

Integrated Bayesian analysis of rare exonic variants to identify risk genes for schizophrenia and neurodevelopmental disorders

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1 Additional file 1 — Supplementary Information

This file (AdditionalFile1.pdf) describes supplementary results, methods, data, figures and short Tables.

1.1 Supplementary Results

1.1.0.1 Simulation case-control data only

To evaluate the performance of the approximate CC model for different parameter values, we simulated a single CC sample with either one or two variant/annotation classes. We tested sample sizes ranging from that of the available data, 1,092 each cases and controls (ASD), and 3,157 cases and controls (SCZ), to larger sample sizes of 10,000 cases and controls, and 20,000 cases and controls.

Overall, high correlations (~ 1) between estimated and simulated parameter values indicate little bias in inference based on CC data (Figure S4 and S6). Slight over estimation was observed for the sample size of 1092, especially for risk-gene proportions.

An additional analysis was carried out to assess the performance of specific simulated values. Correlations were calculated for each mean RR and π value. For one CC class, mean RRs were estimated well by the model with correlations ~ 1 (Figure S5). However, the proportion of risk genes was affected by mean RRs. They were estimated well when mean RRs were between 1.5 and 3.5, but underestimated with smaller mean RRs and slightly overestimated with larger mean RRs (Figure S5). For two CC classes, high correlations (≥ 0.97) between simulated and estimated values were seen for all parameters. In addition, small mean RRs of a given class did not directly affect the estimated values of proportions of risk genes (Figure S7).

The issue of poor estimation for one class, but good estimation for > 1 class was expected. This was an advantage of using multiple classes compared to using only one class in the estimation process when the clustering signal was not very strong. Small mean RRs could result in difficulties in the calculation process to differentiate between a risk gene (mean RR > 1) and a non-risk gene (mean RR ~ 1). If one class was used then many risk genes would be considered to be non-risk genes. If more than one class was used, such risk genes would be assigned as genuine risk genes due to the information available from other classes.

1.2 Supplementary methods

1.2.1 Analysis of SCZ data

1.2.1.1 Obtaining non-heterogeneous population samples for case-control data of SCZ

The case-control data sets were divided into three big populations: Finland, United Kingdom and Sweden. For the Sweden population, this was a large data set and was also sequenced at different centers [Genovese et al. \(2016\)](#), therefore we divided this population as follows.

A simple combination between a clustering process using a multivariate normal mixture model and a data analyzing strategy using linear and generalized linear models was used to divide the Sweden data into non-heterogeneous populations. [Genovese et al. \(2016\)](#) recently analyzed all case-control data sets by adjusting for multiple covariates: genotype gender of individuals (SEX), 20 principal components (PCs), year of birth of individuals (BIRTH), Aligent kit used in wet-labs (KIT) by using linear regression and generalized linear regression models as in Equation 1. They reported significant results for NoExAC LoF and MiD variants; therefore, this information was used in this step. We defined homogeneous populations as populations which were not much affected by the covariates. Thus, for the populations, analyzing results using Equation

1 (adjusting covariates) would not be much different from those results using Equation 2 (not adjusting covariates). The `mclust` package Version 5.2 [Fraley and Raftery \(1999\)](#) which uses a multivariate normal mixture model was used to divide 11,161 samples (4,929 cases and 6,232 controls) into different groups. To see all situations of the grouping process, we used `mclust` with three strategies on 11,161 samples: grouping all 20 PCs, grouping all 20 PCs and total counts, and grouping only the first three PCs. The number of groups were set between 2 and 6. For each clustering time, Equation 1 and 2 were used to calculate p values for each variant category of each group from the clustering results (p1 and p2 respectively); then, Spearman correlation [Spearman \(1904\)](#) between p-value results from the two Equations (cPvalue) was calculated. Next, to filter reliable results from the clustering process, we set criteria:

- cPvalue ≥ 0.85 and p-values for NoExAC ≤ 0.005 .
- Ratio p1/p2 from Equation 1 and 2 had to be between 0.1 and 1.

From results satisfied the above criteria, we manually chose groups which had similar results between Equation 2 and 1.

$$\begin{aligned} \text{logit}(P(SCZ = 1)) &\sim \text{count} + \text{countAll} + \text{sex} + \text{birth} + \text{kit} + \sum_{i=1}^{20} PC_i \\ \text{count} &\sim SCZ + \text{countAll} + \text{sex} + \text{birth} + \text{kit} + \sum_{i=1}^{20} PC_i \end{aligned} \quad (1)$$

$$\begin{aligned} SCZ &\sim \text{count} \\ \text{count} &\sim SCZ \end{aligned} \quad (2)$$

For the data from the UK10K project [Singh et al. \(2016\)](#), we divided the data into two separate populations England and Finland, and tested NoExAC variants in these populations by calculating sample-size-adjusted ratios between cases and controls. The ratios were 0.91 and 0.95 for the UK data. Regarding the Finland data, the ratio for MiD variants was only 0.41 which were extremely low. This could be a special case for the population or might be because of other technical reasons. We did not use this population in the next stage because it showed a different trend with other populations.

1.2.2 extTADA pipeline: extended transmission (case-control) and de novo analysis

This section describes more details the pipeline.

1.2.2.1 extTADA for one *de novo* population and one case/control population

extTADA is summarized in Table S3 and Figure S2. There, $x_{dn} \sim \text{Pois}(2N_d\mu, \gamma_{dn})$, $x_{ca} \sim \text{Pois}(qN_1\gamma_{cc})$, $x_{cn} \sim \text{Pois}(qN_0)$, and $\gamma_{dn} \sim \text{Gamma}(\tilde{\gamma}_{dn}\beta_{dn}, \beta_{dn})$, $\gamma_{cc} \sim \text{Gamma}(\tilde{\gamma}_{cc}\beta_{cc}, \beta_{cc})$, $q \sim \text{Gamma}(\rho, \nu)$.

Let K be the number of categories (e.g., LoF, MiD), and $x_i = (x_{i1}, \dots, x_{iK})$ be the vector of counts at the i^{th} given gene. The Bayes Factor for each j^{th} category to test two hypotheses: $H_0 : \gamma = 1$ versus $H_1 : \gamma \neq 1$ was:

$$\begin{aligned} B_{ij} &= \frac{P(x_{ij}|H_1)}{P(x_{ij}|H_0)} \\ &= \frac{\int P(x_{ij}|\gamma, q)P(q|H_1)P(\gamma|H_1)dq d\gamma}{\int P(x_{ij}|\gamma, q)P(q|H_0)P(\gamma|H_0)dq d\gamma} \\ &\quad \text{Because } \gamma = 1 \text{ for } H_0 \\ &= \frac{\int P(x_{ij}|\gamma, q)P(q|H_1)P(\gamma|H_1)dq d\gamma}{\int P(x_{ij}|q)P(q|H_0)dq} \end{aligned} \quad (3)$$

In Equation 3, $x_{ij} = x_{dn}$ for *de novo* data and $x_{ij} = (x_{ca}, x_{cn})$ for case-control data. In addition, the integral over q was not applicable for *de novo* data because there is no q parameter for *de novo* data.

As in He et al. (2013), the BF for the i^{th} gene combining all categories is:

$$B_i = \prod_{j=1}^K B_{ij} \quad (4)$$

To calculate BFs, hyper parameters in Table S3 need to be inferred. Let ϕ_{1j} and ϕ_{0j} be hyper-parameters for H_1 and H_0 respectively. A mixture model of the two hypotheses was used to infer parameters using information across the number of tested genes (m) as:

$$P(x|\phi_1, \phi_0) = \prod_{i=1}^m \left[\pi \prod_{j=1}^K P(x_{ij}|\phi_{1j}) + (1 - \pi) \prod_{j=1}^K P(x_{ij}|\phi_{0j}) \right] \quad (5)$$

Equation 5 was calculated across categories as in Equation 4.

We used the same approach for the analysis of multiple population samples. Let Ndn_{pop} , Cdn and Ncc_{pop} , Ccc be the number of populations, categories for *de novo* and case-control data respectively. The total Bayes Factor of a given gene was the product of Bayes Factors of all populations as in the main text, and all hyper parameters were estimated using Equation 2 in the main text.

The hyper-parameters $\phi_{1j} = (\gamma_{j(dn)}, \gamma_{j(cc)}, \beta_{j(dn)}, \beta_{j(cc)}, \rho_j, \nu_j)$ were estimated using a Hamiltonian Monte Carlo (HMC) Markov chain Monte Carlo (MCMC) method implemented in the **rstan** package Carpenter et al. (2015); R Core Team (2016). However, the model was first simplified by removing q (see below).

1.2.2.2 Simplified approximate case-control model

For case-control (transmitted) data, $q \sim \text{Gamma}(\rho, \nu)$, and hyper-parameters ρ and ν controlled the mean and dispersion of q ; therefore, as in the previous studies He et al. (2013); De Rubeis et al. (2014), ν was heuristically chosen (200 was used in all analyses) and $\frac{\rho}{\nu}$ = the mean frequency across genes in both cases and controls.

We simplified the case-control model by expressing it as

$$P(x_{ca}, x_{cn}|H_j) = P(x_{ca}|x_{ca} + x_{cn}, H_j)P(x_{ca} + x_{cn}|H_j) \quad (6)$$

Because $x_{ca} \sim Pois(N_1 q \gamma_{cc})$ and $x_{cn} \sim Pois(N_0 q)$, assuming that x_{ca} and x_{cn} were **independent**, the case data could be modeled as:

$$x_{ca}|x_{ca} + x_{cn}, H_j \sim Binomial(x_{ca} + x_{cn}, \theta|H_j)$$

$$\text{with } \theta|H_1 = \frac{N_1 \gamma_{cc}}{N_1 \gamma_{cc} + N_0} \text{ and } \theta|H_0 = \frac{N_1}{N_1 + N_0}$$

The marginal likelihood was

$$P(x_{ca}|x_{ca} + x_{cn}, H_j) = \int P(x_{ca}|x_{ca} + x_{cn}, \gamma_{cc}, H_j)P(\gamma_{cc}|H_j)d\gamma_{cc}$$

Based on simulation results, the first part $P(x_{ca}|x_{ca} + x_{cn}, H_j)$ can be used to infer mean RRs ($\bar{\gamma}_{cc}$); therefore only this part was used in the **extTADA** estimation process. However, to calculate Bayes Factors, we used full case-control models. We changed the order of integrals (Supplementary Methods).

