

1 Divergent evolution of head morphology between marine and
2 freshwater sticklebacks

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13 **Abstract**

14 Intraspecific phenotypic differentiation is of common place occurrence, but the degree to
15 which it reflects phenotypic plasticity or local adaptation remains often unclear. To be
16 considered as adaptive, the differentiation must be genetically based and exceed what could
17 be expected by neutral processes only. Using laboratory reared full-sib family data from
18 replicate nine-spined stickleback (*Pungitius pungitius*) populations, we show that freshwater
19 and marine fish display genetically based adaptive differentiation in head size and shape.
20 Utilising identity-by-descent relationships among full-sibs as estimated with the aid of
21 molecular markers, we further show that the studied traits are also highly heritable in all
22 populations indicating and that they can respond to future episodes of natural selection. The
23 head shape and size of pond fish suggests that observed adaptive differentiation has been
24 driven by selection favoring limnetic feeding strategy among the pond fish. Analyses of gill-
25 raker morphology were less conclusive: genetic differentiation was found in gill-raker length
26 (pond > marine) and number, but the degree of divergence in these traits did not exceed
27 neutral expectations. Yet, the direction of divergence in gill raker traits are suggestive of the
28 limnetic feeding mode of pond fish, aligning with the inference from the head morphology
29 analyses.

30

31 **Keywords:** Geometric morphometrics, adaptation, heritability, Q_{ST} , quantitative genetics,

32 *Pungitius pungitius*

33 **Introduction**

34 Phenotypic differentiation associated with different habitat characteristics is of commonplace
35 occurrence in both animal (Endler 1978; Foster 1999) and plant (Linhart and Grant 1996;
36 Bossdorf et al. 2005) kingdoms. Likewise, phenotypic responses to changing environmental
37 conditions have been frequently reported (Walther et al. 2002; Parmesan 2006; Sheridan &
38 Bickford 2011), but many studies struggle to disentangle phenotypic plasticity and
39 genetically based differentiation as the underlying cause for observed differentiation
40 (Gienapp et al. 2008; Merilä 2012; Merilä & Hendry 2014; Stamp & Hadfield 2020).
41 Moreover, even if evidence for genetically based differentiation is found, additional evidence
42 is required to prove that the observed differentiation is adaptive.

43 Common garden studies combined with neutrality tests (e.g., Lande 1976; Rogers 1986;
44 Spitze 1993; Merilä & Crnokrak 2001) provide a fairly straightforward way to differentiate
45 adaptive vs. non-adaptive causes of phenotypic differentiation (Whitlock 2008; Leinonen et al.
46 2013; Savolainen et al. 2013). In particular, the neutrality tests implemented in the program
47 ‘driftsel’ (Karhunen et al. 2013, 2014) provide a statistically powerful approach able to
48 identify footprints of natural selection even in a fairly small sample of populations
49 (Ovaskainen et al. 2011). Nevertheless, the driftsel-approach has not become widely used in
50 studies of population differentiation, possibly because the the field has moved to use genomic
51 data to identify footprints of natural selection (e.g. Storz 2005; Nosil et al. 2009; Narum &
52 Hess 2011), and because common garden studies are logistically demanding. While genomic
53 studies of natural selection appear to provide less laborious paths towards identifying
54 adaptive differentiation than common garden experiments, they are still grappling with the
55 problem of false positive tests (Mallick et al. 2009; Bieme et al. 2013; Hoban et al. 2016), as
56 well as with the difficulty of linking the loci under selection to their phenotypic targets
57 especially in the case of polygenic traits (McKay & Latta 2002; De Kovel 2006; Le Corre &

58 Kremer 2012). In addition, genome scan approaches become very inefficient in picking up
59 signals of selection in systems having high background levels of differentiation, such as in
60 highly subdivided populations experiencing strong drift (Hoban et al. 2016).

61 Variation in fish head shape and jaw structures have been extensively studied as they are
62 important trophic traits influencing fitness (e.g., Roy et al. 2010; McGee et al. 2013) and even
63 believed to have spearheaded adaptive radiations (Brouwers 2011; Sallan & Friedman 2012).
64 While population differences in them have been sometimes shown to be genetically based
65 (Kimmel et al. 2005; Albert et al. 2008; McGee & Wainwright 2013), there is also evidence
66 that there is a strong plastic component to variation in head shape and feeding structures
67 (Troy et al. 1994; Dingemans et al. 2009). Hence, while there are many reports of head and
68 jaw shape differentiation among fish populations (e.g., Walker and Bell 2000; Tobler et al.
69 2011 Østbye et al. 2016), few studies have firmly established an adaptive (contra plastic)
70 basis for this differentiation.

71 Gill-raker morphology constitutes another set of traits that have been extensively studied in
72 the context of fish foraging ecology (e.g., Berner et al. 2008; Wund et al. 2012; Hosoki et al.
73 2019). Fish species have adapted to various habitats by evolving unique gill-raker
74 characteristics that correspond to their specific feeding strategies. As for the case of three-
75 spined sticklebacks (*Gasterosteus aculeatus*), marine populations typically have more,
76 closely spaced and longer gill rakers than freshwater populations (Gross and Anderson. 1984;
77 Leaver et al. 2012; Glazer et al. 2014; Magalhaes et al. 2021), reflecting their limnetic
78 feeding strategy in marine ecosystems and the benthic feeding strategy in freshwater
79 ecosystems (Ravinet et al. 2014). However, formal tests of the adaptive basis of this
80 differentiation are few (but see: Raeymaekers et al. 2007; Seymour et al. 2019).

81 The aim of this study was to test (1) whether pond and marine populations of nine-spined
82 sticklebacks (*Pungitius pungitius*) differ in head and gill raker morphology and (2) whether
83 this differentiation exceeds neutral expectations meaning that observed divergence has been
84 driven by natural selection. To do this, we conducted a common garden experiment and
85 raised fish from four pond and four marine populations to control for environmental effects
86 on phenotypes, and subjected the data to driftsel analyses to identify footprints of selection.
87 The results revealed genetically based phenotypic differentiation in head and gill-raker traits
88 between marine and pond sticklebacks, and that the differentiation in head traits has been
89 driven by divergent natural selection.

90

91 **Methods**

92 *Study populations and sampling*

93 The parental fish used in this study were collected May-June 2018 from four pond (Rytilampi
94 66.38482°N, 29.31561°E; Pyöreälampi 66.26226°N, 29.42916°E; Kirkasvetinenlampi
95 66.43673°N, 29.13568°E; Bynastjärnen 64.45416°N, 19.44075°E) and four marine
96 (Tvärminne 59.83333°N, 23.24900°E; Pori 61.59111°N, 21.47295°E; Raahe 64.68818°N,
97 24.46189°E; Bölesviken 63.66110°N, 20.20940°E) populations from Fennoscandia using
98 minnow traps or beach seine (Fig. 1a). The captured adults were transported to aquaculture
99 facilities of University of Helsinki, and maintained in 1 m³ flow-through freshwater aquaria
100 (one per population) until used in artificial fertilisations.

