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Title:

Dopaminergic modulation of reward discounting: a systematic review and meta-analysis

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Data Availability:

Data and code used in analyses can be viewed and downloaded on OSF (https://osf.io/27cqw/).

Author Contributions:

J.J.C. and G.R.S.L. conceived and designed the research. J.J.C., J.M., L.G., and K.H. searched and extracted the data. J.J.C. analyzed the data with input from G.R.S.L. J.J.C. and G.R.S.L. wrote the manuscript. All authors approved the final version of the manuscript.

Competing Interests:

The authors declare no competing interests.

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Abstract:

Although numerous studies have suggested that pharmacological alteration of the dopamine (DA) system modulates reward discounting, these studies have produced inconsistent findings. Here, we conducted a systematic review and meta-analysis to evaluate DA drug-mediated effects on reward discounting in studies of healthy rodents. This produced a total of 1,343 articles to screen for inclusion/exclusion. From the literature, we identified 117 effects from approximately 1,549 individual rodents. Using random-effects with maximum-likelihood estimation, we meta-analyzed placebo-controlled drug effects for (1) DA transporters, (2) DA D1-like receptor agonists and (3) antagonists, and (4) D2-like agonists and (5) antagonists. Meta-analytic effects showed that DAT-binding drugs decreased reward discounting. While D1 and D2 antagonists both increased discounting, agonist drugs for those receptors had no significant effect on discounting behavior. These findings suggest a nuanced relationship between DA and discounting behavior and urge caution when drawing generalizations about the effects of pharmacologically manipulating dopamine on reward-based decision making.

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Main Text:

Introduction

Every day, all animals make decisions that involve weighing costs and benefits. Animals regularly devalue rewards that are relatively delayed, uncertain, or require more effort than sooner, more certain, or less effortful ones. This process is known as reward discounting. For example, people often choose to eat at restaurants because it is less effortful or time consuming than cooking a meal. In this scenario, people place a greater value on food that is immediately available or easy-to-acquire.

While most individuals discount to some degree, a range of factors influence whether one discounts rewards more steeply (stronger devaluation) or not at all. In humans, for example, income, IQ, age, smoking, and BMI have all been linked to individual differences in reward discounting [1–3]. Aside from these sociodemographic and physical health factors, discounting is often disrupted in many forms of psychopathology [4,5]. An emergent pattern suggests that disruption to circuits involved in the neurotransmission of dopamine (DA) may account for variation in discounting behavior across specific psychopathologies that are often treated with drugs that primarily act on the dopamine system [5,6].

Importantly, drugs that act on the DA system have different effects depending on their targets and action. The putative DA targets for pharmacology are presynaptic synthesis, DA transporters (DAT), and agonism or antagonism of postsynaptic D1-like or D2-like receptors. Variation in results from studies testing the effect of different DA drug effects across these sites on discounting behavior suggests the need for a quantitative comparison of experiments[7–20]. Here, we conducted a systematic review and meta-analysis to evaluate DA drug-mediated effects on reward discounting in studies of healthy rodents (117 effects from approximately 1,549

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animals). We focused on rodents because of the small number of human and non-human primate studies identified (N = 4 studies).

Methods

Literature Search and Study Identification

A meta-analysis was conducted following the Preferred Reporting Items for Systematic Reviews and Meta-Analysis (PRISMA) guidelines [21]. From an initial in-lab library of 34 papers on pharmacological manipulation of dopamine effects on reward discounting, we developed a database of search terms to identify additional studies. We restricted the search to the PubMed database using Medical Subject Headings (MeSH) terms that are most frequently associated with papers in the library. To identify the most frequent MeSH terms, we used the MeSH on Demand tool (https://meshb.nlm.nih.gov/MeSHonDemand) to identify terms from the abstract text of each of the 34 papers. Frequently associated terms that best described the features of the studies of interest included: "animals," "dopamine," "reward," "impulsive behavior," "choice behavior," and "delay discounting." The terms were then combined to search for original research examining how administration of dopaminergic drugs influence reward discounting behavior using the following PubMed search string: "Dopamine" [Mesh] AND ("reward" [Mesh] OR "delay discounting" [Mesh] OR "choice behavior" [Mesh] OR "impulsive behavior" [Mesh] OR "temporal discounting" OR "probability discounting" OR "effort discounting" OR "intertemporal choice" OR "indifference point") AND ("drug" OR "agonist" OR "antagonist").

