Maternal style does not predict infant growth and survival in wild Guinea baboons

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Abstract

Primate infants depend on their mothers for nursing, transportation, and survival, but also benefit from other behaviours such as contact, protection, and the mother serving as a model for social learning. During infancy, variation in maternal care ("maternal style") can have profound consequences for infant development and survival. We investigated the link between maternal style and infant growth and survival in wild Guinea baboons (*Papio papio*). This species lives in a tolerant social system with female-biased dispersal. We used data from N = 80 infants to assess variation in maternal style. The final statistical models comprised data from N = 50 infants for the growth analysis based on Parallel Laser Photogrammetry (PLP) measures of forearm length, and N = 63 infants for the survival analysis. To account for heterogeneity in data, we applied a recently established statistical framework that controls for variation in maternal behaviour with infant age. We found only moderate variation among mothers and no evidence for a link between maternal style and infant growth or survival. Neither infant sex, maternal age, nor variation in NDVI, a measure of habitat productivity, predicted infant growth or survival. We suspect that, in this population living in a habitat with high carrying capacity, disease and predation have greater effects on infant growth and survival than maternal behaviour. Future studies should consider the interplay between infant behaviour and maternal responses in more detail, as it may be crucial for understanding the variation in early social development.

Keywords

Growth, mother-infant relationship, maternal style, *Papio papio*, Parallel Laser Photogrammetry, survival

Introduction

Primate mothers have a profound influence on their infants' growth and survival, providing essential nutrition, social support, and protection. Individual differences in how mothers behave towards their infants are known as maternal style (Altmann, 1980; Fairbanks, 1996; Nicolson, 1987) and have been related to the mother's personality, age, and parity (reviewed in Fairbanks, 1996). Studies of maternal style typically describe maternal behaviours along two dimensions: protectiveness and rejection (e.g., yellow baboons, *Papio cynocephalus*: Altmann, 1980; rhesus macaques, *Macaca mulatta*: Berman, 1990a, 1990b; vervet monkeys, *Cercopithecus aethiops*: Fairbanks & McGuire, 1987). Protective mothers invest more in caregiving, in maintaining contact with their infants, and in restricting their infants' explorations. Rejecting mothers are more tolerant of separation, less attentive, allow more interactions with other group members than protective mothers, and are more likely to increase the distance to the infant.

Infants of rejecting mothers experience less contact earlier (Altmann, 1980), exhibit more exploratory behaviour, and are less afraid of novelty (Bardi & Huffman, 2002; Fairbanks & McGuire, 1988). Yet, a rejecting maternal style might also be associated with higher infant mortality. In yellow baboons, infants of rejecting mothers experienced slightly higher rates of illness and death (Altmann, 1980): five out of seven infants born to rejecting - here termed "laissez-faire" mothers became ill or died, while two out of five infants born to protective (here termed "restrictive") mothers became ill or died. In vervet monkeys, infant mortality was more pronounced among infants with lower suckling success and early weaning (Lee, 1984), whereas a more protective maternal style was associated with higher survival in captive vervet monkeys (Fairbanks, 1996).

Maternal investment is supposed to vary with maternal age (e.g., in wild mandrills older mothers were more in contact with their infants and groomed them more than younger mothers: Roura-Torres et al., 2025) and maternal experience (e.g., more experienced captive gorilla mothers spent more time in contact with younger than with older offspring: Amici et al., 2025). The parental investment theory (Trivers, 1972, 1974) predicts that younger mothers should terminate investment in their offspring earlier because they have more to gain from future reproduction than older mothers. In Barbary macaques (*Macaca sylvanus*), older mothers spent more time in contact with their offspring than young mothers, supporting the parental investment theory (Paul et al., 1993). The maternal experience hypothesis, in contrast, states that young mothers should be more protective than older mothers because they are more inexperienced. Evidence in favour of the maternal experience hypothesis came from a study of rhesus macaques

(Holley & Simpson, 1981), in which primiparous mothers were less confident than multiparous mothers, and from a study with Japanese macaques (Schino et al., 1995), in which multiparous mothers were less protective than primiparous ones.

Furthermore, social, demographic, and ecological variables may be linked to maternal style. For example, a low maternal dominance rank (Altmann, 1980), the presence of newly immigrated males (Fairbanks & McGuire, 1987), or group members who "steal" the infant (Rowell et al., 1964) may increase maternal protectiveness. Conversely, mothers may be more rejecting when surrounded by close kin that offer social support (Altmann, 1980; Berman, 1980). Additionally, the quality and availability of food, which in turn impact maternal health, are crucial. When food is limited, lactating mothers should wean their infants earlier and therefore be more rejecting (Clutton-Brock, 1991). External risk variables may also increase protectiveness; for example, rhesus macaque mothers became more protective after trapping activities on Cayo Santiago (Berman, 1989).

In recent years, differences in maternal style and the consequences for offspring have attracted increasing attention. Revathe and colleagues (2025) investigated individual differences among orangutan mothers, using longitudinal and cross-sectional behavioural data from 15 mothers in the Suaq Balimbing research area. Mothers differed substantially and consistently in four of the six maternal behaviours studied: contact termination, close-proximity termination, carrying, and feeding in proximity. Rolland and colleagues (2025) studied offspring responses to natural threats to infer variation in attachment in wild chimpanzees. They observed considerable differences among offspring and used these to assign attachment types. Fedurek and colleagues (2025) compared mother–infant proximity in two wild chimpanzee populations with varying levels of infanticide risk. In the study population with the higher infanticide risk, less-gregarious mothers stayed closer to their infants when many other adult females were present; the opposite pattern held for highly gregarious mothers. These studies point to the intricate interplay between maternal personality and infant experiences in primate species. Yet, how maternal style affects growth and survival is less well understood, and previous analyses lacked the statistical means to control for variation in maternal investment with respect to infant age (Mundry et al., 2023).

Here, we investigate how differences in maternal style influence the growth and survival of wild infant Guinea baboons (*Papio papio*). Guinea baboons provide insight into how the social environment influences the mother-infant relationship, as they live in a complex multi-level society (Fischer et al., 2017; Patzelt et al., 2014) characterized by low levels of both within- and

between-group aggression. Genetic and behavioural evidence indicate that Guinea baboons exhibit female-biased dispersal (Kopp et al., 2015), with females transferring relatively freely between different societal levels (Goffe et al., 2016). To test the link between maternal style and infant growth and survival, we first characterized maternal style. We predicted that infants who spend less time and have fewer interactions with their mothers have slower growth trajectories and experience higher mortality than those who associate and interact more with their mothers. We included maternal age to control for the fact that a mother's experience can affect an infant's growth (e.g., Altmann & Alberts, 2005) and survival (e.g., Arlet et al., 2014; Pusey, 2012). Further, we included infant sex to control for potential sex differences in growth and survival (e.g., due to early emergence of sexual dimorphism). We also controlled for unit size following the conjecture that females may compete over access to the primary male for infant protection. Finally, we controlled for variation in ecological conditions.

One challenge when analysing the link between maternal style and infant growth and survival is controlling for variation in maternal behaviour with infant age. In brief, at any given point in time, the variation in maternal behaviour could be due to variation between mothers or due to variation in infant age. Failure to control for infant age can give rise to spurious results (Mundry et al., 2023). Moreover, longitudinal data, particularly from wild animal populations, often exhibit a lack of homogeneity. Animals disappear or die, and it may be impossible to continuously monitor study groups. We therefore employed an analytical framework that enabled us to account for changes in behaviour with infant age and data gaps when analysing ontogenetic processes (Mundry et al., 2023).

Methods

Compliance with ethical standards

Approval and research permission were granted by the DPN and the MEPN de la République du Sénégal. This research complies with the regulations set by the Senegalese agencies and the Animal Care Committee at the German Primate Centre. We followed all applicable international, national, and/or institutional guidelines for the care and use of animals.

Study subjects and field site

We collected behavioural data from Guinea baboons ranging near the Centre de Recherche de Primatologie (CRP) Simenti in the Niokolo-Koba National Park, Senegal (described in Fischer et al., 2017). Guinea baboons live in a multi-level social organisation, with 'units' comprising one reproductively active 'primary' male, one to six females, and their immature offspring (Dal Pesco

et al., 2022; Goffe et al., 2016). Females almost exclusively mate and sire offspring with the primary male of their unit (Dal Pesco et al., 2022; Goffe et al., 2016). Several units associate to form a 'party', and two or more parties group together into 'gangs' (Patzelt et al., 2014). Gangs have overlapping home ranges, and competition over food or sleeping trees is low (Ohrndorf et al., 2025a, 2025b; Patzelt et al., 2014). Senegal has a pronounced seasonality: a rainy season from June to October and a dry season from November to April, with an average annual precipitation of 956 mm (Zinner et al., 2021). Seasonality is related to plant productivity. The study population's home range and daily travel distances increase in the rainy season (Zinner et al., 2021).

Behavioural data collection took place from April 2017 to December 2021. We collected growth data using Parallel Laser Photogrammetry (PLP) from April 2019 to August 2021. We used demographical records from long-term data to assess birth dates, maternal age, unit size, and survival. The study subjects were all infants up to 18 months of age associated with the study parties during the study period. We collected data on two parties, party 5 and party 6, since April 2017. In December 2018, we included party 9 in the data collection. During 2019, a subset of individuals split from party 9; this new party was named party 9B. The rest of party 9 was only present until the beginning of 2020. During 2020, party 6 split into party 6W and party 6I. Lastly, we included the infants of parties 13 and 15, for which we previously had only demographic data, in the data collection starting in February 2021.

