

# Low concordance of short-term and long-term selection responses in experimental *Drosophila* populations

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## 1 **ABSTRACT**

2 Experimental evolution is becoming a popular approach to study genomic selection responses  
3 of evolving populations. Computer simulation studies suggested that the accuracy of the sig-  
4 nature increases with the duration of the experiment. Since some assumptions of the com-  
5 puter simulations may be violated, it is important to scrutinize the influence of the experimen-  
6 tal duration with real data. Here, we use a highly replicated Evolve and Resequence study in  
7 *Drosophila simulans* to compare the selection targets inferred at different time points. At each  
8 time point approximately the same number of SNPs deviated from neutral expectations, but  
9 only 10 % of the selected haplotype blocks identified from the full data set could be detected in  
10 the first 20 generations. Those haplotype blocks that emerged already after 20 generations dif-  
11 fer from the others by being strongly selected at the beginning of the experiment and displaying  
12 a more parallel selection response. Consistent with previous computer simulations, our results  
13 confirm that only Evolve and Resequence experiments with a sufficient number of generations  
14 can characterize complex adaptive architectures.

15

16 **KEYWORDS:** experimental evolution, evolve & resequence, *Drosophila simulans*, early adap-  
17 tation, replicated time series data, concordance of adaptation, window-based analysis

## 18 **INTRODUCTION**

19 Deciphering the adaptive architecture is a long-term goal in evolutionary biology. In contrast  
20 to natural populations, experimental evolution (EE) provides the possibility to replicate exper-  
21 iments under controlled, identical conditions and to study how evolution shapes populations  
22 in real time (Kawecki et al. (2012); Schlötterer et al. (2015)). The combination of EE with next-

23 generation sequencing - Evolve and Resequence (E&R) (Turner et al. (2011); Schlötterer et al.  
24 (2015); Long et al. (2015)) - has become a popular approach to study the genomic response to se-  
25 lection and to identify adaptive loci. E&R has been applied to various selection regimes, such as  
26 virus infection (Martins et al. (2014)), host-pathogen co-adaptation (Papkou et al. (2019)), ther-  
27 mal adaptation (Orozco-Terwengel et al. (2012); Barghi et al. (2019)), or body weight (Johansson  
28 et al. (2010)). A wide range of experimental designs have been used, which vary in census popu-  
29 lation size, replication level, history of the ancestral populations, selection regime, and number  
30 of generations (Garland and Rose (2009); Turner et al. (2011); Kawecki et al. (2012); Lang et al.  
31 (2013); Burke et al. (2014); Huang et al. (2014); Hardy et al. (2018); Castro et al. (2019); Michalak  
32 et al. (2019); Seabra et al. (2019)). The duration of E&R studies ranged from less than 20 (Kelly  
33 and Hughes (2018); Turner and Miller (2012); Rêgo et al. (2019)), over a few dozen (Orozco-  
34 Terwengel et al. (2012); Johansson et al. (2010)), up to hundreds of generations (Burke et al.  
35 (2010)). Computer simulations showed that the number of generations has a strong influence  
36 on the power of E&R studies, and increasing the number of generations typically improved the  
37 results (Baldwin-Brown et al. (2014); Kofler and Schlötterer (2014); Vlachos and Kofler (2019)).  
38 Since simulations make simplifying assumptions, it is important to scrutinize these conclu-  
39 sions with empirical data. Until recently no suitable data-sets were available, which included  
40 multiple time points and replicates. We use an E&R experiment (Barghi et al. (2019)), which  
41 reports allele frequency changes in 10 replicates over 60 generations in 10 generation intervals,  
42 to investigate the impact of the experimental duration on the observed genomic response. The  
43 time resolved genomic data of this experiment allows to contrast putative selection targets in-  
44 ferred at different time points. We show that only a subset of the selection targets are detected  
45 at earlier generations, which are not representative of the underlying adaptive architecture.

## 46 **METHODS**

### 47 **Experimental *Drosophila simulans* populations**

48 A detailed description of the *Drosophila simulans* E&R experiment can be found in Barghi et al.  
49 (2019) and Hsu et al. (2019). The pooled individuals (Schlötterer et al. (2014)) from the evolving  
50 populations were sequenced every 10<sup>th</sup> generation starting with the founder population (gen-  
51 eration 0) until generation 60. This E&R experiment started from 202 isofemale lines, which  
52 were collected in Tallahassee, Florida. 10 replicate populations evolved in the laboratory at a  
53 "cycling hot" temperature regime (12 hours light and 28 °C, 12 hours dark and 18 °C). The cen-  
54 sus size of the replicates was 1,000 individuals with non-overlapping generations (Barghi et al.  
55 (2017, 2019); Hsu et al. (2019)).

