Ancient Admixture into Africa from the ancestors of non-Africans

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Abstract

Genetic diversity across human populations has been shaped by demographic history, making it possible to infer past demographic events from extant genomes. However, demographic inference in the ancient past is difficult, particularly around the out-of-Africa event in the Late Middle Paleolithic, a period of profound importance to our species' history. Here we present **SMCSMC**, a Bayesian method for inference of time-varying population sizes and directional migration rates under the coalescent-with-recombination model, to study ancient demographic events. We find evidence for substantial migration from the ancestors of present-day Eurasians into African groups between 40 and 70 thousand years ago, predating the divergence of Eastern and Western Eurasian lineages. This event accounts for previously unexplained genetic diversity in African populations, and supports the existence of novel population substructure in the Late Middle Paleolithic. Our results indicate that our species' demographic history around the out-of-Africa event is more complex than previously appreciated.

1 Introduction

Methods to characterize demographic history from genetic data alone form a useful and independent complement to archaeological approaches [40], and many methods have been developed for this purpose [11, 14, 20–22, 30, 34, 43, 47, 52, 55, 56]. A particularly interesting time in human history is the end of the Middle Paleolithic (~60 thousand years ago (kya)), which saw the divergence of the most deeply sampled lineages of human genetic variation, introgression from multiple archaic sources, and the expansion of anatomically modern humans Out of Africa (OoA). As archaeological evidence and ancient DNA from this period are scarce, inference of demography from present-day genetic data is potentially very informative, though technically challenging.

The phylogenetic trees over a set of samples as they change along the genome through recombination, collectively referred to as the ancestral recombination graph (ARG) [2, 21], record all information about the samples' evolutionary history. This history itself is shaped by the population's demography, a statistical relationship that is quantified by the coalescent-with-recombination (CwR) model [2]. The ARG is a complex data structure which is only weakly constrained by the observed genetic polymorphisms, making inference of demography difficult. By making approximations to the CwR, for instance by making an independent-sites assumption, efficient parametric inference of demography becomes possible [6, 18]. Methods including **PSMC** [11], **diCal** [30] and **SMC++** [31] allow non-parametric inference of demography under a closer approximation to the CwR, but one that does not include gene flow between populations. **MSMC** [22] introduced the cross-coalescent rate and **MSMC-IM** described how to interpret this rate in the context of a isolation-migration

model to estimate a migration rate between populations[56]. However, these methods are not well suited for estimating directional migration rates.

Here we extend **SMCSMC** (Sequential Monte Carlo inference of the Sequentially Markovian Coalescent, [44]) to allow inference of directional migration. **SMCSMC** is a Bayesian method that uses a particle filter to explicitly sample from the posterior distribution of ARGs over multiple diploid samples under the full CwR model. Since particle filters operate by simulating latent variables (here the ARG) under the statistical model of interest, it becomes possible to handle complex demographic scenarios. We exploit this by extending the CwR model to include time-varying directional migration rates in a two-island demographic model. We use the posterior sample of ARGs including migration events to update the parameters of the demographic model, using either expectation-maximization or a variational Bayes procedure, and iterate these steps until convergence. We apply **SMCSMC** to estimate directional migration rates in whole genome sequencing data from the Simons Genome Diversity Panel (SGDP) [35] and the Human Genome Diversity Panel (HGDP) [48] to investigate population structure around the OoA event.

