

Further confirmation for unknown archaic ancestry in Andaman and South Asia.

Mayukh Mondal¹, Ferran Casals², Partha P. Majumder³, Jaume Bertranpetit¹

1 Institut de Biologia Evolutiva (UPF-CSIC), Universitat Pompeu Fabra, Barcelona, Catalonia, Spain

2 Servei de Genòmica, Universitat Pompeu Fabra, Barcelona, Catalonia, Spain

3 National Institute of BioMedical Genomics, Kalyani, West Bengal 741251, India

In a recent paper¹, we have derived three main conclusions: i) that all Asian and Pacific populations share a single origin and expansion out of Africa, contradicting an earlier proposal of two independent waves; ii) that populations from South and Southeast Asia harbor a small proportion of ancestry from an unknown extinct hominin – different from the Neanderthal and the Denisovan – which is absent in Europeans; and, iii) that the characteristic distinctive phenotypes (including very short stature) of Andamanese do not reflect an ancient African origin, but have resulted from strong natural selection on genes related to human body size. Although the single wave out of Africa² and single origin for Asian and Pacific populations have been confirmed³, the existence of admixture with an extinct hominin has been challenged by Skoglund et al.⁴, as they were unable to replicate our results in their data sets. While we had used a wide variety of statistical methods and data sets from diverse populations to draw our inference, Skoglund et al.⁴ have used only one method (D-stats⁵, for the whole genome, not specifically for the relevant genomic regions) and compared only with the Asians, not even with the Europeans. Skoglund et al.⁴ have alleged that our statistical treatment of the data was faulty and have pointed out some possible sources of error. We have reexamined our data focusing on possible sources of error flagged by Skoglund et al.⁴. We have also performed new analyses. The reexamination and new analyses have bolstered our confidence that our earlier inferences were correct and have resulted in an improved model of introgression of modern humans with a hitherto unknown archaic ancestry. We also propose a possible reason for the inability of Skoglund et al.⁴ to validate our inference.

In our reanalysis, first we test the impact on our inference of possible sources of

bias, including the number of individuals, differences in coverage and batch effects. Second, we specifically address questions related to results in populations other than the Andamanese in relation to the detection of introgressed regions. In all cases we considered the ancestral position for humans as defined in 1000 Genomes Project as outgroup⁶ and not the chimpanzee, although the results are similar.

1.- Number of Individuals. Some of our analyses comprised 70 Indians compared to two Europeans, two East Asians and four Africans. This variation in sample size may have introduced a bias towards Indian-specific derived alleles. To check this, we have now performed variant calling using reads obtained by deep-sequencing of two Europeans, two Andamanese, one Papuan and two Africans together. Estimates obtained from Dstat analysis of these data are similar to those reported by us in Mondal et al¹ and although the estimated values are slightly smaller, these values are similar to our original results and are still statistically significant (Table 1). Therefore, we rule out the possibility that because of heterogeneity in the number of individuals of different ethnicities included in our original analysis, we have made a false positive inference of introgression (at least for the Andamanese).

2.- Coverage. Our data coverage is lower (15x) than in the comparison set (an initial set of the Simon Genome Diversity Project [SGDP]⁷). To test the possible impact of differences in sequencing coverage (as suggested by David Reich in a personal communication to us), we downgraded SGDP data to 15x coverage using -f flag in samtools⁸. The results (Table 2) are again statistically significant and are similar to those previously obtained¹, suggesting that effect of differences in depth of coverage had not resulted in a false positive inference on introgression from an extinct hominin.

