Patient reported distress can aid clinical decision making in idiopathic pulmonary fibrosis: analysis of the PROFILE cohort

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Abstract

Idiopathic pulmonary fibrosis is a progressive and fatal interstitial lung disease. We aimed to determine if patient response to a palliative assessment survey could predict disease progression or death.

We undertook a cross-sectional study in a UK clinical cohort of incident cases. Rasch-based methodology provided a disease distress value from an abridged 11 item model of the original 45 item survey. Distress values were compared with measures of lung function. Disease progression or mortality alone was predicted at twelve months from survey completion, with risk of death assessed at three, six and twelve months.

Disease distress values were negatively correlated with lung function (r=-0.275 percent predicted DLCO). Expected survey scores computed from distress values could distinguish disease progression, 8.8 (p=0.004), and people who died, 10.2 (p=0.002), from those who did not progress, 6.9. Actual survey scores predicted disease progression and mortality with an area under the curve of 0.60 and 0.64, respectively. Each point increment in actual score increased risk of twelve-month mortality by 10%, almost 43% of people scoring above 18 did not survive beyond 105 days.

We show 11 clinically identifiable items can be used to score disease distress and predict prognosis to assist in clinical decision making.

1. Introduction

Idiopathic pulmonary fibrosis (IPF) is a progressive interstitial lung disease of unknown aetiology, causing scarring of the lung, shortness of breath, cough and reduction in lung function. A UK study demonstrated an increase in incidence of IPF by 35% between 2000 and 2008, with higher incidence in men and older age groups [1]. Incidence and mortality continues to rise and is also increasing globally [2]. IPF is fatal with no current cure, though disease-modifying treatments are being introduced. In many cases, earlier palliative intervention could reduce the burden on the individual, as recommended in cancer care [3, 4], yet the variable natural history of IPF makes it difficult to predict care needs.

The Sheffield Profile for Assessment and Referral to Care (SPARC) is a holistic needs tool, comprising 45 items across 9 domains, that can aid health professionals in identifying needs for palliative care [5]. Respondents find the questions easy to understand [6] and the tool has been adapted to international settings [7]. Such systematic assessment provides a useful indication of symptom distress [8]. Yet those who suffer chronic disease may not achieve the SPARC criteria for immediate clinical assessment, defined as a score of 3 in a single question, as habituation or affective comorbidity can affect symptom perception [5, 9, 10]. Regular completion of 45 items can be burdensome for those who suffer most [8], whilst self-assessment may result in interpatient differences despite the same underlying distress.

Rasch-based methodology generates a scaled value for a set of responses as an interaction between question difficulty and the individuals' likelihood of scoring. Initially used to standardise scholastic tests, extension of the Item Response Theory (IRT) class of models to allow multiple choice or graded options has found increased recognition in developing healthcare metrics and patient-reported outcomes [11-15]. Further advantages of the method enables question refinement, banking and assessment of the impact of demographics on item response [16-18].

We aimed to determine whether the SPARC could be used with IRT methodology to generate a tool that could appropriately distinguish disease progression in a UK clinical cohort of IPF patients [19, 20], helping to predict short-term prognosis.

2. Methods

The Prospective Observation of Fibrosis in Lung Clinical Endpoints (PROFILE) Central England (NCT01134822) study is a longitudinal observational clinical trial that has been described previously [19, 20]. Age was grouped at 65 or under, 66-79, 80 or over; comorbidities were grouped as none, 1-2, 3 or more. People were asked to complete the SPARC questionnaire, comprised of dichotomous items in two domains, as well as polytomous items in six domains (Supplementary Doc 1). 243 people from the PROFILE Central England study completed the SPARC and were subsequently included in the analysis.

