**Divergent strategies for learning in males and females**

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**Abstract**

Reported sex differences in decision-making and learning can be inconsistent across studies. One interpretation is that these sex differences are not driven by differences in ability, but by differences in strategy, which interact with task design. Here, we examined the strategies adopted by female and male mice as they learned the value of stimuli that varied across two dimensions. Female mice mastered image-value associations more quickly than male mice, and that they used a fundamentally different strategy to do so. Female mice constrained their decision-space early in learning. Conversely, male strategies changed frequently and were more influenced by the stochastic rewards. Individual strategies were related to sex-gated changes in neuronal activation in early learning. Together, we find that sex drives divergent strategies for learning, revealing substantial unrecognized variability in reward-guided decision-making and learning.

## Introduction

Sex differences have been identified in many studies of reward-guided decision-making, seen across multiple mammalian species, and gender differences have been reported in humans (*1–5*). In some tasks males acquire more total rewards than females, a result that is sometimes interpreted as a male advantage in learning (*6, 7*). However, seemingly minor changes to task paradigms can produce the opposite result. In fact, in some sequential probabilistic choice tasks, there is a female advantage in the number of total rewards earned (*2, 8, 9*). These findings suggest that males and females may not actually differ dramatically in their *ability to learn* from rewards. If so, then an alternative explanation is needed for the differences in sequential decision making that lead to different patterns of reward acquisition in different tasks (*10–12*).

Differences in reward outcomes need not solely arise from differences in ability to learn. Instead, they can also arise from differences in the choice of *what to learn about* (*10, 12–15*): the strategy used for learning. In this view, sex-linked mechanisms influence learning strategies, and, because different strategies pay off differently in different environments, females have an advantage in some environments, while males have an advantage in others. However, this can only be uncovered in tasks that use choices that vary in multiple dimensions (*12, 14*). To illustrate, one of the challenges of moving to a new city is finding a favorite restaurant. In a new dining scene, it is not always clear what dimensions of a restaurant--meaning features like location, price, or type of cuisine--best predict high quality meals. One strategy for learning about a new dining scene could be to try all restaurants at random, sampling the entire environment to simultaneously learn about all the dimensions in which restaurants vary, until a

winner is found. This strategy might find the best option, but would be incredibly time

consuming and may be sensitive to noise in meal quality because each restaurant is sampled less

frequently. Another strategy might be to try to learn about the feature dimensions by constraining

the search in one dimension (like neighborhood) while learning about how restaurants vary in the

other dimensions. To a naive observer, this approach may appear to be unnecessarily risk-averse

or limited, but holding some dimensions constant can facilitate learning about other dimensions,

particularly when feedback is noisy. However, it is not clear what factors influence how

individuals select one of these different strategies.

To examine the possibility that sex differences in decision making may arise from different

learning strategies, we examined male and female mice as they performed a two-dimensional

decision-making task: a two-armed visual bandit (*11*, *16–22*). We considered the possibility that

animals were adopting divergent strategies to solve the task, and that these strategies could be

tuned by sex differences. We found that while males and females eventually reached the same

performance level, female mice learned more rapidly than males and acquired more rewards over

the course of learning. The difference in the rate of learning was not because females learned

more from outcome of each trial, but because of a sex difference in the strategies that governed

the next choice. Female mice adopted a consistent and systematic approach of preferring to

choose options in one spatial location, which constrained the decision-space and accelerated their

learning about image values. Conversely, choices made by males did not follow a single,

straightforward strategic approach. Males appeared to consider information from both image and

spatial dimensions simultaneously, and were highly sensitive to the stochastic experience of

reward. As a result, individual males substantially changed *their own* choice strategies over

learning, differing from *themselves* much more so than did individual females. During early

learning, gene expression for the neuronal activation marker c-fos in the nucleus accumbens and

prefrontal cortex significantly correlated with the extent to which female animals (but not male)

used a systematic strategy. These results suggest that sequential decision making for reward can

be achieved through widely divergent strategies, within and between subjects, and that strategies

employed during learning are a significant source of sex differences in decision making.

**Results**

Age-matched male and female wildtype mice (n=32, 16 per sex, strain B6129SF1/J) were trained

to perform a visually-cued two-armed bandit task in touchscreen operant chambers (**Figure 1a**).

This task design was similar to those employed in humans and nonhuman primates (*11*, *16–24*),

in contrast to the spatial bandit designs frequently employed with rodents (*25–28*). Animals were

presented with a repeating set of two different image cues which were associated with different

probabilistic reward outcomes (**Figure 1b**).The reward schedule (80%/20%) was held constant

throughout the session. In contrast with spatial bandit designs, here reward contingencies were

yoked to image identity, which was randomized with respect to location in the chamber on each

trial. This means that the sides (left/right) where image cues appeared were not informative of

the reward contingencies. We repeated the task with six different sets of image pairs. Two out of

the six tested image pairs were excluded from the study due to extremely high initial preference

(>70%) for one image over another. We included four images pairs with equal initial preference

for each image and quantified behavioral data in bins of 150 trials for each animal.

**Females showed accelerated reinforcement learning, but males and females reached**

**equivalent final performance**

To examine learning performance, we calculated in bins of 150 trials the average percentage of

choosing the high-value image (23 bins in total). Regardless of sex, mice were capable of

eventually learning which image was associated with the higher reward probability (**Figure 1c**,

GLM, main effect of sex, $p = 0.51$, $\beta1 = -0.05$; main effect of number of trials, $p < 0.0001$, $\beta2 =$

0.10, see equation 1 in Methods). However, we repeatedly observed that females learned the

image pair discrimination significantly faster than did males (GLM, interaction term, $p < 0.05$,

$\beta3 = -0.02$). We compared these results to a deterministic version of the task in the same animals,

in which one image was always rewarded (100%) and the other was never rewarded (0%). We

did not find any significant sex difference in rate of learning across trials in the deterministic task

(**Figure 1d**, GLM, interaction term, $p = 0.38$, $\beta1 = -0.004$, see equation 1 in Methods),

suggesting the difference was revealed by the stochastic experience of reward.

