The effect of microbial selection on the occurrence-abundance patterns of microbiomes

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Theoretical models are useful to investigate the drivers of community dynamics. Notable are models that consider the events of death, birth, and immigration of individuals assuming they only depend on their abundance – thus, all types share the same parameters. The community level expectations arising from these simple models and their agreement to empirical data have been discussed extensively, often suggesting that in nature, rates might indeed be neutral or their differences not important. But, how robust are these model predictions to type-specific rates? And, what are the consequences at the level of types? Here, we address these questions moving from simple to diverse communities. For this, we build a model where types are differently adapted to the environment. We adapt a computational method from the literature to compute equilibrium distributions of the abundance. Then, we look into the occurrence-abundance pattern often reported in microbial communities. We observe that large immigration and biodiversity – common in microbial systems – lead to such patterns, regardless of whether the rates are neutral or non-neutral. We conclude by discussing the implications to interpret and test empirical data.

Keywords: ecology, neutral theory, niche theory, migration, biodiversity, mathematical model.

23 1 Introduction

24 Theoretical models have been instrumental in understanding ecological systems. Historically, a
 25 handful of puzzling natural observations have motivated their development – from the limits of
 26 exponential growth by Malthus [1] to the competition of species by Lotka and Volterra [2, 3].

27 The stark difference of the frequencies of species within communities is one such observation.
 28 While few species are very abundant, many others barely appear in community surveys [4]. Two
 29 hypotheses have dominated the scientific discussions. On one hand, it is proposed that biotic
 30 interactions and environmental filtering make trophically similar species occupy different niches,
 31 which allows differences in abundance while preserving diversity. This is known as niche theory
 32 [5]. Alternatively, Hubbell and others [6] have emphasized that even if niche differences are
 33 discounted, so only species’ abundances matter, random fluctuations can lead to the patterns
 34 of abundance and diversity observed in nature. This is known as neutral ecological theory [7].

35 Despite their stringent assumptions, neutral models often predict patterns observed in com-
 36 munities as different as the tropical rainforest of Barro Colorado island [7] and host-associated
 37 microbiomes [8, 9, 10]. With time, neutral models have become null hypotheses used to discard
 38 the need for complex mechanistic explanations in data at the community level [6].

39 But how does a neutral model work? In a neutral model the death and birth of individuals
 40 account for changes in community composition. However, because each rate is identical for
 41 all types, after some time, stochastic drift leads to the extinction of all but one type [11].
 42 Thus, to preserve diversity, an external source of individuals by immigration or speciation is
 43 needed. Here, neutral theory builds upon island biogeography. In this theory, MacArthur and
 44 Wilson [12] have modelled the community composition of small habitats (“islands”) connected
 45 by migration to a larger habitat (“mainland”). In neutral models, a local community commonly
 46 receives individuals from an external and larger community [13]. Such community can itself
 47 undergo internal changes or, by separation of time scales, assumed to be constant [7, 13].

48 Early on, neutral models have been used in macroecology to address the patterns of diversity
 49 and abundance of species [7, 6]. More recently, driven by developments in sequencing tech-
 50 nologies, the study of patterns of occurrence and mean frequency in microbial communities has
 51 become possible [14]. At this scale, ecological drift also seems to greatly influence the commu-
 52 nity dynamics, leading to hypothesize that many microbial taxa could be classified as neutral

[15, 10]. However, few taxa, referred to as non-neutral, have occurrences and frequencies different than neutrally expected. It has been suggested that the last group might include, among others, pathogens and symbionts [10].

At least two possibilities could lead to deviations from neutrality. Either different processes from those in the neutral model are necessary, or, alternatively, not all the parameters of the model are actually neutral. Both of these lead to develop models of selection [11]. Although many such models have been developed from niche theory assumptions, fewer have been developed from a neutral theory basis [6, 16]. A direct connection from neutral to selective models would allow to comparing their patterns while acknowledging that both might be operating simultaneously. Indeed, the role of non-neutral processes can only be rejected after ensuring that they can not produce “neutral” patterns [16], especially in data.

Neutral and niche models have been connected in several ways [17, 18, 13]. Some authors have assumed that the rate of types are solely determined by the environment, finding that neutrality might overshadow the niche structure effect [19], depending on diversity, dispersal, and niche overlap [17]. Alternatively, using Lotka-Volterra models with immigration, the effect of competitive interactions has been studied. Early models focused on intraspecific [20] or interspecific [21] competition. Later on, both were considered simultaneously. Haegeman and Loreau tuned the niche overlap using symmetric interactions to investigate the success behind the neutral assumption [18]. Kessler and Shnerb classified the dynamics emerging from interspecific interactions, finding that the neutral case links all classes [13]. Focusing on intraspecific interactions, Gravel et al. studied the influence of immigration, suggesting a continuum from competitive to stochastic exclusion [17]. Throughout these studies, diversity, community size, and environmental fluctuations seem to have great relevance, as pointed out by Chisholm and Pacala and Fisher and Mehta.