All analyses were performed in Stata (SE15.1). Lung function measures recorded at the initial SPARC assessment (± 30 days) were used as baseline, with an outcome of disease progression within one year of SPARC completion defined as 10% relative decline in FVC or death. Percent predicted DLCO (%DLCO) and FVC (%FVC) was calculated using the suite of Global Lung Function Initiative (GLI) tools available from the European Respiratory Society [21]. Where %FVC (n=21) or %DLCO (n=75) could not be confirmed due to missing data, people were excluded from the relevant statistical analyses.

A two-parameter graded response model was ultimately constructed with items showing good discrimination and model assumptions were verified [11, 22]. Parameters of discrimination and difficulty were assigned to each item (survey question) according to how well it could differentiate people across the scale of the underlying distress trait (theta), as well as the probability of a particular answer. Detailed information on the construct of the IRT model is provided as supplementary information; a final model of 11

items was built, where larger theta values indicate more distress. (Supplementary Doc 2; Supplementary Table 1; SupplementaryFigure 1).

Pearson's Correlation determined whether distress (theta) values correlated with %DLCO or %FVC measures. %DLCO was log transformed to meet normality assumptions. IRT test characteristic curves estimated expected questionnaire scores according to distress values calculated from the concise 11 item model, hereafter termed IPF-abridged Profile for Assessment and Referral to Care (I-PARC). One-way ANOVA assessed differences in mean distress between categories of %DLCO (<40%, 40-60%, >60%). Two-way t-test assessed mean distress between those with disease progression and those without, whilst one-way ANOVA additionally assessed categories of disease progression (no disease progression, lung function decline only, death only). Tukey post-hoc analysis between categories was applied.

We calculated the area under the Receiver Operator Characteristic (ROC) curve for the ability of the cumulative I-PARC Score to predict disease progression within 12 months, as well as overall mortality compared with %FVC or %DLCO. The sensitivity and the specificity was compared using the chi-squared test [23].

Kaplan-Meier curves plotted time to disease progression, or overall mortality, against days since completing the questionnaire according to categories of I-PARC Score. Cox regression estimated the hazard ratio (HR) of disease progression, as well as death in 365 days (12 months), 180 days (6 months) and 90 days (3 months) according to increment or categories of I-PARC Scores; lowest scorers as reference. The risk in comparison groups is contingent upon the proportion of the reference group that fails during the specified timescale. Analyses were initially univariate, and then adjusted for age, comorbidity, gender, and a significant interaction between age and comorbidity. Proportional hazard assumptions were checked using Schoenfeld residuals.

3. Results

Demographics

From 243 people within the cohort, 103 (42.4%) had evidence of disease progression within one year, whilst 140 (57.6%) did not (Table 1). Of those with disease progression, 49 died within one year (47.6%). Within the disease progression subgroup, 80.6% were male but no significant gender interaction was observed, similarly comorbidity count was not associated with progression. Those with disease progression had greater representation by ages 80 or over. We identified no significant relationships between communication and progression status, whilst we observe lower proportions with disease progression wanting information about personal finances. No relationship reached significance following Bonferroni correction.

Table 1: Cohort demographics

		No		
Outcome in 1 year:	All	progression	Progression	(p value)
Number	243	140	103	
Gender (%)				
Male	78.2	76.4	80.6	
Female	21.8	23.6	19.4	0.438
Age group (%)				
65 and under	15.6	19.3	10.7	
66 - 79	70.0	70.7	68.9	
80 and over	14.4	10.0	20.4	0.026
Comorbidity count (%)				
None	23.9	26.4	20.4	
1 to 2	47.7	49.3	45.6	
3 or more	28.4	24.3	34.0	0.219
Percentage (%) who wer	e able to talk to);		
Doctor	80.7	80.0	81.6	0.055
Community nurse	14.4	12.1	17.5	0.150
Hospital nurse	53.5	58.6	46.6	0.185
Religious advisor	2.9	2.1	3.9	0.360
Social worker	3.7	4.3	2.9	0.657
Family	65.0	67.1	62.1	0.977
Percentage (%) who war	nted more infor	mation on their:		
Condition	28.8	32.9	23.3	0.206
Care	13.6	14.3	12.6	0.894
Treatment	18.1	21.4	13.6	0.185
Support	14.0	13.6	14.6	0.653
Finances	9.1	12.9	3.9	0.025

Proportions include missing data. Missing values not presented, Chi-squared test p-values based on non-missing data. Significant p-values before Bonferroni correction (p<0.05) in bold. P-value following correction p=0.0036.