3.- Batch effects due to laboratory of origin and processing of the sequences. To check the impact of batch effects, we first downloaded the 1000 Genomes Project⁹ (KGP) vcf files and performed the D-stats analysis. The result obtained when we used Indian Telugu from the UK (ITU) to detect fewer shared African alleles using Yoruba in Ibadan, Nigeria (YRI) compared to Europeans using Utah Residents with Northern and Western Ancestry (CEU) [$D(\text{CEU}, \text{ITU}; \text{YRI}, \text{Ancestral}) = -.0009$ (Zscore=-0.933)] was not statistically significant. However, since data imputation was done in the KGP, we suspected that the statistically non-significant result that we obtained may be caused by having included imputed data. In fact, if Indian populations admixed with an unknown hominin population, the introgressed haplotypes would be at very low frequencies and therefore imputed results would be unreliable (a similar problem might arise if instead of sequence data, only SNP genotyping data would be used). We downloaded 10 CEU, 10 YRI and 10 ITU BAM files, converted to the same coverage (as the coverage is substantially divergent in this data set) and performed the variant calling without any imputation. The result we obtained was clearly in the same direction originally presented in Mondal et al¹ and was statistically significant [$D(\text{CEU}, \text{ITU}; \text{YRI}, \text{Ancestral}) = 0.0079$ (Zscore=3.747)].

Next, we used two CEU and two YRI (two individuals randomly chosen from KGP) with two Andamanese (AND) from our data, and also obtained a significant result (although the lower coverage of the KGP data markedly decreased the level of statistical significance) [$D(\text{CEU,AND};\text{YRI,Ancestral})=0.0109$ ($Z\text{score}=3.595$)]. Then we retrieved fasta formatted data from the Simon Genome Diversity Project³ including one of each, Papuan, French, Sardinian, Dai, Han, Mandenka, Yoruba and Irula. We noted that variant-calling and filtering were already carried out on the raw data; we did not have the opportunity to carry these out ourselves. We obtained results similar to those presented in Mondal et al. (2016) for Papuan (Table 3), but highly variable and non-significant results for Dai, Han or Irula. While we have no way to identify the cause of these differences in results, we suspect that differences in treatment of raw data may be the contributing factors. We note that Mallick et al³ have used a new method for calling and filtering, which may have strong effects on calls of rare and interesting variants.

4.- Results in other Asian and Pacific populations. Our study was mainly focused on the Andamanese population. We accept the suggestion that a more in depth introgression analysis should be performed in other populations using variant calls derived from a joint analysis of raw data. Regarding the Papuan and Australian populations, we think that the differences between them are related to variant calling artifacts, since only the vcf file was available for the Australian individuals. Further studies will have to consider the larger data set recently studied². In relation to the East Asian populations, we did not get significant evidence of introgression, although the results of our analysis point to the possibility that East Asians also may have some introgression from this unknown hominin, even though the proportion of introgression may be lower than in the Andamanese.

5.- Finding Introgressed Regions. We have now performed analyses to detect introgressed regions in the Andamanese genomes using D-stats by genomic region and S^* in our data. D-stats by region exhibited good statistical power to detect regions with strong divergence (that may have introgressed) from the simulated model (Figure 1) described in the methods section of Mondal et al¹. It is clear that the differences in distributions of the D-stats by position correctly separated the two (introgressed and non-introgressed) distributions. If the results of D-stats and S^{*10} are combined, the extent of false positives decreased even further (Figure 2). The vertical line indicates the significance threshold for S^* . The yellow dots in the graph mark the introgressed regions detected by the combined method, positive for both D-stats by position (defined as regions having -1 value) and S^* (defined as regions having more than 95 percentile of the distribution of the non-introgressed model).

Finally, we analyzed the regions which were more likely to be introgressed for both D-stats and S^* (which accounted for a total of ~15Mb per individual), by comparing to other hominin sequences and resulted clearly in a sequence basal to the Neanderthal-Denisovan clade (Figure 3).

