

# Environment affects specialisation of plants and pollinators

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## 6 Abstract

7 What determines whether or not a species is a generalist or a specialist? Evidence that the  
8 environment can influence species interactions is rapidly accumulating. However, a systematic link  
9 between environment and the number of partners a species interacts with has been elusive so far.  
10 Presumably, because environmental gradients appear to have contrasting effects on species depending  
11 on the environmental variable. Here, we test for a relationship between the stresses imposed by  
12 the environment, instead of environmental gradients directly, and species specialisation using a  
13 global dataset of plant-pollinator interactions. We found that the environment can play a significant  
14 effect on specialisation, even when accounting for community composition, likely by interacting  
15 with species' traits and evolutionary history. Species that have a large number of interactions are  
16 more likely to focus on a smaller number of, presumably higher-quality, interactions under stressful  
17 environmental conditions. Contrastingly, the specialists present in multiple locations are more likely  
18 to broaden their niche, presumably engaging in opportunistic interactions to cope with increased  
19 environmental stress. Indeed, many apparent specialists effectively behave as facultative generalists.  
20 Overall, many of the species we analysed are not inherently generalist or specialist. Instead, species'  
21 level of specialisation should be considered on a relative scale depending on where they are found  
22 and the environmental conditions at that location.

23 **Keywords:** eltonian niche, environmental effects, generalisation & specialisation, species degree,  
24 species interactions, and trophic niche

## 25 Introduction

26 Species interactions are known to vary widely across space and time. There are multiple examples of  
27 species that interact with a large number of partners in a particular community or season, but with  
28 fewer in another. Some of this variation can be attributed to environmental drivers. However, how  
29 exactly the environment, specifically the stress it imposes on species, affects whether two species  
30 interact or not, and ultimately the species' specialisation. Understanding how the environment  
31 drives the number of partners is crucial because it underpins the species' role in its community and  
32 shapes the structure of the network of interactions. This structure, in turn, determines ecosystem  
33 function and stability.

34 Species interactions are determined in part by niche processes (the matching of traits) and partly  
35 by neutral processes (more abundant species are more likely to encounter each other and, thus,  
36 interact). The environment can influence both of these processes. It is, therefore, not surprising  
37 that, despite limitations on the spatial extent or the number of environmental gradients considered,  
38 multiple studies have been able to show how changes to interactions can be related to environmental  
39 change (Tylianakis and Morris 2017). For instance, some studies suggest that the strength of some  
40 trophic interactions, like predation (McKinnon et al. 2010; Vucic-Pestic et al. 2011) and herbivory  
41 (Baskett and Schemske 2018), can increase with temperature but might decrease with precipitation  
42 (Pires et al. 2016). Some other studies, however, have shown either no effect (on average) or  
43 non-linear effects of temperature or precipitation on plant-pollinator interactions (Devoto, Medan,  
44 and Montaldo 2005; Gravel et al. 2018). Overall, while it looks clear that pairwise interactions  
45 respond to environmental drivers, there is high variability in the response (Tylianakis et al. 2008).

46 One possible explanation for the seemingly contradictory evidence is that different bioclimatic factors  
47 (like temperature or precipitation) can have contrasting effects on species and their partners. Here we  
48 attempt to simplify this situation by reducing multiple factors into a single measure of environmental  
49 stress. Previous research suggests that environmental stress may affect the number of partners in  
50 different ways depending on its role in the community (for example its trophic guild) or even the  
51 species itself. Specifically, we propose two alternative hypotheses of how environmental stress may  
52 affect specialisation (Tylianakis and Morris 2017). On the one hand, it is possible that when species

53 are under environmental stress, they might be “pressured” to focus on partners with which they are  
54 best adapted to interact. For instance, Hoiss et al. (2012) found increased phylogenetic clustering  
55 between plants and pollinators at higher altitudes; while Peralta et al. (2015) found that parasitoids  
56 in plantation forest, where environmental stress was higher than in native forests, were constrained  
57 to interact with hosts, they were best adapted to attack. Similarly, Lavandero and Tylianakis (2013)  
58 showed that environmental stress due to higher temperature reduced the trophic niche breadth of  
59 parasitoids suggesting higher specialisation.

60 On the other hand, it is also possible that when species are under environmental stress, they are  
61 forced to be more flexible in their interactions. Higher environmental stress is likely to be reflected  
62 in greater energetic or reproductive costs. Therefore they might not be able to sustain encounter  
63 rates with their preferred partners at sufficient levels. In line with this hypothesis, Hoiss, Krauss,  
64 and Steffan-Dewenter (2015) found that the specialisation of plant-pollinator networks decreased  
65 both with elevation and after extreme drought events. Likewise, Pellissier et al. (2010) found a  
66 positive relationship between niche breadth and environmental stress: disk- or bowl-shaped blossoms  
67 (which allow a large number of potential pollinator species to access pollen and nectar rewards)  
68 dominated at high altitude flower communities.

69 Here, we investigate whether and how environmental stress can systematically affect specialisation.  
70 Our main aim is to test the two hypotheses mentioned above that relate environmental stress and  
71 species’ number of partners and investigate whether this changes across species or between trophic  
72 guilds. We propose that specialist species can become “facultative” generalists to reduce their  
73 vulnerability to the absence of preferred partners (for example, when variations in climate decouple  
74 phenologies; Benadi et al. 2014). In other words, we expect that, as environmental stress increases,  
75 specialists should be more likely to engage with more partners. Species with a large number of  
76 partners, on the other hand, should have a larger pool of available partners and might, therefore,  
77 be more likely to specialise under environmental stress and focus on the most beneficial partners.  
78 Importantly, when testing these hypotheses, we control for the potential effects of the environment  
79 in community composition (which has been previously shown to be a determinant factor; Gravel et  
80 al. 2018). We test these hypotheses using data on plant-pollinator interactions. These interactions  
81 provide a particularly interesting system to test these hypotheses because, due to the multiple

