Detection of allele-specific expression in spatial transcriptomics with spASE

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Abstract

Allele-specific expression (ASE), or the preferential expression of one allele, can be observed in transcriptomics data from early development throughout the lifespan. However, the prevalence of spatial and cell type-specific ASE variation remains unclear. Spatial transcriptomics technologies permit the study of spatial ASE patterns genome-wide at near-single-cell resolution. However, the data are highly sparse, and confounding between cell type and spatial location present further statistical challenges. Here, we introduce spASE (<https://github.com/lulizou/spase>), a computational framework for detecting spatial patterns in ASE within and across cell types from spatial transcriptomics data. To tackle the challenge presented by the low signal to noise ratio due to the sparsity of the data, we implement a spatial smoothing approach that greatly improves statistical power. We generated Slide-seqV2 data from the mouse hippocampus and detected ASE in X-chromosome genes, both within and across cell type, validating our ability to recover known ASE patterns. We demonstrate that our method can also identify cell type-specific effects, which we find can explain the majority of the spatial signal for autosomal genes. The findings facilitated by our method provide new insight into the uncharacterized landscape of spatial and cell type-specific ASE in the mouse hippocampus.

Introduction

In diploid organisms, allele-specific expression (ASE) refers to the imbalanced expression of the two parental alleles for a given gene. ASE has been well-studied in the context of epigenetic phenomena such as genomic imprinting and X-chromosome inactivation (XCI) [1, 2, 3], where expression from one allele is silenced. Spatial patterns of ASE have long been observed as a consequence of XCI in female organisms, where the random silencing of either the maternal or paternal X-chromosome in early development is passed to daughter cells, resulting in visible clusters of ASE [4, 5, 6]. By contrast, although studies in bulk and single-cell RNA-sequencing data have revealed widespread variability in ASE throughout the autosome across tissues and cell types [7, 8, 9, 10, 11, 12, 13, 14, 15, 16, 17, 18, 19, 20], relatively little is known about the prevalence of spatial ASE therein.

31 Spatial transcriptomics technologies now provide the opportunity to study spatial
32 ASE patterns genome-wide. For example, Slide-seqV2 [21, 22] has high resolution
33 which enables near-single-cell quantification of ASE with 2D spatial information.
34 However, these data are limited by highly sparse read counts in comparison to
35 bulk or single-cell sequencing technologies, which is further exacerbated by the
36 requirement that reads align uniquely to one allele. In addition, cell type, which
37 drives the majority of variability observed in single-cell data, is highly correlated
38 with spatial location, especially in solid tissue [23]. Therefore, it is important to
39 distinguish between spatial and cell type-specific ASE, which could arise from and
40 contribute to distinct underlying biological mechanisms.

41 Several statistical and computational methods have been developed for studying
42 ASE in bulk and single-cell RNA-seq data [24, 25, 26, 27, 28, 29, 30, 31]. Some fo-
43 cus on estimating allele-specific transcriptional bursting kinetics for individual genes
44 in homogeneous populations of cells [15, 30, 31]. Here, we instead focus on the
45 problem of estimation and inference for the maternal allele probability p for a given
46 gene across 2D space, and we consider how p may vary with cell type. To model
47 p in bulk and single-cell RNA-seq, multiple methods have used a beta-binomial
48 framework, which can flexibly account for overdispersion from unknown technical
49 and biological variability [26, 27, 28]. An additional advantage of this model is that
50 it can be parameterized as a generalized linear model (GLM) [32, 33], allowing for
51 maximum likelihood estimation of p while incorporating covariates of interest such
52 as cell type.

53 The issue of estimating smooth functions from sparsely sampled data has been
54 well-studied [34, 35, 36, 37, 38], and multiple solutions have been developed and
55 implemented as computational methods [39, 40]. In the case of allele-specific spa-
56 tial transcriptomics data, although the read count measured at individual spatial
57 coordinates may be low, smoothing spline methods can increase power by lever-
58 aging information from local neighborhoods of pixels. Generalized additive models
59 are GLMs that incorporate smoothing splines into a regression framework, enabling
60 estimation of the smooth spatial function as well as hypothesis testing for spatial
61 functions deviating from a constant [38, 40].

62 Here, we present spASE, a computational framework for detecting genes with
63 significant ASE patterns in spatial transcriptomics data. We employ a hierarchical
64 beta-binomial smoothing approach based on thin plate regression splines [36, 41] to
65 estimate 2D allele probability functions and detect spatially significant genes. Given
66 the high correlation between cell type and spatial location in solid tissue, our method
67 permits control for cell type effects as well as any other potential covariates of in-
68 terest. Through simulations, we confirm the power and false positive rate control of
69 our method even in highly sparse settings such as those observed in allele-resolved
70 spatial transcriptomics. Additionally, we generate allele-specific Slide-seqV2 data
71 from the hippocampus of an F1 hybrid mouse and find that we are able to recover
72 known patterns of ASE due to XCI in highly-expressed X-chromosome genes, both
73 within and across cell types. We further show that our method can detect cell type-
74 specific ASE, which we find can explain most of the spatial signal observed in au-
75 toosomal genes such as *Ptgds*. Overall, we report new insights into the uncharac-
76 terized landscape of spatial and cell type-specific ASE in the mouse hippocampus,
77 thus demonstrating the utility of spASE for detecting known and novel patterns of
78 ASE in spatial transcriptomics.

Results

A beta-binomial framework for modeling allele-specific spatial transcriptomics

A statistical challenge for allele-specific spatial transcriptomics is that spatial and cell type effects can be confounded (Figure 1a,b). We therefore developed a computational framework that can account for both these sources of variability. Specifically, we developed a beta-binomial GLM that provides a flexible approach to estimation, inference, and visualization of ASE in spatial transcriptomics. We denoted the counts from the maternal allele for gene g and pixel i with Y_{gi} and assumed it followed the distribution:

$$Y_{gi} \mid \lambda_{gi} \sim \text{Binomial}(n_{gi}, \lambda_{gi}),$$

with n_{gi} the observed total UMI count for gene g and pixel i , summing both alleles, and λ_{gi} the probability that a transcript from gene g is from the maternal allele. We assumed that λ_{gi} follows a beta distribution with mean p_{gi} and variance $\phi_g p_{gi}(1-p_{gi})$. Here, p_{gi} is the mean maternal allele probability and ϕ_g is a gene-specific overdispersion parameter ranging from 0 to 1 that accounts for biological and technical variability not explained by binomial sampling.

To account for spatial and cell type effects, we created a logit-linear GLM,

$$\text{logit}(p_{gi}) = \beta_{g0} + f(x_i, y_i) + \sum_{k=1}^K \beta_{gk} X_{ik}, \quad (1)$$

with x_i and y_i the spatial location of pixel i , $f(x_i, y_i)$ a smooth function of location, the X_{ik} 's indicator functions equal to 1 if pixel i is from cell type k , and the β_{gk} parameters representing gene-specific cell type effects. Note that the β_{gk} can be interpreted as the change in log-odds, compared to the reference cell type, of a maternal allele transcript in gene g and cell type k .