### 56 **Genomic analysis hierarchy**

57 We investigated the genomic response of the experimental *Drosophila* populations on three dif-  
58 ferent levels: candidate SNPs, candidate SNPs in a window of fixed length and candidate SNPs  
59 shared with reconstructed selected haplotype blocks. A detailed description for each level is  
60 given below. Reasoning that the most reliable signal is detected at the most advanced genera-  
61 tion (60), we performed the same analysis at earlier time points and determined to what extent  
62 the same selection targets were identified as in generation 60.

### 63 **Identification of candidate SNPs**

64 Barghi et al. (2019) applied various filtering steps to obtain a robust SNP set from the ancestral  
65 population. In short, SNPs were called applying the following criteria: base quality of 40 in at  
66 least one replicate, a coverage between the 2<sup>nd</sup> and 98<sup>th</sup> percentile, and the minor allele is sup-

67 ported by at least ten reads. Repeats, transposable elements, SNPs specific to Y-translocated  
68 genes (Tobler et al. (2017)) and 5-bp regions around indels were excluded from the analysis to  
69 increase the robustness of the SNP set (for further details, see Barghi et al. (2019)), resulting in  
70 5,096,200 SNPs on chromosome X, 2, 3, and 4.

71 To study the selection response at different time points, we identified "candidate SNPs" based  
72 on the frequency difference between the ancestral and evolved populations for each of the six  
73 time points. Following Barghi et al. (2019), replicates were tested separately (Fisher's exact test)  
74 and jointly (Cochran-Mantel-Haenszel test, CMH) to identify SNPs with pronounced allele fre-  
75 quency change (AFC) using PoPoolation2 (Schlötterer et al. (2011)). Minimum and maximum  
76 coverage restrictions were not imposed because outlier SNPs with extreme coverage had already  
77 been removed. Neither the CMH test nor the Fisher's exact test account for AFC due to genetic  
78 drift. To detect SNPs that show more AFC than expected under drift, we performed neutral  
79 simulations with Nest (Jónás et al. (2016)) using estimates of the effective population size ( $N_e$ )  
80 between generation 0 and the focal time point (Table S1- S3). The simulations further used the  
81 empirical starting allele frequencies and sequencing coverages. For the CMH test,  $N_e$  estimates  
82 were averaged across replicates for autosomes and the X chromosome separately. For Fisher's  
83 exact test, we used replicate-specific  $N_e$  estimates of the autosomes. Based on these neutral  
84 simulations we determined candidate SNPs with a false discovery rate smaller than 5 % (Barghi  
85 et al. (2019)).

86 We identified 56,166 candidate SNPs in generation 60, compared to 55,199 in Barghi et al. (2019).  
87 This small discrepancy can be explained by stochastic differences arising from the neutral sim-  
88 ulations used to determine the significance threshold. We excluded six haplotype blocks (3.17,  
89 2.27, 3.21, 3.48, 3.49 and 3.54) (Barghi et al. (2019)) with less than 90% of the previously reported  
90 candidate SNPs.

## 91 **Identification of candidate windows**

92 The number of candidate SNPs is inflated as a result of linkage disequilibrium in the experi-  
93 mental populations (Nuzhdin and Turner (2013); Tobler et al. (2014)). To account for  
94 non-independence of candidate SNPs we used a window based approach. We split the main  
95 chromosomes (X, 2 and 3) into non-overlapping windows of 5,000 SNPs that are segregating in  
96 all generations and replicates. We chose SNPs instead of base pairs as window size measure to  
97 allow for variation in SNP density along the genome. To determine if a given window contains  
98 more candidate SNPs than expected, we sampled the same number of random SNPs as candi-  
99 date SNPs in this window (1,000 iterations). "Candidate windows" contained at least as many  
100 candidate SNPs as the 99<sup>th</sup> percentile of randomly sampled SNPs. Applying the procedure in-  
101 dependently to candidate SNPs from all time points provides time point specific candidate win-  
102 dows (Figure S1). We evaluated the similarity of two time points with the Jaccard index (for both  
103 candidate SNPs, and candidate windows).