1.2.2.3 Control of an implied proportion of protective variants using the relative risk dispersion hyper-parameter

If $\bar{\gamma}$ and β were small then we could see a high proportion of protective variants when $\bar{\gamma}$ is not large. Although this might be of biological interest, it is not currently accounted for in the model. To control the proportion of protective variants, we tested the relationship between β and $\bar{\gamma}$ in determining $\int Gamma(\bar{\gamma}_{dn}\beta_{dn}, \beta_{dn})$. We set this proportion very low (2%) (Figure S3) and built a nonlinear relationship $\beta = e^{a*\bar{\gamma}^b + c}$. The function *nls* in R was used to estimate a, b and c, as 6.77, -1.79 and -0.22 respectively.

1.2.2.4 Calculate Bayes Factor for case/control data

At a given gene, Bayes Factor for each class was calculated as $BF = \frac{P(x_1, x_0|H_1)}{P(x_1, x_0|H_0)}$. The probability for each model ($H_j, j = 0, 1$) was calculated in order to rely only γ parameters as follows.

$$P(x_{ca}, x_{cn}|H_j) = P(x_{cn}|H_j)P(x_{ca}|x_{cn}, H_j) \quad (7)$$

- The first part $P(x_{cn}|H_j)$ was the same as [De Rubeis et al. \(2014\)](#):

$$P(x_{cn}|H_j) = \int P(x_{cn}|q, H_j)P(q|\rho, \nu, H_j)dq = NegBin(x_{cn}|\rho, \frac{N_0}{\nu + N_0}), j = 0, 1 \quad (8)$$

- The second part:

$$\begin{aligned} P(x_{ca}|H_j, x_{cn}) &= \int P(x_{ca}|q, \gamma_{cc})P(q|H_j, x_{cn})P(\gamma_{cc}|H_j)dq d\gamma_{cc} \\ &= \int [P(x_{ca}|q, \gamma_{cc})P(q|H_j, x_{cn})dq] P(\gamma_{cc}|H_j)d\gamma_{cc} \\ &= \int NegBin(x_{ca}|\rho + x_{cn}, \frac{N_0 + \nu}{N_1 \gamma_{cc} + N_0 + \nu})P(\gamma_{cc}|H_j)d\gamma_{cc} \end{aligned} \quad (9)$$

To identify the lower and upper limits of γ_{CC} for the integral, we randomly sampled 10,000 times values from the $Gamma(\bar{\gamma}_{cc} * \beta_{cc}, \beta_{cc})$ and used the minimum and maximum values for the lower and upper limits respectively.

1.2.3 Infer parameters using MCMC results

The `rstan` package [Carpenter et al. \(2015\)](#) was used to run MCMC processes. For simulation data, 5,000 times and a single chain were used. For real data, 20,000 times and three independent chains were used. In addition, for SCZ data we used two steps to obtain final results. Firstly, 10,000 times were run to obtain parameters. After that, we calculated β values from estimated mean RRs as the Equation described in Table [S3](#). Finally, `extTADA` was re-run 20,000 times on the SCZ data with calculated β values set as constants to re-estimate mean RRs and the proportions of risk genes. For each MCMC process, a burning period = a half of total running times was used to assure that chains did not rely on their initial values. For example, we ran and removed 2,500 burning times before the 5,000 running times for simulation data.

We just chose 1,000 samples of each chain from MCMC results to do further analyses. For example, with a chain with 20,000 run times, the step to obtain a sample was 20 run times. For all estimated parameters from MCMC chains, the convergence of each parameter was diagnosed using the estimated potential scale reduction statistic (\hat{R}) introduced in `Stan` [Carpenter et al. \(2015\)](#). To produce heatmap plots, modes as well as the credible intervals (CIs) of estimated parameters, the `Locfit Loader` [\(2007\)](#) was used. The mode values were used as our estimated values for other calculations.

1.3 Supplementary Figures

This file includes Sup Figures below.

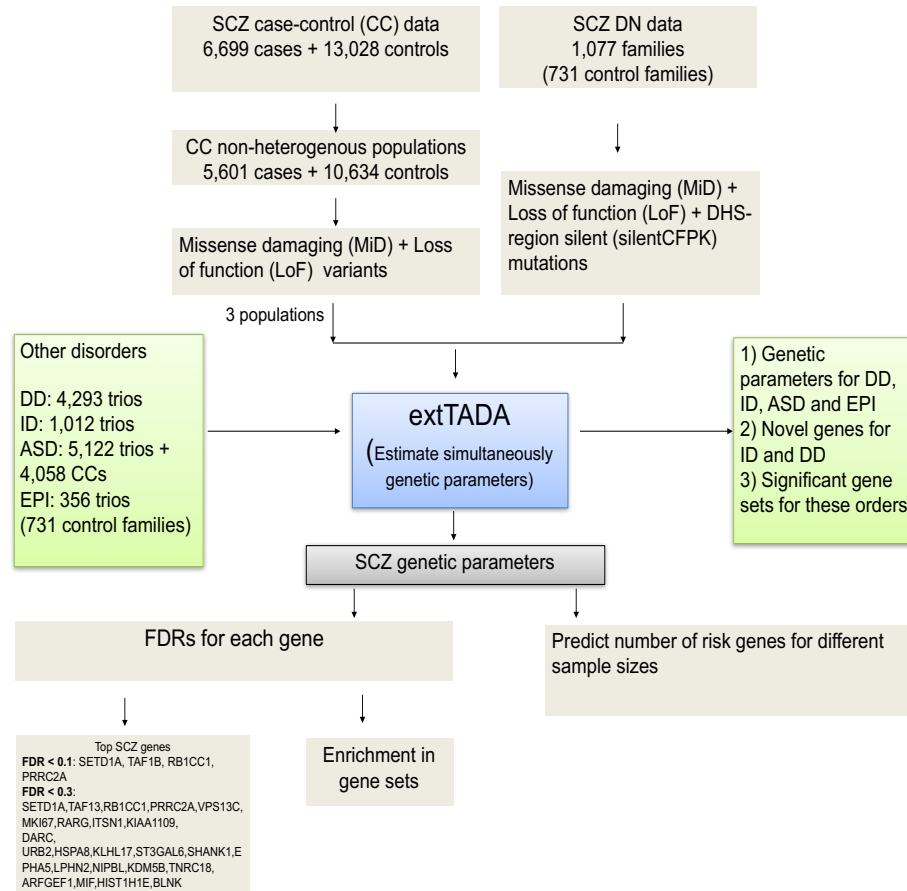


Figure S1: Workflow of data analysis.

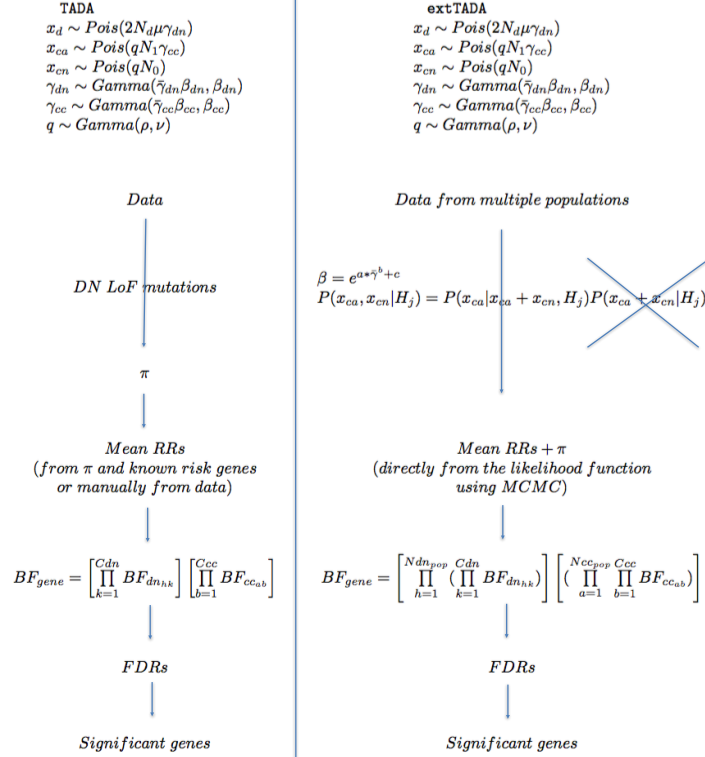


Figure S2: Comparison between **TADA** and **extTADA**. They both use the same model for *de novo* data (x_{dn} and case/control (x_{ca}, x_{cn}) data. **extTADA** combines all categories to obtain parameters and their credible intervals while **TADA** is based on LoF mutations. **extTADA** uses an approximate model for case-control data, and constrains β and γ in the estimation process. **extTADA** is designed to work for multiple populations. **TADA** can be used inside **extTADA**.