Disease distress is associated with lung function and disease progression

A mild negative correlation was observed between distress values and baseline %FVC or %DLCO, indicating that higher distress correlated with worse lung function performance (Figure 1). The Pearson correlation coefficient for distress and %FVC was -0.267 (p=0.0001), whilst for distress and %DLCO it was -0.275 (p=0.0003).

The mean distress value according to severity of %DLCO was calculated and the IRT test characteristic curve estimated the expected score from the continuum of the distress values, calculated from the I-PARC Score model. Mean distress values were significantly different according to severity of %DLCO (p=0.006). People with %DLCO <40% were significantly more distressed than those with a %DLCO >60% (p=0.004), leading to expected scores of 9.1 (θ =0.22) and 5.6 (θ =-0.38), respectively (Figure 2 A).

Mean distress values were also significantly different according to disease progression (p=0.0085), where those who progressed had an expected score of 8.8 (θ =0.18) compared to an expected score of 6.9 for those who did not progress (θ =-0.14) (Figure 2 B-i). Plotting mean distress of patients that died separately provided an expected score of 10.2 (θ =0.39); significance was identified across categories of disease progression (p=0.003), largely driven by the difference between those who died and those who did not progress (p=0.002) (Figure 2 Bii-iii).

Lung function recordings require a suitable fitness to provide acceptable and repeatable measurements, introducing selection bias if individuals struggle to perform them. To understand the influence of missing data we evaluated distress according to completion of lung function (see Supplementary Table 2). Patients with missing %FVC data were more likely to have died (61.9%) than patients with complete %FVC data (16.2%; p<0.0001). However, there was no difference in mean distress between people with missing %FVC compared with complete data (p=0.384, compare scores 8.8 versus 7.6). Similarly, patients with missing %DLCO data were more likely to have died (34.7%

versus 13.7%; p <0.0001), although those missing the data were also more distressed (p=0.002; compare scores 9.5 versus 7.0). In total, 75 people (31%) were missing DLCO measures; analyses based on non-missing DLCO recordings will underestimate actual effect sizes as they exclude a large sample of distressed patients.

I-PARC Score can predict disease progression and mortality

Area under the curve (AUC) of the ROC curve initially assessed the ability of the I-PARC Score to predict disease progression in the complete dataset (Figure 3), resulting in an AUC of 0.60 (95%CI: 0.53-0.68). ROC subsequently assessed the ability of the I-PARC Score to predict death compared with lung function recordings in 167 people with complete data. We identified %DLCO as having the largest AUC of 0.83 (95%CI: 0.75-0.91), whilst %FVC had an AUC of 0.66 (95%CI: 0.56-0.76). The cumulative I-PARC Score provided an AUC of 0.64 (95%CI: 0.52-0.77); a minimum score of 9 resulted in 59.2% sensitivity and 62.9% specificity. No statistical difference was observed between the AUC of I-PARC Score and %FVC (p=0.8), although %DLCO was significantly better at predicting mortality (p=0.013). Ability of I-PARC Score to predict mortality was also assessed independently from lung function recordings in the complete dataset (see SupplementaryFigure 2), resulting in a greater AUC of 0.66 (95%CI: 0.57-0.74). Kaplan Meier curves of time to event were plotted separately for disease progression and death according to observed I-PARC Score (Figure 4), categorised using expected score from mean distress where % DLCO <40% (score of 9). A large proportion of people scoring over 18 from the 11 I-PARC items had rapid disease progression, 51.5% of people scoring at least 9 progressed within one year. Fewer than 75% of people who scored between 9 and 18 survived 365 days (Figure 4 B). A total of 101 people scored 9 or more and 29 died within 365 days, providing a positive predictive value of 28.7%; this is compared with 20 people from the 142 scoring 8 or less, 14.1%. Scoring over 18 increased the positive predictive value to 42.9% from a total of 14 people, all deaths in this group occurred within 105 days.

