Omics-based Hybrid Prediction in Maize

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May 5, 2017

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1 Key message

2 Complementing genomic data with other "omics" predictors can increase the proba-
3 bility of success for predicting the best hybrid combinations using complex agronomic
4 traits.

5 Abstract

6 Accurate prediction of traits with complex genetic architecture is crucial for select-
7 ing superior candidates in animal and plant breeding and for guiding decisions in
8 personalized medicine. Whole-genome prediction (WGP) has revolutionized these
9 areas but has inherent limitations in incorporating intricate epistatic interactions.
10 Downstream "omics" data are expected to integrate interactions within and between
11 different biological strata and provide the opportunity to improve trait prediction.
12 Yet, predicting traits from parents to progeny has not been addressed by a combi-
13 nation of "omics" data. Here, we evaluate several "omics" predictors — genomic,
14 transcriptomic and metabolic data — measured on parent lines at early developmen-
15 tal stages, and demonstrate that the integration of transcriptomic with genomic data
16 leads to higher success rates in the correct prediction of untested hybrid combinations
17 in maize. Despite the high predictive ability of genomic data, transcriptomic data
18 alone outperformed them and other predictors for the most complex heterotic trait,
19 dry matter yield. An eQTL analysis revealed that transcriptomic data integrate
20 genomic information from both, adjacent and distant sites relative to the expressed
21 genes. Together, these findings suggest that downstream predictors capture phys-

22 iological epistasis that is transmitted from parents to their hybrid offspring. We
23 conclude that the use of downstream "omics" data in prediction can exploit impor-
24 tant information beyond structural genomics for leveraging the efficiency of hybrid
25 breeding.

26 **Conflict of Interest**

27 The authors declare that they have no conflict of interest.

28 Introduction

29 Hybrid breeding, which entails crossing of lines from two genetically distant germplasm
 30 collections — called heterotic groups (Melchinger and Gumber, 1998) — has emerged
 31 as a prime strategy to meet demands for a sustainable intensification of agricultural
 32 production (Duvick, 2005). However, unlocking the full potential of hybrid breeding
 33 requires accurate prediction methods to efficiently identify the superior candidates
 34 out of the millions of possible hybrids that could potentially be produced in each cy-
 35 cle of an ordinary-sized breeding program. With the advent of the doubled haploid
 36 (DH) technology (Wedzony et al., 2009) this prediction problem has become even
 37 more challenging because, based on breeder's experience, the vast majority ($\approx 90\%$)
 38 of competing lines in each heterotic group are "new" lines without any phenotypic
 39 records on hybrid progeny from previous breeding cycles. Consequently, among all
 40 hybrid combinations possible between lines from two heterotic groups, about 81%
 41 are T0 hybrids, 18% are T1 hybrids and 1% are T2 hybrids having zero, one or
 42 two parents, respectively, that have been previously tested in other hybrid combina-
 43 tions. Preselection of a few hundred of the most favorable hybrids with high success
 44 rate could significantly reduce the labor-intensive and time-consuming field-testing
 45 (Kadam et al., 2016, Xu et al., 2016). This could greatly impact the efficiency of
 46 hybrid breeding and boost the annual selection gain (Longin, Mi and Würschum,
 47 2015).

48 Whereas yield and other heterotic traits of hybrids are generally poorly predicted
 49 by the performance of their parent lines (Melchinger and Gumber, 1998), WGP has
 50 emerged as a major tool for tackling this challenge (Massman et al., 2013, Technow

et al., 2014). Nevertheless, there is evidence that, even with complete sequence information, genomic prediction may not capture complex interactions between genes and downstream regulation, which act through the entire cascade from genotype to phenotype (Dalchau et al., 2011, Zhu et al., 2012, Rudd et al., 2015, Ritchie, Holzinger, Li, Pendergrass and Kim, 2015). Most studies have evaluated predictive ability by looking at only one kind of endophenotype (intermediaries between genotype and phenotype (Gottesman and Gould, 2003, Mackay, Stone and Ayroles, 2009) such as the transcriptome (Swanson-Wagner et al., 2006, Zenke-Philippi et al., 2016, Xu et al., 2016) or the metabolome (Riedelsheimer et al., 2012, Xu et al., 2016, Dan et al., 2016). The integration of different endophenotypic and genomic data is expected to reflect more closely the variability across genotypes than genomic data alone (Mackay, Stone and Ayroles, 2009, Patti, Yanes and Siuzdak, 2012, Civelek and Lusis, 2014). Two recent studies that integrated multiple biological strata in predicting breast cancer risk (Vazquez et al., 2016) and performance of maize inbred lines (Guo et al., 2016), respectively, demonstrated the benefit of this strategy. However, unlike forecasting clinical or agronomic traits from endophenotypes of the same genotype, hybrid breeding requires the prediction of the genotypic values (GV) of hybrid progeny based on parental information. To achieve this objective, we used the BLUP approach — originally developed in animal breeding (Henderson, 1984) — for the more complex setting of hybrids between parents from two heterotic groups (Bernardo, 1996, Massman et al., 2013). Here, we measured endophenotypes of parent lines to forecast the GV of T0, T1 and T2 hybrid progeny by using prediction equations trained with "omics" information from other parent lines and phenotypic

74 information on their hybrid offspring.