2 Results

Substantial Migration from Eurasian to African Ancestors We use **SMCSMC** to analyse pairs of individuals from the SGDP and simultaneously infer migration rates and effective population sizes (N_e) under a two-island model with directional migration. Population sizes and migration rates are modeled as piece-wise constant across 32 exponentially spaced epochs from 133 to 133016 generations in the past, corresponding to 3.8 thousand to 3.8 million years ago (3.8kya–3.8Mya) using a generation time $g = 29$ years [5]. We find that the method infers high rates of migration from descendants of the OoA event (‘non-Africans’) to Africans, but not in the opposite direction, in the period 30–70kya corresponding to the Late Middle Paleolithic (Fig. 1). In populations from the Niger-Kordofanian and Nilo-Saharan language groups, comprising the majority of the population on the African continent, the peak inferred migration rate from Eurasian populations ($2.5\text{--}3.0 \times 10^{-4}$ and $3.5\text{--}4.0 \times 10^{-4}$, in units of proportion of the target (ancestral African) population replaced per generation) most frequently falls in the epochs spanning 35–45kya, while peak migration rates in the opposite direction are substantially lower ($0.5\text{--}1.0 \times 10^{-4}$) and occur earlier, in the epochs spanning 55–70kya (Supplemental Fig. S1). Populations in the Afroasiatic language group show evidence of large amounts of directional migration in the Holocene (Supplemental Fig. S2), which is consistent with previous findings of relatively recent European introgression into these populations [32, 50].

To assess the impact of errors introduced by statistical phasing, which was used to phase the SGDP data, we repeated the analyses above on a subset of physically phased individuals from the Human Genome Diversity Project (HGDP) [35] (Supplemental Section S3.2). This data set comprises individuals from four African (Yoruban, San, Mbuti, and Biaka) and nine non-African populations (Druze, Han, Karitiana, two Papuan populations, Pathan, Pima, Sardinian, and Yakut). **SMCSMC** results in the HGDP are qualitatively similar to those in the SGDP (Fig. 1b, Supplemental Fig. S1a). Inferred migration rates are, in general, lower in HGDP data than when using matched SGDP samples (Figs. 1a,b and S4, demographic inference in matched sampled in Supplemental Fig./ S3), but in all cases, the migration rates from Eurasia to Africa are substantially higher than in the opposite direction, consistent with the findings in the SGDP (Supplemental Fig. S4).

We asked whether **SMCSMC** has power to detect a large back-migration event in the Late Middle Paleolithic and distinguish it from other demographic scenarios. To answer this we used **SCRM** [29] to simulate a gigabase of sequence data under a two-island demographic model, with effective population sizes chosen to be comparable to typical African and Eurasian populations as inferred from real data. To this we added a 10ky pulse of forward, backward or bidirectional migration of varying strengths, with the midpoint of the migration pulse within the range 40 to 70kya. To quantify the inferred amount of migration we calculate the integrated migration fraction (IMF), defined as one minus the probability that a lineage in the destination (e.g. African) population traced backwards in time remains in that population across a given epoch according to the migration model (see Methods). For the simulations, we chose the most recent 100kya as epoch, and used scenarios with IMFs ranging from 0 to 0.593. For each simulation we report the inferred IMF in both

the forward and backward direction (Fig. 2); full results are given in Supplemental Section S5. We find that **SMCSMC** has good power to detect backward migration pulses up to 60kya (median ratio of inferred and true IMF, 0.91), while power drops off at 70kya (IMF ratio 0.46). In the pure backward migration case, some forward migration is falsely inferred, but this is always substantially less than the inferred backward migration (median ratio inferred forward to true backward IMF, 0.37; true migration peak ≤ 60 kya). However, in the case of true forward migration as well as bidirectional migration, roughly equal mixtures of forward and backward migration are inferred (Fig. 2). We conclude that in the epoch 40–70kya the forward and bidirectional scenarios are difficult to distinguish from each other, but both can be distinguished from backward migration, the only scenario resulting in substantially different inferred backward and forward migration.

To validate the existence of the migration pulse, though not its direction, we next analyzed the same data using **MSMC**, which is widely used to estimate gene flow in the ancient past by estimating the relative cross-coalescent rate (RCCR) between two populations [22, 26, 28, 50]. We use the updated implementation **MSMC2** recommended by the authors and first published in [34]. Each of the **SMCSMC** analyses are repeated using **MSMC2** to estimate effective population size and RCCR (Supplemental Figs. S2, S4, S5). Consistent with previous analyses conducted with **MSMC2**, our estimates show high RCCR in the Late Middle Pleistocene in both the SGDP and the HGDP (Fig. 1c,d) [48, 50]. These observations confirm the existence of a substantial pulse of ancient gene flow between Eurasians (Han Chinese) and Africans.

