

# Assessing performance of pathogenicity predictors using clinically-relevant variant datasets

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## ABSTRACT

**Purpose:** Pathogenicity predictors are an integral part of genomic variant interpretation but, despite their widespread usage, an independent validation of performance using a clinically-relevant dataset has not been undertaken.

**Methods:** We derive two validation datasets: an “open” dataset containing variants extracted from publicly-available databases, similar to those commonly applied in previous benchmarking exercises, and a “clinically-representative” dataset containing variants identified through research/diagnostic exome and diagnostic panel sequencing. Using these datasets, we evaluate the performance of three recently developed meta-predictors, REVEL, GAVIN and ClinPred, and compare their performance against two commonly used *in silico* tools, SIFT and PolyPhen-2.

**Results:** Although the newer meta-predictors outperform the older tools, the performance of all pathogenicity predictors is substantially lower in the clinically-representative dataset. Using our clinically-relevant dataset, REVEL performed best with an area under the ROC of 0.81. Using a concordance-based approach based on a consensus of multiple tools reduces the performance due to both discordance between tools and false concordance where tools make common misclassification. Analysis of tool feature usage may give an insight into the tool performance and misclassification.

**Conclusion:** Our results support the adoption of meta-predictors over traditional *in silico* tools, but do not support a consensus-based approach as recommended by current variant classification guidelines.

**Keywords:** pathogenicity, genomic medicine, meta-predictor, variant interpretation, variant classification

## 45 1. INTRODUCTION

46 As the scale of genomic sequencing continues to increase, the classification of rare genomic variants  
47 is becoming the primary bottle-neck in the diagnosis of rare monogenic disorder. Guidelines  
48 published by the American College of Medical Genetics (ACMG) in 2016<sup>1</sup> have helped bring  
49 consistency to variant classification and have been followed by a number of regional and disorder-  
50 specific publications<sup>2-4</sup>. Common to all guidelines is the recommendation of the use of *in silico*  
51 prediction tools to aid in the classification of missense variants. *In silico* prediction tools are  
52 algorithms designed to predict the functional impact of variation, usually missense changes caused  
53 by single nucleotide variants (SNVs). Though originally designed for the prioritisation of research  
54 variants<sup>5</sup>, the tools are used routinely in clinical diagnostics during variant classification. The tools  
55 integrate a number of features in order to assess the impact of a variant on protein function<sup>6</sup>. Initially,  
56 inter-species conservation formed the bulk of the predictions, with some additional functional  
57 information, such as substitution matrices of physicochemical distances of amino acids (such as  
58 Grantham<sup>7</sup> or PAM<sup>8</sup>), and data derived from a limited number of available X-ray crystallographic  
59 structures<sup>9</sup>. Since the development of the first *in silico* prediction tools over a decade ago<sup>5,9</sup>, large-  
60 scale experiments such as the ENCODE project<sup>10</sup> have generated huge amounts of functional data,  
61 and we now also have access to large-scale databases of clinical and neutral variation<sup>11-13</sup>. These  
62 additional sources of data have led to an explosion of new *in silico* prediction algorithms<sup>14-16</sup> that  
63 purport to increase accuracy.  
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65 However, the large increase in the number of predictors integrated into classification algorithms has  
66 raised concerns about overfitting<sup>17,18</sup>. Overfitting occurs when the prediction algorithm is trained on  
67 superfluous data or features that are irrelevant to the prediction outcome<sup>18</sup>. While it may appear  
68 that an increasingly large feature list leads to improvements in prediction, random variability within  
69 the training dataset may actually result in decreased accuracy when applied to a novel dataset.  
70 Overfitting can be mitigated through the use of increasingly large training datasets, and the usage of  
71 online variant databases, such as the genome aggregation database (gnomAD)<sup>19</sup> and ClinVar<sup>12</sup>,  
72 allows for sufficiently large training datasets. Additionally, reliance on additional information – such  
73 as protein functional data and allele frequency data such as from gnomAD<sup>19</sup> – may be contrary to the  
74 standard assumptions of variant classification methodology, namely that each dataset is  
75 independent and applied only once during classification.  
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77 Current ACMG guidelines recommend the use of a concordance-based approach, where a number of  
78 prediction algorithms are used, and evidence is applied only when there is agreement between  
79 tools. There is no guidance on which *in silico* tools should be used, how many, or on what constitutes  
80 a consensus, and this ambiguity allows for inconsistencies in the application of this piece of evidence  
81 across clinical laboratories. Studies have previously identified the limitations of applying a strict  
82 binary consensus-based approach<sup>20</sup>. In response, multiple groups<sup>14-16</sup> have created meta-predictors;  
83 tools which integrate information from a large number of sources into a machine-learning algorithm.  
84 These tools thereby adhere to the principle of the consensus-based model suggested by ACMG  
85 without the onerous task of determining tool concordance, and reduce discordance when  
86 increasingly large numbers of tools are utilised. Unlike a manual consensus-based model, where  
87 tools are weighted equally, meta-predictors are able to apply weighting to features in order to  
88 maximise accuracy.  
89

