

1 Problem Solving Protocol

2 **Extended Application of Genomic Selection to Screen Multi-Omics Data**  
3 **for Prognostic Signatures of Prostate Cancer**

4 Ruidong Li<sup>1,2†</sup>, Shibo Wang<sup>1†</sup>, Yanru Cui<sup>3†</sup>, Han Qu<sup>1</sup>, John M. Chater<sup>1</sup>, Le Zhang<sup>2</sup>, Julong  
5 Wei<sup>4</sup>, Meiyue Wang<sup>1</sup>, Yang Xu<sup>5</sup>, Lei Yu<sup>1,2</sup>, Jianming Lu<sup>1,6</sup>, Yuanfa Feng<sup>1,6</sup>, Rui Zhou<sup>1,6</sup>,  
6 Yuhan Huang<sup>1,7</sup>, Renyuan Ma<sup>8</sup>, Jianguo Zhu<sup>9\*</sup>, Weide Zhong<sup>6\*</sup>, Zhenyu Jia<sup>1,2\*</sup>

7 **Affiliations:**

8 <sup>1</sup> Department of Botany and Plant Sciences, University of California, Riverside, CA, USA

9 <sup>2</sup> Graduate Program in Genetics, Genomics, and Bioinformatics, University of California,  
10 Riverside, CA, USA

11 <sup>3</sup> College of Agronomy, Hebei Agricultural University, Baoding, China

12 <sup>4</sup> Center for Molecular Medicine and Genetics, Wayne State University, Detroit, MI, USA

13 <sup>5</sup> Agricultural College, Yang Zhou University, Yangzhou, China

14 <sup>6</sup> Department of Urology, Guangdong Key Laboratory of Clinical Molecular Medicine and  
15 Diagnostics, Guangzhou First People's Hospital, School of Medicine, South China  
16 University of Technology, Guangzhou, China

17 <sup>7</sup> Department of Microbiology, Immunology, and Molecular Genetics, University of  
18 California, Los Angeles, CA, USA

19 <sup>8</sup> Department of Mathematics, Bowdoin College, Brunswick, ME, USA

20 <sup>9</sup> Department of Urology, Guizhou Provincial People's Hospital, Guizhou, China

21

22 \* To whom correspondence should be addressed.

23 † These authors contributed equally.

24 **Biographical Note:**

25 Zhenyu Jia is an Associate Professor at University of California, Riverside, USA

26 Weide Zhong is a Full Professor at South China University of Technology, Guangzhou,

27 China

28 Jianguo Zhu is a Urologist at Guizhou Provincial People's Hospital, Guizhou, China

29 Ruidong Li, Han Qu, Le Zhang, Lei Yu and Meiyue Wang are PhD students at University

30 of California, Riverside, USA

31 Shibo Wang and John M. Chater are Postdoctoral Fellows in Dr. Zhenyu Jia's lab at

32 University of California, Riverside, USA

33 Yanru Cui is an Associate Professor at Hebei Agricultural University, China

34 Yang Xu is an Assistant Professor at Yangzhou University, China

35 Julong Wei is a Postdoctoral Fellow at Wayne State University

36 Jianming Lu, Yuanfa Feng, and Rui Zhou are Ph.D. students in Dr. Weide Zhong's lab at

37 South China University of Technology, China

38 Yuhan Huang is a B.S. student at University of California, Los Angeles, USA

39 Renyuan Ma is a B.S. student at Bowdoin College, USA

## 40 **Abstract**

41 Prognostic tests using expression profiles of several dozen genes help provide  
42 treatment choices for prostate cancer (PCa). However, these tests require improvement to  
43 meet the clinical need for resolving overtreatment which continues to be a pervasive  
44 problem in PCa management. Genomic selection (GS) methodology, which utilizes whole-  
45 genome markers to predict agronomic traits, was adopted in this study for PCa prognosis.  
46 We leveraged The Cancer Genome Atlas (TCGA) database to evaluate the prediction  
47 performance of six GS methods and seven omics data combinations, which showed that  
48 the Best Linear Unbiased Prediction (BLUP) model outperformed the other methods  
49 regarding predictability and computational efficiency. Leveraging the BLUP-HAT method,  
50 an accelerated version of BLUP, we demonstrated that using expression data of a large  
51 number of disease-relevant genes and with an integration of other omics data (*i.e.*, miRNAs)  
52 significantly increased outcome predictability when compared with panels consisting of  
53 small numbers of genes. Finally, we developed a novel stepwise forward selection BLUP-  
54 HAT method to facilitate searching multi-omics data for predictor variables with  
55 prognostic potential. The new method was applied to the TCGA data to derive mRNA and  
56 miRNA expression signatures for predicting relapse-free survival of PCa, which were  
57 validated in six independent cohorts. This is a transdisciplinary adoption of the highly  
58 efficient BLUP-HAT method and its derived algorithms to analyze multi-omics data for  
59 PCa prognosis. The results demonstrated the efficacy and robustness of the new  
60 methodology in developing prognostic models in PCa, suggesting a potential utility in  
61 managing other types of cancer.

62 **Key words:** Genomic selection, Best linear unbiased prediction, HAT, Multi-omics data,  
63 Prostate cancer, Prognosis

64

## 65 **Introduction**

66 Prostate cancer (PCa) is the second most common cancer in men worldwide. An  
67 estimated 1,276,106 new cases and 358,989 deaths were reported in 2018 [1]. Three major  
68 challenges need to be better addressed through biomarker studies to improve the  
69 management of the disease and save lives: (I) early detection of the disease, (II) accurate  
70 prediction of tumor progression to avoid overtreatment, and (III) guidance for personalized  
71 therapies for patients carrying different subtypes of PCa. With a focus on the second  
72 challenge, this study adopted the methodology of genomic selection/prediction (GS),  
73 which is commonly applied in agricultural breeding, for an integration of multi-omics to  
74 improve the predictive ability (or predictability, defined in the Methods) for PCa prognosis.

75 The majority of PCa tumors grow slowly and will likely never cause health problems.  
76 A small percentage of patients carry aggressive PCa and require immediate treatment.  
77 Patients with slow growing tumors only require active surveillance. Lacking effective tests  
78 to provide patients with the best choices for treatment based on their individual disease  
79 states, overtreatment continues to be a health issue in PCa management owing to the  
80 associated negative and unnecessary side effects. A few clinically applicable gene  
81 expression signatures have been developed to calculate risk scores for PCa prognosis,  
82 including Prolaris (Myriad Genetics Inc.), a gene expression signature assay that is based  
83 on 31 genes involved in cell cycle progression for cancer risk stratification [2], Decipher  
84 (GenomeDx Biosciences Inc.), a 22-marker expression panel for prediction of systemic

85 progression after biochemical recurrence [3], and OncotypeDX Genomic Prostate Score  
86 (Genomic Health, Inc.), which consists of 17 genes (12 selected genes in four biological  
87 pathways and five reference genes) to predict adverse pathology at the time of radical  
88 prostatectomy [4]. Compared with the clinically applied nomograms [5], these multiple-  
89 gene tests only provide a moderate improvement to disease prognosis, and they all need  
90 further validation by prospective trials [6, 7]. This leaves a wide gap between clinical  
91 practice and its objective for eliminating unnecessary surgeries.

92 Many common human diseases, including cancer, have a polygenic nature, *i.e.*, the  
93 disease phenotypes are controlled by many genetic variants with minor effects. Numerous  
94 studies have indicated that using genome-wide markers as predictors yielded much higher  
95 predictability of complex traits than using a few major Quantitative Trait Loci (QTLs) only  
96 [7-11]. The mediocre predictive abilities of the current prognostic tests are likely due to the  
97 limited number of genes being included in simple linear models, even though some of these  
98 genes are major players of cancer progression. Conventional statistical methods usually  
99 cannot efficiently handle highly saturated models with  $p \gg n$ , where  $p$  is the number of  
100 parameters (selected markers) of the models and  $n$  is the sample size. GS is a powerful  
101 tool in the fields of plant and animal breeding, which estimate genetic effects of thousands  
102 of genome-wide markers simultaneously using whole-genome regression (WGR) models  
103 [12, 13]. Numerous advanced statistical methods, including BLUP [14, 15] and Bayesian  
104 models (*i.e.*, BayesA, BayesB, and BayesC, etc.) [12, 13, 16, 17] have been proposed [18,  
105 19], and the vast success of GS in plant and animal sciences gave an impetus to introduce  
106 this powerful application to human medicine.

