Temperature-dependence of spontaneous mutation rates

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Mutation is the source of genetic variation and the fundament of evolution. At the interphase of ecology and evolution, temperature has long been suggested to have a direct impact on realised spontaneous mutation rates. The question is whether mutation rates can be a species-specific constant under variable environmental conditions, such as variation of the ambient temperature. By combining mutation accumulation with whole genome sequencing in a multicellular organism, we provide empirical support to reject this null hypothesis. Instead mutation rates depend on temperature in a U-shaped manner with increasing rates towards both temperature extremes. This relation has important implications for mutation dependent processes in molecular evolution, processes shaping the evolution of mutation rates and even the evolution of biodiversity as such.

Mutation has been described as the ‘quantum force’ of biology¹: pervasive throughout the tree of life, fundamental basis of evolution and notoriously difficult to measure. Evidence for the variation of mutation rates (μ) has been accumulating for a century, with pioneering investigations in *Drosophila* dating back to the 1930s^{2,3} and more recent studies estimating the variability of μ in microorganisms^{4–6}, plants^{7–9}, invertebrates^{10–14} as well as vertebrates^{15–17}. On the intraspecific level, the spontaneous μ was often claimed to be a species-specific constant¹⁸. The question about which factors directly drive mutational rate variation remains unresolved, particularly in multicellular organisms. Temperature has long been suggested as a major determinant of μ variation. Experiments in the first half of the 20th century (reviewed in Lindgren¹⁹) reported generally large effects of temperature treatments on phenotypically visible mutations. As reported by Muller²⁰ in the first study on this topic, “both the direction of the effect of temperature on the time-rate of mutation, and its approximate magnitude, are the same as in the case of its effect on the time-rate of ordinary chemical reactions”. Interestingly, however, also cold treatments increased the apparent mutation rate in some studies^{21,22}. Apart from these early studies, mainly performed even before DNA was identified as the carrier of genetic information, there is surprisingly little empirical evidence if and to what extent temperature modifies the spontaneous rate of mutation on the molecular level. Higher temperature stress had a significant effect on microsatellite μ in *Caenorhabditis elegans*¹⁰. Studies on *Drosophila melanogaster* suggest that indeed temperature dependent metabolic activities increased the somatic mutation rate as consequence of oxidative stress²³. Stressful temperature conditions increased μ in experimental evolution lines of the seed beetle *Callosobruchus maculatus*¹². And the temperature dependent mass specific metabolic rate seemed to influence the interspecific molecular clock rate that depends at least partially on μ ²⁴.

Despite a century-long research on the potential impact of temperature on mutational variation, we still lack a thorough understanding of mechanisms underlying the temperature-dependence

of μ . Empirical evidence for increased μ under stressful temperature conditions is strong, however, the relation between natural, and thus relevant, temperature ranges and the variability of μ on an intraspecific level is still unclear. It was therefore our aim to test the following null hypothesis (H_0): The spontaneous mutation rate is a species-specific constant and independent of environmental conditions.

To provide a fine-scale resolution of the relation between temperature and μ , we report here directly estimated spontaneous single nucleotide mutation rates in the non-biting midge *Chironomus riparius* under different temperature regimes. *C. riparius* is a fully developed system for multigeneration experiments²⁵ with comprehensive genomic resources²⁶. A haploid spontaneous single nucleotide μ of 2.1×10^{-9} has been previously reported²⁷. Moreover, there is indirect evidence for the temperature-dependence of μ among *C. riparius* populations²⁸. Here, we performed short-term mutation accumulation experiments at temperatures between 12 °C and 26 °C (Fig. 1), representative for the natural range for active development of the species. Temperatures below 12 °C will generally induce developmental pausing for larval overwintering²⁸ and temperatures constantly above 26 °C lead to significant negative fitness effects²⁹. As expected, the generation time of individuals in our experiments decreased with increasing temperature (Fig. 1) and we also observed a reduction of body size with increasing temperature, a well-known phenotypic response for ectotherms³⁰. Whole genome data of the RefPool as ancestral state and 58 single individuals across all mutation accumulation lines (MAL) were passed along the *de novo* mutation calling pipeline established for this species²⁷ (and see methods). Across a total number of $5.5 \cdot 10^9$ analysed genomic positions, we identified 106,149 *de novo* mutations in MAL of all temperatures. Most of these mutations (105,948) were found in five MAL with extraordinary high mutation load (Supplemental Information S1). The occurrence of such so-called mutator lines is a known phenomenon for MAL experiment³¹ and has also been previously observed in *C. riparius*³².