We conducted observations and recorded demographic information daily from 06:30 to 13:00 hours. We identified all group members by their physical characteristics. To record the data, we used handheld devices (Samsung Galaxy Note II GT-7100 or Gigaset GX290) and the Pendragon software version 7.2.21 (Pendragon Software Corporation, Chicago, IL, USA). Observer reliability and identification accuracy were assessed before observations. Researchers left the field site due to the COVID-19 pandemic from April to November 2020. Upon return, all the individuals could be re-identified. Four infants were born during the researchers' absence. For these individuals, we determined age to the nearest month based on patterns of change in the natal coat and skin from black to brown infants.

We did not have complete records for the first 18 months of any infant due to the COVID-19 pandemic evacuation. Furthermore, several infants died, and others we only sampled when they were already older. Thus, our dataset is characterized by many subjects but significant heterogeneity (Figure S1). Overall, we sampled 80 infants (45 females) born to 49 mothers.

Mothers had at least one and up to a maximum of four infants during the study period (mean = 1.63, SD= 0.92). Seventeen infants had young adult mothers, 50 had middle-aged mothers, and 13 had old mothers (Dal Pesco & Fischer, 2022); 10 mothers with more than one infant were assigned to different age categories. Four of the 49 mothers transferred to another unit or party during the study period, and the infants did not follow them (average age at mother transfer = 13.95 months). Two mothers died before their infant was 18 months old (average age at disappearance = 12.75 months). In one of these cases, the mother had already resided in a different unit from her infant for 30 days before her disappearance. Subsamples included in each analysis are reported in each analysis section. All statistical analyses were conducted in the R environment (version 4.5.2: R Core Team, 2022) in the RStudio interface (version 2025.09.2.418: RStudio Team, 2022).

Behavioural data collection

During 2017-2021, we assembled 2361 protocols from 80 infants (45 females) using the focal-animal sampling method (Altmann, 1974). The observations lasted 15 minutes (median; IQR: 13.22-15.25 minutes). We recorded all approaches and leaves by the mother and infant (within 1 m) as well as the overall time an infant spent within 1 m of the mother. We also recorded the durations of nipple contact, contact-sit, maternal grooming of the infant, and carry (Dal Pesco & Fischer, 2022). Contact-sit, grooming, and carry were mutually exclusive. Nipple contact could co-occur with carry, contact-sit, and grooming. We additionally recorded rejection of nipple contact by the mother, rejection of carrying by the mother, and all aggressions against the infant by the mother (e.g., mock bite, push away, threat) as events. Because of their rare occurrence, we combined all rejection behaviours into a single compound 'rejection' measure. Furthermore, we included protective behaviours (e.g., retrieving and restraining) as events. Data collection on rejection and protection behaviours began only in January 2019. All nine maternal behaviours listed above were used in our assessment of inter-individual variation in maternal style (Fig. 1). The detailed ethogram for maternal behaviours is provided in Tables S2.1 and S2.2.

Parallel Laser Photogrammetry

We used Parallel Laser Photogrammetry (PLP) to determine infant growth over time. PLP is a non-invasive method that combines digital photography with parallel lasers to measure body size in wild animals (Anzà et al., 2022; Bergeron, 2007; Durban & Parsons, 2006; Galbany et al., 2017). We used a three-laser system in which the horizontal and vertical lasers were separated by 20 mm to measure the length of the lower arm in infants. We took pictures of each infant monthly

from April 2019 to August 2021, for a total of 52 infants. Detailed explanations for the PLP in our study can be found in the supplement section S.3.

We analysed a total of 1878 pictures (31 median number of images per individual; IQR: 12-57.75) using ImageJ version 1.53k, an open-source software package for scientific image processing and analysis (Schneider et al., 2012). Three researchers, trained in body landmark identification and length assessment, took the measures. The ICC(A,3) for assessing the inter-rater reliability indicated an excellent reliability (N pictures = 243, raters = 3, $0.993 < ICC < 0.995$, $F_{(242,466)} = 172$, $P < 0.001$). For several infants, we measured arm lengths on the same day ($N = 273$ infant-day combinations with a maximum of 14 measures per day). For these, we evaluated the range of arm-length measures obtained for the same infant on the same date. The average maximum difference between two measures was 7.3 mm (median).

Age-corrected determination of maternal style

To determine inter-individual differences in maternal style, we aggregated observations by infant and month, yielding a sample of 582 infant-month data points from $N = 80$ infants (45 females) born to 49 mothers. Since behavioural measures were aggregated by month, while infant age is a continuous measure, we first needed a monthly value of infant age to use in our analyses. To do this, we calculated an average age in days, weighted by the distribution of focal observations within the month. Maternal style for protect/reject behaviours could only be assessed from 2019 and was based on a sample of 425 infant-month data points from $N = 64$ infants (33 females) born to 45 mothers.

To describe mother-infant relationships, we first fitted separate Generalized Linear Mixed Models (GLMMs, Baayen, 2008), one for each of the nine maternal behaviours, with age (z-transformed) as a fixed effect, and mother ID and infant ID as random intercepts. These random intercepts allowed us to account for differences in maternal style originating from the different mothers included in the study, as well as potential variation in maternal style of the same mother toward her multiple infants. Age was included as a random slope within both mother ID and infant ID to account for variation in the effect of age on mothering style between mothers and infants (Barr et al., 2013; Schielzeth & Forstmeier, 2009). Behavioural states (grooming, contact sit, nipple contact, carry, and proximity) were analysed using beta GLMMs fitted with the 'glmmTMB' function of the package 'glmmTMB' (version 1.1.13; Brooks et al., 2017). Count behaviours were analysed using Poisson (maternal approaches, protections, and rejections) or negative binomial (maternal leaves) GLMMs fitted with the 'glmer' and 'glmer.nb' functions of the R package 'lme4'

(version 1.1-37; Bates et al., 2015). All count models included an offset term for observation effort (log transformed). All models, except the 'maternal leave' model, were fitted without correlations between random intercepts and slopes because either convergence problems occurred when they were included or they were estimated to be close to -1 or 1. For similar reasons, in the 'maternal leave' model, we kept correlations among random intercepts and random slopes only for age within mother ID, not for age within infant ID. We fitted models for maternal protection and grooming with an additional observation-level random effect to mitigate overdispersion. In none of the count models was the response overdispersed (dispersion parameter range 0.45-1.10), except for grooming, which we later excluded due to negligible variation in maternal style (see below).

From the GLMMs, we obtained Best Linear Unbiased Predictors (BLUPs; Baayen, 2008). The BLUPs are the estimated deviations of intercepts and slopes from the common average for each level of the grouping factors mother ID and infant ID (Mundry et al., 2023). For each model, we first inspected the estimated standard deviation of the random intercept effects for mother ID and infant ID, not those for random slopes, as random intercept BLUPs provide information about deviations of individual mothers' maternal style from that of the average mother. When both standard deviations were below 0.1, indicating little variation in maternal style among mothers or infants, we did not consider the results of the respective model in further analyses. We excluded grooming from further analysis because there was low variation between mothers and infants, while we retained the other eight variables (Fig. 1). In particular, BLUPs assessment showed variation in the random intercepts corrected for infant age (SD > 0.1) between mothers for the variables carry, maternal leaves, protect, and reject. Maternal rejections also varied among infants, as did contact-sit, proximity, nipple contact, and maternal approaches. Population-average behaviours during the infancy period (0-18 months) are presented in Table S4.

To consider potential correlations among the behavioural variables and avoid the pitfalls of multiple testing, we conducted a principal components analysis on behaviours that showed sufficient variation between mothers and infants. For each one of the eight behaviours, we computed the sum of the BLUPs for each random intercept (BLUPs mother ID + BLUPs infant ID) and used these values in our PCA. This step gave us a dyad-specific deviation from the global intercept. As we wanted to include all behaviours together, and the BLUBs for the protection and rejection models were only available for a subset of the data, our PCA was based on the subsample of 64 infants for whom all dyadic BLUPs were available (correlations between the eight variables are shown in Figure S5). We settled on a two-factor solution and varimax rotation.

The first factor had an Eigenvalue of 2.29 and explained 28.6% of the total variance. The highest loadings on the first factor (hereafter 'proximity score') were 0.834 for contact sit, 0.833 for proximity, and 0.831 for nipple contact (Table S5). The second factor had an Eigenvalue of 1.38 and explained 17.2% of the total variance. The highest loadings on the second factor (hereafter 'activity score') were 0.768 for maternal approach, 0.525 for maternal leave, 0.510 for carry, and 0.492 for maternal protection (Table S5). Maternal rejections did not load highly on either factor (Table SX). We therefore retained it separately as a third dimension of maternal style (Fig. 1).

Age-corrected determination of infant growth

We expressed infant growth as its deviation from the average growth of infants of the same age. To this end, we fitted a non-linear mixed-effects model, in which the response was arm length, and the principal predictor was infant age. We parametrized the link between size and infant age as follows: arm length = $c + a \cdot b^{age}$, where $c > 0$, $a < 0$, and b ranges from 0 to 1. In the model, we let the coefficients c, a, and b vary between infants. We modelled them as fixed and random effects, whereby the latter implies that they are assumed to originate from a normal distribution with a standard deviation to be estimated. The model eventually provides estimates of c, a, and b, and how they vary across infants. This model was implemented in a Bayesian framework using the 'brm' function from the 'brms' package (Bürkner, 2017, 2018, 2021).

