Predictors of intestinal inflammation in asymptomatic first-degree relatives of patients with Crohn's disease

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Abstract

Objective: Relatives of individuals with Crohn's disease (CD) carry an increased number of CD-associated genetic variants and are at increased risk of developing the disease. Multiple environmental and genetic factors contribute to this increased risk. We aimed to estimate the utility of genotype, smoking, family history, and a panel of biomarkers to predict risk in asymptomatic first-degree relatives (FDRs) of CD patients.

Design: We calculated a combined genotype (72 CD-associated genetic markers) and smoking relative risk score in 454 FDRs, and performed capsule endoscopy and collected 22 biomarkers in individuals from the highest and lowest risk quartiles. We then predicted small intestinal inflammation using genetic risk score, smoking status, number of relatives with CD, capsule transit time, and the panel of biomarkers in 124 individuals with complete data. Our principal analysis was to calculate the predictive utility from two machine learning classifiers: an elastic net and a random forest.

Results: Both classifiers successfully predicted FDRs with intestinal inflammation: elastic net (AUC=0.80, 95% CI: 0.62-0.98), random forest (AUC=0.87, 95% CI: 0.75-1.00). The elastic net selected a 3-predictor solution: CD family history (OR=1.31), genetic risk score (OR=1.14), and faecal calprotectin (OR=1.04). The same 3 variables were among the top 5 most important predictors as ranked by the random forest.

Conclusion: A readily collectable panel of genetic risk variants, added to family history and faecal calprotectin, predicts those at greatest risk for developing CD with a good degree of accuracy.

Introduction

Inflammatory bowel disease (IBD), comprising Crohn's disease (CD) and ulcerative colitis (UC), is a chronic inflammatory condition of the gastrointestinal tract associated with significant morbidity. Family members of individuals with CD are at increased risk for developing the disease. Estimates of the sibling relative risk (the ratio of disease risk in siblings compared with the rate in the general population) range from 15 to 42[1]. A retrospective study of the entire Danish population over a period of more than 40 years found that about 12% of new incidents of CD occurred in families already affected by the disease[2]. Disease onset precedes the development of symptoms, which often precede diagnosis by months to years, at which stage many patients have developed complications of the disease (including malnutrition, osteoporosis, and strictures or fistulas requiring surgery)[3]. There is emerging evidence to support the use of immunosuppressive treatment early in the course of CD to reduce the risk of such complications[4]. With an increasing incidence and prevalence of IBD worldwide[5], a clinical tool that facilitates early detection of those at greatest risk opens up the possibility for early intervention to alter or halt aberrant immune and inflammatory responses before the development of overt disease, and perhaps, ultimately, disease prevention.

Although CD pathogenesis remains incompletely understood, genetic risk for the disease is well-established. Heritability, the proportion of population trait variance explained by genetic factors, estimated from identical and non-identical twin concordance rates, is 0.75[6]. Genetic variance explained by CD-associated loci

discovered through genome-wide association studies (GWAS), so-called SNP heritability, is 0.37, indicating that approximately half the total CD heritability is explained by known SNPs[6]. To date, there are 240 single nucleotide polymorphisms (SNPs) robustly associated with the IBD[7]. A recent study of first-degree relatives (FDRs) of patients with IBD showed that they are enriched for IBD-associated risk loci[8].

Familial clustering of disease is likely the result of shared environmental factors, as well as genetic risk factors. The observed increasing incidence of the disease among populations with historically lower rates, and those migrating from regions with low incidence rates to regions with higher rates, is consistent with a substantial environmental risk component to CD[5]. Many lifestyle-related factors have been implicated, including stress, sedentary lifestyle, western diet, poor sleep, and tobacco use[9] [10]. Smoking is the best-studied environmental risk factor; a meta-analysis estimated a two-fold risk increase for CD in current smokers[11].