The spatial effect function was modeled as a thin plate spline [36] defined by

$$f(x_i, y_i) = \sum_j^J \theta_{gj} B_j(x_i, y_i),$$

with $B_j(x, y)$ the smooth basis function for the spline and θ_{gj} the gene-specific parameters that define gene-specific spatial effects. With this definition of f , all the terms in (1) are linear and define a GLM. We can therefore obtain maximum likelihood estimates (MLEs), standard errors, and confidence intervals for all parameters using GLM theory and software. Furthermore, we can test for spatial effects by performing a likelihood ratio test comparing the model with f to a model without space (see Methods for details).

In addition to fitting spatial ASE across all cell types, we can also fit a cell type-specific version of model (1) as

$$\text{logit}(p_{gik}) = \gamma_{g0k} + f_k(x_i, y_i),$$

Figure 1: Schematic of detecting allele-specific expression in spatial transcriptomics using spASE. (a) Input is allele- and cell type-resolved spatial transcriptomics with UMIs. Each shape represents a different cell type, and the color indicates the fraction of observed UMIs that were from the maternal allele. (b) Point estimates and confidence intervals for the estimated maternal allele probability (estimated p maternal) for each cell type. (c) Visualization of the estimated maternal probability function, not controlling for cell type. (d) Visualization of confidence intervals (gray shaded region) around the MLE in a 1D cross-section. The solid line indicates the estimated maternal probability along the black dashed line from c. Light gray dashed line indicates the null of $p = 0.5$. (e) 2D z-score plot visualizing region-level significance of the estimated function from c. (f) Estimated cell type-specific function for the circle cell type from a.

for all pixels i belonging to cell type k , where p and f have been modified to depend on cell type. Note that only certain genes within certain cell types provide enough counts and therefore power to fit our cell type-specific spatial model.

After estimating parameters by maximum likelihood for each gene, we can visualize the smooth maternal probability function across 2D coordinates (Figure 1c), and for any given spatial cross-section, we can additionally visualize confidence intervals (Figure 1d). Region-specific significance can also be assessed using 2D z-score maps (Figure 1e). Running spASE on individual cell types generates spatial ASE maps for cell type-specific estimation (Figure 1f). spASE uses the likelihood ratio to rank genes according to spatial effects variability.

To evaluate the performance of our method, we generated simulated spatial transcriptomics data under a wide variety of sparsity and overdispersion conditions (Supplementary Figure S1, Methods). We calculated the power and false positive rate, and we computed p-values to detect significant spatial ASE. We observed that power decreased as overdispersion increased (Supplementary Figure S1a); however, we found that power is at least 70% even for genes with high overdispersion ($\phi = 0.8$) and as few as 50 pixels with low UMI coverage (e.g. less than 10 UMIs per pixel). With at least 100 pixels for a given gene, the power across all scenarios was at least 85%, even with as low as 1 UMI per pixel. Using a p-value threshold of $p \leq 0.01$, we found that the false positive rate approached the nominal rate of 0.01 as the number of pixels increased, in concordance with the expected asymptotic guarantees of our model (Supplementary Figure S1b,c). We also evaluated confidence interval coverage as a function of sample size and number of UMIs per pixel (Supplementary Figures S2, S3), and we found that the beta-binomial model maintained near-95% coverage across all scenarios.

spASE identifies spatially-significant ASE genes and smooths over sparse allele-specific spatial transcriptomics signal

To test spASE on allele-specific spatial transcriptomics data, we generated Slide-seqV2 data of an F1 hybrid CAST/EiJ x 129S1/SvImJ (CAST x 129) mouse hippocampus and surrounding region (see Methods). We aligned 150bp reads to a pooled CASTx129 transcriptome and only considered reads that uniquely aligned to one allele. We used RCTD [23] to call cell types using a single-cell RNA-sequencing reference of the mouse hippocampus [42], and we filtered to pixels with a high likelihood of sourcing UMIs from a single cell type (Figure 2a). Based on results from our simulations (Supplementary Figure S1), we filtered genes with non-zero UMI counts on at least 100 pixels. Using these filtering criteria resulted in 4,140 genes for downstream analysis, which were expressed on a median of 210 pixels (IQR: 140-384 pixels) with a median number of UMIs per pixel of 1.06 (IQR: 1.04-1.1 UMIs/pixel) (Supplementary Figure S4).

We then fit our model with and without the cell type covariates X_{ik} in the model (see Methods). We found that, compared to autosomal genes, a higher proportion of X-chromosome genes had likelihood-ratio-test significance (Figure 2b), in concordance with the expected patterns of XCI in the X-chromosome. The p-value distribution for autosomal genes was closer to uniform distribution, indicative of less-frequent spatial ASE effects. After controlling for cell type, the autosomal distribution remained similar, while the distribution for the X-chromosome had a lesser skew,

Figure 2: spASE identifies spatially significant ASE genes and smooths over sparse ASE spatial transcriptomics signal. (a) Map of cell types identified by RCTD in the Slide-seqV2 data generated in this study. Each point represents a pixel classified as a singlet of that cell type. (b) Distributions of p -values calculated by spASE for autosomal (A) and X-chromosome (X) genes in the real Slide-seqV2 data, not controlling for cell type ("no cell type") and controlling for cell type ("cell type"). (c) Raw data for *Hpca*, showing higher coverage in the hippocampal formation and sparse coverage in the adjacent regions. (d) Estimated 2D maternal probability function for *Hpca*, with crosshairs indicating the $x = 3$ and $y = 3$ lines, along which point estimates and confidence intervals are plotted in (e) and (f), respectively.

consistent with some the genes appearing to have spatial effects due to confounding with cell type effects.

Using a false discovery rate (FDR) threshold of $q \leq 0.01$, we found ten genes with a spatially significant pattern, of which six were from the X-chromosome (Supplementary Table S1). However, after controlling for cell type, only three genes, two (*Tspan7* and *Plp1*) on the X chromosome and one (*Sst*) autosome, were significant. Other genes, including *Nrip3* and *Ptgds*, were no longer significant after controlling for cell type, indicating that cell type differences were the main driver of spatial ASE for these genes.

spASE accounts for biological and technical noise to avoid detecting false positive ASE. For example, the gene *Hpca* was determined by spASE to not have significant spatial ASE (Figure 2c,d, $p\text{-value} = 0.53$). Although *Hpca* is highly expressed in the hippocampal formation, sparse expression in the adjacent regions resulted in noisier estimates and wide confidence intervals outside the hippocampus (Figure 2e,f). In general, such visualizations enabled by spASE allow for the assessment of both overall significance as well as position-specific significance across space.

spASE detects spatial patterns of XCI across and within cell type in the mouse hippocampus

Next, we used spASE to estimate the maternal allele probability function for X-chromosome genes and found that the patterns for almost all significant X-chromosome genes were similar and anti-correlated with *Xist* expression (Figure 3a-c, Supple-

