## SUPPLEMENTARY METHODS

<sup>2</sup> Simulations: We model adaptive walks in diploid populations with Wright-Fisher

<sup>3</sup> simulations using Fisher's geometric model (FGM) as in Sellis et al. (2011). The

<sup>4</sup> simulations use code modified from Sellis et al. to allow for more than 2 dimensions. We

<sup>5</sup> perform 10,000 replicate simulations with population size N = 5,000. Simulations are

<sup>6</sup> conducted for 10,000 generations. Complete source code is available at

7 https://github.com/sunthedeep/Fisher-Geometric-Model.

In FGM, alleles are represented as a vector in n-dimensional space (Figure 1a). We explore 8 two models, one with two dimensions and one with 25 dimensions. We define the 9 phenotype of a diploid individual as the midpoint of the two vectors of the constituent 10 alleles (Sellis et al., 2011). This amounts to an assumption of phenotypic additivity of 11 alleles, but not necessarily additivity of fitness. The population initially contains a single 12 allele with a distance of 2 units from the optimum, and evolves on a symmetrical Gaussian 13 fitness landscape with single phenotypic optimum at the origin. Fitness is computed using 14 the function: 15

$$w(x) = e^{\frac{-x^2}{2}}$$

where x is the distance of the individuals phenotype to the optimum. The mutation rate is set to  $\mu = 5 * 10^{-6}$  which results in one mutation every 20 generations on average in the population. The angle of the mutation vector is drawn from a spherically uniform distribution. The magnitude of the mutation vector is drawn from an exponential distribution. For the two dimensional regime, the mean of the mutational magnitude is 0.5, while for the 25 dimensional regime, the mean is set to 5. The mutational magnitudes were

chosen to generate sufficient numbers of adaptive walks both with and without
overdominant mutations in simulations of both dimensionalities.

For the remainder of our analysis, we identify the most frequent allele in each simulated 25 population at the end of 10,000 generations of evolution and study the mutations present 26 on that allele. This is the set of mutations typically available for study in experimental 27 systems. We limit our analysis to studying the first five mutations of each adaptive walk 28 and ignore simulations with fewer than 5 mutations in order to control for the length of the 29 adaptive walk when studying predictability. We partition our five-mutation adaptive walks 30 into those that do and those that do not contain overdominant mutations to study the 31 impact of balanced states on predictability. 32

**Partitioning Walks:** Throughout all of our analysis, we have separated walks with and 33 without overdominant mutations. The methodology for this separation is as follows. For 34 each FGM simulation, we have identified the most frequent allele at the end of the 35 simulation, and isolated the first five mutations to occur on this allele. We first determine 36 the time at which the fifth mutation exceeded 5% frequency in the population, which we 37 use as a cutoff for eliminating alleles that have increased in frequency due to drift. All 38 time-points in the simulation after this threshold time are no longer considered for analysis. 39 Throughout the remainder of the simulation, we compute whether the alleles in the 40 population at at least 1% frequency could be maintained as a balanced polymorphism 41 using the method of Kimura (1956). If, at any time before the threshold time, the 42 population contains a set of alleles that could be maintained in a stable polymorphic state, 43 the walk is classified as containing at least one overdominant state. 44

Forward Predictability Analysis: We calculate the forward predictability of the
adaptive trajectory using two metrics. In both of these metrics, we only consider

<sup>47</sup> homozygous phenotypes. Our first metric, maximum pairwise distance, considers pairs of
<sup>48</sup> adaptive walks. We compute the maximum of the phenotypic distances between the
<sup>49</sup> observed single mutant phenotypes of the two adaptive walks, the double mutant
<sup>50</sup> phenotypes, the triple mutant phenotypes etc. Our second metric measures the maximal
<sup>51</sup> deviation from the optimal trajectory. For each adaptive walk, we compute the maximal
<sup>52</sup> phenotypic distance of any encountered (homozygous) phenotype from the line segment
<sup>53</sup> connecting the ancestral phenotype and the optimum.

**Backward Predictability Analysis:** We calculate the probability of all possible 54 mutational trajectories for the given set of mutations in a manner similar to Weinreich 55 et al. (2006), but generalized to allow balanced states. The likelihood of a mutational 56 trajectory is the product of the probabilities of each mutation in the trajectory being 57 generated on the appropriate background and successfully invading the population and 58 reaching equilibrium. The probability of a mutation landing on the appropriate 59 background is proportional to the frequency of the background. For example, if the 60 ancestral allele is balanced with the 1-mutant allele, the probability of generation of a 61 2-mutant allele is proportional to the frequency of the 1-mutant allele. 62