We restricted the meta-analysis to original studies written in English. Studies must have included a healthy animal group (including humans, non-human primates, and rodents) exposed to placebo and/or drug manipulation. Healthy animals that received lesions or other surgical

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manipulation prior to drug manipulation were excluded unless discounting behavior was unaffected by the lesion or surgery. We further limited the analysis to studies using choice tasks or questionnaires with varying levels of temporal delays, probability, or effort expenditure. To reduce the complexity of the impact of various drugs, we limited confirmatory analyses to drugs that exhibit direct primary action on either: D1-like receptors, D2-like receptors, and DAT. Drugs manipulating levels of the dopamine precursor, L-DOPA were also included with the acknowledgement that too few studies may exist to meta-analyze. We excluded studies using healthy controls identified as nicotine users or relatives of patients with Parkinson's Disease. We also excluded studies that only tested drug effects in humans over 30 years old to reduce the influence of strong age-related declines in DA receptors. Additional studies were later excluded from analysis for reasons that would prevent reliable effect size estimation (e.g. unclear or unreported sample sizes, blurry graphs, or unreported measures of variance). All of these search methods were pre-registered on the Open Science Framework prior to the start of any research activity and all literature search materials may be viewed/downloaded at: https://osf.io/27cgw/. These steps taken to exclude studies are presented in the PRISMA flowchart in **Supplementary**

Figure S1.

Data Extraction

Effect size measures were determined using means, standard deviations, standard errors, confidence intervals, and t-statistics whenever available. Effect sizes were calculated using the 'escalc' function provided with the 'metafor' R Statistics package [22]. For studies that employed a between-subjects design, we calculated the standardized mean difference (SMD) in discounting between drug and placebo groups. Since many rodent studies have small sample sizes, we used the unbiased estimator of the sampling variance for between-subject effects to

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account for possible non-normal distributions [23]. For studies that employed a within-subjects (repeated-measures) design, we calculated the standardized mean change score using raw score standardization (SMCR) as this provides a less biased effect size since repeated measures may be correlated [24]. Since these correlations between drug and placebo conditions are rarely reported, they were set to r = .50 to provide a conservative calculation of the variance [25]. To evaluate the robustness of the meta-analytic effects and validate assumptions, we compared the SMCRs assuming r = .60 and compared with effect sizes calculated using the SMD measure for all effects.

For studies that did not explicitly report these values or that used sophisticated study designs, we used a plot digitizer to determine means and standard deviations [26]. For studies that reported effects for multiple doses of the same drug, we only extracted discounting effects from the highest dose. Studies reported different metrics of discounting including: hyperbolic discounting slope "k" parameter, impulsive choice ratio (ICR), proportion of delayed/uncertain/effortful choices, area under the curve (AUC), and indifference point (also referred to as mean adjusted delay (MAD)). Many studies did not report a single discounting parameter, but instead report the proportion of smaller or larger options at varying levels of time, probability, and effort. In these cases, to simplify comparisons across studies, we averaged the reported choice proportions across cost levels. To standardize the directionality of discounting measures, effect sizes were multiplied by either 1 or -1 to ensure that positive values reflect higher discounting (e.g. higher "k", lower AUC).

Random-effects Meta-analysis

Meta-analytic effects were derived using the metafor R package [22] using random effects with restricted maximum likelihood to help account for between-study variance.

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Specifically, we ran five confirmatory models testing the effects of: 1.) D1R agonists, 2.) D1R antagonists, 3.) D2R agonists, 4.) D2R antagonists, and 5.) DAT-binding drugs. We used the Q statistic to test the null hypothesis that the common effect size is zero and *I*² values to assess significance due to variance explained by heterogeneity of the effects [27]. We evaluated publication bias and study precision asymmetry with visual inspection of a funnel plot and Egger's test (p < 0.05).

Exploratory Meta-analyses

Exploratory meta-regressions examined potential interaction effects of rodent strain, discounting cost types, drug delivery location, and drug dose. For details, see Supplementary Information.

Results

Studies Identified and Data Extraction

The literature search, which was run on January 8, 2018, revealed 1,343 articles. After evaluation of exclusion criteria, 42 unique articles with 121 effects published between 1994 and 2017 remained for quantitative analyses (**see Supplementary Figure S1** for a flow chart). Data was extracted using a plot digitizer for nearly all studies as insufficient statistical reporting prevented reliable estimation of effect sizes and variances. The number of effects for each drug type were: D1R agonist (k = 7), D1R antagonist (k = 17), D2R agonist (k = 18), D2R antagonist (k = 45), and DAT (k = 33). As expected, only one effect size could be extracted for drugs acting on presynaptic DA—too few to analyze. Nearly all effect sizes came from within-subjects designs (k = 115) with only 6 using a between-subjects design. It should be noted that several studies reported more than one effect as a result of repeated exposure to multiple drugs or

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multiple samples separately exposed to different drugs. From the included studies, the majority of effect sizes were from rodents (k = 117), with only a few effects in humans (k = 3) and a single effect in non-human primates (rhesus macaques). The number of effects representing discounting of varying costs were somewhat equally distributed between time (k = 46), probability (k = 43), and effort (k = 32) discounting. The most common measure of discounting was the proportion of larger options chosen (k = 91), followed by the indifference point (k = 25), hyperbolic 'k' value (k = 3), and proportion of smaller options chosen (k = 2). Since there were too few effects in primates (k = 4), these effects were not included with rodent effects in analyses but their data are provided on OSF (https://osf.io/27cqw/). See Supplementary Tables S1-5 for details about each effect size included in quantitative meta-analyses.