90 In order to evaluate the accuracy of *in silico* prediction tools, precompiled variant datasets such as  
91 VariBench<sup>21</sup> have been designed to aid in training and benchmarking of pathogenicity predictors.  
92 However, the use of standardised datasets may introduce inherent biases into prediction algorithms,  
93 resulting in false concordance. Typically, prediction software is trained using machine-learning  
94 algorithms, and assessed using variants available from large online public databases<sup>5,6,9,10,14-16,22</sup> such  
95 as ExAC/gnomAD, ClinVar<sup>12</sup>, and SwissProt<sup>23</sup>. It has been previously shown that prediction algorithms  
96 have variable performance when applied to different datasets<sup>6,22,24,25</sup>, and therefore the use of  
97 variant datasets derived from online public databases may not be representative of the performance  
98 of tools when applied in a clinical setting. While studies emphasise the use of 'neutral' variation, the  
99 output from a modern next-generation sequencing pipeline is generally far from neutral, and

100 includes a large number of variant filtering steps in order to reduce the burden of manual variant  
101 assessment<sup>26</sup>.  
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103 Here we evaluate and compare the performance of two traditional *in silico* pathogenicity prediction  
104 tools commonly used for clinical variant interpretation (SIFT<sup>5</sup> and PolyPhen-2<sup>9</sup>), and three meta-  
105 predictors (REVEL<sup>14</sup>, GAVIN<sup>15</sup> and ClinPred<sup>16</sup>) using a publicly available ('open') variant dataset and a  
106 clinically-relevant ('clinical') variant dataset. We show that the tools' performance is heavily affected  
107 by the test dataset, and that all tools may perform worse than expected when classifying novel  
108 missense variants. By assessing the effect of a consensus-based approach, our results support the  
109 use of a single classifier when performing variant classification.  
110

## 111 2. MATERIALS AND METHODS

112 **2.1 Open Dataset** (n=8795, see **Figure S1A**) represents the typical training and validation dataset  
113 used during *in silico* predictor design and benchmarking. Positive ('pathogenic') variants were  
114 downloaded from ClinVar<sup>12</sup> on 13<sup>th</sup> November 2017 and subscription-based HGMD<sup>28</sup> Professional  
115 release 2017.3; neutral ('benign') variants in OMIM<sup>27</sup> morbid genes were downloaded from the  
116 gnomAD<sup>11</sup> database (exomes only data v2.0.1). **ClinVar criteria:** Stringent criteria were used to  
117 increase the likelihood of selected variants being truly pathogenic. Missense SNVs with either  
118 'pathogenic' and/or 'likely pathogenic' classification, multiple submitters and no conflicting  
119 submissions were included; variants with any assertions of 'uncertain', 'likely benign' or 'benign'  
120 were excluded. **HGMD Pro criteria:** Single nucleotide missense variants marked as disease-causing  
121 ('DM') were taken from HGMD Professional release 2017.3. **gnomAD criteria:** Missense SNVs with an  
122 overall minor allele frequency (MAF) between 1% and 5% were selected. These variants were  
123 deemed too common to be disease-causing but are not necessarily filtered out by next-generation  
124 sequencing pipelines depending on the MAF thresholds used. Chromosomal locations with more  
125 than one variant (multiallelic sites) were excluded. Any variants found to be present in the  
126 'pathogenic' and 'neutral' datasets were removed from the both.  
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128 **2.2 Clinical Dataset** (n=1766, see **Figure S1B** and **Supplemental Table S1**) more accurately reflects  
129 variants that might require classification in a clinical diagnostics laboratory following identification in  
130 an exome or genome sequencing pipeline. Variants were selected from three sources. **Group 1**  
131 ('DDD') consists of pathogenic (n=687) and benign (n=533) missense variants identified from 13,462  
132 families in the Deciphering Developmental Disorders (DDD) study that have been through multiple  
133 rounds of variant filtering and clinical evaluation<sup>26,29</sup>. Variants were identified through exome  
134 sequencing and were reported to the patients' referring clinicians for interpretation and  
135 confirmation in accredited UK diagnostic laboratories. All benign variants from this list were assessed  
136 as having no contribution towards the patient's phenotype, and were present in either as  
137 heterozygotes in monoallelic genes or homozygotes in biallelic genes classified according to the  
138 Developmental Disorder Genotype-2-Phenotype database (DDG2P)<sup>30</sup> (data accessed 17/10/2019).  
139 **Group 2** ('Diagnostic') consisted of pathogenic (n = 322) and benign (n=23) missense variants  
140 identified through Sanger sequencing, next-generation sequencing panel analysis or single gene  
141 testing in an accredited clinical diagnostic laboratory. Variants were manually classified according to  
142 the ACMG guidelines on variant interpretation<sup>1</sup> on a 5-point scale (data accessed 23/04/2019).  
143 **Group 3** ('Amish') consisted of benign missense variants (n = 53) identified through a Community  
144 Genomics research study of 220 Amish individuals. Variants were identified through singleton exome  
145 sequencing and were classified as benign based on population frequencies and zygosity within this  
146 study. Two subgroups were manually selected and annotated based on inheritance pattern and  
147 disease penetrance; subgroup (i) consisted of variants in genes that cause a dominantly-inherited  
148 disorder with complete penetrance in childhood, for which the individual was clinically unaffected;  
149 this list was curated by a consultant in clinical genetics; subgroup (ii) consisted of variants in all other  
150 OMIM morbid genes (including those with incompletely penetrant dominant disorders and recessive  
151 and X-linked inheritance), with MAF>5% in the Amish cohort and MAF≤0.01% in gnomAD (data  
152 accessed 18/10/2019).  
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### 2.3 Transcript selection and variant annotation

For the open dataset, the canonical transcript was selected for each variant using the Variant Effect Predictor (VEP)<sup>31</sup>. For the clinical dataset, the HGMD Professional RefSeq transcript was used, unless absent from the database, in which case the MANE primary transcript was selected. Variants were annotated with variant cDNA and protein nomenclature in reference to the selected transcript. PolyPhen-2 and SIFT scores were annotated using VEP. REVEL and ClinPred scores were annotated using flat files containing precomputed scores for all possible single nucleotide substitutions, and in both cases, the combination of nucleotide position, nucleotide change and amino acid change was sufficiently unique to identify a single record, i.e. transcript selection did not affect the scores. GAVIN scores were generated through a batch submission to the GAVIN server.

