## Supplementary Material

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## Supplementary Note

## Supplementary Note 1. Potential Sources of bias <br> 1). Motif-specific error rates

It has been shown that certain sequence motifs may be more susceptible to sequencing error, which could lead to a non-random distribution of false positive singleton calls and subsequently bias our analyses ${ }^{1,2}$. Allhoff et al. $(2013)^{2}$ reported context-specific errors for the Illumina HiSeq platform, noting that the most common of these are strand-specific $T>X$ errors at 5'-GGGT-3' motifs (i.e., there is no evidence of an excess of $A>X$ errors at the reverse complement $5^{\prime}$ - $\operatorname{ACCC}-3$ ' motifs). We reason that if the BRIDGES ERVs are enriched for such context-specific errors, we should see significantly more T>X ERVs at the 5'-GGGI-3' motif than $A>X$ ERVs at the $5^{\prime}$ - $\underline{A C C C}-3$ ' and motif. Of the 115,531 ERVs that occur at this motif, 57,699 were $5^{\prime}-[A>X] C C C-3$ ' variants, and 57,832 were $5^{\prime}-\mathrm{GGG}[\mathrm{T}>X]-3$ ' variants; this difference was not significant, indicating there is no evidence for an enrichment of $\mathrm{T}>\mathrm{X}$ ERVs at this error-prone motif (exact binomial test; $P=0.70$ ). Allhoff et al. (2013) remark that the variants called at error-prone positions tended to have low base quality scores as well as significant strand bias, both of which are detectable with standard filtering protocols ${ }^{2}$. We therefore assume that most motif-specific errors are efficiently filtered by the default strand-bias and quality filters used in our variant calling pipeline, and any undetected errors have a negligible impact on our calculation of relative mutation rates and downstream analyses.

## 2). Mapping error

We also considered the possibility that ERVs occurring on poorly mapped reads might bias our analysis of regional variation in mutation rates. We expect the majority of ERVs in our data are mapped with high confidence, as the pre-filtering steps in our variant calling pipeline remove sites with average phred-scaled mapping quality score $(M Q)<20$ and/or with more than $10 \%$ of reads that are ambiguously mapped (MQ0>10). This filtering strategy is similar to the filters employed by other largescale sequencing projects that have demonstrated well-controlled error rates among singleton calls ${ }^{3,4}$. While a more aggressive mapping quality filter would reduce concerns about region-specific error
biases, doing so would primarily filter out ERVs occurring in repeat-rich pericentromeric regions ${ }^{5}$ thus precluding our ability to assess the mutation spectrum in these regions. Prior research has found that centromeric and pericentromeric regions evolve more rapidly than elsewhere in the genome ${ }^{5-7}$, which is an intriguing phenomenon that would be entirely undetectable if we omit these regions from our analyses.

## Supplementary Note 2. Comparison of 7-mer relative mutation rates with independent estimates

Aggarwala \& Voight (2016) ${ }^{8}$ estimated "substitution rates" using 7,051,667 intergenic variants observed in $\mathrm{N}=379$ Europeans from the 1000 Genomes Phase I study. These substitution rates are analogous to the relative mutation rates used in our study, but are derived from variants across the entire frequency spectrum, encompassing both singletons and common variants. The exact site frequency spectrum for the European intergenic variants is not reported, but Aggarwala \& Voight $(2016)^{8}$ specify $26 \%$ of variants in the 1000G Phase I African sample are singletons or doubletons. Because the BRIDGES sample is $\sim 10$ times larger than the 1000G Phase I European sample, we expect many of the 1000G Phase I European singletons are present in the BRIDGES data in multiple individuals (i.e., nonsingletons), and hence ancestrally older. The rates estimated by Aggarwala \& Voight (2016) ${ }^{8}$ are therefore expected to be more similar to the BRIDGES MAC10+-derived relative mutation rates than they are to the BRIDGES ERV-derived rates. As shown in Supplementary Fig. 3a, the BRIDGES MAC10+-derived rates are more strongly correlated with rates estimated by 1000 Genomes intergenic variants ( $\mathrm{r}=0.995$ ) than with BRIDGES ERVs ( $\mathrm{r}=0.991$ ). Type-specific correlations between MAC10+derived and 1000G-derived rates are also higher for all types except $A>G$ and non- $C p G C>T$ transitions (Supplementary Fig. 3b). Only 129 of the 24,576 7-mer subtypes ( $0.5 \%$ ) have more than a 2 -fold difference between MAC10+-derived and 1000G-derived rates (Supplementary Fig. 3c), compared to 741 (3\%) of 7-mer subtypes with >2-fold difference between MAC10+-derived and ERV-derived rates. The rates estimated by Aggarwala \& Voight $(2016)^{8}$ constitute a benchmark by which we compare our models' ability to predict true de novo mutations; we show that our analogous model based on the