Figure 3: spASE detects spatial ASE in X-chromosome genes across and within cell type in the mouse hippocampus from Slide-seqV2 data. (a) Smoothed maternal allele probability functions for X-chromosome genes highly expressed in the mouse hippocampus and detected as significant (q -value ≤ 0.01): *Tspan7* and (b) *Plp1*. Red color indicates bias towards maternal, blue towards paternal, and white indicates both maternal and paternal alleles are present. The outline of the CA1, CA3, and dentate cell type regions is depicted in the dotted gray areas for reference. (c) Same as a-b for *Xist*. (d) Same as a-c but for *Tspan7*. (e) Same as a-c but only using oligodendrocyte pixels for *Plp1*. (f) Same as a-c for *Tceal3*. (g) 2D z-score plot computed from combining all non-*Xist* X-chromosome genes. (h) 2D z-score plot for *Xist*.

mentary Figure S5), reflecting the expected mosaicism due to XCI by *Xist*. We also found that patterns of XCI were preserved within individual cell types. For example, *Tspan7*, which is relatively highly expressed in astrocytes (Figure 3d, S6), and *Plp1*, which is highly expressed in oligodendrocytes (Figure 3e, Supplementary Figure S7), were both estimated to have maternal probability functions anti-correlated with *Xist*.

One X-chromosome gene, *Tceal3*, exhibited a strong paternal skew unlike the rest of the X-chromosome (Figure 3f, Supplementary Figure S8). However, the estimated *Tceal3* maternal probability still had a similar trend to the observed XCI pattern, with a high paternal bias around the hippocampus and a near-biallelic pattern in the periphery. We investigated the *Tceal3* locus and found that another nearby gene less than 100kb away, *Morf4l2*, also exhibited a strong paternal bias. *Tceal6*, a paralog of *Tceal3*, also showed a paternal bias in a similar pattern to that of *Tceal3*; however, other genes in the *Tceal* family, such as *Tceal5*, did not show the same bias (Supplementary Figure S8).

We then constructed a consensus XCI map by combining the UMI counts of all of the X-chromosome genes excluding *Xist* and fitting the model on the merged spatial profile. We visualized the significance at a region-specific level by computing and plotting z-scores in 2D (Figure 3g, Supplementary Figure S5). We found that the region significant for paternal X-chromosome expression was located precisely around the CA1, CA3, and dentate cell-type layers of the hippocampus as well as around a cluster of interneurons, and was anti-correlated with *Xist* (Figure 3h). *Xist* had fewer spatial regions reaching significance, reflecting its lower UMI coverage and thus wider confidence intervals in most areas (Supplementary Figure S9).

spASE identifies cell type-driven spatial ASE in the autosome of the mouse hippocampus

We next investigated autosomal spatially significant genes with spatial ASE that was explained by cell type-driven ASE. Recall that several genes, including *Nrip3* and *Ptgds*, no longer possessed significant spatial ASE after controlling for cell type, indicating cell type-driven ASE. To quantify such cell type-driven ASE, we used spASE to estimate the overall maternal allele probability for each cell type, revealing several genes previously unknown to exhibit cell type-specific ASE (Figure 4a). For these genes, spASE's estimated spatial ASE patterns were primarily driven by cell type localization distributions (Figure 4b-d). For example, *Nrip3*, one of the most statistically significant autosomal genes, had a high maternal bias in CA1, CA3, and dentate cell types, driving a strong maternal signal observed in the z-score plot (Figure 4b). *Ptgds*, which was highly expressed in both oligodendrocytes and endothelial tip cells, had a strong paternal bias in oligodendrocytes but not endothelial cells (Figure 4c,e). *Sst* exhibited a maternal bias which was enhanced in an interneuron subtype localizing primarily outside of the hippocampus (Figure 4d; Supplementary Figure S10).

We examined whether this high cell type-specificity of ASE for these genes could be explained by genetic differences between the CAST (maternal) and 129 (paternal) mouse strains. Specifically, we investigated if a SNP could alter the binding affinity of a cell type-specific transcription factor at either the promoter or a putative enhancer. We analyzed single-cell ATAC-seq (scATAC) data from the mouse hippocampus [44] and searched for instances of SNPs overlapping known transcription factor binding site (TFBS) motifs in peaks within 50kb upstream and downstream of each gene (Methods). We found a peak in the promoter of *Gm35287* approximately 8kb upstream of *Ptgds* (Supplementary Figure S11), which is predicted to have a TSS-distal with enhancer-like signature for *Ptgds* [45]. Furthermore, we found that this peak has a high co-accessibility ($r^2 = 0.74$) with the *Ptgds* promoter peak for cell types in common between the scATAC data and our spatial data (Figure 4f). In particular, both peaks are preferentially open in oligodendrocytes. We found a SNP, rs8255993, overlapping a known transcription factor motif, PB0044.1, which corresponds to the gene *Mtf1* (Supplementary Figure S11), which is also highly expressed in oligodendrocytes (Figure 4g) [43]. The SNP is A in the paternal strain and C in the maternal, and this position has a strong A signal in the TFBS position weight matrix for *Mtf1* (Figure 4h). Thus, the preferential binding of *Mtf1* to the paternal allele at this distal enhancer is a likely mechanism driving the paternal bias of *Ptgds* observed in oligodendrocytes.

Figure 4: spASE identifies cell type-driven spatial ASE in the autosome. (a) MLEs and associated confidence intervals for the maternal probability p for three of the top autosomal gene hits (q -value ≤ 0.01), *Nrip3*, *Ptgs*, and *Sst*. x-axis: total UMI counts summed across all pixels. (b) 2D z-score plot for *Nrip3*. The hippocampal formation is outlined with dotted black lines. (c-d) Same as b for *Ptgs* and *Sst*, respectively. (e) Raw data for *Ptgs* for endothelial tip (triangles) and oligodendrocyte (circles) pixels. Size of the point indicates the total number of UMI present at that pixel, and color indicates the fraction of the total UMI that were from the CAST (maternal) allele. (f) Average sci-ATAC-seq peak accessibility of the *Ptgs* promoter peak and the nearby (~8kb away) peak in *Gm35287* for the cell types overlapping between the sci-ATAC-seq data set and the Slide-seq data. (g) Single-cell RNA-seq expression for *Mtf1* from the Mouse Brain Atlas [43]. Each point represents a cluster that was classified as one cell type. (h) Position weight matrix for PB0044.1 (*Mtf1*) with 129 (paternal) and CAST (maternal) reference sequences shown on the top and bottom, respectively. Black arrow points to the SNP position of interest.

