Genome wide analysis of gene dosage in 24,092 individuals shows that 10,000 genes modulate cognitive ability

Single sentence summary: CNVs' effect-sizes on intelligence are predicted using measures of intolerance to haploinsufficiency and are distributed across half of the coding genes.

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60 **ABSTRACT**

61 Genomic Copy Number Variants (CNVs) are routinely identified and reported back to patients
62 with neuropsychiatric disorders, but their quantitative effects on essential traits such as cognitive
63 ability are poorly documented. We have recently shown that the effect-size of deletions on
64 cognitive ability can be statistically predicted using measures of intolerance to
65 haploinsufficiency. However, the effect-sizes of duplications remain unknown. It is also
66 unknown if the effect of multigenic CNVs are driven by a few genes intolerant to
67 haploinsufficiency or distributed across tolerant genes as well.
68 Here, we identified all CNVs >50 kilobases in 24,092 individuals from unselected and autism
69 cohorts with assessments of general intelligence. Statistical models used measures of intolerance
70 to haploinsufficiency of genes included in CNVs to predict their effect-size on intelligence.
71 Intolerant genes decrease general intelligence by 0.8 and 2.6 points of IQ when duplicated or
72 deleted, respectively. Effect-sizes showed no heterogeneity across cohorts. Validation analyses
73 demonstrated that models could predict CNV effect-sizes with 78% accuracy. Data on the
74 inheritance of 27,766 CNVs showed that deletions and duplications with the same effect-size on
75 intelligence occur *de novo* at the same frequency.
76 We estimated that around 10,000 intolerant and tolerant genes negatively affect intelligence when
77 deleted, and less than 2% have large effect-sizes. Genes encompassed in CNVs were not enriched
78 in any GOterms but gene regulation and brain expression were GOterms overrepresented in the
79 intolerant subgroup. Such pervasive effects on cognition may be related to emergent properties of
80 the genome not restricted to a limited number of biological pathways.

81

82 **Introduction**

83

84 Copy Number Variants (CNVs) are deletions or duplications larger than 1000 base pairs. The
 85 contribution of CNVs to the etiology of intellectual disability (ID)[1–3], autism[4–6] and
 86 schizophrenia[6–8] is well established. The interpretation of CNVs in research and medical
 87 diagnostics remains essentially binary: benign or pathogenic (contributing to mental illness)[9].
 88 The routine implementation of Chromosomal Micro-Arrays (CMAs) as a first-tier diagnostic test
 89 identifies “pathogenic” CNVs in 10 to 15 % of children with neurodevelopmental disorders
 90 (NDD)[10]. A binary interpretation is however of limited use because patients present a broad
 91 spectrum of cognitive symptoms ranging from severe ID to learning disabilities. The quantitative
 92 effects of CNVs are poorly documented even for important traits such as general intelligence. It
 93 may be available for the most frequently recurrent CNVs but data is often collected in patients
 94 ascertained in the clinic with a bias towards severely affected individuals, leading to potentially
 95 gross overestimation of effect size. Only two studies have been conducted in unselected
 96 populations [11, 12] showing reduced performance on cognitive test for 24 recurrent CNVs.
 97 However, recurrent CNVs only represent a very small fraction of the total amount of ultra-rare
 98 CNVs identified in the neurodevelopmental disorder clinic as well as in the general population.
 99
 100 Intelligence is a major trait assessed in the developmental pediatric and psychiatric clinic. There
 101 is a significant genetic correlation between intelligence and psychiatric disorders and cognitive
 102 impairments represent a major referral criterion to the NDD clinic. The heritability of general
 103 intelligence is estimated at around 50 to 80% [13]. The heritability of variants in linkage
 104 disequilibrium with common SNPs is estimated to be around 22.7%, with variants in poor linkage
 105 disequilibrium with SNPs, including rare CNVs, explaining 31.3% of the phenotypic variation in
 106 intelligence[14]. Two recent GWAS, have identified over 200 loci associated with intelligence
 107 and education[15, 16] , potentially implicating 1000 genes. The latter were largely non-

overlapping with genes previously linked to ID[15]. Contrary to SNPs, there is no ambiguity in the molecular interpretation of a fully deleted or duplicated gene, which invariably decreases or increases transcription respectively. Therefore, CNVs represent a powerful tool to map the effect-sizes of genes (altered by gene dosage) on human traits.

We have previously proposed a framework to estimate and predict the effect-size on intelligence of CNVs. We showed that linear models[17] using the sum of the “probability of being loss-of-function intolerant” (pLI) scores[18] of all genes included in a deletion can predict their effect-size on intelligence quotient (IQ) with 75% accuracy. Our initial study was underpowered to measure the effect-size of duplications. It is also unknown if only a limited number of intolerant genes or a large proportion of genes within CNVs are driving effects on cognitive abilities. More broadly, the number of genes modulating general intelligence remains unknown. The pLI used in our earlier model, ranges from 0 to 1 but has a bimodal distribution and is essentially a categorical variable classifying genes as intolerant (>0.9) or tolerant (≤ 0.9) to protein-loss-of-function (pLoF) [18]. Continuous measures such as the LOEUF[19] (Loss-of-function Observed/Expected Upper bound Fraction) were recently introduced to reflect the full spectrum of intolerance to pLoF. LOEUF range from 0 to 2, and values below 0.35 are suggestive of intolerance.

Our present aims were 1) to test the robustness of effect-size estimates for CNVs across unselected and NDD populations, 2) to establish the effect-size on general intelligence of genomic duplications, 3) to investigate the quantitative relationship between effect-size on general intelligence and the frequency of *de novo* events, and 4) to estimate individual effect-sizes for all protein-coding genes that are intolerant as well as tolerant to pLoF.

We identified CNVs in 24,092 individuals from five general populations, two autism cohorts and one neurodevelopmental cohort. Measures of intolerance to pLoF were used as variables to estimate the effect of CNVs and individual genes on general intelligence. Validation procedures

133 using cognitive data on CNVs from 47 published reports and the UKBB demonstrated a near
134 80% accuracy of model estimated. We implemented an online tool to help clinicians and
135 researchers estimate the effect-size of any CNVs on general intelligence.

RESULTS

1) Deletions and duplications have a 3:1 effect-size ratio on general intelligence

We first sought to replicate our previous estimates for the effect-size of deletions on general intelligence computed using pLI [17]. We performed a meta-analysis on 20,151 individuals from 5 unselected populations (Table 1, Supplementary Fig. 1) showing that the deletion of one point of pLI decreases NVIQ or g-factor by 0.18 z-score (95% CI: -0.23 to -0.14, equivalent to 2.7 points of NVIQ, Fig. 1a, Supplementary Table 1). For duplications, we performed a meta-analysis using the same unselected populations. It shows that duplicating one point of pLI decreases NVIQ or g-factor by 0.04 z-score (95% CI: -0.09 to -0.01), which is equivalent to 0.75 points of IQ. Of notes, our previous study [17] was unable to estimate effect-sizes of duplications on general intelligence, likely due to sample size. There was no heterogeneity across cohorts. Sensitivity analyses showed that methods used for cognitive assessments did not influence these results (Fig. 1, Supplementary Table 2).

2) The effect-size of CNVs on general intelligence is not influenced by ascertainment.