### 2.4 Tool benchmarking

The performance of each of the tools was determined for both datasets. For SIFT, PolyPhen-2, REVEL and ClinPred, the output of the analysis was a numerical score between 0 and 1. Initially, all tools were analysed according to the criteria defined in their original publications, with the thresholds for pathogenicity being  $\leq 0.05$  for SIFT,  $\geq 0.9$  for PolyPhen-2 and  $\geq 0.5$  for ClinPred. For REVEL, where no threshold is recommended, a threshold of  $\geq 0.5$  was used. The categorical classification of GAVIN was used directly (“Benign”, “Pathogenic”; variants of uncertain significance (“VOUS”) were removed). A supplementary analysis was done for those tools with a numerical output (SIFT, PolyPhen-2, REVEL and ClinPred), to more accurately compare their performance. A unique threshold was selected for each tool to calculate the specificity when sensitivity was set to 0.9. In order to include GAVIN in this analysis, a third analysis was performed, whereby each tool's specificity was measured when the threshold was adjusted to set the sensitivity identical to that of GAVIN.

		SIFT (2009)	Polyphen-2 (2010)	REVEL (2016)	ClinPred (2018)	GAVIN (2017)
Conservation	Sequence identity – conservation between proteins with a defined sequence identity.			P, S, MP, V, F	P, S	C
	Orthologues – conservation between orthologous proteins within different species.			V, MT	C, D	C
	Protein domains – conservation between members of protein families.			P, MT, MP, F	P, C, D	C
	Predicted nucleotide mutational rate – between-species conservation corrected for predicted mutational models.			P, MP	P, C, D	C
Genetic Variation	Pathogenic variation – databases of annotated pathogenic variants.			V, MT		
	Benign variation – databases of annotated benign or neutral variants.			V, MT		
Functional (nucleotide)	Epigenetics (CpG) – variation at CpG dinucleotides/islands; histone modification; DNA accessibility; chromatin.				C, D	C
	DNA/RNA sequence context – regulatory; transcription factor binding; sequence motif.				C, D, FC	C
	Gene expression				C, D, FC	C
Functional (protein)	Residue-specific functional evidence – active site, binding, post-transcriptional modification, sequence motif, amino acid composition (tracts), secondary structure, disulphide bond formation.			MP, V, MT		
	Protein-specific functional evidence – flexibility, stability, solvent accessibility, intrinsic disorder.			P, MP, V	P, C, D	C
Amino Acid Properties	Amino acid properties (physicochemical change) – volume, hydrophobicity, Grantham distance, polarity.			P, V	P, C, D	C

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**Figure 1. *In silico* pathogenicity predictor feature usage and source.** Shading indicates that a category of evidence is utilised by the tool. Codes within each box indicate that the feature is inherited from another tool. Feature lists were taken from the tools' original publications, supplementary materials and available online material. C – CADD; D – DANN; F – FATHMM; MP – MutPred; MT – MutationTaster; P – PolyPhen-2; S – SIFT; V – VEST; An extended version is shown in Supplemental Figure S2.

186 **3. RESULTS**

187 **3.1 Classification of variant sources**

188 We compared the feature list of all tools benchmarked in this study (PolyPhen-2, SIFT, REVEL, GAVIN  
 189 and ClinPred) and, in the case of the meta-predictors, the tools that they use as part of their  
 190 algorithm (MPC<sup>32</sup>, MutPred<sup>33</sup>, VEST<sup>34</sup>, CADD<sup>35</sup>, DANN<sup>36</sup>, SNPEff<sup>37</sup>, FATHMM<sup>38</sup>, FitCons<sup>39</sup> and  
 191 MutationTaster<sup>40</sup>). Features were split into five broad categories: Conservation, Genetic variation,  
 192 Functional evidence (nucleotide), Functional evidence (protein) and Amino acid properties (see  
 193 **Figure 1 and Supplemental Figure S2**). In general, the meta-predictors employ a wider variety of  
 194 sources, and are less heavily reliant on conservation alone. CADD/DANN and FitCons, and by  
 195 extension GAVIN and ClinPred, are the only predictors with features within the *Functional*  
 196 (*nucleotide*) category and are therefore able to predict the pathogenicity of a variant in the context  
 197 of its nucleotide change, regardless of whether there is a resultant amino acid change.