BRIDGES MAC10+-derived rates performs similarly to these previously published rates, and the models based on BRIDGES ERV-derived rates consistently predict de novo mutations with greater accuracy (Supplementary Fig. 6).

## Supplementary Note 3. Tests for enrichment/depletion of de novo mutations in feature-associated subtypes.

Our multivariate models identify specific 7-mer subtypes found to be enriched (or depleted) for ERVs when occurring in the presence of a genomic feature. While our model validation results demonstrate that accounting for these features in aggregate improves prediction of de novo mutations, it does not show that, for a given single feature, these subtype-specific effects could also be detected among actual de novo mutations. Because the available catalogs of de novo mutations are relatively sparse, validating each individual feature-associated 7-mer subtype is not feasible. Instead, we looked across all 7-mer subtypes associated with a given feature in the same direction, and tested if the de novo mutations of those subtypes were higher (or lower) than expected under the null assumption that a feature has no effect on those subtypes' mutability.

Hence, for each feature, we identified regions of the genome covered by that feature, and calculated the expected number of ERVs in those regions based on the 7-mer relative mutation rates (i.e., assuming the feature has no effect on mutability) of subtypes significantly associated in the same direction with the feature. Assuming no systematic bias, this number is proportional to the expected number of de novo mutations (e.g., if we expect 36,000 BRIDGES ERVs in those regions [0.1\% of all ERVs], we would expect $\sim 47[0.1 \%]$ of the $G o N L 9 / I n o v a{ }^{10}$ de novo mutations occur in the same regions). We compared the expected number to the observed number of de novo mutations using onesided Pearson's Chi-squared tests, each with 1 degree of freedom (prop.test() function in R). A significant result indicates that observed counts of de novo mutations in the feature vary as predicted (higher or lower) from the expected count. Ten of the 15 tests showed a significant enrichment or
depletion of observed de novo mutations (Supplementary Table 6a). These results are not solely a result of feature-associated DNA methylation, as the associations remained significant when subtypes with CpG dinucleotides were excluded (Supplementary Table 6b). Note that four of the non-significant tests described in Supplementary Table 6a where we predicted an increase in de novo mutations had fewer observed de novo mutations than expected (CpG islands, GC content, H3K27me3, and laminassociated domains). This may indicate false positive sin our model, but is also consistent withThis may a limited ability to confidently call de novo mutations in the GoNL/Inova datasets due to low coverage in GC-rich regions ${ }^{9,10}$. We conclude that most of the mutagenic effects of genomic features inferred by our model are likely operative in the germline and play a role in shaping mutation rate heterogeneity across the genome.

## Supplementary Note 4. Potential mechanisms for TTAAAA hypermutability

Our finding of a 3-fold depletion of TTAAAA AT>TA motifs in DNase hypersensitive sites provides an excellent example of how our results can be leveraged to better understand the origins of certain mutation patterns. We identify two possible mechanisms that might explain the context-dependent mutation probabilities of AT>TA mutations at TTAAAA hexamers. As described in the main text, L1 EN nicking activity has been shown to vary according to the nucleosomal context of its target motifs, usually occurring at a higher rate in nucleosome-free DNA, but in some cases actually decreasing in nucleosome-free DNA ${ }^{11}$. Therefore, under the L1 EN model, it is possible to see either a positive or negative association between TTAAAA mutability and DHS.

Slipped-strand mispairing, also known as replication slippage, is another plausible hypothesis for the hypermutability of this motif ${ }^{8}$. Because the nucleosomal architecture is disrupted ahead of the replication fork ${ }^{12}$, and reassembled almost immediately thereafter ${ }^{13}$, nascent DNA containing unresolved lesions that is packaged in nucleosomes could be inaccessible to mismatch repair machinery, thus preserving any errors caused by slippage. In this case, it is also possible to see a negative association between TTAAAA mutability and DHS.