This previous research has proven useful to bridge neutral and selective theories. The link has been instrumental to consider migration, speciation, and stochastic demography key components in ecology. Along this line and motivated by the particularities of microbial communities, large community size and taxa diversity [15, 23, 14], here we investigate the commonly observed occurrence-abundance pattern in neutral and non-neutral contexts. Similarly to Sloan et al. and Allouche and Kadmon, we model death, birth, and immigration within a community, but in contrast to these neutral models, type-specific growth and death rates are determined by the

84 environment.

85 2 Results

86 2.1 A spatially-implicit death-birth model with immigration

87 We consider a set of local communities connected by immigration to a larger community which
 88 contains multiple types of individuals. While local communities change as a result of the
 89 death, birth, and immigration of individuals, the larger community changes on a much longer
 90 time-scale – so immigration to local communities can be assumed to be constant. To derive a
 91 dynamical equation of a local community composition, we account for the events that change
 92 the frequency x_i of each type $i = 1, \dots, S$ within each local community. Individuals die with a
 93 rate proportional to the product $x_i\phi_i$ of their frequency and their death rate ϕ_i . Additionally,
 94 they are born proportional to the product $x_i f_i$ of their frequency and their growth rate f_i – or
 95 arrive with a fraction of the immigration rate m that reflects their frequency p_i in the external
 96 environment. Combining these processes, we obtain

$$\frac{dx_i}{dt} = f_i x_i - \phi_i x_i + m p_i. \quad (1)$$

97 Assume for now an equal death rate for all types, $\phi_i = \phi$, so only f_i , m , and p_i are free
 98 parameters. To hold the community size constant, we use $\sum_i dx_i/dt = 0$ to find $\phi = \bar{f} + m$,
 99 where $\bar{f} = \sum_j x_j f_j$ is the average growth rate of a randomly selected individual. In this way

$$\frac{dx_i}{dt} = x_i(f_i - \bar{f}) + m(p_i - x_i). \quad (2)$$

100 Without immigration, $m = 0$, Eq. (2) shows that only types whose growth rate is larger
 101 than the average increase. After sufficient time, only the type with the largest growth rate
 102 remains. Coexistence is only possible in the neutral case, where all types have the same growth
 103 rate, $f_i = \bar{f}$. There, the initial frequencies remain unchanged. Immigration, $m > 0$, creates an
 104 equilibrium that resembles the external composition, p_i , that for sufficiently large immigration
 105 might promote coexistence, especially if types with small growth rate migrate more. Similar
 106 results are obtained if we assume equal growth rate for all types, $f_i = f$ in Eq. (1) instead. In
 107 the general case, growth and death rates have opposing effects.

Eq. (1) provides useful insights about the dynamics and equilibria; however, only a stochastic model would allow us to compute observables such as the occurrence frequency and the variance. To develop such a model, we track the vector of absolute abundances instead, $\mathbf{n}$, and list the transition rates that change it. The increase of type i by one individual occurs at the expense of the decrease of type j ,

$$R(\mathbf{n} \rightarrow \mathbf{n} + \mathbf{e}_i - \mathbf{e}_j) = \phi_j \frac{n_j}{N} (f_i n_i + m p_i). \quad (3)$$

Here $\mathbf{e}_i$ and $\mathbf{e}_j$ are vectors whose i -th or j -th element equals one and zero elsewhere. The carrying capacity of the community is given by N . The master equation accounts for changes in the probability of observing the community composition $\mathbf{n}$ through time,

$$\begin{aligned} \frac{\partial P(\mathbf{n}, t)}{\partial t} = & - \underbrace{\sum_{\substack{i,j \\ i \neq j}} P(\mathbf{n}, t) (R(\mathbf{n} \rightarrow \mathbf{n} + \mathbf{e}_i - \mathbf{e}_j) + R(\mathbf{n} \rightarrow \mathbf{n} - \mathbf{e}_i + \mathbf{e}_j))}_{\text{Probability outflux}} \\ & + \underbrace{\sum_{\substack{i,j \\ i \neq j}} (P(\mathbf{n} + \mathbf{e}_i - \mathbf{e}_j, t) R(\mathbf{n} + \mathbf{e}_i - \mathbf{e}_j \rightarrow \mathbf{n}) + P(\mathbf{n} - \mathbf{e}_i + \mathbf{e}_j, t) R(\mathbf{n} - \mathbf{e}_i + \mathbf{e}_j \rightarrow \mathbf{n}))}_{\text{Probability influx}}, \end{aligned} \quad (4)$$

where $P(\mathbf{n}, t)$ is the probability density of community composition $\mathbf{n}$ at time t .

