet it is unclear how different motives interact and how such motive combinations affect the neural computation of the behaviors they drive. To answer this question, we induced two prosocial motives simultaneously (multi-motive condition) and separately (single motive conditions). After the different motive inductions, participants performed the same choice task in which they allocated points in favor of the other person (prosocial choice) or in favor of themselves (egoistic choice). We used fMRI to assess prosocial choice-related brain responses and drift diffusion modelling to specify how motive interactions affect individual components of the choice process. Our results showed that the combination of the two motives in the multi-motive condition increased participants' choice biases prior to the behavior itself. On the neural level, these changes in initial prosocial bias were associated with neural responses in the bilateral dorsal striatum. In contrast, the efficiency of the prosocial decision process was comparable between the multi-motive and the single-motive conditions. These findings provide insights into the computation of prosocial choices in complex motivational states, the motivational setting that drives most human behaviors.

Keywords:

motivation, social decision-making, hierarchical drift-diffusion modeling, fMRI, social neuroscience

1. Introduction

All choice behaviors are incited by motives, which can be complex. Documenting this motivational complexity, many animal (Jennings et al., 2013; Kennedy and Shapiro, 2009) and most human behaviors are driven by multiple motives that are active at the same time, and affect each other (Engel and Zhurakhovska, 2016; Hughes and Zaki, 2015; Jagers et al., 2017; Kruglanski et al., 2018; Lewin et al., 1951; Takeuchi et al., 2015; Terlecki and Buckner, 2015). For example, the decision to help an elderly relative is often driven by empathy with her needs, and at the same time, by the wish to reciprocate help received by this person in the past, i.e, the social norm of reciprocity. Consequently, choice behaviors most commonly originate from the interplay between different motives and cannot be explained by one "motivational force" alone. However, motive-interplay is not directly observable. Thus, to understand and predict choice behaviors, it is crucial to elucidate the neuro-computational mechanisms through which multiple interacting motives affect behavioral choice processes.

The processing of single-motive states and its impact on behavioral choices in animals (e.g., place preferences) (Jennings et al., 2013) have been linked to dopaminergic neurons in the striatum (Kim and Im, 2018; Robinson et al., 2006; Salamone and Correa, 2012). In line with these results, human neuroscience studies have shown that the striatum is involved in the processing of different individual motives, as well as motivated choice behaviors, both in the social (Báez-Mendoza and Schultz, 2013; Bhanji and Delgado, 2014) and non-social domain (Salamone et al., 2016; Shohamy, 2011). In more detail, the ventral striatum has been linked to the learning and encoding of values and the predictions of future rewards (Kable and Glimcher, 2007; Liljeholm and O 'Doherty, 2012; O'Doherty et al., 2004; Strait et al., 2015), whereas the dorsal striatum has been linked to initiating and optimizing

choices based on these encoded values (Balleine et al., 2007; Liljeholm and O

'Doherty, 2012; O'Doherty et al., 2004; Palmiter, 2008; Robinson et al.,

2006).Together, this previous work has provided insights into the neural

underpinnings of individual motivational processes. However, the neural computation

of behaviors that are driven by different motives remains unclear.

To address this issue, we developed a paradigm in which participants made the

same choices (prosocial vs. egoistic) based on different, simultaneously activated

motives (multi-motive condition), or based on each of these motives separately

(single-motive conditions). Specifically, we studied the effect of simultaneously

activated social motives in a social choice paradigm in which participants repeatedly

had the choice between a prosocial and an egoistic option (**Fig. 1A**). Inspired by an

influential model of prosocial motivations (Batson et al., 2011), we induced two key

motives that both incite prosocial behavior - the empathy motive, defined as the

affective response to another person's misfortune (Batson et al., 1995; Hein et al.,

2016a; Lamm et al., 2011), and the reciprocity motive, defined as the desire to

reciprocate perceived kindness with a kind behavior (Gouldner, 1960; Hein et al.,

2016a; McCabe et al., 2003).

In combination with fMRI and hierarchical drift-diffusion modeling (hierarchical

DDM) (Forstmann et al., 2016; Ratcliff et al., 2016; Vandekerckhove et al., 2011;

Wiecki et al., 2013), this paradigm allowed us to specify how the combination of

different motives affects individual components of neural goal-directed (i.e., prosocial)

choice computation, compared to computation of the same choice in a simple

motivational state (i.e., driven by only one of the two motives).

Drift-diffusion models (DDMs) characterize how noisy information is accumulated

to select a choice option (**Fig. 1B**) based on three different parameters (the v, z and a

parameters) (Forstmann et al., 2016; Ratcliff et al., 2016). The v-parameter describes

the speed at which information is accumulated in order to choose one of the options, i.e., the efficiency of the choice process itself. The z-parameter reflects the initial choice bias, i.e., the degree to which an individual prefers one of the choice options prior to making the choice. Thus, in contrast to the v-parameter, which models the choice process itself, the z-parameter models the individual bias with which a person enters the choice process. For example, if a person has a strong initial bias towards prosocial choices (reflected by a large value of the parameter z), the starting point of the choice computation is located closer to the prosocial choice boundary, and thus, this person is more likely to choose the prosocial option. The third component, parameter a, quantifies the amount of relative evidence that is required to choose one of the options.

Previous neuroscience studies have identified brain regions that are associated with changes of these choice parameters. For example, it has been shown that reward-related improvement of perceptual discrimination is driven by changes in the z-parameter, related to changes of frontoparietal activation (Mulder et al., 2012). Another study using a task from the perceptual domain has shown that increased evidence accumulation under time pressure is linked to increased activation in premotor regions (preSMA) and the dorsal striatum (Forstmann et al., 2008). Other studies have used similar modelling approaches to investigate value-based decisions (Gluth et al., 2012; Hare et al., 2011), for example using a buying task in which participants could decide to accept or reject a stock after receiving probabilistic information about the stock from different rating companies (Gluth et al., 2012). As a main result, Gluth and colleagues showed that the amount of relative evidence that participants required for making a choice was related to the neural response in the anterior insula (AI) and the dorsal striatum. The finding in dorsal striatum resembled evidence from DDM studies obtained with perception paradigms (Forstmann et al.,

2008) and indicates that·the striatum is a plausible neural candidate for tracking

changes in choice components in different motivational settings (e.g., induced by

time pressure, Forstmann et al., 2008, or by others' information, Gluth et al., 2012).

In our study, we modeled the three relevant choice parameters (v, z, and a) for

choices that were driven by the combination of the two motives and for the same

choices that were driven by each of the motives separately. It is important to note that

the choice process may also be influenced by other motives than empathy and

reciprocity, i.e., the motives that were experimentally induced. That said, our

paradigm can provide insights into the multi-motive choice process even if other

motives are potentially activated, because multi-motive choices are contrasted with

the same choices that are driven by the respective single motives.

According to one hypothesis, the simultaneous activation of multiple motives may

facilitate the computation of the choice option that is favored by the motives. In the

present paradigm this means that computation of the prosocial choice option should

be facilitated since empathy and reciprocity both drive prosocial behavior. In this

case, we should observe an increase in prosocial behavior in the multi-motive

condition (empathy and reciprocity motive active) compared to the single-motive

conditions (only empathy or only reciprocity active) that cannot be explained by the

difference between the single-motive conditions. Specifying the mechanism

underlying such a facilitation, the DDM proposes that a facilitation of prosocial

choices in the multi-motive condition may originate A) from an increased speed of

information accumulation (v-parameter; **Fig. 1C**, left panel (Flagan et al., 2017;

Janczyk and Lerche, 2019; Krajbich et al., 2015)), B) from an enhancement of

participants' initial bias to choose the prosocial option (z-parameter; **Fig. 1C**, middle

panel; (Chen and Krajbich, 2018; Mulder et al., 2012; Toelch et al., 2018), or C) from

an enhancement of the v- as well as the z-parameter in the multi-motive condition,

compared to the single-motive condition.

Alternatively, we may observe fewer prosocial decisions in the multi-motive

condition compared to the single motive conditions, reflected by a decreased speed

of information accumulation (lower DDM v-parameter) and/or decreased initial bias to

choose the prosocial option (lower DDM z-parameter). Moreover, in the multi-motive

condition, participants are required to process two motives simultaneously, in addition

to the trial-by-trial choice option information (which was constant across all conditions

because participants performed the identical choice task). This additional motive-

related information may cause participants to make more careful responses in the

multi-motive condition and thus increase the a-parameter in the multi-motive

condition compared to the single-motive conditions (**Fig.1C**, right panel).

Regarding the underlying neural mechanisms, we hypothesized that changes in

DDM choice parameters in the multi-motive compared to the single-motive conditions

might be related to changes of activation in the ventral striatum, i.e., a region that is

involved in the integration of different choice values (here the value of empathy-

based and the value of reciprocity-based choices; (Kable and Glimcher, 2007;

Liljeholm and O 'Doherty, 2012; O'Doherty et al., 2004; Strait et al., 2015), and/ or

activation of the dorsal striatum, i.e., a region that is related to integration of choice

preferences that derive from these different choice values (Balleine et al., 2007;

Liljeholm and O 'Doherty, 2012; O'Doherty et al., 2004; Palmiter, 2008; Robinson et

al., 2006).

Figure 1. Example of point allocation during the choice task, schematic illustration of the drift-diffusion model and hypotheses regarding the impact of different drift-diffusion parameters on the choice process in multi-motive and single-motive conditions. **(A)** Participants chose between a prosocial and an egoistic option to allocate points to themselves (in this example shown in green) and a partner (in this example shown in red). Colors were counter-balanced across participants. In this example trial, the participant chose the prosocial option, which maximized the outcome of the partner at a cost to the participant (green box). **(B)** The drift-diffusion model conceptualizes the choice process as noisy accumulation of information (squiggly blue line). The v-parameter describes the speed at which information is accumulated in order to choose one of the options, i.e., the efficiency of the choice process itself. The z-parameter reflects the initial choice bias, i.e., the degree to which an individual prefers one of the choice options prior to making the choice. The third component, parameter a, quantifies the amount of relative evidence that is required to choose one of the options. Once the accumulated information reaches either boundary, the choice is made (upper boundary = prosocial choice; lower boundary = egoistic choice). **(C)** An enhancement of prosocial choice frequency in the multi-motive condition (red) compared to the single motive conditions (i.e., the empathy or the reciprocity condition; blue) may result from an increased speed of information accumulation (v-parameter; left panel), and/or an increased initial bias toward making a prosocial choice (z-parameter; middle panel). On average, the amount of required relative evidence (a-parameter) may be higher in the multi-motive condition compared to the single motive conditions (right panel).

