Detection of allele-specific expression in spatial transcriptomics with spASE

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Abstract

Allele-specific expression (ASE), or the preferential expression of one allele, can be 2 observed in transcriptomics data from early development throughout the lifespan. 3 However, the prevalence of spatial and cell type-specific ASE variation remains un-4 clear. Spatial transcriptomics technologies permit the study of spatial ASE patterns 5 genome-wide at near-single-cell resolution. However, the data are highly sparse, 6 and confounding between cell type and spatial location present further statistical 7 challenges. Here, we introduce spASE (https://github.com/lulizou/spase), a 8 computational framework for detecting spatial patterns in ASE within and across 9 cell types from spatial transcriptomics data. To tackle the challenge presented by 10 the low signal to noise ratio due to the sparsity of the data, we implement a spatial 11 smoothing approach that greatly improves statistical power. We generated Slide-12 seqV2 data from the mouse hippocampus and detected ASE in X-chromosome 13 genes, both within and across cell type, validating our ability to recover known ASE 14 patterns. We demonstrate that our method can also identify cell type-specific ef-15 fects, which we find can explain the majority of the spatial signal for autosomal 16 genes. The findings facilitated by our method provide new insight into the uncharac-17 terized landscape of spatial and cell type-specific ASE in the mouse hippocampus. 18

Introduction

In diploid organisms, allele-specific expression (ASE) refers to the imbalanced ex-20 pression of the two parental alleles for a given gene. ASE has been well-studied in 21 the context of epigenetic phenomena such as genomic imprinting and X-chromosome 22 inactivation (XCI) [1, 2, 3], where expression from one allele is silenced. Spatial pat-23 terns of ASE have long been observed as a consequence of XCI in female organ-24 isms, where the random silencing of either the maternal or paternal X-chromosome 25 in early development is passed to daughter cells, resulting in visible clusters of ASE 26 [4, 5, 6]. By contrast, although studies in bulk and single-cell RNA-sequencing data 27 have revealed widespread variability in ASE throughout the autosome across tis-28 sues and cell types [7, 8, 9, 10, 11, 12, 13, 14, 15, 16, 17, 18, 19, 20], relatively little 29 is known about the prevalence of spatial ASE therein. 30

Spatial transcriptomics technologies now provide the opportunity to study spatial 31 ASE patterns genome-wide. For example, Slide-segV2 [21, 22] has high resolution 32 which enables near-single-cell quantification of ASE with 2D spatial information. 33 However, these data are limited by highly sparse read counts in comparison to 34 bulk or single-cell sequencing technologies, which is further exacerbated by the 35 requirement that reads align uniquely to one allele. In addition, cell type, which 36 drives the majority of variability observed in single-cell data, is highly correlated 37 with spatial location, especially in solid tissue [23]. Therefore, it is important to 38 distinguish between spatial and cell type-specific ASE, which could arise from and 39 contribute to distinct underlying biological mechanisms. 40

Several statistical and computational methods have been developed for studying 41 ASE in bulk and single-cell RNA-seq data [24, 25, 26, 27, 28, 29, 30, 31]. Some fo-42 cus on estimating allele-specific transcriptional bursting kinetics for individual genes 43 in homogeneous populations of cells [15, 30, 31]. Here, we instead focus on the 44 problem of estimation and inference for the maternal allele probability p for a given 45 gene across 2D space, and we consider how p may vary with cell type. To model 46 p in bulk and single-cell RNA-seq, multiple methods have used a beta-binomial 47 framework, which can flexibly account for overdispersion from unknown technical 48 and biological variability [26, 27, 28]. An additional advantage of this model is that 49 it can be parameterized as a generalized linear model (GLM) [32, 33], allowing for 50 maximum likelihood estimation of p while incorporating covariates of interest such 51 as cell type. 52

The issue of estimating smooth functions from sparsely sampled data has been 53 well-studied [34, 35, 36, 37, 38], and multiple solutions have been developed and 54 implemented as computational methods [39, 40]. In the case of allele-specific spa-55 tial transcriptomics data, although the read count measured at individual spatial 56 coordinates may be low, smoothing spline methods can increase power by lever-57 aging information from local neighborhoods of pixels. Generalized additive models 58 are GLMs that incorporate smoothing splines into a regression framework, enabling 59 estimation of the smooth spatial function as well as hypothesis testing for spatial 60 functions deviating from a constant [38, 40]. 61

Here, we present spASE, a computational framework for detecting genes with 62 significant ASE patterns in spatial transcriptomics data. We employ a hierarchical 63 beta-binomial smoothing approach based on thin plate regression splines [36, 41] to 64 estimate 2D allele probability functions and detect spatially significant genes. Given 65 the high correlation between cell type and spatial location in solid tissue, our method 66 permits control for cell type effects as well as any other potential covariates of in-67 terest. Through simulations, we confirm the power and false positive rate control of 68 our method even in highly sparse settings such as those observed in allele-resolved 69 spatial transcriptomics. Additionally, we generate allele-specific Slide-seqV2 data 70 from the hippocampus of an F1 hybrid mouse and find that we are able to recover 71 known patterns of ASE due to XCI in highly-expressed X-chromosome genes, both 72 within and across cell types. We further show that our method can detect cell type-73 specific ASE, which we find can explain most of the spatial signal observed in au-74 tosomal genes such as Ptgds. Overall, we report new insights into the uncharac-75 terized landscape of spatial and cell type-specific ASE in the mouse hippocampus, 76 thus demonstrating the utility of spASE for detecting known and novel patterns of 77 ASE in spatial transcriptomics. 78

79 **Results**

A beta-binomial framework for modeling allele-specific spatial transcriptomics

A statistical challenge for allele-specific spatial transcriptomics is that spatial and cell type effects can be confounded (Figure 1a,b). We therefore developed a computational framework that can account for both these sources of variability. Specifically, we developed a beta-binomial GLM that provides a flexible approach to estimation, inference, and visualization of ASE in spatial transcriptomics. We denoted the counts from the maternal allele for gene g and pixel i with Y_{gi} and assumed it followed the distribution:

$$Y_{gi} \mid \lambda_{gi} \sim \mathsf{Binomial}(n_{gi}, \lambda_{gi}),$$

with n_{gi} the observed total UMI count for gene g and pixel i, summing both alleles, and λ_{gi} the probability that a transcript from gene g is from the maternal allele. We assumed that λ_{gi} follows a beta distribution with mean p_{gi} and variance $\phi_g p_{gi}(1-p_{gi})$. Here, p_{gi} is the mean maternal allele probability and ϕ_g is a gene-specific overdispersion parameter ranging from 0 to 1 that accounts for biological and technical variability not explained by binomial sampling.

⁹⁵ To account for spatial and cell type effects, we created a logit-linear GLM,

$$logit(p_{gi}) = \beta_{g0} + f(x_i, y_i) + \sum_{k=1}^{K} \beta_{gk} X_{ik},$$
(1)

⁹⁶ with x_i and y_i the spatial location of pixel i, $f(x_i, y_i)$ a smooth function of location, ⁹⁷ the X_{ik} 's indicator functions equal to 1 if pixel i is from cell type k, and the β_{gk} ⁹⁸ parameters representing gene-specific cell type effects. Note that the β_{gk} can be ⁹⁹ interpreted as the change in log-odds, compared to the reference cell type, of a ¹⁰⁰ maternal allele transcript in gene g and cell type k.