Asymptomatic FDRs may display phenotypic features in common with CD patients including altered intestinal permeability, positive serological antimicrobial markers, disordered innate and acquired immunity, faecal dysbiosis, and elevated faecal calprotectin (FC)[12]. Overt small intestinal (SI) inflammation has been described in FDRs who have undergone ileocolonoscopy[13], intestinal ultrasound[14] and video capsule endoscopy (VCE)[15].

Evidence of increased risk factors for CD in FDRs raises the possibility of predicting those at risk of developing the disease[12]. We hypothesized that the observed increase in the number of CD-associated risk loci in FDRs relative to healthy controls[8], elevated levels of FC[12], and smoking status, taken together could provide sufficient information to detect at-risk individuals before the development of overt symptoms. In this study we aimed to assess the clinical utility of a disease risk model to predict the presence of SI inflammation as detected by VCE. We applied two machine learning methods to genetic risk (72 genetic markers), smoking status, number of relatives with CD, 22 biomarkers, and assessed the predictive ability of the derived models.

Material and Methods

Participants

We recruited 480 healthy FDRs (full siblings, offspring or parents) between the age of 18 and 55, through patients with CD attending the IBD service at Guy's & St Thomas' NHS Foundation Trust (GSTT) and members of Crohn's & Colitis UK (a charity supporting patients with IBD). CD diagnosis in probands was confirmed by their gastroenterologist or general practitioner. Interested FDRs provided information regarding family history of IBD, medical history (including gastrointestinal symptoms), smoking status, medication use, allergies, primary care provider and ethnicity. In keeping with the population of the meta-analysis of CD GWAS used for calculating the genotype relative risk[16], only FDRs of European ancestry were included. The inclusion criteria were: ability to give written informed

consent, age 18-55 years, confirmed history of Crohn's disease in the proband, and absence of gastrointestinal symptoms in the FDR. The exclusion criteria were: a previous diagnosis of IBD, a previous diagnosis of irritable bowel syndrome, or presence of a major co-morbidity. Additional exclusion criteria for capsule endoscopy were: pregnancy, major bowel surgery, or the use of non-steroidal anti-inflammatory drugs in the 4 weeks prior to capsule endoscopy (low-dose aspirin excluded).

DNA collection and genotyping

480 FDRs met the inclusion criteria and were sent a saliva sampling kit (Oragene™ DNA OG-500) and 455 returned saliva samples. DNA extraction and genotyping protocols are included in the supplementary materials. DNA samples were genotyped on the Immunochip, which is a custom Illumina Infinium array containing 196,524 SNPs and small insertion/deletions selected mainly from GWAS analysis of 12 immune-mediated diseases[17]. It included all SNPs from CD-associated loci at the time of design. The three major *NOD2* variants were analysed separately (rs2066844, rs2066845 and rs2066847), instead of the *NOD2* tagging SNP, rs2076756. Owing to the enhanced risk of homozygotes and compound heterozygotes for *NOD2* variants, the *NOD2* variants were combined, rather than treated as independent loci. Genotypes for a total of 72 SNPs from known CD risk loci were extracted for genetic risk profiling.

The relevant SNP data were extracted and underwent quality control (QC) in PLINK[18]. SNPs with missingness >3% and Hardy-Weinberg equilibrium outliers ($p \leq 1 \times 10^{-4}$) were excluded. X chromosome heterozygosity rates were used to determine sex

empirically. Population structure was assessed by principal components analysis in PLINK as previously described[19], derived from CEU HapMap 3 individuals (Northern and Western European Ancestry). No outliers were identified.

Following QC, 11 saliva samples were repeated due to a low DNA concentration or quality, one of which repeatedly failed QC and was excluded. 454 FDRs were successfully genotyped (61% female; median age 34, range 18-55; 40% were siblings, 46% offspring, 14% parents of probands; 44% were current or ex-smokers).