This slippage mechanism, however, appears to be unlikely for the following reasons. First, replication slippage inherently results in short insertions or deletions rather than point mutations. Mapping error could potentially cause an insertion/deletion to be falsely identified as a single-nucleotide variant, but such errors would need to be extremely prevalent in our data (and also context-dependent) in order to observe a 3-fold depletion of these singletons in DHS. Given the quality metrics we report for the BRIDGES singletons, it seems unlikely that these results are purely a technical artifact.

Furthermore, if slippage were the primary mechanism, we would expect other motifs ending in poly-A 4mers to also show an inverse association with DHS. Among the 13 NNNAAAA subtypes whose mutability is significantly associated with DHS, only five are inversely associated, three of which are NNTAAAA motifs (i.e., conforming closely to the canonical target for L1 EN nicking activity). The other eight subtypes all show higher mutation rates in DHS, which conflicts with the proposed slippage+chromatinization mechanism.

## Supplementary Note 5. Derivation of false discovery rate by Ts/Tv statistics

(1) Let $T S_{o}=T S_{t p}+T S_{f p}$ be the number of observed transitions, consisting of both true positives ( $T S_{t p}$ ), and false positives $\left(T S_{f p}\right)$
(2) Let $T V_{o}=T V_{t p}+T V_{f p}$ be the number of observed transversions.
(3) Based on findings from other large-scale sequencing studies, the true positive Ts/Tv ratio, $T S T V_{T}=\frac{T S_{t p}}{T V_{t p}}$ is between 2.0 and $2.1^{14}$.
(4) Because there are 8 possible transversions and 4 possible transitions, if errors are occurring at random, the $T s / T v$ ratio for random false positive errors $\left(T S T V_{\epsilon}\right)$ should be 0.5 , that is, $\frac{T S_{f p}}{T V_{f p}}=$ 0.5 .

Solving this system of four equations, it follows that $T V_{f p}=\frac{T S T V_{T} \times T V_{o}-T S_{o}}{T S T V_{T}-0.5}$ and $T S_{f p}=0.5 \times T V_{f p}$, so the false discovery rate is estimated as:

$$
\frac{T S_{f p}+T V_{f p}}{T S_{o}+T V_{o}}=\frac{0.5\left(\frac{T S T V_{T} \times T V_{o}-T S_{o}}{T S T V_{T}-0.5}\right)+\frac{T S T V_{T} \times T V_{o}-T S_{o}}{T S T V_{T}-0.5}}{T S_{o}+T V_{o}}
$$

Assuming a true $T S T V_{T}$ between 2.0 and 2.1, we estimate a false discovery rate of $0.6-2.6 \%$ among the BRIDGES ERVs.

## Acknowledgements

The BRIDGES study was supported by R01 MH094145 to Michael Boehnke and Richard M. Myers and U01 MH105653 to Michael Boehnke. The collection and storage of cases and controls from the Centre for Addiction and Mental Health (CAMH) in Toronto and from the Institute of Psychiatry, Psychology and Neuroscience (loPPN), King's College London in London, U.K. was supported by funding from GlaxoSmithKline, from the Canadian Institutes of Health Research to John B. Vincent, MOP-172013 (CAMH), and funding from the National Institute for Health Research (NIHR) Biomedical Research Centre at South London and Maudsley NHS Foundation Trust and King's College London (IoPPN). The views expressed are those of the author(s) and not necessarily those of the UK NHS, the NIHR or the UK Department of Health. Case and control collection was supported by Heinz C. Prechter Bipolar Research Fund at the University of Michigan Depression Center to Melvin G. McInnis (Prechter). Data and biomaterials were collected for the Systematic Treatment Enhancement Program for Bipolar Disorder (STEP-BD), a multi-center, longitudinal project selected from responses to RFP \#NIMH-98-DS-0001, "Treatment for Bipolar Disorder" which was led by Gary Sachs and coordinated by Massachusetts General Hospital in Boston, MA with support from 2N01 MH080001-001. The Genomic Psychiatric Cohort wishes to acknowledge all of the research participants in this cohort; the study was supported by U01 MH105641, R01 MH085548, R01MH104964. The MCTFR study was supported through grants from the National Institutes of Health DA037904, DA024417, DA036216, DA05147, AA09367, DA024417, HG007022, and HL117626.