In this work, we investigate the probability distribution at equilibrium, i.e. the state where the master equation equals zero. In this case, the influx and outflux to each state balance each other, ending up with a system of equations that can be solved to find $P(\mathbf{n})$. For communities composed of two types ($S = 2$), a detailed balance analysis [24] leads to a recurrence equation of the arbitrarily denoted type 1

$$P(n_1) = P(0) \prod_{\mathbf{n}=(0,N)}^{(n_1-1,1)} \frac{R(\mathbf{n} \rightarrow \mathbf{n} + \mathbf{e}_1 - \mathbf{e}_2)}{R(\mathbf{n} + \mathbf{e}_1 - \mathbf{e}_2 \rightarrow \mathbf{n})}, \quad (5)$$

satisfying $\sum_0^N P(n_1) = 1$. In this case, the transition rates are simplified to the single variable n_1 , using $n_2 = N - n_1$ and $p_2 = 1 - p_1$.

For communities with more than two types ($S > 2$) analyses are more challenging, as all possible compositions must be considered. This is particularly true for microbial communities, where many types interact (10^1 to 10^4 taxa are common) in large communities (10^3 to

10¹⁴ individuals). Although a recurrence equation exists [21], the exponential increase in the number of states and transitions with S and N , make its computation unfeasible. This is a problem common to microscopic and even mesoscopic descriptions, which has been deemed “the curse of dimensionality” [25]. In neutral models, the equality of rates allows to reduce analyses to a single dimension – that of a focal type [15]. However, unless density dependence is neglected, non-neutral models are inherently multidimensional, as transitions depend on the current community composition.

A potential way forward is to acknowledge that, typically, rather than being interested in the probability of every possible community, we are interested in marginal probabilities. In other words, the added probabilities over various dimensions. Methods of model reduction have been developed towards this aim. Based on various assumptions, these methods sacrifice “microscopic” information in the interest of specific observables. Jahnke introduced the *model reduction by conditional expectations* (MRCE), where, while selected types are described stochastically, others are modeled using a mean-field approximation [26]. The MRCE is derived from the Bayes theorem, by which $P(\mathbf{n}, t)$ is given by the product of two probabilities, one for some chosen types and one for the conditional probability of the others. Then, the probabilities of the others are replaced by expected abundances. Because of the last point, the method is particularly suited to systems where types have peaked distributions and large populations – a situation that can be akin to some microbial communities.

In this paper we combine the MRCE method [26] with a detailed balance analysis [24] to compute the marginal probability distribution of types within a microbial community. For each distribution at equilibrium, we extract the probability of occurrence, $P(n_i \geq 1)$, the mean frequency $E(n_i)/N$, and compare them in situations of neutrality versus non-neutrality.

To apply the MRCE method, we adapt our model to the convention in [26]. First, we split the vector of abundances $\mathbf{n} \in \mathbb{Z}^S$ into a focal type i , n_i , and the set of others, $\mathbf{n}_j \in \mathbb{Z}^{S-1}, j \neq i$, for which the marginal probability, $\tilde{P}(n_i) \approx P(n_i)$, and the expected abundance conditioned on the focal type, $(\tilde{\mathbf{n}}_j|n_i) \approx \sum_{\mathbf{n}_j} \mathbf{n}_j P(\mathbf{n}_j|n_i)$, are approximated. Then, each transition rate is factored as the product of rates of the focal type and other types,

$$R(\mathbf{n} \rightarrow \mathbf{n} + \mathbf{e}_i - \mathbf{e}_j) = R_i^+(n_i) R_j^-(\tilde{\mathbf{n}}_j|n_i) \quad (6a)$$

$$R(\mathbf{n} \rightarrow \mathbf{n} - \mathbf{e}_i + \mathbf{e}_j) = R_i^-(n_i) R_j^+(\tilde{\mathbf{n}}_j|n_i). \quad (6b)$$

156 In our model, $R_i^+(n_i) = f_i n_i + m p_i$, $R_j^-(\tilde{n}_j|n_i) = \phi_j \frac{\tilde{n}_j}{N}$, $R_i^-(n_i) = \phi_i \frac{n_i}{N}$ and $R_j^+(\tilde{n}_j|n_i) =$
 157 $f_j \tilde{n}_j + m p_j$. With these transformations, the equilibrium is given by the simplified master
 158 equation of the focal type i ,

$$\begin{aligned} 0 = & -\tilde{P}(n_i) \left(R_i^-(n_i) \sum_{j \neq i}^S R_j^+(\tilde{n}_j|n_i) + R_i^+(n_i) \sum_{j \neq i}^S R_j^-(\tilde{n}_j|n_i) \right) \\ & + \tilde{P}(n_i + 1) R_i^-(n_i + 1) \sum_{j \neq i}^S R_j^+(\tilde{n}_j|n_i + 1) \\ & + \tilde{P}(n_i - 1) R_i^+(n_i - 1) \sum_{j \neq i}^S R_j^-(\tilde{n}_j|n_i - 1), \end{aligned} \quad (7a)$$