104 The number of candidate SNPs in a window is a summary statistic which ignores the signifi-  
105 cance of the candidate SNPs. If a signal is robust between two time points, we expect the same  
106 *p*-value based ranking of candidate SNPs. Thus, we also evaluated whether candidate SNPs  
107 in a given window had a similar relative significance. For each candidate window we created  
108 a ROC-like curve (similar to Jakšić and Schlötterer (2016)) by ranking the candidate SNPs by  
109 their *p*-values - the most significant SNP was assigned rank 1 - and calculating the overlap in  
110 top-ranked SNPs between two time points.

## 111 **HAploTYPE block Discovery Rate (HADR)**

112 Barghi et al. (2019) clustered candidate SNPs from F60 into selected haplotype blocks based on  
113 similar allele frequency trajectories over time and replicates

114 (Franssen et al. (2016)). The reconstructed haplotype blocks were further validated with ex-  
115 perimentally phased haplotypes from ancestral and evolved populations (Barghi et al. (2019))  
116 and 96 % of the reconstructed haplotype blocks could be confirmed. This suggests that recon-  
117 structed haplotype blocks provide a reliable set of linked candidate SNPs.  
118 Taking advantage of this additional confirmation of the candidate SNPs in a selected haplotype  
119 block we developed a third measure of similarity between time points. We determined the frac-  
120 tion of candidate SNPs comprising a haplotype block that were also discovered at a given time  
121 point (haplotype block discovery rate, HADR) using the poolSeq R-package (Taus et al. (2017)).  
122 We note that inference of selected haplotype blocks at each generation does not provide a good  
123 alternative to HADR, as the ability to cluster SNPs into haplotype blocks is dependent on the  
124 number of time points (Franssen et al. (2016)), resulting in less power at early time points com-  
125 pared to later ones.

## 126 **Early Detected HApIotype blocks (EDHAs)**

127 We applied hierarchical clustering (Pollard and Laan (2005)), PCA and kmeans (Hartigan and  
128 Wong (1979)) to group haplotype blocks based on their HADR patterns. The hyper-parameter  
129  $k$ , which determines the number of clusters, was set to 5 based on the gap statistic approach  
130 (Tibshirani et al. (2001)). The k-means clustering resulted into a group of 10 haplotype blocks  
131 with elevated HADR in generation 20 (Figure S2). This group of 10 haplotype blocks can also  
132 be separated from other haplotype blocks by the first principal component of a PCA applied  
133 to HADR from F10 to F50 (Figure S3). We refer to the haplotype blocks in this cluster as early  
134 detected haplotype blocks (EDHAs). We evaluated whether EDHAs have distinct characteristics  
135 compared to all other haplotype blocks using the following features: haplotype block length,  
136 median starting allele frequency, average recombination rate (*D. simulans* recombination map

137 from Howie et al. (2019)), selection coefficient ( $s$ , estimated with poolSeq (Taus et al. (2017)))  
138 in generation 20,  $s$  in generation 60, selection coefficient ratio  $r_s = \frac{s_{20}}{s_{60}}$ , and number of rising  
139 replicates in generation 20. Following Barghi et al. (2019), we classified a haplotype block as  
140 replicate specific, if the allele frequency of candidate SNPs from a haplotype block increases on  
141 average by at least 10%. We also used AFC thresholds of 5, 15, and 20% to determine whether  
142 these haplotype blocks were rising in a given replicate. Selection coefficients were averaged  
143 (mean) over replicates that passed the AFC threshold.

## 144 **RESULTS & DISCUSSION**

### 145 **Subsequent time points are more similar for advanced generations**

146 We studied the similarity of selection signatures for different time points using 10 replicates of  
147 a *D. simulans* population, which evolved for 60 generations to a novel hot environment (Barghi  
148 et al. (2019)). With Pool-Seq data from every 10<sup>th</sup> generation, we evaluated the selection sig-  
149 nature on three different levels: candidate SNPs, candidate SNPs in a window of fixed length  
150 and candidate SNPs shared with reconstructed selected haplotype blocks. The similarity of two  
151 time points was determined by the Jaccard index, a dimensionless parameter ranging from 0  
152 (no overlap between two sets) to 1 (sets are identical). We found that all candidate SNP sets are  
153 more similar than expected by chance. The Jaccard index ranged from 0.08 (generation 10 vs  
154 generation 60) to 0.40 (generation 50 vs generation 60), where subsequent time points are more  
155 similar than those separated for more than 10 generations (e.g.  $J=0.15$  (generation 10 vs gen-  
156 eration 20);  $J=0.08$  (generation 10 vs generation 60)). Furthermore, the similarity of candidate  
157 SNP sets from subsequent time points increases with time until it ultimately more than doubles  
158 for the last two generations ( $J=0.15$  (generation 10 vs generation 20);  $J=0.34$  (generation 50 vs