200 **3.2 Benchmarking predictor performance for in the open and clinical datasets**

201 Initially, each of the tools was benchmarked according to the threshold provided by the tools'  
 202 authors. This analysis involved a dichotomisation of scores with no intermediate range, see **Table 1**.  
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Open Dataset		True Positives	True Negatives	False Positives	False Negatives	Count	Sensitivity	Specificity	MCC
Individual	SIFT	2304	4051	1957	444	8756	0.84	0.67	0.48
	Polyphen	2390	4200	1563	358	8511	0.87	0.73	0.56
	REVEL	2396	5754	292	352	8794	0.87	0.95	0.83
	GAVIN	2618	5912	134	130	8794	0.95	0.98	0.93
	ClinPred	2471	6041	5	277	8794	0.90	1.00	0.93
Consensus	SIFT+Polyphen	2121	3351	2695	627	8794	0.77	0.55	0.30
	REVEL+ClinPred	2236	5751	295	512	8794	0.81	0.95	0.78

Clinical Dataset		True Positives	True Negatives	False Positives	False Negatives	Count	Sensitivity	Specificity	MCC
Individual	SIFT	1031	217	410	108	1766	0.91	0.35	0.31
	Polyphen	1021	216	411	118	1766	0.90	0.34	0.29
	REVEL	983	377	250	156	1766	0.86	0.60	0.48
	GAVIN	1100	157	461	39	1757	0.97	0.25	0.33
	ClinPred	1107	174	453	32	1766	0.97	0.28	0.37
Consensus	SIFT+Polyphen	960	139	489	179	1767	0.84	0.22	0.08
	REVEL+ClinPred	973	150	478	166	1767	0.85	0.24	0.12

204 **Table 1. Results of variant classification for individual tool, and two consensus-based combinations, for**  
 205 **datasets A, B and C.** For consensus-based results non-concordant, where tools disagree on the  
 206 classification, were considered incorrect.  
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208 Matthews correlation coefficient (MCC) was calculated as follows:

$$MCC = \frac{(TP \times TN - FP \times FN)}{\sqrt{(TP + FP) \times (TP + FN) \times (TN + FP) \times (TN + FN)}}$$

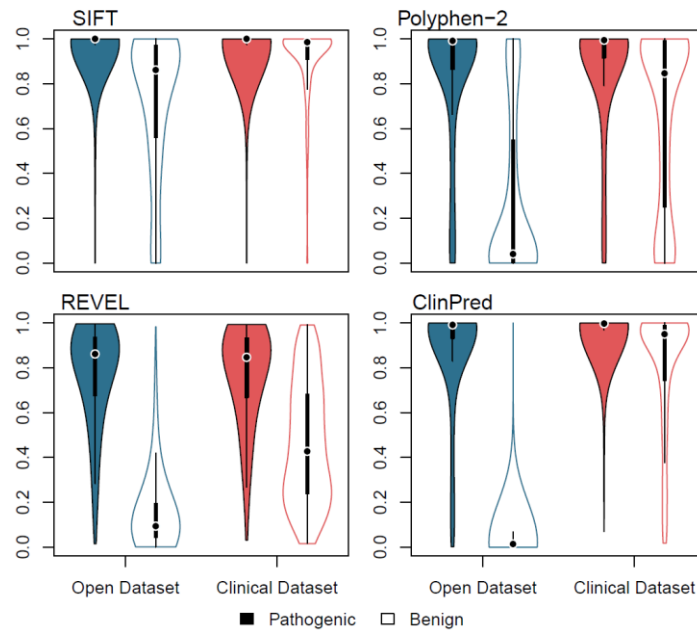
209 TP = True Positives; FP = False Positives; TN = True Negatives; FN = False Negatives;

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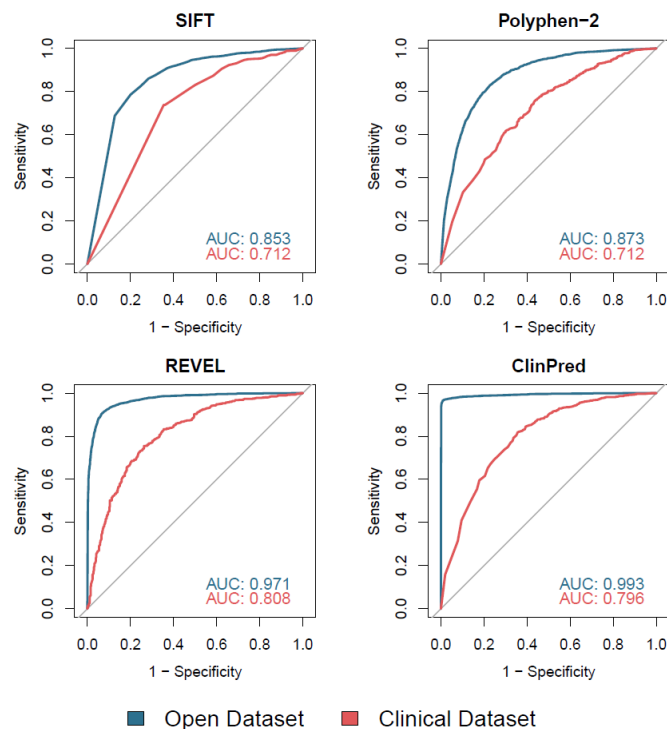
214 The distribution of scores from SIFT, PolyPhen-2, REVEL and ClinPred is shown in **Figure 2** and ROC  
 215 curves are shown in **Figure 3**. Of the tools with numerical outputs, ClinPred has the highest  
 216 discriminatory power for the open dataset with an area under the ROC curve (AUC) of 0.993, while  
 217 REVEL has the highest AUC for the clinical dataset (0.808). The two meta-predictors outperformed  
 218 SIFT and PolyPhen-2 in both datasets. In agreement with tool author benchmarking<sup>14-16</sup> the meta-  
 219 predictors REVEL, ClinPred and GAVIN were highly proficient at classifying the variants in the open  
 220 dataset, achieving sensitivities of 0.87, 0.90 and 0.95, and specificities of 0.95, 1.00 and 0.98,