2. Material and methods

2.1. Participant details

Forty-two right-handed healthy female participants (mean age = 23.1 years, SD = 2.8 years) and four female confederates took part in the experiment. We chose female participants as well as female confederates in order to control for gender and avoid cross-gender effects. The confederates were students who had been trained to serve in all the different conditions counterbalanced across participants. Prior to the experiment, written informed consent was obtained from all the participants. The study was approved by the local ethics committee (BB 023/17). Participants received monetary compensation. Three participants had to be excluded due to technical problems and dropout. Another subject had to be excluded due to excessive head movements (more than 5% of the scans contained rapid head motion with more than 0.5 mm displacement per TR). Five participants had to be excluded as outlier based on their choices (less than ten prosocial choices across all condition; three standard deviations above the mean in central measures). Thus, we analyzed 33 data sets using a within-subjects design. We aimed for 34 data sets, which corresponds to the median sample size of neuroimaging studies determined in a recent review (N = 33; Yeung, 2018). We tested 40 participants to meet this target, accounting for a drop-out rate of about 15% which, based on our experience, is common in fMRI studies. Our final analyses includes 33 data sets, in accordance with the median sample reported by Yeung (2018). Given that it is difficult to collect large data sets with expensive and time-consuming methods like fMRI, the importance of stringent statistical thresholds is highlighted (Carter et al., 2016; Roiser et al., 2016; Woo et al., 2014; Yeung, 2018). To analyse the results of the second level regression we thus used cluster-level family wise error correction at the whole brain level after applying a threshold of

224 p < .001 on an uncorrected level. Neural activations that are thresholded at this level

225 are seen as valid and reliable (Eklund et al., 2016; Woo et al., 2014; Yeung, 2018).

2.2.   Procedure

Prior to the motive induction and choice task, the individual thresholds for pain

stimulation were determined for the participants and all the confederates (see section

2.5. Pain stimulation for details). Next, the participants and confederates were

assigned their different roles by a manipulated lottery (drawing matches). In order to

ensure that each participant was always assigned her designated role as a

participant (pain recipient during motive induction; decider during the decision task),

the drawing of the matches was organized in such a way that she always drew the

last match. The confederates were assigned the roles of the empathy partner, the

reciprocity partner, the multi-motive partner or the baseline partner, and these roles

were counterbalanced across participants. In accordance with these roles, two of the

confederates first went to an ostensible other experiment and the other two waited to

be seated in the scanner room. Each confederate was matched with a specific color

and seating position (to the left vs. to the right of the fMRI scanner), and their color

designation and seating positions were counter-balanced across participants.

Next, the first two confederates (the empathy partner, reciprocity partner, multi-

motive partner, or baseline partner) were seated to the left and the right of the

participant who was lying inside the fMRI scanner and the first motive induction took

place (for overview of an example procedure, see **Fig. 2**). After the motive induction,

image acquisition for the choice task was started, during which the participant

allocated points to her respective partners. This way the participant only had to

remember interactions with two partners at any one time. After the choice task, the

first confederates were replaced by the other two confederates and the second part

commenced. Part 2 had the same structure as part 1: first, the participant underwent

motive induction 2 followed again by the choice task. The order of motive inductions

and the type of partner the confederates represented were counterbalanced across

participants.

At the end of the experiment, all the confederates left and the participant remained in

the scanner until anatomical image acquisition was completed. Finally, participants

were asked to complete a questionnaire measuring trait aspects of empathy (Jordan

et al., 2016) and reciprocity (Perugini et al., 2003). Participants spent approximately

60 min in the scanner and the entire procedure lasted approximately 2.5 hours. To

avoid possible reputation effects, which could influence participants' behavior,

participants were informed that they would not meet the confederates after the

experiment.

Figure 2. Overview of an example experimental procedure. The study consisted of
two parts. In this example, in part 1, the empathy motive was activated towards one
confederate (the empathy partner) and the reciprocity motive was activated towards
the other confederate (the reciprocity partner). In the following choice task,
participants allocated points to the empathy partner (i.e., driven by the empathy
motive) or the reciprocity partner (i.e., driven by the reciprocity motive). Next, the
confederates were replaced by two new individuals that served as partners for part 2.
In part 2, the empathy and the reciprocity motive were activated simultaneously
towards one confederate (multi-motive partner) and no motive was induced towards
the other confederate (baseline partner). In the following choice task, participants

allocated points towards the multi-motive partner (i.e., driven by two motive simultaneously) and towards the baseline partner (i.e., independently of any motive induction). The order of motive induction (empathy, reciprocity, multi-motive, baseline) was counterbalanced across participants and the four confederates. The respective partner was indicated by a cue in one of four counterbalanced colors.

2.3.   Motive inductions

2.3.1. Empathy induction

During the study, participants were paired with four partners (confederates of the experimenter). Participants saw the hand of the respective partners with the attached pain electrode. In the empathy condition, the participants repeatedly observed one of the confederates (the empathy partner) receiving painful shocks in a number of trials, a situation known to elicit an empathic response (Batson et al., 1995; Hein et al., 2016a; Lamm et al., 2011). Each empathy-induction trial started with a colored arrow shown for 1,000 ms, which indicated the empathy partner. After this cue and a jittered (1,000–2,000 ms) fixation cross, the same colored flash was displayed for 1,500 ms. Participants were informed that a dark-colored flash indicated that the corresponding partner received a painful stimulus at that moment; a light-colored flash indicated a non-painful stimulus. During (ostensible) stimulation of the respective partner, participants either saw a dark colored flash (painful stimulation) or a light colored flash (non-painful stimulation). Since all partners were confederates of the experimenter, they did not actually receive painful stimulations. Thus, the trials in which participants saw the dark colored flash were "ostensibly painful" for the partner. To assess the success of the empathy induction, a rating scale was shown for a maximum of 6 s. Participants reported how they felt after observing the partner receive painful or non-painful stimuli ("How do you feel?" in German). The scale ranged from -4 (labeled "very bad") to +4 (labeled "very good") with intervals of one and was visually displayed. Before analysis, the induction ratings were recoded such

that high positive values reflect strong responses to the induction procedure (strong

empathy motive). Participants had to respond within 6 s. The inter-trial interval was

1,500 ms. Empathy induction consisted of 12 trials: nine of which were ostensibly

painful for their partner (i.e., the confederate).

2.3.2. Reciprocity induction

The reciprocity motive is defined as the desire to reciprocate perceived kindness with

kind behavior (Gouldner, 1960; Hein et al., 2016a; McCabe et al., 2003). Therefore,

in the reciprocity condition, we induced the reciprocity motive by instructing one of the

confederates (the reciprocity partner) to give up money in several trials to save the

participant from painful shocks (Hein et al., 2016a). Each reciprocity-induction trial

also started with an arrow colored in the reciprocity partner's color, which pointed

toward the seating position of the reciprocity partner (left or right) and was shown for

1,000 ms. Next, the participants were shown a flash displayed to the right and a

crossed-out flash displayed to the left of a centered fixation cross. Participants were

told that this was the decision screen, which the reciprocity partner also saw while

making her decision to either save or not save the participant from painful stimulation.

After a jittered interval of 2,000 to 4,000 ms, a box appeared around one of the

flashes, indicating the ostensible choice of the reciprocity partner. Depending on

where the box was displayed, the reciprocity partner had either decided to forego a

monetary award of 2 € in order to save the participant from painful stimulation (a box

around the crossed-out flash) or decided to take the money and not save the

participant (a box around the flash that was not crossed-out). After 1000ms

participants rated how they felt about the decision of the partner ("How do you feel?"

in German). The ratings were recoded such that high positive values reflect a strong

positive response to the decision of the partner, indicating a strong reciprocity motive.

After a jittered (1,000 to 2,000 ms) fixation cross, the participant saw an information

on the screen, indicating whether the decision of the reciprocity partner would be

implemented ("decision accepted") or not ("decision declined"), displayed for 1000

ms. This additional stage was included in order to ensure the same amount of painful

stimulations were administered across all conditions (50 %), while at the same time

allowing for the high rate (75 %; 9 out of 12 trials) of the reciprocity partner's

decisions to help. While instructing the participants, it was highlighted that the choice

of the reciprocity partner reflected her willingness (or unwillingness) to help, while a

computer algorithm decided about the implementation of the decision.

Thus, four types of reciprocity trials were possible. When the partner decided to save

the participant from painful stimulation and this decision was accepted, the participant

did not receive a painful stimulus, which was visually represented by a crossed-out

flash (1,500 ms). However, when the reciprocity partner's decision to save the

participant was declined, participants received a painful stimulus, which was

accompanied by the display of a flash (1,500 ms). Similarly, when the partner

decided not to save the participant and this decision was accepted, the participant

received a painful stimulus accompanied by the display of a flash. Finally, when the

partner decided to not save the participant and this decision was declined, the

participant did not receive painful stimulation, which was visually represented by a

crossed-out flash. The inter trial fixation cross was displayed for 1,500 ms before the

next trial started.

2.3.3. Multi-motive induction

In the multi-motive condition, the participants repeatedly observed how one of the

confederates (the multi-motive partner) received painful shocks and also gave up

money to spare the participant from painful shocks. The multi-motive induction

procedure combined the empathy- and reciprocity-induction procedures. As in the

empathy-induction condition, it included 12 empathy induction trials, nine of which

were ostensibly painful for the partner. As in the reciprocity-induction condition, it

included 12 reciprocity trials, of which participants received help in nine out of 12

trials. The stimulation and trial structure were identical to the empathy- and

reciprocity-induction trials described above, except that the relevant colors were

replaced by the colors matched to the multi-motive partner (i.e., the color of the pain

flash in the empathy trials and the color of the box highlighting the decision of the

partner in the reciprocity trials).