Calculation of high and low risk groups

We generated a combined smoking and genotype risk score in 454 FDRs of individuals with CD using the R package REGENT[22]. Summary results were obtained for 72 genetic variants from GWAS meta-analysis in CD[16]. Smoking risk was calculated using ORs from a large case-control study[23]. 147 FDRs in the highest and lowest risk quartiles of the risk score were invited to undergo VCE, and to provide stool and blood samples for biomarker analysis.

Subsequent statistical analyses were performed on the subset of 124 participants for whom we had complete data on all measures (age in years (mean=37.42, sd=10.46), sex (male=42, female=82), and genetic risk score and smoking were considered in the modelling as separate variables.

Smoking status was recorded as "Current", "Ex", or "Never" smoked (29, 33, 62 respectively among 124 individuals with complete data). We used "Never" smoked

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as the reference level. The number of family members with CD (CD family history) was coded as "single" or "multiple" (111, 13 respectively among 124 individuals with complete data).

Video capsule endoscopy (VCE)

Capsule endoscopies were performed in the Endoscopy Unit at GSTT following written informed consent. The MiroCam™ (Intromedic, Seoul, Korea) VCE system was used for all capsule endoscopies (see supplementary materials for full description and protocol). Two validated scoring systems were used to quantify the degree of inflammation within the small intestine: the Lewis Score[20] and the Capsule Endoscopy Crohn's Disease Activity Index (CECDAI)[21]. SI transit time was defined as the passage from the first duodenal to the first caecal image, in minutes.

Biomarkers

Full biomarker description and protocol are included in the supplementary materials.Stool was analysed for faecal calprotectin and serum for high-sensitivity C-reactive protein (hs-CRP), anti-saccharomyces cerevisiae antibodies (ASCA), and cytokines and growth factors. **Table 1** shows the descriptive statistics for the complete list of 22 biomarkers measured. We dropped from the statistical analyses, 6/22 biomarkers with near zero variance among the 124 individuals with complete data.

Table 1. Biomarker descriptive statistics

Biomarker	Description	Mean (SD) / Count	Median	Range

VCAM	Vascular cell adhesion molecule (µg/g)	573.44 (140.77)	559.69	38.49–1034.17
ICAM	Intercellular adhesion molecule (µg/L)	239.26 (53.04)	232.26	33.54–390.27
ESEL	E-selectin (µg/L)	14.35 (4.97)	13.07	5.29–26.26
PSEL	P-selectin (µg/L)	277.70 (135.40)	231.54	134.21–729.24
LSEL	L-selectin (µg/L)	1872.42 (306.43)	1851.77	1242.69–2809.83
*ASCA IgG	Saccharomyces cerevisiae antibodies Immunoglobulin G	Negative: 112, Equivocal: 4, Positive: 8		
*ASCA IgA	Saccharomyces cerevisiae antibodies Immunoglobulin A	Negative: 113, Equivocal: 1, Positive: 10		
IL2	Interleukin 2 (ng/L)	1.17 (0.81)	0.90	0.90–6.07
*IL4	Interleukin 4 (ng/L)	2.13 (0.10)	2.12	2.12–3.18
IL6	Interleukin 6 (ng/L)	0.96 (1.28)	0.58	0.18–8.59
IL8	Interleukin 8 (ng/L)	63.87 (130.17)	4.64	0.80–415.00
*IL10	Interleukin 10 (ng/L)	0.49 (0.41)	0.37	0.37–3.75
*IL1a	Interleukin 1 alpha (ng/L)	0.26 (0.22)	0.19	0.19–2.18
IL1b	Interleukin 1 beta (ng/L)	1.18 (2.84)	0.67	0.26–31.09
IL1RA	Interleukin-1 receptor antagonist (ng/L)	321.47 (335.95)	194.64	24.30–1657.37
FC	Faecal calprotectin (µg/g)	68.58 (112.15)	36.25	10.00–866.50
VEGF	Vascular endothelial growth factor (pg/ml)	60.19 (46.44)	47.50	8.57–283.08
EGF	Endothelial growth factor (ng/L)	56.69 (39.01)	49.52	2.72–228.15
*INFy	Interferon gamma (ng/L)	0.45 (0.09)	0.44	0.44–1.20
TNFa	Tumor necrosis factor alpha (ng/L)	0.78 (0.25)	0.64	0.59–1.48
MCP1	Monocyte chemoattractant protein 1 (ng/L)	92.92 (95.69)	72.11	4.21–646.00
hs-CRP	Highly sensitive C-	2.53 (3.41)	1.30	0.10–18.30

	reactive protein (mg/L)			