Figure 1: Experimental set up of mutation accumulation at different constant temperatures. The succession of circles indicates how many generations (F1-F5) could be passed under the different temperature conditions. Circles are labelled with the total number of MAL that survived until the respective generation. Mean generation time (GT) is decreasing with increasing temperature as indicated in the detail graph. Number of generational passages (GP = MAL • generation) sums up to a total of 205.

Table 1: Overview of results per temperature, excluding mutator lines: **G**enerational **P**assages of successfully sequenced MAL, mean number of **C**allable **S**ites, total number of de novo mutations, number of **S**ingle **N**ucleotide **M**utations, number of **S**ingle **N**ucleotide **I**ndels, mean haploid mutation rates per generation, respectively of total mutations, SNMs and SNIs.

temperature	GP	mean CS	mutations total	SNMs	SNIs	μ_{total}	μ_{SNM}	μ_{SNI}	transitions	transversions
12	20	8.92E+07	27	23	4	7.70E-09	6.60E-09	1.26E-09	17	6
14	11	7.33E+07	6	5	1	4.02E-09	3.42E-09	9.25E-10	3	2
17	45	1.07E+08	26	20	6	2.77E-09	2.13E-09	6.75E-10	8	12
20	50	1.23E+08	52	42	10	4.27E-09	3.28E-09	8.55E-10	18	24
23	33	9.60E+07	40	37	3	6.40E-09	5.90E-09	5.45E-10	26	11
26	26	7.87E+07	50	45	5	1.24E-08	1.12E-08	1.35E-09	28	17

The total rate of spontaneous mutations per haploid genome per generation ranged between $2.77 \cdot 10^{-9}$ and $1.24 \cdot 10^{-8}$ across all temperature experiments (Tab. 1). This range covers the previously reported μ for *C. riparius*²⁷ and the magnitude is comparable to μ estimates of different other insects, as e.g. *D. melanogaster*³³ ($2.8 \cdot 10^{-9}$), *Heliconius melpomene*³⁴ ($2.9 \cdot 10^{-9}$).

Temperature had a significant impact on realised spontaneous μ in our MAL experiments. The lowest rate of total mutations was calculated for the 17 °C experiment with $2.77 \cdot 10^{-9}$ (95% HDI: lower $1.74 \cdot 10^{-9}$ and upper $3.78 \cdot 10^{-9}$; Tab. 1 and Supplemental Information S2). From this minimum, μ increased monotonically with the absolute temperature difference in both directions (Fig. 1). μ at the extreme temperatures 12°C ($7.70\text{E-}09$, 95% HDI: lower $4.90 \cdot 10^{-9}$ and upper $1.06 \cdot 10^{-8}$, $\Delta 5$ °C from minimum) and 26°C ($1.24\text{E-}08$, 95% HDI: lower $8.85 \cdot 10^{-9}$ and upper $1.57 \cdot 10^{-8}$, $\Delta 9$ °C from minimum) was 2.79, respectively 4.54 times higher than the minimum rate (respective posterior probabilities 99.8% and 100%, Tab. 2). In all cases there was predominant to conclusive evidence that the next temperature step in the direction of both extremes induced a higher μ (Tab. 2).

