The geometry and genetics of hybridization

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December 2, 2019

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Running title:

Genetics, geometry and hybridization

Key words:

Fisher's geometric model, quantitative genetics, bounded hybrid superiority, line crosses, speciation

Abstract

We develop an analytical framework for predicting the fitness of hybrid genotypes, based on Fisher's geometric model. We first show that all of the model parameters have a simple geometrical and biological interpretation. Hybrid fitness decomposes into intrinsic effects of hybridity and heterozygosity, and extrinsic measures of the (local) adaptedness of the parental lines; and all of these correspond to distances in a phenotypic space. We also show how these quantities change over the course of divergence, with convergence to a characteristic pattern of intrinsic isolation. Using individual-based simulations, we then show that the predictions apply to a wide range of population genetic regimes, and divergence conditions, including allopatry and parapatry, local adaptation and drift. We next connect our results to the quantitative genetics of line crosses in variable or patchy environments. This relates the geometrical distances to quantities that can be estimated from cross data, and provides a simple interpretation of the "composite effects" in the quantitative genetics partition. Finally, we develop extensions to the model, involving selectively-induced disequilibria, and variable phenotypic dominance. The geometry of fitness landscapes provides a unifying framework for understanding speciation, and wider patterns of hybrid fitness.

Introduction

When genetically distinct populations meet and mate, their divergent alleles are brought together in new combinations. The fitness of these novel genotypes will influence the outcome of the hybridization, and might be a source of information about the divergence history of the populations (e.g., Dobzhansky 1937; Lynch 1991; Rundle and Whitlock 2001; Welch 2004; Gavrilets 2004; Wade and Demuth 2005; Rosas et al. 2010; Fraïsse et al. 2016; Yamaguchi and Otto 2019).

To predict and interpret data from hybrids, various approaches have been used. One approach uses fitness landscapes (Dobzhansky 1937; Hill 1982; Orr 1995; Gavrilets 2004). Here, strong insights can come from simple models, with a few, biologically meaningful parameters; but such models often apply to a limited range of cases, and can be difficult to fit to data. A second approach uses quantitative genetics, applying to line crosses, statistical tools that were developed for single populations (Cockerham 1980; Hill 1982; Lynch 1991; Lynch and Walsh 1998, Chs. 9-10; Rundle and Whitlock 2001; Wade and Demuth 2005). This approach is fully general and widely applied, but the key quantities (the "composite effects"), can be difficult to interpret (this is in contrast to the quantitative genetics of single populations, where a large body of theory can help us to understand the variance components; Hill et al. 2008; Mäki-Tanila and Hill 2014; Barton 2017; Walsh and Lynch 2018).

Previous authors have gained fresh insights by combining these approaches (Lynch 1991; Wade and Demuth 2005; Yamaguchi and Otto 2019). Here, following these authors, we draw an explicit connection between the quantitative genetics of line crosses (Hill 1982; Rundle and Whitlock 2001), and a class of fitness landscapes based on Fisher's geometric model (Fisher 1930, Ch. 2).

Fisher's geometric model has been widely used to study single populations (Orr 1998; Hartl and Taubes 1998; Walsh and Lynch 2018, Ch. 27), and hybridization, and it can account for a wide variety of empirical patterns (Mani and Clarke 1990; Barton 2001; Rosas et al. 2010; Chevin et al. 2014; Fraïsse et al. 2016; Simon et al. 2018; Thompson et al. 2019; Yamaguchi and Otto 2019). But despite these successes, some serious doubts remain. First, the model assumes that fitness is determined by a few quantitative traits, each with additive genetics. This assumption is less restrictive than it seems, because the "traits" need not be identified with standard quantitative traits (such as height or weight), but might emerge from a variety of more complex underlying phenotypic models (Martin 2014; Schiffman and Ralph 2017; Fraïsse and Welch 2019). Nevertheless, additivity leads to some questionable predictions, especially for the initial F1 cross (Fraïsse et al. 2016). Second, while Fisher's model is closely linked with quantitative genetics (Fisher 1930 Ch. 2; Barton 1989; Orr 1998; Rockman 2012), a formal connection cannot be made without analytical approximations (Simon et al. 2018), and it is not clear how widely these approximations apply.

For these reasons, the current paper is in three parts. In Part I, we introduce the analytical predictions of Fisher's model, focussing on the geometrical and biological meanings of its parameters. We also compare the approximations to individual-based simulations, including a wide range of assumptions about environmental change, population demography, and the population genetic regime. In Part II, we connect Fisher's model to quantitative genetics, showing how the composite effects correspond neatly to the parameters of Fisher's model. We then express results for standard line crosses in different environments, unifying results from previous studies (Wright 1922; Lynch 1991; Hatfield and Schluter 1999; Rundle and Whitlock 2001; Chevin et al. 2014; Simon et al. 2018; Yamaguchi and Otto 2019). Finally, in Part III, we introduce two extensions to the model, involving selectively induced associations between heterospecific alleles, and phenotypic dominance. This increases the scope and realism

of the model, but also adds complexity. We end by discussing some implications of our results for understanding the process of speciation.

Model and Results

1   Fisher's geometric model and hybridization

1.1   Basics

We consider hybrids between two diploid parental lines P1 and P2, which differ by d substitutions. For simplicity, we will assume that the genetic variation within each parental line, is negligible compared to the divergence between them, although in principle, all analyses could be extended to include within-line variation (e.g. Lynch and Walsh 1998, Ch. 9; Roze and Blanckaert 2014).

Hybrids will contain some combination of alleles from the two parental lines. We characterize hybrid genotypes in terms of their heterozygosity, p_{12}, and hybrid index, h (see Figure 1). The heterozygosity is the proportion of the d divergent sites where the hybrid carries one allele from each line; the hybrid index is the total proportion of the divergent alleles which come from line P2. As such, h ranges from 0 for a pure P1 genotype, to 1 for a pure P2 genotype. We will also use the notation p_1 and p_2 to refer to the proportion of divergent sites that are homozygous for alleles from P1 and P2, such that $p_1 + p_2 + p_{12} = 1$, and

$$h \equiv p_2 + \frac{1}{2} p_{12} \tag{1}$$

(Turelli and Orr 2000; Simon et al. 2018). Our overall aim is to determine how the distribution of hybrid fitnesses depends on h and p_{12}.

1.2   Fisher's model as a fitness landscape

Under Fisher's geometric model, all genotypes are associated with the values of n continuously varying phenotypic traits. To begin with, we assume that the phenotypic effects of all d substitutions act additively on each trait. In this case, the value of trait i for any given hybrid genotype can be written as

$$z_i = z_{\mathrm{P1},i} + \sum_{j \in J_{hom}} m_{ij} + \sum_{j \in J_{het}} \frac{1}{2} m_{ij} \tag{2}$$

where $z_{\mathrm{P1},i}$ denotes the value of trait i in parental line P1, J_{hom} denotes the subset of $d \times p_2$ loci that are homozygous for the P2 allele, and J_{het} denotes the non-overlapping subset of $d \times p_{12}$ loci that are heterozygous. The m_{ij} describe the effects on trait i of introducing the P2 allele at locus j (whether derived or ancestral) into a P1 background. In the P2 genotype, by definition, all substitutions are present in homozygous state, and so its phenotype can be written as:

$$z_{\mathrm{P2},i} = z_{\mathrm{P1},i} + \sum_{j=1}^{d} m_{ij} \tag{3}$$

All of these quantities are illustrated in Figure 1. Next, we must decide how trait values determine fitness. We will assume that the fitness of a genotype depends on the distance of its phenotype from some optimal phenotype, as

determined by the environment. We use the following weighted distance measure

$$\|\mathbf{z} - \mathbf{o}\|_{\lambda}^2 \equiv \sum_{i=1}^{n} \lambda_i (z_i - o_i)^2 \tag{4}$$

where $\mathbf{o} = (o_1, o_2, \ldots, o_n)$ is the vector of optimal values, and λ_i is the strength of selection on trait i in the current environment. (This metric is equivalent to the "breakdown score" of Turelli and Orr 2000, and Simon et al. 2018.) Equation 4 implies that selection is independent on all traits. However, if correlated selection can be approximated by a multivariate Gaussian function, then, without loss of generality, we can always change the axes, and define n new traits, under independent selection (Waxman and Welch 2005; Martin and Lenormand 2006a). Furthermore, if the distribution of the fixed effects, the m_{ij}, is also sufficiently close to multivariate normality, including arbitrary covariances, then we can choose new traits, under independent selection, which have unit variances and no covariances. In this case, the λ_i capture both differences in selection, and differences in the typical sizes of factors fixed (Martin and Lenormand 2006a; see also Appendix 1 for a full derivation).

