

1 **A healthy childhood environment helps to combat inherited susceptibility to obesity**

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45

46 **Abstract**

47 Objectives: To investigate the degree by which the inherited susceptibility to obesity is modified
48 by environmental factors during childhood and adolescence.

49 Design: Cohort study with repeated measurements of diet, lifestyle factors and anthropometry.

50 Setting: The pan-European IDEFICS/I.Family cohort

51 Participants: 8,609 repeated observations from 3,098 children aged 2 to 16 years, examined
52 between 2007 and 2014.

53 Main outcome measures: Body mass index (BMI) and waist circumference. Genome-wide
54 polygenic risk scores (PRS) to capture the inherited susceptibility of obesity were calculated
55 using summary statistics from independent genome-wide association studies of BMI. Gene-
56 environment interactions of the PRS with sociodemographic (European region, socioeconomic
57 status) and lifestyle factors (diet, screen time, physical activity) were estimated.

58 Results: The PRS was strongly associated with BMI ($r^2 = 0.11$, $p\text{-value} = 7.9 \times 10^{-81}$) and waist
59 circumference ($r^2 = 0.09$, $p\text{-value} = 1.8 \times 10^{-71}$) in our cohort. The associations with BMI
60 increased from $r^2=0.03$ in 3-year olds to $r^2=0.18$ in 14-year olds and associations with waist
61 circumference from $r^2=0.03$ to $r^2=0.14$. Being in the top decile of the PRS distribution was
62 associated with 3.63 times higher odds for obesity (95% confidence interval (CI): [2.57, 5.14]).
63 We observed significant interactions with demographic and lifestyle factors for BMI as well as
64 waist circumference. The risk of becoming obese among those with higher genetic
65 susceptibility was ~38% higher in children from Southern Europe (BMI: $p\text{-interaction} = 0.0066$,
66 Central vs. Southern Europe) and ~61% higher in children with a low parental education (BMI:
67 $p\text{-interaction} = 0.0012$, low vs. high). Furthermore, the risk was attenuated by a higher intake
68 of dietary fiber (BMI: $p\text{-interaction}=0.0082$) and shorter screen times (BMI: $p\text{-}$
69 $\text{interaction}=0.018$).

70 Conclusions: Our results highlight that a healthy childhood environment might partly offset a
71 genetic predisposition to obesity during childhood and adolescence.

72

73

74 **Key words:** Adolescents, BMI, children, gene-environment interaction, nutrition, physical
75 activity, screen time, socio-demographic factors, polygenic risk score, waist circumference

76 **Introduction**

77 Obesity is a complex multifaceted condition and its prevalence has been increasing
78 continuously over previous decades and has reached a high plateau in Western countries [1].
79 In 2015, a total of 107.7 million children and 603.7 million adults were obese. Although the
80 prevalence of obesity among children has been lower than that among adults, the rate of
81 increase in childhood obesity has been greater than the rate of increase in adult obesity, which
82 is most likely due to adverse changes of environmental and demographic factors with a direct
83 impact on children's health [2].

84 With the advent of genome-wide association studies (GWAS), it was shown that multiple
85 genetic loci increase the susceptibility to obesity [3,4]. However, genome-wide significant
86 variants identified in the first large-scale GWAS on body mass index (BMI) only account for a
87 small portion of BMI variation (~2.7%) [3]. A more recent genome-wide meta-analysis extended
88 the number of individuals from ~300,000 [3] to ~700,000 [4], which consequently increased
89 the number of genome-wide significant SNPs from 97 to 751. Even these 751 genome-wide
90 significant SNPs account for only ~6.0% of the variance of BMI [4]. However, genome-wide
91 estimates suggest that common variation accounts for >20% of BMI variation [3], which
92 highlights the polygenic architecture of BMI. More recently, whole genome data even increased
93 the fraction of variance of BMI accounted for by genetic variants, both common and rare, to
94 40% [5]. From twin studies we know that the heritability of BMI also depends on socioeconomic
95 status [6] and physical activity [7], suggesting that when socioeconomic status or physical
96 activity is high, genetic factors become less influential. Using candidate SNPs - either single
97 genotypes or <100 SNPs combined in a polygenic risk score (PRS), which is defined as a
98 weighted sum of BMI-related risk alleles - it was further shown that the genetic predisposition
99 to obesity is attenuated by a healthy lifestyle including physical activity [8,9] and adherence to
100 healthy dietary patterns [9–15]. However, most previous gene-environment (GxE) interaction
101 studies primarily involved adults [8–15] or used only a candidate SNP [16], so that it is unknown
102 whether the inherited susceptibility to obesity is modified by environmental factors already

103 during childhood and adolescence. Another limitation of previous gene-environment
104 interaction analyses is that they were based on <100 SNPs that reached genome-wide
105 significance in previous GWAS on BMI [3], which do not capture the whole polygenic risk profile
106 of obesity due to their low heritability. Khera et al. suggested that the power to predict BMI by
107 PRS can be improved by using lower p-value thresholds or even genome-wide approaches
108 [17]. Using a genome-wide polygenic risk score based on effect estimates from [3], Khera et
109 al. reported that the PRS-effect on weight and BMI z-scores emerges early in life and increases
110 until adulthood and that a high PRS is a strong risk factor for severe obesity and associated
111 diseases [17]. The authors suggested that given that the weight trajectories of individuals in
112 different PRS deciles start to diverge early in childhood, targeted strategies for obesity
113 prevention may have maximal effect when employed early in life. However, because lifestyle
114 factors were not considered in their study, it is not known to which degree the genetic
115 predisposition to obesity is modifiable by a healthy lifestyle early in life. Another limitation of
116 [17] is the use of weight and BMI as only proxies for obesity. Since several studies have shown
117 that classifying obesity using BMI alone misses an increasing proportion of individuals
118 categorized as obese [18,19], it is important to test the performance of BMI-PRS for the
119 prediction of waist circumference, which is proposed to be a better proxy for obesity-associated
120 metabolic abnormalities [20].