239 Discussion

240 Allele-resolved spatial transcriptomics suffers from high sparsity in comparison to
241 bulk and single-cell sequencing, and confounding between cell type and spatial
242 location present further statistical challenges. Here, we describe a statistical ap-
243 proach and software (spASE) which allows for estimating and visualizing 2D allele
244 probabilities for sparsely expressed genes, as well as for testing spatial significance
245 while controlling for user-specified covariates such as cell type. Through simula-
246 tions, we demonstrate that our method maintains high power to detect ASE even
247 with as few as 100 pixels and as low as 1 UMI per pixel for a given gene. We gen-
248 erated Slide-seqV2 data from an F1 female CASTx129 mouse hippocampus and
249 show that our method recovers known patterns of XCI both within and across cell
250 type (Figure 3). We further show that our method can identify cell type-specific ASE,
251 which if not accounted for can be confused with spatial signal (Figure 4).

252 The primary *in situ* validation of our method was in the X-chromosome, where we
253 found the same pattern of XCI both within and across cell types for multiple genes.
254 XCI is thought to occur early in embryonic development in female organisms, be-
255 fore cell type differentiation [4], and the maternal and paternal chromosomes are
256 thought to be equally likely to be inactivated. Thus, the pattern we observed in our
257 data likely reflects randomly determined XCI in the early mouse embryo that propa-
258 gated through to the adult hippocampus. This phenomenon can potentially explain
259 why the X-chromosome p-value distribution was slightly but not fully affected by
260 controlling for cell type (Figure 2b), as some nearby cell types may be derived from
261 the same X-inactivated progenitor cell. Notably, *Tceal3* exhibited a strong paternal
262 bias, but still had a spatial pattern that was similar to the general XCI pattern we
263 observed in other X-chromosome genes (Supplementary Figures S5, S8). Another
264 nearby gene, *Morf4l2*, also exhibited a paternal bias. Thus, the pattern we ob-
265 served in *Tceal3* may be the combined result of XCI and another form of epigenetic
266 imprinting.

267 One limitation of our spatial ASE analysis is that low UMI coverage limits the
268 spatial resolution of ASE estimates. For example, within the XCI analysis, *Xist* was
269 lowly expressed (Figure 3h); however, for genes with higher coverage, such as *Plp1*,
270 it was possible to resolve the spatial ASE function further by increasing the degrees
271 of freedom used to construct the 2D basis functions. Due to our limited spatial
272 resolution, although we detected spatial patterns of XCI between the hippocampal
273 formation (paternal bias) and surrounding areas (maternal bias), it is likely that in-
274 creased statistical power would be achieved and higher-resolution spatial patterns
275 would be uncovered given a higher-coverage dataset.

276 Similarly, although we found multiple instances of differential ASE across cell
277 types as previously observed [20, 46], our analysis did not detect any spatial ASE
278 in autosomal genes not explainable by cell type. We note that the statistical power
279 was lower for the detection of spatial effects compared to the detection of cell-type
280 differences. It is possible that autosomal spatial ASE effects might be detected
281 given increased coverage and sample size.

282 We found that *Sst* exhibited a strong maternal bias for interneurons, particularly
283 for a subtype located outside of the hippocampal formation with high *Sst* expression
284 (Supplementary Figure S10). *Sst* is a well-known neuropeptide expressed through-
285 out the brain which has been studied in the context of various neurological diseases

[47]. However, we were not able to detect a likely cell type-specific transcription factor with a nearby binding site that was affected by strain-specific genetic variation as we did for *Ptgds*, although it is possible that the bias may only affect this subtype which is not represented in the scATAC-seq data set we used. Also, note that *Sst* exhibited low levels of expression in other cell types, which limited statistical power. Overall, these findings demonstrate that our method is broadly applicable for ASE discovery in spatial transcriptomics. Our rigorous computational approach will inform future analyses on the variability and biological mechanisms driving spatial and cell type-specific ASE.

Methods

Slide-seqV2 of CAST/EiJ x 129S1/SvImJ F1 mice

We obtained a female CAST/EiJ x 129S1/SvImJ (CASTx129) mouse from Jackson laboratories. The CASTx129 cross contains ~23 million SNPs, or approximately 1 SNP for every ~110 bp [48, 49]. This SNP density is approximately tenfold the SNP density in human cells and thus provides high resolution to interrogate ASE. Slide-seqV2 was performed as described previously [21, 22] on two adjacent, 10um-thick coronal slices of the hippocampus.

Alignment of Slide-seqV2 data

We generated a pooled CASTx129 transcriptome using the command `create-hybrid` from the EMASE [50] software on the CAST and 129 transcript fasta files downloaded from <ftp://churchill-lab.jax.org/software/g2gtools/mouse/R84-REL1505/>. We then aligned 150bp reads to this pooled transcriptome with bowtie2 [51] using the parameters `-k 4 -p 16 --very-sensitive`. We used a custom script (<https://github.com/lulizou/spASE/blob/master/scripts/processBowtie2.py>) for processing the aligned BAM file [52] to create a gene UMI count matrix only from reads that uniquely aligned to one gene and one allele. We restricted attention to alignments with 3 or fewer mismatches and only considered alignments that had the fewest number of mismatches for that read. We overlaid data from the two slices by rotating and shifting the slices to overlap according to the location of the hippocampal formation.

Beta-binomial model for allele-specific expression in spatial transcriptomics

Let n_{gi} denote the observed total counts of gene g at cell or pixel i and Y_{gi} denote the observed maternal allele UMI counts for gene g at cell or pixel i . Let λ_{gi} denote the unknown mean probability of observing a maternal allele for each transcript of gene g at pixel i . We assume $Y_{gi}|\lambda_{gi} \sim \text{Binomial}(n_{gi}, \lambda_{gi})$, where n_{gi} is observed and λ_{gi} is a random variable, independently distributed (conditional on ϕ_g and p_{gi} , defined below) for each gene g and pixel i . We further assume that λ_{gi} follows a beta distribution with mean p_{gi} and variance $\phi_g p_{gi}(1 - p_{gi})$. The likelihood of this model for a single gene g can be written as

$$\mathcal{L}(p_{gi}, \phi_g; n_{g\cdot}, y_{g\cdot}) = \prod_i \binom{n_{gi}}{y_{gi}} \frac{B(p_{gi}(1 - \phi_g)/\phi_g + y_{gi}, (1 - p_{gi})(1 - \phi_g)/\phi_g) + n_{gi} - y_{gi}}{B(p_{gi}(1 - \phi_g)/\phi_g, (1 - p_{gi})(1 - \phi_g)/\phi_g)} \quad (2)$$

where B denotes the beta function. We used maximum likelihood estimation to obtain the estimates $\hat{p}_{gi}$ and $\hat{\phi}_g$ for each gene and the associated standard errors, determined from the Fisher information. When using this model for single-cell data to estimate ASE for a single gene, p_{gi} is assumed to be the same for all cells i ; when estimating p_{gi} for spatial transcriptomics, p_{gi} is assumed to be dependent on pixel i as described below. Assuming asymptotic normality, we used these point estimates and standard errors to construct 95% confidence intervals for $\hat{p}_{gi}$. Finally, we used the Benjamini-Hochberg [53] procedure to produce q -values to control the false discovery rate.