To determine the origins of this sex difference in early performance, we first considered that sex

differences in early learning might reflect differences in the rate of value-updating and/or the

level of random noise in choice--the typical parameters of reinforcement learning models. We fit

a delta-rule reinforcement model (*18, 25, 29, 30*) to measure individual differences in the

learning rate parameter and noise parameter, based on choices of images. However, the

likelihood surface of the model given by parameters learning rate ($\alpha$) and inverse temperature ($\beta$)

was flat, which prevented parameter optimization. This suggested to us that the basic RL model

based on images as the sole choice dimension could not characterize individual differences in

learning in this task.

Since rodents are generally highly spatial, we hypothesized that mice might have a bias towards

using spatial information earlier in the task, before switching to use image information, as

demonstrated by the high rates of reward late in training. Consistent with our hypothesis, we

observed a short period of high side bias in females early in learning (**Figure 1e**), which could

include a preference for either the left or right side and seemed to precede the acquisition of the

reward contingency. Following the decrease of side bias, female mice improved their percentage

of choosing high-value image. (GLM, main effect of sex, $p < 0.001$, $\beta 1 = -0.129$; main effect of

number of trials, $p < 0.001$, $\beta 2 = -0.017$, see equation 1 in methods). Strength of side bias

independent of left or right side was calculated using methods described in previous behavioral

lateralization literature (*31*).

**Females systematically reduced the dimensions of the task by strongly preferring one side**

A side bias is only one of several local strategies that mice could have been using as they learned

the reward contingencies in the task. For example, another local strategy is a spatial win-stay

strategy, where the side of the last choice is repeated if and only if it was rewarded.

Alternatively, an image win-stay strategy would repeat the last image, if and only if it was

rewarded, or an image bias strategy would simply select the previous image, regardless of

reward. To understand how different animals employed these different local strategies and

processed through them over time, we constructed a generalized linear model (GLM) to predict

each choice, based on a weighted combination of these local strategies. The model had a term to

account for two classes of basic strategies: outcome-independent strategies and outcome-

dependent, win-stay, strategies (**Figure 2a**, see Equation 4 in methods). Outcome-independent

strategies (image repeat and spatial repeat) captured the tendency of repeating either the side or

the image of the previous trial, regardless of the outcome. Conversely, outcome-dependent

strategies followed a win-stay lose-shift policy for either a side or an image, which captured the

tendency  to only repeat the side or the image of the previous trial when it was rewarded. Fitting

the GLM allowed us to estimate how much each of these four strategies was employed within

each animal on each bin of trials. We will call this set of weights--the precise pattern of local

strategies employed over time--the "global strategy" employed by each individual animal.

Across all animals, we found a global strategy where a specific procession through local

strategies was used when learning image pairs (**Figure 2b**). First, animals showed an early

tendency towards outcome-independent spatial repeat, giving way to a later interaction between

reward outcome and image choice, with outcome-insensitive image repeat (the optimal strategy)

increasing in the later stages of testing. To examine whether sex influenced the strength of this

strategy procession, we compared the global strategy used by male and female animals. We

observed a consistent and pronounced pattern of strategy procession only in females (**Figure 2c**).

In contrast, in males, the weight of both image-based strategies simply increased slowly  over

time (**Figure 2d**), with little change in spatial strategies. Thus, neither a procession of multiple

strategies nor a prominent strategy in the early learning stage was observed in male mice.

The sex difference in this strategy selection was intriguing, but it could have been driven by only

a few females. Therefore, we next characterized individual variability in strategy via principal

component analysis. Here, we estimated the major axes of interindividual variability in strategy,

meaning in the unique combinations of the four strategy weight vectors over time and across all

animals, regardless of sex. Principal components (PC) 1 and 2 captured the majority of the

interindividual variance: 59% of the variability between animals (**Figure 2e**). PC1 reflected a

global preference for a side or an image and did not significantly differ between sexes ($p > 0.9$,

AUC = 0.43). The mean principal component scores of PC1 for females and males were 0.03 and

-0.03, respectively. The mean PC score difference between females and males ($mean_{(F-M)}$) was

0.07 (95% CI = [-1.70, 1.80], $t(30) = 0.08$). Critically, PC2 mirrored the same procession of

strategies observed in female, but not male mice  (**Figure 2c-d**). This suggests that the extent to

which individuals used this procession of strategies explained a large fraction (22%) of the

interindividual variability in our animals. The mean PC score of PC2 for females was 0.98 and

for males was -0.98. The mean PC score difference between females and males ($mean_{(F-M)}$) was

1.96 (95% CI = [0.87, 3.05], $t(30) = 3.67$). Moreover, knowing the PC2 score (the projection of

an individual animal's behavior onto PC2) allowed us to discriminate male and female animals

with remarkable accuracy (receiver operating characteristic analysis, AUC = 0.86, significant

discrimination: $p < 0.001$). No other PCs differed between sexes ($p > 0.4$, AUC < 0.6). Together,

these results suggest that the choice of what strategy to follow explains substantial individual

variability in multidimensional decision-making, and that differences in strategy can depend on

sex.

There are two possible explanations why females consistently implemented the strategy

procession captured by PC2. One hypothesis is that this early side bias reflected an *energy saving*

strategy that saved time and/or effort by just repeating the same side. Alternatively, this early

side bias could be an *cognitively effortful* strategy to constrain decision-making to one

dimension. These two views make different predictions of the relationship between the use of

strategy procession captured by PC2 and reaction times (RTs), which were defined as the time

between the onset of two visual stimuli and the registration of a nose poke response on one of the

two stimuli. If side bias was an effort saving strategy, then the animals who score highest on PC2

should also make the fastest decisions. On the other hand, if side bias was a cognitively effortful

strategy, the speed of decision-making should be slower in animals who use this strategy. To test

these two hypotheses, correlation analyses were run to assess the relationship between the use of

PC2 strategy and reaction time. PC2 scores were positively correlated with reaction time (**Figure**

**3a**, Spearman's correlation, $r_s = 0.452$, $p = 0.009$; Pearson's correlation, $r = 0.347$, $p = 0.051$),

suggesting that the animals who used the early side bias strategy made slower decisions. This

suggests that this strategy is effortful, rather than  energy saving. There were no significant

correlative relationships between reaction time and other PCs.