121 In this study, 1) we show the prediction capacity of the PRS proposed in [17] for BMI as well
122 as for waist circumference of European children and adolescents and 2) analyze its interaction
123 with parental education, region of residence, selected dietary variables and physical activity to
124 investigate to which degree the inherited susceptibility to obesity in children is modified by
125 these sociodemographic and lifestyle factors. The analyses are based on 8,609 repeated
126 observations from 3,098 children and adolescents aged 2 to 16 years from the pan-European
127 IDEFICS/I.Family cohort.

128

129 **Methods**

130 *Study Population*

131 The pan-European IDEFICS/I.Family cohort [21,22] is a multi-center, prospective study on the
132 association of social, environmental and behavioral factors with children's health status.
133 Children were recruited through kindergarten or school settings in Belgium, Cyprus, Estonia,
134 Germany, Hungary, Italy, Spain and Sweden. In 2007/2008, 16,229 children aged between 2
135 and 9.9 years participated in the baseline survey. Follow-up surveys were conducted after two
136 (FU1, N = 11,043 plus 2,543 newcomers) and six years (FU2, N = 7,117 plus 2,512 newly
137 recruited siblings). Physical examinations covered a broad spectrum of parameters according
138 to a detailed and standardized study protocol. Questionnaires were completed by parents for
139 children younger than 12 years. In the second follow-up (FU2), adolescents of 12 years of age
140 or older reported for themselves. All questionnaires were developed in English and translated
141 into local languages. The quality of translations was checked by back translation into English.
142 The study was conducted in agreement with the Declaration of Helsinki; all procedures were
143 approved by the local ethics committees and written and oral informed consents were obtained
144 from the parents, their children and adolescents, respectively, as applicable. Children were
145 selected for a whole-genome scan based on their participation in the individual study modules.
146 Children from Cyprus were not included in this initial genotyping to minimize population
147 stratification.

148

149 *Assessment of BMI and Waist Circumference*

150 BMI was calculated as weight divided by height squared [kg/m^2]. Height was measured to the
151 nearest 0.1 cm by a SECA 225 Stadiometer (Seca GmbH & Co. KG., Hamburg, Germany) and
152 body weight was measured in fasting state in light underwear on a calibrated scale accurate
153 to 0.1 kg by a Tanita BC 420 SMA scale (TANITA, Tokyo, Japan). Waist circumference was

154 measured in upright position with relaxed abdomen and feet together using an inelastic tape
155 (Seca 200, Birmingham, UK), precision 0.1 cm, midway between the iliac crest and the lowest
156 rib margin to the nearest 0.1 cm [23]. Age- and sex-specific BMI and waist circumference z-
157 scores for children and adolescents were calculated using reference data from the
158 International Obesity Task Force [24] and from British children [25], respectively.

159

160 *Genotyping and Quality Control*

161 DNA was extracted from saliva or blood samples using established procedures. Genotyping
162 of 3,515 children was performed on the UK Biobank Axiom array (Santa Clara, USA) in two
163 batches (2015 and 2017). Following the recommendations of [26], sample and genotype
164 quality control measures were applied (see supplementary materials for details), resulting in
165 3,099 children and 3,424,677 genotypes after imputation. A genetic relatedness matrix was
166 calculated to account for the degree of relatedness within the study sample and to adjust for
167 population stratification [27,28] by using the program EMMAX
168 (<https://genome.sph.umich.edu/wiki/EMMAX>).

169

170 *Polygenic Risk Score Calculation*

171 We calculated PRS based on genome-wide summary statistics for BMI from European
172 ancestry populations. The PRS (called PRS-Khera) was proposed in [17]. It consists of
173 2,100,302 SNPs and is based on summary statistics from the first large-scale GWAS of BMI
174 (~300,000 samples) [3]. PRS-Khera was calculated in [17] using a computational algorithm
175 called LDpred, which is a Bayesian approach to calculate a posterior mean effect for all
176 variants using external weights with subsequent shrinkage based on linkage disequilibrium
177 [29]. Using LDpred, each variant was reweighted according to the prior GWAS [3], the degree

178 of correlation between a variant and others nearby, and a tuning parameter that denotes the
179 proportion of variants with non-zero effect.

180 In sensitivity analyses, the performance of PRS-Khera was compared to the PRS calculated
181 with PRSice [30] and the PRS based on only genome-wide significant SNPs from two
182 reference populations (same reference population as for PRS-Khera (~300,000 samples) [3]
183 and the largest published GWAS study of BMI to date (~700,000 samples) [4]). More details
184 on the different PRS are given in the supplementary methods and Figures S1 to S3.

