- 552 Supplementary Information
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1. Inference of ECNs from longitudinal data

556 We consider that abundance of *O* bacterial operational taxonomic units (OTUs) are measured 557 over a period of *T* days in *S* subjects. We model the read counts $n_{os}(t)$ of OTUs "*o*" on any given 558 day *t* in subject *s* as a multinomial distribution:

$$p(\{n_{os}(t)\}) = \frac{N_s(t)!}{\prod_o n_{os}(t)!} \prod_o q_{os}(t)^{n_{os}(t)}$$
(S1)

where $N_s(t) = \sum_o n_{os}(t)$ is the total read count on a given day and $q_{os}(t)$ are the underlying propensities for individual OTUs. We model these propensities using the exponential Gibbs-Boltzmann distribution which allows us to capture large variations in OTU abundances³³

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$$q_{os}(t) = \frac{1}{\Omega_{st}} \exp\left(-\sum_{k=1}^{K} z_k(t)\theta_{kos}\right)$$
(S2)

where $z_k(t)$ are time-specific latents that are shared by all OTUs and subjects, and θ_{kos} are OTUand subject-specific loadings that are shared across all time points. The number K of latents/loadings is chosen such that $K \ll 0, T$ thereby achieving a lower dimensional description of the time series data. We obtain the zs and the θs using the maximum likelihood approach.

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569 To that end, we write down the log-likelihood of the data:

- 570 $L = const. + \sum_{t,o,s} n_{os}(t) \log q_{os}(t).$ (S3)
- 571 The constant term of the likelihood does not depend on the parameters and can thus be omitted 572 in likelihood maximization. Simplifying using Eq. S1 and S2, we have

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$$L = -\sum_{t,o,s,k} N_s(t) x_{os}(t) z_k(t) \theta_{kos} - \sum_{t,s} \log \Omega_{st}$$
(S4)

574 Here $x_{os}(t) = n_{os}(t)/N_s(t)$ is the relative abundance of OTU o at time t. We obtain the 575 gradients

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$$\frac{\partial L}{\partial z_k(t)} = -\sum_{o,s} N_s(t) (x_{os}(t) - q_{os}(t)) \theta_{kos} \text{ and}$$
(S5)

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$$\frac{\partial L}{\partial \theta_{kos}} = -\sum_{t} N_s(t) z_k(t) \left(x_{os}(t) - q_{os}(t) \right)$$
(S6)

578 We use gradient ascent algorithm to find the latents and the loadings that maximize the 579 likelihood.

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581 For a given *K*, using the microbiome data $x_{os}(t)$ and starting from random initialization, we first 582 simultaneously infer the latents $z_k(t)$ and the features Θ_{kos} . We observe that the $T \times K$ matrix

z of latents can be multiplied by an invertible matrix $B(z \rightarrow zB)$ and the corresponding matrix 583 $K \times O \times S$ matrix of features can be multiplied by the inverse B^{-1} ($\Theta \rightarrow B^{-1}\Theta$) and the 584 abundance predictions from the model do not change. Therefore, we use the Gram-Schmidt 585 procedure to orthogonalize the matrix of latents such that $z \rightarrow z'$ where $z'^T z' = I_K$ is an identity 586 matrix. For any matrix of latents z, the matrix multiplier B that leads to the orthonormal 587 588 transformation can be found by solving the equation $B^T(z^Tz)B = I_K$. Once B is identified, we also transform the Θ matrix ($\Theta \rightarrow \Theta' = B^{-1}\Theta$). At the end of this procedure, we end up with 589 orthonormal latents \mathbf{z}' and corresponding features $\mathbf{\Theta}'$ that correspond to the same abundances 590 591 as z and Θ . For the sake of simplicity of notation, we drop the primes.

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593 Next, we model the dynamics of the orthonormal latents using a linear dynamical system:

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$$z_k(t) = \sum_{k'} A_{kk'} z_{k'}(t) + u_k + \eta_k(t)$$
(S7)

595 where we assume that $A_{kk'} = A_{k'k}$, and $\eta_k(t)$ are Gaussian distributed uncorrelated noise 596 vectors: $\langle \eta_k(t_1)n_{k'}(t_2) \rangle = \delta_{12}\delta_{kk'}$ where δ_{ab} is the Kronecker delta function. Our task is to find 597 the interaction matrix A and the vector u that fits this model. We achieve this using squared error 598 minimization. We write

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$$E(\boldsymbol{A}, \boldsymbol{u}) = \sum_{t} \left(z_k(t) - z_{k, pred}(t) \right)^2$$
(S7)

600 where $z_k(t)$ is the inferred latent and $z_{k,pred}(t)$ is the corresponding prediction using $z_k(t-1)$ 601 and Eq. S7. We restrict the summation only over time points t such that measurements are 602 available for time points t and t - 1. The squared error is minimized using a simulated annealing 603 approach. Once the matrix A is identified, we transform the orthonormal latents $z_k(t)$ into 604 ecological normal modes $y_k(t)$ as described in the manuscript.