Cox regression estimates of the risk of disease progression and risk of death in one year for each unit increment of the I-PARC Score are reported in Table 2. For every point increase, the risk of disease progression increased by 5% (HR 1.05 95%CI 1.02-1.09) whilst risk of death increased by 10% (HR 1.10 95%CI 1.05-1.15). Disease progression was almost 60% more likely for people scoring 9-18 (HR 1.59 95%CI 1.04-2.44), compared to those scoring less, whilst those scoring highest were at three times the risk (HR 3.19 95%CI 1.53-6.64). People scoring between 9 and 18 were more than twice as likely to die within 365 days as those scoring less than 9 (HR 2.43 95%CI 1.30-4.55). Estimates for those scoring over 18 did not meet proportional hazard assumptions at 365 days, although their risk of death in 180 days was 12 fold that of those scoring less than 9 (HR 12.32 95%CI 3.84-39.55), whilst the risk for those scoring between 9 and 18 more than doubled (HR 2.82 95%CI 1.03-7.68). Similar results were estimated at 90 days for those scoring highest (HR 19.47 95%CI 1.98-191.13), though scoring between 9 and 18 was not statistically different from scoring less (HR 7.69 95%CI 0.89-66.28).

Table 2. Cox regression risk estimates of disease progression and overall mortality by I-PARC Score.

	Number	% fail	HR	95%CI	HR*	95%CI*
365 day disease progression risk		1111	33,001	1113	337001	
I-PARC Score	243	42.39%	1.05	(1.01-1.08)	1.05	(1.02-1.09)
I-PARC category						
Score 0-8	142	35.92%	1		1	
Score 9-18	87	49.43%	1.50	(1.00-2.26)	1.59	(1.04-2.44)
Score 19-33	14	64.29%	2.69	(1.32-5.47)	3.19	(1.53-6.64)
365 day mortality ri	sk					
I-PARC Score	243	20.16%	1.08	(1.03-1.12)	1.10	(1.05-1.15)
I-PARC category						
Score 0-8	142	14.08%	1		1	
Score 9-18	87	26.44%	2.02	(1.11-3.68)	2.43	(1.30-4.55)
Score 19-33	14	42.86%	4.44	(1.78-11.08)	† 5.74	(2.20-14.97)
180 day mortality ri	sk					
I-PARC category						
Score 0-8	142	4.93%	1		1	
Score 9-18	87	11.49%	2.43	(0.92-6.37)	2.82	(1.03-7.68)
Score 19-33	14	42.86%	11.69	(3.92-34.86)	12.32	(3.84-39.55)

90 day mortality risk						
I-PARC category						
Score 0-8	142	0.70%	1		1	
Score 9-18	87	5.75%	8.37	(0.98-71.61)	7.69	(0.89-66.28)
Score 19-33	14	21.43%	34.25	(3.56-329.37)	19.47	(1.98-191.13)

 $[\]hbox{*adjusted for age group, comorbidity count, gender, age comorbidity interaction.}\\$

4. Discussion

The SPARC questionnaire is recognised within the UK National Health Service as a way to help address and improve end-of-life care management [5, 6, 9, 24]. We have utilised a unique clinical cohort of patients with IPF [19], determining that items specified within the I-PARC list can identify patients at high risk of death within 3-6 months, which could be appropriate for triggering referral to palliative care (Box 1).