101 *Common garden experiments and genotype data*

102 Artificial fertilizations among randomly selected adult fish in breeding condition were
103 conducted between May 24th and July 4th 2018. Standard *in vitro* fertilization techniques and
104 egg rearing methods were applied following Arnott and Barber (2000) and as described in

105 detail in Fraimout *et al.* (2022). In brief, eggs from gravid females were obtained by gently
106 squeezing their abdomens over a petri dish. Sperm was retrieved from males after they were
107 over-anesthetized with tricaine methanesulfonate (MS-222) by dissecting the testes and
108 subsequently mincing them in the same petri dish containing the eggs. Eggs and sperm were
109 mixed to ensure fertilization and kept in water until hatching. To avoid fungal infections,
110 water in the petri dishes was changed twice daily and clutches were inspected for signs of
111 fungal infections or death - dead and infected eggs were removed. At hatching each clutch
112 was split in two replicate 11 x 10 x 10 cm plastic containers filled with filtered freshwater in
113 which they were maintained for a four week period during which yolk resorption took place.
114 The larvae were fed *ad libitum* with live brine shrimp (*Artemia sp. nauplii*). After this, all
115 replicated families were transferred to Allentown Zebrafish Rack Systems (Aquaneering Inc.,
116 San Diego, USA). These racks had a closed water circulation system with physical, chemical,
117 biological and UV filters. The fish in racks were fed for the first four weeks with a mixture of
118 *Artemia nauplii* and chopped chironomid larvae, and after this, with chopped chironomid
119 larvae. Part of the water in the rack systems was changed every fortnight to maintain good
120 water quality. Fish were reared in family groups (maximum five individuals per aquarium) in
121 racks for a period of ca. 1 year (mean age: 316.4 days) after which the fish were euthanized
122 with MS-222. Temperature and light conditions were kept constant throughout the
123 experiment (12:12 LD; 15°C) except during an overwintering period of 3 months (November
124 – January) when water temperature was lowered to 10°C and light cycle to 5:19 LD.

125 We used data on molecular variation among populations to estimate a neutral baseline for the
126 tests of selection (see below). Specifically, we obtained genome-wide SNP data for all
127 individuals using a restriction site-associated DNA (RAD) sequencing technique (2b-RAD;
128 Wang *et al.*, 2012). Genomic DNA was extracted from fin clips using a standard salting out
129 protocol (LoperaBarrero *et al.* 2008) and 2b-RAD libraries were built following Momigliano

130 et al (2018) using a slightly modified version of the protocol available online at
131 [https://github.com/z0on/2bRAD_GATK/blob/master/2bRAD_protocol_may15_2017_nnrw.p](https://github.com/z0on/2bRAD_GATK/blob/master/2bRAD_protocol_may15_2017_nnrw.pdf)
132 [df](https://github.com/z0on/2bRAD_GATK/blob/master/2bRAD_protocol_may15_2017_nnrw.pdf)). Briefly, high molecular weight DNA was digested using the *BcgI* enzyme (New England
133 Biolabs), adaptors were ligated and the fragments amplified via PCR following the protocol
134 provided in Momigliano *et al.* (2018). Libraries were pooled and the target fragments were
135 isolated using a BluePippin size selector (Sage Science). Libraries were sequenced using
136 Illumina technology (HiSeq 4000) at the Beijing Genomic Institute (BGI; Hong-Kong). Raw
137 reads were demultiplexed and mapped to the *P. pungitius* genome (v.6; Varadharajan *et al.*
138 2019) using bowtie2 (Langmead and Salzberg, 2012). We used samtools (v.1.10; Li *et al.*
139 2009) to convert SAM files to BAM files.

140 We also used SNP-data to identify sex of all F₁ offspring, as they were measured before
141 reaching sexual maturity. From the genotype data file, we used the *snpRelate* R package
142 (Zheng *et al.* 2012) to perform a Principal Component Analysis (PCA) based on all markers
143 located on the sex chromosome (Linkage group 12; Natri *et al.* 2019) and assigned sex to the
144 F₁ individuals according to their clustering with the parental individuals of known sex.
145 Finally, we pruned the dataset from markers in high linkage disequilibrium (LD) using the
146 *snpGdsLDpruning* function of the *snpRelate* package and retained only markers with LD <
147 0.8 and a minimum allele frequency (MAF) > 0.05 and removed all sex-linked SNPs. This
148 resulted in a total of 2,660 informative SNPs.

149

150 *Analyses of head size and shape*

151 We quantified variation in head size and shape among the study populations following a
152 landmark-based geometric morphometrics approach (Bookstein 1991). A total of 32
153 landmarks (see Fig. 1b for details) and semi-landmarks were digitized from the left side of

154 individuals from digital images using the tpsDig2.1 software (Rohlf 2006). Specifically, 4
155 fixed landmarks were placed as in Yang *et al.* (2016) at: the posterior extent of the
156 supraoccipital (landmark 1; Fig. 1b), the anterior insertion of the premaxilla (landmark 2; Fig.
157 1b), the anterior-ventral extent of the preopercular (landmark 3; Fig. 1b) and the ventral
158 extent of the preopercular bone (landmark 4; Fig. 1b). Semilandmarks were placed by
159 resampling a total of 30 points along an outline starting from landmark 1 and ending on
160 landmark 4 describing head shape using the ‘draw curves’ mode in tpsDig. To avoid
161 redundancy the first and last semi landmarks (overlapping with fixed landmarks 1 and 4)
162 were removed prior to analyses. All subsequent morphometric analyses from the obtained
163 landmark configuration were performed in R using the *geomorph* package (v.4.0.1; Adams *et*
164 *al.* 2013). First, we used the plotOutliers function in geomorph to identify outlier individuals
165 based on their Procrustes distance from the mean head shape of the samples. We removed 16
166 individuals outside the upper quartile range and retained a total of 425 individuals for
167 subsequent head shape analyses. Landmarks were superimposed following the Procrustes
168 generalized least square superimposition (Dryden & Maria 1992) and semi-landmarks were
169 slid to the mean shape configuration and their position optimized using Procrustes distance
170 by setting the *ProcD* option of the *gpagen* function to “TRUE” in geomorph. Head centroid
171 size obtained from the Procrustes alignment was used as measurement for head size in all
172 subsequent analyses. For head shape, we performed a PCA on the procrustes coordinates
173 using the *gm.prcomp* function. The non-null principal components (PC) describing shape
174 variation were used as shape data in the subsequent analyses and standardized for allometric
175 variation (Rolshausen *et al.* 2015). Specifically, we followed the allometric approach of
176 Leonart *et al.* (2000) where the standardized trait measurement PC_S is defined such that:

177
$$PC_S = PC_0 (HS_m / HS_0)^b \quad (1)$$

178 where PC_0 is the original (i.e., unstandardized) trait measurement, HS_M is the mean head
179 centroid size per population; HS_0 is the head centroid size of each individual and b is the
180 population-specific coefficient of within-slope regression of head centroid size on the focal
181 PC.