Finally, fitness is a function of the distance metric, such that fitness decreases with distance (Turelli and Moyle 2007; Simon et al. 2018). This could take a form such as

$$\ln w = -\alpha \|\mathbf{z} - \mathbf{o}\|_{\lambda}^k \tag{5}$$

where k denotes the curvature of the fitness landscape, i.e. how quickly fitness declines with the distance from the optimum (Peck et al. 1997; Tenaillon et al. 2007; Fraïsse et al. 2016; Fraïsse and Welch 2019). In the remainder of the paper, following common practice (Turelli and Orr 2000; Wade and Demuth 2005; Turelli and Moyle 2007), we will not work with fitness directly. Instead, we follow Simon et al. (2018), and use the following scaled squared distance (or scaled transformed fitness):

$$r^2 \equiv 4 \frac{\|\mathbf{z} - \mathbf{o}\|_{\lambda}^2}{d \sum_{i=1}^{n} \lambda_i} \tag{6}$$

The denominator in eq. 6 can be understood in various ways. First, it is simply a convenient scaling to remove nuisance parameters, so that the results depend on a handful of simple geometric quantities. Second, when the parental lines are close to the optimum, applying the scaling is equivalent to dividing all fitness measures by the measures for some reference class of hybrid (Hill 1982; Lynch and Walsh 1998 Ch. 9; Simon et al. 2018; Yamaguchi and Otto 2019). Third, the denominator can itself be written as a "reference distance":

$$d \sum_{i=1}^{n} \lambda_i = \left\| \sqrt{d} \mathbf{1} \right\|_{\lambda}^2 \tag{7}$$

(where $\mathbf{1}$ is the unit vector). Equation 7 is the amount of phenotypic change that would be expected for a single population that wandered away from its ancestral state via an undirected random walk of d steps, where the step size is given by the λ_i. This is equivalent to the amount of phenotypic change expected by fixing d mutations via random drift, but conditioned on the observed distribution of effect sizes.

For brevity in what follows, "distance" will refer to scaled quantities such as eq. 6, and we return to the interpretation of the scaling below.

1.3   The Brownian bridge approximation

Equations 2-3 suggest an approximation that can help us to understand the distribution of hybrid fitnesses. In this section, we describe this approximation in geometric terms, and discuss its biological interpretation in the following section. If each hybrid genotype contains an effectively random sets of alleles, we can treat the m_{ij} as random variables that are statistically independent on each trait. This means that, if we consider a hybrid genotype with a given level of heterozygosity and hybridity, its expected distance from the optimum is

$$E\left(r^2\right) = 4\frac{\sum \lambda_i \mathrm{Var}\left(z_i\right)}{d \sum \lambda_i} + 4\frac{\sum \lambda_i E^2\left(z_i - o_i\right)}{d \sum \lambda_i} \tag{8}$$

$$\equiv V + M \tag{9}$$

where the means and variances are defined over all hybrid genotypes with the same heterozygosity and hybrid index. To derive the quantities V and M, we will now assume that the m_{ij} on each trait are the increments of a Brownian bridge, i.e., a random walk constrained at each end by the parental phenotypes, and split into d steps (Revuz and Yor 1999; Simon et al. 2018). This approximation relies on the fact that alleles are placed into hybrids in an effectively random combinations, but it does not require that the true process of divergence resembled a random walk. The approximation also depends on the normality of the m_{ij}; but if the divergence, d, is sufficiently large, then the summations in eq. 2 lead to central-limit-type behaviour. This can lead to approximate normality, regardless of the true distribution of the m_{ij} (see also Barton et al. 2017 section 3.2).

Simon et al. (2018) presented this Brownian bridge approximation, and in Appendix 1, we rederive the same result with less haste. For the term V, which captures the variation in hybrid phenotypes, we show that

$$V = 4p_2(1-p_2) + p_{12}(1-p_{12}) - 4p_2 p_{12} \tag{10}$$

$$= 4h(1-h) - p_{12} \tag{11}$$

where the three terms in eq. 10, come from variation in the effects of homozygous P2 alleles, variation in the effects of heterozygous alleles, and a negative covariance term (because a given allele cannot appear in both homozygous and heterozygous state in the same genome).

For the term M, we note that, with additivity, the expected hybrid phenotype will lie on the line connecting the parental phenotypes. M will therefore depend on the distance of the parental phenotypes from the optimum, and from each other. In particular, we find

$$M \equiv h^2 r_{2O}^2 + (1-h)^2 r_{1O}^2 + 2h(1-h)r_{2O}r_{1O}\rho \tag{12}$$

$$= r_{.O}^2 + \left(h - \tfrac{1}{2}\right)\left(r_{2O}^2 - r_{1O}^2\right) - h(1-h)r_{12}^2 \tag{13}$$

Equation 12 depends on three geometrical quantities, which are illustrated in Figure 2a. These quantities are the scaled squared distance to the optimum of two parental phenotypes:

$$r_{1O}^2 \equiv 4\frac{\|\mathbf{z}_{P1} - \mathbf{o}\|_\lambda^2}{\left\|\sqrt{d}\mathbf{1}\right\|_\lambda^2}, \qquad\qquad r_{2O}^2 \equiv 4\frac{\|\mathbf{z}_{P2} - \mathbf{o}\|_\lambda^2}{\left\|\sqrt{d}\mathbf{1}\right\|_\lambda^2} \qquad (14)$$

and ρ, which is the "cosine similarity" of the vectors connecting these phenotypes to the optimum. ρ can vary between 1, when these vectors point in the same direction, and -1, when they point in opposite directions (see also eq. 44). In equation 13, we use the cosine rule ($r_{12}^2 = r_{1O}^2 + r_{2O}^2 - 2r_{1O}r_{2O}\rho$), to give the same result in terms of the scaled distance between the parental phenotypes

$$r_{12}^2 \equiv 4\frac{\|\mathbf{z}_{P1} - \mathbf{z}_{P2}\|_\lambda^2}{\left\|\sqrt{d}\mathbf{1}\right\|_\lambda^2} \qquad (15)$$

(Figure 2a) where we have also used the notation $r_{.O}^2 \equiv \frac{1}{2}\left(r_{1O}^2 + r_{2O}^2\right)$ for the mean distance from the optimum of the parental phenotypes.

1.4 Biological interpretation

Let us now consider a group of hybrids that might vary in their values of h and p_{12}. Putting the results above together, we have

$$\begin{aligned}
E\left(r^2\right) = & r_{.O}^2 \\
& + \left(\bar{h} - \frac{1}{2}\right)\left(r_{2O}^2 - r_{1O}^2\right) \\
& - \bar{p}_{12} \\
& + 4\left[\bar{h}(1-\bar{h}) - \mathrm{Var}(h)\right]\left(1 - \frac{1}{4}r_{12}^2\right)
\end{aligned} \qquad (16)$$

where $\bar{p}_{12}$ is the mean level of heterozygosity in hybrids of interest, and $\bar{h}$ and $\mathrm{Var}(h)$ are the mean and variance of their hybrid indexes (Simon et al. 2018). All of the terms in eq. 16 have a clear biological interpretation. First, and simplest, $r_{.O}^2$ is the mean transformed fitness of the parental lines; it tells us that hybrids will be fitter, on average, if their parents are fitter, on average. Second, there is a term that depends on the difference in the parental fitnesses: $\left(\bar{h} - \frac{1}{2}\right)\left(r_{2O}^2 - r_{1O}^2\right)$; it tells us that hybrids will be fitter if they contain more alleles from the fitter parent. The third term, $-\bar{p}_{12}$, is an intrinsic benefit of heterozygosity; it states that, for any given value of h, hybrids are fitter when they are more heterozygous. This is a form of heterosis (Frankel 1983). The effect is "intrinsic" because, unlike the two previous terms, it does not depend on the position of the environmental optimum. The final term captures the intrinsic effects of hybridity. Because it depends on $\bar{h}(1-\bar{h}) - \mathrm{Var}(h)$, its contribution is smallest when hybrids are close to the parental types, either because $\bar{h}$ is close to 0 or 1, or because $\mathrm{Var}(h)$ is close to its maximal value: $\bar{h}(1-\bar{h})$. This is also an "intrinsic" effect because it depends on the distance r_{12}^2, whose value is not affected by the position of the optimum (Figure 2a). The sign of this term changes with the size of r_{12}^2. When $r_{12}^2 < 4$, hybridity

brings a net fitness cost. This reflects the breaking up of coadapted gene complexes in the parental lines (Wallace 1991; Lynch 1991; Simon et al. 2018). When $r_{12}^2 > 4$, hybridity brings a net benefit. This reflects the potential benefits of transgressive variation in hybrids (Yakimowski and Rieseberg 2014).