¹⁰¹ The spatial effect function was modeled as a thin plate spline [36] defined by

$$f(x_i, y_i) = \sum_{j}^{J} \theta_{gk} B_j(x_i, y_i),$$

with $B_j(x, y)$ the smooth basis function for the spline and θ_{gk} the gene-specific parameters that define gene-specific spatial effects. With this definition of f, all the terms in (1) are linear and define a GLM. We can therefore obtain maximum likelihood estimates (MLEs), standard errors, and confidence intervals for all parameters using GLM theory and software. Furthermore, we can test for spatial effects by performing a likelihood ratio test comparing the model with f to a model without space (see Methods for details).

¹⁰⁹ In addition to fitting spatial ASE across all cell types, we can also fit a cell type-¹¹⁰ specific version of model (1) as

$$\mathsf{logit}(p_{gik}) = \gamma_{g0k} + f_k(x_i, y_i)$$



Figure 1: Schematic of detecting allele-specific expression in spatial transcriptomics using spASE. (a) Input is alleleand cell type-resolved spatial transcriptomics with UMIs. Each shape represents a different cell type, and the color indicates the fraction of observed UMIs that were from the maternal allele. (b) Point estimates and confidence intervals for the estimated maternal allele probability (estimated p maternal) for each cell type. (c) Visualization of the estimated maternal probability function, not controlling for cell type. (d) Visualization of confidence intervals (gray shaded region) around the MLE in a 1D cross-section. The solid line indicates the estimated maternal probability along the black dashed line from c. Light gray dashed line indicates the null of p = 0.5. (e) 2D z-score plot visualizing region-level significance of the estimated function from c. (f) Estimated cell type-specific function for the circle cell type from a.

for all pixels *i* belonging to cell type k, where p and f have been modified to depend on cell type. Note that only certain genes within certain cell types provide enough counts and therefore power to fit our cell type-specific spatial model.

After estimating parameters by maximum likelihood for each gene, we can visualize the smooth maternal probability function across 2D coordinates (Figure 1c), and for any given spatial cross-section, we can additionally visualize confidence intervals (Figure 1d). Region-specific significance can also be assessed using 2D z-score maps (Figure 1e). Running spASE on individual cell types generates spatial ASE maps for cell type-specific estimation (Figure 1f). spASE uses the likelihood ratio to rank genes according to spatial effects variability.

To evaluate the performance of our method, we generated simulated spatial 121 transcriptomics data under a wide variety of sparsity and overdispersion conditions 122 (Supplementary Figure S1, Methods). We calculated the power and false positive 123 rate, and we computed p-values to detect significant spatial ASE. We observed that 124 power decreased as overdispersion increased (Supplementary Figure S1a); how-125 ever, we found that power is at least 70% even for genes with high overdispersion 126 $(\phi = 0.8)$ and as few as 50 pixels with low UMI coverage (e.g. less than 10 UMIs 127 per pixel). With at least 100 pixels for a given gene, the power across all scenarios 128 was at least 85%, even with as low as 1 UMI per pixel. Using a p-value threshold 129 of $p \leq 0.01$, we found that the false positive rate approached the nominal rate of 130 0.01 as the number of pixels increased, in concordance with the expected asymp-131 totic guarantees of our model (Supplementary Figure S1b,c). We also evaluated 132 confidence interval coverage as a function of sample size and number of UMIs per 133 pixel (Supplementary Figures S2, S3), and we found that the beta-binomial model 134 maintained near-95% coverage across all scenarios. 135

spASE identifies spatially-significant ASE genes and smooths over sparse allele-specific spatial transcriptomics signal

To test spASE on allele-specific spatial transcriptomics data, we generated Slide-138 segV2 data of an F1 hybrid CAST/EiJ x 129S1/SvImJ (CAST x 129) mouse hip-139 pocampus and surrounding region (see Methods). We aligned 150bp reads to a 140 pooled CASTx129 transcriptome and only considered reads that uniquely aligned to 141 one allele. We used RCTD [23] to call cell types using a single-cell RNA-sequencing 142 reference of the mouse hippocampus [42], and we filtered to pixels with a high like-143 lihood of sourcing UMIs from a single cell type (Figure 2a). Based on results from 144 our simulations (Supplementary Figure S1), we filtered genes with non-zero UMI 145 counts on at least 100 pixels. Using these filtering criteria resulted in 4,140 genes 146 for downstream analysis, which were expressed on a median of 210 pixels (IQR: 147 140-384 pixels) with a median number of UMIs per pixel of 1.06 (IQR: 1.04-1.1 148 UMIs/pixel) (Supplementary Figure S4). 149

¹⁵⁰ We then fit our model with and without the cell type covariates X_{ik} in the model ¹⁵¹ (see Methods). We found that, compared to autosomal genes, a higher propor-¹⁵² tion of X-chromosome genes had likelihood-ratio-test significance (Figure 2b), in ¹⁵³ concordance with the expected patterns of XCI in the X-chromosome. The *p*-value ¹⁵⁴ distribution for autosomal genes was closer to uniform distribution, indicative of less-¹⁵⁵ frequent spatial ASE effects. After controlling for cell type, the autosomal distribution ¹⁵⁶ remained similar, while the distribution for the X-chromosome had a lesser skew,



Figure 2: spASE identifies spatially significant ASE genes and smooths over sparse ASE spatial transcriptomics signal. (a) Map of cell types identified by RCTD in the Slide-seqV2 data generated in this study. Each point represents a pixel classified as a singlet of that cell type. (b) Distributions of *p*-values calculated by spASE for autosomal (A) and X-chromosome (X) genes in the real Slide-seqV2 data, not controlling for cell type ("no cell type") and controlling for cell type ("cell type"). (c) Raw data for *Hpca*, showing higher coverage in the hippocampal formation and sparse coverage in the adjacent regions. (d) Estimated 2D maternal probability function for *Hpca*, with crosshairs indicating the x = 3 and y = 3 lines, along which point estimates and confidence intervals are plotted in (e) and (f), respectively.

¹⁵⁷ consistent with some the genes appearing to have spatial effects due to confound-¹⁵⁸ ing with cell type effects.