182 We tested for population and sex differences for head size and shape using procrustes
183 analyses of variance as implemented in the *geomorph* package. Differences in head centroid
184 sizes were tested using the *procD.lm* function and fitting a model including body length, sex,
185 habitat and population as explanatory variables. Similarly, we tested for differences in head
186 shape by fitting a model including log-transformed head centroid size, sex, habitat and
187 population as explanatory variables. We further tested for differentiation in head shape data
188 using multivariate analysis of variance (MANOVA) based on the standardized PC described
189 above.

190

191 *Analyses of gill raker number and size*

192 Gill raker traits were measured from a total of 126 individuals from 64 families across eight
193 populations — four pond ($n=51$) and four marine ($n=75$) populations. These samples
194 included 62 wild-caught individuals and 64 of their laboratory-raised F_1 generation offspring,
195 comprising 32 females, 92 males, and two individuals of unknown sex.

196 Four gill raker traits were measured: number of gill rakers, length, width, and spacing
197 between rakers. The number of gill rakers were counted on the left side of the first gill arch
198 unless the left gill rakers were damaged, in which case the right side counts were used instead.
199 As the dorsal arch is prone to damage during dissection, only gill rakers on the ventral arch
200 were counted. We measured the length and width of the longest gill raker along the arch, and

201 defined spacing as the center-to-center spacing between the base of a pair of adjacent rakers,
202 specifically between the longest raker and the adjacent one on the upper side (Fig. 1c).

203 To ensure high repeatability of the measurements, we followed the laboratory protocol of
204 Nicholas and Craig (2016). In summary, we first dissected the first branchial gill arch,
205 digested it in a 10% Potassium Hydroxide (KOH) solution for two hours, and stained it with
206 0.008% Alizarin Red S in 1% KOH in water for at least two hours, as the staining time varies
207 depending on the gill arch size. We then rinsed the raker to remove excess stain, placed the
208 specimen in 89% glycerol, and flat-mounted the skeleton on bridged cover slips.
209 Measurements were obtained using digital images taken with a fluorescent stereomicroscope
210 equipped with an Olympus SZX16 or a Nikon FI3 camera. Length-related measurements
211 were accurate to the nearest 0.01 mm using ImageJ software v. 1.51 (Schneider et al. 2012).
212 To validate the measurement data, we independently measured the same set of traits for 20%
213 of randomly selected samples and calculated the repeatability (Nakagawa and Schielzeth,
214 2010). The repeatability of these traits from two independent measurements ranged from
215 0.855 to 0.967 ($p < 0.001$), indicating a high level of precision in the measured traits.

216 As body size (standard length) varied between habitats and sexes, we size-standardized all
217 linear gill raker traits prior to statistical analysis. This was done by constructing linear mixed
218 effect models, incorporating habitat and sex as fixed factors and population and pedigree
219 information as random factors, to determine if marine and pond populations differed in mean
220 values of gill raker traits and whether these traits displayed sexual dimorphism. This analysis
221 was performed using the 'mmer' function in the 'sommer' R package (Giovanny, 2016) using
222 R version 4.2.2 (R Core Team, 2022).

223

224 *Tests of natural selection*

225 To test whether observed population differentiation in morphology exceeded neutral
226 expectations, the approach of Ovaskainen et al. (2011) as implemented in the R packages
227 *driftsel* and *rafm* (Karhunen *et al.* 2013) was adopted. We followed the workflow from
228 Karhunen *et al.* (2013) and first estimated the matrix of population-level coancestry from
229 genomic data. Due to the computational burden of the MCMC-based algorithm, we sampled
230 2000 random markers from the pruned SNP dataset (Li *et al.* 2019) which were used as input
231 for the *do.all* function of the *rafm* package. We ran the model for 15,000 MCMC iterations
232 with a burn-in period of 5000 and sampling every 10th iteration. The resulting coancestry
233 matrix was then used as a neutral baseline for the *driftsel* analysis. We ran the *MH* function of
234 the *driftsel* package using the posterior samples of the coancestry matrix as prior information
235 and used either head centroid size or the three first size-standardized-PC most representative
236 of habitat shape variation as input traits. Sex of the individuals was used as covariate and
237 body length was added as covariate for the model using head centroid size only. We
238 subsequently used the *S.test* and *H.test* functions to calculate the *S* and *H* statistics,
239 respectively, as indicators of signal of selection. The rationale behind each statistic is similar
240 to that of classical Q_{ST} - F_{ST} comparisons with values of *S* or *H* > 0.95 indicating signal of
241 divergent selection, values equal to 0.5 indicates a scenario compatible with neutral evolution
242 and values < 0.05 imply stabilizing selection. Contrary to the *S.test*, the *H.test* allows to
243 incorporate habitat information into the *driftsel* framework to test whether habitat similarity
244 correlates with phenotypic similarity among populations (Karhunen *et al.* 2014).

245 To verify the robustness of the selection test results, we also applied the method of Martin *et*
246 *al.* (2008) to our genetic and phenotypic data. This method is a multivariate approach to the
247 Q_{ST} - F_{ST} comparison and tests for proportionality between the genetic (co)variance matrix (**G**)
248 and the among-population divergence (**D**). Under a scenario of purely neutral evolution, the
249 proportionality between **G** and **D** is defined as:

$$250 \quad \mathbf{D} = 2F_{ST} / (1 - F_{ST}) \mathbf{G} \quad (2)$$

251 where F_{ST} is estimated from putatively neutral molecular markers (Lande 1979, Lofsvold
252 1988, Marroig and Cheverud, 2004, McGuigan et al 2005, Martin et al. 2008, Berner et al
253 2010). We applied the framework of Martin *et al.* (2008) to calculate the coefficient of
254 proportionality ρ_{ST} between \mathbf{D} and \mathbf{G} , and compared it to the value of $2F_{ST} / (1 - F_{ST})$ obtained
255 from the pruned SNP dataset. The rationale behind this test is that values of $\rho_{ST} > 2F_{ST} / (1 -$
256 $F_{ST})$ would indicate a signal divergent selection. We used the *pairwise.prop* R function of the
257 *neutrality* package from Martin et al. (2008) to calculate ρ_{ST} and the *hierfstat* R package
258 (v.0.5.11; Goudet & Jombart 2022) to get bootstrapped estimates of F_{ST} from neutral loci.
259 Statistical difference between ρ and $2F_{ST} / (1 - F_{ST})$ was assessed based on the overlapping of
260 confidence intervals around the two estimates.