107 In this study, we established a novel method, named Stepwise Forward Selection using  
108 BLUP-HAT (SFS-BLUPH), and applied this method to data from the TCGA Prostate  
109 Adenocarcinoma (TCGA-PRAD) project to develop a multi-omics signature for PCa  
110 prognosis. At first, the pre-radical prostatectomy nomogram developed by Memorial Sloan  
111 Kettering Cancer Center (MSKCC) was used to derive six quantitative disease traits,  
112 including progression-free probability in five years (PFR5YR), progression-free  
113 probability in ten years (PFR10YR), organ-confined disease (OCD), extracapsular  
114 extension (ECE), lymph node involvement (LNI), and seminal vesicle invasion (SVI).  
115 These six traits were then used to evaluate six GS models and three types of omics data  
116 including mRNA transcriptome (TR), miRNAs (MI), and methylome (ME) as well as all  
117 possible combined data (TR+MI, TR+ME, MI+ME, TR+MI+ME) to identify the best  
118 combination of model and omics data for predicting PCa outcomes. The six GS models  
119 included BLUP [14, 15], Least Absolute Shrinkage and Selection Operator (LASSO) [20],  
120 Partial Least Squares (PLS) [21], BayesB [13], Support Vector Machines (SVM) [22] using  
121 the radial basis function (SVM-RBF), and the polynomial kernel function (SVM-POLY).  
122 The results indicated that the most widely used GS model, BLUP, outperformed the other  
123 models in terms of predictability and computational efficiency. The computational  
124 efficiency was further boosted by adopting the BLUP-HAT method, an optimized version  
125 of BLUP [23]. With the BLUP-HAT method and the TCGA-PRAD data, we demonstrated  
126 that: (I) prediction models using expression profiles of a large number of genes selected  
127 from the transcriptome outperformed three clinically employed tests which only considered  
128 the expression of a small number of major genes. (II) The predictability for disease traits  
129 can be further increased if the selective predictors from other omic types (*i.e.*, miRNAs in

130 this study) were also factored into the prognostic models. Finally, we utilized the new SFS-  
131 BLUPH method to screen the gene and miRNA expression data in the TCGA-PRAD  
132 training dataset for the optimal signatures of predictor variables in predicting RFS followed  
133 by a rigorous validation in six independent PCa cohorts. The new SFS-BLUPH  
134 methodology demonstrated its translational potential and may be widely adopted for  
135 management of other types of cancer.

136

## 137 **Methods**

### 138 **TCGA-PRAD dataset**

139 Multi-omics data (including HTSeq-Counts of RNA-seq, BCGSC miRNA Profiling  
140 of miRNA-seq, and Beta value of Illumina Human Methylation 450 array) and clinical data  
141 for 495 PCa patients from the TCGA-PRAD project were downloaded and processed by a  
142 series of functions in the R package *GDCRNATools* [24]. The mRNAs and miRNAs with  
143 counts per million reads (CPM) < 1 in more than half of the patients as well as the  
144 methylation probes with any missing values were filtered out before subsequent analysis.  
145 Certain clinical characteristics, such as pre-operative PSA, which were not available in the  
146 Genomic Data Commons (GDC) data portal were retrieved from Broad GDAC Firehose  
147 (<https://gdac.broadinstitute.org/>). The TCGA-PRAD dataset was used for two purposes: (1)  
148 to compare the performance of GS models and different omics data in predicting PCa  
149 outcomes and evaluate the predictabilities of tens of thousands of BLUP-HAT models with  
150 various numbers of genes or miRNAs, and (2) to serve as a training dataset for the  
151 development of a multi-omics signature for RFS prediction. The clinical characteristics for  
152 495 patients were summarized in Table 1.

153

**Table 1: Clinical characteristics of the patients in TCGA-PRAD project**

		Patients ( <i>N</i> = 495)
Age at diagnosis (years)	≤ 65	353
	> 65	142
Clinical tumor stage	T1a	1
	T1b	2
	T1c	172
	T2a	54
	T2b	54
	T2c	50
	T3a	36
	T3b	17
	T4	2
Gleason score	≤ 6	45
	7 (3+4)	149
	7 (4+3)	98
	≥ 8	203
Pre-operative PSA (ng/mL)	0-3.9	52
	4-9.9	273
	10-19.9	99
	≥ 20	55

154

155 **Independent validation datasets**

156 The profiling data of mRNAs and/or miRNAs as well as clinical data (with available  
157 RFS data) in six public datasets (GSE70769, DKFZ2018, GSE116918, GSE107299,  
158 GSE54460, and MSKCC2010) were used to validate the prognostic signatures [25-30].  
159 MSKCC2010 had both mRNA and miRNA data, while the other five datasets only had  
160 mRNA data. Detailed information for these six datasets was summarized in Table 2.



161 Processed microarray data for GSE70769 and GSE116918 were downloaded from GEO  
 162 (<https://www.ncbi.nlm.nih.gov/geo/>) using R package *GEOquery* [31]; Reads per kilobase  
 163 per million mapped reads (RPKM) data for DFKZ2018 and processed microarray datasets  
 164 for MSKCC2010 were downloaded from cBioPortal (<https://www.cbioportal.org/>) [32].  
 165 Raw data of GSE107299 were downloaded from GEO and normalized with the Robust  
 166 Multichip Average (RMA) method implemented in the R package *oligo* [33]. Raw  
 167 sequencing data for GSE54460 were downloaded from SRA  
 168 (<https://www.ncbi.nlm.nih.gov/sra>) under the accession number SRP036848. The raw  
 169 sequencing data were aligned using *STAR (version 2.7.2a)* software [34], quantified using  
 170 *featureCounts (version 2.0.0)* software [35], and normalized using the Trimmed Mean of  
 171 M-values (TMM) normalization method implemented in the R package *edgeR* [36].

172 **Table 2: Information of the six publicly available independent validation datasets**

Dataset	Sample Size	Transcriptome Platform	miRNA Platform	Tissue
GSE70769	85	Illumina HumanHT-12 V4.0	×	Fresh frozen
DKFZ2018	32	Illumina HiSeq 2000 (RNAseq)	×	Fresh frozen
GSE116918	229	ADXPCv1a520642	×	FFPE
GSE107299	94	Affymetrix Human Gene 2.0 ST Array	×	Fresh frozen
GSE54460	90	Illumina HiSeq 2000 (RNAseq)	×	FFPE
MSKCC2010	61 (40)*	Affymetrix Human Exon 1.0 ST Array	Agilent-019118 Human miRNA Microarray 2.0	Fresh frozen

173 \* For MKSCC2010 dataset, 61 patients have gene expression data, and 40 of them have both gene expression  
 174 and miRNA expression data.

## 175 **Pre-radical prostatectomy nomograms**

176 The pre-radical prostatectomy nomogram (<https://www.mskcc.org/nomograms/>),  
177 developed by the MSKCC, utilizes pre-treatment clinical data to predict the extent of the  
178 cancer and long-term outcomes following radical prostatectomy, which can be analyzed as  
179 quantitative traits by genomic prediction models. We used this tool to predict six post-  
180 surgery disease traits, including progression-free probability in five years (PFR5YR),  
181 progression-free probability in ten years (PFR10YR), organ-confined disease (OCD),  
182 extracapsular extension (ECE), lymph node involvement (LNI), and seminal vesicle  
183 invasion (SVI). The pre-surgery clinical characteristics used for nomogram calculation  
184 included age, preoperative PSA level, Gleason score (primary Gleason and secondary  
185 Gleason), and clinical tumor stage based on the American Joint Committee on Cancer  
186 (AJCC) version 7 staging system [37].