In the spatial setting, for each gene g , we model p_{gi} as a smooth spatially-varying function. Specifically, we used thin plate regression splines [36, 40, 41] to estimate smooth maternal allele probability surfaces for each gene. Thin plate regression splines allow estimation of a smooth function of 2D coordinates. The number of basis functions d determines the smoothness, with lower values of d corresponding to smoother functions. Choice of d depends on the sparsity of the data, since lower values of d reduce the variance, but at the risk of introducing bias. In the analysis presented here, we kept d constant across genes to ensure comparability. For our sparse Slide-seqV2 data, we found that $d = 10$ provided enough complexity to model allelic patterns in the hippocampus sample examined here while also maintaining power to detect significant differences, and we used $d = 15$ when plotting estimated maternal allele probability functions using all pixels. We also demonstrate reproducibility of results (i.e. genes detected as having a significant spatial pattern) across a range of values for d (Supplementary Tables S1-S4). In practice, we recommend visualization of the estimated probability function and confidence intervals (e.g. Figure 2c-f) to guide selection of d .

In the model for cell type-specific spatial ASE detection, the term θ_{gk} corresponds to the effect size for cell type k for gene g . If no cell type annotations are available, or if a cell type effect does not exist (see likelihood ratio test below), then θ_{gk} can be assumed to be the same for all k .

To test whether there was a significant spatial pattern beyond cell type, we assumed a baseline model with only cell types as covariates and performed a likelihood ratio test comparing model (1) to the baseline model, i.e. for each gene g , we compute

$$\Lambda_g = -2(\ell(\mathbf{p}_g, \phi_g; \mathbf{n}_g, \mathbf{Y}_g, \mathbf{X}) - \ell(\mathbf{p}_g, \phi_g; \mathbf{n}_g, \mathbf{Y}_g, \mathbf{X}, \mathbf{x}, \mathbf{y})) \quad (3)$$

where ℓ is the log-likelihood computed from (1), $\mathbf{p}_g$ is the vector of maternal probabilities at each pixel i for gene g , $\mathbf{n}_g$ is the total number of UMIs at each pixel i for gene g , $\mathbf{Y}_g$ is the total number of maternal-derived UMIs at each pixel i for gene g , $\mathbf{X}$ is the $i \times k$ matrix of indicators of each cell type k at each pixel i , and $\mathbf{x}, \mathbf{y}$ are the vectors of 2D spatial coordinates for each pixel i . This test statistic has an asymptotic χ^2_1 distribution, which we use to compute p-values for each gene.

For visualization, we plotted both the estimated smooth function for the MLE $\hat{p}_{gi}$ (Figure 3a-c) as well as 2D z -score plots (Figure 3d-e, Figure 4e). Z -score plots were calculated on an evenly spaced grid of points over the sample by taking the

point estimate $\text{logit}(\hat{p}_{gi})$ at each location and its associated standard error s_{gi} and computing $z_{gi} = \text{logit}(\hat{p}_{gi})/s_{gi}$.

Computational implementation

spASE is implemented as an R package (<https://github.com/lulizou/spase>). We generated thin plate regression splines using the R package `mgcv` [40]. Specifically, we used the `smoothCon` function to construct spline basis functions. As basis functions can depend on the scale of the spatial covariates, we used normalized coordinates and also normalized the basis functions after construction by subtracting the mean and dividing by the standard deviation. We used the implementation of the beta-binomial likelihood from the R package `aod` [54].

We ran spASE in multiple modes: 1) not controlling for cell type, 2) controlling for cell type by allowing each cell type to have a different intercept term, and 3) allowing for each cell type to have a different spatial pattern. For 1), we experimented with using all pixels or only pixels confidently called as single cells (singlets). We found that using all pixels allowed us to increase our power and resolution for estimation of p_{gi} and ϕ_i , and for evaluating significance of spatial fits; thus, for visualization in our figures, we use all pixels, unless the figure is specifically denoted to be a single cell type. For results directly comparing to cell type models from 2) and 3), e.g. significant genes detected when controlling and not controlling for cell type, we use only pixels confidently classified in both cases to ensure comparable sample sizes of pixels.

Simulation details

We simulated beta-binomial count data to evaluate the power, false positive rate, and p -values calculated using spASE. For each simulated gene, to construct random spatial ASE patterns, we used a random linear combination of basis functions calculated from the pixel locations of the Slide-seqV2 hippocampus data set using degrees of freedom $d = 15$. We first sampled a random number of total pixels $N \in \{50, 100, 250, 500\}$ and used a fixed number of UMIs per pixel. The range of average UMIs per pixel for a single gene reached up to 14 in the real Slide-seqV2 hippocampus data set; for testing purposes, we used values ranging from 1 to 50. We drew the coefficients θ_g of a random linear combination of basis functions from a standard normal distribution. Then, we chose a fixed overdispersion parameter $\phi \in \{0.1, 0.3, 0.5, 0.8\}$, sampled the true binomial probabilities λ_{gi} for each pixel location i from the beta distribution, and simulated counts from the binomial model.

To evaluate the asymptotic confidence interval coverage properties, we simulated ground truth data generated under the beta-binomial model without spatial covariates under a range of values for overdispersion, total coverage, and number of cells or pixels (Supplementary Figure S2). Specifically, we tested total UMI conditions based on three dataset-specific distributions for total UMI for a single gene across cells or pixels: Smart-seq3, which had generally high total UMI for each cell [31]; Slide-seq high, which had low counts with a high right skew; and Slide-seq low, which had mostly low counts (less than 5 per pixel, Supplementary Figure S3). We evaluated a range of values for the total number of cells or pixels the gene was captured on. We also tested a range of overdispersion values, namely

411 $\phi \in \{0, 0.1, \dots, 1\}$, and noticed the trend was the same as ϕ increased, thus we only
 412 show results here for $\phi = 0, 0.1, 0.8$. We set the maternal probability at $p = 0.5$ and
 413 simulated 5,000 iterations for each condition.

414 Prediction of cell types in spatial transcriptomics

415 We ran RCTD [23] to predict cell types in our Slide-seq data using a previously
 416 published single-cell RNA-seq reference for mouse hippocampus [42]. RCTD is a
 417 supervised approach that learns cell type profiles from a single cell reference and
 418 predicts cell type labels for spatial transcriptomics pixels using a Poisson log-normal
 419 model that accounts for platform effects between single-cell RNA-seq and Slide-seq
 420 data. We used a threshold of a likelihood difference of 100 between the minimum
 421 score and singlet score to classify singlets. After running RCTD, we filtered to pixels
 422 that were predicted to contain a single cell-type and added these cell type labels as
 423 the covariates in model (1) to perform all cell type-specific analyses.