Next, we extracted the individual-specific infant growth parameters from the fitted model. We then determined the instantaneous fitted growth (i.e., the derivative of the fitted model) for each infant on each day between its first and last arm length measurement. Then we subtracted the population-level growth on the same day from each of these values. Finally, we averaged the deviations. These values, which are the slope differences, indicate how fast an infant grew relative to an average infant of the same age (i.e., relative infant growth). We then used the slope differences as a response in our growth models. We used infant growth rather than infant size because growth reflects changes in size over time. Differences in size between individuals of the same age may be due to genetic factors rather than to the conditions affecting infant development. Assessing individual growth trajectories in comparison to the average population growth trajectory for infants of the same age is a more accurate indicator of infant development.

Modelling the effect of maternal style on growth and survival

Of the 64 infants included in the principal components analysis to determine inter-individual differences in maternal style, one was excluded from both the growth and the survival models because full information on unit size, one of our control predictors, was unavailable due to

uncertainty about its unit association. Thus, we included 63 infants (32 females) in the survival analysis. This sample included 44 mothers. Fourteen infants had young adult mothers, 41 had middle-aged mothers, and eight had old mothers.

The growth model included a smaller subset of infants because growth information was collected only from 2019 onwards. We included 52 infants for whom we had PLP pictures for the assessment of relative infant growth (Fig. 1). Of those, two had to be excluded from the growth model: one due to uncertain unit association (see above), the other because the infant died shortly after birth and no behavioural data could be collected. Therefore, for the growth model, we included 50 infants (25 females) born to 35 mothers. Nine infants had young adult mothers, 34 had middle-aged mothers, and seven had old mothers.

To analyse the link between maternal style and infant growth, we fitted a Linear Mixed Effects Model (LMM) using the 'lmer' function of the R package 'lme4' (version 1.1-37; Bates et al., 2015). For the link between maternal style and infant survival, we fitted a Cox proportional hazards model using the 'coxme' function of the package 'coxme' (version 2.2-22; Therneau, 2022). Mother ID and unit ID (i.e., the unit with which the infant was associated for most of its observed presence – from birth/first sighting to 18 months/last sighting) were included as random intercepts. No random slopes were theoretically identifiable. We used maternal age, infant sex, unit size, and the Normalized Difference Vegetation Index (as a proxy for ecological conditions; hereafter NDVI) as fixed-effect control factors (Fig. 1). Maternal age was defined as the mother's age category on the day of each infant's birth. During the study period, five unidentified females transferred to our study parties with an infant; for these cases, maternal age was assigned on their first observation day using our age category definitions (Dal Pesco & Fischer, 2022). Unit size for each infant was calculated as the weighted average number of unit co-residents (i.e., the primary male and all associated adult and subadult females) during the infant observed period. If an infant changed units during the period considered, we assigned it to the unit it belonged to for the longest period.

We obtained the NDVI values from the Copernicus Global Land Service (CGLS) (https://land.copernicus.eu/global/), which is part of the European Commission's Earth Observation programme. See supplementary section S6 for details regarding NDVI extraction and processing. To control for the fact that we were forced to use two different NDVI versions, we created an additional indicator variable indicating whether the NDVI value for a given day was based on version 1 or version 2. Since some of the daily NDVI values were interpolated between

NDVI versions 1 and 2, the indicator variable denoted the relative contribution of NDVI version 2 to a daily NDVI value. We deemed this step necessary because the two NDVI versions were not entirely comparable. NDVI was used as a fixed-effect control, together with its interaction with the indicator variable denoting the NDVI version.

Before fitting the two models, we inspected whether the distribution of the response variable (for the LMM) and all quantitative predictors was roughly symmetrical. Maternal rejections were square-root transformed in both models to improve symmetry; to ensure all values were positive, we added 1 in the growth model and subtracted the minimum value in the survival model. To ease model convergence and the comparison of estimates, we z-transformed all covariates to a mean of 0 and a standard deviation of 1 before fitting each model (Schielzeth, 2010). Before inference, we performed several diagnostic validations. Residuals showed no deviations from the assumptions of normality and homoscedasticity, as assessed by visually inspecting the residuals' QQ-plot (Field, 2005) and a scatterplot of residuals versus fitted values (Quinn & Keough, 2002). Model stability was evaluated by comparing estimates obtained when running the models with the levels of the random effects excluded one at a time with those obtained for the complete data set (Nieuwenhuis et al., 2012). Both models were of acceptable stability.

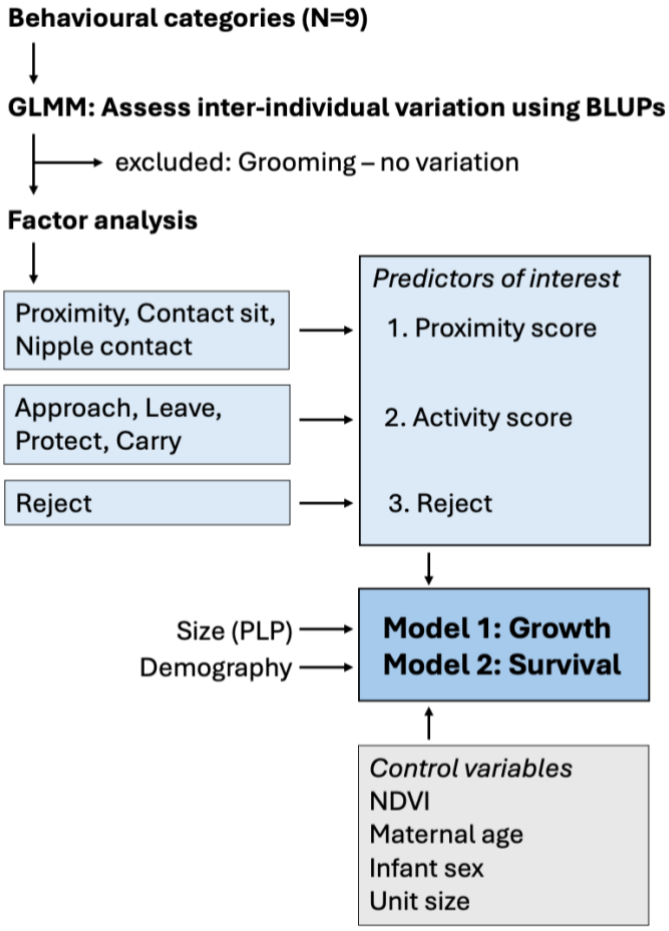

Figure 1. Workflow. We first screened the predictor variables for variation between mothers and infants using the estimated standard deviation of the BLUPs. We then used a principal components analysis to reduce the number of variables used to estimate variation in maternal style, yielding two factors and one separate variable ('reject'). Lastly, we fitted the growth and survival models using the three resulting variables as predictors of interest, along with several control variables.

As a general test of the statistical significance of maternal style and to avoid "cryptic multiple testing" (Forstmeier & Schielzeth, 2011), we conducted a full-null model comparison, whereby the null model lacked the three fixed effect predictors of maternal style (proximity factor, activity factor, and reject), but aside from that, was identical to the full model. We used a likelihood ratio test to compare the full and null models (Dobson, 2002). We employed the Satterthwaite approximation to test for the effects of the individual fixed effects in the infant growth LMM (Luke, 2017). To this end, we refitted the models with restricted maximum likelihood. For the survival analysis, we assessed the significance of the individual fixed effects by dropping them from the model one at a time (R function "drop1"). We obtained bootstrapped confidence intervals for model estimates and fitted values using the function 'bootMer' in the package 'lme4';

we also used the function 'simulate' in the package 'glmmTMB' to obtain bootstrapped confidence intervals for duration behaviours included in Figure 2. We determined degrees of freedom and *P*-values for the LMMs using the function 'lmerTest' of the package 'lmerTest' (Kuznetsova et al., 2017).

Results

General development

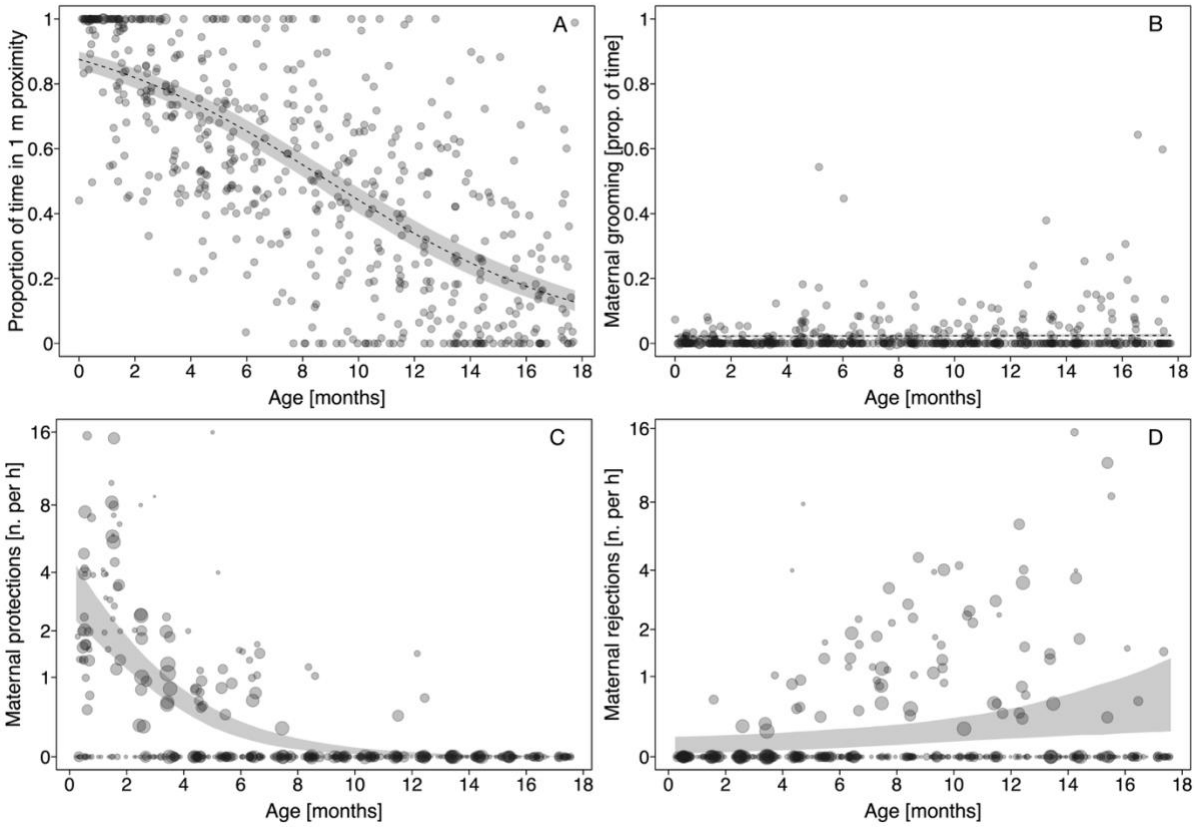

Figure 2. Raw data and estimates (mean and confidence intervals) for A: proportion of time in 1-m proximity; B: proportion of time the mother grooms the infant; C: number of maternal protections per hour. D: number of maternal rejections per hour.