Migration Pre-dates East-West Eurasian Divergence To assess whether the inferred back-migration shows variation across the descendants of the OoA event, we repeated the analyses using three representative non-African groups in the SGDP: Han Chinese, French European, and Papuans. Since simulations show that **SMCSMC** has little power to detect migration predating 70kya, and to exclude Holocene migration, the epoch we use to calculate real-data IMFs comprise the period of peak inferred migration up to the period of diminishing power (30–70kya); we use this epoch for all subsequent analyses. Inferred IMFs are not significantly different between Han Chinese and European populations in non-Afroasiatic populations ($p=0.14$, two-tailed paired t-test; Figs. 1h and S6, Table S1), consistent with migration occurring before the European-East Asian split approximately 40kya [20]. The contribution of this admixture event to extant African genetic variation is substantial; the estimated IMFs indicate that for individuals in the major African language groups, approximately a third of ancestral lineages trace their ancestry through the proto-Eurasian population (Niger-Kordofian group, 0.35 ± 0.04 ; Nilo-Saharan groups, 0.41 ± 0.03 ; Table S1). When we estimate these proportions using a Papuan sample to represent non-African descendants we find slightly but significantly smaller values compared to estimates using either the Han Chinese or European populations (mean difference of 0.029 ± 0.002 , $p=9.2 \times 10^{-15}$, and 0.025 ± 0.004 , $p=2.3 \times 10^{-10}$, paired t-tests, Supplemental Table S1, S2). Similarly, in the HGDP, inferred migration in both Papuan groups (Sepik and Highlands) was 0.025 ± 0.004 ($p=1.4 \times 10^{-6}$) lower than French and Han (Supplemental Table S3). We comment on this observation in the Discussion.

Directional Migration Explains Excess Inferred African Genetic Diversity 100kya Previous studies looking at effective population sizes (N_e) in human ancestral populations have consistently reported inflated inferences in African populations approximately 100kya, often hypothesized to be due to unaccounted-for population substructure within Africa [11, 22]. We use **SMCSMC** to analyze African individuals paired with an individual from one of three non-African populations (Han Chinese, French European, and Papuans) and infer N_e for the African ancestral population under a two-island model with directional migration. Each analysis was repeated three times to assess the contribution of stochastic sampling to the inferences (Figs. 3, S5, per population N_e in Supplemental Fig. S7). **SMCSMC** infers substantially lower African N_e than **MSMC** in the period 80kya–300kya. In addition, while **MSMC** inferences show convergence of African and Eurasian ancestral N_e estimates only around 300kya, inferences from **SMCSMC** indicate convergence at 150kya (Fig. 3a), closer to the hypothesized time of the diversification of the ancestral lineages prior to the main out-of-Africa migration episode [34, 38]. The same analysis on physically phased samples from HGDP show that these results are not driven by errors due to statistical phasing (Fig. S4 and Supplemental Section S3.2). When we

used **SMCSMC** to infer both African and European N_e under a single-population model without migration, N_e estimates were comparable to those from **MSMC** (Fig. 3b), indicating that the **SMCSMC** inferences are not driven by methodological biases particular to **SMCSMC**.

To more directly support the interpretation that the lower African N_e inferred by **SMCSMC** is due to appropriate modeling of directional migration, we again used coalescent simulation with **SCRM** to investigate various migration scenarios and their effects on inferred African N_e . Using the simulation framework as above, we examine N_e estimates inferred under a two-island model with migration, and in addition N_e separately inferred for each of the two simulated populations under a single-population model (Supplemental Section S5). Focusing on single-population inferences, we found that for simulated African populations that had received substantial migration from the simulated Eurasian population either through backward or bidirectional migration, inferred N_e values indeed were substantially inflated compared to true values (Fig. 3c,d), while this effect was not seen when forward (African-to-Eurasian) migration was simulated (Fig. 3e). Similarly, single-population Eurasian N_e estimates were inflated in the presence of forward and bidirectional migration, but not backward migration (Supplemental Figs. S8–S10). In contrast, when using a model that includes migration, inferred African N_e do not show inflation in any of the three scenarios (Fig. 3c-e). We conclude that the inferences from **SMCSMC** and **MSMC** are compatible with substantial back-migration from ancestral Eurasians into Africans, but not substantial bidirectional or forward migration.

