

# Enrichment analyses identify shared associations for 25 quantitative traits in over 600,000 individuals from seven diverse ancestries

Samuel Pattillo Smith<sup>1,2</sup>, Sahar Shahamatdar<sup>1,2</sup>, Wei Cheng<sup>1,2</sup>, Selena Zhang<sup>1</sup>, Joseph Paik<sup>1</sup>, Misa Graff<sup>3</sup>, Christopher Haiman<sup>4</sup>, T.C. Matise<sup>5</sup>, Kari E North<sup>3</sup>, Ulrike Peters<sup>6</sup>, Eimear Kenny<sup>7,8,9,10</sup>, Chris Gignoux<sup>11</sup>, Genevieve Wojcik<sup>12</sup>, Lorin Crawford<sup>1,13,14,\*</sup>, and Sohini Ramachandran<sup>1,2,\*,†</sup>

<sup>1</sup>Center for Computational Molecular Biology, Brown University, Providence RI

<sup>2</sup>Department of Ecology and Evolutionary Biology, Brown University, Providence RI

<sup>3</sup>Department of Epidemiology, University of North Carolina - Chapel Hill, Chapel Hill NC

<sup>4</sup>Department of Preventative Medicine, University of Southern California, Los Angeles CA

<sup>5</sup>Department of Genetics, Rutgers University, Piscataway NJ

<sup>6</sup>Public Health Sciences Division, Fred Hutchinson Cancer Research Center, Seattle WA

<sup>7</sup>The Center for Genomic Health, Icahn School of Medicine at Mount Sinai, New York City NY

<sup>8</sup>The Charles Bronfman Institute for Personalized Medicine, Icahn School of Medicine at Mount Sinai, New York City NY

<sup>9</sup>Department of Medicine, Icahn School of Medicine at Mount Sinai, New York City NY

<sup>10</sup>Department of Genetics and Genomic Sciences, Icahn School of Medicine at Mount Sinai, New York City NY

<sup>11</sup>Division of Biomedical Informatics and Personalized Medicine, University of Colorado, Denver CO

<sup>12</sup>Department of Epidemiology, Johns Hopkins University, Baltimore MD

<sup>13</sup>Department of Biostatistics, Brown University, Providence RI

<sup>14</sup>Microsoft Research New England, Cambridge MA

\*indicates these authors contributed equally

†To whom correspondence should be addressed: [sramachandran@brown.edu](mailto:sramachandran@brown.edu)

‡Present address: 164 Angell Street, Floor 3, Providence, RI 02906

Last edited November 3, 2021

## Abstract

Since 2005, genome-wide association (GWA) datasets have been largely biased toward sampling European ancestry individuals, and recent studies have shown that GWA results estimated from self-identified European individuals are not transferable to non-European individuals due to various confounding challenges. Here, we demonstrate that enrichment analyses which aggregate SNP-level association statistics at multiple genomic scales—from genes to genomic regions and pathways—have been underutilized in the GWA era and can generate biologically interpretable hypotheses regarding the genetic basis of complex

36 trait architecture. We illustrate examples of the robust associations generated by enrichment analy-  
37 ses while studying 25 continuous traits assayed in 566,786 individuals from seven diverse self-identified  
38 human ancestries in the UK Biobank and the Biobank Japan, as well as 44,348 admixed individuals  
39 from the PAGE consortium including cohorts of African-American, Hispanic and Latin American, Na-  
40 tive Hawaiian, and American Indian/Alaska Native individuals. We identify 1,000 gene-level associations  
41 that are genome-wide significant in at least two ancestry cohorts across these 25 traits, as well as highly  
42 conserved pathway associations with triglyceride levels in European, East Asian, and Native Hawaiian  
43 cohorts.

## 44 Introduction

45 Over the past two decades, funding agencies and biobanks around the world have made enormous invest-  
46 ments to generate large-scale datasets of genotypes, exomes, and whole-genome sequences from diverse  
47 human ancestries that are merged with medical records and quantitative trait measurements<sup>1-8</sup>. However,  
48 analyses of such datasets are usually limited to the application of standard genome-wide association (GWA)  
49 SNP-level association analyses, in which SNPs are tested one-by-one for significant association with a phe-  
50 notype<sup>9-11</sup> (Table 1). Yet, even in the largest available multi-ancestry biobanks, GWA analyses fail to offer  
51 a comprehensive view of genetic trait architecture among human ancestries.

52 SNP-level GWA results are difficult to interpret across multiple human ancestries due to a litany of  
53 confounding variables, including: (*i*) ascertainment bias in genotyping<sup>2,5</sup>, (*ii*) varying linkage disequilib-  
54 rium (LD) patterns<sup>12,13</sup>, (*iii*) variation in allele frequencies due to different selective pressures and unique  
55 population histories<sup>13-17</sup>, and (*iv*) the effect of environmental factors on phenotypic variation<sup>18-21</sup>. These  
56 confounders and the observed low transferability of GWA results across ancestries<sup>2,22,23</sup> have generated  
57 an important call for increasing GWA efforts focused on populations of diverse, non-European ancestry  
58 individuals.

59 We also note, as other studies have<sup>6,24</sup>, that the GWA SNP-level test of association is rarely applied to  
60 non-European ancestry individuals<sup>25</sup>. There are two likely explanations for leaving non-European ancestry  
61 individuals out of analyses: (*i*) researchers are electing to not analyze diverse cohorts due to a lack of  
62 statistical power and concerns over other confounding variables (recently covered in Ben-Eghan et al.<sup>24</sup>); or  
63 (*ii*) the analyses of non-European cohorts yield no genome-wide significant SNP-level associations. In either  
64 case, valuable information is being ignored in GWA studies or going unreported in resulting publications<sup>24-26</sup>.  
65 Even when diverse ancestries are analyzed, GWA studies usually condition on GWA results identified using  
66 European ancestry cohorts to detect and give validity to other SNP-level associations<sup>6</sup>. In our own analysis

67 of abstracts of publications between 2012 and 2020 using UK Biobank data, we found that only 33 out of 166  
68 studies (19.87%) reported genome-wide significant associations in any non-European ancestry cohort (Figure  
69 S1 - Figure S2). Focusing energy and resources on increasing GWA sample sizes without intentional focus  
70 on sampling of non-European populations will thus likely perpetuate an already troubling history of leaving  
71 non-European ancestry samples out of GWA analyses of large-scale biobanks such as the UK Biobank<sup>24</sup>  
72 (but see Sinnott-Armstrong et al.<sup>27</sup>). However, we note that non-European ancestry GWA studies have  
73 — and will continue to have — smaller sample sizes than existing and emerging European-ancestry GWA  
74 cohorts, limiting the precision of effect size estimates in these studies. What has received less attention than  
75 the need to improve GWA study design is the potential of enrichment analyses to characterize genetic trait  
76 architecture in multi-ancestry datasets while accounting for variable statistical power to detect, estimate,  
77 and replicate genetic associations among cohorts.

78 In this analysis, we illustrate that focusing solely on  $p$ -values from the standard GWA framework is  
79 insufficient to capture the genetic architecture of complex traits. Specifically, we propose that expansion of  
80 association analyses to the genomic scale of genes and pathways generates robust and interpretable hypothe-  
81 ses about trait architecture in multi-ancestry cohorts. These enrichment analyses can increase the power to  
82 detect genes through the aggregation of SNPs of small effect (which explain the majority of the heritability  
83 of most traits<sup>28,29</sup>). Mathieson<sup>30</sup> recently highlighted the pattern of homogeneity of direction of effect in  
84 multi-ancestry studies even when individual SNPs are not categorized as genome-wide significant in multiple  
85 ancestries. Gene and pathway level analyses can prioritize biological regions where there is homogeneity in  
86 the direction of effect, generating biologically interpretable hypotheses for the genetic architecture of complex  
87 traits in multiple ancestry cohorts.

88 In this study of 25 quantitative traits and more than 600,000 diverse individuals from the UK Biobank  
89 (UKB), BioBank Japan (BBJ), and the PAGE study data (Table S1 - Table S9), we detail biological insights  
90 gained from the application of gene and pathway level enrichment analyses to seven diverse ancestry cohorts.  
91 We perform genetic association tests for SNPs, genes, and pathways across multiple ancestry groups with  
92 a trait of interest. We test for significantly mutated subnetworks of genes using known protein-protein in-  
93 teraction networks and the Hierarchical HotNet software<sup>31</sup>. Enrichment analyses do not require generating  
94 additional information beyond standard GWA inputs (or outputs, for methods that take in GWA summary  
95 statistics). We demonstrate that moving beyond SNP-level associations allows for a biologically comprehen-  
96 sive prioritization of shared and ancestry-specific mechanisms underlying genetic trait architecture.

## 97 **Materials and Methods**

### 98 **Data overview**

99 We performed statistical tests of association at the SNP, gene, and pathway-level for 25 quantitative traits.  
100 These analyses were performed on data seven ancestry cohorts drawn from the UK Biobank, BioBank Japan  
101 (BBJ), and PAGE consortia (Table S3). Descriptions of the samples studied, including sample size, and  
102 number of SNPs for each ancestry cohort are given in Table S1 and Table S4 - Table S9. For an extensive  
103 description of each cohort from the three biobanks that we analyze in this study, see the Supplemental  
104 Information.

### 105 **SNP-level GWA analyses**

106 In the European, African, and South Asian ancestry cohorts from the UK Biobank, we performed GWA  
107 studies for each ancestry-trait pair in order to test whether the same SNP(s) are associated with a given  
108 quantitative trait in different ancestries. SNP-level GWA effect sizes were calculated using plink and the  
109 `--glm` flag<sup>32</sup>. Age, sex, and the first twenty principal components were included as covariates for all traits  
110 analyzed<sup>7</sup>. Principal component analysis was performed using flashpca 2.0<sup>33</sup> on a set of independent markers  
111 derived separately for each ancestry cohort using the plink command `--indep-pairwise 100 10 0.1` .  
112 Summary statistics for the 25 quantitative traits in the Biobank Japan, as well as available ancestry-trait  
113 pairs in the PAGE study data, were then compared with the results from the association analyses in the UK  
114 Biobank cohorts (same traits as listed in Table S4). In each analysis of an ancestry-trait pair, a separate  
115 Bonferonni-corrected significance threshold was calculated using the number of SNPs tested in that particular  
116 ancestry-trait pair (Table S10).

### 117 **Comparison of SNP level fine-mapping methods and effect size direction**

118 We compared the results of two fine-mapping methods, SuSie<sup>34</sup> and PESCA<sup>35</sup>, when applied to SNP-level  
119 summary statistics in the European (UKB) and East Asian (BBJ) cohorts. SuSie is an iterative Bayesian  
120 stepwise selection method that identifies a credible set of SNPs that contribute to a phenotype of interest<sup>34</sup>.  
121 Using the effect sizes and standard errors generated from the standard GWA framework for each trait in each  
122 ancestry, we applied SuSiE in order to identify probable sets of causal SNPs. We then found the correlation  
123 between the posterior inclusion probabilities of each SNP in the European and East Asian cohorts (Table  
124 S11 - Table S13).

125 Unlike SuSiE, the PESCA framework is explicitly designed for identifying shared SNP-level association

126 signals between multiple ancestry cohorts versus ancestry-specific associations<sup>35</sup>. In addition to GWA sum-  
127 mary statistics, PESCA uses information about the correlation structure between SNPs (i.e., LD) to identify  
128 SNPs that are likely to be causal in two cohorts of interest. Shi et al.<sup>35</sup> analyzed seven continuous traits in  
129 the European (UKB) and East Asian (BBJ) cohorts using PESCA and produced posterior probabilities that  
130 individual SNPs were: (i) associated with a phenotype in both the both cohorts, (ii) associated with the  
131 trait of interest only in the European cohort, or (iii) associated with the trait of interest only in the East  
132 Asian cohort. For each trait, we calculated the number of SNPs that were nominally significant ( $p$ -value  
133  $< 10^{-5}$ , as in the original PESCA analysis) in the standard GWA framework in both the European and  
134 East Asian cohorts and had a PESCA posterior probability of being associated in both ancestries  $> 0.8$  (see  
135 Table S14). We also found the number of SNPs that had a PESCA posterior probability of being associated  
136 in both ancestries  $> 0.8$  that were only identified as significant in one ancestry using the GWA framework.

137 Finally, we explored the recent proposition of Mathieson<sup>30</sup> that the direction of effect for SNP-level  
138 summary statistics might be conserved among ancestry cohorts even if those variants are not genome-wide  
139 significant in either cohort. To that end, for each of the 25 traits that we analyzed, we compared the direction  
140 of SNP effect sizes between the European and East Asian ancestry cohorts. We were only able to carry this  
141 analysis out for variants that were genotyped in both cohorts (Table S14). For each remaining nominally  
142 significant variant, we stored the direction of the effect size and checked the direction of effect size in the other  
143 ancestry. When zero was included within the range of the effect size plus or minus one standard deviation,  
144 we assumed the SNP did not have the same direction in both cohorts. The results of our comparison are  
145 shown in Table S14. We note that this test may be confounded by the precision of effect size estimation and  
146 warrants further exploration, including an analysis of local false sign discovery rates (see<sup>36,37</sup>).

## 147 Gene-level association tests

148 In order to test aggregated sets of SNP-level GWA effect sizes for enrichment of associated mutations with  
149 a given quantitative trait, we applied gene- $\varepsilon$ <sup>38</sup> to each ancestry cohort we studied for each trait of interest,  
150 resulting in 125 sets of gene-level association statistics (Table S3, Table S15). The gene- $\varepsilon$  method takes two  
151 summary statistics as input: (i) SNP-level GWA marginal effect size estimates  $\hat{\beta}$  derived using ordinary least  
152 squares and (ii) an LD matrix  $\Sigma$  empirically estimated from external data (e.g., directly from GWA study  
153 genotype data, a matrix estimated from a population with similar genomic ancestry to that of the samples  
154 analyzed in the GWA study). It is well-known that SNP-level effect size estimates can be inflated due to  
155 various correlation structures among genome-wide genotypes. gene- $\varepsilon$  uses its inputs to derive regularized  
156 effect size estimates through elastic net penalized regression.

157 In practice, gene- $\epsilon$  and other enrichment analyses<sup>39–41</sup> can be applied to any user-specified set of genomic  
158 regions, such as regulatory elements, intergenic regions, or gene sets. These gene-level association tests enable  
159 identifying traits in which genetic architecture may be heterogeneous among individuals at the SNP-level  
160 across individuals; applying gene- $\epsilon$  in multiple ancestry cohorts allows researchers to further test whether  
161 genes associated with a trait of interest are the same, or vary, across ancestries. gene- $\epsilon$  takes as input a list  
162 of boundaries for all regions to be tested for enrichment of associations. In our study, we applied gene- $\epsilon$  to all  
163 genes and transcriptional elements defined in Gusev et al.<sup>42</sup> for human genome build 19. Throughout this  
164 study, we refer to the resulting effect size estimates produced by gene- $\epsilon$  as “gene-level association statistics”.

165 In our gene-level analysis, SNP arrays included both genotyped and high-confidence imputed SNPs (in-  
166 formation score  $\geq 0.8$ ) for each ancestry-trait pair. To compute the LD matrix, we first pruned highly linked  
167 SNPs so that the number of SNPs included for any chromosome was less than 35,000 SNPs — the computa-  
168 tional limit of gene- $\epsilon$  due to the size of the LD matrix — using the plink command `--indep-pairwise 100 10 0.5`.  
169 For the UK Biobank European, South Asian, and African ancestry cohorts, we then derived empirical LD es-  
170 timates between each pair of SNPs for each chromosome in each cohort using plink flag `--r square` applied  
171 to the empirical genotype and high-confidence imputed data. The ancestry-specific SNP arrays were then  
172 used to calculate 23,603 gene-level association statistics for the European ancestry cohort, 23,671 gene-level  
173 association statistics for the South Asian ancestry cohort, and 23,575 gene-level association statistics for the  
174 African ancestry cohort.

175 To calculate gene-level association statistics using Biobank Japan summary statistics, we first found the  
176 intersection between SNPs included in the analysis of each trait (Table S4) and SNPs included in the 1000  
177 Genomes Project phase 3 data for the sample of 93 JPT individuals. Note, this intersection was different  
178 among some traits as the genotype data in the Biobank Japan were from different studies, which in turn used  
179 different genotyping arrays. We then pruned highly linked markers for each trait separately using the plink  
180 flag `--indep-pairwise 100 10 X` where `X` was determined by finding the highest possible value that led  
181 to less than 35,000 SNPs being included on each chromosome for the trait. Because of the increased density  
182 of SNPs with effect size estimates for Height, `X` was set to prune more conservatively at `X = 0.15`. For all  
183 other traits, `X` was set to 0.5. The number of regions for which a gene- $\epsilon$  gene-level association statistic was  
184 calculated for each trait is given in Table S4.

185 GWA summary statistics for the five cohorts in the PAGE study data were used as input to gene- $\epsilon$   
186 for each available ancestry-trait combination. The array of markers for each ancestry cohort in the PAGE  
187 study data was pruned using plink flag `--indep-pairwise 100 10 X`. `X` was set to the maximum value  
188 in each ancestry that ensured no chromosome contained more than 35,000 markers: `X` was set to 0.05 for  
189 the African-American cohort, 0.08 for the Hispanic and Latin American and AIAN cohorts, and 0.25 for

190 the Native Hawaiian cohorts. Finally, for each ancestry-trait combination, genes that passed the Bonferroni  
191 corrected  $p$ -value ( $p = 0.05/\text{number of genes tested}$ ) were labeled as “significant” throughout this study (see  
192 Table S15 for specific thresholds).

## 193 **Pathway analysis and network propagation using Hierarchical HotNet**

194 We tested for shared and divergent gene-level association results among interacting genes for each trait  
195 of interest using network propagation of gene- $\epsilon$  gene-level association statistics as input to Hierarchical  
196 HotNet<sup>31</sup>. Hierarchical HotNet identifies significantly altered subnetworks using gene-level scores as input;  
197 in this study, these gene scores were set to  $-\log_{10}$ -transformed  $p$ -values of gene- $\epsilon$  gene-level association test  
198 statistics (see also Nakka et al.<sup>41,43</sup>). For each ancestry-trait combination, we assigned  $p$ -values of 1 to genes  
199 with  $p$ -values greater than 0.1 to make the resulting networks both sparse and more interpretable (again see  
200 Nakka et al.<sup>41,43</sup>). In addition, ancestry-trait pairs in which less than 25 genes produced gene- $\epsilon$   $p$ -values  
201 less than 0.1 were discarded as there were an insufficient number of gene-level statistics to populate the  
202 protein-protein interaction networks.

203 We used three protein-protein interaction networks: ReactomeFI 2016<sup>44</sup>, iRefIndex 15.0<sup>45</sup>, and HINT+HI<sup>46,47</sup>.  
204 For the ReactomeFI 2016 interaction network, interactions with confidence scores less than 0.75 were dis-  
205 carded. The HINT+HI interaction network consists of the combination of all interactions in HINT binary,  
206 HINT co-complex, and HuRI HI interaction networks. We ran Hierarchical HotNet ( $10^3$  permutations) on  
207 the thresholded  $-\log_{10}$ -transformed gene-level  $p$ -values for each ancestry-trait combination. We restricted  
208 our further investigation to the largest subnetwork identified in each significant ancestry-trait-interaction  
209 network combination ( $p < 0.05$ ).

## 210 **Results**

### 211 **SNP-level replication of GWA results among ancestries is the exception, not the** 212 **rule**

213 Multiple recent studies have interrogated the extent to which SNP-level associations for a given trait repli-  
214 cate across ancestries, both empirically and under a variety of models (see Wojcik et al.<sup>6</sup>, Durvasula and  
215 Lohmueller<sup>22</sup>, Shi et al.<sup>35</sup>, Carlson et al.<sup>48</sup>, Liu et al.<sup>49</sup>, Eyre-Walker<sup>50</sup>, Shi et al.<sup>51</sup>). To extensively  
216 compare variant-level associations among the seven ancestry cohorts that we analyzed, we first examined  
217 the number of genome-wide significant SNP-level associations that replicated exactly based on chromosomal  
218 position and rsID in multiple ancestries (see Figure S3a and Figure S3c, with Bonferroni-corrected thresholds



219 provided in Table S10). Exact replication of at least one SNP-level association across two or more ancestries  
220 occurs in all 25 traits that were studied. The C-reactive protein (CRP) trait had the highest proportion of  
221 replicated SNP associations in multiple ancestries, with 18.95% replicating using the standard GWA frame-  
222 work in at least two ancestries, but has a relatively low number of unique GWA significant SNPs (2,734)  
223 when compared to other traits (Figure 1). This is likely because the genetic architecture of CRP is sparse  
224 and highly conserved across ancestries, as is shown in Figure 2. We note that the concordance of genome-  
225 wide significant SNP-level association statistics for CRP among five ancestry cohorts is exceptional. In the  
226 other 24 traits we analyzed, we did not observe any SNP-level replication among five cohorts. C-reactive  
227 protein, which is encoded by the gene of the same name located on chromosome 1, is synthesized in the  
228 liver and released into the bloodstream in response to inflammation. In our standard GWA analysis of SNP-  
229 level association signal in each ancestry cohort with CRP, rs3091244 is genome-wide significant in a single  
230 ancestry cohort. rs3091244 has been functionally validated as influencing C-reactive protein levels<sup>52,53</sup> and  
231 is linked to genome-wide significant SNPs in the other two ancestries for which genotype data is available.  
232 Interestingly, all GWA significant SNP-level associations for CRP in the Native Hawaiian ancestry cohort  
233 replicate in both the African-American (PAGE) and the Hispanic and Latin American cohorts (these three  
234 cohorts were all genotyped on the same array).

235 In the other 24 traits, the proportion of genome-wide SNP-level replications was below 10% (Figure  
236 1a). For polygenic traits, replication of SNP-level GWA results is challenging to interpret considering the  
237 large number of GWA significant associations for the trait overall. For example, height contains the largest  
238 number of replicated SNP-level associations in our multi-ancestry analysis — but these only represent 8.90%  
239 of all unique SNP-level associations with height discovered in any ancestry cohort. A more comprehen-  
240 sive discussion of previously associated SNPs is available for both height and CRP in the Supplemental  
241 Information.

## 242 **Fine-mapping methods have variable efficacy in identification of SNP level asso-** 243 **ciations among ancestry cohorts**

244 Often, replication of GWA results across cohorts is tested using genomic regions centered on a SNP. Scans  
245 across the region surrounding the SNP of interest are usually defined arbitrarily — using physical windows  
246 (or “clumps”) to smooth over ascertainment bias and varying LD across cohorts or ancestries instead of  
247 using regions that are biologically annotated such as genes or transcriptional elements. While clumping  
248 presents an easy way to scan for regional replication of a given GWA finding, the corresponding results are  
249 not readily interpretable when prioritizing GWA results for downstream validation. We performed clumping



250 using windows of size 1Mb centered around significant SNP-level associations (see Materials and Methods).  
251 Height had the largest proportion of windows that contain a SNP-level association that replicated in at least  
252 two ancestries (Figure S3b and Figure S3e). In the three traits with the greatest proportion of windows  
253 containing SNP-level replications — height (77.09% of clumps), urate (65.89%), and low density lipoprotein  
254 (54.40%) — we then recorded the number of genes and transcriptional elements within the window that  
255 contained GWA significant SNP-level associations. We found that for all three traits, the vast majority of  
256 1Mb windows that were used to clump SNP-level associations contained multiple genes and transcriptional  
257 elements with significantly associated SNPs: height (94.04%, 17.93 genes in clump (mean)  $\pm$  15.71 (standard  
258 deviation)), urate (97.47%,  $18.44 \pm 13.72$ ), and low density lipoprotein (99.12%,  $14.85 \pm 12.89$ ). Thus, we  
259 find window-based clumping does not easily produce biologically interpretable hypotheses for downstream  
260 validation.

261 Recent analyses of multi-ancestry GWA cohorts have also tested for effect size heterogeneity<sup>6,34,51,54,55</sup>.  
262 We applied the fine-mapping method SuSiE<sup>34</sup> to identify signals of effect size heterogeneity in the three  
263 ancestry cohorts for which we had access to raw genotype data (UK Biobank European ancestry, African  
264 ancestry, and South Asian ancestry individuals; see Table S1). We find little evidence of correlated SuSiE  
265 effect size estimates among ancestry cohorts, including among independent subsamples of the UK Biobank  
266 European ancestry individuals Table S11 - Table S13. In addition, we applied PESCA (a method developed  
267 by Shi et al.<sup>51</sup>) to the results of our SNP-level analysis to understand how the modeling of LD affected the  
268 power to identify probably causal SNPs shared in the European and East Asian ancestry cohorts. PESCA  
269 improves upon standard clumping approaches by modeling the LD in a region to identify SNPs that are  
270 likely to be causal for the same trait in multiple ancestries. In a comparison with the results from seven  
271 continuous traits analyzed in the original study<sup>51</sup>, we found that the vast majority of SNPs identified by  
272 PESCA as causal (posterior probability  $> 0.8$ ) in both ancestries were also nominally significant in our SNP-  
273 level association results (see Table S14). Both SuSiE<sup>34</sup> and PESCA<sup>51</sup> demonstrate the utility of modeling  
274 variation in LD structure among ancestries when conducting multi-ancestry GWA studies.

275 Recently, Mathieson<sup>30</sup> proposed the hypothesis that the direction of effect sizes is the same among  
276 ancestries, even when the effects are not genome-wide significant. To test this, we compared the direction  
277 of effect in SNPs that were significant in either the European or East Asian ancestry cohort to the direction  
278 of the effect in the other ancestry where the SNP was tested using the standard GWA framework. We limit  
279 the comparison to the European and East Asian cohorts due to their large sample sizes which increases the  
280 precision of effect size estimates. Table S14 shows the number of variants that were significantly associated  
281 with each trait in at least one of the European and East Asian ancestry cohorts, and also displays the number  
282 of those variants that have the same direction of effect as the significant variant in the other ancestry. In

283 the 25 traits that we analyzed, the direction of effect was conserved in both the European and East Asian  
284 ancestry cohorts (between effect direction concordance from 55.87% and 76.56% of SNPs across 25 traits).  
285 The remaining SNPs where the direction of the effect size was not conserved represent those SNPs that:  
286 (i) had different direction of effect size, (ii) were not tested in both ancestry cohorts, or (iii) had effect  
287 size estimates within one standard error of zero (Table S14). The observed conservation of effect size  
288 direction in multiple ancestry cohorts, even when SNPs are non-significant in one or more cohorts, is a  
289 primary assumption of regional enrichment methods and supports Mathieson<sup>30</sup>'s hypothesis and findings.  
290 This suggests that regional enrichment methods, which are sensitive to shared patterns of effect size direction  
291 among cohorts, are a natural approach to apply to GWA summary statistics even in the absence of replication  
292 SNP-level GWA signals among cohorts.

### 293 **Valuing biological mechanism over statistical significance**

294 Enrichment analyses aggregate SNP-level association statistics using predefined SNP sets, genes, and path-  
295 ways to identify regions of the genome enriched for trait associations beyond what is expected by chance.  
296 Published enrichment analyses have demonstrated the ability to identify trait associations that go unidenti-  
297 fied when using standard SNP-level GWA analysis<sup>38,40,56-60</sup>. The standard GWA method is known to have  
298 a high false discovery rate (FDR)<sup>36,61</sup>, which enrichment analyses can mitigate. Our analyses in Figure S4  
299 and Figure S5 illustrate that two methods — regression with summary statistics (RSS)<sup>60</sup>, a fully Bayesian  
300 method; and gene- $\epsilon$ <sup>38</sup> — control FDR particularly well both in the presence and absence of population struc-  
301 ture. Enrichment methods also increase power for identifying biologically interpretable trait associations in  
302 studies with smaller sample sizes than present-day GWA studies. For example, Nakka et al.<sup>41</sup> identified  
303 an association between *ST3GAL3* and attention deficit hyperactivity disorder (ADHD) using methods that  
304 aggregated SNP-level signals across genes and networks. ADHD is a trait with heritability estimates as high  
305 as 75% which had no known genome-wide significant SNP-level associations at the time; Nakka et al.<sup>41</sup>  
306 studied genotype data from just 3,319 individuals with cases, 2,455 controls and 2,064 trios<sup>62</sup>. A study by  
307 Demontis et al.<sup>63</sup> later found a SNP-level association in the *ST3GAL3* gene, but was only able to do so  
308 with a cohort an order of magnitude larger (20,183 individuals diagnosed with ADHD and 35,191 controls,  
309 totaling 55,374 individuals).

310 Because non-European GWA ancestry cohorts usually have much smaller sample sizes compared to  
311 studies with individuals of European ancestry, enrichment analyses offer a unique opportunity to boost  
312 statistical power and identify biologically relevant genetic associations with traits of interest using multi-  
313 ancestry datasets. In a simulation study using synthetic phenotypes generated from the European and

314 African ancestry cohorts in the UK Biobank, we show that gene- $\epsilon$  is able to identify significantly associated  
315 genes even in smaller cohorts ( $N = 10,000$  and  $N = 4,967$  in the European and African ancestry cohorts,  
316 respectively) without the inflated false discovery rate that is often exhibited by the standard GWA framework  
317 (Figure S6 and Figure S7). Additionally, in these simulations, gene- $\epsilon$  correctly identifies “causal” genes that  
318 are commonly associated in both cohorts (Figure S8 and Figure S9). These simulations illustrate the utility  
319 of modeling LD (and in the case of gene- $\epsilon$ , additionally shrinking inflated effect sizes) information to identify  
320 enrichment of SNP-level associations in predefined SNP sets.

321 In an analysis performed by Ben-Eghan et al.<sup>24</sup> on 45 studies analyzing UK Biobank data, the second  
322 most commonly stated reason for omitting non-European cohorts in applied GWA analyses was due to lack of  
323 power for identifying SNP-level GWA signals. We tested for gene-level associations in each of the 25 complex  
324 traits in each ancestry cohort for which we had data (Table S1 - Table S9), and identified associations in  
325 genes and transcriptional elements shared across ancestries for every trait. All of our analyses discussed here  
326 used gene- $\epsilon$  (see performance comparison with other enrichment analyses in Cheng et al.<sup>38</sup> and Figure S6  
327 - Figure S9), an empirical Bayesian approach that aggregates SNP-level GWA summary statistics, where  
328  $p$ -values for each gene are derived by constructing an empirical null distribution based on the eigenvalues  
329 of a gene-specific partitioning of the LD matrix (for more details, see Cheng et al.<sup>38</sup>). Our analyses show  
330 that several hematological traits have a higher rate of significant gene-level associations that replicate across  
331 multiple ancestry cohorts than SNP-level associations that replicate across ancestry cohorts (Figure 1b).  
332 These include platelet count (PLC), mean corpuscular hemoglobin (MCH), mean corpuscular hemoglobin  
333 concentration (MCHC), hematocrit, hemoglobin, mean corpuscular volume (MCV), red blood cell count  
334 (RBC), and neutrophil count (Figure S3f). Focusing on platelet count as an example, we identify 65 genes  
335 that are significantly enriched for associations in multiple ancestries when tested using gene- $\epsilon$  (see Table S15  
336 for details on Bonferroni thresholds used to correct for the number of genes tested)<sup>38</sup>. Fifty-five of these  
337 genes are significantly associated in both the European and East Asian ancestry cohorts, and the remaining  
338 ten all replicate in other pairs of ancestry cohorts. Overall, each of the six ancestry cohorts in our analysis  
339 share at least one significant gene with another ancestry cohort, as shown in Figure S10.

340 Results from gene-level enrichment analyses can be further propagated on protein-protein interaction  
341 networks to identify interacting genes enriched for association signals<sup>64</sup>. Often, studies use network propa-  
342 gation as a way to incorporate information from multiple “omics” databases in order to identify significantly  
343 mutated gene subnetworks or modules contributing to a particular disease<sup>65</sup>. An unexplored extension  
344 of network propagation is how it can be used with multi-ancestry GWA datasets to identify significantly  
345 mutated subnetworks that are shared or ancestry-specific<sup>43</sup>.

346 To conduct network propagation of gene-level association results in our analyses, we applied the Hier-

347 archical HotNet method<sup>31</sup> to gene- $\epsilon$  gene-level association statistics for each trait-ancestry data set. In 3,  
348 we display the significant ( $p$ -value  $< 0.01$ ) network results for triglyceride levels in three ancestry cohorts:  
349 European, East Asian, and Native Hawaiian (networks separated by ancestry are available in Figure S11).  
350 In both the European and East Asian cohorts, we identify enrichment of mutations in a highly connected  
351 subnetwork of genes in the apolipoprotein family. In addition, we identify a gene subnetwork enriched for  
352 mutations in the East Asian and Native Hawaiian cohorts that interacts with the significantly mutated sub-  
353 network identified in both the European and East Asian cohorts. For instance, beta-secretase 1 (*BACE1*),  
354 is a genome-wide significant gene-level association in the East Asian cohort but does not contain SNPs  
355 previously associated with triglycerides in any ancestry cohort in the GWAS catalog. Additionally, both  
356 *APOL1* and *HBA1* were identified as significantly associated with triglycerides using gene- $\epsilon$  in our analysis  
357 of the Native Hawaiian ancestry cohort, and both genes were part of significant subnetworks identified by  
358 Hierarchical HotNet in the European and Native Hawaiian ancestry cohorts. Details on replicated SNP-level  
359 and gene-level associations among ancestries for triglyceride levels are shown in Figure S12 and Figure S13,  
360 respectively. SNP-level and gene-level association results are further discussed for both platelet count and  
361 triglyceride levels in the Supplemental Information.

## 362 Discussion

363 Many recent studies have proposed changes to multi-ancestry GWA study design<sup>2,5,6,22,24–27,66,67</sup>. In this  
364 analysis, we have focused on the potential of *methods* to increase the insight gained into complex trait archi-  
365 tecture from multi-ancestry GWA datasets via the generation of biologically interpretable hypotheses. We  
366 demonstrate the potential gains of moving beyond standard SNP-level GWA analysis using 25 quantitative  
367 complex traits among seven human ancestry cohorts in three large biobanks: BioBank Japan and the UK  
368 Biobank, and the PAGE consortium database (Table S1 - Table S9). Ultimately, we believe that complex  
369 traits demand analysis across multiple genomic scales and ancestries in order to gain biological insight into  
370 complex trait architecture and ultimately achieve personalized medicine.