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Supplementary Figures


## Supplementary Figure 1

Densities of $\log _{10}$-scaled 7-mer relative mutation rates, estimated using the downsampled BRIDGES ERVs (red) and BRIDGES MAC10+ variants (blue). P-values from the Kolmogorov-Smirnov test for distributional equivalence are shown in the upper left corner of each panel.

\#G/C bases in flanking region $\circ\langle 4 \Delta\rangle=4$ MAC10+:ERV ratio -〈2 • 2

Supplementary Figure 2 Detailed comparison between ERV-derived and MAC10+-derived A>G transition rates. Points are colored by the ratio between the two rates for that subtype (orange: $M A C 10+: E R V>2$; green: $M A C 10+: E R V<2$ ). The shape of each point indicates the number of $G$ or $C$ bases in the $+/-3$ nucleotides flanking the variant site. Among the 103 7-mer motifs with a MAC10+:ERV ratio $>2,100$ have 4 or more G/C bases in the flanking region. (inset) Sequence logo for these 103 7-mer subtypes with MAC10+:ERV ratio >2 shows flanking regions are enriched for G/C bases.
a.


| - |  |  | $\bullet$ | C>T (non-CpG) | - |  | pG>TpG |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| - | A> |  | - | $\mathrm{C}>\mathrm{A}($ non-CpG) | - |  | pG 7 ApG |
| - | A) |  | - | C>G (non-CpG) | - |  | pG $>\mathrm{GpG}$ |

b.

c.

(a) Relationship between 7-mer relative mutation rates estimated using down-sampled BRIDGES ERVs (x-axis) and variants with a minor allele count >= 10 (MAC10+; y-axis), excluding subtypes with $<50$ variants in either dataset. (b) Type-specific 2D-density plots, as situated in the scatterplot of a. The dashed line indicates an expected least-squares regression line if there is no bias present. (c) Heatmap shows ratio between relative mutation rates calculated on MAC10+ variants and ERVs for each 7-mer mutation subtype. Subtypes with higher MAC10+-derived rates relative to ERV-derived rates are shaded gold, and subtypes with lower MAC10+-derived rates relative to ERV-derived rates are shaded green.
a.

b.

C.





## Supplementary Figure 4

(a) Relationship between 7-mer relative mutation rates estimated using BRIDGES variants with a minor allele count >= 10 (MAC10+; x-axis), and 7-mer rates calculated from intergenic variants in the European 1000G phase I sample (y-axis) (b) Type-specific 2D-density plots, as situated in the scatterplot of $\mathbf{a}$. The dashed line indicates an expected least-squares regression line if there is no bias present. (c) Heatmap shows ratio between relative mutation rates calculated on MAC10+ variants and 1000G variants for each 7-mer mutation subtype. Subtypes with higher 1000G-derived rates relative to MAC10+-derived rates are shaded gold, and subtypes with lower 1000G-derived rates relative to MAC10+-derived rates are shaded green. 1000G-derived rates shown here are scaled relative to the MAC10+-derived rates.


Supplementary Figure 5 Distributions of effect sizes (including non-significant effects) on mutability for the 14 genomic features considered in the logistic regression model. For each feature, we plotted the empirical distributions of these subtype-specific odds ratios for each basic mutation type.
*Replication timing is coded with negative values indicating later replicating regions, so an OR<1 means mutation rate increases in late-replicating regions. Note that effects in CpG islands are shown on a wider scale than other features.
a.

b.


Supplementary Figure 6 Comparison of variance explained by all models for (a) all mutation types combined, and (b) stratified by mutation type.