159 and a set of equations for the expected abundance of the others conditioned on the abundance
 160 of the focal type $(\tilde{\mathbf{n}}_j|n_i)$,

$$\begin{aligned} \mathbf{0} = & -(\tilde{\mathbf{n}}_j|n_i) \tilde{P}(n_i) \left(R_i^-(n_i) \sum_{j \neq i}^S R_j^+(\tilde{n}_j|n_i) + R_i^+(n_i) \sum_{j \neq i}^S R_j^-(\tilde{n}_j|n_i) \right) \\ & + \tilde{P}(n_i + 1) R_i^-(n_i + 1) \left(\sum_{j \neq i}^S R_j^+(\tilde{n}_j|n_i + 1) ((\tilde{\mathbf{n}}_j|n_i + 1) + \mathbf{e}_j) \right) \\ & + \tilde{P}(n_i - 1) R_i^+(n_i - 1) \left(\sum_{j \neq i}^S R_j^-(\tilde{n}_j|n_i - 1) ((\tilde{\mathbf{n}}_j|n_i - 1) - \mathbf{e}_j) \right). \end{aligned} \quad (7b)$$

161 We solve this system of equations in the range of $n_i = 0, \dots, N$, starting from $n_i = N$. By
 162 definition $\tilde{P}(N + 1) = 0$, so no probability flux to or from $N + 1$ occurs. Then, the influx from
 163 $n_i = N$ implies $R_j^-(\tilde{n}_j|N) = 0$, specifically $(\tilde{\mathbf{n}}_j|N) = \mathbf{0}$. We end up with a simplified system of
 164 equations for $n_i = N$. To compute $\tilde{P}(N - 1)$ and $(\tilde{\mathbf{n}}_j|N - 1)$ from this, we assume without loss
 165 of generality $\tilde{P}(N) = c_p$, where c_p is a positive constant. Consecutive $\tilde{P}(n_i - 1)$ and $(\tilde{\mathbf{n}}_j|n_i - 1)$
 166 are computed iteratively. Finally, the normalization $\sum_0^N \tilde{P}(n_i) = 1$ is enforced.

167 A reliable numerical method is needed to solve Eq. (7a-7b). The large difference between
 168 the magnitudes of $\tilde{P}(n_i - 1)$ and $(\tilde{\mathbf{n}}_j|n_i - 1)$ can cause numerical problems. To avoid them, we
 169 extract $\tilde{P}(n_i - 1)$ from Eq. (7a) and substitute it in Eq. (7b) – note that all else are known
 170 values. The resulting system of equations is solved for $(\tilde{\mathbf{n}}_j|n_i - 1)$, and these substituted in
 171 Eq. (7a) to compute $\tilde{P}(n_i - 1)$. Caution is needed in cases that lead to a normalized $\tilde{P}(N) \approx 0$,
 172 especially if computations are performed in a machine with limited float representation. In this
 173 case, we find the $n_i = n_i^*$ closest to $n_i = N$ that while declaring $P(n_i > n_i^*) = 0$ and $P(n_i^*) = c_p$
 174 allows for the iterative solution.

Figure 1: **Expected equilibrium of a type if rates in the community are neutral.** (A) If the immigration is very small, the population either goes extinct or reaches fixation. A larger immigration reduces the variation in frequency, centered at its fraction of immigrants (here $p_1 = 0.5$). (B) The mean frequency increases with the fraction of immigrants p_1 , but is independent of the immigration rate m . (C) Also the occurrence frequency increase with the fraction of immigrants (p_1), but in an S-shaped manner that depends on m . Deviations from these patterns have been suggested to indicate non-neutral rates [10]. The community size is $N = 10^3$.

175 Compared to the fully stochastic model that scales with $2^{-S}N^S$, here, we solve $N(S - 1)$
 176 equations for the marginal probability of each type, i.e. $N(S^2 - S)$ equations for the community.
 177 This model reduction allows us to approximate the equilibrium of large communities with many
 178 interacting types more rapidly.

179 2.2 The neutral expectation

180 We start by considering the neutral case – a situation where the rates of all types are equal
 181 ($f_i = \phi_i = 1$ for all i in $\{1, \dots, S\}$). In contrast to the deterministic model at equilibrium,
 182 Eq. (2), the frequencies of single stochastic realizations change through time, driven by the
 183 probabilistic nature of events. As a result, a distribution of frequencies centered at the value
 184 set by the source of immigrants (p_i) emerges. The spread of this distribution inversely depends
 185 on the magnitude of the immigration, m .