When decomposing total μ into SNM (μ_{SNM}) and SNIs (μ_{SNI}), the distribution of the data in relation to temperature was almost identical for the former (Fig. 2, bottom). Due to the low number of *de novo* SNIs, μ_{SNI} were comparably low (Tab.1, Fig. 2 bottom) and, at first sight, no clear U-shape distribution in relation to temperature appeared. However, fitting second order polynomial regression models to the data found a U-shape relation between temperature and all measured rates, with a respective goodness of fit of $r^2=0.98$, $r^2=0.99$ and $r^2 = 0.73$ (Fig. 2).

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Figure 2: Spontaneous mutation rates per haploid genome per generation in relation to temperature of the MAL experiment. **Top:** Mutation rate of the total sum of de novo mutations. **Bottom:** Separate mutation rates of single nucleotide mutations (SNM) and single nucleotide indels (SNI). Nonlinear fit as second order polynomial regression, R^2 values are given to describe the goodness of model fit. Corresponding test statistics listed in Supplemental Information S2.

Table 2: Selected μ_{total} comparisons of rate ratios between temperatures (cf. Supplemental Information S2 for full table). Posterior probabilities indicate the support of ratios > 1 between compared temperatures.

Comparison	Rate Ratio	Post. Prob. Rate > 1 (%)
12°C vs. 17°C	2.79	99.8
26°C vs. 17°C	4.54	100
12°C vs. 14°C	1.92	99.9
14°C vs. 17°C	1.46	53.1
20°C vs. 17°C	1.54	99.3
23°C vs. 20°C	1.49	78.7
26°C vs. 23°C	1.92	97.5

Our results thus deliver unequivocal empirical evidence to reject the H_0 : Single nucleotide mutation rates are not constant within *C. riparius* but depend on temperature with a minimal rate at an optimum temperature. Such a relation was theoretically assumed³⁵ but empirical evidence was largely lacking. Before the present study, only Ogur and colleagues³⁶ documented the effects of a fine scaled, wider temperature range on spontaneous mutations in *Saccharomyces* species sixty years ago. They used respiration deficiency as phenotypic mutation proxy and found a U-shaped relation, similar to what we found for nucleotide mutations in our whole-genome scale approach in a multicellular organism (Fig. 2). Other studies were restricted either to only partially covering the respective species' physiological temperature range and/or extreme temperatures that would a priori evoke stressful conditions for the species of choice e.g. 10,37.

There are a few mechanistic hypotheses that could explain the observed pattern^{11,35,37}. The metabolic rate hypothesis (MH) suggests that increased metabolic activity leads to the production of free oxygen radicals (ROS) faster than can be effectively eliminated by the organismal antioxidative stress response which act as endogenous mutagens³⁸. As ambient temperature determines the metabolic rate of ectotherm organisms, μ should rise with temperature as a consequence of increased oxidative stress, which causes DNA damage and as a result, sometimes erroneous repair³⁹. The effect of high temperatures leading to increased mutation rates has been documented for almost a century^{2,10,19,20,24} and recent experiments with *Escherichia coli* underline this³⁷. However, it is well established that also low temperatures can induce oxidative stress, even though the mechanisms are not well known⁴⁰. Antioxidative stress response to both high and low temperatures was e.g. observed in *C. riparius*⁴¹. It was speculated that low temperatures either weakens the systems of ROS elimination, and/or enhances ROS production³⁸. In our experiments, the number of successful MAL and their respective amount of generational passages (Fig. 2) reveals the

increased stress-level at both of the two extreme temperatures. At 12 and 26 °C, all MAL were lost after the second generation. Whereas the majority of lines reached the fifth generation of mutation accumulation at the intermediate temperatures 17, 20, and 23 °C. Metabolic effects alone could therefore explain the observed pattern.