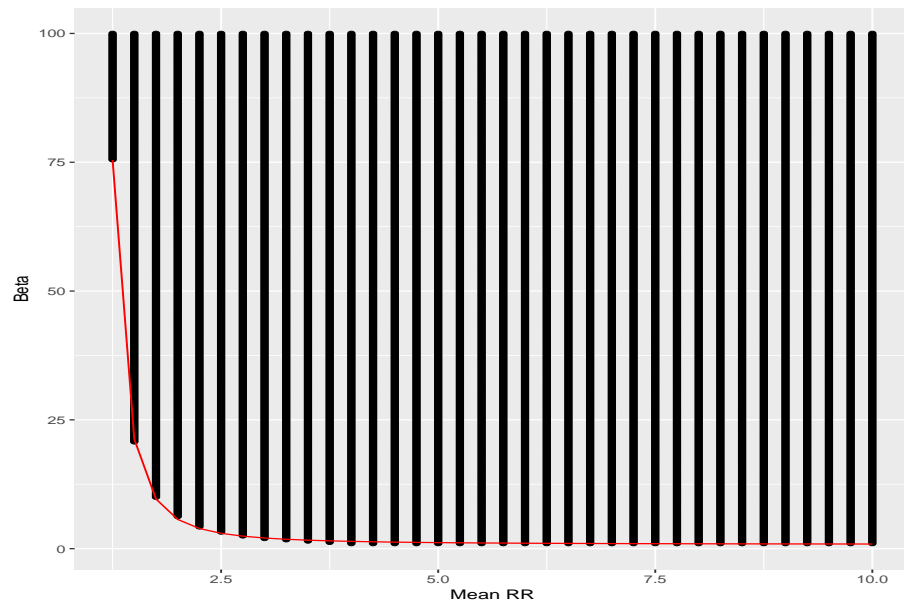


Figure S3: A grid of β and $\bar{\gamma}$ values. Points on the red line are corresponding with the proportion of protective variants less than 2%.

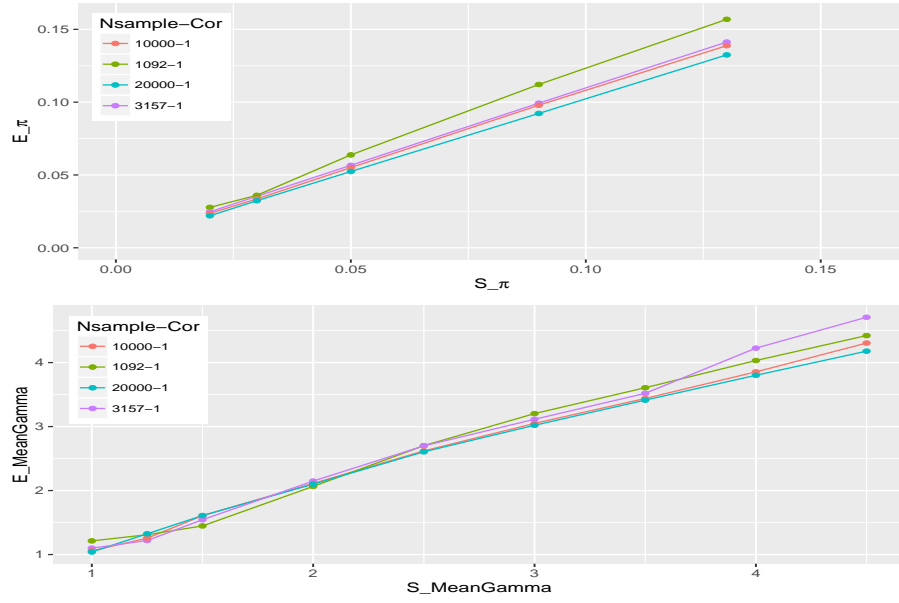


Figure S4: Correlations between estimated and simulated values for one CC class with different sample sizes. X and Y axes describe simulated (S) and estimated (E) values respectively. The top picture is for mean relative risks (MeanRRs) while the bottom picture is for the proportion of risk genes (π). Legends show sample sizes and correlations. These estimated values were averaged across simulation results. Detailed values are presented in Figure S5.

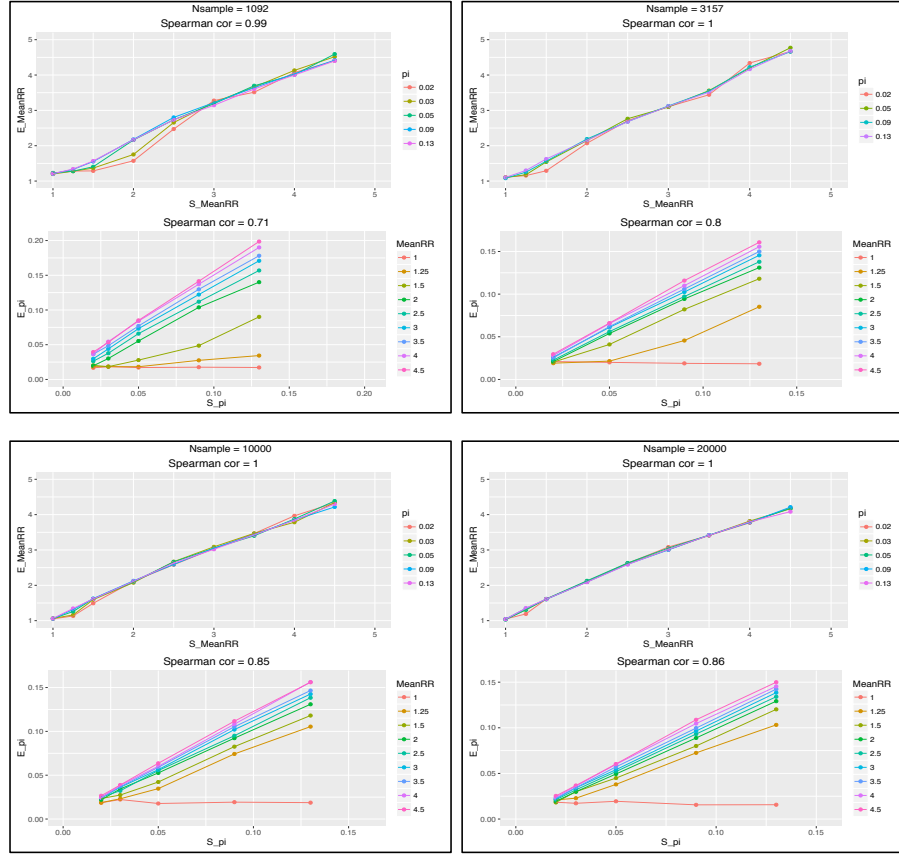


Figure S5: Correlation between simulated and estimated values for one-category case/control data.

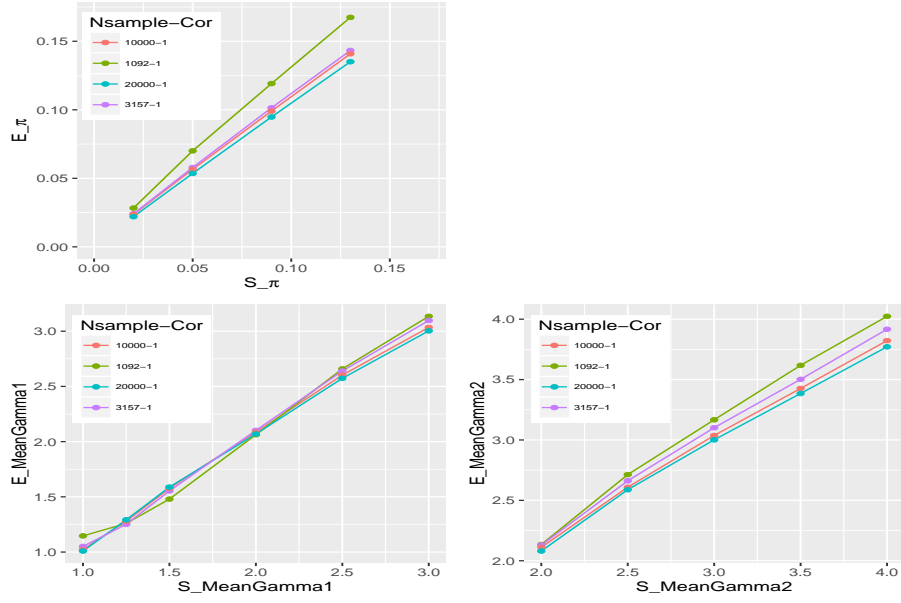


Figure S6: Correlations between estimated and simulated values for two CC class with different sample sizes. X and Y axes describe simulated (S) and estimated (E) values respectively. A range of mean relative risks for two classes (MeanGamma1 and MeanGamma2) and risk-gene proportions (π) were used in the simulation process. Legends show sample sizes and correlations. These estimated values were averaged across simulation results. Detailed values are presented in Figure S7.

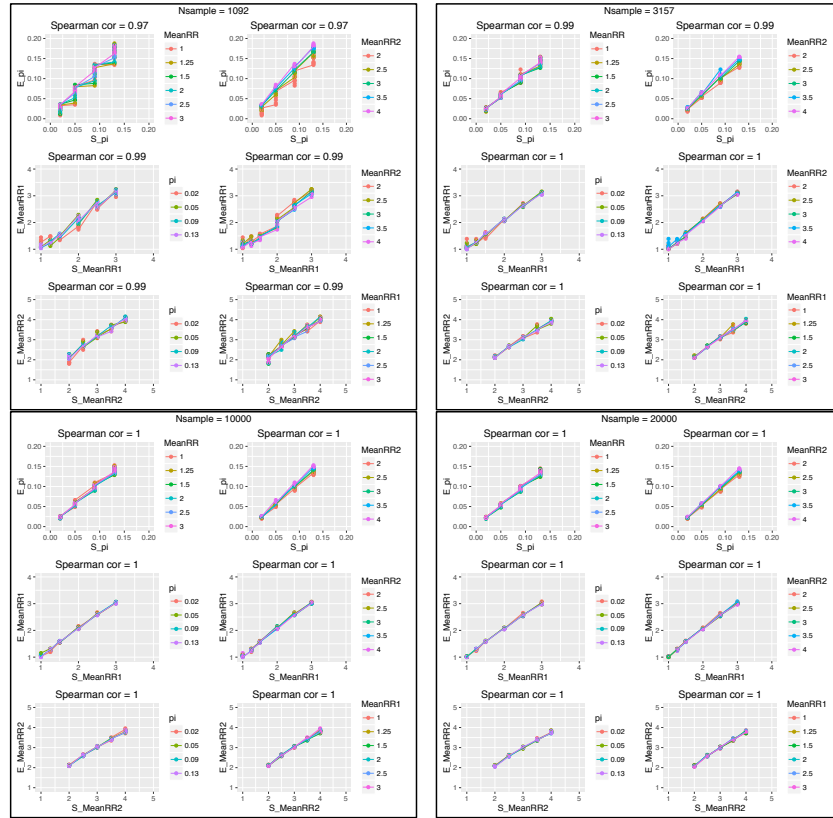


Figure S7: Correlation between simulated and estimated values for two-category case/control data.