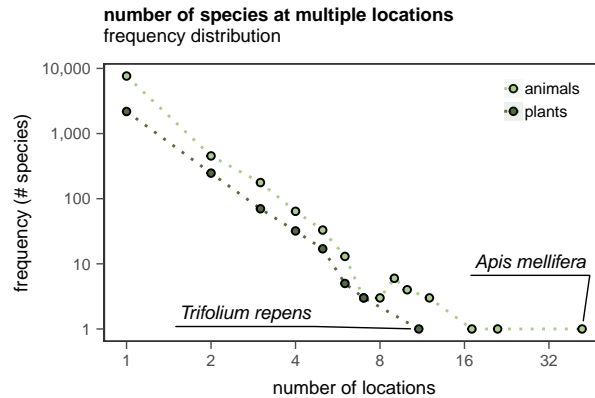


Figure 1: Frequency distribution of the number of locations in which a species is present. The most common pollinator species was *Apis mellifera*, which was sampled at 42 locations, while the most common plant species was *Trifolium repens*, which was sampled at 11 locations.

82 trade-offs involved in the pollination service, there are multiple intuitive ways in which we could  
83 imagine species respond to environmental stress given the available partners. We estimate the  
84 stress species might experience in their community by calculating the bioclimatic suitability of their  
85 communities given the species' patterns of global occurrence.

## 86 Methods

87 We retrieved plant-pollinator networks from the Web of Life database (Fortuna, Ortega, and  
88 Bascompte 2014). This database contains datasets originating from 57 studies published in the  
89 primary literature between 1923 and 2016. Calculating the environmental stress of species in  
90 their community and their potential partners required us to reduce both the taxonomic and  
91 distributional/location uncertainty. A critical step towards reducing this uncertainty is to ensure  
92 that the names used to identify species are valid and unambiguous, which in turn allow us to  
93 obtain further information from biological databases and accurately match species across studies.  
94 Therefore, our first step was to ensure consistent spelling and standardisation of species names  
95 synonyms (see Supplementary Methods). The cleaning process resulted on a total of 2,555 plants  
96 and 8,406 pollinator species distributed across 73 locations around the globe (Figure 1 and S1).

97 After matching species across studies as accurate as possible, we carried on two more steps. First,  
98 we calculated the environmental stress of species in their communities. Second, we relate the species

99 stress in their community with the number of partner species it has as a metric of their level of  
100 specialisation.

## 101 **Environmental stress**

102 We calculated the environmental stress of species in their communities. We assume that stress  
103 a species experiences in a particular location is inversely related to the suitability of the average  
104 environmental conditions in that place. As we aim to compare specialisation levels for different levels  
105 of environmental stress, we only calculate bioclimatic suitability for species that were present in at  
106 least two communities. To calculate the bioclimatic suitability of a species in a particular location,  
107 we used a niche-factor analysis (Hirzel et al. 2002; Broennimann et al. 2012). This approach is  
108 based on the probability density function of species distribution in an environmental variable space.  
109 Habitats are characterised by a collection of environmental variables. In a nutshell, those habitats  
110 in which the species occurs more often are deemed to be more suitable for the species than habitats  
111 in which the species has never been observed. As bioclimatic suitability is calculated in a scale from  
112 zero to one following the niche-factor analysis, for simplicity, we define environmental stress as one  
113 minus suitability.

114 The niche factor analysis requires two critical pieces of information. First, it requires information  
115 about the occurrences of the species of interest. Second, the method requires information about the  
116 environmental conditions for all the locations in which the species occurs. We retrieved 38.1 million  
117 occurrences from the Global Biodiversity Information Facility (GBIF; <https://www.gbif.org>). Issues  
118 with data quality are a central issue hampering the use of publicly available species occurrence  
119 GBIF data in ecology and biogeography (Jetz et al. 2019). We, therefore, followed a series of filters  
120 and geographic heuristics to correct or remove erroneous and imprecise referencing records (see  
121 supplementary methods; Zizka et al. 2019) which allowed us to identify and remove 7.5 million  
122 potentially problematic occurrences from further analysis. We integrated the occurrences from our  
123 plant-pollinator communities to the cleaned occurrences retrieved from GBIF.

124 We retrieved environmental data from WorldClim V2.0, which includes 19 bioclimatic variables  
125 commonly used in species distribution modelling (Fick and Hijmans 2017). We then complemented

126 data obtained from WorldClim with data from Envirem (Title and Bemmels 2017), which includes  
127 16 extra bioclimatic and two topographic variables. The additional set of variables from Envirem  
128 are relevant to ecological or physiological processes and thus have the potential to improve our  
129 suitability estimation (Title and Bemmels 2018). We obtained all environmental data as rasters  
130 composed by cells of 2.5 arc-minutes. We chose this resolution because it provides a reasonable  
131 match to the locational accuracy of the species occurrences found in GBIF, particularly those that  
132 originate from preserved specimens in museum collections.

133 After obtaining information about species occurrence and the environment, we then merged these two  
134 datasets such that a vector with details of our 37 bioclimatic and topographic variables characterised  
135 the location of each occurrence. Sets of occurrence data tend to be spatially aggregated due to  
136 sample bias (tendency to collect close to cities, certain countries). Moreover, spatial autocorrelation  
137 arises in ecological data because geographically clumped records tend to be more similar in physical  
138 characteristics and/or species abundances than do pairs of locations that are farther apart. To  
139 account for such spatial dependency in occurrence data, we only included one occurrence record  
140 if a species had more than one within a cell of the bioclimatic raster. We did this to avoid giving  
141 more weight to areas with a high number of occurrences, a common scenario in occurrence records  
142 collected opportunistically as the ones we use here. In this step we removed 85.4% of the occurrences  
143 which resulted in a total of 4.5 million occurrences used in our niche analysis.