If the side bias in females was a cognitively effortful strategy, we would expect to see that

female mice are slowest at making their choices when they are most engaging this strategy:

during early learning. To test this hypothesis, we computed average RTs across 23 bins of 150

trials for males and females. Consistent with our hypothesis, females responded slower during

early learning (bin 1-15) (GLM, interaction term, $\beta3 = 0.03$, $p = 0.0007$) and significantly slower

than males across all trials (GLM, main effect of sex, $\beta1 = -0.62$, $p < 0.0001$). The mean RT

across all trials for males was 1.89 seconds with standard deviation of 0.13, and the group

average for females was 2.04 seconds with standard deviation of 0.21. The reaction time

decreased as the animals ran more trials in both males and females (**Figure 3b**, GLM, main

effect of number of trials, $\beta2 = -0.04$, $p<0.0001$, see equation 1 in methods). Thus, female mice

were slowest during the period in which they were using the side bias strategy the most, again

consistent with the idea that this is a cognitively effortful strategy, rather than an energy saving

one.

**Males varied strategies over time in response to immediate past reward**

Although our analyses captured the procession of strategies employed across essentially all

female mice that learned the task, they provided little insight into what the males were doing.

One likely explanation is that males were more inconsistent than females. Males could be more

inconsistent than females for any one of three reasons: hypothesis 1) males were more *random*

(and thus each choice would be unpredictable within an animal), hypothesis 2) males were more

*idiosyncratic* and less uniform as a group (and thus responses would differ across individuals,

but still be predictable within an individual), or hypothesis 3) that males were more *changeable*

(and thus a given male was predictable in the sense that he was not random, but he was still more

likely to change his strategic approach from one epoch to the next).

To test hypothesis 1 (randomness), we asked whether male choices were more or less predictable

than female choices. We reasoned that if males were just choosing randomly, it would be

impossible to predict their choices. Therefore, within each block of trials in each animal, we

calculated conditional mutual information (*32*, *33*), which quantifies the dependence of current

choices (side, image) on the choice of previous trial, given the outcome of the previous trial. If

the current choice an animal made was random, it would be independent of the choice and

outcome of the previous trial, and we would expect to see low mutual information, shown as

uniform "bands" on the probability heatmap (**Figure 4a**). Conversely, if the current choice was

heavily influenced by the content of the previous trial, we would expect to see high mutual

information, shown in a more checkered or selective pattern on the probability heatmap. We

calculated conditional mutual information for each trial bin across sexes. We found that mutual

information decreased over time in both sexes, reflecting the gradual acquisition of the strategy

of choosing high reward probability cue regardless of the outcome of the previous trial (**Figure**

**4a**, GLM, main effect of number of trials, $\beta 2 = -0.001$, p $=0.0002$, see equation 1&5 in

Methods). However, surprisingly, the mutual information of male mice was *higher* than that of

females (main effect of sex, $\beta 1 = 0.043$, p $< 0.0001$), particularly early in learning (interaction

term, $\beta 3 = -0.002$, p $< 0.0001$). Thus, males were, if anything, less random than females.

One possibility is that the high mutual information in males was a result of increased sensitivity

to the last  reward, meaning that the last reward had a bigger effect on the males' idiosyncratic

decisions about what to do next. To estimate sensitivity to reward, we examined how reward

outcomes affected the reaction time (RT) on the next trial  in both males and females. If the

decision of an animal was not affected by the outcome of the past trial, then we would expect to

see no difference between reaction time for last rewarded and last unrewarded trial ($RT_{reward} -$

$RT_{no\ reward} = 0$). Males responded significantly faster when they had just received a reward from

the previous trial (**Figure 4b and 4c,** one-sample t-test, mean RT effect $= -0.14$, 95% CI $= [-$

0.23, -0.05], t(15) $= -3.38$, p $= 0.004$). Conversely, the reaction times of females were not

affected by the outcome of the last trial (one-sample t-test, mean RT effect $= -0.03$, 95% CI $= [-$

0.13, 0.06], t(15) $= -0.75$, p $= 0.47$). These results again suggested that female decisions were not

affected on a trial-by-trial basis by the outcome of each trial because they followed a global

strategy while male choices that were heavily affected by recent rewards.

Although males were, if anything, less random in their decision-making than females, it

remained possible that the apparent lack of local strategies occurred because male strategies were

inconsistent--either because of idiosyncratic differences between males (hypothesis 2) or

changeability within males (hypothesis 3). To do this, we developed a technique for comparing

how similar one set of choices was to other set of choices. We expressed the choices in each bin

as a probability vector, with each element of the vector reflecting the probability of that unique

combination of behaviors {last choice, last outcome, this choice}. The average angle between

any two of probability vectors across animals, trial bins, or image pairs is then a measure of the

variability in choices between those two conditions. Males were not more idiosyncratic than

females on a population level; that is, the choices of any male were not more variable from other

males than any female's choices were from other females (**Figure 4d**, GLM, main effect of sex,

$\beta 1 = -1.47$, p = 0.11, see equation 1 in Methods). However, the males were more variable *within*

*themselves,* both across bins within one image pair (**Figure 4e**, GLM, main effect of sex, $\beta 1 =$

4.24, p < 0.0001, see equation 1 in Methods) and across multiple image pairs (**Figure 4f**, GLM,

main effect of sex, $\beta 1 = 4.54$, p = 0.047, see equation 1 in Methods). Overall, the variability in

choices decreased across time, as the divergent strategies used by individual animals started to

converge to the optimal strategy in this task, which is to choose the high-value image

consistently (GLM, main effect of number of trials, within sex between suject: $\beta 2 = -0.78$, p <

0.0001; within subject between bins: $\beta 2 = -0.359$, p < 0.0001). Together, these results suggest

that individual males displayed divergent choice patterns and were changing between complex

strategies over time and the repetition of the same task, while females largely adopted a shared,

systematic approach to learning.

To visualize animals' patterns of choices expressed in the probability vectors, we used

multidimensional scaling (MDS) (*34–36*) to reduce the dimensionality of strategy space,

allowing us to project the high-dimensional "strategy path" throughout learning onto a 2

dimensional space. This allows us to easily visualize the similarity between patterns of choice

across animals over time and across repetitions. Each color path represents a strategy path used

in a different repetition of the task (4 repetitions in total). **Figure 4g** shows examples of strategy

paths of males and females. The optimal strategy in this bandit task, which is to choose the high

value image consistently regardless of the outcome, is represented by a star in the low

dimensional space. Both males and females "strategy path" showed gradual approximation to the

optimal strategy over time. Consistent with the quantification described above, the strategy path

of males are visibly more variable and different across repetitions of the task, whereas the

strategy path of females were more unified and consistent across repetitions.