261

262 ***Heritability of head morphology***

263 We estimated the heritability of head shape variation using animal models under a Bayesian
264 framework using the MCMCglmm R package (Hadfield 2010). We partitioned the
265 phenotypic variance V_P of head shape into its additive V_A genetic component and estimated
266 heritability as $h^2 = V_A / V_P$. For each population, we fitted a model including the first PC (PC1)
267 describing habitat variation in head as a response variable and added head centroid size as
268 covariate and sex of the individuals as fixed effect. We used the same approach to estimate h^2
269 of head centroid size using sex of the individuals as a fixed effect. For both models, to
270 account for actual relatedness among individuals in each population, we appended the
271 Genomic Relationship Matrix (GRM) estimated from SNP data as a random effect in the
272 model. The GRM was estimated from the pruned SNP data using the *G.matrix* function of the
273 *snpReady* R package (Granatto *et al.* 2018). Each model was run with 303000 MCMC

274 iteration with a burn-in period of 3000 and sampling every 100th iteration. Model checking
275 was performed by visually inspecting the trace plots of the MCMC chains and by inspecting
276 the effect sizes of variance components using the *effecSize* function of the *coda* R package
277 (Plummer *et al.* 2006). All h^2 values are reported as the median of the posterior samples along
278 with their 95% Highest Posterior Density (HPD) intervals.

279

280 **Results**

281

282 ***Population differentiation in head morphology***

283 We found significant effects of sex, population of origin and habitat on head centroid size
284 (Table 1). Freshwater fish had larger heads than marine fish (Fig. 2a) and head centroid size
285 varied to a greater extent among pond than among marine populations (Fig. S1). Comparison
286 between sexes revealed sexual dimorphism in head centroid size, males having bigger
287 relative head centroid size than females (Fig. 2a). We also found differences in head shape
288 among habitats with the first PC of shape variation discriminating between pond and marine
289 individuals (Fig. 2b). This difference was reflected by a significant habitat effect in the
290 procrustes Anova on head shape ($p < 0.001$, Table 2).

291

292 ***Population differentiation in gill raker traits***

293 Individuals from pond populations were larger than those from marine populations ($X_{\text{pond}} =$
294 53.39 ± 2.39 [S.E.] mm, $X_{\text{marine}} = 42.78 \pm 1.12$ mm; $F_{1,124} = 27.75$, $p < 0.001$), and females
295 were bigger than males ($X_{\text{females}} = 57.16 \pm 2.52$ mm; $X_{\text{males}} = 43.20 \pm 1.23$ mm; $F_{1,123} = 51.79$,
296 $p < 0.001$). Marine and pond populations differed significantly in both the number ($X_{\text{pond}} =$

297 9.94 ± 0.13 ; $X_{\text{marine}} = 8.59 \pm 0.08$; $t_{3.02} = 5.572$, $p = 0.01$) and length ($X_{\text{pond}} = 1.45 \pm 0.04$ mm;
298 $X_{\text{marine}} = 1.09 \pm 0.02$ mm; $t_{5.91} = 2.630$, $p = 0.039$) of gill rakers, as did the sexes ($X_{\text{females}} =$
299 9.11 ± 0.20 ; $X_{\text{males}} = 9.14 \pm 0.10$; $t_{122.79} = 2.794$, $p = 0.006$) for number and ($X_{\text{females}} = 1.31 \pm$
300 0.06 mm; $X_{\text{males}} = 1.21 \pm 0.03$ mm; $t_{119.68} = 4.549$, $p < 0.001$) for length. Sticklebacks from
301 pond populations had more and longer rakers (Fig. 3), and males possessed shorter rakers for
302 their size relative to females. However, no significant differences were observed in the width
303 or spacing of rakers across habitats or sexes ($p > 0.05$ in all tests).

304

305 *Tests of selection*

306 The three first PCs from the PCA of size-standardized-PC of shape variation were used as
307 input for driftsel which revealed a signal of divergent selection ($S = 0.953$ and $H = 0.993$; Fig.
308 4). All populations from both habitats were indicated to have diverged more than expected by
309 chance from their common ancestor (Fig. 4). The direction of divergence differed
310 consistently for marine and pond populations, with the exception one pond population (KRK;
311 Fig. 4b). Individuals from this population had an intermediate head morphology compared to
312 other populations and were not significantly different from the reconstructed ancestral
313 population (Fig. 4). We found similar results when analyzing head centroid size differentiation
314 with $S = 0.978$ and $H = 1$.

315

316 The multivariate $Q_{\text{ST}}\text{-}F_{\text{ST}}$ approach indicated that the proportionality coefficient ρ was higher
317 than $2F_{\text{ST}}/(1-F_{\text{ST}})$ obtained from neutral markers ($\rho_{\text{ST}} = 1.8$ [1.1; 5.3]; $2F_{\text{ST}}/(1-F_{\text{ST}}) = 0.828$
318 [0.766; 0.890]), thus indicating a signal of divergent selection for head shape. Neither X^2 or
319 Bartlett-corrected X^2 test rejected the null hypothesis of proportionality between **D** and **G** (p
320 $= 0.07$; $p=0.159$, respectively) thus indicating that both matrices are proportional.

321

322 We did not find evidence for diversifying natural selection for any gill raker trait ($S = 0.58$, H
323 $= 0.86$ for the univariate model for raker number; $S = 0.69$, $H = 0.89$ for length; $S = 0.49$, $H =$
324 0.62 for width; and $S = 0.54$, $H = 0.74$ for gap), suggesting that pond and marine populations
325 have not evolved significantly further away from the ancestral mean than would be expected
326 under neutrality. The same conclusion was reached combining all four gill raker traits in a
327 multivariate analysis ($S = 0.55$; $H = 0.91$). This was also confirmed by the proportionality
328 coefficient ($p_{ST} = 0.33$ [0.21; 0.80]) being not significantly higher than $2F_{ST} / (1 - F_{ST})$ of 0.828
329 [0.766; 0.890].

330

331 *Heritability of head size and shape*

332 We found moderate to high h^2 values for head shape and size in all populations (Table 3).
333 Estimates for different populations were not significantly different as indicated by the
334 overlapping 95% HPD intervals (Table 3).