371 As has been previously noted<sup>5,25</sup>, non-European ancestry cohorts are often excluded from GWA analyses  
372 of multi-ancestry biobanks; complementing the analyses of Ben-Eghan et al.<sup>24</sup>, we find that 80.13% of UK  
373 Biobank studies over the last 9 years only report significant SNP-level associations in the white British cohort  
374 (Figure S1 - Figure S2), despite the tens of thousands of individuals of non-European ancestry sampled in  
375 that data set. Unless this practice is curbed by the biomedical research community, it will exacerbate already  
376 existing disparities in healthcare across diverse communities. There are undoubted benefits from increased  
377 sampling in a given ancestry for association mapping using the standard GWA framework, but it is still

378 unknown the extent to which results from larger GWA and fine-mapping studies using European-ancestry  
379 genomes will generalize to the entire human population<sup>2,20</sup>.

380 Here, we have not addressed the downstream consequences of using self-identified ancestry to define  
381 cohorts in large-scale GWA studies (but see Urbut et al.<sup>37</sup>, Willer et al.<sup>68</sup>, Lin et al.<sup>69</sup>, Yang et al.<sup>70</sup>).  
382 Each sample we analyzed has also experienced environmental exposures that may influence the statistical  
383 detection of genetic associations, and some of those environmental exposures may be correlated with genomic  
384 ancestry<sup>13,71–73</sup>. Interrogation of the influence of gene by environment interactions on complex traits must  
385 be done with highly controlled experiments, which can in turn help prioritize traits in which association  
386 studies will be interpretable and useful. Increasing sample size in GWA studies alone will not resolve these  
387 fundamental biological questions: the proportion of phenotypic variance explained by associations discovered  
388 as sample sizes increase in GWA studies has largely reached diminishing returns<sup>39</sup>, and gene by environment  
389 interactions are increasingly influential, and estimable, in large biobanks with cryptic relatedness<sup>74,75</sup>.

390 Many recent methodological advances that leverage GWA summary statistics have focused on: testing  
391 the co-localization of causal SNPs (e.g., fine mapping<sup>34,54,76,77</sup>); the non-additive effects of SNP-level inter-  
392 actions (i.e., epistasis<sup>78,79</sup>); and multivariate GWA tests<sup>37,79–81</sup>. While these methods can be extended and  
393 applied to multi-ancestry GWA analyses, they still focus on SNP-level signals of genetic trait architecture  
394 (see also Brown et al.<sup>82</sup>, Galinsky et al.<sup>83</sup>). Unlike the traditional GWA method, enrichment analyses  
395 increase statistical power by aggregating SNP-level signals of genetic associations and allowing for genetic  
396 heterogeneity in SNP-level trait architecture across samples, as well as offering the opportunity for immedi-  
397 ate insights into trait architecture using existing datasets. However, these methods have been comparatively  
398 underused in multi-ancestry GWA studies.

399 While many studies note that differences in LD across ancestries affect transferability of effect size  
400 estimates<sup>6,48,84–86</sup>, recent studies in population genetics have additionally debated how various selection  
401 pressures and genetic drift may hamper transferability of GWA results across ancestries (see for example,  
402 Edge and Rosenberg<sup>15,16</sup>, Novembre and Barton<sup>18</sup>, Harpak and Przeworski<sup>20</sup>, Durvasula and Lohmueller<sup>22</sup>,  
403 Mostafavi et al.<sup>23</sup>). Future GWA studies should be coupled with approaches from studies of how evolutionary  
404 processes shape the genetic architecture of complex traits<sup>20,28,50,87</sup>.

405 Two open questions must be tackled when studying complex trait architecture in the multi-ancestry  
406 biobank era: (i) to what extent is the true genetic trait architecture (causal SNPs and/or their effects on  
407 a trait of interest) heterogeneous across cohorts?<sup>6,88</sup> and (ii) which components of GWA results (e.g.  $p$ -  
408 values, estimated effect size, direction of effect sizes) are transferable across ancestries, at any genomic scale?  
409 Continued application of the standard SNP-level GWA approach will not answer these questions. However,  
410 enrichment methods that aggregate SNP-level effects, test for effect size heterogeneity, leverage genomic

411 annotations and gene interaction networks offer opportunities to directly test these fundamental questions.  
412 Methods can and should play an important role as biomedical research shifts current paradigms to extend  
413 the benefits of personalized medicine beyond people of European ancestry.

414 Additionally, biomedical researchers should continue to pressure both funding agencies and institutions  
415 to diversify their sampling efforts in the name of inclusion and addressing—instead of exacerbating—genomic  
416 health disparities. In addition to those efforts, we believe existing and new methods can increase the return  
417 on investment in multi-ancestry biobanks, ensure that every bit of information from these datasets is studied,  
418 and prioritize biological mechanism above SNP-level statistical association signals by identifying associations  
419 that are robust across ancestries.

## 420 **Appendices**

421 Detailed information about the SNP-level results for both C-reactive protein and height, as well as, gene  
422 and pathway level associations for platelet count and triglyceride levels can be found in the Appendix.

## 423 **Declaration of interests**

424 C.G. owns stock in 23andMe. E.E.K. and C.G. are members of the scientific advisory board for Encompass  
425 Bioscience. E.E.K. consults for Illumina.

## 426 **Acknowledgments**

427 We thank Kirk Lohmueller and Alicia R. Martin for helpful comments on an earlier version of this manuscript,  
428 as well as the Crawford and Ramachandran Labs for helpful discussions. This research was conducted in  
429 part using computational resources and services at the Center for Computation and Visualization at Brown  
430 University as well as, using the UK Biobank Resource under Application Number 22419. The Population  
431 Architecture Using Genomics and Epidemiology (PAGE) program is funded by the National Human Genome  
432 Research Institute (NHGRI) with co-funding from the National Institute on Minority Health and Health Dis-  
433 parities (NIMHD). The WHI program is funded by the National Heart, Lung, and Blood Institute, National  
434 Institutes of Health, U.S. Department of Health and Human Services through contracts 75N92021D00001,  
435 75N92021D00002, 75N92021D00003, 75N92021D00004, 75N92021D00005. The HCHS/SOL study was car-  
436 ried out as a collaborative study supported by contracts from the National Heart, Lung and Blood Institute  
437 (NHLBI) to the University of North Carolina (N01-HC65233), University of Miami (N01-HC65234), Albert

438 Einstein College of Medicine (N01-HC65235), Northwestern University (N01-HC65236) and San Diego State  
439 University (N01-HC65237). S.P.S. is a trainee supported under the Brown University Predoctoral Training  
440 Program in Biological Data Science (NIH T32 GM128596). L.C. acknowledges the support of an Alfred  
441 P. Sloan Research Fellowship and a David & Lucile Packard Fellowship for Science and Engineering. This  
442 work was also supported by US National Institutes of Health R01 GM118652 to S.R., and S.R. acknowledges  
443 additional support from National Science Foundation CAREER Award DBI-1452622.

## 444 **Web Resources**

445 The methods applied in this paper include: PESCA (<https://github.com/huwenboshi/pesca>), RSS(<https://github.com/stephenslab/rss>), gene- $\epsilon$  (<https://github.com/ramachandran-lab/geneepsilon>), and Hierar-  
446 chical HotNet (<https://github.com/raphael-group/hierarchical-hotnet>).

## 448 **Data and code availability**

449 All scripts, publicly available data, and outputs from GWA, gene, and pathway association tests are avail-  
450 able at [https://github.com/smithsap/redefining\\_replication](https://github.com/smithsap/redefining_replication). Results from PESCA analyses were  
451 provided through personal correspondence with Huwenbo Shi.



## 452 Figures and Tables

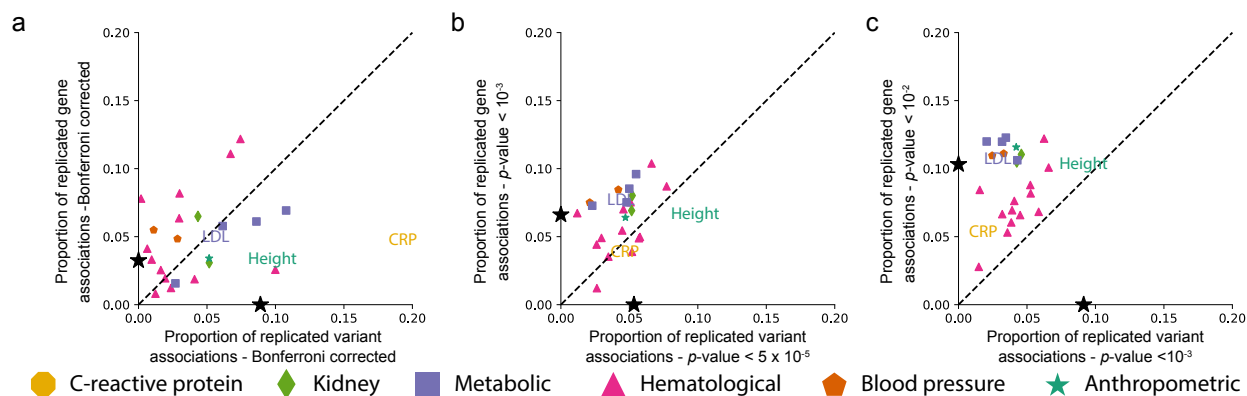
Genomic Scale	Association Test	Model of Genetic Trait Architecture	Relevant Example
SNPs	Standard univariate genome-wide association (GWA) test	The true mutation-level trait architecture is the same for all individuals.	Many inflammatory bowel disease mutations replicate across ancestries <sup>49</sup> .
SNP-Sets/Genes	Gene-level association tests (e.g., gene- $\epsilon$ <sup>38</sup> , SKAT <sup>40</sup> )	Core genes are the same across all ancestries, with potentially varying causal SNPs.	Late-onset Alzheimer's disease risk from <i>ApoE4</i> allele is lower in African ancestry individuals <sup>89</sup> .
Pathways/Networks	Pathway analysis and network propagation (e.g., Hierarchical HotNet <sup>31</sup> , RSS <sup>60</sup> ).	Core genes differ across ancestries, but are all in the same annotated pathway.	Skin pigmentation architecture in the same pathway differs between African and European ancestry individuals <sup>2</sup> .

**Table 1: The three genomic scales and corresponding association tests used in this study.**

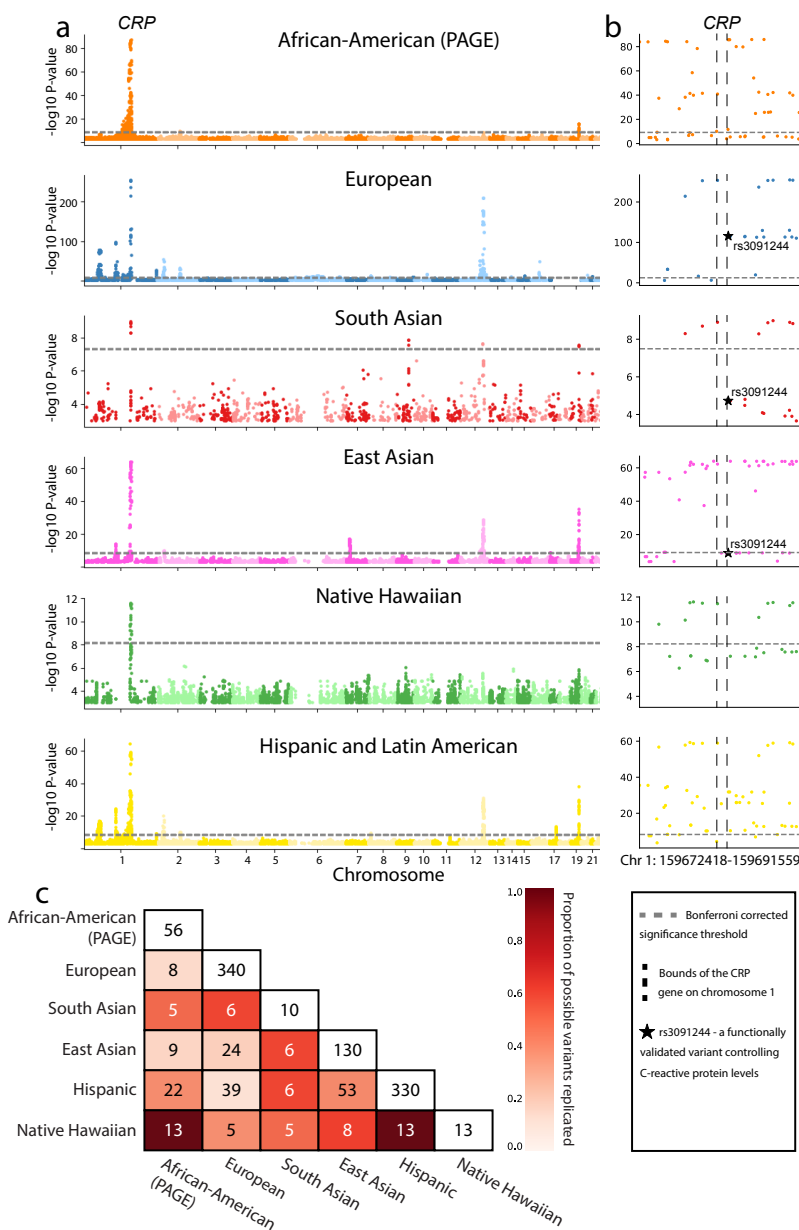
The models of genetic trait architecture corresponding to each genomic scale and statistical method that have been previously invoked in the literature (including relevant examples cited in the last column). These nested genomic scales should routinely be leveraged in multi-ancestry GWA studies to generate biologically interpretable hypotheses of trait architecture across ancestries.

Trait	Number of significant SNPs in at least the European or East Asian cohort	Number of SNPs with same direction of effect	Percentage of SNPs with same direction of effect
BMI	6,374	4,456	69.91
Basophil	884	520	58.82
CRP	694	463	66.71
Cholesterol	3,451	2,379	68.94
DBP	1,962	1,320	67.28
EGFR	7,129	5,038	70.67
Eosinophil	5,890	3,608	61.26
HBA1C	3,030	2,206	72.81
HDL	5,995	4,077	68.01
Height	33,577	22,090	65.79
Hematocrit	5,382	3,602	66.93
Hemoglobin	5,280	3,601	68.20
LDL	2,521	1,929	76.56
Lymphocyte	2,214	1,237	55.87
MCH	6,763	4,674	69.11
MCHC	1,760	1,124	63.86
MCV	7,489	5,208	69.54
Monocyte	3,929	2,565	65.28
Neutrophil	5,431	3,850	70.89
PLC	11,014	7,161	65.02
RBC	9,211	6,263	67.99
SBP	1,807	1,182	65.41
Triglyceride	3,743	2,686	71.76
WBC	6,017	4,105	68.22
Urate	5,864	3,787	64.58

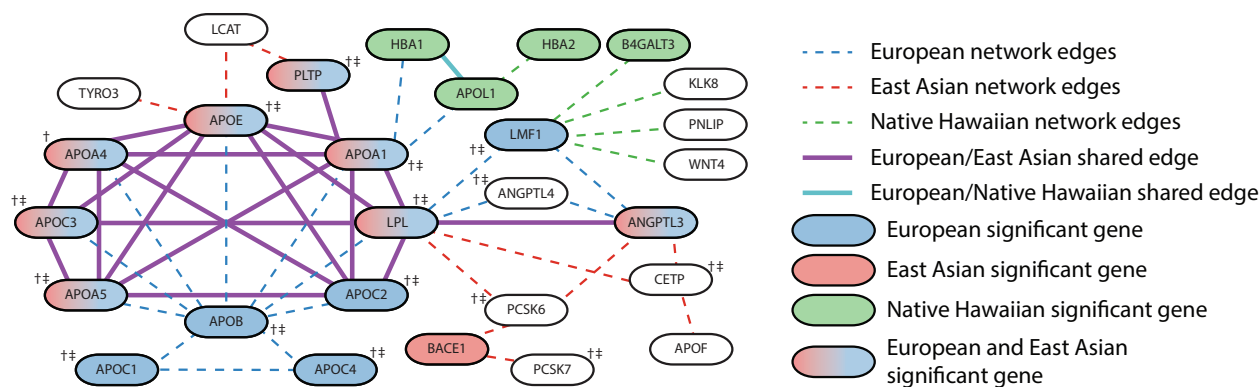
**Table 2: Effect size homogeneity in variants identified as significant in the European or East Asian cohorts.** In each of the 25 traits analyzed in this study a majority of variants that are significant in at least the European or East Asian cohorts had the same direction of effect in the other ancestry cohort.



**Figure 1: Less stringent significance thresholds lead to a decrease in the proportion of replicated SNP-level associations and an increase in the proportion of gene-level associations among ancestries for each of the 25 traits analyzed.** **a.** Proportion of all SNP-level Bonferroni-corrected genome-wide significant associations in any ancestry that replicate in at least one other ancestry is shown on the x-axis (see Table S10 for ancestry-trait specific Bonferroni corrected  $p$ -value thresholds). On the y-axis we show the proportion of significant gene-level associations that were replicated for a given phenotype in at least two ancestries (see Table S15 for Bonferroni corrected significance thresholds for each ancestry-trait pair). The black stars on the x- and y-axes represent the mean proportion of replicates in SNP and gene analyses, respectively. C-reactive protein (CRP) contains the greatest proportion of replicated SNP-level associations of any of the phenotypes. **b.** The x-axis indicates the proportion of SNP-level associations that surpass a nominal threshold of  $p$ -value  $< 10^{-5}$  in at least one ancestry cohort that replicate in at least one other ancestry cohort. The y-axis indicates the proportion of gene-level associations that surpass a nominal threshold of  $p$ -value  $< 10^{-3}$  in at least one ancestry cohort and replicate in at least one other ancestry cohort. Nominal  $p$ -value thresholds tend to decrease the proportion of replicated SNP-level associations and tend to increase the proportion of replicated gene-level associations. The number of unique SNPs and genes that replicated in each cohort is given in Figure S15. **c.** The x-axis indicates the proportion of SNP-level associations that surpass a nominal threshold of  $p$ -value  $< 10^{-3}$  in at least one ancestry cohort that replicate in at least one other ancestry cohort. The y-axis indicates the proportion of gene-level associations that surpass a nominal threshold of  $p$ -value  $< 10^{-2}$  in at least one ancestry cohort and replicate in at least one other ancestry cohort. The number of unique SNPs and genes that replicated in each cohort is given in Figure S16. As shown in panel **b**, nominal  $p$ -value thresholds tend to decrease the proportion of replicated SNP-level associations and tend to increase the proportion of replicated gene-level associations. Expansion of three letter trait codes are given in Table S2 and a version of this plot with all trait names displayed as text is shown in Figure S14. Figure S14 shows the same set of plots with all traits represented as text.

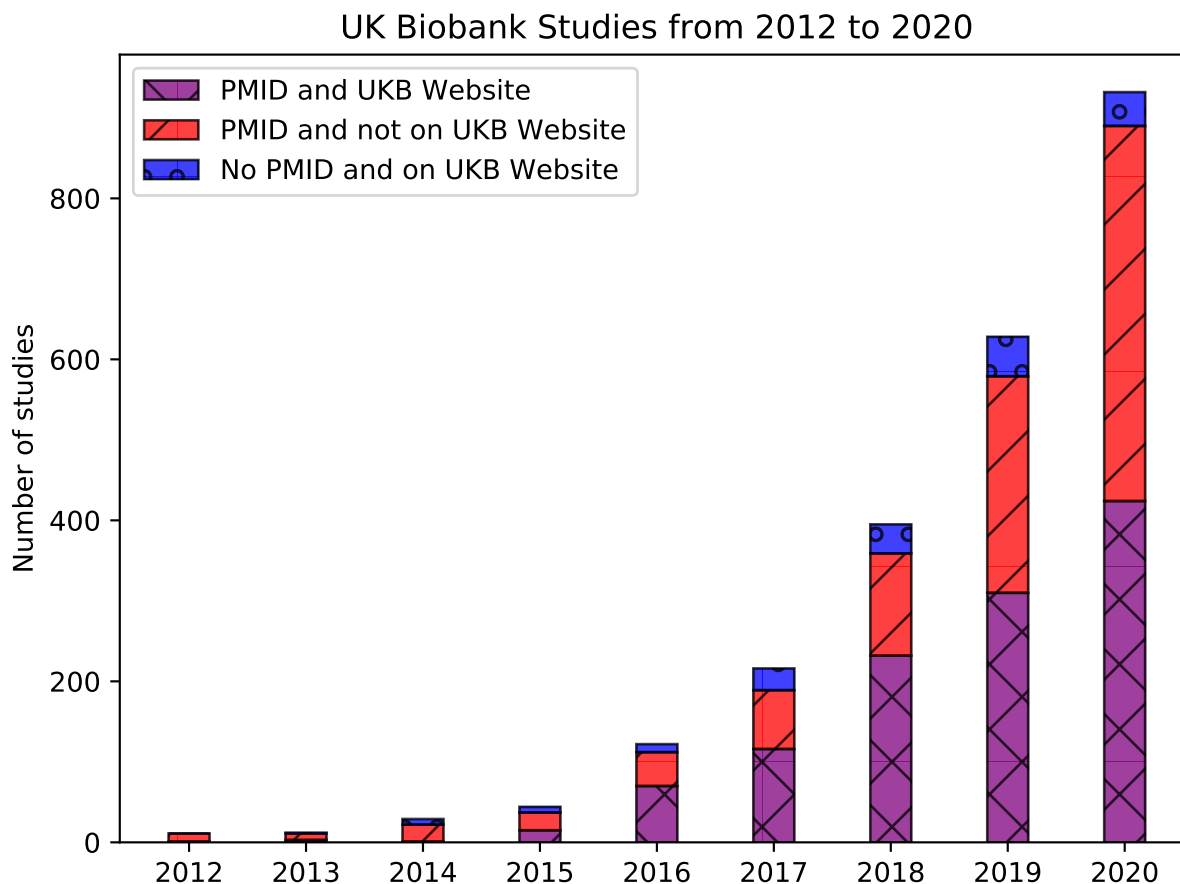


**Figure 2: C-reactive protein is an exceptional trait where standard GWA analyses may be sufficient to identify shared genetic architecture among ancestry cohorts.** **a.** Manhattan plot for SNP-level associations with C-reactive protein levels. Ancestry-specific Bonferroni-corrected significance thresholds are shown with dashed horizontal grey lines and listed in Table S10. Note that the scale of the  $-\log_{10}$ -transformed  $p$ -values on the y-axis is different for each ancestry for clarity. **b.** Manhattan plot of SNP-level associations around the *CRP* gene located on chromosome 1 for each ancestry (zoomed in from panel **a**). Boundaries of the *CRP* gene are shown with vertical dashed black lines. All six ancestries contain genome-wide significant SNPs in the region. Black stars in the European, South Asian, and East Asian plots represent rs3091244, a SNP that has been functionally validated as contributing to serum levels of C-reactive protein<sup>52,53</sup>. **c.** Heatmap of Bonferroni-corrected significant genotyped SNPs replicated between each pair of ancestries analyzed. Here, we focus on SNPs in the 1MB region surrounding the *CRP* gene. Entries along the diagonal represent the total number of SNP-level associations in the 1MB region surrounding the *CRP* gene for each ancestry. The color of each cell is proportional to the percentage of SNP-level associations replicated out of all possible replications in each ancestry pair (i.e., the minimum of the diagonal entries between the pairs being considered). For example, the maximum number of genome-wide significant SNPs that can possibly replicate between the Hispanic and East Asian is 25, and 20 replicate resulting in the cell color denoting 80% replication. A similar matrix, computed including imputed SNPs is shown in Figure S17.

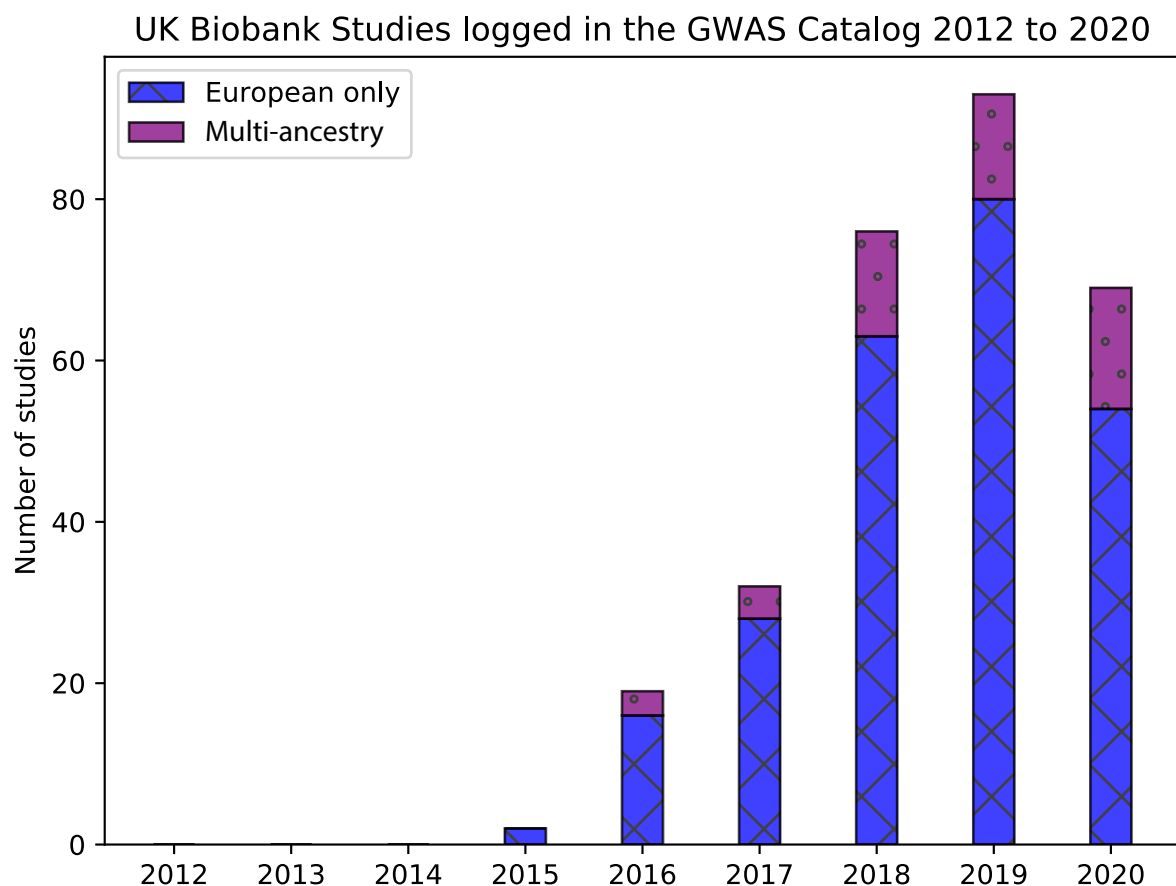


**Figure 3: A subnetwork of apolipoprotein genes is significantly enriched for mutations in European, East Asian, and Native Hawaiian ancestries associated with triglyceride levels.** The largest significantly altered subnetwork ( $p$ -value  $< 0.05$ ) for triglyceride levels contains overlapping gene subnetworks for each of the European, East Asian, and Native Hawaiian ancestries when analyzed independently with Hierarchical HotNet<sup>31</sup>. Each node in the network represents a gene. The shading of each node indicates the statistical significance of the association of that gene with triglyceride levels in a particular cohort. Two genes are connected if their protein products interact based on the ReactomeFI 2016<sup>44</sup> (European, East Asian) or iRefIndex 15.0<sup>45</sup> (Native Hawaiian) protein-protein interaction networks. Several genes from the apolipoprotein gene family are significantly associated with triglyceride levels in both the European and East Asian cohorts (see Data Availability). Additionally, the interactions between them form a highly connected subnetwork. Smaller subnetworks identified in the Native Hawaiian cohort are distal modules that are connected to the subnetwork detected in the European cohort. Not all genes in the largest significantly altered subnetwork for the Native Hawaiian ancestry group are shown for visualization purposes (127 not pictured here). Genes that contain SNPs previously associated to triglyceride levels in a European cohort in the GWAS catalog are indicated with †. Similarly, genes that contain SNPs previously associated with triglyceride levels in a non-European cohort in the GWAS Catalog are indicated with ‡. The studies identifying these associations are given in Table S16.

453 Supplemental Figures

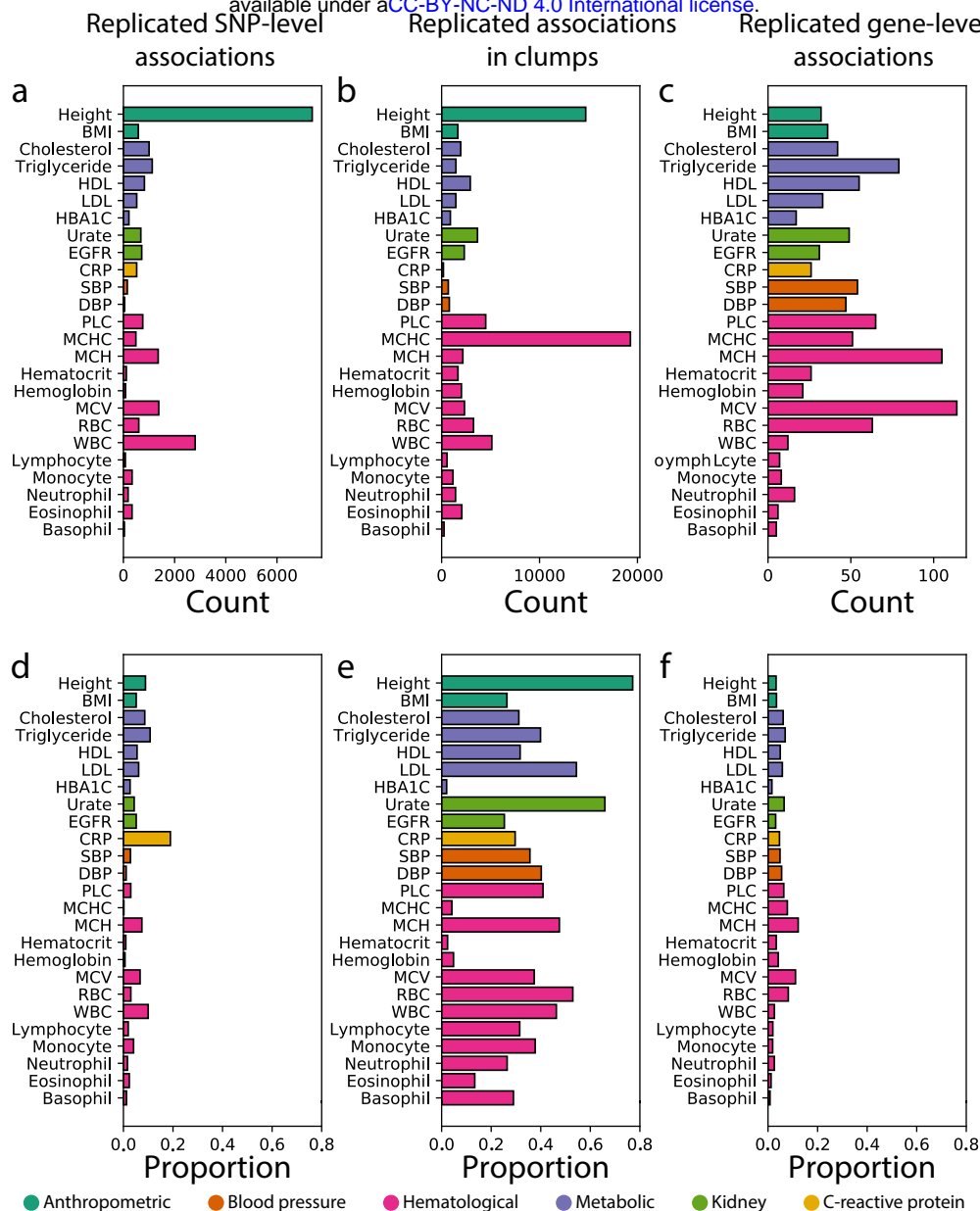


**Figure S1: Number of publications we identified using UK Biobank data from 2012 to 2020.** Studies identified using PMIDs as described in the Supplemental Information. Studies that are displayed on the UK Biobank website (<https://www.ukbiobank.ac.uk/>) and identified on PubMed are shown in purple. Studies listed on the UK Biobank website but do not have a PMID are shown in blue, and studies only identified using PubMed but not listed on the UK Biobank website are shown in red. The protocols for identifying studies both on PubMed and the UK Biobank website are detailed in the Supplemental Information. Data from both the UK Biobank website and PubMed were accessed on January 12, 2021.