**Sex mediated the ability of neuronal activity to explain strategy selection**

The ability to learn and perform bandit tasks is highly sensitive to alterations in corticolimbic

structures. However, it remains unclear how alterations in these structures predict choice

strategy, much less sex differences in choice strategy. To address this question, we examined

neuronal activity in several corticolimbic brain regions by examining the expression of *c-fos,* an

immediate early gene often used as a marker of neuronal activation. The animals from the

previous figures were sacrificed after the second day of a new, final image-reward pairing (each

animal having completed 400-500 total trials), corresponding to when the female side bias was

greatest. We compared mRNA expression level for *c-fos* across five brain regions, including

nucleus accumbens (NAc), dorsal medial striatum (DMS), amygdala (AMY), hippocampus

(HPC), and prefrontal cortex (PFC), using quantitative real-time PCR (**Figure 5a**). In each of the

five brain regions, females had a higher c-fos expression level than did males (unpaired t-test,

NAc: $t(30) = 2.41$, $p = 0.02$; DMS: $t(30) = 2.31$, $p = 0.03$; AMY: $t(30) = 4.05$, $p < 0.001$; HPC:

$t(30) = 2.74$, $p = 0.01$; PFC: $t(29) = 3.163$, $p = 0.003$).

To understand whether activation of any of these brain resions correlated with the side bias

strategy, we constructed a GLM to predict PC2 from c-fos expression level in each brain region

and sex. The results suggested that only two regions, the  NAc and PFC predicted strategy use, as

indexed by PC2 score  (**Figure 5b**, GLM, NAc: $\beta_1 = 0.72$, $p = 0.02$; DMS: $\beta_2 = 0.48$, $p = 0.14$;

AMY: $\beta_3 = 0.52$, $p = 0.10$; HPC: $\beta_4 = 0.55$, $p = 0.08$; PFC: $\beta_5 = 0.75$, $p = 0.02$; sex was included

as a variable in the model and was also significant: $\beta_6 = 0.99$, $p = 0.0009$, see equation 2 in

Methods). Because each region was also correlated with sex to differing extents (and sex

independently predicted PC2), we next asked whether NAc and PFC were the best predictors of

PC2 because these regions were the most strongly correlated with sex (**Figure 5c**). However,  the

predictive effect of NAc and PFC c-fos expression was not because NAc and PFC were the most

highly with sex. Instead, sex was most strongly correlated with AMY, which was not a

significant predictor of PC2. To confirm that these correlations between regional activation and

early side bias strategy was meaningful, we fit the same GLM to predict PC1, and none of the

predictor variables were significant. We confirmed these results with a Pearson product-moment

correlation coefficient, which again suggested a significant positive correlation between c-fos

expression in NAc/PFC and PC2 scores (**Figure 5d**, NAc: $r = 0.40$, $n = 32$, $p < 0.03$; PFC: $r =$

$0.41$, $n = 32$, $p < 0.02$; averages across a median split of PC2 within each sex are illustrated in

**Figure 5e**; main effect sex: NAc: $F(1,28) = 12.87$, $p = 0.001$; PFC: $F(1,28) = 13.47$, $p = 0.001$).

Next, we asked whether an animals' sex altered the relationship between NAc and PFC c-fos

activity PC2 scores. To do this, we used a structural equation modeling (SEM) approach (*37, 38*)

to analyze the structural relationship between sex, gene expression, and PC2 and latent

constructs (**Figure 5f**). First we used a direct model and regressed c-fos expression of either NAc

or PFC on strategy selection, both NAc and PFC were significant direct predictors of PC2 scores

(NAc: $\beta = 0.72$, $p = 0.022$; PFC: $\beta = 0.75$, $p = 0.019$, see equation 6 in Methods). Then we fit a

mediation model that allows us to understand how sex influences neural activation in NAc and

PFC, which in turn influences strategy selection. Regressing the mediator variable sex on c-fos

expression in NAc/PFC confirmed that neural activation is a significant predicor of the mediator

sex (NAc: $\alpha = 0.20$, $p = 0.024$; PFC: $\alpha = 0.26$, $p < 0.004$, see equation 7 in Methods). When we

regressed strategy selection on both the mediator variable (sex) and independent variable (neural

activation in NAc/PFC), the result showed that the mediator sex was a significant predictor of

strategy selection (NAc: $\beta 1 = 1.66$, $p = 0.008$; PFC: $\beta 1 = 1.64$, $p = 0.012$), and the strength of the

direct model is now greatly reduced and became non-significant when accounted for the

mediating effect of sex (NAc: $\beta' = 0.38$, $p = 0.20$; PFC: $\beta' = 0.33$, $p = 0.31$). The Sobel (1982)

first-order test was used to assess the presence of mediation (*38*). The indirect effect was

calculated as the product of coefficients and was significant for both NAc and PFC (NAc: $\alpha\beta' =$

0.34, $z = 1.836$, $p < 0.039$; PFC: $\alpha\beta' = 0.42$, $z = 2.035$, $p < 0.026$). Together, these results

suggest that the relationship between PFC and NAc c-fos and PC2 differed, depending on the

animals' sex. This suggests that sex-linked mechanisms gate the relationship between these

circuits and strategic decision-making and highlight these regions as promising targets for future

studies looking at the effects of sex on the neural circuits responsible for implementing strategic

learning.

**Discussion**

By training male and female mice on a stochastic two-dimensional decision-making task, we

were able to evoke a range of problem solving strategies across individuals. In this task, each cue

has two dimensions - the identity of the image and the location of the image. Animals had to

explore the reward value associated with both cue dimensions to determine which were most

predictive of reward. Although both male and female mice eventually learned the right strategy,

choosing the high-value image, female mice learned faster. The richness of this task allowed us

to uncover sex differences in *how* the animals achieved the associations across time. We

discovered that female mice adopted a consistent and systematic approach where they processed

through different strategies over time. Early in learning, they constrained their search space by

only sampling the outcomes of images on one side (left or right). This approach, which occurred

when animals were most uncertain about the best choice, reduced the number of dimensions they

were learning about and permitted more rapid acquisition of the image-value association. In

contrast, males employed a strategy of decisions that seemed to combine both image and spatial

location, changed frequently,  and was strongly influenced by the immediate prior experience of

reinforcement. While both sexes eventually reached equivalent levels of performance, our data

reveal that the journeys individual animals take to get there can vary dramatically, implicating

the potential for wide divergence in neural circuit mechanisms in normal decision making.