2.3.4. Additional control trials for empathy and reciprocity induction

In order to equalize the number and types of trials (i.e., the length and structure of the

interaction with each motive partner) across conditions, the empathy-induction

procedure also included trials that were identical to the reciprocity trials, except that

the computer decided whether the participant would be saved from a painful stimulus

and not the empathy partner. This computer's decision was visually represented by a

white-colored box appearing either around the crossed-out flash (saving the

participant) or the normal flash (not saving the participant). It was clearly explained to

each participant that the color white was not matched with any of the partners but

indicated the computer's choice. The empathy-induction procedure consisted of 12

control trials, in addition to the 12 empathy trials described above, resulting in 24

trials, i.e., the identical number of trials as the multi-motive induction procedure.

Similarly, the reciprocity-induction procedure included trials that were identical to the

empathy-induction trials, except that the reciprocity partner only received non-painful

stimulation on these trials, as visually represented by a light-colored flash. In total,

the reciprocity-induction procedure consisted of 12 of these control trials and 12

reciprocity trials (see above), i.e., 24 trials (identical to the other conditions).

2.3.5. Baseline induction

The baseline procedure consisted of 24 trials in total, 12 trials in which the baseline

partner only received non-painful stimulation and 12 trials in which the computer

decided whether the participant would be saved from a painful stimulus or not. This

computer's decision was visually represented by a white-colored box either

appearing around the crossed-out flash (saving the participant) or the normal flash

(not saving the participant). It was clearly explained to the participant that the white

box did not represent the decision of a person but indicated the computer's choice.

2.4.   Choice task

After the motive inductions, participants performed a social choice task inside the

fMRI scanner. The choice task was a two-alternative-forced-choice adaptation of the

commonly used Dictator Game (Forsythe et al., 1994), which has been successfully

used in previous studies (e.g., Chen and Krajbich, 2018; Hein et al., 2016b; Krajbich

et al., 2015). In each trial of this choice task, participants allocated money to

themselves and one of the partners (**Fig. 1A**) and could choose between maximizing

the relative outcome of the other person by reducing their own relative outcome

(prosocial choice) and maximizing their own relative outcome at a cost to the partner

(egoistic choice). The outcome was relative to the outcome that the participant would

have gained when choosing the other option. The initial number of points was always

higher for the participant compared to the partners. This measure was inspired by

previous behavioral economics research, showing that prosocial behaviors depend

on the initial payoff allocation between the participant and the participant's partner

(Bolton and Ockenfels, 2000; Charness and Rabin, 2002; Fehr and Schmidt, 1999).

In particular, if subjects have a lower initial payoff than their partner

("disadvantageous initial inequality"), they are much less willing to behave

altruistically toward the partner compared to a situation with advantageous initial

inequality (i.e., when the participant has a higher initial payoff than the partner). The

choice options used in the present study created advantageous inequality to optimize

the number of prosocial choices, which was the main focus of our study. The exact

point distributions are provided in **Table A1**.

Depending on the type of partner the participants faced in the choice task, there were

four conditions – the empathy condition, the reciprocity condition, the multi-motive

condition, and the baseline condition. Importantly, the choice task was identical in all

the conditions.

In more detail, participants were asked repeatedly to choose between two

different distributions of points that each represented different amounts of monetary

pay-offs for themselves and one of the partners (see **Fig. 1A**). Each choice-trial

started with a colored arrow shown for 1,000 ms, indicating the next interaction

partner. After this cue, participants saw the two possible distributions of points in

different colors, indicating the potential gain for the participant or for the current

partner. Colors were counterbalanced across participants. Participants had to choose

one of the distributions within 4,000 ms. A green box appeared around the

distribution that was selected by the participant at 4,000 ms after distribution onset.

The box was shown for 1,000 ms. The length of the inter-trial interval, as indicated by

a fixation cross, was jittered between 4,000 and 6,000 ms. At the end of the

experiment, two of the distributions chosen by the participant were randomly selected

for payment (100 points = 50 cents). We analyzed 38 choice trials in each motive-

induction condition, i.e., 152 trials in total. In addition to the 38 trials, each condition

contained four trials in which the same choice option maximized the outcome of the

participant and the partner (non-competitive trials). These trials were included to

increase the variability of choices and thus to keep the participants engaged. They

were excluded from the analyses because they could not be classified as prosocial or

egoistic choice trial. For each condition and participant, the same distributions were

used and presented in random order.

2.5.   Pain stimulation

For pain stimulation, we used a mechano-tactile stimulus generated by a small

plastic cylinder (513 g). The projectile was shot against the cuticle of the left index

finger using air pressure (Impact Stimulator, Labortechnik Franken, Release

1.0.0.34). The criterion for painful stimulation was a subjective value of 8 on a pain

scale ranging from 1 (no pain at all, but a participant could feel a slight touch of the

projectile) to 10 (extreme, hardly bearable pain). The participants were told that a

value of 8 corresponded to a painful, but bearable stimulus, and a non-painful

stimulus corresponded to a value of 1 on the same subjective pain scale. These

subjective pain thresholds were determined using a stepwise increase of air pressure

(stepsize of 0.25 mg/s), starting with the lowest possible pressure (0.25 mg/s), which

caused the projectile to barely touch the cuticle, and increasing in stimulus intensity

until it reached a level that corresponded to the individual's value of 8 (range = 2.75–

3.5 mg/s).

2.6.   Experimental design and statistical analyses

2.6.1. Regression analyses

Regression analyses were conducted using the R-packages "lme4 and "car"

(R Core-Team, 2018). For mixed models, we report the chi-square values derived

from Wald chisquare tests using the "Anova" (car package) function. For predefined

contrasts we report the t-values derived from the summary() function. When more

than one predictor was included in the model, the function emmeans was used in

order to compute contrasts between factor levels.

To test the differences in induction ratings and the relationship between

induction ratings and frequencies of prosocial choice, the mean induction ratings and

frequencies of prosocial choices were calculated for each participant for each

condition (empathy, reciprocity, multi-motive, and baseline) and entered as a

dependent variable into mixed models with conditions (empathy, reciprocity, multi-

motive, and baseline) and induction ratings as fixed effects and participants as

random effects.

To test whether prosocial behavior was influenced by trial-by-trial point

information, condition and their interaction, a logistic mixed model regression was

conducted with the possible gain for the partner (i.e. difference in points between the

two options for the partner, |partner's gain option 1 – partner's gain option 2|), the

possible loss for the participant (i.e., the difference in points between the two options

for the participant, |participant's gain option 1 – participant's gain option 2|) and

condition as predictor variables. The binary choice outcome (prosocial vs. egoistic

choice) was used as dependent variable. To investigate the differences in prosocial

behavior between the social motives (multi-motive > reciprocity and multi-motive >

empathy) more closely, contrasts were calculated using the emmeans function.

To specifically test whether prosocial behavior was differentially influenced by

inequity aversion, a logistic mixed model regression was conducted with the predictor

variables condition and the difference in point equality of the participant's and the

partner's outcome between the two choice options. To compute this variable, we first

calculated calculating the difference between the gains for each option (i.e., |partner's

gain option 1 – participant's gain option 1| for each choice option). Second, these

differences were subtracted from each other in order to obtain a measure of point

equality for each choice trial. Again, the binary choice outcome (prosocial vs. egoistic

choice) was used as dependent variable.

To test whether the frequency of prosocial choices and reaction times were

equally distributed across conditions, we conducted pairwise Kolmogorov-Smirnov

tests.

Additionally, we investigated whether the relationship between the possible

gain for the partner and participants' probability to make a prosocial choice can be

described in terms of a psychometric function. For the estimation of the psychometric

functions we used the R-package "quickpsy" which implements a Maximum-

Likelihood-Estimation procedure to fit the cumulative normal distribution. To test

whether the points of subjective equality (PSEs) differed between conditions, we

conducted a linear mixed model with the condition as fixed effect, participant as

random effect and PSE as dependent variable.

To test whether the relative difference between empathy and reciprocity in the

z-parameter and a-parameter could explain the percent changes of these parameters

in the multi-motive condition compared to the reciprocity condition, the percent

change values ($\Delta z_{\text{multi-motive/reciprocity}}$ and $\Delta a_{\text{multi-motive/reciprocity}}$) were entered as

dependent variables in a linear regression model. The respective relative differences

($\Delta z_{\text{empathy/reciprocity}}$ and $\Delta a_{\text{empathy/reciprocity}}$) and one regressor modeling the parameter

type (z-parameter, a-parameter) were included as predictors.

2.6.2. Drift-diffusion modeling

We used hierarchical drift-diffusion modeling (HDDM) (Vandekerckhove et al.,

2011; Wiecki et al., 2013), which is a version of the classical drift-diffusion model that

exploits between-subject and within-subject variability using Bayesian parameter

estimation methods, because it is ideal for use with relatively small sample sizes. The

analyses were conducted using the python implementation of HDDM version 0.8.0

(Wiecki et al., 2013).

Based on binary choices, the HDDM approach provides detailed insights into

the computation of egoistic and prosocial choices, because it uses all the raw data

that is available (trial-by-trial reaction times and choice outcome information of all

choices, irrespective of point distributions) to estimate sub-components of the

underlying decision process. The v, z and a-parameters for each participant capture

how each person maneuvers between the egoistic and the prosocial choice options,

and finally approaches a decision boundary (i.e., the boundary for an egoistic or a

prosocial choice). In line with previous studies that have used a similar procedure in

the realm of social decision making (e.g., Chen and Krajbich, 2018; Gallotti and

Grujić, 2019), we believe that the HDDM results provide a sensitive and fine-grained

proxy for individual differences in prosociality.

Since we did not have prior hypotheses about which and how many of the

three central DDM parameters may reflect motive complexity, we estimated 11

possible variants of the DDM model, ranging from the most simple model (no

parameter is modulated by condition) to the full model with v, z, and a possibly being

modulated by our four conditions (baseline, empathy, reciprocity, and multi-motive).

Since the point information varied between trials, which may influence drift-rate, we

allowed the drift rate to vary by the trial-by-trial possible gain for the partner (see

section 2.6.1. for computation of this value). We performed model comparison based

on the deviance information criterion (DIC) and extracted the parameters of the

winning model (lowest DIC value). Apart from the three parameters of interest for our

research question (v, z, a), additional parameters are included in the estimation

procedure. We also estimated the non-decision time t and allowed for trial-by-trial

variations of the initial bias (sz), he drift rate (sv) and the non-decision time (st).

