

Extreme climatic events but not environmental heterogeneity shape within-population genetic variation in maritime pine

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1 Abstract

2 How evolutionary forces interact to maintain quantitative genetic variation within populations has
3 been a matter of extensive theoretical debates. While mutation and migration increase genetic
4 variation, natural selection and genetic drift are expected to deplete it. To date, levels of ge-
5 netic variation observed in natural populations are hard to predict without accounting for other
6 processes, such as balancing selection in heterogeneous environments. We aimed to empirically
7 test three hypotheses: (i) admixed populations have higher quantitative genetic variation due to
8 introgression from other gene pools, (ii) quantitative genetic variation is lower in populations from
9 harsher environments (i.e. experiencing stronger selection), and (iii) quantitative genetic variation
10 is higher in populations from spatially heterogeneous environments. We used phenotypic mea-
11 surements of five growth, phenological and functional traits from three clonal common gardens,
12 consisting of 523 clones from 33 populations of maritime pine (*Pinus pinaster* Aiton). Populations
13 from harsher climates (mainly colder areas) showed lower genetic variation for height in the three
14 common gardens. Surprisingly, we did not find any association between within-population genetic
15 variation and environmental heterogeneity or population admixture for any trait. Our results sug-
16 gest a predominant role of natural selection in driving within-population genetic variation, and
17 therefore indirectly their adaptive potential.

18 1 Introduction

19 Most complex traits show substantial heritable variation in natural populations. How evolution-
20 ary forces interact to maintain such variation remains a long-standing dilemma in evolutionary
21 biology and quantitative genetics (Johnson and Barton 2005). While mutation and genetic drift
22 have straightforward roles, generating and eliminating variation respectively, the effect of natural
23 selection is more complicated (Walsh and Lynch 2018). Stabilizing selection, i.e. the selection of
24 intermediate phenotypes, is often strong in natural populations (Hereford et al. 2004). This type
25 of selection is expected to deplete genetic variation (Fisher 1930), either directly on the focal trait
26 or indirectly via pleiotropic effects (Johnson and Barton 2005). Theoretical models based on the
27 balance between mutation, drift and stabilizing selection support this idea, but they suggest lower
28 fitness heritability values than those generally observed in empirical studies (Johnson and Barton
29 2005). Balancing selection encompasses various evolutionary processes that can maintain greater
30 than neutral genetic variation within populations (Mitchell-Olds et al. 2007). The most widely
31 studied of these processes are heterozygote advantage, frequency-dependent selection, e.g. in dis-
32 ease resistance or self-incompatibility systems (Bergelson et al. 2001, Charlesworth et al. 2005),
33 and temporally or spatially fluctuating selection pressures (Felsenstein 1976). The maintenance
34 of stable polymorphism in spatially heterogeneous environments was first theorized by Levene’s
35 archetypal model (1953), under the assumptions of random mating within generations and soft
36 selection. Since then, a large corpus of single-locus and polygenic models, most often determinis-
37 tic, have generally concluded that genetic polymorphisms can only be maintained under restrictive
38 conditions (Spichtig and Kawecki 2004, Byers 2005). In this line, McDonald and Yeaman (2018)
39 showed with stochastic individual-based simulations that substantial within-population genetic
40 variation can be maintained in spatially heterogeneous environments at intermediate migration
41 rates, regardless of population size. However, the relative importance of the different evolutionary
42 forces driving within-population genetic variation remains largely unknown.

43 Long-dating empirical work has addressed the evolutionary processes underlying the maintenance
44 of genetic and discrete-trait polymorphisms (reviewed in Hedrick 1986, 2006), e.g. plant-pathogen
45 interactions (Karasov et al. 2014), antagonistic pleiotropy (Carter and Nguyen 2011), environmen-
46 tal heterogeneity (Chakraborty and Fry 2016), and temporal fluctuations (Bergland et al. 2014).
47 Genomics have allowed the broad application of genome-wide scans for signatures of selection.
48 Overall these scans suggest that many loci are under adaptive directional selection (Barreiro et
49 al. 2008, Fu and Akey 2013) and that the proportion of genetic polymorphisms maintained by
50 environmental heterogeneity tends to be low (Hedrick 2006). However those scans typically have
51 low power to detect signatures of balancing selection or local adaptation (Fijarczyk and Babik
52 2015). Far fewer empirical studies have focused on assessing the distribution and extent of the
53 quantitative genetic variation within populations, and its underlying causes (Lynch and Walsh
54 1998). Traits more closely related to fitness, such as life-history traits, have generally higher
55 additive genetic variance, but lower heritabilities, than morphometric traits (Price and Schluter
56 1991, Houle 1992, Kruuk et al. 2000). The hypothesis that populations evolving under strong
57 selection pressures display lower levels of genetic variation has been supported in experimentally

58 evolving quail populations under unfavorable vs favorable treatments (Marks 1978), in controlled
59 experiments (Colautti et al. 2010; but see Merilä et al. 2004, Stock et al. 2014), in natural popu-
60 lations of *Drosophila birchii* subject to climatic selection (but see *D. bennata* and *D. serrata*; van
61 Heerwaarden et al. 2009) and in some natural populations of great tits subject to varying levels
62 of food availability (Charmantier et al. 2004). Higher genetic variation in populations evolving
63 under spatially varying selection pressures is supported by experimental evolution of *Drosophila*
64 populations (Mackay 1981, Huang et al. 2015; but not Yeaman et al. 2010) and in forest trees
65 evaluated in common gardens (Yeaman and Jarvis 2006). The lack of general trends from these
66 empirical studies can be explained by method-specific pitfalls to accurately estimate quantitative
67 genetic variation, e.g. the genetic and environmental variances are hard to disentangle in the wild,
68 and when estimated in common gardens, their environment-dependent nature does not allow for
69 wide generalization of estimates (Hoffmann and Parsons 1991, Merilä et al. 2001, Charmantier
70 et al. 2004). In addition, gene flow has been hypothesized to have either a positive effect on the
71 adaptive potential, by increasing standing genetic variation, or a negative effect via gene swamping
72 (Kremer et al. 2012, Tigano and Friesen 2016), which may depend on the spatial scale considered
73 (Bridle et al. 2009).