Random-effects meta-analysis – confirmatory analyses

D1R agonists. A meta-analysis across D1R agonists did not identify a significant common effect of drug over placebo on discounting (Q = 5.26, p = .511, $I_2 < 0.00\%$; Cohen's d =.136, SE = .118, 95% CI [-.095, .368]. Egger's test for plot asymmetry did not suggest the presence of publication bias (z = .169, p = .865). See forest plot on Figure 1 and funnel plot on Supplementary Figure S2.

Figure 1. Meta-analysis of placebo-controlled effect of D1 agonism on reward discounting. Higher values indicate increased discounting on drug.

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		D1 Agonist		
Author, Year	Drug	_	Ν	Effect Size [95% CI]
St. Onge, 2009	SKF81297		8	-0.17 [-0.87, 0.53]
St. Onge, 2011	SKF81297		12	-0.08 [-0.65, 0.48]
Larkin, 2016.2	SKF81297		10	0.03 [-0.59, 0.65]
Stopper, 2013	SKF81297	·¦∳i	11	0.08 [-0.52, 0.67]
Larkin, 2016.1	SKF81297	·	10	0.13 [-0.50, 0.75]
Simon, 2011	SKF81297	·····	12	0.22 [-0.35, 0.80]
Koffarnus, 2011	SKF81297	·	12	0.77 [0.12, 1.41]
RE Model	I		1	0.14 [-0.10, 0.37]
	-2.0 -	1.0 0.0 1.0 2	.0	
		Effect Size		

D1R antagonists. A meta-analysis across D1R antagonists identified a significant common effect of drug over placebo on discounting (Q = 47.9, p < .001, I_2 = 55.6%; Cohen's d = .532, SE = .120, 95% CI [.296, .767]. Across effects, D1R antagonists increased discounting over placebo. Egger's test for plot asymmetry suggested the presence of publication bias (z = 6.11, p < .001). See forest plot on Figure 2 and funnel plot on Supplementary Figure S3.

Figure 2. Forest plot of placebo-controlled effect of D1 antagonism on reward discounting. Higher values indicate increased discounting on drug.

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D1 Antagonist							
Author, Year	Drug	C	Ν	Effect Size [95% CI]			
Simon, 2011	SCH23390		12	0.01 [-0.55, 0.58]			
Hosking, 2014.1	SCH23390	i∳i	28	0.03 [-0.34, 0.40]			
Wade, 2000	SCH23390	i i i i i i i i i i i i i i i i i i i	17	0.04 [-0.44, 0.51]			
Hosking, 2014.2	SCH23390	H ∲ ⊷i	22	0.22 [-0.20, 0.64]			
St. Onge, 2011	SCH23390	r i ∉-i	12	0.25 [-0.33, 0.82]			
Larkin, 2016.1	SCH23390	H-A-A	8	0.31 [-0.40, 1.02]			
Larkin, 2016.2	SCH23390	⊢ ∔ ♦ −1	8	0.32 [-0.39, 1.03]			
Stopper, 2013	SCH23390	H.	13	0.37 [-0.19, 0.94]			
St. Onge, 2009		+++++	8	0.39 [-0.33, 1.11]			
Koffarnus, 2011	SCH23390	⊷ ••	12	0.57 [-0.04, 1.18]			
Pattij, 2014	SCH23390	! — ◆ _1	13	0.72 [0.11, 1.32]			
Li, 2015	SCH23390	H A H	16	0.81 [0.25, 1.38]			
Pardey, 2013.2	SCH23390	⊢ •1	10	1.10 [0.31, 1.88]			
Cousins, 1994	SCH23390		9	1.25 [0.38, 2.12]			
Pardey, 2013.1	SCH23390	· • • · ·	10	1.44 [0.56, 2.33]			
Bardgett, 2009	SCH23390	; • • • •	9	2.48 [1.16, 3.80]			
Sink, 2007	SCH39166	· · · · · ·	► 8	7.92 [3.98, 11.87]			
RE Model		•		0.53 [0.30, 0.77]			
	-2.0	1.0 3.0 5.0 7.0					
		Effect Size					

D2R agonists. A meta-analysis across D2R agonists did not identify a significant common effect of drug over placebo on discounting (Q = 55.1, p < .001, I_2 = 74.4%; Cohen's d = .044, SE = .151, 95% CI [-.251, .339]. Egger's test for plot asymmetry did not suggest the presence of publication bias (z = -1.67, p = .096). See forest plot on Figure 3 and funnel plot on Supplementary Figure S4.