## 187 **Genomic selection methodologies**

188 In this study, we compared the predictive ability of six widely used GS methods,  
189 including BLUP, LASSO, PLS, BayesB, SVM-POLY, and SVM-RBF. The BLUP method  
190 was implemented using a custom R script [38]. LASSO, PLS, and BayesB were  
191 implemented in the R packages *glmnet* [39], *pls* [40], and *BGLR* [41], respectively. The  
192 two SVM methods, SVM-RBF and SVM-POLY, were implemented in the R *kernelab*  
193 package [42].

194 The mRNA, miRNA, and methylation features, which were initially profiled in  
195 different ranges, were rescaled by z-score transformation, allowing for an objective  
196 comparison among these multi-omics profiles and for integrated analyses.

197 The predictability of a model, defined as the squared correlation coefficient ( $r^2$ )  
198 between the observed and predicted trait values, was calculated through a 10-fold cross  
199 validation (CV) procedure. In a 10-fold CV, the sample was arbitrarily partitioned into ten  
200 portions with approximately equal size. In each iteration, nine portions were used as the  
201 training data to develop the model and the remaining one portion was used as the test data  
202 for model evaluation. This process was repeated ten times with each portion having been  
203 used as the test data exactly once. The entire 10-fold CV was then replicated ten times to  
204 reduce the variation caused by random partitioning.

### 205 **BLUP-HAT method**

206 The BLUP-HAT model [23], which produces the same results as BLUP but enjoys  
207 much more computational efficiency due to the avoidance of the time-consuming CV, was  
208 used in place of the conventional BLUP method to compare the predictabilities of many  
209 thousands of models with various numbers of predictors. The linear mixed model that  
210 accounts for the relationship between each trait and predictor variables can be expressed  
211 as

$$212 \quad \mathbf{y} = [y_1 \cdots y_n]^T = \mathbf{1}\beta + \sum_{k=1}^m \mathbf{Z}_k \gamma_k + \boldsymbol{\varepsilon} \quad (1)$$

213 where  $\mathbf{y}$  is the vector of trait values for  $n$  patients,  $\mathbf{1}$  is a vector of 1's,  $\beta$  is the intercept  
214 (overall mean),  $\mathbf{Z}_k$  is a numerical vector for the  $k^{\text{th}}$  predictor variable,  $\gamma_k$  is the effect of  
215  $k^{\text{th}}$  variable,  $m$  is the number of predictor variables in the model, and  $\boldsymbol{\varepsilon}$  is an  $n \times 1$  vector  
216 of random errors. We assume that  $\gamma_k \sim N(0, \sigma_\gamma^2)$  for all  $k = 1, \dots, m$ , and  
217  $\boldsymbol{\varepsilon} \sim N(\mathbf{0}, \mathbf{I}\sigma^2)$  so that

218 
$$\text{Var}(\mathbf{y}) = \mathbf{V} = \frac{1}{m} \sum_{k=1}^m \mathbf{Z}_k \mathbf{Z}_k^T (m\sigma_\gamma^2) + \mathbf{I}\sigma^2 \quad (2)$$

219 
$$= \mathbf{K}\sigma_A^2 + \mathbf{I}\sigma^2,$$

219 where

220 
$$\mathbf{K} = \frac{n}{m} \cdot \frac{\sum_{k=1}^m \mathbf{Z}_k \mathbf{Z}_k^T}{\text{tr}\left(\frac{1}{m} \sum_{k=1}^m \mathbf{Z}_k \mathbf{Z}_k^T\right)} \quad (3)$$

221 is a relatedness matrix which is equivalent to the kinship matrix in GS [38]. Let us define

222 
$$\xi = \sum_{k=1}^m \mathbf{Z}_k \gamma_k$$
 as the poly-predictor effect, and  $\sigma_A^2 = m\sigma_\gamma^2$  as the poly-predictor variance,

223 we can rewrite the mixed model (1) as

224 
$$\mathbf{y} = \beta + \xi + \varepsilon \quad (4)$$

225 Thence, the Henderson's equation for the mixed model (4) can be derived as

226 
$$\begin{bmatrix} \mathbf{1}^T \mathbf{1} & \mathbf{1}^T \\ \mathbf{1} & \mathbf{I} + \mathbf{K}^{-1} / \lambda \end{bmatrix} \begin{bmatrix} \beta \\ \xi \end{bmatrix} = \begin{bmatrix} \mathbf{1}^T \mathbf{y} \\ \mathbf{y} \end{bmatrix} \quad (5)$$

227 where  $\mathbf{1}$  is an identity matrix and  $\lambda = \frac{\sigma_A^2}{\sigma^2}$ . The best linear unbiased estimation (BLUE) of

228 the fixed effects and the best linear unbiased prediction (BLUP) of the random poly-

229 predictor effect are obtained via

230 
$$\begin{bmatrix} \hat{\beta} \\ \hat{\xi} \end{bmatrix} = \begin{bmatrix} \mathbf{1}^T \mathbf{1} & \mathbf{1}^T \\ \mathbf{1} & \mathbf{I} + \mathbf{K}^{-1} / \lambda \end{bmatrix}^{-1} \begin{bmatrix} \mathbf{1}^T \mathbf{y} \\ \mathbf{y} \end{bmatrix} \quad (6)$$

231 The variance-covariance matrix of the BLUE and BLUP is

$$232 \quad \text{Var} \begin{bmatrix} \hat{\beta} \\ \hat{\xi} \end{bmatrix} = \begin{bmatrix} \mathbf{1}^T \mathbf{1} & \mathbf{1}^T \\ \mathbf{1} & \mathbf{I} + \mathbf{K}^{-1} / \lambda \end{bmatrix}^{-1} \sigma^2 \quad (7)$$

233 Following the BLUP-HAT method described by Xu [23], the predicted poly predictor

234 effect can be expressed using a linear function of the observed poly-predictor effect

235 involving the hat matrix  $\mathbf{H}$ , *i.e.*,

$$236 \quad \hat{\xi} = \mathbf{K} \sigma_A^2 \mathbf{V}^{-1} \xi = \mathbf{H} \xi \quad (8)$$

237 with  $\mathbf{H} = \mathbf{K} \sigma_A^2 \mathbf{V}^{-1}$ . Let  $\hat{y} = \hat{\beta} + \hat{\xi}$  be the predicted trait values and let  $\hat{e} = y - \hat{y}$  be the

238 residuals, with  $\hat{e}_i$  being the  $i^{\text{th}}$  element of the residual vector  $\hat{e}$ . The predicted residual

239 for individual  $i$  becomes

$$240 \quad \tilde{e}_i = \frac{1}{1 - h_{i,i}} \hat{e}_i \quad (9)$$

241 where  $h_{i,i}$  represents the  $i^{\text{th}}$  diagonal entry on  $\mathbf{H}$ . The total sum of squares is defined as

$$242 \quad SS = \sum_{i=1}^n (y_i - \bar{y})^2 \quad (10)$$

243 where  $\bar{y} = \sum_{i=1}^n y_i / n$ .

244 The predicted sum of squares is

245 
$$PRESS = \sum_{i=1}^n \tilde{e}_i^2 \quad (11)$$

246 The trait predictability of the BLUP-HAT version is

247 
$$r^2 = 1 - \frac{PRESS}{SS} \quad (12)$$

248

249 **Commercial panels for PCa prognosis**

250 Three commercial gene expression panels for PCa prognosis were compared in this  
251 study, including:

252 (I) Prolaris® (Myriad Genetics Inc., Salt Lake City, US): The Prolaris gene signature  
253 consists of 31 cell cycle genes and 15 house-keeping genes. All of the 31 genes can map  
254 to Ensembl gene IDs in the TCGA gene expression dataset (Supplementary Table S1). The  
255 15 house-keeping genes were not included in the panel for prediction.