75 Materials and Methods

76 Genetic material and phenotyping

77 The entire genetic material consisted of a set of 1,536 hybrids, denoted as H_{Tot} ,
78 produced in 16 factorial mating designs between 142 Dent and 103 Flint lines from
79 the maize breeding program at the University of Hohenheim, on which agronomic
80 data for silage maize production, as well as pedigree and genomic data were available.
81 A subset of this material, albeit from different trials, has been used for genomic
82 prediction of traits related to grain maize production (Technow et al., 2014). For
83 hybrid prediction, we used a core set $H \subset H_{Tot}$ of 617 hybrids, produced in six
84 factorials with hybrid sets $H_{FAC(i)} (i = 1, 2, \dots, 6; H = \bigcup_{i=1}^6 H_{FAC(i)})$ from crosses
85 between 57 Dent and 41 Flint inbred lines, denoted as $D = \{1, 2, \dots, 57\}$ and $F =$
86 $\{1, 2, \dots, 41\}$ (File S1). All hybrids were evaluated in field experiments at three
87 or more agro-ecologically diverse locations across Germany. In the trials of each
88 factorial, which included at least five common check genotypes, the entries were
89 randomized in α lattice designs and planted in 2-row plots. Dry matter yield (DMY,
90 t/ha) and dry matter content (DMC, %) of whole-plant aboveground biomass were
91 determined by established procedures (Riedelsheimer et al., 2012). For quality traits,
92 contents of fiber (ADF, %), fat (FAT, %), protein (PRO, %), starch (STA, %), and
93 sugars (SUG, %) in dry matter were measured in the harvested plant material using
94 calibrated near-infrared spectroscopy (NIRS; Grieder et al. (2011), File S1).

95 Pedigree-based relationship coefficients

96 Coancestry coefficients were calculated using SAS (version 9.4, SAS Institute) for all
 97 possible pairs of lines in each heterotic group according to established rules (Falconer
 98 and Mackay, 1996) under the following assumptions (Cox, Murphy and Rodgers,
 99 1986): (i) all lines in a pedigree are genetically homogeneous and homozygous, (ii)
 100 pairs of genotypes with no known common parentage are unrelated, and (iii) a line
 101 derived from a cross or backcross obtained a proportional fraction of the genome from
 102 each parent, as expected under Mendelian inheritance in the absence of selection.

103 Genotyping

104 Genotyping of all inbred lines was performed with the Illumina SNP chip MaizeSNP50
 105 (Ganal et al., 2011). After performing a commonly used quality check (Technow
 106 et al., 2014) and imputation of missing data (Browning and Browning, 2009), a total
 107 of 21,565 polymorphic SNPs was available and used for all further analyses.

108 Metabolite profiling

109 Seedlings of all parental inbred lines were grown under controlled conditions inside
 110 climate chambers to quantify the metabolite profiles of their roots 3.5 days after
 111 sowing, as detailed by de Abreu e Lima et al. (2017). The experiment was laid out
 112 as a randomized incomplete block design with replicated germination boxes. For leaf
 113 metabolic profiles (known metabolites and unannotated chromatographic peaks), a
 114 field experiment was carried out in an α lattice design with two replications at one

location in southern Germany in the spring of 2012. Excision of leaves at the third leaf stage was performed according to an established protocol (Riedelsheimer et al., 2012) 28 days after sowing, in the afternoon of a cloudy day, and finalized within 45 minutes for the entire experiment. For both profiling procedures, all material was transferred directly into containers with dry ice and then into liquid nitrogen to quench metabolic activity.

Transcriptome profiling

For transcriptome profiling, five seeds per parent line were taken from the same seed lot as used for metabolite profiling and laid out inside a climate chamber in a randomized complete block design with five replications. Seedlings were sampled seven days after sowing, snap-frozen in liquid nitrogen, and stored at -80°C until use. Prior to mRNA extraction, roots from all replicates of a genotype were pooled and homogenized. A custom 2K-microarray (GPL22267) was assembled from a subset of the 47K maize oligonucleotide array (GPL6438). Two-color hybridizations were carried out separately for each of the six factorials using interwoven loop designs (Kerr and Churchill, 2001). The average number of shared genotypes between factorials was 4.5 and ranged from 2 to 10.

Statistical analysis of agronomic traits

Agronomic data were analyzed in two stages, following Technow et al. (2014) by accounting for year, location, field replication, block and genotype effects as well as their interactions (File S1). In the first stage, and separately for each environment,

best linear unbiased estimates (BLUEs) of the α designs were computed for every hybrid using REML-based linear mixed-model analyses. In the second stage, BLUEs were computed for all hybrids in H_{Tot} . The BLUEs of hybrids in the core set H served as response variables in our hybrid prediction models and cross-validation routines. For all predictions we used computationally efficient best linear unbiased predictor (BLUP) models, which have the same properties as those of a selection index because we previously accounted for fixed effects (Mrode (2014), pp. 34, 311, 312). For general and specific combining abilities (GCA and SCA) of parent lines, we used ASReml (Butler et al., 2009) to compute best linear unbiased predictors (BLUPs), variance components ($\sigma_{GCA^D}^2$, $\sigma_{GCA^F}^2$, σ_{SCA}^2) and entry-mean heritabilities (H^2) of all hybrids in H_{Tot} , treating all effects in the model as random. The covariance matrices of the GCA and SCA effects were defined by multiplying the variance components with their respective genomic relationship matrices (File S1).