186 As shown in Fig. 1A, large immigration drives the equilibrium distribution towards its mean
 187 value, p_i . On the contrary, without or little immigration, the distribution splits. Thus, the
 188 frequencies zero (no individuals of the i -th type) and one (only individuals of the i -th type) are
 189 the most probable, decaying towards intermediate frequencies. This is a consequence of noisy

190 fluctuations that, for a single realization, lead to the extinction of all but one type. Whether
191 the frequency one or zero is most likely depends on the proximity of the initial state.

192 The mean frequency of the stochastic model identically corresponds to the frequency of
193 the deterministic model. As shown in Fig. 1B, regardless of the total immigration, the mean
194 frequency of a type increases linearly with the fraction of migrants of its kind.

195 Besides the mean frequency, one of the simplest, but most informative observables is the
196 occurrence frequency of individuals of a given type in the community. In other words, the
197 probability of observing at least one individual of that type, $P(n_i \geq 1)$. Immigration increases
198 this probability up to the point where the type is always observed in the community (Fig. 1C).
199 Importantly, this probability does not increase linearly with the fraction of migrants. Instead,
200 an S-shaped curve is observed, where changes of immigration of rare or abundant types do not
201 modify their occurrence.

202 Using two simple observables, the mean frequency and the occurrence frequency, we can
203 describe the state of types within a community. In the following, we relax the assumption of
204 neutrality – not enforcing equal growth and death rates. Then, we contrast both observables
205 to their neutral expectation.

206 **2.3 Immigration lessens the effect of growth and death differences**

207 To understand the effect of non-neutral rates, we start from a community composed of only
208 two types. Furthermore, we assume only one of them has a non-neutral rate, either f_i or ϕ_i . In
209 this way, we aim to see the effect in the neutral and non-neutral fractions of the community.

210 For a growth rate below one ($f_1 < 1$) or a death rate above one ($\phi_1 > 1$), the non-neutral type
211 has a reduced mean frequency that preserves its linear relationship to the fraction of immigrants
212 (Fig. 2A-B and Fig. 3A-B). However, in contrast to the neutral expectation, immigration does
213 play a role, as large migration can reduce the changes occurring in the internal community
214 dynamics (compare panels A to B in Fig. 2-3). In this context, the neutral type ($f_2 = \phi_2 = 1$)
215 benefits from the reduced proliferation of its partner, thus, gaining in frequency, especially if
216 most immigrants belong to the neutral type.

217 A similar picture arises for the occurrence pattern. While the non-neutral type occurs less
218 frequently, the neutral type thrives, occurring more often than when both types are neutral
219 (Fig. 2C-D and Fig. 3C-D). The change can be as severe as losing all non-neutral individuals

Figure 2: **Effect of non-neutral growth rates on the equilibrium of a community with two types.** One of two types has non-neutral growth rate ($f_1 \neq f_2 = 1$), but the death rate is neutral ($\phi_1 = \phi_2 = 1$). In contrast to its all-neutral ($f_1 = f_2 = 1$) expectation, a lower growth rate of the non-neutral type ($f_1 < f_2$) reduces its mean frequency and occurrence. The change can be of several orders of magnitude. Inversely, a larger growth rate of the non-neutral type ($f_1 > f_2$) increases its mean frequency and occurrence. The effect of growth rate differences on the internal dynamics is reduced if immigration is larger, especially for slowly growing types. Immigration is (A, C) $m/N = 10^{-3}$ and (B, D) $m/N = 10^{-1}$, with community size $N = 10^3$.

from the community (panel C in Fig. 2-3). Crucially, large total immigration can prevent this (compare panels C to D in Fig. 2-3), even if most migrants are of the neutral type.

Once the roles are reversed, so the non-neutral growth rate is above one ($f_1 > 1$) or the death rate below one ($\phi_1 < 1$), the mean frequency and occurrence patterns mirror the previous results (Fig. 2 and Fig. 3). Although changes produced by non-neutrality in growth ($f_1 \neq 1$) or death ($\phi_1 \neq 1$) rates are qualitatively similar, they show quantitative differences.

Figure 3: **Effect of non-neutral death rates on the equilibrium of a community with two types.**

One of two types has non-neutral death rate ($\phi_1 \neq \phi_2 = 1$), but neutral growth rate ($f_1 = f_2 = 1$). Differences in death rates modify the mean frequency and occurrence of both types. A larger immigration reduces differences to the all-neutral ($\phi_1 = \phi_2 = 1$) expectation in a similar fashion to differences in growth rate (Fig. 2). Immigration is (A, C) $m/N = 10^{-3}$ and (B, D) $m/N = 10^{-1}$, with community size $N = 10^3$.

226 We conclude that even for the simplest community (one with two types), just one non-neutral
 227 rate is enough to change the community occurrences and abundances substantially from their
 228 all-neutral expectation. This is more visible through the mean frequency (as changes of several
 229 orders of magnitude are possible) and for communities with little external migration – where
 230 the internal dynamics is more important.