424 Mouse hippocampus scATAC-seq, cis-regulatory elements, and TFBS motifs

425 We analyzed previously published mouse hippocampus sci-ATAC-seq count ma-
 426 trices (GSE118987) [44]. The mice used in the study were the wild-type C57/B6
 427 strain, and the data were aligned to the mm10 genome. We extracted the called
 428 peak annotations and counts using the command `scitools split`. To quantify ac-
 429 cessibility of peaks, we computed the average count within cell types. We searched
 430 for known TFBS motifs within peaks with the command `matchMotifs` from the R
 431 package `motifmatchr` [55] using all motifs for *Mus musculus* in the JASPAR 2020
 432 database [56]. We overlaid annotations of ENCODE cis-regulatory elements (cCREs)
 433 [45] for mm10 downloaded from the UCSC Table Browser [57] at the *Ptgds* locus
 434 and visualized the annotations using IGV [58].

435 Mouse strain SNP data

436 We obtained gene-specific SNP annotations for 129S1/SvImJ and CAST/EiJ with
 437 respect to the mm10 reference using REL1505 of the Mouse Genomes Project
 438 (https://www.sanger.ac.uk/sanger/Mouse_SnpViewer/rel-1505) [59, 60]. For each
 439 gene of interest, we searched for SNPs overlapping the gene, within 50kb upstream
 440 of the gene start, and 50kb downstream of the gene end.

441 Data availability

442 Slide-seqV2 data generated in this study will be made available upon publication.

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449 Author contributions

450 RAI, FC, and LSZ conceived of the study. TZ and EM generated the Slide-seqV2
451 data. LSZ, RAI, DMC, FC, TZ, and MJA analyzed the data. LSZ, RAI, DMC, TZ,
452 and FC wrote the manuscript. All authors read and approved the final manuscript.

453 Competing interests

454 The authors declare no competing interests.

455 References

- 456 [1] Ferguson-Smith, A. C. & Surani, M. A. Imprinting and the epigenetic asymme-
457 try between parental genomes. *Science* **293**, 1086–1089 (2001).
- 458 [2] Reik, W., Dean, W. & Walter, J. Epigenetic reprogramming in mammalian
459 development. *Science* **293**, 1089–1093 (2001).
- 460 [3] Knight, J. C. Allele-specific gene expression uncovered. *Trends Genet* **20**,
461 113–116 (2004).
- 462 [4] Lyon, M. F. Gene action in the X-chromosome of the mouse (*Mus musculus*
463 L.). *Nature* **190**, 372–373 (1961).
- 464 [5] Disteche, C. M. & Berletch, J. B. X-chromosome inactivation and escape. *J*
465 *Genet* **94**, 591–599 (2015).
- 466 [6] Wu, H. *et al.* Cellular resolution maps of X-chromosome inactivation: impli-
467 cations for neural development, function, and disease. *Neuron* **81**, 103–119
468 (2014).
- 469 [7] Yan, H., Yuan, W., Velculescu, V. E., Vogelstein, B. & Kinzler, K. W. Allelic
470 variation in human gene expression. *Science* **297**, 1143 (2002).
- 471 [8] Cowles, C. R., Hirschhorn, J. N., Altshuler, D. & Lander, E. S. Detection of
472 regulatory variation in mouse genes. *Nat Genet* **32**, 432–437 (2002).
- 473 [9] Bray, N. J., Buckland, P. R., Owen, M. J. & O'Donovan, M. C. Cis-acting vari-
474 ation in the expression of a high proportion of genes in human brain. *Hum*
475 *Genet* **113**, 149–153 (2003).
- 476 [10] Gimelbrant, A., Hutchinson, J. N., Thompson, B. R. & Chess, A. Widespread
477 monoallelic expression on human autosomes. *Science* **318**, 1136–1140
478 (2007).

- 479 [11] Zwemer, L. M. *et al.* Autosomal monoallelic expression in the mouse. *Genome*
480 *Biol* **13**, R10 (2012).
- 481 [12] Chess, A. Mechanisms and consequences of widespread random monoallelic
482 expression. *Nat Rev Genet* **13**, 421–428 (2012).
- 483 [13] Deng, Q., Ramsköld, D., Reinius, B. & Sandberg, R. Single-cell RNA-seq re-
484 veals dynamic, random monoallelic gene expression in mammalian cells. *Sci-*
485 *ence* **343**, 193–196 (2014).
- 486 [14] Reinius, B. & Sandberg, R. Random monoallelic expression of autosomal
487 genes: stochastic transcription and allele-level regulation. *Nat Rev Genet* **16**,
488 653–664 (2015).
- 489 [15] Larsson, A. J. M. *et al.* Genomic encoding of transcriptional burst kinetics.
490 *Nature* **565**, 251–254 (2019).
- 491 [16] Rv, P., Sundaresh, A., Karunyaa, M., Arun, A. & Gayen, S. Autosomal Clonal
492 Monoallelic Expression: Natural or Artifactual? *Trends Genet* (2020).
- 493 [17] Lee, C., Kang, E. Y., Gandal, M. J., Eskin, E. & Geschwind, D. H. Profiling
494 allele-specific gene expression in brains from individuals with autism spectrum
495 disorder reveals preferential minor allele usage. *Nat Neurosci* **22**, 1521–1532
496 (2019).
- 497 [18] Chess, A. Monoallelic Gene Expression in Mammals. *Annu Rev Genet* **50**,
498 317–327 (2016).
- 499 [19] Pickrell, J. K. *et al.* Understanding mechanisms underlying human gene ex-
500 pression variation with RNA sequencing. *Nature* **464**, 768–772 (2010).
- 501 [20] Kerimov, N. *et al.* A compendium of uniformly processed human gene expres-
502 sion and splicing quantitative trait loci. *Nat Genet* **53**, 1290–1299 (2021).
- 503 [21] Rodriques, S. G. *et al.* Slide-seq: A scalable technology for measuring
504 genome-wide expression at high spatial resolution. *Science* **363**, 1463–1467
505 (2019).
- 506 [22] Stickels, R. R. *et al.* Highly sensitive spatial transcriptomics at near-cellular
507 resolution with Slide-seqV2. *Nature Biotechnology* 1–7 (2020).
- 508 [23] Cable, D. M. *et al.* Robust decomposition of cell type mixtures in spatial tran-
509 scriptomics. *Nat Biotechnol* 1–10 (2021).
- 510 [24] Knowles, D. A. *et al.* Allele-specific expression reveals interactions between
511 genetic variation and environment. *Nat Methods* **14**, 699–702 (2017).
- 512 [25] Fan, J. *et al.* ASEP: Gene-based detection of allele-specific expression across
513 individuals in a population by RNA sequencing. *PLOS Genetics* **16**, e1008786
514 (2020).
- 515 [26] Santoni, F. A. *et al.* Detection of Imprinted Genes by Single-Cell Allele-Specific
516 Gene Expression. *Am J Hum Genet* **100**, 444–453 (2017).