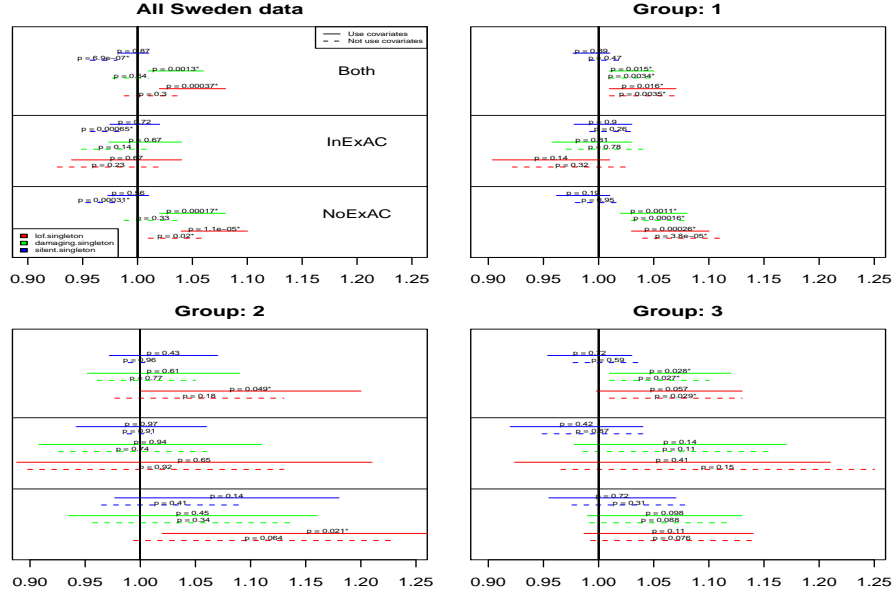


Figure S8: Odds ratios for the analysis of all case-control samples. Top left picture shows odds ratios for all Sweden samples while the three other pictures show odds ratios for three groups after the clustering process. Only group 1 and 3 are used in the current analysis because there are strong differences between results using covariates and not using covariates in group 2. P values were calculated for variants in (InExAC), not in (NoExAC) the ExAC database, and all variants (Both).

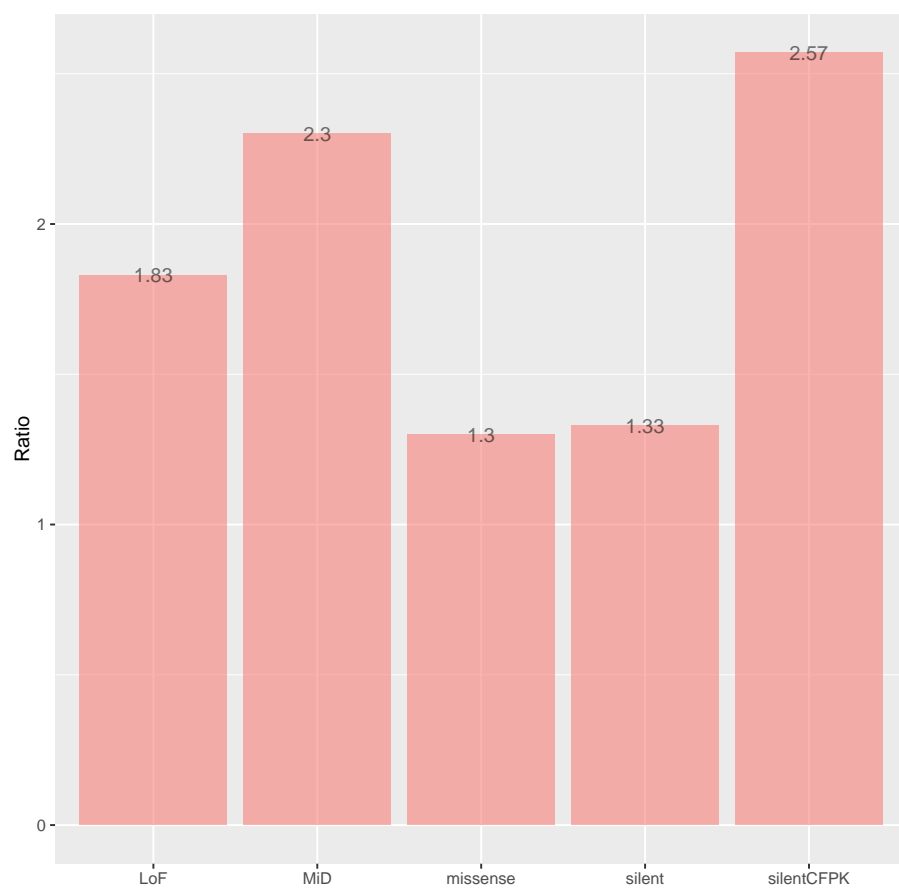


Figure S9: Ratios of *de novo* mutations between SCZ probands and controls (unaffected siblings). "silentFCPk" describes for silent mutations within frontal cortex-derived DHS (silentCerebrumfrontalocPk.narrowPeak). MiD mutations are missense mutations derived from 7 methods.

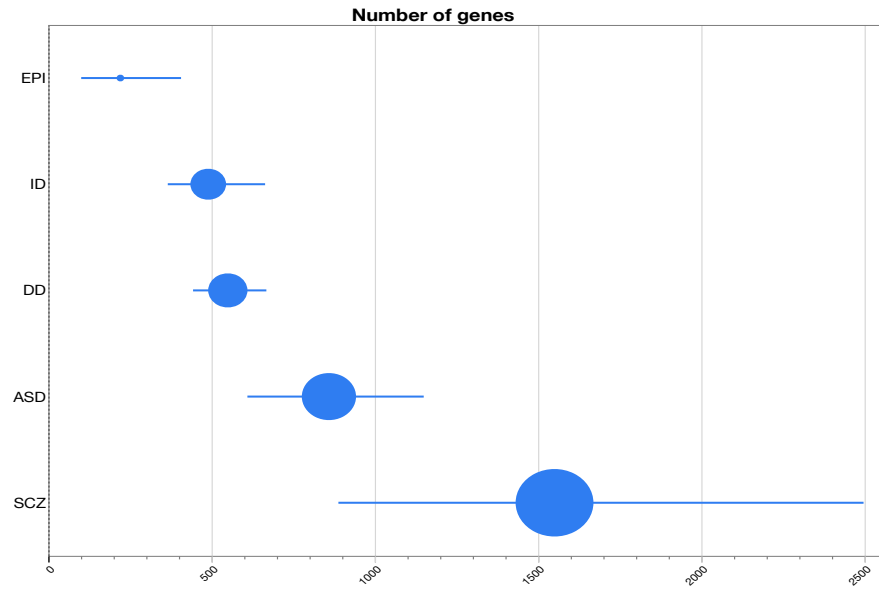


Figure S10: Estimated gene counts for all disorders, with 95% CIs. Point sizes are proportional to sample sizes.

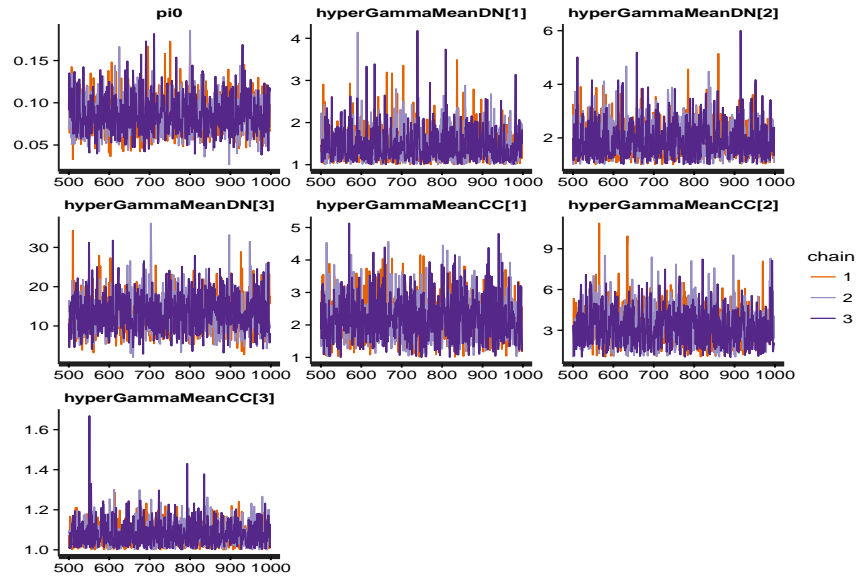


Figure S11: MCMC results for SCZ data.