## Supplementary Tables

| Partition | \# Singletons | Ts/Tv ratio | \%dbSNP (b142) | \% of Full Set |
| :---: | :---: | :---: | :---: | :---: |
| Full Set | 36,087,319 | 2.02 | 17.4 | 100 |
| Filter 1 (QUAL>=30)* | 20,796,900 | 2.03 | 17.8 | 58 |
| Filter 2 (MQ>56) | 33,550,098 | 2.01 | 17.3 | 93 |
| Filter 3 (passed 1000G strict mask) | 28,958,837 | 1.94 | 17.5 | 80 |
| All Filters (MQ>56, QUAL>=30, 1000G strict mask) | 16,535,856 | 2.00 | 17.6 | 46 |

## Supplementary Tables 2a-2d Relative mutation rate estimates for 1-mers, 3-mers, 5-mers, and 7-mers <br> [see separate spreadsheet, table_S2_K-mer_relative_rates.xlsx]

Each table contains data used to calculate relative mutation rates for K-mers of a given length. Each row in the table contains the following columns: 1) basic mutation type; 2) K-mer motif corresponding to a reference base $A$ or $C$ at the central mutated position (the reverse complement of each motif, corresponding to reference base T or G is given in parentheses); 3) number of singletons observed in the BRIDGES data of the K-mer subtype defined by columns 1 and $2 ; 4$ ) total number of times the motif in column 2 is observed in the reference genome; 5) relative mutation rate of singletons in that subtype (column 3 divided by column 4). For 7-mer subtypes (Supplementary Table 2d), we include four additional columns: 6) number of singletons in that subtype, after downsampling to 12M; 7) relative mutation rate of downsampled singletons in that subtype (column 6 divided by column 4); 8) number of MAC10+ variants observed in the BRIDGES data of that subtype; 9) relative mutation rate of polymorphisms of that subtype (column 8 divided by column 4).

Supplementary Table 3a Summary of overall model fit statistics for de novo testing data

| Model | Nagelkerke's <br> $\mathbf{R}^{2}$ | AIC | P-value <br> (likelihood <br> ratio test) |
| :--- | :---: | :---: | :---: |
| 1-mers | $0.082(0.082)$ | $326076(326076)$ | -- |
| 3-mers | $0.136(0.111)$ | $309619(317089)$ | $<2.2 \mathrm{e}-308$ |
| 5-mers | $0.143(0.117)$ | $307405(315331)$ | $<2.2 \mathrm{e}-308$ |
| 7-mers | $0.145(0.119)$ | $306738(314705)$ | $1.56 \mathrm{e}-148$ |
| 7-mers+features <br> 7-mers (downsampled | $0.147(0.119)$ | $306146(314943)$ | $3.08 \mathrm{e}-147$ |
| BRIDGES ERVs) | 0.119 | 314723 | -- |
| 7-mers (BRIDGES MAC10+ <br> variants) <br> 7-mers (intergenic 1000G <br> polymorphisms) | 0.114 | 316400 | -- |

Due to the nested structure of the first 5 models in this table (described in Materials and Methods), Nagelkerke's $\mathrm{R}^{2}$ is slightly biased upwards for models with more parameters. For a more direct comparison with the other 3 models, we repeated each these models with only 1 composite predictor (as was done for the downsampled ERV, MAC10+, and 1KG polymorphism models), and we include Nagelkerke's $\mathrm{R}^{2}$ and AIC values for these models in parentheses. Note that the relative differences in Nagelkerke's $R^{2}$ between non-nested K-mer and ( $\mathrm{K}+2$ )-mer models are nearly identical to what we observe in the nested modeling framework. Also, because all models are applied to the same testing data, AIC is a valid means of comparison between all models, regardless of number of predictors; the nested 7 -mer+features model achieves the lowest AIC, indicating this model provides the best overall fit. The last column of P -values come from likelihood ratio test between each nested model and the corresponding model in the preceding row, where such nested models exist.

Supplementary Table 3b Summary of type-specific model fit statistics for de novo testing data. Each type is shown in a sub-table, with the number of de novo mutations and non-mutated sites used in the partitioned testing data indicated in the subheading.

A>C (2920 de novo mutations; 198481 non-mutated sites)

| Model | Nagelkerke's R | AIC | P-value <br> (likelihood <br> ratio test) |
| :--- | :---: | :---: | :---: |
| 3-mers | 0.0023 | 30447 | -- |
| 5-mers | 0.0072 | 30309 | $4.20 \mathrm{e}-32$ |
| 7-mers | 0.0094 | 30248 | $1.94 \mathrm{e}-15$ |
| 7-mers+features | 0.0095 | 30249 | 0.385 |
| 7-mers (downsampled BRIDGES ERVs) | 0.0079 | 30288 | -- |
| 7-mers (BRIDGES MAC10+ variants) | 0.0035 | 30413 | -- |
| 7-mers (intergenic 1000G polymorphisms) | - | 0.0043 | 30388 |