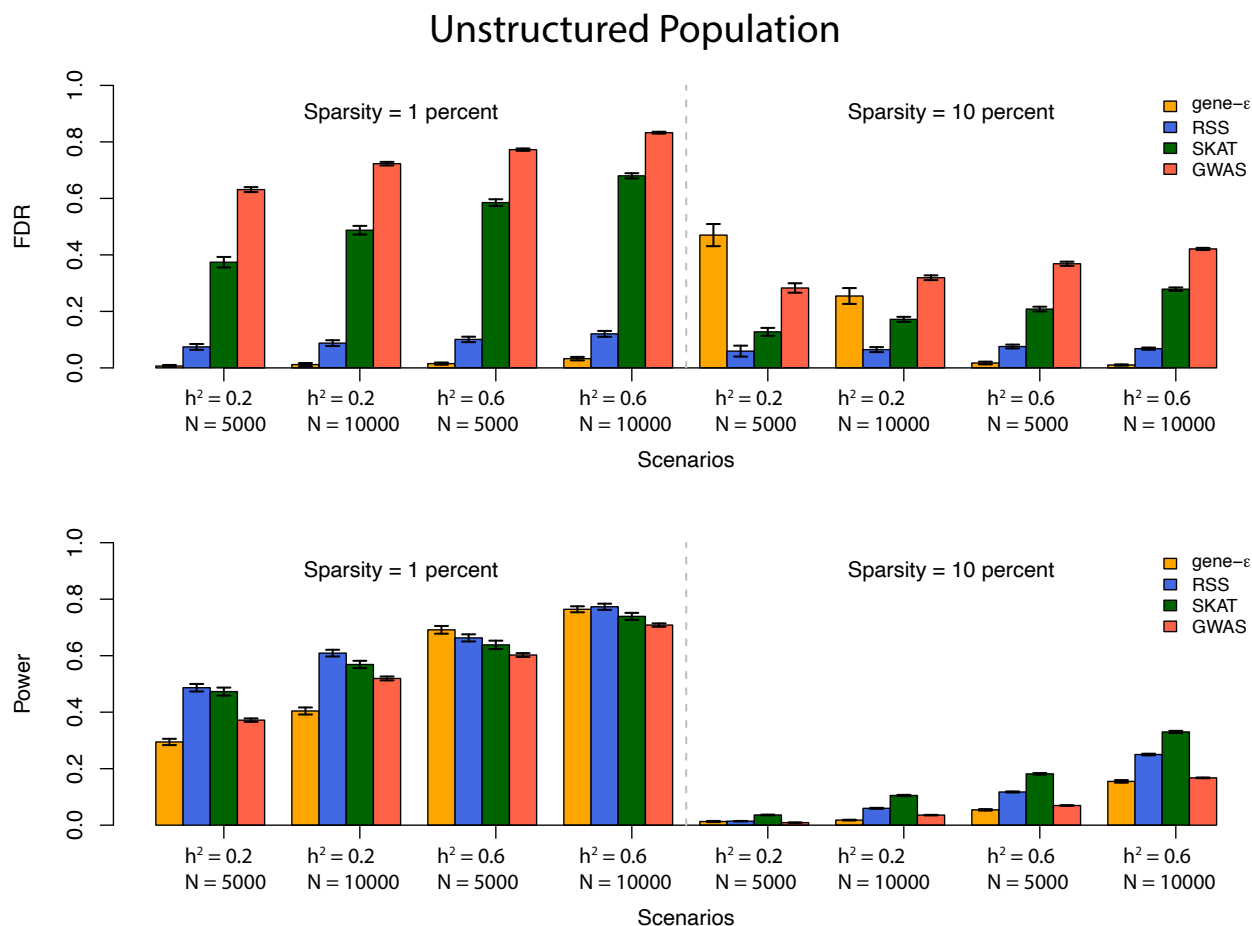


**Figure S2: Number of studies published using UK Biobank data from 2012 to 2020 that have available metadata in the GWAS Catalog.** Our protocols for identifying studies from the GWAS Catalog are detailed in the Supplemental Information. Multi-ancestry studies are shown in purple and include those that list samples of more than one ancestral group in the GWAS catalog (as defined according to the protocol using Popejoy and Fullerton<sup>25</sup>, available on the GitHub page [https://github.com/ramachandran-lab/redefining\\_replication](https://github.com/ramachandran-lab/redefining_replication)). Studies that only list samples of European ancestry in the GWAS catalog are shown in blue. Every multi-ancestry analysis includes samples of European ancestry and of at least one other ancestry. GWAS Catalog data was accessed on January 10, 2021 from the website <https://www.ebi.ac.uk/gwas/docs/file-downloads> using the final release file of 2020 (see file named `gwas_catalog.v1.0.2-associations_e100_r2020-12-15.tsv`).



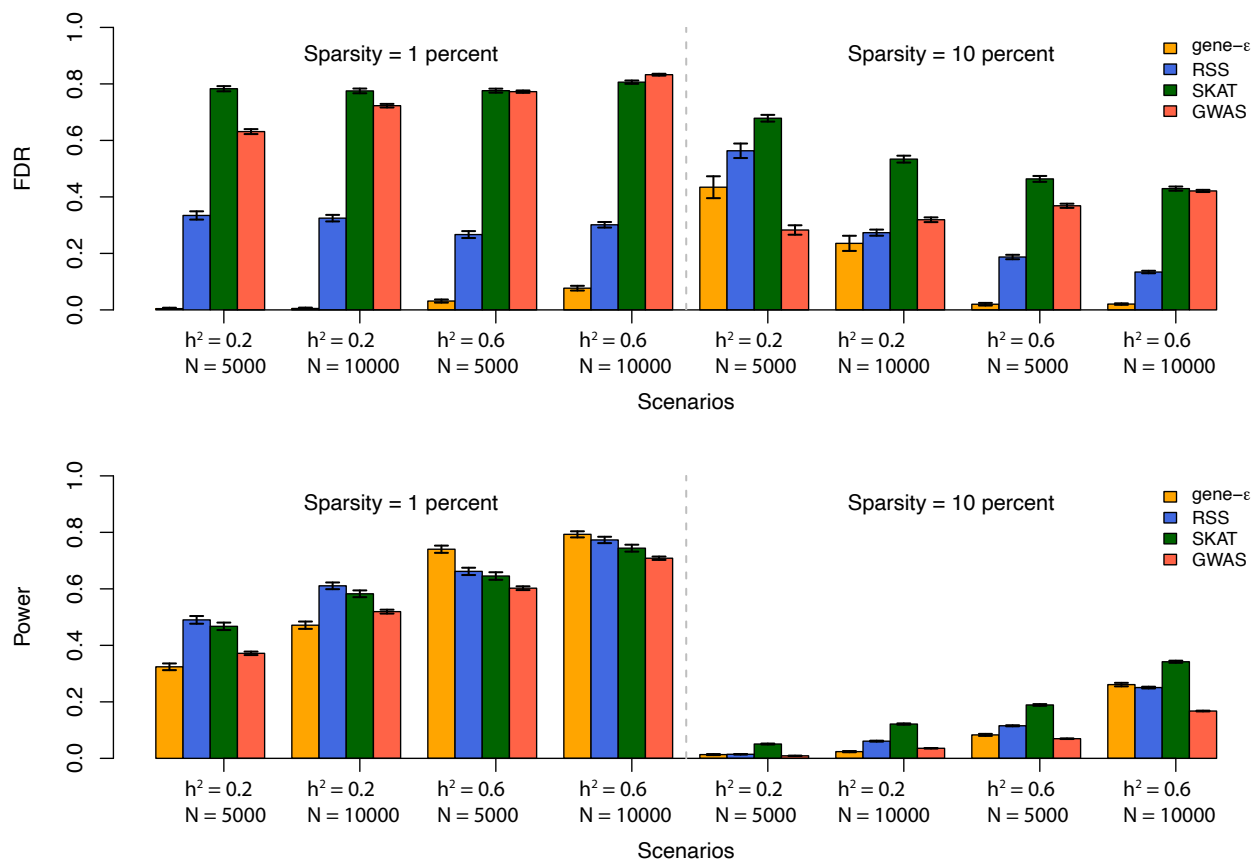


**Figure S3: Summaries of replicated associations at multiple genomic scales among ancestry cohorts for all 25 traits analyzed using Bonferroni-corrected thresholds.** Expansion of three letter trait codes are given in Table S2. (a) Number and (d) proportion of genome-wide significant SNPs associated with a phenotype in at least one ancestry cohort that were replicated in at least two ancestry cohorts. In all 25 traits, genome-wide significant SNPs replicate in at least two ancestry cohorts. Height contains over 7,000 replicated SNPs among the seven ancestry cohorts analyzed, illustrating its highly polygenic architecture. For many traits across all categories, with the exception of other biochemical (i.e., CRP), the replication rate of genes is higher in gene-level associations than at the SNP-level. (b) Number and (e) proportion of 1Mb windows, or “clumps”, that contain at least one genome-wide significant SNP-level associations for a given phenotype in at least two ancestry cohorts. (c) Number and (f) proportion of genome-wide significant gene-level associations that replicate among ancestry cohorts. Replicated associations in hematological are common at the gene-level in hematological and metabolic traits. For instance, in three of the four cohorts with mean corpuscular hemoglobin (MCH) measurements *HBA1* and *HBA2* were identified as significant associated with MCH in the African, European, and East Asian ancestry cohorts Table S3. The denominator of the proportion is calculated as the total number of unique SNPs, clumps, or genes that are significantly associated with a trait in at least on ancestral cohort. Note that d and f correspond to Figure 1a and b, with an altered x-axis upper limit of 0.8 in this figure.

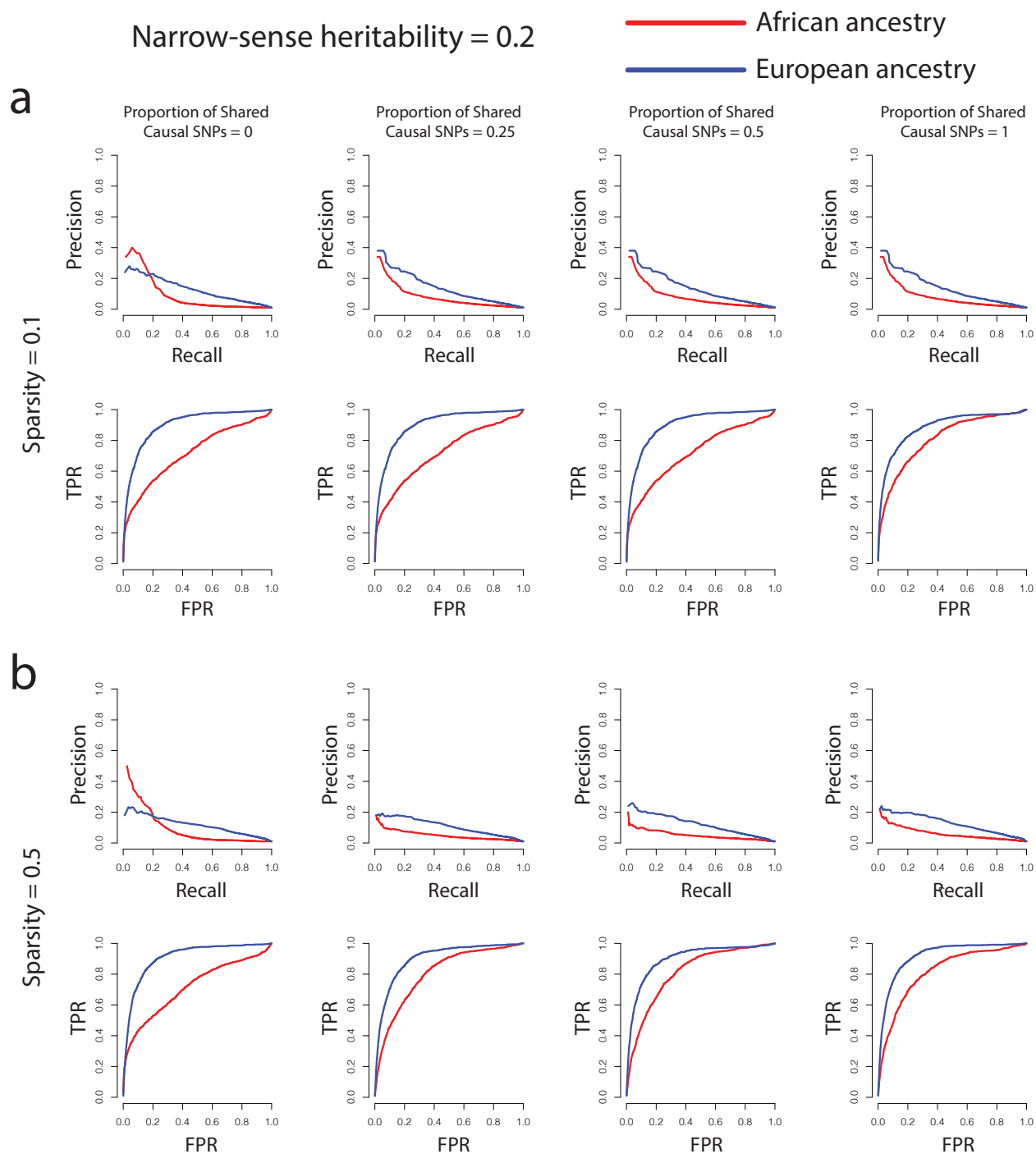


**Figure S4: gene- $\epsilon$  outperforms and controls false discovery rate (FDR) better than other association methods in simulations with varying heritability and sample size.** Simulations were designed to assess gene versus SNP-level association false discovery rate (FDR) and power in an unstructured population as described by the protocols in the Supplemental Information. The top and bottom panels show the FDR and power of four different association methods on 100 simulated datasets, respectively. We compared performance of three gene-level association test methods (gene- $\epsilon$ <sup>38</sup>, RSS<sup>60</sup>, SKAT<sup>40</sup>) with outputs from the standard GWA association test under different simulation parameters (sample size  $N$ , narrow-sense heritability  $h^2$ , and sparsity). We define sparsity as the proportion of SNPs that are ground-truth causal. Standard errors across the simulated replicates are shown using black whisker plots. Simulation protocol is described in the Supplemental Information.

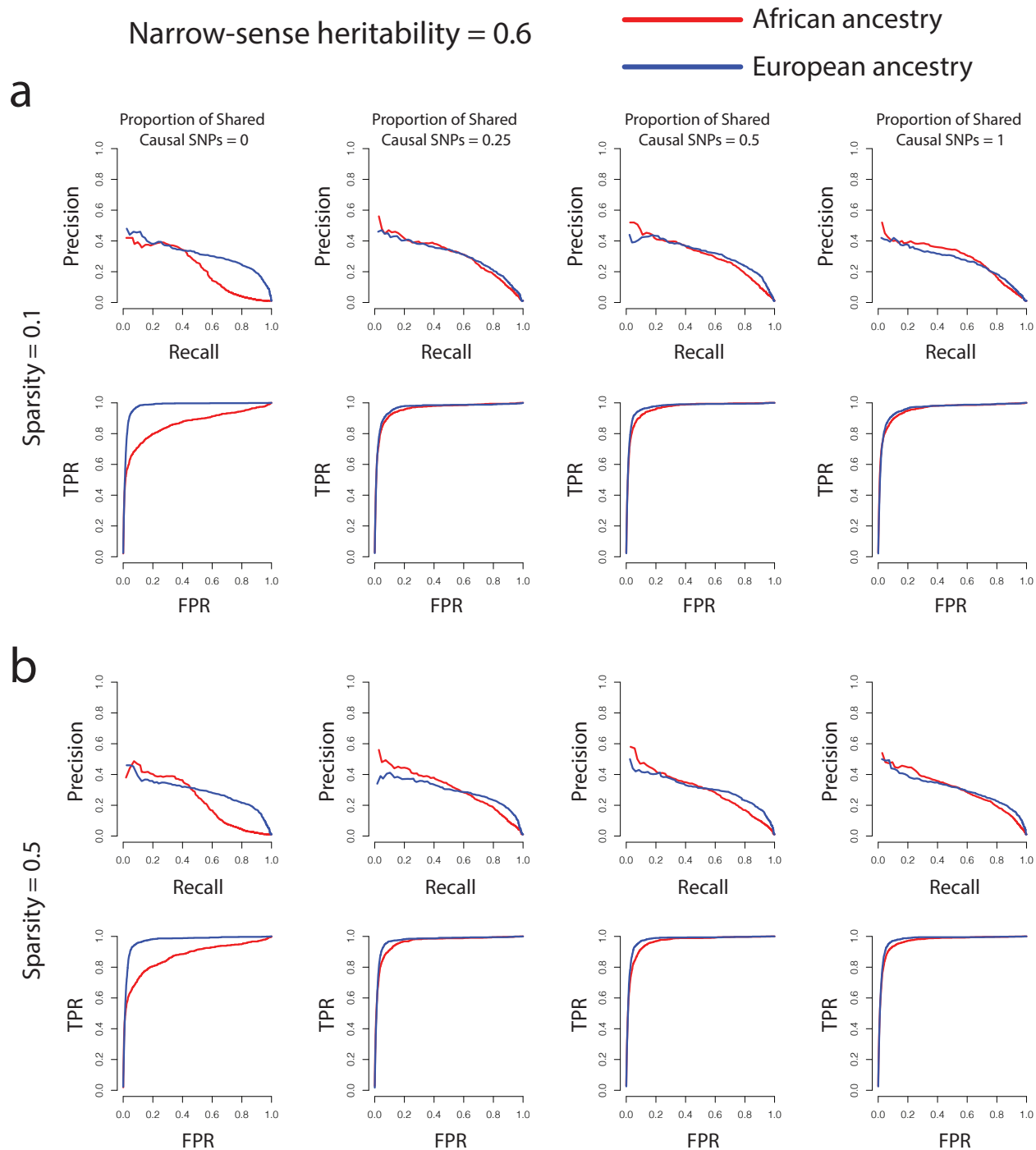
## Structured Population



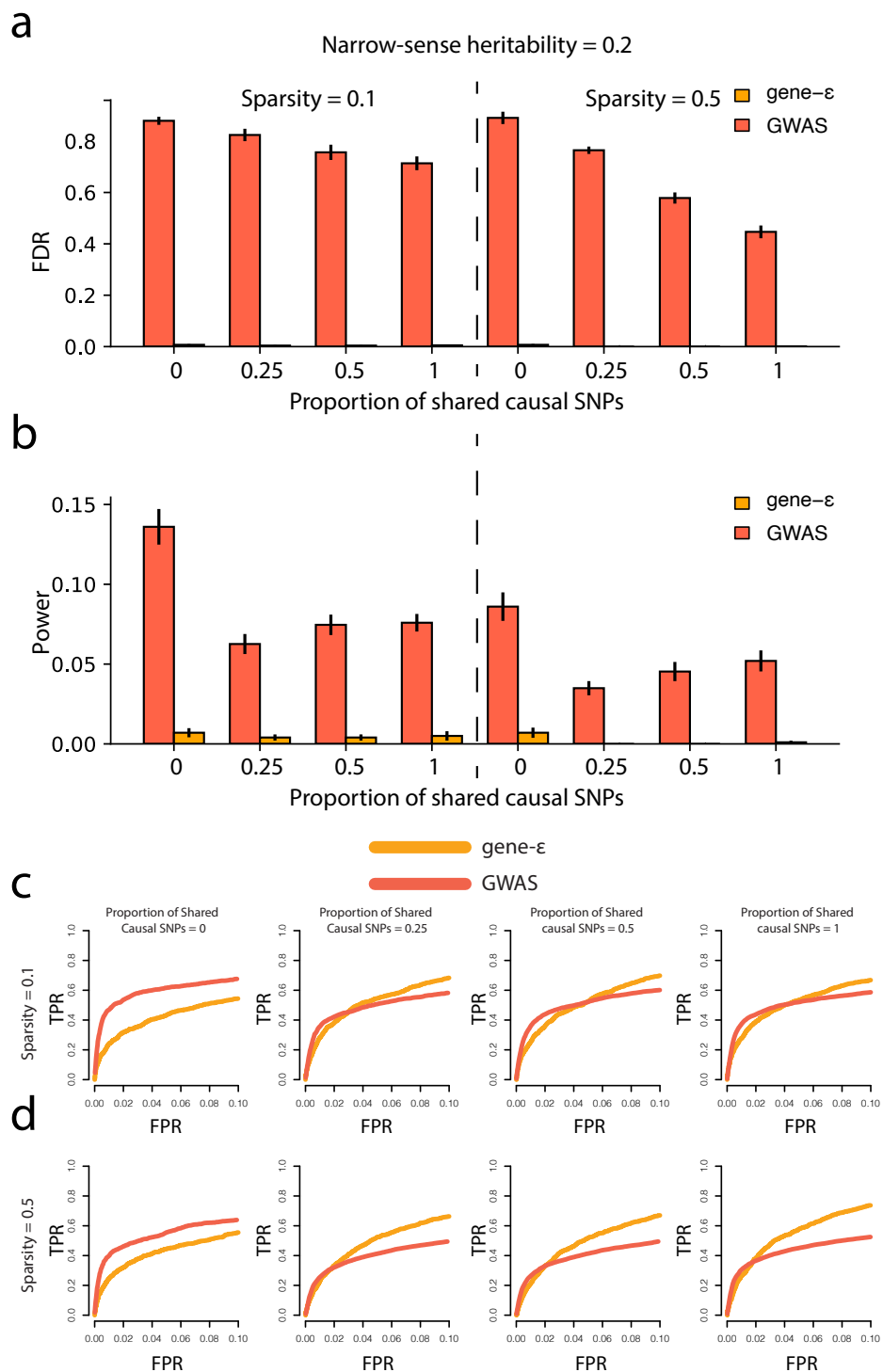
**Figure S5:  $gene-\epsilon$  outperforms and controls false discovery rate (FDR) better than other association methods in simulations with varying heritability and sample size.** Simulations are designed to assess gene versus SNP-level association false discovery rate (FDR) and power in an structured population as described by the protocols in the Supplemental Information. The top and bottom panels show the FDR and power of four different association methods on 100 simulated datasets, respectively. We compared performance of three gene-level association test methods ( $gene-\epsilon$ <sup>38</sup>, RSS<sup>60</sup>, SKAT<sup>40</sup>) with outputs from the standard GWA association test under different simulation parameters (sample size  $N$ , narrow-sense heritability  $h^2$ , and sparsity). We define sparsity as the proportion of SNPs that are designated to be causal. Standard errors across the simulated replicates are shown using black whisker plots.



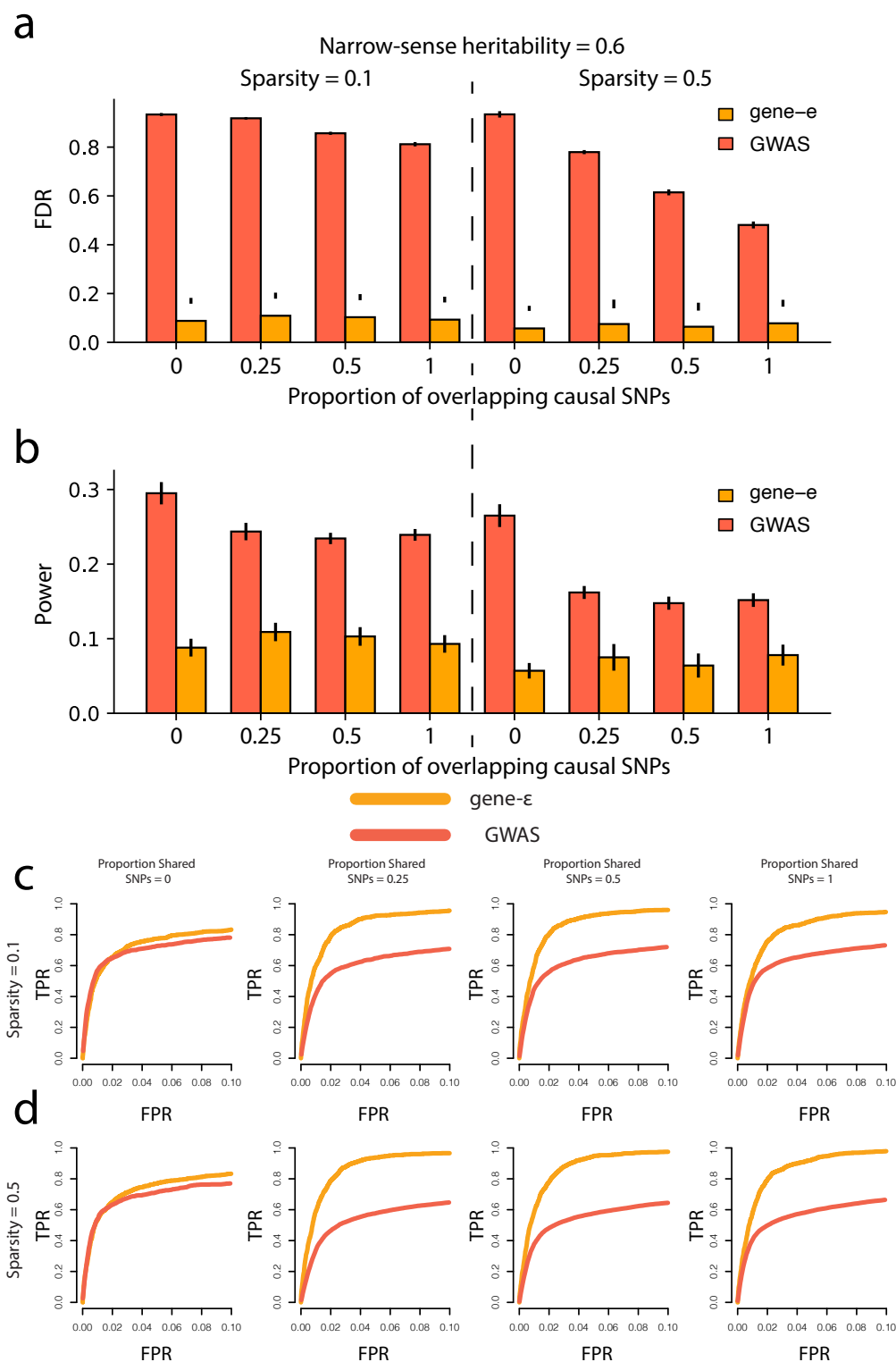
**Figure S6: gene- $\epsilon$  identifies associated genes in two simulated ancestry cohorts under a variety of genetic architectures with low narrow sense heritability.** **a.** Precision-recall (top row) and receiver operating curves (bottom row) for gene- $\epsilon$  analysis of cohorts simulated using genotypes from individuals of European ( $N = 10,000$ ; blue line) and African ( $N = 4,967$ ; red line) ancestry, respectively. Narrow-sense heritability was set to  $h^2 = 0.2$  in each simulation. Sparsity of causal SNPs was set to a proportion of 0.1 and the proportion of causal SNPs shared was tested at different values. 50 replicates of each set of simulations under each parameter were performed. **b.** Precision-recall (top row) and receiver operating curves (bottom row) for gene- $\epsilon$  analysis of 50 replicated simulations of a European and African cohort using a causal SNP sparsity of 0.5.



**Figure S7: gene- $\epsilon$  identifies associated genes in two simulated ancestry cohorts under a variety of genetic architectures with high narrow sense heritability. a.** Precision-recall (top row) and receiver operating curves (bottom row) for gene- $\epsilon$  analysis of cohorts simulated using genotypes from individuals of European ( $N = 10,000$ ; blue line) and African ( $N = 4,967$ ; red line) ancestry, respectively. Narrow-sense heritability was set to  $h^2 = 0.6$  in each simulation. Sparsity of causal SNPs was set to a proportion of 0.1 and the proportion of causal SNPs shared was tested at different values. 50 replicates of each set of simulations under each parameter were performed. **b.** Precision-recall (top row) and receiver operating curves (bottom row) for gene- $\epsilon$  analysis of 50 replicated simulations of a European and African cohort using a causal SNP sparsity of 0.5.

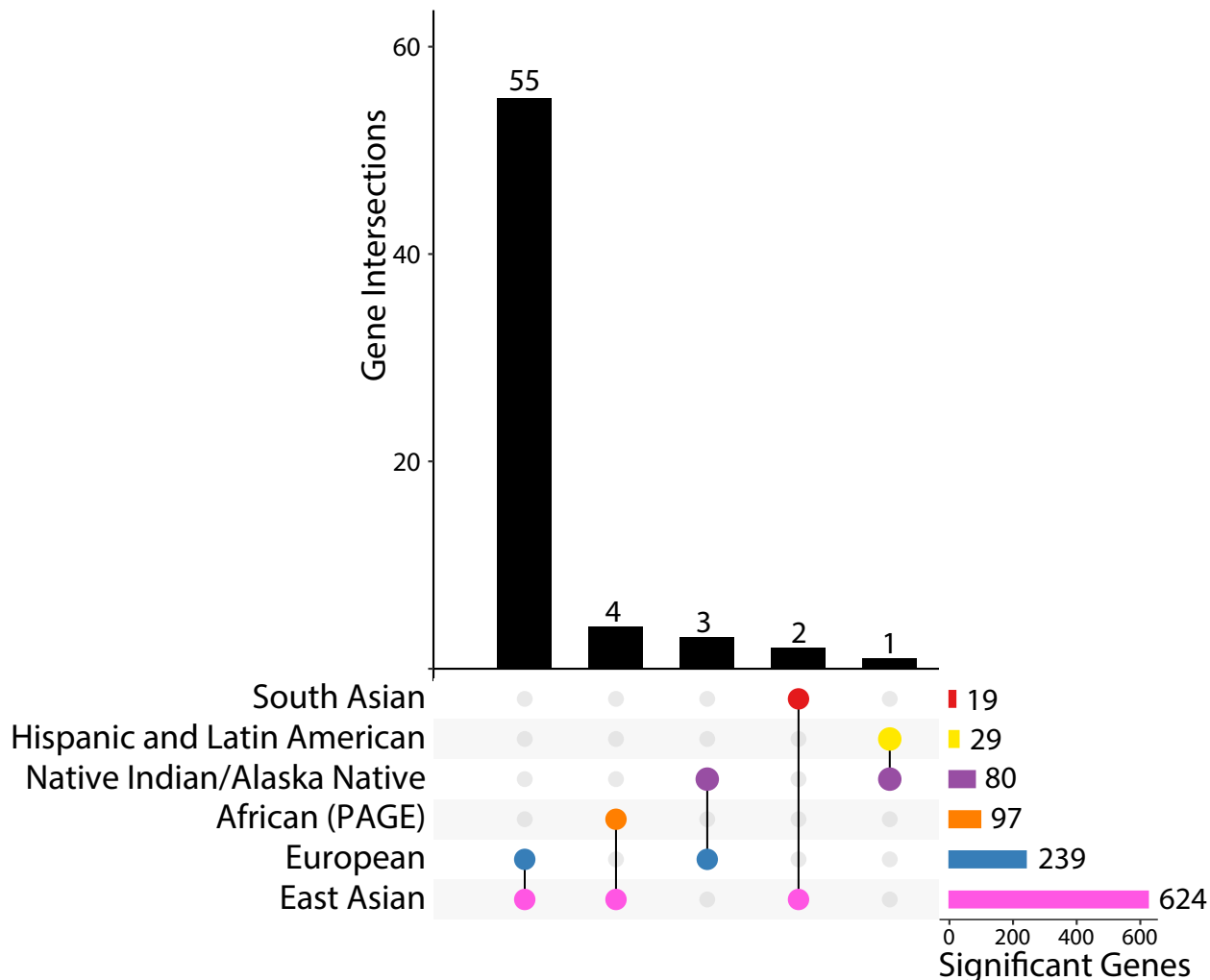


**Figure S8: gene- $\epsilon$  (orange) has a lower false discovery rate for identification of shared genetic determinants between cohorts than the standard GWA framework (red).** Narrow-sense heritability (percent variance explained by the genotype matrix) was set to  $h^2 = 0.2$  for all simulations. **a.** False discovery rate of shared genetic determinants between two ancestry cohorts using varying levels of causal SNP sparsity and proportion of shared causal SNPs between the cohorts. **b.** Power of gene- $\epsilon$  and the standard GWA framework to detect shared genetic determinants between two cohorts. Error bars were calculated using the results from 50 simulations of each parameter set of sparsity and proportion of shared causal SNPs for both FDR(a) and Power(b). **c.** Receiver operating curves corresponding to simulations of genetic architecture when causal SNP sparsity is equal to 0.1 (corresponding to the left-hand panels of a and b). **d.** Receiver operating curves corresponding to simulations of genetic architecture when causal SNP sparsity is equal to 0.5 (corresponding to the right-hand panels of a and b).

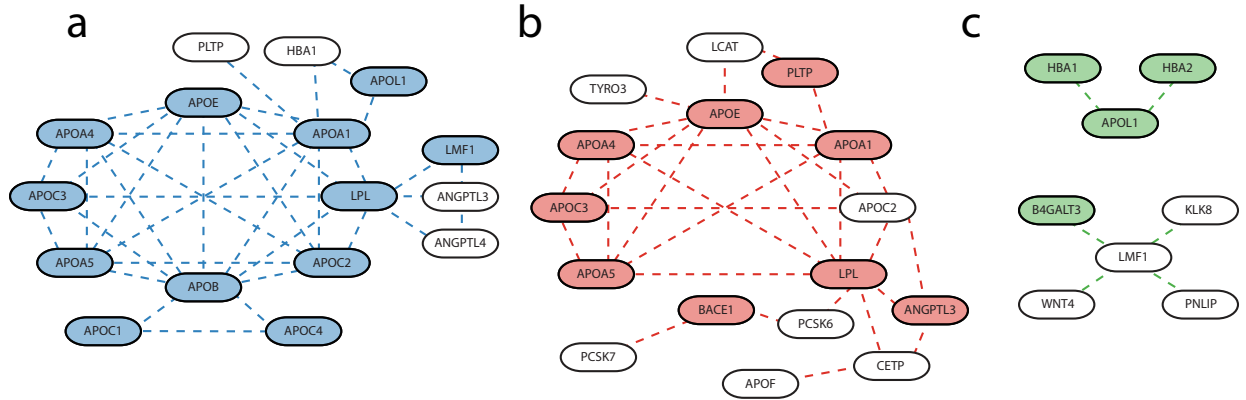


**Figure S9: gene-ε (orange) has a lower false discovery rate for identification of shared genetic determinants between cohorts than the standard GWA framework (red).** Narrow-sense heritability (percent variance explained by the genotype matrix) was set to  $h^2 = 0.6$  for all simulations. **a.** False discovery rate of shared genetic determinants between two ancestry cohorts using varying levels of causal SNP sparsity and proportion of shared causal SNPs between the cohorts. **b.** Power of gene-ε and the standard GWA framework to detect shared genetic determinants between two cohorts. Error bars were calculated using the results from 50 simulations of each parameter set of sparsity and proportion of shared causal SNPs for both FDR(a) and Power(b). **c.** Receiver operating curves corresponding to simulations of genetic architecture when causal SNP sparsity is equal to 0.1 (corresponding to the left-hand panels of **a** and **b**). **d.** Receiver operating curves corresponding to simulations of genetic architecture when causal SNP sparsity is equal to 0.5 (corresponding to the right-hand panels of **a** and **b**).