Note. Biomarker descriptive statistics for 124 individuals with complete data. * = variables with near-zero variance not included in statistical modelling

Statistical analysis

All statistical analyses were performed in the R programming language and environment[24]. For the predictive modelling we used the R package caret[25], and its dependencies glmnet[26] and ranger[27].

Explanatory modelling

We first fitted multivariate logistic regression models with dichotomized Lewis score (<135 = Normal i.e. without intestinal inflammation, ≥135 = Abnormal, i.e. with intestinal inflammation) as response variable, and age, sex, genetic risk score, smoking status, CD family history, VCE capsule transit time and the biomarkers as explanatory variables. Applying generalised linear regression models to a dataset with a large number of predictor variables relative to the number of samples is not optimal for several reasons including over-fitting (explaining noise as well as signal in the sample), multi-collinearity (correlation and potential redundancy among predictors), and low interpretability[28]. We used stepwise model selection to address this overfitting, multicollinearity, and model complexity. At each step in the iterative selection procedure, a variable is considered for addition to (or subtraction from) the current set of explanatory variables based on a model comparison criterion. We compared two criteria: Akaike's information criterion (AIC) and the Bayesian

information criterion (BIC). However, this approach does not eliminate the problems of overfitting, multicollinearity, and model complexity, especially with a small sample and large number of measured variables. In addition, explanatory models can quantify the relationship between predictors and outcome in the particular sample collected (e.g. regression coefficients, and total variance explained), but they give no measure of the predictive performance of the putative predictors in unseen cases. Given that the ultimate aim of the study was to estimate the utility of the genetic, environmental and biomarker variables to identify high risk individuals, we went on to derive a series of predictive models. Machine learning is specifically designed to deal with the large number of variables relative to sample size problem, and includes techniques to address the low events-per-variable ratio (small number of cases).

Predictive modelling

Our goal was to develop a predictive model for SI inflammation, achieving a balance between interpretability (by reducing the number of predictor variables) and predictive ability. Thus we used machine learning, which has techniques for variable subset selection and estimation of how well a given model will perform at predicting future data[29]. Machine learning finds structure in data and addresses over-fitting when there are a large number of predictors. More generally, it is a set of techniques that improve performance on a specified task (e.g. classifying absence/presence of inflammation) with experience (exposure to data). We divided the data into a *training* sample (2/3) for model building and a *test* sample (1/3) for evaluation of model predictive performance and compared two machine learning techniques:

Elastic net. The elastic net is an extension of the basic regression framework that allows selection of the most important subset of predictors. It mixes a ridge penalty (which shrinks the coefficients of correlated predictors towards each other) and a LASSO penalty (which selects one among a group of correlated predictors and shrinks the coefficients of the others to zero) to perform variable selection[28]. We used 20 repeats of 5-fold cross validation to estimate model parameters (penalty mixing factor (α) and penalty strength (λ)) that optimised the model's prediction performance, i.e., its ability to correctly classify individuals with and without inflammation. We measured prediction performance with the area under that receiver-operator characteristic curve (AUC: a plot of the true positive rate against the false positive rate for different cut-offs of the model estimated probability of being Abnormal) and accuracy (the fraction of test sample predictions that are true). The absolute value of the t-statistic for each parameter in the model is used to judge relative variable importance.