The probability of a new allele (generated through mutation) invading and reaching a 63 stable intermediate frequency or fixing from a single copy is calculated empirically through 64 10,000 Wright-Fisher (Fisher, 1930; Wright, 1931) simulations. These simulations are 65 entirely separate from the FGM simulations used to generate the adaptive walks used 66 throughout this work. In order to compute this probability, one first needs to define the 67 expected equilibrium frequency of a new allele. This is complicated because new alleles can 68 potentially balance with any of the existing alleles or fix in the population. We first ask 69 whether the new allele can balance with any of preexisting alleles by determining whether 70 the fitness of the heterozygous genotype is greater than the fitness of both homozygous 71

<sup>72</sup> genotypes. If the allele can balance, we compute the equilibrium frequency of the new

 $\frac{t}{s+t}$ 

73 (derived) allele as

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from Gillespie (2004) when  $0 \le s \le 1$  and  $0 \le t \le 1$ , where an individual homozygous for 75 the ancestral allele has a relative fitness of 1 - t, the heterozygote has fitness 1 and the 76 derived homozygote has fitness 1 - s. If one of the two homozygous genotypes is the most 77 fit state, we determine whether the new allele is either fixed or lost and set the expected 78 equilibrium frequency of the new allele to 1 or 0, respectively. Using this equilibrium 79 frequency, we can compute the mean fitness of the population at equilibrium for this pair 80 of alleles. Through this process, we make a simplifying assumption that balanced states 81 with more than two alleles are unlikely. This process is repeated to compute the mean 82 fitness of every possible balanced state involving the new allele. We choose the state with 83 the highest mean fitness as the equilibrium condition for this new allele. This can be either 84 a balanced state or fixation or loss of this new allele. If the new allele is not present in the 85 computed equilibrium condition with the highest mean fitness, we determine that the new 86 allele cannot invade the population. Otherwise, if the new allele is able to invade the 87 population, we compute the likelihood of reaching this equilibrium condition by asking how 88 frequently the new allele (starting at a single copy) can get to 90% of its equilibrium 89 frequency in our Wright-Fisher simulations. 90

We are forced to utilize empirical estimations through simulations and not the classical analytic solutions (Haldane, 1927; Kimura, 1962) for invasion and fixation probability as many of the observed mutations have a selective advantage exceeding 100%, violating the assumptions of the analytic solutions that the mutations are weakly beneficial. Our

simulations (not shown) suggest that the analytic solutions significantly overestimate the
invasion probability under these conditions.

We validate this method of computing the likelihood of a particular adaptive trajectory by 97 testing whether the high probability trajectories are more likely to have been observed in 98 our FGM simulations. We sort all trajectories by their computed probability of occurrence 99 (excluding those with zero probability) and bin them into 40 equally-sized bins. We found 100 that the median of the trajectory probabilities within a bin is significantly positively 101 correlated with the number of trajectories in that bin that were observed in our original 102 FGM simulations (Pearson  $r^2 = 0.997$ ,  $p \ll 10^{-10}$ ), suggesting that our method is truly 103 capturing the likelihood of a trajectory taking place in the FGM simulations. 104

Note that in a traditional haploid model, where each successive mutation fixes in the 105 population, there are 5! = 120 possible orders of the five mutations observed in the 106 simulated walk to generate the five-mutant allele observed at the end of the simulation. In 107 our diploid model, however, each mutation can occur on any allele in the population where 108 it has not already been introduced. Therefore, the same mutation can be introduced into 109 the population multiple times, but it can occur only once on each allele. The probabilities 110 of all viable mutational orders (where the 5-mutant allele is successfully reached) are then 111 rescaled to add up to 1, to give the probability of a trajectory conditional on reaching the 112 final 5-mutant allele. Mutation orders where the mutations were introduced on the final 113 adapted allele in the same order, but different balanced intermediate alleles were 114 encountered, are not distinguished from each other as these would be indistinguishable in 115 natural systems. 116

<sup>117</sup> We define the effective number of trajectories as

 $\frac{1}{\sum p^2}$ 

where p is the rescaled probability for each viable trajectory in a given five-mutation 119 system. The effective number of trajectories is defined to be 0 when there are no viable 120 trajectories. This is similar to the effective number of alleles in a population (Kimura1964), 121 the predictability metric of Roy (2009) and the entropy metric of Palmer et al. (2013). 122 Thus, when a single trajectory dominates the probability density, the effective number of 123 trajectories is close to 1, even if there are many other trajectories with nonzero probability. 124 This provides a single metric of the diversity of mutational orders that are possible and 125 summarizes the backward predictability of the adaptive walk. 126

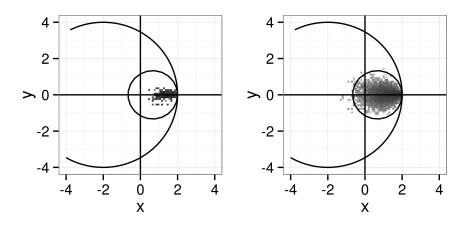




Figure S1. Phenotypic distribution of homozygous alleles in the first two phenotypic
dimensions as in Figure 1b,c for FGM simulations conducted using 25 dimensions and a
mean mutational magnitude of 5.