Discussion

To the extent feasible, we have examined all putative sources of bias⁴ – such as batch effects, sample size and coverage differences, data processing artifacts – that may have impacted on our earlier inference¹ of the presence of an unknown hominin ancestry in the Andamanese population. We have shown that our earlier inference was robust and that we had not made a false-positive inference. We have replicated the results originally presented by us¹ using additional data, independently of the origin of the data, when we were able to obtain raw data and analyze the data using a common set of methods. We wonder whether the new method of calling and filtering raw data introduced by Mallick et al³ and likely used by Skoglund et al⁴ may be the reason for the differences in results obtained by Skoglund et al⁴ and Mondal et al¹. In addition, results of the present analyses of the introgressed regions further support our earlier conclusions¹.

Finally, since in population genetics studies, disparate data sets generated by various laboratories are used, we propose that, in a manner similar to what has been done for somatic mutation detection in cancer¹¹, the population genetics research community should initiate a joint effort to study the impact of using different variant-calling and filtering procedures on population genetic inferences. We are, of course, willing to contribute data and computing effort to such an endeavor (which could comprise a subset of the sequences of the SGDP³).

Methods

Variant Calling

The BAM files used were previously generated¹ or downloaded from 1000 Genomes Project site⁹. In some of the analysis we downgraded the coverage of BAM files using -f flag from samtools-1.2⁸. Variant calling was done by GATK-3.5¹² closely following “best practices of GATK”. Calling was done with default options and using --max-alternate_alleles 20 (to capture all genetic diversity present in the populations) by running HaplotypeCaller in GVCF mode on each sample separately. Then GenotypeGVCFs was used to produce the joint genotyping of the gVCFs on all of the samples together to generate a raw SNP and indel Variant Calling File (VCF). The raw VCF was filtered using post variant calling recalibration steps as listed in GATK “Best Practices”. VariantRecalibration from GATK were used to calculate various statistics for novel variants (both for SNPs and indels) and then recalibrated according to their needs using ApplyRecalibration. As different data sets have differing amounts of false positives, it was necessary to give the appropriate importance to the different data sets for the training algorithm. We took the value for prior likelihood of true sites defined in the GATK website for the corresponding data sets.

For the recalibration steps, the following datasets were used, with the following commands: SNPs

1. dbSNP version 137: -resource:dbsnp, known=true, training=false, truth=false, prior=2.0.
2. hapmap version 3.3: -resource:hapmap, known=false, training=true, truth=true, prior=15.0 (International Hapmap3 Consortium 2010).
3. Omni genotyping array 2.5 million 1000G: -resource:omni, known=false, training=true, truth=true, prior=12.0.

4. 1000G phase 1 high confidence: -resource:1000G, known=false, training=true, truth=false, prior=10.0.
5. With the flags -an QD -an MQRankSum -anReadPosRankSum -an FS -an DP -an InbreedingCoeff. All other parameters were set to default values.

Indels

1. Mills 1000G high confidence indels: -resource:mills, known=false, training=true, truth=true, prior=12.0.
2. dbSNP version 137: -resource:dbsnp, known=true, training=false, truth=false, prior=2.0.

With the flags --maxGaussians 4 -an FS -an ReadPosRankSum -an MQRankSum -an DP -an InbreedingCoeff. All other parameters were set to default values.

D-stats Calculation

We first converted the vcf file to plink format using vcftools¹³ and we added ancestral information, which was downloaded from 1000 Genomes Project site, using plink¹⁴. We converted plink to eigensoft format using convertf and D-stats was calculated using qpDstat both of them came together with Admixtools 1.1⁵. All the D-stats calculations were done on SNPs which are present in every Individual thus not biasing for allele frequency which present in low covered regions.

SGDP 3rd dataset

We downloaded SGDP files from the site

http://sharehost.hms.harvard.edu/genetics/reich_lab/cteam_lite_public3.tar, 11/07/2016

Individual information was extracted using cpoly³. Individuals used in this analysis were: Chimp, B_French-3, B_Mandenka-3, B_Mbuti-4, B_Yoruba-3, S_Irula-1, B_Papuan-15, B_Sardinian-3, B_Dai-4 and B_Han-3.