1.4.1 Conditions for hybrid advantage

To gain further insight, let us consider the conditions for hybrid genotypes to be fitter than the parental lines. This is clearest if we consider balanced hybrids, with equal contributions from the two parents. Setting $h = 1/2$ in eq. 16, we find:

$$E\left(r^2\right) = r_{.O}^2 - \frac{1}{4}r_{12}^2 + 1 - \bar{p}_{12}, \qquad h = 1/2 \tag{17}$$

This expression demonstrates the importance of the distance r_{12}^2. Hybrid advantage, i.e., $E\left(r^2\right) < r_{.O}^2$, is only possible if $r_{12}^2 > 0$. An advantage for homozygous hybrids - which might, in principle, breed true - is only possible if $r_{12}^2 > 4$. These conditions do not depend on the position of the optimum. Nevertheless, hybrids will be fitter in some environments than others, and to see this, let us write eq. 17 as

$$E\left(r^2\right) = \frac{r_{12}^2}{4}\frac{1-\varepsilon}{\varepsilon} + 1 - \bar{p}_{12}, \qquad h = 1/2 \tag{18}$$

Here, $0 \geq \varepsilon \geq 1$ describes the position of the optimum relative to the parental phenotypes, and is defined by

$$\varepsilon \equiv \frac{r_{12}^2}{4r_{.O}^2} = \frac{1}{2}\left(1 - \Delta\rho\right) \tag{19}$$

where ρ is defined above (Figure 2a), and $\Delta \equiv (2r_{1O}r_{2O})/(r_{1O}^2 + r_{2O}^2)$, parameterises the difference in the distances of the two parents to the optimum. The quantity ε is illustrated in Figure 2b.

It follows from eq. 18, that the potential for hybrid advantage depends on a property of the parental lines: r_{12}^2, and a property of the environment where the hybridization takes place: ε. For a given value of r_{12}^2, hybrid fitness will increase with ε, and will be maximized at $\varepsilon = 1$ (Figure 2b). This restates a result of Yamaguchi and Otto (2019): hybrids will enjoy the greatest advantage in intermediate environments, to which the parental lines are equally maladapted ($\Delta = 1$), and maladapted in opposite directions in phenotypic space ($\rho = -1$). This result also follows intuitively: intermediate hybrid phenotypes will best match intermediate environments (Moore 1977).

1.4.2 Hybrid fitness and the process of divergence

Results above show that r_{12}^2 is a key quantity for determining the outcomes of hybridization. By definition, r_{12}^2 depends on the amount of phenotypic change that accrued during the genomic divergence between the parental lines; and this is scaled by the expectations under a random walk (eq. 15). For this reason, as shown in Figure 3, the value of r_{12}^2 contains some information about the actual mode of divergence between the populations.

For example, if $r_{12}^2 > 4$, then the parents show more phenotypic divergence than expected under a random walk. This implies that the chain of d substitutions, which connect the parental phenotypes, do so with relatively little meandering or changing of direction. As such, an observation of $r_{12}^2 \gg 4$ suggests that the parental lines diverged via positive selection, and in a consistent direction in phenotypic space. This is illustrated in Figure 3a.

Similarly, an observation of $r_{12}^2 \approx 4$ is expected if the parental phenotypes really did diverge by random-walk-like evolution (Figure 3b). This might occur if parental lines fixed mutations regardless of their fitness costs, e.g., under severe inbreeding. It is notable that the intrinsic effects of hybridity vanish when $r_{12}^2 \approx 4$ (eq. 16). This agrees with the observation that heterozygosity, rather than hybridity, is the major determinant of fitness in crosses between inbred lines (Wright 1922; Neal 1935; Simon et al. 2018).

Finally, if $r_{12}^2 < 4$, then populations have accrued less phenotypic divergence than would be expected under a random walk. This could occur in two quite distinct ways. First, stabilizing selection might maintain the phenotype at a (more-or-less) stationary optimum, while still allowing for divergence at the genomic level, perhaps by nearly neutral evolution. This is sometimes called "system drift" (Barton 1989; Mani and Clarke 1990; Hartl and Taubes 1998; Rosas et al. 2010; Schiffman and Ralph 2017), and it is illustrated in Figure 3c. Alternatively, divergence could involve adaptation to a moving optimum, but with a complex pattern of environmental change, e.g., with an optimum that oscillated back and forth; this is illustrated in Figure 3d. In both cases, the result would be a chain of substitutions that went less far than expected under a random walk, such that $r_{12}^2 < 4$.

Importantly, over large periods of time, at least one of these two processes - system drift, or complex environmental change - is very likely to occur. So at very large divergences, it becomes increasingly likely that $r_{12}^2 \approx 0$ will hold, with the loss of any potential for hybrid advantage. In fact, it follows directly from eqs. 14-15 that all of the key distances shown in Figure 2a will tend to vanish at large divergences, and we have the limit:

$$\lim_{d \to \infty} E\left(r^2\right) = 4\bar{h}(1 - \bar{h}) - 4\text{Var}\left(h\right) - \bar{p}_{12} \tag{20}$$

Biologically, eq. 20 implies that, as populations diverge genetically, both extrinsic fitness effects, and any intrinsic benefits of hybridity, will tend to become less and less important. The model predicts convergence to a characteristic pattern of intrinsic isolation between the parental lines, where a fixed cost of hybridity is mitigated by a fixed benefit of heterozygosity.

1.5 Individual-based simulations

To illustrate the points above, and test the robustness of our approximations, we used individual-based simulations of Fisher's model. In particular, we simulated the divergence between pairs of populations, followed by hybridization in the ancestral environment. We simulated under $2^7 = 128$ different combinations of population-genetic parameters (Table 1). These parameters were chosen to encompass a range of conditions, including varying levels of standing variation, recombination, distributions of mutant effects, and other properties of the fitness landscape. Justifications of our parameter choices, and full details of all simulations, are given in Appendix 2.

For each of the 128 parameter combinations, we simulated 15 distinct scenarios of population demography, and environmental change. The purpose of the scenarios was to generate predictably different patterns of genomic divergence (Figure 3), and predictably different patterns of hybrid fitness in the ancestral environment (Figure 2b).

Figure 4 describes these divergence scenarios, and plots the realized values of the key quantities r_{12}^2 and ε (see eqs. 15, 18 and 19). First, let us consider the results for scenarios 1-4. In these scenarios, the divergence between the populations involved local adaptation, with optima that moved in opposite directions, with respect to the ancestral habitat (see top right-hand panel). As Figure 4 shows, the adaptive divergence in these simulations led to some very high values of r_{12}^2, conducive to hybrid advantage. The r_{12}^2 values are highest when the divergence

took place in the face of ongoing gene flow (scenarios 1-2). This is because the gene flow tended to keep the overall genomic divergence low (Yeaman and Whitlock 2011), such that paths of substitutions resemble Figure 3a more than Figure 3c. The r_{12}^2 values are also higher when the optima moved in single jump (scenarios 1-3), than when it moved back and forth via smooth cycles (scenario 4). This reflects the meandering in the chains of substitutions when environmental conditions fluctuate (compare Figure 3d to Figure 3a). For all four scenarios, values of ε were close to their maximal value of 1, implying that the optimum in the ancestral habitat - which was exactly intermediate between the two new optima - was ideally placed to maximize hybrid advantage (Figure 2b). There was also some variation in ε, either because the cycling optima returned to their ancestral state (scenario 4), or because high levels of maladaptive gene flow moved the parental populations away from their local optima (scenario 1).

Scenarios 5-6 resemble scenarios 3-4, except that population P1 evolved, while population P2 remained in their common ancestral state (see second right-hand panel). As expected, this had little effect on r_{12}^2 values, but reduced ε from ~1 to exactly $1/2$. This is because the ancestral optimum no longer maximized hybrid advantage: in the ancestral habitat, hybrids were often fitter than parental line P1, but never fitter than the optimal P2 (Figure 2b). Very similar results were obtained for scenarios 7-8. Here, populations adapted to moving optima on different traits (see third right-hand panel). In these cases, the value of $\varepsilon \approx 1/2$ reflected the fact that hybrids could gain a fitness advantage over both parental lines, but less so than in an intermediate habitat (Figure 2b).