$A>G$ (11400 de novo mutations; 198793 non-mutated sites)

| Model | Nagelkerke's <br> $\mathbf{R}^{2}$ | AIC | P-value <br> (likelinood <br> ratio test) |
| :--- | :---: | :---: | :---: |
| 3-mers | 0.037 | 85999 | -- |
| 5-mers | 0.063 | 84087 | $<2.2 \mathrm{e}-308$ |
| 7-mers | 0.066 | 83829 | $2.22 \mathrm{e}-58$ |
| 7-mers+features | 0.067 | 83792 | $3.68 \mathrm{e}-10$ |
| 7-mers (downsampled BRIDGES ERVs) | 0.063 | 84065 | -- |
| 7-mers (BRIDGES MAC10+ variants) | 0.060 | 84244 | -- |
| 7-mers (intergenic 1000G polymorphisms) ${ }^{8}$ | 0.060 | 84278 | -- |

A>T (2455 de novo mutations; 198320 non-mutated sites)

| Model | Nagelkerke's R² | AIC | P-value <br> (likelihood <br> ratio test) |
| :--- | :---: | :---: | :---: |
| 3-mers | 0.015 | 26123 | -- |
| 5-mers | 0.016 | 26103 | $3.27 \mathrm{e}-06$ |
| 7-mers | 0.017 | 26092 | $3.16 \mathrm{e}-04$ |
| 7-mers+features | 0.017 | 26090 | 0.038 |
| 7-mers (downsampled BRIDGES ERVs) | 0.008 | 26307 | -- |
| 7-mers (BRIDGES MAC10+ variants) | 0.001 | 26466 | -- |
| 7-mers (intergenic 1000G polymorphisms) | 0.002 | 26440 | -- |

non-CpG C>A (3620 de novo mutations; 128765 non-mutated sites)

| Model | Nagelkerke's R | AIC | P-value <br> (likelihood <br> ratio test) |
| :--- | :---: | :---: | :---: |
| 3-mers | 0.011 | 32901 | -- |
| 5-mers | 0.021 | 32606 | $9.73 \mathrm{e}-67$ |
| 7-mers | 0.030 | 32340 | $3.75 \mathrm{e}-60$ |
| 7-mers+features | 0.032 | 32290 | $4.60 \mathrm{e}-13$ |
| 7-mers (downsampled BRIDGES ERVs) | 0.030 | 32351 | -- |
| 7-mers (BRIDGES MAC10+ variants) | 0.024 | 32523 | -- |
| 7-mers (intergenic 1000G polymorphisms) |  |  |  |

non-CpG C>G (3561 de novo mutations; 128746 non-mutated sites)

| Model | Nagelkerke's R | AIC | P-value <br> (likelihood <br> ratio test) |
| :--- | :---: | :---: | :---: |
| 3-mers | 0.006 | 32603 | -- |
| 5-mers | 0.018 | 32271 | $1.25 \mathrm{e}-75$ |
| 7-mers | 0.023 | 32102 | $4.79 \mathrm{e}-39$ |
| 7-mers+features | 0.024 | 32093 | $1.10 \mathrm{e}-03$ |
| 7-mers (downsampled BRIDGES ERVs) | 0.022 | 32127 | -- |
| 7-mers (BRIDGES MAC10+ variants) | 0.018 | 32251 | -- |
| 7-mers (intergenic 1000G polymorphisms) |  |  |  |

non-CpG C>T (10321 de novo mutations; 128774 non-mutated sites)

| Model | Nagelkerke's R² | AIC | P-value <br> (likelihood <br> ratio test) |
| :--- | :---: | :---: | :---: |
| 3-mers | 0.006 | 73240 | -- |
| 5-mers | 0.012 | 72902 | $5.40 \mathrm{e}-76$ |
| 7-mers | 0.014 | 72771 | $9.49 \mathrm{e}-31$ |
| 7-mers+features | 0.015 | 72728 | $1.92 \mathrm{e}-11$ |
| 7-mers (downsampled BRIDGES ERVs) | 0.014 | 72779 | -- |
| 7-mers (BRIDGES MAC10+ variants) | 0.012 | 72905 | -- |
| 7-mers (intergenic 1000G polymorphisms) |  |  |  |