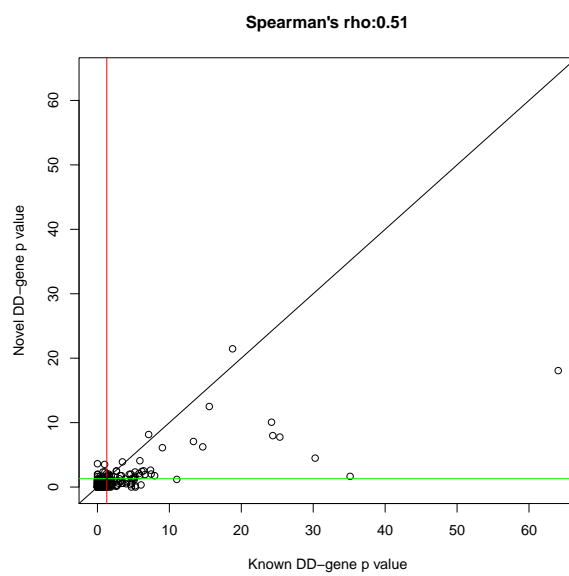


Figure S12: The correlation of gene-set p values (values are $-\log_{10}(\text{p values})$) between known and novel genes from **extTADA** results for DD

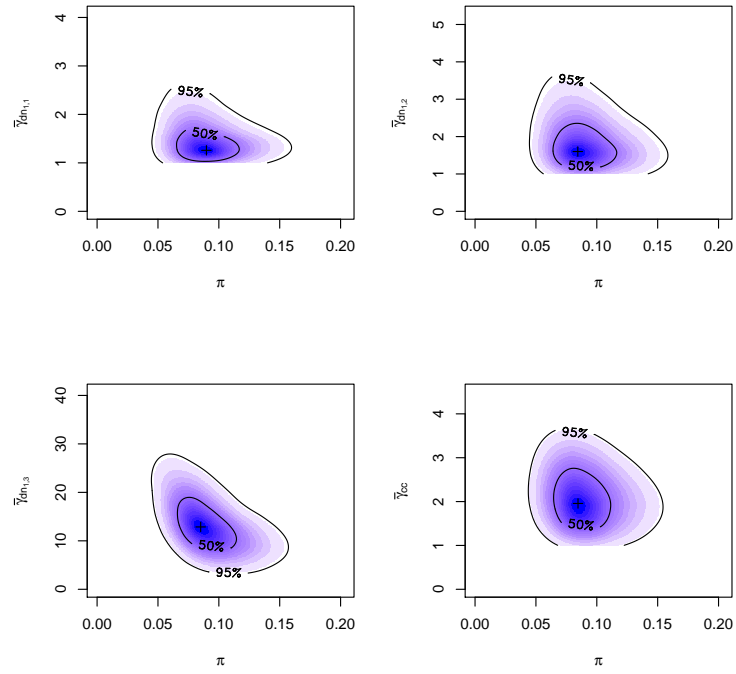


Figure S13: SCZ genetic parameters when mean RRs of case-control data are equal.

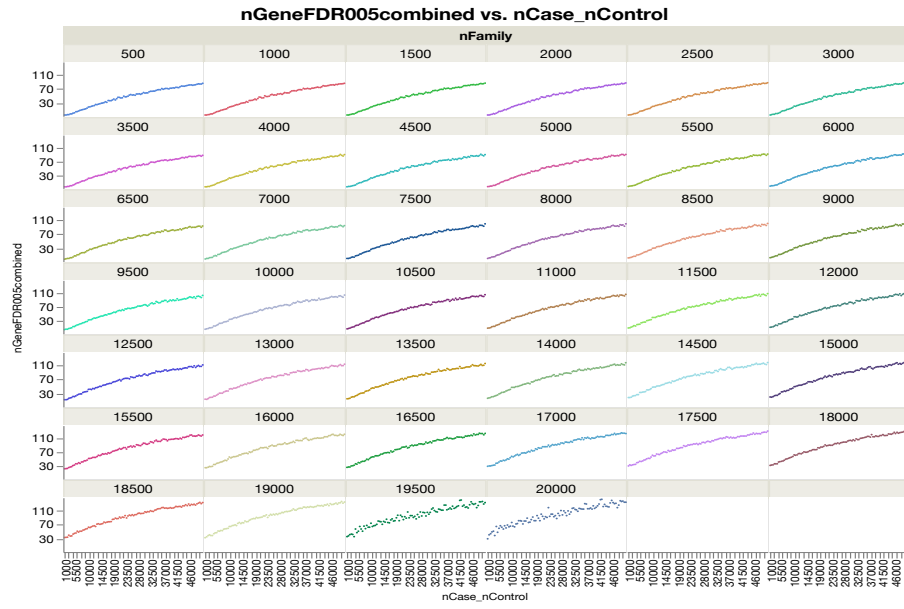


Figure S14: Number of risk genes with different sample sizes based on genetic architecture predicted by extTADA. Case/control number is only for cases (with equal controls) (x-axis); each panel shows results for a given number of trios.

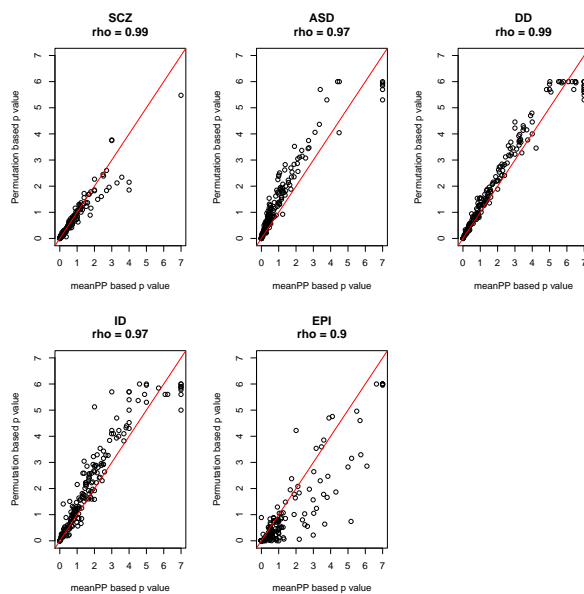


Figure S15: The correlation of gene-set p values ($-\log(\text{p value})$) between mean posterior profitability (meanPP) based method and permutation based methods.

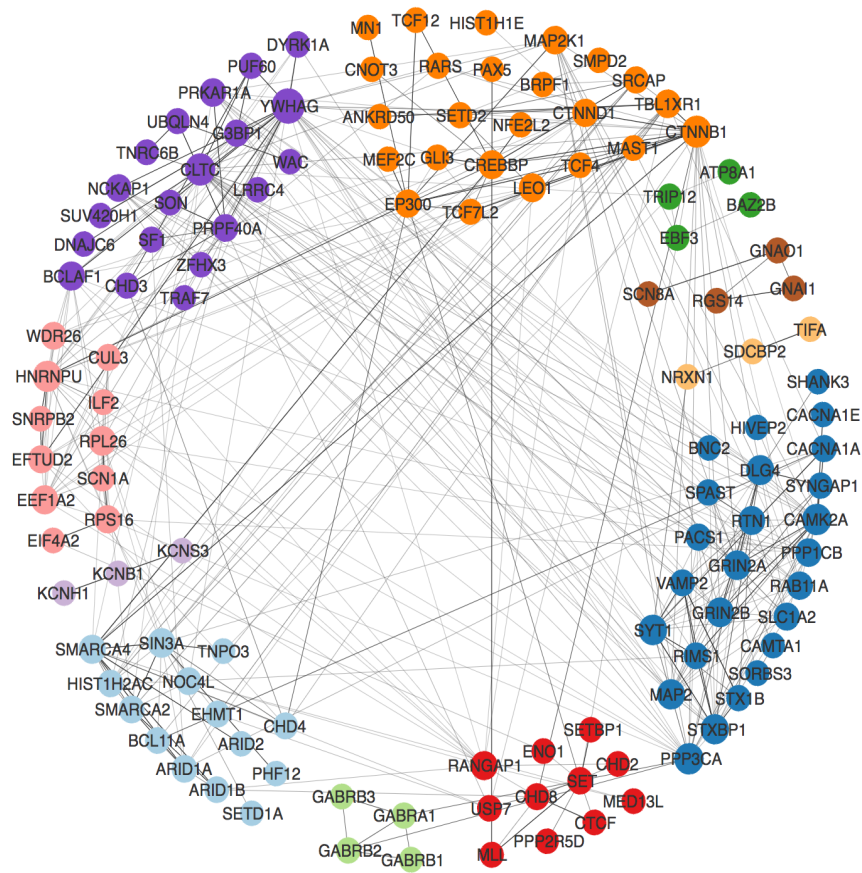


Figure S16: GeNets InWeb PPI network for 288 NDD genes, with direct edges only.