Statistical analysis of endophenotypes

Raw data were normalized using established procedures for metabolites (van den Berg et al., 2006) and transcripts (Smyth and Speed, 2003, Ritchie et al., 2007). From these data, we obtained BLUEs for metabolite levels and transcript abundance of each line using REML-based mixed-model analyses. The statistical models for the analysis of metabolite profiles accounted for various experimental effects as detailed by de Abreu e Lima et al. (2017). After applying quality checks and computing BLUEs, 92 leaf metabolic analytes and 283 root metabolic analytes remained for further analyses. BLUEs for transcriptomic data were computed using the R-package *limma* (Ritchie,

158 Phipson, Wu, Hu, Law, Shi and Smyth, 2015) in reference to established protocols
 159 (Smyth and Speed, 2003, Ritchie et al., 2007, Frisch et al., 2010). The design matrix
 160 for the linear model was based on the dye-labeling of a reference genotype. To account
 161 for possible differences between the microarrays, replicates of some genotypes across
 162 at least two factorials were included, and modeled through a fixed effect term. All
 163 gene expression values were subsequently computed, based on the log-ratio relative
 164 to this common genotype (Smyth, 2004). In total, 1,323 gene expression profiles were
 165 available. Repeatabilities (w^2) were estimated for each endophenotype at the inbred
 166 line level using the same models as for the computation of BLUEs, but treating the
 167 genotype effect as random. This analysis was performed jointly for the Dent and
 168 Flint lines allowing for different means and heterogeneous genotypic variances of the
 169 heterotic groups, but assuming a common error variance. Variance components were
 170 estimated by Gibbs sampling using the *R* package *MCMCglmm* (Hadfield, 2010).

171 Prediction models and model evaluation

172 Predictions of hybrid performance were compared on the basis of the core set of
 173 hybrids H and the corresponding sets of parent lines D and F on which data for
 174 all five predictors (P, pedigree; G, genomic; T, transcriptomic; L, leaf metabolic;
 175 R, root metabolic data) were available with the exception of data on a few lines
 176 missing at random for R due to fungal contamination. The matrices $\mathbf{W}_D$ and $\mathbf{W}_F$
 177 are matrices of standardized feature measurements for the various predictors (G, T,
 178 L, R). The matrix $\mathbf{W}$ has dimension 'number of parent lines in the corresponding
 179 heterotic group' ($n_D = 142$, $n_F = 103$) times 'number of features' ($w_G = 21,565$,

180 $w_T = 1,323$, $w_L = 92$, $w_R = 283$). The columns in $\mathbf{W}_D$ and $\mathbf{W}_F$ are centered and
181 standardized to unit variance, respectively.

The kernels pertaining to each predictor and lines from each heterotic group — corresponding to genomic relationship matrices in the case of SNPs — can then be defined as

$$\mathbf{G}_D = \frac{1}{W} \mathbf{W}_D \mathbf{W}_D^\top, \mathbf{G}_F = \frac{1}{W} \mathbf{W}_F \mathbf{W}_F^\top, \quad (1)$$

182 where W denotes the number of features (VanRaden, 2008). In the case of pedi-
183 gree data (P), coancestry coefficients were used directly for $\mathbf{G}_D$ and $\mathbf{G}_F$, respectively.

184 The universal model for GCA and SCA effects was:

$$\mathbf{y} = \mu + \sum_{c=1}^C \mathbf{Z}_D \mathbf{g}_{Dc} + \sum_{c=1}^C \mathbf{Z}_F \mathbf{g}_{Fc} + \sum_{c=1}^C \mathbf{Z}_S \mathbf{s}_c + \epsilon, \quad (2)$$

185 where $\mathbf{y}$ is the vector of observed hybrid performance (BLUEs), μ is the fixed
186 model intercept, $\mathbf{Z}_D$ is the corresponding design matrix associating the random GCA
187 effects of the lines in D ($\mathbf{g}_{Dc}$) with $\mathbf{y}$, $\mathbf{Z}_F$ is the corresponding design matrix asso-
188 ciating the random GCA effects of the lines in F ($\mathbf{g}_{Fc}$) with $\mathbf{y}$ and $\mathbf{Z}_S$ is a design
189 matrix associating the SCA effects ($\mathbf{s}_c$), pertaining to hybrid combinations for the
190 c -th predictor data type with the corresponding hybrid measurements in $\mathbf{y}$. Thus,
191 the model in Eq. 2 can accomodate just one ($C = 1$) or multiple ($C > 1$) predictors
192 simultaneously. The random effects ($\mathbf{g}_{Dc}$ and $\mathbf{g}_{Fc}$) have expectation zero and covari-
193 ance matrices equal to $\mathbf{G}_{Dc} \sigma_{\text{GCA}^{Dc}}^2$ and $\mathbf{G}_{Fc} \sigma_{\text{GCA}^{Fc}}^2$ for the GCA effects of the Dent
194 and Flint lines, respectively, $\mathbf{S}_c \sigma_{\text{SCA}^c}^2$ for the SCA effects and $\mathbf{I} \sigma_\epsilon^2$ for the residual er-
195 ror. For each combination between crosses of lines $i \times k$ and $j \times l$, the corresponding

elements in $\mathbf{S}_c$ were obtained as the product of the respective elements f_{ij} in $\mathbf{G}_{Dc}$ and f_{kl} in $\mathbf{G}_{Fc}$, respectively (Schnell, 1965, Henderson, 1985, Bernardo, 1996, Massman et al., 2013, Technow et al., 2014, Jiang and Reif, 2015) (File S1). Note that, in the majority of cases, only GCA effects were considered. In the absence of epistasis, this model is equivalent to a feature model accounting for dually defined additive effects in each heterotic group and dominance effects between them. Extensions of the single-predictor models were made by adding GCA and SCA effects for any additional predictor assuming stochastic independence of effects. In order to obtain unbiased estimates of the predictive ability and to compare different models and predictor combinations, following Technow et al. (2014), we devised a cross-validation (CV) scheme, stratified by the parent lines and using 1,000 runs (CV1000, File S1). All prediction models were implemented using the *R* package *BGLR* (Pérez and de Los Campos, 2014).

