

An Adaptive Detection for Automatic Spike Sorting Based on Mixture of Skew-t distributions

Ramin Toosi¹, Mohammad Ali Akhaee¹, and Mohammad-Reza A. Dehaqani^{1,2}

¹School of Electrical and Computer Engineering, College of Engineering, University of Tehran, Tehran, Iran

²School of Cognitive Sciences, Institute for Research in Fundamental Sciences, Tehran, Iran

Appendix S1: Zero-Phase Filtering

In this method, first the signal is filtered, and then the time reversed version of the filtered signal would be filtered again. Assuming h is the filter response, in the z-transform domain we have:

$$\begin{aligned} D(z) &= R(z)H(z) \\ R_F(z) &= D(z^{-1})H(z) \end{aligned} \quad (\text{S.1})$$

where R is the z-transform of the raw signal (r). When $|z| = 1$ or $z = e^{j\omega}$, the output reduces to $R_F(e^{j\omega})|H(e^{j\omega})|^2$. Thus, the output has zero-phase distortion.

Appendix S2: Expectation-Maximization for Mixture of Skew-t distribution

In the EM algorithm, it is usual to define a membership variable, $Z_i = (Z_{i1}, \dots, Z_{ig})$, where

$$Z_{ij} = \begin{cases} 1, & \text{if } s_i \text{ belongs to the component } j \\ 0, & \text{otherwise} \end{cases} \quad (\text{S.2})$$

The completed log-likelihood function can be written as follows:

$$\begin{aligned} \ell_c(\theta|s, t, u, z) &= C - \sum_{i=1}^N \sum_{j=1}^g Z_{ij} \left[\log(\pi_j) - \frac{1}{2} \log(\gamma_j) - \frac{u_i}{2} (s_i - \mu_j - \Delta_j t_i)^T \gamma_j^{-1} (s_i - \mu_j - \Delta_j t_i) \right. \\ &\quad \left. + \frac{\nu}{2} \log\left(\frac{\nu}{2}\right) + \left(\frac{\nu}{2} - 1\right) \log(u_i) - \log\left(\frac{\nu}{2}\right) - \frac{\nu}{2} u_i \right] \end{aligned} \quad (\text{S.3})$$

where C is a constant, and

$$\delta = \frac{\lambda_j}{\sqrt{1 + \lambda_j^T \lambda_j}}, \quad \Delta = \Sigma_j^{-\frac{1}{2}} \delta_j, \quad \gamma_j = \Sigma_j - \Delta_j \Delta_j^T, \quad t_i \sim N(\mu + \Delta t_i, u_i^{-1} \gamma_j) \quad (\text{S.4})$$

The E-step includes calculating of $Q(\theta|\theta^{(k)}) = E_{\theta^{(k)}}(\ell_c(\theta|s, t, u, z)|s)$, where the superscript (k) indicates the value of the parameters in the k 'th iteration. The EM algorithm which finds the parameters of a mixture of p-dimensional skew-t distributions can be summarized as follows:

E-Step: First, consider the following auxiliary parameters.

$$\begin{aligned}
M^2(\boldsymbol{\theta}) &= (1 + \Delta^T \boldsymbol{\gamma}^{-1} \Delta)^{-1} \\
m(\boldsymbol{\theta}, s) &= M^2(\boldsymbol{\theta}) \Delta^T \boldsymbol{\gamma}^{-1} (s - \boldsymbol{\mu}) \\
A(\boldsymbol{\theta}, s) &= \lambda^T \Sigma^{-\frac{1}{2}} (s - \boldsymbol{\mu}) \\
\beta(\boldsymbol{\theta}, s) &= \frac{2\Gamma\left(\frac{v+p+2}{2}\right) (v + d_\Sigma(s, \boldsymbol{\mu}))^{-1} T\left(\sqrt{\frac{v+p+2}{v-d_\Sigma(s, \boldsymbol{\mu})}} A(\boldsymbol{\theta}, s) | v + p + 2\right)}{\Gamma\left(\frac{v+p}{2}\right) T\left(\sqrt{\frac{v-p}{v+d_\Sigma(s, \boldsymbol{\mu})}} A(\boldsymbol{\theta}, s) | v + p\right)} \\
\tau(\boldsymbol{\theta}, s) &= \frac{1}{\left(\sqrt{\frac{v+p}{v+d_\Sigma(s, \boldsymbol{\mu})}} A(\boldsymbol{\theta}, s) | v + p\right)} \frac{\Gamma\left(\frac{v+p+1}{2}\right)}{\pi^{\frac{1}{2}} \Gamma\left(\frac{1}{2}\right)} \frac{(v + d_\Sigma(s, \boldsymbol{\mu}))^{(v-p)/2}}{(v + d_\Sigma(s, \boldsymbol{\mu}) + A(\boldsymbol{\theta}, s)^2)^{(v-p+1)/2}} \\
\xi(\boldsymbol{\theta}, s) &= \beta(\boldsymbol{\theta}, s) m(\boldsymbol{\theta}, s) + M(\boldsymbol{\theta}) \tau(\boldsymbol{\theta}, s) \\
\omega(\boldsymbol{\theta}, s) &= \beta(\boldsymbol{\theta}, s) (m(\boldsymbol{\theta}, s))^2 + (M(\boldsymbol{\theta}))^2 + M(\boldsymbol{\theta}) m(\boldsymbol{\theta}, s) \tau(\boldsymbol{\theta}, s)
\end{aligned} \tag{S.5}$$