These parameters were not estimated to vary by condition. They were nonetheless

included based on the results by Lerche and Voss (2016), who showed that in most

cases, it is beneficial to include these parameters in order to improve model fit. In the

estimation procedures we used the default values for the priors and hyperpriors

provided by the HDDM package. In more detail, the "informative group mean priors

are created to roughly match parameter values reported in the literature and collected

by Matzke and Wagenmakers (2009)." cited from (Wiecki et al., 2013, Supplementary

Material, page 1). Model convergence was checked by visual inspection of the

estimation chain of the posteriors, as well as computing the Gelman-Rubin Geweke

statistic for convergence (all values < 1.01) (Gelman and Rubin, 1992). To assess

model fit, we conducted posterior predictive checks by comparing the observed data

with 500 datasets simulated by our model (Wiecki et al., 2013). This approach allows

for the computation of intervals within which the parameter falls with 95 % probability.

If the observed data falls within the 95 % credibility interval of the simulated data, the

model can describe the data well. The present results revealed a good match

between the observed data and the modeled data. Parameters of interest from the

winning model were extracted for further analysis. Specifically, for each participant,

the condition-specific v-parameters, z-parameters, and a-parameters were extracted

(resulting in 12 parameters per participant). In HDDM, the z-parameter is always

relative to a. The reported values of z thus range between 0 and 1 and correspond to

the absolute value of z divided by the a-parameter (z/a)

For closer investigation of processing differences in complex vs. more simple

motivational states, we compared the posterior distributions of the conditions for each

parameter by computing the probabilities for the multi-motive parameter being larger

than the single motive parameters. This was done by calculating the densities of the

differences distributions that are larger than zero (Wiecki et al., 2013). Additionally,

we used the plausible value approach to estimate the corresponding t-value. This

approach consists of repeatedly sampling participants' individual parameters from the

winning model's posterior distribution. Extracting these parameters and comparing

between the different conditions using frequentist statistics results in distributions of t-

values whose means are a plausible proxy for the actual underlying t-value (Ly et al.,

2017; Marsman et al., 2016).

2.6.3. *fMRI data acquisition*

Imaging data was collected at a 3T MRI-scanner (Verio, Siemens, Erlangen,

Germany) with a 32-channel head coil. Functional imaging was performed with a

multiband EPI sequence of 72 transversal slices oriented along the subjects' AC-PC

plane (multi-band acceleration factor of 6). The in plane resolution was 2.5 x 2.5 mm²

and the slice thickness was 2.5 mm. The field of view was 210 x 210 mm²,

corresponding to an acquisition matrix of 84 x 84. The repetition time was 1 s, the

echo time was 33.6 ms, and the flip angle was 54°. Structural imaging was

conducted using a sagittal T1-weighted 3D MPRAGE with 176 slices, and a spatial

resolution of 1 x 1 x 1 mm³. The field of view was 250 x 250 mm², corresponding to

an acquisition matrix of 256 x 256. The repetition time was 1,690 ms, the echo time

was 2.52 ms, the total acquisition time was 3:50 min, and the flip angle was 9°. For

the T1-weighted images, GRAPPA with a PAT factor of 2 was used. We obtained, on

average, 1,911 (SD = 5.6 volumes) EPI-volumes during the choice task of each

participant. We used a rubber foam head restraint to avoid head movements.

2.6.4. *fMRI Preprocessing*

Preprocessing and statistical parametric mapping were performed with SPM12

(Wellcome Department of Neuroscience, London, UK) and Matlab version 9.2

(MathWorks Inc; Natick, MA). Spatial preprocessing included realignment to the first

scan, and unwarping and coregistration to the T1 anatomical volume images.

Unwarping of geometrically distorted EPIs was performed using the FieldMap

Toolbox. T1-weighted images were segmented to localize grey and white matter, and

cerebro-spinal fluid. This segmentation was the basis for the creation of a DARTEL

Template and spatial normalization to Montreal Neurological Institute (MNI) space,

including smoothing with a 6 mm (full width at half maximum) Gaussian Kernel filter

to improve the signal-to-noise-ratio. To correct for low-frequency components, a high-

pass filter with a cut-off of 128 s was used.

2.6.5. *fMRI statistical analysis*

Our participants made prosocial choices in the majority of the trials (Mean = 74

%, SD = 19 %) with more than half of the participants making prosocial choices in 80

% or more of the trials in at least one of the four conditions (see **Table A2**). Given the

lack of egoistic choices and given that our study focused on the computation of

prosocial choices, egoistic choices trials were not included in the imaging analyses

and we also refrained from computing direct contrasts between prosocial and egoistic

choices.

First-level analyses were performed with the general linear model (GLM),

using a canonical hemodynamic response function (HRF) and its first derivative (time

derivative). Regressors were defined from cue onset until the individual response

was made by pressing a button (resulting in a time window of 1,000 ms + individual

response time). For each of the four conditions (the three motive conditions and

baseline condition), the respective regressors of prosocial choice trials were included

as regressors of interest. The respective regressors of all the other trials (e.g.,

egoistic choice trials and trials with missed button presses) were included as

regressors of no interest. Given that more than half of our participants (64 %) made

fewer than five egoistic decisions in at least one of the conditions, we refrained from

computing direct contrasts between prosocial and egoistic choices and included

egoistic choices in this regressor of no interest (see **Table A2** for the number of trials

per participant and condition). The residual effects of head motions were corrected

for by including the six estimated motion parameters for each participant and each

session as regressors of no interest. To allow for modeling all the conditions in one

GLM, an additional regressor of no interest was included, which modeled the

potential effects of session.

For the second-level analyses, contrast images for comparisons of interest (empathy

> reciprocity, multi-motive > empathy, reciprocity > empathy, and multi-motive >

reciprocity) were initially computed on a single-subject level. In the next step, the

individual images of the main contrast of interest (multi-motive > reciprocity) were

regressed against the percent change in the z-parameter ($\Delta z_{\text{multi-motive/reciprocity}}$) and a-

parameter ($\Delta a_{\text{multi-motive/reciprocity}}$) in the multi-motive condition, relative to the reciprocity

condition, using second-level regressions. Second-level results were corrected for

multiple comparisons, using cluster-level family wise error (FWE) correction on a

whole brain level. We also report results at a threshold of $P_{\text{uncorrected}} < 0.001$ and a

cluster threshold of $k > 10$ in Appendix A: Supplementary Material.

To test if the neural response in the dorsal striatum was related to the relative

difference in z between empathy and reciprocity ($\Delta z_{\text{empathy/reciprocity}}$), the (multi-motive

> reciprocity) contrast was regressed against the empathy vs reciprocity z-differences

($\Delta z_{\text{empathy/reciprocity}}$) and the multi-motive z-enhancement ($\Delta z_{\text{multi-motive/reciprocity}}$) in the

same model. Additionally, the individual beta-estimates of the neural multi-motive

condition > reciprocity and empathy > reciprocity contrasts were extracted from an

independent anatomical ROI of bilateral putamen based on the aal nomenclature

(Tzourio-Mazoyer et al., 2002), using MarsBaR (Brett et al., 2002) and the WFU

PickAtlas Tool (Maldjian et al., 2003).

In order to clarify the commonly shared influence of the partner's possible gain

on the neural prosocial choice process, we added the partner's gain as trial-by-trial

parametric modulator of the decision phase to a first level GLM in which all conditions

are collapsed into one single regressor. On the second level, we conducted a one

sample t-test on this parametric modulator corrected for multiple comparisons, using

cluster-level family wise error (FWE) correction on a whole brain level.

The reported anatomical regions were identified using the SPM anatomy

toolbox (Eickhoff et al., 2005).

2.7.   Data and code availability

Behavioral data and scripts are available at github.com

(https://github.com/AnneSaulin/complex_motivations).

Imaging data are available at neurovault.org

(https://www.neurovault.org/collections/5879/).

3. Results

3.1.   Motive ratings

During the empathy induction, participants indicated how they felt after

observing the person in pain. During the reciprocity induction, they indicated how

they felt after receiving a favor from the other person. In the multi-motive condition,

participants provided both of these ratings. Strong empathy is indicated by negative

feelings when seeing the partner in pain, indicated by negative ratings. Strong

reciprocity is indicated by positive feelings when observing the decision of the

partner, indicated by positive ratings. To allow the comparison of the ratings in all

conditions, empathy ratings were recoded such that positive ratings now reflect

strong empathy, i.e., multiplied by -1. The results of linear mixed models (lmms)

showed that the induction ratings in the motive conditions were significantly higher

than those in the baseline condition (χ^2 = 515.15, P < .000001, rating$_{baseline}$= -1.02 ±

1.00, rating$_{empathy}$= 1.57 ± 0.77, rating$_{reciprocity}$= 1.50 ± 0.89, rating$_{multi-motive}$ = 1.54 ±

0.91, (*M ± SEM*)). There were no significant differences in the induction ratings

between the motive conditions (χ^2 = 0.14, P = .93). The induction ratings in the

motive conditions were significantly associated with the frequency of prosocial

choices (χ^2 = 6.38, P = .01). This effect held to a comparable extent across all three

motive conditions (motive condition × rating interaction, χ^2 = 3.61, P = .16).

Specifically, the two single-motive conditions yielded similar induction ratings (χ^2 =

0.23, P = .64) and had a comparable effect on the frequency of prosocial choices (χ^2

= 4.77, P = .03, condition × rating interaction, χ^2 = 2.06, P = .15). These results show

that the strength of motive induction and the link to prosocial choices was

comparable for the empathy and the reciprocity motives (**Fig. A1**).

3.2.   Frequency of prosocial choices

Pairwise Kolmogorov-Smirnov tests revealed that the frequency of prosocial

decisions was comparably distributed across conditions (all Ds < 0.24, all Ps > 0.29,

for detailed statistics, please see **Table A3**) as were the reaction times (all Ds < 0.27,

all Ps > .17, **Fig. 3C**).

The frequency of prosocial choice was significantly influenced by condition (χ^2

= 56.99, P < .0001, see **Fig. 3A** and **Fig. A2**, prosoc$_{baseline}$= 67.7 ± 20.6 %,

prosoc$_{empathy}$= 77.0 ± 16.8, prosoc$_{reciprocity}$= 73.1 ± 18.4, prosoc$_{multi-motive}$= 77.4 ± 18.0

(*M ± SEM*)), indicating that the motive inductions had a differential effect on later

prosocial choices. Moreover, prosocial choices were influenced by the possible gain

for the partner (χ^2 = 6.68.64, P < .0001). However, model comparison revealed that

neither including the possible gain for the participant (χ^2 = .23, P = .63), nor its

interaction with condition significantly improved the model fit (χ^2 = .86, P = .84). Thus,

for the analyses reported below condition and possible gain for the partner were used

as additional predictors.