Random forest. A random forest is a collection of classification or regression trees. A tree is a series of splitting rules. At each split in a particular tree, from a random subset of predictors a single predictor is chosen that produces branches with the best split of Normal and Abnormal samples. Each resulting branch is then split until it ends in a "leaf" containing only (or mainly) Normal or Abnormal training samples. A test sample is then pass through each tree, and is assigned the (majority) class of the leaf on which it lands. Every tree in the forest produces a prediction and these predictions are combined to give a single consensus prediction for the individual[28 29]. We used 20 repeats of 5-fold cross validation on the training data to select the

optimum number of randomly selected predictors (*mtry*) that maximized the AUC. Overall variable importance was determined by permuting predictors one at a time and measuring the mean decrease in accuracy averaged over all trees.

These two predictive modelling techniques, as well as more general statistical modelling and machine learning considerations including pre-processing, test-train dataset split, cross validation, variable importance, and class imbalance are described in detail in the supplementary material.

Results

Capsule endoscopy findings

147 FDRs underwent VCE: in 144 the caecum was reached. There was one capsule retention, managed conservatively with prokinetic agents. The commonest abnormal finding was of small aphthous ulceration in the distal ileum, and no strictures were identified. Marked inflammation typical of CD (>150 aphthous ulcers throughout the small intestine) was found in only one FDR, in the high risk group. 94% of the lowest risk group had no SI inflammation (Lewis score <135) compared with 68% of the highest risk group (p=0.0001). In the highest risk group, 9% had moderate- severe inflammation (Lewis score ≥790) compared with none in the low risk group (p=0.016).

Characteristics of the highest and lowest risk quartiles

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In total, 124 of the 147 FDRs who underwent VCE had complete study data and are included in all analyses. **Table 2** shows the distribution of proband relationship, sex, age, smoking status, CD family history, and VCE-determined SI inflammation by combined genotype and smoking relative risk. Among these variables, only smoking and Lewis score showed a significant difference by risk quartile. The Lewis and CECDAI capsule scores were highly correlated (Pearson's r=0.89, 95% CI=0.85-0.92, p<0.01) and when each measure was dichotomized (Normal, Abnormal) only 3 samples were classified differently. For all modelling we used the dichotomized Lewis score as our outcome, as this has previously been shown to correlate well with FC where the CECDAI did not[30].

Table 2. Distribution of demographics, relationship and Lewis score by risk quartile

	Highest risk quartile	Lowest risk quartile	Difference
N = 124	67	57	p = 0.25
Relationship to proband			p = 0.90
sibling	26 (39%)	23 (40%)	
offspring	29 (43%)	22 (39%)	
parent	12 (18%)	11 (19%)	
Sex			p = 0.94
female	45 (67%)	37 (65%)	
male	22 (33%)	20 (35%)	
Age			p = 0.94
mean	37.35	37.49	
median	*37.4*	*36.2*	
range	*19.5 – 55.6*	*20.5 – 56.7*	
Smoking status			p = 1.1 x 10^{-6}
current	23 (34%)	6 (11%)	
ex	25 (37%)	8 (14%)	
never	19 (28%)	43 (75%)	
CD family history			p = 0.78
single	59 (88%)	52 (91%)	

multiple	8 (12%)	5 (9%)	
Lewis score			$p = 3.7 \times 10^{-4}$
Normal (<135)	45 (67%)	53 (93%)	
Abnormal (≥135)	22 (33%)	4 (7%)	

Difference = test of significant differences by risk quartile. Percentages are column percentages within each subheading. p-values relate to a test for equality of proportions, a *t*-test of mean differences, a Pearson's chi-squared test, or a Fisher's exact test as appropriate.