One fundamental unanswered question is *why* females as a group employed a highly similar and

consistent strategy. Zador (2019) recently proposed that much of animal behavior is not dictated

by supervised or unsupervised learning algorithms, but are largely innate, shaped by biological

constraints (*39*). The biological constraints and organization imposed by the multiple

mechanisms of sexual differentiation are known to drive a tuning of the circuits important for

reward-guided decisions (*40–44*). Among these are well-known effects of testosterone in driving

exploratory and impulsive behavior (*43*, *45*, *46*). Conversely, energy-conserving and habitual

behaviors are more prevalent in female animals, including during foraging (*1*, *2*, *47*). Gonadal

hormones, such as ovarian hormone estradiol (E2), are thought to exert modulatory control over

cost/benefit decision-making that increased E2 resulted in reduction of high-effort choices (*5*, *48*,

*49*). In addition, dissociable impacts of sex chromosomes on reward-guided behaviors (*50*) that

have been described as promoting habit in XX carriers and increased effort in XY carriers. Of

course, most impacts of sexual differentiation are graded, rather than dichotomous across the

sexes. Indeed, here we found that a small number of males showed some tendency to use the

female strategy, implicating graded mechanisms of masculinization in sex-gated strategy

selection. Previous evidence of sex-differences in decision-making has been  interpreted as

evidence that females avoid unnecessary effort in the pursuit of food. However, our results

suggest that this strategy may be effortful, not effortless. Further, in many circumstances, it may

be a better strategy than the male pattern of indiscriminate exploration. Thus, it is possible that

these effects are due to differences in the behavioral ecology of male and female animals, which

creates different biological constraints on learning across sexes.

What kinds of environments might advantage the typical female strategy over the typical male

strategy or vice versa? Our result suggests that the systematic approach employed by female

mice may have greater success with high-dimensional learning tasks. By reducing or eliminating

choices in one dimension, females decreased the number of dimensions that varied at a time, and

were able to learn other associations more quickly. However, in sparser environment, a strategy

of exhausting all options in one dimension could become less effective. Having a range of

divergent strategies and changing choice patterns in response to the reward outcome, as seen in

male mice, may lead to a greater chance of success in varying and volatile environments (*17*, *51*,

*52*) at the cost of greater risk to an individual male.  Indeed, it is possible that these differences in

the match between sex-specific strategies and the environment may be a major contributing

factor in the inconsistent gender and sex differences across tasks with different levels of

volatility (*1*, *2*). An intriguing possibility is that the unified, consistent, and systematic strategy

we observed in female mice, as well as the volatile and diverse strategy we observed in male

mice in the same task, may emerge from evolved sex-biased strategies for foraging in the wild

that were critical to survival for the species as a whole, by dividing risk and reward across the

population.

Sequential decision-making and learning is often studied with spatial bandit tasks, in which

reward probabilities are linked to left and right levers or sides that are visually identical (*17*, *18*,

*25*, *53–55*), particularly in rodent models. In these spatial bandit tasks, side bias in choice has

sometimes been reported in rodent operant work as a behavioral artifact and animals displaying

such bias were often excluded from experiments (*56–58*). However, in the current task, both the

side and the identity of the image cues could have been informative of the reward probabilities.

In principle, animals could simultaneously sample both dimensions to learn side values and

image values at the same time. However, in practice, it appears that the early side bias in female

mice "jump-started" their learning by controlling for space while exploring choice-outcome

values of the images, which in this task happened to be the more informative dimension of

reward. Intriguingly, this suggests that females were covertly learning about the correct cue

dimension while behaviorally selecting the wrong item, and were able to convert this to

successful learning due to the stability of the task structure. This view suggests that , we should

be able to design tasks that prevent the successful use of this strategy, and which might therefore

shift the presentation of the sex difference in decision-making.

Our data implicate the prefrontal cortex (PFC) and nucleus accumbens (NAc), part of the ventral

striatum, in the differences in strategy between males and females. These regions have been

widely implicated in reward-guided decision making, but so have the other regions we tested for

which we didn't find a significant relationship to behavior (*11*, *12*, *15*). One possibility is that the

PFC and accumbens are particularly engaged in strategic decision-making. This resonates with

previous studies that have implicated the PFC in implementing strategies and rule-guided

behaviors (*51*, *55*, *59–63*) and the NAc in selecting and implementing learning strategies (*12*,

*14*). Implementing different strategies produces changes in how different choice dimensions are

represented in the PFC and  NAc (*64*), and lesions in the NAc can drive animals towards a low-

dimensional action-based strategy or prevent animals from switching between strategies (*11*, *14*).

The PFC is also sensitive to gonadal hormones during risky decision making (*65*), and

dopaminergic function in the accumbens regulates risky decision making in a sex-specific

manner (*66*), perhaps due to sex differences in dopamine neurons (*44*). Here, the relationship

between both PFC and NAc and strategy use was mediated by sex, suggesting that whatever the

relationship between these regions and strategic decision-making, it is likely to be sex-specific.

**Methods**

Animals. Thirty-two BL6129SF1/J mice (16 males and 16 females) were obtained from Jackson Laboratories (stock #101043). Mice arrived at the lab at 7 weeks of age, and were housed in groups of four with ad libitum access to water while being mildly food restricted (85-95% of free feeding weight) for the experiment. Animals engaging in operant testing were housed in a 0900–2100 hours reversed light cycle to permit testing during the dark period, between 09:00 am and 5:00 pm. Before operant chamber training, animals were food restricted to 85%-90% of free feeding body weight and had been pre-exposed to the reinforcer (Ensure). Pre-exposure to the reinforcer occurred by providing an additional water bottle containing Ensure for 24 hours in the home cage and verifying consumption by all cagemates. Operant testing occurred five days per week (Monday-Friday), and the animals were fed after training with ad lib food access provided on Fridays. All animals were cared for according to the guidelines of the National Institution of Health and the University of Minnesota.

Apparatus. Sixteen identical triangular touchscreen operant chambers (Lafayette Instrument Co., Lafayette, IN) were used for training and testing. Two walls black were acrylic plastic. The third wall housed the touchscreen and was positioned directly opposite the magazine. The magazine provided liquid reinforcer (Ensure) delivered by a peristaltic pump, typically 7ul (280 ms pump duration). ABET-II software (Lafayette Instrument Co., Lafayette, IN) was used to program operant schedules and to analyze all data from training and testing.