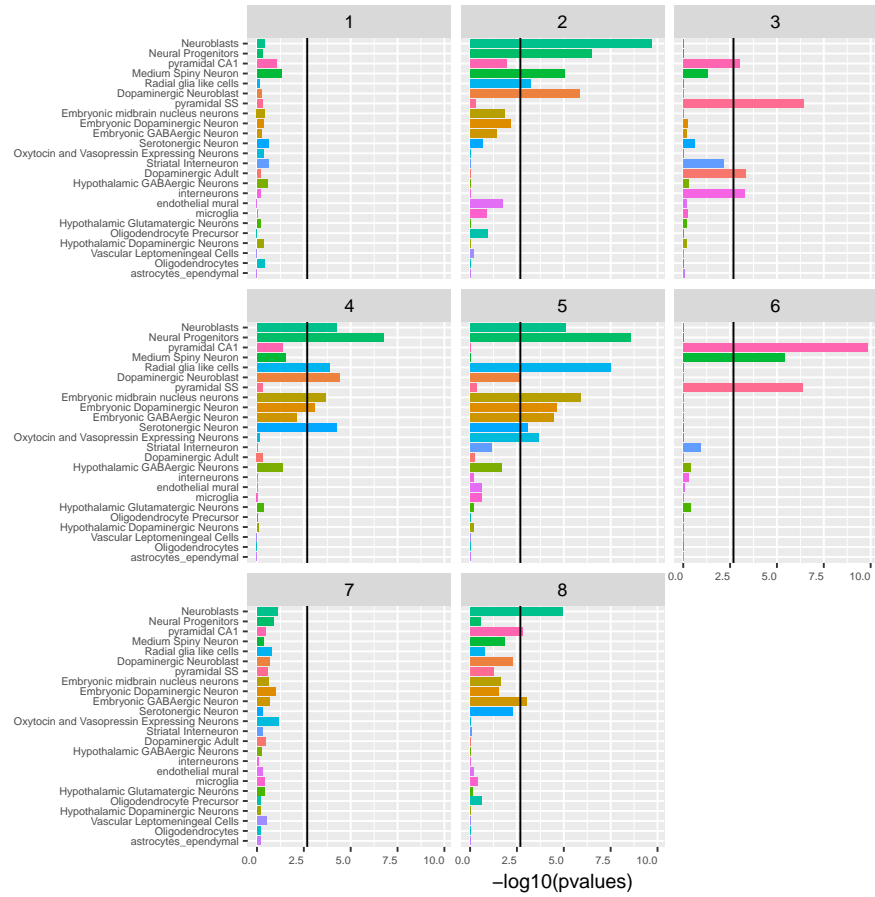


Figure S17: Evaluation of enrichment of each community from GeNets results in brain scRNAseq datasets from mouse.

1.4 Supplementary Tables

This part includes Sup Tables below.

Source	Disease	DN	DN control	Case	Control
Fromer et al. (2014)	SCZ	617	84	4954/4248 1745/1353	6239/5865 6789/4769
Girard et al. (2011)	SCZ	14			
Gulsuner et al. (2013)	SCZ	105			
McCarthy et al. (2014)	SCZ	57			
Xu et al. (2012)	SCZ	231			
Guipponi et al. (2014)	SCZ	53			
Genovese et al. (2016)	SCZ				
Singh et al. (2016)	SCZ				
Deciphering Developmental Disorders Study (2017)	DD	4293			
EuroEPINOMICS-RES Consortium et al. (2014)	EPI	356			
De Ligt et al. (2012)	ID	100	20		
Hamdan et al. (2014)	ID	41			
Rauch et al. (2012)	ID	51			
Lelieveld et al. (2016)	ID	820			
Turner et al. (2016)	ASD	5122	343 50 200	404	3654
De Rubeis et al. (2014)	ASD				
Iossifov et al. (2012)	ASD				
ORoak et al. (2012)	ASD				
Sanders et al. (2012)	ASD				

Table S1: *de novo* and case/control data. For ASD studies, Turner et al. (2016) integrated previous results in their study; therefore only *de novo* meta data in their study are shown in the table. In addition, for ASD case-control data, only one homogeneous Sweden population from De Rubeis et al. (2014) was used. For case-control data of SCZ, after correcting for the population stratification, only 4,248 cases (3,157 + 1,091) + 5,865 (4,672 + 1,193) controls from Genovese et al. (2016) and 1,353 cases + 4,769 controls from Singh et al. (2016) are used in this study.

Gene set name	Abbreviation	Author
Missense constrained genes	constrained	Samocha et al. (2014)
Loss-of-function tolerance genes	pLI90	Lek et al. (2015)
RBFOX2 and RBFOX1/3 genes	rbfox2, rbfox13	Weyn-Vanhentenryck et al. (2014)
FMRP genes	fmrp	Darnell et al. (2011)
CELF4 genes	celf4	Wagnon et al. (2012)
synaptic genes	synaptome	Pirooznia et al. (2012)
microRNA-137	mir137	Robinson et al. (2015)
PSD-95 complex genes	psd95	Bayés et al. (2011)
ARC and NMDA receptors genes	nmdarc	Kirov et al. (2012)
Essential genes	essential	Ji et al. (2016)
Human accelerated regions and primate accelerated regions	HARs, PARS	Lindblad-Toh et al. (2011)
Known ID gene sets	IDallKnownGenes	Lelieveld et al. (2016)
Voltage-gated Calcium Channel Genes	vacc	
CHD8 promoter targets	chd8 hNSC, chd8 hNSC specific, chd8 human brain, chd8 hNSC human brain, chd8 hNSC human mouse	Cotney et al. (2015)
Allelic-biased expression genes in neurons	AlleleBiasedExpression.Neuron	Lin et al. (2012)
24 gene sets from 24 modules	Module.M1..M24	Johnson et al. (2016)
<i>de novo</i> copy number variants		Genovese et al. (2016)
ASD	CNV.denovo.gain/loss.asd	
Bipolar	CNV.denovo.gain/loss.bd	
SCZ	CNV.denovo.gain/loss.scz	
MiD and LoF <i>de novo</i> mutations		
DD	DD.allDenovoMiDandLoF	
ASD	ASD.allDenovoMiDandLoF	
EPI	EPI.allDenovoMiDandLoF	
ID	ID.allDenovoMiDandLoF	

Table S2: Abbreviations of known gene sets used in this study.

<i>Data model</i>	<i>Parameter prior</i>	<i>Hyper prior</i>
$x_{dn} \sim P(2N_{dn}\mu\gamma_{dn})$	$\gamma_{dn} \sim \text{Gamma}(\bar{\gamma}_{dn} * \beta_{dn}, \beta_{dn})$ $\beta_{dn} = e^{a*\bar{\gamma}_{dn}^b + c}$	$\bar{\gamma}_{dn} \sim \text{Gamma}(\bar{\bar{\gamma}}_{dn}, \bar{\bar{\beta}}_{dn})$
$x_{ca} \sim P(N_1 q \gamma_{cc})$	$\gamma_{cc} \sim \text{Gamma}(\bar{\gamma}_{cc} * \beta_{cc}, \beta_{cc})$ $\beta_{cc} = e^{a*\bar{\gamma}_{cc}^b + c}$ $q \sim \text{Gamma}(\rho, \nu)$	$\bar{\gamma}_{cc} \sim \text{Gamma}(\bar{\bar{\gamma}}_{cc}, \bar{\bar{\beta}}_{cc})$ $\frac{\rho}{\nu} = \text{mean}(\sum(x_{cn} + x_{ca}))$ $\nu = 200$
$x_{cn} \sim P(N_0 q)$	$q \sim \text{Gamma}(\rho, \nu)$	$\frac{\rho}{\nu} = \text{mean}(\sum(x_{cn} + x_{ca}))$ $\nu = 200$
	$\pi \sim \text{Beta}(1, 5)$	

Table S3: Parameter information used in all analyses. N_{dn}, N_1, N_0 are sample sizes of families, cases and controls respectively. $\bar{\gamma}$ is mean RRs and β controls the dispersion of γ . $\bar{\bar{\gamma}}$ and $\bar{\bar{\beta}}$ are priors for $\bar{\gamma}$ and are set in advance (they are inferred from simulation data). β is inferred from the equation $e^{a*\bar{\gamma}^b + c}$ inside the estimation process with $a = 6.83$, $b = -1.29$ and $c = -0.58$.

Parameter		Q50	Q5	Q95
π	0.02	0.0224	0.0125	0.0253
	0.05	0.0535	0.0351	0.0611
	0.09	0.0965	0.0752	0.1063
	0.13	0.1381	0.11	0.149
$\bar{\gamma}_{DN}$	5	4.265	3.5608	4.947
	10	8.575	5.7255	10.4417
	15	13.23	9.9955	15.925
	20	17.07	14.2005	20.3087
$\bar{\gamma}_{CC}$	1.5	1.64	1.5938	1.7888
	2	2.21	2.1638	2.2662
	2.5	2.76	2.7138	2.8575
	3	3.225	3.14	3.31
	3.5	3.675	3.5812	3.7663

Table S4: Simulated and estimated values of *de novo* (DN) and case-control (CC) parameters. Q50, Q5 and Q95 are for quantile values of 0.5, 0.05 and 0.95 respectively.