256 (II) Decipher® (GenomeDX Inc., Vancouver, Canada): The Decipher is a 22-marker  
257 panel involving 19 genes because two markers may be derived from the same gene (e.g.,  
258 one in the coding region, and the other one in the intronic region). One of the 19 genes,  
259 Prostate Cancer Associated Transcript 32 (PCAT-32) does not have a unique ID in the  
260 Ensembl genome annotation, so expression of 18 genes with unique Ensembl IDs were  
261 used to represent this panel (Supplementary Table S2).

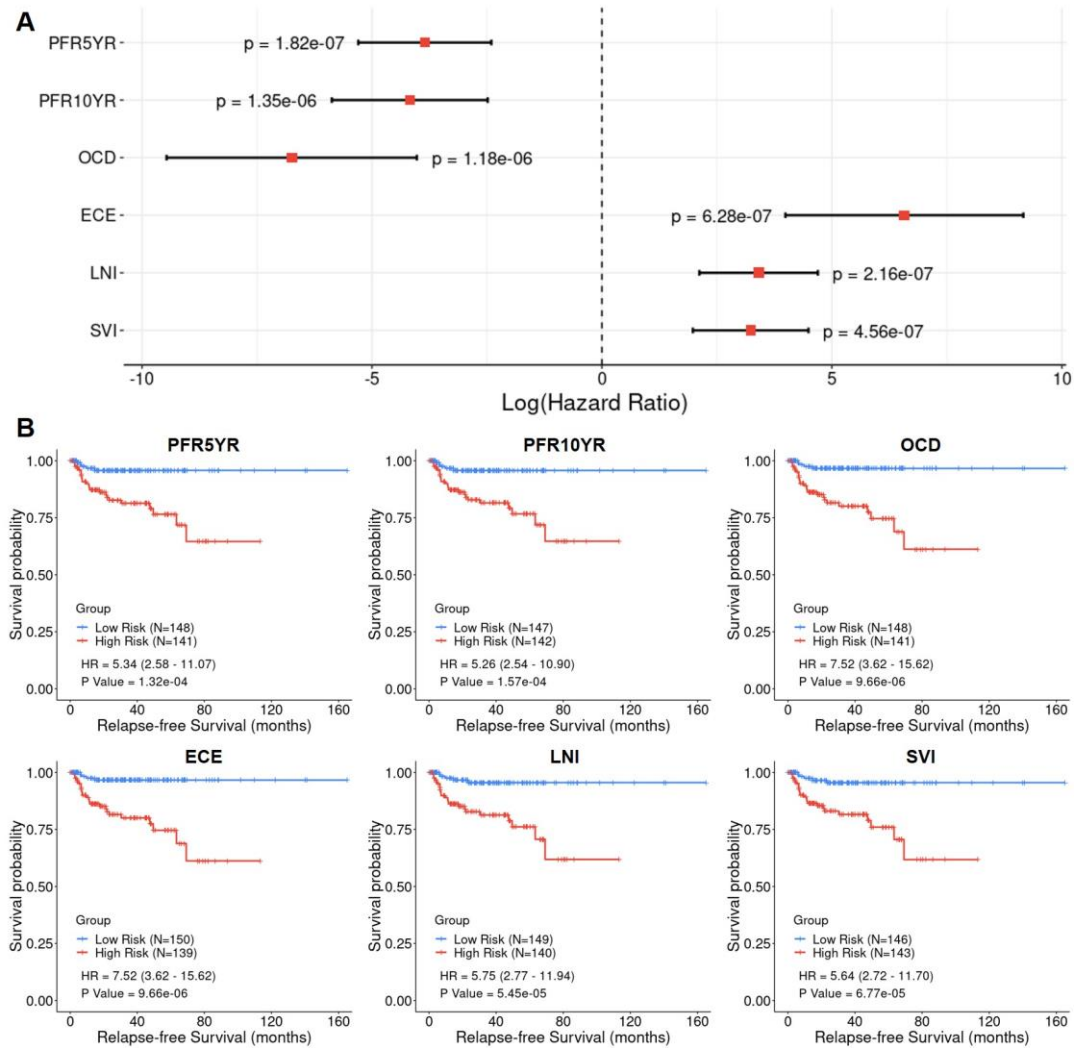
262 (III) OncotypeDX GPS® (Genomic Health Inc., Redwood City, USA): OncotypeDX  
263 GPS consists of 17 genes (12 genes in four biological pathways and five reference genes).  
264 Expression of the 12 genes were all quantified in the TCGA dataset and were used for  
265 prediction (Supplementary Table S3).

266

## 267 **Results**

### 268 **Comparison of GS methodologies using various omics data for PCa outcome** 269 **prediction**

270 We first used the six nomogram-derived traits to systematically evaluate six different  
271 GS methods with combinations of various types of omics datasets in full loads (*i.e.*, entire  
272 mRNA transcriptome, and/or entire set of miRNAs, and/or entire methylome). Although  
273 the most important trait of interest for PCa prognosis is the observed clinical outcome (*i.e.*,  
274 RFS), the nomogram-derived traits can represent collective characteristics of a patient's  
275 disease status and are much less affected by post-surgery therapies compared to the  
276 observed outcomes that are sometimes biased and complicated by incorrectly documented  
277 treatment history. The MSKCC pre-radical prostatectomy nomogram predicts the extent of  
278 the cancer and long-term results following radical prostatectomy, which can be treated as  
279 quantitative traits by the GS models. From the TCGA-PRAD dataset, 289 of the 495  
280 primary tumor patients with the available clinical data required for nomogram calculation  
281 were used for the analyses. Cox Proportional-Hazards (CoxPH) survival analysis was  
282 performed to measure the association between each nomogram-derived trait and RFS. We  
283 also performed Kaplan Meier (KM) survival analysis by classifying patients into two risk  
284 groups based on the median value for each trait. For PFR5YR, PFR10YR, and OCD, the  
285 higher the nomogram values, the lower the risk according to the definitions of the traits.  
286 On the contrary, the higher the nomogram values for ECE, LNI, and SVI, the higher the  
287 risk. Both CoxPH and KM survival analyses indicated that all the six nomogram-derived  
288 traits were significantly associated with RFS (Figure 1), indicating that they were ideal  
289 substitutes for the target traits and could be used for evaluating prognostic models.