Comparison of predictive abilities

Predictive abilities were obtained by calculating Pearson correlations between predicted ($\hat{y}$) and observed phenotypes (y), separately for three test set partitions (T0, T1 and T2 hybrids). For each CV run, the training and validation sets were stored to ensure the validity of comparisons between any predictor and combinations thereof. For any two predictors, say A and B , we then have orthogonal vectors with predictive abilities r_A and r_B of length 'number of cross validation runs'.

216 Evaluation of a pre-selection bias in transcriptomic data

217 A custom 2K-microarray (GPL22267) was assembled from a subset of the 47K maize
 218 oligonucleotide array (GPL6438), based on association of genes with hybrid perfor-
 219 mance or mid-parent heterosis for grain yield and grain dry matter content of maize.
 220 These two traits were evaluated in separate grain-yield trials with hybrids from fac-
 221 torial $H_{\text{FAC}(1)}$ (Frisch et al. (2010), Thiemann et al. (2010), File S1). To ensure that
 222 no pre-selection bias was introduced in hybrid prediction using these transcriptomic
 223 data, we compared predictive abilities among the various predictors when excluding
 224 $H_{\text{FAC}(1)}$ from the entire set H .

225 Association mapping

226 For each of the seven agronomic traits, we performed a genome-wide association study
 227 (GWAS) with GCA effects of all 142 Dent and 103 Flint parent lines as response
 228 variables using the EMMAX-method (Kang et al., 2010) as implemented in *cpgen*
 229 (Heuer, 2015). To avoid using the marker data twice, GCA effects were calculated
 230 using only pedigree information. Furthermore, an eQTL analysis was carried out
 231 to examine statistically significant associations between genomic and transcriptomic
 232 data for the parent lines (D and F) of the core set H plus five additional lines. This
 233 was accomplished in the same way as in the GWAS for agronomic traits, but here the
 234 BLUPs of the transcriptomic data of each mRNA were used as the response variables.
 235 Associations in each GWAS were declared statistically significant at $\alpha = 0.05$ after
 236 Bonferroni correction.

237 Probability of success

238 Following Robson, Powers and Urquhart (1967), we calculated the probability of suc-
 239 cess ($P[r, \beta]$) that a hybrid, selected at random from the upper β percent fraction of
 240 the distribution of predicted values for predictor A, has a phenotypic value contained
 241 in the upper β percent of the distribution of observed values. Denoting the predictive
 242 ability of a given predictor by r , this conditional probability was calculated assuming
 243 a bivariate normal distribution

$$\begin{bmatrix} \hat{y} \\ y \end{bmatrix} \sim N \left(\begin{bmatrix} 0 \\ 0 \end{bmatrix}, \begin{bmatrix} 1 & r \\ r & 1 \end{bmatrix} \right). \quad (3)$$

244 The required integrals were solved within the *R* statistical environment using the
 245 *mvtnorm* package (Genz et al., 2017).

246 Principal component analysis

247 Principal component analyses (PCA) were carried out to examine whether differ-
 248 ent predictors can distinguish between Dent and Flint parent lines and to explore
 249 whether subpopulations exist within either hetrotic group. Prior to each PCA, all
 250 variables were scaled and centered. Clusters represent two component mixtures of
 251 bivariate t-distributions, which were estimated using Maximum Likelihood. Ellipses
 252 were drawn based on the 0.95 quantiles of the respective bivariate t-distributions.
 253 Unless stated otherwise, all statistical analyses were carried out inside the *R* envi-

254 ronment for statistical computing (R Core Team, 2016).

255 **Data availability**

256 The data and the code used to analyze the data are available upon request.

Results

Agronomic data

Mean values of the 1,536 hybrids for the seven evaluated agronomic traits, relevant for animal feed and biogas production, were of the same magnitude as reported by Riedelsheimer et al. (2012) and Grieder et al. (2011). For all traits, $\sigma_{GCA^D}^2$ and $\sigma_{GCA^F}^2$, describing the main effects of the parents from each heterotic group, together explained more than 93% of the genotypic variance among hybrids (Table 1). Heritabilities were moderate to high for all agronomic traits, indicating a high precision of field experiments and data collection (Table 1).

[Table 1 about here.]

Predictor data

Repeatabilities (w^2) for endophenotypes varied considerably in both groups of parents (Fig. S1a) with average values ranging from 0.31 to 0.41, except for transcriptomic data in Flint material where the average repeatability was only 0.18. Nevertheless, in the latter case, 291 out of 1,323 transcripts still exceeded a threshold of $w^2 = 0.4$.