However, a negative correlation between temperature and μ could also arise from the positive relation between ambient temperature and generation times in the ectotherm *C. riparius*⁴². Whatever mutagens are acting (endogenous ROS production, cosmic rays), their time to induce mutations in the germline should increase with generation time. One could term this the generation length hypothesis (GL). Several studies are lending support for the GL: in primates μ can be predicted via the reproductive longevity¹⁵ and the mammalian male mutation bias is higher for species with long generation times⁴³. There is evidence that this effect does not depend on the absolute number of cell divisions⁴⁴. Also in invertebrates there is significant evidence for a correlation between rates of molecular evolution and generation times¹¹. Even though these studies investigated only the relation between temperature independent generation time and μ , their results implied that the temperature-dependent changes in generation time of ectotherms could impact μ . Direct support comes from temperature experiments impacting on life-history and μ in seed beetles¹². Martin and Palumbi⁴⁵ already concluded their comparative analysis on the relation of body size, metabolic rate, generation time and the molecular clock that our understanding of molecular evolution could be improved when reconsidering “the generation time hypothesis to include physiological effects such as the metabolic rate”.

There is empirical evidence for all of these processes to play a role³⁵, leading to potentially complex interactions in each particular case. The here discovered U-shaped relation between μ and temperature could therefore be either due to increased oxidative stress towards physiological extreme temperatures (MH) and/or an interaction with inversely increasing generation times with decreasing temperatures (GL). As generation time and metabolism are intricately linked in ectotherm organisms, additional studies separating these effects are necessary to infer the acting processes.

We additionally disentangled temperature-effects on the mutation spectrum. Excluding mutator lines, we detected 201 de novo mutations during a total of 177 generational passages with 172 SNMs (86 %) and 29 indels (Tab. 1). SNMs divided in 100 transitions and 72 transversions (ratio=1.38, Tab. 1) with a non-random distribution of mutation types among temperatures. Overall, transitions occurred significantly more often than transversions (Mann-Whitney U=53.5, p=0.0015). This pattern provides statistical support for the tendency that we have observed in a previous study with the same species²⁷. The transition bias is a well-known pattern of molecular evolution⁴⁶, however, the cause has not yet been resolved. There are two alternative hypotheses

relating the bias to either a mutational^{47,48} or selective^{49,50} mechanism. With regard to our experimental design of mutation accumulation, we can neglect selection in favour of “less severe” biochemical amino-acid changes due to transitions, and thus reject the ‘transition-bias due to selection’ hypothesis. More importantly, our data allows a more fine-scaled resolution of the mutation spectrum and we found temperature to have a significant effect on mutation types ($F=4.94$, $p=0.0028$), explaining 28 % of the total variation (Fig. 2). Towards the temperature extremes, frequency of all four transition types increased, whereas the effect on transversions was unidirectional with increasing temperatures and mostly due to an increase of A↔T mutations (Fig. 2). This temperature-dependent mutation spectrum provides important evidence that temperature is more than a general stress factor acting on mutational processes, what has been claimed otherwise in a study comparing the effect of temperature stress on spontaneous mutations in nematods¹⁰. The observed differences in the mutation spectrum suggest that temperature as a bipolar factor is shaping mutational variation through different molecular mechanisms.

Figure 3: Heatmap of detected SNM types per experimental temperature. Transitions in rows 1 and 2, transversions in rows 3 to 6. Different mutation types explain 43 % ($F=7.52$, $p=0.0002$) and temperatures 28 % ($F=4.94$, $p=0.0028$) of the total variance.

Our observation that μ varies significantly in response to environmental temperature has important implications for mutation dependent processes in molecular evolution, processes shaping the evolution of mutation rates and even the evolution of biodiversity as such. Temperature as an abiotic environmental factor is considered to hold a key position at the interphase of ecology and

evolution for its influence on organismal physiology⁵¹, population divergence⁵² and now also mutational variability. The substantial variation of μ in response to ambient temperature, and perhaps also other environmental factors, is especially relevant for our understanding of evolutionary processes under natural conditions. For the multivoltine *C. riparius*, for example, μ in a natural population is expected to vary by a factor of more than four among the generations throughout a year (Fig. 4). This suggests that the mean effective μ over the generations in a year under natural conditions is five times higher than measured in the laboratory²⁷. Most interestingly, the effective μ is rather driven by cold temperatures (Fig. 4), with respective consequences for the expected mutational spectrum.