Since genomic variants with large effects on general intelligence are thought to be removed from the general population as a result of negative selective pressure, this may have led to an underestimation of the effect-size of CNVs in unselected populations. To examine this possibility, we analyzed 3,941 individuals (Table 1, Supplementary Fig. 1) from two autism cohorts, which include individuals with ID and *de novo* CNVs. Effect-sizes of pLI on general intelligence were the same than those observed in unselected populations for deletions and duplications and we did not observe any heterogeneity across cohorts (Fig. 1, Supplementary Table 1). Finally, we asked if effect-sizes of pLI were the same in large CNVs rarely observed in the general population or in autism cohorts. We tested 226 CNV carriers and 325 intrafamilial controls from 132 families ascertained in the clinic (Table 1). Effect-sizes of pLI on IQ were very

similar with a decrease of 0.147 z-score, 95% CI: -0.18 to -0.11 ($P=1.1\times 10^{-15}$) in deletions and 0.069 z-score, 95% CI: -0.1 to -0.04 ($P=8.7\times 10^{-6}$) in duplications (Supplementary Table 3).

3) Mega-analysis suggests additive effects of constraint scores on general intelligence

We pooled samples after adjusting for variables including cognitive test and cohorts to perform a mega-analysis of 24,092 individuals carrying 13,001 deletions and 15,856 duplications encompassing 36% of the coding genome (Fig. 1b, Supplementary Fig. 2a). The effect-size of pLI was unchanged, decreasing general intelligence by 0.175 z-score ($SE=0.016$, $P=1.25\times 10^{-28}$) and 0.054 z-score ($SE=0.009$, $P=1.90\times 10^{-9}$) for deletions and duplications, respectively (Supplementary Table 4). The partial R^2 shows that deletions and duplications measured by pLI explain respectively 0.5% and 0.1% of the total variance of intelligence in the complete dataset; in line with the fact that large effect-size CNVs are rare in the general population.

Among 11 variables, the 2 main constraint scores (pLI and 1/LOEUF) best explained (based on AIC) the variance of general intelligence (Supplementary Table 4). For the remainder of the study, we transitioned to using LOEUF because it is a continuous variable (the pLI is essentially binary) and is now recommended as the primary constraint score by gnomAD. Analyses using pLI are presented in supplemental results.

There was no interaction between constraint scores and age or sex (Supplementary Table 5 to 8). Non-linear models did not improve model fit (Supplementary Table 9 to 10), suggesting an additive effect of constraint scores.

4) The effect-size of 1/LOEUF on intelligence is the same in recurrent neuropsychiatric CNVs and non-recurrent CNVs

We show that removing 608 individuals carrying any of the 121 recurrent CNV previously associated with neuropsychiatric conditions [17] does not influence the effect-size of 1/LOEUF

on general intelligence (Supplementary Table 11). It has been posited that the deleteriousness of large psychiatric CNVs may be due to interactions between genes encompassed in CNVs. We therefore asked if the effect-size of 1/LOEUF is the same for CNVs encompassing small and large numbers of genes. We recomputed the linear model 6 times after incrementally excluding individuals with a total sum of $1/\text{LOEUF} \geq 60, 40, 20, 10, 4$ and 2.85 for deletions and duplications separately. Effect-sizes remain similar whether deletions encompass >10 or >60 points of 1/LOEUF (Fig. 1d, Supplementary Fig. 2b).

5) Gene dosage of 1% of coding genes shows extreme effect-size on general intelligence.

Our ability to estimate large effect sizes is likely hampered by the explanatory variable (1/LOEUF) used in the model because there is only a 60-fold difference between the smallest and largest value. To improve model accuracy for large effect-size genes, we used a list of 256 ID-genes[2, 20], previously identified with an excess of *de novo* mutations in NDD cohorts. We identified 126 CNVs encompassing at least one ID-gene (Fig. 2).

We recomputed the model by integrating 4 explanatory variables: the sum of 1/LOEUF for ID and non-ID-genes encompassed in deletions and duplications. The effect-size on intelligence of 1/LOEUF for ID-genes was 7 to 11-fold higher than the effect-size of non-ID genes which remained unchanged (Supplementary Table 12, 13 and Fig. 3). The mean effect of ID-genes intolerant to pLoF ($\text{LOEUF} < 0.35$) was a decrease of 20 points of IQ for deletions and 9 points for duplications (Supplementary Table 13).

6) Model explains nearly 80% of the effect-size of CNVs.

As a validation procedure, we compared model estimates to published observations for 47 recurrent CNVs reported in clinical series and in the UKBB¹⁷ (Supplementary Table 14 and 15). When cognitive data was available from both clinical and the UKBB ($n=13$), we used the mean of both effect-sizes. Concordance between model estimates and previously published measures

was 0.78 for all CNVs (95% CI, 0.66-0.86, $P=4.3 \times 10^{-11}$, Fig. 4). Accuracy was similar for deletions (ICC=0.71 [0.5;0.84], $P=1.8 \times 10^{-5}$) and duplications (ICC=0.85 [0.7;0.93], $P=3 \times 10^{-7}$) as well as for small and large CNVs including trisomy 21 (Fig. 3a and 3b, Supplementary Fig. 5).

7) CNVs with the same impact on intelligence have the same *de novo* frequency.

Because measures of intolerance to haploinsufficiency explain equally well the effect-sizes of deletions and duplications on intelligence, we investigated the relationship between effects on intelligence and *de novo* frequency for deletions and duplications. We established inheritance for 26,437 CNVs in 6 cohorts (Supplementary Table 16). There was a strong relationship between effects on general intelligence estimated by the model and the frequency of *de novo* observations for deletions ($P=1.9 \times 10^{-65}$) and duplications ($P=4.6 \times 10^{-24}$, Fig. 3c).

Deletions and duplications with the same impact on general intelligence show similar *de novo* frequency CNVs (Fig. 3c).

The concordance between the probability of occurring *de novo* estimated by the model (after removing recurrent CNVs) and *de novo* frequency reported in the DECIPHER database on 31 recurrent CNVs was 0.81 ([0.67-0.9]; $P=8.2 \times 10^{-8}$) (Fig. 3d, Supplementary Table 17 and Fig. 6).

8) Estimating effect-sizes of individual genes using LOEUF

Since we were underpowered to perform a gene-based GWAS, we first divided all genes in 4 categories: highly intolerant genes (LOEUF<0.2; n=980), moderately intolerant genes ($0.2 \leq \text{LOEUF} < 0.35$ n=1,762), tolerant genes ($0.35 \leq \text{LOEUF} < 1$; n=7,442) and highly tolerant genes (LOEUF ≥ 1 ; n=8,267). This dichotomization of LOEUF values also allowed to test whether the previous linear models were driven by subgroups of genes. The sum of genes in each category was used as four explanatory variables to explain general intelligence in the same linear model. For deletions, highly, moderately intolerant and tolerant genes showed negative effects on

general intelligence (Fig. 4a, Supplementary Table 18). For duplications only moderately intolerant genes showed negative effects (Supplementary Fig. 7 and Table 18). We were underpowered to further subdivide these LOEUF categories, so we tested 38 overlapping LOEUF categories in 38 linear models. Each model used 2 explanatory variables: number of genes within and outside the LOEUF category (size = 0.15 LOEUF). For haploinsufficiency, negative effects on general intelligence were observed for genes within 13 categories across intolerant and tolerant LOEUF values. For duplications, only 2 categories had negative effects (Fig. 4a, Supplementary Fig. 7 and Table 19).

9) Most biological functions affect cognition.