Dent and Flint lines were clearly separated in principal component analyses of genomic and transcriptomic data (Fig. 1a) without signs of subpopulations within either group. However, they overlapped for leaf metabolic and, to an even greater extent, for root metabolic data. Off-diagonal elements of the kernels $\mathbf{G}_D$ and $\mathbf{G}_F$, respectively, showed moderate correlations between genomic and transcriptomic data

($\rho_D \approx 0.56$, $\rho_F \approx 0.44$, Fig. S2). Correlations between the off-diagonal elements of the **G**-matrices were highest for the comparison between genomic and pedigree data ($\rho_D \approx 0.72$, $\rho_F \approx 0.63$). Intriguingly, the associations between the **G**-matrices for the root and leaf metabolic data were very low ($\rho_D \approx 0.12$, $\rho_F \approx 0.06$).

[Figure 1 about here.]

We observed high median pairwise linkage disequilibrium (LD) between SNP markers ($r^2 \approx 0.39$ in Dent and $r^2 \approx 0.37$ in Flint material) at a distance of $\Delta\text{Mbp} \leq 0.125$ (Fig. 1b). After an initial drop in r^2 for $\Delta > 0.125$, substantial long-range LD remained. Large differences in allele frequencies in the two heterotic groups were present for 57% of SNPs (Fig. 2a,b) — particularly in the telomeric regions of the genome. An eQTL analysis performed with the parent lines suggests that transcript abundance integrates variegated genetic information given the fact that i) on the same chromosome, significant associations not only occurred between adjacent but also between distant pairs of expressed genes and SNPs and ii) 50% of the significant associations ($\alpha = 0.05$, Bonferroni-corrected) occurred between expressed genes and SNPs on different chromosomes (Fig. 2).

[Figure 2 about here.]

Predictive abilities

Assuming a polygenic architecture for all traits, as suggested by results from a GWAS (Fig. S3), we chose the best linear unbiased predictor (BLUP) method as a baseline for prediction of T0, T1 and T2 hybrids. Given that we corrected for fixed effects

in advance, this method corresponds to a selection index. A cross-validation scheme with 1,000 runs (CV1000), stratified by the parent lines, was devised (File S1, Fig. S4). Our main emphasis was on predicting T0 hybrids given the fact that they constitute the majority of possible hybrids in practical breeding programs (Kadam et al., 2016).

For predictive abilities (r) of T0 hybrids, transcriptomic data alone were the best predictor for the most complex and highly heterotic trait, DMY, as well as for PRO (Fig. 3a). With transcriptomic data, the predictive ability r for DMY was 14.9% higher than for genomic data, resulting in an 85% increase in the probability of successfully selecting the best hybrid candidates $P[r, \beta]$ for $\beta = 0.01\%$ (Fig. 3b). This selection intensity corresponds to picking the top 100 out of 10^6 predicted hybrids for production and intensive testing in field trials.

[Figure 3 about here.]

Compared to other individual predictors, r obtained with genomic data alone were higher for FAT and SUG. Root metabolites displayed moderate to high predictive abilities for DMY and FAT, but did not perform well otherwise. Leaf metabolites performed relatively poorly for all traits. Regardless of the trait, combinations of genomic and transcriptomic information displayed robust and consistently high predictive abilities. Except for PRO, incorporating additional endophenotypes as predictors into our models did not yield notable improvements but remained at the same level compared to combining genomic and transcriptomic data. Incorporating SCA effects into our models did not further improve predictive abilities (Fig. S5). Results for the combination of other predictors with metabolic data are not presented

322 because no improvement of predictive abilities over the combination of genomic data
 323 with transcriptomic data and pedigree data could be achieved. Finally, we assessed
 324 the influence of the number of SNPs and mRNAs on predictive abilities. For genomic
 325 data, a subset of 5,000 SNPs already yielded the same predictive ability as when us-
 326 ing the entire available set. For transcriptomic data, the predictive ability improved
 327 only marginally with subsets larger than 50% of the available transcripts (Fig. S6).

328 Discussion

329 A paradigm shift in hybrid breeding

330 Hybrid breeding programs are generally based on genetically divergent heterotic
 331 groups. Their use enables a better exploitation of heterosis when conducting crosses
 332 between them (Melchinger and Gumber, 1998) and is expected to reduce the ratio of
 333 specific to general combining ability variance ($\sigma_{SCA}^2 : \sigma_{GCA}^2$) in the crosses, thereby
 334 allowing for the selection of hybrids largely on the basis of GCA of their parent lines
 335 (Reif et al., 2007). However, obtaining accurate estimates of GCA requires the eval-
 336 uation of new lines in combinations with testers from the opposite heterotic group
 337 in multi-environment field trials. The promise of hybrid prediction is to accelerate
 338 breeding programs by skipping a large share of these tests in favor of selecting the
 339 most promising hybrids before they are even produced (Technow et al., 2014). This
 340 approach involves the prediction of an impressive number of putative hybrid candi-
 341 dates (n^2) using predictor data collected on only $2n$ parent lines. Crucial for hybrid
 342 prediction are predictors, which not only reflect the relationship between parental
 343 inbred lines but also the interaction of the two parental genomes in their hybrid
 344 progeny.

345 **Heterotic groups** Because of genetic drift and selection for hybrid performance,
 346 allele frequencies are expected to diverge in the two heterotic groups, thereby en-
 347 larging their genetic distance (Falconer and Mackay, 1996, Reif et al., 2007, Laripe
 348 et al., 2017). Consistent with this hypothesis and two pilot studies with U.S. maize

lines (Gerke et al., 2015, Hall et al., 2016), Dent and Flint lines in our study were clearly separated in principal component analyses of genomic and transcriptomic data. With large differences in allele frequencies p^I and p^{II} in the two heterotic groups, as observed for 57% of SNPs, dominance variance σ_D^2 becomes very small because it is a function of the product $p^I(1-p^I)p^{II}(1-p^{II})$ (Stuber and Cockerham (1966), File S1).