Less Gene Flow to Central and South African Hunter-Gatherers We infer substantial Eurasian back-migration into all African groups, however the inferred IMFs for individuals from Khoe-San populations are significantly lower than for any other group (difference with Niger-Kordofians, 0.14 ± 0.02 , $p = 4.4 \times 10^{-14}$; difference with Nilo-Saharans, 0.20 ± 0.03 , $p = 6.9 \times 10^{-9}$, two-tailed t-test, Table S4). To further support this observation we used **MSMC** to estimate the relative cross-coalescent rate (RCCR) for several populations, and find evidence for gene flow between Yorubans and Eurasians that is not shared with the Khoe-San individuals in either the SGPD and the HGDP (Fig. 1c,d). These results are consistent across Eurasian donor populations (Fig. S3). The Khoe-San individuals are particular outliers, whose ancestors are inferred to have experienced approximately half the amount of admixture seen in Nilo-Saharan and Niger-Kordofanian groups (Fig. S6). In addition, we find that the Mbuti and Biaka, both Central African hunter-gatherer populations, show levels of Eurasian gene flow that are intermediate between levels observed in the Khoe-San and Yorubans (Fig. 1a,b, Supplemental Table S1). This is mirrored by inferred IMFs for Central African Hunter Gatherers, which are significantly lower than other Niger-Kordofanian groups (difference -0.08 ± 0.03 , $p = 1.2 \times 10^{-3}$, Table S4), possibly reflecting the proposed early split times of the Mbuti and Biaka from the remainder of ancestral African populations between 60 and 200kya [41, 53].

No Evidence for Excess Neanderthal Ancestry Previous studies have proposed that a backflow from Eurasia may have brought Neanderthal ancestry into African populations [57]. To assess whether the proposed Late Middle Paleolithic back migration might have introduced Neanderthal material, we analyzed a Yoruban and a French individual using **SMCSMC** to draw a sample from the posterior distribution of ARGs, isolated the marginal trees containing an inferred back-migration event in the epoch 30–70kya, and reported the inferred admixture tracts (“segments”, Supplementary Section S4). To assess whether the identified segments are plausible, we confirmed that their length distribution is consistent with IMF and timing of the migration inferred by **SMCSMC** (Supplemental Section S4.1, S11), and, as expected, we found that these African segments with putative Eurasian ancestry tend to be more closely related to a Eurasian sample than another representative of the same African population (Table S5, Supplemental Fig. S12, S13) in a global dataset of modern and ancient individuals compiled by the Reich group (see URLs). Within these African segments that are likely enriched for material with Eurasian ancestry, we then used D statistics [13] to identify enrichment for Neanderthal material compared to an African background. We find no evidence for gene flow with a Vindija Neanderthal on the Mbuti baseline, or when compared to a different Yoruban (Table S6, S7). We additionally find no evidence for increased affinity to the Vindija Neanderthal when compared to the Altai, as would be expected if the material were descended from admixing Eurasians (Table S8). However, we find that restricted to the identified segments, D statistics have power to detect evidence

for the known admixture from Vindija into a French individual (Fig. 4), suggesting that lack of power does not explain the lack of evidence we find for Neanderthal admixture into Africans. In addition, we find no differences in affinity to Neanderthals or Denisovans between the variants which fall in segments and the whole genome (Fig. 4d). Taken together, this suggests that Eurasian-derived segments of the African genomes are not enriched with Neanderthal material.

3 Discussion

We have developed an approach for estimating demographic parameters and ARGs from whole genome sequence data, which can handle inference in complex demographic models, and implemented this in the software program **SMCSMC** [44]. We used **SMCSMC** to investigate ancient migration rates and population substructure, and found evidence for a substantial admixture from ancestors of present-day Eurasian populations into African populations in the Late Middle Paleolithic.