74 Forest trees have specific life-history traits and genomic features making them interesting model
75 species in population and quantitative genetic studies (Petit and Hampe 2006, Savolainen et al.
76 2007). Compared to crop species, they remain largely undomesticated (Neale and Savolainen 2004).
77 Most forest trees are outcrossing, have high lifetime reproductive output and long generation times.
78 They often display important gene flow among populations through long-distance pollen dispersal
79 (Kremer et al. 2012). They show slow rates of macroevolution (i.e. low nucleotide substitution
80 rates and low speciation rates; Petit and Hampe 2006), generally have large effective population
81 sizes, with distributions often covering a wide range of environmental conditions (Alberto et al.
82 2013). Extensive work has revealed strong clines at large geographical scales in the population-
83 specific mean values of phenotypic traits (reviewed in Savolainen et al. 2007, Benito Garzón et al.
84 2019), e.g. phenological traits with latitude or altitude (Alberto et al. 2011, Thibault et al. 2020) or
85 height growth with cold hardiness (Rehfeldt et al. 1999, Leites et al. 2012). Genetic differentiation
86 at microgeographic spatial scales has also been repeatedly observed (reviewed in Linhart and
87 Grant 1996, Jump and Peñuelas 2005, Scotti et al. 2016), suggesting rapid rates of microevolution
88 (Petit et al. 2004, Petit and Hampe 2006). Possible explanations include the fact that forest
89 trees have high levels of genetic diversity and that most of their quantitative and neutral genetic
90 variation is within populations (Hamrick 2004). To our knowledge, only two empirical studies
91 investigated the potential causes underlying the maintenance of quantitative trait variation within
92 forest tree populations. Yeaman and Jarvis (2006) showed that 20% of growth genetic variation in
93 lodgepole pine populations was attributable to regional heterogeneity, suggesting an important role
94 of gene flow and varying selection pressures. In the neotropical oak *Q. oleoides*, Ramírez-Valiente
95 et al. (2019) found lower quantitative genetic variation in harsher environments, but not higher
96 quantitative genetic variation in temporally fluctuating environments. They also suggested only a
97 marginal effect of genetic structure and diversity on the maintenance of within-population genetic
98 variation.

99 In this study, we aimed to test competing hypotheses regarding the relationship between quan-
100 titative genetic variation within maritime pine populations and the potential underlying drivers
101 that maintain this variation. We used phenotypic measurements of growth (height), phenological
102 (bud burst and duration of bud burst) and functional ($\delta^{13}\text{C}$ and specific leaf area, SLA) traits
103 from three clonal common gardens, consisting of 522 clones (i.e. genotypes) from 33 populations,
104 spanning all known gene pools in the species (Jaramillo-Correa et al. 2015) and genotyped for
105 5,165 SNPs. For each trait, we compared Bayesian hierarchical models that estimate the relation-
106 ship between the total genetic variances within populations and some potential drivers, namely
107 climate's harshness at the locations of origin of the populations (i.e. drought intensity and se-
108 vere cold events), environmental heterogeneity in the forested areas surrounding the populations,
109 and the level and origin of admixture in the populations, as estimated with SNP markers. The
110 competing, but not mutually exclusive, hypotheses tested are: i) the most admixed populations
111 have higher quantitative genetic variation due to introgression from other gene pools, and this
112 relationship is proportional to the divergence between sink and source gene pools; ii) quantitative
113 genetic variation is lower in populations that have evolved in harsher environments, as a result
114 of higher selection pressures in these regions; and iii) quantitative genetic variation is higher in
115 populations that have evolved in spatially heterogeneous environments. Importantly, the last two
116 hypotheses require the action of natural selection, while the first does not. Therefore, we expect
117 the last two hypotheses to be mostly supported for fitness-related traits, while the first hypothesis
118 may apply uniformly to all traits. Determining the patterns of within-population quantitative ge-
119 netic variation across species' ranges and the relative importance of the evolutionary forces driving
120 the maintenance of such variation is necessary to assess the evolutionary potential of forest tree
121 populations. Empirical studies tackling these questions remain extremely rare in forest trees (but
122 see Yeaman and Jarvis 2006, Ramírez-Valiente et al. 2019), yet they are much needed to anticipate
123 forest tree responses to ongoing global change and therefore develop adaptive management and
124 conservation strategies.

125 **2 Materials & Methods**

126 **2.1 Maritime pine, a forest tree growing in heterogeneous environ-** 127 **ments**

128 Maritime pine (*Pinus pinaster* Ait., Pinaceae) is a wind-pollinated, outcrossing and long-lived
129 tree species with large ecological and economical importance in western Europe and North Africa.
130 Maritime pine is largely appreciated for its wood, for stabilizing coastal and fossil dunes and, as
131 a keystone species, for supporting biodiversity (Viñas et al. 2016). The distribution of maritime
132 pine natural populations is scattered and covers a wide range of environmental conditions. Several
133 studies have provided evidence of genetic differentiation for adaptive traits in this species, suggest-
134 ing local adaptation (e.g. González-Martínez et al. 2002, de Miguel et al. 2020). Maritime pine can
135 grow in widely different climates: the dry climate along the northern coasts of the Mediterranean

136 Basin (from Portugal to western Italy), the mountainous climates of south-eastern Spain and Mo-
137 rocco, the wetter climate of the Atlantic region (from the Spanish Iberian region to the western
138 part of France) and the continental climate of central Spain. Maritime pine can also grow on a wide
139 range of substrates, from sandy and acidic soils to more calcareous ones. Maritime pine presents a
140 strong population genetic structure with occasional admixture, suggesting gene flow among gene
141 pools. Six gene pools have been described by previous literature, located in the French Atlantic
142 region, Iberian Atlantic region, central Spain, south-eastern Spain, Corsica and Northern Africa
143 (Fig. 1; Alberto et al. 2013, Jaramillo-Correa et al. 2015). These gene pools probably result from
144 the expansion of different glacial refugia (Bucci et al. 2007, Santos-del-Blanco et al. 2012).

145 2.2 Phenotypic data

146 Phenotypic data was obtained from three clonal common gardens (Table S1 and Fig. 1), planted
147 in 2011 and located in environments considered favorable to maritime pine, as evidenced by the
148 high survival rate at these sites (Table S1). The common gardens of Asturias (Spain, Iberian
149 Atlantic region) and Bordeaux (France, French Atlantic region) have very similar climates, with
150 mild winters, no severe cold events, high annual rainfall and relatively wet summers (Tables S3-S5
151 and Fig. 1). The common garden of Portugal (planted in Fundão) shows slightly colder winters
152 and lower summer precipitation than in Asturias (Table S4 and Fig. 1). In each of these common
153 gardens, trees belonging to 522 clones (i.e. genotypes) from 33 populations, including the six
154 known gene pools in the species, were planted following a randomized complete block design with
155 8 blocks, 8 trees per clone and from 2 to 28 clones per population (with an average of 15). To
156 obtain the clones, trees at least 50 m apart were sampled in natural stands, and one seed per tree
157 was planted in a nursery and vegetatively propagated by cuttings (see Rodríguez-Quilón et al.
158 2016 for details). Clones were therefore considered unrelated.