The frequency of empathy-based prosocial choices was increased compared

to reciprocity-based choices (z ratio = 2.94, P = .02), whereas the frequency of

prosocial choices between the multi-motive condition and the empathy condition was

comparable (z-ratio = .56, P = .94). However, the multi-motive condition yielded

significantly more prosocial choices compared to the reciprocity condition (z-ratio =

3.49, P = .003).

To clarify this effect, we calculated the percent change in prosocial choices in

the multi-motive condition relative to each single motive condition

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$$\Delta prosoc_{\text{multi-motive/reciprocity}} = \frac{prosoc_{\text{multi-motive}} - prosoc_{\text{reciprocity}}}{prosoc_{\text{reciprocity}}} \times 100$$

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$$\Delta prosoc_{\text{multi-motive/empathy}} = \frac{prosoc_{\text{multi-motive}} - prosoc_{\text{empathy}}}{prosoc_{\text{empathy}}} \times 100$$

where $prosoc_{\text{multi-motive}}$ equals the frequency of the prosocial choices in the multi-

motive condition, $prosoc_{\text{reciprocity}}$ equals the frequency of prosocial choices in the

reciprocity condition, and $prosoc_{\text{empathy}}$ equals the frequency of prosocial choices in

the empathy condition.

The percent change of the multi-motive condition relative to reciprocity was

significantly positive ($t(32)$ = 2.07, P = .047, $\Delta prosoc_{\text{multi-motive/reciprocity}}$ = 8.61 ± 4.17

($M \pm SEM$)), demonstrating that prosocial choices were enhanced when reciprocity

was combined with empathy, relative to reciprocity alone. The percent change in the

multi-motive condition relative to the empathy condition was not significantly different

from zero ($t(32)$ = 0.42, P = .674, $\Delta prosoc_{\text{multi-motive/empathy}}$ = 1.05 ± 2.47 ($M \pm SEM$)),

indicating that the simultaneous activation of the reciprocity motive did not enhance

the empathy motive.

3.3. Reaction times

Reaction times were significantly influenced by conditions ($\chi^2 = 27.89$, P < .0001).

That is, participants were faster in the motive conditions compared to the baseline

condition (baseline vs. empathy: t(32) = 5.03, P < .0001, baseline vs. reciprocity:

t(32) = 3.62, P = .002, baseline vs. multi-motive: t(32) = 3.70, P = .001). There were

no differences in reaction times for prosocial choices between the motive conditions

(all Ps > .49) (**Fig. 3B**), and the reaction time distributions were comparable (**Fig.**

3C).

Figure 3. Descriptive statistics, distributions and psychometric function of the choice and reaction time data. **(A)** Mean proportion of prosocial choices per condition. Error bars denote standard errors of means. **(B)** Mean reaction times per condition. Error bars denote standard errors of means. **(C)** Distribution of reaction times across participants per condition. **(D)** Psychometric functions for the different conditions of the probability to make a prosocial choice depending on the amount of points the participants' partner could possibly gain in each trial, that is, the point value for the partner in case of a prosocial choice minus the point value for the partner in

case of an egoistic choice. Vertical dashed lines indicate the points of subjective
equality in the different conditions (for exact values and spread, see **Table A4**).
Please note that the probability of making a prosocial decision never reached 0
because of the high frequency of prosocial choices in our data (participants made
prosocial choices even if the gain for the other person was low). Thus, our results do
not yield data points much lower than the respective points of subjective equality
(PSEs).
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3.4.   Point equality and prosocial behavior

In a next step, we tested whether considerations of equity differentially

influenced participants' prosocial behavior in the different conditions.

We considered the difference in point equality of the participant's and the

partner's outcome between the two choice options to test whether inequity aversion

differentially influenced participants' prosocial choice behavior. The results showed a

main effect of difference in point equality (χ^2 = 65.87, P < .0001) and a main effect of

condition (χ^2 = 46.91, P < .0001). However, no interaction effect was observed (χ^2 =

0.19, P = .98). Based on these results we conclude that inequity aversion does not

differentially affect the different motive conditions and thus cannot explain behavioral

differences in the different conditions.

In line with these results, the relationship between the frequency of prosocial

choices and the partner's possible gain was comparable between the different motive

conditions, as reflected by comparable values for the points of subjective equality

based on the psychometric functions estimated for the different conditions (χ^2 = 2.89,

P = .41, **Fig. 3D** and **Table A4**).

Since we were most interested in the underlying prosocial decision processes

in more complex as compared to simpler motivational states, we used hierarchical

drift-diffusion modeling (HDDM) (Vandekerckhove et al., 2011; Wiecki et al., 2013) to

understand prosocial choice behavior in the multi-motive condition relative to the

reciprocity condition and relative to the empathy condition.

3.5. Hierarchical drift-diffusion modeling

We estimated the three aforementioned DDM parameters (v, z, a) for every condition and participant. We also estimated the non-decision time t (0.94 ± 0.04 (M ± SEM)). However, this parameter was not estimated to vary by condition and was thus not further analyzed. Comparing the observed data with 500 datasets simulated by our model (Wiecki et al., 2013) showed that the winning model fit the data with 95% credibility (see **Table A5** for overview of all models and DIC values, and see **Table A6** for quantile comparison and 95% credibility). Based on the hypotheses depicted in **Fig. 1C**, we tested whether the observed percent change in the multi-motive condition can be explained by an increase in the speed of information accumulation (v-parameter, **Fig. 1C**, left panel), and/or an increase in initial prosocial bias (z-parameter, **Fig. 1C**, middle panel). Additionally, we tested whether the induction of both motives enhanced the amount of relative evidence that participants required during the choice process, relative to the two single-motive conditions (a-parameter, **Fig. 1C**, right panel).

Testing the first hypothesis (**Fig. 1C**, left panel), we observed no significant differences between the motive conditions ($p(v_{multi-motive} > v_{empathy}) = 46.38\ \%$, plausible t(32) = -.31, P = .76; $p(v_{multi-motive} > v_{reciprocity}) = 72.25\ \%$, plausible t(32) = 1.03, P = .31, see **Fig. A4** for distribution of t-values). Further, there was a slight percent change in v-parameters in the multi-motive condition relative to the reciprocity condition ($\Delta v_{multi-motive/reciprocity} = \frac{v_{multi-motive} - v_{reciprocity}}{v_{reciprocity}} \times 100 = -1.54 \pm 0.92\ \%$ ($M \pm SEM$), $t(32) = -1.71$, P = .09) and no percent change relative to the empathy condition ($\Delta v_{multi-motive/empathy} = \frac{v_{multi-motive} - v_{empathy}}{v_{empathy}} \times 100 = -0.13 \pm 0.83\ \%$ ($M \pm SEM$), $t(32) = -0.16$, P = .88). This result showed that the speed of information accumulation, i.e., the

efficiency of the choice process itself, was mainly unaffected by the combination of

the two motives, relative to the single-motive conditions.

Testing the second hypothesis (**Fig. 1C**, middle panel), we observed an

increase in initial prosocial bias (z-parameter) in the multi-motive condition compared

to the reciprocity condition ($p(z_{multi\text{-}motive} > z_{reciprocity})$) = *93.55 %*, plausible t(32) = 3.66, P

< .001) (**Fig. 4A**), but not compared to the empathy condition ($p(z_{multi\text{-}motive} > z_{empathy})$) =

70.33 %, plausible t(32) = 1.67, P = .10). The percent change in the *z*- parameter of

the multi-motive condition was significantly positive relative to the reciprocity

condition ($z_{multi\text{-}motive/reciprocity} = \dfrac{z_{mulit\text{-}motive} - z_{reciprocity}}{z_{reciprocity}} \times 100 = 6.40 \pm 1.21$ % (*M ± SEM*),

($t(32)$ = 5.36, P < .001) and marginally larger than zero relative to the empathy

condition ($\Delta z_{multi\text{-}motive/empathy} = \dfrac{z_{multi\text{-}motive} - z_{empathy}}{z_{empathy}} \times 100 = 2.54 \pm 1.39$ % (*M ± SEM*),

($t(32)$ = 1.85, P = .07).

In addition, we had hypothesized that the combination of the two motives may

increase the amount of relative evidence that participants required in order to reach a

decision (captured by the *a*-parameter; **Fig. 1C,** right panel). The *a*-parameter was

not significantly higher in the multi-motive condition compared to the reciprocity

condition ($p(a_{multi\text{-}motive} > a_{reciprocity})$) = 84.70 %, plausible t(32) = 1.73, P = .09) and the

empathy condition ($p(a_{multi\text{-}motive} > a_{empathy})$) = 82.35 %, plausible t(32) = 1.43, P = .16).

However, there was a significantly positive relative percent change in *a*-parameters in

the multi-motive condition relative to the reciprocity condition ($\Delta a_{multi\text{-}motive/reciprocity}$ =

$\dfrac{a_{multi\text{-}motive} - a_{reciprocity}}{a_{reciprocity}} \times 100 = 9.77 \pm 4.36$ % (*M ± SEM*), $t(32)$ = 2.28, P = .03) and also

relative to the empathy condition ($\Delta a_{multi\text{-}motive/empathy} = \dfrac{a_{multi\text{-}motive} - a_{empathy}}{a_{empathy}} \times 100 = 9.57\pm$

4.63 % (*M ± SEM*), $t(32)$ = 2.10, P = .04).

Figure 4. Increase in initial prosocial bias in the multi-motive condition relative to the

reciprocity condition and related neural activity. **(A)** Initial prosocial bias (z-parameter)

were significantly stronger in the multi-motive compared to the reciprocity condition

(plausible t(32) = 3.66, P < .001). Individual values are depicted for the multi-motive

condition (red) and the reciprocity condition (blue). Means and standard errors of the

mean are depicted in black. **(B)** The individual changes of initial prosocial choice

biases in the multi-motive condition relative to the reciprocity condition were tracked

by an increase in neural responses in the bilateral dorsal striatum (P(whole-brain

$FWE_{cluster-corrected}$) = .001; MNI peak coordinates; right hemisphere: x = 30, y = 2, z = -

2, left hemisphere: x = -28, y = -9, z = 1; visualized at P < .001 uncorrected; **Table A8**

and **Fig. A6**).