144 A common issue of terrestrial bioclimatic datasets is that the boundaries of the cells with information  
145 do not precisely match the landmass boundaries. The result of this mismatch is that not all  
146 environmental variables were available for 3,273 of the raster cells with occurrences (0.8% of the  
147 total). As expected, the vast majority of these problematic cells were close to the shore. To address  
148 this issue, we calculated the average value of environmental variables within a 5km buffer of the  
149 centre of the cell where the variable was missing and used it to approximate the value of the variable  
150 in that cell. Using this procedure, we were able to fill environmental variables for 89.3% of the cells  
151 where they were missing. To fill the remaining 350 cells, we repeated the aforementioned procedure  
152 but instead using a 10km buffer. We removed from further analysis occurrences located within the  
153 135 cells for which we were unable to fill environmental variables (0.03% of the total).

154 Next, we calculated the probability density function of the species distribution in environmental space.

155 To determine the environmental space, we used the first two components from a principal component  
156 analysis of the 37 bioclimatic variables associated with the species occurrences. Specifically we used  
157 the `dudi.pca` function from the R package `ade4` 1.7.13 (Dray and Dufour 2007) and center and  
158 scale all bioclimatic variables to have a mean of zero and a unit variance. We then determined the  
159 position of species occurrences in the environmental space and estimate their bivariate probability  
160 density function. We used a kernel method to estimate this density and normalised it such that  
161 it ranges between zero and one. We used the kernel density method in the niche-factor analysis  
162 (Broennimann et al. 2012) rather than the distance from the mode (Hirzel et al. 2002) (as it has  
163 been proposed earlier) because it has been shown to reduce the procedure's sensitivity to sampling  
164 effort and the resolution of the environmental space. Specifically, to calculate the probability density  
165 function we used `ecospat.grid.clim.dyn` from the R package `ecospat` 3.0 (Broennimann, Di  
166 Cola, and Guisan 2018) with a grid resolution of 200. We then determined the location in the  
167 environmental space of the plant-pollinator communities using the function `suprow` from `ade4`. The  
168 normalised density at that particular location (which we calculated using the R package `raster`  
169 2.8.19; Hijmans 2019) corresponds the bioclimatic suitability. The result of all these steps is the  
170 environmental stress which corresponds to one minus the bioclimatic suitability for a species of a  
171 particular location.

172 We used a sensitivity analysis to determine the minimum number of occurrences that are necessary  
173 to have robust environmental stress estimations. For that we used the species with most occurrences  
174 available, *Archilochus colubris*, and calculated the mean absolute error of the bioclimatic suitability  
175 values obtained with one thousand subsamples from the 74,791 occurrences available from GBIF.

## 176 **Data analysis**

177 We then used a set of Bayesian multilevel models to evaluate the impact of environmental stress  
178 on species specialisation. Specifically, we use the normalised degree of species as our response  
179 variable; that is, the number of species it interacts with given the number of species in the opposite  
180 guild (Martín González, Dalsgaard, and Olesen 2010). The normalised degree was modelled using  
181 a logit link function, and a binomial distribution in which the number of partner species a focal  
182 interacts with is the number of successes, and the number of species in the opposite guild is the



183 number of trials. We are aware that whether species interact or not is not a Bernoulli process as  
184 species interactions are not strictly independent from each other. However, the use of a binomial  
185 distribution allows us to account for the differences in species richness across communities indirectly.  
186 Importantly, results are qualitatively similar when we model species degree directly using a Poisson  
187 distribution and a logarithmic link function.

188 We evaluated four models to assess the relative importance of suitability. A first model, our baseline  
189 model, included five variables. The predictors in the baseline model were the environmental stress,  
190 its number of known possible partners in the community, and both the species guild (plant or a  
191 pollinator) and its interaction with environmental stress. We included the number of known possible  
192 partners as a predictor in our models as it allows us to control for the effects of the environment  
193 on community composition, effectively accounting for species co-occurrence. We calculated this  
194 metric by determining the number of partners with which the species is known to interact in any  
195 other community. Controlling for the number of potential partners makes our model a particularly  
196 stringent test of our environmental-stress hypotheses because this variable could explain a large  
197 proportion of variance. Often, the potential and the actual number of partners is the same or very  
198 close to each other, especially for rare species present only in a few communities.

199 We allowed the intercept and slope of the stress-specialisation relationship to vary among species.  
200 This approach allowed us to investigate two questions. First, it allows us to inspect the extent to  
201 which environmental stress affects species similarly. Second, by investigating the correlation between  
202 the intercept and the slope as a model parameter, it allowed us to inspect the extent by which  
203 species with a small or large number of partner species respond to increasing levels of environmental  
204 stress. To account for unmeasured differences between communities, like sampling effort, sampling  
205 method, or diversity, we also allowed the model intercept to be different for each community in our  
206 study. To facilitate model interpretation and convergence, we scaled all continuous variables to have  
207 a mean of zero and a unit variance.

208 We compared this baseline model with three alternative models in which we removed one predictor  
209 at a time. To quantify the difference between models, in terms of their expected out-of-sample  
210 performance, we use the Wanatabe-Akaike information criterion (WAIC). All models were fitted  
211 under a Bayesian framework using the R package `brms` 2.8.0 (Bürkner 2017, 2018) as an interface

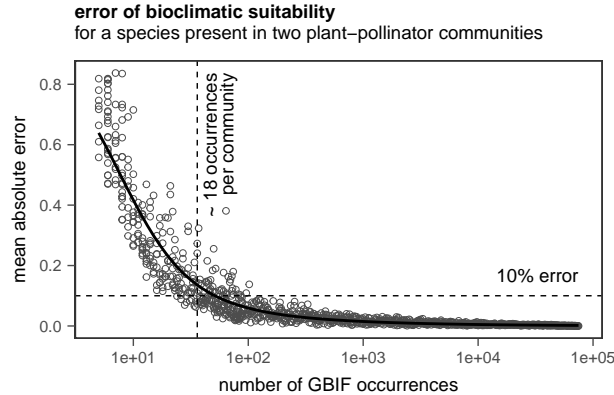


Figure 2: Sensitivity analysis of environmental stress error. The number of independent occurrences retrieved from GBIF is inversely related to the error of bioclimatic suitability for our plant-pollinator networks. The sensitivity analysis was performed by subsampling occurrences of *Archilochus colubris* the species in our dataset with the largest number of occurrences in GBIF, which was recorded in two of our communities.