Figure 3. Forest plot of placebo-controlled effect of D2 agonism on reward discounting. Higher values indicate increased discounting on drug.

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		D2 Agonist		
Author, Year	Drug	Dirigemet	Ν	Effect Size [95% CI]
Rokosik, 2012 St. Onge, 2009.1 Tremblay, 2017 Bardgett, 2009 Pes, 2017 St. Onge, 2011 Stopper, 2013.1 St. Onge, 2009.3 Koffarnus, 2011.4 Larkin, 2016.1 Larkin, 2016.2 Koffarnus, 2011.3 Stopper, 2013.3 Simon, 2011 St. Onge, 2009.2 Koffarnus, 2011.2 Koffarnus, 2011.1	Quinpirole Quinpirole B Sumanirole PD 128,907 Bromocriptine PD 128,907 P ramipexole		10 8 12 9 30 12 11 12 8 12 9 9 12 11 12 8 12 12	$\begin{array}{c} -2.09 \left[-3.20, -0.98\right] \\ -1.07 \left[-1.93, -0.20\right] \\ -0.95 \left[-1.63, -0.27\right] \\ -0.30 \left[-0.97, \ 0.37\right] \\ -0.13 \left[-0.49, \ 0.23\right] \\ -0.09 \left[-0.66, \ 0.47\right] \\ -0.04 \left[-0.63, \ 0.55\right] \\ -0.00 \left[-0.57, \ 0.57\right] \\ 0.10 \left[-0.60, \ 0.79\right] \\ 0.14 \left[-0.43, \ 0.71\right] \\ 0.20 \left[-0.46, \ 0.86\right] \\ 0.23 \left[-0.43, \ 0.89\right] \\ 0.28 \left[-0.30, \ 0.86\right] \\ 0.32 \left[-0.29, \ 0.92\right] \\ 0.49 \left[-0.11, \ 1.09\right] \\ 0.80 \left[0.01, \ 1.60\right] \\ 1.08 \left[0.37, \ 1.79\right] \\ 1.09 \left[0.38, \ 1.81\right] \end{array}$
RE Model		-		0.04 [-0.25, 0.34]
		-3.0 -1.0 1.0 3.	0	
		Effect Size		

D2R antagonists. A meta-analysis across D2R antagonists identified a significant common effect of drug over placebo on discounting (Q = 166.3, p < .001, I_2 = 76.2%; Cohen's *d* = .505, SE = .097, 95% CI [.315, .696]. Across effects, D2R antagonists increased discounting over placebo. Egger's test for plot asymmetry did suggest the presence of publication bias (z = 8.45, p < .001). See forest plot on Figure 4 and funnel plot on Supplementary Figure S5.

Figure 4. Forest plot of placebo-controlled effect of D2 antagonism on reward discounting. Higher values indicate increased discounting on drug.