335

336 **Discussion**

337 The results revealed consistent differences in head and gill-raker morphology between marine
338 and pond populations of nine-spined sticklebacks reared in common garden conditions. This
339 together with the results of the selection tests provide firm evidence that the genetically based
340 differentiation in head size and shape (but not in gill-raker morphology) has been driven by
341 divergent natural selection favouring different head morphology in marine and pond habitats.
342 Since the studied pond populations have evolved independently in total isolation since their
343 colonisation after the last glaciation (Herczeg et al. 2009; Shikano et al. 2010), the results
344 also provide strong evidence for parallel phenotypic evolution in head morphology: random
345 processes are not expected to result in evolution of similar morphologies in different

346 localities (Schluter 2000). However, even if the heritabilities of head traits were similar in
347 different populations, the question whether the genetic underpinnings of the similar head
348 morphology in different pond populations are underlied by the same or different genetic loci
349 remains to be investigated. Earlier studies of other traits in these populations suggest that the
350 similar phenotypic outcomes have been the result of recruitment of different genetic loci to
351 underlie similar phenotypic adaptations (Kemppainen et al. 2021; Fraimout et al. 2022; Yi et
352 al. 2023). In the following sections, we discuss how the results advance our understanding of
353 local adaptation in general, and that of sticklebacks in particular.

354 The driftsel approach circumvents many shortcomings associated with traditional Q_{ST} - F_{ST}
355 approaches aiming to test whether trait differentiation among populations exceeds that to be
356 expected due to stochastic processes alone (Karhunen et al. 2013). Here, it provided evidence
357 for adaptive nature of head size and shape differentiation among pond and marine stickleback
358 populations. As such, the results align with earlier evidence for adaptive differentiation
359 among the studied pond and marine populations (e.g. Herczeg et al. 2009; Ab Ghani et al.
360 2013; Karhunen et al. 2014; Fraimout et al. 2022). Particularly noteworthy is that despite the
361 high degree of background differentiation (mean $F_{ST} = 0.286$ across all populations) in our
362 data, driftsel was able to pick up a signal of divergent natural selection. The approach of
363 Martin *et al.* (2008) also suggested that differentiation in head shape was greater than the
364 neutral expectation but that the magnitude of divergence was moderate. Nevertheless, it is
365 conceivable that attempts to achieve the same with any F_{ST} outlier tests in data with this high
366 neutral background level of differentiation would likely fail (e.g., Hoban et al. 2016; Liet al.
367 2019). Besides, as pointed out by Bierne *et al.* (2013), outlier tests are not expected to work
368 well when genetic architecture of local adaptation has a polygenic basis, as is likely to be the
369 case for most quantitative traits of ecological interest.

370 While the differences in head size and shape were very clear and consistent between marine
371 and freshwater populations, the driftsel analysis revealed one anomaly: head traits in one of
372 the pond populations (KRK) seemed to diverge in an orthogonal direction from marine and
373 pond populations. While it might be tempting to ascribe such a deviation to random genetic
374 drift from the hypothetical ancestor, it is also possible that there might be a biological
375 explanation for this. Namely, in earlier analyses of behavioural differences between the same
376 marine and freshwater populations, the KRK population showed a similar deviation in
377 response to predation risk by piscine predators (Framout et al. 2021). What differentiates the
378 KRK population from other studied pond populations is that it has a history of artificial
379 introduction of brown trout (*Salmo trutta*) to this locality (Herczeg et al. 2009, 2010). Hence,
380 while sticklebacks from the three other pond populations have evolved in absence of piscine
381 predators since the ponds became isolated, the fish in KRK have been faced with recent
382 brown trout predation. Hence, it is conceivable that this might have selected their behaviour
383 and head morphology to converge back towards that of marine ancestors. Nonetheless, no
384 piscine predators were observed in the field at the time of sampling and further studies should
385 investigate the link between predation pressure and head shape morphology in *P. pungitius*.

386 While the evidence for adaptive differentiation in head size and shape was clear-cut, the
387 results regarding the gill-raker morphology were not. We observed significant genetically
388 based differentiation in gill-raker number and length between pond and marine fish, this
389 differentiation did not exceed neutral expectations. This result does not exclude the
390 possibility that natural selection could be behind the observed differentiation, only that
391 statistical evidence for this is lacking. The consistent results for head gill-raker traits
392 confirmed that the pond fish have evolved to become better suited for feeding on planktonic
393 prey (longer snouts, upturned jaws, more and longer rakers) and marine fish better at feeding
394 benthic prey with a blunt head, short snouts, fewer and shorter rakers. While this might at the

395 first sight seem counterintuitive and go against expectations grounded on the work done in
396 the related three-spined stickleback in which freshwater fish have benthic like heads and gill-
397 rakers and marine fish limnetic like heads and gill-rakers (Rundle et al. 2003; Reimchen and
398 Nosil, 2006; Raeymaekers et al. 2007), one has to keep in mind that there are subtle
399 differences in ecologies of the three- and nine-spined stickleback species (Merila 2013).
400 Specifically, the lack of piscine predators in isolated ponds opens up the pelagic niche for
401 nine-spined sticklebacks (Karlsson & Byström 2005) and could favor development of more
402 numerous and longer gill-rakers. Conversely, marine nine-spined sticklebacks (and those
403 living in large lakes with predatory fish) living with piscine predators differ from three-
404 spined sticklebacks in that they are not as gregarious as three-spined sticklebacks and tend to
405 hide and feed in protection vegetation. Hence, the lack of marine three-spined stickleback-
406 like gill-raker morphology in marine nine-spined sticklebacks makes sense if they rely less on
407 planktonic diet and feed more benthically compared to syntopic three-spined sticklebacks.