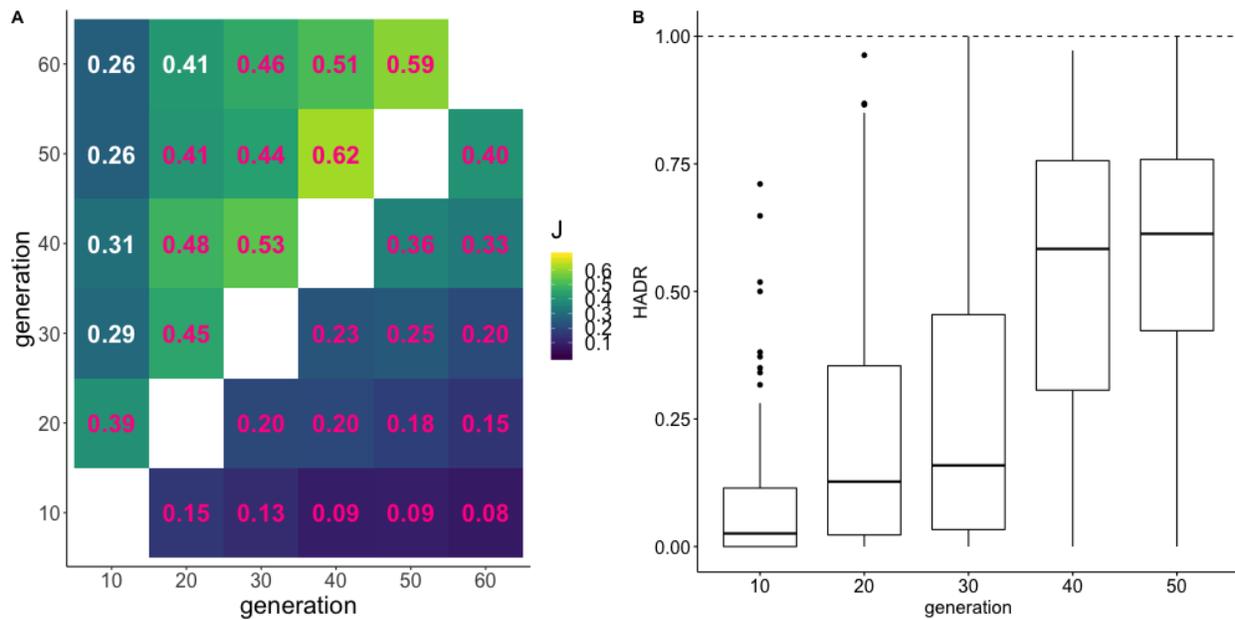
159 generation 60), Figure 1A). The monotonic increase in similarity with time shows that selection  
160 patterns are more reliably detected at later generations.

161 Since the analysis of single SNPs suffers from considerable stochasticity, and neighboring SNPs  
162 are not independent (Tobler et al. (2014); Howie et al. (2019)), we also repeated the analysis of  
163 different time points using non-overlapping windows of 5,000 SNPs. Reasoning that windows  
164 containing a target of selection will harbor multiple candidate SNPs, we defined selected win-  
165 dows as those, which harbor more candidate SNPs than expected by chance. Consistent with  
166 higher stochasticity on the SNP level, a higher similarity was observed for candidate windows  
167 (from  $J=0.26$  (generation 10 vs generation 60) to  $J=0.62$  (generation 40 vs generation 50)). Again,  
168 adjacent time points have a higher Jaccard index than time points farther apart ( $J=0.26$  (genera-  
169 tion 10 vs generation 60)  $J=0.39$  (generation 10 vs generation 20)). The similarity of subsequent  
170 time points also increases with the duration of the experiment ( $J=0.39$  (generation 10 vs gen-  
171 eration 20) ;  $J=0.59$  (generation 50 vs generation 60), Figure 1A). In contrast to the SNP level,  
172 the set of selected windows after 10 generations is only significantly similar to generation 20,  
173 but not to any other generation. Thus, the pattern of less reliable selection targets in the early  
174 generations is confirmed on the window level, albeit with different significance levels.