Operant Training

*Pretraining.* animals were exposed daily to a 30-min session of initial touch training, during

which a blank white square (cue) was presented on one side of the touchscreen, counterbalancing

left and right between trials. This schedule provided free reinforcement every 30 seconds, during

which the cue was on. If animals touched the cue during this period,  a reward three times the

size of the regular reward was dispensed (840 ms). This led to rapid acquisition. Following this,

animals were exposed daily to a 30-min session of must touch training. This schedule followed

the same procedure as the initial touch training, but free reinforcers were terminated and animals

were required to nose poke the image in order to obtain a regular reward (7-uL, 280 ms).

*Deterministic pairwise discrimination training.* Animals were exposed to 10 days of pairwise

discrimination training, during which animals were presented with two highly discriminable

image cues ("marbles" and "fan"). One image was always rewarded and the other one was not.

Within each session, animals completed either 250 trials or spent a maximum of two hours in the

operant chamber (typically these mice completed ~200 trials/day).

*Two-armed bandit task.* Animals were trained to perform a two-arm visual bandit task in the

touchscreen operant chamber. Each trial, animals were presented with a repeating set of two

different images on the left and right side of screen, counterbalancing left and right across the

session. Nose poke to one of the displayed images on the touchscreen was required to register a

response. Nose poke on one image triggered a reward 80% of the time (high payoff image),

whereas the other image was only reinforced 20% of the time (low payoff image). Following the

reward collection, which was registered as entry and exit of the feeder hole, the magazine would

illuminate again and the mouse must re-enter and exit the feeder hole to initiate the next image

trial. If the previous trial was unrewarded, a 3-second time-out was triggered, during which no

action could be taken. Following the timeout, the magazine would illuminate and the mouse must

enter and exit the feeder hole to initiate the next image trial. The ABET II system recorded trial

to trial image chosen history, reward history, grid position of the images with time-stamp. Within

each day, animals completed either 250 trials or spent a maximum of two hours in the operant

chamber. Animals were given 14 days to learn about the probabilistic reward schedule of one

image pair, before moving onto the next image pair. A total of six image pairs were trained, but

two image pairs were eliminated from analyses due to very high initial preference (>70%) for

one novel image over another, indicating that (to the mice) these images appeared unexpectedly

similar to previously experienced images with learned reward values.

*RNA quantification.* At the end of training, animals were sacrificed after the second day of

learning a new image pair (around 400-500 trials of experience per mouse), when we expected to

see the biggest difference in learning performance and strength of lateralization. Animal brains

were extracted and targeted brain regions were dissected. We extracted RNA from targeted brain

areas and assessed gene expression for the *fos* genes in the nucleus accumbens (NAc), dorsal

medial striatum (DMS), amygdala (AMY), and hippocampus (HPC), using quantitative Real

Time PCR system (BioRad, USA). Fos expression normalized to the housekeeper gene

glyceraldehyde 3-phosphate dehydrogenase (*gapdh*) was calculated using the comparative delta

Ct method.

Data analysis

*Generalized Linear Models (GLMs).* In order to determine whether sex and number of trials

(bins) predicts the accuracy of the task, strength of lateralization, reaction time, mutual

information (MI), or angle between probability vectors, we fit a series of generalized linear

models of the following form:

$$Y = \beta_0 + \beta_1(sex) + \beta_2(trials) + \beta_3(sex)(trials)$$    [1]

Where Y is the dependent variable (accuracy, laterality, reaction time, MI, or angle). In this

model, β1 described the main effect of sex and β2 described the main effect of number of trials

(bins). β3 captures any interaction effect between sex and number of trials (bins).

To determine whether c-fos expression in NAc, DMS, AMY, HPC, PFC, and sex predicted the

weights of Principal Component (PC) 2, we fit the following generalized linear model.

$$PC2 = \beta_0 + \beta_1(NAc) + \beta_2(DMS) + \beta_3(AMY) + \\ \beta_4(HPC) + \beta_5(PFC) + \beta_6(sex)$$    [2]

In this model, β1- β5 captures the predictive effect of gene expression in five regions on the use

of PC2 strategy. β6 described the effect of sex on the weights of PC2.

*Degree of lateralization.* As a measure of the strength of side bias, we used the absolute

percentage of laterality (*31*), calculated for each mouse according to the following formula:

$$Degree\ of\ laterality = \left| \frac{right - left}{right + left} \right|$$    [3]

*Generalized Logistic Regression Model.* Mice could base their decisions on reward history in the

spatial or image domains or on choice history in the spatial or image domains. To determine how

these four aspects of previous experience affected choice and how these effects changed over

time, we estimated the effect of the last trials' reward outcome (O), image choice (I), and chosen

side (S) using logistic regression. If image (image 1) was on the left side of the screen, we could

predict the probability of choosing that image as a linear combination of the following four

terms:

$$log(\frac{p(I_{1,t})}{p(I_{2,t})}) = \beta_0 + \beta_1 * (I_{1,t-1} - I_{2,t-1}) +$$

$$\beta_2 * O_{t-1} * (I_{1,t-1} - I_{2,t-1}) +$$
$$\beta_3 * (S_{L,t-1} - S_{R,t-1}) +$$
$$\beta_4 * O_{t-1} * (S_{R,t-1} - S_{L,t-1}) \qquad\qquad [4]$$

Where each term (O, I, and S) is a logical, indicating whether or not that event occurred on the

last trial. As a result, the term (I_{1, t-1} - I_{2, t-1}) is 1 if image 1 was chosen on the last trial,

but -1 otherwise. The term $\beta 1$ thus captures the tendency to either repeat the previous image

(when positive) or choose the other image (when negative). The term $\beta 2$ O_{t-1} accounts for

any additional effect of the previous image on choice, when that previous choice was rewarded..

If image 1 was on the left side, (S_{L, t-1} denotes the probability of repeating the left side

where image 1 appeared. However, because image 1 could be either on the left or the right side

of the screen (which allowed us to dissociably estimate the probability of choosing it based on

side bias or image bias), we expanded the (S_{L, t-1} - S_{R, t-1}) term to account for the

current position of image 1 as follows:

$$...((I_{1,t} = L)(S_{L,t-1} - S_{R,t-1}) + (I_{1,t} = R)(S_{R,t-1} - S_{L,t-1}))...$$

Meaning that the current position of image 1 determined the sign of the side bias term. This

model was fit individually to each bin of 150 trials, within each animal and image pair, via cross-

entropy minimization with a regularization term (L2/ridge regression).