The 6,114 different genes encompassed in the CNVs of our dataset did not show any GOterm enrichment except for olfactory related terms (Supplementary Tables 20). We asked if intolerant ($\text{LOEUF} < 0.35$) and tolerant genes ($0.35 < \text{LOEUF} < 1$), which negatively affect IQ in the analysis above were enriched in GOterms. All intolerant and tolerant genes genome-wide, were enriched in 365 and 30 GOterms respectively (Fig. 4b, Supplementary Tables 21, 22). The largest group of GOterms enriched in intolerant genes represented gene regulation (RNA polymerase II transcription factor activity, chromatin organization; Fig. S11), cell death regulation and neuronal function (dendrite and synapse). Among 23 tissues overrepresented in intolerant genes, adult brain and epithelium showed the strongest enrichment (Supplementary Table 21). Top enriched pathways included those in cancer, focal adhesion, Wnt signaling and MAPK (Supplementary Table 21). For tolerant genes, milder enrichments included translation (tRNA) and cytoskeletal structure. Among the 7 significant tissues adult brain showed the strongest enrichment (Fig. 4b, Supplementary Table 22 and Fig. 12). The 2,862 intolerant and tolerant genes encompassed in the CNVs of our dataset showed the same GOterm distribution observed above for the full intolerant and tolerant coding genome. Genes encompassed in CNVs were therefore represented well all

261 molecular functions observed for each LOEUF group at the genome-wide level (Supplementary
262 Table 23).

DISCUSSION

Deletions and duplications have effect-sizes on cognitive ability that are robust across cohorts, clinical diagnoses, and general intelligence assessments. The effect-size ratio on cognitive ability of deletions to duplications is 3:1. The linear sum of pLI or 1/LOEUF predicted the effect-size on intelligence of deletions and duplications with equal accuracy (78%). Using categories of LOEUF values, we provide the first estimates for the individual effect-sizes of protein-coding genes, suggesting that half of the coding genome affects intelligence. The 2,862 genes encompassed in CNVs of our dataset show the same GOterm distribution observed in the intolerant and tolerant coding genome.

Model validation and ascertainment biases

Models show 78% concordance with effect-size of CNVs on IQ from previous literature reports. Estimates are discordant for several CNVs, which may be due to either 1) unidentified large effect-size genes with unreliable LOEUF measures due to the small size of the protein coding region, and 2) ascertainment bias. However, biases from clinically referred individuals can be adjusted for using intrafamilial controls [21, 22]. This is confirmed by effect-sizes using the Ste Justine family genetic cohort. Also, our results suggest that the effect-size of pathogenic CNVs are underestimated in the UKBB[21] while those of small CNVs are largely overestimated in clinical series. The maximum effect size measured in UKBB was only 0.4 z-score including pathogenic CNVs such as 16p11.2, 2q11.2 deletions and 10q11.21-q11.23 deletion containing an ID-gene (*WDFY4*). On the other hand, the effect size of variants such as the 16p13.11 duplications and 1q21.1 CNVs are likely overestimated in clinical series[23]. Therefore, statistical models using a variety of disease and unselected cohorts are likely to provide the most accurate estimates. Surprisingly, an autism diagnosis is not associated with a different impact of CNVs on cognitive ability. A recent study characterizes this finding showing that CNVs similarly

decrease IQ in autism and in unselected populations but are nevertheless more frequent in autism than in controls with same intelligence[24].

Individual effect-sizes of genes, and go their GOterm enrichments

Our study is based on CNVs encompassing intolerant and tolerant genes with the same GOterm distribution observed in those LOEUF categories genome-wide. Only one percent of coding genes with the highest intolerance to pLoF has large effects on cognitive ability (20 and 9 IQ points for deletions and duplications of ID genes). The rest of the intolerant genes (15% of coding genes) have moderate to mild effect-sizes. The group of all intolerant genes is enriched in many GOterms including brain expression and gene regulation as previously reported for this group[2, 25]. Genes considered tolerant to pLoF ($0.35 < \text{LOEUF} < 1$; 40% of coding genes) impact intelligence with small effect-size and are only mildly enriched in GOterms. This is reminiscent of GWAS results for schizophrenia showing that most GOterms contribute to it's heritability [26].

Potential clinical application

Models developed in this study provide a translation of gnomAD constraint scores into cognitive effect-sizes. Model outputs are implemented in a prediction tool (<https://cnvprediction.urca.ca/>), which is designed to estimate the population-average effect-size of any given CNV on general intelligence, not the cognitive ability of the individual who carries the CNV. If the cognitive deficits of an individual are concordant with the effect-size of the CNV they carry, one may conclude that the CNV contributes substantially to those deficits. When discordant (ie. The observed IQ drop is ≥ 15 points (1SD) larger than the model estimate), the clinician may conclude that a substantial proportion of the contribution lies in additional factors which should be investigated, such as additional genetic variants and perinatal adverse events (e.g. neonatal

hypoxic ischemic injury, seizure disorders etc). If IQ cannot be reliably measured (ie. ≤ 4 years or in the case of severe behavioral disorders), the cognitive impact of the CNV predicted by the model may allow to anticipate the need for potential interventions. Overall, the output of this tool can help interpret CNVs in the clinic, but estimates should be interpreted with caution. The model can provide an estimate for the effect size on intelligence of individual genes when deleted. Therefore, one may use this information to estimate the effect size on intelligence of any SNV resulting in a loss of function. However, larger datasets are required to refine the estimates for individual gene.

The relationship between genetic fitness and cognitive abilities

The reasons underlying the tight relationship between general intelligence and epidemiological measures of intolerance to pLoF, is unclear. This relationship is further highlighted by the fact that deletions and duplications with the similar impact on intelligence occur *de novo* with similar frequencies. Behavioral interpretations are intuitive for severe ID but do not apply for CNVs with much milder effects. In other words, individuals with moderate or severe ID have limited offspring due to behavioral deficits but it is unclear how small changes in intelligence may lead to behavioral issues resulting in decreased fitness. Our results also suggest that genes considered as “tolerant” with LOEUF < 1 affect cognitive abilities and are likely under “mild constraint”. Larger samples are required to better characterize the effect of this broad category of “mildly intolerant” genes on cognitive ability.

Limitations

The model relies on constraint scores (LOEUF or pLI), which are epidemiological measures of genetic fitness in human populations, without any consideration of gene function[18, 19]. It is likely that some genes decrease fitness (eg. genes involved in fertility) without affecting general

intelligence. Further studies combining intolerance scores with functional categories are required to investigate this question. While LOEUF was designed to measure intolerance to loss of function, we used it to assess both deletions and duplications. However, our results and a recent report suggest that it also measures the intolerance to increased gene expression [27]. Noise in the model may be related to unreliable constraint scores computed for small genes with a limited number of pLoF variants observed in the gnomAD database. Bias in the model may be introduced by ID genes observed in our dataset. Indeed, they may reflect a less severe subgroup and model outputs should be interpreted with caution when CNVs encompass ID-genes. Another potential bias is related to the fact that models were trained on CNVs encompassing 36% of the coding genome. Projections suggest that 500K individuals from an unselected population would cover 78% (Fig. S8).