Original SPARC item x	I-PARC Score item	In the past month, have you been distressed or bothered by:	Notatall	A little bit	Quite a bit	Very much
7	1	Shortness of breath	0	1	2	3
13	2	Feeling weak	0	1	2	3
14	3	Feeling tired	0	1	2	3
16	4	Feeling sleepy in the day	0	1	2	3
17	5	Loss of appetite	0	1	2	3
21	6	Feeling restless and agitated	0	1	2	3
22	7	Uncontrolled symptoms	0	1	2	3
41	8	Side effects of treatment	0	1	2	3
34	9	Losing independence	0	1	2	3
35	10	Ability to carry out daily activities	0	1	2	3
36	11	Ability to carry out household tasks	0	1	2	3
		TOTAL		<9	9-18	>18

Items from final model presented. Highest scores, 19 or above, may be considered an indicator for follow-up within three months; intermediate scores, 9-18, support review within six months; lower scores, <9, support review within 12 months.

Proportional hazard assumptions met unless stated (†). HR: hazard ratio

We use Item Response Theory as a novel application of the methodology to assess patient reported outcome measures [11], identifying a concise list of 11 items from an original 45 questions that capture the majority of information distinguishing people with high distress. Disease distress correlated with percent predicted lung function and could characterise disease progression. Furthermore, scores could predict mortality with those scoring highest also at greatest risk of death. The associations of disease distress with lung function, disease progression and mortality support the I-PARC Score as a valid tool to inform prognoses.

Strengths and limitations of Item Response Theory application

IRT is an increasingly applied methodology that lends itself to parametric tests, providing an evidence base for reusing questions that can be used to distinguish traits of interest [16, 18, 25]. This study supports the use of IRT methodology in optimising questions within healthcare surveys to predict patient outcomes. Questions that did not distinguish levels of distress, such as those which scored low for the vast majority (e.g. O33: Religious or spiritual needs not being met), or those that can show clustering of responses (e.g. Q30: Thoughts about ending it all; Q32: Worrying thoughts about death or dying) can be dropped in order to retain items offering the most discriminatory value. IRT methodology can also assess the way in which an item may be answered differently according to demographic traits that are unrelated to the latent trait of distress, termed differential item functioning (DIF)[16, 17]. We identified no DIF in the final model according to gender, whilst the cohort was not sufficiently diverse to assess DIF by ethnicity. DIF by demographics of age category and comorbidity count was identified for item 21 (Feeling restless and agitated), however the analyses indicate an underlying relationship between these demographics and disease distress, so are not defined as DIF.

The original SPARC questionnaire collects distress information from an extensive list of possible issues, with a response of 'very much' (score of 3) in any of the 41 scaled items

acting as a flag that the individual would benefit from an immediate palliative care assessment. Within clinical settings, the SPARC questionnaire could be considered too sensitive, or excessive depending on health status [6, 9], whilst under-reporting of distress is an issue in progressive disease [10, 26]. Use of the cumulative score over 11 items provides greater opportunity to capture distress when under-reported, whilst shorter health questionnaires can provide similar insight to lengthier ones [27].

The mean distress for those with lowest %DLCO equated to a score of 9.1, with a score of 9 being subsequently used as a threshold point for grouping people who may undergo disease progression within one year. Lung function measures are a valuable tool in defining disease progression, although limited by the scheduling of measurement recordings that may not reflect the actual timing of progression. Those who progressed had an average score of 8.8, which supports an I-PARC Score of 9 as a suitable benchmark, particularly as those who died had an average score of 10.2.

Other self-reported surveys exist to measure health status in those with progressive lung disease, including King's-Brief Interstitial Lung Disease (K-BILD) questionnaire[15]. The authors of K-BILD used patient interviews to define a series of pertinent questions and Rasch analysis to refine the item list to 15 with 4 domains and a 7-point Likert scale. We use similar methodology to refine the list of predetermined SPARC items, as answered by a sample of 243 IPF patients, to a set of 11 within a single domain and a 4-point Likert scale. Whilst K-BILD offers an empowering way to self-monitor respiratory health, I-PARC should be completed within clinical settings where rapid disease progression is a factor.