Dominance variance (σ_D^2) is the main component contributing to the variance of the specific combining ability effects (σ_{SCA}^2), describing all types of interactions among the parental genomes in hybrid combinations. It was therefore not surprising that the variances of the general combining ability (GCA) effects ($\sigma_{GCA^D}^2$ and $\sigma_{GCA^F}^2$), describing the main effects of the parents from each heterotic group, together explained more than 93% of the genotypic variance among hybrids for agronomic traits, which is consistent with earlier studies on silage maize of the Dent $\times$ Flint heterotic pattern (Geiger, Melchinger and Schmidt, 1986, Argillier, Méchin and Barrière, 2000). While the magnitude of SCA effects was trait specific, it was low for all observed traits, which is in agreement with previously reported values for yield and quality traits in silage maize (Grieder et al., 2012). The importance of GCA in our material was further corroborated by merely marginal differences in predictive abilities between models using only GCA effects and those that additionally incorporated SCA effects (Fig. S5). Nevertheless, in crops such as wheat, with yet no clearly defined heterotic groups (Zhao et al., 2015) and greater importance of SCA, inclusion of SCA effects in the model should improve predictive abilities.

Properties of well-established predictors While pedigree data reflect the expected relationship between genotypes, they do not necessarily depict their realized relationship. Genomic data and downstream endophenotypes offer to improve upon this pedigree-based approximation by more closely mirroring the transmission of genes between genotypes and their interactions. Genomic data have the advantage of reliably capturing Mendelian sampling, thereby improving pedigree-based prediction for many traits. However, genomic data alone may not be the final answer for the prediction of complex traits for two major reasons: First, the number of samples in most studies is considerably smaller than the number of genetic markers or even nucleotides of a genome. This implies that just modeling additive effects already necessitates shrinkage of effects. More importantly, however, interactions between loci throughout the genome can be frequent (Brem et al., 2005, Brown et al., 2014), but attempts to incorporate this epistasis for the prediction of heterotic traits using genomic data have been disappointing when the prediction and training set did not share the same or closely related parents (Jiang and Reif, 2015). This was true even when using recently developed, efficient models (Jarquín et al., 2014, Martini et al., 2016) and suggests that genomic data capture only statistical epistasis, referring to genetic variation at the population level (Sackton and Hartl, 2016), which is generally of negligible magnitude (Hill, Goddard and Visscher, 2008, Mackay, 2014, Guo et al., 2016, Vazquez et al., 2016).

391 **Complementation of predictors**

392 **Flow of biological information** It is well-known that genetic effects on the phe-
 393 notype are mediated through multiple layers of endophenotypes (Civelek and Lusis,
 394 2014, Ritchie, Holzinger, Li, Pendergrass and Kim, 2015) with information mainly
 395 flowing from the genome toward the phenotype via the transcriptome, the proteome
 396 and the metabolome with metabolite fluxes ultimately governing energy production
 397 and growth (Fiévet, Dillmann and de Vienne, 2010). For most traits in our material,
 398 metabolite- and pedigree-based predictive abilities were lower than those obtained
 399 with either transcriptomic or genomic information. However, consistently high pre-
 400 dictive abilities across multiple traits could be realized when combining multiple
 401 predictors, as has been reported previously in humans (Vazquez et al., 2016) and
 402 maize inbred lines (Guo et al., 2016). This suggests complementary properties of the
 403 different predictors resulting in better proxies for the complex interplay in gene net-
 404 works than genomic information alone. Such an advantage is particularly important
 405 for hybrid prediction when parents of prediction set hybrids are not closely related
 406 to parents of training set hybrids (File S1) as was shown by the relative excellence
 407 of transcriptomic data and the use of multiple predictors for the prediction of traits
 408 in T0 hybrids compared to T1 and T2 hybrids.

409 **Tapping new sources of information** Whereas pedigree and genomic informa-
 410 tion are static, subsequent endophenotypes are characterized by pervasive interac-
 411 tions among and between each other (Dalchau et al., 2011, Zhu et al., 2012) and are,
 412 to varying degrees, influenced by biotic (Rudd et al., 2015, Tzin et al., 2015) and abi-

otic perturbations (Caldana et al., 2011, Witt et al., 2012). So while endophenotypes do not exclusively report on physiological epistasis but also on non-heritable effects, they seem to capture important information not represented by the genome given their intermediate position in the genotype-phenotype cascade. We get support for this hypothesis from (i) merely low to moderate correlations between off-diagonal elements of the kernels of different predictors in our study, (ii) mounting evidence for further improvements of predictive abilities when complementing genomic prediction with other endophenotypes despite sufficient marker densities (Fig. S6, Guo et al. (2016)) and (iii) the integration of SNP information from close and distant eQTL in the transcripts analyzed in our study. However, we concede that the number of parental genotypes in our mRNA assays was too small to warrant a reasonable statistical power for detecting epistasis in the expression of transcripts. In breeding programs, predictive abilities are largely driven by relationships — including Mendelian sampling — among genotypes compared to LD between SNP markers and causal QTL (Schopp et al., 2017). Increasing marker densities therefore have limited utility for improving genomic predictions as observed in our material, where SNP-based predictive abilities reached a plateau after using 5,000 equally spaced SNPs (Fig. S6). While two other studies also attempted to model interactions between different predictors, we refrained from this approach given that their reported predictive abilities based on interactions were not different from those in additive models despite using much larger sample sizes (Vazquez et al., 2016, Guo et al., 2016).