Our analysis suggests that a population ancestral to present-day Eurasians contributed as much as a third of the genetic material in many modern African populations. We find no difference in inferred admixture proportions when using French Europeans or Han Chinese as extant representatives of the donor population, indicating that the admixing population must have split from the out-of-Africa population before the East/West Eurasian divergence, implying a lower bound on the timing of the admixture of approximately 40kya [20]. It appears that our results suggest that the migrating population was more similar to present-day French and Chinese populations than to Papuans. However, up to 5% of the genomes of some present-day Papuans have been suggested to derive from archaic introgressions [37], and these contributions will have reduced the inferred levels of admixture into Africans when using Papuans as a representative of the Eurasian ancestors. The alternative explanation, of an earlier divergence of Papuans and Eurasian ancestors, is possible but contested; in light of documented Eurasian admixture into Oceania, the effects of this early isolation are likely to be small relative to the large confounding effects of Denisovan admixture [34, 40].

The proposed period of admixture has biased previous inferences of the African population sizes. We show that including directional migration into the model resolves previously unexplained high inferred N_e in the period 80 to 300kya. It is well known that effective population size estimates are biased in the presence of population substructure and migration [11, 43]. We use simulations to show that the proposed admixture event indeed causes an increase in estimated N_e in analyses that do not explicitly model migration. Correctly modeling of directional migration recovers the correct N_e , and allows us to infer a more recent split time between the two populations than indicated by previous analyses, although we did not attempt to formally estimate this time of divergence.

We found that not all populations in Africa have been equally affected by the proposed migration event. While the ancestors of Niger-Kordofanian and Nilo-Saharan populations show evidence of similar levels of Eurasian admixture, the ancestors of Central African and South African hunter-gatherer populations show markedly lower levels. The date of genetic diversification of both the Central Hunter Gatherers and Khoe-San (SAHG) is contested [53], but a date of 100kya has been proposed [17], providing a putative upper bound on the main admixture event. Our simulations indicate that **SMCSMC** has little power to detect the impact of migration events occurring more than 70kya, providing an additional upper bound on the time of the migration episode, or the fraction of it that left a sufficiently distinct imprint on extant genetic material.

Compared to the remainder of the Niger-Kordofanians and Nilo-Saharans on the one hand, and the SAHG populations on the other, the Mbuti and Biaka show intermediate levels of admixture. Of these populations, the Biaka show slightly higher levels of admixture than the Mbuti, which is likely due to the well-documented admixture from Western African groups not shared with the Mbuti [10]. The lower levels of admixture in Mbuti and Biaka compared to Niger-Kordofian and Nilo-Saharan populations imply at least partial diversification of the former at the time of the migration, placing an upper bound on the timing. However, dating the diversification of these groups is difficult. Recent estimates using f statistics place the split concurrent with the San in a large-scale early expansion 200-250kya [53], while older data consistently report an earlier split time between 50 and 90 kya [45]. Further clarity on the early structure and diversification of hunter-gatherer populations are necessary to interpret their interactions with Eurasian

migrants. The Afroasiatic populations on the other hand show high levels of admixture, which also appears to be of much more recent origin, and it appears likely that this is the result of extensive admixture from Eurasian populations during the Holocene [32, 50].

It has previously been suggested that Eurasian back-migration may be responsible for Neanderthal material in Africans [57]; however, we find no evidence for enrichment of Neanderthal-like material in putatively Eurasian-derived genomic segments in Africans, indicating that Neanderthal introgression into Eurasians occurred after the African introgression event we study here, or that further population structure in the Eurasian ancestral population precluded substantial transmission of Neanderthal material into Africa.