The probability of i 'th spike belongs to the j 'th component, i.e. p_{ij} , is

$$p_{ij} = \frac{\pi_j S T_p(s_i | \boldsymbol{\theta}_j)}{\sum_{l=1}^g \pi_l S T_p(s_i | \boldsymbol{\theta}_l)} \tag{S.6}$$

Given $\boldsymbol{\theta} = \boldsymbol{\theta}^{(k)}$, first we calculate $p_{ij}^{(k)}$, then,

$$\beta_{ij} = p_{ij}^{(k)} \beta(\boldsymbol{\theta}_j^{(k)}, s_i) \quad \xi_{ij} = p_{ij}^{(k)} \xi(\boldsymbol{\theta}_j^{(k)}, s_i) \quad \omega_{ij} = p_{ij}^{(k)} \omega(\boldsymbol{\theta}_j^{(k)}, s_i) \tag{S.7}$$

M-Step: In this step, by maximizing $Q(\boldsymbol{\theta} | \boldsymbol{\theta}^{(k)})$, the update rules for the parameters would be obtained as follows:

$$\begin{aligned}
\pi_j^{(k+1)} &= n^{-1} \sum_{i=1}^N p_{ij}^{(k)} \\
\boldsymbol{\mu}_j^{(k+1)} &= \sum_{i=1}^N (\beta_{ij}^{(k)} s_i - \xi_{ij}^{(k)} \Delta_j^{(k)}) / \sum_{i=1}^N \beta_{ij}^{(k)} \\
\Delta_j^{(k+1)} &= \left[\sum_{i=1}^N \xi_{ij}^{(k)} (s_i - \boldsymbol{\mu}_j^{(k+1)}) \right] / \sum_{i=1}^N \omega_{ij}^{(k)} \\
\boldsymbol{\gamma}_j^{(k+1)} &= \left(\sum_{i=1}^N p_{ij} \right)^{-1} \sum_{i=1}^N \left(\beta_{ij}^{(k)} (s_i - \boldsymbol{\mu}_j^{(k+1)}) (s_i - \boldsymbol{\mu}_j^{(k+1)})^T \right. \\
&\quad \left. - \left[(s_i - \boldsymbol{\mu}_j^{(k+1)}) (\Delta_j^{(k+1)})^T + (\Delta_j^{(k+1)}) (s_i - \boldsymbol{\mu}_j^{(k+1)})^T \right] \xi_{ij}^{(k)} + \Delta_j^{(k+1)} (\Delta_j^{(k+1)})^T \omega_{ij}^{(k)} \right) \\
\nu^{(k+1)} &= \arg \max_{\nu} \sum_{i=1}^N \log \left(\sum_{j=1}^g \pi_j S T_p(s_i | \boldsymbol{\mu}_j^{(k+1)}, \Sigma_j^{(k+1)}, \boldsymbol{\lambda}_j^{(k+1)}, \nu^{(k)}) \right)
\end{aligned} \tag{S.8}$$

The update procedure is based on Δ , and $\boldsymbol{\gamma}$ instead of $\boldsymbol{\lambda}$ and Σ . These parameters could be recovered as follows:

$$\boldsymbol{\lambda}_j = \frac{(\boldsymbol{\gamma}_j + \Delta_j \Delta_j^T)^{-\frac{1}{2}} \Delta_j}{\left[1 - \Delta_j^T (\boldsymbol{\gamma}_j + \Delta_j \Delta_j^T)^{-1} \Delta_j \right]^{\frac{1}{2}}} \quad \Sigma_j = \boldsymbol{\gamma}_j + \Delta_j \Delta_j^T \tag{S.9}$$