159 One growth trait, height, was measured in all common gardens and at different tree ages (Table
160 S1). Two phenology-related traits, the mean bud burst date over four years and the mean duration
161 of bud burst over three years, were measured in Bordeaux and were averaged over several years
162 to suppress differences across years and approximate a normal distribution of their trait values
163 (Table S1). Bud burst corresponds to the date of brachyblast emergence in accumulated degree-
164 days (with base temperature 0°C) from the first day of the year to account for between-year
165 variability in temperature. The duration of bud burst corresponds to the number of degree-days
166 between the beginning of bud elongation and the total elongation of the needles (see Hurel et al.
167 2019). Last, two functional traits, $\delta^{13}\text{C}$ and the specific leaf area (SLA) were measured in Portugal
168 (Table S1). These traits were selected because they showed broad-sense heritabilities that were
169 mostly low but with credibility intervals not crossing zero (> 0.08 in de Miguel et al. 2020). For
170 each trait, phenotypic means and variances across populations are shown in section 1.1 of the
171 Supplementary Information. Prior to analyses, some traits were log-transformed to get closer to
172 normality or mean-centered to help model convergence (Table S1).

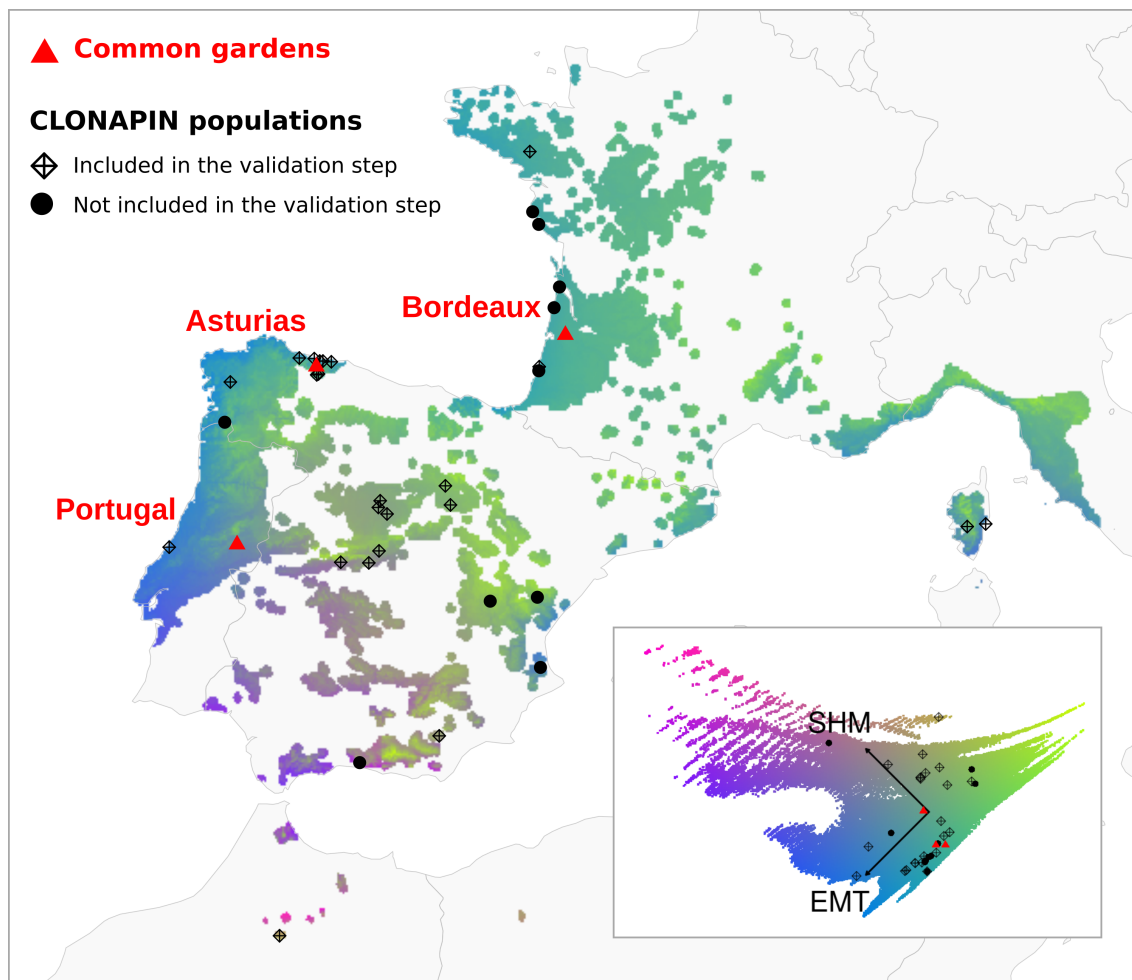


Figure 1. Location of the three common gardens and the 33 populations used in the study. The colors represent the gradients of the extreme minimum temperature (EMT) and summer heat moisture index (SHM) over the period 1901-1950 within the maritime pine range. The climatic gradients were obtained by performing a centered and scaled principal component analysis (shown in the inset on the bottom right) based on EMT and SHM values. The maritime pine distribution combines the EUFORGEN distribution (<http://www.euforgen.org/>) and 10-km radius areas around the National Forest Inventory plots with maritime pines. However, this remains a rough approximation of the actual distribution of maritime pine and therefore probably includes areas experiencing more intense cold or drought episodes than the climatic range of maritime pine.

173 2.3 SNP genotyping and population admixture

174 The 522 clones planted in the Asturias common garden were genotyped with the Illumina Infinium
175 assay described in Plomion et al. (2016), resulting in 5,165 high-quality polymorphic SNPs. There
176 were on average only 3.3 missing values per genotype (ranging between 0 and 142). For each clone,
177 the proportions of ancestry from each of the six known gene pools were estimated in Jaramillo-
178 Correa et al. (2015) using the Bayesian approach available in Structure (Pritchard et al. 2000), and
179 were then averaged by population. Populations were assigned to the gene pool that contributed
180 more than 50% ancestry and the other gene pools were considered as ‘foreign’ gene pools. First,
181 we calculated a population admixture score A, as the proportion of ancestry from foreign gene
182 pools (Table S6). Second, we calculated a population admixture score D that considers both the
183 proportion of foreign ancestries and the divergence between the main and foreign gene pools (Table
184 S6). For that, we weighted the proportions of ancestry from foreign gene pools by the sum of the

185 allele frequency divergence of the main and foreign gene pool from the common ancestral one (F_k ,
186 which should be numerically similar to F_{ST} ; Falush et al. 2003). We developed D considering
187 that some gene pools are more divergent than others and thus may bring higher genetic diversity
188 to an admixed population at the same level of introgression. A and D were highly correlated
189 (Pearson correlation coefficient of 0.91; Table S8). We also calculated a score D_{fst} by weighting the
190 proportions of ancestry from foreign gene pools by the pairwise F_{ST} between the main and foreign
191 gene pools (Table S8). D_{fst} was highly correlated to A and D (Pearson correlation coefficients of
192 0.91 and 0.96, respectively) and we therefore did not keep it in the following analyses.