185

186 *Assessment of Dietary Intake*

187 We used long-term and short-term dietary measurements assessed by food frequency
188 questionnaires (FFQs) and repeated 24 hour dietary recalls, respectively [31]. A fruit and
189 vegetable score was calculated from FFQs (for more details on the FFQs and calculation of
190 the fruit and vegetable score, see supplemental material). We expressed the fruit and
191 vegetable consumption as the relative frequency in relation to all foods reported in the FFQs
192 [32]. The FFQs were self-reported by adolescents 12 years and older and proxy-reported by a
193 parent or other caregiver for children below the age of 12 years.

194 Energy and dietary fiber intake were assessed by repeated 24 hour dietary recalls [33,34].
195 Usual intakes for fiber were estimated based on the validated National Cancer Institute (NCI)
196 method, which is one of the most widely accepted methods for this purpose [35,36]. This
197 method allows for the inclusion of covariates such as age and accounts for different intakes on
198 weekend days vs. weekdays, and further corrects for the day-to-day variation in energy and
199 fiber intakes. Usual intakes were estimated for each child stratified by sex and considering age
200 as a covariate. Fiber intake was here expressed in relation to total energy intake in mg/kcal.
201 See supplemental material for more details.

202

203 *Assessment of Physical Activity*

204 Physical activity was objectively measured by using Actigraph's uniaxial or three-axial
205 accelerometers [37,38]. At baseline and FU1, children were asked to wear the accelerometer
206 for three days (including one weekend day) and at FU2 for a full week during waking hours
207 (except when swimming or showering). The accelerometers were attached to the right hip with
208 an elastic belt. Participants (either the parents or the adolescents themselves) were given
209 written instructions on how to use the accelerometer and were asked to complete diaries to
210 record non-wear times of the device. The daily average cumulative duration of time spent in
211 moderate-to-vigorous physical activity (MVPA) was expressed as minutes per day according
212 to previously defined cut-off values [39]. Especially for children, accelerometer measurements
213 are far less prone to measurement errors than self-reported activities through questionnaires
214 [40,41]. See supplementary material for more details.

215

216 *Assessment of Screen Time*

217 Screen time was assessed by asking how many hours per day the child/adolescent usually
218 spends watching television (including videos or DVDs) and by another question on the time
219 sitting in front of a computer and game console [42,43]. Responses were weighted and
220 summed across weekdays and weekend days and the quantified frequencies from both
221 questions were added to create a continuous variable of total screen time in hours per day.
222 Parents reported for children younger than 12 years, while older children (≥ 12 years) reported
223 for themselves. See supplemental material for more details.

224

225 *Assessment of Sociodemographic Variables*

226 Parental education was retrieved from questionnaires and coded according to the International
227 Standard Classification of Education (ISCED) [44]. For the analyses, the highest parental
228 education of both parents was coded as low (ISCED levels 1 and 2; ≤ 9 years of education),

229 medium (ISCED levels 3 and 4) and high (ISCED levels 5 and 6; ≥ 2 years of education after
230 high school). The region of residence was coded as Northern Europe (Estonia, Sweden),
231 Central Europe (Belgium, Germany, and Hungary) and Southern Europe (Italy, Spain).

232

233 *Statistical Analyses*

234 Our data consist of up to three repeated measurements of individuals, some of which were
235 siblings. We used generalized linear mixed models where the covariance matrix of the random
236 intercept is proportional to a genetic relatedness matrix. We applied the generalized linear
237 mixed model approach of Chen et al. [27] that jointly controls for relatedness and population
238 stratification. All models were adjusted for sex, age, region of residence and parental
239 education. All models that did not include fiber intake were additionally adjusted for the
240 vegetable score. When testing associations with categorical variables (sex, region of residence
241 and parental education), we used the category with the largest sample size as reference
242 category.

243 All p-values from the gene-environment interaction analyses were adjusted according to the
244 number of tested environmental factors using the false-discovery rate (FDR). We reported 95%
245 confidence intervals (95% CI) and two-sided p-values, and considered p-values less than 0.05
246 statistically significant. We used R 3.5.1 [45] for all statistical analyses.

247

248 **Results**

249 The study sample included 8,609 repeated BMI measurements from at maximum three time
250 points (baseline, FU1, FU2) of 3,098 children aged 2 to 16 years (Table 1). The number of
251 participants decreased only slightly between the follow-up investigations from $n = 3,016$ at
252 baseline (mean age 6 years) to $n = 2,656$ at FU2 (mean age 12 years). Half of the children

253 were girls, most children came from families with a medium or high level of education and the
254 majority lived in Central European countries. The distributions of the dietary variables
255 (vegetable score and fiber intake) and time spent in MVPA were similar between baseline and
256 the two follow-up samples, whereas children and adolescents spent more time in front of
257 screens at FU1 and FU2 as compared to baseline. On average, BMI and waist circumference
258 of our analysis group were higher than in the reference populations [24,25] (mean z-scores >
259 0).

260 We found that the PRS-Khera provided the best prediction of BMI (see Table S1 for details on
261 the characteristics of the other PRS). PRS-Khera was strongly associated with BMI ($r^2 = 0.11$,
262 $p\text{-value} = 7.9 \times 10^{-81}$) and waist circumference ($r^2 = 0.09$, 1.8×10^{-71}) in our study population
263 (Table 2). Being in the top decile of the distribution of PRS-Khera was associated with 3.63
264 times higher odds for obesity (95% CI: [2.57, 5.14]) and with 3.09 (95% CI: [2.37, 4.03]) higher
265 odds for being in the top quartile of waist circumference.