Our findings are consistent with several other published observations. Migration rate estimates using **MSMC-IM** revealed high levels of admixture at times comparable to our results [56]. The coalescent intensity function additionally shows similar histories between sub-Saharan African and Eurasian groups with high coalescent intensity in epochs consistent with our inference and those of **MSMC-IM**, supporting both an early split between the groups and a substantial replacement of genetic material more recently than $\sim 100\text{kya}$ [47]. Evidence has been mounting for multiple migrations into the Eurasian continent, possibly mediated by climatic drivers [36, 38]. Eurasian backflow during the Holocene has been well established [24, 25], but earlier migrations have also been proposed before based on observations of the spatial distribution of Y chromosome and mitochondrial haplogroups [1, 3, 4, 8, 33, 42, 51]. At the same time, evidence has been mounting for extreme heterogeneity in the history of sub-Saharan Africans, with several unsampled population theorised to have contributed at various points in the past [49, 53, 55]. In light of these recent studies, the observations in this paper add to a growing body of evidence for complex population structure and migration surrounding the Out of Africa event leading to a substantial replacement of the African population in the Late Middle Paleolithic.

4 Methods

A Particle Filter for Demographic Inference Details of the Sequential Monte Carlo for the Sequentially Markovian Coalescent (**SMCSMC**) algorithm have been previously published [44] (see the URLs for an implementation). Briefly, **SMCSMC** builds an approximation of the posterior distribution of genealogical trees along the genome using a particle filter, a method also known as sequential Monte Carlo. It does so by simulating a number of sequences of genealogical trees (particles) under a fixed set of demographic parameters θ , using the sequential coalescent sampler **SCRM** [29]. Simulated recombination events may change the local trees along the sequence. Particles are then weighted according to their conditional likelihood given observed polymorphisms. To avoid sample depletion, the set of particles is regularly resampled, which tends to remove and duplicate particles with low and high weight respectively. To further increase the efficiency of the procedure, the resampling procedure targets not the partial posterior distribution that includes polymorphisms up to the current location, but also includes a "lookahead likelihood" term that approximates a particle's likelihood's dependence on subsequent polymorphisms, while ensuring that the estimate of the posterior tree distribution remains asymptotically exact. From a sample of trees from the posterior distribution, Variational Bayes (VB) or Stochastic Expectation Maximization (SEM) is used to update the estimates of demographic parameters θ . This is repeated over a given number of iterations, or until θ have converged.

To add the ability to infer time-varying migration rates, we exploit the capabilities of **SCRM** to simulate ARGs under complex demographic scenarios, and collect sufficient statistics (migration opportunity, and number, time and direction of simulated migration events) for each particle.

We use **SMCSMC** to infer effective population sizes and migration matrices in pairs of unrelated individuals from the phased release of the Simons Global Diversity Panel. We set a uniform recombination rate of 3×10^{-9} and a neutral mutation rate of 1.25×10^{-8} , both in units of events per nucleotide per generation; previous results indicate that modeling recombination hotspots minimally affects results [11]. To reduce the number of iterations to convergence, we initialise the particle filter with an approximation of human demographic history (Supplemental Fig. S14). We seed the model with an initial constant symmetric migration rate of 0.0092 ($M_{i,j}$; proportion per generation of the sink population replaced by migrants from the source backwards in

time). We arrive at this value through simulation (Supplemental Section S5, Supplemental Figs. S15, S16).

Multiply Sequential Markovian Coalescent We use **MSMC2** to estimate the effective population size of pairs of African and Eurasian individuals using default configurations and scripts provided in **msmc-tools** (see URLs) [22, 56]. We use a fixed recombination rate in line with our **SMCSMC** analysis and skip ambiguously phased sites. Twenty iterations are performed by default. We additionally compute the relative cross-coalescent rate to examine relative gene flow by transforming the coalescent rates generated by **MSMC2** as indicated in the software documentation.

Coalescent Simulation Coalescent simulations were performed under the sequential coalescent with recombination model (**SCRM**) [29]. Full details of the simulation procedure are detailed in Supplemental Section S5. 1 gigabase (Gb) of sequence was simulated. In addition to branches in local genealogical trees, **SCRM** retains non-local branches in the ancestral recombination graph (ARG) within a user-specified sliding window. In the limit of a chromosome-sized windows **SCRM** is equivalent to the coalescent with recombination, while for a zero-length window it is equivalent to the sequentially Markovian coalescent (SMC') [6, 7]; we use a 100kb sliding window to approximate the CwR and improve accuracy over SMC' while retaining tractable inference.