In the first month of life, infants were mainly in close proximity to their mothers (Fig. 2A); the time spent in close proximity declined with infant age, averaging 20% of the observation time at 18 months of age. Grooming did not vary with age and averaged around 2.2% of the observation time (Fig. 2B). Mothers showed high rates of protection behaviours in the first two months of the infants' lives, with a decreasing rate until about 8 months of age, when they largely ceased to

protect their offspring (Fig. 2C). With increasing infant age, maternal rejection became more frequent (Fig. 2D).

Infant growth and survival

Of the 50 infants for which we had growth data (Fig. 3), 29 grew faster than the population average, and 21 grew more slowly. Infant death occurred similarly in the group of infants that grew faster than the average population (four infants) and in the group that grew more slowly (three infants). Due to the small sample size, we refrained from further statistical analysis. We extracted the estimated forearm length at 30 days to assess differences between males and females. For both sexes, the estimated forearm length at an age of 30 days was 78 mm. Similarly, the estimated forearm length at 30 days was 78 mm for both infants who survived and those who died.

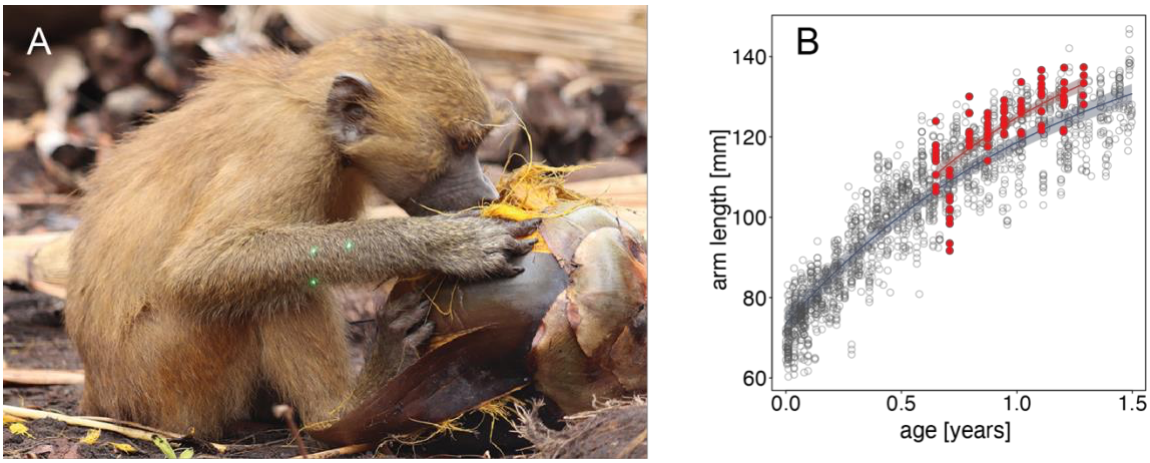

Figure 3. A. Measurement of arm length using Parallel Laser Photogrammetry of subject PIF. Note the three green spots on the infant's forearm. B. Growth curve of the forearm length. Grey circles represent the arm-length measurements (N=1878) for all the infants of the population. The blue line represents the fitted model for the population average, and the red line depicts the fitted model for subject PIF. The shaded areas in blue and in red represent the 95% credible intervals of the fitted model for the average population and the subject PIF, respectively.

Infant survival

Of the 80 infants in our initial sample, 63 survived until 18 months of age. Of the 17 infants who died, mortality occurred throughout the infancy period, with infants dying from an age of 13 days to 16.6 months. The highest infant mortality occurred in the first three months of life, during which seven infants died. Three infants died before reaching six months, four between six and nine months, and just one between nine months and one year. Two infants died after the first year: one

was 13 months old, and the other was 16 months old. Of the 63 infants that entered the survival model (after the subsets described in the method section), 51 survived up to 18-months.

Link between maternal style, growth, and survival

The growth model revealed no apparent relationship between maternal style and infant growth (Fig. 4). The full model with the three test predictors did not explain the data better than the null model (full-null model comparison: $\chi^2 = 3.922$, $df = 3$, $P = 0.27$). None of the test or control (infant sex, maternal age, unit size, and NDVI) predictors showed a noticeable effect on infant growth (Table 1).

The survival model showed no apparent effect of maternal style on infant survival. The full model, which included the three test predictors, revealed no statistically significant relationships between the predictors and the dependent variables (full-null model comparison: $\chi^2 = 0.608$, $df = 3$, $P = 0.895$). None of the test or control (infant sex, maternal age, unit size, and NDVI) predictors showed a noticeable effect on infant survival (Table 2). Although male infants were almost twice as likely to die as female infants (10/35 vs. 7/44), mortality did not differ statistically between the sexes.

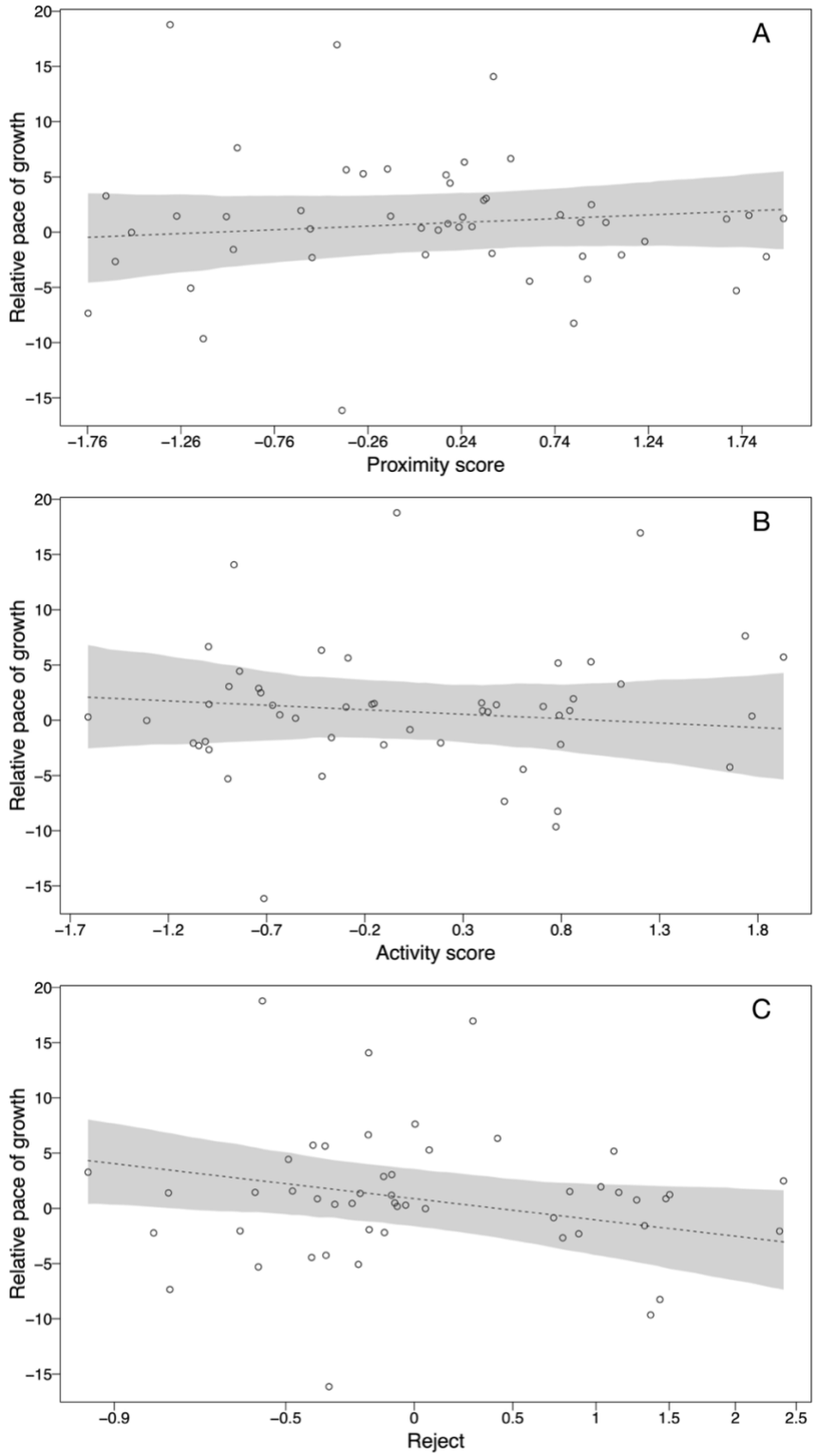

Figure 4. Relationships between the relative pace of growth of infants up to 18 months of Guinea baboons and (A) proximity score, activity score, and (C) maternal rejections. Each circle represents an individual. The dashed line depicts the fitted model, and the shaded areas represent the bootstrapped 95% confidence intervals for all other terms in the model, assuming they are at their average.