193 **2.4 Population-specific environmental heterogeneity and climate** 194 **harshness indexes**

195 To describe the climate under which the populations have evolved, we used the climatic variables
196 at 1-km resolution and averaged over the period 1901-1950 from the *ClimateEU* database (Marchi
197 et al. 2020). Topographic data were generated from NASA's Shuttle Radar Topography Mission
198 (SRTM) at 90-m resolution and then aggregated at 1-km resolution. We used the SAGA v 2.3.1
199 (Conrad et al. 2015) to calculate the topographic ruggedness index (TRI) which quantifies the
200 terrain heterogeneity, i.e. differences in elevation between adjacent cells (Riley et al. 1999). Soil
201 variables were extracted from the European Soil Database at 1-km resolution (Hiederer et al. 2013).
202 All environmental variables used are listed in Table S7 and were mean-centered and divided by
203 their standard deviation prior to analyses.

204 To calculate the environmental heterogeneity around each population location, we extracted raster
205 cell values of the climatic, topographic and soil variables within a 20-km radius around each
206 population location, and kept only raster cells that fell within forested areas, to avoid including
207 environmental data from non-suitable areas (e.g. lakes, mountain peaks; section 1.3.2 of the
208 Supplementary Information). We then performed a principal component analysis (PCA) on the
209 raster cell values and extracted the PC1 and PC2 scores of each cell, accounting for 45.2% and
210 34.1% of the variance, respectively (Fig. S10). To obtain the four indexes of environmental
211 heterogeneity, we calculated the variances of the PC1 and PC2 scores in a 20-km and 1.6-km
212 radius around each population location. The environmental heterogeneity indexes were only very
213 weakly correlated (Pearson correlation coefficients lower than 0.36) with the number of forested
214 cells (i.e. the area considered to calculate the indexes), ensuring that the estimated effects of
215 environmental heterogeneity in further analyses were not due to the area per se (Triantis et al.
216 2003, Stein et al. 2014).

217 To describe the climate harshness at each population location, we used a drought index (the
218 summer heat moisture index averaged over the period 1901-1950, SHM, Table S7) and an index
219 related to severe cold events (the inverse of the extreme minimum temperature during the period
220 1901-1950, invEMT, Table S7). These two indexes were selected as maritime pine shows local
221 adaptation patterns associated with cold tolerance (Grivet et al. 2011) and because detecting
222 changes in the within-population genetic variation along a drought gradient would be key to

223 anticipate tree population responses to ongoing climate change.

224 2.5 Bayesian statistical modeling

225 We modeled the eight phenotypic traits with the same Bayesian statistical model, in which we
226 estimate the linear relationship between the within-population genetic variance and each of the
227 potential drivers successively (i.e. one model per driver): the two admixture scores, the four
228 environmental heterogeneity indexes and the two climate harshness indexes. Each trait y followed
229 a normal distribution (Fig. S1), such as:

$$\begin{aligned} y_{bpcr} &= \mathcal{N}(\mu_{bpc}, \sigma_r^2) \\ \mu_{bpc} &= \beta_0 + B_b + P_p + C_{c(p)} \end{aligned} \quad (1)$$

230 where σ_r^2 is the residual variance, β_0 the global intercept, and B_b , P_p and $C_{c(p)}$ are the block, pop-
231 ulation and clone (nested within population) varying intercepts, which are drawn from a common
232 distribution, such as:

$$\begin{aligned} \begin{bmatrix} B_b \\ P_p \end{bmatrix} &\sim \mathcal{N}\left(0, \begin{bmatrix} \sigma_B^2 \\ \sigma_P^2 \end{bmatrix}\right) \\ C_{c(p)} &\sim \mathcal{N}(0, \sigma_{C_p}^2) \end{aligned} \quad (2)$$

233 where σ_B^2 and σ_P^2 are the variance among blocks and populations and $\sigma_{C_p}^2$ are the population-
234 specific variances among clones (i.e. the within-population genetic variation). To estimate the
235 association between $\sigma_{C_p}^2$ and its potential underlying drivers, we expressed σ_{C_p} as follows:

$$\sigma_{C_p} \sim \mathcal{LN}\left(\ln(\overline{\sigma_{C_p}}) - \frac{\sigma_K^2}{2} + \beta_X X_p, \sigma_K^2\right) \quad (3)$$

236 where $\overline{\sigma_{C_p}}$ is the mean of the population-specific standard deviation among clones σ_{C_p} and X_p is
237 the potential driver considered (see section 2 in the Supplementary Information for more details).

238 To test the accuracy of the model estimates for σ_K^2 and β_X , we simulated data based on two traits
239 (height in Portugal and Bordeaux at 20 and 25-month old, respectively). For each trait, we ran
240 100 simulations and extracted the mean standard error and bias error of the estimates and the
241 coverage of the 80% and 95% credible intervals.

242 Model specification and fit were performed using the Stan probabilistic programming language
243 (Carpenter et al. 2017), based on the no-U-turn sampler algorithm. Models were run with four
244 chains and between 2,500 iterations per chain depending on the models (including 1,250 warm-up

245 samples not used for the inference). All analyses were undertaken in R version 3.6.3 (R Core Team
246 2020) and scripts are available at <https://github.com/JulietteArchambeau/H2Pinpin>.

247 **2.6 Validation step on independent data**

248 To validate our results for height, we used an independent dataset provided by Ricardo Alia in
249 which 23 populations shared with the CLONAPIN network were planted in a progeny test near
250 Asturias (thus in a similar environment). As the progeny test is based on families, we were able to
251 estimate the additive genetic variance within populations. We applied the same model as in our
252 study (replacing clones by families) to height measurements when the trees were 3 and 6-year old
253 (see section 7 of the Supplementary Information for more details).