290

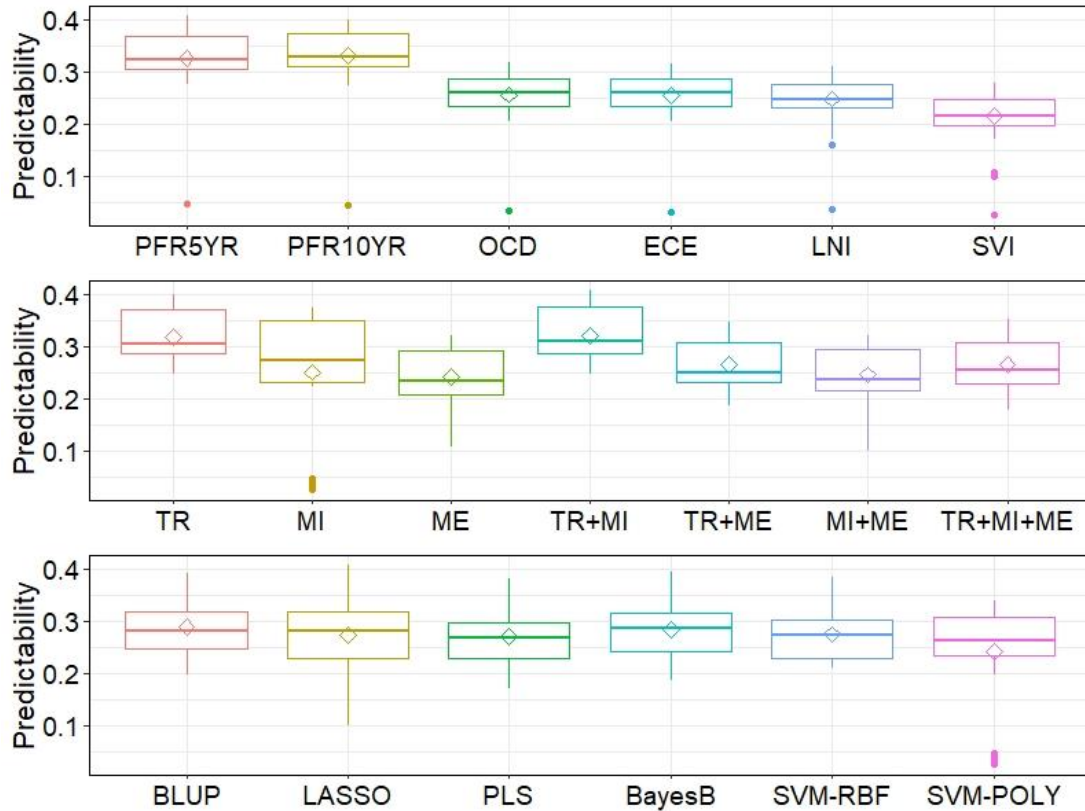
291 **Figure 1.** Cox Proportional-Hazards (CoxPH) and Kaplan-Meier (KM) survival analyses of  
 292 relapse-free survival (RFS) using the six nomogram-derived traits as variables in the TCGA-PRAD  
 293 dataset. (A) Forest plot visualizing the hazard ratio (HR) in log scale, 95% confidence intervals in  
 294 log scale, and p value of CoxPH survival analysis (B) KM curves visualizing the survival  
 295 probabilities over time for high and low risk groups classified based on the median value of the  
 296 nomogram-derived scores for each trait.

297

298



299 In total, 285 out of the 289 patients with all the omics data available were used to  
300 evaluate the performance of different GS methods and combinations of various types of  
301 omics data in predicting nomogram-derived traits. A total of 15,536 genes, 388 mature  
302 miRNAs, and 381,602 methylation probes were included for the comparison. The  
303 predictabilities of six nomogram-derived traits for the 285 patients were evaluated using  
304 six statistical methods and seven omics data combinations *via* 10-fold CV. The results  
305 indicated that the predictabilities of different traits varied substantially (Figure 2), with  
306 PFR5YR and PFR10YR having the greatest predictabilities. Prediction using mRNA  
307 transcriptomic data (TR) outcompeted prediction using either miRNA predictors (MI) or  
308 methylome predictors (ME). The combined use of TR and MI in a single model predicted  
309 disease outcomes slightly better than the model of using TR alone. In general, prediction  
310 models using ME had lower predictabilities than those using TR, MI, and other data  
311 combinations. Among the six GS methods, the conventional BLUP method generally  
312 outperformed the other methods in terms of trait predictability. In addition, BLUP appeared  
313 to be much more efficient in computation time than other methods, especially when a large  
314 number of features were included in the models (Table 3). Therefore, the BLUP method as  
315 well as the gene and miRNA expression data were selected for the subsequent analyses.



316

317 **Figure 2.** Comprehensive evaluation of the performance of six different genomic selection models  
318 (BLUP, LASSO, PLS, BayesB, SVM-POLY, and SVM-RBF) with three omics data (TR:  
319 Transcriptome; MI: miRNAs; ME: methylome) and their combinations (TR+MI, TR+ME, MI+ME,  
320 and TR+MI+ME) using the six nomogram post-surgery traits (PFR5YR: progression-free  
321 probability in 5 years; PFR10YR: progression-free probability in 10 years; OCD: organ-confined  
322 disease; ECE: extracapsular extension; LNI: lymph node involvement; SVI: seminal vesicle  
323 invasion).

324

325

326

327

328

329 **Table 3. Computational times in seconds for the six GS models using different**  
330 **omics data (DELL desktop with 16 cores × 2G memory)**

Method	TR	MI	ME
BLUP	5	1	63
LASSO	34	3	333
PLS	104	1	1,738
BayesB	385	15	9,343
SVM-RBF	145	3	3,965
SVM-POLY	149	47	3,837

331 TR: Transcriptome (15,536 genes); MI: miRNAs (388 mature miRNAs); ME: Methylome (381,602 probes)

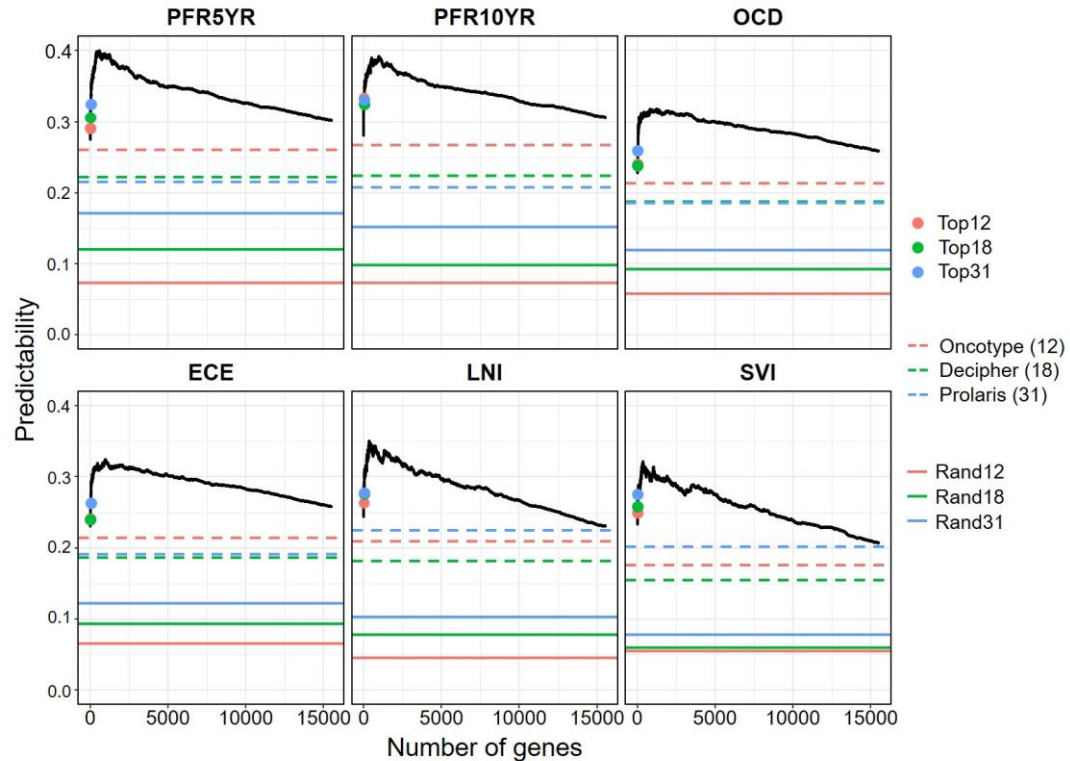
332

### 333 **Evaluation of prognostic models with different numbers of genes and/or miRNAs**

334 Enlightened by the report that HAT method yielded the approximate calculation of  
335 predictability as the conventional CV in the mixed model analysis but with much improved  
336 computational efficiency [23], a BLUP-HAT method was adopted to test tens of thousands  
337 of models to test the two proposed hypotheses: (I) using a large number of genes selected  
338 from the transcriptome to predict the outcomes of PCa patients will outperform the  
339 clinically employed prognostic tests which only rely on several dozen major genes, and (II)  
340 the predictive power will be further increased if other omics predictors are also factored  
341 into the prognostic models.