Author, Year	Drug D2	Antagonist	N	Effect Size [95% CI]
St. Onge, 2011 Li, 2015.1 Ostlund, 2012.3 Koffarnus, 2011.4 St. Onge, 2009.2 Li, 2015.3 Simon, 2011 Stopper, 2013 Ostlund, 2012.1 Larkin, 2016.2 Ostlund, 2012.2 Li, 2015.2 Larkin, 2016.1 Hosking, 2014.1 Pattij, 2014 Shafiei , 2012 Hosking, 2014.2 Koffarnus, 2011.3 St. Onge, 2010.3 Bardgett, 2009.2 St. Onge, 2010.2 Wade, 2000.1 St. Onge, 2010.1 Van Gaalen, 2006 Floresco, 2008.1 Floresco, 2008.1 Floresco, 2008.1 Floresco, 2008.1 Floresco, 2008.1 Floresco, 2008.2 Denk, 2005 Olmstead, 2006 Pardey, 2013.1 Randall, 2012 Sink, 2007 Yohn, 2017 Cousins, 1994.3 Cousins, 1994.1 Bardgett, 2009.1 Causins, 1994.2 RE Model	Eticlopride Flupenthixol L-741,626 Flupenthixol Haloperidol	4 4 4 4	12166 12 8 12 8 10 12 10 12 12 12 12 12 12 12 12 12 12	$\begin{array}{c} -0.41 \ [-1.00, \ 0.18]\\ -0.32 \ [-0.82, \ 0.19]\\ -0.17 \ [-0.50, \ 0.16]\\ -0.14 \ [-0.70, \ 0.43]\\ -0.12 \ [-0.82, \ 0.57]\\ -0.02 \ [-0.51, \ 0.47]\\ -0.01 \ [-0.58, \ 0.55]\\ -0.01 \ [-0.71, \ 0.68]\\ 0.06 \ [-0.26, \ 0.39]\\ 0.06 \ [-0.26, \ 0.39]\\ 0.11 \ [-0.32, \ 0.34]\\ 0.06 \ [-0.26, \ 0.39]\\ 0.11 \ [-0.34, \ 0.60]\\ 0.11 \ [-0.34, \ 0.60]\\ 0.11 \ [-0.34, \ 0.60]\\ 0.11 \ [-0.34, \ 0.60]\\ 0.11 \ [-0.34, \ 0.60]\\ 0.11 \ [-0.34, \ 0.60]\\ 0.11 \ [-0.34, \ 0.60]\\ 0.11 \ [-0.34, \ 0.60]\\ 0.11 \ [-0.34, \ 0.60]\\ 0.11 \ [-0.34, \ 0.60]\\ 0.11 \ [-0.34, \ 0.60]\\ 0.11 \ [-0.34, \ 0.60]\\ 0.11 \ [-0.34, \ 0.60]\\ 0.11 \ [-0.34, \ 0.60]\\ 0.11 \ [-0.34, \ 0.60]\\ 0.50 \ [-0.32, \ 0.83]\\ 0.30 \ [-0.28, \ 0.88]\\ 0.39 \ [-0.29, \ 1.66]\\ 0.52 \ [-0.22, \ 1.26]\\ 0.52 \ [-0.22, \ 1.26]\\ 0.52 \ [-0.00, \ 1.04]\\ 0.53 \ [-0.01, \ 1.32]\\ 0.58 \ [-0.03, \ 1.20]\\ 0.58 \ [-0.03, \ 1.20]\\ 0.58 \ [-0.03, \ 1.20]\\ 0.65 \ [-0.13, \ 1.18]\\ 0.69 \ [-0.88, \ 1.61]\\ 1.23 \ [-0.48, \ 1.98]\\ 1.40 \ (-0.53, \ 2.92]\\ 2.77 \ [-1.41, \ 4.14]\\ 2.98 \ [-5.4, \ 4.43]\\ 4.24 \ [-1.90, \ 6.58]\\ 6.17 \ [-3.24, \ 9.09]\\ 7.08 \ [-3.30, \ 10.86]\\ 0.51 \ [-0.32, \ 0.70]\\ \hline\end{array}$
		Effect Size		

DA transporters. A meta-analysis across DAT-binding drugs identified a significant common effect of drug over placebo on discounting (Q = 163.1, p < .001, $I_2 = 87.2\%$; Cohen's d = -.340, SE = .159, 95% CI [-.651, -.028]. Across effects, DAT-binding drugs decreased discounting over placebo. Egger's test for plot asymmetry did not suggest the presence of publication bias (z = -.414, p = .679). See forest plot on Figure 5 and funnel plot on Supplementary Figure S6.

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Figure 5. Forest plot of placebo-controlled effect of DAT modulation on reward discounting.

Higher values indicate increased discounting on drug.