Explanatory modelling

A logistic regression model for SI inflammation (based on Lewis score) was built including all variables. Using stepwise selection and Akaike's information criterion (AIC) we reduced the number of predictors to a best explanatory subset of 14 (R^2=0.72). Stepwise selection using the Bayesian information criterion (BIC), which seeks a more parsimonious model, yielded a 3 variable model (R^2=0.47) including FC, CD family history (binary-coded "single" or "multiple"), and genetic risk score (supplementary **Table S4**). Although stepwise selection achieves a reduction in the number of predictor variables, and can quantify the variance in the outcome explained by the predictors, these classic statistical approaches do not account for correlation among predictors (supplementary **Figure S1**) and provide no indication of the predictive ability of the derived model.

Predictive modelling

Predictive models were built on a training sample of 83 FDRs (2/3 of the data), and predictive ability assessed on the test sample of 41 FDRs (remaining 1/3 of the data). The predictive performance of the random forest (AUC=0.87, 95% CI=0.75 − 1.00; Accuracy=0.73, 95% CI=0.57 − 0.86) was slightly better than the elastic net

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(AUC=0.80, 95%CI=0.62 − 0.98; Accuracy=0.68, 95% CI=0.52 − 0.82). This superior

performance was the result of correctly classifying an additional 2 test samples (1

Normal, 1 Abnormal) among the 41 unseen test samples (**Table 3**). A full set of

model performance metrics is included in supplementary **Table S5**.

Table 3. Confusion matrices: cross-tabulations of observed (reference) and predicted classification

		Elastic net				Random forest	
		Reference				*Reference*	
		Normal	Abnormal			Normal	Abnormal
Prediction	Normal	22	2	*Prediction*	Normal	23	1
	Abnormal	11	6		Abnormal	10	7

Note. Final model parameter values: elastic net α=0.25, λ=0.50, and the random forest *mtry* = 14 (number of randomly selected predictors).

Figure 1 provides a visual summary of the predictive performance of our two

classifiers. In a receiver-operator characteristic (ROC) curve, a good classifier moves

away from the diagonal line (which indicates chance prediction) closer to the top left

corner; in a precision-recall (PR) curve a good classifier bows closer to the top right

hand corner. Taken together, the ROC and PR curves in **Figure 1** suggest that the

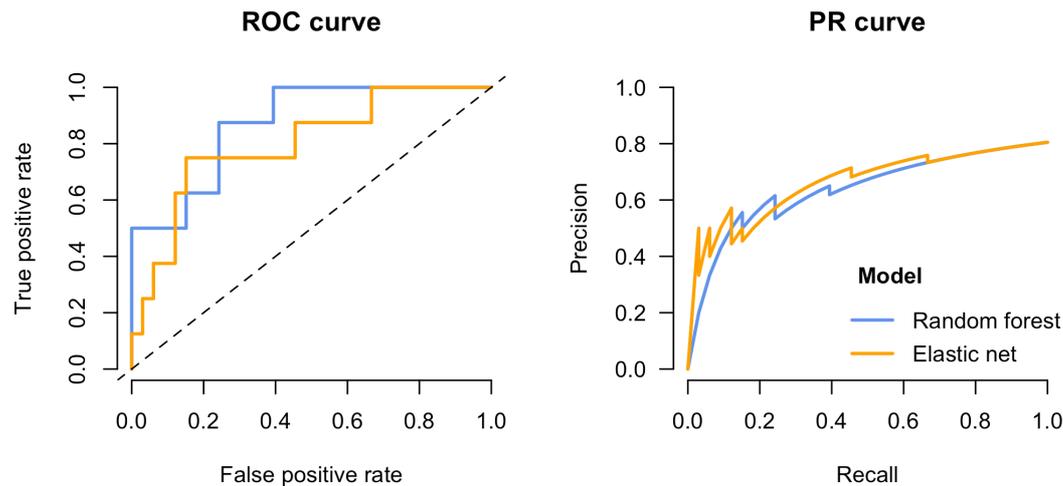models performed similarly, and above chance.