342 The transcriptomic data were used to test the first hypothesis. For each nomogram-  
343 derived trait, genes were sorted in descending order based on their absolute Pearson's  
344 correlation coefficients with the trait. Top  $N$  genes ( $N$  ranges from 5 to 15,536) selected  
345 from the sorted list were sequentially included in the mixed model to calculate the HAT  
346 value (predictability, defined in Equation 12 in the Methods section). In each plot of Figure  
347 3, the predictabilities for the models with the top 12, top 18, and top 31 genes, respectively,

348 and the predictabilities for the models consisting of genes in the three commercial tests  
349 were marked. We also included a set of control models with 12, 18, and 31 random genes,  
350 respectively. For each control model, the random genes were repeatedly selected from the  
351 transcriptome ten times, and the average predictability was calculated and labeled by solid  
352 lines with different colors in Figure 3. The results indicated that, as expected, the  
353 predictabilities of the three commercial panels were significantly higher than the randomly  
354 selected genes, confirming the prognostic abilities of those gene panels. It was observed  
355 that all the evaluated models with sorted genes being sequentially added had better  
356 predictabilities than the three commercial gene panels. The predictabilities rose as more  
357 and more genes had been included in the model until they reached the maximum value,  
358 where thereafter the predictability values started decreasing. Generally, a few hundred  
359 genes were required to have the maximum predictability for each trait, which supported  
360 our first hypothesis that the outcome predictability may be substantially boosted by  
361 including hundreds of the genes on the top of the sorted gene list when compared with the  
362 models using only a small number of the top ‘major’ genes.



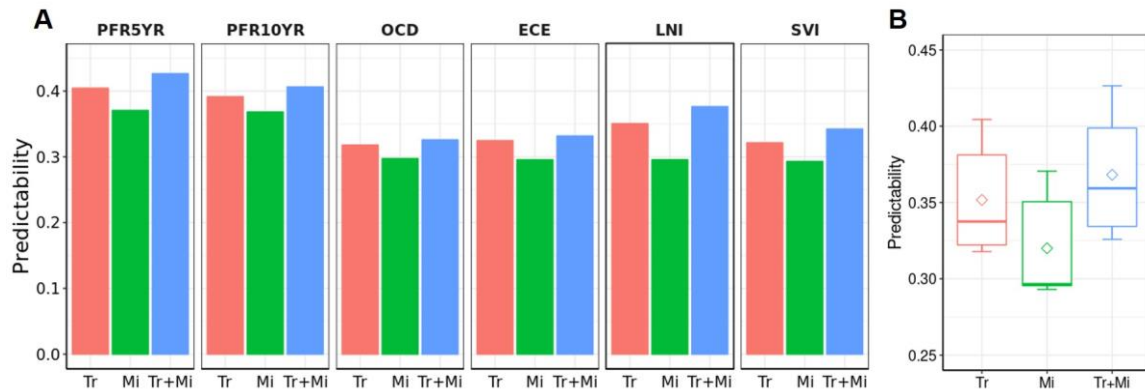
363

364 **Figure 3.** Evaluation of prediction models using different number of genes selected from the  
365 transcriptome in predicting six nomogram-derived traits by the BLUP-HAT method. (Top12,  
366 Top18, and Top31 represent the top 12, 18, and 31 genes in the ranked gene list, respectively.  
367 Rand12, Rand18, and Rand31 represent randomly selected 12, 18, and 31 genes from the  
368 transcriptome, respectively). The numbers of genes that achieved the maximum predictabilities for  
369 PFR5YR, PFR10YR, OCD, ECE, LNI, and SVI are 470, 995, 1246, 989, 366, and 363, respectively.

370

371 To test the second hypothesis that the predictability can be further improved by  
372 integrating panels from other omics data, BLUP-HAT was also used to identify the optimal  
373 set (top  $N$ ) of miRNAs that reached the maximum predictability. Then the predictabilities  
374 of the optimal gene set, the optimal miRNA set, and their combinations were compared for  
375 the six traits. The results indicated that: (1) the models using gene expression data  
376 outperformed the models using expression data of miRNAs, and (2) the models with

377 combined expression of genes and miRNAs had greater predictabilities than those using  
378 genes only, supporting our second hypothesis (Figure 4). To this point, we have used PCa  
379 data to successfully provide strong evidence supporting the two hypotheses, which would  
380 generally hold in other types of cancers and may help guide the development of improved  
381 cancer prognostic models leveraging multi-omics data.



382 **Figure 4.** The performance of different expression panels in predicting the six nomogram-derived  
383 traits using BLUP-HAT. (A) Bar plot visualizing the predictability of each panel for predicting a  
384 trait. (B) Box plot visualizing the overall predictabilities of panels with different omics data across  
385 the six traits. (Tr: a panel of top genes with the highest predictability selected from the ranked gene  
386 list; Mi: a panel of top miRNAs with the highest predictability selected from the ranked miRNAs  
387 list; Tr+Mi: a combined panel of Tr and Mi. Genes/miRNAs in the Tr/ Mi panels for different traits  
388 are different)

389

### 391 Development of multi-omics prognostic models by the SFS-BLUPH methodology

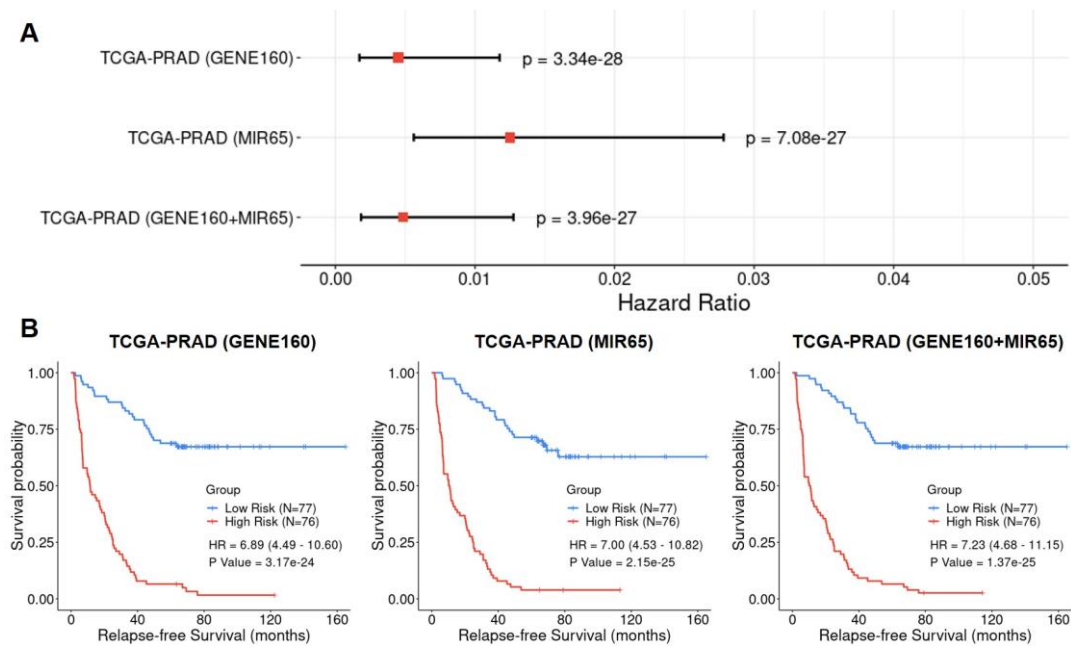
392 The predictive power and computational efficiency of the BLUP-HAT method have  
393 been demonstrated using six PCa outcome traits calculated by nomogram. We then  
394 leveraged this method to select a multi-omics signature for the prediction of RFS, the  
395 disease phenotype of interest. Patients with limited post-surgery follow-up data were

396 eliminated from the initial 495 patients, leaving a total of 153 patients in this analysis, of  
397 which 95 underwent disease relapse or biochemical recurrence (BCR) within five years  
398 after prostatectomy. The outcome phenotypic value for a patient was defined as 1 if either  
399 this patient had not relapsed within five years or the time to first BCR was more than five  
400 years; otherwise, the outcome phenotypic value was calculated by dividing the time to first  
401 BCR by five, yielding a continuous score variable. Note that the greater the RFS score, the  
402 higher the probability of RFS (or the better the outcome). The newly defined outcome trait,  
403 which represented the probability of being RFS in five years (RFS5YR) after surgery, was  
404 most clinically relevant to disease prognosis.