Finally, all models imply additive effects and massive datasets would be required to test for gene-gene and gene-environment interactions. However, the fact that very large CNVs (such as trisomy 21) are accurately estimated by the model suggests that genetic interactions within large genomic segments or even chromosomes cannot be readily observed. There is long standing discordance between observations made at the microscopic and macroscopic level. Indeed, molecular studies provide unequivocal evidence that gene-gene interactions are common but quantitative genetic theory suggests that contributions from non-additive effects to phenotypic variation in the population are small. Reconciling these two observations, polygenic models assume that interactions are the rule rather than the exception. Interactions are, in fact, accounted for in the additive models[28]. For example, LOEUF values are correlated with the number of protein-protein interactions[19] and our results also show that the intolerant genes are enriched in GOterms linked to “gene regulation”. In other words, the level of interactions for a given gene is directly related to its “individual” effect size on intelligence (ie. chromatin remodelers have a very broad interaction network, low LOEUF values and high effect sizes on intelligence).

364

365 **Conclusions**

366 The effect-size of deletions or duplications on intelligence can be accurately estimated with
367 additive models using constraint scores. The same relationship between gene dosage and
368 cognition apply to small benign CNVs as well as extreme CNVs such as Down syndrome. We
369 provide a map of effect-sizes at the individual gene level but to move beyond this rough outline,
370 much larger sample sizes are required. Nonetheless, these results suggest that a large proportion
371 (56%) of the coding genome covering all molecular functions influences cognitive abilities. One
372 may therefore view the genetic contribution to cognitive difference as an emergent property of
373 the entire genome not restricted to a limited number of biological pathways.

Materials and Methods

1. Cohorts

We included five cohorts from the general population, two autism cohorts and one familial cohort with at least one CNV-carrier child recruited for a neurodevelopmental disorder (Table1). Studies for each cohort were reviewed by local institutional review boards. Parents/guardians and adult participants gave written informed consent and minors gave assent.

2. Measures of general intelligence

General intelligence was assessed using the neurocognitive tests detailed in table 1. Measures of non-verbal intelligence quotient (NVIQ) were available in five cohorts and general intelligence factor (g-factor)[29] was computed in four cohorts, based on cognitive tests, primarily assessing fluid non-verbal reasoning (Table1, Supplementary Fig. 1). Intelligence measures were normalized using z-score transformations to render them comparable. The concordance between z-scored NVIQ and g-factor available for three cohorts ranged from 60 to 77% (Supplementary Table 24).

3. Genetic information

CNV calling and filtering

For all SNP array data, we called CNVs with PennCNV and QuantiSNP using previously published methods [17]. For the MSSNG dataset[30], we used CNVs called on whole genome sequencing by Trost *et al.* [31].

CNV filtering steps were previously published (Supplemental material). For the mega-analysis, we applied an additional filtering criterion, selecting CNVs encompassing at least 10 probes for all array technologies used across all cohorts.

The Sainte-Justine CNV-family cohort included participants on the basis of one pathogenic CNV identified in the diagnostic cytogenetic laboratory using an Agilent 180K array.

Annotation of CNVs

We annotated the CNVs using Gencode V19 (hg19) with ENSEMBL (<https://grch37.ensembl.org/index.html>). Genes with all transcripts fully encompassed in CNVs were annotated using 12 variables present in previous article[17]. Non-coding regions were annotated with the number of expression quantitative trait loci (eQTLs) regulating genes expressed in the brain[32]. CNV scores were derived by summing all scores of genes within CNVs.[17]. Also, we used a list of 256 ID-genes[2, 20], previously identified with an excess of *de-novo* mutations in NDD cohorts.

4. Statistical analyses

Modelling the effect of CNVs on intelligence

General intelligence was adjusted within each cohort for age and sex when required ($Z_{adj\ Intell.}$; see supplemental material and Supplementary Fig. 9 and 10). To estimate the effect of CNVs on general intelligence, we fit the model developed by Huguet et al. [17] where the sum of pLI (or any of the 10 other scores) for all genes encompassed in deletions or duplications, respectively, is the variable used to predict the adjusted Z-score of general intelligence:

$$\text{Model for deletion } (\mathcal{M}1_{DEL}): Z_{adj\ intell.} \sim \beta_{0,DEL} + \beta_{1,DEL} \times \sum_{gene} pLI$$

where $\beta_{0,DEL}$, $\beta_{1,DEL}$ are the regression coefficients. The same model was applied to duplications. First, models $\mathcal{M}1_{DEL}$ and $\mathcal{M}1_{DUP}$ were fitted independently and adjusted for each cohort and results were used in the meta-analyses. Second, in the mega-analysis, $\mathcal{M}1_{DEL}$ and $\mathcal{M}1_{DUP}$ were fitted after pooling all samples and adjusting on the type of cognitive measure and cohort.

To take into account ID-genes that have a greater impact on intelligence, we used a model including 4 predictive variables ($\mathcal{M}2$):

$$Z_{adj\ intell.} \sim \beta_0 + \beta_1 \times \sum_{ID\ gene\ in\ deletion} \frac{1}{LOEUF} + \beta_2 \times \sum_{ID\ gene\ in\ duplication} \frac{1}{LOEUF} + \beta_3 \times \sum_{non-ID\ gene\ in\ deletion} \frac{1}{LOEUF} + \beta_4 \times \sum_{non-ID\ gene\ in\ duplication} \frac{1}{LOEUF}$$

where β_0 , β_1 , β_2 , β_3 and β_4 are the regression coefficients.

The variance explained by deletions and duplications (measured by pLI) was computed using partial R^2 in the full dataset as well as the subgroup (n=14,874) of unrelated individuals.

Sensitivity analyses

We tested non-linearity of the effect of haploinsufficiency scores on general intelligence by using polynomial regression model and by exploring a smooth function of the effect of haploinsufficiency scores using a Gaussian kernel regression method (<https://cran.r-project.org/web/packages/KSPM/index.html>) flexible enough to account for various types of effects (Supplementary material).

Model Validation

To validate our models, we computed the concordance between model predictions and loss of IQ measured for 47 recurrent CNVs obtained in previous publications (supplementary material). The concordance was computed using the intraclass coefficient correlation of type (3,1) ($ICC_{(3,1)}$) [33].

Modelling the probability to be de novo

We performed logistic regressions to estimate the probability of a CNV being *de novo* ($P_{de\ novo}$) as a function of the haploinsufficiency scores:

Model for deletions ($\mathcal{M}3_{DEL}$):

$$\text{logit}(P_{de novo}) \sim \beta_{0,DEL} + \beta_{1,DEL} \times Z_{adj\text{intell.estimated by } \mathcal{M}2\text{ deletion.}}$$

where $\beta_{0,DEL}$, $\beta_{1,DEL}$ are the regression coefficients. The same model was applied to duplications

($\mathcal{M}3_{DUP}$)

For these analyses, we added two clinical populations (Decipher, decipher.sanger.ac.uk/) and the cytogenetic database of Sainte-Justine Hospital, where genetic data could be compared between the child and their parents, and applied the same filtering as for the previous CNV selection leading to a total of 26,437 CNVs. (Supplementary Table 16). The binary outcome variable was the type of transmission (1=*de novo*, 0=inherited).

To validate these models, we computed the concordance between model estimates and percentage of *de novo* variants computed with Decipher for 27 recurrent CNVs.