Using a false discovery rate (FDR) threshold of $q \le 0.01$, we found ten genes with a spatially significant pattern, of which six were from the X-chromosome (Supplementary Table S1). However, after controlling for cell type, only three genes, two (*Tspan7* and *Plp1*) on the X chromosome and one (*Sst*) autosome, were significant. Other genes, including *Nrip3* and *Ptgds*, were no longer significant after controlling for cell type, indicating that cell type differences were the main driver of spatial ASE for these genes.

spASE accounts for biological and technical noise to avoid detecting false positive ASE. For example, the gene *Hpca* was determined by spASE to not have signifcant spatial ASE (Figure 2c,d, *p*-value = 0.53). Although *Hpca* is highly expressed in the hippocampal formation, sparse expression in the adjacent regions resulted in noisier estimates and wide confidence intervals outside the hippocampus (Figure 2e,f). In general, such visualizations enabled by spASE allow for the assessment of both overall significance as well as position-specific significance across space.

spASE detects spatial patterns of XCI across and within cell type in the mouse hippocampus

Next, we used spASE to estimate the maternal allele probability function for X chromosome genes and found that the patterns for almost all significant X-chromosome
 genes were similar and anti-correlated with *Xist* expression (Figure 3a-c, Supple-



Figure 3: spASE detects spatial ASE in X-chromosome genes across and within cell type in the mouse hippocampus from Slide-seqV2 data. (a) Smoothed maternal allele probability functions for X-chromosome genes highly expressed in the mouse hippocampus and detected as significant (q-value ≤ 0.01): *Tspan7* and (b) *Plp1*. Red color indicates bias towards maternal, blue towards paternal, and white indicates both maternal and paternal alleles are present. The outline of the CA1, CA3, and dentate cell type regions is depicted in the dotted gray areas for reference. (c) Same as a-b for *Xist.* (d) Same as a-c but only using astrocyte pixels for *Tspan7*. (e) Same as a-c but only using oligodendrocyte pixels for *Plp1.* (f) Same as a-c for *Tceal3.* (g) 2D *z*-score plot computed from combining all non-*Xist* X-chromosome genes. (g) 2D *z*-score plot for *Xist.*

¹⁷⁸ mentary Figure S5), reflecting the expected mosaicism due to XCI by *Xist*. We also ¹⁷⁹ found that patterns of XCI were preserved within individual cell types. For exam-¹⁸⁰ ple, *Tspan7*, which is relatively highly expressed in astrocytes (Figure 3d, S6), and ¹⁸¹ *Plp1*, which is highly expressed in oligodendrocytes (Figure 3e, Supplementary Fig-¹⁸² ure S7), were both estimated to have maternal probability functions anti-correlated ¹⁸³ with *Xist*.

One X-chromosome gene, *Tceal3*, exhibited a strong paternal skew unlike the 184 rest of the X-chromosome (Figure 3f, Supplementary Figure S8). However, the es-185 timated Tceal3 maternal probability still had a similar trend to the observed XCI 186 pattern, with a high paternal bias around the hippocampus and a near-biallelic pat-187 tern in the periphery. We investigated the Tceal3 locus and found that another 188 nearby gene less than 100kb away, Morf412, also exhibited a strong paternal bias. 189 Tceal6, a paralog of Tceal3, also showed a paternal bias in a similar pattern to that 190 of Tceal3; however, other genes in the Tceal family, such as Tceal5, did not show 191 the same bias (Supplementary Figure S8). 192

We then constructed a consensus XCI map by combining the UMI counts of 193 all of the X-chromosome genes excluding Xist and fitting the model on the merged 194 spatial profile. We visualized the significance at a region-specific level by computing 195 and plotting z-scores in 2D (Figure 3g, Supplementary Figure S5). We found that 196 the region significant for paternal X-chromosome expression was located precisely 197 around the CA1, CA3, and dentate cell-type layers of the hippocampus as well as 198 around a cluster of interneurons, and was anti-correlated with Xist (Figure 3h). Xist 199 had fewer spatial regions reaching significance, reflecting its lower UMI coverage 200 and thus wider confidence intervals in most areas (Supplementary Figure S9). 201

spASE identifies cell type-driven spatial ASE in the autosome of the mouse hippocampus

We next investigated autosomal spatially significant genes with spatial ASE that 204 was explained by cell type-driven ASE. Recall that several genes, including Nrip3 205 and Ptgds, no longer possessed significant spatial ASE after controlling for cell 206 type, indicating cell type-driven ASE. To quantify such cell type-driven ASE, we 207 used spASE to estimate the overall maternal allele probability for each cell type, 208 revealing several genes previously unknown to exhibit cell type-specific ASE (Figure 209 4a). For these genes, spASE's estimated spatial ASE patterns were primarily driven 210 by cell type localization distributions (Figure 4b-d). For example, Nrip3, one of the 211 most statistically significant autosomal genes, had a high maternal bias in CA1, 212 CA3, and dentate cell types, driving a strong maternal signal observed in the z-213 score plot (Figure 4b). *Ptgds*, which was highly expressed in both oligodendrocytes 214 and endothelial tip cells, had a strong paternal bias in oligodendrocytes but not 215 endothelial cells (Figure 4c,e). Sst exhibited a maternal bias which was enhanced 216 in an interneuron subtype localizing primarily outside of the hippocampus (Figure 217 4d: Supplementary Figure S10). 218

We examined whether this high cell type-specificity of ASE for these genes could 219 be explained by genetic differences between the CAST (maternal) and 129 (pater-220 nal) mouse strains. Specifically, we investigated if a SNP could alter the binding 221 affinity of a cell type-specific transcription factor at either the promoter or a putative 222 enhancer. We analyzed single-cell ATAC-seq (scATAC) data from the mouse hip-223 pocampus [44] and searched for instances of SNPs overlapping known transcription 224 factor binding site (TFBS) motifs in peaks within 50kb upstream and downstream of 225 each gene (Methods). We found a peak in the promoter of Gm35287 approximately 226 8kb upstream of *Ptads* (Supplementary Figure S11), which is predicted to have a 227 TSS-distal with enhancer-like signature for *Ptgds* [45]. Furthermore, we found that 228 this peak has a high co-accessibility ($r^2 = 0.74$) with the *Ptgds* promoter peak for 229 cell types in common between the scATAC data and our spatial data (Figure 4f). 230 In particular, both peaks are preferentially open in oligodendrocytes. We found a 231 SNP, rs8255993, overlapping a known transcription factor motif, PB0044.1, which 232 corresponds to the gene *Mtf1* (Supplementary Figure S11), which is also highly ex-233 pressed in oligodendrocytes (Figure 4g) [43]. The SNP is A in the paternal strain 234 and C in the maternal, and this position has a strong A signal in the TFBS position 235 weight matrix for *Mtf1* (Figure 4h). Thus, the preferential binding of *Mtf1* to the pa-236 ternal allele at this distal enhancer is a likely mechanism driving the paternal bias 237 of *Ptads* observed in oligodendrocytes. 238



Figure 4: spASE identifies cell type-driven spatial ASE in the autosome. (a) MLEs and associated confidence intervals for the maternal probability p for three of the top autosomal gene hits (q-value ≤ 0.01), Nrip3, Ptgds, and Sst. x-axis: total UMI counts summed across all pixels. (b) 2D z-score plot for Nrip3. The hippocampal formation is outlined with dotted black lines. (c-d) Same as b for Ptgds and Sst, respectively. (e) Raw data for Ptgds for endothelial tip (triangles) and oligodendrocyte (circles) pixels. Size of the point indicates the total number of UMI present at that pixel, and color indications the fraction of the total UMI that were from the CAST (maternal) allele. (f) Average sci-ATAC-seq peak accessibility of the Ptgds promoter peak and the nearby (~8kb away) peak in Gm35287 for the cell types overlapping between the sci-ATAC-seq data set and the Slide-seq data. (g) Single-cell RNA-seq expression for Mtf1 from the Mouse Brain Atlas [43]. Each point represents a cluster that was classified as one cell type. (h) Position weight matrix for PB0044.1 (Mtf1) with 129 (paternal) and CAST (maternal) reference sequences shown on the top and bottom, respectively. Black arrow points to the SNP position of interest.