Figure 4: Expected variation of mutation rate calculated per generation for a natural *C. riparius* field population (left y-axis) with regard to the annual variation in water temperature (Hasselbach in Hesse, Germany) averaged for each generation period (right y-axis). Due to the temperature-dependence of generation times, *C. riparius* can pass eight generations per year (x-axis). The mean effective μ is 5.29-fold higher than μ measured under laboratory conditions²⁷ (indicated by arrow).

Different other population genetic parameters depend on μ , in particular the estimation of the long-term effective population size (N_e) from theta⁵³. Given the variation of μ with regard to the natural temperature range (Fig. 4), it appears advisable to use a mean effective rate to obtain accurate estimates of N_e . The same applies for the estimation of divergence times from sequence divergence data, with the additional complication that e.g. changes in N_e over time may have additionally influenced the base line μ with regard to the cost-fidelity hypothesis⁵⁴. Last but not least, the evolutionary speed hypothesis (ESH), aiming to explain the latitudinal biodiversity gradient⁵⁵, assumed as one molecular underpinning for faster evolutionary processes in lower latitudes that

the higher temperatures in these regions increase μ , leading thus to faster mutation accumulation²⁸. If, however, the finding of increased μ towards both temperature extremes is a general phenomenon, it would make this particular aspect of the ESH hypothesis obsolete.

Our scenario of mutation variation to be expected in natural populations also raises the question whether the position of the here observed μ optimum is in itself a species-specific constant or whether this position on the temperature scale can evolve in response to local temperature conditions. Answers to this question will enhance our understanding of the interplay between climate variability and molecular evolution⁵⁶.

Methods

Mutation accumulation experiments. The experimental set up for mutation accumulation followed the exact procedure described in Oppold & Pfenninger²⁷. In this previous study, we performed the mutation accumulation experiment with the ‘Laufer’ strain of *C. riparius* under constant conditions at 20 °C. Simultaneously and now presented in this study, we established mutation accumulation lines (MAL) at five additional temperatures: 12 °C, 14 °C, 17 °C, 23 °C and 26 °C. Ten MAL were initiated for each temperature condition, with additional backup MAL to compensate for the loss of lines due to lethal inbreeding effects.

Whole genome resequencing. To establish the genetic baseline of the ancestral state against which to compare de novo mutations, we sequenced the parental individuals as pool of their offspring, hereafter called RefPool. Head capsules of adult midges were pooled for DNA extraction and pooled sequencing on a Illumina HiSeq platform as 150 bp paired-end library to an expected coverage of 60-80X. After mutation accumulation across generations, a single female midge of each MAL was whole genome sequenced on an Illumina HiSeq platform as 150 bp paired-end library to an expected mean coverage of 25X (see Oppold and Pfenninger 2017 for details).

Whole genome sequencing data of RefPools and individuals per MAL was cleaned from adapters and quality trimmed using the automatic wrapper script *autotrim* (available at <https://github.com/schell/autotrim>). Genomic data of the 20 °C mutation accumulation experiment (RefPool R1 and 20 °C MAL) has already been analysed in our previous study²⁷, however mapped to an older assembly version. To make use of the more complete and less fragmented genome assembly, cleaned reads of RefPools and MAL (including the data published in Oppold & Pfenninger, 2017) were mapped against the latest version of the *C. riparius* genome assembly²⁶ with masked repeat- and TE-regions (see Schmidt et al. 2020 for TE- and repeat-library). We used the Burrows-Wheeler Aligner *bwa mem*⁵⁷ (0.7.17-r1188) for mapping and followed the GATK best

practices pipeline⁵⁸ to remove duplicates⁵⁹ (picard tools v2.20.8), realignment around indels (insertions and deletions, supported by *GATK* version 4) and recalibration of bases.