212 for Stan (Carpenter et al. 2017). For each model, we used four Markov chains of 4,000 iterations  
213 each; we used half of the iterations for warmup. We used weakly informative priors for all model  
214 parameters. Specifically we used normal priors of mean zero and standard deviation ten for the  
215 population-level effects and the intercepts, a half-Cauchy prior with a location of zero and a scale of  
216 two for the standard deviations, and, when applicable, an LKJ-correlation prior with parameter  
217  $\zeta = 1$  for the correlation matrix between group-level parameters.

## 218 Results

219 After performing our sensitivity analysis, we found that, for a species, we need roughly 18 independent  
220 occurrences for each community for which we aim to estimate the environmental stress. This is  
221 the number of occurrences necessary to maintain the mean absolute error of bioclimatic suitability  
222 below 0.1 (Fig. 2). We therefore removed from further analyses 286 species for which we did not  
223 have enough occurrences to obtain robust estimates.

224 Our models performed relatively well. The Bayesian  $R^2$  for our baseline model was 0.89, which  
225 indicates our models were able to capture a large proportion of the variability on the data. Overall,  
226 we found that environmental stress does not have a consistent effect across species. Indeed, when  
227 looking at the fixed effects, stress has virtually no relationship with the normalised degree—our

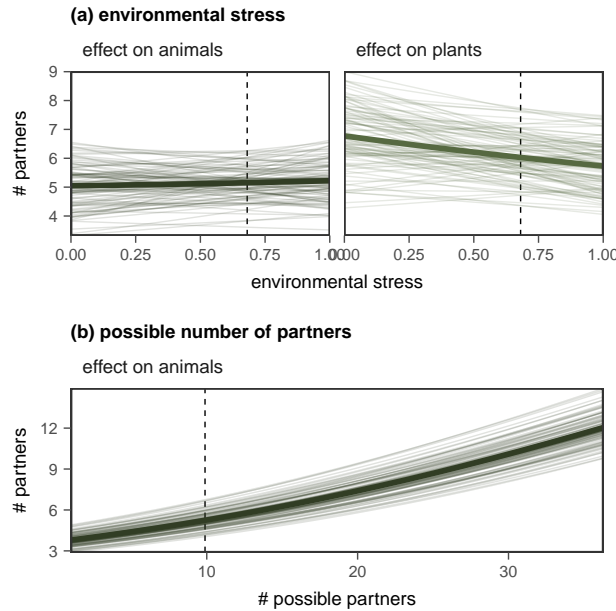


Figure 3: Conditional effects of predictors in our baseline model. The shown values are based on predictions for a hypothetical community with 76 plants and 33 pollinators. These values correspond to the median number of species in each guild across communities. In each panel, we condition on the mean value of the other predictor in the model. We indicate mean values for each predictor with a vertical dashed line. For model fitting, we scaled all predictors to have a mean of zero and unit variance; however, here we show the unscaled predictors to facilitate interpretation. To illustrate the uncertainty around the fitted estimates, we plot the fits of 100 independent draws from the posterior distribution. The thick lines indicate the mean values of the response distribution. As there was no interaction between the guild and the number of possible interactions, we only show the conditional effect of pollinators.

228 metric of specialisation (Figure 3a). However, environmental stress was still an important predictor  
229 in our model. The difference in WAIC between our baseline model and the model that did not  
230 include environmental stress was  $489 \pm 94$  (Table 1). This apparent discrepancy can be explained  
231 by the variability of the specialisation-stress relationship across species.

232 For some species, there is a strong negative relationship between stress and specialisation, while for  
233 others, there is a strong positive relationship (Figure 4a). Interestingly, the slope of this relationship  
234 correlates with the species' intercept in the model (Figure 4b and c). Recall that the model estimates  
235 the intercept at the mean value for stress across communities (0.68). The mean correlation coefficient  
236 was 0.52 [0.33, 0.67]. Therefore, the slope of the stress-specialisation relationship was more likely to  
237 be positive for species with a large number of partners under average stress conditions (and negative  
238 for species with a smaller number of partners). Extrapolating to no-stress conditions: species that

Table 1: Comparison in out of sample predictive power of the baseline model (bold) and their alternatives. We rank models by their expected log predictive density based on their Wanatabe-Akaike information criterion (WAIC).

predictors	WAIC	SE
<b>stress x guild + # possible partners</b>	6,592	170
stress + # possible partners	6,595	166
guild + # possible partners	7,081	202
stress x guild	8,041	290

239 would interact with a small number of partners under no stress are more likely to interact with  
240 more partners as stress increases, whereas those that would interact with a large number of partners  
241 are more likely to interact with less.

242 As expected, we found a strong and positive relationship between the number of possible interactions  
243 and the number of realised interactions in the community. There was also a large difference of WAIC  
244 between the model that included this predictor and that that excluded it. This result indicates  
245 that the availability of potential partners—this is, community composition—accounts for a large  
246 proportion of the variability in species degree. Importantly, our findings relating to the variability of  
247 the stress-specialisation relationship were qualitatively unchanged, whether we included this variable  
248 or not.

249 The standard deviation (in the parameters scale) of the community intercepts was 1.02 [0.85, 1.23]  
250 which indicates the importance of the local context when determining specialisation. The standard  
251 deviation of the species intercept was 0.54 [0.48, 0.61], and that of the species' stress slope was 0.38  
252 [0.32, 0.44] (95% credible intervals shown within square brackets).