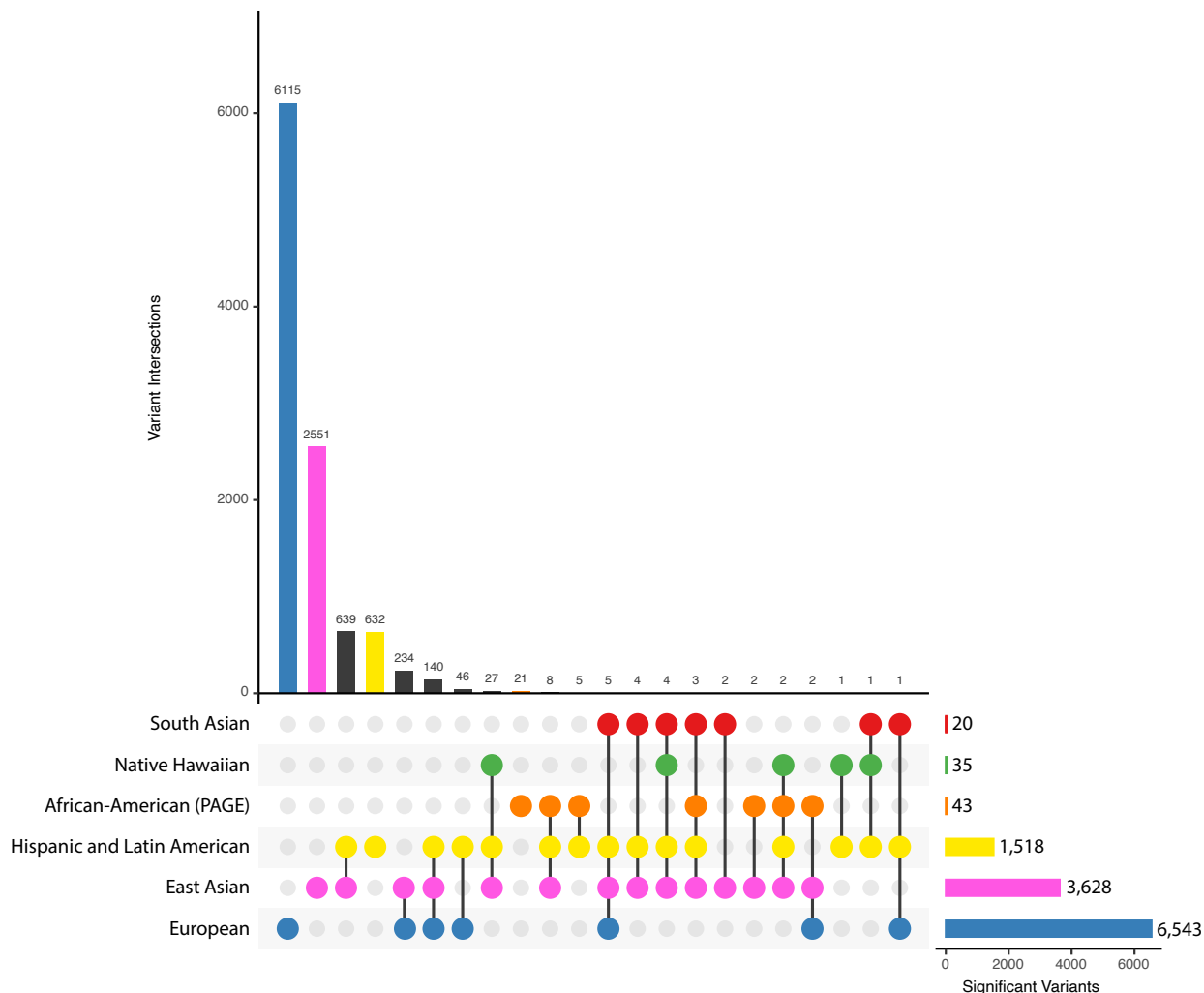




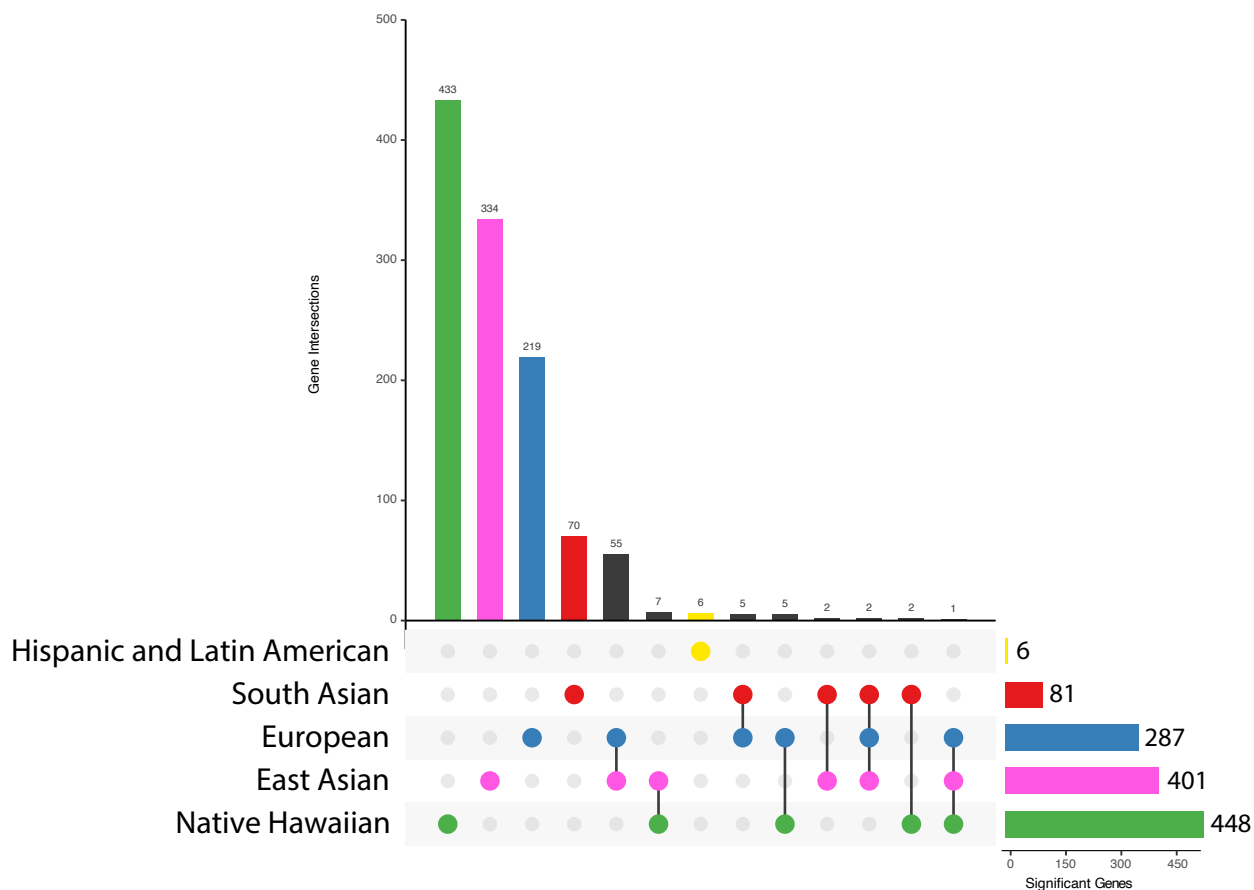
**Figure S10: All six ancestries have gene-level associations with platelet count that replicate in at least one other ancestry.** Total number of genome-wide significant genes in each ancestry, after correcting for total number of regions tested, are given in the bar plot located in the bottom right (significance thresholds are given in Table S15, sample sizes are given in Table S5 - Table S8). Shared gene-level association statistics between pairs of ancestries are shown in the vertical bar plot; the pair of ancestries represented by each bar can be identified using the dots and links below the barplot. Of the 65 genome-wide significant gene-level association statistics that replicate in at least two ancestry cohorts, only 25 contain SNPs that have been previously associated with platelet count in at least one ancestry in at least one study in the GWAS catalog (<https://www.ebi.ac.uk/gwas/home>) This plot was generated using the UpSetR package<sup>90</sup>.



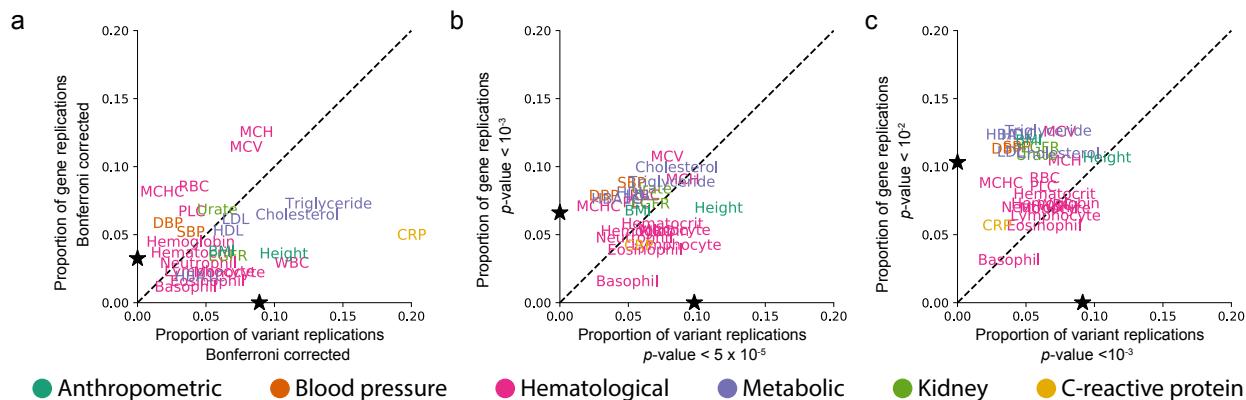
**Figure S11: Significantly mutated subnetworks associated with triglyceride levels identified in the (a) European, (b) East Asian, and (c) Native Hawaiian ancestry cohorts.** Significantly mutated subnetworks were identified using the Hierarchical HotNet method<sup>31</sup>. Genes that were identified in each ancestry as significantly associated with triglyceride levels using the gene- $\epsilon$  method are shaded using ancestry-specific color coding (also used in Figure 3, European—blue, East Asian—pink, Native Hawaiian—green). Significantly mutated subnetworks in the (a) European and (b) East Asian cohorts were identified using the ReactomeFI<sup>44</sup> protein-protein interaction network, and the significantly mutated subnetwork in the (c) the Native Hawaiian cohort was identified using the iRefIndex 15.0<sup>45</sup> protein-protein interaction networks. Genes that are present in any of the significantly mutated subnetworks that contain SNPs previously associated with triglyceride levels in the GWAS Catalog are listed with corresponding citations in Table S16.



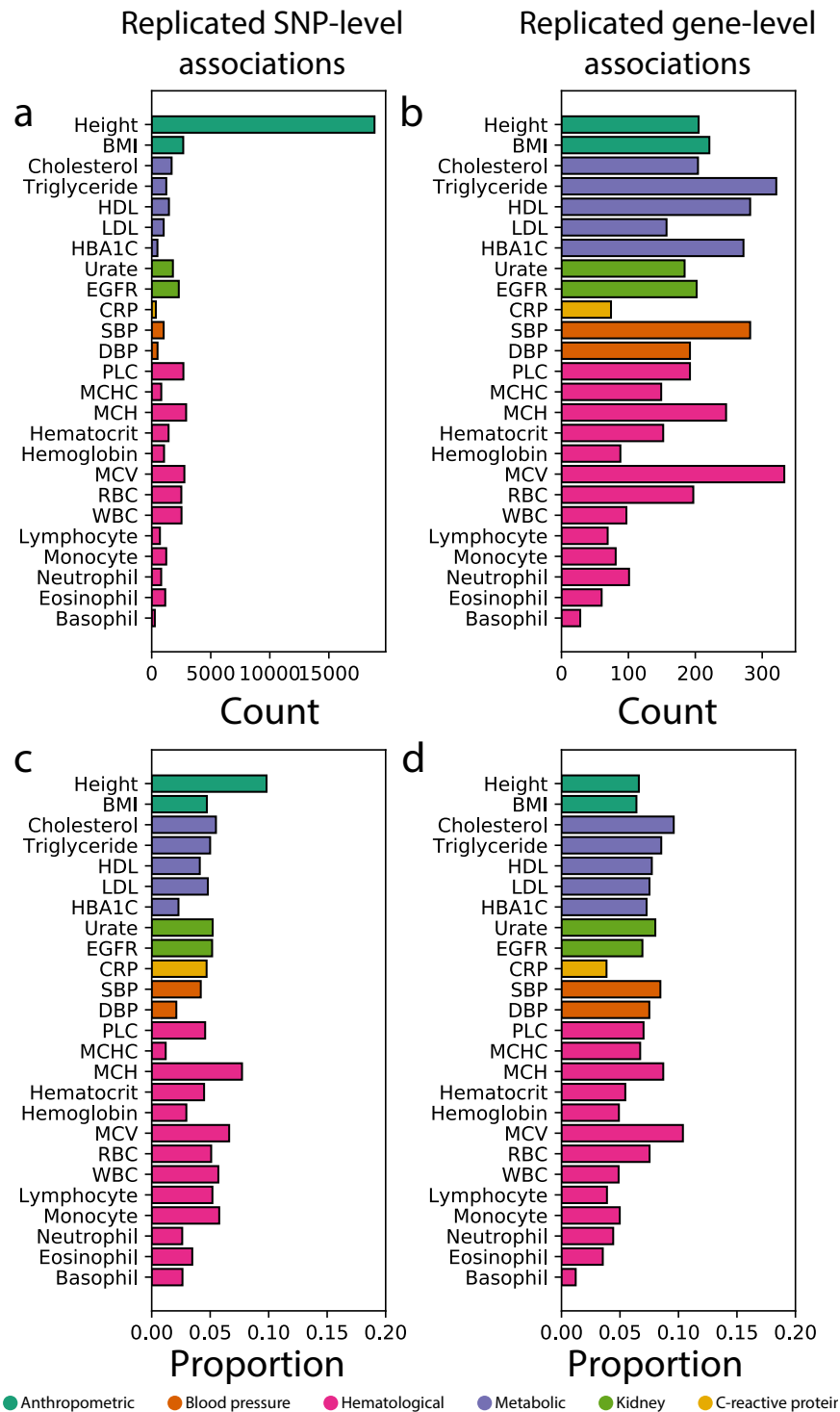
**Figure S12: Shared SNP-level associations with triglyceride levels in six ancestral cohorts.** Total number of genome-wide significant genes in each ancestry, after correcting for total number of regions tested, are given in the bar plot located in the bottom right (significance thresholds are given in Table S10 and sample sizes are given in Table S5 - Table S8). Shared SNP-level association statistics between pairs of ancestries are shown in the vertical bar plot. The pair of ancestries represented by each bar can be identified using the dots and links below the vertical barplot. This plot was generated using the UpSetR package<sup>90</sup>.



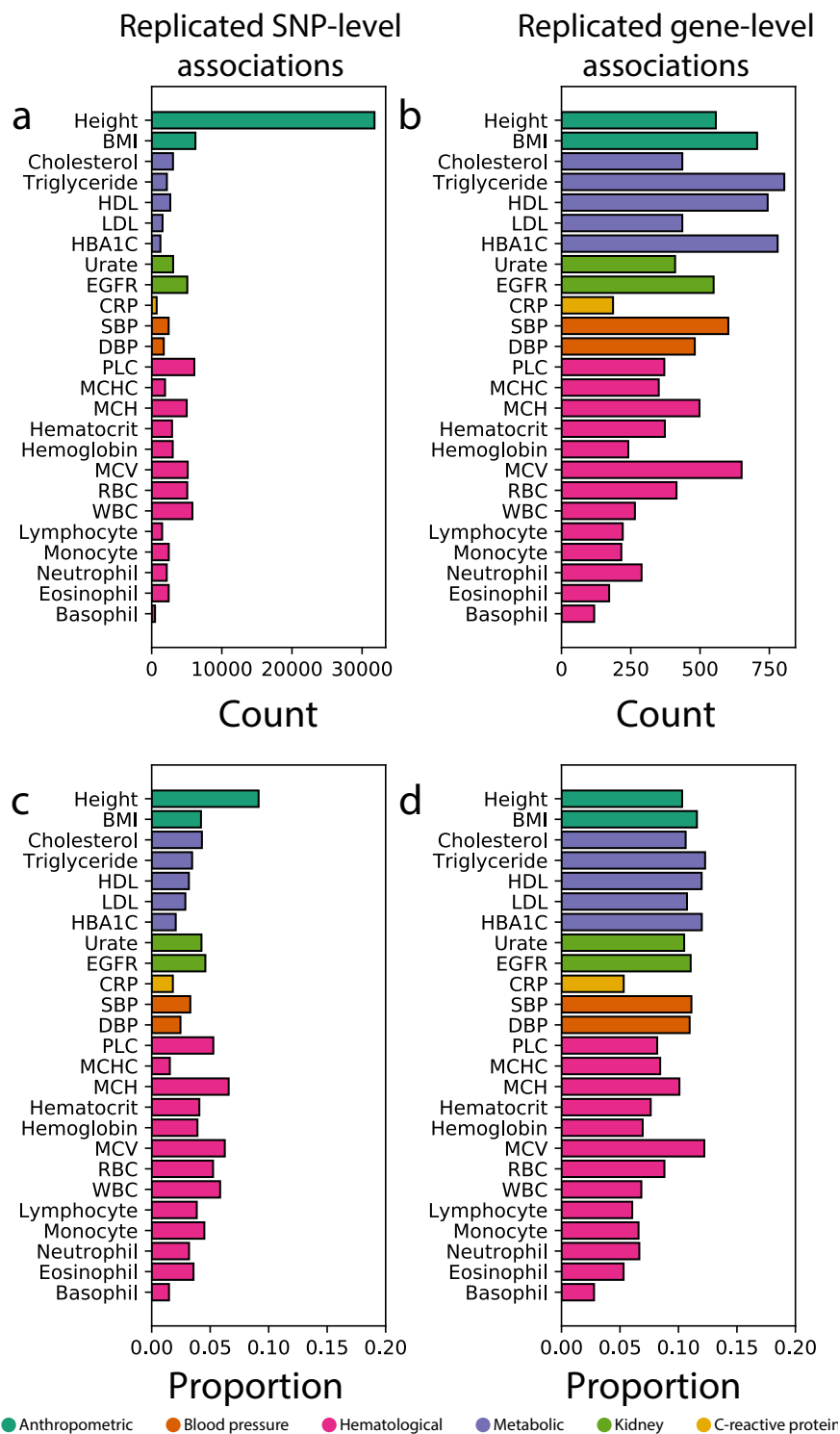
**Figure S13: Shared gene-level associations with triglyceride levels in five ancestral cohorts.** Total number of genome-wide significant genes in each ancestry, after correcting for total number of regions tested, are given in the bar plot located in the bottom right (significance thresholds are given in Table S15 and sample sizes are given in Table S5 - Table S8). Shared gene-level association statistics between pairs of ancestries are shown in the vertical bar plot. The pair of ancestries represented by each bar can be identified using the dots and links below the vertical barplot. This plot was generated using the UpSetR package<sup>90</sup>.



**Figure S14: Less stringent significance thresholds lead to a decrease in the proportion of replicated SNP-level associations and an increase in the proportion of gene-level associations among ancestries for each of the 25 traits analyzed.** **a.** Proportion of all SNP-level Bonferroni-corrected genome-wide significant associations in any ancestry that replicate in at least one other ancestry is shown on the x-axis (see Table S10 for ancestry-trait specific Bonferroni corrected  $p$ -value thresholds). On the y-axis we show the proportion of significant gene-level associations that were replicated for a given phenotype in at least two ancestries (see Table S15 for Bonferroni corrected significance thresholds for each ancestry-trait pair). The black stars on the x- and y-axes represent the mean proportion of replicates in SNP and gene analyses, respectively. C-reactive protein (CRP) contains the greatest proportion of replicated SNP-level associations of any of the phenotypes. **b.** The x-axis indicates the proportion of SNP-level associations that surpass a nominal threshold of  $p$ -value  $< 10^{-5}$  in at least one ancestry cohort that replicate in at least one other ancestry cohort. The y-axis indicates the proportion of gene-level associations that surpass a nominal threshold of  $p$ -value  $< 10^{-3}$  in at least one ancestry cohort and replicate in at least one other ancestry cohort. Nominal  $p$ -value thresholds tend to decrease the proportion of replicated SNP-level associations and tend to increase the proportion of replicated gene-level associations. The number of unique SNPs and genes that replicated in each cohort is given in Figure S15. **c.** The x-axis indicates the proportion of SNP-level associations that surpass a nominal threshold of  $p$ -value  $< 10^{-3}$  in at least one ancestry cohort that replicate in at least one other ancestry cohort. The y-axis indicates the proportion of gene-level associations that surpass a nominal threshold of  $p$ -value  $< 10^{-2}$  in at least one ancestry cohort and replicate in at least one other ancestry cohort. The number of unique SNPs and genes that replicated in each cohort is given in Figure S16. As shown in panel **b**, nominal  $p$ -value thresholds tend to decrease the proportion of replicated SNP-level associations and tend to increase the proportion of replicated gene-level associations. Expansion of three letter trait codes are given in Table S2.

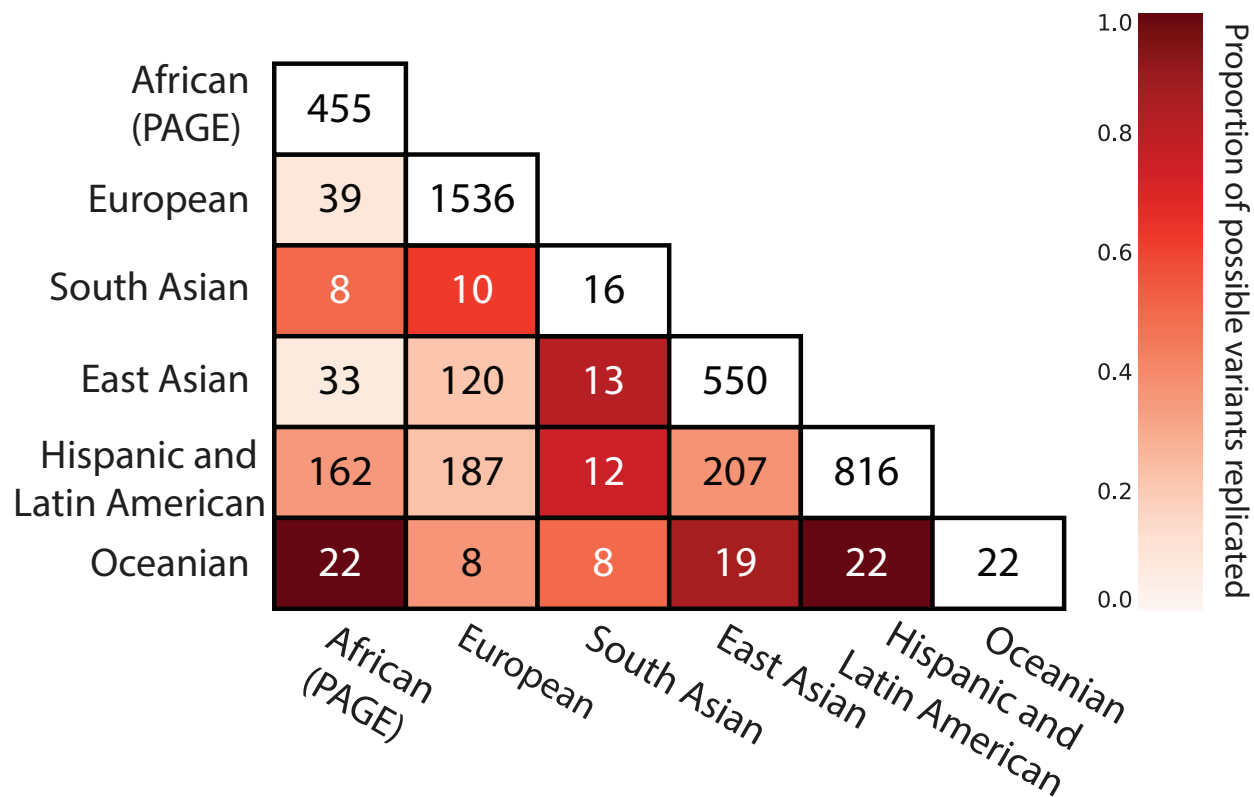


**Figure S15: Summaries of replicated associations at multiple genomic scales among ancestry cohorts for all 25 traits analyzed using nominal  $p$ -value thresholds.** (a) Number and (c) proportion of genome-wide significant SNPs associated with a phenotype in at least one ancestry cohort that were replicated in at least two ancestry cohorts using a nominal  $p$ -value threshold of  $5 \times 10^{-5}$ . (b) Number and (d) proportion of genome-wide significant gene-level associations that replicate among ancestry cohorts using a nominal  $p$ -value threshold of  $10^{-3}$ . Expansion of three letter trait codes are given in Table S2.

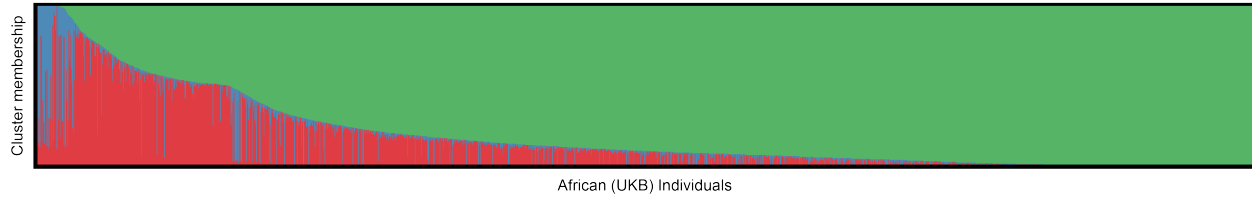


**Figure S16: Summaries of replicated associations at multiple genomic scales among ancestry cohorts for all 25 traits analyzed using nominal  $p$ -value thresholds.** (a) Number and (c) proportion of genome-wide significant SNPs associated with a phenotype in at least one ancestry cohort that were replicated in at least two ancestry cohorts using a nominal  $p$ -value threshold of  $10^{-3}$ . (b) Number and (d) proportion of genome-wide significant gene-level associations that replicate among ancestry cohorts using a nominal  $p$ -value threshold of  $10^{-2}$ . Expansion of three letter trait codes are given in Table S2. Expansion of three letter trait codes are given in Table S2.

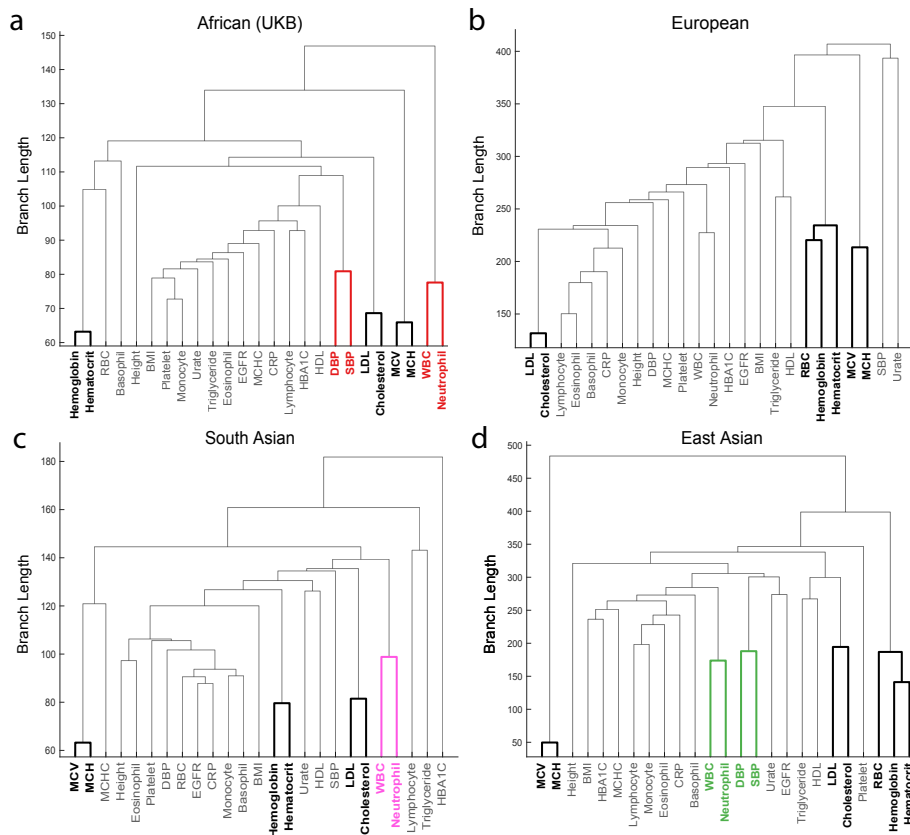




**Figure S17: Pairwise replication of SNP-level association signals for C-reactive protein in six ancestral cohorts using genotype and imputed data.** Imputed data was available and included in this analysis for each of the six cohorts. The inclusion of imputed SNPs in GWA analysis of C-reactive protein increases both the number of significant SNPs in each ancestry (along the diagonal) as well as the number of replicating significant SNP-level associations among pairs of ancestry cohorts (lower triangular). The corresponding analysis of pairwise SNP-level replication using only genotype data from each cohort is shown in Figure 2c.



**Figure S18: ADMIXTURE results from self-identified African individuals in the UK Biobank.** We performed 10 runs of unsupervised ADMIXTURE<sup>91</sup> setting  $K = 3$  on 4,967 self-identified African individuals in the UK Biobank. We included YRI and CEU individuals from the 1000 Genomes Project<sup>92</sup> to identify European and African ancestry components in the ADMIXTURE results. We then filtered out all individuals with less than 5% membership in the African component (identified as the component shown in green). The AIAN and European ancestry components are shown in blue and red, respectively. Our ADMIXTURE pipeline used the same protocol described in Bitarello and Mathieson<sup>84</sup>. All scripts for ADMIXTURE runs as well as filtration steps are available on the GitHub page given in the Data Availability section.



**Figure S19: Multiple prioritized trait clusters with shared core genetic trait architecture replicate in the (a) African (UKB), (b) European, (c) South Asian, and (d) East Asian ancestry cohorts.** The WINGS algorithm identified prioritized phenotype clusters in each of these ancestry cohorts, denoted in each dendrogram as clades with emboldened lines. Three clusters of phenotypes were found in all ancestries (shown and labeled in black), comprising: mean corpuscular volume (MCV) and mean corpuscular hemoglobin (MCH), hemoglobin and hematocrit, and the metabolic traits low-density lipoprotein (LDL) and cholesterol levels. In both the European and East Asian ancestry cohorts, red blood cell count (RBC) was also a member of the hemoglobin and hematocrit phenotype cluster. Two other phenotype clusters were identified in at least two ancestry cohorts. One of these clusters contains white blood cell count (WBC) and neutrophil count, and the other contains diastolic and systolic blood pressure (DBP and SBP). These two clusters are color-coded according to the ancestry cohorts in which they are prioritized. The WINGS algorithm was applied to traits from each ancestry cohort separately as described in the Supplemental Information.

<sup>454</sup> **Supplemental Tables**

Ancestry cohort label in this study	Study	Label from original study	Sample Size	Number of SNPs
European	UK Biobank	Self-identified white british	349,411	1,933,096
East Asian	Biobank Japan	East Asian	19,190 - 206,692	4,823,101 - 6,628,005
Hispanic and Latin American	PAGE	Admixed Hispanic and Latin American	15,522 - 21,955	8,576,622 - 8,822,607
African-American (PAGE)	PAGE	African-American	3,750 - 17,280	12,107,345 - 12,274,127
South Asian	UK Biobank	Self-identified South Asian	5,716	958,375
African (UKB)	UK Biobank	African	4,967	578,320
Native Hawaiian	PAGE	Native Hawaiian	1,777 - 3,938	6,656,996 - 6,966,169
American Indian/Alaska Native	PAGE	AIAN	574 - 645	3,970,247 - 8,504,923

**Table S1: Ancestry cohorts analyzed in this study.** In studies where GWA summary statistics were available to us, sample size and number of SNPs differ due to original study design. The specific sample size and number of SNPs for each trait in studies from Biobank Japan and PAGE are provided in Table S4-Table S9.

Trait Name	Code
Body mass index	BMI
High density lipoprotein	HDL
Low density lipoprotein	LDL
Hemoglobin A1c	HBA1C
Estimated glomerular filtration rate	EGFR
C-reactive protein	CRP
Systolic blood pressure	SBP
Diastolic blood pressure	DBP
Platelet count	PLC
Mean corpuscular hemoglobin concentration	MCHC
Mean corpuscular hemoglobin	MCH
Mean corpuscular volume	MCV
Red blood cell count	RBC
White blood cell count	WBC

**Table S2: Abbreviations used throughout this study for 14 quantitative traits analyzed in this study.** The remaining 11 traits analyzed were Basophil count, Cholesterol, Eosinophil count, Height, Hematocrit, Hemoglobin, Lymphocyte count, Monocyte count, Neutrophil count, and Triglyceride levels, respectively. These are not abbreviated in the main text.

Trait Name or Code	AIAN	Native Hawaiian	Hispanic
BMI	Yes	Yes	Yes
Basophil count	No	No	No
CRP	Yes	Yes	Yes
Cholesterol	Yes	Yes	Yes
DBP	Yes	No	Yes
EGFR	Yes	No	Yes
Eosinophil count	No	No	No
HBA1C	Yes	Yes	Yes
HDL	Yes	Yes	Yes
Height	Yes	Yes	Yes
Hematocrit	No	No	No
Hemoglobin	No	No	No
LDL	Yes	Yes	Yes
Lymphocyte count	No	No	No
MCH	Yes	No	Yes
MCHC	No	No	No
MCV	No	No	No
Monocyte count	No	No	No
Neutrophil count	No	No	No
PLC	Yes	No	Yes
RBC	No	No	No
SBP	Yes	No	Yes
Triglyceride	No	Yes	Yes
Urate	No	No	No
WBC	Yes	No	Yes

**Table S3: Traits assayed in the PAGE study data by ancestry cohort.** Data were available for each of the 25 listed traits in the UK Biobank European, South Asian, and African cohorts, as well as, the East Asian cohort from the Biobank Japan. Thus, each trait was analyzed in a minimum of four ancestries and a maximum of seven ancestries.