221 respectively. For variants in the clinical dataset, although the sensitivity each tool remained largely  
222 constant, the specificity of all tools dropped considerably. For REVEL, ClinPred and GAVIN, specificity  
223 is reduced to 0.62, 0.28 and 0.25, respectively [Table 1].  
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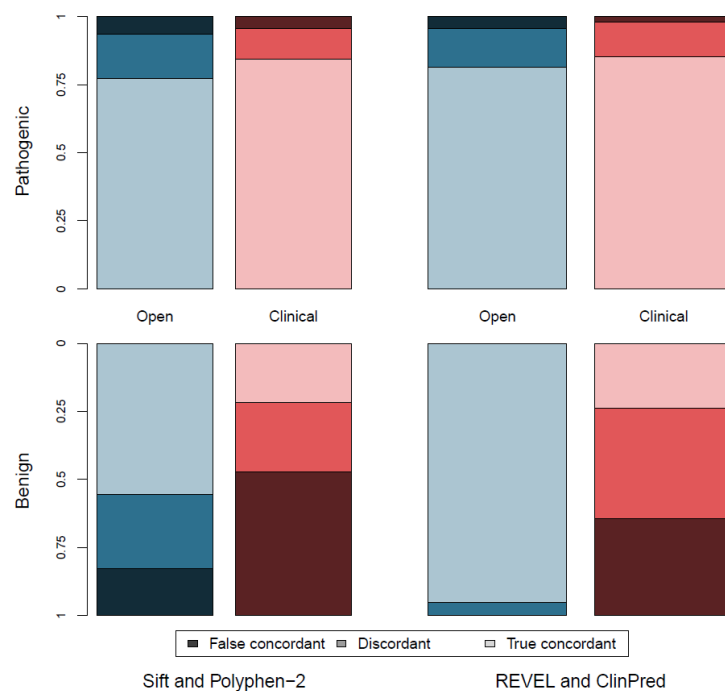


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227 **Figure 2. Violin plot showing variant scores for SIFT, PolyPhen-2, REVEL and ClinPred using two datasets.**  
228 Open dataset – blue; clinical dataset – red; pathogenic variants – filled; benign variants – unfilled. Plot was  
229 generated in R using the 'vioplot' function in the 'vioplot' library. For ease of comparison, SIFT scores have  
230 been inverted.  
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234 **Figure 3. Receiver operating characteristic (ROC) curves for SIFT, PolyPhen-2, REVEL and ClinPred using**  
235 **two datasets.** Open dataset – blue; clinical dataset – red. Generated in R using the 'roc' and 'plot.roc'  
236 functions in the 'pROC' library. Area under the ROC curve (AUC) was calculated in R using the 'roc'  
237 function. For ease of comparison, SIFT scores have been inverted.  
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241 It was apparent that the threshold suggested by the tools' authors was not well-suited to both  
242 datasets, given the tools' very high sensitivity but low specificity in the clinical dataset. In order to  
243 correct for this we performed a supplementary analysis for those predictors which gave a numerical  
244 output (SIFT, PolyPhen-2, REVEL and ClinPred). Here, a variable threshold was allowed for each tool  
245 to give a common sensitivity of 0.9 (i.e. pathogenic variation is called correctly 90% of the time). The  
246 threshold required to give a sensitivity of 0.9 in each tools is shown in **Table S2**. The specificity of  
247 each tool at the determined threshold is shown in **Figure S3**. When allowed a variable threshold the  
248 tools' specificity increased significantly, with PolyPhen-2, SIFT, REVEL and ClinPred achieving a  
249 specificity of 0.67, 0.63, 0.93 and 0.99 for the open dataset, and 0.34, 0.33, 0.52 and 0.52 for the  
250 clinical dataset, respectively. In order to include GAVIN in this analysis, a third analysis was  
251 performed in which each tool was given a threshold to match the sensitivity achieved by GAVIN in  
252 each of the datasets. The specificity of all five tools is shown in **Figure S4**, and the sensitivity and  
253 threshold for each tool is shown in **Table S3**.  
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255 **Figure 4. Concordance between tools separated by dataset and classification (pathogenic and benign).**  
256 Open dataset – blue; clinical dataset – red; pathogenic variants – top graph; benign variants – bottom  
257 graph. True concordance indicates that the tools agree, and were correct. False concordance indicates  
258 that the tools agree but were incorrect. Discordance indicates that the tools disagreed on the  
259 classification.  
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### 262 3.3 Use of individual tools versus a consensus-based approach between multiple tools

263 In accordance with current variant classification guidelines, we investigated the effect of performing  
264 a consensus-based analysis, using two commonly-used tools, SIFT and PolyPhen-2, and two meta-  
265 predictors, REVEL and ClinPred, to determine whether this combined approach has improved  
266 sensitivity/specificity over the individual tools. **Figure 4** shows the true concordance rate (variants  
267 classified correctly by both tools), false concordance rate (variants classified incorrectly by both  
268 tools) and discordance rate (variants for which the tools disagreed) for each of these tool pairings for  
269 the pathogenic and benign variants in both datasets. Within the clinically-relevant dataset, the tools  
270 are either falsely concordant or discordant for ~15% of pathogenic variants but ~77% of benign  
271 variants. The sensitivity and specificity of this approach is shown in **Table 1**. Use of a consensus-  
272 based approach introduces a third "discordance" category to the classification where no *in silico*  
273 evidence can be used, which applied to 24% and 16% of variants when considering the concordance  
274 of PolyPhen-2 and SIFT, and 8% and 23% when considering the concordance between REVEL and  
275 ClinPred, for the open and clinical datasets, respectively.  
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#### 277 4. DISCUSSION