## 253 Discussion

254 We set out to explore whether and how environmental stress can systematically affect specialisation.  
255 After accounting for the pool of potential partners, we found that environmental conditions contribute  
256 to determining whether a species is a generalist or a specialist *in their community*. We also found  
257 that the particular effect of the environment is strongly dependent on the species. Based on existing  
258 literature, we proposed two alternative hypotheses of how environmental stress may affect species'  
259 specialisation, and we found evidence for both of them. Species with a large number of partners

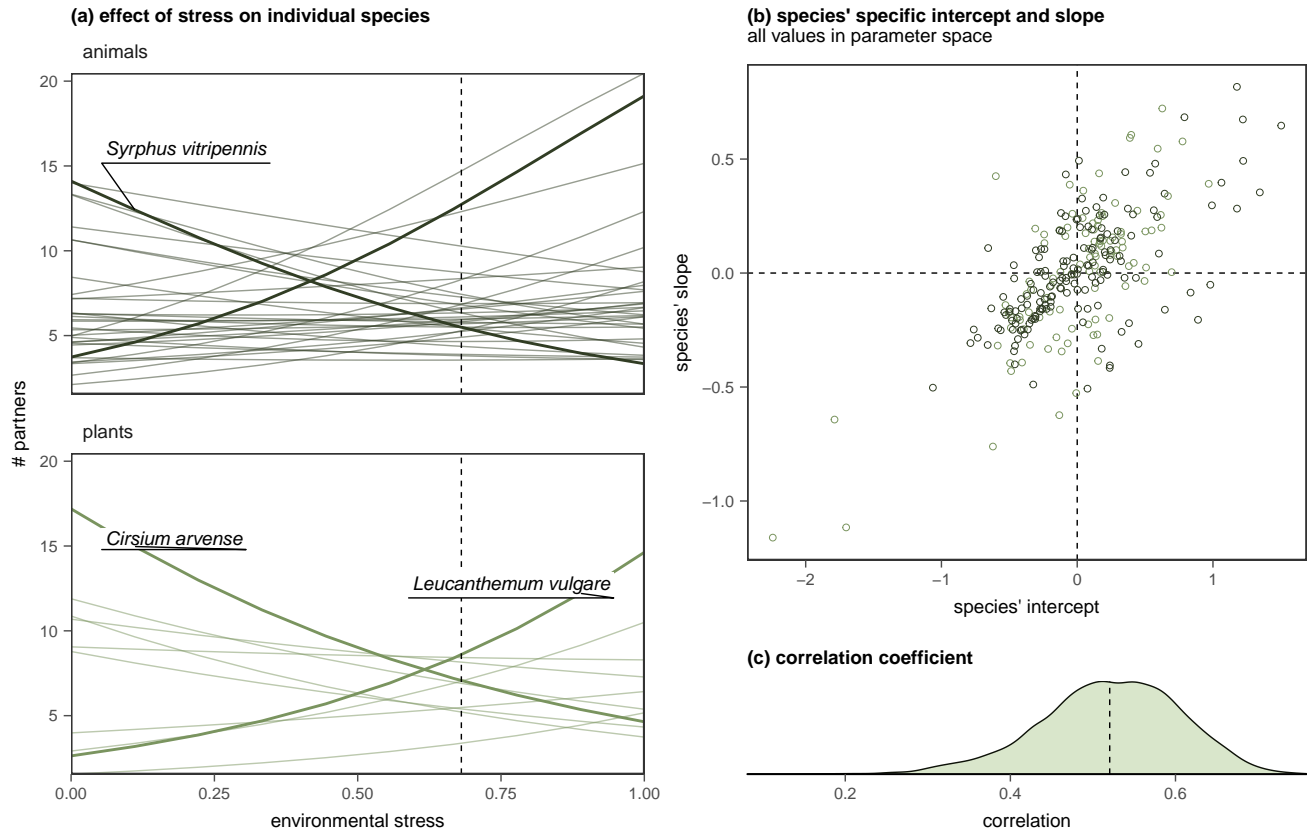


Figure 4: Species-level effects of environmental stress (a) Conditional effect of stress for individual species. Each line corresponds to the median relationship for each species. Although we included in the analysis of all species that are present in two or core communities, to facilitate visualisation here, we show only species for which there is suitability information in at least six communities (10 plants and 33 pollinators). As in the previous figure, fitted values assume a hypothetical community of median size. In each panel, we highlight two species for which the relationship between environmental suitability and the normalised degree was particularly strong. (b and c) The correlation between the species' intercept and the species' slope of suitability was negative. The species' intercept can be interpreted as the relative difference between the number of partners a species has under mean levels of environmental stress and the mean number of partners across all species. Positive values of species' slope indicate a positive relationship between stress and the number of partners and vice-versa.

260 in low-stress communities were more likely to have a negative relationship and hence reduce the  
261 number of partners as stress increases. Contrastingly, species in our datasets with a small number  
262 of partners in low-stress communities were more likely to have a larger number of partners in more  
263 stressful communities. In summary, environmental stress pushes species that are flexible enough to  
264 change their interaction partners towards intermediate levels of specialisation, a so-called “regression  
265 towards the mean”.

266 Our results suggest that changes in community composition are indeed the primary channel through  
267 which the environment determines changes interaction probability. However, they also show that,  
268 for a large number of species, the environment may also play a substantial role in determining  
269 their level of specialisation. Previous research has recognised that environmental factors may help  
270 explain the changes in network structure along environmental gradients that cannot be explained by  
271 community composition (Tylianakis, Tscharntke, and Lewis 2007). However, how these two factors  
272 were linked had been elusive so far (Gravel et al. 2018). We believe that part of this difficulty  
273 could have arisen because species, and ultimately network structure, can respond in multiple, and  
274 contrasting, ways depending on the particular bioclimatic variable examined (e.g. temperature or  
275 precipitation). Using stress to summarise the effect on species of multiple environmental gradients  
276 allowed us to detect a clear signal of the environment in species’ interaction patterns.

277 Although both niche and neutral processes are relevant in determining species interactions, our model  
278 suggests that niche processes may be the predominant mechanism through which the environment  
279 *systematically* affects specialisation. First, it is unlikely that environmental stress correlates to local  
280 species abundances (Pearce and Ferrier 2001; Sagarin, Gaines, and Gaylord 2006). Second, even if  
281 there is a relationship between stress and abundances, a particular environmental gradient could  
282 have a positive effect on the abundance of some species and a negative effect on others. Indeed, we  
283 find that within a community there is a wide range of stress values, even for the relatively limited  
284 number of species we were able to include in our analysis.