DA Transporter				
Author, Year	Drug		Ν	Effect Size [95% CI]
0				4 0 0 1 5 0 2 0 5 4 1
Sommer, 2014.1	MRZ 9547		10	-4.23 [-5.93, -2.54]
Li, 2015	Cocaine	⊢● −1	16	-1.50 [-2.22, -0.79]
Sommer, 2014.4	Methylphenidate		11	-1.47 [-2.46, -0.48]
Baarendse, 2012.1			14	-1.29 [-2.00, -0.58]
Wiskerke, 2011	d-Amphetamine		14	-1.17 [-1.84, -0.49]
Mai, 2003	d-Amphetamine		11	-1.14 [-1.90, -0.38]
Sommer, 2014.2	MRZ 9546	⊢ ◆ −1	10	-1.13 [-2.09, -0.17]
Mai, 2015.1	d-Amphetamine	⊢ •−1;	10	-1.12 [-1.91, -0.33]
Sommer, 2014.5	d-Amphetamine	⊢ ∙ -i	10	-0.93 [-1.86, 0.00]
van Gaalen, 2006	Methylphenidate	⊢♦ -1	16	-0.88 [-1.46, -0.31]
Yohn, 2016.1	GBR12909	H+H :	18	-0.88 [-1.42, -0.33]
Wade, 2000	d-Amphetamine	H+H	17	-0.82 [-1.37, -0.27]
Randall, 2015	Bupropion	H o H	42	-0.73 [-1.07, -0.39]
Baarendse, 2012.2	d-Amphetamine	⊢ • · ·	14	-0.51 [-1.07, 0.04]
Mai, 2015.2	Cocaine	ı– ♦ -i	10	-0.48 [-1.13, 0.18]
St. Onge, 2009	d-Amphetamine	⊢	8	-0.40 [-1.13, 0.32]
Orsini, 2017.1	d-Amphetamine	⊢ ♦ ;I	15	-0.30 [-0.82, 0.22]
St. Onge, 2010.2	d-Amphetamine	⊢♦H	15	-0.29 [-0.81, 0.23]
St. Onge, 2010.3	d-Amphetamine	⊢ • ;⊣	12	-0.14 [-0.71, 0.43]
Barbelivien, 2008	d-Amphetamine	I∳I	31	-0.10 [-0.45, 0.25]
Floresco, 2008.2	d-Amphetamine	⊢ ∔ – I	8	-0.04 [-0.73, 0.66]
Siemian, 2017 c	I-Methamphetamine	⊢ ∳ -1	8	0.03 [-0.67, 0.72]
Koffarnus, 2011.2	GBR 12909	⊢ ∳ –I	12	0.14 [-0.43, 0.71]
Koffarnus, 2011.1	d-Amphetamine	i. i.t. ← -i	12	0.35 [-0.23, 0.93]
Sommer, 2014.3	Modafinil	⊢∔ ⊷-1	11	0.36 [-0.51, 1.22]
St. Onge, 2010.1	d-Amphetamine	i , i ♦ I	16	0.37 [-0.14, 0.88]
Zeeb, 2016	d-Amphetamine	i ✦✦I	30	0.44 [0.07, 0.82]
Orsini, 2017.2	d-Amphetamine	i . ← 1	11	0.51 [-0.12, 1.14]
Floresco, 2008.1	d-Amphetamine	i <u>+</u> ● - 1	8	0.55 [-0.20, 1.29]
Hernandez, 2014	Cocaine		6	1.13 [0.11, 2.16]
Cousins, 1994	Amphetamine		7	2.01 [0.72, 3.30]
Yohn, 2016.2	Atomoxetine	⊢ +−−1	8	2.31 [0.98, 3.64]
RE Model		•		-0.34 [-0.65, -0.03]
	Г—			
	-6.0	-4.0 -2.0 0.0 2.0 4.0		
	0.0			
		Effect Size		

Random-effects meta-analysis – exploratory analyses

Reward cost type. A model that included the interaction between reward cost type and drug type suggested a significant moderation effect ($Q_{Moderator} = 56.0$, p < .001, $I_2 = 77.6\%$). Inspection of the coefficients revealed a main effect within studies of effort discounting (Cohen's d = .915, SE = .338, p = .007, 95% CI [.252, 1.58]) and time discounting (Cohen's d = .742, SE

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= .252, p = .003, 95% CI [.248, 1.24]), but not probability discounting (Cohen's d = .153, SE = .190, p = .423, 95% CI [-.220, .526]), and a main effect of DAT-binding drugs (Cohen's d = - 1.30, SE = .413, p = .002, 95% CI [-2.11, -.488]). There were no significant interactions between reward cost type and drug type.

Rodent strain. A model that included the interaction between rodent strain group and drug type suggested a significant moderation effect ($Q_{Moderator} = 71.6$, p < .001, $I_2 = 71.96\%$). Inspection of the coefficients revealed a main effect of DAT-binding drugs (*Cohen's d* = -3.76, SE = .552, p < .001, 95% CI [-3.15, -.988]) and a significant interaction between DAT-binding drugs and Long Evans rodent strain (*Cohen's d* = 1.97, SE = .599, p < .01, 95% CI [.793, 3.14]). A follow-up model that tested the effect of rodent strain within DAT-binding drugs suggested a significant effect (Q = 14.3, p < .007, $I_2 = 83.1\%$) and revealed that discounting for Lister Hooded (Cohen's *d* = -.896, SE = .352, *p* = .011, 95% CI [-1.59, -.207]) and Sprague-Dawley (Cohen's *d* = -.436, SE = .220, *p* = .048, 95% CI [-.868, -.004]) rats was significantly reduced on drug over placebo. DAT effects for Wistar (Cohen's *d* = -1.02, SE = .546, *p* = .061, 95% CI [-2.09, .049]) and Long Evans (Cohen's *d* = .134, SE = .232, *p* = .562, 95% CI [-.320, -.588])

Drug injection location. From studies that directly injected a DA drug in the brain (k=28), a model that included the interaction between drug type and injection location (coded as cortical or within the nucleus accumbens) suggested a weak significant moderation effect (Q_{Moderator} = 11.3, p = .046, I_2 = 38.7%). Inspection of the coefficients revealed a main effect within studies injecting a drug into the cortex (Cohen's d = .398, SE = .141, p = .005, 95% CI [.121, .674]) but not the nucleus accumbens (Cohen's d = .229, SE = .275, p = .404, 95% CI [-

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.309, .767]). There were no significant interactions between injection location and drug type. See forest plot on Supplementary Figure S8.