Figure 4: **Occurrence-abundance pattern in general non-neutral communities.** (A) The non-neutral pattern of a diverse community largely resembles neutral patterns, see Fig. 1C. (B) However, the change from neutrality of each type can be large (blue arrows), shown here for $m/N = 10^{-3}$. In general, the mean frequency does not equal the fraction of immigrants p_i , assuming otherwise underestimates the change from neutrality (yellow arrows). (C) Similar to a community with two types, Fig. 2-3, the overlap to the neutral expectation increases when immigration, m , is increased to $m/N = 10^{-1}$. The growth and death rates, f_i and ϕ_i , were sampled from a normal distribution with mean 1 and standard deviation 0.1, where $P(f_i < 0.8) = P(f_i > 1.2) \approx 0.023$ and $P(\phi_i < 0.8) = P(\phi_i > 1.2) \approx 0.023$. The fractions of migrants p_i range from 10^{-4} to 10^{-1} and have a $G \approx 0.6$, Eq. (8), indicating intermediate immigration asymmetry. Except from the immigration rate m , all rates in (B-C) are equal. The community size is $N = 10^3$.

2.4 Neutral and non-neutral patterns are similar at the community level but full of differences at the level of types

Communities with two types might occur *in vitro*. However, in nature, communities are much more diverse, especially for microbes. We have produced random instances of such diverse communities, sampling growth and death rates, f_i and ϕ_i , from a normal distribution with mean one and a desired standard deviation. Similarly, we have produced random fractions of migrants, p_i , just conditioned on the Gini index of the community,

$$G = \frac{1}{S-1} \sum_{i,j} |p_i - p_j|. \quad (8)$$

This number that indicates the asymmetry in immigration between types from zero to one, allow us to compare communities quantitatively, regardless of their number of types S . As an example, for $G = 0$ the fractions of migrants are identical for each type, while for $G = 1$ the

Figure 5: **Occurrence-abundance pattern for different levels of asymmetry in the parameters.** A pattern is robust to various asymmetries in the immigration, p_i , and growth and death rates, f_i and ϕ_i . Each community has forty types. For low symmetry in immigration, the types span the range more widely. Colors from dark to light indicate how non-neutral a type is, quantified as the geometric distance from $(f_i, \phi_i) = (1, 1)$. Types overlap regardless of their non-neutrality. The fractions of immigrants, p_i , have a $G \approx 0.3$ (A-B) or $G \approx 0.6$ (C-D). The growth and death rates, f_i and ϕ_i , were sampled from a normal distribution with mean 1 and standard deviations 0.1 (A, C) or 0.2 (B, D). In the last case, $P(f_i < 0.8) = P(f_i > 1.2) \approx 0.159$ and $P(\phi_i < 0.8) = P(\phi_i > 1.2) \approx 0.159$. Immigration is $m/N = 10^{-1}$, with community size $N = 10^3$.

241 source pool only contains a single type.

242 Using these parameters, we have computed the occurrence and abundance frequency of all
 243 types in a certain community. Interestingly, the community patterns that we observe are very
 244 similar to those expected from neutrality (Fig. 4A compared to Fig. 1C) – even if asymmetries
 245 of growth, death, and immigration increase (Fig. 5). In particular, large immigration together
 246 with high biodiversity consistently lead to these patterns (Fig. 5). This indicates that neither

neutrality nor non-neutrality, but large immigration and biodiversity are behind these patterns.

Even when neutral and non-neutral patterns are similar at the community level, we observe large differences at the level of types. While in the “all-neutral” case, the mean frequency equals the fraction of migrants, $E(n_i)/N = p_i$, this is not the case in a non-neutral scenario (Fig. 4B-C). Neither is for the occurrence frequency. The distance from the neutral expectation of each type is not simply related to the level of non-neutrality of its own parameters. Rather, neutral and non-neutral types fall on, above, or below the neutral expectation (Fig. 5), highlighting the inherent multidimensionality determining the equilibrium of these communities.

To investigate the effect of single parameters at the level of types, we chose two representative types – one close to the neutral expectation, and another one distant from it (Fig. 6A). Our results show that types do not remain on or far from the neutral expectation. Rather, the relative magnitude of their growth and death rate, f_i and ϕ_i , is crucial to observe simultaneous decrease or increase in occurrence and mean frequency (Fig. 6C-D). In particular, types with a smaller fraction of immigrants, p_i , experience more abrupt changes. Only large fractions of immigrants allow to overcome the effect of growth and death rate differences, leading to large occurrence and mean frequency at the level of types (Fig. 6B).

2.5 To test neutrality the niche structure must be known first

So far we have used our model to compute observables based on known parameters. However, we can invert this process to infer parameters from simulations or experimental data.