- 517 [27] Choi, K., Raghupathy, N. & Churchill, G. A. A Bayesian mixture model for the
518 analysis of allelic expression in single cells. *Nature Communications* **10**, 5188
519 (2019).
- 520 [28] Zitovsky, J. P. & Love, M. I. Fast effect size shrinkage software for beta-binomial
521 models of allelic imbalance. *F1000Research* (2020).
- 522 [29] Kim, J. K. & Marioni, J. C. Inferring the kinetics of stochastic gene expression
523 from single-cell RNA-sequencing data. *Genome Biology* **14**, R7 (2013).
- 524 [30] Jiang, Y., Zhang, N. R. & Li, M. SCALE: modeling allele-specific gene expres-
525 sion by single-cell RNA sequencing. *Genome Biology* **18**, 74 (2017).
- 526 [31] Larsson, A. J. M. *et al.* Transcriptional bursts explain autosomal random
527 monoallelic expression and affect allelic imbalance. *PLOS Computational Bi-*
528 *ology* **17**, e1008772 (2021).
- 529 [32] McCullagh, P. & Nelder, J. *Generalized Linear Models, 2nd Edition* (Chapman
530 and Hall, 1989).
- 531 [33] Agresti, A. *Categorical Data Analysis, 3rd Edition* (Wiley, 2012).
- 532 [34] Reinsch, C. H. Smoothing by spline functions. *Numer. Math.* **10**, 177–183
533 (1967).
- 534 [35] Wahba, G. Smoothing noisy data with spline functions. *Numer. Math.* **24**,
535 383–393 (1975).
- 536 [36] Duchon, J. Splines minimizing rotation-invariant semi-norms in Sobolev
537 spaces. In Schempp, W. & Zeller, K. (eds.) *Constructive Theory of Functions*
538 *of Several Variables*, Lecture Notes in Mathematics, 85–100 (Springer, Berlin,
539 Heidelberg, 1977).
- 540 [37] Craven, P. & Wahba, G. Smoothing noisy data with spline functions. *Numer.*
541 *Math.* **31**, 377–403 (1978).
- 542 [38] Hastie, T. & Tibshirani, R. Generalized Additive Models. *Statistical Science* **1**,
543 297 – 310 (1986).
- 544 [39] Chambers, J. & Hastie, T. *Statistical Models in S* (Wadsworth & Brooks/Cole,
545 1992).
- 546 [40] Wood, S. Fast stable restricted maximum likelihood and marginal likelihood
547 estimation of semiparametric generalized linear models. *Journal of the Royal*
548 *Statistical Society* **73**, 3–36 (2011).
- 549 [41] Wood, S. N. Thin plate regression splines. *Journal of the Royal Statistical*
550 *Society: Series B (Statistical Methodology)* **65**, 95–114 (2003).
- 551 [42] Saunders, A. *et al.* Molecular Diversity and Specializations among the Cells of
552 the Adult Mouse Brain. *Cell* **174**, 1015–1030.e16 (2018). Publisher: Elsevier.
- 553 [43] Yao, Z. *et al.* A taxonomy of transcriptomic cell types across the isocortex and
554 hippocampal formation. *Cell* **184**, 3222–3241.e26 (2021).

- 555 [44] Sinnamon, J. R. *et al.* The accessible chromatin landscape of the murine
556 hippocampus at single-cell resolution. *Genome Res.* **29**, 857–869 (2019).
- 557 [45] Moore, J. E. *et al.* Expanded encyclopaedias of DNA elements in the human
558 and mouse genomes. *Nature* **583**, 699–710 (2020).
- 559 [46] GTEx Consortium. The GTEx Consortium atlas of genetic regulatory effects
560 across human tissues. *Science* **369**, 1318–1330 (2020).
- 561 [47] Song, Y.-H., Yoon, J. & Lee, S.-H. The role of neuropeptide somatostatin in the
562 brain and its application in treating neurological disorders. *Exp Mol Med* **53**
563 (2021).
- 564 [48] Gendrel, A.-V. *et al.* Developmental dynamics and disease potential of random
565 monoallelic gene expression. *Dev Cell* **28**, 366–380 (2014).
- 566 [49] Xu, J. *et al.* Landscape of monoallelic DNA accessibility in mouse embryonic
567 stem cells and neural progenitor cells. *Nat Genet* **49**, 377–386 (2017).
- 568 [50] Raghupathy, N. *et al.* Hierarchical analysis of RNA-seq reads improves the
569 accuracy of allele-specific expression. *Bioinformatics* **34**, 2177–2184 (2018).
- 570 [51] Langmead, B. & Salzberg, S. L. Fast gapped-read alignment with Bowtie 2.
571 *Nat Methods* **9**, 357–359 (2012).
- 572 [52] Li, H. *et al.* The Sequence Alignment/Map format and SAMtools. *Bioinformatics*
573 **25**, 2078–2079 (2009).
- 574 [53] Benjamini, Y. & Hochberg, Y. Controlling the False Discovery Rate: A Practical
575 and Powerful Approach to Multiple Testing. *Journal of the Royal Statistical*
576 *Society: Series B (Methodological)* **57**, 289–300 (1995).
- 577 [54] Lesnoff, M., Lancelot & R. *aod: Analysis of Overdispersed Data* (2012). R
578 package version 1.3.1.
- 579 [55] Schep, A. *motifmatchr: Fast Motif Matching in R* (2021). R package version
580 1.14.0.
- 581 [56] Fornes, O. *et al.* JASPAR 2020: update of the open-access database of tran-
582 scription factor binding profiles. *Nucleic Acids Research* **48**, D87–D92 (2020).
- 583 [57] Karolchik, D. *et al.* The UCSC Table Browser data retrieval tool. *Nucleic Acids*
584 *Res* **32**, D493–496 (2004).
- 585 [58] Robinson, J. T. *et al.* Integrative genomics viewer. *Nat Biotechnol* **29**, 24–26
586 (2011).
- 587 [59] Keane, T. M. *et al.* Mouse genomic variation and its effect on phenotypes and
588 gene regulation. *Nature* **477**, 289–294 (2011).
- 589 [60] Yalcin, B. *et al.* Sequence-based characterization of structural variation in the
590 mouse genome. *Nature* **477**, 326–329 (2011).

Supplementary information

591

Figure S1: Spatial transcriptomic simulation results. (a) Power as function of number of pixels and number of UMI per pixel. x-axis is number of UMIs per pixel. Numbers in the gray panels indicate the number of total pixels. Curves are colored by the amount of overdispersion (ϕ) in the true model. (b) False positive rate as a function of number of pixels and number of UMI per pixel. (c) Expected p -values generated under a Uniform(0,1) distribution vs. observed p -values computed by spASE for the null case of no spatial ASE.

Figure S2: Simulation results for beta-binomial coverage probabilities as compared to the binomial and quasibinomial models. Data was generated from a beta-binomial model where each cell or bead had a total UMI count drawn from the total UMI count distribution from one of three settings, Smart-seq3 (a,d,g), Slide-seq lowly expressed gene (b,e,h), or Slide-seq highly expressed gene (c,f,i). We also tested a range of values for overdispersion (ϕ): 0 (a,b,c), 0.1 (d,e,f) and 0.8 (g,h,i).