254 **3 Results**

255 In the data simulation, σ_K^2 (the standard deviation of the logarithm of the within-population
256 genetic variation) and β_X (the coefficient of the potential drivers of the within-population genetic
257 variation) were properly estimated by the models (Table S9 and S10). Across 100 simulations, the
258 mean standard error was around 0.066 for σ_K^2 and 0.054 for β_X , the mean bias error was around
259 0.018 for σ_K^2 and -0.004 for β_X , the coverage of the 80% credible interval was around 93% for σ_K^2
260 and 80% for β_X , and the coverage of the 95% credible interval was around 98% for σ_K^2 and 96%
261 for β_X (Table S9 and S10). These simulations therefore showed that, under the assumption that
262 the statistical model reflects the processes at work, our model displayed a satisfactory accuracy to
263 be used in the following analyses.

264 The proportion of variance explained by the models (i.e. the sum of the among-population, among-
265 clone and among-block variances) and the variance partitioning varied broadly across traits (Fig.
266 S11 and section 5.3 in the Supplementary Information). More specifically, the models explained
267 between 40% and 50% of the variance for phenology-related traits, between 30% and 40% for
268 functional traits, and from 20% for height in Portugal to almost 60% for height in Bordeaux at
269 85-month old (Fig. S11). Residual variance explained most of the variance for all traits, except
270 for height in Bordeaux at 85-month old, where 40% of the variance came from variation among
271 populations, 40% from residuals and the remaining 20% from variation among clones (Fig. S18).
272 Variation among populations was higher than variation among clones for height and $\delta_{13}\text{C}$ (Figs.
273 S14, S16, S18, S20 and S28), but not for SLA and phenology-related traits (Figs. S26, S22 and
274 S24).

275 Environmental heterogeneity indexes and population admixture scores were not associated with
276 within-population genetic variation for any trait (Figs. 2 and S12). In contrast, we found a
277 consistent negative association with the inverse of the extreme minimum temperature across the
278 three common gardens for height, indicating that populations undergoing severe cold events display

279 less genetic variation (Fig. 2). Interestingly, in the Bordeaux common garden, this negative
 280 relationship was found at 25-month old, but not at 85-month old (Fig. 2). A negative association
 281 with the summer heat moisture index was also detected for height in Asturias, and less markedly
 282 but still with a high probability in Bordeaux at 25-month old (Fig. 2). Holding all other parameters
 283 constant, a one-standard deviation increase in the inverse of the extreme minimum temperature
 284 was associated, on average, with a 32.6%, 21.6% and 17.9% decrease of σ_{C_p} for height in Portugal,
 285 Bordeaux at 25-month old and Asturias, respectively. Similarly, a one-standard deviation increase
 286 in the summer heat moisture index was associated, on average, with 15.6% and 23.8% decrease of
 287 σ_{C_p} for height in Bordeaux at 25-month old and Asturias, respectively (see details of the calculation
 288 in the section 4 of the Supplementary Information). Unexpectedly, populations experiencing severe
 289 cold events showed higher genetic variation for SLA (Fig. 2). Within-population genetic variation
 290 was not correlated with the number of clones per population for any trait (maximum Pearson
 291 correlation coefficient = 0.57; Table S11).

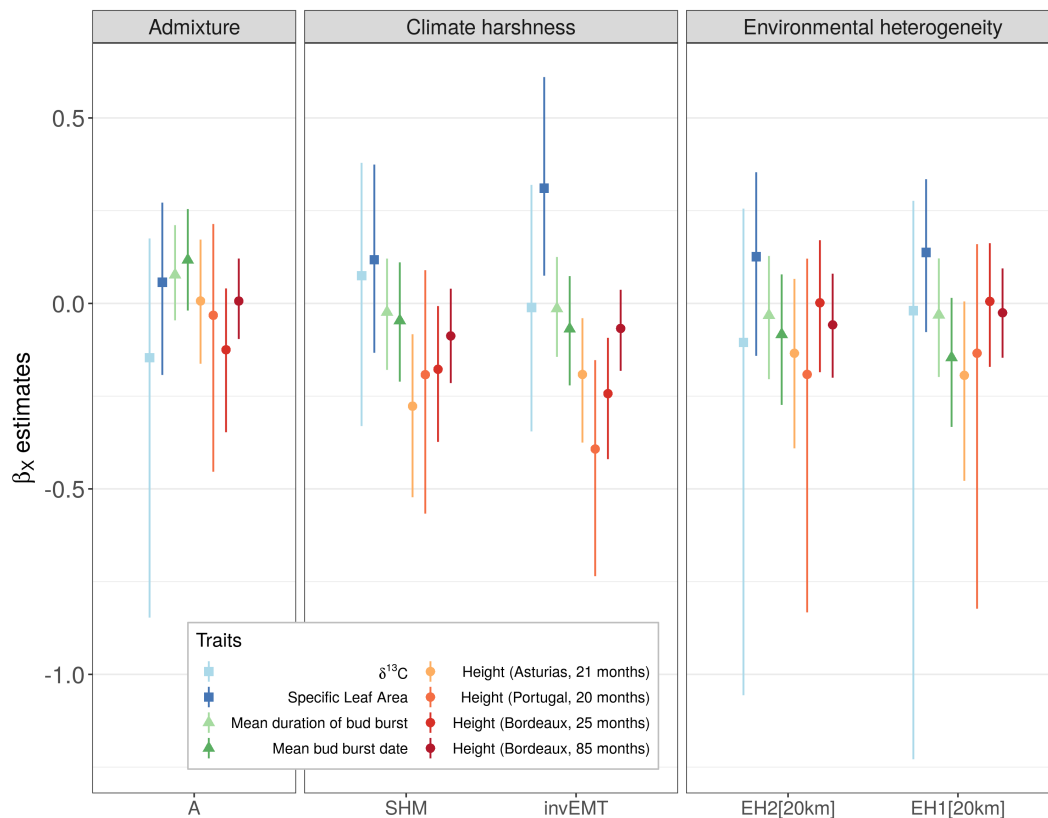


Figure 2. Median and 95% credible intervals of the β_X posterior distributions. β_X coefficients stand for the association between the within-population genetic variation and its potential underlying drivers on the x-axis: the inverse of the extreme minimum temperature during the studied period (invEMT), the summer heat moisture index (SHM), an admixture score (A), the environmental heterogeneity in a 20-km radius around the population location (EH1[20km] and EH2[20km]) calculated based on the projection of the PC1 and PC2 scores. Colors stand for the different traits under study and the shapes for the different types of traits, i.e. functional traits (squares), phenology-related traits (triangles) and height (circles).