*Principal component analysis.* In order to determine how decision-making strategies differed

across animals and bins, we looked for the major dimensions of interindividual variability in

decision-making strategies. To do this, we took advantage of the fact that the coefficients of the

generalized linear model provided a simplified description of how decision-making depended on

image, side, and outcome for each subject in each bin. Because the generalized logistic

regression model estimated 4 terms per image pair and there were 23 independent bins per image

pair, this meant that each animals' behavior for a given image pair could be described as a 4*23

by 1 dimensional vector. We then used principal component analysis to identify the linear

combinations of model parameters that explained the most variance across subjects and

repetitions of image pairs (across 32 (animals) x 4 (repetitions) = 128 total strategy vectors). The

first two principal components, which explained the majority of the variance (59%), are

illustrated in Figure 2e.

*Conditional mutual information and model-free analyses.* To account for idiosyncratic strategies,

which could vary across animals or image pairs, we used a model-free approach to quantify the

extent to which behavior was structured without making strong assumptions about what form

this structure might take. We quantified the extent to which choice history was informative about

current choices as the conditional mutual information between the current choice (C) and the last

choice ($C\_t-1$), conditioned on the reward outcome of the last trial (R):

$$I(C_t; C_{t-1}|R) = \sum_{r \in R} \sum_{c_{t-1} \in C} \sum_{c_t \in C} P_{C_t, C_{t-1}, R}(c_t, c_{t-1}, r) log \frac{P_R(r) P_{C_t, C_{t-1}, R}(c_t, c_{t-1}, r)}{P_{C_t, R}(c_t, r) P_{C_{t-1}, R}(c_{t-1}, r)}$$

579            [5]

Where the set of choice options (C) represented the unique combinations of each of the 2 images

and 2 sides (4 combinations). To account for observed differences in overall probability of

reward for male and female animals, the mutual information was calculated independently for

trials following reward delivery and omission, and then summed across these two conditions.

We used a similar approach to provide a model-free description of the animals' choice patterns.

Briefly, instead of finding the set of beta weights that best described reliance on various history-

dependent strategies over time, we directly calculated the joint probability of each possibility

combination of last choice (image and side), last outcome (reward and unrewarded), and current

choice (image and side). This means that we represented the animals' history-dependent choice

pattern for each image pair as an 32-dimensional vector (4 (last choice) x 2 (last outcome) x 4

(current choice) = 32) of joint probabilities. Via a geometric interpretation of a multinomial

distribution, we considered the animal's pattern of behavior within any bin of trials as a point on

the 32-1 dimensional simplex formed by length-1 vectors. This geometric approach allowed us to

map strategies over time or across bins as a diffusion process across this simplex, where the

angle between two vectors (between animals/between bins/between repetitions) is proportional to

step between them on a strategy simplex. The bigger the step between two vectors, the more

variable the behavior pattern is.

*Mediation Analysis.* First we used a direct model and regressed c-fos expression of either NAc or

PFC on weights of PC2. When assessing a mediation effect, three regression models are

examined:

Model 1 (direct):

$$PC2 = \gamma_1 + \beta(NAc) + \epsilon1$$    [6]

Model 2 (mediation):

$$Sex = \gamma_2 + \alpha(NAc) + \epsilon2$$    [7]

Model 3 (indirect)

$$PC2 = \gamma_3 + \beta'(NAc) + \beta_1(sex) + \epsilon3$$    [8]

In these models, $\gamma1$, $\gamma2$, and $\gamma3$ represent the intercepts for each model, while $\epsilon1$, $\epsilon2$, and $\epsilon3$

represent the error term. $\beta$ denotes the relationship between dependent variable (PC2 weights)

and independent variable (NAc c-fos expression) in the first model, and $\beta'$denotes the same

relationship in the third model. $\alpha$ represents the relationship between independent variable (NAc

c-fos expression) and mediator (sex) in the second model. The mediation effect is calculated

using the product of coefficients ($\alpha\beta1$). The Sobel test is used to determine whether the

mediation effect is statistically significant (*38*).

*Reinforcement Learning Model.*

We considered a basic reinforcement learning model, following the Rescorla Wagner rule. In this

model, subjects first learn the expected value of each image based on the history of its previous

outcome value Q and use these Q values to decide what to do next. The expected value of arm $k$

on the $t$ th trial, $Q_t^k$, is updated based on the reward outcome of each trial:

$$Q_{t+1}^k = Q_t^k + \alpha(r_t - Q_t^k)$$

In each trial, $r_t - Q_t^k$ captures the reward prediction error (RPE), which is the diffe-rence

between expected outcome and the actual outcome. The parameter $\alpha$ is the learning rate, which

determines the rate of updating RPE. Action selection was performed based on a Softmax

probability distribution:

$$P(a_{t+1} = k) = \frac{e^{\beta Q_t^k}}{\sum_j e^{\beta Q_t^j}}$$

where inverse temperature $\beta$ determines the level of random exploration.

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**Figures and Figure Captions**

**Figure 1**. **Females showed accelerated acquisition of the high reward probability image in a**

**stochastic two-armed visual bandit task.** A) Schematic of the mouse touch-screen operant

chamber used in our task. B) Schematic of two-armed Visual Bandit task. Images varied between

the two locations across trials. The reward probabilities for two images are 80% and 20%,

respectively. C) Average learning performance (percent correct) across four repetitions of the

task with four pairs of images. While both males and females reached the same final

performance, females displayed an accelerated learning curve. D) No sex difference in learning

performance was observed in deterministic reward schedule (100%/0%). E) Females displayed

stronger side bias for item selection on the touchscreen, regardless of the direction of

lateralization, early on in learning. This behavior lateralization disappeared as female mice

learned the task. Data shown as bins of 150 trials. * indicates $p < 0.05$. Bars ± SEM. N=16/sex,

wildtype F1 strain B6129SF1/J.