Taken together, the DDM results showed that the combination of the two

motives enhanced participants' bias for choosing the prosocial option, relative to the

initial prosocial choice bias biases induced by the reciprocity motive (captured by the

percent change in the z-parameter). The combination of empathy and reciprocity also

led to a relative increase in the amount of relative evidence that people required to

make a choice relative to the reciprocity motive, and also relative to empathy

(captured by the percent change in the a-parameter). In contrast, the speed of

information accumulation, i.e., the efficiency of the choice process itself, was

comparable between multi-motive and single-motive conditions (no change in v-

parameter).

These results may indicate that the observed percent changes in the multi-

motive condition relative to the reciprocity condition (in the z- and the a-parameters)

originate from the simultaneous activation of the two motives in the multi-motive

condition. Alternatively, as we observed no significant difference between the multi-

motive condition and the empathy-condition, it is also conceivable that the empathy

motive replaced the reciprocity motive when the two motives were activated

simultaneously. In this case, the observed percent changes in the multi-motive

condition would reflect the dominance of empathy over reciprocity, instead of a multi-

motive effect. If in fact empathy replaced the co-activated reciprocity motive, the

relative difference in the z-parameters and a-parameters between the empathy and

the reciprocity conditions should predict the individual extent of the percent changes

in the multi-motive condition relative to the reciprocity condition. To test this

explanation, we calculated the relative differences in the z-parameters and a-

parameters between empathy and reciprocity ($\Delta z_{\text{empathy/reciprocity}} =$

$\frac{z_{\text{empathy}} - z_{\text{reciprocity}}}{z_{\text{reciprocity}}} \times 100$ and $\Delta a_{\text{empathy/reciprocity}} = \frac{a_{\text{empathy}} - a_{\text{reciprocity}}}{a_{\text{reciprocity}}} \times 100$), entered them as

predictors in a regression analysis, and tested their effects on the observed percent

changes in the multi-motive condition ($\Delta z_{\text{multi-motive/reciprocity}}$; $\Delta a_{\text{multi-motive/reciprocity}}$). This

analysis revealed no significant effects ($\beta = -0.20$, $P = .42$; interaction with parameter

type (z vs a): $\beta = .002$, $P = 1.00$). These results demonstrate that the difference

between the two single motives cannot account for the changes in choice parameters

in the multi-motive condition relative to the reciprocity condition, bolstering the claim

that the observed effects are driven by the simultaneous activation of the two

motives. The three DDM parameters of interest for each condition and the relative

differences between the baseline condition and the motive conditions are provided in

Figure A3 and **Table A7**.

3.6. Imaging results

Next, we investigated the neural underpinnings of the prosocial decision process

comparing the multi-motive and the single motive conditions. The main contrasts of

mean neural activation during the prosocial decision phase did not show significant

neural activations (neither whole-brain nor small-volume corrected).

We investigated how the simultaneous activation of the two motives, and the

resulting changes in initial prosocial bias and amount of required relative evidence

affected the neural computation of prosocial choices. To do so, we regressed

participants' individual percent change in initial prosocial biases ($\Delta z_{\text{multi-motive/reciprocity}}$)

and the amount of relative evidence ($\Delta a_{\text{multi-motive/reciprocity}}$) on the neural contrast in

prosocial choices between the multi-motive condition and the reciprocity condition,

using second-level regressions. As a main result, the first analysis revealed

activations in the bilateral dorsal striatum that were related to the individual change in

prosocial bias (right hemisphere: $P(\text{FWE}_{\text{cluster-corrected}}) = 0.001$; center co-ordinates: x

= 30, y = 2, z = -2; k = 143 voxels, t(31) = 5.49; left hemisphere: $P(\text{FWE}_{\text{cluster-corrected}})$

= 0.003; center co-ordinates: x = -28, y = -9, z = 1; k = 121 voxels, t(31) = 5.36; **Fig.**

4B, Fig. A6, Table A8). The stronger the percent increase in initial prosocial bias in

the multi-motive condition relative the reciprocity condition, the stronger the neural

response in bilateral dorsal striatum.

To test the alternative hypothesis that the increase in dorso-striatal activity may

reflect the dominance of empathy (captured by the relative difference in z-parameters

between empathy and reciprocity, $\Delta z_{\text{empathy/ reciprocity}}$), instead of a multi-motive effect,

we also compared the relationship between $\Delta z_{\text{multi-motive/reciprocity}}$ and $\Delta z_{\text{empathy/reciprocity}}$

on extracted beta values of the multi-motive vs reciprocity contrast using an

independent anatomical mask of bilateral putamen based on the aal nomenclature

(Tzourio-Mazoyer et al., 2002). The results showed that neural activation in dorsal

striatum is associated with $\Delta z_{\text{multi-motive/reciprocity}}$, but not with $\Delta z_{\text{empathy/reciprocity}}$

(significant interaction between index type and neural activation: $\beta = 0.69$, $P = .003$,

Fig. A5B). Thus, empathy dominance is not likely to explain the results.

To test whether the increase in striatal activation in the multi-motive compared to

the reciprocity condition is driven by outliers, we extracted the individual beta-

estimates of the multi-motive vs reciprocity contrast from bilateral dorsal striatum and

plotted its relationship with the percent change in the z-parameter in the multi-motive

condition relative to the reciprocity condition (**Fig. A5A**). The inspection of the plot

shows that the relationship between the percent signal change in the z-parameter in

the multi-motive condition relative to the reciprocity condition was not driven by

outliers.

The respective second-level regression with the percent change in the *a*-

parameter revealed neural activity in bilateral anterior insula on a lower, uncorrected

threshold ($P_{\text{uncorrected}} < 0.001$; center co-ordinates right hemisphere: $x = 33$, $y = 32$, z

$= 1$, $P(FWE_{\text{cluster-corrected}}) = .970$, $k = 9$ voxels; center co-ordinates left hemisphere: $x =$

-30, $y = 27$, $z = -2$, $P(FWE_{\text{cluster-corrected}}) = .902$, $k = 13$ voxels).

Additionally, we tested whether trial-by-trial changes in the partner's gain

modulate neural activation during the prosocial choice process. In line with the

behavioral results, the partner's gain did not differentially influence neural activation

in the different conditions. However, neural activation during the prosocial choice

process in bilateral insula was significantly associated with trial-by-trial changes in

the partner's gain across all four conditions (right insula peak-coordinates: $x = 43$, $y =$

-6, $z = 18$, $k = 108$ voxels, $P(FWE$ whole-brain cluster corrected$) = .009$; left insula

peak-coordinates: $x = -38$, $y = -9$, $z = 16$, $k = 517$ voxels, $P(FWE$ whole-brain cluster

corrected$) < .001$; see Table A9 for all clusters $k > 10$). The same analysis including

prosocial as well as egoistic choice trials replicated this result (right insula peak-

coordinates: x = 43, y = -6, z = 16, k = 109 voxels, P(FWE whole-brain cluster

corrected) = .009; left insula peak-coordinates: x = -38, y= -9, z = 18, k = 294 voxels,

P(FWE whole-brain cluster corrected) < .001). Hence, the insular activation appears

to track trial-by-trial changes of the partner's gain and does so to comparable extent

in all conditions.

4. Discussion

Many behaviors derive from complex motivational states that are

characterized by different, simultaneously activated motives (Engel and

Zhurakhovska, 2016; Hughes and Zaki, 2015; Jagers et al., 2017; Takeuchi et al.,

2015; Terlecki and Buckner, 2015). However, the mechanisms through which

combinations of motives affect behaviors, e.g., the computation of prosocial choices,

are poorly understood.

Our results showed that the simultaneous activation of two motives changes

participants' choices compared to activation of a single motive condition. In more

detail, a combination of two prosocial motives (the empathy and the reciprocity

motive) elicited more prosocial choices than the reciprocity motive alone (**Fig. 3A**).

This multi-motive increase occurred although the two single motives were activated

with comparable strength (indicated by the induction ratings, **Fig. A1**). Moreover, the

different motive conditions had no effect on reaction times (**Fig. 3B** and **3C**),

inequality aversion, or the subjective value assigned to the partner's gains (**Fig. 3D**).

Specifying the change in prosocial choice behavior in the multi-motive

condition, the drift-diffusion-modeling analyses showed that the combination of the

two motives enhanced participants' initial bias for making a prosocial choice,

compared to the reciprocity condition (reflected by the increase in z-parameter; **Fig.**

4A) and with a similar trend, relative to the empathy condition. Moreover, the

combinations of the two motives increased the relative amount of relative evidence

that participants required during the choice process (reflected by the relative increase

of the a-parameter). This indicates that participants assessed their choices more

carefully if they made them based on two different motives. In contrast, the speed of

information accumulation, i.e., the efficiency of the decision process itself (reflected

by the v-parameter), remained unchanged. The observed change in initial prosocial

bias (the z-parameter) is in line with previous findings that reported a shift of choice

biases due to the prior likelihood of one of the choice options or a higher reward

value associated with one option (Mulder et al., 2012), personal predispositions

(Chen and Krajbich, 2018), or prior information about how other people decided

(Toelch et al., 2018). Extending these results, our findings reveal that initial choice

biases are altered by simultaneously activated motives, and thus characterize how

complex motivational states change the choice process compared to single-motive

states.

We hypothesized that changes in DDM choice parameters in the multi-motive

compared to the single-motive conditions might be related to changes in activation in

the ventral or dorsal striatum, inspired by evidence associating the ventral striatum

with the processing of choice values (Kable and Glimcher, 2007; Liljeholm and O

'Doherty, 2012; O'Doherty et al., 2004; Strait et al., 2015), and/or the dorsal striatum

with encoding of choice preferences (Balleine et al., 2007; Liljeholm and O 'Doherty,

2012; O'Doherty et al., 2004; Palmiter, 2008; Robinson et al., 2006).Our results show

that the combination of different motives is associated with an increase in activation

in bilateral dorsal striatum, reflecting an enhancement of individual prosocial choice

biases in the multi-motive condition relative to the reciprocity condition (**Fig. 4**), and,

based on extracted beta-values from an independent anatomical region, also relative

to the empathy condition (**Fig. A5**). The increase in activation in dorsal striatum is in

line with previous neuroscience studies showing that motivation-related changes in

decision parameters are captured by dorsal striatal responses (Forstmann et al.,

2008; Gluth et al., 2012). Extending this previous evidence, we show that the dorsal

striatum integrates choice biases that are elicited by multiple motivational forces, and

thus provides a plausible neural candidate for the generation of complex motivational

states.