266 The correlation between PRS-Khera and BMI increased along the age range, from a squared
267 correlation with BMI of $r^2 = 0.02$ [0.01, 0.12] in the 2-year olds to $r^2 = 0.18$ [0.11, 0.27] in the
268 14-year olds (Figure 1 and Table S2). Similar trends were found for waist circumference, for
269 which the squared correlation with PRS-Khera was $r^2 = 0.03$ [0.01, 0.07] in 3-year olds and r^2
270 $= 0.14$ [0.08, 0.22] in 14-year olds (Figure 1 and Table S2). This increase of correlation by age
271 group was confirmed in our sensitivity analyses using other genome-wide PRS (Figure S4 and
272 Table S3).

273 We found a significant gene-environment interaction of PRS-Khera with parental education
274 (low vs. high) as well as with the European region of residence (Central vs. Southern) for BMI
275 as well as for waist circumference (Figure 2, Tables S4). Children and adolescents from
276 families with a low level of education were at a higher risk of becoming obese among those
277 with higher genetic susceptibility than children from families with a high level of education (low:
278 beta estimate from education-stratified analysis for association between PRS-Khera and BMI

279 = 0.48; 95% CI: [0.38, 0.59], high: beta estimate = 0.30; 95% CI: [0.26, 0.34], adjusted p-value
280 interaction = 0.0106, Figure 2 and Table S4). Furthermore, children and adolescents from
281 Southern European countries showed an increased genetic susceptibility to a high BMI in
282 comparison to children and adolescents from Central Europe (Central Europeans: beta
283 estimate from region-stratified analysis for association between PRS-Khera and BMI = 0.29;
284 95% CI: [0.23, 0.34], Southern Europeans: beta estimate = 0.40; 95% CI: [0.34, 0.45], adjusted
285 p-value interaction = 0.0246, Figure 2 and Table S4). Interactions were confirmed in our
286 sensitivity analyses using other genome-wide PRS (Figure S5). We did not find significant
287 interactions between PRS-Khera and sex, the comparison of low vs. medium parental
288 education, nor the comparison of Central vs. Northern European region of residence (Figure
289 2, Table S4).

290 The genetic susceptibility to a high BMI was further modified by intake of dietary fiber and
291 screen time (Figure 3, Tables S4). Children and adolescents with a higher fiber intake showed
292 an attenuated risk of becoming obese despite their genetic susceptibility (adjusted p-values
293 for interaction: 0.025 for BMI and 0.023 for waist circumference). Furthermore, the more time
294 the children and adolescents spent in front of screens, the higher was their risk of becoming
295 obese among those with higher genetic susceptibility (adjusted p-value interaction = 0.042).
296 Interactions between PRS-Khera and the fruit and vegetable score or MVPA were not
297 significant.

298

299 **Discussion**

300 In our pan-European cohort of children aged 2 to 16 years, we found a strong association of a
301 polygenic risk score of obesity with BMI as well as with waist circumference and this
302 association increased by age. We observed a prediction r^2 of 18% in 14-year olds, which is
303 even higher than in the original study containing mainly adults [4]. We further found significant
304 interactions with socioeconomic and behavioral factors for BMI as well as waist circumference:

305 we observed gene-environment interactions with (1) the European region of residence, which
306 most likely reflect cultural lifestyle differences, (2) education, (3) dietary fiber intake and (4) the
307 time children spent in front of screens. Of note, all of these interactions would have remained
308 undetected in this sample of children when only focusing on genome-wide significant variants
309 as was done in previous studies (compare Figures S5 and S6) [8–15].

310

311 *Comparison with Previous Studies*

312 Although obesity is known to be highly polygenic, most previous gene-environment interaction
313 analyses focused on <100 genome-wide significant variants that account for <3% of BMI
314 variation. In this study we used a genome-wide PRS proposed in [17], which provides a more
315 comprehensive measurement of the inherited susceptibility to obesity. Using this PRS (called
316 PRS-Khera), we observed a prediction r^2 of 10.8% for BMI, which is almost 5 times higher than
317 the prediction accuracy obtained using the <100 genome-wide significant SNPs from the
318 ~300,000 samples in [3] and twice the prediction accuracy obtained using the <1,000 genome-
319 wide significant SNPs from the ~700,000 samples in [4] (Table S1). PRS-Khera reached a
320 similar prediction accuracy for BMI than it has been reported from large-scale PRS in previous
321 studies (~10.2% using the summary statistics from the ~700,000 samples and a p-value
322 threshold of 10^{-3} (6,781 SNPs) [4] and ~8.5% [17] using a genome-wide PRS from the
323 ~300,000 samples in [3]).

324 Of note, in our study, the prediction accuracy of the PRS strongly depended on age, reaching
325 a prediction r^2 of 18% in 14-year olds, which is in accordance with Khera et al. who showed
326 that the association between the PRS and weight emerges early in life and increases into
327 adulthood [17]. This surprisingly high prediction accuracy in adolescents from our study might
328 be explained by the age difference between our study and the GIANT Consortium / UK
329 Biobank, which was used in [4]. The GIANT Consortium / UK Biobank included mainly adults,
330 whereas we analyzed data from children aged 2 to 16 years. In contrast to the positive
331 correlation between age and prediction accuracy during childhood shown in this manuscript as

332 well as in previous studies [17,46], a weak negative correlation could be observed in adults
333 >45 years of age from the UK Biobank, an age group in which aging-related diseases become
334 more prevalent (Table S3 in [17]). Therefore, we hypothesize that the highest prediction
335 accuracy of the PRS for BMI might be reached in adolescents and young adults.