Simulations

We used ms to simulate demographic models¹⁵. The Andamanese demographic model is based on Mondal et al¹.

To calculate efficiency, two different models were used. 1) The null model: we used Andamanese demography with modern humans without using any introgression from hominin. 2) The H1 model: we simulated a hominin population which has diverged from modern humans 300 kya ago (as mutation rate 2.3×10^{-8} was used). This is the best scenario. In both cases we simulated around 50 kb of 60,000 regions. We chose a specific number of segregating sites of 100 and recombination rate of 1.3×10^{-8} .

```
Null=ms 20 60000 -s 100 -r 26 50000 -l 3 4 4 2 -n 1 1.4474 -n 2 3.3814 -n 3 .1055 -m 1 2 1.244 -m 2 1 1.244 -m 2 3 0.88 -m 3 2 0.88 -g 2 179 -ej .023 3 2 -en .023 2 .1861 -em .023 1 2 6 -em .023 2 1 6 -em .023 1 3 0 -em .023 3 1 0 -em .023 2 3 0 -em .023 3 2 0 -ej .051 2 1 -en .148 1 .73 -p 5
```

```
H1=ms 10 60000 -s 100 -r 26 50000 -l 3 4 4 2 -n 1 1.4474 -n 2 3.3814 -m 1 2 1.244 -m 2 1 1.244 -g 2 179 -en .023 2 .1861 -em .023 1 2 6 -em .023 2 1 6 -em .023 1 3 0 -em .023 3 1 0 -em .023 2 3 0 -em .023 3 2 0 -ej .051 2 1 -en .148 1 .73 -ej .4 3 1 -p 5
```

D-stats⁵ and S^{*10} was calculated as mentioned in their respective.

Data repository:

vcf files from Andamanese, Indian and a set of comparison from an initial set of the SGDP⁷ are available temporarily in:

https://www.dropbox.com/s/rk6jijocefhwns1/GreatApe_Indiafinal_Autosome.vcf.gz?dl=0

and will be uploaded soon in a general server.

References

1. Mondal, M. *et al.* Genomic analysis of Andamanese provides insights into ancient human migration into Asia and adaptation. *Nat. Genet.* 1–102 (2016). doi:10.1038/ng.3621
2. Malaspina, A.-S., *et al.*. The genomic history of Indigenous Australia. *Nature*, (2016), in press.
3. Mallick, S. *et al.* The Simons Genome Diversity Project: 300 genomes from 142 diverse populations. *Nature* (2016), in press.
4. Skoglund, P., Mallick, S., Patterson, N. & Reich, D. No evidence for unknown archaic ancestry in South Asia. *Nature Genetics* (2016) this issue.
5. Patterson, N. *et al.* Ancient admixture in human history. *Genetics* **192**, 1065–1093 (2012).
6. Abecasis, G. R. *et al.* An integrated map of genetic variation from 1,092 human genomes. *Nature* **491**, 56–65 (2012).
7. Meyer, M. *et al.* A High-Coverage Genome Sequence from an Archaic Denisovan Individual. *Science* **338**, 222–226 (2012).
8. Li, H. *et al.* The Sequence Alignment/Map format and SAMtools. *Bioinformatics* **25**, 2078–2079 (2009).
9. Auton, A. *et al.* A global reference for human genetic variation. *Nature* **526**, 68–74 (2015).
10. Vernot, B., Akey, J. M. & Vernot B., A. J. M. Resurrecting surviving Neandertal lineages from modern human genomes. *Science* **343**, 1017–1021 (2014).
11. Alioto, T. S. *et al.* A comprehensive assessment of somatic mutation detection in cancer using whole-genome sequencing. *Nat Commun* **6**, 10001 (2015).
12. McKenna, A. *et al.* The genome analysis toolkit: A MapReduce framework for analyzing next-generation DNA sequencing data. *Genome Res.* **20**, 1297–1303 (2010).
13. Danecek, P. *et al.* The variant call format and VCFtools. *Bioinformatics* **27**, 2156–2158 (2011).
14. Purcell, S. *et al.* PLINK: a tool set for whole-genome association and population-based linkage analyses. *Am. J. Hum. Genet.* **81**, 559–575 (2007).
15. Hudson, R. R. Generating samples under a Wright-Fisher neutral model of genetic variation. *Bioinformatics* **18**, 337–338 (2002).