CpG>ApG (304 de novo mutations; 6108 non-mutated sites)

| Model | Nagelkerke's R² | AIC | P-value <br> (likelihood <br> ratio test) |
| :--- | :---: | :---: | :---: |
| 3-mers | 0.013 | 2424 | -- |
| 5-mers | 0.027 | 2397 | $5.82 \mathrm{e}-08$ |
| 7-mers | 0.029 | 2394 | $3.43 \mathrm{e}-02$ |
| 7-mers+features | 0.030 | 2395 | 0.18 |
| 7-mers (downsampled BRIDGES ERVs) | 0.027 | 2395 | -- |
| 7-mers (BRIDGES MAC10+ variants) | 0.020 | 2408 | -- |
| 7-mers (intergenic 1000G polymorphisms) | 0.024 | 2400 | -- |

CpG>GpG (270 de novo mutations; 6292 non-mutated sites)

| Model | Nagelkerke's R | AIC | P-value <br> (likelihood <br> ratio test) |
| :--- | :---: | :---: | :---: |
| 3-mers | 0.010 | 2218 | -- |
| 5-mers | 0.013 | 2216 | 0.037 |
| 7-mers | 0.021 | 2203 | $9.90 \mathrm{e}-05$ |
| 7-mers+features | 0.022 | 2202 | 0.083 |
| 7-mers (downsampled BRIDGES ERVs) | 0.015 | 2208 | -- |
| 7-mers (BRIDGES MAC10+ variants) | 0.015 | 2208 | -- |
| 7-mers (intergenic 1000G polymorphisms) |  |  |  |

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| CpG>TpG (6960 de novo mutations; 6289 non-mutated sites) |  |  |  |
| :--- | :---: | :---: | :---: |
| Model | Nagelkerke's R | AIC | P-value <br> (likelihood <br> ratio test) |
| 3-mers | 0.010 | 18067 | -- |
| 5-mers | 0.017 | 18005 | $8.36 \mathrm{e}-16$ |
| 7-mers | 0.027 | 17900 | $7.45 \mathrm{e}-25$ |
| 7-mers+features | 0.075 | 17415 | $7.23 \mathrm{e}-108$ |
| 7-mers (downsampled BRIDGES ERVs) | 0.026 | 17905 | -- |
| 7-mers (BRIDGES MAC10+ variants) | 0.024 | 17928 | -- |
| 7-mers (intergenic 1000G polymorphisms) ${ }^{8}$ | 0.024 | 17935 | -- |


| Type | Mean MAC10+/ERV ratio ( $\leq 3$ C/G bases) | Mean MAC10+/ERV ratio ( $\geq 4$ C/G bases) | P-value |
| :---: | :---: | :---: | :---: |
| A $>C$ | 0.97 | 1.12 | $8.00 \mathrm{e}-30$ |
| $A>G$ | 1.00 | 1.28 | 2.37e-161 |
| A $>$ T | 0.89 | 0.89 | 0.81 |
| $C>A(n o n-C p G)$ | 0.76 | 0.72 | $2.61 \mathrm{e}-09$ |
| $C>G($ non-CpG) | 0.89 | 0.93 | $2.98 \mathrm{e}-04$ |
| C>T (non-CpG) | 0.93 | 0.85 | $1.75 \mathrm{e}-39$ |
| CpG>ApG | 1.15 | 0.96 | 4.97e-22 |
| CpG>GpG | 1.46 | 1.33 | $2.80 \mathrm{e}-04$ |
| CpG>TpG | 1.02 | 0.98 | 1.01e-09 |

For each mutation subtype, we calculated the ratio between MAC10+-derived and ERV-derived relative mutation rates. Then, for each of the 9 basic types, we grouped 7-mer subtypes into low C/G subtypes ( $\leq 3 \mathrm{C} / \mathrm{G}$ bases in the +/-3 flanking positions) and high C/G subtypes ( $\geq 4 \mathrm{C} / \mathrm{G}$ bases in the +/-3 flanking positions) and performed t-tests for differences in the mean MAC10+/ERV ratios of these two groups.