Trait Name or Code	Sample Size	Total SNPs	Pruned SNPs	Regions	Citations
Basophil count	62,076	5,653,566	410,465	23,106	93
BMI	158,284	5,653,566	410,465	23,085	94
CRP	75,391	5,608,701	408,166	23,108	93
DBP	136,615	5,653,566	410,465	23,085	93
eGFR	143,658	5,608,701	408,166	23,108	93
Eosinophil count	62,076	5,653,566	410,465	23,106	93
HDL	70,657	5,608,701	408,166	23,108	93
Height	159,095	6,296,332	354,647	23,679	95
Hematocrit	108,757	5,653,566	410,465	23,106	93
Hemoglobin	108,769	5,653,566	410,465	23,085	93
HbA1c	75,391	5,608,701	408,166	23,108	93
LDL	72,866	5,608,701	408,166	23,108	93
Lymphocyte count	62,076	5,653,566	410,465	23,106	93
MCH	108,054	5,653,566	410,465	23,106	93
MCHC	108,738	5,653,566	410,465	23,106	93
MCV	108,526	5,653,566	410,465	23,085	93
Monocyte count	62,076	5,653,566	410,465	23,106	93
Neutrophil count	62,076	5,653,566	410,465	23,106	93
PLC	108,208	5,653,566	410,465	23,085	93
RBC	108,794	5,653,566	410,465	23,085	93
SBP	136,597	5,653,566	410,465	23,085	93
Cholesterol	128,305	5,608,701	408,166	23,108	93
Triglyceride	105,597	5,608,701	410,465	23,108	93
Urate	109,029	5,608,701	408,166	23,108	93
WBC	107,694	5,653,566	408,166	23,085	93

**Table S4: Number of individuals, total SNPs, pruned SNPs used for gene- $\epsilon$ , and genes and transcription factors (regions) included in the analysis for each trait in Biobank Japan data. Regions were defined using the hg19 list provided in Gusev et al. <sup>42</sup>.**



Trait Name or Code	Sample Size	Total SNPs	Pruned SNPs	Regions
BMI	17,127	12,139,115	404,401	24,216
CRP	8,349	12,274,126	404,572	24,206
DBP	11,380	12,148,801	405,188	24,218
eGFR	8,261	12,128,273	403,371	24,207
Hemoglobin A1c	17,127	12,139,115	404,401	24,215
HDL	10,085	12,114,827	404,089	24,201
Height	17,280	12,139,907	404,522	24,201
LDL	9,720	12,107,344	403,740	24,218
MCHC	3,750	12,132,232	405,558	24,217
PLC	8,850	12,131,935	404,497	24,193
SBP	11,380	12,148,801	405,188	24,218
Cholesterol	10,137	12,110,337	403,674	24,222
Triglyceride	9,980	12,110,879	403,455	24,206
WBC	8,802	12,126,732	404,579	24,219

**Table S5: Number of individuals, total SNPs, pruned SNPs used for gene- $\epsilon$ , and genes and transcription factors (regions) included in the analysis for each trait in the African-American cohort of the PAGE study data.**

Trait Name or Code	Sample Size	Total SNPs	Pruned SNPs	Region Count
BMI	21,955	8,812,436	432,762	24,138
CRP	15,912	8,576,621	397,941	24,118
DBP	21,549	8,784,112	430,360	24,126
Estimated glomerular filtration rate	18,548	8,702,426	422,598	24,123
HbA1c	21,955	8,812,436	432,762	24,138
HDL	17,751	8,583,603	412,771	24,122
Height	22,187	8,822,606	433,604	24,132
LDL	17,373	8,588,800	413,074	24,116
MCHC	15,522	8,763,739	427,208	24,132
PLC	18,949	8,612,804	415,201	24,115
SBP	21,549	8,784,112	430,360	24,126
Cholesterol	17,802	8,586,887	412,830	24,115
Triglyceride	17,856	8,594,121	413,546	24,104
WBC	18,206	8,603,503	414,462	24,123

**Table S6: Number of individuals, total SNPs, pruned SNPs used for gene- $\epsilon$ , and genes and transcription factors (regions) included in the analysis for each trait in the Hispanic and Latin American cohort of the PAGE study data.**

Trait Name or Code	Sample Size	Total SNPs	Pruned SNPs	Regions
BMI	645	8,374,976	421,826	24,124
CRP	574	8,504,922	417,287	24,136
DBP	636	8,376,521	421,528	24,136
eGFR	602	8,336,044	417,540	24,132
Hemoglobin A1c	645	8,374,976	421,826	24,124
HDL	604	8,315,912	415,939	24,121
Height	645	8,375,624	421,750	24,117
LDL	591	8,360,719	419,544	24,123
MCHC	20	3,970,246	62,339	17,381
PLC	603	8,294,302	414,530	24,133
Systolic blood pressure	636	8,376,521	421,528	24,136
Cholesterol	604	8,586,887	415,939	24,121
WBC	602	8,289,567	414,462	24,133

**Table S7: Number of individuals, total SNPs, pruned SNPs used for gene- $\epsilon$ , and genes and transcription factors (regions) included in the analysis for each trait in the AIAN cohort of the PAGE study data.**

Trait Name or Code	Sample Size	Total SNPs	Pruned SNPs	Regions
BMI	3,936	6,664,738	415,221	23,885
CRP	1,777	6,966,169	428,517	23,834
Hemoglobin A1c	3,936	6,664,738	415,221	23,885
HDL	1,912	6,656,996	416,255	23,894
Height	3,938	6,660,920	415,172	23,878
LDL	1,900	6,662,802	416,810	23,895
Cholesterol	1,915	6,660,807	416,425	23,899
Triglycerides	1,915	6,660,807	416,425	23,899

**Table S8: Number of individuals, total SNPs, pruned SNPs used for gene- $\epsilon$ , and genes and transcription factors (regions) included in the analysis for each trait in the Native Hawaiian (Native Hawaiian) cohort of the PAGE study data.**

Trait Name or Code	Sample Size	Total SNPs	Pruned SNPs	Regions
BMI	4,647	15,362,633	433,356	24,085
CRP	1,811	14,374,461	428,656	24,116
DBP	1,086	12,470,507	416,273	24,112
eGFR	150	8,314,417	337,167	24,017
HbA1c	4,647	15,362,633	433,356	24,085
HDL	2,378	13,413,244	428,598	24,072
Height	4,679	15,366,710	433,005	24,103
LDL	2,316	13,327,313	428,741	24,075
MCHC	128	8,089,136	315,583	23,946
PLC	541	10,528,072	421,929	24,098
SBP	1,086	12,470,507	416,273	24,112
Cholesterol	2,387	13,436,190	428,656	24,078
Triglyceride	2,381	13,423,953	429,246	24,073
WBC	543	10,570,051	421,776	24,095

**Table S9: Number of individuals, total SNPs, pruned SNPs used for gene- $\epsilon$ , and genes and transcription factors (regions) included in the analysis for each trait in the Asian cohort of the PAGE study data.**

Trait Name or Code	African or African-American ( $\times 10^{-8}$ )	East Asian ( $\times 10^{-8}$ )	AIAN ( $\times 10^{-8}$ )	Native Hawaiian ( $\times 10^{-9}$ )	Hispanic ( $\times 10^{-8}$ )
Basophil count	8.646*	8.844	NA	NA	NA
BMI	0.412	8.844	0.597	7.502	5.674
CRP	0.407	8.915	0.588	7.176	5.830
Cholesterol	0.414	8.915	0.601	7.507	5.823
DBP	0.412	8.844	0.597	NA	5.692
EGFR	0.412	8.915	0.600	NA	5.746
Eosinophil count	8.646*	8.844	NA	NA	NA
HBA1C	0.412	8.915	0.597	7.502	5.674
HDL	0.413	8.915	0.601	7.511	5.825
Height	0.412	7.941	0.597	7.506	5.667
Hematocrit	8.646*	8.844	NA	NA	NA
Hemoglobin	8.646*	8.844	NA	NA	NA
LDL	0.413	8.914	0.598	7.504	5.822
Lymphocyte count	8.646*	8.843	NA	NA	NA
MCH	8.646*	8.844	NA	NA	NA
MCHC	0.412	8.844	1.259	NA	5.705
MCV	8.646*	8.844	NA	NA	NA
Monocyte count	8.646*	8.844	NA	NA	NA
Neutrophil count	8.646*	8.844	NA	NA	NA
PLC	0.412	8.844	0.603	NA	5.805
RBC	8.646*	8.844	NA	NA	NA
SBP	0.412	8.844	0.597	NA	5.692
Triglyceride	0.413	8.915	NA	7.507	5.818
Urate	8.646*	8.915	NA	NA	NA
WBC	0.412	8.844	NA	NA	5.812

**Table S10: Bonferroni  $p$ -value threshold corrected for number of SNPs tested for each ancestry-trait pair.** Thresholds are calculated as 0.05 divided by the number of SNPs tested in each ancestry-trait pair. The term "NA" indicates that there was no data for that ancestry-trait pair. The threshold for Bonferroni-corrected significance was the same for every trait in the European ( $p$ -value  $< 2.587 \times 10^{-8}$ ) and South Asian ( $p$ -value  $< 5.217 \times 10^{-8}$ ) cohorts from the UK Biobank. Traits for which the UK Biobank African cohort was used are denoted with a \*; otherwise, the African-American cohort from the PAGE study data was used. See Table S2 for expansion of trait codes.

Trait Name or Code	Median Effect Size Correlation	Effect Sizes > 0.1 (Mean)	Median PIP Correlation	PIPs > 0.01 (Mean)
BMI	0.0143	0	0.073	67.3
Basophil	$3.69 \times 10^{-6}$	0	0.002	211.8
CRP	0.096	0.7	0.159	142.6
Cholesterol	0.966	1	0.232	53.1
DBP	0.001	0.6	0.012	173.1
EGFR	$4.71 \times 10^{-6}$	0	0.005	149
Eosinophil	0.0179	0.2	0.042	325.8
HBA1C	0.012	0.2	0.025	26.8
HDL	0.325	0.1	0.219	57.6
Height	$9.20 \times 10^{-5}$	1.6	0.010	187.7
Hematocrit	$-4.16 \times 10^{-6}$	0.5	0.015	40.3
Hemoglobin	$-1.13 \times 10^{-5}$	0.6	0.014	45.8
LDL	0.949	1	0.331	60.9
Lymphocyte	$-1.01 \times 10^{-6}$	1.5	0.001	1368.5
MCH	$7.098 \times 10^{-6}$	0.6	0.007	60.9
MCHC	$-2.00 \times 10^{-6}$	0.9	0.01	58.7
MCV	$1.71 \times 10^{-5}$	0.6	0.01	70
Monocyte	0.002	0.1	0.017	349.4
Neutrophil	0.038	0.1	0.071	66.5
PLC	0.609	0.9	0.154	112.5
RBC	$9.650 \times 10^{-5}$	0.9	0.028	49.7
SBP	0.001	0.6	0.019	166.8
Triglyceride	0.348	0.0	0.247	164.7
Urate	0.252	0.5	0.24	39.9
WBC	0.0002	0.1	0.01	347.8

**Table S11: Replication of effect sizes and posterior inclusion probabilities (PIPs) among ten independent subsamples of the UK Biobank European ancestry cohort using SuSiE<sup>34</sup> for fine-mapping.** The sample size of the ten independent, non-overlapping subsamples of the UK Biobank European ancestry cohort was set to 10,000. For the 1,895,051 SNPs that were analyzed in every European ancestry cohort subsample ( Table S1) and the effect sizes and PIPs (columns 2 and 4, respectively) generated using the SuSiE method<sup>34</sup>, we calculated the median correlation coefficient between all possible pairwise comparisons (10 choose 2) of the European ancestry cohort subsamples. Column 3 reports the mean number of SNPs with effect sizes greater than 0.1 across all ten European ancestry cohort subsamples for each trait. Column 5 reports the mean number of SNPs with a posterior inclusion probability greater than 0.01 across the ten European ancestry cohort subsamples for each trait.

Trait Name or Code	Number of Shared Effect Sizes > 0.1	Median Effect Size Correlation	African Effect Sizes > 0.1	European Effect Sizes > 0.1 (Mean)	Number of Shared PIPs > 0.01	Median PIP Correlation	African PIPs > 0.01	European PIPs > 0.01 (Mean)
BMI	0	$5.12 \times 10^{-5}$	2	0	1	0.027	1206	8.9
Basophil	0	$-3.01 \times 10^{-5}$	0	0	0	0.001	403	40.3
CRP	0	$-3.68 \times 10^{-6}$	0	0.4	0	$5.88 \times 10^{-5}$	95	33.6
Cholesterol	1	0.5	1	0.9	5	0.467	1144	8.4
DBP	0	$1.58 \times 10^{-5}$	10	0.3	0	0.008	1249	20.3
EGFR	0	$2.41 \times 10^{-6}$	1	0	0	0.001	984	24.7
Eosinophil	0	$-1.97 \times 10^{-5}$	0	0	0	0.003	979	29.3
HBA1C	0	$-1.15 \times 10^{-6}$	5	0.2	0	0.004	959	6.5
HDL	0	0.014	0	0.1	2	0.014	31	11.6
Height	0	$-1.22 \times 10^{-5}$	19	0.7	0	0.003	208	33.7
Hematocrit	0	$-1.14 \times 10^{-5}$	1	0	0	0.016	199	12.2
Hemoglobin	0	$-6.31 \times 10^{-6}$	1	0	0	0.022	568	12.5
LDL	1	0.674	1	1	8	0.497	995	13.5
Lymphocyte	0	$2.71 \times 10^{-5}$	0	0	1	0.001	60	115.5
MCH	0	$1.14 \times 10^{-5}$	1	0.2	0	0.005	1332	11.8
MCHC	0	$-1.56 \times 10^{-6}$	1	0.2	0	0.008	646	13.6
MCV	0	$2.58 \times 10^{-6}$	1	0.2	0	0.005	917	11.8
Monocyte	0	$-2.44 \times 10^{-5}$	0	0	2	0.002	2215	27.7
Neutrophil	0	-0.0001	1	0	0	0.001	2386	10.2
PLC	0	-0.017	1	0.6	3	0.035	2088	13.7
RBC	0	$-6.35 \times 10^{-6}$	1	0.1	0	-0.002	3	12.7
SBP	0	$4.31 \times 10^{-5}$	13	0.1	0	0.008	1133	20
Triglyceride	0	0.004	0	0.1	8	0.009	2231	25.6
Urate	1	0.01	2	0.3	3	0.051	334	9.8
WBC	0	-0.0001	1	0	5	0.0006	2862	48.3

**Table S12: Replication of effect sizes and posterior inclusion probabilities (PIPs) between the UK Biobank African ancestry cohort and ten independent subsamples of the UK Biobank European ancestry cohort using SuSiE<sup>34</sup> for fine-mapping.** Each of the ten independent, non-overlapping subsamples of the UK Biobank European ancestry cohort was set to be equal in size to the sample size of the African ancestry cohort ( $N = 4,967$ ), Table S1. Column headers containing "(mean)" indicate a mean is generated averaging over the ten independent European ancestry cohort subsamples. For the 496,997 SNPs that were analyzed in both the African ancestry cohort and every European ancestry cohort subsample, we compared the SuSiE<sup>34</sup> effect size estimates and PIPs. For both effect sizes and PIPs, the median correlation coefficient between the African ancestry cohort and the pairwise comparison to each European ancestry cohort subsample is reported in the third and seventh columns, respectively. Column 3 reports the total number of SNPs with effect sizes greater than 0.1 in the African cohort. Column 4 gives the mean number of effect sizes greater than 0.1 in the European ancestry cohort subsamples for each trait. We performed the same comparison for the PIPs using a threshold of 0.01. Column 2 reports the number of SNPs that surpassed an effect size of 0.1 in both the African ancestry cohort and at least one of the European ancestry cohort subsamples. Column 6 reports the number of SNPs that surpasses a PIP of 0.01 in the African ancestry cohort and at least one European ancestry cohort subsample.

Trait Name or Code	Number of Shared Effect Sizes > 0.1	Median Effect Size Correlation	South Asian Effect Sizes > 0.1	European Effect Sizes > 0.1 (Mean)	Number of Shared PIPs > 0.01	Median PIP Correlation	South Asian PIPs > 0.01	European PIPs > 0.01 (Mean)
BMI	0	$7.07 \times 10^{-5}$	1	0	0	0.0173	180	18
Basophil	0	$-1.10 \times 10^{-5}$	0	0	0	0.0008	210	78.4
CRP	0	0.0309	0	0.4	17	0.121	181	64.7
Cholesterol	0	0.004	0	0.9	2	0.0127	20	9
DBP	0	$-2.12 \times 10^{-5}$	0	0.6	0	0.013	115	40
EGFR	0	$-1.26 \times 10^{-5}$	0	0	0	$5.12 \times 10^{-5}$	452	44.2
Eosinophil	0	$2.16 \times 10^{-5}$	0	0	0	0.006	63	69.1
HBA1C	0	$-3.33 \times 10^{-5}$	2	0.3	0	0.011	133	13.3
HDL	0	0.525	0	0.1	3	0.505	78	12.2
Height	0	$-2.13 \times 10^{-5}$	12	1.1	1	0.001	754	66.8
Hematocrit	0	$-4.78 \times 10^{-6}$	1	0.3	0	$6.24 \times 10^{-5}$	24	17.4
Hemoglobin	0	$-4.19 \times 10^{-6}$	1	0.3	0	0.011	19	16.1
LDL	0	0.491	1	1	14	0.386	44	15.6
Lymphocyte	0	$-4.40 \times 10^{-6}$	0	1.1	0	0.0001	113	343.1
MCH	0	$5.29 \times 10^{-6}$	2	0.5	0	0.008	78	22.3
MCHC	0	$-5.77 \times 10^{-6}$	1	0.9	0	0.007	45	20.3
MCV	0	$1.74 \times 10^{-5}$	3	0.5	0	0.005	81	26.3
Monocyte	0	$-4.05 \times 10^{-6}$	0	0.1	0	0.001	156	76.8
Neutrophil	0	0.221	0	0	16	0.172	70	25.7
PLC	0	0.577	0	0.8	1	0.679	126	22.7
RBC	0	$2.79 \times 10^{-6}$	0	0.3	0	-0.011	41	20.9
SBP	0	-0.0004	0	0.3	0	0.037	116	43.2
Triglyceride	0	0.158	0	0	53	0.173	189	46.4
Urate	0	0.137	1	0.4	6	0.150	52	15.8
WBC	0	0.001	0	0	18	0.008	52	93

**Table S13: Replication of effect sizes and posterior inclusion probabilities (PIPs) between the UK Biobank South Asian ancestry cohort and ten independent subsamples of the UK Biobank European ancestry cohort using SuSiE<sup>34</sup> for fine-mapping.** Each of the ten independent, non-overlapping subsamples of the UK Biobank European ancestry cohort was set to be equal in size to the sample size of the South Asian ancestry cohort ( $N = 5,660$ ), Table S1. Column headers containing "(mean)" indicate a mean is generated averaging over the ten independent European ancestry cohort subsamples. For the 863,569 SNPs that were analyzed in both the South Asian ancestry cohort and every European ancestry cohort subsample, we compared the SuSiE<sup>34</sup> effect size estimates and PIPs. For both effect sizes and PIPs, the median correlation coefficient between the South Asian ancestry cohort and the pairwise comparison to each European ancestry cohort subsample is reported in the third and seventh columns, respectively. Column 3 reports the total number of SNPs with effect sizes greater than 0.1 in the South Asian cohort. Column 4 gives the mean number of effect sizes greater than 0.1 in the European ancestry cohort subsamples for each trait. We performed the same comparison for the PIPs using a threshold of 0.01. Column 2 reports the number of SNPs that surpassed an effect size of 0.1 in both the South Asian ancestry cohort and at least one of the European ancestry cohort subsamples. Column 6 reports the number of SNPs that surpasses a PIP of 0.01 in the South Asian ancestry cohort and at least one European ancestry cohort subsample.

Trait	PESCA EUR only and EUR GWA nominal significance	PESCA EAS only and EAS GWA nominal significance	PESCA both and EUR GWA nominal significance	PESCA both and EAS GWA nominal significance	PESCA both and GWA nominal significance in both
BMI	0	0	97	539	94
Cholesterol	2	13	14	54	11
HDL	5	8	29	74	27
LDL	0	6	16	43	13
MCH	3	14	39	180	39
MCV	1	11	44	215	44
Triglyceride	4	8	15	49	15

**Table S14: Overlap between SNPs identified by PESCA and GWA analyses in the European ancestry cohort from the UK Biobank and the East Asian ancestry cohort from the Biobank Japan.** For seven continuous traits, we compared SNP-level association  $p$ -values from our analysis to the posterior probabilities calculated in Shi et al.<sup>51</sup> using the PESCA framework. For each of the seven traits, there were SNPs that had a posterior probability  $> 0.8$  of being causal in both ancestries and were also nominally significant ( $p$ -value  $5 < 10^{-6}$ ) using the standard GWA SNP-level framework.



Trait	African or African-American ( $\times 10^{-6}$ )	East Asian ( $\times 10^{-6}$ )	AIAN ( $\times 10^{-6}$ )	Native Hawaiian ( $\times 10^{-6}$ )	Hispanic and Latin American ( $\times 10^{-6}$ )
Basophil count	1.121	2.197	NA	NA	NA
BMI	2.096	2.197	2.073	2.093	2.072
CRP	2.097	2.085	2.072	2.100	2.073
Cholesterol	2.096	2.198	2.073	2.092	2.073
DBP	2.096	2.197	2.072	NA	2.073
EGFR	2.097	2.198	2.073	NA	2.073
Eosinophil count	2.121*	2.197	NA	NA	NA
HBA1C	2.096	2.198	2.073	2.093	2.072
HDL	2.097	2.198	2.073	2.093	2.073
Height	2.097	2.131	2.073	2.094	2.072
Hematocrit	2.121*	2.197	NA	NA	NA
Hemoglobin	2.121*	2.197	NA	NA	NA
LDL	2.096	2.198	2.073	2.093	2.073
Lymphocyte count	2.121*	2.197	NA	NA	NA
MCH	2.121*	2.197	NA	NA	NA
MCHC	2.096	2.197	2.877	NA	2.072
MCV	2.121*	2.197	NA	NA	NA
Monocyte count	2.121*	2.197	NA	NA	NA
Neutrophil count	2.121*	2.197	NA	NA	NA
PLC	2.097	2.197	2.072	NA	2.073
RBC	2.121*	2.197	NA	NA	NA
SBP	0.412	8.844	0.597	NA	5.692
Triglyceride	2.096	2.197	NA	2.072	2.073
Urate	2.121*	2.197	NA	NA	NA
WBC	2.092	2.197	2.072	NA	2.073

**Table S15: Bonferroni  $p$ -value threshold corrected for number of gene-level association tests performed for each ancestry-trait pair.** Thresholds are calculated as 0.05 divided by the number of SNPs tested in each ancestry-trait pair. The term "NA" indicates that there was no data for that ancestry-trait pair. The threshold for Bonferroni-corrected significance was the same for every trait in the European ( $p$ -value  $< 2.092 \times 10^{-6}$ ) and South Asian ( $p$ -value  $< 1.085 \times 10^{-6}$ ) cohorts from the UK Biobank. Traits for which the UK Biobank African cohort was used are denoted with a \*; otherwise, the African-American cohort from the PAGE study data was used. See Table S2 for expansion of trait codes.

Gene	African or African-American	European	East Asian	Hispanic
<i>ANGPTL4</i>	96	97	NA	97
<i>APOA1</i>	98	98	NA	98
<i>APOA4</i>	NA	99	NA	NA
<i>APOA5</i>	100	100–103	104	100
<i>APOB</i>	105	105	NA	105
<i>APOC1</i>	98	99	NA	98
<i>APOC2</i>	100	100	NA	100
<i>APOC3</i>	98	98,99	NA	98
<i>APOC4</i>	99	99	NA	99
<i>APOE</i>	99	99,100	NA	99
<i>CETP</i>	98	99	NA	98
<i>LMF1</i>	NA	99	NA	NA
<i>LPL</i>	98	99	104	98
<i>PCSK6</i>	96	99	NA	96
<i>PCSK7</i>	100	100	104	100
<i>PLTP</i>	97,98,100,105	97–99,105–107	NA	97,98,100,105

**Table S16: Genes shown in Figure 3 as associated with triglyceride levels are supported by publications in the GWAS Catalog.** Each of the genes listed is present in the significantly mutated subnetworks identified using Hierarchical HotNet<sup>31</sup> as enriched for associations with triglyceride levels in the European, East Asian, or Native Hawaiian ancestry cohorts. We mapped SNP-level associations from the GWAS Catalog to the 29 genes present in the significantly mutated subnetworks shown in Figure 3 (using the gene list provided by Gusev et al.<sup>42</sup>) to generate the results for the 16 genes shown here.

Gene	African-American	European	South Asian	East Asian	AIAN	Hispanic and Latin American
<i>RDH13</i>	$4.14 \times 10^{-10}$	$9.95 \times 10^{-1}$	$8.80 \times 10^{-2}$	$1.76 \times 10^{-6}$	1	$7.88 \times 10^{-1}$
<i>AGPAT5</i>	1	$1.30 \times 10^{-6}$	$7.83 \times 10^{-1}$	$5.00 \times 10^{-1}$	$7.33 \times 10^{-8}$	1
<i>GP6</i>	$7.20 \times 10^{-10}$	$9.93 \times 10^{-1}$	$1.47 \times 10^{-1}$	$9.07 \times 10^{-7}$	1	$6.92 \times 10^{-1}$
<i>ALDH2</i>	1	$1.00 \times 10^{-20}$	$1.13 \times 10^{-2}$	$1.00 \times 10^{-20}$	1	1
<i>RAB8A</i>	$9.57 \times 10^{-1}$	$1.00 \times 10^{-20}$	1	$5.76 \times 10^{-6}$	1	$9.97 \times 10^{-1}$
<i>CUX2</i>	1	$5.13 \times 10^{-7}$	$1.16 \times 10^{-1}$	$3.44 \times 10^{-11}$	1	1
<i>ACAD10</i>	1	$1.47 \times 10^{-10}$	$1.10 \times 10^{-2}$	$2.00 \times 10^{-10}$	1	1

**Table S17: Gene-level association  $p$ -values for seven genes associated with platelet count in at last two ancestry cohorts.** Of the 65 genes that were associated with platelet count in at least two ancestry cohorts, these seven contained previously identified SNP-level associations in studies submitted to the GWAS Catalog. Previous associations in the GWAS Catalog are discussed in the Supplemental Information. Ancestry-specific Bonferroni corrected significance thresholds for gene-level association analysis of platelet count are shown in Table S15.

Gene	African-American (PAGE)	European	South Asian	East Asian	Native Hawaiian	Hispanic and Latin American
<i>APOA1</i>	$4.99 \times 10^{-1}$	$1.00 \times 10^{-20}$	$9.91 \times 10^{-1}$	$1.00 \times 10^{-20}$	$7.52 \times 10^{-1}$	$4.99 \times 10^{-1}$
<i>APOA4</i>	$4.99 \times 10^{-1}$	$1.00 \times 10^{-20}$	$2.51 \times 10^{-5}$	$1.00 \times 10^{-20}$	$9.15 \times 10^{-1}$	$4.99 \times 10^{-1}$
<i>APOA5</i>	$4.99 \times 10^{-1}$	$1.42 \times 10^{-11}$	$1.60 \times 10^{-6}$	$9.95 \times 10^{-1}$	$3.67 \times 10^{-12}$	$4.99 \times 10^{-1}$
<i>APOC3</i>	$4.99 \times 10^{-1}$	$1.00 \times 10^{-20}$	$9.82 \times 10^{-1}$	$9.83 \times 10^{-1}$	$3.05 \times 10^{-15}$	$4.99 \times 10^{-1}$
<i>APOE</i>	$4.99 \times 10^{-1}$	$1.00 \times 10^{-20}$	$8.65 \times 10^{-1}$	$1.00 \times 10^{-20}$	1	1
<i>PLTP</i>	$4.99 \times 10^{-1}$	$4.29 \times 10^{-9}$	$9.66 \times 10^{-1}$	$6.66 \times 10^{-15}$	$1.00 \times 10^{-2}$	$4.99 \times 10^{-1}$
<i>LPL</i>	$4.99 \times 10^{-1}$	$4.08 \times 10^{-13}$	$3.00 \times 10^{-3}$	$1.00 \times 10^{-20}$	$6.59 \times 10^{-1}$	$4.99 \times 10^{-1}$
<i>ANGPTL3</i>	$4.99 \times 10^{-1}$	$8.86 \times 10^{-8}$	$2.00 \times 10^{-3}$	$1.00 \times 10^{-20}$	$4.00 \times 10^{-3}$	1
<i>ANGPTL4</i>	$4.99 \times 10^{-1}$	$1.00 \times 10^{-20}$	$9.78 \times 10^{-1}$	$9.99 \times 10^{-1}$	$9.89 \times 10^{-1}$	1
<i>APOC1</i>	$4.99 \times 10^{-1}$	$1.67 \times 10^{-16}$	$4.99 \times 10^{-1}$	$1.00 \times 10^{-20}$	$9.81 \times 10^{-1}$	$4.99 \times 10^{-1}$
<i>APOC2</i>	$4.99 \times 10^{-1}$	$3.57 \times 10^{-13}$	$7.71 \times 10^{-1}$	$1.11 \times 10^{-1}$	$9.11 \times 10^{-1}$	$4.99 \times 10^{-1}$
<i>APOC4</i>	$4.99 \times 10^{-1}$	$3.72 \times 10^{-13}$	$7.36 \times 10^{-1}$	$2.58 \times 10^{-14}$	$9.73 \times 10^{-1}$	$4.99 \times 10^{-1}$
<i>APOB</i>	$4.99 \times 10^{-1}$	$1.00 \times 10^{-20}$	$9.99 \times 10^{-1}$	$7.32 \times 10^{-12}$	$1.00 \times 10^{-3}$	1
<i>LMF1</i>	$9.98 \times 10^{-1}$	$8.03 \times 10^{-7}$	1	$3.21 \times 10^{-2}$	$3.79 \times 10^{-5}$	$9.98 \times 10^{-1}$
<i>APOL1</i>	$4.99 \times 10^{-1}$	$5.30 \times 10^{-2}$	$6.40 \times 10^{-2}$	1	$8.89 \times 10^{-11}$	$4.99 \times 10^{-1}$
<i>HBA1</i>	$4.99 \times 10^{-1}$	$3.75 \times 10^{-5}$	$9.99 \times 10^{-1}$	$4.51 \times 10^{-1}$	$2.46 \times 10^{-10}$	1
<i>HBA2</i>	$4.99 \times 10^{-1}$	$1.30 \times 10^{-5}$	$9.99 \times 10^{-1}$	$4.51 \times 10^{-1}$	$3.93 \times 10^{-10}$	$4.99 \times 10^{-1}$
<i>B4GALT3</i>	$4.99 \times 10^{-1}$	$6.80 \times 10^{-2}$	$7.21 \times 10^{-1}$	$4.99 \times 10^{-1}$	$1.23 \times 10^{-6}$	1
<i>KLK8</i>	1	1	$1.62 \times 10^{-6}$	$9.89 \times 10^{-1}$	$1.00 \times 10^{-3}$	1
<i>PNLIP</i>	$4.99 \times 10^{-1}$	$9.99 \times 10^{-1}$	$9.26 \times 10^{-1}$	$7.75 \times 10^{-1}$	$1.00 \times 10^{-3}$	$4.99 \times 10^{-1}$
<i>WNT4</i>	$4.99 \times 10^{-1}$	$9.61 \times 10^{-1}$	$9.99 \times 10^{-1}$	$4.99 \times 10^{-1}$	$3.29 \times 10^{-5}$	$4.99 \times 10^{-1}$
<i>BACE1</i>	$4.99 \times 10^{-1}$	$5.55 \times 10^{-17}$	$2.20 \times 10^{-2}$	$9.99 \times 10^{-16}$	$6.69 \times 10^{-1}$	$4.99 \times 10^{-1}$
<i>CETP</i>	$4.99 \times 10^{-1}$	$1.00 \times 10^{-3}$	$9.99 \times 10^{-1}$	$1.41 \times 10^{-6}$	$9.99 \times 10^{-1}$	$4.99 \times 10^{-1}$
<i>PCSK6</i>	$4.99 \times 10^{-1}$	1	$9.99 \times 10^{-1}$	$1.83 \times 10^{-5}$	$1.00 \times 10^{-3}$	$4.99 \times 10^{-1}$
<i>PCSK7</i>	$4.99 \times 10^{-1}$	$1.66 \times 10^{-8}$	$9.97 \times 10^{-1}$	$1.00 \times 10^{-20}$	$9.99 \times 10^{-1}$	$4.99 \times 10^{-1}$
<i>LCAT</i>	$4.99 \times 10^{-1}$	$5.00 \times 10^{-1}$	1	$6.24 \times 10^{-3}$	$4.38 \times 10^{-1}$	$4.99 \times 10^{-1}$
<i>APOF</i>	$4.99 \times 10^{-1}$	$5.78 \times 10^{-1}$	$7.79 \times 10^{-1}$	$4.10 \times 10^{-3}$	$9.64 \times 10^{-1}$	$4.99 \times 10^{-1}$
<i>TYRO3</i>	$4.99 \times 10^{-1}$	$9.28 \times 10^{-1}$	$9.99 \times 10^{-1}$	$1.20 \times 10^{-2}$	$8.57 \times 10^{-1}$	$4.99 \times 10^{-1}$

**Table S18: Gene- $\epsilon$   $p$ -values for the 28 genes present in the significantly mutated sunnetworks associated with triglyceride level in the European, East Asian, and Native Hawaiian cohorts.**

Each of these genes is present in Figure 3 which depicts the overlapping significantly mutated subnetworks identified using Hierarchical HoNet<sup>31</sup> identified in an analysis of triglyceride levels in the European, East Asian, and Native Hawaiian cohorts. Known SNP-level associations identified within the bounds of these genes in previous studies submitted to the GWAS Catalog are discussed in the Supplemental Information. Ancestry-specific Bonferroni corrected significance thresholds for gene-level association analysis of triglyceride levels are shown in Table S15.

## 455 S1 Supplemental Information

### 456 Supplemental Subjects and Methods

#### 457 UK Biobank Data

458 We downloaded individual genotype data using the UK Biobank's (UKB) `ukbgene` resource, [https://](https://biobank.ctsu.ox.ac.uk/crystal/download.cgi)  
459 [biobank.ctsu.ox.ac.uk/crystal/download.cgi](https://biobank.ctsu.ox.ac.uk/crystal/download.cgi). European individuals from the UK Biobank data were  
460 selected using the self-identified ancestry (data field 21000) using values outlined at [https://biobank.](https://biobank.ctsu.ox.ac.uk/crystal/field.cgi?id=21000)  
461 [ctsu.ox.ac.uk/crystal/field.cgi?id=21000](https://biobank.ctsu.ox.ac.uk/crystal/field.cgi?id=21000). Using the relatedness file provided by the UK Biobank,  
462 one individual from each related pair was then randomly removed. This process was repeated for individuals  
463 whose self-identified ancestry was South Asian.

464 We performed unsupervised genome-wide ancestry estimation using ADMIXTURE by setting  $K = 3^{91}$  on  
465 the self-identified African ancestry cohort. We also included YRI and CEU individuals in the ADMIXTURE  
466 runs from the 1000 Genomes Project, to identify the ancestry components corresponding to African and Euro-  
467 pean ancestry. We removed individuals containing less than 5% membership in the African ancestry compo-  
468 nent and more than 5% membership in the third component, which corresponds to American Indian/Alaskan  
469 Native (AIAN) ancestry (Figure S18). We downloaded imputed SNP data from the UK Biobank for all re-  
470 maining individuals and removed SNPs with an information score below 0.8. Information scores for each SNP  
471 are provided by the UK Biobank (<http://biobank.ctsu.ox.ac.uk/crystal/refer.cgi?id=1967>). The  
472 remaining genotype and high-quality imputed SNPs were put through a stringent quality control pipeline in  
473 each ancestry cohort to obtain cohort-specific SNPs to be used for further analysis as detailed in the main  
474 text (detailed below).