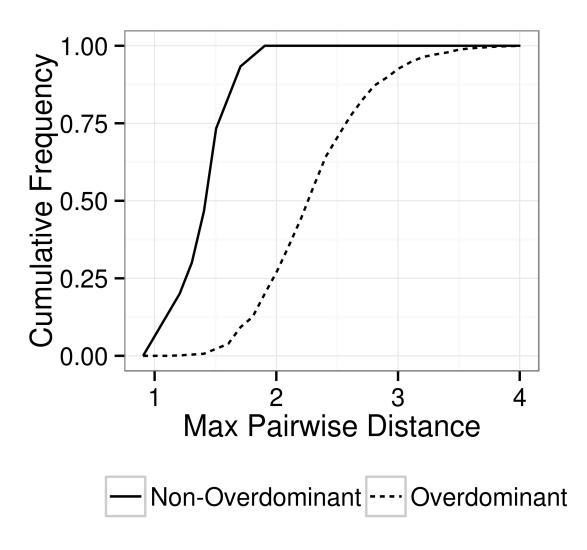
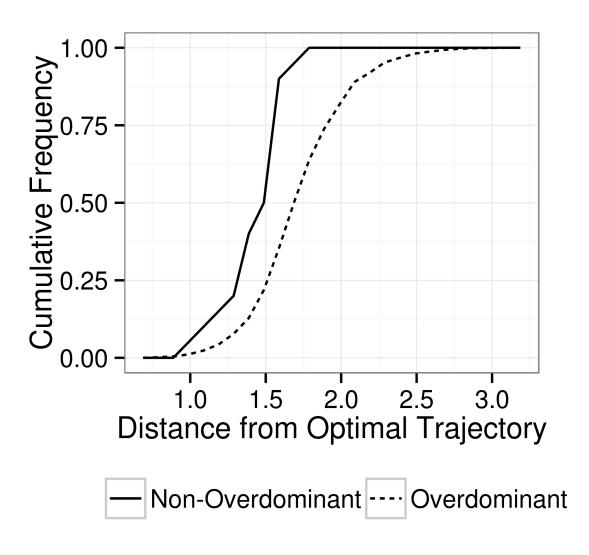


Figure S2. Maximum pairwise distance metric to study forward predictability as in Figure
2 for FGM simulations conducted using 25 dimensions and a mean mutational magnitude

<sup>134</sup> 2 for FGM simulations conducted using 25 dimensions and a mean mutationa

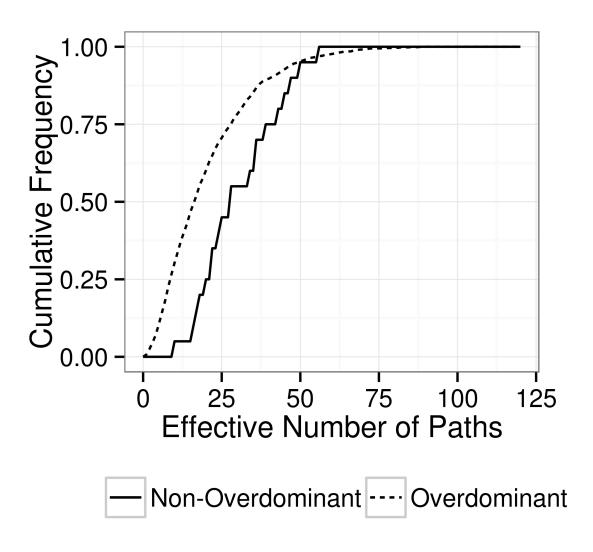
135 of 5. Kolmogorov-Smirnov test  $p \ll 10^{-10}$ 



<sup>137</sup> Figure S3. Maximum distance from the optimal trajectory metric to study forward

<sup>138</sup> predictability as in Figure 3 for FGM simulations conducted using 25 dimensions and a

<sup>139</sup> mean mutational magnitude of 5. Kolmogorov-Smirnov test  $p = 10^{-6}$ 



<sup>141</sup> Figure S4. Effective number of paths statistic to study backward predictability as in Figure

<sup>142</sup> 4 for FGM simulations conducted using 25 dimensions and a mean mutational magnitude

<sup>143</sup> of 5. Kolmogorov-Smirnov test p = 0.03

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