In scenarios 9-10, results were qualitatively different. Here, the populations adapted, independently, to identical environmental change (see fourth right-hand panel). As a result, the parental phenotypes were very close to each other (such that $r_{12}^2 \ll 1$), and the ancestral optimum was poorly placed to generate hybrid advantage ($\varepsilon \approx 0$).

The next group of scenarios, 11-13, involved stable environmental conditions, so that all of the divergence accumulated via genetic drift in allopatry (bottom right-hand panel). Here, values of ε varied greatly, reflecting the fact that parental populations deviated from their shared optimum in independent, random directions. High values of r_{12}^2 were obtained only in one case: when the population size was very small (scenario 11). In this case, selection was often ineffective, at least initially, so that the population phenotypes wandered away from their optimal values (see Figure 3b, and lighter lines in the lower right-hand panel of Figure 4). By contrast, for scenarios 12-13, parental phenotypes remained close to their shared optimum at all times (see darker lines in lower right-hand panel of Figure 4), and this kept values of r_{12}^2 consistently low.

Finally, for scenarios 14-15, we reanalysed the simulations from scenarios 3-4, but scored hybrids in a "benign environment" (Hatfield and Schluter 1999). This means that fitness was affected only by the $n-1$ phenotypic traits that had not been subject to diversifying selection. As such, the fitness effects in hybrids reflect the pleiotropic effects of adaptive substitutions (Thompson 2019), and results closely resemble those under "system drift".

Figure 4 confirms that there are consistent differences between hybridization scenarios, which reflect both the history of divergence (captured in r_{12}^2) and the current environmental conditions (captured in ε). Nevertheless, as shown in Figure 5a-b, the Brownian bridge approximation applied well to all of the simulations reported above. Figure 5a shows results for hybrids that were formed by randomly assembling combinations of homozygous alleles from the parental lines, in different proportions (i.e., with different values of h). Results show that $\overline{r^2 - M}$ is well predicted by eq. 11 for all simulations. In many of these simulations, the distribution of fixed effects was highly non-normal (see Figure S1 for details), but the Brownian bridge approximation applied nonetheless. Figure 5b illustrates another feature common to all of the simulations: M tended to vanish if the genomic divergence grew large. With large d, the outcome of hybridization was always predicted by eq. 20, regardless of the divergence

scenario, or the current position of the optimum.

2   Fisher's model and the quantitative genetics of line crosses

2.1   The composite effects under Fisher's model

We now relate the fitness landscape results, derived above, to some standard empirical tools for investigating hybrid fitness. These tools are usually applied to controlled crosses, including the initial F1 (P1×P2), the F2 (random mating among the F1), and backcrosses of the F1 to the parental lines. Fitness differences between these crosses, within and between environments, are key data for understanding the outcomes of hybridization (Lynch 1991; Hatfield and Schluter 1999; Rundle and Whitlock 2001; Rosas et al. 2010; Simon et al. 2018; Yamaguchi and Otto 2019).

For each cross type, the quantities $\bar{p}_{12}$ and $\bar{h}$ are simple to calculate, but $\mathrm{Var}(h)$ will depend on the levels of segregation and recombination (Lynch and Walsh 1998, Ch. 9). If $\bar{c}$ is the mean rate of recombination among pairs of loci, then $\mathrm{Var}(h) \approx (1 - 2\bar{c})/4$ among F1 gametes (Zeng et al. 1990; Lynch and Walsh 1998, Ch. 9); and so, with random union of gametes, $\mathrm{Var}(h)$ will be half of this value for the F2, and a quarter of this value for backcrosses. Furthermore, $\bar{c} \approx 0.5$ for many species (Lynch and Walsh 1998, Ch. 9), and in those cases, $\mathrm{Var}(h)$ can often be neglected.

Now, let us follow Hill (1982; see also Lynch 1991, Lynch and Walsh 1998, Ch. 9), and write the expected value of a trait in a hybrid as

$$\mu = \mu_0 + \theta_S \{\alpha_1\} + \theta_S^2 \{\alpha_2\} + \theta_H \{\delta_1\} + \theta_H^2 \{\delta_2\} + \theta_S \theta_H \{\alpha_1 \delta_1\} + ... \tag{21}$$

where $\theta_S \equiv 1 - 2\bar{h}$ and $\theta_H \equiv 2\bar{p}_{12} - 1$. The curly brackets contain the "composite effects", which are defined in Table 2 (and noting that, in this standard notation, $\{\alpha_1 \delta_1\}$ describes an interaction term, and not a product). Equation 21 includes only pairwise effects, but the model might also be expanded to include higher-order terms.

Both eqs. 16 and 21 are polynomials in $\bar{h}$ and $\bar{p}_{12}$, and so we can apply eq. 21 to the transformed hybrid fitness, setting $\mu = E(r^2)$, and solve for the composite effects. The results, found in Table 2 (column "single environment"), show that Fisher's model predicts three non-zero composite effects. Their values reflect the biological distinctions discussed above. In particular, the additive effect, $\{\alpha_1\} = \frac{1}{2}(r_{1O}^2 - r_{2O}^2)$, captures the benefits of carrying alleles from the fitter parent, while the dominance effect, $\{\delta_1\} = -\frac{1}{2}$, captures the intrinsic benefits of heterozygosity. The pairwise epistatic effect, $\{\alpha_2\} = \frac{1}{4}r_{12}^2 - 1$, balances the intrinsic costs and benefits of hybridity.

Rundle and Whitlock (2001) presented a useful extension of eq. 21 for traits scored in two environments. Introducing an indicator variable, I, which is 0 for individuals scored in "environment A", and 1 for individuals scored in "environment B", and defining $\theta_E \equiv 2I - 1$, their model contains the additional terms:

$$\mu_{AB} = \mu + \theta_E \{\varepsilon\} + \theta_E \theta_S \{\alpha_1 \varepsilon\} + \theta_E \theta_H \{\delta_1 \varepsilon\} + ... \tag{22}$$

Fisher's geometric model is trivially extended in the same way, by adding a second environment, with a distinct optimum. This is illustrated in Figure 2c. Again, we can solve for the composite effects, and these are shown in

Table 2 (column "two environments"). Results show that adding a second environment leaves the dominance and epistatic effects unchanged, confirming that they represent the intrinsic effects of heterozygosity and hybridity. Of the remaining quantities, the additive effect, $\{\alpha_1\}$, is now averaged across environments ($r_{1\cdot}^2 - r_{2\cdot}^2 \equiv \frac{1}{2}(r_{1A}^2 - r_{2A}^2 + r_{1B}^2 - r_{2B}^2)$), while the main environmental effect, $\{\varepsilon\}$, is simply the difference in fitness between environments, averaged across the parental lines ($r_{\cdot B}^2 - r_{\cdot A}^2 \equiv \frac{1}{2}(r_{1B}^2 - r_{1A}^2 + r_{2B}^2 - r_{2A}^2)$).

Finally, the additive-by-environment interaction is

$$\{\alpha_1 \varepsilon\} = \tfrac{1}{4}\left[\left(r_{1B}^2 - r_{1A}^2\right) - \left(r_{2B}^2 - r_{2A}^2\right) \right] \tag{23}$$

$$= \tfrac{1}{2} r_{12} r_{AB} \rho^* \tag{24}$$

Here, $-1 \leq \rho^* \leq 1$ is the cosine similarity of the vector linking the parental phenotypes, and the vector linking the two optima (Figure 2c): $\{\alpha_1 \varepsilon\}$ will be large when the difference between the phenotypes of P1 and P2 resembles the difference between optima A and B. Indeed, the predicted values of $\{\alpha_1\}$, $\{\varepsilon\}$ and $\{\alpha_1 \varepsilon\}$ are equivalent to the quantities described by Blanquart et al. (2013) for measuring the extent of local adaptation (see their eqs. 1-2), but applied to fitness values that have been transformed and scaled (eq. 6). As shown in Appendix 3, the same framework is also extendable to hybrids that are formed in patchy ecotones, or in habitats that are ecologically intermediate between A and B.

2.2 Local adaptation and ecological isolation

While $\{\alpha_1 \varepsilon\}$ is a possible measure of local adaptation, it does not describe the extent of ecological isolation between the parental lines. For example, $\{\alpha_1 \varepsilon\}$ might be very high, even if P1 were fitter than P2 in both habitats (Kawecki & Ebert 2004; Blanquart et al. 2013). However, we do have a measure of isolation in an important special case. If the two parental lines are well adapted to different local optima, then r_{12}^2 describes the distance between these optima, as well as the distance between the parental phenotypes. As a consequence, results for hybrids depend on r_{12}^2 alone; and this is shown in Table 2 (column "local adaptation").