Table 1. Results of the Linear Mixed Model (LMM) analysis of the influence of the test predictors of maternal style on infant growth (growth model). Control predictors are infant sex, maternal age, unit size, and NDVI values. Estimates are shown, along with standard errors (SE), 95% confidence intervals, significance tests, and minimum and maximum estimates from model stability analysis.

Term	**Est.**	**SE**	**CI_{lower}**	**CI_{upper}**	***df***	***t***	***P***	**min**	**max**
Intercept	1.886	1.652	-1.127	4.938	/	/	/	0.899	3.265
Proximity score [(1)]	0.679	0.828	-0.771	2.059	27.252	0.820	0.419	0.198	1.465
Activity score [(2)]	-0.623	1.064	-2.678	1.320	38.473	-0.585	0.562	-8.108	0.188
Reject [(3)]	-1.631	0.831	-3.228	-0.218	32.159	-1.963	0.058	-5.275	-0.972
Maternal age (OAF) [(4)]	-3.868	2.440	-8.199	0.247	[(5)]		0.206 [(5)]	-8.721	-1.231
Maternal age (YAF) [(4)]	0.834	1.998	-2.905	4.564				-1.795	8.037
Infant sex [(6)]	-1.427	1.586	-4.120	1.664	34.866	-0.900	0.374	-3.381	1.572
Unit size [(7)]	0.614	1.362	-1.796	3.432	18.902	0.451	0.657	-0.129	4.489
NDVI [(8)]	-1.857	0.801	-3.550	-0.327	/	/	/	-2.545	-0.854
NDVI version [(9)]	1.992	0.864	0.374	3.604	/	/	/	1.249	4.838
NDVI*NDVI version	0.531	0.841	-0.919	2.072	29.592	0.632	0.533	0.133	2.321

[(1)] *z*-transformed to a mean of zero and a standard deviation of 1; the original mean and standard deviation were 0.078 and 0.981, respectively

[(2)] *z*-transformed to a mean of zero and a standard deviation of 1; the original mean and standard deviation were 0.018 and 0.899, respectively

[(3)] square-root transformed and then *z*-transformed to a mean of zero and a standard deviation of 1; the original mean and standard deviation of the square-rooted variable were 1.022 and 0.377, respectively

[(4)] Dummy coded with mature adult female (MAF) being the reference level; OAF = old adult female and YAF = young adult female.

[(5)] To get an overall effect of maternal age we run the function anova on the model with restricted maximum likelihood obtaining the following F and df values: F(2, 30.081) = 1.668.

[(6)] Dummy coded with female being the reference level.

[(7)] *z*-transformed to a mean of zero and a standard deviation of 1; the original mean and standard deviation were 3.882 and 1.142, respectively.

[(8)] *z*-transformed to a mean of zero and a standard deviation of 1; the original mean and standard deviation were 0.461 and 0.135, respectively

[(9)] *z*-transformed to a mean of zero and a standard deviation of 1; the original mean and standard deviation were 0.376 and 0.469, respectively

Table 2. Results of the Cox proportional hazard model analysing the influence of the test predictors of maternal style on infant survival. Control predictors are infant sex, maternal age, unit size, and NDVI values. Estimates are shown, along with standard errors (SE), significance tests, and minimum and maximum estimates from model stability analysis.

Term	**Estimate**	**SE**	χ^2	***df***	***P***	**min**	**max**
Proximity score [1]	-0.078	0.349	0.050	1	0.823	-0.237	0.169
Activity score [2]	-0.237	0.373	0.419	1	0.518	-0.570	0.248
Reject [3]	0.110	0.322	0.118	1	0.731	-0.084	0.479
Maternal age (OAF) [4]	0.730	0.811	2.117	2	0.347	0.022	1.982
Maternal age (YAF) [4]	-0.982	1.171				-21.515	-0.514
Infant sex [5]	0.866	0.652	1.847	1	0.174	0.483	1.622
Unit size [6]	-0.097	0.381	0.064	1	0.800	-0.524	0.351
NDVI [7]	0.183	0.423	-	-	-	-0.665	0.782
NDVI version [8]	0.144	0.517	-	-	-	-0.510	0.816
NDVI*NDVI version	0.915	0.477	3.289	1	0.070	0.146	1.920

[1] *z*-transformed to a mean of zero and a standard deviation of 1; the original mean and standard deviation were 0.04 and 0.888, respectively

[2] *z*-transformed to a mean of zero and a standard deviation of 1; the original mean and standard deviation were 0.014 and 0.849, respectively

[3] square-root transformed and then *z*-transformed to a mean of zero and a standard deviation of 1; the original mean and standard deviation of the square-rooted variable were 1.013 and 0.361, respectively

[4] Dummy coded with mature adult female (MAF) being the reference level; OAF = old adult female and YAF = young adult female; the indicated significance test is an overall test of the effects of maternal age

[5] Dummy coded with female being the reference level

[6] *z*-transformed to a mean of zero and a standard deviation of 1; the original mean and standard deviation were 4.031and 1.073, respectively.

[7] *z*-transformed to a mean of zero and a standard deviation of 1; the original mean and standard deviation were 0.529 and 0.118, respectively

[8] *z*-transformed to a mean of zero and a standard deviation of 1; the original mean and standard deviation were 0.433 and 0.455, respectively

Qualitative assessment of protection and rejection behaviours

To facilitate comparisons with previous data on yellow baboons (Altmann, 1980), we additionally considered only those mother-infant dyads for which we had data on the first four months (i.e., cases recorded from the time we began to record protection and rejection behaviours; N = 15). Our qualitative assessment showed that two of the mothers exhibited no protective behaviours at all, while the other 13 mothers protected their infants on average 2.06 times per hour (SD = 1.43) over the first four months. While six of these mothers consistently showed protective behaviours at least once a month, the other seven did not. Of the 15 infants in this sample, 14 survived to 18 months of age. All infants for whom we had no records of protection in the first four months survived, while of the six infants with the highest rates of protection, only three survived.

Discussion

The primary aim of this study was to characterise maternal styles in wild Guinea baboons and examine their effects on infant growth and survival up to 18 months. Although mothers differed in their social investment, neither the maternal contact score (including time in contact-sit, 1-m proximity, and nipple contact) nor the maternal activity score (including maternal approaches, leaves, protections, and carry durations) nor maternal rejections were related to infant growth or survival. Although infant sex was only a control predictor in our analysis, we did not find differences in maternal investment in relation to infant sex (see also Arbaiza-Bayona et al., 2025 for similar findings in Assamese macaques, *Macaca assamensis*).

The finding that maternal style had no evident effects on infant growth or survival may be due to methodological or sample limitations, or it may reflect a genuine absence of an effect. The idea that there might be no noticeable effect would not be surprising, given that the study population lives in a habitat with a high carrying capacity (Ohrndorf et al., 2025a). Thus, there is likely only minor variation in the mothers' energetic state, which could translate into variation in investment into infant growth. Alternatively, prenatal effects (e.g., Berghänel et al., 2016) may be more important in determining infant health and growth than postnatal care. Either way, we found no support for a strong link between maternal style and growth and survival in this species. We suspect that mortality is mainly related to disease or injuries in young infants and predation, once young become more independent. Whether infant disappearances are indeed due to predation is hard to ascertain, however, as key predators such as leopards, *Panthera pardus*, are mostly active at night, and predation by snakes or birds of prey would be difficult to detect.

Infant sex had no noticeable impact on infant growth rate and survival in our sample. In a comparative analysis of neonatal body mass in 109 primate species, neonatal sexual dimorphism positively correlated with adult dimorphism (Smith & Leigh, 1998). Despite the striking differences in adult body mass among Guinea baboons (Fischer et al., 2017), we found no evidence of sex-related differences in growth or estimated size during infancy. Similarly, Anzà and colleagues (2022), using the PLP method in Assamese macaque infants, reported no differences in forearm length during the first year of life; sexual dimorphism in forearm length emerged only once males reached 5-7 years of age. The late emergence of sexual dimorphism in Guinea baboon body size (at least post-18 months) occurs in the absence of detectable sex-biased maternal investment, arguing against early maternal allocation as a primary driver of sexual dimorphism later in life.

To account for potential ecological conditions that may influence growth and survival in our sample of infants, we also included NDVI as a fixed control effect. Again, we found no apparent effect on infant growth and survival. However, we now know that NDVI is a poor proxy for food availability in our population, because several key feeding tree species produce ripe fruits at the peak of the dry season when NDVI greenness is lowest (Ohrndorf, 2025). In a one-year long study, food availability scores were also the highest during this period (Ohrndorf et al., 2025a). In hindsight, we would not include this variable again.

Surprisingly, maternal age was not a prominent determinant of infant growth and survival, suggesting that maternal inexperience does not seem to entail consequences for infants in our population. This finding stands in stark contrast to observations for other primate species, where younger mothers often experience a higher infant mortality than older ones (e.g., four groups of habituated wild grey-cheeked mangabeys, *Lophocebus albigena*, in Uganda: Arlet et al., 2014; six out of seven studies from captive and wild primate populations across different species: Pusey, 2012). One possibility is that favourable ecological conditions, along with social support from unit members (i.e., the primary male and co-resident females) within a generally tolerant environment, may buffer inexperienced mothers from resource and skill-related constraints, allowing infants of young mothers to develop just as successfully as those of more experienced mothers.