336 In our study, we found significant interactions between PRS-Khera and sociodemographic as
337 well as lifestyle factors for BMI and waist circumference. Interactions with socioeconomic
338 status [9], physical activity [8,9], and dietary factors [9–15] have been reported previously.
339 However, all of these studies included only <100 genome-wide significant SNPs (e.g. from [3]).
340 By using a genome-wide PRS we were able to detect interactions with sociodemographic and
341 with lifestyle factors which would have remained undetected when using only genome-wide
342 significant SNPs (Figures S5 and S6).

343 Furthermore, previous GxE interaction studies [8–15] were mainly based on adults whereas in
344 our study we analyzed data from children aged 2 to 16 years. Therefore, our results provide
345 new insights about how a healthy childhood environment might partly offset a genetic
346 predisposition to obesity during childhood and adolescence. In our study, we identified children
347 from families with low levels of education as being about 61% more susceptible to the
348 polygenic burden of obesity than children from families with a high level of education. In
349 addition, we found that children from Southern Europe had a higher genetic susceptibility to
350 obesity in comparison to children from Central Europe. Parental education and region of
351 residence reflect a variety of social and cultural differences and many of them are difficult to
352 capture by questionnaires. Since a previous analysis of the same cohort showed that low
353 parental education was associated with higher intakes of unhealthy food among children, e.g.
354 sugar-rich and fatty foods [47,48], part of the effect modification might be due to dietary habits.
355 The differences in the risk of becoming obese among children with a higher genetic
356 susceptibility across different European regions might be explained by differences in dietary or
357 cultural habits [49,50].

358 Furthermore, we found an interaction between PRS-Khera and intake of fiber, with children
359 with a higher intake of fiber having a reduced risk for obesity despite their genetic susceptibility.
360 This finding is in line with many other studies that have shown that a healthy diet can attenuate
361 the genetic burden of obesity [9–15]. Interactions between PRS-Khera and physical activity
362 (MVPA) were not significant, but the direction of interaction effect was in line with previous
363 studies [8,9]. An explanation for this might be that MVPA was only assessed in ~40% of our
364 analysis group (Table 1), which reduced the statistical power to detect interactions between
365 MVPA and PRS.

366

367 *Strengths and Limitations of this Study*

368 Important strengths of this study include: detailed and repeated phenotyping of participants in
369 this cohort with partly objective measures (MVPA), inclusion of thousands of children from
370 diverse regions in Europe and the longitudinal approach across key developmental periods
371 [22]. Dietary assessment in children is a challenging task [51], and different dietary assessment
372 have different strengths and limitations. We used two different dietary assessment methods –
373 a fruit and vegetable score derived from FFQs and fiber intake calculated from the more
374 detailed 24-hour dietary recalls. The harmonized protocol in all countries that was enforced by
375 a central quality control and a central data management ensures comparability of
376 measurements across study centers. Another major strength of our study is the application of
377 genome-wide PRS for obesity, which has an almost 5 times higher prediction accuracy than
378 previously used PRS [9–15] and with which we identified interactions that would have
379 remained undetected when only focusing on genome-wide significant variants (compare
380 Figures S5 and S6).

381 Our study also has several limitations. First, measurement errors of self-reported lifestyle
382 behaviors are inevitable. However, measurement error in environmental exposure typically
383 biases the interaction effect toward the null [52], which does not increase the risk for false

384 positive findings but reduces the statistical power to detect subtle interactions. Second, the
385 use of PRS derived from associations with BMI in the analyses of waist circumference led to
386 slightly lower prediction accuracy for waist circumference than for BMI. However, since PRS-
387 Khera is known to be a strong risk factor for severe obesity and associated health outcomes
388 [17], we decided to use this PRS for both obesity measurements.

389

390 *Conclusions*

391 Our study showed significant interactions between the polygenic risk for an increased BMI and
392 sociodemographic and behavioral factors that affect BMI as well as waist circumference.
393 Among children with a high genetic risk, we identified children from Southern Europe, children
394 from families with a low level of education, children with a low intake of fiber and children who
395 spend more time in front of screens as being particularly susceptible to obesity. These results
396 provide evidence that the risk for obesity among children with a high genetic susceptibility
397 varies by environmental and sociodemographic factors during childhood. This has important
398 implications for future public health prevention efforts, because it suggests that children at a
399 high genetic risk may benefit even more from prevention measures than children with a low
400 genetic risk.

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406 **Competing financial interests declaration**

407 The authors have nothing to declare.

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Tables

Table 1. Study characteristics of the 8,609 repeated observations from 3,098 children.