We modelled migration as a 10ky pulse of constant migration rate resulting in an integrated migration fraction (IMF) of 0 to 0.593. The migration pulse was centered at various times between 40 and 70 kya. Due to the amount of compute required, we then used **SMCSMC** to infer the demographic parameters using a reduced set of 5000 particles and 5 iterations of the VB procedure. To aid convergence, we started inference at a reasonable approximation of human demographic history (see Supplemental Section S5.1, Supplemental Fig. S17). We modelled N_e and migration rates as piecewise continuous functions and set 32 exponentially spaced epochs from 133 to 133016 generations in the past. To convert evolutionary rates to years we set a generation time of 29 years [5]. For computational efficiency, individual genomes were split into 120 chunks and processed in parallel, with sufficient statistics collected and processed together in the VB steps.

Isolating Anciently Admixed Segments We sampled genealogical trees with migration events from the posterior distribution estimated by the particle filter under the final, converged, demographic parameters. We scan along the sequence and identified marginal trees with migration events from the source (Eurasian) population to the sink (African) population (forward in time) within the desired time period along with the beginning and end position of that tree in the genome sequence. In this process, we ignore recombination event that alter a tree in such a way that the migration event is retained.

Sequence Data and Preparation We downloaded whole genome sequence (WGS) data from the phased release of the Simons Genome Diversity Panel and converted it to **.seg** file format using scripts provided (See URLs). We apply two masks to the data. First, we mask the data with the strict accessibility mask provided by the 1000 genomes project (see URLs). Second, we mask any sites absent chimpanzee ancestry, to address a known variant issue in the data that resulted in artificially long runs of homozygosity [56]. We develop a **Snakemake** [12] pipeline for efficiently analysing sequence data with both **SMCSMC** and **MSMC2**. We assume a mutation rate of 1.25×10^{-8} and a recombination rate of 3×10^{-9} (events per nucleotide per generation), in line with recent literature [16, 23]. The number of particles, and the number of VB iterations, are set per analyses, and are reported in figure captions. Unless otherwise noted, the names of individuals used in this paper are the first in their population (e.g. an individual named Yoruban is **S.Yoruba-1** in the SGDP nomenclature); a complete list of sample identifiers is provided in Supplemental Table S9.

Formal Statistics Patterson's formal statistics were calculated with **ADMIXTOOLS** [13] and the **admixr** package [54] in R. We converted the above sequence data to Eigenstrat format with **vcf2eigenstrat** formerly distributed with **admixr**. We merged SGDP and archaic Eigenstrat datasets with **convertf** and **mergeit** implemented in **ADMIXTOOLS**.

Integrated Migration Fraction The IMF, the total fraction of a particular population A replaced during a particular time period from T_0 to T_1 generations in the past is found as follows. Let $\rho(t)$ be the instantaneous rate of migration out of A per unit of time in the backward direction (i.e. into A forwards in time), and $F(t)$ the fraction not migrated in the epoch $[T_0, t]$, then $\frac{d}{dt}F(t) = -\rho(t)F(t)$ with solution $F(t) = e^{-\int_{T_0}^t \rho(t)dt}$, so that the IMF is given by $1 - F(T_1)$. The integral is calculated as a finite sum since ρ is piecewise constant.

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URLs

Simons Genome Diversity Panel phased release, https://sharehost.hms.harvard.edu/genetics/reich_lab/sgdp/phased_data/.
 Human Genome Diversity Panel, ftp://ngs.sanger.ac.uk/production/hgdp/hgdp_wgs.20190516/.
 Ancient DNA, <http://cdna.eva.mpg.de/andertal/>.
 Strict 1000 Genomes accessibility mask, ftp://ftp.1000genomes.ebi.ac.uk/vol1/ftp/release/20130502/supporting/accessible_genome_masks/.
 SMCSMC, <https://github.com/luntergroup/smcsmc>.
 MSMC2, <https://github.com/stschiff/msmc2>.
 msmc-tools, <https://github.com/stschiff/msmc-tools>.
 vcf2eigenstrat, <https://github.com/bodkan/vcf2eigenstrat>.
 ADMIXTOOLS, <https://github.com/DReichLab/AdmixTools>.
 admixr, <https://github.com/bodkan/admixr>.