**Figure 2. Female mice use a procession of strategies, initially using a spatial bias followed**

**by a switch to responding based on image domain.** A) Schematic of four basic local strategies

based on choice and reward history of image and spatial dimensions of the task. B) A

generalized logistic regression model revealed a global strategy - a clear procession of four local

strategies - that mice started repeating one side before switching to choosing the reinforced

image. C) female mice displayed more pronounced global strategy procession from spatial-based

strategy to image outcome-based strategy. D) male mice displayed increased image win-stay

lose-shift over time but no prominent global strategy in the early learning stage was observed. E)

A principal component analysis (PCA) was conducted on the estimates of global strategy

strength over time across all animals regardless of sex. Principal component (PC) 1 and 2

accounted for about 60% of the variance. PC 1 described a general preference for responding

based on image value and did not differ between sexes. PC 2 captured the same global strategy

procession reflected in the generalized logistic model - a strong contribution of the "side repeat"

behavior early in training, followed by a rapid transition to "image outcome", indicative of a

sudden shift away from "where" and towards "what" in solving the task. Projecting each animal

onto this PC 2 showed a clear separation between the sexes (blue male, pink female), AUC =

0.86, p < 0.001. This suggests that the strategy procession from spatial repeat to image outcome

is a female-specific strategy. Note that a few males are positive for Principal Component 2, and

their individual behavior supports that these males also employed this strategy to a weaker

extent. In contrast, the few females that are negative for this Principal Component did not show

evidence of having learned the task. Data shown as bins of 150 trials. Bars ± SEM.

**Figure 3. The side-to-image strategy procession captured by Principal Component (PC) 2 is**

**a cognitively demanding strategy but not a time-saving strategy.** A) Correlation analyses

revealed a significant positive correlation between PC2 scores and reaction time. The decision-

making time was longer within animals primarily used PC2 strategy. B) Predominantly using

PC2, females responded slower during early learning (GLM, interaction term, $\beta3 = 0.03$, p =

0.0007). Note that, the slow reaction time during early learning in females matches with the time

period (bin 1-10) during which females relied on side-bias "heuristics" for decision-making.

Data shown as bins of 150 trials. * indicates $p < 0.05$. Bars ± SEM.

**Figure 4. Male mice were more likely to differ from themselves over time, with choice patterns dependent on past outcomes.** We expressed each animal's history-dependent choice pattern as an 32-dimensional vector of joint probabilities and measured the angle between vectors, which is proportional to the step between them on a strategy simplex. A) Illustration of choice patterns of low mutual information and high mutual information. If choice on trial $t$ is independent of choice on the previous trial ($t-1$), probability heatmap should show band-like pattern (choosing the same choice regardless of the previous choice). Conversely, high mutual information has more checkered choice patterns. Conditioned mutual information is higher in males, indicating that responses are more uniquely affected by the previous trial variables than

they are in females. B) One-sample t-test was conducted across bins to compare the difference in

reaction time (RT) between rewarded and unrewarded trials to 0 (when there is no effect of past

outcome on the reaction time). Male mice have significant RT effects on the last reward. There

was no difference in reaction time between rewarded and unrewarded trials in female mice. The

bins marked by asterisks have $p < 0.05$ for the one-sample t-test. C) average RT effect of last

reward across all trials. Overall, male responded faster when the last trial was rewarded than

unrewarded. D). Males and females were equally variable between animals within sex. An

individual male is no more different from other males in behavior than a female is from other

females. E) Choice patterns of a given male compared to himself over bins of 150 trials were

more variable than in a given female compared to herself . F) Choice patterns of a given male to

himself were more variable and divergent across repetitions of the same task than in females

compared to themselves across repetitions. G) Multidimensional scaling (MDS) was used to

visualize animal's strategy path across trials and repetitions by reducing the dimensionality of

the strategy space. Each of the four colors within one sex represents one repetition of the task.

The star represents the point of optimal strategy for this task, which is to choose the high reward

probability image. In both males and females, the strategy paths showed a gradual approach to

the optimal strategy point, indicating that both sexes were able to learn the optimal strategy. The

strategy paths of females are consistent and similar across repetitions, suggesting that female

mice used a similar strategy every time to solve the problem. On the other hand, in male mice,

the steps between each bins of 150 trials in the strategy space were larger, suggesting higher

variability in choice patterns. Thus, male mice used divergent strategies throughout learning and

used different approaches each time to learn the same task. Data shown as bins of 150 trials. *

indicates $p < 0.05$. Bars ± SEM.

**Figure 5. Both sex and neuronal activity can account for strategy selection, but sex**

**mediated the ability of neural activity to explain strategy selection.**

**A**) cFos gene expression (qRT-PCR) in five brain regions: nucleus accumbens (NAc), dorsal

medial striatum (DMS), amygdala (AMY), hippocampus (HPC), and prefrontal cortex (PFC).

Female mice showed elevated c-fos expression across all five brain regions. Asterisks marked

significant difference between sexes (*: $p < 0.05$ **: $p < 0.01$ ***: $p<0.001$). **B**) Heatmap of

correlation matrix of c-fos expression level among five brain regions. **C**) c-fos expression in

NAc and PFC, and sex, predict the use of PC2 strategy (GLM, NAc: $\beta_1 = 0.72$, $p = 0.02$; PFC:

$\beta_5 = 0.75$, $p = 0.02$; sex: $\beta_6 = 0.99$, $p = 0.0009$). Asterisks marked significant beta weights ($p <$

$0.05$). **D**) cFos gene expression in NAc and PFC is significantly correlated with the weight of

PC2. **E**) The use of PC2 strategies procession was analyzed with a 2 (sex: male versus female) x

2 (c-fos expression quartile in NAc/PFC: bottom versus top) between-subjects ANOVA. The

main effect of sex was significant for both NAc and PFC (NAc: $F_{(1,28)} = 12.87$, $p = 0.001$;

PFC: $F_{(1,28)} = 13.47$, $p = 0.001$). F) Causal modeling of the relationship between gene

expression level in NAc and PFC and the weight of PC2. The models on top are direct models,

indicating that c-fos expression levels in both NAc and PFC are significant predictors of the

weights of PC2. The bottom models are the mediation models, in which sex mediated the

relationship between neural activity (c-fos expressions in NAc and PFC) and strategy selection

(weights of PC2). The arrows are regressions. Paths are labeled with estimated coefficients and

significant coefficients are marked by asterisks. The strength of the direct model is greatly

reduced and became non-significant when accounted for the mediating effect of sex. This

suggests that sex mediated neural measures in explaining strategy selection. Bars ± SEM.

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