pi	dn_RR	cc_RR	e.pi	Q50	Q5	Q95	e.dn_RR	Q50	Q5	Q95	e.cc_RR	Q50	Q5	Q95
0.02	5	1.5	0.0126	0.0051	0.0224	4.72	2.65	9.22	1.83	1.70	2.53	1.83	1.70	2.53
0.02	5	2	0.0158	0.0054	0.0394	4.08	2.33	11.88	2.27	1.88	3.01	2.27	1.88	3.01
0.02	5	2.5	0.0230	0.0093	0.0419	3.49	2.05	10.25	2.83	2.22	3.43	2.83	2.22	3.43
0.02	5	3	0.0219	0.0123	0.0355	3.52	2.06	8.49	3.31	2.74	4.25	3.31	2.74	4.25
0.02	5	3.5	0.0269	0.0171	0.0373	3.76	2.10	9.59	3.80	3.15	4.73	3.80	3.15	4.73
0.02	10	1.5	0.0185	0.0085	0.0280	5.72	2.82	10.08	1.80	1.54	2.57	1.80	1.54	2.57
0.02	10	2	0.0176	0.0076	0.0373	5.22	2.55	13.43	2.24	1.85	3.05	2.24	1.85	3.05
0.02	10	2.5	0.0218	0.0118	0.0342	5.13	2.39	17.29	2.80	2.27	3.57	2.80	2.27	3.57
0.02	10	3	0.0227	0.0137	0.0344	7.29	2.35	15.88	3.28	2.64	4.23	3.28	2.64	4.23
0.02	10	3.5	0.0255	0.0163	0.0328	7.87	2.79	14.28	3.73	3.07	4.69	3.73	3.07	4.69
0.02	15	1.5	0.0152	0.0046	0.0315	11.63	4.65	28.84	1.77	1.48	2.67	1.77	1.48	2.67
0.02	15	2	0.0213	0.0091	0.0348	10.19	3.32	23.63	2.25	1.86	2.84	2.25	1.86	2.84
0.02	15	2.5	0.0230	0.0118	0.0377	9.73	4.00	20.84	2.69	2.06	3.50	2.69	2.06	3.50
0.02	15	3	0.0226	0.0128	0.0370	10.51	3.53	21.52	3.18	2.55	4.11	3.18	2.55	4.11
0.02	15	3.5	0.0240	0.0140	0.0364	10.56	3.58	20.62	3.62	3.04	4.70	3.62	3.04	4.70
0.02	20	1.5	0.0138	0.0062	0.0398	14.96	5.47	47.26	1.68	1.35	2.29	1.68	1.35	2.29
0.02	20	2	0.0188	0.0079	0.0363	14.10	4.79	36.13	2.28	1.81	3.20	2.28	1.81	3.20
0.02	20	2.5	0.0233	0.0110	0.0343	13.61	5.67	26.36	2.76	2.12	3.54	2.76	2.12	3.54
0.02	20	3	0.0243	0.0140	0.0371	13.89	6.44	23.95	3.41	2.63	4.30	3.41	2.63	4.30
0.02	20	3.5	0.0240	0.0146	0.0352	14.87	6.90	24.81	3.77	2.96	4.65	3.77	2.96	4.65
0.05	5	1.5	0.0343	0.0120	0.0688	4.55	2.26	14.06	1.71	1.51	2.17	1.71	1.51	2.17
0.05	5	2	0.0479	0.0279	0.0699	4.86	2.32	10.44	2.21	1.88	2.60	2.21	1.88	2.60
0.05	5	2.5	0.0556	0.0351	0.0743	4.59	2.06	8.03	2.81	2.31	3.17	2.81	2.31	3.17
0.05	5	3	0.0558	0.0427	0.0722	4.36	2.08	8.18	3.35	2.91	3.74	3.35	2.91	3.74
0.05	5	3.5	0.0621	0.0435	0.0727	3.65	1.89	7.36	3.73	3.22	4.48	3.73	3.22	4.48
0.05	10	1.5	0.0381	0.0161	0.0723	9.20	3.92	15.41	1.74	1.45	2.18	1.74	1.45	2.18
0.05	10	2	0.0531	0.0293	0.0801	8.71	3.70	12.99	2.26	1.91	2.71	2.26	1.91	2.71
0.05	10	2.5	0.0528	0.0386	0.0727	8.76	4.15	14.48	2.74	2.47	3.11	2.74	2.47	3.11
0.05	10	3	0.0569	0.0416	0.0737	8.22	4.47	13.57	3.25	2.83	3.72	3.25	2.83	3.72
0.05	10	3.5	0.0615	0.0491	0.0733	8.06	3.97	13.17	3.66	3.30	4.29	3.66	3.30	4.29
0.05	15	1.5	0.0406	0.0182	0.0877	13.51	6.94	24.71	1.67	1.43	1.98	1.67	1.43	1.98
0.05	15	2	0.0489	0.0311	0.0723	14.04	8.16	22.70	2.19	1.90	2.61	2.19	1.90	2.61
0.05	15	2.5	0.0522	0.0327	0.0734	13.13	8.28	20.66	2.72	2.38	3.11	2.72	2.38	3.11
0.05	15	3	0.0577	0.0449	0.0732	12.37	7.27	18.59	3.19	2.83	3.75	3.19	2.83	3.75
0.05	15	3.5	0.0607	0.0465	0.0756	11.97	8.23	18.55	3.61	3.11	4.29	3.61	3.11	4.29
0.05	20	1.5	0.0418	0.0205	0.0814	18.37	9.74	32.56	1.63	1.37	1.97	1.63	1.37	1.97
0.05	20	2	0.0482	0.0325	0.0697	17.08	10.14	29.26	2.27	1.91	2.60	2.27	1.91	2.60
0.05	20	2.5	0.0537	0.0406	0.0733	16.59	10.57	23.23	2.77	2.29	3.06	2.77	2.29	3.06
0.05	20	3	0.0569	0.0424	0.0770	16.15	10.37	24.32	3.23	2.84	3.75	3.23	2.84	3.75
0.05	20	3.5	0.0596	0.0449	0.0765	15.50	10.23	21.45	3.75	3.19	4.61	3.75	3.19	4.61
0.09	5	1.5	0.0767	0.0404	0.1207	4.46	2.17	9.59	1.66	1.51	1.97	1.66	1.51	1.97
0.09	5	2	0.0904	0.0666	0.1115	4.52	2.04	7.33	2.23	2.03	2.54	2.23	2.03	2.54
0.09	5	2.5	0.0963	0.0753	0.1256	4.70	2.52	7.54	2.79	2.49	3.11	2.79	2.49	3.11
0.09	5	3	0.1040	0.0879	0.1217	3.90	2.08	6.71	3.19	2.85	3.68	3.19	2.85	3.68
0.09	5	3.5	0.1039	0.0876	0.1211	4.22	2.34	7.93	3.70	3.35	4.13	3.70	3.35	4.13
0.09	10	1.5	0.0778	0.0423	0.1208	10.01	5.56	17.73	1.64	1.46	1.93	1.64	1.46	1.93
0.09	10	2	0.0925	0.0660	0.1196	9.26	5.85	13.45	2.16	1.96	2.49	2.16	1.96	2.49
0.09	10	2.5	0.0963	0.0729	0.1170	9.30	7.16	12.50	2.82	2.40	3.18	2.82	2.40	3.18
0.09	10	3	0.0992	0.0831	0.1189	9.25	6.11	12.76	3.22	2.95	3.61	3.22	2.95	3.61
0.09	10	3.5	0.1070	0.0885	0.1222	8.29	5.81	10.94	3.67	3.36	4.20	3.67	3.36	4.20
0.09	15	1.5	0.0822	0.0507	0.1257	14.59	9.22	22.62	1.61	1.43	1.89	1.61	1.43	1.89
0.09	15	2	0.0911	0.0668	0.1217	14.35	9.39	20.13	2.16	1.94	2.45	2.16	1.94	2.45
0.09	15	2.5	0.0978	0.0754	0.1202	13.77	10.40	17.99	2.72	2.40	3.00	2.72	2.40	3.00
0.09	15	3	0.0997	0.0844	0.1206	13.50	10.60	16.88	3.13	2.82	3.49	3.13	2.82	3.49
0.09	15	3.5	0.1036	0.0861	0.1229	12.95	9.89	16.86	3.60	3.20	4.15	3.60	3.20	4.15
0.09	20	1.5	0.0804	0.0495	0.1236	19.92	13.06	31.58	1.60	1.38	1.82	1.60	1.38	1.82
0.09	20	2	0.0920	0.0694	0.1205	18.18	12.71	24.69	2.21	1.95	2.51	2.21	1.95	2.51
0.09	20	2.5	0.0958	0.0742	0.1166	18.28	13.76	22.90	2.75	2.49	3.05	2.75	2.49	3.05
0.09	20	3	0.0974	0.0816	0.1202	17.55	13.32	22.38	3.28	2.95	3.59	3.28	2.95	3.59
0.09	20	3.5	0.1067	0.0925	0.1171	16.49	13.66	20.83	3.68	3.32	4.21	3.68	3.32	4.21
0.13	5	1.5	0.1163	0.0720	0.1671	4.87	2.51	8.11	1.65	1.49	1.83	1.65	1.49	1.83
0.13	5	2	0.1250	0.0991	0.1603	5.15	2.82	7.72	2.22	2.04	2.53	2.22	2.04	2.53
0.13	5	2.5	0.1387	0.1173	0.1654	4.65	2.51	7.04	2.77	2.52	3.08	2.77	2.52	3.08
0.13	5	3	0.1469	0.1220	0.1649	4.40	2.83	6.17	3.20	2.93	3.47	3.20	2.93	3.47
0.13	5	3.5	0.1467	0.1293	0.1747	4.45	2.46	6.13	3.69	3.36	4.25	3.69	3.36	4.25
0.13	10	1.5	0.1094	0.0707	0.1660	10.69	7.35	17.96	1.68	1.53	1.84	1.68	1.53	1.84
0.13	10	2	0.1306	0.1113	0.1529	9.40	6.89	13.22	2.17	2.01	2.36	2.17	2.01	2.36
0.13	10	2.5	0.1432	0.1197	0.1595	9.15	7.21	11.97	2.73	2.52	3.01	2.73	2.52	3.01
0.13	10	3	0.1457	0.1308	0.1682	8.89	6.64	11.08	3.23	3.00	3.54	3.23	3.00	3.54
0.13	10	3.5	0.1497	0.1320	0.1728	8.54	6.62	10.61	3.60	3.26	3.97	3.60	3.26	3.97
0.13	15	1.5	0.1180	0.0778	0.1677	15.08	10.41	22.63	1.60	1.46	1.80	1.60	1.46	1.80
0.13	15	2	0.1277	0.1044	0.1593	14.45	11.56	18.45	2.15	1.96	2.40	2.15	1.96	2.40
0.13	15	2.5	0.1380	0.1124	0.1625	14.34	11.23	17.93	2.72	2.45	3.04	2.72	2.45	3.04
0.13	15	3	0.1432	0.1254	0.1667	13.41	11.03	16.81	3.13	2.87	3.63	3.13	2.87	3.63
0.13	15	3.5	0.1488	0.1281	0.1674	13.00	10.35	16.30	3.56	3.20	3.98	3.56	3.20	3.98
0.13	20	1.5	0.1203	0.0862	0.1765	19.72	13.93	26.77	1.61	1.48	1.83	1.61	1.48	1.83
0.13	20	2	0.1325	0.1093	0.1546	18.54	15.11	23.43	2.21	1.99	2.39	2.21	1.99	2.39
0.13	20	2.5	0.1351	0.1130	0.1601	18.38	14.63	22.97	2.79	2.50	3.00	2.79	2.50	3.00
0.13	20	3	0.1434	0.1256	0.1645	18.43	14.94	22.27	3.24	2.94	3.60	3.24	2.94	3.60
0.13	20	3.5	0.1488	0.1320	0.1637	16.81	13.99	20.18	3.64	3.34	4.11	3.64	3.34	4.11

Table S5: Estimated values for the cases in Table S4, for each unique set of parameter values. The first three columns are simulated values. The following columns show estimated π , *de novo* mean relative risk (dn RR) and case-control (cc) RR; for each parameter, shown are median (Q50) and 5th and 95th %-iles (Q5 and Q95) estimates over 100 simulation replicates.