475 We performed the following quality control filters in the European, South Asian, and African cohorts  
476 from the UK Biobank (Application number 22419). Genotype data for 488,377 individuals in the UK  
477 Biobank were downloaded using the UK Biobank's `ukbgene` ([https://biobank.ctsu.ox.ac.uk/crystal/](https://biobank.ctsu.ox.ac.uk/crystal/download.cgi)  
478 [download.cgi](https://biobank.ctsu.ox.ac.uk/crystal/download.cgi)) tool and converted using the provided `ukbconv` tool ([https://biobank.ctsu.ox.ac.uk/](https://biobank.ctsu.ox.ac.uk/crystal/refer.cgi?id=149660)  
479 [crystal/refer.cgi?id=149660](https://biobank.ctsu.ox.ac.uk/crystal/refer.cgi?id=149660)). Phenotype data was also downloaded for those same individuals using  
480 the `ukbgene` tool. Individuals identified by the UK Biobank to have high heterozygosity, excessive related-  
481 ness, or aneuploidy were removed (1,550 individuals). After then separating individuals into self-identified  
482 ancestral groups using data field 21000. Within these cohorts, unrelated individuals were then selected by  
483 randomly selecting an individual from each pair of related individuals. This resulted in 349,469 European  
484 individuals, 5,716 South Asian individuals, and 4,967 African individuals.

485 Genotype quality control was then performed on each cohort separately using the following steps. All

486 structural variants were first removed, leaving only single nucleotide polymorphisms in the genotype data.  
487 Next, all AT/CG SNPs were removed to avoid possible confounding due to sequencing errors. Then, SNPs  
488 with minor allele frequency less than 1% were removed using the plink2<sup>32</sup> `--maf 0.01`. We then removed  
489 all SNPs found to be in Hardy-Weinberg equilibrium, using the plink `--hwe 0.000001` flag to remove all  
490 SNPs with a Fisher's exact test  $p$ -value  $> 10^{-6}$ . Finally, all SNPs with missingness greater than 1% were  
491 removed using the plink `--mind 0.01` flag.

## 492 **Biobank Japan Data**

493 We downloaded summary statistics for 25 quantitative traits from the Biobank Japan (BBJ) website (<http://jenger.riken.jp/en/result>)<sup>1,93-95</sup>. The sample descriptions and number of SNPs included in our  
494 analyses are given in Table S4. The number of SNPs included in the analysis of each trait represent those  
495 SNPs that: (i) contained an rsid number that could be mapped to the hg19 genome build, (ii) overlapped  
496 with SNPs contained within the 1000 Genomes Project phase 3 genotype data, and (iii) had a minor allele  
497 frequency greater than 0.01 in Japanese (JPT) individuals in the 1000 Genomes Project. We used the 1000  
498 Genomes phase 3 data from 93 JPT individuals to estimate the linkage disequilibrium (LD) between SNPs  
499 in BioBank Japan for which we had the summary statistic data; LD was estimated separately for each of  
500 the 25 quantitative traits using the trait specific SNP arrays. LD estimates were used in the calculation of  
501 regional association statistics.  
502

## 503 **Population Architecture using Genomics and Epidemiology (PAGE) Study Data**

504 Summary statistics for genotyped and imputed SNPs in five admixed populations were downloaded from the  
505 Population Architecture using Genomics and Epidemiology (PAGE)<sup>6</sup> with permission granted via approval of  
506 manuscript proposal. We included summary statistics for up to 14 quantitative traits for African-American,  
507 Hispanic and Latin American, Native Hawaiian, American Indian/Alaska Native, and Asian ancestry cohorts  
508 when available. All AT/CG SNPs were omitted, and SNPs with an IMPUTE2 information score greater than  
509 0.8 were included in this analysis. Number of individuals and SNPs varied across ancestry-trait combinations  
510 and are given in Table S4 - Table S9.

511 Individuals from the 1000 Genomes Project phase 3<sup>92</sup> and the Human Genome Diversity Panel (HGDP)<sup>108</sup>  
512 were used to obtain LD estimates between SNPs for each ancestry cohort. To construct the LD reference  
513 panel for PAGE summary statistics from the African-American ancestry cohort, unrelated individuals from  
514 the 1000 Genomes Americans of African Ancestry in SW USA (ASW) and African Caribbeans in Barbados  
515 (ACB) were used. Only SNPs found in both the 1000 Genomes imputed data and PAGE summary statistics

516 files were used in gene-level association and heritability analyses. We used the same approach to compute  
517 reference LD estimates between SNPs for the Hispanic and Latin American, AIAN, and Asian ancestry  
518 cohorts, with the following 1000 Genomes reference population, respectively: Mexican Ancestry from Los  
519 Angeles USA (MXL) and Puerto Ricans from Puerto Rico (PUR); Colombians from Medellin, Colombia  
520 (CLM) and Peruvians from Lima, Peru (PEL); and the East Asian superpopulation (EAS). For the Native  
521 Hawaiian individuals from the PAGE study, there were no appropriate reference populations in the 1000  
522 Genomes data. In order to construct a reference LD matrix for the Native Hawaiian ancestry cohort, we  
523 randomly sampled individuals from populations in the most recent release of the HGDP proportional to the  
524 global ancestry of the Native Hawaiian cohort. The Native Hawaiian cohort's global ancestry proportions  
525 were determined using ADMIXTURE runs to be 47.89% Oceanian, 25.16% East Asian, 25.51% European,  
526 0.90% African, and 0.54% AIAN in a separate publication (Wojcik preprint - in prep.). We did not sam-  
527 ple from populations with less than 1% of the total ancestry in the admixture analysis referenced above.  
528 The resulting sample from which LD was estimated included 39 individuals from the Papuan Sepik in New  
529 Guinea and Melanesian in Bougainville, 14 individuals from the French in France, and 14 individuals from  
530 the Yoruba in Nigeria.

### 531 **WHI study cohort description**

532 The Women's Health Initiative (WHI) is a long-term, prospective, multi-center cohort study investigating  
533 post-menopausal women's health in the US. WHI was funded by the National Institutes of Health and the  
534 National Heart, Lung, and Blood Institute to study strategies to prevent heart disease, breast 124 cancer,  
535 colon cancer, and osteoporotic fractures in women 50-79 years of age. WHI involves 161,808 women recruited  
536 between 1993 and 1998 at 40 centers across the US. The study consists of two parts: the WHI Clinical Trial  
537 which was a randomized clinical trial of hormone therapy, dietary modification, and calcium/Vitamin D  
538 supplementation, and the WHI Observational Study, which focused on many of the inequities in women's  
539 health research and provided practical information about incidence, risk factors, and interventions related  
540 to heart disease, cancer, and osteoporotic fractures. For this project, women who self identified as European  
541 were excluded from the study sample (dbGaP study accession number: phs000227).

### 542 **HCHC/SOL study cohort description**

543 The Hispanic Community Health Study / Study of Latinos (HCHS/SOL) is a multi center study of His-  
544 panic/Latino populations with the goal of determining the role of acculturation in the prevalence and devel-  
545 opment of diseases, and to identify other traits that impact Hispanic/Latino health. The study is sponsored  
546 by the National Heart, Lung, and Blood Institute (NHLBI) and other institutes, centers, and offices of the

547 National Institutes of Health (NIH). Recruitment began in 2006 with a target population of 16,000 persons  
548 of Cuban, Puerto Rican, Dominican, Mexican or Central/South American origin. Household sampling was  
549 employed as part of the study design. Participants were recruited through four sites affiliated with San Diego  
550 State University, Northwestern University in Chicago, Albert Einstein College of Medicine in Bronx, New  
551 York, and the University of Miami. Researchers from seven academic centers provided scientific and logistical  
552 support. Study participants who were self-identified Hispanic/Latino and aged 18-74 years underwent ex-  
553 tensive psycho-social and clinical assessments during 2008-2011. A re-examination of the HCHS/SOL cohort  
554 is conducted during 2015-2017. Annual telephone follow-up interviews are ongoing since study inception to  
555 determine health outcomes of interest. (dbGaP study accession number: phs000555).

#### 556 **BioMe Biobank study cohort description**

557 The Charles Bronfman Institute for Personalized Medicine at Mount Sinai Medical Center (MSMC), BioMe™  
558 54 BioBank (BioMe) is an EMR-linked bio-repository drawing from Mount Sinai Medical Center consented  
559 patients which were drawn from a population of over 70,000 inpatients and 800,000 outpatients annually.  
560 The MSMC serves diverse local communities of upper Manhattan, including Central Harlem (86% African  
561 American), East Harlem (88% Hispanic/Latino), and Upper East Side (88% Caucasian/White) with broad  
562 health disparities. BioMe™ 58 enrolled over 26,500 participants from September 2007 through August  
563 2013, with 25% African American, 36% Hispanic/Latino (primarily of Caribbean origin), 30% Caucasian,  
564 and 9% of Other ancestry. The BioMe™ 60 population reflects community-level disease burdens and health  
565 disparities with broad public health impact. Biobank operations are fully integrated in clinical care pro-  
566 cesses, including direct recruitment from clinical sites waiting areas and phlebotomy stations by dedicate  
567 Biobank recruiters independent of clinical care providers, prior to or following a clinician standard of care  
568 visit. Recruitment currently occurs at a broad spectrum of over 30 clinical care sites. Study participants  
569 of self-reported European ancestry were not included in this analysis. (dbGaP study accession number:  
570 phs000925).

#### 571 **MEC study cohort description**

572 The Multiethnic Cohort (MEC) is a population-based prospective cohort study including approximately  
573 215,000 men and women from Hawaii and California. All participants were 45-75 years of age at baseline, and  
574 primarily of 5 ancestries: Japanese Americans, African Americans, European Americans, Hispanic/Latinos,  
575 and Native Hawaiians. MEC was funded by the National Cancer Institute in 1993 to examine lifestyle risk  
576 factors and genetic susceptibility to cancer. All eligible cohort members completed baseline and follow-up  
577 questionnaires. Within the PAGE II investigation, MEC proposes to study: 1) diseases for which we have



578 DNA available for large numbers of cases and controls (breast, prostate, and colorectal cancer, diabetes,  
579 and obesity); 2) common traits that are risk factors for these diseases (e.g., body mass index / weight,  
580 waist-to-hip ratio, height), and 3) relevant disease-associated biomarkers (e.g., fasting insulin and lipids,  
581 steroid hormones). The specific aims are: 1) to determine the population-based epidemiologic profile (al-  
582 lele frequency, main effect, heterogeneity by disease characteristics) of putative causal SNPs in the five  
583 racial/ethnic groups in MEC; 2) for SNPs displaying effect heterogeneity across ethnic/racial groups, we will  
584 utilize differences in LD to identify a more complete spectrum of associated SNPs at these loci; 3) investi-  
585 gate gene x gene and gene x environment interactions to identify modifiers; 4) examine the associations of  
586 putative causal SNPs with already measured intermediate phenotypes (e.g., plasma insulin, lipids, steroid  
587 hormones); and 5) for SNPs that do not fall within known genes, start to investigate their relationships with  
588 gene expression and epigenetic patterns in small genomic studies. For this project, MEC contributed African  
589 American, Japanese American, and Native Hawaiian samples.(dbGaP study accession number: phs000220).

## 590 **Fine-mapping analyses using the SuSiE framework**

591 To perform SNP-level fine mapping analyses on a given quantitative trait, we applied Sum of Single Effects  
592 (SuSiE) variable selection software<sup>34</sup>. SuSiE implements a Bayesian linear regression model on individual  
593 level data where sparse prior distributions are placed on the effect size of SNP and posterior inclusion  
594 probabilities (PIPs) are used to summarize their statistical relevance to the trait of interest. The software  
595 for SuSiE requires an input  $\ell$  which fixes the maximum number of causal SNPs to include in the model. In  
596 this work, we consider results when this parameter is chosen conservatively ( $\ell = 3000$ ). We used the three  
597 cohorts for which we had genotype data from the UK Biobank (African, European, and South Asian) to  
598 test whether there was effect size heterogeneity among ancestries in the 25 traits analyzed in this study. We  
599 first selected ten independent, non-overlapping subsamples of 10,000 individuals from the European ancestry  
600 cohort and filtered out any SNPs that had a minor allele frequency of less than 0.01. For each subsample, we  
601 then applied SuSiE to each of the 25 traits and compared the effect sizes and posterior inclusion probabilities.  
602 The average number of SNPs with an effect size greater than 0.01 and average number of SNPs with a PIP  
603 greater than 0.01 for each trait across the ten cohorts are reported in Table S11. Table S11 also reports the  
604 median correlation coefficient of effect sizes and PIPs among the 45 pairwise comparisons between the 10  
605 subsample cohorts.

606 We then applied SuSiE to the African and South Asian ancestry cohorts and compared their resulting  
607 effect sizes and PIPs to ten independent, non-overlapping subsamples of the European ancestry cohort. The  
608 number of SNPs with an effect size greater than 0.1 and PIPs greater than 0.01 in both the focal cohort

609 (either African or South Asian) and at least one of the ten European ancestry subsamples of the same size  
610 are reported in Table S12 and Table S13. Also reported in these tables, are the mean number of effect sizes  
611 greater than 0.01 and PIPs greater than 0.01 across the European ancestry subsamples for each trait and the  
612 number of unique effect sizes greater than 0.01 and PIPs greater than 0.01 that were only identified in the  
613 African or South Asian ancestry cohorts. Finally, Table S12 and Table S13 report the median correlation  
614 coefficient of the African or South Asian ancestry cohort effect sizes and PIPs with the ten European ancestry  
615 subsample cohorts of the same size.

## 616 S1 Description of the gene- $\varepsilon$ framework

617 A unique feature of gene- $\varepsilon$  is that it treats SNPs with spuriously associated nonzero effects as non-associated.  
618 gene- $\varepsilon$  assumes a reformulated null distribution of SNP-level effects  $\tilde{\beta}_j \sim \mathcal{N}(0, \sigma_\varepsilon^2)$ , where  $\sigma_\varepsilon^2$  is the SNP-  
619 level null threshold and represents the maximum proportion of phenotypic variance explained (PVE) by a  
620 spurious or non-associated SNP. This leads to the reformulated SNP-level null hypothesis  $H_0: \mathbb{E}[\beta_j^2] \leq \sigma_\varepsilon^2$ .  
621 To infer an appropriate  $\sigma_\varepsilon^2$ , gene- $\varepsilon$  fits a  $K$ -mixture of normal distributions over the regularized effect sizes  
622 with successively smaller variances (i.e.,  $\sigma_1^2 > \dots > \sigma_K^2 = 0$ ). In this study as in Cheng et al.<sup>38</sup>, we  
623 assume that associated SNPs will appear in the first set, while spurious and non-associated SNPs appear  
624 in the latter sets. As a final step, gene- $\varepsilon$  computes its gene-level association test statistic for the  $g$ -th gene  
625 by conformably partitioning the regularized GWA effect size estimates and computing the quadratic form  
626  $\tilde{Q}_g = \tilde{\beta}_g^T \tilde{\beta}_g$ . Corresponding  $p$ -values are then derived using Imhof's method. This assumes the common gene-  
627 level null  $H_0: Q_g = 0$ , where the null distribution of  $Q_g$  is dependent upon the eigenvalues from the scaled  
628 LD matrix  $\sigma_\varepsilon^2 \mathbf{\Sigma}$ . For details on implementation, validation and performance comparison with simulations,  
629 and empirical application to UK Biobank white British individuals in six traits, see Cheng et al.<sup>38</sup>.

## 630 S2 Regression with Summary Statistics (RSS) Enrichment.

Consider a GWA study with  $N$  individuals typed on  $P$  SNPs. For the  $j$ -th SNP, assume that we are given  
corresponding effect sizes  $\hat{\beta}_j$  and standard error  $\hat{s}_j$  via a single-SNP linear model fit using OLS. RSS then  
implements the following likelihood to model the GWA summary statistics<sup>60</sup>

$$\hat{\beta} \sim \mathcal{N}(\hat{\mathbf{S}}\mathbf{\Sigma}\hat{\mathbf{S}}^{-1}\beta, \hat{\mathbf{S}}\mathbf{\Sigma}\hat{\mathbf{S}}) \quad (1)$$

where  $\hat{\mathbf{S}} = \text{diag}(\hat{\mathbf{s}})$  is a  $J \times J$  diagonal matrix of standard errors,  $\mathbf{\Sigma}$  is again used to represent some empirical  
estimate of the LD matrix (i.e., using some external reference panel with ancestry matching the cohort of  
interest), and  $\beta$  are the true (unobserved) SNP-level effect sizes. To model gene-level enrichment, RSS

assumes the following hierarchical prior structure on the true effect sizes

$$\beta_j \sim \pi_j \mathcal{N}(0, \sigma_\beta^2) + (1 - \pi_j) \delta_0, \quad (2)$$

$$\sigma_\beta^2 = h^2 \left( \sum_{j=1}^J \pi_j N^{-1} \widehat{s}_j^{-2} \right)^{-1}, \quad (3)$$

$$\pi_j = \left( 1 + 10^{-(\theta_0 + a_j \theta)} \right)^{-1}, \quad (4)$$

where  $\delta_0$  is point mass centered at zero,  $h^2$  denotes the narrow-sense heritability of the trait,  $a_j$  is an indicator detailing whether the  $j$ -th SNP is inside a particular gene,  $\theta_0$  is the background proportion of trait-associated SNPs, and  $\theta$  reflects the increase in probability (on the  $\log_{10}$ -odds scale) when a SNP within a gene has non-zero effect. Here, the authors follow earlier works<sup>109</sup> and place independent uniform grid priors on the hyper-parameters  $\{h^2, \theta_0, \theta\}$ . Note that, unlike other methods, RSS does not calculate a  $P$ -value for assessing gene-level association. Instead, RSS produces a posterior enrichment probability that at least one SNP in a given gene boundary is associated with the trait

$$P_g := 1 - \Pr[\beta_j = 0, \forall j \in \mathcal{J}_g | \mathbf{D}] \quad (5)$$

631 where  $\mathbf{D}$  represents all of the input data including the GWA summary statistics  $\{\widehat{\boldsymbol{\beta}}, \widehat{\mathbf{s}}\}$ , the estimated LD  
 632 matrix  $\boldsymbol{\Sigma}$ , and any applicable SNP annotations or weights  $\mathbf{a} = (a_1, \dots, a_J)$ . See<sup>60,110</sup> for more details on  
 633 preferred hyper-parameter settings. As noted in the main text, RSS is relies on a Markov chain Monte Carlo  
 634 (MCMC) scheme for sampling posterior distributions and estimating model parameters. As a result, its  
 635 algorithm can be subject to convergence issues if these (or the random seed) are not chosen properly.

### 636 S3 SNP-set (Sequence) Kernel Association Test (SKAT).

The implementation of SKAT required access to raw phenotype  $\mathbf{y}$  and genotype  $\mathbf{X}$  information for  $N$  individuals typed on  $J$  SNPs. To assess enrichment of the  $|\mathcal{J}_g|$  variants within gene  $g$ , consider the linear model with sub-matrix  $\mathbf{X}_g$

$$\mathbf{y} = \beta_0 + \mathbf{X}_g \boldsymbol{\beta}_g + \mathbf{e}, \quad \mathbf{e} \sim \mathcal{N}(\mathbf{0}, \tau^2 \mathbf{I}) \quad (6)$$

where  $\beta_0$  is an intercept term,  $\boldsymbol{\beta}_g = (\beta_1, \dots, \beta_{|\mathcal{J}_g|})$  is a vector of regression coefficients for the SNPs within the gene of interest, and  $\mathbf{e}$  is a normally distributed error term with mean zero and scaled variance  $\tau^2$ . For model flexibility, gene-specific SNP effects  $\beta_j$  are assumed to follow an arbitrary distribution with mean

zero and marginal variances  $a_j\sigma_\beta^2$ , where  $\sigma_\beta^2$  is a variance component and  $a_j$  is a pre-specified weight for the  $j$ -th SNP. To this end, SKAT uses a variance component scoring approach and tests the null hypothesis  $H_0: \beta = \mathbf{0}$ , or equivalently  $H_0: \sigma_\beta^2 = 0$ . The corresponding gene-level test statistic  $\widehat{Q}_g$  then takes on the familiar quadratic form

$$\widehat{Q}_g = (\mathbf{y} - \widehat{\beta}_0)^\top \mathbf{K}_g (\mathbf{y} - \widehat{\beta}_0) \quad (7)$$

where  $\widehat{\beta}_0$  is the predicted mean of trait under the null hypothesis, and is computed by projecting  $\mathbf{y}$  onto the column space of the intercept (i.e., a vector of ones). The term  $\mathbf{K}_g = \mathbf{X}_g \mathbf{A}_g \mathbf{A}_g \mathbf{X}_g^\top$  is commonly referred to as an  $N \times N$  kernel matrix, where  $\mathbf{A}_g = \text{diag}(a_1, \dots, a_{|\mathcal{J}_g|})$  is used to denote a diagonal weight matrix that changes for each gene  $g$ . Each element of  $\mathbf{K}_g$  is computed via the linear kernel function

$$k(\mathbf{x}_i, \mathbf{x}_{i'}) = \sum_{j=1}^{|\mathcal{J}_g|} a_j x_{ij} x_{i'j}. \quad (8)$$

637 While implementing SKAT in this work, we follow previous works and set each weight to be  $\sqrt{a_j} =$   
638  $\text{Beta}(\text{MAF}_j, 1, 25)$  — the beta distribution density function with pre-specified parameters evaluated at the  
639 sample minor allele frequency (MAF) for the  $j$ -th SNP in the gene region. For more details, see<sup>40,111–113</sup>.

## 640 **Clustering traits sharing a core set of associated genes using the WINGS algo-** 641 **rithm**

642 We used the WINGS algorithm<sup>114</sup> to identify clusters of traits sharing a core set of genes enriched for  
643 associated mutations. WINGS takes as input a gene ( $M$ ) by trait ( $N$ ) matrix and uses the Ward distance  
644 metric to find the distance among vectors of gene scores for different phenotypes; in this study, we used  
645 gene- $\varepsilon$  gene-level association statistics as the input to WINGS. The more significantly associated genes that  
646 two traits share, the closer they will be in the gene-dimensional space. Applying WINGS to a matrix of  
647 gene scores for each ancestry separately, we examined whether the same traits clustered together, separately  
648 in each ancestry. We constructed matrices of gene- $\varepsilon$  gene-level association statistics for the UK Biobank  
649 European, African, South Asian (from the UK Biobank) and East Asian (Biobank Japan) ancestry cohorts.  
650 Each of these matrices contained gene-level association statistics for all 25 quantitative traits of interest.  
651 The total number of genes and regulatory regions included were: European (23,603), African (23,575),  
652 South Asian(23,671), and East Asian (21,435). For the East Asian ancestry cohort, we limited the genes  
653 to the intersection of genes with gene- $\varepsilon$  gene-level association statistics across all 25 traits. The number  
654 of gene scores calculated for each trait in the East Asian ancestry cohort varies due to the heterogeneity

655 in imputed and genotype SNP arrays in the Biobank Japan studies (Table S4 and Table S15). Figure S19  
656 shows the resulting dendrograms displaying prioritized phenotypes identified using the WINGS algorithm on  
657 each cohort's gene score matrix. The WINGS algorithm is designed to run on 25 phenotypes or more (see  
658 McGuirl et al.<sup>114</sup> for details), and we therefore did not apply the WINGS algorithm to the AIAN, Native  
659 Hawaiian, or Hispanic and Latin American cohorts as there was not data for enough phenotypes (Table  
660 S5-Table S9).

## 661 **Analysis of GWAS Catalog Metadata and Previous GWA Publications**

662 We cross-referenced our results from association testing at multiple genomic scales against previously pub-  
663 lished results in the GWAS catalog (<https://www.ebi.ac.uk/gwas/>) and in PubMed using the following  
664 processes.

665 In order to collect PubMed IDs (PMIDs) for publications associated with the UK Biobank, a two-part  
666 data collection process was used. The first process was to directly search for publications with variations  
667 of the term "UK Biobank" (e.g., U.K. Biobank, United Kingdom Biobank) from PubMed using the Entrez  
668 Programming Utilities (E-Utilities) API. The E-Utilities API is the public API to the NCBI Entrez sys-  
669 tem and allows direct access to all Entrez databases including PubMed. Search queries were formulated by  
670 narrowing publications using year published and then further narrowing to those publications with varia-  
671 tions of the search term "UK Biobank" in either the title or abstract. The open-source Python package  
672 Entrez (<https://biopython.org/DIST/docs/api/Bio.Entrez-module.html>) from the Biopython Project  
673 was used to facilitate interaction with the E-Utilities API.

674 The second data collection process was to gather information from publications listed directly on the UK  
675 Biobank website (<https://www.ukbiobank.ac.uk/>). Since the majority of publications on the website did  
676 not have an easily accessible PMID, identifying information including publication title and year was scraped  
677 and used to retrieve a publication's corresponding PMID (again using the E-Utilities API). The HTML/XML  
678 document parsing Python library BeautifulSoup ([https://www.crummy.com/software/BeautifulSoup/  
679 bs4/doc/](https://www.crummy.com/software/BeautifulSoup/bs4/doc/)) was used to parse the HTML of the various UK Biobank webpages, and the Python Requests  
680 library (<https://requests.readthedocs.io/en/master/>) was used to programatically send HTTP calls  
681 to the server hosting the website. PMIDs were retrieved directly from the XML output of the E-Utilities  
682 API calls.

683 The PMIDs retrieved from both processes were aggregated into a single set of unique PMIDs, as some  
684 publications were identified by both processes. Publications that could not get associated PMIDs from the  
685 second data collection process were flagged for manual processing. The PMIDs that were retrieved from

686 PubMed directly but could not be found based on the publication information provided on the UK Biobank  
687 website were noted. Conversely, the PMIDs that could be retrieved from publication information found on  
688 the UK Biobank website but not directly from PubMed were also noted.

689 Using the compiled list of PMIDs, analyses of the UK Biobank data set reported in the GWAS cat-  
690 alog association data were compiled. Previous genotype-to-phenotype association data and sample an-  
691 cestry descriptions were downloaded from <https://www.ebi.ac.uk/gwas/docs/file-downloads>. Unique  
692 genotype-to-phenotype associations were parsed using a set of custom python scripts. All scripts used  
693 in the curation of PMIDs, parsing of GWAS catalog summary data, and determination of previously  
694 published genotype-to-phenotype associations from UK Biobank studies are available on GitHub ([https://github.com/ramachandran-lab/redefining\\_replication](https://github.com/ramachandran-lab/redefining_replication)).

## 696 **Simulation design to test the power and false discovery rate of GWA and gene-** 697 **level association analyses**

### 698 **Simulations of a single population**

699 In our simulation studies, we used the following general simulation scheme to generate quantitative traits  
700 using real genotype data on chromosome 1 from  $N$  randomly sampled individuals of European ancestry in  
701 the UK Biobank. This pipeline follows from previous studies<sup>38,78</sup>. We will use  $\mathbf{X}$  to denote the  $N \times J$   
702 genotype matrix, with  $J$  denoting the number of single nucleotide polymorphisms (SNPs) encoded as 0, 1, 2  
703 copies of a reference allele at each locus and  $\mathbf{x}_j$  representing the genotypic vector for the  $j$ -th SNP. Following  
704 quality control procedures detailed in the Supplemental Information, our simulations included  $J = 36,518$   
705 SNPs distributed across genome. We used the NCBI's RefSeq database in the UCSC Genome Browser to  
706 assign SNPs to genes which resulted in  $G = 1,408$  genes in the simulation studies.

707 After the annotation step, we simulated phenotypes by first assuming that the total phenotypic variance  
708  $\mathbb{V}[\mathbf{y}] = 1$ , and that all observed genetic effects explained a fixed proportion of this value (i.e., narrow-sense  
709 heritability,  $h^2$ ). Next, we randomly selected a certain percentage of genes to be enriched for associations  
710 and denoted the sets of SNPs that they contained as  $\mathcal{C}$ . Within  $\mathcal{C}$ , we selected causal SNPs in a way such  
711 that each associated gene at least contains one SNP with non-zero effect size. Quantitative continuous traits  
712 were then generated under the following two general linear models:

713 • Standard Model:  $\mathbf{y} = \sum_{c \in \mathcal{C}} \mathbf{x}_c \beta_c + \mathbf{e}$

714 • Population Structure Model:  $\mathbf{y} = \mathbf{W}\mathbf{b} + \sum_{c \in \mathcal{C}} \mathbf{x}_c \beta_c + \mathbf{e}$

715 where  $\mathbf{y}$  is an  $N$ -dimensional vector containing all the phenotype states;  $\mathbf{x}_c$  is the genotype for the  $c$ -th

716 causal SNP;  $\beta_c$  is the additive effect size for the  $c$ -th SNP; and  $\mathbf{e} \sim \mathcal{N}(0, \tau^2 \mathbf{I})$  is an  $N$ -dimensional vector  
717 of normally distributed environmental noise. Additionally, in the model with population structure,  $\mathbf{W}$  is an  
718  $N \times M$  matrix of the top  $M = 10$  principal components (PCs) from the genotype matrix and represents  
719 additional population structure with corresponding fixed effects  $\mathbf{b}$ . The effect sizes of SNPs in genes enriched  
720 for associations are randomly drawn from standard normal distributions and then rescaled so they explain  
721 a fixed proportion of the narrow-sense heritability  $\mathbb{V}[\sum \mathbf{x}_c \beta_c] = h^2$ . The coefficients for the genotype PCs  
722 are also drawn from standard normal distributions and rescaled such that  $\mathbb{V}[\mathbf{W}\mathbf{b}] = 10\%$  of the total  
723 phenotypic variance, with the variance of all non-genetic effects contributing  $\mathbb{V}[\mathbf{W}\mathbf{b}] + \mathbb{V}[\mathbf{e}] = (1 - h^2)$ . For  
724 any simulations conducted under the population structure model, genotype PCs are not included in any of  
725 the model fitting procedures, and no other preprocessing normalizations were carried out to account for the  
726 additional population structure. More specifically, GWA summary statistics are then computed by fitting a  
727 single-SNP univariate linear model via ordinary least squares (OLS):

$$\hat{\beta}_j = (\mathbf{x}_j^T \mathbf{x}_j)^{-1} \mathbf{x}_j^T \mathbf{y}; \quad (9)$$

728 for every SNP in the data  $j = 1, \dots, J$ . These OLS effect size estimates, along with an empirically LD matrix  
729  $\Sigma$  computed directly from the full  $N \times J$  genotype matrix  $\mathbf{X}$ , are given to gene- $\varepsilon$ . We also retain standard  
730 errors and  $p$ -values for the implementation of competing methods: RSS<sup>60</sup>, SKAT<sup>40</sup>, and the standard GWA  
731 SNP-level association test. Given the simulation procedure above, we simulate a wide range of scenarios for  
732 comparing the performance of gene-level association approaches by varying the following parameters:

- 733 • Number of individuals:  $N = 5,000$  and  $10,000$ ;
- 734 • Narrow-sense heritability:  $h^2 = 0.2$  and  $0.6$ ;
- 735 • Percentage of enriched genes: 1% (sparse) and 10% (polygenic);

736 Furthermore, we set the number of causal SNPs with non-zero effects to be some fixed percentage of all SNPs  
737 located within the designated genes enriched for associations. We set this percentage to be 0.125% in the  
738 1% associated SNP-set case, and 3% in the 10% associated SNP-set case. All performance comparisons are  
739 based on 100 different simulated runs for each parameter combination. Lastly, for each simulated dataset,  
740 we also selected some number of intergenic SNPs (i.e., SNPs not mapped to any gene) to have non-zero  
741 effect sizes. This was done to mimic genetic associations in unannotated regulatory elements. Specifically,  
742 five randomly selected intergenic SNPs were given non-zero contributions to the trait heritability in the 1%  
743 enriched genes case, and 30 intergenic SNPs were selected in the 10% enriched genes case.

744 All performance comparisons are based on 100 different simulated runs for each parameter combina-  
745 tion. We computed gene-level  $p$ -values for gene- $\epsilon$ , SKAT, and the single-SNP GWAS. For evaluating the  
746 performance of RSS, we compute posterior enrichment probabilities. For all approaches, we assessed the  
747 power and false discovery rates when identifying enriched genes at either a Bonferroni-corrected threshold  
748 ( $p = 0.05/1,408$  genes =  $3.55 \times 10^{-5}$ ) or according to the median probability model (posterior enrichment  
749 probability  $> 0.5$ )<sup>115</sup>. Figure S4 and Figure S5 show the mean performances (and standard errors) across  
750 all simulated replicates.

### 751 **Simulations of genetic trait architecture in two populations**

752 We used the African (UKB) cohort and a subset of the European cohort and simulation studies to test the  
753 ability of GWAS and gene- $\epsilon$  to detect shared causal SNPs (in the case of gene- $\epsilon$ , genes containing causal  
754 SNPs) in a multi-ancestry study. Using the same simulation protocol as that described for testing power of  
755 different enrichment analysis methods, described in *Simulations in a single population*, we labeled all genes  
756 containing at least one causal SNP as "causal". We first determined the power of gene- $\epsilon$  to identify SNPs or  
757 genes that are causal in each cohort under a variety of genomic architectures. The total amount of variance  
758 explained in the phenotype by the causal SNPs (i.e. the narrow-sense heritability) to be equal to 0.2 or  
759 0.6. In each of these contexts, the sparsity of causal variants as a function of the total number of variants  
760 was set to either 0.1 or 0.5. These values of causal SNP sparsity were selected in order to ensure that an  
761 ample number of SNPs were associated with the phenotype in both cohorts. Finally, the overlap in causal  
762 SNPs between the two cohorts was tested at proportions equal to 0 (no overlap in causal between SNPs  
763 cohorts) 0.25, 0.5, and 1 (complete overlap in causal SNPs between cohorts). For each of these parameter  
764 sets, 50 replicate simulations were performed of two cohorts derived from 10,000 European individuals and  
765 4,967 African individuals, respectively. We summarize the performance of the standard GWA framework  
766 and gene- $\epsilon$  across the parameter space. Generally, gene- $\epsilon$  performs better on the European cohort than it  
767 does in the African cohort, but is better powered in the African cohort when the causal SNPs are the same  
768 in both cohorts (Figure S6 and Figure S7). Additionally, gene- $\epsilon$  performs better when identifying causal  
769 genes that are shared between the two cohorts - particularly when traits have high heritability Figure S8 -  
770 Figure S9.