435 Transcriptomic data

436 **Utility of transcriptomic data for trait predictions** Of particular note was
 437 the excellent performance of transcriptomic data in predicting dry matter yield and
 438 protein. Evidence that parental gene expression patterns might be predictive of
 439 hybrid performance is given by (i) prevailing additive expression patterns in maize
 440 hybrids (Springer and Stupar, 2007*a*, Stupar et al., 2008), (ii) a positive correlation of
 441 the proportion of additive gene expression with the yield of hybrids (Guo et al., 2006),
 442 and (iii) co-localization of additively expressed genes with heterotic QTL (Thiemann
 443 et al., 2014). According to metabolic flux theory, gene expression in hybrids at the
 444 mid-parent level can generate hybrid vigor by counterbalancing opposing detrimental
 445 expression levels in their parent lines on a genome-wide scale (Kacser and Burns,
 446 1981, Springer and Stupar, 2007*b*). The same concept is expected to apply to other
 447 quantitative endophenotypes (Lisec et al., 2011).

448 **Pre-selection bias** As pointed out earlier, our transcripts were pre-selected based
 449 on associations with grain dry matter yield and grain dry matter content in hybrids,
 450 using a subset of the data included in our study ($H_{\text{FAC}(1)}$). Hence, genotypes used
 451 for the pre-selection (*i.e.* $H_{\text{FAC}(1)}$) could be regarded as a training set. By combining
 452 this "training set" and genotypes from the remaining five factorials, we might have
 453 introduced a bias by using predictors that have already seen the response variable
 454 in the $H_{\text{FAC}(1)}$ genotypes. To rule out the existence of such a bias, we have com-
 455 pared the predictive abilities of different predictors for the complement of $H_{\text{FAC}(1)}$.
 456 Two findings indicate that no bias in the comparison of predictive abilities was in-

457 introduced: (i) Relative differences in predictive abilities between transcriptomic and
 458 pedigree or genomic data did not change when excluding genotypes from $H_{\text{FAC}(1)}$
 459 from the data (Fig. S7) and (ii) transcriptomic data performed rather poorly in
 460 predicting dry matter content although this trait was also among the criteria for the
 461 pre-selection procedure. Finally, an independent study using RNA-Seq data for the
 462 prediction of traits in maize inbred lines also reported exceptionally good perfor-
 463 mance of transcriptomic data in the prediction of multiple yield-related traits (Guo
 464 et al., 2016).

465 **Relative excellence of predictors for different traits**

466 **Tissue and sampling time** Despite the great prospects of using endophenotypes
 467 for trait predictions, some aspects require careful consideration when using this ap-
 468 proach. A particular challenge in endophenotype-based prediction efforts is the choice
 469 of a suitable tissue and sampling time. Tissue-related effects regarding gene expres-
 470 sion were found in studies on humans (Yang et al., 2015, Mele et al., 2015, Searle
 471 et al., 2016) and *A. thaliana* (Schmid et al., 2005) and in maize hybrids with respect
 472 to metabolome composition and metabolite abundance (Witt et al., 2012). Moreover,
 473 the age of an organism can selectively influence the expression of genes as observed in
 474 studies on humans (Mele et al., 2015, Yang et al., 2015) and *C. elegans* (Vinuela et al.,
 475 2010, Francesconi and Lehner, 2014). The low correlations between the off-diagonal
 476 elements of the kernels calculated from root and leaf metabolites might therefore
 477 be a reflection of highly dynamic processes differing between tissues and during
 478 different developmental stages. Whereas root metabolic data and transcriptomic

479 data were obtained from seedlings germinated in standard controlled conditions, leaf
480 metabolic data were derived from field-grown plants at a much later developmental
481 stage, thereby increasing the possibility of environmentally-induced modifications.
482 One might hypothesize, that the choice of sampling time and tissue could influence
483 the chances of successful trait prediction if such age- or tissue-dependent transcripts
484 and metabolites are associated with a phenotypic or clinical trait.

485 **Feature selection** Another explanation for trait-dependent excellence of any pre-
486 dictor might lie in the sampling of features. In this study, only a small subset of
487 metabolites was sampled and even very recent technologies (Xu et al., 2016, Dan
488 et al., 2016) capture only a fraction of the estimated set of metabolites (Fernie,
489 2007). Moreover, the smaller differences in metabolite levels between both heterotic
490 groups (Fig. 1) were most likely not conducive to capturing basic components un-
491 derlying complex heterotic traits. It is also possible that transcriptomic data are
492 associated with more biological processes than metabolite data and better capture
493 the genetic effects relevant for the prediction of T0 hybrids.

494 **Prospects for metabolites** Previously observed moderate metabolite-based pre-
495 dictive abilities for T1 hybrids (Riedelsheimer et al., 2012) were confirmed in our
496 study (Fig. S8), but for the majority of traits, root metabolites reached only medium
497 and leaf metabolites even lower predictive abilities when predicting T0 hybrids. De-
498 spite the aforementioned shortcomings of metabolites, they have shown to be intrigu-
499 ing predictors due to their physiological proximity to the phenotype, which provides
500 information that is impossible to infer from DNA or proteins (Fernie and Stitt, 2012),

as well as encouraging results from other studies (Guo et al., 2016, Dan et al., 2016).
A recently introduced technology, allowing for live-measurements of small molecules
in the blood of living and awake animals (Arroyo-Currás et al., 2017), might overcome
the problem of poorly time-resolved snap-shots of some metabolites with extremely
fast turnover rates (Arrivault et al., 2009) if modified to properly work in plants.