Anti-Parkinson dose effects. For drugs commonly prescribed to treat Parkinson's Disease, the levodopa equivalent dose (LED) was estimated for effect sizes associated with D2 agonists (k = 6) and presynaptic DA agonists (k = 1). A model testing the effect of the continuous LED covariate suggested a significant moderation effect (QModerator = 11.54, p < .001, I_2 = 75.6%). There was a significant negative correlation between LED and reported effect-size (Cohen's d = -.016, SE = .005, p < .001, 95% CI [-.025, -.007]). See meta-regression correlation plot on

Supplementary Figure S9.

Anti-psychotic dose effects. For drugs commonly prescribed to treat psychotic symptoms typically present in schizophrenia and bipolar disorder, the chlorpromazine equivalent dose (CPZ) was estimated for effect sizes associated with D2 antagonists (k = 21). A model testing the effect of the continuous CPZ covariate did not suggest a significant moderation effect (Q_{Moderator} = 1.08, p = .299, I_2 = 89.9%). There was a non-significant positive correlation between CPZ and reported effect-size (Cohen's *d* = .017, SE = .017, *p* = .299, 95% CI [-.015, .050]). See meta-regression correlation plot on Supplementary Figure S10.

Discussion

Across 117 effects in rodents, a confirmatory quantitative meta-analysis suggested that: (1.) DAT-modulating drugs decrease discounting, (2.) D1R and D2R agonists do not impact discounting, and (3.) D1R and D2R antagonist moderately increase discounting. **See summarized effect size comparison in Figure 6**.

D1-like and D2-like receptors

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The D1-like and D2-like receptor-mediated effects challenge long-held views about DA receptor function. Traditionally it has been presumed that activation of D1 and D2 receptors produces opposing effects (upregulation versus downregulation of intracellular signaling to increase cAMP levels, respectively) in order to support different behaviors [28]. These results suggest, however, that DA may influence discounting behavior via these receptors in a similar manner. This is consistent with an emerging perspective that D1 and D2 receptors engage dissociable processes in the dorsal striatum but not the ventral striatum [29–32]. A number of convergent findings indicate these receptors are not functionally dissociable in the accumbens for specific motivated behaviors in the ventral striatopallidal pathway [29,33,34]. It is therefore possible that reward valuation mechanisms that support discounting behavior are reflected in ventral striatal signaling where activation of D1 or D2 receptor types play similar roles. On balance, the meta-analytic effects for D1 and D2 receptors do not support a view of opposing direct/indirect pathway function in mediating discounting behavior.

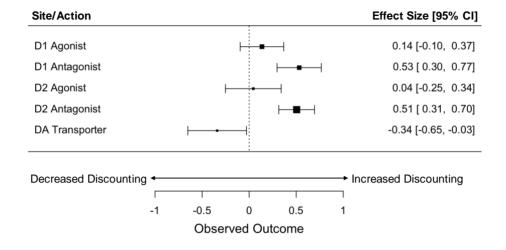
DA transporters

The meta-analysis showed that modulation of dopamine transporters moderately decreased discounting behavior. This effect is consistent with therapeutic management of ADHD symptoms from DAT-binding drugs and provides additional insight on the role of DAT-mediated extracellular DA increases on motivated behavior [35–38]. Specifically, the data suggest increased DAT-mediated DA efflux from presynaptic terminals is associated with greater patience, risk aversion, and willingness to expend effort for rewards. These effects are consistent with observations that acute increases in dopamine availability increases motivational vigor for rewards [39].

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Figure 6. Summary plot comparing effect sizes for the various dopamine drug targets and action



on discounting. Higher values indicate increased discounting on drug.

Reward cost type

Lesion and pharmacological studies in rodents have shown that mesolimbic DA similarly impacts probability [7] and effort [15] discounting. Complicating this, though, one study has shown that physical and not cognitive effort discounting is modulated by pharmacological stimulation of mesolimbic DA [8]. Accordingly, while discounting may exhibit some domain general value processing across cost types, there may be subtle differences in how DA function uniquely accounts for effort requirements in reward preferences. Although it has been assumed that probability and time are discounted similarly [40], there is evidence that increasing reward magnitude contributes to decreased discounting over time but increased discounting over probabilities [41]. Prior work in humans suggests that the relationship between dopamine function and discounting may vary by cost (time delay, probability, or effort) [6]. Although there was not evidence for differential drug effects across different cost domains in the present meta-

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analysis, these analyses may have been underpowered. The null effects here do not add substantial evidence to the question of overlap or dissociation in discounting different cost types. *Rodent Strain*

An exploratory meta-regression identified an interaction between drug binding site and strain for DAT-binding drugs. Specifically, within DAT-binding drugs, Wistar, Lister Hooded, and Sprague Dawley, but not Long Evans rats decreased discounting. The order of rodent strain effects (from decreased to increased discounting) indicated that Wistar > Lister Hooded > Sprague Dawley > Long Evans. It has been reported that dopaminergic differences exist between strains [42–45]. Consistent with the meta-analytic effect, Wistar rats have been shown to exhibit higher levels of DAT than Sprague Dawley rats [42]. In addition, inter-strain differences in traits that have been known to covary with dopamine function and motivation like body fat distribution may account for the observed effects [46].