With local adaptation of this kind, large values of r_{12}^2 imply pure ecological isolation (where P2 and P1 are kept distinct solely by environment-dependent selection against their divergent alleles), while small values imply pure intrinsic isolation (where hybrids are less fit than the parents in both habitats). By changing the value of r_{12}^2 we can interpolate between these two extremes, and this is shown in Supplementary Figure S2. With local adaptation, r_{12}^2 is both a measure of the amount of meandering in the chain of fixed differences, and a measure of the relative strengths of ecological isolation versus intrinsic isolation.

To illustrate this, Figure 6 shows results from a single simulation run. In this run, parental populations adapted in allopatry, to distinct optima. The adaptation involved fixing ~50 substitutions, and after this, the populations continued to diverge via system drift. We consider hybridization between these populations in five distinct environments (Figure 6a), tracking both the composite effects over the complete course of the divergence (Figure 6b-e), and the fitness of crosses generated by standard amphimixis, at two time points (Figure 6f-o)

As is clear from Figure 6b-e, several of the composite effects change in similar ways, reflecting their common dependence on r_{12}^2. In the earliest stages of divergence, r_{12}^2 values were high, reflecting the adaptation of the parental lines to distant, and distinct optima. After 100 substitutions, this was manifest in the clear signatures of

ecological isolation in the parental habitats (Figure 6f-g). In the environment to which P1 is adapted, hybrids tended to be fitter when they carry more P1 alleles, and vice versa (Rundle and Whitlock 2001). For the same reason, in an ecologically intermediate habitat, there was a clear signal of bounded hybrid advantage (Figure 6c). Hybrids tended to have the favoured intermediate phenotype (Moore 1977; Yamaguchi and Otto 2019). Hybrid advantage, at a lower level, also occurred in the ancestral habitat (Figure 6i); but this had nothing to do with the habitat being ancestral, and the same patterns were observed in an entirely novel habitat, characterized by similar values of r_{1O}^2, r_{2O}^2 and r_{12}^2 (Figure 6j).

As the genomic divergence increased, r_{12}^2 decreased, and so all of these differences between environments were greatly reduced (Figure 6b-e). In the example shown, after ~2000 substitutions, hybrid fitnesses were already converging towards the characteristic pattern of intrinsic isolation (Figure 6k-o), with the fixed cost of hybridity, $\{\alpha_2\} \approx -1$, and the fixed benefit of heterozygosity $\{\delta_1\} = -\frac{1}{2}$ (eq. 20).

The results shown in Figure 6 assume free recombination between all loci, but comparable results apply with limited recombination (Supplementary Figure S2). Convergence to the characteristic pattern of intrinsic isolation occurs whatever the mode of divergence (Supplementary Figures S4-S5).

3 Two extensions

3.1 Later crosses, and the Bulmer effect

In principle, the results above apply to any hybrid genome, including later generation crosses. However, this ignores an important complication. The results rely heavily on heterospecific alleles appearing in random combinations, but selection on the early generation hybrids can induce non-random associations between the alleles in their gametes. This can increase the fitness of their offspring, without changing allele frequencies (Bulmer 1971; Walsh and Lynch 2018, Ch. 16). For example, with random union of gametes, the distributions of h and p_{12} will often remain unchanged between the F2 and F3 generations, and so eqs. 21 and 16 make the same predictions for both. However, selection on the F2 parents can upset this expectation.

To understand the strength of this effect, let us consider the case of free recombination ($\bar{c} = 1/2$), and optimal parental lines ($r_{.O} = 0$). With these assumptions, the variance in trait values (the z_i) among F3 offspring is

$$\text{Var}\left(z_{i,\text{F3}}\right) = \frac{1}{2}\left[\text{Var}\left(z_{i,\text{F2}}\right) + \text{Var}\left(z_{i,\text{F2}(sel)}\right)\right] \tag{25}$$

where $z_{i,\text{F2}}$ and $z_{i,\text{F2}(sel)}$ are the trait values for the total F2 population, and for the subset of selected parents (see Walsh and Lynch 2018, Ch. 16, assuming complete heritability). In the extreme case, if only optimal F2 reproduce, then $\text{Var}\left(z_{i,\text{F2}(sel)}\right) = 0$ and $\text{Var}\left(z_{i,\text{F3}}\right) = \frac{1}{2}\text{Var}\left(z_{i,\text{F2}}\right)$. Because the trait variance halves, we have the general bounds

$$\frac{1}{2} \leq \frac{E\left(r^2\right)}{4\bar{h}(1-\bar{h}) - 4\text{Var}\left(h\right) - \bar{p}_{12}} \leq 1 \tag{26}$$

where the upper bound comes from eq. 20, and the lower bound applies if only optimal individuals reproduce. Later crosses will fall somewhere between these bounds, depending on the strength of selection acting on the earlier generation hybrids. This is illustrated in Figure 7a-b. We chose two simulation runs where populations diverged despite a fixed optimum, via system drift. If we generated an F3 cross using a random selection of F2 parents, then the upper bound of eq. 26 applies well (see blue lines in Figure 7). If we selected parents with a

probability proportional to their fitness (eq. 5), the same results continued to apply for later crosses, but only when the fitness function was Gaussian ($k = 2$), so that there was very little inter-individual variation in fitness (Figure 7a). When we imposed a "table-like" fitness function ($k = 6$; Fraïsse et al. 2016), equivalent to strong truncation selection, and implying high variation in parental fitness, then results were closer to the lower bound of eq. 26 (see red lines in Figure 7b). These patterns continued unchanged for other late generation crosses, including the F4 and F5, and also applied to repeated backcrosses to the P1 line (Figure 7a-b).

Figure 7 shows cases where selection on earlier generations of hybrids either has little effect (Figure 7a), or yields simple, predictable effects (Figure 7b). However, it remains a challenge to apply these methods to late-generation hybrids, especially when selection changes allele frequencies (Walsh and Lynch 2018, Chs. 16 and 24; see also Discussion).

3.2 Phenotypic dominance

So far, we have ignored an implausible prediction of Fisher's model. Equation 20, predicts that fully heterozygous hybrids will always be as fit as their parents, because, when $p_{12} = 1$ and $h = 1/2$, the benefits of heterozygosity exactly cancel the costs of hybridity (Barton 2001; Fraïsse et al. 2016; Schiffman and Ralph 2017).

This prediction is worrying because - with strictly biparental inheritance and expression - the initial F1 cross will be fully heterozygous. While many intrinsically isolated species do produce fit F1 (Wallace 1991; Price and Bouvier 2002; Fraïsse et al. 2016), F1 fitness tends to decline with high levels of parental divergence, even in environments where both parents are well adapted (Endler 1977; Bateson 1978; Waser 1993; Price and Bouvier 2002; Edmands 2002; Fraïsse et al. 2016). The model also struggles to explain a second widespread pattern in the F1: when some loci have uniparental inheritance or expression, the reciprocal F1 often have very different fitnesses, even if the parents are both well adapted (Bolnick and Near 2005; Turelli and Moyle 2007; Escobar et al. 2008; Brandvain et al. 2014; Sato et al. 2014; Fraïsse et al. 2016; Bouchemousse et al. 2016). Fisher's model cannot account for these asymmetries unless the global heterozygote is suboptimal (Fraïsse et al. 2016).

In this section, we show that these features of Fisher's model are improved by adding phenotypic dominance (Manna et al. 2011). To see this, let us replace eq. 2 with

$$z_i = z_{\mathrm{P1},i} + \sum_{j \in J_{hom}} m_{ij} + \sum_{j \in J_{het}} \left(\tfrac{1}{2} m_{ij} + \delta_{ij} \right) \tag{27}$$

where δ_{ij} is the deviation from semi-dominance on trait i, caused by introducing the P2 allele at locus j in heterozygous form. We now assume that the δ_{ij} can be treated as the increments of a new, and independent Brownian bridge, linking the midparental value of trait i, to the trait value of the global heterozygote (see the phenotypes labelled P and H in Figure 2d). In Appendix 1, we show that this assumption leads to

$$E\left(r^2\right) \approx V + V_\delta + M + M_\delta \tag{28}$$

where V and M are the additive results (eqs. 10-13), and V_δ and M_δ are the new contributions from variable dominance. In Appendix 2, we show that these new contributions are:

$$V_\delta = 4vp_{12}(1 - p_{12}) \tag{29}$$

$$
\begin{aligned}
M_\delta = &\, p_{12}\left(r_{HO}^2 - r_{PO}^2\right) \\
&- p_{12}(1 - p_{12})r_{PH}^2 \\
&- p_{12}\left(h - \tfrac{1}{2}\right)\left(r_{2H}^2 - r_{1H}^2\right)
\end{aligned} \tag{30}
$$

In equation 29, a new parameter, v, describes the scaled variance of the δ_{ij}. Equation 30, depends on a number of scaled distances, and these are illustrated in Figure 2d. The corresponding changes in the composite effects are listed in Table 2 (column "phenotypic dominance"), and this makes it clear that introducing dominance adds two new composite effects: $\{\delta_2\}$ and $\{\alpha_1\delta_1\}$, and alters the value of a third, $\{\delta_1\}$, which is no longer a constant.