Figure 1: Efficiency of detecting putative introgressed regions by D-stats. The blue bars represent the standard distribution without introgression, and the red bars represent the divergent distribution that may have introgressed; both cases refer to a 50Kb region simulated 60,000 times. Not drawn to scale, as the red distribution would be much less frequent than the blue one in a scenario of small amount of introgression.

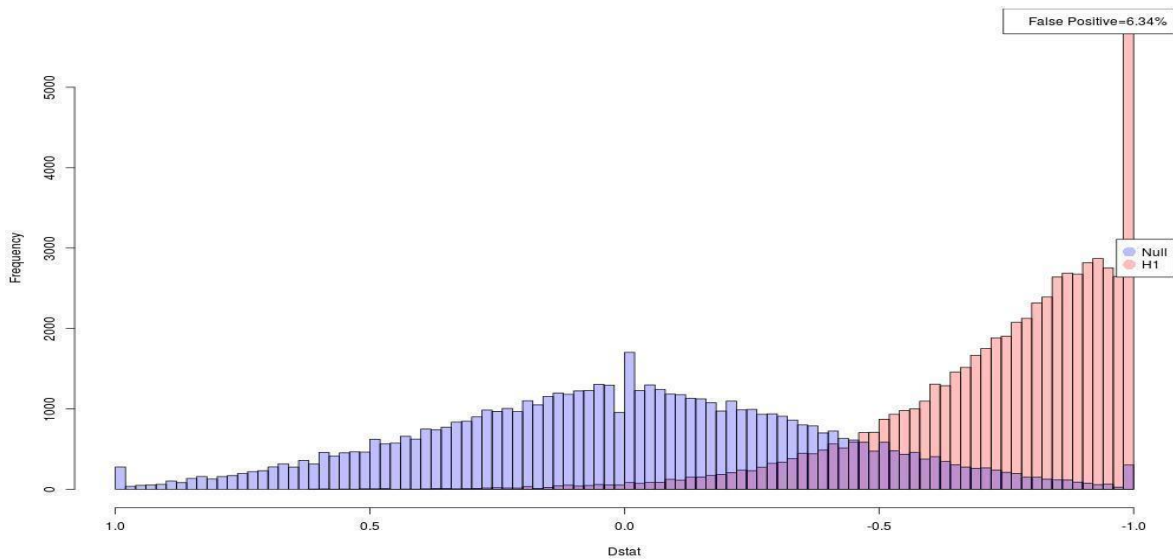


Figure 2: Efficiency of detecting putative introgressed regions by D-stats and S^* in the simulations as in Figure 1. The blue points are the non-introgressed regions, the red points are the regions with high divergence, and yellow points are the regions that are extreme for both D-stats and S^* and are thus considered introgressed.