Although it would be attractive to assess potential differences in maternal style across species and their links to other facets of their social systems, such comparisons remain challenging. Such comparative studies can be hindered by differences in ethograms and data-collection protocols. We illustrate these differences by comparing our results to those reported

for yellow baboons. This species lives in multi-male-multi-female groups with female philopatry, resulting in societies organised around matrilines (Alberts, 2019 for a synthesis). At six months of age, mothers and infants were in 1-m proximity around 48% of the time in yellow baboons (Altmann, 1980), while in Guinea baboons, mothers and infants were in 1-m proximity 65% of the time. The high overall gregariousness of the Guinea baboons (Goffe et al., 2016) may account for this difference between species, but more comprehensive comparisons would be needed to test this conjecture.

Altmann and Samuels (1992) reported the time mothers carried their infants. Differences in carrying time were similar between the two species just after birth, with around 30% of the time in yellow baboons and 42% in Guinea baboons. However, in yellow baboons, the proportion of time spent carrying decreased to 5% by month six, and infants were barely carried by eight months, whereas in Guinea baboons, it remained 24% at six months, and was 6% at 18 months of age. It isn't easy to disentangle whether these differences in carrying time are due to maternal style or environmental conditions. If females need to roam more widely to forage, they will likely spend more time carrying the infants when travelling over long distances. In conjunction, reliable comparisons require means to assess the influence of other variables (e.g., information on time spent moving). Ideally, they are also based on identical data collection protocols (Kalbitzer et al., 2015).

Regarding the characterization of maternal styles, Altmann (1980) classified yellow baboon mothers as "laissez-faire" and "restrictive". Laissez-faire mothers stopped restraining the infants earlier (≤ 1 month) than restrictive mothers, who stopped restricting the infant when they were 1.5 months or older. Mothers in our sample whose infants reached at least an age of 4 months and whose infants were sampled throughout this period also showed marked variation. Following Altmann's classification scheme, some of the mothers would be considered as "laissez-faire", with relatively little overt protection. Other mothers with high levels of protection could be classified as "restrictive". Protection, however, is not solely driven by maternal style; it can also be considered as a response to the infant's behaviour. Infants who continue to cling to their mothers do not prompt maternal intervention, while more active infants may more readily evoke protective behaviour in mothers. Infants are not only passive recipients of maternal behaviour but actively shape the relationship (Fairbanks & Hinde, 2013). Further, when mothers are more rejecting, infants may respond by increasing their effort to maintain proximity (Fairbanks & Mcguire, 1995). Future studies should consider the interplay between infant behaviour and

maternal responses in more detail, as it may be crucial for understanding variation in early social development.

Regarding methodological limitations, there is a trade-off between dense sampling of a few infants or more sparse behavioural sampling of a relatively large number of infants, as in our study. Altmann (1980) initially sampled 12 yellow baboon infants. However, the yellow baboon population experienced high infant mortality, and only seven infants were present at the end of the twelve-month age window. In contrast, at twelve months of age, we still had 29 mother-infant dyads. A very dense sampling of a few infants allows solid inferences for these individuals. Paul (1984) observed seven infant Barbary macaques every other day, totalling 544 observation hours per infant over 2 years (except for one infant that died). Yet, in such a small sample, one or two extreme cases can distort the overall picture, and statistical inferences remain limited.

In contrast, sparse sampling of many infants can better approximate the population average. Though, because limited information is available for any given individual, it becomes difficult to assess how strongly individuals deviate from the population mean. The structure of the data may have affected our ability to estimate maternal style variation and its contribution to infant growth and survival. Therefore, given our broad infant sample but limited data per mother–infant dyad, we are more confident about the average developmental trajectories of the infants than the variation between mothers.

In addition, we found that the PLP method for assessing growth is associated with significant measurement error. While the interobserver reliability in measuring the inter-laser distance, recognising body landmarks in the picture, and measuring arm length in pixels was excellent, there was considerable variation across photographs taken on the same day: the average error (8.0 mm) was equivalent to 2 months of growth, with a maximum error (24.2 mm) equivalent to 8 months of development and a minimum error (0.2 mm) equivalent to ca. two days of growth. In sum, the high uncertainty in both the predictor (maternal style) and the response variable may have contributed to our failure to detect effects on growth. In the survival analysis, however, the response variable was unambiguous.

In summary, we found variation in maternal style across a wide range of maternal behaviours, particularly regarding protection, rejection, carrying behaviour, and maternal approaches and leaves. Nevertheless, we found no link between variation in maternal style and infant growth and survival. Although the comparison between species – notably Guinea and

yellow baboons – must be taken with a pinch of salt, the observed differences appear to be in degree and not in kind. The observed mother-infant relationships follow a common pattern across baboon (and other catarrhine) species. We propose that differences between baboon species emerge only later in life, as part of the physical and physiological maturation, which ultimately leads to the pronounced differences in bonding patterns observed between baboon species (Fischer et al., 2019).

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Author Contributions

Conceptualization: AAdD, FDP, JF; Investigation: AAdD; Methodology: AAdD, FDP, RM, JF; Software: FDP, RM; Formal analysis: AAdD, FDP, RM, JF; Data curation: AAdD, FDP, JF; Writing – Original Draft: AAdD; Writing – Review & Editing: AAdD, FDP, RM, JF; Visualization: AAdD, FDP, RM, JF; Funding acquisition: JF.

Data availability statement

The data and code for the statistical analyses that support the findings will be made publicly available on OSF upon acceptance of the manuscript.

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Supplementary material for "Maternal style does not predict infant growth and survival in wild Guinea baboons"

S1. Behavioural data availability

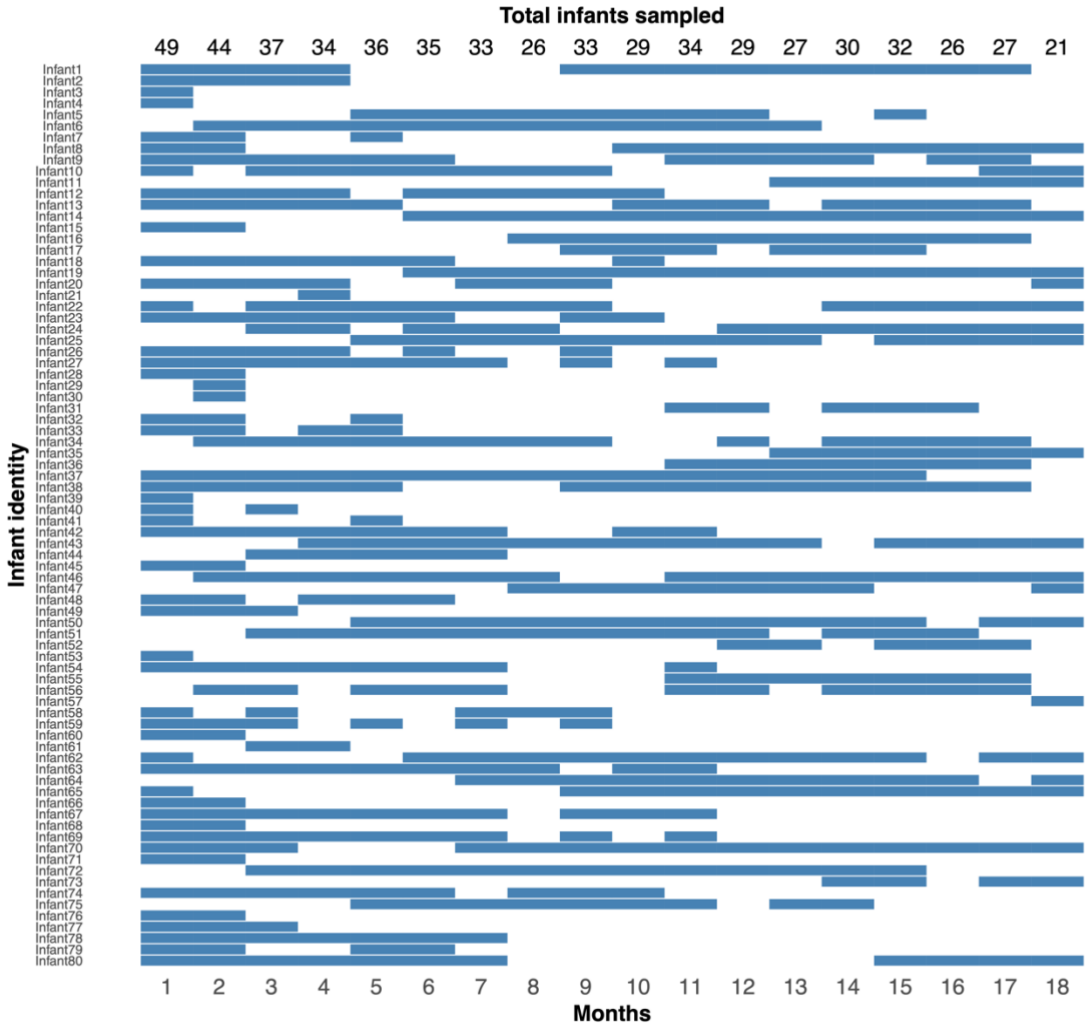

Figure S1 Overview of the sample sizes for behavioural data collection during the study period (April 2017 – December 2021). The blue cells depict monthly data availability for the 80 studied infants from birth until 18-months of age. On the top the total number of infants sampled per month is reported (mean= 32.3, range= 21-49). Data collection gaps are due to multiple reasons: infants were born before the study started (i.e., no data are available from birth), infants were born less than 18-months from the study period end (i.e., no data collection until 18-months of age), infants transferred into/out of our study parties (i.e., no data available before/after the transfer), study parties or individuals not found for a time period, COVID-19 data collection break (from April to November 2020)