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606 The scripts for obtaining ECNs y and corresponding loadings Φ from read count data can be 607 found at: https://github.com/mayar-shahin/EMBED.

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609 2. Generating in silico data

610 **Out of Phase Sinusoids.** We generated 40 OTU abundances for 30 time points. The un-611 normalized abundances of 20 OTUs followed sinusoidal oscillation: $X(t) = A_1(B_1 \sin(0.5t) + 1)$ 612 and the un-normalized abundances of the other 20 OTUs followed phase-shifted oscillation with 613 the same frequency $X(t) = A_1(B_1 \cos(0.5t) + 1)$. As and Bs are uniform random numbers 614 between 0 and 1. We normalize the generated abundances to produce the underlying probability 615 distribution of the data. We used multinomial sampling with a sequencing depth of 25000 to 616 generate relative OTU abundances on each day (**SI Fig. 1**, panels **A** and **B**).

Exponentials and Sinusoids. We generated 40 OTU abundances for 30 time points. 20 OTUs followed an exponential decay $X(t) = 10A_1 \exp(-0.1t)$, 10 OTUs oscillated according to $X(t) = A_2(B_2 \sin(0.5t) + 1)$ and 10 OTUs oscillated with double the frequency X(t) = $A_3(B_3 \sin(t) + 1)$. As above, *A*s and *B*s are uniform random numbers between 0 and 1. We sampled the abundances using the multinomial distribution as above (**SI Fig. 1**, panels **C** and **D**).

Sum of Sinusoids. We generated 40 OTU abundances for 30 time points. 20 OTUs followed a single high frequency oscillation $X(t) = A_1(B_1 \cos(1.5t) + 1)$. The remaining OTU abundances were generated by the addition of two different sinusoids: $X(t) = A_2(B_2 \sin(0.5t) + B_3 \sin(t) + 1)$. As above, As and Bs are uniform random numbers between 0 and 1. We sampled the abundances using the multinomial distribution as above (**SI Fig. 1**, panels **E** and **F**).

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3. Obtaining the microbiome time series from sequencing data

Murine gut microbiome response to oscillating diet. We downloaded the microbiome abundance time series data on mice fed an alternating diet of high fat high sugar chow (HFHS) and low-fat plant polysaccharide chow (LFPP) from Carmody et al.²⁵ as described previously¹⁴. Each mouse that was subjected to an oscillatory diet was treated separately. Based on our previous work on technical noise in 16s measurements, we only analyzed OTUs with mean abundances > 0.1%¹³ averaged across all time points and mice. On every day, the abundances of the rest of the OTUs were lumped together in a single meta-species.

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639 **Murine gut microbiome response to antibiotics.** We downloaded microbiome abundance data 640 from Ng et al.¹⁰. We focused on the data where mice were administered the antibiotic 641 ciprofloxacin. Out of the 10 cages in which the mice were housed, we omitted data from cages 642 2, 4, 5, and 8 where many time points were missing. As above, we analyzed OTUs with mean 643 abundance > 0.1% and combined the rest of the OTUs in a meta-species.