	Baseline	First follow-up (FU1)	Second follow-up (FU2)
n	3016	2937	2656
Age (years)			
Mean (SD)	6.19 (1.77)	8.12 (1.80)	11.75 (1.83)
Median (IQR)	6.60 (3.10)	8.50 (3.20)	11.90 (3.20)
Range	2.0-9.7	3.4-11.9	6.6-16.2
Sex			
Female (%)	1510 (50.07)	1472 (50.12)	1331 (50.11)
Male (%)	1506 (49.93)	1465 (49.88)	1325 (49.89)
Parental education			
Low (%)	180 (5.97)	166 (5.65)	156 (5.87)
Medium (%)	1337 (44.33)	1204 (40.99)	1172 (44.13)
High (%)	1463 (48.51)	1476 (50.26)	1310 (49.32)
European region of residence			
Central (%)	1250 (41.45)	1218 (41.47)	1114 (41.94)
North (%)	743 (24.64)	721 (24.55)	682 (25.68)
South (%)	1023 (33.92)	998 (33.98)	860 (32.38)
Fruit and vegetable score (%)			
Mean (SD)	1.47 (0.75)	1.54 (0.80)	1.47 (0.78)
Median (IQR)	1.38 (0.96)	1.47 (1.02)	1.37 (0.97)
Range	0.00-5.71	0.00-5.83	0.00-6.07
Missing	58	154	106
Fiber intake (mg/kcal)			
Mean (SD)	8.17 (1.31)	8.23 (0.90)	8.22 (1.27)
Median (IQR)	8.13 (1.79)	8.13 (1.48)	8.07 (1.61)
Range	3.87-15.76	5.76-11.56	4.74-13.89
Missing	826	1100	660
MVPA (hours/day)			
Mean (SD)	0.67 (0.36)	0.67 (0.36)	0.64 (0.37)
Median (IQR)	0.61 (0.46)	0.62 (0.47)	0.57 (0.47)
Range	0.02-2.29	0.03-2.74	0.00-2.42
Missing	1240	1297	871
Screen time (hours/day)			
Mean (SD)	1.60 (1.00)	1.89 (1.08)	2.34 (1.50)
Median (IQR)	1.50 (1.07)	1.75 (1.43)	2.02 (1.79)
Range	0.00-8.00	0.00-8.00	0.00-8.00
Missing	93	132	150
BMI z-scores			
Mean (SD)	0.34 (1.16)	0.41 (1.18)	0.51 (1.12)
Median (IQR)	0.23 (1.48)	0.32 (1.67)	0.45 (1.62)
Range	-5.42-5.80	-5.76-4.65	-2.96-3.83
Obese	204	214	179
Waist circumference z-scores			
Mean (SD)	0.24 (1.45)	0.59 (1.29)	0.78 (1.25)
Median (IQR)	0.16 (1.61)	0.46 (1.72)	0.71 (1.77)
Range	-27.98-5.65	-6.79-5.33	-7.75-4.38
Top quartile	461	443	316
Missing	76	22	55

Z-scores for BMI and waist circumference were calculated according to [24,25]. Boys with a BMI z-score > 2.29 and girls with a BMI z-score > 2.19 were defined as obese [24,25].

Table 2. Associations of PRS-Khera with BMI, obesity and waist circumference in IDEFICS/I.Family.

A) BMI						
Scale of PRS	BMI			Obesity		
	Est., 95% CI	p-value	R²	OR, 95% CI	p-value	AUC
Continuous	0.33 [0.30, 0.37]	7.9e-81	0.108	2.33 [2.01, 2.70]	2.0e-29	0.736
Top decile	0.61 [0.49, 0.73]	5.4e-24	0.036	3.63 [2.57, 5.14]	2.7e-13	0.598

B) Waist circumference						
Scale of PRS	Waist circumference			Waist top quartile		
	Est., 95% CI	p-value	R²	OR, 95% CI	p-value	AUC
Continuous	0.36 [0.32, 0.40]	1.8e-71	0.088	1.97 [1.78,2.17]	1.5e-40	0.683
Top decile	0.69 [0.55, 0.82]	8.8e-24	0.032	3.09 [2.37,4.03]	6.1e-17	0.569

Associations adjusted for region of residence, sex, age, parental education, vegetable score. Z-scores for BMI and waist circumference were calculated according to [24,25]. Boys with a BMI z-score > 2.29 and girls with a BMI z-score > 2.19 were defined as obese [24,25].

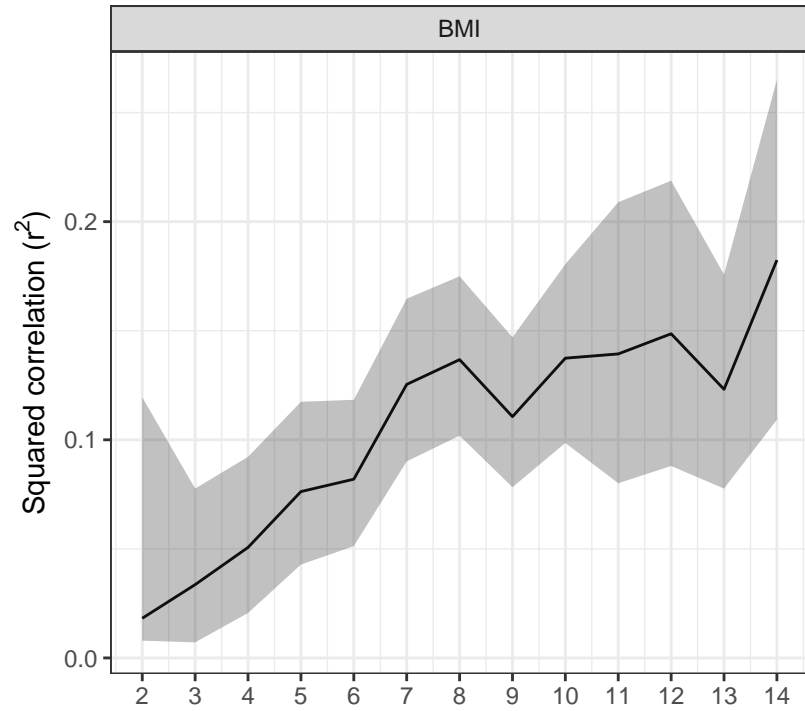
Figure Legends

Figure 1. Squared correlation (r^2 with 95% confidence intervals) of PRS-Khera with BMI and waist circumference in dependence of age. Squared correlations could not be calculated for ≥ 15 -year old children due to the small sample size in these age groups (see Tables S1 & S2). Waist circumference was not measured in 2-year old children.