Estimating the effect-size of individual genes based on LOEUF values

We used 4 categories of LOEUF values to estimate the effect-size of genes classified as highly intolerant (LOEUF <0.2, n=980), moderately intolerant (0.2 ≤ LOEUF <0.35 n=1,762), tolerant (0.35 ≤ LOEUF <1, n=7,442), and highly tolerant to haploinsufficiency (LOEUF ≥ 1, n=8,267). For deletions, model 4 is as follow:

($\mathcal{M}4_{del}$):

$$Z_{adj\text{intell.}} \sim \beta_0 + \beta_1 \times \sum (\text{highly intolerant genes } i) + \beta_2 \times \sum (\text{moderately intolerant genes } i) + \beta_3 \times \sum (\text{tolerant genes } i) + \beta_4 \times \sum (\text{highly tolerant genes } i)$$

where $\beta_{0,CNV\text{ type}}$, $\beta_{1,CNV\text{ type}}$, $\beta_{2,CNV\text{ type}}$, $\beta_{3,CNV\text{ type}}$ and $\beta_{4,CNV\text{ type}}$ are the regression coefficients. The same model was applied for duplications.

To explore smaller categories of LOEUF values, we slid a window of size 0.15 LOEUF units, in increments of 0.05 units thereby creating 38 categories across the range of LOEUF values. We performed 38 linear models:

($\mathcal{M}5_{del}$):

$$Z_{\text{adj intell.}} \sim \beta_{0,CNV \text{ type}} + \beta_{1,CNV \text{ type}} \times \sum (\text{genes } i \text{ inside the window}) \\ + \beta_{2,CNV \text{ type}} \times \sum (\text{genes } i \text{ outside the window})$$

461 where $\beta_{0,CNV \text{ type}}$, $\beta_{1,CNV \text{ type}}$ and $\beta_{2,CNV \text{ type}}$ are the regression coefficients.

462 The same models were performed for duplications. Estimates were corrected for multiple testing

463 (38 tests) using FDR.

464 GOterms Enrichment

465 For the GOterms enrichment for the tolerant and intolerant genes with all a genome and CNVs

466 between unselected, ASD and both populations, we used DAVID release 6.8[34] ([https://david-](https://david-d.ncifcrf.gov)

467 [d.ncifcrf.gov](https://david-d.ncifcrf.gov)). We kept the defaults parameters and save only the terms with Bonferroni

468 corrected p-values <0.05. We then passed the list to REVIGO[35] (<http://revigo.irb.hr/>) to

469 summarize and group the redundant GO.

470

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REFERENCES

1. Coe BP, Witherspoon K, Rosenfeld JA, van Bon BWM, Vulto-van Silfhout AT, Bosco P, et al. Refining analyses of copy number variation identifies specific genes associated with developmental delay. *Nat Genet.* 2014;46:1063–1071.
2. Coe BP, Stessman HAF, Sulovari A, Geisheker MR, Bakken TE, Lake AM, et al. Neurodevelopmental disease genes implicated by de novo mutation and copy number variation morbidity. *Nat Genet.* 2019;51:106–116.
3. Wilfert AB, Sulovari A, Turner TN, Coe BP, Eichler EE. Recurrent de novo mutations in neurodevelopmental disorders: properties and clinical implications. *Genome Med.* 2017;9.
4. Huguet G, Ey E, Bourgeron T. The genetic landscapes of autism spectrum disorders. *Annu Rev Genomics Hum Genet.* 2013;14:191–213.
5. Pinto D, Delaby E, Merico D, Barbosa M, Merikangas A, Klei L, et al. Convergence of Genes and Cellular Pathways Dysregulated in Autism Spectrum Disorders. *Am J Hum Genet.* 2014;94:677–694.
6. Maillard AM, Ruef A, Pizzagalli F, Migliavacca E, Hippolyte L, Adaszewski S, et al. The 16p11.2 locus modulates brain structures common to autism, schizophrenia and obesity. *Mol Psychiatry.* 2015;20:140–147.
7. Sakai M, Watanabe Y, Someya T, Araki K, Shibuya M, Niizato K, et al. Assessment of copy number variations in the brain genome of schizophrenia patients. *Mol Cytogenet.* 2015;8.
8. Szatkiewicz JP, O'Dushlaine C, Chen G, Chambert K, Moran JL, Neale BM, et al. Copy number variation in schizophrenia in Sweden. *Mol Psychiatry.* 2014;19:762–773.
9. Riggs ER, Andersen EF, Cherry AM, Kantarci S, Kearney H, Patel A, et al. Technical standards for the interpretation and reporting of constitutional copy-number variants: a joint consensus recommendation of the American College of Medical Genetics and Genomics (ACMG) and the Clinical Genome Resource (ClinGen). *Genet Med.* 2019:1–13.
10. Miller DT, Adam MP, Aradhya S, Biesecker LG, Brothman AR, Carter NP, et al. Consensus Statement: Chromosomal Microarray Is a First-Tier Clinical Diagnostic Test for Individuals with Developmental Disabilities or Congenital Anomalies. *Am J Hum Genet.* 2010;86:749–764.
11. Kendall KM, Bracher-Smith M, Fitzpatrick H, Lynham A, Rees E, Escott-Price V, et al. Cognitive performance and functional outcomes of carriers of pathogenic copy number variants: analysis of the UK Biobank. *Br J Psychiatry.* 2019;214:297–304.
12. Stefansson H, Meyer-Lindenberg A, Steinberg S, Magnusdottir B, Morgen K, Arnarsdottir S, et al. CNVs conferring risk of autism or schizophrenia affect cognition in controls. *Nature.* 2014;505:361–366.
13. Posthuma D, de Geus EJC, Boomsma DI. Perceptual Speed and IQ Are Associated Through Common Genetic Factors. *Behav Genet.* 2001;31:593–602.
14. Hill WD, Arslan RC, Xia C, Luciano M, Amador C, Navarro P, et al. Genomic analysis of family data reveals additional genetic effects on intelligence and personality. *Mol Psychiatry.* 2018;23:2347–2362.
15. Savage JE, Jansen PR, Stringer S, Watanabe K, Bryois J, de Leeuw CA, et al. Genome-wide association meta-analysis in 269,867 individuals identifies new genetic and functional links to intelligence. *Nat Genet.* 2018;50:912–919.
16. Hill WD, Marioni RE, Maghziyan O, Ritchie SJ, Hagenaars SP, McIntosh AM, et al. A combined analysis of genetically correlated traits identifies 187 loci and a role for neurogenesis and myelination in intelligence. *Mol Psychiatry.* 2019;24:169–181.
17. Huguet G, Schramm C, Douard E, Jiang L, Labbe A, Tihy F, et al. Measuring and Estimating the Effect Sizes of Copy Number Variants on General Intelligence in Community-

Based Samples. *JAMA Psychiatry*. 2018;75:447–457.

18. Lek M, Karczewski KJ, Minikel EV, Samocha KE, Banks E, Fennell T, et al. Analysis of protein-coding genetic variation in 60,706 humans. *Nature*. 2016;536:285–291.

19. Karczewski KJ, Francioli LC, Tiao G, Cummings BB, Alföldi J, Wang Q, et al. Variation across 141,456 human exomes and genomes reveals the spectrum of loss-of-function intolerance across human protein-coding genes. *BioRxiv*. 2019:531210.

20. McRae JF, Clayton S, Fitzgerald TW, Kaplanis J, Prigmore E, Rajan D, et al. Prevalence and architecture of de novo mutations in developmental disorders. *Nature*. 2017;542:433–438.

21. D’Angelo D, Lebon S, Chen Q, Martin-Brevet S, Snyder LG, Hippolyte L, et al. Defining the Effect of the 16p11.2 Duplication on Cognition, Behavior, and Medical Comorbidities. *JAMA Psychiatry*. 2016;73:20–30.