Equations 29 and 30 are both proportional to p_{12} and so they alter predictions only for heterozygous hybrids. The predictions are altered in two ways. First, a non-zero value of $\{\alpha_1\delta_1\}$ (which corresponds to the third term in eq. 30), now allows for "directional dominance". For example, in Figure 2d, the global heterozygote, H, is much closer to the P2 phenotype than to the P1 phenotype ($r_{2H} < r_{1H}$), which implies that P2 alleles are dominant on average. This sort of asymmetry is required to model "dominance drive" in hybrid zones, where alleles can spread due to their dominance relations alone (Mallet and Barton 1989; Barton 1992). Second, by changing $\{\delta_1\}$ (via the first term in eq. 30) variable dominance can reduce or reverse the beneficial effects of heterozygosity. For example, for the global heterozygote, we have

$$E\left(r^2\right) \approx r_{HO}^2, \qquad p_{12} = 1 \tag{31}$$

which, with phenotypic dominance, can differ from the midparental value: $r_{.O}^2 - r_{12}^2/4 \equiv r_{PO}^2$ (Figure 2c). Most importantly, the distance r_{HO}^2, unlike the midparent, is not expected to decline to zero as divergence increases. This is because the global heterozygote - unlike the parental genotypes - may never be exposed to natural selection, and cannot, in any case, breed true. As such, it can continue to wander away from the optimum as divergence increases, even if effective stabilizing selection acts on the parental phenotypes. In addition, the effects of heterozygosity are now extrinsic, since they depend on the position of the optimum.

To test the predictions of eqs. 28-31, we repeated our complete set of simulations (Figure 4; Table 1), but now including variable phenotypic dominance. To do this, the heterozygous effect of each new mutation on each trait, was generated by multiplying its homozygous effect by a uniformly-distributed random number. Results are summarised in Figure 5c-f, and in more detail in Supplementary Figure S6.

Figure 5c corresponds to Figure 5a, and shows that the standard Brownian bridge continues to hold for homozygous hybrids. Figure 5e shows equivalent results for heterozygous hybrids, i.e., random assemblies of P2 alleles, added in heterozygous state to a P1 background. In this case, when $h = p_{12}/2$, we have the prediction $V + V_\delta = (1 + 4v)p_{12}(1 - p_{12})$. Results show that these hybrids are usually less fit than predicted under the additive model (additive predictions are indicated by the dotted white line in Figure 5e). In most cases, results agree with predictions after setting $v = 1/12$ (shown by the black line in Figure 5e), which matches the variance of the uniform distribution, used to generate the heterozygous effects of new mutations. However, the parameter v applies to

fixed effects, and not to new mutations, and so there are some clear differences between the values of v realised in different parameter regimes and divergence scenarios. This includes a few runs where these heterozygous hybrids were fitter than the additive predictions, suggesting that the P2 alleles were recessive on average (see Figure S6 for more details).

Results relevant to the F1 are shown in Figure 5d and 5f. These panels confirm that, even when divergence has increased, and the parental distances has declined towards zero (Figure 5d), the phenotype of the global heterozygote can remain a non-negligible distance from the optimum (Figure 5f). This implies that, with phenotypic dominance, optimal but divergent parents may generate an unfit F1 (eq. 31). This is confirmed in Figure 7c-d. Here, results closely match the additive results shown in Figure 7a-b, except for the F1, which is noticeably less fit. Similar results for a complete set of crosses are shown in supplementary Figure S7.

Discussion

Fisher's model, and selection and drift in speciation

Using fitness landscapes, based on Fisher's geometric model, we have developed analytical predictions for the fitness of hybrids between divergent lines. The predictions involve a small number of quantities with an intuitive biological meaning. For example, with the simplest model (additive phenotypes, and a single optimum), rank-order fitness differences between classes of hybrid, depend solely on three distances: r_{1O}^2, r_{2O}^2 and r_{12}^2 (Figure 2a). Two of these distances are "extrinsic", in that they depend on the position of optimum, while the third, r_{12}^2, is "intrinsic", i.e., a property of the parental phenotypes. The result is that we can classify a large number of ecological scenarios, according to the values of these distances.

This classification cuts across many of the distinctions that have been most important in the theoretical study of phenotypic divergence, and in the study of speciation. For example, many studies have investigated the distributions of fixed effects on quantitative traits (Fisher 1930; Orr 1998; Rockman 2002; Griswold 2006; Yeaman and Whitlock 2011; Matuszewski et al. 2014; Débarre et al. 2015; Dittmar et al. 2016), and many more studies have investigated the relative roles of divergent selection, mutation order and drift, in driving reproductive isolation (Mani and Clarke 1990; Schluter 2000; Coyne and Orr 2004; Gavrilets 2004). By contrast, results presented here highlight properties of hybrid fitness that apply to a wide range of effect size distributions (Figure S1), and which are relatively insensitive to the forces driving the divergence (Figures 3-5).

This way of characterizing our results comes with two caveats. First, much of the generality follows from the scaling applied in eq. 6. As stressed by several previous authors, positive selection will often fix larger and more variable effects, and lead to divergence that is more rapid, and more resistant to the swamping effects of gene flow. These facts alone will lead to selectively-driven changes having a greater influence on hybrid fitness (Chevin et al. 2014; Fraïsse et al. 2016; Yamaguchi and Otto 2019), and implicate natural selection, rather than drift, in the early stages of divergence (Jiggins and Mallet 2000; Coyne and Orr 2004, Ch. 11; Yamaguchi and Otto 2019). In the results presented here, all of these factors are absorbed into the scaling factor (eqs. 6-7), with selection leading to larger values of d and larger values of the λ_i. Second, our results do highlight one important difference between selection and drift. Adaptive divergence - especially in the face of gene flow - is more likely to lead to high values of the scaled distance, r_{12}^2 (Figures 3-4). This implies that selection is far more likely than drift to generate bounded hybrid advantage (Figure S2). It also implies that - in some cases at least - something about the mode of divergence might be inferred from the value of r_{12}^2.

Nevertheless, despite these points, our results do suggest that reproductive isolation can evolve in similar ways, under a wide variety of conditions. For example, the value of the additive-by-additive epistasis, $\{\alpha_2\}$, tells us nothing at all about the role of epistatic genetic variation during the divergence process (Lynch 1991; Welch 2004; Wade and Demuth 2005). Even the timescale of divergence, which is certain to vary, probably does so in a way that is quite sensitive to many details of the model (e.g., Whitlock 1997; Orr 1998; Barton 2005; Griswold 2006; Yeaman and Whitlock 2011; Rockman 2012; Débarre et al. 2015; Matuszewski et al. 2014, 2015; Dittmar et al. 2016; Thompson et al. 2019; Yamaguchi and Otto 2019). But in all cases, if the class of fitness landscape studied here is sufficiently close to reality, intrinsic isolation - and eventually speciation - is an inevitable consequence of divergence.

The connection to quantitative genetics

In the second part of this work, we drew a formal connection between Fisher's model, and the quantitative genetics of line crosses (Hill 1982; Lynch 1991; Rundle and Whitlock 2001; Wade and Demuth 2005; Yamaguchi and Otto 2019). Here, the main goal was to shed light on the meaning of the composite effects. However, results also show that Fisher's model makes testable predictions about the values of these effects, and that the key geometric quantities might be estimated from cross data.