405 In order to refine an optimal multi-omics signature for the prediction of RFS, we  
406 developed a novel stepwise forward selection strategy by leveraging the highly efficient  
407 BLUP-HAT method and the TCGA-PRAD multi-omics datasets. Similarly, we sorted all  
408 of the genes in descending order based on their absolute Pearson's correlation coefficients  
409 with RFS. The initial BLUP-HAT model included the top two genes from the sorted list.  
410 In each following step, the next gene in the list was added to the current model for a  
411 calculation of the RFS predictability; this gene was retained if the addition of it increased  
412 the RFS predictability, otherwise, this gene was discarded. This selection process was  
413 repeated until all genes in the sorted list were sequentially tested, which yielded a refined  
414 160-gene signature (GENE160) for predicting RFS. The same selection strategy was  
415 applied to the miRNA data to derive a refined 65-miRNA signature (MIR65) for predicting  
416 RFS.

417 In the TCGA-PRAD training set, three BLUP prognostic models (GENE160, MIR65,  
418 and GENE160+MIR65) were built using the selected genes and/or miRNAs for the

419 prediction of the RFS scores. An RFS score was calculated for each patient *via* Leave-one-  
420 out cross validation (LOOCV), and the median value of these RFS scores was used to  
421 dichotomize the TCGA-PRAD cohort into a high-risk group (RFS scores less than the  
422 median value) and a low-risk group (RFS scores greater than the median value). The  
423 CoxPH regression analysis indicated that the scores calculated using all of the three  
424 signatures were significantly associated with RFS in the TCGA-PRAD training set (Figure  
425 5A). The KM survival analysis showed that the patients in the low-risk group had  
426 significantly higher survival probability than those in the high-risk group (Figure 5B).



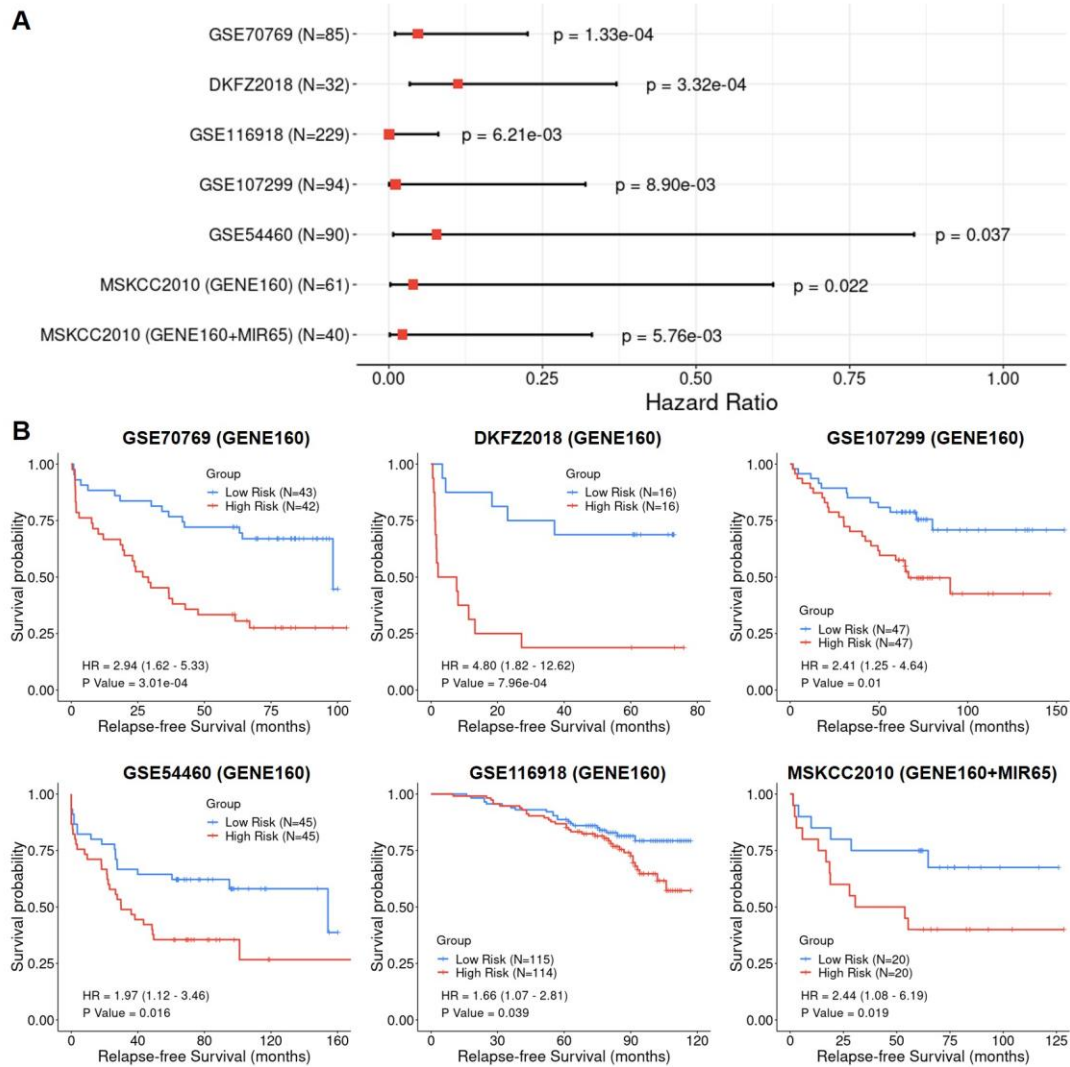
427

428 **Figure 5.** Cox Proportional-Hazards (CoxPH) and Kaplan-Meier (KM) survival analyses of  
429 relapse-free survival (RFS) using the GENE160, MIR65, and GENE160+MIR65 signatures in the  
430 TCGA-PRAD training dataset. (A) Forest plot visualizing the hazard ratio (HR), 95% confidence  
431 intervals, and p value of CoxPH survival analysis. (B) KM curves visualizing the survival  
432 probabilities over time for high and low risk groups classified based on the median predicted RFS  
433 scores in the cohort

434



435 We further validated the prognostic performance of the GENE160 and  
436 GENE160+MIR65 signatures using six independent cohorts. Note that these additional six  
437 datasets were not created using the same platform as the TCGA-PRAD data; thus, certain  
438 predictor variables of small number, either from 160 genes or from 65 miRNAs, were  
439 missing in some datasets (Supplementary Table S4). While validating the signatures and  
440 the methodology with each dataset, we only employed the available genes and/or miRNAs  
441 in a BLUP regression analysis. LOOCV was also used to calculate the RFS scores for the  
442 patients in each validation cohort. The CoxPH regression analysis and the KM analysis  
443 were then utilized to evaluate the association between the calculated RFS scores and the  
444 observed RFS outcomes. Although the RNAs were collected from different types of tissues  
445 (*i.e.*, fresh frozen tumor tissue or FFPE) and the RNA abundance data were profiled using  
446 a variety of platforms (*i.e.*, four different gene microarrays and RNAseq), the CoxPH  
447 regression analysis and the KM survival analyses indicated that the GENE160 signature  
448 alone was able to robustly predict RFS or differentiate high-risk patients from low-risk  
449 patients in these six datasets (Figure 6). Note that for the cohort of MSKCC2010, the  
450 CoxPH regression analysis rendered a significant result ( $p = 0.02$ ), while the KM analysis  
451 only showed prognostic tendency ( $p = 0.15$ ). Since the miRNA data is available for the  
452 MSKCC2010 dataset, we tested the multi-omics model with the integration of GENE160  
453 and MIR65 signatures, which showed a significantly increased prognostic ability in this  
454 validation set. The  $p$  value for the CoxPH regression analysis has been improved from 0.02  
455 (GENE160) to  $5.76e-03$  (GENE160+MIR65), while the result for the KM analysis became  
456 statistically significant ( $p = 0.019$ ).