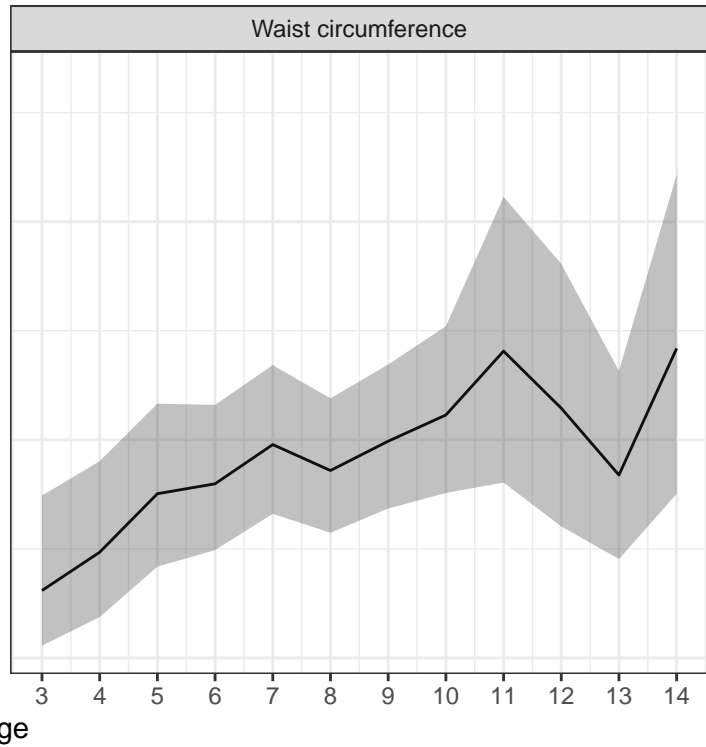
Figure 2. Interactions between PRS-Khera and sociodemographic factors on BMI and waist circumference. Associations between PRS and BMI / waist circumference are shown in different strata (beta estimates and 95% CIs) as well as in the whole study population (**red** line). Raw p-values (p_r) and FDR-adjusted p-values (p_a) are given for the test of deviations of the association between PRS and obesity in one subgroup in comparison to the reference category (interaction). The category without p-values is the reference category.

Figure 3. Interactions between PRS-Khera and lifestyle factors on BMI and waist circumference. Associations between PRS and obesity are shown in dependence of the PRS (beta estimates and 95% CIs) as well as in the whole study population (**red** line). The distributions of the lifestyle factors are shown in histograms. Raw p-values (p_r) and FDR-adjusted p-values (p_a) are given for the interaction terms.