I-PARC Score offers prognosis estimates for clinical decisions

ROC analysis confirmed that %DLCO was the best measure for predicting one year mortality in this cohort [28]. I-PARC Score appeared to perform slightly better at identifying true positives compared with %FVC when the accepted proportion of false positives was restricted to 50%. Lung function measurements can be challenging for

patients with severe disease to provide, and this appears particularly apparent in people missing DLCO who were significantly more distressed than those without missing DLCO values. These data illustrate the considerable survival bias associated with lung function data in studies of patients with progressive lung disease. As a result, these analyses may underestimate I-PARC performance relative to lung function measures, yet demonstrate the value of patient-reported measures.

Survival curves indicated that whilst the majority of the cohort survived one year after completing their SPARC questionnaire, those with the highest I-PARC scores had poorer life expectancy than those scoring lowest. The positive predictive value of mortality in one year was 29% for those scoring above the threshold of 9, and increased to 43% for those scoring 19 or more. The findings provide evidence that patient reported distress can predict early mortality with similar accuracy to predictions made with FVC lung function recordings.

Adjusted survival analyses further show that risk of one-year mortality increased by 10% with each incremental point, whilst risk of progression increased by 5%. Intermediate scoring of 9 to 18 heightened the likelihood of progression and more than doubled the risk of death, compared to those scoring lower. It can be observed that a large proportion of the highest scorers did not survive beyond 105 days, although these estimates are based on low numbers. This relatively large sample of people with confirmed IPF supports the benefit of patient reported distress in predicting disease progression and death.

Given that it is challenging to perform lung function in patients with progressive disease it is important to develop clinical markers of prognosis as well as assessing health status. The I-PARC Score requires no specialist equipment or training to complete, takes consideration of the individual's concerns, can be undertaken at any time without a requirement for repeated measures, and can aid decisions to review patients earlier. Combined with its associations with lung function and mortality, this patient-reported

outcome is a potential component of composite endpoints in clinical trials, although further validation and studies on treatment sensitivity are essential [29].

The predictive capacity could not exclude all false positives, yet we recognise that early integration of palliative care has improved outcomes in other progressive diseases [3, 4, 30]. Future studies may include I-PARC Score as part of a composite scoring system for accurately predicting IPF prognosis, adding value in treatment recommendations [28]. We recommend utilising it in combination with available lung function recordings when making clinical decisions (see Box 1). We welcome further study to validate the tool in separate cohorts of patients with progressive lung disease, including IPF.

Conclusion

The I-PARC Score, developed using the SPARC holistic tool, offers an encouraging method to assess prognosis and recognise palliative care needs for those with progressive disease. In producing a standardised model based on the latent trait of distress for people with IPF, we generated an abridged list of items that could distinguish those who died within one year, with higher scores having worse prognosis. Future study should address the validity of this concise tool in other progressive lung diseases.