278 We have compared the performance of five *in silico* pathogenicity predictors – two tools used  
279 routinely in variant classification (SIFT and PolyPhen-2) and three recently developed meta-  
280 predictors (REVEL, ClinPred and GAVIN) – using two variant datasets: an open dataset collated using  
281 the selection strategy commonly employed when benchmarking tool performance, and a clinically-  
282 representative dataset composed of rare and novel variants identified through high-throughput  
283 research and clinical sequencing and manual classification. Overall, the data herein show that meta-  
284 predictors have a greater sensitivity and specificity than the classic tools in both variant datasets.  
285 However, despite the increased accuracy of the meta-predictors, all tools performed substantially  
286 worse in the clinical dataset compared with the open dataset. This difference in tool performance  
287 illustrates the importance of considering the provenance of variants when benchmarking tools and  
288 how overfitting of a classifier to the training dataset can occur when increasingly large sets of variant  
289 features are utilised. Our analysis suggests that REVEL performs best when classifying rare variants  
290 routinely identified in clinical sequencing pipelines, with an AUC for our clinical dataset of 0.808,  
291 followed closely by ClinPred with an AUC of 0.796 [Figure 3] and with a higher specificity than GAVIN  
292 in a direct (albeit suboptimal) comparison [Figure S4]. While the REVEL team does not suggest a  
293 strict threshold for categorisation, in our analysis for the clinical dataset, a threshold of 0.43 gave a  
294 sensitivity of 0.9, and a specificity of 0.52, which is comparable to previous studies' threshold of  
295 0.5<sup>16</sup>.

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297 Current guidelines on the classification of variants indicate that evidence should only apply when  
298 multiple tools are concordant<sup>1</sup>. However, the use of concordance introduces a third category to  
299 variants classification (discordance), where there is disagreement between tools and therefore the  
300 tools cannot be used as evidence to categorise the variant as either benign or pathogenic. Our data  
301 show that the use of concordance between multiple tools gives a lower sensitivity and specificity  
302 than the use of either of these tools in isolation, and furthermore that their performance is much  
303 below that of the meta-predictors.

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305 As with all similar studies, we were limited by the availability of novel variants not present in online  
306 databases such as gnomAD. The use of under-represented and genetically isolated populations, such  
307 as the Amish, allowed for the identification of a number of novel benign variants and suggests that  
308 such populations may be a rich source for future studies. We also identified a number of both  
309 pathogenic and benign variants in a clinical population through a translational research study (DDD).  
310 While steps were taken to ensure that the benign variants attained from this group were indeed  
311 benign (all variant were present within either monoallelic genes, or in biallelic genes in a  
312 homozygous state, and were annotated by the referring clinician as having no contribution towards  
313 the patient's clinical phenotype), nonetheless it cannot be guaranteed that the variants had no  
314 impact of protein function. The study highlights the need for improved data-sharing between clinical  
315 laboratories. While a number of online repositories exist for the sharing of rare pathogenic variants,  
316 no such resource is available for the sharing of rare benign variants.

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318 The study supports the adoption of *in silico* meta-predictors for use in variant classification according  
319 to the ACMG guidelines, but recommends the use of a single meta-predictor over the application of  
320 a consensus-based approach. Each of the tools utilises different though heavily overlapping data  
321 sources and the feature list utilised by a tool should be carefully considered before the tool is  
322 utilised. Our results also suggests that tools that utilise gnomAD data directly may have low  
323 specificity when classifying rare or novel variants and that care should be taken when utilising these  
324 tools in conjunction with the ACMG guidelines. Although use of a meta-predictor tools offers notable  
325 advantages to the use of the previously available and widely adopted *in silico* tools, the remaining  
326 issues to be addressed before they can be used as more at a level greater than supporting evidence  
327 for clinical variant interpretation.



328 **Supplemental Materials:**

329 File S1: PDF file containing supplemental Figures S1, S2, S3, S5 and supplemental Tables S2 and S3.

330 File S2: Microsoft Excel file containing Supplemental Table S1.

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335 **Web Resources**

336 CADD: <https://cadd.gs.washington.edu/>

337 dbSNFP: <https://sites.google.com/site/jpopgen/dbNSFP>

338 GAVIN: <https://molgenis20.gcc.rug.nl/menu/main/gavin-app>

339 gnomAD: <https://gnomad.broadinstitute.org/>

340 HGMD Professional: <https://portal.biobase-international.com/hgmd/pro/start.php>

341 OMIM: <https://www.omim.org/>

342 PolyPhen-2: <http://genetics.bwh.harvard.edu/pph2/>

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359 **CONFLICT OF INTEREST**