Discussion

Allele-resolved spatial transcriptomics suffers from high sparsity in comparison to 240 bulk and single-cell sequencing, and confounding between cell type and spatial 241 location present further statistical challenges. Here, we describe a statistical ap-242 proach and software (spASE) which allows for estimating and visualizing 2D allele 243 probabilities for sparsely expressed genes, as well as for testing spatial significance 244 while controlling for user-specified covariates such as cell type. Through simula-245 tions, we demonstrate that our method maintains high power to detect ASE even 246 with as few as 100 pixels and as low as 1 UMI per pixel for a given gene. We gen-247 erated Slide-seqV2 data from an F1 female CASTx129 mouse hippocampus and 248 show that our method recovers known patterns of XCI both within and across cell 249 type (Figure 3). We further show that our method can identify cell type-specific ASE, 250 which if not accounted for can be confused with spatial signal (Figure 4). 251

The primary *in situ* validation of our method was in the X-chromosome, where we 252 found the same pattern of XCI both within and across cell types for multiple genes. 253 XCI is thought to occur early in embryonic development in female organisms, be-254 fore cell type differentiation [4], and the maternal and paternal chromosomes are 255 thought to be equally likely to be inactivated. Thus, the pattern we observed in our 256 data likely reflects randomly determined XCI in the early mouse embryo that propa-257 gated through to the adult hippocampus. This phenomenon can potentially explain 258 why the X-chromosome p-value distribution was slightly but not fully affected by 259 controlling for cell type (Figure 2b), as some nearby cell types may be derived from 260 the same X-inactivated progenitor cell. Notably, *Tceal3* exhibited a strong paternal 261 bias, but still had a spatial pattern that was similar to the general XCI pattern we 262 observed in other X-chromosome genes (Supplementary Figures S5, S8). Another 263 nearby gene, Morf412, also exhibited a paternal bias. Thus, the pattern we ob-264 served in *Tceal3* may be the combined result of XCI and another form of epigenetic 265 imprinting. 266

One limitation of our spatial ASE analysis is that low UMI coverage limits the 267 spatial resolution of ASE estimates. For example, within the XCI analysis, Xist was 268 lowly expressed (Figure 3h); however, for genes with higher coverage, such as *Plp1*, 269 it was possible to resolve the spatial ASE function further by increasing the degrees 270 of freedom used to construct the 2D basis functions. Due to our limited spatial 271 resolution, although we detected spatial patterns of XCI between the hippocampal 272 formation (paternal bias) and surrounding areas (maternal bias), it is likely that in-273 creased statistical power would be achieved and higher-resolution spatial patterns 274 would be uncovered given a higher-coverage dataset. 275

Similarly, although we found multiple instances of differential ASE across cell types as previously observed [20, 46], our analysis did not detect any spatial ASE in autosomal genes not explainable by cell type. We note that the statistical power was lower for the detection of spatial effects compared to the detection of cell-type differences. It is possible that autosomal spatial ASE effects might be detected given increased coverage and sample size.

We found that *Sst* exhibited a strong maternal bias for interneurons, particularly for a subtype located outside of the hippocampal formation with high *Sst* expression (Supplementary Figure S10). *Sst* is a well-known neuropeptide expressed throughout the brain which has been studied in the context of various neurological diseases

[47]. However, we were not able to detect a likely cell type-specific transcription 286 factor with a nearby binding site that was affected by strain-specific genetic varia-287 tion as we did for Ptgds, although it is possible that the bias may only affect this 288 subtype which is not represented in the scATAC-seq data set we used. Also, note 289 that Sst exibited low levels of expression in other cell types, which limited statistical 290 power. Overall, these findings demonstrate that our method is broadly applicable for 291 ASE discovery in spatial transcriptomics. Our rigorous computational approach will 292 inform future analyses on the variability and biological mechanisms driving spatial 293 and cell type-specific ASE. 294

295 Methods

²⁹⁶ Slide-seqV2 of CAST/EiJ x 129S1/SvImJ F1 mice

²⁹⁷ We obtained a female CAST/EiJ x 129S1/SvImJ (CASTx129) mouse from Jackson ²⁹⁸ laboratories. The CASTx129 cross contains \sim 23 million SNPs, or approximately 1 ²⁹⁹ SNP for every \sim 110 bp [48, 49]. This SNP density is approximately tenfold the SNP ³⁰⁰ density in human cells and thus provides high resolution to interrogate ASE. Slide-³⁰¹ seqV2 was performed as described previously [21, 22] on two adjacent, 10um-thick ³⁰² coronal slices of the hippocampus.

303 Alignment of Slide-seqV2 data

We generated a pooled CASTx129 transcriptome using the command create-hybrid 304 from the EMASE [50] software on the CAST and 129 transcript fasta files down-305 loaded from ftp://churchill-lab.jax.org/software/g2gtools/mouse/R84-REL1505/. 306 We then aligned 150bp reads to this pooled transcriptome with bowtie2 [51] using 307 the parameters -k 4 -p 16 --very-sensitive. We used a custom script (https: 308 //github.com/lulizou/spASE/blob/master/scripts/processBowtie2.py) for pro-309 cessing the aligned BAM file [52] to create a gene UMI count matrix only from reads 310 that uniquely aligned to one gene and one allele. We restricted attention to align-311 ments with 3 or fewer mismatches and only considered alignments that had the 312 fewest number of mismatches for that read. We overlaid data from the two slices by 313 rotating and shifting the slices to overlap according to the location of the hippocam-314 pal formation. 315

Beta-binomial model for allele-specific expression in spatial transcriptomics

Let n_{qi} denote the observed total counts of gene g at cell or pixel i and Y_{qi} denote 317 the observed maternal allele UMI counts for gene g at cell or pixel i. Let λ_{ai} denote 318 the unknown mean probability of observing a maternal allele for each transcript of 319 gene g at pixel i. We assume $Y_{gi}|\lambda_{gi} \sim \text{Binomial}(n_{gi}, \lambda_{gi})$, where n_{gi} is observed 320 and λ_{qi} is a random variable, independently distributed (conditional on ϕ_q and p_{qi} , 321 defined below) for each gene g and pixel i. We further assume that λ_{gi} follows a 322 beta distribution with mean p_{qi} and variance $\phi_q p_{qi}(1-p_{qi})$. The likelihood of this 323 model for a single gene q can be written as 324

$$\mathcal{L}(p_{gi},\phi_g;n_{g\cdot},y_{g\cdot}) = \prod_i \binom{n_{gi}}{y_{gi}} \frac{B(p_{gi}(1-\phi_g)/\phi_g + y_{gi},(1-p_{gi})(1-\phi_g)/\phi_g) + n_{gi} - y_{gi})}{B(p_{gi}(1-\phi_g)/\phi_g,(1-p_{gi})(1-\phi_g)/\phi_g)}$$
(2)

where B denotes the beta function. We used maximum likelihood estimation to 325 obtain the estimates \hat{p}_{gi} and $\hat{\phi}_g$ for each gene and the associated standard errors, 326 determined from the Fisher information. When using this model for single-cell data 327 to estimate ASE for a single gene, p_{qi} is assumed to be the same for all cells *i*; 328 when estimating p_{qi} for spatial transcriptomics, p_{qi} is assumed to be dependent on 329 pixel i as described below. Assuming asymptotic normality, we used these point 330 estimates and standard errors to construct 95% confidence intervals for \hat{p}_{ai} . Finally, 331 we used the Benjamini-Hochberg [53] procedure to produce q-values to control the 332 false discovery rate. 333