408 What would be the selective advantage of the large head with strongly upward protruding jaw
409 for pond sticklebacks? In absence of data on diet composition and/or trophic positioning data
410 derived from stable isotopes from the study populations, we can only speculate to this effect.
411 In the absence of native predatory fish, sticklebacks colonising freshwater habitats tend to
412 evolve large heads with strongly upward protruding jaws (Walker & Bell 2000; Bell &
413 Aguirre 2013; but see: Voje et al. 2013; Østbye et al. 2016). Such head morphology is
414 thought be an adaptation to feed on limnetic diets and improved feeding performance on
415 zooplankton (Lavin & McPhail 1986; Walker & Bell 2000). On the other hand, the head
416 shapes of freshwater three-spined sticklebacks feeding primarily either on a limnetic or
417 benthic diet (Wootton 1984; Gow et al. 2007) resemble closely those of pond and marine
418 nine-spined sticklebacks, respectively. This could indicate that pond nine-spined sticklebacks
419 released from predation pressure from piscine predators have shifted their diet towards

420 feeding more zooplankton. Data from one of our study ponds (Rytilampi) indicates that
421 although sticklebacks in this pond tend to feed mostly benthic prey, they do also consume
422 pelagic zooplankton (Merilä & Eloranta 2017). Furthermore, data from Swedish lakes with
423 predatory fish show that large nine-spined sticklebacks tend to be more planktivorous than
424 their smaller conspecifics as they experience reduced predation risk in the pelagic open-water
425 areas (Karlsson & Byström, 2005). The results of the analyses of gill-raker morphology align
426 with these inferences: pond sticklebacks had long gill-rakers typical for fish adapted on
427 feeding zooplankton, whereas marine ones had gill-rakers typical for benthivorous fish
428 (Ingram et al. 2012). Hence, while the information gleaned from literature gill-rakers
429 suggests that the observed divergence - clearly driven by divergent natural selection - is likely
430 associated with differences in food acquisition, further studies looking into diets and feeding
431 modes are needed to establish this firmly. Nevertheless, what is clear is that the observed
432 divergence in head morphology cannot be explained by neutral process or phenotypic
433 plasticity, and that the observed differences are not subtle, but apparent also for the naked eye
434 (Fig.1b) and corresponding to substantial differences with large effect sizes.

435 The results further confirm that not only are freshwater stickleback populations adapted to
436 their respective habitats, but also possess ample heritable variation in both relative head size
437 and shape. Heritability estimates were fairly similar for marine and pond populations,
438 suggesting that the latter have not suffered from massive loss of adaptive variation in spite of
439 having lost significant amounts of neutral genetic variation due to population size bottlenecks
440 and drift (Shikano et al. 2010; Kivikoski et al. 2023). Although these heritability estimates
441 derive from full-sib analyses and hence potentially include maternal effect and dominance
442 variance, earlier analyses of body size variation from two of these populations suggest that
443 maternal effects dissipate quickly (Shimada et al. 2011) and there are no quantifiable
444 dominance effects (Ab Ghani et al. 2012; Fraimout et al. 2021). Hence, we conclude that

445 head size and shape are heritable, and hence able to respond to natural selection in all studied
446 populations also in future.

447 Finally, although the mean head size and shape differed between marine and pond
448 populations, one should note that there was also some degree of overlap among pond and
449 marine populations, and heterogeneity in mean values of these variables as for instance
450 reflected in variance in the degree that the different ponds were indicated to have diverged
451 from the reconstructed ancestral form. This is not surprising given that freshwater habitats are
452 heterogenous ecosystems in terms of their biotic and abiotic conditions and therefore, the
453 direction and strength of natural selection on phenotypic traits are unlikely to be equal in all
454 ponds. In fact, this kind of heterogeneity in mean trait values in replicate freshwater
455 populations of sticklebacks has been seen also in earlier studies (e.g., Kaeuffer et al. 2012;
456 Østbye et al. 2016).

457 In line with earlier results (Herczeg et al. 2010), we observed strong sexual dimorphism in
458 head size and shape, males having larger differently shaped heads than females. While this
459 might be indicative of niche partitioning among the sexes, it is also possible that sex
460 differences are related to different sex roles. Nine-spined sticklebacks males use their jaws to
461 build nests into which females lay their eggs, and males also defend their nests and hatchling
462 offspring with their mouths (Wootton 1984). The degree of divergence among the two sexes
463 was less compared to that among habitats - this may suggest that strength of habitat
464 associated selection exceeds that of sex-specific selection.

465

466 ***Conclusions***

467 To sum up, the results provide evidence for adaptive divergence in head size and shape
468 between pond and marine nine-spined sticklebacks, as well as genetically based divergence in

469 gill-raker length and number. However, we failed to find evidence to indicate that the latter
470 differentiation has been driven by natural selection. We further show the male and female
471 sticklebacks display pronounced sexual dimorphism in head size and shape, and that further
472 studies would be needed to understand the functional significance of this dimorphism. The
473 result further suggests moderate to high heritability of relative head size and head shape
474 suggesting that these traits have high potential to evolve in response to natural selection also
475 in future.

476

477 **Ethical statement**

478 All experiments were conducted under a permit from the Animal Experiment Board in
479 Finland (permit reference ESAVI/4979/2018). Permission to collect fish from PYO, KRK
480 and RYT ponds was obtained from Metsähallitus (license # MH794/2018). Authorisation to
481 catch fish from the marine populations was provided with personal National Fishing Licenses.

482 **Data Availability**

483 The raw data underlying this article including pictures, TPS files and genotype data will be
484 made available in the Dryad Digital Repository. R scripts to replicate all analyses are publicly
485 available at: <https://github.com/afraimout/>

486 **Author contributions**

487 J.M. and A.F. conceived and designed the research. A.F. conducted field work and collected
488 field samples. A.F. and Y.C. performed wet lab and data analyses. J.M. and A.F. led the
489 writing effort, K.R. and Y.C. provided input and reviewed the manuscript.

490

491 **Acknowledgements**

492 We thank Jacquelin DeFaveri, Niko Björkell, Karlina Ozolina, Niina Nurmi, and
493 Miinastiina Issakainen for their help in sampling and rearing sticklebacks. M. Issakainen also
494 helped with DNA extractions. LEE Jae Hyun kindly drew the gill raker diagram. We thank
495 Michael Bell for guidance on staining and measuring gill rakers. Our research was funded by
496 the Academy of Finland (grant # 218343 to JM). The authors have no conflict of interest to
497 declare.

498

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720 **TABLES AND FIGURES**

721 **Table 1. Results from the Procrustes ANOVA on head centroid size.** The analysis of
722 variance (ANOVA) table is shown for the model testing for differences in head centroid size.
723 Effect: the fixed effect included in the models. Df: degrees of freedom. SS: sums of squares;
724 MS: mean squares; Rsq: correlation coefficient (R^2); F: values from the F-distribution; Z:
725 effect sizes. *p*: *p*-value with bold font indicating statistical significance ($p < 0.05$).

Effect	Df	SS	MS	Rsq	<i>F</i>	<i>Z</i>	<i>p</i>
Body length	1	14.735	14.735	0.716	3227.578	12.631	0.001
Habitat	1	2.966	2.966	0.144	649.773	9.425	0.001
Population	6	0.157	0.026	0.007	5.737	3.975	0.001
Sex	1	1.424	1.424	0.055	250.226	7.018	0.001
Residuals	348	1.589	0.005	0.077			

726

727 **Table 2. Results from the Procrustes ANOVA on head shape.** The analysis of variance
728 (ANOVA) table is shown for the model testing for differences in head shape. Effect: the fixed
729 effect included in the models with semi-columns indicating interaction. Df: degrees of
730 freedom. SS: sums of squares; MS: mean squares; Rsq: correlation coefficient (R^2); F : values
731 from the F-distribution; Z : effect sizes. p : p -value with bold font indicating statistical
732 significance ($p < 0.05$).