References

- 1. Navaratnam, V., et al., *The rising incidence of idiopathic pulmonary fibrosis in the U.K.* Thorax, 2011. **66**(6): p. 462-7.
- 2. Hutchinson, J., et al., Global incidence and mortality of idiopathic pulmonary fibrosis: a systematic review. Eur Respir J, 2015. **46**(3): p. 795-806.
- 3. Hoerger, M., et al., *Defining the Elements of Early Palliative Care That Are Associated With Patient-Reported Outcomes and the Delivery of End-of-Life Care.* J Clin Oncol, 2018. **36**(11): p. 1096-1102.
- 4. Temel, J.S., et al., *Early Palliative Care for Patients with Metastatic Non–Small-Cell Lung Cancer.* New England Journal of Medicine, 2010. **363**: p. 733-42.
- 5. Gardiner, C., et al., Extent of palliative care need in the acute hospital setting: a survey of two acute hospitals in the UK. Palliat Med, 2013. **27**(1): p. 76-83.
- 6. Hughes, P., et al., Consumer views on a new holistic screening tool for supportive and palliative-care needs: Sheffield Profile for Assessment and Referral for Care (SPARC): a survey of self-help support groups in health care. Health Expect, 2015. **18**(4): p. 562-77.
- 7. Leppert, W., M. Majkowicz, and S.H. Ahmedzai, *The adaptation of the Sheffield Profile for Assessment and Referral for Care (SPARC) to the Polish clinical setting for needs assessment of advanced cancer patients*. J Pain Symptom Manage, 2012. **44**(6): p. 916-22.
- 8. Homsi, J., et al., *Symptom evaluation in palliative medicine: patient report vs systematic assessment.* Support Care Cancer, 2006. **14**(5): p. 444-53.
- 9. Ahmed, N., et al., A Pilot Randomized Controlled Trial of a Holistic Needs Assessment Questionnaire in a Supportive and Palliative Care Service. J Pain Symptom Manage, 2015. **50**(5): p. 587-98.
- 10. Petersen, S., et al., *Illness and symptom perception: a theoretical approach towards an integrative measurement model.* Clin Psychol Rev, 2011. **31**(3): p. 428-39.
- 11. Nguyen, T.H., et al., *An introduction to item response theory for patient-reported outcome measurement.* Patient, 2014. **7**(1): p. 23-35.
- Belvedere, S.L. and N.A. de Morton, *Application of Rasch analysis in health care is increasing and is applied for variable reasons in mobility instruments.* J Clin Epidemiol. 2010. **63**(12): p. 1287-97.
- 13. Tesio, L., Measuring Behaviours and Perceptions: Rasch Analysis as a Tool for Rehabilitation Research. Journal of Rehabilitation Medicine, 2003. **35**(3): p. 105-115.
- 14. Gomez, R.V., A., Item Response Theory Analysis of the Anxiety and Mood Disorders in Clinic Referred Children. Assessment, 2016. **25**(2): p. 235-246.
- 15. Patel, A.S., et al., *The development and validation of the King's Brief Interstitial Lung Disease (K-BILD) health status questionnaire.* Thorax, 2012. **67**(9): p. 804-10.
- 16. Teresi, J.A.O.-W., K.; Kleinman, M.; Eimicke, J. P.; Crane, P. K.; Jones, R. N.; Lai, J.S.; Choi, S.; Hays, R.; Reeve, B.; Reise, S.; Pilkonis, P.; Cella, D., *Analysis of differential item functioning in the depression item bank from the Patient Reported Outcome Measurement Information System (PROMIS): An item response theory approach.* Psychology Science Quarterly, 2009. **51**(2): p. 148-180.
- 17. Teresi, J.R., M.; Lai, J.S.; Silver, S., Occurrences and sources of Differential Item Functioning (DIF) in patient-reported outcome measures: Description of DIF methods, and review of measures of depression, quality of life and general health. Psychol Sci Q, 2008. **50**(4).
- 18. Stucky, B.D., et al., *Developing an item bank and short forms that assess the impact of asthma on quality of life.* Respir Med, 2014. **108**(2): p. 252-63.