Particularly relevant is the possibility of testing niche structure in data [15, 8, 9, 10]. Our model indicates care is needed to quantify the true difference from neutrality (Fig. 4B-C). In fact, the comparison of the selective case to the neutral case can only be inferred after fitting all parameters of the general model (m , p_i , f_i , and ϕ_i for all i). This is in contrast to the – often used – method by Sloan et al. for neutral conditions, where only the immigration rate m is fitted, while all growth and death rates are assumed $f_i = \phi_i = 1$, and the fraction of migrants p_i equalled to the mean frequency $E(n_i)/N$. Our results indicate these assumptions on the data are unfounded and lead to underestimate niche structure (Fig. 4B-C), especially in large communities with many types. Moreover, the consistent occurrence-abundance pattern that we observe (Fig. 5), and often reported in data [8, 9, 10], emerges from a general death-birth processes with immigration, Eq. (3), not just from a neutral process (where $f_i = \phi_i = 1$ for

Figure 6: **Effect of growth, death, and immigration at the level of types.** (A) The community shown corresponds to Fig. 5C, with $G \approx 0.6$ for p_i , and f_i and ϕ_i drawn from $\mathcal{N}(1, 0.1)$. Two types are spotted by circles, one that falls on the neutral expectation and the other distant from it. Single parameters are modified in (B-D) for both types. Arrows in the colorbars indicate their original values. (B) For large fractions of migrants, p_i , non-neutral types are indistinguishable from the neutral expectation; only for small fractions they are below it. (C) Different growth rates, f_i , lead non-neutral types to fall on, above, or below the neutral expectation. Changes are especially abrupt for the type with less immigration. (D) Different death rates, ϕ_i , mirror the effect of changing growth rates qualitatively. Immigration is $m/N = 10^{-1}$, with community size $N = 10^3$.

all i). Niche structure – and thus neutrality – can not be discarded or confirmed if certain parameters are fixed *a priori* [15].

The large number of parameters to be fitted requires large datasets. For a community with S types, $3S + 1$ parameters must be fitted, thus requiring at least $3S + 1$ data points. The $2S$ data points obtained from the occurrence and mean frequencies are not sufficient. We propose to include additional observables that can be readily computed from data [27]. These might include, but not be limited to, raw and central moments of the frequency. From this set of observables, available Bayesian methods [28] can be used to infer the parameters using Eq. (7a-7b).

In Fig. 7, we show two potential observables, the variance and the second moment of the distribution. In a community with two types, $S = 2$, both observables reflect the differences in growth, f_i , and death rates, ϕ_i . Only some variances overlap for distinct rates. In this sense,

Figure 7: **Variance and second raw moment of the frequency.** A community with two types is considered. (A-B) One type has non-neutral growth rate ($f_1 \neq f_2 = 1$) but neutral death rate ($\phi_1 = \phi_2 = 1$), or (C-D) a neutral growth rate ($f_1 = f_2 = 1$) but non-neutral death rate ($\phi_1 \neq \phi_2 = 1$). (B, D) $f_1 > 1$ and $f_1 < 1$ lead to a second raw moment above or below the neutral expectation, respectively. This moment increases continuously with the fraction of migrants, p_i , while the variance reaches a maximum at intermediate p_i (A, C). In contrast to the second raw moment, the variance of different growth and death rates overlaps. Differences in death rates mirror the effect of growth rate differences qualitatively. Immigration is $m/N = 10^{-1}$, with community size $N = 10^3$.

the second raw moment might provide more information to discriminate them. A set of similar observables could allow to characterize the rates of empirical communities.

3 Discussion

Understanding the drivers of communities is one of the main objectives of ecological research. In this work, we have used a stochastic death-birth model with immigration to investigate

the equilibrium distribution of communities. Comparing cases where changes only depend on the abundances to cases where types have different birth or death rates, we have identified conditions leading to a robust occurrence-abundance pattern – often reported empirically.

Our approach acknowledges the intrinsic density dependence of communities, Eq. (3), but simultaneously allow us to compute the equilibrium distribution of large and diverse communities, Eq. (7a-7b). Combining a method of model reduction [26] and a detailed balance analysis [24], we asked questions directly linked to empirical observations. In contrast to studies emphasizing biotic interactions [20, 21, 18, 13], our model can be classified with studies that focus on the differential adaptation to the environment [17, 19]. As some of these studies, our results highlight the central role of immigration and biodiversity in community ecology [19, 22].

We tested the reliability of our approach by reproducing known results of neutral adaptation [15]. Namely, that the mean frequency of a type equals its immigration and that the occurrence frequency increases in an S-shaped manner with the mean frequency, Fig. 1. These results already capture the important role of immigration but discard the frequency dependent effects of other types – for which biodiversity might be important.

The match between community level patterns of neutral models and empirical data has been documented extensively [6, 8, 9, 10]. Still, some empirical evidence is at odds with neutral theory [29, 30]. The mismatch with evolutionary history – including phylogenetic trees [29, 30], is one of them. It has been observed that mild differences in adaptation lead to full agreement [31] – indicating the need to consider models with differential adaptation, even if this is mild.