360 The authors declare no conflict of interest.

## 361 REFERENCES

- 362 1. Richards S, Aziz N, Bale S, et al. Standards and guidelines for the interpretation of sequence  
363 variants: A joint consensus recommendation of the American College of Medical Genetics and  
364 Genomics and the Association for Molecular Pathology. *Genet Med*. 2015;17(5):405-424.  
365 doi:10.1038/gim.2015.30
- 366 2. Romanet P, Odou M-F, North M-O, et al. Proposition of adjustments to the ACMG-AMP  
367 framework for the interpretation of MEN1 missense variants. *Hum Mutat*. 2019;40(6):661-  
368 674. doi:10.1002/humu.23746
- 369 3. Maxwell KN, Hart SN, Vijai J, et al. Evaluation of ACMG-Guideline-Based Variant Classification  
370 of Cancer Susceptibility and Non-Cancer-Associated Genes in Families Affected by Breast  
371 Cancer. *Am J Hum Genet*. 2016;98(5):801-817. doi:10.1016/j.ajhg.2016.02.024
- 372 4. Sian Ellard, Emma L Baple, Alison Callaway, Ian Berry, Natalie Forrester, Clare Turnbull,  
373 Martina Owens, Diana M Eccles, Stephen Abbs, Richard Scott, Zandra C Deans, Tracy Lester,  
374 Jo Campbell, William G Newman SR and DJM. *ACGS Best Practice Guidelines for Variant  
375 Classification in Rare Disease 2020.*; 2019.
- 376 5. Sim NL, Kumar P, Hu J, Henikoff S, Schneider G, Ng PC. SIFT web server: Predicting effects of  
377 amino acid substitutions on proteins. *Nucleic Acids Res*. 2012;40(W1).  
378 doi:10.1093/nar/gks539
- 379 6. Thusberg J, Olatubosun A, Vihinen M. Performance of mutation pathogenicity prediction  
380 methods on missense variants. *Hum Mutat*. 2011. doi:10.1002/humu.21445
- 381 7. Grantham R. Amino acid difference formula to help explain protein evolution. *Science (80- )*.  
382 1974. doi:10.1126/science.185.4154.862
- 383 8. Dayhoff MO, Schwartz RM, Orcutt BC. A Model of Evolutionary Change in Proteins. In: *Atlas  
384 of Protein Sequence and Structure*. ; 1978:345-352.
- 385 9. Adzhubei IA, Schmidt S, Peshkin L, et al. A method and server for predicting damaging  
386 missense mutations. *Nat Methods*. 2010;7(4):248-249. doi:10.1038/nmeth0410-248
- 387 10. Dunham I, Kundaje A, Aldred SF, et al. An integrated encyclopedia of DNA elements in the  
388 human genome. *Nature*. 2012;489(7414):57-74. doi:10.1038/nature11247
- 389 11. Karczewski KJ, Francioli LC, Tiao G, et al. Variation across 141,456 human exomes and  
390 genomes reveals the spectrum of loss-of-function intolerance across human protein-coding  
391 genes. *bioRxiv*. 2019:531210. doi:10.1101/531210
- 392 12. Landrum MJ, Lee JM, Benson M, et al. ClinVar: improving access to variant interpretations  
393 and supporting evidence. *Nucleic Acids Res*. 2018;46(D1):D1062-D1067.  
394 doi:10.1093/nar/gkx1153
- 395 13. Stenson PD, Mort M, Ball E V., Shaw K, Phillips AD, Cooper DN. The Human Gene Mutation  
396 Database: Building a comprehensive mutation repository for clinical and molecular genetics,  
397 diagnostic testing and personalized genomic medicine. *Hum Genet*. 2014.  
398 doi:10.1007/s00439-013-1358-4
- 399 14. Ioannidis NM, Rothstein JH, Pejaver V, et al. REVEL: An Ensemble Method for Predicting the  
400 Pathogenicity of Rare Missense Variants. *Am J Hum Genet*. 2016;99(4):877-885.  
401 doi:10.1016/j.ajhg.2016.08.016
- 402 15. van der Velde KJ, de Boer EN, van Diemen CC, et al. GAVIN: Gene-Aware Variant  
403 INterpretation for medical sequencing. *Genome Biol*. 2017;18(1). doi:10.1186/s13059-016-  
404 1141-7
- 405 16. Alirezaie N, Kernohan KD, Hartley T, Majewski J, Hocking TD. ClinPred: Prediction Tool to  
406 Identify Disease-Relevant Nonsynonymous Single-Nucleotide Variants. *Am J Hum Genet*.  
407 2018;103(4):474-483. doi:10.1016/j.ajhg.2018.08.005
- 408 17. Subramanian J, Simon R. Overfitting in prediction models - Is it a problem only in high  
409 dimensions? *Contemp Clin Trials*. 2013. doi:10.1016/j.cct.2013.06.011
- 410 18. Hawkins DM. The Problem of Overfitting. *J Chem Inf Comput Sci*. 2004.  
411 doi:10.1021/ci0342472
- 412 19. Karczewski KJ, Francioli LC, Tiao G, et al. Variation across 141,456 human exomes and  
413 genomes reveals the spectrum of loss-of-function intolerance across human protein-coding