$\bar{\beta}_{DN}$	$\bar{\beta}_{CC}$	$e.\pi$	$e.\bar{\beta}_{DN}$	$e.\bar{\beta}_{CC}$	$e.\beta_{DN}$	$e.\beta_{CC}$	FDR0.01	FDR0.05	FDR0.1	FDR0.25	FDR0.5
0.01	0.11	0.0008985	12.16	2.37	0.82	1.38	0	0	0	0	0
0.01	0.14	0.0013925	7.76	2.02	0.84	2.08	0	0	0	0	0
0.01	0.2	0.0011444	7.38	1.66	0.86	3.2	0	0	0	0	0
0.01	0.33	0.0014319	10.32	1.46	0.83	5.06	0	0	0	0	0
0.01	1	0.0010192	6.12	1.26	0.87	24.04	0	0	0	0	0
0.02	0.11	0.0012389	5.7	1.72	0.88	2.07	0	0	0	0	0
0.02	0.14	0.00339	6.25	1.6	0.88	5.1	0	0	0	0	0
0.02	0.2	0.0036757	12.62	1.53	0.83	4.77	0	0	0	0	0
0.02	0.33	0.0040126	3.34	1.32	1.14	15.47	0	0	0	0	0
0.02	1	0.0057346	5.27	1.15	0.92	51.7	0	0	0	0	0
0.03	0.11	0.0012311	7.23	1.63	0.87	2.43	0	0	0	0	0
0.03	0.14	0.0009967	6.37	1.61	0.87	3.88	0	0	0	0	0
0.03	0.2	0.0022818	5.16	1.55	0.92	5.4	0	0	0	0	0
0.03	0.33	0.0110319	4.16	1.35	1.02	16.06	0	0	0	0	2
0.03	1	0.004111	3.75	1.19	1.03	42.34	0	0	0	0	0
0.05	0.11	0.0018204	5.78	1.38	0.9	5.92	0	0	0	0	0
0.05	0.14	0.0015779	7.84	2.04	0.86	2.14	0	0	0	0	0
0.05	0.2	0.0034645	4.75	1.34	0.94	9.15	0	0	0	0	0
0.05	0.33	0.0123621	1.75	1.24	2.27	24.09	0	0	0	0	0
0.05	1	0.0035687	3.63	1.18	1.03	47.33	0	0	0	0	0

Table S6: Estimated values in the case $\pi = 0$ and $\bar{\gamma} = 1$. The first two columns are $\bar{\beta}$ values (prior information of $\bar{\gamma}$: $\bar{\gamma} \sim \text{Gamma}(1, \bar{\beta})$). The third to the seventh columns are genetic parameters estimated from **extTADA**. Next columns are the number of risk genes estimated with the corresponding FDR values in the header.

Disease	Mutation	Count	Sample size	Mutation count per sample size
SCZ	silentFCPk	50	1077	0.05
	MiD	105	1077	0.1
	LoF	116	1077	0.11
ASD	MiD	618	5122	0.12
	LoF	638	5122	0.12
ID	MiD	223	1022	0.22
	LoF	225	1022	0.23
EPI	MiD	69	356	0.19
	LoF	52	356	0.15
DD	MiD	1041	4293	0.24
	LoF	1066	4293	0.25

Table S7: *de novo* mutation counts of categories and their mutation counts per sample size for schizophrenia (SCZ), autism spectrum disorder (ASD), epilepsy (EPI), intellectual disorder (ID) and developmental disorder (DD).

Parameters	Estimated mode	lCI	uCI
SCZ_pi_silentFCPkdn	0.0056	0	0.1977
SCZ_hyperGammaMean_silentFCPkdn	1.5802	1.001	21.5139
SCZ_pi_MiDdn	0.012	0	0.2368
SCZ_hyperGammaMean_MiDdn	1.7486	1	17.8548
SCZ_pi_LoFdn	0.0548	0.0124	0.2062
SCZ_hyperGammaMean_LoFdn	11.1857	3.3973	31.3602
SCZ_pi_MiD+LoFcc	0.069	0.0296	0.1359
SCZ_hyperGammaMean_MiD+LoFcc	2.0176	1.2133	5.3694
SCZ_hyperGammaMean_MiD+LoFcc	3.2288	1.2372	17.1478
SCZ_hyperGammaMean_MiD+LoFcc	1.0691	1.0002	2.9574

Table S8: Genetic parameters for SCZ data if single class is used in the analysis.

Table S9: extTADA results of SCZ risk gene identification (See LongSupTables.xlsx Download).

Table S10: extTADA risk gene identification results of ID data (See LongSupTables.xlsx Download).

Table S11: extTADA risk gene identification results of DD data (See LongSupTables.xlsx Download).

Table S12: extTADA risk gene identification results of ASD data (See LongSupTables.xlsx Download).

Table S13: extTADA risk gene identification results of EPI data (See LongSupTables.xlsx Download).

Parameters	Estimated mode	lCI	uCI
SCZ_pi0	9.37	5.47	15.12
SCZ_meanRR_silentFCPkdenovo	1.3068	1.0005	2.7489
SCZ_meanRR_MiDdenovo	2.2246	1.0006	5.3491
SCZ_meanRR_LoFdenovo	15.1491	5.8606	27.3941
SCZ_meanRR_MiD+LoFccPop1	1.8677	1.0374	3.0736
SCZ_meanRR_MiD+LoFccPop2	2.2632	1.003	4.9168
SCZ_meanRR_MiD+LoFccPop3	1.0372	1.0002	1.1807
ASD_pi	9.47	7.61	12.27
ASD_meanRR_MiDdenovo	5.09	2.47	10.51
ASD_meanRR_LoFdenovo	20.23	12.21	32.31
ASD_meanRR_LoFcc	2.48	1.48	5.95
ID_pi	3.53	2.63	4.56
ID_meanRR_MiDdenovo	35.29	21.46	51.62
ID_meanRR_LoFdenovo	105.44	74.58	143.02
DD_pi	1.91	1.57	2.37
DD_meanRR_MiDdenovo	22.72	13.91	34.36
DD_meanRR_LoFdenovo	99.94	75.39	127.18
EPI_pi	1.67	0.96	3.1
EPI_meanRR_MiDdenovo	71.77	37.15	125.14
EPI_meanRR_LoFdenovo	94.98	51.73	176.16

Table S14: SCZ and NDD genetic parameters after adjusting mutation rates.

Parameter	Mode	lCI	uCI
pi0	0.0821	0.0487	0.1398
hyperGammaMeanDN[1]	1.2199	1.0001	2.2
hyperGammaMeanDN[2]	1.4407	1.0043	2.9893
hyperGammaMeanDN[3]	11.9591	4.1894	23.9414
hyperGammaMeanCC	1.9498	1.0845	3.2072

Table S15: Estimated genetic parameters for SCZ data with the same mean RRs for case-control data.

Parameters	Estimated mode	lCI	uCI
SCZ_pi	0.0732	0.0306	0.1506
SCZ_meanRR_silentFCPkdenovo	1.2353	1.0021	3.6086
SCZ_meanRR_MiDdenovo	1.4459	1.0008	4.7004
SCZ_meanRR_LoFdenovo	12.0403	4.6136	25.8786
SCZ_meanRR_MiD+LoFccPop1	1.5856	1.1255	4.0881
SCZ_meanRR_MiD+LoFccPop2	1.7361	1.0438	4.8856
SCZ_meanRR_MiD+LoFccPop3	1.0698	1.0001	2.9991

Table S16: SCZ genetic parameters using all variants in and not in ExAC database (InExAC + NoExAC).

Table S17: extTADA results of SCZ risk gene identification after adjusting mutation rates (See LongSupTables.xlsx Download).

Gene set	GN0	GN0	P value	FDR
pLI09	3488	3241	1.00e-05	8.45e-03
rbfox2	3068	2895	1.33e-05	8.45e-03
GGGAGGRR_V\$_MAZ_Q6	2274	2114	3.50e-05	1.33e-02
ACAGGGT,MIR-10A,MIR-10B	123	116	3.00e-05	1.33e-02
chd8.human.brain	2798	2601	5.00e-05	1.58e-02
rbfox13	3445	3230	1.70e-04	4.62e-02
FMRP_targets	839	792	2.10e-04	4.99e-02

Table S18: Enrichment of gene sets from different databases with SCZ genes from **extTADA** results. These p values were obtained by 10,000,000 simulations, and then adjusted by using the method of [Benjamini and Hochberg \(1995\)](#).

Table S19: The p values of enrichment tests for known gene sets in SCZ, DD, ID, ASD and EPI (See LongSupTables.xlsx Download).

Table S20: The p values of enrichment tests for all gene sets in SCZ, DD, ID, ASD and EPI (See LongSupTables.xlsx Download).

Table S21: Community memberships of the GeNets InWeb PPI 288 NDD genes network.

Table S22: Enrichment results of GeNets. These are enrichment results of 6 communities obtained from GeNets.

References

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