We found that the simultaneous activation of the empathy motive and the

reciprocity motive in the multi-motive condition enhanced the participants' initial

prosocial biases relative to the reciprocity condition. This indicates that the empathy

motive enhanced the reciprocity motive, but not vice versa. Given this result, we

argued that the observed changes in the multi-motive condition may reflect the

dominance of one motive over the other motive (i.e., a dominance of empathy over

reciprocity). If this were true, the multi-motive induced changes in the choice process

would reflect a motivation that is similar to the state induced by the dominant motive,

instead of a motivational state that was incited by the combination of different

motives. Our results show that the multi-motive induced changes in the choice

process (i.e., DDM and neural choice parameters) are in fact related to differences

between the multi-motive condition and the reciprocity condition and cannot be

explained by mere dominance of the empathy motive over the reciprocity motive.

This finding supports the conclusion that the simultaneous activation of two motives

alters the prosocial choice process compared to single motive states.

Because we were mainly interested in participants' choices under the different

motive conditions, our paradigm was designed to optimize the number of trials in the

allocation task. The motive induction procedures only included the minimal number of

trials required for inducing the different motives (twelve trials per motive induction).

Due to the small number of trials, an analysis of neural responses during the multi-

motive and single-motive induction procedures would not be meaningful.

Participants made prosocial choices also in the baseline condition which

indicates that participants are motivated to behave prosocially independently of the

induced motives. It is thus important to note that prosocial choices can be driven by

other motives apart from empathy and reciprocity. However, these additional motives

should be the same across conditions since participants perform the same social

choice task. Hence, contrasting behavior between the different conditions should

carve out the effects that were experimentally manipulated, i.e., how the combination

of empathy and reciprocity influences the prosocial choice process relative to

empathy or reciprocity alone.

Likewise, because the "basic" social choice tasks were identical in all

experimental conditions, task-specific effects should average out if the different

conditions are contrasted. This means that the observed effects are driven by the

different motive inductions and should be independent of the social choice task that

was used in the present study. In other words, our behavioural and neural findings

should generalize to other behaviours that are elicited by the combination of the

empathy and the reciprocity motives. However, in how far the present effects are

scaled depending on the exact task affordances (e.g., relative importance of the

single motives for the respective task, how much time participants have to deliberate

their decision) is a question for future research. Likewise, future studies need to test if

the observed increase in striatal activation due to a multi-motive alteration of initial

choice bias also applies to other (e.g., non-social) motivational states. Food choices,

for example, are often driven by more than one motive such as the motive to eat

healthy and the motive to eat sweet food, maximizing calorie intake. The dorsal

striatum has previously been associated with food choice preferences in healthy

(Small et al., 2003; Wallace et al., 2014) as well as pathological participants (Foerde

et al., 2015). It is thus possible that the interplay between the non-social motives

during food choices is associated with neural activation in the dorsal striatum.

To avoid cross-gender effects, which are likely to occur if female participants

are paired with male confederates and vice versa, we only tested females. Future

studies are required to show if our results generalize to male participants.

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5. Conclusions

Based on our current findings we conclude that the simultaneous activation of

two different prosocial motives changes the computation of prosocial choices.

According to our results, choices that are made in a more complex motivational state,

i.e., driven by multiple motives, are characterized by a change in initial choice bias,

which is associated with an increased neural response in dorsal striatum. Moreover,

choices were made more carefully relative to simple motivational states. Together,

these findings show how the human brain combines different prosocial motives, and

how this motive combination affects the computation of prosocial choices.

Authors' Contributions

G.H. and A.S. designed the research with input from J.K., M.L., and U.H.; A.S. and

U.H. performed the research; A.S. programmed the experiment; A.S. and U.H.

analyzed the data with input from G.H., M.L., and J.K.; G.H. and A.S. wrote the paper

with input from U.H., M.L., and J.K..

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References

Báez-Mendoza, R., Schultz, W., 2013. The role of the striatum in social behavior.
Front. Neurosci. 7. https://doi.org/10.3389/fnins.2013.00233

Balleine, B.W., Delgado, M.R., Hikosaka, O., 2007. The Role of the Dorsal Striatum
in Reward and Decision-Making. J. Neurosci.
https://doi.org/10.1523/JNEUROSCI.1554-07.2007

Batson, C.D., Ahmad, N., Stocks, E.L., 2011. Four forms of prosocial motivation
egoism, altruism, collectivism, and principlism, in: Dunning, D. (Ed.), Social
Motivation. Psychology Press, New York, pp. 103–126.
https://doi.org/10.4324/9780203833995

Batson, C.D., Turk, C.L., Shaw, L.L., Klein, T.R., 1995. Information Function of
Empathic Emotion. J. Pers. Soc. Psychol. 68, 300–313.

Bhanji, J.P., Delgado, M.R., 2014. The social brain and reward: Social information
processing in the human striatum. Wiley Interdiscip. Rev. Cogn. Sci. 5, 61–73.
https://doi.org/10.1002/wcs.1266

Bolton, G.E., Ockenfels, A., 2000. ERC: A theory of equity, reciprocity, and
competition. Am. Econ. Rev. https://doi.org/10.1257/aer.90.1.166

Brett, M., Anton, J.L., Valabregue, R., Poline, J.B., 2002. Region of interest analysis
using an SPM toolbox - abstract presented at the 8th International Conference
on Functional Mapping of the Human Brain, June 2-6, 2002, Sendai, Japan.
Neuroimage. https://doi.org/http://dx.doi.org/10.1016/S1053-8119(02)90010-8

Carter, C.S., Lesh, T.A., Barch, D.M., 2016. Thresholds, Power, and Sample Sizes in
Clinical Neuroimaging. Biol. Psychiatry Cogn. Neurosci. Neuroimaging 1, 99–
100. https://doi.org/10.1016/j.bpsc.2016.01.005

Charness, G., Rabin, M., 2002. Understanding social preferences with simple tests.

Q. J. Econ. https://doi.org/10.1162/003355302760193904

Chen, F., Krajbich, I., 2018. Biased sequential sampling underlies the effects of time pressure and delay in social decision making. Nat. Commun. 9, 1–10. https://doi.org/10.1038/s41467-018-05994-9

Eickhoff, S.B., Stephan, K.E., Mohlberg, H., Grefkes, C., Fink, G.R., Amunts, K., Zilles, K., 2005. A new SPM toolbox for combining probabilistic cytoarchitectonic maps and functional imaging data. Neuroimage 25, 1325–1335. https://doi.org/10.1016/j.neuroimage.2004.12.034

Eklund, A., Nichols, T.E., Knutsson, H., 2016. Cluster failure: Why fMRI inferences for spatial extent have inflated false-positive rates. Proc. Natl. Acad. Sci. U. S. A. https://doi.org/10.1073/pnas.1602413113

Engel, C., Zhurakhovska, L., 2016. When is the risk of cooperation worth taking? The prisoner's dilemma as a game of multiple motives. Appl. Econ. Lett. 23, 1157–1161. https://doi.org/10.1080/13504851.2016.1139672

Fehr, E., Schmidt, K.M., 1999. A theory of fairness, competition, and cooperation. Q. J. Econ. https://doi.org/10.1162/003355399556151

Flagan, T., Mumford, J.A., Beer, J.S., 2017. How do you see me? The neural basis of motivated meta-perception. J. Cogn. Neurosci. 29, 1908–1917. https://doi.org/10.1162/jocn_a_01169

Foerde, K., Steinglass, J.E., Shohamy, D., Walsh, B.T., 2015. Neural mechanisms supporting maladaptive food choices in anorexia nervosa. Nat. Neurosci. https://doi.org/10.1038/nn.4136

Forstmann, B.U., Dutilh, G., Brown, S., Neumann, J., Von Cramon, D.Y., Ridderinkhof, K.R., Wagenmakers, E.J., 2008. Striatum and pre-SMA facilitate decision-making under time pressure. Proc. Natl. Acad. Sci. U. S. A. https://doi.org/10.1073/pnas.0805903105

Forstmann, B.U., Ratcliff, R., Wagenmakers, E.-J., 2016. Sequential Sampling Models in Cognitive Neuroscience: Advantages, Applications, and Extensions. Annu. Rev. Psychol. 67. https://doi.org/10.1146/annurev-psych-122414-033645

Forsythe, R., Horowitz, J.L., Savin, N.E., Sefton, M., 1994. Fairness in simple bargaining experiments. Games Econ. Behav. 6, 347–369.

Gallotti, R., Grujić, J., 2019. A quantitative description of the transition between intuitive altruism and rational deliberation in iterated Prisoner's Dilemma experiments. Sci. Rep. https://doi.org/10.1038/s41598-019-52359-3

Gelman, A., Rubin, D.B., 1992. Inference from Iterative Simulation Using Multiple Sequences. Stat. Sci. 7, 457–472.

Gluth, S., Rieskamp, J., Büchel, C., 2012. Deciding when to decide: Time-variant sequential sampling models explain the emergence of value-based decisions in the human brain. J. Neurosci. https://doi.org/10.1523/JNEUROSCI.0727-12.2012

Gouldner, A.W., 1960. The Norm of Reciprocity: A Preliminary Statement. Am. Sociol. Rev. https://doi.org/10.2307/2092623

Hare, T.A., Schultz, W., Camerer, C.F., O'Doherty, J.P., Rangel, A., 2011. Transformation of stimulus value signals into motor commands during simple choice. Proc. Natl. Acad. Sci. U. S. A. https://doi.org/10.1073/pnas.1109322108

Hein, G., Morishima, Y., Leiberg, S., Sul, S., Fehr, E., 2016a. The brain's functional network architecture reveals human motives. Science (80-.). 351, 1074–1078.

Hein, G., Morishima, Y., Leiberg, S., Sul, S., Fehr, E., 2016b. The brain's functional network architecture reveals human motives. Science (80-.). 351, 1074–1078.