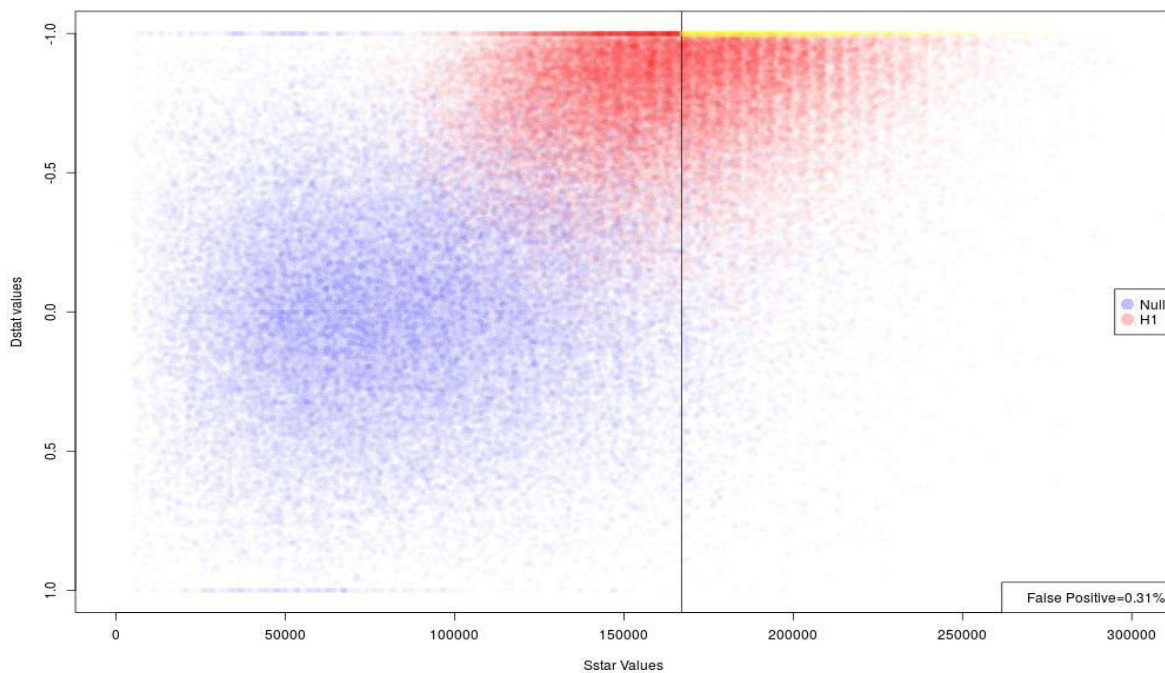


Figure 3: Hierarchical cluster analysis on the dissimilarity matrix of individual sequences. Modern humans: YRI, Yoruba, DAI from China, French, French. Other sequences: JAR-61, putative introgressed regions in a Jarawa, AltaiNea, Neanderthal from Altai, DenisovaPinky, Denisova.

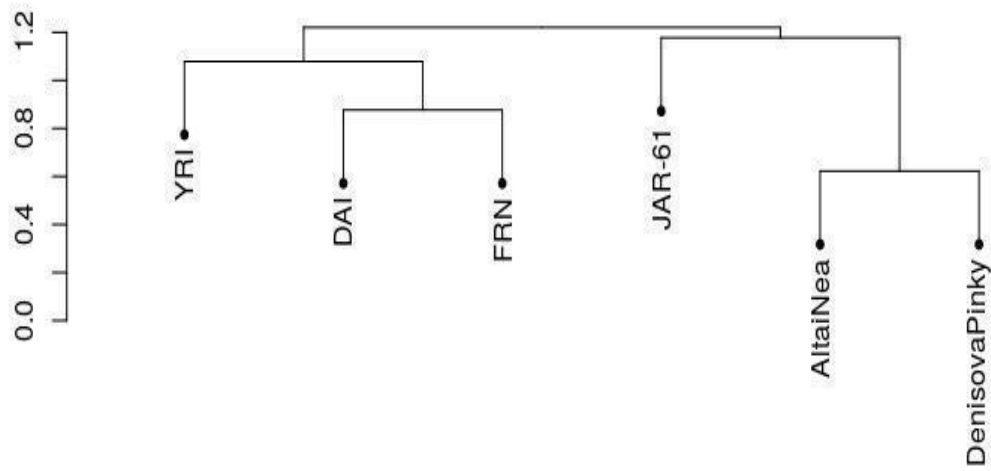


Table 1: Results of D-stats analysis using two individuals from each, Europe (French and Sardinian) and Africa (Yoruba and Mandenka), from an initial set of the SGDP⁷ and Andaman¹ (Jarawa and Onge) and one Papuan⁷.