In the spatial setting, for each gene g, we model p_{qi} as a smooth spatially-varying 334 function. Specifically, we used thin plate regression splines [36, 40, 41] to estimate 335 smooth maternal allele probability surfaces for each gene. Thin plate regression 336 splines allow estimation of a smooth function of 2D coordinates. The number of 337 basis functions d determines the smoothness, with lower values of d corresponding 338 to smoother functions. Choice of d depends on the sparsity of the data, since lower 339 values of d reduce the variance, but at the risk of introducing bias. In the analy-340 sis presented here, we kept d constant across genes to ensure comparability. For 341 our sparse Slide-seqV2 data, we found that d = 10 provided enough complexity to 342 model allelic patterns in the hippocampus sample examined here while also main-343 taining power to detect significant differences, and we used d = 15 when plotting 344 estimated maternal allele probability functions using all pixels. We also demonstrate 345 reproducibility of results (i.e. genes detected as having a significant spatial pattern) 346 across a range of values for d (Supplementary Tables S1-S4). In practice, we rec-347 ommend visualization of the estimated probability function and confidence intervals 348 (e.g. Figure 2c-f) to guide selection of d. 349

In the model for cell type-specific spatial ASE detection, the term θ_{gk} corresponds to the effect size for cell type k for gene g. If no cell type annotations are available, or if a cell type effect does not exist (see likelihood ratio test below), then θ_{gk} can be assumed to be the same for all k.

To test whether there was a significant spatial pattern beyond cell type, we assumed a baseline model with only cell types as covariates and performed a likelihood ratio test comparing model (1) to the baseline model, i.e. for each gene g, we compute

$$\Lambda_g = -2(\ell(\mathbf{p}_g, \phi_g; \mathbf{n}_g, \mathbf{Y}_g, \mathbf{X}) - \ell(\mathbf{p}_g, \phi_g; \mathbf{n}_g, \mathbf{Y}_g, \mathbf{X}, \mathbf{x}, \mathbf{y}))$$
(3)

where ℓ is the log-likelihood computed from (1), \mathbf{p}_g is the vector of maternal probabilities at each pixel *i* for gene *g*, \mathbf{n}_g is the total number of UMIs at each pixel *i* for gene *g*, \mathbf{Y}_g is the total number of maternal-derived UMIs at each pixel *i* for gene *g*, \mathbf{X} is the *i* × *k* matrix of indicators of each cell type *k* at each pixel *i*, and \mathbf{x}, \mathbf{y} are the vectors of 2D spatial coordinates for each pixel *i*. This test statistic has an asymptotic χ_1^2 distribution, which we use to compute p-values for each gene.

For visualization, we plotted both the estimated smooth function for the MLE \hat{p}_{gi} (Figure 3a-c) as well as 2D *z*-score plots (Figure 3d-e, Figure 4e). *Z*-score plots were calculated on an evenly spaced grid of points over the sample by taking the point estimate logit(\hat{p}_{gi}) at each location and its associated standard error s_{gi} and computing $z_{gi} = \text{logit}(\hat{p}_{gi})/s_{gi}$.

369 Computational implementation

spASE is implemented as an R package (https://github.com/lulizou/spase). We generated thin plate regression splines using the R package mgcv [40]. Specifically, we used the smoothCon function to construct spline basis functions. As basis functions can depend on the scale of the spatial covariates, we used normalized coordinates and also normalized the basis functions after construction by subtracting the mean and dividing by the standard deviation. We used the implementation of the beta-binomial likelihood from the R package aod [54].

We ran spASE in multiple modes: 1) not controlling for cell type, 2) controlling for 377 cell type by allowing each cell type to have a different intercept term, and 3) allowing 378 for each cell type to have a different spatial pattern. For 1), we experimented with 379 using all pixels or only pixels confidently called as single cells (singlets). We found 380 that using all pixels allowed us to increase our power and resolution for estimation 381 of p_{ai} and ϕ_i , and for evaluating significance of spatial fits; thus, for visualization in 382 our figures, we use all pixels, unless the figure is specifically denoted to be a single 383 cell type. For results directly comparing to cell type models from 2) and 3), e.g. 384 significant genes detected when controlling and not controlling for cell type, we use 385 only pixels confidently classified in both cases to ensure comparable sample sizes 386 of pixels. 387

388 Simulation details

We simulated beta-binomial count data to evaluate the power, false positive rate, 389 and *p*-values calculated using spASE. For each simulated gene, to construct ran-390 dom spatial ASE patterns, we used a random linear combination of basis functions 391 calculated from the pixel locations of the Slide-seqV2 hippocampus data set us-392 ing degrees of freedom d = 15. We first sampled a random number of total pixels 393 $N \in \{50, 100, 250, 500\}$ and used a fixed number of UMIs per pixel. The range of 394 average UMIs per pixel for a single gene reached up to 14 in the real Slide-segV2 395 hippocampus data set; for testing purposes, we used values ranging from 1 to 50. 396 We drew the coefficients θ_q of a random linear combination of basis functions from 397 a standard normal distribution. Then, we chose a fixed overdispersion parameter 398 $\phi \in \{0.1, 0.3, 0.5, 0.8\}$, sampled the true binomial probabilities λ_{ai} for each pixel lo-399 cation *i* from the beta distribution, and simulated counts from the binomial model. 400

To evaluate the asymptotic confidence interval coverage properties, we simu-401 lated ground truth data generated under the beta-binomial model without spatial 402 covariates under a range of values for overdispersion, total coverage, and number 403 of cells or pixels (Supplementary Figure S2). Specifically, we tested total UMI con-404 ditions based on three dataset-specific distributions for total UMI for a single gene 405 across cells or pixels: Smart-seq3, which had generally high total UMI for each 406 cell [31]; Slide-seg high, which had low counts with a high right skew; and Slide-407 seq low, which had mostly low counts (less than 5 per pixel, Supplementary Figure 408 S3). We evaluated a range of values for the total number of cells or pixels the 409 gene was captured on. We also tested a range of overdispersion values, namely 410

 $\phi \in \{0, 0.1, \dots, 1\}$, and noticed the trend was the same as ϕ increased, thus we only show results here for $\phi = 0, 0.1, 0.8$. We set the maternal probability at p = 0.5 and simulated 5,000 iterations for each condition.