S2. Ethogram

Table S2.1. Ethogram of state behaviours between a mother and her infant that were considered for the FAs. Adapted from the CRP Simenti work manual (Dal Pesco & Fischer, 2022; version August 2022)

Behaviour	Description
Maternal grooming	The mother moves her fingers and/or lips through the fur of her infant. A grooming bout is considered finished only once the mother has stopped grooming her infant for more than 5 s. Grooming, contact-sit, and carry are mutually exclusive.
Contact-sit	Mother and infant are sitting, lying, or standing within < 10 cm of one another, meaning that the individuals either touch skin-to-skin or are so near that their furs touch, with at least one individual sitting or lying down. This applies to the body, the hairs, and the different body parts (e.g., feet or hands) of the two individuals. However, if one individual is in contact with another one only with its tail, body contact is recorded only if it is the base of the tail that is in contact. A contact-sit bout is considered finished only once contact stopped for more than 5 s. Grooming, contact-sit, and carry are mutually exclusive.
Proximity	Total time that the mother and infant are within 1 m proximity during the observation time. This is calculated using information on: - proximity at the beginning and end of the focal protocol - detailed recordings of movements during the observation time (i.e., approaches into 1 m and leaves out of 1 m between the mother and infant)
Nipple contact	The infant has its mother's nipple in its mouth, either holding onto it or suckling. This behaviour is not mutually exclusive to contact sit, carry, and grooming.
Carry	The infant is carried by its mother, either ventrally or dorsally. If the mother stops walking but the infant is still holding on to the mother's fur with all four extremities, it is still carrying behaviour. However, when the mother is not walking and the immature is grabbing the mother's fur only with the forelimbs, this is not coded as carried, but as contact sit. Grooming, contact-sit, and carry are mutually exclusive.

Table S2.2. Ethogram of event behaviours between a mother and her infant that were considered for the FAs. Adapted from the CRP Simenti work manual (Dal Pesco & Fischer, 2022; version August 2022)

Behaviour	Description
Protection	Aggregate of two protective behaviours: retrieve and restrain. *Retrieve*: the mother recovers the infant (including from other individuals) by grabbing or holding it with one or both arms impeding that the infant moves away from her. *Restrain*: the mother stops the infant from moving away by grabbing and holding the infant.
Rejection	Aggregate of rejecting behaviours and aggressive behaviours directed from the mother to the infant. *Rejection of nipple contact:* the mother avoids her infant from starting nipple contact. *Rejection of carrying:* the mother prevents her infant from being carried. *Mock-bite*: the mother bites her infant with moderate strength. *Push away*: the mother presses or shoves her infant away from her, avoiding either contact or proximity between them. *Threat*: the mother stares at the infant with lifted eyebrows or stands on all four limbs with her hands placed slightly in front of the body. The mother can also conduct a small but fast movement forward while all legs and arms remain on the ground, or one arm loses ground (ground slap). *Physical fight*: The mother hits or bites or throws the infant on the ground. *Submissive behaviour*: the infant displays submissive behaviour such as fear grin, crouch, or tail-raise towards the mother or via vocalisations such as keck or scream.
Maternal approaches	The mother decreases her distance to the infant to less than 1 m.
Maternal leaves	The mother increases her distance to the infant to more than 1 m.

S3. Parallel Laser Photogrammetry

The apparatus, or Parallel Laser Photogrammetry (PLP) system, consisted of a digital single-lens reflex (DSLR) Camera Canon EOS 200D and an objective Canon EFS 55-250mm, along with a laser box that contained three parallel lasers (laser class 2 520 nm 1mW DI520-1-3) (Fig. S3.1). We chose a three-laser system because a third laser creates a right isosceles triangle with the other two lasers, which allows to become aware and correct whenever the photographed object deviates from the perpendicular orientation of the camera-object axis, or to detect whether the photographed object has tilts on the surface (Anzà et al., 2022; Galbany et al., 2016). In both cases, the right angle of the triangle would be lost, as one of the laser dots would appear displaced (see Anzà et al., 2022). We used green lasers, as in daylight they are more visible than red lasers (Bergeron, 2007; Durban & Parsons, 2006). Three PLP systems were used during the duration of the study: PLP1, PLP2, and PLP3.

Figure S3.1. Apparatus for Parallel Laser Photogrammetry (PLP) used in the study, seen from different angles. The apparatus, also referred to as PLP system, consisted of a camera and a laser box with three lasers and a switch. Batteries were located inside the laser box. Horizontal and vertical lasers were separated by 20 millimetres

The number of pixels between the paired lasers projected on the surface of the body (inter-laser distance) that appear in the photo are used as a scale (Richardson et al., 2022) since a relationship can be made between inter-laser distance and the already-known inter-beam distance (see formula below). The scale is then used to convert the arm size in pixels (determined by the distance between body landmarks) into arm size in millimetres (see formula below). Unfortunately, due to calibration issues in the field, only the horizontal paired lasers could be used as a scale (see below for details about calibration issues).

$$\text{Arm length (mm)} = \frac{\text{inter-beam distance (mm)}}{\text{inter-laser distance (pixels)}} \times \text{Arm length (pixels)}$$

The individuals were photographed in their natural habitat, in lateral view. Observer distance was about 2 to 5m. To determine the length of the lower arm, we defined two landmark endpoints for our measurements (Fig. S3.2, A-B): the olecranon and the ulnar styloid process.

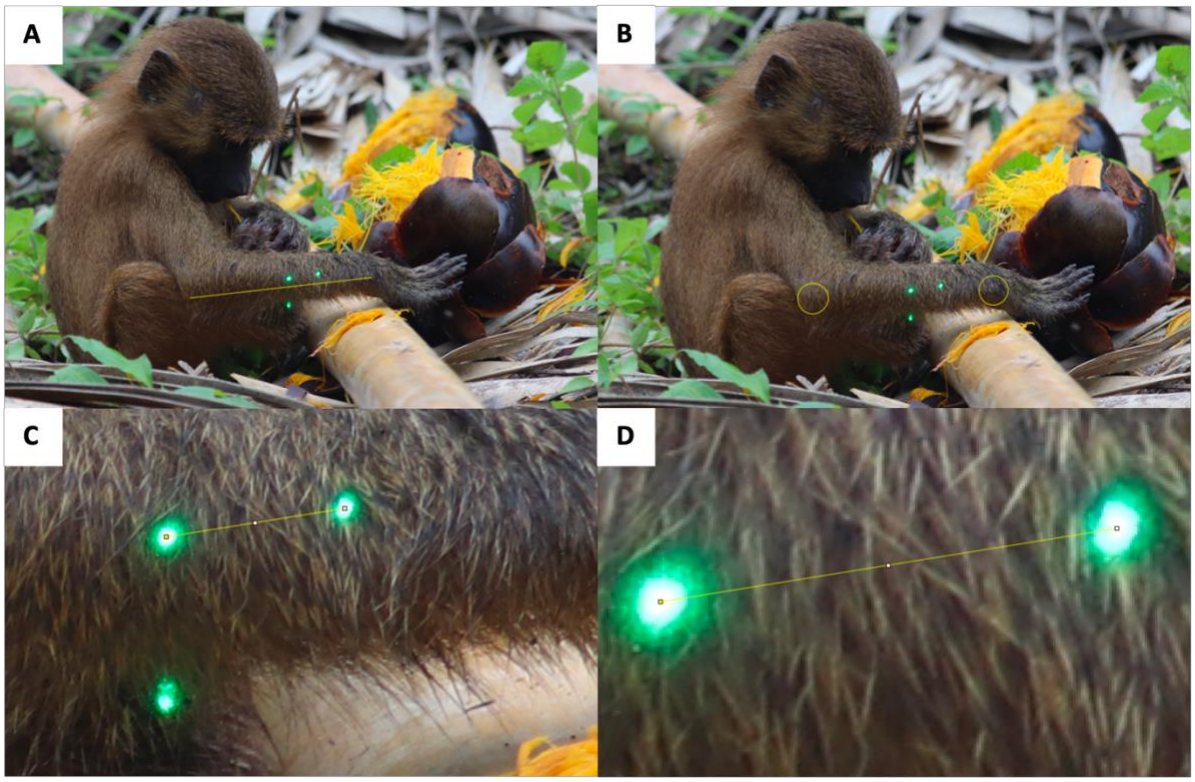

Figure S3.2 Lower arm measurement (A) in infant Guinea baboons is determined by body landmarks: olecranon (B, left yellow circle) and ulnar styloid process (B, right yellow circle). The horizontal lasers were used as a scale (inter-beam distance: 20 mm). Inter-laser distance was measured in ImageJ from centre to centre of the green laser dots (C, D). In our study, only horizontal lasers were used, due to a calibration issue with the vertical lasers

To minimize error, when we selected pictures to be measured, as those which met the following criteria (adapted from Galbany et al., 2017): landmarks were identifiable in the picture, and in focus; individuals had to be either sitting or standing, but always flexing their arm (Fig. S3.2, A-B); an ideal position was one in which the individual had the upper and lower arm forming a right angle. Laser points were on the same plane as the body landmarks of the infant. This plane was visually perpendicular to the optical axis of the camera objective lens and to the projection of the laser beams, so that the body of the infant was not rotated. Further, preferably the infant was situated in the centre of the picture.