Figure S3: Total coverage distribution scenarios used for confidence interval coverage simulations shown in Figure S1. These were taken from genes to represent different sampling distributions for n_{gi} , the total number of UMI per gene per cell (or pixel).

Figure S4: Histogram of total pixels that each gene had non-zero UMI counts for. The filtering threshold of 100 pixels per gene is shown with the dashed black line.

Figure S5: 2D z -score plots for a sample of 16 highly expressed X-chromosome genes. Red color indicates bias towards the maternal (CAST) allele; blue indicates bias towards the paternal (129) allele.

Figure S6: Within-astrocyte ASE for *Tspan7*. (a) Raw data for astrocyte singlets plotted using 2D coordinates for each pixel. The size of the point indicates the total UMI count for the gene *Tspan7* at that pixel. The color indicates the fraction of total UMIs that were from the maternal (CAST) allele. (b) Smoothed 2D maternal allele probability function (fitted p), estimated from the raw data shown in a using 5 degrees of freedom. (c) Overlay of data from a on the smoothed surface in b. (d) 2D z-score plot generated for the smoothed surface shown in b.

Figure S7: Within-oligodendrocyte ASE for *Plp1*. (a) Raw data plotted using 2D coordinates for each pixel. The size of the point indicates the total UMI count at that pixel. The color indicates the fraction of total UMIs that were from the maternal (CAST) allele. (b) Smoothed 2D maternal allele probability function (fitted p), estimated from the raw data shown in a. (c) Overlay of data from a on the smoothed surface in b. (d) 2D z-score plot generated for the smoothed surface shown in b.

Figure S8: X-chromosome genes detected with a paternal bias. (a) IGV view of the *Tceal3* locus (coordinates are mm10). (b) Raw data for *Tceal3*, which was detected as having a significant spatial pattern ($q \leq 0.01$). (c-d) Same as b for *Morf4l2* and *Tceal6*, which did not have a significant spatial pattern, but had paternal bias. (e-g) Smoothed maternal probability functions for *Tceal3*, *Morf4l2*, and *Tceal6*, respectively.

Figure S9: Confidence interval visualization for genes displayed in Figure 3. (a-d) Estimated maternal probability functions for *Tspan7*, *Plp1*, *Xist*, and *Tceal3* as shown in Figure 3. (e-h) Confidence interval visualizations in cross-sections along the $x = 3$ and $y = 3$ lines for each gene shown in a-d.

Figure S10: Interneuron ASE for *Sst*. (a) Raw data for each non-zero measurement pixel; color indicates fraction of total UMI that were maternal for each pixel, size of point indicates total UMI for that pixel. (b) Smoothed maternal allele probability surface. (c) Raw data for only interneuron singlets. Average expression for each boxed region is shown. (d) Confidence intervals from cross-hair slices in b.

Figure S11: *Ptgd* locus with annotations. (a) IGV view (mm10) of the *Ptgd* locus showing the upstream *Gm35287* locus. Dark blue indicates Refseq gene annotation, cyan indicates peaks called from sci-ATAC-seq data from the mouse hippocampus, yellow denotes cis-regulatory elements (cCREs) from the ENCODE database (lighter yellow indicates distal-TSS enhancer-like signatures, darker yellow indicates proximal-TSS enhancer like signatures), magenta indicates predicted transcription factor binding site motifs within sci-ATAC-seq peaks, black indicates SNP locations for the CAST/EiJ and 129S1/SvmJ strains relative to mm10. (b) Zoomed-in genome browser view of the PB0044.1 motif (*Mtf1* gene) located in the peak overlapping *Gm35287*.

Table S1: Genes detected as spatially significant (q -value ≤ 0.01) in Slide-seqV2 of the mouse hippocampus, not controlling for cell type, degrees of freedom $d = 10$, restricting to pixels with a confident singlet classification by RCTD.

	Gene	Total UMI	χ^2 p -value	q -value	X-chr
1	Ptgds	1584	0.00e+00	0.00e+00	FALSE
2	Tspan7	4744	0.00e+00	0.00e+00	TRUE
3	Plp1	12850	3.33e-16	4.53e-13	TRUE
4	Nrip3	1963	8.39e-11	8.56e-08	FALSE
5	Sst	837	2.45e-07	2.00e-04	FALSE
6	Pcsk1n	1523	3.83e-07	2.60e-04	TRUE
7	Rgs4	679	4.02e-06	2.34e-03	FALSE
8	Atrx	1096	5.37e-06	2.74e-03	TRUE
9	Mageh1	362	9.04e-06	4.10e-03	TRUE
10	Gpm6b	1601	2.07e-05	8.44e-03	TRUE

Table S2: Genes detected as spatially significant (q -value ≤ 0.01) in Slide-seqV2 of the mouse hippocampus, not controlling for cell type, $d = 5$, restricting to pixels with a confident singlet classification by RCTD.

	Gene	Total UMI	χ^2 p -value	q -value	X-chr
1	Tspan7	4744	8.30e-14	3.41e-10	TRUE
2	Nrip3	1963	4.82e-11	6.60e-08	FALSE
3	Ptgds	1584	4.18e-11	6.60e-08	FALSE
4	Sst	837	4.19e-09	4.31e-06	FALSE
5	Rgs4	679	6.43e-07	5.29e-04	FALSE
6	Lypd1	241	1.40e-05	9.61e-03	FALSE

Table S3: Genes detected as spatially significant (q -value ≤ 0.01) in Slide-seqV2 of the mouse hippocampus, not controlling for cell type, $d = 15$.

	Gene	Total UMI	χ^2 p -value	q -value	X-chr
1	Plp1	12850	0.00e+00	0.00e+00	TRUE
2	Ptgds	1584	0.00e+00	0.00e+00	FALSE
3	Tspan7	4744	0.00e+00	0.00e+00	TRUE
4	Nrip3	1963	4.89e-11	4.92e-08	FALSE
5	Gstm7	326	1.51e-06	1.14e-03	FALSE
6	Pcsk1n	1523	1.70e-06	1.14e-03	TRUE
7	Sst	837	3.05e-06	1.76e-03	FALSE
8	Gpm6b	1601	4.94e-06	2.49e-03	TRUE

Table S4: Genes detected as spatially significant ($q\text{-value} \leq 0.01$) in Slide-seqV2 of the mouse hippocampus, not controlling for cell type, $d = 20$.

	Gene	Total UMI	χ^2 $p\text{-value}$	$q\text{-value}$	X-chr
1	Ptgds	1584	0.00e+00	0.00e+00	FALSE
2	Tspan7	4744	0.00e+00	0.00e+00	TRUE
3	Nrip3	1963	1.38e-10	1.80e-07	FALSE
4	Gpm6b	1601	1.08e-07	9.42e-05	TRUE
5	Pcsk1n	1523	1.21e-07	9.42e-05	TRUE
6	Gstm7	326	4.06e-06	2.63e-03	FALSE
7	Magt1	127	8.99e-06	5.00e-03	TRUE
8	H1f2	105	1.66e-05	8.10e-03	FALSE