Predictor requirements Besides improving upon predictions based on pedigree
relationships by capturing Mendelian sampling, the widespread use of genomic in-
formation in trait prediction has been driven by the ease of its application. In
order to compete with genomic data, other 'omics' data therefore require the use
of standardized sampling conditions to obtain large repeatabilities and the possibil-
ity of season-independent sample extraction from seeds, seedlings or young roots to
achieve high throughput.

Conclusions

The use of whole-genome information has considerably advanced trait prediction over
traditional pedigree-based BLUP by incorporating previously unobservable Mendelian
sampling. Combining variegated sources of information promises to capture complex
interactions between genes and endophenotypes, leading to stable predictions across
traits. Especially if an extremely small fraction of the candidates is selected from the
millions of possible new hybrids from each breeding cycle, the success of forecasts
is a strongly convex function of predictive ability (Fig. 3b). Therefore, consider-
ing endophenotypes could have a substantial effect on the success and economics

522 of hybrid breeding. Given the anticipated technological improvements in RNA-Seq
523 and metabolite profiling, as well as the forthcoming adoption of the DH-technology
524 for many crops (Kelliher et al., 2017), a paradigm shift from exclusively genomic
525 prediction models to more inclusive approaches seems imminent.

Author Contributions

W.S. and A.E.M developed the lines and hybrids, W.S. and A.E.M designed the field experiments, T.A.S analyzed the agronomic and pedigree data, L.W., M.S., A.S., A.E.M and M.W. designed the metabolic experiments, A.S. conducted the metabolic experiments, S.S. designed the transcriptomic experiments, F.S. and A.T. conducted the transcriptomic experiments, M.W. analyzed the metabolic and transcriptomic data, M.W., T.A.S, G.T., C.H. and A.E.M. devised the prediction models, M.W., C.H. and T.A.S. implemented the prediction models and developed software, H.F.U. and A.E. contributed to the statistical analysis. M.W., A.E.M., S.S, G.T., Z.N. and C.C.S. wrote the manuscript.

Acknowledgements

The authors acknowledge support by the state of Baden-Württemberg through bwHPC. This project was funded by the German Federal Ministry of Education and Research (BMBF) within the projects OPTIMAL (FKZ: 0315958B,0315958F), SYNBREED (FKZ: 0315528D) and by the German Research Foundation (DFG, Grants No. ME 2260/5-1 and SCHO 764/6-1). Financial support for M.W. was provided by the Fiat

Panis foundation, Ulm, Germany. We thank the staff of the Agricultural Experimental Research station, University of Hohenheim, for excellent technical assistance in conducting the field experiments. We are indebted to the group of R. Fries from Technische Universität München for the SNP genotyping of the parent inbred lines, to X. Mi for his assistance in preparing auxiliary figures based on the Mathematica software, to C. Zenke for advice on the computation of transcriptomic BLUEs and to P. Schopp for advice on prediction models.

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Figure 1: Properties of predictor data for Dent (red) and Flint (teal) parent lines. (a) Principal component analysis (PCA) of both groups for genomic (G), transcriptomic (T), leaf metabolic (L) and root metabolic (R) data. The variance explained by PC 1 (x axis) and PC 2 (y axis) are shown in the caption of each facet. (b) Linkage disequilibrium decay as a function of the distance between two loci using 40 bins of 0.125 Mbp width, each. The median r^2 is depicted as a horizontal bar whereas the mean r^2 is depicted as a white diamond.

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1	Summary of agronomic traits. Traits are characterized by overall mean (μ), variance components of GCA effects for Dent ($\sigma_{GCA^D}^2$) and Flint lines ($\sigma_{GCA^F}^2$) and SCA effects (σ_{SCA}^2) (followed by s.e.m.) as well as entry mean heritabilities (H^2).	51
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Table 1: Summary of agronomic traits. Traits are characterized by overall mean (μ), variance components of GCA effects for Dent ($\sigma^2_{\text{GCA}^D}$) and Flint lines ($\sigma^2_{\text{GCA}^F}$) and SCA effects (σ^2_{SCA}) (followed by s.e.m.) as well as entry mean heritabilities (H^2).

Trait	μ	$\sigma^2_{\text{GCA}^D}$	$\sigma^2_{\text{GCA}^F}$	σ^2_{SCA}	H^2
DMY (t/ha)	19.00	1.51 ± 0.22	1.00 ± 0.18	0.17 ± 0.03	0.82
DMC (%)	34.13	4.17 ± 0.59	5.03 ± 0.79	0.49 ± 0.07	0.91
ADF (%)	20.93	0.27 ± 0.06	0.40 ± 0.09	0.02 ± 0.02	0.43
FAT (‰)	30.02	1.01 ± 0.19	2.10 ± 0.37	0.17 ± 0.05	0.73
PRO (‰)	69.65	3.11 ± 0.53	2.77 ± 0.51	0.29 ± 0.10	0.70
STA (%)	35.56	2.95 ± 0.51	3.33 ± 0.63	0.24 ± 0.10	0.69
SUG (‰)	38.12	31.33 ± 5.15	31.90 ± 5.73	4.12 ± 1.02	0.77