⁴¹⁴ Prediction of cell types in spatial transcriptomics

We ran RCTD [23] to predict cell types in our Slide-seq data using a previously 415 published single-cell RNA-seg reference for mouse hippocampus [42]. RCTD is a 416 supervised approach that learns cell type profiles from a single cell reference and 417 predicts cell type labels for spatial transcriptomics pixels using a Poisson log-normal 418 model that accounts for platform effects between single-cell RNA-seg and Slide-seg 419 data. We used a threshold of a likelihood difference of 100 between the minimum 420 score and singlet score to classify singlets. After running RCTD, we filtered to pixels 421 that were predicted to contain a single cell-type and added these cell type labels as 422 the covariates in model (1) to perform all cell type-specific analyses. 423

⁴²⁴ Mouse hippocampus scATAC-seq, cis-regulatory elements, and TFBS motifs

We analyzed previously published mouse hippocampus sci-ATAC-seq count ma-425 trices (GSE118987) [44]. The mice used in the study were the wild-type C57/B6 426 strain, and the data were aligned to the mm10 genome. We extracted the called 427 peak annotations and counts using the command scitools split. To quantify ac-428 cessibility of peaks, we computed the average count within cell types. We searched 429 for known TFBS motifs within peaks with the command matchMotifs from the R 430 package motifmatchr [55] using all motifs for *Mus musculus* in the JASPAR 2020 431 database [56]. We overlaid annotations of ENCODE cis-regulatory elements (cCREs) 432 [45] for mm10 downloaded from the UCSC Table Browser [57] at the Ptads locus 433 and visualized the annotations using IGV [58]. 434

435 Mouse strain SNP data

We obtained gene-specific SNP annotations for 129S1/SvImJ and CAST/EiJ with respect to the mm10 reference using REL1505 of the Mouse Genomes Project (https://www.sanger.ac.uk/sanger/Mouse_SnpViewer/rel-1505) [59, 60]. For each gene of interest, we searched for SNPs overlapping the gene, within 50kb upstream of the gene start, and 50kb downstream of the gene end.

441 Data availability

⁴⁴² Slide-seqV2 data generated in this study will be made available upon publication.

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Author contributions

RAI, FC, and LSZ conceived of the study. TZ and EM generated the Slide-seqV2
data. LSZ, RAI, DMC, FC, TZ, and MJA analyzed the data. LSZ, RAI, DMC, TZ,
and FC wrote the manuscript. All authors read and approved the final manuscript.

453 Competing interests

⁴⁵⁴ The authors declare no competing interests.

References

- ⁴⁵⁶ [1] Ferguson-Smith, A. C. & Surani, M. A. Imprinting and the epigenetic asymmetry between parental genomes. *Science* **293**, 1086–1089 (2001).
- ⁴⁵⁸ [2] Reik, W., Dean, W. & Walter, J. Epigenetic reprogramming in mammalian development. *Science* **293**, 1089–1093 (2001).
- [3] Knight, J. C. Allele-specific gene expression uncovered. *Trends Genet* **20**, 113–116 (2004).
- ⁴⁶² [4] Lyon, M. F. Gene action in the X-chromosome of the mouse (Mus musculus ⁴⁶³ L.). *Nature* **190**, 372–373 (1961).
- ⁴⁶⁴ [5] Disteche, C. M. & Berletch, J. B. X-chromosome inactivation and escape. *J* ⁴⁶⁵ *Genet* **94**, 591–599 (2015).
- [6] Wu, H. *et al.* Cellular resolution maps of X-chromosome inactivation: impli cations for neural development, function, and disease. *Neuron* 81, 103–119 (2014).
- ⁴⁶⁹ [7] Yan, H., Yuan, W., Velculescu, V. E., Vogelstein, B. & Kinzler, K. W. Allelic ⁴⁷⁰ variation in human gene expression. *Science* **297**, 1143 (2002).
- [8] Cowles, C. R., Hirschhorn, J. N., Altshuler, D. & Lander, E. S. Detection of regulatory variation in mouse genes. *Nat Genet* **32**, 432–437 (2002).
- [9] Bray, N. J., Buckland, P. R., Owen, M. J. & O'Donovan, M. C. Cis-acting variation in the expression of a high proportion of genes in human brain. *Hum Genet* **113**, 149–153 (2003).
- [10] Gimelbrant, A., Hutchinson, J. N., Thompson, B. R. & Chess, A. Widespread
 monoallelic expression on human autosomes. *Science* **318**, 1136–1140
 (2007).

- ⁴⁷⁹ [11] Zwemer, L. M. *et al.* Autosomal monoallelic expression in the mouse. *Genome* ⁴⁸⁰ *Biol* **13**, R10 (2012).
- ⁴⁸¹ [12] Chess, A. Mechanisms and consequences of widespread random monoallelic ⁴⁸² expression. *Nat Rev Genet* **13**, 421–428 (2012).
- [13] Deng, Q., Ramsköld, D., Reinius, B. & Sandberg, R. Single-cell RNA-seq re veals dynamic, random monoallelic gene expression in mammalian cells. *Sci- ence* 343, 193–196 (2014).
- [14] Reinius, B. & Sandberg, R. Random monoallelic expression of autosomal
 genes: stochastic transcription and allele-level regulation. *Nat Rev Genet* 16, 653–664 (2015).
- ⁴⁸⁹ [15] Larsson, A. J. M. *et al.* Genomic encoding of transcriptional burst kinetics. ⁴⁹⁰ *Nature* **565**, 251–254 (2019).
- ⁴⁹¹ [16] Rv, P., Sundaresh, A., Karunyaa, M., Arun, A. & Gayen, S. Autosomal Clonal ⁴⁹² Monoallelic Expression: Natural or Artifactual? *Trends Genet* (2020).
- [17] Lee, C., Kang, E. Y., Gandal, M. J., Eskin, E. & Geschwind, D. H. Profiling
 allele-specific gene expression in brains from individuals with autism spectrum
 disorder reveals preferential minor allele usage. *Nat Neurosci* 22, 1521–1532
 (2019).
- ⁴⁹⁷ [18] Chess, A. Monoallelic Gene Expression in Mammals. *Annu Rev Genet* **50**, 317–327 (2016).
- ⁴⁹⁹ [19] Pickrell, J. K. *et al.* Understanding mechanisms underlying human gene ex-⁵⁰⁰ pression variation with RNA sequencing. *Nature* **464**, 768–772 (2010).
- ⁵⁰¹ [20] Kerimov, N. *et al.* A compendium of uniformly processed human gene expres-⁵⁰² sion and splicing quantitative trait loci. *Nat Genet* **53**, 1290–1299 (2021).
- [21] Rodriques, S. G. *et al.* Slide-seq: A scalable technology for measuring
 genome-wide expression at high spatial resolution. *Science* 363, 1463–1467
 (2019).
- ⁵⁰⁶ [22] Stickels, R. R. *et al.* Highly sensitive spatial transcriptomics at near-cellular ⁵⁰⁷ resolution with Slide-seqV2. *Nature Biotechnology* 1–7 (2020).
- ⁵⁰⁸ [23] Cable, D. M. *et al.* Robust decomposition of cell type mixtures in spatial tran-⁵⁰⁹ scriptomics. *Nat Biotechnol* 1–10 (2021).
- ⁵¹⁰ [24] Knowles, D. A. *et al.* Allele-specific expression reveals interactions between ⁵¹¹ genetic variation and environment. *Nat Methods* **14**, 699–702 (2017).
- [25] Fan, J. *et al.* ASEP: Gene-based detection of allele-specific expression across
 individuals in a population by RNA sequencing. *PLOS Genetics* 16, e1008786
 (2020).
- ⁵¹⁵ [26] Santoni, F. A. *et al.* Detection of Imprinted Genes by Single-Cell Allele-Specific ⁵¹⁶ Gene Expression. *Am J Hum Genet* **100**, 444–453 (2017).