175 For an alternative measure of similarity we used the ranking of candidate SNPs in a specific win-  
176 dow based on their  $p$ -values and compared it between different time points. If a signal is robust  
177 between two time points, we expect the same SNP ranking in a selected window. Consistent  
178 with the other tests, we found that the congruence in candidate SNP ranking increases with  
179 time (Figure 2). To rule out that rare SNPs are responsible for the dissimilarity between early  
180 and late time points, we calculated similarity measures based on SNPs that are segregating at  
181 all generations and time points. Nevertheless, including SNPs which were lost in at least one  
182 replicate during the experiment did not result in a pronounced decrease in similarity (Figure

183 S4).

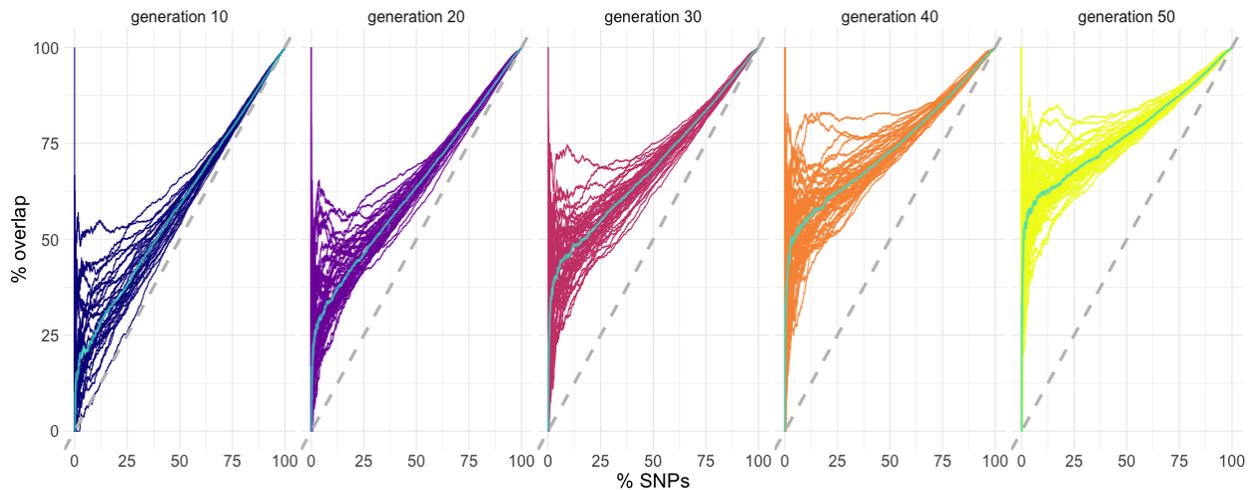
184 The analysis of selected haplotype blocks provides another possibility to control for non-independence  
185 of single candidate SNPs. We calculated the haplotype block discovery rate (HADR) - the frac-  
186 tion of candidate SNPs in a haplotype block that are rediscovered at a given time point. Similar  
187 to the other analyses, we observe higher similarity between later time points (Figure 1B), with a  
188 pronounced increase of median HADR between generation 30 (< 25 %) and 40 (> 50 %).  
189 Independent of the measure of similarity between time points, we consistently find that selec-  
190 tion signatures at early time points are less reliable than those from later time points. With the  
191 limitation that the true targets of selection are not known, this observation highlights that a  
192 more reliable identification of selection targets strongly benefits from additional generations of  
193 selection. Given that fewer "real" selection targets were identified in the first time points, it is  
194 remarkable that a similar number of candidate SNPs was detected at each time point (Table S4).  
195 This may imply that earlier time points harbor more false positives, but it is also possible that  
196 these targets were only selected during the first generations. To distinguish between these two  
197 alternative explanations, the analysis of haplotypes will be required to separate linked hitchhik-  
198 ers from selected sites. Furthermore, experimental validation of selection targets in secondary  
199 E&R studies (Burny et al. (2019)) may be another route to confirm selection signatures beyond  
200 statistical testing.



**Figure 1:** Similarity measures for candidate SNPs, candidate SNPs in a window of fixed length and candidate SNPs shared with reconstructed selected haplotype blocks

**(A):** Jaccard index for pairwise comparisons of candidate sets. The top triangle shows candidate window sets, the bottom triangle candidate SNP sets. Significant similarities ( $p$ -value  $< 0.05$  after multiple testing correction, 10,000 bootstraps) are marked in pink.