22. Moreno-De-Luca A, Evans DW, Boomer KB, Hanson E, Bernier R, Goin-Kochel RP, et al. The Role of Parental Cognitive, Behavioral, and Motor Profiles in Clinical Variability in Individuals With Chromosome 16p11.2 Deletions. *JAMA Psychiatry*. 2015;72:119–126.

23. Bernier R, Steinman KJ, Reilly B, Wallace AS, Sherr EH, Pojman N, et al. Clinical phenotype of the recurrent 1q21.1 copy-number variant. *Genet Med*. 2016;18:341–349.

24. Douard E, Zeribi A, Schramm C, Tamer P, Loum MA, Nowak S, et al. Effect Sizes of Deletions and Duplications on Autism Risk Across the Genome. *Am J Psychiatry*. 2020:appi.ajp.2020.19080834.

25. Satterstrom FK, Kosmicki JA, Wang J, Breen MS, De Rubeis S, An J-Y, et al. Large-Scale Exome Sequencing Study Implicates Both Developmental and Functional Changes in the Neurobiology of Autism. *Cell*. 2020;180:568–584.e23.

26. Boyle EA, Li YI, Pritchard JK. An expanded view of complex traits: from polygenic to omnigenic. *Cell*. 2017;169:1177–1186.

27. An open resource of structural variation for medical and population genetics | *bioRxiv*. <https://www.biorxiv.org/content/10.1101/578674v1.full>. Accessed 31 December 2019.

28. Wray NR, Wijmenga C, Sullivan PF, Yang J, Visscher PM. Common Disease Is More Complex Than Implied by the Core Gene Omnigenic Model. *Cell*. 2018;173:1573–1580.

29. Deary IJ. Intelligence. *Annu Rev Psychol*. 2011;63:453–482.

30. Yuen RKC, Merico D, Bookman M, Howe JL, Thiruvahindrapuram B, Patel RV, et al. Whole genome sequencing resource identifies 18 new candidate genes for autism spectrum disorder. *Nat Neurosci*. 2017;20:602–611.

31. Trost B, Walker S, Wang Z, Thiruvahindrapuram B, MacDonald JR, Sung WWL, et al. A Comprehensive Workflow for Read Depth-Based Identification of Copy-Number Variation from Whole-Genome Sequence Data. *Am J Hum Genet*. 2018;102:142–155.

32. Ramasamy A, Trabzuni D, Guelfi S, Varghese V, Smith C, Walker R, et al. Genetic variability in the regulation of gene expression in ten regions of the human brain. *Nat Neurosci*. 2014;17:1418–1428.

33. Shrout PE, Fleiss JL. Intraclass correlations: uses in assessing rater reliability. *Psychol Bull*. 1979;86:420–428.

34. Huang DW, Sherman BT, Lempicki RA. Systematic and integrative analysis of large gene lists using DAVID bioinformatics resources. *Nat Protoc*. 2009;4:44–57.

35. Supek F, Bošnjak M, Škunca N, Šmuc T. REVIGO Summarizes and Visualizes Long Lists of Gene Ontology Terms. *PLOS ONE*. 2011;6:e21800.

36. Schumann G, Loth E, Banaschewski T, Barbot A, Barker G, Büchel C, et al. The IMAGEN study: reinforcement-related behaviour in normal brain function and psychopathology. *Mol Psychiatry*. 2010;15:1128–1139.

37. Pausova Z, Paus T, Abrahamowicz M, Bernard M, Gaudet D, Leonard G, et al. Cohort Profile: The Saguenay Youth Study (SYS). *Int J Epidemiol*. 2017;46:e19.

38. Deary IJ, Gow AJ, Pattie A, Starr JM. Cohort Profile: The Lothian Birth Cohorts of 1921

571 and 1936. *Int J Epidemiol.* 2012;41:1576–1584.

572 39. Awadalla P, Boileau C, Payette Y, Idaghdour Y, Goulet J-P, Knoppers B, et al. Cohort
573 profile of the CARTaGENE study: Quebec’s population-based biobank for public health and
574 personalized genomics. *Int J Epidemiol.* 2013;42:1285–1299.

575 40. Smith BH, Campbell A, Linksted P, Fitzpatrick B, Jackson C, Kerr SM, et al. Cohort
576 Profile: Generation Scotland: Scottish Family Health Study (GS:SFHS). The study, its
577 participants and their potential for genetic research on health and illness. *Int J Epidemiol.*
578 2013;42:689–700.

579 41. Fischbach GD, Lord C. The Simons Simplex Collection: A Resource for Identification of
580 Autism Genetic Risk Factors. *Neuron.* 2010;68:192–195.

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Acknowledgments

Conflict of interest: The authors declare that they have no conflict of interest.

Funding/Support: This research was enabled by support provided by Calcul Quebec

(<http://www.calculquebec.ca>) and Compute Canada (<http://www.computeCanada.ca>).

Sebastien Jacquemont is a recipient of a Bursary Professor fellowship of the Swiss National Science Foundation, a Canada Research Chair in neurodevelopmental disorders, and a chair from the Jeanne et Jean Louis Levesque Foundation. Catherine Schramm is supported by an Institute for Data Valorization (IVADO) fellowship. Petra Tamer is supported by a Canadian Institute of Health Research (CIHR) Scholarship Program. Guillaume Huguet is supported by the Sainte-Justine Foundation, the Merit Scholarship Program for foreign students, and the Network of Applied Genetic Medicine fellowships. Thomas Bourgeron is a recipient of a chair of the Bettencourt-Schueler foundation. This work is supported by a grant from the Brain Canada Multi-Investigator initiative and CIHR grant 159734 (Sebastien Jacquemont, Celia Greenwood, Tomas Paus). The Canadian Institutes of Health Research and the Heart and Stroke Foundation of Canada fund the Saguenay Youth Study (SYS). SYS was funded by the Canadian Institutes of Health Research (Tomas Paus, Zdenka Pausova) and the Heart and Stroke Foundation of Canada (Zdenka Pausova). Funding for the project was provided by the Wellcome Trust. This work was also supported by an NIH award U01 MH119690 granted to Laura Almasy, Sebastien Jacquemont and David Glahn and U01 MH119739. The authors wish to acknowledge the resources of MSSNG (www.mss.ng), Autism Speaks and The Centre for Applied Genomics at The Hospital for Sick Children, Toronto, Canada. We also thank the participating families for their time and contributions to this database, as well as the generosity of the donors who supported this program. We are grateful to all the families who participated in the Simons

Variation in Individuals Project (VIP) and the Simons VIP Consortium (data from Simons VIP are available through SFARI Base). We thank the coordinators and staff at the Simons VIP and SCC sites. We are grateful to all of the families at the participating SSC sites and the principal investigators (A. Beaudet, M.D., R. Bernier, Ph.D., J. Constantino, M.D., E. Cook, M.D., E. Fombonne, M.D., D. Geschwind, M.D., Ph.D., R. Goin-Kochel, Ph.D., E. Hanson, Ph.D., D. Grice, M.D., A. Klin, Ph.D., D. Ledbetter, Ph.D., C. Lord, Ph.D., C. Martin, Ph.D., D. Martin, M.D., Ph.D., R. Maxim, M.D., J. Miles, M.D., Ph.D., O. Ousley, Ph.D., K. Pelphrey, Ph.D., B. Peterson, M.D., J. Piggot, M.D., C. Saulnier, Ph.D., M. State, M.D., Ph.D., W. Stone, Ph.D., J. Sutcliffe, Ph.D., C. Walsh, M.D., Ph.D., Z. Warren, Ph.D., and E. Wijsman, Ph.D.). We appreciate obtaining access to phenotypic data on SFARI base.