Figure 2: **Simulation study.** SCRM was used to simulate 1 gigabase sequence data for two diploid individuals under three different migration models. Migration was simulated backwards (from a Eurasian-like population to an African-like population), forward (the reverse), and symmetrically (equal migration in both directions). The amount of migration indicates the proportion of the sink population replaced by the source over a 10ky period centered at 40, 50, 60, or 70kya. The total IMF inferred by SMCSMC over the last 100ky is plotted and compared to the true simulated amount. For reference, the inferred IMF in either direction across 0-100kya for a Yoruban and Han individual is given in dashed lines. 5 iterations of variational Bayes and 5000 particles were used for inference. The effective population size model and additional details are given in Supplemental Section S5.

Figure 3: Effective population size inference. **a.** Analyzing a Nigerian Yoruban and a Han Chinese individual from the Simons Genome Diversity Panel jointly in a two-island model with directional migration using SMCSMC yields markedly lower N_e estimates and a more recent apparent split time, than when the same data are analyzed using MSMC with a model that does not explicitly include migration. Analyses for SMCSMC repeated three times; range of the estimates shaded. **b.** When each individual is analysed separately, using a model not including migration, N_e estimates from SMCSMC are similar to those of MSMC2. (Joint estimate from **a** included for comparison.) **c, d, and e.** Inferred Eurasian and African N_e from data simulated under a two-island model with, from left to right, a backward Eurasia-to-Africa, a bidirectional, and a forward migration pulse lasting 10ky (dashed vertical lines; same data as for Fig. 1e-g). Particularly for the backward migration case, inferred N_e under a two-island model tracks the true values (black) well, while inferred N_e under a single-population model are inflated around the split time. All SMCSMC analyses used 10000 particles and 25 variational Bayesian iterations; MSMC analyses used 40 iterations (Supplemental Section S5).

Figure 4: African introgressed segments are more similar to Eurasians but show no Neanderthal or Denisovan enrichment. **a.** $D(\text{Test Yoruban, Comparison Yoruban; Test population, Chimpanzee})$ calculated for all populations in the Reich Human Origins dataset (see URLs). D statistics in the putatively migrated segments are higher across the board in 3589 ancient, 6472 present day individuals. **b.** The same D statistic but computed for all African populations and individuals sampled from the Paleolithic. Neanderthal and Denisovan samples (marked with red bar) show low affinity to a Yoruban in putatively migrated segments. **c.** Histogram of D statistics computed in **a.** showing clear inflation of statistics calculated in segments (red) versus all markers (blue). **d.** Subset of individuals from **a.** involving Neanderthal ($n = 6$), Denisovan ($n = 1$), and a unique mixture individual ($n = 1$) with statistics calculated in segments (red) and all markers (blue) for all $n = 17$ African individuals indicating no difference in this population.

Figure 5: Proposed model for the diversification of modern human lineages. The proposed admixture from ancestors of present-day non-Africans is indicated in green. Several elements of this model remain unresolved, including the timing of the proposed admixture into Central and Southern Hunter Gatherer populations and these population's branching order, gene flow from an unsampled archaic African human lineage, and the different episodes of Neanderthal and Denisovan admixture. Date ranges for the diversification of Central and Southern African Hunter Gatherers [35], Out of Africa migration [53], Eastern / Western Eurasian diversification [22], Eastern / Western African diversification [50], Oceanian / Eastern Eurasian diversification [39], the split of the Mbuti from Biaka [35], Neanderthal gene flow [15], and Denisovan Gene flow [34] may be found in references given.

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