285 Recent research suggests that species are continuously changing their interaction partners wherever  
286 environmental conditions change in space or time (Raimundo, Guimarães, and Evans 2018). So far  
287 it appears that this rewiring is primarily driven by generalist species (Ponisio, Gaiarsa, and Kremen  
288 2017; Burkle, Marlin, and Knight 2013), presumably because generalist species are less sensitive

289 to trait matching of their interaction partners (CaraDonna et al. 2017). Our results add two  
290 important nuances to these findings. First, because “generalists” seem to focus on a smaller number  
291 of partners as environmental conditions deteriorate, we show that trait matching might still play a  
292 role in determining the interactions of generalist species. Second, and most importantly, our results  
293 suggest that only a small proportion of species are “true generalists” or “true specialists” this is,  
294 species that interact with a large or small number of partners regardless of the environmental stress,  
295 respectively. This pattern implies that rewiring is not exclusive of species with a large number of  
296 partners. Instead, at least a fraction of the species that appear to be specialist *in their communities*  
297 might be as flexible, if not more, than those with a large number of partners, effectively behaving  
298 as facultative generalists in the face of environmental change. These flexible “specialists” might  
299 therefore have a more significant role in network persistence than previously expected.

300 In our model, we can roughly divide species between true specialists, true generalists, and flexible  
301 species. However, there is a fourth group that remained invisible to our model but has important  
302 implications for network persistence and stability. Species that can vary their interaction partners  
303 flexibly and their role in the network are more likely to persist in their community as environmental  
304 conditions vary (Gaiarsa, Kremen, and Ponisio 2019). We propose this fourth group is composed of  
305 true specialists that are constrained to interact with partners of high trait-matching and therefore  
306 were not likely to be found in more than one community. If species that are not flexible are  
307 unlikely to persist over temporal or spatial environmental gradients, we can expect specialised  
308 communities that are highly constrained by trait-matching (like some plant-hummingbird networks;  
309 Vizenin-Bugoni, Maruyama, and Sazima 2014; Maruyama et al. 2014) to be far more vulnerable to  
310 increased climate change-induced environmental stress and habitat degradation than communities  
311 where role and interaction flexibility are more prevalent.

312 Similarly, if the patterns we see in our models have also played a role during the evolutionary  
313 history of pollination communities, our results also help explain why only a small fraction of  
314 plant-pollinator interactions show a strong signature of deep co-evolutionary history (Hutchinson,  
315 Cagua, and Stouffer 2017). The increases in the stress that species are predicted to experience due  
316 to rapid environmental change might further erode the co-evolutionary history of specialist species.  
317 Communities as a whole might be in a trajectory of even more diffuse co-evolution. For specialists,

318 at least, the longer-term benefits of being able to interact with multiple partners might be more  
319 important than the shorter-term benefits of interacting with partners of high trait matching.

320 The structural implications of the “regression towards the mean” that environmental stress promotes  
321 are less clear. However, it is plausible to expect that nestedness, and therefore network stability,  
322 might be reduced in the face of rapid environmental change. Determining exactly how the changes in  
323 degree caused by environmental stress reflect on systematic changes in network structure would be  
324 an interesting avenue of research. Answering this question would require expanding our suitability  
325 analysis to all species in the community and compare the degree distribution of networks along a  
326 gradient of stress for the community as a whole.

327 In conclusion, we show that the environment can affect the specialisation level of plants and  
328 pollinators in systematic ways beyond community composition. Species that are inflexible with their  
329 interaction partners are unlikely to persist under more stressful environmental conditions. However,  
330 we show that many species are flexible in regards to their specialisation levels and therefore are  
331 not inherently generalists or specialists. Instead, the species’ level of specialisation/generalisation  
332 should be considered on a relative scale depending on where they are found and the environmental  
333 conditions at that location.

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## 341 References

- 342 Baskett, Carina A., and Douglas W. Schemske. 2018. “Latitudinal Patterns of Herbivore Pressure  
343 in a Temperate Herb Support the Biotic Interactions Hypothesis.” Edited by Vojtech Novotny.  
344 *Ecology Letters* 21 (4): 578–87. <https://doi.org/10.1111/ele.12925>.
- 345 Benadi, Gita, Thomas Hovestadt, Hans-Joachim Poethke, and Nico Blüthgen. 2014. “Specialization  
346 and Phenological Synchrony of Plant-Pollinator Interactions Along an Altitudinal Gradient.” Edited  
347 by Thomas Ings. *Journal of Animal Ecology* 83 (3): 639–50. [https://doi.org/10.1111/1365-2656](https://doi.org/10.1111/1365-2656.12158).  
348 12158.
- 349 Broennimann, Olivier, Valeria Di Cola, and Antoine Guisan. 2018. *Ecospat: Spatial Ecology*  
350 *Miscellaneous Methods*.
- 351 Broennimann, Olivier, Matthew C. Fitzpatrick, Peter B. Pearman, Blaise Petitpierre, Loïc Pellissier,  
352 Nigel G. Yoccoz, Wilfried Thuiller, et al. 2012. “Measuring Ecological Niche Overlap from Occurrence  
353 and Spatial Environmental Data: Measuring Niche Overlap.” *Global Ecology and Biogeography* 21  
354 (4): 481–97. <https://doi.org/10.1111/j.1466-8238.2011.00698.x>.
- 355 Burkle, L. A., J. C. Marlin, and T. M. Knight. 2013. “Plant-Pollinator Interactions over 120 Years:  
356 Loss of Species, Co-Occurrence, and Function.” *Science* 339 (6127): 1611–5. [https://doi.org/10](https://doi.org/10.1126/science.1232728).  
357 1126/science.1232728.
- 358 Bürkner, Paul-Christian. 2017. “Brms: An R Package for Bayesian Multilevel Models Using Stan.”  
359 *Journal of Statistical Software* 80 (1). <https://doi.org/10.18637/jss.v080.i01>.
- 360 ———. 2018. “Advanced Bayesian Multilevel Modeling with the R Package Brms.” *The R Journal*  
361 10 (1): 395. <https://doi.org/10.32614/RJ-2018-017>.
- 362 CaraDonna, Paul J., William K. Petry, Ross M. Brennan, James L. Cunningham, Judith L.  
363 Bronstein, Nickolas M. Waser, and Nathan J. Sanders. 2017. “Interaction Rewiring and the Rapid  
364 Turnover of Plant-Pollinator Networks.” Edited by Ferenc Jordan. *Ecology Letters* 20 (3): 385–94.  
365 <https://doi.org/10.1111/ele.12740>.
- 366 Carpenter, Bob, Andrew Gelman, Matthew D. Hoffman, Daniel Lee, Ben Goodrich, Michael