Figure 1. Classifier evaluation. Test sample performance was similar for the elastic net and random forest: elastic net sensitivity=0.75, specificity=0.67; random forest sensitivity=0.88, specificity=0.70. ROC curve: True positive rate (or Sensitivity) = TP / (TP+FN); True negative rate (or 1-Specificity) = 1 - (TN / (TN+FP)); PR curve: Precision (or Positive Predictive Value) = TP / (TP+FP) and incorporates the type I error rate; Recall = Sensitivity

Predictor importance

Table 4 shows the predictors selected by the elastic net and their relative importance in both classifiers. The elastic net reduced the full set of predictors to CD family history (OR=1.31), genetic risk score (OR=1.14), and FC (OR=1.04), which are the same 3 predictors selected by stepwise BIC (supplementary **Table S4**). There is no acceptable way to assign p-values or confidence intervals to individual predictors in a cross-validated elastic net solution as standard approaches ignore "the complex selection procedure for defining the reduced model in the first place"[31]. Beyond the standard clinical predictors (family history and FC), genetic risk provided additional predictive utility.

Table 4. Relative predictor importance

Rank	Elastic net		Random forest	
	Predictor	*Importance*	*Predictor*	*Importance*
1	CD family history	0.27	faecal calprotectin	0.0967
2	genetic risk score	0.13	genetic risk score	0.0264
3	faecal calprotectin	0.04	hs-CRP	0.0151
4			IL6	0.0043
5			CD family history	0.0023

Note. Only top 5 predictors shown for the random forest. The random forest importance is the mean decrease in accuracy given by the difference in error rate after permuting the particular variable, averaged over all trees. The elastic net importance is the absolute value of the regression coefficient *t*-statistic.

Genetic risk score improves predictive performance

Genetic risk score was one of three variables selected by the elastic net and ranked the second most important predictor by both classifiers. However, it is still relevant to determine how much additional predictive value genetic risk adds. A model built on all variables excluding genetic risk score produced a lower AUC (0.78, 95% CI: 0.55 − 1.00, p=0.0071) based on two variables: FC (OR=1.09) and CD family history (OR=1.63), compared to the full elastic net model AUC (0.80, 95% CI=0.62 − 0.98, p=0.0039 (Mann-Whitney *U* test derived[32])).

Conclusion

In our sample of 124 asymptomatic FDRs, we found that the combination of number of CD-affected relatives, genetic risk score, and level of FC made a good predictor of SI inflammation. The elastic net and random forest classifiers performed similarly, with significant above-chance prediction of SI inflammation.

The two models assume a different relationship between predictors and outcome: elastic net combines predictors linearly, whereas random forest makes multiple binary splits of the observations on one predictor at a time. In the elastic net, it is the combined set of selected predictors that captures the variance explained by correlated factors not selected. For example, the elastic net did not select smoking. This does not suggest smoking is any less important as a risk factor for the disease but is simply a product of the study design and the correlational pattern of the particular predictors included. The random forest, which by design does not exclude predictors, ranked smoking among the top 10 most important variables. Given the different design of the classifiers, there is no expectation that they should agree. However, that their results converge – CD family history, genetic risk score, and FC appeared among the top 5 predictors of the random forest – increases confidence in the utility of the particular set of predictors included.

CD family history was the strongest predictor in the elastic net model, and ranked 5th most important predictor in the random forest. A retrospective study of the entire Danish population followed for a 44-year period found that the incidence rate for IBD was increased among individuals with two or more affected relatives[2]. By design, all our study participants had at least one relative with CD; we found that having a stronger family history (two or more relatives with CD) increased the risk for SI inflammation. It is worth noting that although studies often include the raw number of affected family members as a risk factor, it is a crude measure of familial risk. Ideally, total family size, structure, and age of affected individuals as well as raw count should be incorporated in a measure of "family history"[33].