292 Importantly, in the validation analysis, we also found a negative association between the inverse of
 293 the extreme minimum temperature and the within-population additive genetic variation for height
 294 at 3-year old, but not at 6-year old, and we did not find any association with the other potential
 295 drivers (Fig. 3).

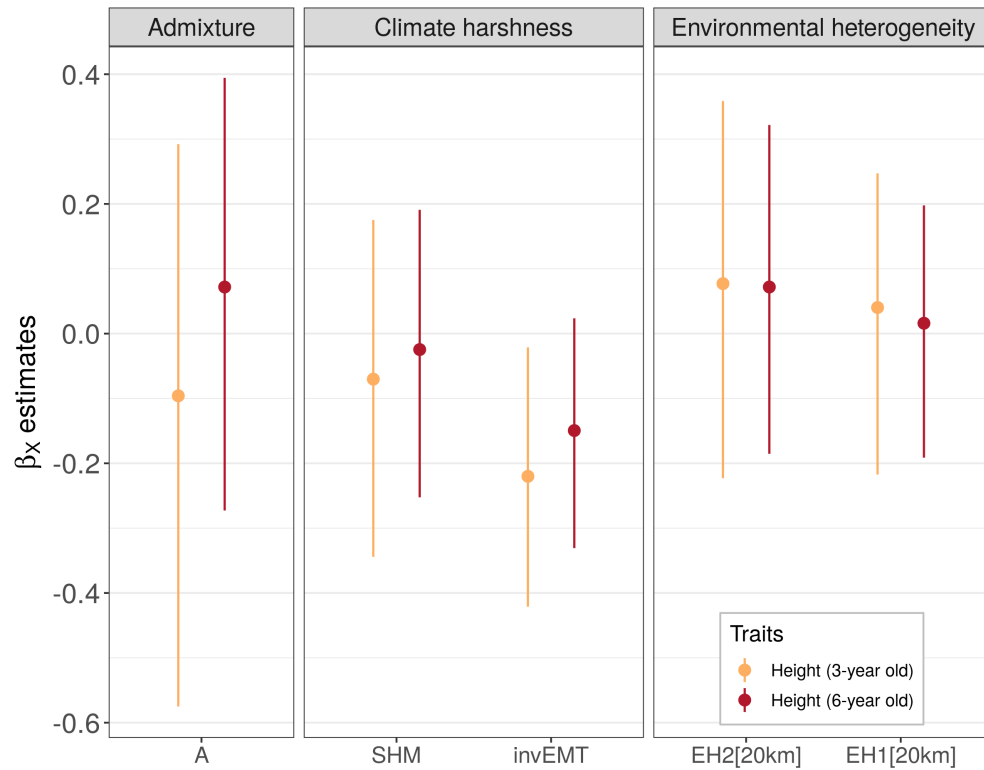


Figure 3. Validation step using independent height measurements from a common garden near Asturias. Median and 95% credible intervals of the β_X posterior distributions are shown. In the validation analysis, β_X coefficients stand for the association between the within-population additive genetic variation and its potential underlying drivers on the x-axis. A description of the drivers can be found in the legend of Fig. 2.

296 4 Discussion

297 How quantitative genetic variation is maintained within populations remains a long-standing open
298 question that has been extensively explored in theoretical work but lacks empirical evidence to
299 date (Johnson and Barton 2005). Our study suggests that genetic variation for height in maritime
300 pine is lower in populations exposed to severe cold events, thus supporting the hypothesis that
301 quantitative genetic variation in fitness-related traits is lower in populations under strong selection
302 (Fisher 1930). Across all traits studied, we did not find higher genetic variation in populations
303 located in heterogeneous landscapes, which goes against the predictions of some theoretical models
304 (McDonald and Yeaman 2018, Walsh and Lynch 2018) and an empirical study in lodgepole pine
305 (Yeaman and Jarvis 2006). Admixed populations did not show higher genetic variation, suggesting
306 that the observed patterns are not confounded by gene flow between distinct gene pools increasing
307 genetic variation. Empirically-based detection of the footprints of natural selection on within-
308 population genetic variation is much needed to understand how populations are adapted to their
309 current environments and will evolve under changing conditions.

310 4.1 Severe cold events may decrease within-population genetic varia- 311 tion

312 Height genetic variation was lower in populations experiencing harsher climates, especially severe
313 cold events (invEMT; Fig. 2). This result supports the hypothesis that strong stabilizing selection
314 in harsh environments depletes quantitative genetic variation within populations (Fisher 1930) and
315 echoes similar results in another forest tree, *Quercus oleoides*. For this Mesoamerican white oak
316 species, Ramírez-Valiente et al. (2019) found lower genetic variation averaged over functional and
317 growth traits in populations experiencing low precipitation and high temperatures during the dry
318 season. The importance of severe cold events as a driver of height genetic variation in maritime
319 pine is supported by the association between candidate-gene allele frequency and temperature
320 gradients (Grivet et al. 2011, Jaramillo-Correa et al. 2015), suggesting a major role of minimum
321 temperatures in the species adaptive evolution. Indeed, lower genetic variation in areas subject
322 to cold events may enhance adaptation to local conditions, but it may also hamper the adaptive
323 potential of populations under new climates. Noticeably, severe cold events were highly correlated
324 with altitude in our study (Pearson’s correlation of 0.9), and adaptation patterns along altitudinal
325 gradients are common in forest trees (e.g. Kurt et al. 2012). Therefore, we cannot exclude that the
326 association between height genetic variation and severe cold events is triggered by more complex
327 environmental factors typical of high altitude conditions (e.g. reduced vapor pressure deficit,
328 higher maximum solar radiation; Körner 1995). Estimating selection strength directly in natural
329 populations, as in Bontemps et al. (2016), would be highly valuable, albeit challenging in forest
330 trees.

331 Within-population genetic variation was unlikely to be influenced by demographic history and
332 gene flow across gene pools, as we did not find any association between within-population genetic
333 variation for height (and other traits) and population admixture indexes (Fig. 2). Another po-
334 tential scenario that may explain our results is related to the joint effect of environment on the
335 short-term expression of quantitative genetic variation and the strength of natural selection in the
336 novel environments of the common gardens (Hoffmann and Merilä 1999, Wood and Brodie 2016;
337 see examples for natural populations in Wilson et al. 2006 for wild sheep and Husby et al. 2011
338 for great tits). Noticeably, the expression of hidden genetic variation in novel environments (i.e.
339 ‘cryptic genetic variation’; Schlichting 2008) may be as large as the genetic variation resulting from
340 the long-term divergent evolution of populations (Wood and Brodie 2015). However, in our study,
341 this scenario is unlikely as the negative associations between height genetic variation and climate
342 harshness were consistent across the three common gardens (i.e. across distinct environmental
343 conditions; partially reflected in Fig. 1; see also Tables S3-S5), which suggests that the lower
344 height genetic variation in populations from harsher climates is independent from the environmen-
345 tal conditions in the common gardens and thus likely to be intrinsic to the populations. Last,
346 the sampled populations may not fully cover the climatic range of maritime pine (Fig. 1), which
347 reduces our ability to detect an association between some climatic drivers and within-population
348 genetic variation, and therefore may explain why no association was detected for the summer heat
349 moisture index (SHM).