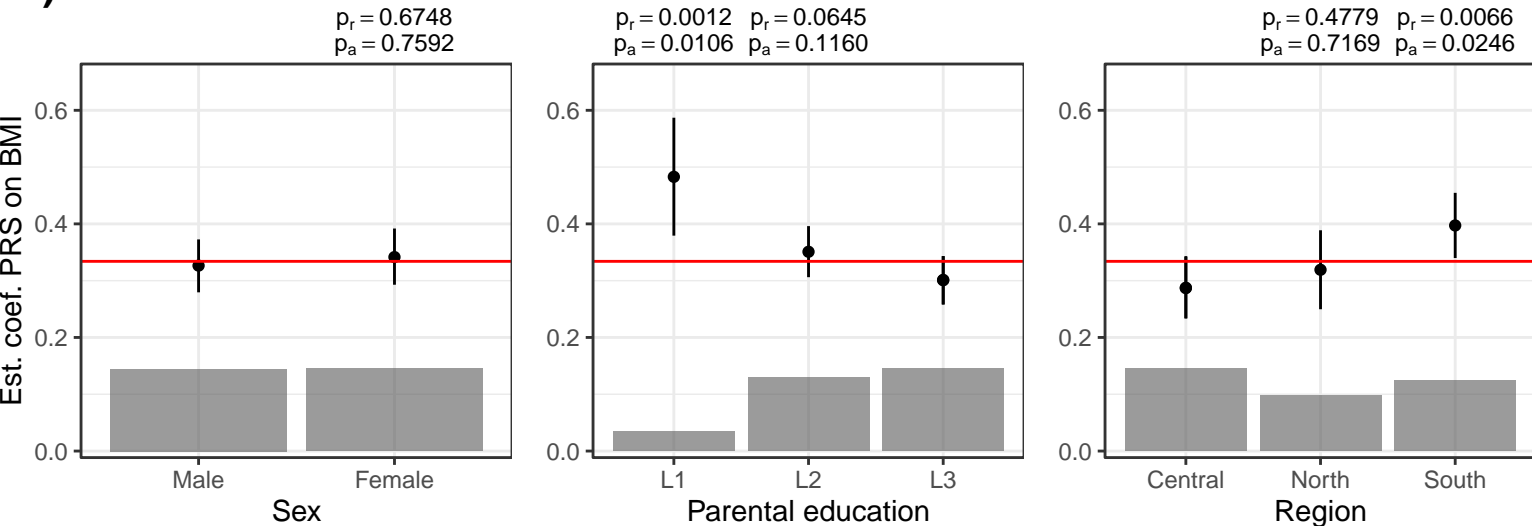
BMI



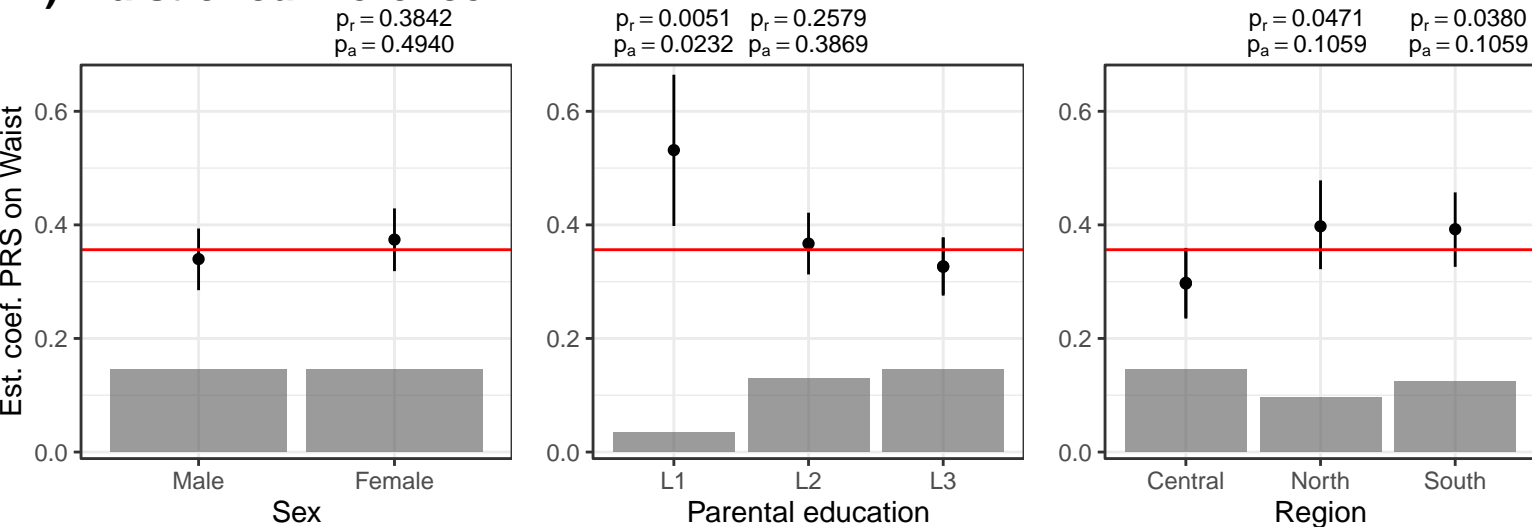
Waist circumference



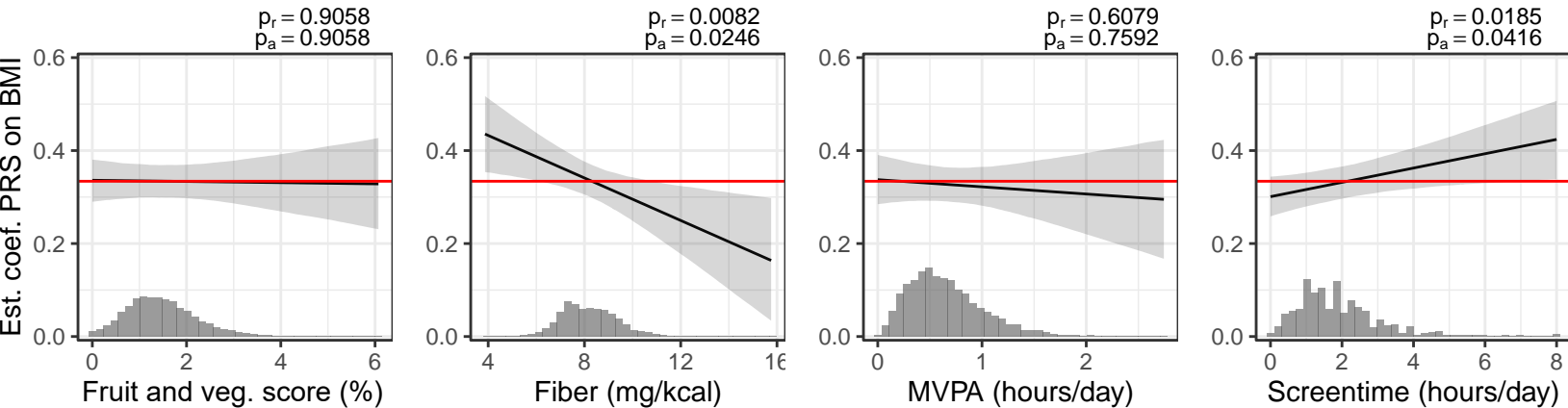
A) BMI



B) Waist circumference



A) BMI



B) Waist circumference