We found that 44 of 124 FDRs (35%) had abnormally elevated FC (≥50 μg/g, usually indicative of intestinal pathology[12]). Asymptomatic FDRs with increased FC have previously been observed[12]. Given that FC is part of the diagnostic work-up for CD[10], and increased levels predict relapse[34], our finding of the predictive utility of FC for SI inflammation suggest it is a promising biomarker for detecting those at greatest risk for CD. However, using a cut-off of ≥50 μg/g might include too many false positives - a cut-off of >250 μg/g may be a more appropriate level for a screening tool in asymptomatic individuals as this correlates with active CD[35].

Genetic risk score was a significant predictor of SI inflammation. The predictive accuracy of our elastic net classifiers with or without genotype were both within the 0.7–0.8 range generally considered to be "acceptable discrimination"[36]. One way to evaluate the added value of genetic risk score is to consider what difference it might make in clinical practice. Our estimates showed a 2-percentage point increase in AUC on inclusion of genetic risk. Given that AUC is the probability that in a randomly selected case-control pair the case will be assigned a higher risk score, the model including genetic risk score accurately predicts an additional 2 cases in every 100 randomly selected pairs over the model without genetic risk. Updating the genetic risk panel as more risk loci are discovered is likely to increase the performance of the model further, and would be feasible now that low cost genome-wide SNP arrays are available which provide good coverage of known common IBD risk variants.

Mild SI inflammation found at VCE has been reported in 24% of asymptomatic FDRs, but the subsequent development of CD was not determined[37]. A study of 38 FDRs who underwent ileocolonoscopy found mild endoscopic and histological inflammation in 26% and CD in 13%. Those with mild inflammation were followed up with repeat ileocolonoscopy after a mean of 53 months without endoscopic or histological progression of inflammation[13]. In our study 26/124 (21%) of FDRs had abnormal Lewis scores, which is broadly consistent with these data. Long term follow-up of our FDR cohort is planned (out to 10 years), and may provide further information regarding risk of developing overt CD as opposed to asymptomatic SI inflammation. It is not yet certain whether these features are predictive of future development of CD, or are simply a spectrum of the "at-risk" individual who has not had the necessary environmental exposure(s) required to develop disease. Progression to CD in cases of isolated terminal ileitis found at ileocolonoscopy in the general population is not clear-cut, largely due to small study sample sizes, short duration of follow-up, heterogeneity of symptoms, and retrospective study design. Nevertheless, the rate appears to be low: 1% in a study with a mean follow-up of 29.9 months[38],and 5% in another with a median follow-up of 97.5 months.[39].

Although we used an unseen subset of samples (set aside from the full sample) to test the predictive performance of our model, this test sample was subject to the same design decisions as the training samples, as well as any study-specific idiosyncrasies. Ultimately our finding will require external validation, i.e., replication in a completely independent sample. We included 72 CD-associated genetic variants known at the time of clinical assessment[16], and the single best-understood lifestyle

factor, smoking. A replication could benefit from the inclusion of the up-to-date list of 240 IBD-associated SNPs[7] (given that most IBD loci confer risk for both CD and UC[40]), and a more comprehensive picture of the environmental risk (i.e., in addition to smoking, medication, diet, stress, sleep, physical activity[9]). Finally, this study did not assess the gut microbiome, a likely predictor of risk for CD and potential target for intervention. The Genetic, Environmental and Microbiome (GEM) Project, currently recruiting 5000 CD FDRs internationally, will hopefully bring further insights into pre-clinical CD with the combination of all of these factors[41].

In parallel with our increasing understanding of the specific genetic and environmental risk factors that combine to make an individual's immune system hostile to commensal gut flora, a clinically useful tool for detecting those at greatest risk for CD before the presentation of overt symptoms would prioritise patient screening and follow up. Early detection opens up the possibility for early intervention, additional targets for drug development, and disease prevention. Our study suggests that a CD prediction tool can be built from a small set of biomarkers, known genetic risk variants, and family history of CD. Inclusion of risk factors from the most recent findings will improve prediction accuracy.

Ethical approval & consent

Ethical approval was granted by Health Research Authority, UK (10/H0710/011) and all participants provided their written informed consent.

Funding

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