Supplementary Table 5 Genomic features used in mutation models

| Feature | Source | Cell Type | Resolution |
| :--- | :--- | :--- | :--- |
| H3K4me1, H3K4me3, <br> H3K9ac, H3K9me3, <br> H3K27ac, H3K27me3, <br> H3K36me3 | Roadmap Epigenomics <br> Project $^{15}$ | Peripheral Blood <br> Mononuclear <br> Primary Cells | 1bp (inside vs. outside <br> of broad peak) |
| Replication timing | Koren et al., 2012 |  |  |

A script to download the exact external data files used in this paper is available at https://github.com/carjed/smaug-genetics

Supplementary Table 6a Univariate tests for enrichment or depletion of de novo mutations occurring in feature-associated subtypes identified by logistic regression models.

| Feature | Expected direction of effect | \# mutations in featureassociated subtypes ${ }^{\text {a }}$ | \# mutations in feature |  | p-value | Consistent direction? |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | expected ${ }^{\text {b }}$ | observed ${ }^{\text {c }}$ |  |  |
| CpG Islands | Increased | 648 | 17 | 14 | 0.65 | No |
|  | Decreased | 7072 | 331 | 84 | $7.5 \times 10^{-35}$ | Yes |
| GC content | Increased | 256 | 22 | 8 | 0.99 | No |
|  | Decreased | 2350 | 56 | 23 | $1.1 \times 10^{-4}$ | Yes |
| H3K36me3 | Increased | 3731 | 589 | 682 | $2.4 \times 10^{-3}$ | Yes |
|  | Decreased | 898 | 162 | 98 | $1.3 \times 10^{-5}$ | Yes |
| H3K9me3 | Increased | 7361 | 1165 | 1905 | $3.8 \times 10^{-51}$ | Yes |
| H3K27me3 | Increased | 896 | 274 | 252 | 0.86 | No |
| H3K4me1 | Decreased | 2839 | 557 | 463 | $1.2 \times 10^{-3}$ | Yes |
| H3K4me3 | Decreased | 3406 | 566 | 487 | $8.1 \times 10^{-3}$ | Yes |
| DHS | Increased | 1091 | 177 | 184 | 0.38 | Yes |
|  | Decreased | 2898 | 701 | 645 | 0.04 | Yes |
| Laminassociated domains | Increased | 485 | 187 | 171 | 0.85 | No |
| Recombination rate | Increased | 2190 | 306 | 377 | $1.9 \times 10^{-3}$ | Yes |
| Replication timing | Increased | 2359 | 278 | 321 | 0.03 | Yes |

For each feature for a given effect direction, we included all 7-mer subtypes with significant association of the feature with relative mutation rate (Fig. 5). We counted: athe total number of de novo mutations of those subtypes, bthe number of these mutations expected to occur in the feature under the null expectation that the feature has no impact on mutability (i.e., assuming only an effect of sequence context), and 'the number of these mutations that were observed to occur in the feature. Significant associations are indicated by a one-sided $p$-value (observed numbers are consistent with model predictions) in bold.

Supplementary Table 6b Univariate tests for enrichment or depletion of de novo mutations occurring in feature-associated subtypes (excluding CpG subtypes) identified by logistic regression models.

| Feature | Expected <br> direction of <br> effect | \# mutations in <br> feature- <br> associated <br> subtypes $^{\text {a }}$ | \# mutations in feature |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | expected $^{\text {b }}$ | observed $^{\text {c }}$ |  |  |  |  |

For each feature, we identified all 7-mer subtypes where our model estimated a significant association (Fig. 3), separated into either an increased or decreased direction of effect, and counted: athe total number of non-CpG de novo mutations of those subtypes, ${ }^{\text {b }}$ the number of these mutations that would occur in regions of the genome where that feature was present, under the null expectation that the feature has no impact on mutability (i.e., assuming only an effect of sequence context), and 'the number of these mutations that were observed in the presence of that feature. Significant associations are indicated by a one-sided p-value in bold. Note that only 5 of the 15 groups described in Supplementary Table 6a contained sufficient numbers of non-CpG de novo mutations to perform these tests.

Supplementary Table 7 Parameter estimates for genomic features model
[see separate spreadsheet, table_S7_feature_parameter_estimates.xlsx]
This table contains effect size estimates and standard errors of 16 parameters ( 14 features, plus intercept and read depth) for each of the 24,489 7-mer subtypes with at least 10 singletons in the BRIDGES data.