This claim also comes with an important caveat: the predictions in Table 2 do not apply to raw fitness values, but to values that have been suitably transformed and scaled; in our notation, they apply to r^2 and not to w (see eqs. 5-6). This means that the tests cannot be applied "off-the-shelf" to many published estimates. While data transforms are an inherent part of quantitative genetics (Lynch and Walsh 1998, Ch. 11), there is also the problem of estimating the scaling factor of eq. 7, which forms the denominator of eq. 6. This quantity is relatively easy to estimate from a diverse collection of hybrids (see Simon et al. 2018), but the extra parameter is more difficult to estimate from a limited number of controlled crosses (Yamaguchi and Otto 2019). We ducked this issue in Figure 6, by estimating eq. 7 directly from the simulated fixed effects (see Appendix 1) - but this would not be possible for real-world data. While this is a real limitation, there are many special cases where the geometric quantities can be straightforwardly estimated from fitness values that are transformed but unscaled (i.e., from the numerator of eq. 6). For example, with two locally-adapted populations (Table 2), the key distance r_{12}^2 can be estimated from a ratio of composite effects:

$$r_{12}^2 = \frac{4\{\alpha_1 \varepsilon\}}{\{\alpha_1 \varepsilon\} - 2\{\alpha_2\}} \tag{32}$$

so that the scaling factor cancels. Simon et al. (2018) give other, similar examples.

Extensions to the model

The third and final part of this work explored two extensions to Fisher's model. First, we highlighted the effects of disequilibria caused by selection on early generation hybrids (as opposed to selection during the divergence). This will be an important consideration if the model is applied to entire swarms (Jiggins and Mallet 2000; Allendorf et al. 2001; Vernesi et al. 2003; Simon et al. 2018). If some, but not all hybrids have strong selectively-induced associations between heterospecific alleles, or between unlinked homospecific alleles, then h and p_{12} alone will be insufficient to predict their fitness. Observations of strong incompatibilities, involving small genomic regions, suggests that such associations might be widespread (Barton 2001; Coyne and Orr 2004, Ch. 8; Fraïsse et al. 2016).

Second, we have incorporated phenotypic dominance into Fisher's model. Our treatment ignored one known feature of dominance: the differences between large- and small-effect changes (Wright 1929, Manna et al. 2011; Fraïsse et al. 2016). Nevertheless, even our simple treatment improved the realism of Fisher's model in a qualitative way. By reducing the beneficial effects of heterozygosity, the extended model can account for low fitness F1 between highly divergent, but equally fit parental lines (Fraïsse et al. 2016). This extension also shows how Fisher's model can incorporate other modelling approaches as special cases (Simon et al. 2018). For example, when the parental lines have high fitness, eq. 28 is identical to eq. A37 of Simon et al. (2018), which was derived from a model of Dobzhansky-Muller incompatibilities, with variable dominance relations (Turelli and Orr 2000).

Finally, both of our extensions illustrate how existing work on quantitative traits can be used with Fisher's model - even when, as here, the model is treated as a fitness landscape, and the "traits" are not taken too literally (Martin 2014; Fraïsse and Welch 2019). In future, the approach might be extended to landscapes with multiple phenotypic optima (Whitlock 1997), or to populations that diverged phenotypically via linkage disequilibria, rather than allelic substitutions (Kremer & Le Corre 2012; Savolainen et al. 2013).

Acknowledgements

We are very grateful to Jean-Baptiste Grodwohl, Sally Otto, Aylwyn Scally and Alexis Simon. Some of the simulations used the Montpellier Bioinformatics Biodiversity platform supported by the LabEx CeMEB, an ANR "Investissements d'avenir" program (ANR-10-LABX-04-01).

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Figures

Figure 1

Fisher's geometric model as a model of hybridization. Each genotype is associated with the values of n quantitative traits (illustrated with $n = 2$), and so mutations and substitutions are vectors of change in this n-dimensional space. Fitness depends on the distance of the phenotype from an optimum. Shown are two parental lines, P1 and P2, which differ by $d = 7$ substitutions. The arrows represent the effects of each substitution in heterozygous and homozygous state. They are defined relative to the P1 genotype, regardless of whether the alleles are derived or ancestral (the m_{ij} in eq. 2 are the components of these vectors). Also shown is a hybrid genotype. In this hybrid, 1/7 of the divergent sites contains an allele from each line ($p_{12} = 1/7$), and two further P2 alleles are present as homozygotes (one ancestral and one derived), yielding a hybrid index of $h = 2.5/7$.

Figure 2

The predictions of Fisher's geometric model for hybrids, depend on a small number of geometric quantities. These are defined in some multi-dimensional trait space, but are estimable, in principle. (a) With additive phenotypes, and a single environmental optimum, predictions depend on the scaled distance of the two parental phenotypes from the optimum (r_{1O} and r_{2O}), and the angle between the vectors linking these phenotypes to the optimum. The cosine of this angle, ρ, measures the extent to which the parental populations are maladapted to the environment in similar ways. Results can also be given in terms of the scaled distance of the parental phenotypes from each other (r_{12}), and this does not depend on the position of the optimum. (b) For balanced hybrids (with $h = 1/2$), results depend solely on r_{12}, and a quantity ε, which quantifies the potential of the current environment to generate hybrid advantage. The maximum value of $\varepsilon = 1$ is obtained with an intermediate optimum, which best matches the intermediate hybrid phenotypes. Lower values of ε correspond to optima placed on circles of increasing diameter. (c) With two environments, A and B, characterized by different optima, one measure of local adaptation is ρ^*: the cosine similarity between the vectors linking the optima, and the parental phenotypes. When the two parental phenotypes are very close to the two optima, scaled hybrid fitness depends only on r_{12}. (d) With variable phenotypic dominance, results depend on the phenotype of the global heterozygote, H, which is equivalent to the F1 cross under strictly biparental inheritance and expression. A consequence of variable dominance is that H may differ from the midparental phenotype, P. In the example shown, H is closer to the P2 phenotype than the P1 phenotype, this implies directional dominance, with P1 alleles being recessive on average.

Figure 3

The scaled distance r_{12}^2 plays a key role in determining patterns of hybrid fitness, and it can vary systematically with the mode of divergence between the parents. The variation depends on the chain of d substitutions that differentiate the parental lines, and compares their trajectory to a random walk with the same number of steps, and distribution of effect sizes. (a) When the substitutions form a more-or-less direct path between the parental phenotypes, the observed phenotypic difference will be greater than would be predicted under a comparable random walk; this implies that $r_{12}^2 > 4$. (b) when the true path of divergence really did resemble a random walk, then $r_{12}^2 \approx 4$ is expected. This might happen if stabilizing selection on the phenotype was ineffective, or if the optimum value wandered erratically. Systematically smaller values of r_{12}^2 are predicted under two conditions. Either (c) genomic divergence continued, despite effective stabilizing selection on the phenotype, leading to "system drift". Or (d) populations successfully tracked environmental optima that moved in complex ways, e.g., via oscillations.

Figure 4

Individual-based simulations of population divergence under Fisher's geometric model. We simulated 15 distinct divergence scenarios, summarized in the left-hand panel. These scenarios differed in the rate of gene flow during divergence (captured in the expected number of migrants, Nm), and the pattern of environmental change. In particular, optima moved in a single abrupt jump ("J" and triangular points), in smooth repeating cycles ("C" and circular points), or remained stationary throughout the simulation ("-" and square points). Patterns of environmental change also varied between the scenarios as indicated by colour, and illustrated in the right-hand panels. The two central columns show the realized values of two important determinants of hybrid fitness; with points and intervals showing the median and complete range of 128 simulations runs, each with the different set of population genetic parameters (see Table 1). The second column shows r_{12}^2: the scaled difference between the parental phenotypes (Figure 2b), whose value can reflect the process of divergence (Figure 3). The vertical dotted line shows the value $r_{12}^2 = 4$, above which homozygous hybrids have an expected fitness advantage over their parents. The third panel shows ε; this describes the position of the optimum in a given habitat, such that expected hybrid fitness increases with ε (Figure 2b). Values shown here are for the ancestral habitat, which matches the common ancestral phenotype of the diverging populations. For allopatric simulations (scenarios 3-15) runs terminated after one of the populations had fixed 500 substitutions. For parapatric simulations (scenarios 1-2), runs terminated after $50/U$ generations, and mutations were treated as "fixed" if allele frequencies differed by more than 50% between the demes; 11 runs, for which $d < 10$, were excluded. Full details of the simulations are found in Appendix 2.