**B:** The rate at which selected SNPs of 93 haplotype blocks from generation 60 were already discovered at earlier generations (HADR).



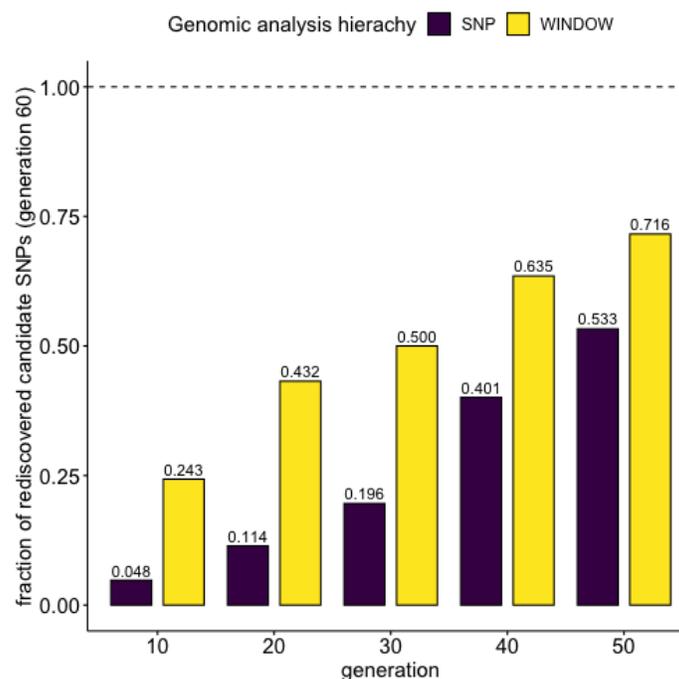
**Figure 2:** The rank of candidate SNPs becomes more congruent with time. In this ROC-like graph the ranking of all candidate SNPs in candidate windows is compared. Each panel shows one intermediate time point compared to generation 60. The percent overlap for each candidate window is indicated by a separate line. The median overlap (turquoise line) monotonically increases with experimental duration, demonstrating that the ranking of candidate SNPs is more robust for advanced generations. The expected overlap is shown as dashed, grey line.

## 201 **Only few selection targets are shared across all generations**

202 More than 27,000 candidate SNPs can be identified at each time point (Table S4), but only a  
203 small (4.8 %) subset is consistently detected at every generation (Figure 3; including rare SNPs  
204 see Figure S5). Independent of the importance of more reliable selection signatures with an in-  
205 creasing number of generations, this analysis raises an important concern about the usefulness  
206 of meta-analyses on the SNP level. With less than 5% of the SNPs being shared in the same se-  
207 lection experiment, it will be extremely difficult to compare studies that started from different  
208 founder populations and were selected for a different number of generations.

209 We repeated the analysis for windows and determined the number of selected windows that  
210 are shared across all generations. With 18 out of 74 candidate windows in generation 60 (24.3  
211 %, Figure 3) being detected at all generations, the window analysis shows more consistency