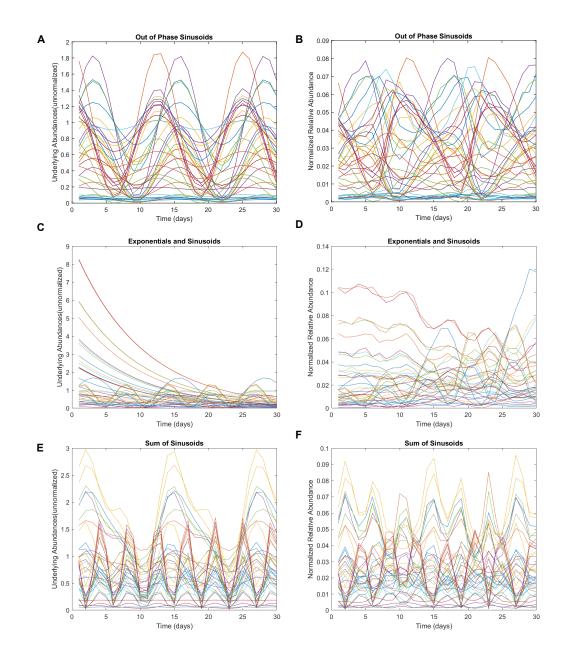
644 645

4. Performing CLR and SSVD on 4 data sets

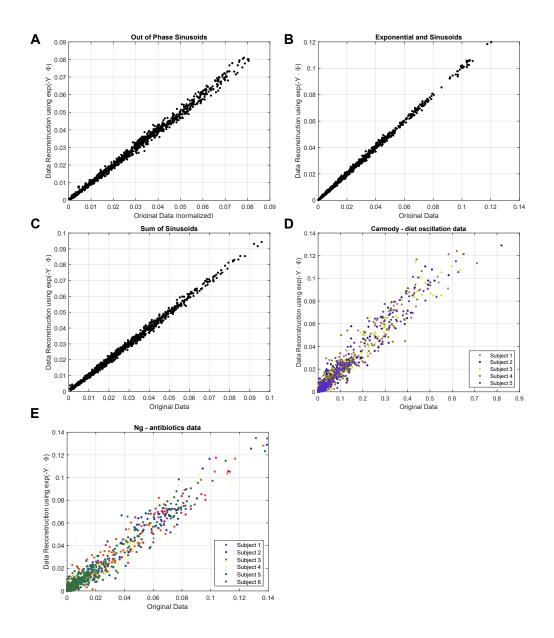
We downloaded 4 publicly available data sets from 4 different studies^{10–12,25}. Each data set 646 647 comprised microbiome abundance tables for multiple subjects, see SI Table 1. We use the package released by Martino et al.³² (https://github.com/biocore/gemelli) to perform Robust 648 Centered-Log Ratio transform (CLR) on the abundances followed by sparse singular value 649 650 decomposition. To test how well the dimensionality reduced version capture the data, we 651 calculate an approximate reconstruction of the abundance time series using the first K singular 652 values. As suggested by Martino et al. ³¹, the resulting approximation was re-exponentiated and 653 then normalized. We Then calculated the KL-divergence with the true abundances.

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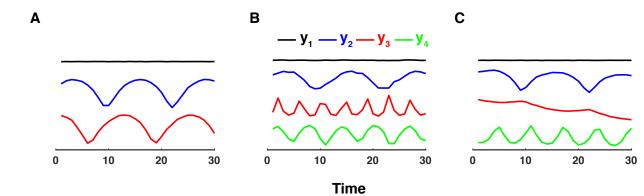
656 Supplementary Figures



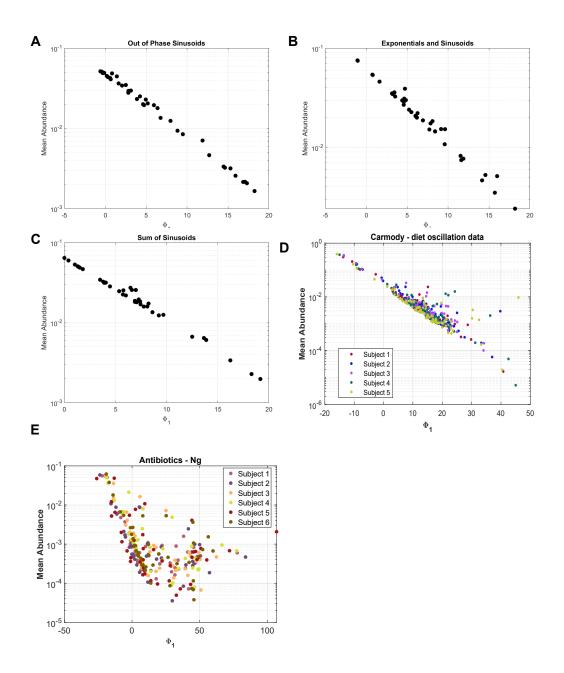
Supplementary Figure 1. Collective abundance variation of bacterial species in *in silico* data sets
 representing out of phase oscillations (A and B), exponential decays and oscillations (C and D), and sum
 of sinusoids (E and F).



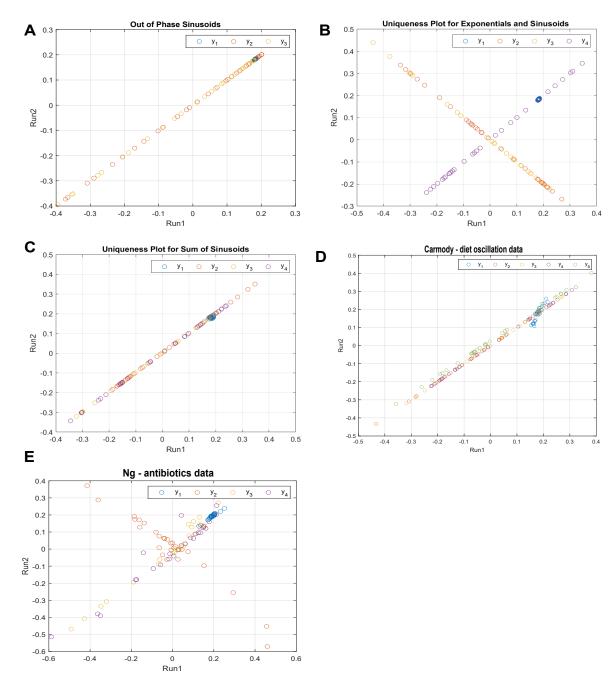
Supplementary Figure 2. EMBED-based reconstruction of time series (y axis) compared to the microbiome
 time series data (x-axis) for the three *in silico* data sets (A, B, and C) and the two experimental data sets
 (D and E) considered in this study.