Here, we considered a general death-birth model where large immigration consistently led to a robust occurrence-abundance pattern. Interestingly, evidence suggests that large immigration might indeed be common in various environmental and host-associated microbiomes [10]. Others that deviate from the occurrence-abundance pattern have small immigration [10]. Such seems to be the case in *Caenorhabditis elegans*, where active destruction of microbes during feeding results in reduced immigration to the gut microbiome [32].

A second observation is that with differential adaptation, biodiversity takes a central role. In contrast to the simplest community of two interacting types (Fig. 2-3), diverse communities promote an occurrence-abundance pattern that resembles the neutral case (Fig. 5). With biodiversity, less extreme occurrences and mean frequencies are observed (compare Fig. 2-3 to Fig. 5). Our results agree with research showing that in the limit of high biodiversity, various

neutral and non-neutral patterns converge at the community level [19].

Previous research has speculated about the ecological role of types based on their location in the occurrence-abundance curve [10] – the motivation being the possibility to identify microbial taxa actively involved in biotic interactions. Our results indicate that such direct identification from occurrence-abundance curves remains challenging, mainly because neutral and non-neutral types can overlap (Fig. 6). We propose a way forward, based on the inclusion of new observables computed from data [27] (Fig. 7) combined with robust fitting approaches [28].

Our focus at the level of types revealed the difficulty of assessing niche structure and neutrality from empirical data. While niche and neutral patterns can be indistinguishable at the community level, at the level of types, big differences are observed (Fig. 5-6). Commonly, in microbial ecology, models have been tested at the community level, where, embraced by a principle of parsimony, neutral interpretations have been suggested [8, 9, 10]. Our model suggests this is indeed sensible for community level questions. However, for questions at the level of types – including that of ecological roles – general models including differential adaptation can not be avoided. In this case, no parsimonious preference can be given to neutral hypotheses.

The last observation calls for a broader discussion on terminology. As defined by Fisher and Mehta, a community is “statistically neutral” if its distribution can not be distinguished from a distribution constructed under the assumption of ecological neutrality. We must note, however, that ecological neutrality implies statistical neutrality, but statistical neutrality does not necessarily imply ecological neutrality [22]. As our results indicate, a reference to large immigration and biodiversity, rather than neutrality, is more accurate and prevents misleading interpretations, that in their worst form, could lead to unfounded generalizations or hold research questions back. On the contrary, our results suggest that numerous questions about neutrality, adaptation, and ecological roles, in microbial ecology and elsewhere are yet to be answered.

Although we mainly focused on microbial communities, our work can be framed in the larger macro-ecological literature. There, a substantial number of models have linked neutral and niche theories [20, 21, 18, 13, 17, 19]. Heated debates have occurred; however, they have benefited from a close revision of the assumptions on the models and a careful discussion of their implications [19, 6, 31]. The observation of asymptotically equivalent patterns for neutral and non-neutral rates is one of their main results [19]. We believe microbial research

can be guided along this line while offering powerful methods to investigate general ecological questions [27]. In particular, the possibility to work, *in vivo* and *in vitro*, with large and diverse communities in much shorter time scales [33].

Finally, we should mention some limitations of our work. A limitation of origin is that we considered a differential adaptation to the environment as the sole source of non-neutrality. Certainly, this is not true in nature, where types take part in numerous symbiotic interactions [13]. Therefore, any empirical application of our model should be preceded by evidence of little to no symbiosis. A technical limitation is that we have only approximated the stochastic dynamics [26]. Our results should be more robust in large communities where types have limited variance [26]. Interestingly, large immigration – which appears to be common in microbial communities [10] – might lead to satisfying this condition.

Although we provided a focused analysis of the occurrence-abundance pattern at equilibrium, future work could study its dynamics [34] and derive exact equations for these and other observables [27]. In addition, identifying neutral and non-neutral types remains an open problem. The development of methods for parameter inference from data [27] seems the way forward.

4 Conclusion

Here, we presented a general death-birth model with immigration. Using a method of reduction for the stochastic model, we analysed the equilibrium distribution of abundances for communities equally or differently adapted to the environment. We observe that the community pattern of occurrence-abundance, often reported empirically, is consistently observed in conditions of large immigration and high diversity, regardless of the adaptation to the environment. However, at the level of types, differences in adaptation still lead to large changes.

Availability of code

The data generated and analysed during the current study can be simulated from the Python code available via GitHub at <https://github.com/romanzapien/occurrence-abundance.git>.

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389 **Authors’ contributions**

390 The original model was developed in discussions between RZC and AT. RZC analysed the
 391 model, programmed the code, and wrote the initial draft. All authors interpreted the results,
 392 reviewed the manuscript, and approved the final version.

393 **Competing interests**

394 We declare no competing interests.

395 **References**

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