Estimation of spontaneous mutation rates. We define de novo mutations to be single base exchanges (hereafter called single nucleotide mutations, SNM) and single base insertions and deletions (hereafter called single nucleotide indels, SNI) that happen in the germline and are thus inherited to the next generation. After a defined number of generational passages (mutation accumulation) and when compared against an ancestral genotype, de novo mutations can be identified to be novel and unique to a respective MAL. While these criteria are obviously straightforward, their application to whole genome data required a stringent control for base and read quality to avoid false positives as well as false negatives. We therefore combined the statistical approach for genome-wide detection of de novo mutations^{27,33} with a probabilistic mutation calling approach⁶⁰. Whilst SNIs can only be detected via the statistical approach, the probabilistic mutation calling implemented in the tool *accuMulate*⁶⁰ allowed for a more sensitive detection of SNMs lowering the false negative rate. Each set of MAL samples per temperature was compared against its respective RefPool (R1 or R2). For the statistical approach, each MAL sample was analysed as separate bam file, whereas *accuMulate* requires the input of one overall bam file into which all MAL samples per temperature including the respective RefPool were merged. *accuMulate* output was further filtered according to the following criteria (cf. Winter et al. 2018 for description of categories): probability of a mutation ≥ 0.9 , probability of one mutation ≥ 0.9 , probability of correct descendant genotype ≥ 0.9 , minimum depth $= n_{\text{MAL}} \cdot 15X + 15X$, maximum depth $= n_{\text{MAL}} \cdot 47X + 195X$, number of mutant allele in RefPool $= 0$, mapping quality difference ≤ 1.96 , insert size difference ≤ 1.96 , strand bias ≥ 0.05 , pair-mapping rate difference ≥ 0.05 . For the calculation of mutation rates and to compensate for coverage variation among samples (Supplementary Information S1), we used the separate bam files to more precisely define the number of callable sites for each sample details described in 27 as direct divisor, instead of using the *denominator* of the *accuMulate* package. Finally, the combination of the two mutation detection approaches gave us the most robust and comprehensive set of de novo mutation candidates, all of which were visually curated in *IGV* (v2.6.3),⁶¹ excluding SNMs in the 23 °C mutator lines (see below). A MAL was defined to be a mutator line as soon as we observed a 4-fold increase of its mutation rate when compared against the rates of the remaining MAL of the same temperature.

Test on temperature dependence of rates. To account for the Poisson distribution of our data (mutation counts), we used a Bayesian framework (R package Bayesian First Aid⁶²) to infer credible intervals (95% highest density intervals of posterior distributions, HDI) and to test differences between rates of different temperatures. To allow for the observation of zero mutations, the model uses Jeffreys prior on lambda. The posterior distribution was sampled with default values

(3 chains, 5000 iterations). Posterior probability indicates support of rate ratios. Non-overlapping HDIs of two rates give decisive support for differences between temperatures.

Estimation of mutational variation for natural population.

Daily water temperature data for a *C. riparius* population in a small river (Hasselbach, Hessen, Germany 50.167562°N, 9.083542°E) of a nearby waste water treatment plant was analysed for the years 2009-2018. A population model of the temperature dependence of the generation time was fitted to the averaged data to obtain the number and length of the seasonal generations. The mean water temperature for each generation was then used to infer the expected μ from the second order polynomial function fitted on the relation between experimental temperature and μ (Fig 2, top).

Data availability

Genomic data generated and analysed for this study are available at European Nucleotide Archive (ENA project number: pending). Detailed lists of de novo mutations and Bayesian estimates can be found in the Supplemental Information linked to this article.

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Author contributions

A.-M.W. and M.P. conceived the study, A.M.W. performed the experiments and collected the data, A.-M.W. and M.P. analysed the data and equally contributed to writing the manuscript.

Supplemental Information

Supplemental Information S1: Detailed lists of de novo mutations per replicate

Supplemental Information S2: Details on statistics of Bayesian estimation and second order polynomial regression