457

458 **Figure 6.** Cox Proportional-Hazards (CoxPH) and Kaplan-Meier (KM) survival analyses of  
 459 relapse-free survival (RFS) using the GENE160 and GENE160+MIR65 panels in six independent  
 460 validation datasets. (A) Forest plot visualizing the hazard ratio (HR), 95% confidence intervals,  
 461 and p value of CoxPH survival analysis (B) KM curves visualizing the survival probabilities over  
 462 time for high and low risk groups classified based on the median predicted RFS scores in each  
 463 cohort.

464

465

466

## 467 **Discussion**

468       Due to the cost of gene testing and the convenience of modeling, establishment of a  
469 prognostic test only using dozens of gene expression profiles has been the rule of thumb in  
470 the past decades. In our study, the predictabilities of three commercial panels of PCa  
471 prognosis were significantly higher than those of randomly selected gene sets, suggesting  
472 that the genes in these panels are indeed associated with disease progression. For example,  
473 Prolaris consists of 31 cell cycle progression (CCP) genes, many of which are functionally  
474 relevant to PCa recurrence [2]. Genes representing multiple biological pathways in  
475 Decipher are associated with PCa progression and have been reported to be differentially  
476 expressed throughout PCa progression [3]. The selected genes in Oncotype have also been  
477 verified to be related to PCa aggressiveness [4]. These several dozens of genes included in  
478 the commercial panels are no doubt biologically critical in PCa. However, these genes,  
479 even with major effects, may not be the best or complete set of predictors for PCa prognosis,  
480 which may be indicated by the results shown in Figure 3, *i.e.*, all the sequentially evaluated  
481 models had better predictabilities than the three commercial gene panels. This may be  
482 ascribed to two major reasons: (1) due to the heterogeneity of PCa tumors, the major genes  
483 in one cohort may not necessarily be major players in another cohort, and (2) models with  
484 a large number of genes, including both major players and minor genes, may render a better  
485 prediction of outcomes than a panel with only ‘major genes’.

486       The rapid advancement in biotechnology has significantly reduced operational cost,  
487 allowing us to develop improved tests by including a large number of genes, a practice  
488 previously limited by economic constraints. However, conventional statistical methods  
489 cannot efficiently handle highly saturated models with  $p \gg n$ , *i.e.*, the number of predictor

490 variables is much larger than the number of observations. Robust GS models such as BLUP  
491 and Bayesian methods (*i.e.*, BayesA, BayesB, and BayesC, etc.) have been proposed and  
492 applied to handle saturated linear regression models in plant and animal breeding. However,  
493 the computational advantages of these advanced methods have been rarely applied to  
494 cancer prognosis and warrant investigation. In this study, we took advantage of  
495 transdisciplinary expansion to adapt these powerful GS methodologies from agricultural  
496 sciences to human cancer research. The results indicated that BLUP outcompeted other  
497 rival methods in both predictive ability and computational efficiency. When many  
498 thousands of prediction models need to be compared, BLUP-HAT may further reduce the  
499 computational cost by avoiding lengthy CV.

500 The computationally efficient BLUP-HAT model was utilized to evaluate tens of  
501 thousands of models in regard to their performance in predicting clinical outcomes of PCa.  
502 The results from these comparisons demonstrated that, when compared with the currently  
503 used commercial panels with a limited number of genes, inclusion of many more genes  
504 with minor effects on the disease may collectively improve the overall RFS predictability.  
505 The BLUP-HAT model also enjoyed the easiness of combining multi-omics data into a  
506 single model, which allowed for a further improvement of the predictive ability.

507 We established a novel stepwise forward selection BLUP-HAT method to facilitate  
508 searching available multi-omics data for predictor variables with predictive potential.  
509 Using the TCGA data as a training set, we developed a 160-gene signature and a 65-  
510 miRNA signature for predicting the RFS of PCa. The GENE160 signature alone was  
511 successfully validated in all six independent cohorts, and the GENE160+MIR65 multi-  
512 omics signature showed significantly improved predictability compared with GENE160

513 signature in the only test set where miRNA data was available. Certain genes or miRNAs  
514 were missing in some validation sets because different platforms were used for generating  
515 these independent datasets. The RFS predictabilities in these validation analyses might  
516 have been increased if the missing genes/miRNAs were added back to the prognostic  
517 models. The validation was also successful when FFPE samples were analyzed  
518 (GSE116918 and GSE54460). These results indicated that the signatures and the  
519 methodology were robust even when the quality of RNA samples was relatively low,  
520 suggesting a great potential in clinical application. A limitation of the study is that the size  
521 of the training set ( $n = 153$ ) and six validation sets ( $n < 100$  in general) were small, which  
522 was quite different from studies of plants or animals. An improved prognostic model for  
523 an accurate prediction of RFS for PCa patients can be developed when data for large  
524 cohorts become available in the future.

525 In summary, we demonstrated that (1) a large number of disease-relevant genes render  
526 better prediction of PCa outcomes than a few dozen major genes, and (2) the combination  
527 of multi-omics predictor variables can further increase the predictability. We developed a  
528 novel SFS-BLUPH methodology which can efficiently search multi-omics data for  
529 predictor variables with prognostic potential. This method may be applied to any private  
530 database for the development of clinically useful tests for PCa prognosis. The new method  
531 may also be extendedly applied to different cancers or other types of human diseases.

532

533

534

535

## 536 **Key points**

- 537 • We adopted genomic selection methods from the agricultural sciences and applied  
538 these to cancer research.
- 539 • We systematically evaluated the performance of six genomic selection methods  
540 using three omics data and their combinations in predicting prostate cancer  
541 outcomes, and found that the Best Linear Unbiased Prediction (BLUP) method  
542 outperformed the other models in terms of trait predictability and computational  
543 efficiency.
- 544 • With the more computationally efficient BLUP-HAT methodology, we  
545 demonstrated that (1) prediction models using expression data of a large number of  
546 genes selected from the transcriptome outperformed the clinically employed tests  
547 which only considered a small number of major genes, and (2) the integration of  
548 other omics data (*i.e.*, miRNAs) in the model will further increase the predictability.
- 549 • We developed a novel stepwise forward selection BLUP-HAT (SFS-BLUPH)  
550 method to search multi-omics data for predictor variables to predict relapse-free  
551 survival of prostate cancer patients. The methodology has been successfully  
552 validated using six independent cohorts.

553

## 554 **Data Access**

555 All the scripts used in this study, including data preprocessing, genomic selection model  
556 evaluation, implementation of BLUP-HAT method, development and validation of the  
557 SFS-BLUPH model, as well as data visualization are freely available at  
558 <https://github.com/rli012/BLUPHAT>.

## 559 **Funding**

560 The work was supported by UC Riverside Faculty Start-up Fund, UC Riverside Hellman  
561 Fellowship, UC Academic Senate Regents Faculty Fellowship and Faculty Development  
562 Award, and UC Cancer Research Coordinating Committee Competition Award to Zhenyu  
563 Jia. Jianguo Zhu was supported by the National Natural Science Foundation of China  
564 81660426 and 81873608. Weide Zhong was supported by the National Key Basic Research  
565 Program of China (2015CB553706), the National Natural Science Foundation of China  
566 (81571427), and Guangzhou Municipal Science and Technology Project (201803040001;  
567 201707010291).

568

## 569 **Acknowledgments**

570 We would like to thank Professor Shizhong Xu at University of California, Riverside, for  
571 discussing this project with us and providing insightful comments on the study.

572

## 573 **Disclosure Declaration**

574 The authors declare that they have no competing interests.

575

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