Additional Contributions: Julien Buratti (Institute Pasteur), and Vincent Frouin, Ph.D. (Neurospin), acquired data for IMAGEN. Manon Bernard, BSc (database architect, The Hospital for Sick Children), and Helene Simard, MA, and her team of research assistants (Cégep de Jonquière) acquired data for the Saguenay Youth Study. Antoine Main, M.Sc. (UHC Sainte-Justine Research Center, HEC Montreal), Lionel Lemogo, M.Sc. (UHC Sainte-Justine Research Center), and Claudine Passo, Pg.D. (UHC Sainte-Justine Research Center), provided bioinformatical support. Maude Auger, Pg.D.; and Kristian Agbogba, B.Sc. (UHC Sainte-Justine Research Center), provided website development. Dr. Paus is the Tanenbaum Chair in Population Neuroscience at the Rotman Research Institute, University of Toronto, and the Dr. John and Consuela Phelan Scholar at Child Mind Institute, New York.

Role of the Funder/Sponsor: The funder had no role in the design and conduct of the study; collection, management, analysis, or interpretation of the data; preparation, review, or approval of the manuscript; or decision to submit the manuscript for publication.

636 **Tables and Figures** 637

Ascertainment	Cohort	Array type	n=	Females, n (%)	Age in years Mean (SD)	Type of intelligence measures	Z-scored intelligence measure Mean (SD)
Unselected (n=20,151)	IMAGEN [36]	610Kq; 660Wq	1,744	891 (51%)	14.4 (0.4)	WISC-IV (and g-factor, similarities score, vocabulary score, block design score, matrix reasoning score)	0.44 (0.98) ***
	SYS children[37]	610Kq; HOE-12V	967	505 (52%)	15.0 (1.8)	WISC-III (and g-factor using 63 cognitive measures†)	0.30 (0.87) ***
	SYS parents[37]	HOE-12V	602	321 (53%)	49.5 (4.9)	g-factor, 12 cognitive measures‡	0 (1)
	LBC1936[38]	610Kq	504	247 (49%)	70.0 (-)*	Moray House Test (and g-factor)	0.05 (0.96) ***
	CaG-GSA[39]	GSA	2,074	1,094 (53%)	54.3 (7.6)	g-factor, Reasoning, Memory, Reaction time	-0.02 (1.03)
	CaG-Omni2.5[39]	Omni2.5	515	281 (55%)	52.4 (8.6)		-0.10 (1.02)
	CaG (all)[39]	GSA; Omni2.5	2,589	1,375 (53%)	53.9 (7.8)		-0.03 (1.03)
	G-Scot[40]	610Kq	13,745	8,101 (59%)	46.7 (15.0)	g-factor, Logical Memory, Digit Symbol, Verbal fluency, Mill Hill Vocabulary	0.00 (0.99)
Autism (n=3,941)	SSC-1Mv1[41]	1Mv1	332	44 (13%)	9.5 (3.2)	WISC-IV n=19; DAS-II E-Y n=96; DAS-II S-A n=179; Mullen n=12; WASI-I n=26	-0.55 (1.59)
	SSC-1Mv3[41]	1Mv3	1,182	157 (13%)	8.8 (3.5)	WISC-IV n=16; DAS-II E-Y n=531; DAS-II S-A n=539; Mullen n=77; WASI-I n=19	-0.98 (1.66)
	SSC-Omni2.5[41]	Omni2.5	1,048	140 (13%)	9.2 (3.7)	WISC-IV n=10; DAS-II E-Y n=403; DAS-II S-A n=494; Mullen n=124; WASI-I n=17	-1.25 (1.87)
	SSC (all)[41]	1Mv1; 1Mv3; Omni2.5	2,562	341 (13%)	9.03 (3.6)	WISC-IV n=45; DAS-II E-Y n=1,030; DAS-II S-A n=1,212; Mullen n=213; WASI-I n=62	-1.03 (1.75)
	MSSNG[30]	WGS	1,379	275 (20%)	9.2 (4.4)	WISC-IV n=46; WASI-II n=338; Leiter n=372; Raven n=214; Stanford Binet n=281; WPPSI n=128	-0.47 (1.58)
NDD** (n=551)	Ste-Justine-probands	Agilent 180 K array	132	52 (39%)	7.23 (5.46)	WISC-V n=36; WASI-II n=8; WPPSI-IV n=38; Leiter-R n=18; Mullen n=32	-1.31 (1.02)
	Ste-Justine-siblings		87	44 (50%)	7.75 (5.72)	WISC-V n=28; WASI-II n=13; WPPSI-IV n=31; Leiter-R n=3; Mullen n=12	-0.29 (0.98)
	Ste-Justine-parents		310	180 (58%)	37.80 (7.13)	WASI-II	-0.10 (1.16)
	Ste-Justine-other members		22	12 (54%)	43 (21.27)	WASI-II	-0.04 (1.32)

638

639 **Table 1. Cohort descriptions**

640 Cohorts include 24,092 individuals, including 14,874 unrelated individuals. SSC and CaG

641 cohorts were broken down into sub-samples based on array technology (Supplementary

642 methods). †63 and ‡ 12 cognitive measures were respectively used to compute the g-factor in

643 SYS children and parents (Supplementary methods). NDD: neurodevelopmental disorders, SYS:

644 Saguenay Youth Study, CaG: CARTaGEN, LBC: Lothian Birth Cohort, SSC: Simons Simplex

645 Collection; n=number of individuals remaining for analysis after quality control. The mean and
 646 Standard Deviation (SD) for g-factor slightly deviate from 0 and 1 in some cohorts since they
 647 were computed on all available data (before the exclusion of some individuals for poor quality
 648 array) and summarized here only for individuals included in the analyses. *All individuals from
 649 LBC1936 were assessed at 70 years old explaining the absence of SD computation. **The NDD
 650 cohort was used only in the replication analysis and was not included in meta- or mega-analyses.
 651 *** We displayed the Z-scores of IQ, because IQ was preferred to g-factor for all analyses, even
 652 if results were similar (Supplementary Table 1 and 3).
 653

Fig. 1. Effect of intolerant score on general intelligence measured for deletions and duplications.

Meta-analysis estimating the effect of deletions **a.** and duplications **b.**, measured by sum of pLI, on general intelligence (Table S1). X-axis values represent z-scores of general intelligence. Deleting one point of pLI decreases the general intelligence by 0.18 z-scores (2.7 points of IQ). Duplicating one point of pLI decreases the general intelligence by 0.05 z-scores (0.75 points of IQ). The squares represent the effect-size computed for each sample. Their size negatively correlated to variance. Diamonds represent the summary effect across cohorts. Their lengths correspond to the 95% confidence intervals of the mean effect-size. **c.** Estimated proportion of the coding genome within each category defined by LOEUF, encompassed in CNVs present in the mega-analysis according to sample size (randomly selected within the mega-analysis). We

666 observed $N_{\text{CNVs gene}}=6,315$ with $N_{\text{Del. gene}}=2,282$ and $N_{\text{Dup. gene}}=5,223$). **d.** Estimated effect of
 667 $1/\text{LOEUF}$ on general intelligence after removing individuals with a sum of $1/\text{LOEUF}$ larger than
 668 60, 40, 20, 10, 4 and 2.85 (2.85 corresponds to $1/0.35$, the cut-off for intolerance to pLoF
 669 gnomAD). n: number of individuals with a total sum of $1/\text{LOEUF} > 0$.