- [27] Choi, K., Raghupathy, N. & Churchill, G. A. A Bayesian mixture model for the analysis of allelic expression in single cells. *Nature Communications* **10**, 5188 (2019).
- ⁵²⁰ [28] Zitovsky, J. P. & Love, M. I. Fast effect size shrinkage software for beta-binomial ⁵²¹ models of allelic imbalance. *F1000Research* (2020).
- ⁵²² [29] Kim, J. K. & Marioni, J. C. Inferring the kinetics of stochastic gene expression ⁵²³ from single-cell RNA-sequencing data. *Genome Biology* **14**, R7 (2013).
- ⁵²⁴ [30] Jiang, Y., Zhang, N. R. & Li, M. SCALE: modeling allele-specific gene expres-⁵²⁵ sion by single-cell RNA sequencing. *Genome Biology* **18**, 74 (2017).
- [31] Larsson, A. J. M. *et al.* Transcriptional bursts explain autosomal random monoallelic expression and affect allelic imbalance. *PLOS Computational Biology* **17**, e1008772 (2021).
- ⁵²⁹ [32] McCullagh, P. & Nelder, J. *Generalized Linear Models, 2nd Edition* (Chapman and Hall, 1989).
- [33] Agresti, A. Categorical Data Analysis, 3rd Edition (Wiley, 2012).
- ⁵³² [34] Reinsch, C. H. Smoothing by spline functions. *Numer. Math.* **10**, 177–183 (1967).
- ⁵³⁴ [35] Wahba, G. Smoothing noisy data with spline functions. *Numer. Math.* **24**, 383–393 (1975).
- [36] Duchon, J. Splines minimizing rotation-invariant semi-norms in Sobolev
 spaces. In Schempp, W. & Zeller, K. (eds.) *Constructive Theory of Functions of Several Variables*, Lecture Notes in Mathematics, 85–100 (Springer, Berlin,
 Heidelberg, 1977).
- ⁵⁴⁰ [37] Craven, P. & Wahba, G. Smoothing noisy data with spline functions. *Numer. Math.* **31**, 377–403 (1978).
- [38] Hastie, T. & Tibshirani, R. Generalized Additive Models. *Statistical Science* 1, 297 310 (1986).
- ⁵⁴⁴ [39] Chambers, J. & Hastie, T. *Statistical Models in S* (Wadsworth & Brooks/Cole, ⁵⁴⁵ 1992).
- [40] Wood, S. Fast stable restricted maximum likelihood and marginal likelihood
 estimation of semiparametric generalized linear models. *Journal of the Royal Statistical Society* **73**, 3–36 (2011).
- ⁵⁴⁹ [41] Wood, S. N. Thin plate regression splines. *Journal of the Royal Statistical* ⁵⁵⁰ *Society: Series B (Statistical Methodology)* **65**, 95–114 (2003).
- ⁵⁵¹ [42] Saunders, A. *et al.* Molecular Diversity and Specializations among the Cells of ⁵⁵² the Adult Mouse Brain. *Cell* **174**, 1015–1030.e16 (2018). Publisher: Elsevier.
- ⁵⁵³ [43] Yao, Z. *et al.* A taxonomy of transcriptomic cell types across the isocortex and ⁵⁵⁴ hippocampal formation. *Cell* **184**, 3222–3241.e26 (2021).

- ⁵⁵⁵ [44] Sinnamon, J. R. *et al.* The accessible chromatin landscape of the murine ⁵⁵⁶ hippocampus at single-cell resolution. *Genome Res.* **29**, 857–869 (2019).
- ⁵⁵⁷ [45] Moore, J. E. *et al.* Expanded encyclopaedias of DNA elements in the human and mouse genomes. *Nature* **583**, 699–710 (2020).
- ⁵⁵⁹ [46] GTEx Consortium. The GTEx Consortium atlas of genetic regulatory effects across human tissues. *Science* **369**, 1318–1330 (2020).
- [47] Song, Y.-H., Yoon, J. & Lee, S.-H. The role of neuropeptide somatostatin in the
 brain and its application in treating neurological disorders. *Exp Mol Med* 53 (2021).
- ⁵⁶⁴ [48] Gendrel, A.-V. *et al.* Developmental dynamics and disease potential of random ⁵⁶⁵ monoallelic gene expression. *Dev Cell* **28**, 366–380 (2014).
- ⁵⁶⁶ [49] Xu, J. *et al.* Landscape of monoallelic DNA accessibility in mouse embryonic ⁵⁶⁷ stem cells and neural progenitor cells. *Nat Genet* **49**, 377–386 (2017).
- ⁵⁶⁸ [50] Raghupathy, N. *et al.* Hierarchical analysis of RNA-seq reads improves the ⁵⁶⁹ accuracy of allele-specific expression. *Bioinformatics* **34**, 2177–2184 (2018).
- ⁵⁷⁰ [51] Langmead, B. & Salzberg, S. L. Fast gapped-read alignment with Bowtie 2. ⁵⁷¹ *Nat Methods* **9**, 357–359 (2012).
- ⁵⁷² [52] Li, H. *et al.* The Sequence Alignment/Map format and SAMtools. *Bioinformatics* ⁵⁷³ **25**, 2078–2079 (2009).
- [53] Benjamini, Y. & Hochberg, Y. Controlling the False Discovery Rate: A Practical and Powerful Approach to Multiple Testing. *Journal of the Royal Statistical Society: Series B (Methodological)* 57, 289–300 (1995).
- ⁵⁷⁷ [54] Lesnoff, M., Lancelot & R. *aod: Analysis of Overdispersed Data* (2012). R ⁵⁷⁸ package version 1.3.1.
- ⁵⁷⁹ [55] Schep, A. *motifmatchr: Fast Motif Matching in R* (2021). R package version ⁵⁸⁰ 1.14.0.
- ⁵⁸¹ [56] Fornes, O. *et al.* JASPAR 2020: update of the open-access database of tran-⁵⁸² scription factor binding profiles. *Nucleic Acids Research* **48**, D87–D92 (2020).
- ⁵⁸³ [57] Karolchik, D. *et al.* The UCSC Table Browser data retrieval tool. *Nucleic Acids* ⁵⁸⁴ *Res* **32**, D493–496 (2004).
- ⁵⁸⁵ [58] Robinson, J. T. *et al.* Integrative genomics viewer. *Nat Biotechnol* **29**, 24–26 (2011).
- ⁵⁸⁷ [59] Keane, T. M. *et al.* Mouse genomic variation and its effect on phenotypes and ⁵⁸⁸ gene regulation. *Nature* **477**, 289–294 (2011).
- ⁵⁸⁹ [60] Yalcin, B. *et al.* Sequence-based characterization of structural variation in the ⁵⁹⁰ mouse genome. *Nature* **477**, 326–329 (2011).

Supplementary information



Figure S1: Spatial transcriptomic simulation results. (a) Power as function of number of pixels and number of UMI per pixel. x-axis is number of UMIs per pixel. Numbers in the gray panels indicate the number of total pixels. Curves are colored by the amount of overdispersion (ϕ) in the true model. **(b)** False positive rate as a function of number of pixels and number of UMI per pixel. **(c)** Expected *p*-values generated under a Uniform(0,1) distribution vs. observed *p*-values computed by spASE for the null case of no spatial ASE.



Figure S2: Simulation results for beta-binomial coverage probabilities as compared to the binomial and quasibinomial models. Data was generated from a beta-binomial model where each cell or bead had a total UMI count drawn from the total UMI count distribution from one of three settings, Smart-seq3 (a,d,g), Slide-seq lowly expressed gene (b,e,h), or Slide-seq highly expressed gene (c,f,i). We also tested a range of values for overdispersion (phi): 0 (a,b,c), 0.1 (d,e,f) and 0.8 (g,h,i).