- 367 Betancourt, Marcus Brubaker, Jiqiang Guo, Peter Li, and Allen Riddell. 2017. “Stan : A  
368 Probabilistic Programming Language.” *Journal of Statistical Software* 76 (1). <https://doi.org/10.18637/jss.v076.i01>.
- 370 Devoto, Mariano, Diego Medan, and Norberto H. Montaldo. 2005. “Patterns of Interaction  
371 Between Plants and Pollinators Along an Environmental Gradient.” *Oikos* 109 (3): 461–72. <https://doi.org/10.1111/j.0030-1299.2005.13712.x>.
- 373 Dray, Stéphane, and Anne-Béatrice Dufour. 2007. “The Ade4 Package: Implementing the Duality  
374 Diagram for Ecologists.” *Journal of Statistical Software* 22 (4). <https://doi.org/10.18637/jss.v022.i04>.
- 375 Fick, Stephen E., and Robert J. Hijmans. 2017. “WorldClim 2: New 1-Km Spatial Resolution  
376 Climate Surfaces for Global Land Areas: NEW CLIMATE SURFACES FOR GLOBAL LAND  
377 AREAS.” *International Journal of Climatology* 37 (12): 4302–15. <https://doi.org/10.1002/joc.5086>.
- 378 Fortuna, Miguel A., Raul Ortega, and Jordi Bascompte. 2014. “The Web of Life.” *arXiv:1403.2575*  
379 *[Q-Bio]*, March. <http://arxiv.org/abs/1403.2575>.
- 380 Gaiarsa, Marilia P, Claire Kremen, and Lauren C Ponisio. 2019. “Interaction Flexibility Predicts  
381 Pollinator Population Dynamics.” *Manuscript Submitted for Publication*, 15.
- 382 Gravel, Dominique, Benjamin Baiser, Jennifer A Dunne, Jens-Peter Kopelke, Neo D Martinez,  
383 Tommi Nyman, Timothee Poisot, et al. 2018. “Bringing Elton and Grinnell Together: A Quantitative  
384 Framework to Represent the Biogeography of Ecological Interaction Networks.” *Ecography* 41: 1–15.  
385 <https://doi.org/10.1111/ecog.04006>.
- 386 Hijmans, Robert J. 2019. *Raster: Geographic Data Analysis and Modeling*.
- 387 Hirzel, A H, J Hausser, D Chessel, and N Perrin. 2002. “Ecological-Niche Factor Analysis: How to  
388 Compute Habitat-Suitability Maps Without Absence Data?” 83 (7): 10.
- 389 Hoiss, Bernhard, Jochen Krauss, Simon G. Potts, Stuart Roberts, and Ingolf Steffan-Dewenter.  
390 2012. “Altitude Acts as an Environmental Filter on Phylogenetic Composition, Traits and Diversity  
391 in Bee Communities.” *Proceedings of the Royal Society B: Biological Sciences* 279 (1746): 4447–56.  
392 <https://doi.org/10.1098/rspb.2012.1581>.

- 393 Hoiss, Bernhard, Jochen Krauss, and Ingolf Steffan-Dewenter. 2015. “Interactive Effects of Elevation,  
394 Species Richness and Extreme Climatic Events on Plant-Pollinator Networks.” *Global Change Biology*  
395 21 (11): 4086–97. <https://doi.org/10.1111/gcb.12968>.
- 396 Hutchinson, Matthew C., Edgar Fernando Cagua, and Daniel B. Stouffer. 2017. “Cophylogenetic  
397 Signal Is Detectable in Pollination Interactions Across Ecological Scales.” *Ecology* 98 (10): 2640–52.  
398 <https://doi.org/10.1002/ecy.1955>.
- 399 Jetz, Walter, Melodie A. McGeoch, Robert Guralnick, Simon Ferrier, Jan Beck, Mark J. Costello,  
400 Miguel Fernandez, et al. 2019. “Essential Biodiversity Variables for Mapping and Monitor-  
401 ing Species Populations.” *Nature Ecology & Evolution* 3 (4): 539–51. <https://doi.org/10.1038/>  
402 [s41559-019-0826-1](https://doi.org/10.1038/s41559-019-0826-1).
- 403 Lavandero, Blas, and Jason M. Tylianakis. 2013. “Genotype Matching in a Parasitoid-Host  
404 Genotypic Food Web: An Approach for Measuring Effects of Environmental Change.” *Molecular*  
405 *Ecology* 22 (1): 229–38. <https://doi.org/10.1111/mec.12100>.
- 406 Martín González, Ana M., Bo Dalsgaard, and Jens M. Olesen. 2010. “Centrality Measures and  
407 the Importance of Generalist Species in Pollination Networks.” *Ecological Complexity* 7 (1): 36–43.  
408 <https://doi.org/10.1016/j.ecocom.2009.03.008>.
- 409 Maruyama, Pietro K., Jeferson Vizentin-Bugoni, Genilda M. Oliveira, Paulo E. Oliveira, and Bo  
410 Dalsgaard. 2014. “Morphological and Spatio-Temporal Mismatches Shape a Neotropical Savanna  
411 Plant-Hummingbird Network.” *Biotropica* 46 (6): 740–47. <https://doi.org/10.1111/btp.12170>.
- 412 McKinnon, L., P. A. Smith, E. Nol, J. L. Martin, F. I. Doyle, K. F. Abraham, H. G. Gilchrist, R. I.  
413 G. Morrison, and J. Bêty. 2010. “Lower Predation Risk for Migratory Birds at High Latitudes.”  
414 *Science* 327 (5963): 326–27.
- 415 Pearce, Jennie, and Simon Ferrier. 2001. “The Practical Value of Modelling Relative Abundance of  
416 Species for Regional Conservation Planning: A Case Study.” *Biological Conservation* 98 (1): 33–43.  
417 [https://doi.org/10.1016/S0006-3207\(00\)00139-7](https://doi.org/10.1016/S0006-3207(00)00139-7).
- 418 Pellissier, Loïc, Julien Pottier, Pascal Vittoz, Anne Dubuis, and Antoine Guisan. 2010. “Spatial  
419 Pattern of Floral Morphology: Possible Insight into the Effects of Pollinators on Plant Distributions.”