Hughes, B.L., Zaki, J., 2015. The neuroscience of motivated cognition. Trends Cogn. Sci. 19, 62–64. https://doi.org/10.1016/j.tics.2014.12.006

Jagers, S.C., Linde, S., Martinsson, J., Matti, S., 2017. Testing the Importance of Individuals' Motives for Explaining Environmentally Significant Behavior. Soc.

Sci. Q. 98, 644–658. https://doi.org/10.1111/ssqu.12321

Janczyk, M., Lerche, V., 2019. A diffusion model analysis of the response-effect

compatibility effect. J. Exp. Psychol. Gen. 148, 237–251.

Jennings, J.H., Sparta, D.R., Stamatakis, A.M., Ung, R.L., Pleil, K.E., Kash, T.L.,

Stuber, G.D., 2013. Distinct extended amygdala circuits for divergent

motivational states. Nature 496, 224. https://doi.org/10.1038/nature12041

Jordan, M.R., Amir, D., Bloom, P., 2016. Are empathy and concern psychologically

distinct? Emotion 16, 1107–1116. https://doi.org/10.1037/emo0000228

Kable, J.W., Glimcher, P.W., 2007. The neural correlates of subjective value during

intertemporal choice. Nat. Neurosci. https://doi.org/10.1038/nn2007

Kennedy, P.J., Shapiro, M.L., 2009. Motivational states activate distinct hippocampal

representations to guide goal-directed behaviors. Proc. Natl. Acad. Sci.

https://doi.org/10.1073/pnas.0903259106

Kim, B., Im, H.-I., 2018. The role of the dorsal striatum in choice impulsivity. Ann. N.

Y. Acad. Sci.

Krajbich, I., Hare, T., Bartling, B., Morishima, Y., Fehr, E., 2015. A Common

Mechanism Underlying Food Choice and Social Decisions. PLoS Comput. Biol.

https://doi.org/10.1371/journal.pcbi.1004371

Kruglanski, A.W., Shah, J.Y., Fishbach, A., Friedman, R., Chun, W.Y., Sleeth-

Keppler, D., 2018. A theory of goal systems, in: The Motivated Mind. Routledge,

pp. 215–258. https://doi.org/10.4324/9781315175867

Lamm, C., Decety, J., Singer, T., 2011. Meta-analytic evidence for common and

distinct neural networks associated with directly experienced pain and empathy

for pain. Neuroimage 54, 2492–2502.

https://doi.org/10.1016/j.neuroimage.2010.10.014

Lerche, V., Voss, A., 2016. Model complexity in diffusion modeling: Benefits of

making the model more parsimonious. Front. Psychol.

https://doi.org/10.3389/fpsyg.2016.01324

Lewin, K., Cartwright, D., Price, D., 1951. Field theory in social science: Selected

theoretical papers., American Sociological Review. https://doi.org/10.1037/0021-

9010.69.1.85

Liljeholm, M., O 'Doherty, J.P., 2012. Contributions of the striatum to learning,

motivation, and performance: an associative account Anatomical and functional

delineations of the striatum. Trends Cogn. Sci.

https://doi.org/10.1016/j.tics.2012.07.007

Ly, A., Boehm, U., Heathcote, A., Turner, B.M., Forstmann, B., Marsman, M.,

Matzke, D., 2017. A Flexible and Efficient Hierarchical Bayesian Approach to the

Exploration of Individual Differences in Cognitive-model-based Neuroscience, in:

Computational Models of Brain and Behavior.

https://doi.org/10.1002/9781119159193.ch34

Maldjian, J.A., Laurienti, P.J., Kraft, R.A., Burdette, J.H., 2003. An automated method

for neuroanatomic and cytoarchitectonic atlas-based interrogation of fMRI data

sets. Neuroimage. https://doi.org/10.1016/S1053-8119(03)00169-1

Marsman, M., Maris, G., Bechger, T., Glas, C., 2016. What can we learn from

Plausible Values? Psychometrika 81, 274–289. https://doi.org/10.1007/s11336-

016-9497-x

Matzke, D., Wagenmakers, E.J., 2009. Psychological interpretation of the ex-

gaussian and shifted wald parameters: A diffusion model analysis. Psychon.

Bull. Rev. https://doi.org/10.3758/PBR.16.5.798

McCabe, K.A., Rigdon, M.L., Smith, V.L., 2003. Positive reciprocity and intentions in

trust games. J. Econ. Behav. Organ. 52, 267–275.

https://doi.org/10.1016/S0167-2681(03)00003-9

Mulder, M.J., Wagenmakers, E.-J., Ratcliff, R., Boekel, W., Forstmann, B.U., 2012. Bias in the Brain: A Diffusion Model Analysis of Prior Probability and Potential Payoff. J. Neurosci. 32, 2335–2343. https://doi.org/10.1523/jneurosci.4156-11.2012

O'Doherty, J., Dayan, P., Schultz, J., Deichmann, R., Friston, K., Dolan, R.J., 2004. Dissociable Roles of Ventral and Dorsal Striatum in Instrumental Conditioning. Science (80-.). https://doi.org/10.1126/science.1094285

Palmiter, R.D., 2008. Dopamine signaling in the dorsal striatum is essential for motivated behaviors: Lessons from dopamine-deficient mice. Ann. N. Y. Acad. Sci. 35. https://doi.org/10.1196/annals.1417.003

Perugini, M., Gallucci, M., Presaghi, F., Ercolani, A.P., 2003. The Personal Norm of Reciprocity. Eur. J. Pers. https://doi.org/10.1002/per.474

Ratcliff, R., Smith, P.L., Brown, S.D., McKoon, G., 2016. Diffusion Decision Model: Current Issues and History. Trends Cogn. Sci. 20, 260–281. https://doi.org/10.1016/j.tics.2016.01.007

Robinson, S., Sotak, B.N., During, M.J., Palmiter, R.D., 2006. Local dopamine production in the dorsal striatum restores goal-directed behavior in dopamine-deficient mice. Behav. Neurosci. 120, 196. https://doi.org/10.1037/0735-7044.120.1.000

Roiser, J.P., Linden, D.E., Gorno-Tempinin, M.L., Moran, R.J., Dickerson, B.C., Grafton, S.T., 2016. Minimum statistical standards for submissions to Neuroimage: Clinical. NeuroImage Clin. 12, 1045–1047. https://doi.org/10.1016/j.nicl.2016.08.002

Salamone, J.D., Correa, M., 2012. The Mysterious Motivational Functions of Mesolimbic Dopamine. Neuron 76, 470–485. https://doi.org/10.1016/j.neuron.2012.10.021

Salamone, J.D., Pardo, M., Yohn, S.E., López-Cruz, L., Sanmiguel, N., Correa, M., 2016. Mesolimbic dopamine and the regulation of motivated behavior, in: Behavioral Neuroscience of Motivation. Springer, Cham, pp. 231–257. https://doi.org/10.1007/7854_2015_383

Shohamy, D., 2011. Learning and motivation in the human striatum. Curr. Opin. Neurobiol. 21, 408–414. https://doi.org/10.1016/j.conb.2011.05.009

Small, D.M., Jones-Gotman, M., Dagher, A., 2003. Feeding-induced dopamine release in dorsal striatum correlates with meal pleasantness ratings in healthy human volunteers. Neuroimage. https://doi.org/10.1016/S1053-8119(03)00253-2

Strait, C.E., Sleezer, B.J., Hayden, B.Y., 2015. Signatures of value comparison in ventral striatum neurons. PLoS Biol. https://doi.org/10.1371/journal.pbio.1002173

Takeuchi, R., Bolino, M.C., Lin, C.C., 2015. Too many motives? The interactive effects of multiple motives on organizational citizenship behavior. J. Appl. Psychol. 100, 1239. https://doi.org/10.1037/apl0000001

Team, R.C., 2018. R: A Language and Environment for Statistical Computing [WWW Document]. Vienna, Austria.

Terlecki, M.A., Buckner, J.D., 2015. Social anxiety and heavy situational drinking: Coping and conformity motives as multiple mediators. Addict. Behav. 40, 77–83. https://doi.org/10.1016/j.addbeh.2014.09.008

Toelch, U., Panizza, F., Heekeren, H.R., 2018. Norm compliance affects perceptual decisions through modulation of a starting point bias. R. Soc. Open Sci. 5, 171268. https://doi.org/10.1098/rsos.171268

Tzourio-Mazoyer, N., Landeau, B., Papathanassiou, D., Crivello, F., Etard, O., Delcroix, N., Mazoyer, B., Joliot, M., 2002. Automated anatomical labeling of activations in SPM using a macroscopic anatomical parcellation of the MNI MRI

single-subject brain. Neuroimage 15, 273–289.

https://doi.org/10.1006/nimg.2001.0978

Vandekerckhove, J., Tuerlinckx, F., Lee, M.D., 2011. Hierarchical Diffusion Models

for Two-Choice Response Times. Psychol. Methods 16, 44–62.

https://doi.org/10.1037/a0021765

Wallace, D.L., Aarts, E., Dang, L.C., Greer, S.M., Jagust, W.J., D'Esposito, M., 2014.

Dorsal striatal dopamine, food preference and health perception in humans.

PLoS One. https://doi.org/10.1371/journal.pone.0096319

Wiecki, T. V., Sofer, I., Frank, M.J., 2013. HDDM: Hierarchical Bayesian estimation of

the Drift-Diffusion Model in Python. Front. Neuroinform. 7, 14.

https://doi.org/10.3389/fninf.2013.00014

Woo, C.-W., Krishnan, A., Wager, T.D., 2014. Cluster-extent based thresholding in

fMRI analyses: Pitfalls and recommendations. Neuroimage 91, 412–419.

https://doi.org/10.1016/j.neuroimage.2013.12.058

Yeung, A.W.K., 2018. An updated survey on statistical thresholding and sample size

of fMRI studies. Front. Hum. Neurosci. 12, 1–7.

https://doi.org/10.3389/fnhum.2018.00016

Zhang, Y., Fishbach, A., Kruglanski, A.W., 2007. The dilution model: How additional

goals undermine the perceived instrumentality of a shared path. J. Pers. Soc.

Psychol. https://doi.org/10.1037/0022-3514.92.3.389

1232