Drug injection location

An exploratory meta-regression showed that regardless of drug site and action, administration of substances directly into the cortex had a greater impact on increasing discounting than those injected in ventral striatum. In both humans and rodents, lesion studies support the importance of both the vmPFC and ventral striatum in discounting behavior. Rodents with lesions to the mOFC and NAcc and human patients with vmPFC/OFC lesions discount monetary and food rewards more steeply [47–49]. Future work should evaluate whether appreciable differences exist in cortical versus striatal dopaminergic signaling on discounting behavior.

Dose-dependent effects

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Prescription of antipsychotic and anti-Parkinson medications, have known side effects including changes in impulse control which may be the result of variation in reward discounting functions [50]. One source of these effects might arise from the prescribed dose. Exploratory meta-regressions testing whether drug dose moderates the effect size of antipsychotic medications and anti-Parkinson medications on discounting revealed that variation in presynaptic rescue drug doses were negatively association with effect sizes. Specifically, studies using lower levodopa equivalent doses (LED) reported higher discounting on drug while studies using higher LED reported lower discounting on drug. This dose-dependency enhances our understanding of the linearity of dopamine effects. It should be cautioned, however, that these dose equivalencies are based on clinical use in humans. Nevertheless, since only some drugs have known LED or CPZ conversion rates, future work should seek to identify an expanded set of dose equivalencies. *Caveats, Limitations, and Concluding Remarks*

Although the meta-analysis yielded important insight about dopaminergic drug effects on discounting behavior, features of the literature limit our ability to make exact claims about function. First and foremost, although we have a clear understanding of where dopaminergic binding sites are across the brain, drugs do not naturally bind to specific regions. This may be problematic since dopamine receptor signaling varies between the striatum, midbrain, and cortex [51,52]. In addition, although the meta-analysis can explain choice behavior, it cannot speak to the specific value function that supports the underlying behavior. More specifically, it is unclear whether differential modulation of the dopamine system shifts how much weight an animal places on reward magnitudes or costs. This is important because work in rodents and humans suggests that costs and magnitudes are anatomically dissociable [53,54] and one study has shown that methylphenidate effects on humans' subjective value of cognitive effort depends on caudate

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DA synthesis capacity [39]. Future work should therefore evaluate whether dopaminergic drugs are more strongly modulating preferences by altering computations supporting integration of costs, magnitudes, and subjective value.

Since many dopaminergic drugs exhibit non-negligible binding to serotonin, norepinephrine, and adrenergic receptors and transporters, effects cannot be exclusively attributed to dopamine function [55,56]. In addition, it is important to acknowledge that acute administration of dopaminergic drugs have been demonstrated to have different effects on neurotransmission from chronic administration. For example, acute haloperidol administration contributes to higher spontaneous firing of dopamine neurons than chronic administration [57]. Whereas acute administration of psychostimulants increase dopamine release, PET studies of humans with psychostimulant addictions have shown that chronic use contributes to reduced dopamine release, transporter availability, and D2 receptor availability [58].

The meta-analysis is further limited by the scarcity of human studies. Moreover, the literature search and meta-analysis was limited to studies of discounting in healthy animals. While this decision was made to isolate drug effects from disruptions in behavior due to lesions or psychopathology, it is possible that drugs may impact animals depending on systemic alterations to circuits from diseases such as Parkinson's Disease or schizophrenia [5,6]. An additional caveat raised by the exploratory meta-regressions suggests that drug doses may partially account for reported effect sizes. Since the meta-analysis data extraction protocol was limited to selection of the highest dose effect when multiple doses were available, the effects cannot adequately account for non-linear effects of drug dose on discounting.

While this meta-analysis was limited to discounting paradigms, future ones should evaluate how dopamine pharmacology effects differ for other reward-related behaviors. For example,

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many studies have evaluated dopaminergic drug effects on reinforcement and probabilistic reversal learning. It is unknown whether consistent cross-study patterns of results with respect to D1 and D2 receptors would emerge given that discounted value representations may reflect the same updated valuation process in reinforcement learning, relying on signaling in the ventral striatum and medial prefrontal cortex [59]. In general, we hope this meta-analysis encourages additional meta-analytic work in behavioral pharmacology and provision of publicly available data. The present meta-analysis here contributes to our understanding of how dopamine signaling mediates preferences for delayed, effortful, or uncertain reward outcomes.

Supplementary Information accompanies this paper at: https://osf.io/27cqw/

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