- 420 *Oikos* 119 (11): 1805–13. <https://doi.org/10.1111/j.1600-0706.2010.18560.x>.
- 421 Peralta, Guadalupe, Carol M. Frost, Raphael K. Didham, Arvind Varsani, and Jason M. Tylianakis.  
422 2015. “Phylogenetic Diversity and Co-Evolutionary Signals Among Trophic Levels Change Across  
423 a Habitat Edge.” Edited by Frank van Veen. *Journal of Animal Ecology* 84 (2): 364–72. <https://doi.org/10.1111/1365-2656.12296>.
- 425 Pires, Aliny P. F., Nicholas A. C. Marino, Diane S. Srivastava, and Vinicius F. Farjalla. 2016.  
426 “Predicted Rainfall Changes Disrupt Trophic Interactions in a Tropical Aquatic Ecosystem.” *Ecology*  
427 97 (10): 2750–9. <https://doi.org/10.1002/ecy.1501>.
- 428 Ponisio, Lauren C., Marilia P. Gaiarsa, and Claire Kremen. 2017. “Opportunistic Attachment  
429 Assembles Plant-Pollinator Networks.” Edited by Dominique Gravel. *Ecology Letters* 20 (10):  
430 1261–72. <https://doi.org/10.1111/ele.12821>.
- 431 Raimundo, Rafael L.G., Paulo R. Guimarães, and Darren M. Evans. 2018. “Adaptive Networks for  
432 Restoration Ecology.” *Trends in Ecology & Evolution* 33 (9): 664–75. <https://doi.org/10.1016/j.tree.2018.06.002>.
- 434 Sagarin, Raphael D., Steven D. Gaines, and Brian Gaylord. 2006. “Moving Beyond Assumptions to  
435 Understand Abundance Distributions Across the Ranges of Species.” *Trends in Ecology & Evolution*  
436 21 (9): 524–30. <https://doi.org/10.1016/j.tree.2006.06.008>.
- 437 Title, Pascal O., and Jordan B. Bemmels. 2017. “ENVIREM: ENVironmental Rasters for Ecological  
438 Modeling Version 1.0.” University of Michigan. <https://doi.org/10.7302/Z2BR8Q40>.
- 439 ———. 2018. “ENVIREM: An Expanded Set of Bioclimatic and Topographic Variables Increases  
440 Flexibility and Improves Performance of Ecological Niche Modeling.” *Ecography* 41 (2): 291–307.  
441 <https://doi.org/10.1111/ecog.02880>.
- 442 Tylianakis, Jason M., Raphael K. Didham, Jordi Bascompte, and David A. Wardle. 2008. “Global  
443 Change and Species Interactions in Terrestrial Ecosystems.” *Ecology Letters* 11 (12): 1351–63.  
444 <https://doi.org/10.1111/j.1461-0248.2008.01250.x>.
- 445 Tylianakis, Jason M., and Rebecca J. Morris. 2017. “Ecological Networks Across Environmental  
446 Gradients.” *Annual Review of Ecology, Evolution, and Systematics* 48 (1): 25–48. <https://doi.org/>

447 10.1146/annurev-ecolsys-110316-022821.

448 Tylianakis, Jason M., Teja Tscharntke, and Owen T. Lewis. 2007. “Habitat Modification Alters the  
449 Structure of Tropical HostParasitoid Food Webs.” *Nature* 445 (7124): 202–5. <https://doi.org/10.1038/nature05429>.

451 Vizentin-Bugoni, Jeferson, Pietro Kiyoshi Maruyama, and Marlies Sazima. 2014. “Processes  
452 Entangling Interactions in Communities: Forbidden Links Are More Important Than Abundance in  
453 a HummingbirdPlant Network.” *Proceedings of the Royal Society B: Biological Sciences* 281 (1780):  
454 20132397. <https://doi.org/10.1098/rspb.2013.2397>.

455 Vucic-Pestic, Olivera, Roswitha B. Ehnes, Björn C. Rall, and Ulrich Brose. 2011. “Warming up the  
456 System: Higher Predator Feeding Rates but Lower Energetic Efficiencies: Warming and Functional  
457 Responses.” *Global Change Biology* 17 (3): 1301–10. <https://doi.org/10.1111/j.1365-2486.2010.02329.x>.

459 Zizka, Alexander, Daniele Silvestro, Tobias Andermann, Josué Azevedo, Camila Duarte Ritter,  
460 Daniel Edler, Harith Farooq, et al. 2019. “CoordinateCleaner: Standardized Cleaning of Occurrence  
461 Records from Biological Collection Databases.” Edited by Tiago Quental. *Methods in Ecology and*  
462 *Evolution* 10 (5): 744–51. <https://doi.org/10.1111/2041-210X.13152>.