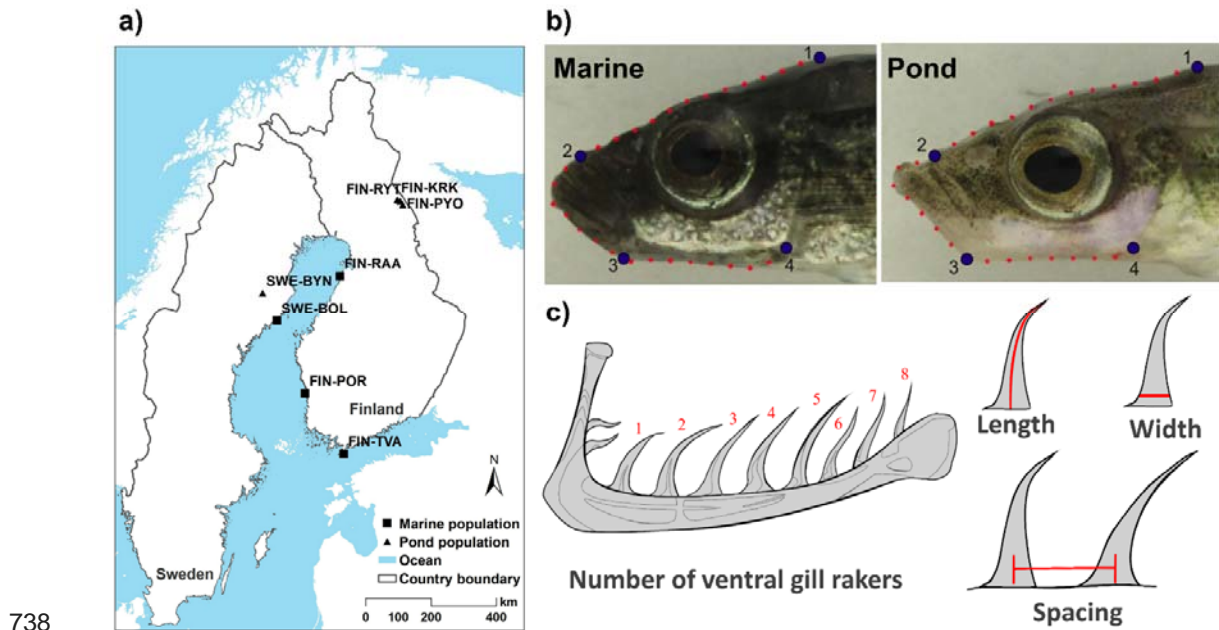
Effect	Df	SS	MS	Rsq	F	Z	p
Log(Csize)	1	0.072	0.072	0.123	69.743	7.479	0.001
Habitat	1	0.085	0.085	0.144	81.928	7.176	0.001
Population	6	0.059	0.009	0.100	9.156	7.307	0.001
Sex	1	0.012	0.012	0.019	11.302	4.508	0.001
Log(Csize):Habitat	1	0.001	0.001	0.002	1.284	0.724	0.241
Residuals	347	0.359	0.001	0.610			

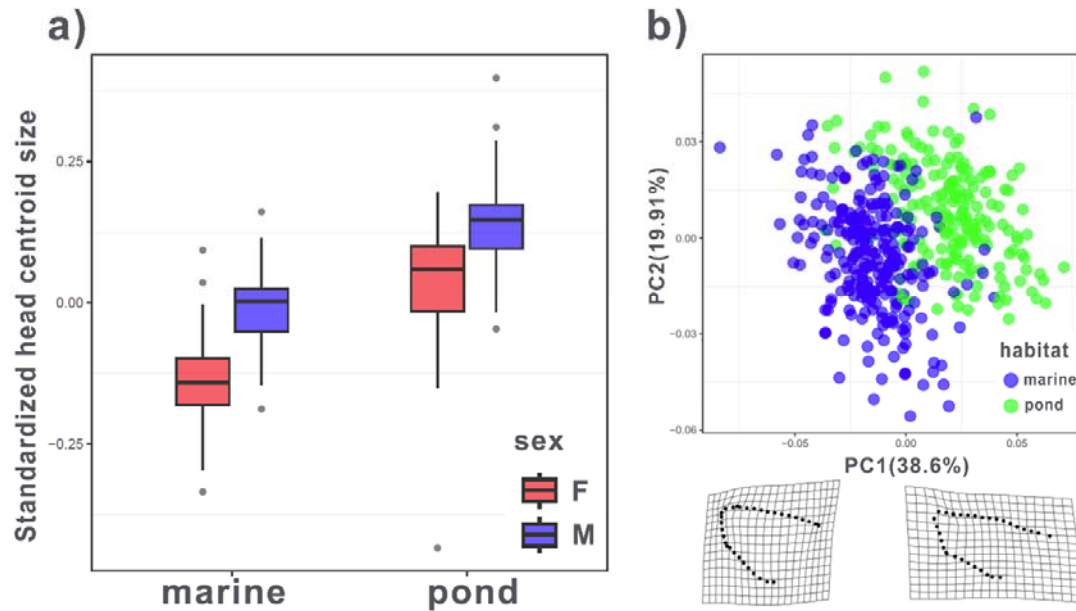
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734 **Table 3. Estimates of heritability for head centroid size and head shape.** For each
735 population, the posterior median of h^2 is shown for head centroid size (h^2_{size}) and the PC1 of
736 head shape variation (h^2_{shape}) along with their 95% highest posterior density intervals.