## 771 Appendix

### 772 S4 SNP-level results for height and C-reactive protein

773 In Figure S3a and Figure S3d, we found that, across 25 traits analyzed, height had the greatest number  
774 of genome-wide significant SNP-level associations (76,910 unique associations) in at least one ancestry. Of  
775 these SNP-level associations, 8.90% (7,377 SNPs) replicate based off of rsID in at least two ancestry cohorts.  
776 Height is not the only trait in which the standard GWA SNP-level association test detects associations that  
777 replicate extensively across ancestries. In fact, SNP-level associations replicate in each of the 25 continuous  
778 traits that we analyze in this study.

779 We analyzed SNP-level associations with C-reactive protein in six ancestry cohorts: African-American  
780 (PAGE), European, South Asian, East Asian, Native Hawaiian, and Hispanic and Latin American cohorts.  
781 C-reactive protein is an example of a trait with a sparse and highly conserved genetic architecture across  
782 ancestries, as shown in Figure 2. Many SNPs within the *CRP* gene have been previously associated with  
783 C-reactive protein plasma levels<sup>116–118</sup>. In our analysis, rs3091244 is genome-wide significant in only the  
784 European ancestry cohort, and has been functionally validated as influencing C-reactive protein levels<sup>52,53</sup>.  
785 The SNP rs3091244 is located in a promoter region slightly upstream of *CRP*, and it has clinical implications  
786 for both atrial fibrillation<sup>119</sup> and lupus erythematosus<sup>120</sup> (European  $p = 1.54 \times 10^{-116}$ ; East Asian  $p =$   
787  $1.15 \times 10^{-9}$ ).

788 We expanded our search for replicated GWA SNP-level association signals across ancestry cohorts by  
789 scanning for 1 Mb regions that contained associations to the same phenotype in two or more ancestries—  
790 a process often referred to as “clumping”. These windows were centered at every unique genome-wide  
791 significant SNP in any ancestry for a given trait (we refer the 1Mb window around the significant SNP as a  
792 “clump”, Figure S3b and Figure S3e). In addition to the largest number of unique SNP-level associations,  
793 height also had the largest proportion of clumps containing a significant SNP-level GWA association signal  
794 that replicated in at least two ancestry cohorts (see Figure S3b and Figure S3e). The three traits with the  
795 greatest proportion of clumps containing SNP-level GWA signals that replicate in multiple ancestry cohorts  
796 were height (77.09% of clumps), urate (65.89%), and low density lipoprotein (54.40%).

797 In addition to the SNP-level associations on chromosome 1 surrounding the *CRP* gene across all six  
798 ancestry cohorts (displayed in Figure 2), there are other regions of the genome that contain significant GWA  
799 associations with C-reactive protein that replicate in multiple ancestry cohorts. On chromosome 2, there is a  
800 cluster of four SNPs significantly associated with C-reactive protein levels in the European, East Asian, and  
801 Hispanic and Latin American ancestry cohorts. Of these, rs1260326 (European  $p = 1.01 \times 10^{-55}$ ; East Asian  
802  $p = 1.70 \times 10^{-9}$ ; Hispanic and Latin American  $p = 1.24 \times 10^{-20}$ ), rs780094 (European  $p = 9.95 \times 10^{-51}$ ;

803 East Asian  $p = 1.70 \times 10^{-9}$ ; Hispanic and Latin American  $p = 1.14 \times 10^{-16}$ ), and rs6734238 (African-  
804 American (PAGE)  $p = 3.04 \times 10^{-10}$ ; European  $p = 8.38 \times 10^{-34}$ ; South Asian  $p = 2.17 \times 10^{-9}$ ) were  
805 statistically significant in three of the six ancestry cohorts that we analyzed. Each of these three SNPs has  
806 been previously associated with C-reactive protein levels in a European ancestry cohort<sup>121-123</sup>. Of these  
807 three SNPs, only one (rs6734238) had previously been replicated in other ancestries (in African-American,  
808 and Hispanic and Latin American cohorts<sup>124</sup>).

809 On chromosome 19 there are 23 SNPs that are associated with CRP in the African-American PAGE,  
810 European, and Hispanic and Latin American ancestry cohorts. Two other SNPs are associated with C-  
811 reactive protein in the African-American (PAGE), European, and Hispanic and Latin American cohorts,  
812 as well as the East Asian ancestry cohort. One of these two SNPs, rs7310409 (African-American (PAGE)  
813  $p = 8.57 \times 10^{-9}$ ; European  $p = 3.57 \times 10^{-210}$ ; East Asian  $p = 2.72 \times 10^{-27}$ ; Hispanic and Latin American  $p =$   
814  $5.35 \times 10^{-29}$ ) located in the HNF1 homeobox A (*HNF1A*) gene, has been previously associated with C-reactive  
815 protein levels in only a European ancestry cohort<sup>122,123</sup>. Three additional significant SNPs in our analysis  
816 have been previously associated with European ancestry cohorts in previous studies, including: rs1169310<sup>124</sup>  
817 (European  $p = 1.52 \times 10^{-172}$ ; East Asian  $p = 1.28 \times 10^{-18}$ ; Hispanic and Latin American  $p = 1.17 \times 10^{-27}$ ),  
818 rs1183910<sup>121,125</sup> (European  $p = 5.50 \times 10^{-177}$ ; East Asian  $p = 3.16 \times 10^{-29}$ ; Hispanic and Latin American  
819  $p = 7.47 \times 10^{-29}$ ), and rs7953249<sup>126</sup> (European  $p = 1.19 \times 10^{-177}$ ; East Asian  $p = 1.10 \times 10^{-19}$ ; Hispanic  
820 and Latin American  $p = 4.80 \times 10^{-29}$ ). Two SNPs, rs2259816 (European  $p = 2.77 \times 10^{-172}$ ; East Asian  
821  $p = 9.33 \times 10^{-18}$ ; Hispanic and Latin American  $p = 1.90 \times 10^{-27}$ ) and rs7979473 (African  $p = 1.49 \times 10^{-9}$ ;  
822 East Asian  $p = 6.06 \times 10^{-29}$ ; Hispanic and Latin American  $p = 1.56 \times 10^{-30}$ ), have been previously associated  
823 with C-reactive protein in both African-American and Hispanic and Latin American ancestry cohorts<sup>124</sup>.  
824 There is one final group of three SNPs associated with C-reactive protein in the African-American (PAGE),  
825 European, East Asian, and Hispanic and Latin American ancestry cohorts on chromosome 19. One of them,  
826 rs4420638 (East Asian  $p = 9.93 \times 10^{-29}$ ; Hispanic and Latin American  $p = 2.03 \times 10^{-30}$ ), has been previously  
827 associated in a European ancestry cohort<sup>121,123,125</sup>. These four regions indicate a highly conserved SNP-  
828 level architecture of C-reactive protein across six ancestry cohorts. Interestingly, we were unable to replicate  
829 associations with C-reactive protein across ancestries at the gene or pathway levels.

## 830 Gene and pathway association results

831 Three genes, *GP6*, *RDH13*, and *AGPAT5*, were significantly associated with platelet count (PLC) in the  
832 African-American (PAGE) ancestry cohort and the East Asian ancestry cohort (Figure S10). Of these, no  
833 significant SNPs in the glycoprotein VI platelet (*GP6*) gene have been reported in the GWAS catalog for

834 either ancestry cohort. However, a single SNP within *GP6*, rs1613662, has previously been associated with  
835 mean platelet volume in a GWA study analyzing a European ancestry cohort<sup>127</sup>. *GP6* plays a critical  
836 role in platelet aggregation, and mutations have been previously associated with fetal loss<sup>128</sup>. The retinol  
837 dehydrogenase 13 (*RDH13*) gene has no reported GWAS catalog associations with platelet count, but is  
838 within 60kb of a SNP significantly associated with platelet aggregation<sup>129</sup>. Of the three genes significantly  
839 associated with PLC in both the European and AIAN cohorts, 1-Acylglycerol-3-Phosphate O-Acyltransferase  
840 5 (*AGPAT5*) is a member of a gene family known to play a role in immunity and inflammation response<sup>130</sup>.

841 Alcohol dehydrogenase 2 (*ALDH2*) has additionally been associated with hypertension in an elderly  
842 Japanese cohort<sup>131</sup>. A member of the RAS oncogene family (*RAB8A*) has been shown to play a role in the  
843 inhibition of inflammatory response. In contrast, the cut like homeobox 2 *CUX2* gene contains a significantly  
844 associated SNP in the array used in this study for the East Asian ancestry cohort, but it has no previous  
845 associations in a European ancestry cohort. However, *CUX2* is significantly associated at the gene-level in  
846 both the European and East Asian ancestry cohorts. Although not reported as being associated with PLC  
847 in the GWAS Catalog, a single SNP, rs61745424 which encodes a missense mutation, has been previously  
848 identified as being related to the trait<sup>132</sup>. The gene- $\varepsilon$  association statistics for the seven genes significantly  
849 associated with PLC are available in Table S17.

850 Finally, a single gene, acyl-CoA dehydrogenase family member 10 (*ACAD10*) associated in our gene-level  
851 analysis of PLC, was significant in both the European and East Asian ancestry cohorts (European gene-  
852  $\varepsilon$   $p = 1.47 \times 10^{-10}$ ; East Asian gene- $\varepsilon$   $p = 2.00 \times 10^{-10}$ ) but contained no previous associations in the GWAS  
853 catalog. The African-American and Hispanic and Latin American ancestry cohorts analyzed in Qayyum  
854 et al.<sup>133</sup> both contain SNPs within *ACAD10* that are significantly associated with PLC.

855 In our analysis of triglyceride levels in six ancestry cohorts (African-American (PAGE), European, East  
856 Asian, South Asian, Hispanic and Latin American, and Native Hawaiian), we identified shared genetic  
857 architecture at the SNP, gene, and subnetwork level. Replicated SNPs and genes between the six ancestry  
858 cohorts are shown in Figure S12-Figure S13. We focus our discussion of results at the network level in  
859 the European, East Asian, and Native Hawaiian ancestry cohorts (Figure 3). In the European and East  
860 Asian ancestry cohorts, we identified 55 shared genome-wide significant associations at the gene-level. Of  
861 these results, eight genes lie in the same significantly mutated subnetwork (Hierarchical HotNet  $p < 10^{-3}$ )  
862 when analyzing each ancestry cohort independently. Five of those eight genes belong to the apolipoprotein  
863 family of genes, including: apolipoprotein A1 (*APOA1*), apolipoprotein A4 (*APOA4*), apolipoprotein A5  
864 (*APOA5*), apolipoprotein C3 (*APOC3*), apolipoprotein E (*APOE*). Specifically, the apolipoprotein play a  
865 central role in lipoprotein biosynthesis and transport. All of these genes contain SNPs previously associated  
866 with triglyceride levels in a European ancestry cohort<sup>98-103</sup>. All five genes also contain SNPs previously

867 associated with triglyceride levels in non-European ancestry cohorts. Specifically, *APOA1*, *APOC3*, and  
868 *APOE* each contain SNPs previously associated with triglyceride levels in African-American and Hispanic  
869 and Latin American ancestry cohorts<sup>98,99</sup>. The *APOA5* gene has previously been associated to triglyceride  
870 levels in an East Asian, African-American, and Hispanic and Latin American ancestry cohorts<sup>100,104</sup>.

871 The other three genes that were significantly associated with triglyceride levels in the European and  
872 East Asian ancestry cohorts are members of the largest significantly mutated subnetwork including phos-  
873 pholipid transfer protein (*PLTP*; European gene- $\epsilon$   $p = 4.29 \times 10^{-9}$ ; East Asian gene- $\epsilon$   $p = 6.66 \times 10^{-15}$ ),  
874 lipoprotein lipase (*LPL*; European gene- $\epsilon$   $p = 4.08 \times 10^{-13}$ ; East Asian gene- $\epsilon$   $p = 1.00 \times 10^{-20}$ ), and  
875 angiopoietin like 3 (*ANGPTL3*; European gene- $\epsilon$   $p = 8.86 \times 10^{-8}$ ; East Asian gene- $\epsilon$   $p = 1.00 \times 10^{-20}$ ).  
876 *PLTP* has previously been associated with triglyceride levels in European, African-American, and Hispanic  
877 and Latin American ancestry cohorts<sup>97-100,105-107,134</sup>. *LPL* is one of the most well-studied genes in the  
878 regulation of triglyceride levels. It has previously been associated with triglyceride levels in European an-  
879 cestry cohorts<sup>96-103,105-107,134,134-146</sup>, East Asian ancestry cohorts<sup>104,147</sup>, and African ancestry cohorts as  
880 well as Hispanic and Latin American ancestry cohorts<sup>96-98,100,105,145,146,148,149</sup>. The final gene that was  
881 genome-wide significant in both the European and East Asian ancestry cohorts, *ANGPTL3*, has no previ-  
882 ous associations in the GWAS catalog and presents a novel candidate gene within the network. While not  
883 significant in any gene-level analysis, the gene *ANGPTL4* (European gene- $\epsilon$   $p = 1.00 \times 10^{-20}$ ; East Asian  
884 gene- $\epsilon$   $p = 9.99 \times 10^{-1}$ ) from the same family is present in the largest subnetwork in the European cohort  
885 and also has also been previously identified as having associations in European, African, and Hispanic and  
886 Latin American ancestry cohorts<sup>96,97,100,145,146,150</sup>.

887 In our analysis of the European ancestry cohort from the UK Biobank, we additionally identified a set of  
888 eight genes that are connected to the core network discussed above. One of these genes is *ANGPTL4*, which  
889 we discussed above. Five of these genes were significant at the gene-level in the European ancestry cohort,  
890 including four apolipoprotein genes (*APOC1*; European gene- $\epsilon$   $p = 1.67 \times 10^{-16}$ , *APOC2*; European gene-  
891  $\epsilon$   $p = 3.57 \times 10^{-13}$ , *APOC4*; European gene- $\epsilon$   $p = 3.72 \times 10^{-13}$ , and *APOB*; European gene- $\epsilon$   $p = 1.00 \times 10^{-20}$ )  
892 and lipase maturation factor 1 (*LMF1*; European gene- $\epsilon$   $p = 8.03 \times 10^{-7}$ ). Each of these genes have been  
893 previously associated with triglyceride levels in a European ancestry cohort<sup>100</sup>. Additional associations  
894 were also found in that same study which conducted a meta-analysis of European, African-American, and  
895 Hispanic and Latin American ancestry cohorts. The final two genes included in the significantly mutated  
896 subnetwork of the European ancestral cohort, *APOL1* and *HBA1*, were not were not identified as genome-  
897 wide significant by gene- $\epsilon$  and have no previous SNP-level associations with triglyceride levels in the GWAS  
898 Catalog. Interestingly, both *APOL1* (Native Hawaiian gene- $\epsilon$   $p = 8.89 \times 10^{-11}$ ) and *HBA1* (Native Hawaiian  
899 gene- $\epsilon$   $p = 2.46 \times 10^{-10}$ ) were both identified as genome-wide significant by gene- $\epsilon$  in our analysis of the Native

900 Hawaiian ancestry cohort and the interaction between them was identified in our Hierarchical HotNet<sup>31</sup>  
901 analysis as present in both the European and Native Hawaiian ancestry cohorts.

902 In addition to *APOL1* and *HBA1*, six more genes are connected to the core network of genes that overlap  
903 in the East Asian and European significantly mutated subnetworks. Of these, both *HBA2* and *B4GALT3* are  
904 significant at the gene-level in the Native Hawaiian ancestry cohort alone. They are each connected to genes  
905 identified in both the European and Native Hawaiian ancestry cohorts as members of the largest significantly  
906 mutated subnetwork. The final three genes include kallikrein related peptidase 8 (*KLK8*), pancreatic lipase  
907 *PNLIP*, and wnt family member 4 (*WNT4*) which were not significant at the gene-level and did not contain  
908 previous SNP-level associations in the GWAS catalog.

909 In the largest significantly mutated subnetwork identified in our analysis of the East Asian ancestry  
910 cohort, we identified seven genes that were not shared by the networks in other ancestry cohorts. One of these  
911 genes, beta-secretase 1 (*BACE1*; East Asian gene- $\epsilon$   $p = 3.57 \times 10^{-13}$ ; European gene- $\epsilon$   $p = 5.55 \times 10^{-17}$ ),  
912 was significant at the gene-level but contained no previously associated SNPs in any cohort in the GWAS  
913 catalog. *BACE1* plays a role in the metabolism of amyloid beta precursor protein<sup>151</sup>. Three of the genes  
914 within the network identified in the East Asian ancestry cohort contain previously associated SNPs in  
915 both European and non-European ancestry cohorts, including: cholesteryl ester transfer protein (*CETP*),  
916 proprotein convertase subtilisin/kexin type 6 (*PCSK6*), and proprotein convertase subtilisin/kexin type 7  
917 (*PCSK7*)<sup>98,105</sup>. The final three genes in the significantly mutated subnetwork identified in the East Asian  
918 ancestry cohort were not significant at the gene-level and do not contain previously associated SNPs in the  
919 GWAS catalog in any ancestral cohort. Lecithin-cholesterol acyltransferase (*LCAT*) is involved in cholesterol  
920 biosynthesis and apolipoprotein F (*APOF*) encodes one of the minor apolipoprotein genes present in plasma.  
921 Finally, tyrosine-protein kinase receptor 3 (*TYRO3*) plays a role in ligand recognition and cell metabolism<sup>152</sup>.  
922 The gene- $\epsilon$   $p$ -values in each ancestry cohort for each of the 28 genes discussed here are shown in Table S18.

## 923 References

- 924 [1] Akiko Nagai, Makoto Hirata, Yoichiro Kamatani, Kaori Muto, Koichi Matsuda, Yutaka Kiyohara,  
925 Toshiharu Ninomiya, Akiko Tamakoshi, Zentaro Yamagata, Taisei Mushiroda, et al. Overview of the  
926 biobank japan project: study design and profile. *Journal of epidemiology*, 27(Supplement.III):S2–S8,  
927 2017.
- 928 [2] Alicia R Martin, Christopher R Gignoux, Raymond K Walters, Genevieve L Wojcik, Benjamin M  
929 Neale, Simon Gravel, Mark J Daly, Carlos D Bustamante, and Eimear E Kenny. Human demographic

- 930 history impacts genetic risk prediction across diverse populations. The American Journal of Human  
931 Genetics, 100(4):635–649, 2017.
- 932 [3] Pamela L Sankar and Lisa S Parker. The precision medicine initiative’s all of us research program: an  
933 agenda for research on its ethical, legal, and social issues. Genetics in Medicine, 19(7):743–750, 2017.
- 934 [4] Clare Bycroft, Colin Freeman, Desislava Petkova, Gavin Band, Lloyd T Elliott, Kevin Sharp, Allan  
935 Motyer, Damjan Vukcevic, Olivier Delaneau, Jared O’Connell, et al. The uk biobank resource with  
936 deep phenotyping and genomic data. Nature, 562(7726):203–209, 2018.
- 937 [5] Alicia R Martin, Masahiro Kanai, Yoichiro Kamatani, Yukinori Okada, Benjamin M Neale, and Mark J  
938 Daly. Clinical use of current polygenic risk scores may exacerbate health disparities. Nature genetics,  
939 51(4):584, 2019.
- 940 [6] Genevieve L Wojcik, Mariaelisa Graff, Katherine K Nishimura, Ran Tao, Jeffrey Haessler, Christo-  
941 pher R Gignoux, Heather M Highland, Yesha M Patel, Elena P Sorokin, Christy L Avery, et al. Genetic  
942 analyses of diverse populations improves discovery for complex traits. Nature, 570(7762):514–518, 2019.
- 943 [7] Mashaal Sohail, Robert M Maier, Andrea Ganna, Alex Bloemendal, Alicia R Martin, Michael C  
944 Turchin, Charleston WK Chiang, Joel Hirschhorn, Mark J Daly, Nick Patterson, et al. Polygenic  
945 adaptation on height is overestimated due to uncorrected stratification in genome-wide association  
946 studies. Elife, 8:e39702, 2019.
- 947 [8] Jeremy J Berg, Arbel Harpak, Nasa Sinnott-Armstrong, Anja Moltke Joergensen, Hakhamanesh  
948 Mostafavi, Yair Field, Evan August Boyle, Xinjun Zhang, Fernando Racimo, Jonathan K Pritchard,  
949 et al. Reduced signal for polygenic adaptation of height in uk biobank. Elife, 8:e39725, 2019.
- 950 [9] Pak C Sham and Shaun M Purcell. Statistical power and significance testing in large-scale genetic  
951 studies. Nature Reviews Genetics, 15(5):335–346, 2014.
- 952 [10] Alkes L Price, Chris CA Spencer, and Peter Donnelly. Progress and promise in understanding the  
953 genetic basis of common diseases. Proceedings of the Royal Society B: Biological Sciences, 282(1821):  
954 20151684, 2015.
- 955 [11] Peter M Visscher, Naomi R Wray, Qian Zhang, Pamela Sklar, Mark I McCarthy, Matthew A Brown,  
956 and Jian Yang. 10 years of gwas discovery: biology, function, and translation. The American Journal  
957 of Human Genetics, 101(1):5–22, 2017.

- 958 [12] Jonathan K Pritchard and Molly Przeworski. Linkage disequilibrium in humans: models and data.  
959 The American Journal of Human Genetics, 69(1):1–14, 2001.
- 960 [13] Jeremy J Berg and Graham Coop. A population genetic signal of polygenic adaptation. PLoS Genet,  
961 10(8):e1004412, 2014.
- 962 [14] Mattias Jakobsson, Michael D Edge, and Noah A Rosenberg. The relationship between  $f_{st}$  and the  
963 frequency of the most frequent allele. Genetics, 193(2):515–528, 2013.
- 964 [15] Michael D Edge and Noah A Rosenberg. Upper bounds on  $f_{st}$  in terms of the frequency of the most  
965 frequent allele and total homozygosity: the case of a specified number of alleles. Theoretical population  
966 biology, 97:20–34, 2014.
- 967 [16] Michael D Edge and Noah A Rosenberg. A general model of the relationship between the apportionment  
968 of human genetic diversity and the apportionment of human phenotypic diversity. Human biology, 87  
969 (4):313–337, 2015.
- 970 [17] Farhad Hormozdiari, Anthony Zhu, Gleb Kichaev, Chelsea J-T Ju, Ayellet V Segrè, Jong Wha J  
971 Joo, Hyejung Won, Sriram Sankararaman, Bogdan Pasaniuc, Sagiv Shifman, et al. Widespread allelic  
972 heterogeneity in complex traits. The American Journal of Human Genetics, 100(5):789–802, 2017.
- 973 [18] John Novembre and Nicholas H Barton. Tread lightly interpreting polygenic tests of selection. Genetics,  
974 208(4):1351, 2018.
- 975 [19] Noah A Rosenberg, Michael D Edge, Jonathan K Pritchard, and Marcus W Feldman. Interpreting  
976 polygenic scores, polygenic adaptation, and human phenotypic differences. Evolution, medicine, and  
977 public health, 2019(1):26–34, 2019.
- 978 [20] Arbel Harpak and Molly Przeworski. The evolution of group differences in changing environments.  
979 PLoS Biology, 19(1):e3001072, 2021.
- 980 [21] Luisa Pereira, Leon Mutesa, Paulina Tindana, and Michèle Ramsay. African genetic diversity and  
981 adaptation inform a precision medicine agenda. Nature Reviews Genetics, 2021.
- 982 [22] Arun Durvasula and Kirk E Lohmueller. Negative selection on complex traits limits phenotype pre-  
983 diction accuracy between populations. The American Journal of Human Genetics, 2021.
- 984 [23] Hakhamanesh Mostafavi, Arbel Harpak, Ipsita Agarwal, Dalton Conley, Jonathan K Pritchard, and  
985 Molly Przeworski. Variable prediction accuracy of polygenic scores within an ancestry group. Elife, 9:  
986 e48376, 2020.



- 987 [24] Chief Ben-Eghan, Rosie Sun, Jose Sergio Hleap, Alex Diaz-Papkovich, Hans Markus Munter, Audrey V  
988 Grant, Charles Dupras, and Simon Gravel. Don't ignore genetic data from minority populations, 2020.
- 989 [25] Alice B Popejoy and Stephanie M Fullerton. Genomics is failing on diversity. Nature News, 538(7624):  
990 161, 2016.
- 991 [26] Carlos D Bustamante, M Francisco, and Esteban G Burchard. Genomics for the world. Nature, 475  
992 (7355):163–165, 2011.
- 993 [27] Nasa Sinnott-Armstrong, Yosuke Tanigawa, David Amar, Nina J Mars, Matthew Aguirre, Guhan Ram  
994 Venkataraman, Michael Wainberg, Hanna M Ollila, James P Pirruccello, Junyang Qian, et al. Genetics  
995 of 38 blood and urine biomarkers in the uk biobank. BioRxiv, page 660506, 2019.
- 996 [28] Evan A Boyle, Yang I Li, and Jonathan K Pritchard. An expanded view of complex traits: from  
997 polygenic to omnigenic. Cell, 169(7):1177–1186, 2017.
- 998 [29] Nasa Sinnott-Armstrong, Sahin Naqvi, Manuel A Rivas, and Jonathan K Pritchard. Gwas of three  
999 molecular traits highlights core genes and pathways alongside a highly polygenic background. eLife,  
1000 2021.
- 1001 [30] Iain Mathieson. The omnigenic model and polygenic prediction of complex traits. The American  
1002 Journal of Human Genetics, 2021.
- 1003 [31] Matthew A Reyna, Mark DM Leiserson, and Benjamin J Raphael. Hierarchical hotnet: identifying  
1004 hierarchies of altered subnetworks. Bioinformatics, 34(17):i972–i980, 2018.
- 1005 [32] Christopher C Chang, Carson C Chow, Laurent CAM Tellier, Shashaank Vattikuti, Shaun M Purcell,  
1006 and James J Lee. Second-generation plink: rising to the challenge of larger and richer datasets.  
1007 Gigascience, 4(1):s13742–015, 2015.
- 1008 [33] Gad Abraham, Yixuan Qiu, and Michael Inouye. Flashpca2: principal component analysis of biobank-  
1009 scale genotype datasets. Bioinformatics, 2017.
- 1010 [34] Gao Wang, Abhishek Sarkar, Peter Carbonetto, and Matthew Stephens. A simple new approach  
1011 to variable selection in regression, with application to genetic fine mapping. Journal of the Royal  
1012 Statistical Society: Series B (Statistical Methodology), 82(5):1273–1300, 2020.
- 1013 [35] Huwenbo Shi, Steven Gazal, Masahiro Kanai, Evan M Koch, Armin P Schoech, Katherine M Siewert,  
1014 Samuel S Kim, Yang Luo, Tiffany Amariuta, Hailiang Huang, et al. Population-specific causal disease



- 1015 effect sizes in functionally important regions impacted by selection. Nature Communications, 12(1):  
1016 1–15, 2021.
- 1017 [36] Matthew Stephens. False discovery rates: a new deal. Biostatistics, 18(2):275–294, 2017.
- 1018 [37] Sarah M Uebert, Gao Wang, Peter Carbonetto, and Matthew Stephens. Flexible statistical methods  
1019 for estimating and testing effects in genomic studies with multiple conditions. Nature genetics, 51(1):  
1020 187–195, 2019.
- 1021 [38] Wei Cheng, Sohini Ramachandran, and Lorin Crawford. Estimation of non-null snp effect size dis-  
1022 tributions enables the detection of enriched genes underlying complex traits. bioRxiv, page 597484,  
1023 2020.
- 1024 [39] Yan Zhang, Guanghao Qi, Ju-Hyun Park, and Nilanjan Chatterjee. Estimation of complex effect-size  
1025 distributions using summary-level statistics from genome-wide association studies across 32 complex  
1026 traits. Nature genetics, 50(9):1318–1326, 2018.
- 1027 [40] Michael C Wu, Seunggeun Lee, Tianxi Cai, Yun Li, Michael Boehnke, and Xihong Lin. Rare-variant  
1028 association testing for sequencing data with the sequence kernel association test. The American Journal  
1029 of Human Genetics, 89(1):82–93, 2011.
- 1030 [41] Priyanka Nakka, Benjamin J Raphael, and Sohini Ramachandran. Gene and network analysis of  
1031 common variants reveals novel associations in multiple complex diseases. Genetics, 204(2):783–798,  
1032 2016.
- 1033 [42] Alexander Gusev, Arthur Ko, Huwenbo Shi, Gaurav Bhatia, Wonil Chung, Brenda WJH Penninx,  
1034 Rick Jansen, Eco JC De Geus, Dorret I Boomsma, Fred A Wright, et al. Integrative approaches for  
1035 large-scale transcriptome-wide association studies. Nature genetics, 48(3):245, 2016.
- 1036 [43] Priyanka Nakka, Natalie P Archer, Heng Xu, Philip J Lupo, Benjamin J Raphael, Jun J Yang, and  
1037 Sohini Ramachandran. Novel gene and network associations found for acute lymphoblastic leukemia  
1038 using case-control and family-based studies in multiethnic populations. Cancer Epidemiology and  
1039 Prevention Biomarkers, 26(10):1531–1539, 2017.
- 1040 [44] Antonio Fabregat, Konstantinos Sidiropoulos, Phani Garapati, Marc Gillespie, Kerstin Hausmann,  
1041 Robin Haw, Bijay Jassal, Steven Jupe, Florian Korninger, Sheldon McKay, et al. The reactome  
1042 pathway knowledgebase. Nucleic acids research, 44(D1):D481–D487, 2016.

- 1043 [45] Sabry Razick, George Magklaras, and Ian M Donaldson. irefindex: a consolidated protein interaction  
1044 database with provenance. BMC bioinformatics, 9(1):405, 2008.
- 1045 [46] Jishnu Das and Haiyuan Yu. Hint: High-quality protein interactomes and their applications in under-  
1046 standing human disease. BMC systems biology, 6(1):92, 2012.
- 1047 [47] Thomas Rolland, Murat Taşan, Benoit Charloteaux, Samuel J Pevzner, Quan Zhong, Nidhi Sahni,  
1048 Song Yi, Irma Lemmens, Celia Fontanillo, Roberto Mosca, et al. A proteome-scale map of the human  
1049 interactome network. Cell, 159(5):1212–1226, 2014.
- 1050 [48] Christopher S Carlson, Tara C Matise, Kari E North, Christopher A Haiman, Megan D Fesinmeyer,  
1051 Steven Buyske, Fredrick R Schumacher, Ulrike Peters, Nora Franceschini, Marylyn D Ritchie, et al.  
1052 Generalization and dilution of association results from european gwas in populations of non-european  
1053 ancestry: the page study. PLoS Biol, 11(9):e1001661, 2013.
- 1054 [49] Jimmy Z Liu, Suzanne Van Sommeren, Hailiang Huang, Siew C Ng, Rudi Alberts, Atsushi Takahashi,  
1055 Stephan Ripke, James C Lee, Luke Jostins, Tejas Shah, et al. Association analyses identify 38 sus-  
1056 ceptibility loci for inflammatory bowel disease and highlight shared genetic risk across populations.  
1057 Nature genetics, 47(9):979, 2015.
- 1058 [50] Adam Eyre-Walker. Genetic architecture of a complex trait and its implications for fitness and genome-  
1059 wide association studies. Proceedings of the National Academy of Sciences, 107(suppl 1):1752–1756,  
1060 2010.
- 1061 [51] Huwenbo Shi, Kathryn S Burch, Ruth Johnson, Malika K Freund, Gleb Kichaev, Nicholas Mancuso,  
1062 Astrid M Manuel, Natalie Dong, and Bogdan Pasaniuc. Localizing components of shared transethnic  
1063 genetic architecture of complex traits from gwas summary data. The American Journal of Human  
1064 Genetics, 106(6):805–817, 2020.
- 1065 [52] AJ Szalai, J Wu, EM Lange, MA McCrory, CD Langefeld, A Williams, SO Zakharkin, Varghese  
1066 George, DB Allison, GS Cooper, et al. Single-nucleotide polymorphisms in the c-reactive protein (crp)  
1067 gene promoter that affect transcription factor binding, alter transcriptional activity, and associate with  
1068 differences in baseline serum crp level. Journal of Molecular Medicine, 83(6):440–447, 2005.
- 1069 [53] Shi-Chao Zhang, Ming-Yu Wang, Jun-Rui Feng, Yue Chang, Shang-Rong Ji, and Yi Wu. Reversible  
1070 promoter methylation determines fluctuating expression of acute phase proteins. Elife, 9:e51317, 2020.
- 1071 [54] Farhad Hormozdiari, Emrah Kostem, Eun Yong Kang, Bogdan Pasaniuc, and Eleazar Eskin. Identi-  
1072 fying causal variants at loci with multiple signals of association. Genetics, 198(2):497–508, 2014.