- 19. Jenkins, R.G., et al., Longitudinal change in collagen degradation biomarkers in idiopathic pulmonary fibrosis: an analysis from the prospective, multicentre PROFILE study. The Lancet Respiratory Medicine, 2015. **3**(6): p. 462-472.
- 20. Maher, T.M., et al., An epithelial biomarker signature for idiopathic pulmonary fibrosis: an analysis from the multicentre PROFILE cohort study. The Lancet Respiratory Medicine, 2017. **5**(12): p. 946-955.
- 21. ERS, Global Lung Function Initiative Tools. European Respiratory Society E-Learning Resources: http://www.ers-education.org/guidelines/global-lung-function-initiative.aspx: p. last accessed July 2018.
- Yang, F.M. and S.T. Kao, *Item response theory for measurement validity*. Shanghai Arch Psychiatry, 2014. **26**(3): p. 171-7.
- 23. DeLong, E., D. DeLong, and D. Clarke-Pearson, *Comparing the Areas under Two or More Correlated Receiver Operating Characteristic Curves: A Nonparametric Approach.* Biometrics, 1988. **44**(3): p. 837-845.
- 24. Richardson, A., et al., *Patients' needs assessment in cancer care: a review of assessment tools.* Support Care Cancer, 2007. **15**(10): p. 1125-44.
- Paap, M.C., et al., *The COPD-SIB: a newly developed disease-specific item bank to measure health-related quality of life in patients with chronic obstructive pulmonary disease.* Health Qual Life Outcomes, 2016. **14**: p. 97.
- 26. Langsetmo, L., et al., *Underreporting exacerbation of chronic obstructive pulmonary disease in a longitudinal cohort.* Am J Respir Crit Care Med, 2008. **177**(4): p. 396-401.
- 27. Morishita-Katsu, M., et al., *The COPD assessment test and St George's Respiratory Questionnaire: are they equivalent in subjects with COPD?* Int J Chron Obstruct Pulmon Dis, 2016. **11**: p. 1543-51.
- 28. Raghu, G., et al., An official ATS/ERS/JRS/ALAT statement: idiopathic pulmonary fibrosis: evidence-based guidelines for diagnosis and management. Am J Respir Crit Care Med, 2011. **183**(6): p. 788-824.
- 29. Raghu, G., et al., *Idiopathic pulmonary fibrosis: clinically meaningful primary endpoints in phase 3 clinical trials.* Am J Respir Crit Care Med, 2012. **185**(10): p. 1044-8.
- Yoong, J., et al., *Early palliative care in advanced lung cancer: a qualitative study.* JAMA Intern Med, 2013. **173**(4): p. 283-90.

Figure 1. Negative correlation of distress with percent predicted lung function.

A) Plotted %FVC against distress value generated from I-PARC Score model resulted in a Pearson correlation coefficient (r) of -0.267 (p=0.0001) from 222 people (220 degrees freedom); %FVC explains 7.1% of variation in disease distress. B) Plotted Log transformed %DLCO against distress value resulted in r=-0.275 (p=0.0003) from 168 people (166 degrees freedom); %DLCO explains 7.6% of variation in disease distress.

Figure 2. Expected I-PARC Scores from mean distress in lung function categories. Item response theory test characteristic curve from 11 items in final model plots distress (theta) values against expected score. Scale of theta has a mean of 0 and the arbitrary range (4) represents incremental standard deviations. A) Mean theta values plotted for %DLCO category: >60% (long dash), 40-60% (short dash), <40% (solid); i. one-way ANOVA p=0.006; ii. Post-hoc Tukey analysis p=0.004 <40 %DLCO compared with >60. B) Mean theta values plotted for disease progression: no progression (long dash), disease progression (solid); i. Two way t-test p=0.0085. Death only (short dash); ii. One-way ANOVA death, decline only, no decline p=0.003; iii. Post-hoc Tukey analysis p=0.002 death compared with no decline.

Figure 3. Receiver-Operator Curve (ROC) assesses sensitivity and specificity of measures in predicting progression. A) Area under curve (AUC) for I-PARC Score in predicting disease progression = 0.60 (95% confidence interval 0.53-0.68) B. AUC for I-PARC Score (solid line) in predicting one year mortality =0.64 (0.52-0.77); %DLCO (long dash) AUC=0.83 (0.75-0.91); %FVC (short dash) AUC=0.66 (0.56-0.76). %DLCO AUC greater than I-PARC AUC, p=0.013.

Figure 4. Kaplan Meier analyses of time to event. A) Kaplan Meier Curves plot proportion of people with no disease progression against time since completing questionnaire: I-PARC Score low (0-8, solid line), intermediate (8-18, dashed), high (9-33, dotted). B) Kaplan Meier Curves plot proportion of people surviving against time since completing questionnaire: I-PARC Score low (0-8, solid line), intermediate (8-18, dashed), high (9-33, dotted).