Population	h^2_{size}	h^2_{shape}
SWE-BYN	0.466 [0.223; 0.766]	0.566 [0.393; 0.718]
FIN-KRK	0.509 [0.168; 0.826]	0.519 [0.286; 0.778]
FIN-PYO	0.449 [0.222; 0.723]	0.543 [0.363; 0.721]
FIN-RYT	0.389 [0.125; 0.701]	0.674 [0.538; 0.785]
SWE-BOL	0.464 [0.221; 0.732]	0.529 [0.387; 0.655]
FIN-RAA	0.424 [0.201; 0.682]	0.630 [0.489; 0.741]
FIN-TVA	0.504 [0.250; 0.270]	0.562 [0.439; 0.700]
FIN-POR	0.499 [0.221; 0.798]	0.559 [0.415; 0.710]

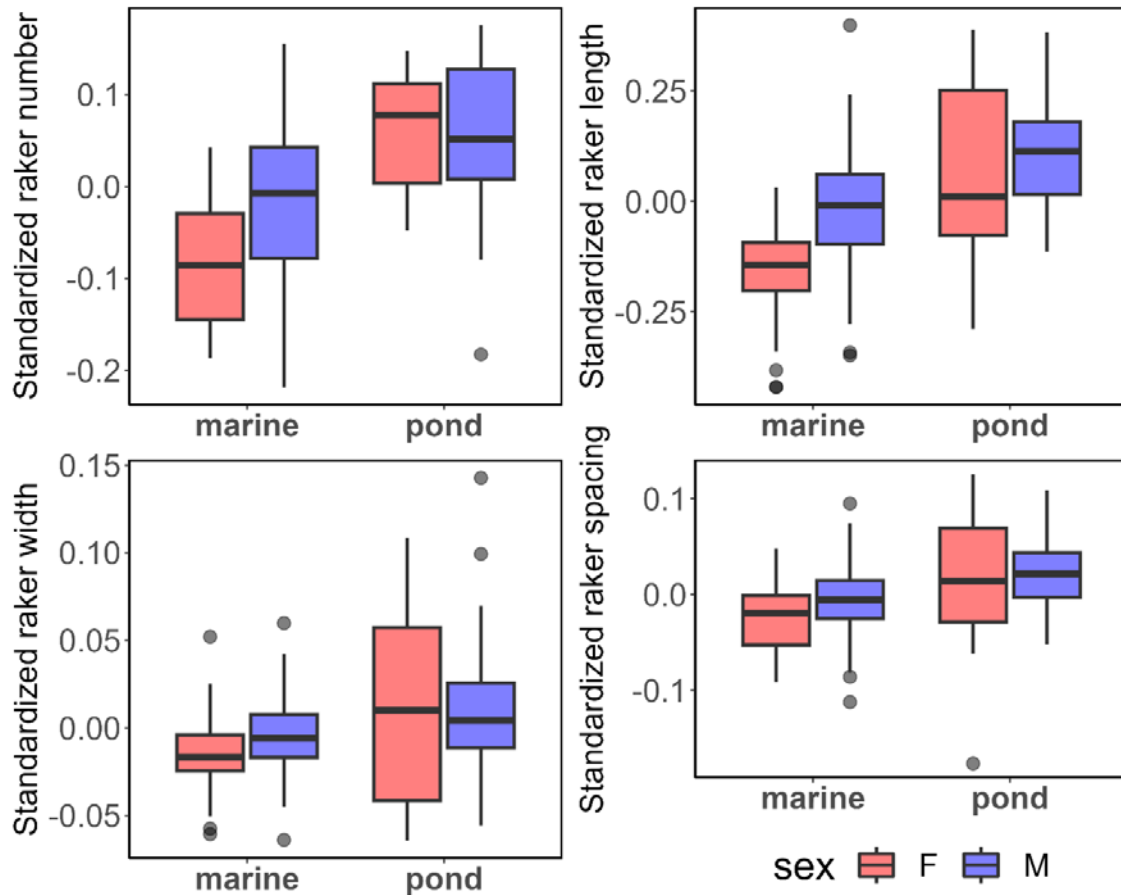
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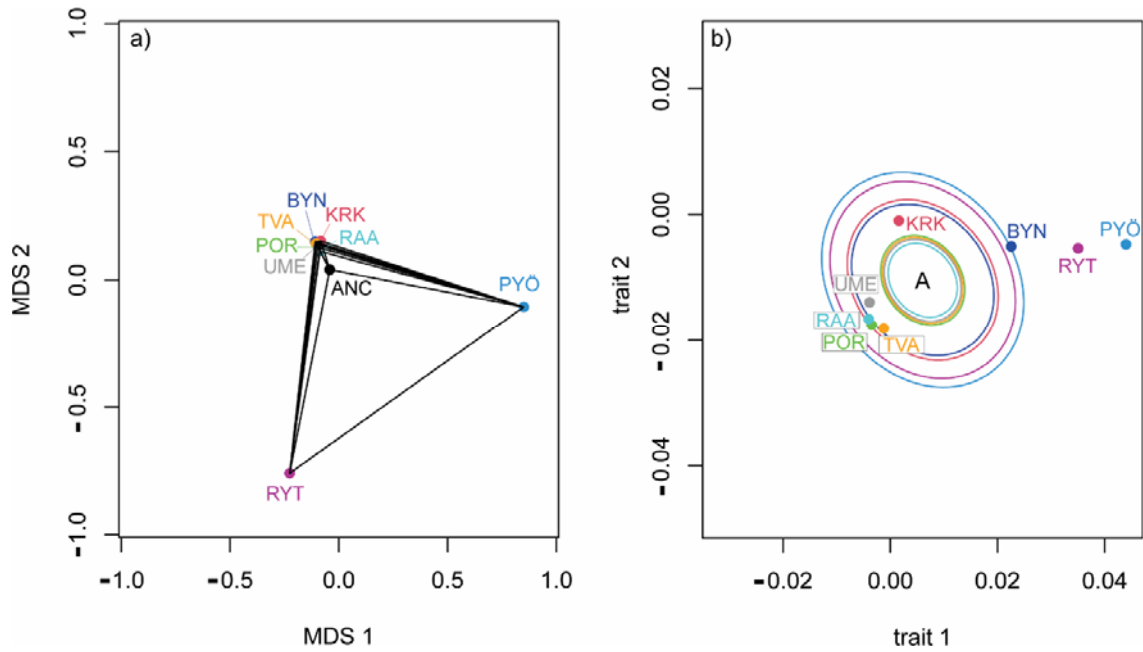
748 **Fig. 2.** Head centroid size and shape differentiation. a) The residuals from a regression of
749 head centroid size on body length were used to describe standardized head size and are
750 plotted for each habitat (marine or pond) and the two sexes separately (males = M, females =
751 F). b) Principal component (PC) analysis of head shape variation. Individuals from marine
752 (blue filled circles) and pond (green filled circles) habitats are plotted along the two first PC
753 axes. Deformation grids depict the shape change associated with PC1 of head shape variation.



754

755 **Fig. 3.** Differences in mean size standardized values of four raker traits in marine and pond

756 environments, as well as in two sexes (females = F, males = M).



757

758 **Fig. 4.** Results of the driftsel analysis. a) Multidimensional scaling (MDS) patterns of neutral genetic
759 differentiation estimated from SNP data obtained from the *viz.theta* function in program driftsel. b)
760 Population means in the trait space from the *viz.trait* function. Ellipses represent the expected median
761 distance from the inferred ancestral population (ANC) under neutral evolution.