W	X	Y	Z	D score	Z score
French	Jarawa	Yoruba	Ancestral	0.0181	4.3
French	Onge	Yoruba	Ancestral	0.0158	3.427
French	Jarawa	Mandenka	Ancestral	0.0188	4.419
French	Onge	Mandenka	Ancestral	0.0135	2.936
Sardinian	Jarawa	Yoruba	Ancestral	0.0188	4.223
Sardinian	Onge	Yoruba	Ancestral	0.0166	3.543
Sardinian	Jarawa	Mandenka	Ancestral	0.0194	4.691
Sardinian	Onge	Mandenka	Ancestral	0.0142	3.166
French	Papuan	Yoruba	Ancestral	0.0336	6.893
French	Papuan	Mandenka	Ancestral	0.0269	5.439
Sardinian	Papuan	Yoruba	Ancestral	0.0344	6.847
Sardinian	Papuan	Mandenka	Ancestral	0.0277	5.782

Table 2: Results of D-stats analysis using similar coverage for all individuals (15x). French, Sardinian, Yoruba and Mandenka are from an initial set of the SGDP⁷; Jarawa and Onge are from Mondal et al.¹.

W	X	Y	Z	D score	Z score
French	Jarawa	Yoruba	Ancestral	0.015	3.945
French	Onge	Yoruba	Ancestral	0.0127	3.285
French	Jarawa	Mandenka	Ancestral	0.0129	3.202
French	Onge	Mandenka	Ancestral	0.0118	2.889
Sardinian	Jarawa	Yoruba	Ancestral	0.0167	4.011
Sardinian	Onge	Yoruba	Ancestral	0.0145	3.462
Sardinian	Jarawa	Mandenka	Ancestral	0.0144	3.677
Sardinian	Onge	Mandenka	Ancestral	0.0134	3.373

Table 3: Results of D-stats analysis using data from SGDP (3rd data set)⁷: one Papuan, one French, one Sardinian, one Dai, one Han and one Irula.

W	X	Y	Z	D score	Z score
French	Papuan	Mandenka	Chimp	0.0403	7.313
French	Papuan	Mbuti	Chimp	0.0229	4.526
French	Papuan	Yoruba	Chimp	0.0303	5.681
Sardinian	Papuan	Mandenka	Chimp	0.0376	6.847
Sardinian	Papuan	Mbuti	Chimp	0.0323	6.492
Sardinian	Papuan	Yoruba	Chimp	0.0401	7.626
French	Dai	Mandenka	Chimp	0.0124	2.755
French	Dai	Mbuti	Chimp	-0.0139	-3.228
French	Dai	Yoruba	Chimp	-0.0046	-0.986
Sardinian	Dai	Mandenka	Chimp	0.01	2.175
Sardinian	Dai	Mbuti	Chimp	-0.0039	-0.839
Sardinian	Dai	Yoruba	Chimp	0.0058	1.188
French	Han	Mandenka	Chimp	0.0107	2.224
French	Han	Mbuti	Chimp	-0.0024	-0.544
French	Han	Yoruba	Chimp	0	-0.007
Sardinian	Han	Mandenka	Chimp	0.0082	1.593
Sardinian	Han	Mbuti	Chimp	0.0074	1.563
Sardinian	Han	Yoruba	Chimp	0.0103	2.077
French	Irula	Mandenka	Chimp	0.007	1.491
French	Irula	Mbuti	Chimp	-0.0081	-1.799
French	Irula	Yoruba	Chimp	-0.0012	-0.279
Sardinian	Irula	Mandenka	Chimp	0.0045	0.988
Sardinian	Irula	Mbuti	Chimp	0.0019	0.445
Sardinian	Irula	Yoruba	Chimp	0.0093	1.986