Figure S3: Total coverage distribution scenarios used for confidence interval coverage simulations shown in Figure S1. These were taken from genes to represent different sampling distributions for n_{gi} , the total number of UMI per gene per cell (or pixel).



Figure S4: Histogram of total pixels that each gene had non-zero UMI counts for. The filtering threshold of 100 pixels per gene is shown with the dashed black line.



Figure S5: 2D *z*-score plots for a sample of 16 highly expressed X-chromosome genes. Red color indicates bias towards the maternal (CAST) allele; blue indicates bias towards the paternal (129) allele.



Figure S6: Within-astrocyte ASE for *Tspan7*. (a) Raw data for astrocyte singlets plotted using 2D coordinates for each pixel. The size of the point indicates the total UMI count for the gene *Tspan7* at that pixel. The color indicates the fraction of total UMIs that were from the maternal (CAST) allele. (b) Smoothed 2D maternal allele probability function (fitted p), estimated from the raw data shown in a using 5 degrees of freedom. (c) Overlay of data from a on the smoothed surface in b. (d) 2D *z*-score plot generated for the smoothed surface shown in b.



Figure S7: Within-oligodendrocyte ASE for *Plp1*. (a) Raw data plotted using 2D coordinates for each pixel. The size of the point indicates the total UMI count at that pixel. The color indicates the fraction of total UMIs that were from the maternal (CAST) allele. (b) Smoothed 2D maternal allele probability function (fitted p), estimated from the raw data shown in a. (c) Overlay of data from a on the smoothed surface in b. (d) 2D z-score plot generated for the smoothed surface shown in b.



Figure S8: X-chromosome genes detected with a paternal bias. (a) IGV view of the *Tceal3* locus (coordinates are mm10). (b) Raw data for *Tceal3*, which was detected as having a significant spatial pattern ($q \le 0.01$). (c-d) Same as b for *Morf4l2* and *Tceal6*, which did not have a significant spatial pattern, but had paternal bias. (e-g) Smoothed maternal probability functions for *Tceal3*, *Morf4l2*, and *Tceal6*, respectively.

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Figure S9: Confidence interval visualization for genes displayed in Figure 3. (a-d) Estimated maternal probability functions for *Tspan7, Plp1, Xist,* and *Tceal3* as shown in Figure 3. (e-h) Confidence interval visualizations in cross-sections along the x = 3 and y = 3 lines for each gene shown in a-d.



Figure S10: Interneuron ASE for *Sst*. (a) Raw data for each non-zero measurement pixel; color indicates fraction of total UMI that were maternal for each pixel, size of point indicates total UMI for that pixel. (b) Smoothed maternal allele probability surface. (c) Raw data for only interneuron singlets. Average expression for each boxed region is shown. (d) Confidence intervals from cross-hair slices in b.



Figure S11: *Ptgds* **locus with annotations. (a)** IGV view (mm10) of the *Ptgds* locus showing the upstream *Gm35287* locus. Dark blue indicates Refseq gene annotation, cyan indicates peaks called from sci-ATAC-seq data from the mouse hippocampus, yellow denotes cis-regulatory elements (cCREs) from the ENCODE database (lighter yellow indicates distal-TSS enhancer-like signatures, darker yellow indicates proximal-TSS enhancer like signatures), magenta indicates predicted transcription factor binding site motifs within sci-ATAC-seq peaks, black indicates SNP locations for the CAST/EiJ and 129S1/SvmJ strains relative to mm10. (b) Zoomed-in genome browser view of the PB0044.1 motif (*Mtf1* gene) located in the peak overlapping *Gm35287*.

	-				
	Gene	Total UMI	$\chi^2 p$ -value	<i>q</i> -value	X-chr
1	Ptgds	1584	0.00e+00	0.00e+00	FALSE
2	Tspan7	4744	0.00e+00	0.00e+00	TRUE
3	Plp1	12850	3.33e-16	4.53e-13	TRUE
4	Nrip3	1963	8.39e-11	8.56e-08	FALSE
5	Sst	837	2.45e-07	2.00e-04	FALSE
6	Pcsk1n	1523	3.83e-07	2.60e-04	TRUE
7	Rgs4	679	4.02e-06	2.34e-03	FALSE
8	Atrx	1096	5.37e-06	2.74e-03	TRUE
9	Mageh1	362	9.04e-06	4.10e-03	TRUE
10	Gpm6b	1601	2.07e-05	8.44e-03	TRUE

Table S1: Genes detected as spatially significant (*q*-value ≤ 0.01) in Slide-seqV2 of the mouse hippocampus, not controlling for cell type, degrees of freedom d = 10, restricting to pixels with a confident singlet classification by RCTD.

Table S2: Genes detected as spatially significant (*q*-value ≤ 0.01) in Slide-seqV2 of the mouse hippocampus, not controlling for cell type, d = 5, restricting to pixels with a confident singlet classification by RCTD.

	Gene	Total UMI	$\chi^2 p$ -value	<i>q</i> -value	X-chr
1	Tspan7	4744	8.30e-14	3.41e-10	TRUE
2	Nrip3	1963	4.82e-11	6.60e-08	FALSE
3	Ptgds	1584	4.18e-11	6.60e-08	FALSE
4	Sst	837	4.19e-09	4.31e-06	FALSE
5	Rgs4	679	6.43e-07	5.29e-04	FALSE
6	Lypd1	241	1.40e-05	9.61e-03	FALSE

	Gene	Total UMI	$\chi^2 p$ -value	<i>q</i> -value	X-chr
1	Plp1	12850	0.00e+00	0.00e+00	TRUE
2	Ptgds	1584	0.00e+00	0.00e+00	FALSE
3	Tspan7	4744	0.00e+00	0.00e+00	TRUE
4	Nrip3	1963	4.89e-11	4.92e-08	FALSE
5	Gstm7	326	1.51e-06	1.14e-03	FALSE
6	Pcsk1n	1523	1.70e-06	1.14e-03	TRUE
7	Sst	837	3.05e-06	1.76e-03	FALSE
8	Gpm6b	1601	4.94e-06	2.49e-03	TRUE

Table S3: Genes detected as spatially significant (*q*-value ≤ 0.01) in Slide-seqV2 of the mouse hippocampus, not controlling for cell type, d = 15.

	Gene	Total UMI	$\chi^2 p$ -value	<i>q</i> -value	X-chr
1	Ptgds	1584	0.00e+00	0.00e+00	FALSE
2	Tspan7	4744	0.00e+00	0.00e+00	TRUE
3	Nrip3	1963	1.38e-10	1.80e-07	FALSE
4	Gpm6b	1601	1.08e-07	9.42e-05	TRUE
5	Pcsk1n	1523	1.21e-07	9.42e-05	TRUE
6	Gstm7	326	4.06e-06	2.63e-03	FALSE
7	Magt1	127	8.99e-06	5.00e-03	TRUE
8	H1f2	105	1.66e-05	8.10e-03	FALSE

Table S4: Genes detected as spatially significant (*q*-value ≤ 0.01) in Slide-seqV2 of the mouse hippocampus, not controlling for cell type, d = 20.