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Fig. 2. Effect-size of intellectual disability (ID) genes on general intelligence.

a. Venn diagram of ID genes in ASD and in general population cohorts. We identified 66 CNVs encompassing at least one ID-gene in ASD cohorts (31 deletions and 35 duplications) and 60 in the general population (13 deletions and 47 duplications) (Supplementary methods). Genes were previously defined as harboring an excess of *de novo* loss of function (bold) or missense mutations in neurodevelopmental cohorts: (a) *DYNC1H1*, *PHF21A*, *SHANK3*, *TRA2B*, *FOXP1*, *SETD5*, *NR4A2*, *TCF7L2*, *SOX5*, *POU3F3*, *ARID1B*, *EBF3*, *HNRNPU*; (b) *SET*, *ZBTB18*, *DLG4*, *CHAMP1*, *CNOT3*, *U2AF2*, *HIST1H2AC*, *DNM1*, *RAI1*, *CREBBP*, *HIST1H1E*, *ASXL1*, *CABP7*; (c) *PRPF18*, *PPP2R1A*, *EEF1A2*; (d) *TRAF7*, *DEAF1*, *STC1*, *MYT1L*, *BRPF1*,

682 *CBL, SPAST, WDR87, NFE2L3, STARD9, TCF20, KMT2C, FAM200B, KDM5B, CHD2,*
683 *BTF3, ITPR1, HMGXB3. b.* Effect-size of 1/LOEUF on general intelligence estimated in a model
684 using two explanatory variables (sum of 1/LOEUF of deleted and duplicated genes) or 4
685 explanatory variables (sum of 1/LOEUF of ID genes and non-ID genes for deletions and
686 duplication).
687
688

Fig. 3. Concordance between model predictions and published observations for CNV effects on general intelligence and for *de novo* frequency.

a. and **b.** Concordance between model estimates (with 1/LOEUF and ID-genes) and literature of clinical data and UKBB reports for general intelligence loss observed in respectively 27 and 33 recurrent CNVs for a total of ascertained carriers of 47 recurrent CNVs (Supplementary Table 15). X- and Y-values: effect size of CNVs on z-scored general intelligence. **b.** Zoom of the rectangle drawn in the lower left section of panel **a**. We represented values from clinical data by a circle and those from UKBB data by a square. The cross represents the mean value of z-scored IQ loss for the 13 recurrent CNVs observed both in literature and in UKBB. **c.** and **d.** The model

699 uses 2 explanatory variables (1/LOEUF of non-ID-genes and ID-genes). **c.** Probability of *de novo*
700 estimated by our *de novo* model (Y-axis) according to the loss of IQ estimated by a model using
701 1/LOEUF for ID and non-ID genes as two explanatory variables (X-axis). The *de novo* model
702 was fitted on 13,114 deletions (red) and 13,323 duplications (blue) with available inheritance
703 information observed in DECIPHER, CHU Sainte-Justine, SSC, MSSNG, SYS and G-Scot. **d.**
704 Concordance between *de novo* frequency observed in DECIPHER (X-axis) and the probability of
705 being *de novo* estimated by models when excluding recurrent CNVs of the training dataset (Y-
706 axis) 1/LOEUF for ID and non-ID genes as an explanatory variable for 27 recurrent CNVs. The
707 first bisector represents the perfect concordance. Deletions are in red and duplications in blue.
708 Empty circles or square are CNVs encompassing ID-genes. ICC indicates intraclass correlation
709 coefficient (3, 1). Each point represents a recurrent CNV: (1) TAR Deletion; (2) 1q21.1 Deletion;
710 (3) 2q11.2 Deletion; (4) 2q13 Deletion; (5) *NRXNI* Deletion; (6) 2q13 (*NPHPI*) Deletion; (7)
711 3q29 (*DLGI*) Deletion; (8) 7q11.23 (William-Beuren) Deletion; (9) 8p23.1 Deletion; (10)
712 10q11.21q11.23 Deletion; (11) 13q12.12 Deletion; (12) 13q12 (*CRYLI*) Deletion; (13) 15q13.3
713 (BP4-BP5) Deletion; (14) 15q11.2 Deletion; (15) 16p11.2-p12.2 Deletion; (16) 16p13.3 ATR-16
714 syndrome Deletion; (17) 16p11.2 Deletion; (18) 16p11.2 distal Deletion; (19) 16p13.11 Deletion;
715 (20) 16p12.1 Deletion; (21) 17p11.2 (Smith-Magenis) Deletion; (22) 17q12 Deletion; (23)
716 17q21.31 Deletion; (24) NF1-microdeletion syndrome Deletion; (25) 17p12 (*HNPP*) Deletion;
717 (26) 22q11.2 Deletion; (27) TAR Duplication; (28) 1q21.1 Duplication; (29) 2q21.1 Duplication;
718 (30) 2q13 Duplication; (31) 2q13 (*NPHPI*) Duplication; (32) 7q11.23 Duplication; (33)
719 10q11.21q11.23 Duplication; (34) 13q12.12 Duplication; (35) 15q11q13 (BP3-BP4) Duplication;
720 (36) 15q11.2 Duplication; (37) 15q13.3 Duplication; (38) 15q13.3 (*CHRNA7*) Duplication; (39)
721 16p11.2 Duplication; (40) 16p11.2 distal Duplication; (41) 16p13.11 Duplication; (42) 16p12.1
722 Duplication; (43) 17p11.2 Duplication; (44) 17q12 (*HNF1B*) Duplication; (45) 17p12 (*CMT1A*)
723 Duplication; (46) Trisomic 21 Duplication; (47) 22q11.2 Duplication.

Fig. 4. Effect-size on general intelligence of individual genes encompassed in CNVs and their GOterms enrichment.

The light grey histogram represents the distribution of LOEUF values for 18,451 autosomal genes. The blue line represents the estimates for a gene in each of the 4 categories of LOEUF included in the model (Supplementary methods): highly intolerant genes (LOEUF < 0.2, n=980), moderately intolerant genes (0.2 ≤ LOEUF < 0.35 n=1,762), tolerant genes (0.35 ≤ LOEUF < 1, n=7,442) and genes highly tolerant to pLoF (LOEUF ≥ 1, n=8,267). The orange line represents the estimated effect-size of 37 categories of genes based on their LOEUF values (sliding windows=0.15) in the model (Supplementary methods). Genes with a LOEUF below 0.35 (vertical red line) are considered to be intolerant to pLoF by gnomAD. Left Y-axis values: z-scored general intelligence (1 z-score is equivalent to 15 points of IQ) for deletion. Right Y-axis values: number of genes represented in the histogram. For **Fig. 6b** each point represents a GOterm for which enrichment was observed for all intolerant (n=2,742) or tolerant genes (n=7,442) (**b.**), for all intolerant (n=609) or tolerant genes (n=2,251) encompassed in CNVs (**C**) when compared to the whole coding genome (Bonferroni). **b.** X-axis: % of genes included in the

741 GTerm genome-wide; Y-axis: % of genes included in the GTerm for all intolerant
742 ($0 < \text{LOEUF} < 0.35$) and tolerant genes ($0.35 \leq \text{LOEUF} < 1$).