350 Most importantly, the validation analysis provided independent evidence that additive within-
351 population genetic variation for height was lower in populations experiencing extreme cold events
352 for young trees but not for older trees (Fig. 3). This supports the robustness of our study and
353 suggests that our results were unlikely to be biased by considering the total variance instead of
354 the additive one, which was somehow expected as two previous studies in maritime pine found low
355 non-additive effects for growth (Gaspar et al. 2013), and height and diameter (Lepoittevin et al.
356 2011).

357 With respect to specific leaf area (SLA), where only a single common garden (i.e. a single environ-
358 ment) was assessed, cryptic genetic variation (as defined above) may indeed underlie the higher
359 genetic variation found in populations experiencing severe cold events. A study in maritime pine
360 suggests that SLA depends strongly on environmental conditions (Alía et al. 2014), which is sup-
361 ported in our study by its weak genetic variation in the Portugal common garden (less than 10%
362 of the phenotypic variance is explained by the population or clone effects), with a large part of the
363 variance associated with the block effect (Fig. S26). Cryptic genetic variation is more likely to be
364 expressed when the differences between original and current environments are large (Paaby and
365 Rockman 2014). Some populations experiencing severe cold events (and high altitude conditions)
366 may therefore have reached the threshold inducing a release of cryptic variation in the Portugal
367 common garden. However, this is not a general pattern as we did not find any association between
368 the climatic transfer distances (i.e. the absolute difference between the climate in the population
369 and the climate in the test site) and the within-population genetic variation for SLA (see section
370 6 of the Supplementary Information). Replicating SLA measurements in common gardens at high
371 altitude or experiencing extreme cold episodes would be highly valuable to test this hypothesis.

372 **4.2 Environmental heterogeneity is not associated with higher genetic** 373 **variation**

374 Populations from heterogeneous environments did not show higher genetic variation for any trait
375 (Fig. 2), which was also the case for the independent height data from the validation analysis
376 (Fig. 3). This goes against a previous estimate in lodgepole pine suggesting that up to 20% of
377 the genetic variation in growth within populations is explained by environmental heterogeneity
378 (Yeaman and Jarvis 2006). A potential explanation of this discrepancy is the smaller experiment
379 size in our study compared to that of Yeaman and Jarvis (103 populations with an average of 28
380 planting sites per population). However, in our study, we obtained reasonable credible intervals
381 for most traits (allowing the detection of associations with other drivers) and data simulations
382 suggested that our models have adequate power, rendering this explanation unlikely.

383 Another explanation is that genetic variation within populations is not affected by the environ-
384 mental heterogeneity at the regional scale imposed by the 1×1 km resolution of our climate
385 dataset but at finer spatial scales (also discussed in Yeaman and Jarvis 2006). Indeed, popula-
386 tions can adapt along microgeographic environmental gradients despite the homogenizing effect
387 of gene flow (Richardson et al. 2014), even for forest tree populations with their long-generation

388 times and large effective population sizes (Scotti et al. 2016). However, a correlation between
389 regional and microgeographic environmental heterogeneity across the maritime pine range is very
390 likely: populations showing the highest environmental heterogeneity in our study were located in
391 mountainous areas in which we also expect higher microgeographic variation, e.g. the C6mpeta
392 population (COM) located in the Tejada and Almjara mountains (southern Spain), the Arenas
393 de San Pedro population (ARN) located in the Sierra de Gredos (central Spain) or the Pineta
394 population (PIE) located close to the Punta di Forchelli (Corsican mountains), while populations
395 with the lowest environmental heterogeneity were located on flat plateaus, e.g. populations from
396 the Landes plateau and the Atlantic coastal regions in France (HOU, MIM, PET, VER, OLO,
397 STJ, PLE), and populations from the central Spain plateau near to Segovia (CUE, COC, CAR).
398 Thus, even if genetic variation was maintained by migration-selection balance at microgeographic
399 scales, we would have been able to detect the effect of environmental heterogeneity at the regional
400 scale. Nevertheless, more studies characterizing adaptation at microgeographic scales are needed
401 to assess the spatial scale of genetic adaptation in maritime pine.

402 Another explanation of the discrepancy with Yeaman and Jarvis (2006) could be that we used
403 young trees (between 20 and 85-month old) while they used 20-year old trees. Indeed, the pro-
404 cesses generating within-population genetic variation might be age-dependent, as shown for climate
405 harshness in Bordeaux, where the association was present when the trees were 25-month old but
406 not in older trees. In forest trees, genetic parameters often vary with age; e.g. heritability generally
407 increases with age until reaching a plateau, especially for height-related traits (Balocchi et al. 1993,
408 Johnson et al. 1997, Sierra-Lucero et al. 2002, Jansson et al. 2003, Kroon et al. 2011), but may also
409 decrease in some cases (Lu and Charrette 2008, Kroon et al. 2011). In maritime pine, an increase
410 in heritability with age was found in Costa and Durel (2011) but not in Kusnandar et al. (1998).
411 To our knowledge, the drivers of heritability changes with age remain unclear. Competition among
412 trees in common gardens might play a role in the expression of age-dependent heritabilities for
413 diameter growth, but not for height in *Pinus radiata* (Lin et al. 2013). Replicating our analysis in
414 older trees would be interesting to further assess patterns of association between within-population
415 genetic variation and environmental heterogeneity, and their underlying causes.