- 1073 [55] Alicia R Martin, Meng Lin, Julie M Granka, Justin W Myrick, Xiaomin Liu, Alexandra Sockell,  
1074 Elizabeth G Atkinson, Cedric J Werely, Marlo Möller, Manjinder S Sandhu, et al. An unexpectedly  
1075 complex architecture for skin pigmentation in africans. Cell, 171(6):1340–1353, 2017.
- 1076 [56] Brian L Browning and Sharon R Browning. Efficient multilocus association testing for whole genome as-  
1077 sociation studies using localized haplotype clustering. Genetic Epidemiology: The Official Publication  
1078 of the International Genetic Epidemiology Society, 31(5):365–375, 2007.
- 1079 [57] Jimmy Z Liu, Allan F Mcrae, Dale R Nyholt, Sarah E Medland, Naomi R Wray, Kevin M Brown,  
1080 Nicholas K Hayward, Grant W Montgomery, Peter M Visscher, Nicholas G Martin, et al. A versatile  
1081 gene-based test for genome-wide association studies. The American Journal of Human Genetics, 87  
1082 (1):139–145, 2010.
- 1083 [58] Christiaan A de Leeuw, Joris M Mooij, Tom Heskes, and Danielle Posthuma. Magma: generalized  
1084 gene-set analysis of gwas data. PLoS Comput Biol, 11(4):e1004219, 2015.
- 1085 [59] Qian S Zhang, Brian L Browning, and Sharon R Browning. Genome-wide haplotypic testing in a  
1086 finnish cohort identifies a novel association with low-density lipoprotein cholesterol. European Journal  
1087 of Human Genetics, 23(5):672–677, 2015.
- 1088 [60] Xiang Zhu and Matthew Stephens. Bayesian large-scale multiple regression with summary statistics  
1089 from genome-wide association studies. The annals of applied statistics, 11(3):1561, 2017.
- 1090 [61] John D Storey and Robert Tibshirani. Statistical significance for genomewide studies. Proceedings of  
1091 the National Academy of Sciences, 100(16):9440–9445, 2003.
- 1092 [62] Benjamin M Neale, Sarah E Medland, Stephan Ripke, Philip Asherson, Barbara Franke, Klaus-Peter  
1093 Lesch, Stephen V Faraone, Thuy Trang Nguyen, Helmut Schäfer, Peter Holmans, et al. Meta-analysis  
1094 of genome-wide association studies of attention-deficit/hyperactivity disorder. Journal of the American  
1095 Academy of Child & Adolescent Psychiatry, 49(9):884–897, 2010.
- 1096 [63] Ditte Demontis, Raymond K Walters, Joanna Martin, Manuel Mattheisen, Thomas D Als, Esben  
1097 Agerbo, Gísli Baldursson, Rich Belliveau, Jonas Bybjerg-Grauholm, Marie Bækvad-Hansen, et al.  
1098 Discovery of the first genome-wide significant risk loci for attention deficit/hyperactivity disorder.  
1099 Nature genetics, 51(1):63–75, 2019.
- 1100 [64] Mark D Leiserson, Fabio Vandin, Hsin-Ta Wu, Jason R Dobson, and Benjamin R Raphael. Pan-cancer  
1101 identification of mutated pathways and protein complexes, 2014.

- 1102 [65] Lenore Cowen, Trey Ideker, Benjamin J Raphael, and Roded Sharan. Network propagation: a universal  
1103 amplifier of genetic associations. Nature Reviews Genetics, 18(9):551, 2017.
- 1104 [66] Lucia A Hindorff, Vence L Bonham, Lawrence C Brody, Margaret EC Ginoza, Carolyn M Hutter,  
1105 Teri A Manolio, and Eric D Green. Prioritizing diversity in human genomics research. Nature Reviews  
1106 Genetics, 19(3):175, 2018.
- 1107 [67] Roseann E Peterson, Karoline Kuchenbaecker, Raymond K Walters, Chia-Yen Chen, Alice B Popejoy,  
1108 Sathish Periyasamy, Max Lam, Conrad Iyegbe, Rona J Strawbridge, Leslie Brick, et al. Genome-wide  
1109 association studies in ancestrally diverse populations: opportunities, methods, pitfalls, and recommen-  
1110 dations. Cell, 179(3):589–603, 2019.
- 1111 [68] Cristen J Willer, Yun Li, and Gonalo R Abecasis. Metal: fast and efficient meta-analysis of  
1112 genomewide association scans. Bioinformatics, 26(17):2190–2191, 2010.
- 1113 [69] Dan-Yu Lin, Ran Tao, William D Kalsbeek, Donglin Zeng, Franklyn Gonzalez II, Lindsay Fernandez-  
1114 Rhodes, Mariaelisa Graff, Gary G Koch, Kari E North, and Gerardo Heiss. Genetic association analysis  
1115 under complex survey sampling: the hispanic community health study/study of latinos. The American  
1116 Journal of Human Genetics, 95(6):675–688, 2014.
- 1117 [70] Jian Yang, Noah A Zaitlen, Michael E Goddard, Peter M Visscher, and Alkes L Price. Advantages  
1118 and pitfalls in the application of mixed-model association methods. Nature genetics, 46(2):100–106,  
1119 2014.
- 1120 [71] Scott I Vrieze, William G Iacono, and Matt McGue. Confluence of genes, environment, development,  
1121 and behavior in a post-gwas world. Development and Psychopathology, 24(4):1195, 2012.
- 1122 [72] Suzanne H Gage, George Davey Smith, Jennifer J Ware, Jonathan Flint, and Marcus R Munafo. G=  
1123 e: What gwas can tell us about the environment. PLoS Genetics, 12(2):e1005765, 2016.
- 1124 [73] Luisa N. Borrell, Jennifer R. Elhawary, Elena Fuentes-Afflick, Jonathan Witonsky, Nirav Bhakta,  
1125 Alan H.B. Wu, Kirsten Bibbins-Domingo, Jos R. Rodrguez-Santana, Michael A. Lenoir, James R.  
1126 Gavin, Rick A. Kittles, Noah A. Zaitlen, David S. Wilkes, Neil R. Powe, Elad Ziv, and Esteban G.  
1127 Burchard. Race and genetic ancestry in medicine - a time for reckoning with racism. New England  
1128 Journal of Medicine, 0(0):null, 0. doi: 10.1056/NEJMms2029562. URL [https://doi.org/10.1056/](https://doi.org/10.1056/NEJMms2029562)  
1129 [NEJMms2029562](https://doi.org/10.1056/NEJMms2029562).

- 1130 [74] Alexander I Young, Michael L Frigge, Daniel F Gudbjartsson, Gudmar Thorleifsson, Gyda Bjornsdot-  
1131 tir, Patrick Sulem, Gisli Masson, Unnur Thorsteinsdottir, Kari Stefansson, and Augustine Kong. Re-  
1132 latedness disequilibrium regression estimates heritability without environmental bias. Nature genetics,  
1133 50(9):1304–1310, 2018.
- 1134 [75] Alexander I Young, Stefania Benonisdottir, Molly Przeworski, and Augustine Kong. Deconstructing  
1135 the sources of genotype-phenotype associations in humans. Science, 365(6460):1396–1400, 2019.
- 1136 [76] Farhad Hormozdiari, Martijn Van De Bunt, Ayellet V Segre, Xiao Li, Jong Wha J Joo, Michael Bilow,  
1137 Jae Hoon Sul, Sriram Sankararaman, Bogdan Pasaniuc, and Eleazar Eskin. Colocalization of gwas and  
1138 eqtl signals detects target genes. The American Journal of Human Genetics, 99(6):1245–1260, 2016.
- 1139 [77] Nathan LaPierre, Kodi Taraszka, Helen Huang, Rosemary He, Farhad Hormozdiari, and Eleazar Eskin.  
1140 Identifying causal variants by fine mapping across multiple studies. In International Conference on  
1141 Research in Computational Molecular Biology, pages 257–258. Springer, 2020.
- 1142 [78] Lorin Crawford, Ping Zeng, Sayan Mukherjee, and Xiang Zhou. Detecting epistasis with the marginal  
1143 epistasis test in genetic mapping studies of quantitative traits. PLoS genetics, 13(7):e1006869, 2017.
- 1144 [79] Michael C Turchin and Matthew Stephens. Bayesian multivariate reanalysis of large genetic studies  
1145 identifies many new associations. PLoS genetics, 15(10):e1008431, 2019.
- 1146 [80] Xiang Zhou and Matthew Stephens. Genome-wide efficient mixed-model analysis for association stud-  
1147 ies. Nature genetics, 44(7):821–824, 2012.
- 1148 [81] Matthew Stephens. A unified framework for association analysis with multiple related phenotypes.  
1149 PloS one, 8(7):e65245, 2013.
- 1150 [82] Brielin C Brown, Chun Jimmie Ye, Alkes L Price, Noah Zaitlen, Asian Genetic Epidemiology Network  
1151 Type 2 Diabetes Consortium, et al. Transethnic genetic-correlation estimates from summary statistics.  
1152 The American Journal of Human Genetics, 99(1):76–88, 2016.
- 1153 [83] Kevin J Galinsky, Yakir A Reshef, Hilary K Finucane, Po-Ru Loh, Noah Zaitlen, Nick J Patterson,  
1154 Brielin C Brown, and Alkes L Price. Estimating cross-population genetic correlations of causal effect  
1155 sizes. Genetic epidemiology, 43(2):180–188, 2019.
- 1156 [84] Barbara Domingues Bitarello and Iain Mathieson. Polygenic scores for height in admixed populations.  
1157 bioRxiv, 2020.

- 1158 [85] Davide Marnetto, Katri Pärna, Kristi Läll, Ludovica Molinaro, Francesco Montinaro, Toomas Haller,  
1159 Mait Metspalu, Reedik Mägi, Krista Fischer, and Luca Pagani. Ancestry deconvolution and par-  
1160 tial polygenic score can improve susceptibility predictions in recently admixed individuals. Nature  
1161 communications, 11(1):1–9, 2020.
- 1162 [86] Hailiang Huang, Yunfeng Ruan, Yen-Chen Anne Feng, Chia-Yen Chen, Max Lam, Akira Sawa, Alicia  
1163 Martin, Shengying Qin, and Tian Ge. Improving polygenic prediction in ancestrally diverse popula-  
1164 tions. 2021.
- 1165 [87] Laura K Hayward and Guy Sella. Polygenic adaptation after a sudden change in environment. BioRxiv,  
1166 page 792952, 2019.
- 1167 [88] Jon McClellan and Mary-Claire King. Genetic heterogeneity in human disease. Cell, 141(2):210–217,  
1168 2010.
- 1169 [89] Farid Rajabli, Briseida E Feliciano, Katrina Celis, Kara L Hamilton-Nelson, Patrice L Whitehead,  
1170 Larry D Adams, Parker L Bussies, Clara P Manrique, Alejandra Rodriguez, Vanessa Rodriguez, et al.  
1171 Ancestral origin of apoe  $\epsilon$ 4 alzheimer disease risk in puerto rican and african american populations.  
1172 PLoS genetics, 14(12), 2018.
- 1173 [90] Jake R Conway, Alexander Lex, and Nils Gehlenborg. Upsetr: an r package for the visualization of  
1174 intersecting sets and their properties. Bioinformatics, 33(18):2938–2940, 2017.
- 1175 [91] David H Alexander, John Novembre, and Kenneth Lange. Fast model-based estimation of ancestry in  
1176 unrelated individuals. Genome research, 19(9):1655–1664, 2009.
- 1177 [92] Peter H Sudmant, Tobias Rausch, Eugene J Gardner, Robert E Handsaker, Alexej Abyzov, John  
1178 Huddleston, Yan Zhang, Kai Ye, Goo Jun, Markus Hsi-Yang Fritz, et al. An integrated map of  
1179 structural variation in 2,504 human genomes. Nature, 526(7571):75–81, 2015.
- 1180 [93] Masahiro Kanai, Masato Akiyama, Atsushi Takahashi, Nana Matoba, Yukihide Momozawa, Masashi  
1181 Ikeda, Nakao Iwata, Shiro Ikegawa, Makoto Hirata, Koichi Matsuda, et al. Genetic analysis of quan-  
1182 titative traits in the japanese population links cell types to complex human diseases. Nature genetics,  
1183 50(3):390–400, 2018.
- 1184 [94] Masato Akiyama, Yukinori Okada, Masahiro Kanai, Atsushi Takahashi, Yukihide Momozawa, Masashi  
1185 Ikeda, Nakao Iwata, Shiro Ikegawa, Makoto Hirata, Koichi Matsuda, et al. Genome-wide association  
1186 study identifies 112 new loci for body mass index in the japanese population. Nature genetics, 49(10):  
1187 1458, 2017.

- 1188 [95] Masato Akiyama, Kazuyoshi Ishigaki, Saori Sakaue, Yukihide Momozawa, Momoko Horikoshi, Makoto  
1189 Hirata, Koichi Matsuda, Shiro Ikegawa, Atsushi Takahashi, Masahiro Kanai, et al. Characterizing rare  
1190 and low-frequency height-associated variants in the japanese population. Nature communications, 10  
1191 (1):1–11, 2019.
- 1192 [96] Derek Klarin, Scott M Damrauer, Kelly Cho, Yan V Sun, Tanya M Teslovich, Jacqueline Honerlaw,  
1193 David R Gagnon, Scott L DuVall, Jin Li, Gina M Peloso, et al. Genetics of blood lipids among  
1194 300,000 multi-ethnic participants of the million veteran program. Nature genetics, 50(11):1514–1523,  
1195 2018.
- 1196 [97] Paul S De Vries, Michael R Brown, Amy R Bentley, Yun J Sung, Thomas W Winkler, Ioanna Ntalla,  
1197 Karen Schwander, Aldi T Kraja, Xiuqing Guo, Nora Franceschini, et al. Multiancestry genome-  
1198 wide association study of lipid levels incorporating gene-alcohol interactions. American journal of  
1199 epidemiology, 188(6):1033–1054, 2019.
- 1200 [98] Raymond Noordam, Maxime M Bos, Heming Wang, Thomas W Winkler, Amy R Bentley, Tuomas O  
1201 Kilpeläinen, Paul S de Vries, Yun Ju Sung, Karen Schwander, Brian E Cade, et al. Multi-ancestry  
1202 sleep-by-snp interaction analysis in 126,926 individuals reveals lipid loci stratified by sleep duration.  
1203 Nature communications, 10(1):1–13, 2019.
- 1204 [99] Tom G Richardson, Eleanor Sanderson, Tom M Palmer, Mika Ala-Korpela, Brian A Ference, George  
1205 Davey Smith, and Michael V Holmes. Evaluating the relationship between circulating lipoprotein lipids  
1206 and apolipoproteins with risk of coronary heart disease: A multivariable mendelian randomisation  
1207 analysis. PLoS medicine, 17(3):e1003062, 2020.
- 1208 [100] Thomas J Hoffmann, Elizabeth Theusch, Tanushree Haldar, Dilrini K Ranatunga, Eric Jorgenson,  
1209 Marisa W Medina, Mark N Kvale, Pui-Yan Kwok, Catherine Schaefer, Ronald M Krauss, et al. A  
1210 large electronic-health-record-based genome-wide study of serum lipids. Nature genetics, 50(3):401–  
1211 413, 2018.
- 1212 [101] Alexander M Kulminski, Jian Huang, Yury Loika, Konstantin G Arbeev, Olivia Bagley, Arseniy  
1213 Yashkin, Matt Duan, and Irina Culminskaya. Strong impact of natural-selection-free heterogeneity in  
1214 genetics of age-related phenotypes. Aging (Albany NY), 10(3):492, 2018.
- 1215 [102] Aldi T Kraja, Dhananjay Vaidya, James S Pankow, Mark O Goodarzi, Themistocles L Assimes,  
1216 Iftikhar J Kullo, Ulla Sovio, Rasika A Mathias, Yan V Sun, Nora Franceschini, et al. A bivariate

- 1217 genome-wide approach to metabolic syndrome: Stampeed consortium. *Diabetes*, 60(4):1329–1339,  
1218 2011.
- 1219 [103] Cassandra N Spracklen, Peng Chen, Young Jin Kim, Xu Wang, Hui Cai, Shengxu Li, Jirong Long,  
1220 Ying Wu, Ya Xing Wang, Fumihiko Takeuchi, et al. Association analyses of east asian individuals  
1221 and trans-ancestry analyses with european individuals reveal new loci associated with cholesterol and  
1222 triglyceride levels. *Human molecular genetics*, 26(9):1770–1784, 2017.
- 1223 [104] Makoto Kurano, Kazuhisa Tsukamoto, Shigeo Kamitsuji, Naoyuki Kamatani, Masumi Hara, Toshio  
1224 Ishikawa, Bong-Jo Kim, Sanghoon Moon, Young Jin Kim, and Tamio Teramoto. Genome-wide asso-  
1225 ciation study of serum lipids confirms previously reported associations as well as new associations of  
1226 common snps within pcsk7 gene with triglyceride. *Journal of human genetics*, 61(5):427–433, 2016.
- 1227 [105] Amy R Bentley, Yun J Sung, Michael R Brown, Thomas W Winkler, Aldi T Kraja, Ioanna Ntalla,  
1228 Karen Schwander, Daniel I Chasman, Elise Lim, Xuan Deng, et al. Multi-ancestry genome-wide  
1229 gene-smoking interaction study of 387,272 individuals identifies new loci associated with serum lipids.  
1230 *Nature genetics*, 51(4):636–648, 2019.
- 1231 [106] Ida Surakka, Momoko Horikoshi, Reedik Mägi, Antti-Pekka Sarin, Anubha Mahajan, Vasiliki Lagou,  
1232 Letizia Marullo, Teresa Ferreira, Benjamin Miraglio, Sanna Timonen, et al. The impact of low-  
1233 frequency and rare variants on lipid levels. *Nature genetics*, 47(6):589–597, 2015.
- 1234 [107] Tanya M Teslovich, Kiran Musumuru, Albert V Smith, Andrew C Edmondson, Ioannis M Stylianou,  
1235 Masahiro Koseki, James P Pirruccello, Samuli Ripatti, Daniel I Chasman, Cristen J Willer, et al.  
1236 Biological, clinical and population relevance of 95 loci for blood lipids. *Nature*, 466(7307):707–713,  
1237 2010.
- 1238 [108] Anders Bergström, Shane A McCarthy, Ruoyun Hui, Mohamed A Almarri, Qasim Ayub, Petr Danecek,  
1239 Yuan Chen, Sabine Felkel, Pille Hallast, Jack Kamm, et al. Insights into human genetic variation and  
1240 population history from 929 diverse genomes. *Science*, 367(6484), 2020.
- 1241 [109] Peter Carbonetto and Matthew Stephens. Integrated enrichment analysis of variants and pathways in  
1242 genome-wide association studies indicates central role for il-2 signaling genes in type 1 diabetes, and  
1243 cytokine signaling genes in crohn’s disease. *PLoS genetics*, 9(10):e1003770, 2013.
- 1244 [110] Xiang Zhu and Matthew Stephens. Large-scale genome-wide enrichment analyses identify new trait-  
1245 associated genes and pathways across 31 human phenotypes. *Nature communications*, 9(1):1–14, 2018.



- 1246 [111] Michael C Wu, Peter Kraft, Michael P Epstein, Deanne M Taylor, Stephen J Chanock, David J  
1247 Hunter, and Xihong Lin. Powerful snp-set analysis for case-control genome-wide association studies.  
1248 The American Journal of Human Genetics, 86(6):929–942, 2010.
- 1249 [112] Seunggeun Lee, Mary J Emond, Michael J Bamshad, Kathleen C Barnes, Mark J Rieder, Deborah A  
1250 Nickerson, ESP Lung Project Team, David C Christiani, Mark M Wurfel, Xihong Lin, et al. Opti-  
1251 mal unified approach for rare-variant association testing with application to small-sample case-control  
1252 whole-exome sequencing studies. The American Journal of Human Genetics, 91(2):224–237, 2012.
- 1253 [113] Iuliana Ionita-Laza, Seunggeun Lee, Vlad Makarov, Joseph D Buxbaum, and Xihong Lin. Sequence  
1254 kernel association tests for the combined effect of rare and common variants. The American Journal  
1255 of Human Genetics, 92(6):841–853, 2013.
- 1256 [114] Melissa R McGuirl, Samuel Pattillo Smith, Björn Sandstede, and Sohini Ramachandran. Detecting  
1257 shared genetic architecture among multiple phenotypes by hierarchical clustering of gene-level associ-  
1258 ation statistics. Genetics, 2020.
- 1259 [115] Maria Maddalena Barbieri, James O Berger, et al. Optimal predictive model selection. The annals of  
1260 statistics, 32(3):870–897, 2004.
- 1261 [116] Terhi Kettunen, Carita Eklund, Mika Kähönen, Antti Jula, Hannu Päivä, Leo-Pekka Lyytikäinen,  
1262 Mikko Hurme, and Terho Lehtimäki. Polymorphism in the c-reactive protein (crp) gene affects crp  
1263 levels in plasma and one early marker of atherosclerosis in men: The health 2000 survey. Scandinavian  
1264 journal of clinical and laboratory investigation, 71(5):353–361, 2011.
- 1265 [117] Amy Z Fan, Ajay Yesupriya, Man-huei Chang, Meaghan House, Jing Fang, Renée Ned, Donald Hayes,  
1266 Nicole F Dowling, and Ali H Mokdad. Gene polymorphisms in association with emerging cardiovascular  
1267 risk markers in adult women. BMC medical genetics, 11(1):6, 2010.
- 1268 [118] E Komurcu-Bayrak, N Erginel-Unaltuna, A Onat, B Ozsait, C Eklund, M Hurme, N Mononen, R Laak-  
1269 sonen, G Hergenc, and T Lehtimäki. Association of c-reactive protein (crp) gene allelic variants with  
1270 serum crp levels and hypertension in turkish adults. Atherosclerosis, 206(2):474–479, 2009.
- 1271 [119] S-N Chang, C-T Tsai, C-K Wu, J-K Lee, L-P Lai, S-W Huang, L-Y Huang, C-D Tseng, J-L Lin, F-T  
1272 Chiang, et al. A functional variant in the promoter region regulates the c-reactive protein gene and  
1273 is a potential candidate for increased risk of atrial fibrillation. Journal of internal medicine, 272(3):  
1274 305–315, 2012.

- 1275 [120] Hyoun-Ah Kim, Hye-Young Chun, Seung-Hyun Kim, Hae-Sim Park, and Chang-Hee Suh. C-reactive  
1276 protein gene polymorphisms in disease susceptibility and clinical manifestations of korean systemic  
1277 lupus erythematosus. The Journal of rheumatology, 36(10):2238–2243, 2009.
- 1278 [121] Abbas Dehghan, Josée Dupuis, Maja Barbalić, Joshua C Bis, Gudny Eiriksdottir, Chen Lu, Niina Pel-  
1279 likka, Henri Wallaschofski, Johannes Kettunen, Peter Henneman, et al. Meta-analysis of genome-wide  
1280 association studies in 80 000 subjects identifies multiple loci for c-reactive protein levels. Circulation,  
1281 123(7):731–738, 2011.
- 1282 [122] Paul M Ridker, Guillaume Pare, Alex Parker, Robert YL Zee, Jacqueline S Danik, Julie E Buring,  
1283 David Kwiakowski, Nancy R Cook, Joseph P Miletich, and Daniel I Chasman. Loci related to  
1284 metabolic-syndrome pathways including lepr, hnf1a, il6r, and gckr associate with plasma c-reactive  
1285 protein: the women’s genome health study. The American Journal of Human Genetics, 82(5):1185–  
1286 1192, 2008.
- 1287 [123] Symen Ligthart, Ahmad Vaez, Urmo Vösa, Maria G Stathopoulou, Paul S De Vries, Bram P Prins, Pe-  
1288 ter J Van der Most, Toshiko Tanaka, Elnaz Naderi, Lynda M Rose, et al. Genome analyses of 200,000  
1289 individuals identify 58 loci for chronic inflammation and highlight pathways that link inflammation  
1290 and complex disorders. The American Journal of Human Genetics, 103(5):691–706, 2018.
- 1291 [124] Alex P Reiner, Sandra Beleza, Nora Franceschini, Paul L Auer, Jennifer G Robinson, Charles Kooper-  
1292 berg, Ulrike Peters, and Hua Tang. Genome-wide association and population genetic analysis of  
1293 c-reactive protein in african american and hispanic american women. The American Journal of Human  
1294 Genetics, 91(3):502–512, 2012.
- 1295 [125] Paul Elliott, John C Chambers, Weihua Zhang, Robert Clarke, Jemma C Hopewell, John F Peden,  
1296 Jeanette Erdmann, Peter Braund, James C Engert, Derrick Bennett, et al. Genetic loci associated  
1297 with c-reactive protein levels and risk of coronary heart disease. Jama, 302(1):37–48, 2009.
- 1298 [126] Deog Kyeom Kim, Michael H Cho, Craig P Hersh, David A Lomas, Bruce E Miller, Xiangyang Kong,  
1299 Per Bakke, Amund Gulsvik, Alvar Agustí, Emiel Wouters, et al. Genome-wide association analysis  
1300 of blood biomarkers in chronic obstructive pulmonary disease. American journal of respiratory and  
1301 critical care medicine, 186(12):1238–1247, 2012.
- 1302 [127] William J Astle, Heather Elding, Tao Jiang, Dave Allen, Dace Ruklisa, Alice L Mann, Daniel Mead,  
1303 Heleen Bouman, Fernando Riveros-Mckay, Myrto A Kostadima, et al. The allelic landscape of human  
1304 blood cell trait variation and links to common complex disease. Cell, 167(5):1415–1429, 2016.

- 1305 [128] Juraĵ Sokol, Kamil Biringer, Maria Skerenova, Miroslav Hasko, Lenka Bartosova, Jan Stasko, Jan  
1306 Danko, and Peter Kubisz. Platelet aggregation abnormalities in patients with fetal losses: the gp6  
1307 gene polymorphism. Fertility and sterility, 98(5):1170–1174, 2012.
- 1308 [129] Andrew D Johnson, Lisa R Yanek, Ming-Huei Chen, Nauder Faraday, Martin G Larson, Geoffrey  
1309 Tofler, Shioh J Lin, Aldi T Kraja, Michael A Province, Qiong Yang, et al. Genome-wide meta-analyses  
1310 identifies seven loci associated with platelet aggregation in response to agonists. Nature genetics, 42  
1311 (7):608–613, 2010.
- 1312 [130] Takao Shimizu. Lipid mediators in health and disease: enzymes and receptors as therapeutic targets  
1313 for the regulation of immunity and inflammation. Annual review of pharmacology and toxicology, 49:  
1314 123–150, 2009.
- 1315 [131] Yuji Shimizu, Kazuhiko Arima, Yuko Noguchi, Shin-Ya Kawashiri, Hiroto Yamanashi, Mami Tamai,  
1316 Yasuhiro Nagata, and Takahiro Maeda. Potential mechanisms underlying the association between  
1317 single nucleotide polymorphism (brap and aldh2) and hypertension among elderly japanese population.  
1318 Scientific Reports, 10(1):1–9, 2020.
- 1319 [132] John D Eicher, Nathalie Chami, Tim Kacprowski, Akihiro Nomura, Ming-Huei Chen, Lisa R Yanek,  
1320 Salman M Tajuddin, Ursula M Schick, Andrew J Slater, Nathan Pankratz, et al. Platelet-related  
1321 variants identified by exomechip meta-analysis in 157,293 individuals. The American Journal of Human  
1322 Genetics, 99(1):40–55, 2016.
- 1323 [133] Rehan Qayyum, Beverly M Snively, Elad Ziv, Michael A Nalls, Yongmei Liu, Weihong Tang, Lisa R  
1324 Yanek, Leslie Lange, Michele K Evans, Santhi Ganesh, et al. A meta-analysis and genome-wide  
1325 association study of platelet count and mean platelet volume in african americans. PLoS Genet, 8(3):  
1326 e1002491, 2012.
- 1327 [134] Cristen J Willer, Ellen M Schmidt, Sebanti Sengupta, Gina M Peloso, Stefan Gustafsson, Stavroula  
1328 Kanoni, Andrea Ganna, Jin Chen, Martin L Buchkovich, Samia Mora, et al. Discovery and refinement  
1329 of loci associated with lipid levels. Nature genetics, 45(11):1274, 2013.
- 1330 [135] Rita PS Middelberg, Manuel AR Ferreira, Anjali K Henders, Andrew C Heath, Pamela AF Madden,  
1331 Grant W Montgomery, Nicholas G Martin, and John B Whitfield. Genetic variants in lpl, oasl and  
1332 tomm40/apoe-c1-c2-c4 genes are associated with multiple cardiovascular-related traits. BMC medical  
1333 genetics, 12(1):123, 2011.

- 1334 [136] Sekar Kathiresan, Olle Melander, Candace Guiducci, Aarti Surti, Noël P Burt, Mark J Rieder, Gre-  
1335 gory M Cooper, Charlotta Roos, Benjamin F Voight, Aki S Havulinna, et al. Six new loci associated  
1336 with blood low-density lipoprotein cholesterol, high-density lipoprotein cholesterol or triglycerides in  
1337 humans. Nature genetics, 40(2):189–197, 2008.
- 1338 [137] Jaspal S Kooner, John C Chambers, Carlos A Aguilar-Salinas, David A Hinds, Craig L Hyde, Gre-  
1339 gory R Warnes, Francisco J Gómez Pérez, Kelly A Frazer, Paul Elliott, James Scott, et al. Genome-  
1340 wide scan identifies variation in mlxpl associated with plasma triglycerides. Nature genetics, 40(2):  
1341 149, 2008.
- 1342 [138] Sekar Kathiresan, Cristen J Willer, Gina M Peloso, Serkalem Demissie, Kiran Musunuru, Eric E  
1343 Schadt, Lee Kaplan, Derrick Bennett, Yun Li, Toshiko Tanaka, et al. Common variants at 30 loci  
1344 contribute to polygenic dyslipidemia. Nature genetics, 41(1):56–65, 2009.
- 1345 [139] Maria Keller, Dorit Schleinitz, Julia Förster, Anke Tönjes, Yvonne Böttcher, Antje Fischer-Rosinsky,  
1346 Jana Breitfeld, Kerstin Weidle, Nigel W Rayner, Ralph Burkhardt, et al. Thoc5: a novel gene involved  
1347 in hdl-cholesterol metabolism. Journal of lipid research, 54(11):3170–3176, 2013.
- 1348 [140] Richa Saxena, Benjamin F Voight, Valeriya Lyssenko, Noël P Burt, Paul IW de Bakker, Hong Chen,  
1349 Jeffrey J Roix, Sekar Kathiresan, Joel N Hirschhorn, Mark J Daly, et al. Genome-wide association  
1350 analysis identifies loci for type 2 diabetes and triglyceride levels. Science, 316(5829):1331–1336, 2007.
- 1351 [141] Yurii S Aulchenko, Samuli Ripatti, Ida Lindqvist, Dorret Boomsma, Iris M Heid, Peter P Pramstaller,  
1352 Brenda WJH Penninx, A Cecile JW Janssens, James F Wilson, Tim Spector, et al. Loci influencing  
1353 lipid levels and coronary heart disease risk in 16 european population cohorts. Nature genetics, 41(1):  
1354 47, 2009.
- 1355 [142] Rubina Tabassum, Joel T Rämö, Pietari Ripatti, Jukka T Koskela, Mitja Kurki, Juha Karjalainen,  
1356 Priit Palta, Shabbeer Hassan, Javier Nunez-Fontarnau, Tuomo TJ Kiiskinen, et al. Genetic architecture  
1357 of human plasma lipidome and its link to cardiovascular disease. Nature communications, 10(1):1–14,  
1358 2019.
- 1359 [143] Liang He, Yelena Kernogitski, Irina Kulminskaya, Yury Loika, Konstantin G Arbeev, Elena Loiko,  
1360 Olivia Bagley, Matt Duan, Arseniy Yashkin, Svetlana V Ukraintseva, et al. Pleiotropic meta-analyses  
1361 of longitudinal studies discover novel genetic variants associated with age-related diseases. Frontiers  
1362 in genetics, 7:179, 2016.

- 1363 [144] Dawn M Waterworth, Sally L Ricketts, Kijoung Song, Li Chen, Jing Hua Zhao, Samuli Ripatti, Yurii S  
1364 Aulchenko, Weihua Zhang, Xin Yuan, Noha Lim, et al. Genetic variants influencing circulating lipid  
1365 levels and risk of coronary artery disease. Arteriosclerosis, thrombosis, and vascular biology, 30(11):  
1366 2264–2276, 2010.
- 1367 [145] Gina M Peloso, Paul L Auer, Joshua C Bis, Arend Voorman, Alanna C Morrison, Nathan O Stitzel,  
1368 Jennifer A Brody, Sumeet A Khetarpal, Jacy R Crosby, Myriam Fornage, et al. Association of low-  
1369 frequency and rare coding-sequence variants with blood lipids and coronary heart disease in 56,000  
1370 whites and blacks. The American Journal of Human Genetics, 94(2):223–232, 2014.
- 1371 [146] Tuomas O Kilpeläinen, Amy R Bentley, Raymond Noordam, Yun Ju Sung, Karen Schwander,  
1372 Thomas W Winkler, Hermina Jakupović, Daniel I Chasman, Alisa Manning, Ioanna Ntalla, et al.  
1373 Multi-ancestry study of blood lipid levels identifies four loci interacting with physical activity. Nature  
1374 communications, 10(1):1–11, 2019.
- 1375 [147] Yoichiro Kamatani, Koichi Matsuda, Yukinori Okada, Michiaki Kubo, Naoya Hosono, Yataro Daigo,  
1376 Yusuke Nakamura, and Naoyuki Kamatani. Genome-wide association study of hematological and  
1377 biochemical traits in a japanese population. Nature genetics, 42(3):210, 2010.
- 1378 [148] Marc A Coram, Qing Duan, Thomas J Hoffmann, Timothy Thornton, Joshua W Knowles, Nicholas A  
1379 Johnson, Heather M Ochs-Balcom, Timothy A Donlon, Lisa W Martin, Charles B Eaton, et al.  
1380 Genome-wide characterization of shared and distinct genetic components that influence blood lipid  
1381 levels in ethnically diverse human populations. The American Journal of Human Genetics, 92(6):  
1382 904–916, 2013.
- 1383 [149] Deepti Gurdasani, Tommy Carstensen, Segun Fatumo, Guanjie Chen, Chris S Franklin, Javier Prado-  
1384 Martinez, Heleen Bouman, Federico Abascal, Marc Haber, Ioanna Tachmazidou, et al. Uganda genome  
1385 resource enables insights into population history and genomic discovery in africa. Cell, 179(4):984–1002,  
1386 2019.
- 1387 [150] Elisabeth M van Leeuwen, Aniko Sabo, Joshua C Bis, Jennifer E Huffman, Ani Manichaikul, Albert V  
1388 Smith, Mary F Feitosa, Serkalem Demissie, Peter K Joshi, Qing Duan, et al. Meta-analysis of 49  
1389 549 individuals imputed with the 1000 genomes project reveals an exonic damaging variant in angptl4  
1390 determining fasting tg levels. Journal of medical genetics, 53(7):441–449, 2016.
- 1391 [151] Robert Perneczky, Panagiotis Alexopoulos, and Alzheimer’s Disease euroimaging Initiative. Cere-

- 1392           brospinal fluid bace1 activity and markers of amyloid precursor protein metabolism and axonal degen-  
1393           eration in alzheimer's disease. Alzheimer's & Dementia, 10:S425–S429, 2014.
- 1394 [152] Christoph Heiring, Björn Dahlbäck, and Yves A Muller. Ligand recognition and homophilic interactions  
1395           in tyro3 structural insights into the axl/tyro3 receptor tyrosine kinase family. Journal of Biological  
1396           Chemistry, 279(8):6952–6958, 2004.