Supplementary Figure 3. EMBED-based inference of ecological normal modes for the three in silico data
 sets. A: out of frequency oscillations, B: sum of sinusoids, and C: exponential decay and sinusoids

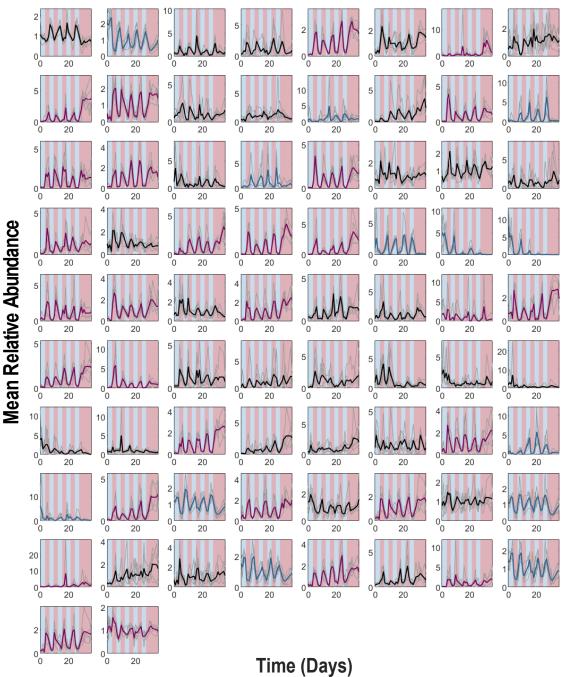


Supplementary Figure 4. Correlation between mean abundance of OTUs and their weight in the first 688 loading Φ_1 for the *in silico* data sets (A, B, and C) and the two experimental data sets (D and E) considered 689 in this study.

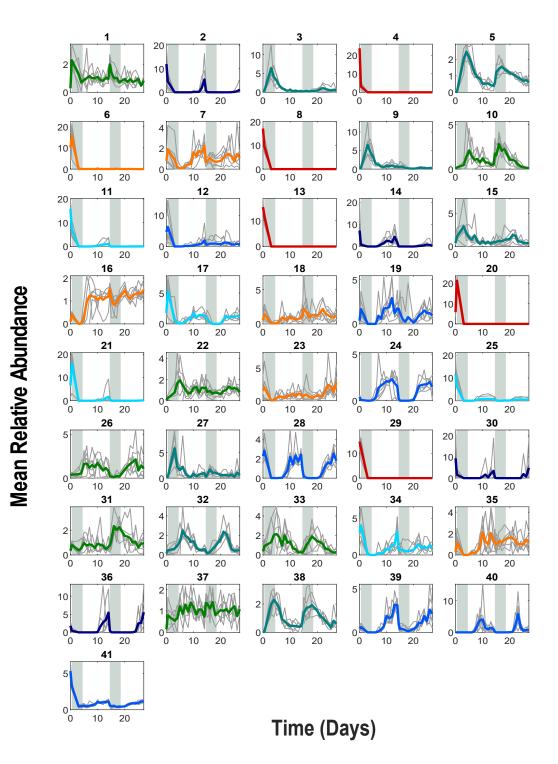


697 Supplementary Figure 5. Correlation plot showing that the ecological normal modes (ECNs) inferred using
 698 EMBED are unique (up to a sign). The x- and the y-axis represent the ECNs inferred in two independent
 699 runs.

HFHS LFPP



Supplementary Figure 6. Abundance time series of individual OTUs in the diet oscillation study. The gray
 lines represent abundances in individual subjects. The dark lines represent averages over subjects. The
 colors represent the cluster identities in main text Figure 2. HFHS: High Fat High Sugar diet and LFPP: Low
 Fat Plat Polysaccharide diet.



Supplementary Figure 7. Abundance time series of individual OTUs in the antibiotics-treatment study.
The gray lines represent abundances in individual subjects. The dark lines represent averages over
subjects. The colors represent the cluster identities in main text Figure 3. The gray bars represent the
duration of time when the antibiotic was administered.