Figure 5

The Brownian bridge approximation (eqs. 8-13) compared to individual-based simulations. (a) Predictions for V (eq. 11), apply well to all divergence scenarios and population genetic parameter regimes. The lines plotted are the mean of $r^2 - M$, across 10,000 homozygous hybrids, created from hd P2 alleles, and $(1-h)d$ P1 alleles, chosen at random. Results are shown for all of the simulations reported in Figure 4, and use the same colour scheme. (b) For all divergence scenarios, M tends to decline over time, as long as genetic divergence increases (eq. 20). Results are shown for balanced hybrids ($h = 1/2$), with an illustrative set of parameters ($k = 6$, $n = 2$, $NU = 1$, $Ns = -0.1$, bottom-up mutations, and free recombination). Similar results were obtained for all parameter regimes, except for a subset of simulations of divergence in parapatry, where divergence levels reached a quasi-equilibrium value, instead of increasing steadily. (c)-(f) Adding variable phenotypic dominance to the model yields qualitatively different results for heterozygous hybrids (eqs. 28-30). (c) The Brownian bridge approximation continues to apply to homozygous hybrids, but (e) a second Brownian bridge applies to heterozygous alleles placed in a P1 background. The white dotted line shows the prediction of $p_{12}(1 - p_{12})$ that applies under additivity; while the black line shows the prediction of $(1 + 4v)p_{12}(1 - p_{12})$ with $v = 1/12$. (f) $M + M_\delta$ declines over time for homozygous hybrids, but (f) not for heterozygous hybrids. For balanced hybrids, with $h = 1/2$, this reflects the difference between the midparent (with $p_{12} = 0$, $M + M_\delta = r_{.O}^2 - \frac{1}{4}r_{12}^2 = r_{PO}^2$) and the global heterozygote (with $p_{12} = 1$, $M + M_\delta = r_{HO}^2$). In most cases, the parental phenotypes (and therefore the midparent) are directly subject to selection, while the global heterozygote is not. Results imply that - with phenotypic dominance - F1 fitness can be low at high divergences (eq. 31).

Figure 6

Hybridization with local adaptation under Fisher's geometrical model. Panel (a) shows a cartoon of the divergence process that was simulated, with two populations adapting to abruptly shifting optima, and then continuing to accumulate divergence in allopatry. In the common ancestral population, the initial values of all traits were set to zero, but the optimal value of one of the traits was set to one, choosing a different trait for each population (i.e., "scenario 7" from Figure 4). Panels b-e show the change, with increasing divergence, in the composite effects (Table 2), as measured with respect to both parental environments (panel b), or in other single environments (panels c-e). The remaining panels show results for simulated hybrids, plotting r^2 on a reversed axis, such that fitter genotypes are higher up on the plots. Points show simulated values for the three fixed genotypes (the parental lines, P1 and P2 , and the F1 = P1 × P2). Points with error bars show mean and 95% quantiles for 10,000 recombinant hybrids, from the reciprocal backcrosses: BC(P1) = F1 × P1 and BC(P2) = F1 × P2; and the F2 = F1 × F1. The dark point at the centre of each plot shows results for homozygous hybrids, derived as from automictic selfing among F1 gametes. As such the three central crosses (F1, F2 and "Fn", are all balanced hybrids, with $\bar{h} = \frac{1}{2}$, but have maximally different heterozygosities, $\bar{p}_{12} = 1, \frac{1}{2}$ and 0). All points are compared to analytical predictions (blue and red lines). These use eq. 16, with the measured values of r_{1O} and r_{2O} and the assumption that $\rho = -1$ (for the "intermediate" environment) or $\rho = 0$ (all other cases). We note that these predictions relied on our knowing the scaling factor (eq. 7) which can be measured directly from simulated data, but not from real-world data (see Appendix 1 and Discussion). Finally, hybrids were scored in the early stages of divergence ($d = 100$; red lines, panels f-j), where the signal of local adaptation is strong, and at later times ($d = 2000$; blue lines, panels k-o), where the composite effects tend to decline, with convergence towards a characteristic pattern of intrinsic isolation (eq. 20). Divergence was simulated with the following parameters: $N = 1000, N\bar{s} = -10, 2NU = 1, n = 2, k = 2,$

free recombination, and "bottom-up" mutations (see Table 1).

Figure 7

Hybrid fitness for later-generation crosses, showing the effects of selection acting on the earlier hybrids. Plots show the squared distance from the optimum, r^2, on a reverse axis, so that higher points are fitter. After the initial F1 cross, we simulated either random union of gametes among the hybrids (F2-F5), or repeated backcrossing to parental line P1 (BC1-BC4). For the later crosses, we chose parents either wholly at random (black points and lines), or with a probability proportional to their fitness (asterisks and red points and lines). In each case, results for 10,000 simulated hybrids (mean and 95% quantiles), are compared to analytical predictions from eq. 16 (blue lines), assuming optimal parental phenotypes ($r_{.O} = 0$). This is equivalent to the upper bound of eq. 26. When the variance in parental fitness was low (panel (a)), selection had little effect. When variance was high (panel (b)), results for the later generations approach predictions for extreme truncation selection, such that only optimal parents reproduce (red lines; lower bound of eq 26). Panels (c)-(d) show equivalent results, when populations were simulated with variable phenotypic dominance. The clearest consequence is that the F1 are suboptimal, even when the parental lines are optimal (blue and red lines used eqs. 29 and 31 with the observed r_{HO}^2, and $v = 1/12$, from the variance of a uniform distribution). All panels used simulations with the following parameters: $N = 1000$, $N\bar{s} = -0.1$, $2NU = 1$, $n = 2$, free recombination, "bottom-up" mutations, stationary optima matching the ancestral state; hybrids were formed as soon as one of the diverging populations had fixed 1000 substitutions.

Table 1: Parameter values used in the simulations

Parameter	#Conditions	Description
1. Population genetic parameters:		
Recombination rate	2	Free recombination ($\bar{c} = 0.5$), or a single chromosome of map length 100 cM ($\bar{c} = 0.216..$)
Mutation rate, U	4	Values chosen so that $2NU = 10, 1, 0.1$ or 0.01 when $N = 1000$
2. Fitness landscape parameters:		
Curvature, k	2	$k = 2$ or 6 (Gaussian or "table-like"; see eq. 5)
Number of "traits", n	2	$n = 2$ or 20
3. Distribution of new mutations:		
Homozygous effects:	2	Either "bottom-up" (i.i.d. normal on each trait), or "top-down" (random orientations and leptokurtic magnitudes)
	2	Variance such that $E(N\bar{s}) = -10$, or -0.1, in an optimal phenotype with $N = 1000$ (effective/ineffective selection)

Table 2: Composite effects under Fisher's geometric model

Composite effect		Geometric model prediction			
		one environment	two environments	local adaptation	phenotypic dominance
$\{\alpha_1\}$	additive	$\frac{1}{2}(r_{1O}^2 - r_{2O}^2)$	$\frac{1}{2}\left(r_{1\cdot}^2 - r_{2\cdot}^2\right)$	0	$\frac{1}{2}(r_{1O}^2 - r_{2O}^2) - \frac{1}{4}(r_{1H}^2 - r_{2H}^2)$
$\{\delta_1\}$	dominance	$-\frac{1}{2}$	$-\frac{1}{2}$	$-\frac{1}{2}$	$-\frac{1}{2}\left(1 + r_{PO}^2 - r_{HO}^2\right)$
$\{\alpha_2\}$	additive-by-additive epistasis	$\frac{1}{4}r_{12}^2 - 1$	$\frac{1}{4}r_{12}^2 - 1$	$\frac{1}{4}r_{12}^2 - 1$	$\frac{1}{4}r_{12}^2 - 1$
$\{\delta_2\}$	dominance-by-dominance	0	0	0	$\frac{1}{4}r_{HP}^2 - v$
$\{\alpha_1\delta_1\}$	additive-by-dominance	0	0	0	$-\frac{1}{4}(r_{1H}^2 - r_{2H}^2)$
$\{\varepsilon\}$	environmental	-	$\frac{1}{2}(r_{\cdot B}^2 - r_{\cdot A}^2)$	0	-
$\{\alpha_1\varepsilon\}$	additive-by-environmental	-	$\frac{1}{4}\left[(r_{1B}^2 - r_{1A}^2) - (r_{2B}^2 - r_{2A}^2)\right]$	$\frac{1}{2}r_{12}^2$	-
$\{\delta_1\varepsilon\}$	dominance-by-environmental	-	0	0	-
$\{\mu_0\}$	(intercept)	$\frac{1}{2} + r_{\cdot O}^2 - \frac{1}{4}r_{12}^2$	$\frac{1}{2} + r_{\cdot\cdot}^2 - \frac{1}{4}r_{12}^2$	$\frac{1}{2} + \frac{1}{4}r_{12}^2$	$\frac{1}{2} + v + r_{\cdot O}^2 - \frac{1}{4}r_{12}^2 - \frac{1}{4}r_{HP}^2 + \frac{1}{2}\left(r_{HO}^2 - r_{PO}^2\right)$