416 Finally, a last explanation is related to the different biological features between lodgepole pine and
417 maritime pine. Lodgepole pine has extensive gene flow and low population structure ($F_{ST} = 0.016$
418 in Yeaman et al. 2016) while maritime pine shows restricted gene flow with strong population struc-
419 ture (at least six distinct gene pools and $F_{ST} = 0.112$; Jaramillo-Correa et al. 2015; our study) and
420 fragmented distribution (Alberto et al. 2013). Pollen dispersal kernels in maritime pine are highly
421 leptokurtic, as for other wind-pollinated pines (Schuster and Mitton 2000, Robledo-Arnuncio and
422 Gil 2005), with estimated mean dispersal distances from 78.4 to 174.4m (de-Lucas et al. 2008).
423 Interestingly, McDonald and Yeaman (2018) showed that high levels of quantitative genetic vari-
424 ance can be maintained when a trait is under stabilizing selection only at intermediate levels of
425 migration. Migration rates in maritime pine may therefore not be strong enough to compensate
426 for the purifying effect of natural selection in heterogeneous environments, especially in moun-
427 tainous areas which may represent barriers to gene flow and where populations are more isolated
428 (see Gonz6lez-Mart6nez et al. 2007 for maritime pine). Meanwhile, in the homogeneous plateaus

429 of the Landes forest and central Spain, natural selection may be low because conditions are more
430 favorable, and these populations are less isolated, which may maintain genetic variation at lev-
431 els similar to those of populations in heterogeneous landscapes. Investigating local adaptation
432 and gene flow at microgeographic scales in natural populations of maritime pine located in both
433 homogeneous and heterogeneous environments would be highly valuable to understand why en-
434 vironmental heterogeneity does not seem to play a major role in maintaining genetic variation
435 in this species. Moreover, conducting similar analyses in sister species such as Scots pine, with
436 low population genetic structure and continuous populations (Alberto et al. 2013), could help to
437 determine whether genetic variation in forest tree populations experiencing higher migration rates
438 are more prone to be impacted by environmental heterogeneity.

439 **4.3 Link to fitness and genetic constraints may explain the different** 440 **patterns across traits**

441 Height was the only trait that showed a consistent association between within-population genetic
442 variation and climate harshness. This pattern supports the hypothesis that natural selection
443 mainly depletes genetic variation of traits most directly related to fitness. Indeed, height can be
444 seen as the end-product of multiple ecophysiological processes (Grattapaglia et al. 2009). Taller
445 trees perform better in the competition for light, water and nutrients, and are therefore more likely
446 to have higher fecundity (Rehfeldt et al. 1999, Wu and Ying 2004, Aitken and Bemmels 2015) and
447 lower mortality (Wyckoff and Clark 2002, Zhu et al. 2017). However, taller trees are also more
448 susceptible to spring and fall cold injury (Howe et al. 2003) and to drought (Bennett et al. 2015,
449 McDowell and Allen 2015, Stovall et al. 2019). In maritime pine, effective reproductive success (i.e.
450 the number of successfully established offspring) is related to tree size. Indeed, González-Martínez
451 et al. (2006) found a significant positive female selection gradient for diameter (height was not
452 tested, but diameter and height are strongly correlated in conifers; see, for example, Fig. 1 in
453 Castedo-Dorado et al. 2005 for maritime pine) and suggested that offspring mothered by bigger
454 trees could have a selective advantage due to better quality seeds favouring resilience in the face
455 of severe summer droughts and microsite variation. This evidence also supports the idea of height
456 as a relevant fitness component in maritime pine.

457 Although less directly related to fitness than height, leaf phenology-related traits exhibit steep
458 adaptation gradients in forest trees and have a relatively high heritability, e.g. 0.15-0.51 for bud
459 burst in pedunculate oak (Scotti-Saintagne et al. 2004), 0.45-1 in Sitka spruce (Alfaro et al. 2000)
460 and 0.54 for bud burst and 0.30 for the duration of bud burst in our study in maritime pine.
461 Gauzere et al. (2020) showed that both the mean and the variance of leaf phenology-related traits
462 varied along an altitudinal gradient in natural oak populations, with populations at high altitude
463 having a narrower fitness peak. We might therefore have expected lower genetic variation for leaf
464 phenology-related traits in populations experiencing severe cold events (and at higher altitude), as
465 found along an altitudinal gradient in sessile oak for bud phenology (Alberto et al. 2011). However,
466 such association may be hidden in common gardens with different climates from those of the

467 populations' location, because of the release of high levels of cryptic genetic variation (Schlichting
468 2008). Moreover, phenology-related traits can show opposite genetic clines in common gardens and
469 natural populations (e.g. Vitasse et al. 2009). Estimating genetic parameters of phenology-related
470 traits directly in the field, which is now technically possible by using large genomic datasets and
471 advanced statistical methodologies (Gienapp et al. 2017), may therefore be necessary to investigate
472 potential associations between within-population genetic variation and climate harshness, or other
473 selective pressures.

474 Importantly, theoretical work suggests that much of the genetic variation associated with a trait
475 is likely maintained by pleiotropic effects, which are independent of the selection on that trait,
476 implying that stabilizing selection can only act on a reduced number of independent dimensions in
477 the trait space (Barton 1990, Walsh and Lynch 2018). As we used univariate models, we cannot
478 exclude that the likely associations with height genetic variation originate from genetic correlation
479 with other traits under selection, or that the lack of association with other traits (notably functional
480 traits such as $\delta^{13}\text{C}$) does not originate from genetic constraints (Walsh and Blows 2009). For
481 example, in maritime pine, trait canalisation and genetic constraints may explain low quantitative
482 genetic differentiation for hydraulic traits (e.g. P50, the xylem pressure inducing 50% loss of
483 hydraulic conductance; Lamy et al. 2014), and sapling height was found to be either positively or
484 negatively associated with disease susceptibility depending on the pathogen (e.g. necrosis length
485 caused by *Diplodia sapinea* or *Armillaria ostoyae*, respectively; Hurel et al. 2019). Trade-offs
486 between traits may also explain the unexpected association between minimum temperatures and
487 high genetic variation for SLA, as, for instance, SLA is known to be positively correlated with
488 leaf life span, low assimilation rates and nutrient retention, i.e. traits linked to conservation of
489 acquired resources (Ackerly et al. 2002).

490 5 Conclusion

491 Our manuscript contributes to the current debate on the maintenance of quantitative genetic
492 variation within populations by providing empirical support for the role of natural selection in
493 decreasing genetic variation. Indeed, our results consistently showed that genetic variation for
494 height is lower in maritime pine populations experiencing severe cold events (i.e. experiencing
495 stronger selection). Surprisingly, we found no association between environmental heterogeneity at
496 the regional scale and within-population genetic variation for several traits; whether for technical
497 reasons (e.g. sample size, spatial scale considered) or for genuine biological reasons (e.g. too
498 low migration), it would be worth further exploration. Indeed, understanding the evolutionary
499 forces shaping within-population genetic variation could shed light on how populations adapt to
500 their local environment, thereby providing insight into how they may respond to future changes in
501 environmental conditions.

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