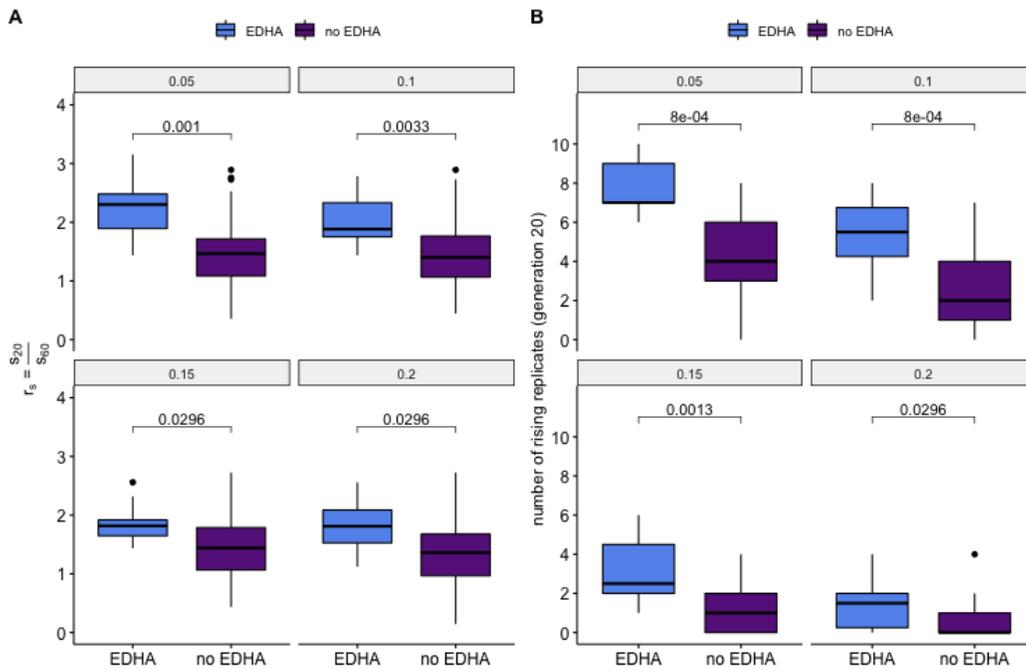
212 across time points than a SNP-based analysis. This observation is independent of window size  
213 (Table S5- S6) and the inclusion of rare SNPs into the analysis (Figure S5). We propose that  
214 meta-analyses of E&R data should be performed on the level of windows, or probably based on  
215 selected haplotype blocks to avoid false negatives due to the high stochasticity of SNP-based  
216 analyses.



**Figure 3:** Less than 5% of candidate SNPs in generation 60 are detected consistently at every generation. The bars depict the fraction candidate SNPs (purple) and candidate windows (yellow) at generation 60, which are candidates in all subsequent generations (e.g. 40.1% of generation 60 candidate SNPs are candidates in generation 50 and 40). Candidate windows are more consistent than candidate SNPs. Figure S5 depicts the ratios for candidate sets that are not restricted to SNPs segregating in all generations and time points.

217 **Selection signatures detected early in the experiment are not representative**  
218 **of the underlying adaptive architecture**

219 This study focused on the comparison of selection targets detected at early and late time points.  
220 Since analyses based on single SNP are very stochastic, we investigated the fraction of candidate  
221 SNPs comprising a haplotype block that were also discovered at earlier time points (HADR). We  
222 detected 10 haplotype blocks with elevated HADR in generation 20 (EDHAs, Figure S3). 10 ED-  
223 HAs were detected based on kmeans clustering (see Material & Methods and Figures S2-S3). We  
224 found that EDHAs do not differ in their starting allele frequency, haplotype block length, aver-  
225 age recombination rate or absolute selection coefficients from other haplotype blocks (Figure  
226 S6). EDHAs are, however, more strongly selected at the beginning of the experiment, but do not  
227 differ from the remaining haplotype blocks at later generations. The comparison of the rela-  
228 tive selection intensity of early and late time points identified significant differences of EDHAs  
229 from the other haplotype blocks (Figure 4A). Consistent with stronger selection at earlier time  
230 points, the selection signature of EDHAs is significantly more parallel across replicates after 20  
231 generations of adaptation. (Figure 4B). All statistical tests, which are evaluating a parallel selec-  
232 tion signature across replicates, are more likely to detect selection signatures, which are shared  
233 across replicates, even with only moderate allele frequency changes. This could result in a bi-  
234 ased picture of the underlying genetic architecture. The analysis of selection signatures in repli-  
235 cated experiments running for only a moderate number of generations is more likely to detect  
236 parallel than replicate specific selection signatures. This bias is not restricted to our study, but  
237 also an experimental study of *D. simulans* populations adapting 10 to 20 generations to a new  
238 temperature regime (Kelly and Hughes (2018)) found more parallel selection responses. We pro-  
239 pose that additional analyses contrasting selection signatures of early and late time points are  
240 needed to confirm the enrichment of parallel selection signatures in short-term experiments.



**Figure 4:** Early Detectable HAploTYPE blocks (EDHAs) differ from the other selected haploTYPE blocks. The ratio of selection coefficients determined for early generations (generation 20) and late generations (generation 60) is significantly higher for EDHAs (A). EDHAs rise in more replicates (B) than other haploTYPE blocks. Both observations are robust to different AFC thresholds. Values above the boxplots represent the two-tailed Mann-Whitney test  $p$ -values corrected for multiple testing with the Benjamini-Hochberg procedure.

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<sup>246</sup> **DATA AVAILABILITY**

<sup>247</sup> All scripts will be uploaded and available on GitHub upon publication.

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