To control for possible irregularities in the laser alignment and because we could not adjust and calibrate the lasers (i.e., inter-beam distance) during the field season, we projected them onto a millimetre paper. We took pictures at 2 and 5 m before and after each field day on which we took pictures. The pictures were later measured with ImageJ. This way, we could know the exact daily inter-laser distance and, therefore, the exact scale, which we used as a daily scale, since we had calibration problems during the study period: over time, the vertical lasers deviated from their original distance considerably, while the horizontal ones deviated from it only slightly. The vertical lasers were ultimately not used for measuring growth, as the differences were too large when increasing the distance between the photographed object and the camera. The calibration issues might have occurred because the apparatus was carried during field work and occasionally walked through dense terrain or due to temperature changes, since the inter-laser distance was shorter at the peak of the dry, hottest season for two years.

S4. Mother-infant behaviours

Table S4 Fitted values of the GLMMs for the eight behaviours that showed variation in maternal style. These values represent the population average behaviours throughout the infancy period (from birth to 18-months of age). For maternal grooming, contact-sit, carry, time 1m proximity, and nipple contact fitted values indicate proportion of time; for maternal approaches, maternal leaves, maternal protections, and maternal rejections they indicate number events per hour

	Month																		
Behaviour	0	1	2	3	4	5	6	7	8	9	10	11	12	13	14	15	16	17	18
Maternal grooming	0.022	0.022	0.022	0.023	0.023	0.023	0.023	0.023	0.023	0.023	0.024	0.024	0.024	0.024	0.024	0.024	0.024	0.025	0.025
Contact-sit	0.503	0.467	0.430	0.395	0.360	0.327	0.296	0.266	0.239	0.213	0.190	0.168	0.149	0.131	0.115	0.101	0.089	0.077	0.068
Carry	0.415	0.381	0.349	0.318	0.289	0.261	0.236	0.211	0.189	0.169	0.150	0.133	0.118	0.105	0.092	0.081	0.071	0.063	0.055
Time 1m proximity	0.875	0.849	0.819	0.784	0.745	0.702	0.654	0.603	0.550	0.496	0.442	0.389	0.338	0.291	0.249	0.210	0.176	0.147	0.121
Nipple contact	0.495	0.457	0.420	0.383	0.348	0.314	0.283	0.253	0.225	0.200	0.177	0.156	0.137	0.120	0.105	0.092	0.080	0.069	0.060
Maternal approaches	0.013	0.013	0.013	0.013	0.012	0.012	0.012	0.012	0.012	0.012	0.011	0.011	0.011	0.011	0.011	0.011	0.011	0.010	0.010
Maternal leaves	0.022	0.023	0.025	0.026	0.028	0.030	0.032	0.034	0.037	0.039	0.042	0.045	0.048	0.051	0.054	0.058	0.062	0.066	0.071
Maternal protections	0.047	0.030	0.019	0.012	0.008	0.005	0.003	0.002	0.001	0.001	0.001	0	0	0	0	0	0	0	0
Maternal rejections	0.002	0.002	0.002	0.002	0.002	0.003	0.003	0.003	0.003	0.004	0.004	0.004	0.005	0.005	0.006	0.007	0.007	0.008	0.009

S5. Factor analysis

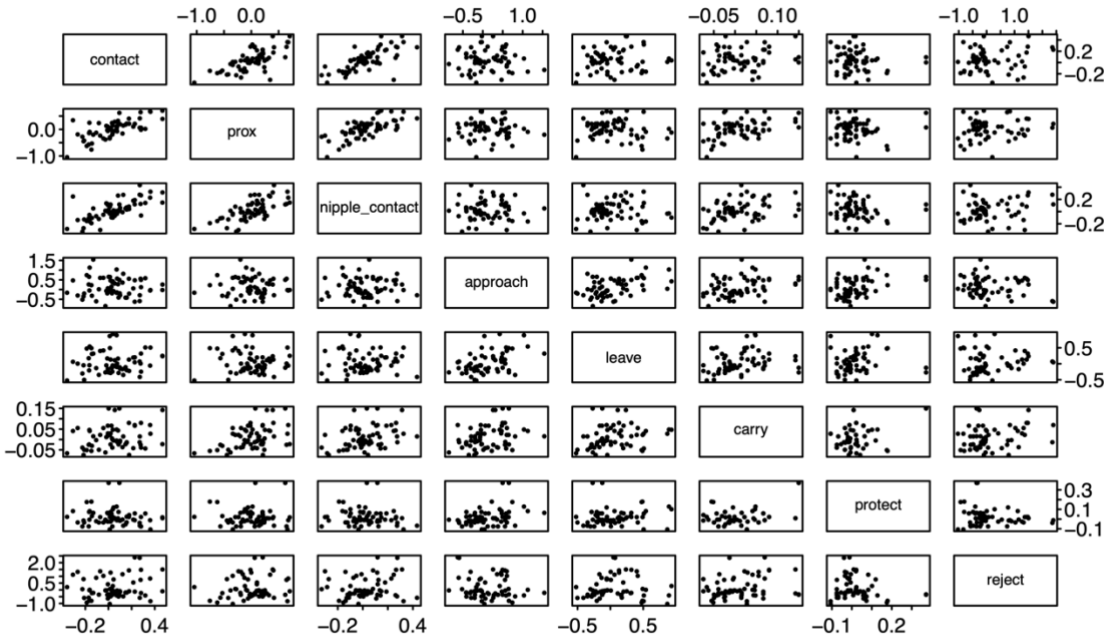

Figure S5 Scatter plot depicting the correlations between the eight behaviours that showed sufficient variation between mothers (i.e., >0.1 Standard Deviation). Each point represents the estimated BLUP for a given mother-infant pair

Table S5 Table with detailed factor loadings of the considered behaviours on the two-factors

	Factor1	Factor2
Contact-sit	0.834	-0.029
Time 1m proximity	0.833	0.02
Nipple contact	0.831	0.062
Maternal approaches	-0.045	0.768
Maternal leaves	-0.007	0.525
Carry	0.389	0.51
Maternal protections	-0.112	0.492
Maternal rejections	0.199	-0.048

S6. Ecological conditions (NDVI)

We controlled for environmental conditions due to the potential effect of food availability on the infant survival and growth. The Normalized Difference Vegetation Index (NDVI) was used as a proxy for food availability. The NDVI is a widely used bio-geophysical indicator of the land's green plant coverage - or vegetation density - and, therefore, it is considered an indirect measure of rainfall. The NDVI values were obtained from the Copernicus Global Land Service (CGLS) (https://land.copernicus.eu/global/), which is part of the Earth Observation programme of the European Commission. We downloaded the NDVI values at 300 m spatial resolution from 2016-2022 (accessed in July 2022); although the behavioural data for the infants of the study were first collected in 2017, we included NDVI data since 2016 because some of the infants in the survival model were born during that year. Values of NDVI were based on a 10-day maximum composite value derived from daily acquisitions of Earth land surface reflectance observations from the sensor PROVA-B (product version 1, from 2016-2020: Swinnen & Carolien, 2016) and sensor Sentinel-3 OLCI (version 2, from 2021: Swinnen & Carolien, 2022).

We limited values of NDVI to our study region (Lat 13º0’14.252'' to 13º5'39.647''; Long -13°13'24.813'' to -13°18'54.438'') and used the function ncvar_get of the package ncdf4 (version 1.19: Pierce, 2021) in R (R Core Team 2022) to extract the NDVI values from the downloaded files. We excluded non-optimal observations (digital value >250), which created interferences due to their origin in cloud coverage and ocean. In case of missing NDVI values, we computed a linear interpolation between the last and next days for which we had available data (Fig. S6). The daily NDVI values were averaged across the study area and later used as a control fixed effect for our main growth and survival models.

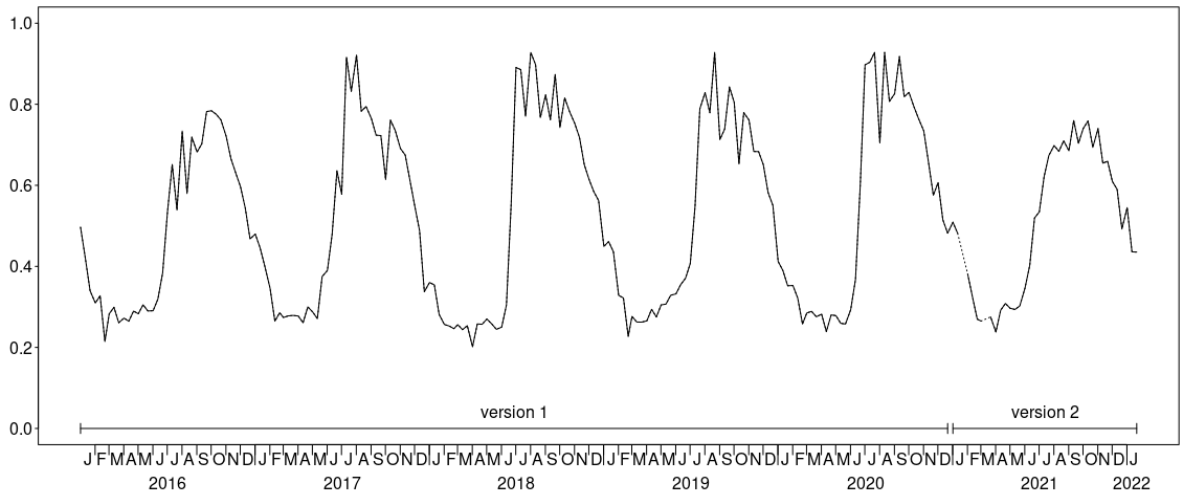

Figure S6 NDVI values in the study region during the study period extracted from versions 1 and 2 of the Copernicus Global Land Service (CGLS). Note the pronounced seasonal variation in plant productivity, being lowest in the driest months (March-May) and most pronounced in the rainy season (June-October). The dashed line represents interpolated values.