- 414 genes. *bioRxiv*. 2019. doi:10.1101/531210
- 415 20. Ghosh R, Oak N, Plon SE. Evaluation of in silico algorithms for use with ACMG/AMP clinical  
416 variant interpretation guidelines. *Genome Biol*. 2017;18(1). doi:10.1186/s13059-017-1353-5
- 417 21. Sasidharan Nair P, Vihinen M. VariBench: A Benchmark Database for Variations. *Hum Mutat*.  
418 2013;34(1):42-49. doi:10.1002/humu.22204
- 419 22. Grimm DG, Azencott CA, Aicheler F, et al. The evaluation of tools used to predict the impact  
420 of missense variants is hindered by two types of circularity. *Hum Mutat*. 2015.  
421 doi:10.1002/humu.22768
- 422 23. D506-D515. UniProt: a worldwide hub of protein knowledge The UniProt Consortium. *Nucleic*  
423 *Acids Res*. 2019. doi:10.1093/nar/gky1049
- 424 24. Dong C, Wei P, Jian X, et al. Comparison and integration of deleteriousness prediction  
425 methods for nonsynonymous SNVs in whole exome sequencing studies. *Hum Mol Genet*.  
426 2015. doi:10.1093/hmg/ddu733
- 427 25. Niroula A, Vihinen M. How good are pathogenicity predictors in detecting benign variants?  
428 *PLoS Comput Biol*. 2019. doi:10.1371/journal.pcbi.1006481
- 429 26. Wright CF, Fitzgerald TW, Jones WD, et al. Genetic diagnosis of developmental disorders in  
430 the DDD study: a scalable analysis of genome-wide research data. *Lancet (London, England)*.  
431 2015;385(9975):1305-1314. doi:10.1016/S0140-6736(14)61705-0
- 432 27. Hamosh A, Scott AF, Amberger JS, Bocchini CA, McKusick VA. Online Mendelian Inheritance in  
433 Man (OMIM), a knowledgebase of human genes and genetic disorders. *Nucleic Acids Res*.  
434 2005;33(DATABASE ISS.). doi:10.1093/nar/gki033
- 435 28. Stenson PD, Mort M, Ball E V, et al. The Human Gene Mutation Database: towards a  
436 comprehensive repository of inherited mutation data for medical research, genetic diagnosis  
437 and next-generation sequencing studies. *Hum Genet*. 2017;136(6):665-677.  
438 doi:10.1007/s00439-017-1779-6
- 439 29. Wright CF, McRae JF, Clayton S, et al. Making new genetic diagnoses with old data: iterative  
440 reanalysis and reporting from genome-wide data in 1,133 families with developmental  
441 disorders. *Genet Med*. 2018;20(10):1216-1223. doi:10.1038/gim.2017.246
- 442 30. Thormann A, Halachev M, McLaren W, et al. Flexible and scalable diagnostic filtering of  
443 genomic variants using G2P with Ensembl VEP. *Nat Commun*. 2019. doi:10.1038/s41467-019-  
444 10016-3
- 445 31. McLaren W, Gil L, Hunt SE, et al. The Ensembl Variant Effect Predictor. *Genome Biol*.  
446 2016;17(1):122. doi:10.1186/s13059-016-0974-4
- 447 32. Samocha KE, Kosmicki JA, Karczewski KJ, et al. Regional missense constraint improves variant  
448 deleteriousness prediction. *bioRxiv*. 2017:148353. doi:10.1101/148353
- 449 33. Li B, Krishnan VG, Mort ME, et al. Automated inference of molecular mechanisms of disease  
450 from amino acid substitutions. *Bioinformatics*. 2009;25(21):2744-2750.  
451 doi:10.1093/bioinformatics/btp528
- 452 34. Carter H, Douville C, Stenson PD, Cooper DN, Karchin R. Identifying Mendelian disease genes  
453 with the variant effect scoring tool. *BMC Genomics*. 2013;14 Suppl 3:S3. doi:10.1186/1471-  
454 2164-14-S3-S3
- 455 35. Kircher M, Witten DM, Jain P, O'roak BJ, Cooper GM, Shendure J. A general framework for  
456 estimating the relative pathogenicity of human genetic variants. *Nat Genet*. 2014;46(3):310-  
457 315. doi:10.1038/ng.2892
- 458 36. Quang D, Chen Y, Xie X. DANN: A deep learning approach for annotating the pathogenicity of  
459 genetic variants. *Bioinformatics*. 2015;31(5):761-763. doi:10.1093/bioinformatics/btu703
- 460 37. Cingolani P, Platts A, Wang LL, et al. A program for annotating and predicting the effects of  
461 single nucleotide polymorphisms, SnpEff: SNPs in the genome of *Drosophila melanogaster*  
462 strain w1118; iso-2; iso-3. *Fly (Austin)*. 2012;6(2):80-92. doi:10.4161/fly.19695
- 463 38. Shihab HA, Gough J, Cooper DN, et al. Predicting the Functional, Molecular, and Phenotypic  
464 Consequences of Amino Acid Substitutions using Hidden Markov Models. *Hum Mutat*.  
465 2013;34(1):57-65. doi:10.1002/humu.22225
- 466 39. Gulko B, Hubisz MJ, Gronau I, Siepel A. A method for calculating probabilities of fitness

- 467 consequences for point mutations across the human genome. *Nat Genet.* 2015;47(3):276-  
468 283. doi:10.1038/ng.3196
- 469 40. Schwarz JM, Rödelberger C, Schuelke M, Seelow D. MutationTaster evaluates disease-  
470 causing potential of sequence alterations. *Nat Methods.* 2010;7(8):575-576.  
471 doi:10.1038/nmeth0810-575
- 472 41. Jaganathan K, Kyriazopoulou Panagiotopoulou S, McRae JF, et al. Predicting Splicing from  
473 Primary Sequence with Deep Learning. *Cell.* 2019;176(3):535-548.e24.  
474 doi:10.1016/j.cell.2018.12.015